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## Maternal Exposure to Traffic-Related Air Pollution and Birth Defects in Massachusetts

Mariam S. Girguis<sup>a</sup>, Matthew J. Strickland<sup>b</sup>, Xuefei Hu<sup>b</sup>, Yang Liu<sup>b</sup>, Scott M. Bartell<sup>a,c,d</sup>, and Verónica M. Vieira<sup>a</sup>

<sup>a</sup>Program in Public Health, University of California, Irvine, CA, USA

<sup>b</sup>Department of Environmental Health, Emory University Rollins School of Public Health, Atlanta, GA, USA

<sup>c</sup>Department of Statistics, Donald Bren School of Information and Computer Sciences, University of California, Irvine, CA, USA

<sup>d</sup>Department of Epidemiology, School of Medicine, University of California, Irvine, CA, USA

### Abstract

Exposures to particulate matter with diameter of 2.5  $\mu\text{m}$  or less ( $\text{PM}_{2.5}$ ) may influence risk of birth defects. We estimated associations between maternal exposure to prenatal traffic-related air pollution and risk of cardiac, orofacial, and neural tube defects among Massachusetts births conceived 2001 through 2008. Our analyses included 2,729 cardiac, 255 neural tube, and 729 orofacial defects. We used satellite remote sensing, meteorological and land use data to assess  $\text{PM}_{2.5}$  and traffic-related exposures (distance to roads and traffic density) at geocoded birth addresses. We calculated adjusted odds ratios (OR) and confidence intervals (CI) using logistic regression models. Generalized additive models were used to assess spatial patterns of birth defect risk. There were positive but non-significant associations for a  $10\mu\text{g}/\text{m}^3$  increase in  $\text{PM}_{2.5}$  and perimembranous ventricular septal defects (OR = 1.34, 95% CI: 0.98, 1.83), patent foramen ovale (OR = 1.19, 95% CI: 0.92, 1.54) and patent ductus arteriosus (OR = 1.20, 95% CI: 0.95, 1.62). There was a non-significant inverse association between  $\text{PM}_{2.5}$  and cleft lip with or without palate (OR = 0.76, 95% CI: 0.50, 1.10), cleft palate only (OR = 0.89, 95% CI: 0.54, 1.46) and neural tube defects (OR = 0.77, 95% CI: 0.46, 1.05). Results for traffic related exposure were similar. Only ostium secundum atrial septal defects displayed significant spatial variation after accounting for known risk factors.

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*Corresponding author and full address:* Verónica Vieira, D.Sc., Program in Public Health, University of California, 653 E. Peltason Dr., Irvine, CA, 92697-3957, USA.

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The Institutional Review Boards of the University of California at Irvine and the MA Department of Public Health approved this research.

## Keywords

air pollution; birth defects; near-roadway pollution; satellite-based PM<sub>2.5</sub>; traffic-density

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## 1. Introduction

Birth defects are prevalent in 3% of US live births (1), with cardiac, orofacial, and neural tube defects among the most common defects observed (2). Exposure to air pollution during pregnancy has been suggested to increase risk of birth defects (3–8) in some studies. The time between conception and birth is a sensitive and critical time for fetal development due to rapid cell proliferation and rapid development of various organ systems, thus understanding the influence of ambient exposures on fetal development may elucidate the mechanisms behind abnormal fetal development. Studies of fetal exposure to traffic-related air pollution including particulate matter with a diameter of 2.5 µm or less (PM<sub>2.5</sub>) have shown associations with adverse birth outcomes such as intrauterine growth retardation and preterm births (9–10), but investigations of the association of PM<sub>2.5</sub> on birth defect risk have been inconclusive (11–18).

Exposure estimates for earlier studies were constrained to individuals living near air monitoring stations without daily assessments, limiting both spatial and temporal resolution of the exposure assessment resulting in exposure misclassification (19, 20). Earlier studies were unable to adjust for important confounders such as individual-level socioeconomic status (SES) and may have been limited by case ascertainment over a short study period (19, 20). Only one other study (17) to our knowledge has accounted for road density and residential distance to roadways, local measures of traffic-related air pollution, in addition to PM<sub>2.5</sub> estimates when assessing risk of birth defects and exposure to ambient air pollution. Satellite-based PM<sub>2.5</sub> prediction models can provide additional spatial and temporal information for exposure assessment. Models have evolved from using one single predictor (21) to multiple predictors (22–24) and from one-stage models (25) to multi-stage non-linear models (26–29).

The objective of this study is to examine the relationship between cardiac, orofacial, and neural tube defects and traffic-related air pollution using satellite-based PM<sub>2.5</sub> exposure estimates and eight years of birth defects data for Massachusetts. To further assess the influence of PM<sub>2.5</sub> on birth defect risk, our study includes an analysis of geographic patterns of birth defects across Massachusetts.

## 2. Methods

### 2.1 Study population

We obtained all live and still births from the Massachusetts state birth registry with an estimated conception date from January 1, 2001 through December 31, 2008. All births in the Massachusetts Birth Defects Registry having cardiac, orofacial, and neural tube defects (*International Classification of Diseases*, Ninth Revision, Clinical Modification (ICD-9-CM) codes 740.0–743.0, 745.0–748.0 and 749.0–749.3) were identified as cases. The

Massachusetts Birth Defects Monitoring Program conducts active surveillance to collect diagnoses made before age one. We randomly selected 1,000 infants conceived each year to serve as a common control group among all live births without the defects of interest. We excluded birth addresses that could not be successfully geocoded to x and y coordinates (2%) and we excluded syndromic births (5%) that were associated with the outcomes of interest (eTable 1) among cases and controls. The Institutional Review Boards of the University of California at Irvine and the Massachusetts Department of Public Health approved this research.

Due to differing etiologies, we divided cardiac defects into anatomical groupings based on ICD-9-CM groups. We only included groups with more than 70 cases to ensure sufficient numbers for model convergence. We also assessed the five most common single ICD-9-CM code cardiac defect diagnoses as their own outcome group. In total, there were 17 outcome groups for cardiac anomalies: transposition of great vessels, tetralogy of fallot, ventricular septal defect, ostium secundum atrial septal defect, endocardial cushion defect, pulmonary valve atresia and stenosis, aortic valve stenosis, hypoplastic left heart syndrome, patent ductus arteriosus, coarctation of aorta, pulmonary artery anomalies, insufficiency of the aortic valve, atrial septal defect- not otherwise specified (a subset of ostium secundum atrial septal defect), perimembranous ventricular septal defect (a subset of ventricular septal defect), muscular ventricular septal defect (a subset of ventricular septal defect), single common atrium (a subset of endocardial cushion defects), and patent foramen ovale (a subset of ostium secundum atrial septal defect). Cases of patent ductus arteriosus and patent foramen ovale were excluded if the infant was preterm (<36 weeks, 4% of patent ductus arteriosus and patent foramen ovale cases). Infants with more than one birth defect were categorized into multiple defect groups unless diagnoses were similar (eTable 2). Because the majority of cases had multiple birth defect diagnosis (74%), to assess if there was a difference between infants with multiple defects and isolated defects, these two groups were analyzed separately.

For neural tube defects, spina bifida was the most common defect and was analyzed separately; all other neural tube defects were analyzed together due to small numbers. Anencephaly was excluded (13% of neural tube defects) as those included in the registry may not be representative of all anencephaly cases due to early termination. Orofacial defects were divided into two categories: cleft lip with or without palate and cleft palate only.

## 2.2 Exposure assessment

Our primary analysis examined the relationship between birth defects and PM<sub>2.5</sub> exposures modeled using satellite remote sensing, meteorological and land use data. Aerosol optical depth (AOD) is the integral of particle light extinction coefficients from the surface to the top of the atmosphere. It is a measure of the degree to which aerosols prevent light from penetrating the atmosphere and retrieved using wavelengths most sensitive to particles with sizes from 0.1 to 2 μm (30). Thus, AOD is related to the loadings of fine particles in the atmosphere and is a strong predictor of ground-level PM<sub>2.5</sub> concentrations as most fine particles are emitted and confined in the boundary layer. The number of stationary ambient

monitors is limited and the distribution is sparse, while AOD-estimated PM<sub>2.5</sub> concentrations have the potential to expand the spatiotemporal coverage of ground networks and improve the accuracy of estimates of personal exposure to PM<sub>2.5</sub> (31). The Geostationary Operational Environmental Satellite (GOES) is the major weather satellite operated by the National Oceanic and Atmospheric Administration (NOAA). GOES provides an aerosol and smoke product (GASP) with AOD retrievals every 30 minutes from sunrise to sunset at 4 km nominal spatial resolution. We obtained GASP AOD data from the NOAA National Environmental Satellite, Data, and Information Service. In this study, AOD measurements (available from 9 am to 3 pm local time) were averaged to generate daily AOD estimates (24).

The 24-hour average PM<sub>2.5</sub> concentrations from 2001 to 2008 collected from 35 U.S. Environmental Protection Agency (EPA) Federal Reference Monitors (FRM) were downloaded from the EPA’s Air Quality System Technology Transfer Network (<http://www.epa.gov/ttn/airs/airsaqs/>). Meteorological fields, including temperature and wind speed, were provided by the North American Land Data Assimilation System (NLDAS) Phase 2 and downloaded from the NLDAS website (<http://ldas.gsfc.nasa.gov/nldas/>). Elevation data were obtained from the National Elevation Dataset (NED) (<http://ned.usgs.gov>). Major roads were extracted from ESRI StreetMap USA (Environmental Systems Research Institute, Inc., Redland, CA). Forest cover data were derived from 2001 and 2006 land cover maps downloaded from the National Land Cover Database (NLCD) (<http://www.mrlc.gov>). Primary PM<sub>2.5</sub> emissions were obtained from the 2002, 2005, and 2008 EPA National Emission Inventory (NEI) facility emissions reports. We developed a linear mixed effects model with 24-hour average EPA PM<sub>2.5</sub> measurements from 2001 to 2008 as the dependent variable and AOD, meteorological fields and land use variables as predictors. The model incorporates day-specific random intercepts and slopes for AOD, temperature, and wind speed to account for the temporally varying relationship between PM<sub>2.5</sub> (based on fixed ground monitors) and AOD. (32). This model was run annually for a 4 km modeling grid covering the spatial extent of Massachusetts to estimate daily PM<sub>2.5</sub> concentrations from 2001 to 2008. The model structure can be expressed as

$$PM_{2.5,st} = (\beta_0 + \beta_{0,t}) + (\beta_1 + \beta_{1,t}) AOD_{st} + (\beta_2 + \beta_{2,t}) Temperature_{st} + (\beta_3 + \beta_{3,t}) Wind Speed_{st} + \beta_3 Elevation_s + \beta_4 Major Roads_s + \beta_5 Forest Cover_s + \beta_6 Point Emissions_s + \varepsilon_{st}$$

$$(\beta_{0,t}, \beta_{1,t}, \beta_{2,t}, \beta_{3,t}) \sim N[(0, 0, 0, 0), \Psi]$$

where  $PM_{2.5,st}$  is the measured ground level PM<sub>2.5</sub> concentration (µg/m<sup>3</sup>) at site  $s$  in day  $t$ ;  $\beta_0$  and  $\beta_{0,t}$  (day-specific) are the fixed and random intercept, respectively;  $AOD_{st}$  is the GASP AOD value (unitless) at site  $s$  in day  $t$ ;  $\beta_1$  and  $\beta_{1,t}$  (day-specific) are the fixed and random slopes for AOD, respectively;  $Temperature_{st}$  is the air temperature (K) at site  $s$  in day  $t$ ;  $\beta_2$  and  $\beta_{2,t}$  (day-specific) are the fixed and random slopes for temperature, respectively;  $Wind Speed_{st}$  is the 2 meters above ground wind speed (m/sec) at site  $s$  in day  $t$ ;  $\beta_3$  and  $\beta_{3,t}$  (day-specific) are the fixed and random slopes for wind speed, respectively;  $Elevation_s$  is elevation values (m) at site  $s$ ;  $Major Roads_s$  is road length (m) at site  $s$ ;  $Forest Cover_s$  is percentage of forest cover (unitless) at site  $s$ ;  $Point Emissions_s$  is point emissions (tons per year) at site  $s$ ; and  $\Psi$  is an unstructured variance-covariance matrix for the random

effects. A ten-fold cross validation (CV) was conducted to evaluate the performance of the model, and statistical indicators including the coefficient of determination ( $R^2$ ) and square root of the mean squared prediction errors (RMSPE) were calculated between predicted  $PM_{2.5}$  concentrations and observations. The results show that CV RMSPE ranges from 2.42 to 3.50  $\mu\text{g}/\text{m}^3$ , and CV  $R^2$  ranges from 0.78 to 0.88 for years 2001 through 2008, indicating a good performance of the model.

The clinically estimated gestational age of infants was subtracted from date of birth to calculate conception date. The exposure assessment was performed for varying gestational weeks depending on the outcome of interest. For cardiac, neural tube, and orofacial defects, the windows of exposure are weeks 3–7, 1–4, and 6–12 of pregnancy, respectively, as these periods are considered to be the most critical exposure window for the development of the specified birth defects (33–35). Daily  $PM_{2.5}$  estimates were averaged for each exposure assessment interval. Infants were assigned a  $PM_{2.5}$  exposure measure if there was at least a single daily  $PM_{2.5}$  estimate for each week of the relevant gestational period. Satellite measures provide extensive spatial coverage throughout Massachusetts allowing us to assign exposure estimates to 95% of Massachusetts births included in our study. As a sensitivity analysis, we also assessed the effects of average daily  $PM_{2.5}$  over the first trimester of pregnancy.

To understand the influence of local traffic-related pollution on a near-roadway spatial scale, we considered the role of distance to major roadways and traffic density near birth residence (36–39). Using geographic information system software (ArcGIS, version 10.0; ESRI), we calculated the shortest distance (m) between each maternal address at birth and the nearest Class 1 (limited access highways) or 2 (multilane highways without limited access) road segment, and traffic density was calculated by summing the annual average daily traffic (AADT) for all Class 1 and 2 road segments within a 200 meter grid of Class 1 and 2 road segments (38), as air pollution from traffic reaches background levels around 200 meters (37). To estimate density for road segments with missing AADT counts, the AADT from the segments nearest to the missing segment of the same class with available data was used.

### 2.3 Statistical analysis

We used logistic regression models to calculate crude and adjusted odds ratios (ORs) and 95% confidence intervals (CI) for each birth defect outcome group and  $PM_{2.5}$  exposure. Exposure-response relations were also investigated using cubic splines. We considered adjustment for the following covariates obtained from Massachusetts birth records: plurality, maternal race/ethnicity, maternal education, maternal language preference, delivery payment source, smoking during pregnancy, alcohol consumption during pregnancy, adequacy of prenatal care (measured by the Adequacy of Prenatal Care Utilization Index), marital status, and maternal age. Infant covariates included birth year, parity, and season of conception. We used geocoded addresses to determine the median household income and median home value of census block groups. Analyses that included local traffic measures (traffic density and distance to roadways) were modeled continuously using cubic spline models. For each defect group, we used the change of estimate criterion (10%) to determine which covariates would be included in the model (40). We excluded observations with missing covariates in

our analysis. Although the percent of missing information was less than 5% for each variable of interest, we applied multivariate imputation for variables with missing values using the predictive mean matching, logistic regression, and polytomous logistic regression imputation method for continuous, binary, and categorical variables, respectively, to evaluate the influence of missing data on our estimates. We also assessed effect modification of PM<sub>2.5</sub> exposure by maternal education. All analyses were conducted using R (R Software Version 3.0.3). The R packages mgcv and mice were used for the cubic spline models and multiple imputations, respectively.

We examined spatial patterns of birth defects using generalized additive models (GAMs) for each outcome group (41, 42) to determine the residual influence of geographic location after accounting for PM<sub>2.5</sub> as a confounder and potential mediator. The optimal amount of smoothing was determined by minimizing the Akaike Information Criterion and permutations were used to test for the significance of the smooth term for location. The underlying spatial pattern of the defect was first assessed using an unadjusted model with only the smooth term for location. Two adjusted GAMS were also modeled. The first was adjusted for potential confounders including maternal education, age, race, adequacy of prenatal care, and season of conception. The second included all the indicated covariates from the first model with the addition of satellite-based PM<sub>2.5</sub> measures to assess residual spatial patterns after accounting for the contribution of PM<sub>2.5</sub> and other risk factors. The R package MapGAM was used to run the spatial GAMs and create the maps.

### 3. Results

We obtained records for 2,729 cardiac, 726 orofacial, 255 neural tube defect cases, and 7,816 controls geocoded, non-syndromic births from the 611,854 births in Massachusetts conceived from January 1, 2001 to December 31, 2008. The percentage of cases with mothers 30–34 years old ranged from 27–31%, depending on the defect. Between 61–72% of mothers were of non-Hispanic White race/ethnicity, and between 56–74% of mothers reported education levels greater than high school (Table 1). Eight percent of control mothers reported smoking during pregnancy, while up to 10% of mothers with orofacial defect reported smoking. The proportion of mothers that reported drinking during pregnancy ranged from 0.5–2% among the defect groups. Approximately 75% of all cases reported “adequate” prenatal care.

#### 3.1 PM<sub>2.5</sub> and traffic-related exposures

Our PM<sub>2.5</sub> analysis included 2,610 (95.6%) cardiac, 692 (95.3%) orofacial, and 247 (96.8%) neural tube defects with a measure of prenatal PM<sub>2.5</sub> exposure for the relevant gestational period. Of the 7,816 controls, 278 (4%), 352 (5%), and 254 (3%) were excluded from the cardiac, orofacial, and neural tube analyses, respectively, due to missing PM<sub>2.5</sub> data during their respective critical window of exposure. Traffic density and residential distance to nearest major roadway were successfully calculated for all cases and controls.

Cubic splines relating continuous satellite-based PM<sub>2.5</sub> exposure to log odds of birth defects were approximately linear. As such, we report results from logistic regression models of the association between birth defects and PM<sub>2.5</sub> exposure modeled continuously for a 10 µg/m<sup>3</sup>

increase of  $PM_{2.5}$  (Table 2). Non-imputed results are presented since estimates were not altered when missing values were imputed. The adjusted ORs for perimembranous ventricular septal defects (OR = 1.34, 95% CI: 0.98, 1.83), patent foramen ovale (OR = 1.18, 95% CI: 0.91, 1.53) and patent ductus arteriosus (OR = 1.24, 95% CI: 0.94, 1.62) were elevated and approaching significance. Odds ratios for the remaining cardiac defects were generally null with wide confidence intervals. Results for the general neural tube defects group suggested an inverse association (OR = 0.70, 95% CI: 0.46, 1.05) whereas for spina bifida the OR was slightly elevated (OR = 1.22, 95% CI: 0.61, 2.30), although neither was statistically significant. Non-significant inverse associations were also observed for orofacial defects (cleft lip with or without palate: OR = 0.76, 95% CI: 0.50, 1.10; cleft palate: OR = 0.89, 95% CI: 0.54, 1.46). We found that crude estimates were similar to adjusted estimates for all defects except for endocardial cushion and all orofacial defects (Table 2). When assessing a wider exposure window of the first trimester of pregnancy, we found similar results for most defects (eTable3). Compared to estimates using narrow exposure windows, estimates obtained using first trimester exposure windows were closer to the null for patent ductus arteriosus and atrial septal defect and further away from the null for tetralogy of fallot and endocardial cushion defect. We found evidence for effect modification of  $PM_{2.5}$  exposure by maternal education level for endocardial cushion defects ( $P = 0.017$ ), perimembranous VSD ( $P = 0.030$ ), and single common atrium ( $P = 0.020$ ) (eTable 4).

For comparison, we also examined the relationship between birth defects and local traffic measures. Results for residential proximity to major roadways, modeled continuously with cubic splines are shown in Figure 1 for defects that yielded significant associations (see eResults1 for further details). Cubic splines indicated that traffic density and risk of defects was approximately linear, therefore odds ratios are presented for an interquartile range increase of AADT (eTable 5). Results show similar associations with traffic density as many of the same defects associated with  $PM_{2.5}$  and distance to major roadways. Ostium secundum atrial septal defects (OR = 1.03, 95% CI: 1.00, 1.06), patent foramen ovale (OR = 1.05, 95% CI: 1.01, 1.08), and insufficiency of the aortic valve (OR = 1.07, 95% CI: 1.01, 1.12) were positively associated with traffic density. Patent ductus arteriosus was positively associated with traffic density and approaching significance (OR = 1.03, 95% CI: 0.99, 1.07). Cleft lip with or without palate was negatively associated with traffic density (OR = 0.92, 95% CI: 0.85, 0.98). Results were similar for imputed datasets (data not shown).

### 3.2 Geographic location

Spatial analyses using GAMs showed significant associations between geographic location and certain birth defects (eTable 6). Birth location remained statistically significant ( $P = 0.004$ ) for ostium secundum atrial septal defects after adjusting for demographic, socioeconomic, behavioral risk factors, and  $PM_{2.5}$  exposure (Figure 2). The relationship between birth location and hypoplastic left heart syndrome was borderline statistically significant ( $P = 0.067$ ) and was fully explained only after adding  $PM_{2.5}$  to the model ( $P = 0.144$ , Figure 2). To determine the influence of missing data on outcomes in the spatial analysis, five imputed data sets were generated to run the GAMs and permutation tests for hypoplastic left heart syndrome and ostium secundum atrial septal defect. Results were similar to those generated using the original dataset (data not shown).

Although unadjusted models indicated significant spatial variation across Massachusetts for other defects, adjusted models suggested that the patterns were due to socioeconomic, demographic, and behavioral risk factors, and including PM<sub>2.5</sub> did not alter the geographic pattern.

#### 4. Discussion

We examined the spatial relationship between PM<sub>2.5</sub> and other traffic-related measures using anatomical groupings of cardiac, neural tube and orofacial birth defects. There is evidence to support the hypothesis that exposure to PM<sub>2.5</sub> and traffic-related air pollution increases risk of patent foramen ovale and patent ductus arteriosus, as these defects displayed a positive but non-significant association with PM<sub>2.5</sub>. Patent foramen ovale displayed positive significant associations with both distance to major roadways and traffic density.

We assessed traffic-related air pollution at various spatial scales, as our PM<sub>2.5</sub> measures represent larger scale pollution whereas traffic density and distance to major roads represent more local measures of pollution. Perimembranous ventricular septal defects demonstrated a positive but non-significant association with PM<sub>2.5</sub> exposure but not with local measures of pollution. Ostium secundum atrial septal defects and insufficiency of the aortic valve defects were significantly associated with local measures of air pollution, but not PM<sub>2.5</sub>, indicating that there may be a specific local pollutant affecting the development of these defects.

There was a noticeable rise in risk for ostium secundum type atrial septal defect, patent ductus arteriosus, patent foramen ovale, insufficiency of the aortic valve, and cleft lip with or without palate among infants with birth addresses around 1000–1500 meters away from a major roadway (Figure 1). It is not clear that the rise and decline in risk around 1000–1500 meters is statistically significant as it is contained within the CI at shorter distances for most outcomes. It is therefore difficult to tell if the higher risk at that distance is meaningful or an effect of random variation for most birth defects. However, the CIs are narrower for insufficiency of the aortic valve, suggesting that the pattern is meaningful for that outcome. Moreover, the shared pattern of increased risk estimates around 1000–1500 meters suggests that some unmeasured risk factor for birth defects may be more common in that range. We examined the confounders of infants whose birth address was between 1000–1500 meters away from the nearest major roadway and found they were similar to births residing at other distances. Further investigation is needed to better understand the shared pattern of increased risk estimates around 1000–1500 meters.

Interestingly, we observed inverse associations between PM<sub>2.5</sub> and traffic density and cleft lip with or without palate, although the association with PM<sub>2.5</sub> was not significant. Other studies have found no association between cleft lip with or without palate and PM<sub>2.5</sub> (16, 18), yielding null inverse associations. One study also found a consistent, but non-significant inverse association with PM<sub>2.5</sub> and cleft lip with or without palate across all PM<sub>2.5</sub> quartiles (17).



In our secondary analysis of effect modification, we found that women with the lowest education level (less than high school) had the highest effect estimate compared to women with at least high school or college education for a majority of birth defects. Although there is some evidence for effect modification, this stratified analysis contains small numbers of cases, especially in the less than high school category, and therefore the effect estimates have wide confidence intervals.

The mapped results of our crude GAMs indicated there was a statistically significant association between geographic location and certain birth defects. Residential location remained significantly associated with ostium secundum atrial septal defects in the fully adjusted spatial model, suggesting that the persistent areas of increased risk are not fully explained by individual risk factors included in the model. These spatial patterns may be due to unidentified environmental or social determinants. Adjusting for PM<sub>2.5</sub> exposure influenced the spatial patterns for hypoplastic left heart syndrome only, suggesting that spatial patterns of increased risk for all other defects are not strongly associated with PM<sub>2.5</sub>. Furthermore, despite analyzing eight years of births for an entire state, rare birth defects are prone to small case numbers, limiting the power to detect significant spatial associations.

We investigated three groups of atrial septal defects; ostium secundum atrial septal defects, atrial septal defects, NOS, and patent foramen ovale. To our knowledge no other studies have examined the influence of PM<sub>2.5</sub> exposure on risk of ostium secundum atrial septal defects. We did not find any studies that investigated the role of patent foramen ovale and air pollutants. We found an overall positive significant association between atrial septal defects, NOS and residential distance to major roadways ( $p=0.030$ ) and increased risk with PM<sub>2.5</sub> (OR= 1.23, 95% CI: 0.78, 1.90). Gilboa et al. also found a positive significant association between PM<sub>10</sub> and atrial septal defects in Texas (5). We found a non-significant positive association between patent ductus arteriosus and PM<sub>2.5</sub> while Agay-Shay et al. found a significant inverse association with isolated patent ductus arteriosus and PM<sub>2.5</sub> (11). Another study (43) found significant increased odds of patent ductus arteriosus and PM<sub>10</sub> while utilizing further restrictive criteria for this outcome group based on active surveillance records, but we were unable to replicate this method because of limited information from the Massachusetts birth defects registry. Although there are differences in the composition of PM<sub>10</sub> and PM<sub>2.5</sub>, this may indicate that general particulate matter exposure may be associated with risk of patent ductus arteriosus. We did not find strong evidence supporting an association between neural tube defects and traffic-related measures of air pollution. Similarly, Padula et al. found no association with PM<sub>2.5</sub> or traffic density and neural tube defects (17).

To better understand the association of PM<sub>2.5</sub> and traffic-related exposure on birth defects, our study includes improved exposure assessment methods, standardized definitions of cases, and important confounders. We believe we have captured several important SES measures, but we are less confident about our ability to capture alcohol and smoking behaviors from self-reported measures. We used satellite data to obtain fine spatial distribution estimates of PM<sub>2.5</sub> during pregnancy for all of Massachusetts. The use of measured PM<sub>2.5</sub> for each 4 km square grid cell allows for more fine-scale measures of exposure compared to previous studies that have relied on measures from stationary

monitoring stations. We believe that using this finer spatial resolution of PM<sub>2.5</sub> reduced residential exposure misclassification compared to using monitoring stations alone. By including distance to major roadways and traffic density, we were able to assess the influence of local road networks on risk of birth defects.

Although we were able to control for important covariates, we were unable to directly control for folic acid supplementation, a major contributor of neural tube defects (44). As a proxy for folic acid supplementation we used the adequacy of prenatal care index. This index measures prenatal care based on initiation and adequacy of received services. It has been found that adequacy of prenatal care is a useful proxy for folic acid supplementation during pregnancy (45).

We utilized critical windows of exposure specific to birth defect anatomical location. Previous studies have used average exposures over the first trimester of pregnancy or gestational weeks 3–8. Organ development is time specific, therefore the use of very specific exposure windows can strongly influence the effect estimates and allow for evaluation of exposure-response relationships to be more readily observed (46, 47). We noticed variability in effect estimates when assessing exposure during the first trimester of pregnancy compared to narrow critical windows of exposure for transposition of the great vessel, tetralogy of fallot, endocardial cushion defect, patent ductus arteriosus, coarctation of the aorta, pulmonary artery anomalies, atrial septal defects, NOS, and spina bifida (eTable 3).

This analysis includes birth defect diagnoses among still births and diagnoses made up to one year of life. However, we are unable to more fully investigate the effects of fetal toxicity as we do not have records of fetal deaths. Fetal deaths may be more sensitive to the effects of air pollution than those who survive to birth with birth defects. Therefore, we may not have captured the conceptions most at risk. We hypothesize that the detected association would be weaker than the true association had all at risk subjects been included in this analysis.

This study only assessed the relationship between birth defects and traffic-related exposures using residential location of mother at time of delivery. Pregnant women are a mobile population and many women may not have been living in the same home during the early prenatal period as they were at the time of birth of the infant. We believe that this may lead to exposure misclassification but it is likely to be non-differential among cases and controls (48). Residential mobility was estimated to be between 3% and 14% in recent American studies based on birth cohorts (48, 49) and mobility rates did not differ among mothers of infants born with birth defects and mothers of infants born without birth defects (48). Additionally, maternal time-activity patterns were not accounted for in our investigation. This may result in exposure misclassification, although it is also likely to be non-differential among cases and controls. Recent studies have found that outdoor residential levels of exposure act as a good surrogate for personal exposure (50). Both sources of exposure misclassification may increase type 2 error so our null results should be interpreted with caution.

Given the multiple comparisons involved in testing a large range of birth defects with three different metrics for traffic-related air pollution, this analysis may result in significant associations due to chance (e.g., type 1 errors). Therefore we have emphasized associations that are observed for more than one measure of traffic-related air pollution. Furthermore, we acknowledge that PM<sub>2.5</sub> is a complex mixture of particles with varying toxicity and we do not have measures of composition.

## 5. Conclusions

In summary, we found evidence to suggest that PM<sub>2.5</sub> exposure during pregnancy may be associated with risk of patent foramen ovale, patent ductus arteriosus, and perimembranous ventricular septal defects. Our findings also support a possible relationship between ostium secundum atrial septal defects and insufficiency of the aortic valve with local traffic-related air pollutants. Our spatial analyses show that there were geographic regions with increased risk of ostium secundum atrial septal defects in Massachusetts, even after accounting for PM<sub>2.5</sub> exposure. There is limited evidence to suggest that cleft lip with or without palate may have an inverse association with PM<sub>2.5</sub> exposure and traffic-related air pollution in general. To fully understand the influence of PM<sub>2.5</sub> on birth defects, larger studies of these specific birth defect groups using valid PM<sub>2.5</sub> estimates are needed.

## Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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## Abbreviations

<b>AADT</b>	annual average daily traffic
<b>AOD</b>	aerosol optical depth
<b>CI</b>	confidence interval
<b>GAMs</b>	generalized additive models
<b>GOES</b>	Geostationary Operational Environmental Satellite
<b>ICD-9-CM</b>	International Classification of Diseases, Clinical Modification, Ninth Revision
<b>NOS</b>	not otherwise specified
<b>NOAA</b>	National Oceanic and Atmospheric Administration
<b>OR</b>	odds ratio
<b>PM<sub>2.5</sub></b>	particulate matter with diameter of 2.5 μm or less

SES socioeconomic status

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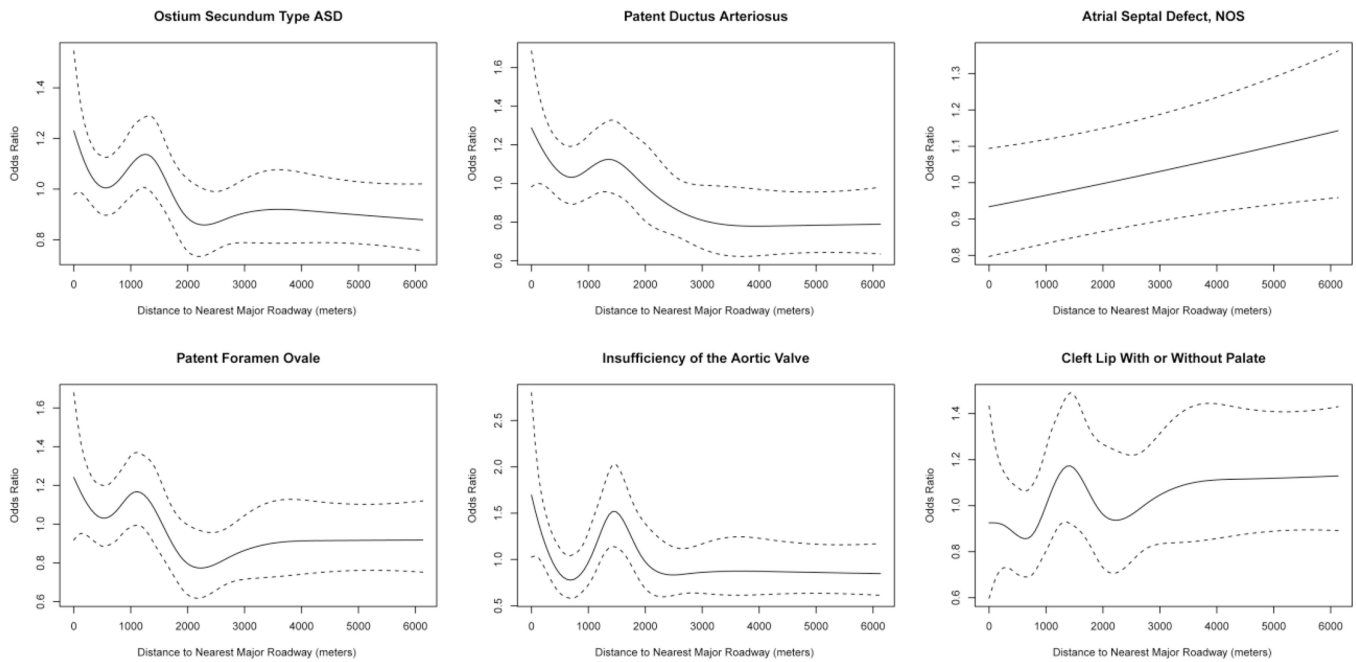
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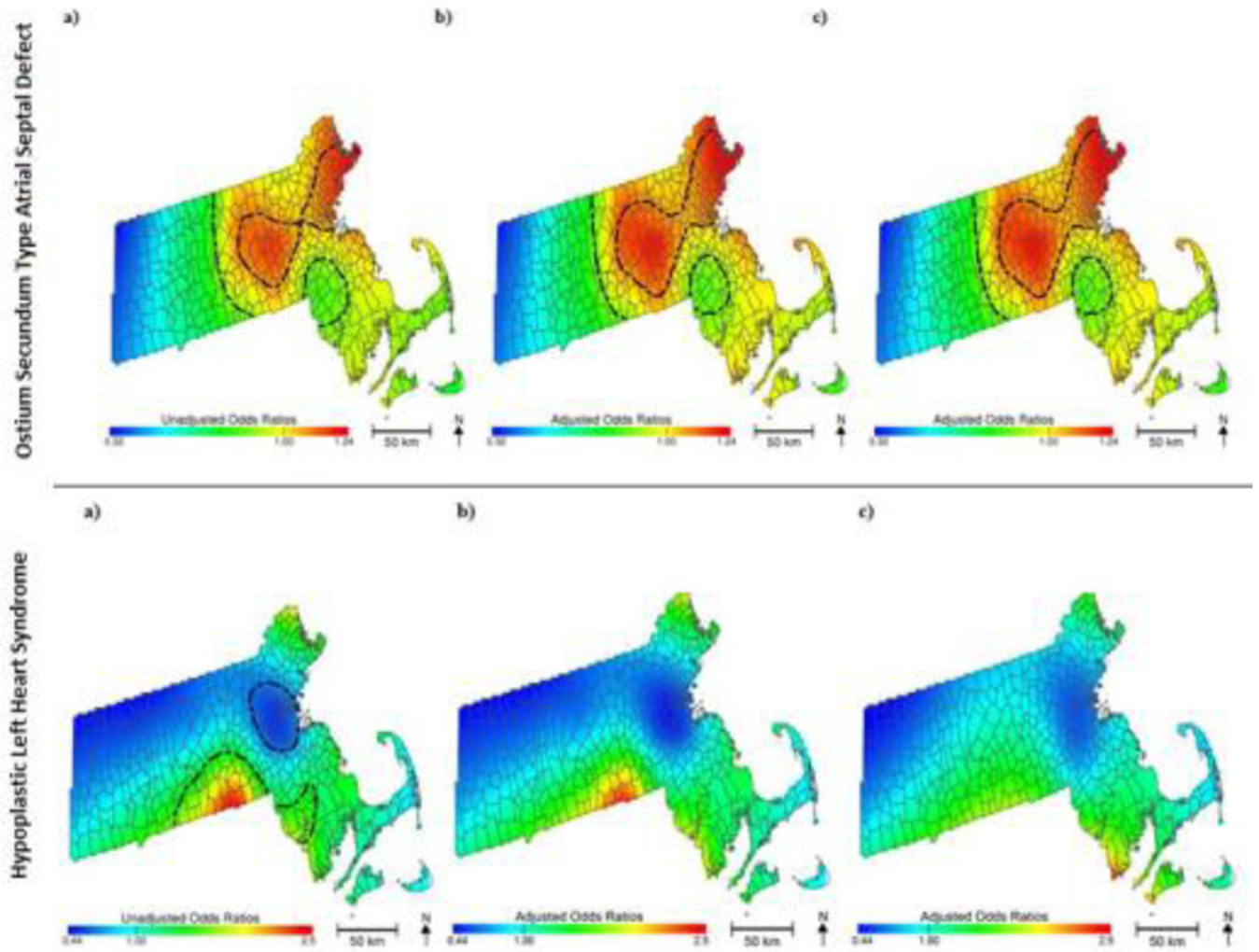
**Highlights**

- Improved PM<sub>2,5</sub> satellite measures enable full advantage of a statewide birth cohort.
- Some evidence of association of patent foramen ovale and patent ductus arteriosus with traffic related air pollution.
- Limited evidence of association of neural tube defects with traffic related air pollution.
- Limited evidence of association of orofacial defects with traffic related air pollution.
- Atrial septal defects display significant spatial variation.



**Figure 1.** Association of Residential Distance to Major Roadways and Risk of Defects. Exposure response curve showing the adjusted odds ratio (solid line) and the 95% confidence interval (dashed line) using cubic splines to model the association of residential distance to a major roadway and birth defects among infants conceived in Massachusetts, 2001–2008. Only birth defects with significant associations are displayed. Estimates are only presented for residential addresses within continental Massachusetts. All models adjusted for maternal race, education, median household income of block group, alcohol consumption during pregnancy, and plurality. Cardiac defects further adjusted for maternal age, language preference, parity, adequacy of prenatal care. Orofacial defects further adjusted for season of conception, infant sex, adequacy of prenatal care, and smoking during pregnancy. Major roadways defined as limited access highways and multi-lane highways (class 1 and 2 roads). Abbreviations: ASD – atrial septal defect; NOS - not otherwise specified.





**Figure 2.** Geographic Patterns of Ostium Secundum Atrial Septal Defect and Hypoplastic Left Heart Syndrome Risk among infants conceived in Massachusetts, 2001–2008: (a) Unadjusted, (b) Adjusted without  $PM_{2.5}$  Exposure, and (c) Adjusted with  $PM_{2.5}$  Exposure. Statistically significant geographic areas of increased or decreased risk of birth defect are indicated using black contour lines. Maps adjusted for maternal education, race, smoking during pregnancy, alcohol consumption during pregnancy, prenatal care, infant sex, and season of conception.

**Table 1**

Description of Birth Defect Cases and Randomly Selected Controls in Massachusetts, Conceived 2001–2008

n(%) <sup>a</sup>	Controls (n=7,816)	Cardiac (n=2,729)	Orofacial (n=726)	Neural Tube (n=255)
<b>Infant Sex</b>				
Male	4,032 (51.6)	1,412 (51.7)	411 (56.7)	107 (42.0)
Female	3,784 (48.4)	1,317 (48.3)	314 (43.3)	148 (58.0)
<b>Maternal Age</b>				
<20 years	462 (5.9)	148 (5.4)	46 (6.3)	19 (7.4)
21–24 years	1,227 (15.7)	417 (15.2)	118 (16.2)	38 (14.9)
25–29 years	1,878 (24.0)	586 (21.4)	181 (24.9)	71 (27.8)
30–34 years	2,418 (30.9)	853 (31.2)	211 (29.1)	70 (27.4)
35+ years	1,831 (23.4)	725 (26.5)	169 (23.3)	57 (22.3)
<b>Parity</b>				
0	3,521 (45.0)	1,225 (44.8)	320 (44.1)	123 (48.2)
1	2,665 (34.1)	884 (32.3)	247 (34.0)	84 (32.9)
2	1,612 (20.6)	617 (22.6)	157 (21.6)	48 (18.8)
<b>Adequacy of Prenatal Care</b>				
Adequate	6,098 (78.0)	2,121 (77.7)	529 (72.9)	190 (74.5)
Intermediate	1,372 (17.5)	462 (16.9)	149 (20.5)	45 (17.6)
Inadequate	247 (3.1)	103 (3.7)	36 (4.9)	14 (5.4)
Unknown	74 (0.9.0)	34 (1.2)	10 (1.3)	5 (1.9)
None	25 (0.3)	9 (0.3)	1 (0.1)	1 (0.3)
<b>Smoking During Pregnancy</b>				
Yes	623 (7.9)	205 (7.5)	74 (10.2)	22 (8.6)
No	7193 (92.0)	2524 (92.4)	651 (89.7)	233 (91.3)
<b>Drinking During Pregnancy</b>				
Yes	137 (1.7)	36 (1.3)	11 (1.5)	1 (0.4)
No	7,674 (98.1)	2,691 (98.6)	713 (98.3)	254 (99.6)
<b>Season of Conception</b>				
Winter	1,883 (24.0)	646 (23.6)	158 (21.7)	55 (21.5)
Spring	1,869 (23.1)	651 (23.8)	209 (28.8)	62 (24.3)
Summer	2,008 (25.6)	657 (24.0)	151 (20.8)	63 (24.7)
Fall	2,056 (26.3)	775 (28.4)	207 (28.5)	75 (29.4)
<b>Gestational Age</b>				
<37 weeks	692 (8.8)	646 (23.7)	99 (13.6)	66 (25.8)
37 weeks	7,124 (91.1)	2,083 (76.3)	626 (86.4)	189 (74.1)
<b>Small for Gestational Age</b>				
Yes	845 (10.8)	651 (23.8)	153 (21.1)	95 (37.2)
No	6,961 (89.1)	2,076 (76.0)	572 (78.9)	159 (62.3)

n(%) <sup>a</sup>	Controls (n=7,816)	Cardiac (n=2,729)	Orofacial (n=726)	Neural Tube (n=255)
<b>Maternal Race/Ethnicity</b>				
Non-Hispanic White	5,379(68.8)	1,837 (67.3)	522 (72.0)	156 (61.1)
Non-Hispanic Black	682(8.7)	284 (10.4)	35 (4.8)	33 (12.9)
Hispanic	1,059(13.5)	375 (13.7)	102 (14.0)	46 (18.0)
Asian/Pacific Islander	530 (6.7)	157 (5.7)	49 (6.7)	11 (4.3)
Other	162 (2.0)	75 (2.7)	17 (2.3)	9 (3.5)
<b>Maternal Education</b>				
<12 <sup>th</sup> grade	793 (10.1)	302 (11.1)	86 (11.8)	30 (11.8)
High school graduation	2,010 (25.7)	710 (26.0)	2,15 (29.6)	81 (31.8)
Some college	5,002 (64.0)	1,712 (62.7)	424 (58.4)	144 (56.4)
<b>Maternal Language Preference</b>				
English	6,945 (88.8)	2,407 (88.2)	642 (88.5)	209 (81.9)
Spanish	400 (5.1)	164 (6.0)	39 (5.3)	22 (8.6)
Portuguese	181 (2.3)	59 (2.1)	15 (2.0)	10 (3.9)
Other	275 (3.5)	86 (3.1)	25 (3.4)	13 (5.1)
<b>Household Income</b>				
<\$20,000	339 (4.4)	126 (4.6)	32 (4.4)	11 (4.3)
\$20,000–\$69,999	3899 (49.8)	1431 (52.4)	386 (39.4)	158 (62.0)
\$70,000	3578 (45.8)	1172 (42.9)	307 (42.3)	86 (33.7)
<b>Delivery Source of Payment</b>				
HMO	1,365 (17.4)	437 (16.0)	125 (17.2)	51 (20.0)
Medicaid/Common Health	1,988 (25.4)	741 (27.1)	207 (28.5)	79 (30.9)
Other	4,463 (57.1)	1,550 (56.8)	392 (54.0)	125 (49.0)

<sup>a</sup> Percentages may not sum to 100% due to rounding.

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**Table 2**

Crude and Adjusted Odds Ratios<sup>a</sup> and 95% Confidence Interval for Prenatal Exposure to PM<sub>2.5</sub> in Massachusetts, and Cardiac, Neural Tube, and Orofacial Defects Among Infants Conceived Between 2001–2008.

Defect Type <sup>b</sup>	n	Crude OR(95% CI) <sup>c</sup>	Adjusted OR (95% CI) <sup>c</sup>
Isolated Birth Defects	890	0.92 (0.71, 1.21)	0.97 (0.94, 1.00)
Multiple Birth Defects	2571	1.01 (0.84, 1.20)	1.01 (0.85, 1.21)
<b>Total Cases</b>	3,461		
<i>Cardiac</i>			
Transposition of the Great Vessel	233	0.90 (0.59, 1.37)	0.92 (0.60, 1.41)
Tetralogy of Fallot	153	0.99 (0.58, 1.68)	1.00 (0.59, 1.71)
Ostium Secundum ASD	1457	0.96 (0.80, 1.16)	0.98 (0.81, 1.19)
Endocardial Cushion Defect	139	0.99 (0.58, 1.68)	1.18 (0.67, 2.09)
Pulmonary Valve Atresia/Stenosis	436	1.01 (0.73, 1.39)	1.05 (0.76, 1.45)
Aortic Valve Stenosis	93	1.15 (0.58, 2.30)	1.18 (0.58, 2.38)
Hypoplastic Left Heart Syndrome	69	0.70 (0.34, 1.45)	0.73 (0.35, 1.42)
Patent Ductus Arteriosus	675	1.24 (0.95, 1.62)	1.24 (0.94, 1.62)
Coarctation of Aorta	205	1.03 (0.65, 1.63)	1.03 (0.65, 1.64)
Pulmonary Artery Anomalies	172	1.10 (0.67, 1.83)	1.04 (0.64, 1.68)
VSD	864	1.08 (0.86, 1.37)	1.09 (0.86, 1.37)
Perimembranous VSD	494	1.32 (0.98, 1.81)	1.34 (0.98, 1.83)
Muscular VSD	328	0.87 (0.61, 1.25)	0.89 (0.62, 1.27)
Single Common Atrium, Cor Trioculare	335	1.12 (0.78, 1.62)	1.19 (0.82, 1.72)
Atrial Septal Defect, NOS	235	1.24 (0.80, 1.94)	1.23 (0.78, 1.90)
Patent Foramen Ovale	725	1.15 (0.89, 1.50)	1.18 (0.91, 1.53)
Insufficiency of Aortic Valve	262	1.11 (0.73, 1.68)	1.16 (0.76, 1.76)
<i>Neural Tube</i>			
Neural Tube Defects	199	0.72 (0.47, 1.09)	0.77 (0.46, 1.05)
Spina Bifida	89	1.22 (0.63, 2.37)	1.18 (0.61, 2.30)
<i>Orofacial</i>			
Cleft Lip with & without Palate	406	1.02 (0.71, 1.47)	0.76 (0.50, 1.10)
Cleft Palate Only	251	1.24 (0.78, 1.95)	0.89 (0.54, 1.46)

Abbreviations: ASD, atrial septal defect; CI, confidence interval; NOS, not otherwise specified; OR, odds ratio; PM<sub>2.5</sub>, particulate matter with a diameter of 2.5 μm or less; VSD, ventricular septal defect.

<sup>a</sup>Control group n=7538,7464,7562 for cardiac, neural tube, and orofacial defect analysis, respectively.

<sup>b</sup>All models adjusted for maternal race, education, median household income of block group, alcohol consumption during pregnancy, and plurality. Isolated and multiple birth defects further adjusted for maternal age, language preference, parity, and adequacy of prenatal care. Cardiac defects further adjusted for maternal age, language preference, parity, and adequacy of prenatal care. Neural tube defects further adjusted for maternal age, language preference, parity, adequacy of prenatal care, and smoking during pregnancy. Orofacial defects further adjusted for season of conception, infant sex, adequacy of prenatal care, and smoking during pregnancy.

<sup>c</sup>ORs and 95% CIs correspond to a 10  $\mu\text{g}/\text{m}^3$  increase of average  $\text{PM}_{2.5}$  measured at the 4 km grid cell of infant's birth address during weeks 3–7, 1–4, and 6–12 of pregnancy for cardiac, neural tube, and orofacial defects, respectively.

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