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Genetic Associations of Periodic Limb Movements of Sleep in the Elderly, for the MrOS Sleep Study

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Abstract

Objective—Assess the relationship of single nucleotide polymorphisms associated with restless legs syndrome to periodic limb movements of sleep in a population cohort of elderly individuals.

Methods—Single nucleotide polymorphisms previously associated with periodic limb movements of sleep or restless legs syndrome were analyzed in 2,356 white male participants in the osteoporotic fractures in men study sleep cohort. Associations of single nucleotide polymorphisms to polysomnographic measured periodic limb movement index 15 were examined with logistic regression adjusted for age, ancestry markers, and periodic limb movements of sleep risk factors.

Results—Of the men in this cohort, 61% had a periodic limb movement index 15. Significant associations were observed between a periodic limb movement index 15 and the number of risk alleles for the two BTBD9 single nucleotide polymorphisms (rs9357271[T], OR=1.38; 95%CI 1.20-1.58, and rs3923809[A], OR=1.43; 95%CI 1.26-1.63), one of the MEIS1 single nucleotide polymorphisms (rs2300478[G], OR=1.31; 95%CI 1.14-1.51) and the MAP2K5/SKOR1 single

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nucleotide polymorphism (rs1026732[G], OR=1.16; 95%CI 1.02-1.31). In a multivariable model controlling for each of the two MEIS1 single nucleotide polymorphisms, the rs6710341[A] single nucleotide polymorphism became a significant risk allele (OR=1.59; 95%CI 1.26-2.00).

Conclusions—Our findings confirm an association of the BTBD9, MEIS1 and MAP2K5/SKOR1 single nucleotide polymorphisms and periodic limb movements of sleep in an elderly cohort not selected for the presence of restless legs syndrome.

Keywords

periodic limb movements of sleep; single nucleotide polymorphism; restless legs syndrome; BTBD9; MEIS1

Periodic limb movements of sleep (PLMS) are contractions of the lower extremities which occur roughly every 20 seconds during sleep, may produce brief arousals from sleep(1) and are associated with dramatic increases in blood pressure (2). They are most prominent in the elderly and in those taking antidepressant medications. In addition, they are observed in roughly 90% of individuals with restless legs syndrome (RLS), in which they are considered the motor sign of the waking urge to move which characterizes that disorder. RLS is associated with a number of specific single nucleotide polymorphisms (SNPs) based on genome wide association scans (GWAS) (3, 4). Two studies have also demonstrated that these SNPs are associated with PLMS (3, 5). The present study further attempts to determine the relationship of these SNPs to PLMS in a population cohort of elderly individuals not selected for the presence of RLS.

METHODS

Study sample

During the Osteoporotic Fractures in Men Study (MrOS) baseline examination from 2000 to 2002, 5,994 community-dwelling men 65 years or older were enrolled at six clinical centers in the United States: Birmingham, Alabama; Minneapolis, Minnesota; Palo Alto, California; the Monongahela Valley near Pittsburgh, Pennsylvania; Portland, Oregon; and San Diego, California (6, 7). In order to participate, men needed to be able to walk without assistance and must not have had a bilateral hip replacement.

The MrOS Sleep Study, an ancillary study of the parent MrOS cohort, recruited 3,135 participants for a comprehensive sleep assessment who did not require positive airway pressure or nocturnal oxygen therapy during polysomnography (PSG). Because the candidate SNPs examined were established in Caucasian populations, 319 were excluded because they were non-white. A further 175 who did not have PSG were excluded. Of these 2,641 participants, 285 did not have SNP data, leaving 2,356 in this analysis cohort. Compared to this analysis subset, those 460 white men who were not included in the analysis were similar in age, body mass index (BMI), PLMS and RLS, and lifestyle covariates. They did have a higher rate of use of dopaminergic medication (2.4% vs. 1.2%, $p = 0.03$).

All men provided written informed consent, and the study was approved by the Institutional Review Board at each site.

Polysomnography

Sleep testing was conducted with unattended in-home polysomnography (Safiro, Compumedics, Inc, Melbourne, Australia). The recording montage included the following: C₃/A₂ and C₄/A₁ electroencephalography, bilateral electrooculography, a bipolar submental electromyography, thoracic/abdominal respiratory inductance plethysmography, naso-oral thermistery, nasal pressure transduction, finger pulse oximetry, lead I ECG, body position, and bilateral anterior tibialis piezoelectric movement sensors. Centrally trained staff performed home visits for unit setup and impedance value verification for each channel as previously described (8). Data were downloaded to a central server at the Central Sleep Reading Center (Cleveland, OH) and scored by certified research polysomnologists using standard criteria (9, 10). Apnea was defined as complete or near complete reduction in thermistor amplitude for at least 10 seconds; hypopnea was defined as at least 30% reduction in nasal pressure signal or summed inductance bands for at least 10 seconds. The apnea-hypopnea index (AHI) was calculated as the total number of apneas and hypopneas per hour of sleep, each associated with $\geq 4\%$ desaturation (10). Arousals were scored according to the American Academy of Sleep Medicine criteria (9).

PLMS were scored according to the 1993 American Academy of Sleep Medicine criteria in which individual leg movements were scored if the duration was between 0.5 and 5 seconds and there was a clear amplitude increase from baseline in leg channels (11). To be considered periodic, at least 4 movements needed to occur in succession no less than 5 seconds and no more than 90 seconds apart. The periodic limb movement index (PLMI) was the total number of periodic leg movements per hour of sleep. Leg movements after respiratory events were excluded unless they were part of a ≥ 4 movement cluster with at least 2 movements occurring independently of respiratory events. The PLM parameter were expressed as both a continuous and categorical variable (PLMI ≥ 15 vs. <15). This value was chosen for the categorical threshold as it is the index above which a diagnosis of periodic limb movement disorder (PLMD) can be made (12). It is also the threshold at which a recent study found both reduced night to night variability and optimal specificity/sensitivity (13). Similarly, this threshold has been used in previous studies of this cohort (14).

Genetics

SNPs within the candidate gene regions previously associated with PLMS or RLS were analyzed (MEIS1 SNP -rs2300478, MEIS1 SNP -rs6710341, BTBD9 SNP-rs3923809, BTBD9 SNP-rs9357271, and MAP2K5/SKOR1 SNP-rs1026732). The California Pacific Medical Center investigators and the University of California, San Diego group collaborated to develop a custom Illumina Golden Gate assay (Illumina, San Diego, CA) to genotype polymorphisms in sleep-related genes (15). Genotype concordance rate was > 0.99 (8% of MrOS samples were plated in duplicate). Samples with $< 90\%$ SNP call rate were excluded. All SNPs summarized here were in Hardy–Weinberg equilibrium ($p > 0.001$).

Other measures

Other covariates examined included age, BMI, medication use (16), smoking, education, alcohol use, physical activity (17), depression (18), and other PSG-derived parameters (sleep staging, AHI, total sleep time, time awake after sleep onset). Ancestry markers were calculated using Autosomal SNPs in multidimensional scaling analysis (MDS) as implemented in PLINK. The first two MDS components were used as covariates (19).

Statistical analysis

Characteristics of participants were compared across PLMI status using chi-square tests for categorical variables, t-tests for normally distributed continuous variables, and Wilcoxon rank sum tests for continuous variables with skewed distributions. A linear trend in PLMI across number of risk alleles for each SNP was examined with a linear regression model. Associations of SNPs and PLMI ≥ 15 were examined with logistic regression models and presented as odds ratios (ORs) and their 95% confidence intervals (CIs). Odds ratios were expressed as dose of the risk allele used as a linear variable (0,1,2). Models were first adjusted for age and ancestry markers, then multivariable models were further adjusted for those covariates shown to be associated with PLMS at $p < 0.10$.

Exploratory analyses examining the association of combinations of these SNPs and PLMI were performed to determine if the associations were independent. Conditional models for the BTBD9 and MEIS1 genes were performed, with the two SNPs in each of these genes as variables in each of the models to determine if these SNPs were independently associated with the PLMI outcome. The SNPs from each gene with the highest association to PLMI were also combined in one predictor (0-6). In this logistic model, the odds ratio was expressed as number of risk alleles used as a linear variable

RESULTS

The overall prevalence of a PLMI ≥ 15 was 61% (Table 1). Compared to those with a PLMI < 15 , those with a PLMI ≥ 15 were on average older, were less likely to have an AHI ≥ 15 , had less total sleep time, more likely to be non-drinkers, and had more depressive symptoms. There were no differences in the use of antidepressants, anticonvulsants, dopaminergics or benzodiazepines between those with and without a PLMI ≥ 15 (Table 1).

Allele frequencies for each of the five SNPs based on PLMI category (PLMI > 15) are shown in Table 2. PLMI as a continuous measure was significantly associated with the number of risk alleles for 4 of the 5 SNPs examined (Figure 1). After multivariable adjustment, significant associations were observed between the PLMI outcome and the number of risk alleles for the two BTBD9 SNPs (rs9357271[T], OR=1.38; 95% CI 1.20- 1.58, and rs3923809[A], OR=1.43; 95% CI 1.26-1.63), one of the MEIS1 SNPs (rs2300478[G], OR=1.31; 95% CI 1.14-1.51) and the MAP2K5/SKOR1 SNP (rs1026732[G], OR=1.16; 95% CI 1.02-1.31) (Table 3). Analyses using PLMI cutoffs of > 10 per hour (which constituted 67% of the cohort) were very similar, with slightly stronger associations of the PLMI outcome and the SNPs.

In further analyses, we included SNPs on the same gene in a multivariable model. In the multivariable model controlling for each of the BTBD9 SNPs, rs3923809[A] continued to be a significant predictor of a PLMI ≥ 15 but rs9357271[T] did not (rs3923809[A], OR=1.35; 1.13, 1.61 and rs9357271[T], OR=1.10; 0.90, 1.33) (Table 4). In the multivariable model controlling for each of the two MEIS1 SNPs, the rs6710341[A] SNP became a significant risk allele (OR=1.59; 95% CI 1.26-2.00) and rs2300478[G] maintained its significance as a risk allele (Table 4). When a SNP from each gene was combined in a multivariable model with the risk alleles as a continuous measure (0-6), the combined gene score was associated with a 1.28-fold increase in odds of having a PLMI ≥ 15 per each additional risk allele (OR=1.28; 95% CI 1.19-1.38) (Table 4).

DISCUSSION

Our findings in a general cohort of older Caucasian men confirm a previous report of an association of the BTBD9 SNP rs3923809 (A allele) and PLMS in two groups selected for the presence of RLS (3). It further extends the association of PLMS to SNPs in the MEIS1 and MAP2K5/SKOR1 genes as well as to another BTBD9 SNP, all originally identified from RLS populations using GWAS (4, 20-22), and all in the elderly who have the highest rates of PLMS. We also establish these genetic associations using PLMS derived from full polysomnographic recordings, which, unlike the original GWAS study, allow us to distinguish leg movements occurring during sleep from those during recumbent wake.

Our point estimates for the association of BTBD9 and MEIS1 SNPs with a PLMI ≥ 15 are very similar to those recently determined by Moore et al.(5), who like us, also found that the independent effects of each BTBD9 SNP were stronger than those of the MEIS1 SNPs for PLMI. However, in conditional analyses (assessing the independent effects of the two SNPs in each gene), we found that the MEIS1 SNPs had individually stronger effects than any of the other SNPs examined. Further, the BTBD9 SNPs had substantially lower point estimates when examined independently than when tested individually. Future investigations should consider such analyses when assessing the role of these SNPs into both PLMS and RLS. Finally, the similarity of the ORs in our study and Moore's is notable given that they used different methods for leg movement detection (surface EMG vs accelerometry) and scoring (automatic scoring vs hand scoring).

An important distinction between the Moore et al study (5) and ours is that they studied subjects in a very narrow middle-age range (59.7 ± 0.17 yrs) whereas our subjects were substantially older (76.5 ± 5.58 yrs). As PLMS become more prominent with age (23) it is not surprising that whereas only 33% of subjects in the Moore paper had a PLMI >15 , 61% met this criterion in our study. The fact that these SNPs predicted PLMI in this older group is significant as it demonstrates that genetics continues to play an important role in PLMS even at an age when leg movements develop in the majority of the population.

BTBD9 is broadly expressed in the brain as well as in visceral organs though its function is not well understood. It has been associated with both primary as well as secondary RLS (24) as well as Tourette's syndrome (25). BTBD9 knockouts produce abnormal sleep and motor activity in both mice (26) and *Drosophila* (27). MEIS1 is a transcription factor known to be

involved in neuronal development and connectivity, and rare variants have recently been associated with RLS (28). The rs1026732 SNP is between the MAP2K5 and SKOR1 genes, both of which have been implicated in neuronal function.

As we have previously proposed, PLMS may be an endophenotype for RLS (29). Endophenotypes are stable, biologically heritable, quantifiable aspects of a more complex disorder. In this way, PLMS may provide insights into the pathophysiology of RLS and/or identify subtypes of RLS that may respond to specific therapeutic interventions. PLMS and RLS may also be part of a disease continuum in which PLMS without RLS represents subclinical disease and increasing leg movement frequency and intensity may be a marker for a more clinically virulent form of RLS.

There are a number of limitations to this study. Our data is limited to older Caucasian men and may not be generalizable to other populations. Piezoelectric sensors and not the gold standard anterior tibialis electromyography were used to measure PLMS. However, PLMI scored using our approach (i.e., 1993 scoring criteria with piezoelectric sensors) is highly correlated ($r=0.81$) with the PLMI derived using the 2013 (30) scoring criteria and use of anterior tibialis EMG (14). Further, the similarity of our BTBD9 and MEIS1 SNP ORs to those of Moore et al (5) who used surface EMG suggests that the two recording methods define substantially similar movements. The assessment of PLMS was performed on a single night, and given night to night variability of leg movements, we may have misclassified some participants (31). However, this would likely bias the associations of these SNPs and PLMS toward the null. Further, analysis of the data using a PLMI cutoff of >10 rather than >15 (which minimizes the effects of night to night variability) did not change the associations with these SNPs. Finally, we did not examine all RLS-related SNPs and future studies should also include other SNPs such as TOX3/BC034767 and PTPRD(A).

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Highlights

- PLMS are associated with single nucleotide polymorphisms (SNPs).
- Logistic regression was used to associate SNPs to polysomnographic measured PLMS.
- Two BTBD9, one MEIS1 and one MAP2K5/SKOR1 SNPs were associated with PLMS>15
- Genetics continues to play an important role in PLMS even in the elderly

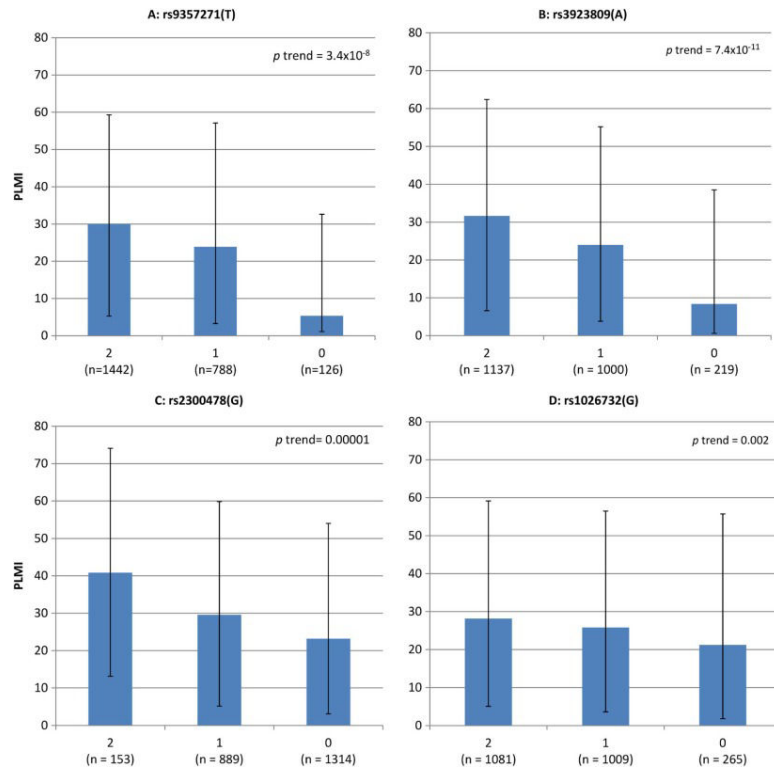


Figure 1. Periodic limb movement index (PLMI) in the four associated SNPs by number of risk alleles, median (inter quartile range)

Table 1

Baseline Characteristics by Periodic Leg Movement Index (PLMI)

Characteristic	All (n = 2,356)	PLMI 15 (n = 1,446)	PLMI<15 (n = 910)	<i>p</i> value
Age, y	76.54 ± 5.58	76.91 ± 5.67	75.95 ± 5.38	<.0001
Body mass index, kg/m ²	27.23 ± 3.76	27.27 ± 3.69	27.16 ± 3.88	0.47
Apnea-hypopnea index ^a	11.86 ± 13.25	11.44 ± 12.9	12.53 ± 13.76	0.14
Apnea-hypopnea index 15	621 (26.36)	358 (24.76)	263 (28.9)	0.03
Self-reported diagnosis of restless legs syndrome	56 (2.41)	39 (2.74)	17 (1.9)	0.20
Total sleep time, min	355.14 ± 69.04	352.71 ± 70.78	358.99 ± 66.04	0.03
Percent of sleep time in nREM	80.90 ± 6.69	80.95 ± 6.81	80.82 ± 6.50	0.13
Percent of sleep time in REM	19.1 ± 6.69	19.05 ± 6.81	19.17 ± 6.5	0.67
Minutes awake after sleep onset	114.6 ± 66.28	114.68 ± 67.31	114.48 ± 64.66	0.94
Current benzodiazepine use	110 (4.67)	70 (4.84)	40 (4.4)	0.62
Current antidepressant use	195 (8.28)	126 (8.71)	69 (7.58)	0.33
Current dopaminergic use	27 (1.15)	16 (1.11)	11 (1.21)	0.82
Current use of anticonvulsants	79 (3.35)	45 (3.11)	34 (3.74)	0.41
Education				0.57
< High school	112 (4.75)	65 (4.5)	47 (5.16)	
High school	402 (17.06)	254 (17.57)	148 (16.26)	
Some college or more	1,842 (78.18)	1,127 (77.94)	715 (78.57)	
Smoking				0.45
Never	925 (39.28)	578 (39.97)	347 (38.17)	
Past	1,387 (58.9)	839 (58.02)	548 (60.29)	
Current	43 (1.83)	29 (2.01)	14 (1.54)	
Alcohol intake, drinks/week				0.02
<1	1,084 (46.21)	697 (48.37)	387 (42.76)	
1-13	1,136 (48.42)	665 (46.15)	471 (52.04)	
14	126 (5.37)	79 (5.48)	47 (5.19)	
PASE physical activity score	145.46 ± 70.66	143.82 ± 68.73	148.05 ± 73.59	0.16
Depression symptom score ^b	1.75 ± 2.11	1.82 ± 2.08	1.64 ± 2.14	0.005

Abbreviations: PLMI = periodic leg movement index; nREM = non-rapid eye-movement; REM=rapid eye movement; PASE = The Physical Activity Scale for the Elderly.

Data shown as mean ± SD or n (%).

P values for continuous variables are from a t-test if normally distributed, a Wilcoxon rank-sum test if skewed. *P* values for categorical data are from a chi-square test.

^a Apneas and hypopneas at 4% oxygen desaturation per hour of sleep.

^b Geriatric Depression Scale score, range 0-1

Table 2

Allele Frequencies by Periodic Leg Movement Index (PLMI)

Gene	SNP	All (N= 2356)	PLMI \geq 15 (N= 1446)	PLMI<15 (N= 910)	<i>p</i> value
BTBD9	rs9357271				9.6E-07
	T_T	1442 (61.21)	926 (64.04)	516 (56.7)	
	C_T	788 (33.45)	468 (32.37)	320 (35.16)	
	C_C	126 (5.35)	52 (3.6)	74 (8.13)	
	rs3923809				2.02E-08
	A_A	1137 (48.26)	750 (51.87)	387 (42.53)	
	G_A	1000 (42.44)	597 (41.29)	403 (44.29)	
G_G	219 (9.3)	99 (6.85)	120 (13.19)		
MEIS1	rs6710341				0.15
	G_G	53 (2.25)	26 (1.8)	27 (2.97)	
	G_A	614 (26.06)	385 (26.63)	229 (25.16)	
	A_A	1689 (71.69)	1035 (71.58)	654 (71.87)	
	rs2300478				0.00063
	G_G	153 (6.49)	110 (7.61)	43 (4.73)	
	G_T	889 (37.73)	569 (39.35)	320 (35.16)	
T_T	1314 (55.77)	767 (53.04)	547 (60.11)		
MAP2K5/SKOR1	rs1026732				0.05
	G_G	1081 (45.9)	687 (47.51)	394 (43.34)	
	A_G	1009 (42.85)	611 (42.25)	398 (43.78)	
	A_A	265 (11.25)	148 (10.24)	117 (12.87)	

P-values from a chi-square test.

Table 3Associations of SNPs with Periodic Leg Movement Index (PLMI ≥ 15 vs. <15)

Gene	SNP	Risk Allele Frequency	Age+Ancestry Marker Adjusted		Multivariable ^a Adjusted	
			OR (95% CI)	p value	OR (95% CI)	p value
BTBD9	rs9357271(T)	0.779	1.41 (1.23-1.63)	1.14×10^{-6}	1.38 (1.20-1.58)	9.22×10^{-6}
	rs3923809(A)	0.695	1.45 (1.28-1.65)	1.34×10^{-8}	1.43 (1.26-1.63)	6.45×10^{-8}
MEIS1	rs6710341(A)	0.847	1.04 (0.88-1.22)	0.68	1.04 (0.88-1.22)	0.68
	rs2300478(G)	0.254	1.33 (1.15-1.52)	6.91×10^{-5}	1.31 (1.14-1.51)	0.0001
MAP2K5/SKOR1	rs1026732(G)	0.673	1.16 (1.02-1.31)	0.02	1.16 (1.02-1.31)	0.02

Abbreviations: PLMI = periodic leg movement index; SNP = single nucleotide polymorphism; OR= odds ratio; CI = confidence interval.

^a Adjusted for age, clinic, ancestry markers, apnea-hypopnea index ≥ 15 , total sleep time, alcohol use, and number of depressive symptoms.

Table 4Associations of Combinations of SNPs with Periodic Leg Movement Index (PLMI ≥ 15 vs. <15)

Gene/SNP	Multivariable ^a Adjusted	
	OR (95% CI)	<i>p</i> value
BTBD9, SNP rs9357271(T) ^b	1.10 (0.90-1.33)	0.35
BTBD9, SNP rs3923809(A) ^b	1.35 (1.13-1.61)	0.001
MEIS1, SNP rs6710341(A) ^b	1.59 (1.26-2.00)	7.40×10^{-5}
MEIS1, SNP rs2300478(G) ^b	1.70 (1.40-2.06)	7.50×10^{-8}
Number of risk alleles for 1 SNP from each gene (0-6) ^c	1.28 (1.19-1.38)	1.69×10^{-10}

Abbreviations: SNP = single nucleotide polymorphism; PLMI = periodic leg movement index; OR= odds ratio; CI = confidence interval.

^a Adjusted for age, clinic, ancestry markers, apnea-hypopnea index ≥ 15 , total sleep time, alcohol use, and number of depressive symptoms.

^b Results from a model containing both SNPs from this gene as separate predictors (0,1,2).

^c Number of A alleles for rs3923809, number of G alleles for rs2300478, number of G alleles for rs1026732.