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# Palliative Care Teams' Cost-Saving Effect Is Larger For Cancer Patients With Higher Numbers Of Comorbidities

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# Abstract

Patients with multiple serious conditions account for a high proportion of health care spending. Such spending is projected to continue to grow substantially because of increased insurance eligibility, the ever-rising cost of care, the continued use of nonbeneficial high-intensity treatments at the end of life, and demographic changes. We evaluated the impact of palliative care consultation on hospital costs for adults with advanced cancer, excluding those with dementia. We found that compared to usual care, the receipt of a palliative care consultation within two days of admission was associated with 22 percent lower costs for patients with a comorbidity score of 2–3 and with 32 percent lower costs for those with a score of 4 or higher. Earlier consultation was also found to be systematically associated with a larger cost-saving effect for all subsamples defined by multimorbidity. Given ongoing workforce shortages, targeting early specialist palliative care to hospitalized patients with advanced cancer and higher numbers of serious concurrent conditions could improve care while complementing strategies to curb the growth of health spending.

Improving care for people with cancer is a US health care priority. Forty percent of Americans will develop cancer in their lifetime, and cancer remains the second leading cause of death in the United States—accounting for almost 600,000 deaths annually.[1] Multimorbidity (the presence of more than one chronic condition) is common in cancer patients because key risk factors for cancer, including aging and unhealthy behaviors such as alcohol and tobacco use, are also major risk factors for other serious chronic conditions.[2]

Patients with multimorbidity are a well-established policy priority in the United States and other high-income countries.[3, 4] Ten-year projections estimate that annual Medicare expenditures will have increased 98 percent by 2024, reaching \$1.2 trillion, and that total annual national health spending will have grown 76 percent, reaching \$5.4 trillion.[5] These estimates are strongly driven by the cost of treating patients with multiple chronic conditions.

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Two-thirds of Medicare beneficiaries have multimorbidity, and there is a strong association between the number of co-occurring conditions and cost: The 16 percent of beneficiaries with six or more chronic conditions account for 47 percent of total program expenditures.[6] The economic burden of treatment for patients with serious illnesses such as cancer and multimorbidity is projected to grow because of expanding insurance eligibility through the Affordable Care Act, demographic changes, and the limited capacity of health systems originally designed to provide acute and episodic care.[3, 7–9]

Moreover, expenditure often does not equate to value: Patients with serious illness continue to receive fragmented care of poor quality,[10] and end-of-life care is becoming more aggressive—with more use of the intensive care unit, more transitions between sites of care, and shorter hospice stays.[11] Multiple chronic conditions act synergistically to increase difficulties in finding appropriate medications and treatment regimens that work for all conditions.[12] Among patients with advanced cancer and other serious illnesses, aggressive treatments often are inconsistent with patients' preferences,[13, 14] have limited efficacy, [15, 16] and are associated with worse quality of life, compared to other treatments.[17]

Studies have demonstrated the beneficial effects on patient and family outcomes when palliative care is introduced into routine cancer care. The effects include improvements in pain and other symptoms; improved family outcomes; reduced hospital costs and readmissions; increased hospice use; and enhanced survival.[18] Palliative care is a teambased specialty (incorporating medicine, nursing, social work, and chaplaincy) focused on improving quality of life for people with serious illness such as cancer by adding a layer of support for patients, their families, and health care providers. Palliative interventions are not focused only on people at end of life and are increasingly available earlier in the care trajectory with observable benefits, and the American Society of Clinical Oncology has recommended the integration of palliative care into standard oncology care.[19]

Although multiple studies have shown that palliative care reduces average costs of care,[20] little is known about how the treatment effect of palliative care on costs varies according to diagnosis or comorbidity.[21] The effect of palliative care is likely not homogeneous and may vary according to a multifaceted interaction of individual and service factors.[22, 23]

In this article we report the effect of palliative care consultation teams—the dominant model of palliative care delivery in US hospitals—on direct hospital costs for advanced cancer patients with multiple comorbidities. Evidence on the relationship among palliative care, multimorbidity, and cost will inform decision making as policy makers seek to improve care for patients with serious illness while curbing cost growth.

# Study Data And Methods

#### **Study Population**

Clinical and hospital cost data were collected between 2007 and 2011 using a prospective and observational multisite study design to evaluate the effect of palliative care on patients with an advanced cancer diagnosis. Patients were recruited from six hospitals—two tertiary care academic medical centers (Mount Sinai Medical Center, in New York City; and

Froedtert Hospital, in Milwaukee, Wisconsin), a specialty cancer center (Virginia Commonwealth University Massey Cancer Center, in Richmond), and three community teaching hospitals (Mount Carmel East, Mount Carmel West, and Mount Carmel St. Ann's, all part of the Mount Carmel Health System, in Columbus, Ohio). Study sites were geographically and structurally diverse and represented ethnically and socioeconomically diverse patient populations. The study was approved by each facility's Institutional Review Board. All participants provided written informed consent at enrollment.

Patients were eligible to participate in the study if they were older than eighteen years and fluent in English and had been admitted to a hospital with an advanced cancer diagnosis. Eligible diagnoses were the following: stage 3 or 4 laryngeal, throat, nasopharyngeal, mouth, or head and neck cancer; non-small-cell lung cancer; mesothelioma, esophageal, stomach or gastric, pancreatic, gallbladder, bile duct, cholangio, ampullary, liver, hepatic, hepatocellular, or ovarian cancer; stage 4 breast, kidney, renal cell, endometrial, uterine, cervical, sarcoma, prostate, or melanoma cancer; Dukes' stage D colon cancer; extensive stage small-cell lung cancer; transplant-ineligible multiple myeloma; relapsed or transplant-ineligible lymphoma; and glioblastoma multiforme.

Patients were excluded if their primary physicians refused to approve their participation in the study or if they were unresponsive or nonverbal, had a diagnosis of dementia, or had previously received a hospital palliative care consultation.

#### Methodology

Patients were not randomly assigned to the treatment or comparison group: Those who were seen by a palliative care consultation team were placed in the treatment group, and those who received usual care only were placed in the comparison group. We controlled for differences using propensity score weights.[24] Treatment and comparison groups were matched based upon multiple potential confounders (Exhibit 1). Full details of the propensity score generation and balancing are provided in the online Appendix.[25]

To examine the cost effect of palliative care for patients by multimorbidity, we stratified our sample using the Elixhauser comorbidity index.[26] This is an additive index that counts the presence of thirty-one serious conditions, including the following three cancer diagnoses: lymphoma, metastatic cancer, and solid tumor without metastasis. It was therefore possible for a patient to have an advanced cancer diagnosis (for example, myeloma) that made him or her eligible for the study but whose diagnosis was not reflected in his or her comorbidity score. We created three subsamples according to multimorbidity at hospital admission: patients with scores on the Elixhauser comorbidity index of 0–1, 2–3, and 4 or higher (Exhibit 2).

We separated patients discharged alive from those who died in the hospital to reduce the amount of unobserved heterogeneity in clinical status and underlying treatment decisions and preferences.[27] The number of patients with multimorbidity who died during the hospitalization (n = 54; survival data missing for 3 patients) was too small to support a separate weighted analysis. Therefore, we performed our primary analysis on those discharged alive only and conducted sensitivity analyses with patients pooled irrespective of

discharge status. Our treatment variable was the receipt of a consultation by a palliative care team within two days of hospital admission. Such a timing-sensitive specification of treatment (as opposed to receiving a consultation at any time) reduces the risk of a type 2 error and improves model performance (details available from the authors). Patients seen by a palliative care team after more than two days in hospital were excluded from our primary analyses and incorporated into our secondary and sensitivity analyses.[22]

We examined several different approaches to propensity score matching (for example, oneto-one matching, one-to-many matching within specified calipers, and inverse probability matching). Kernel weights achieved the best balance across observed confounders with the least amount of bias and were selected for analyses. Treated patients received a weight of one. Individuals in the comparison group with a propensity score within a bandwidth of 0.06 of a treated individual's propensity score were weighted based on their distance from the treated individual. A detailed description of the sample construction, matching methodology selection, and propensity score weight calculation for this study has been published previously.[24] Separate weights were calculated within each subsample.[28]

The primary outcome of interest was total direct hospital costs for the index hospitalization —specifically, the estimated mean treatment effect, or the mean estimated change in total direct hospital costs if a patient in the comparison group was moved to the treatment group, with all other covariates held constant at their original values. Direct costs are those attributable to a specific utilization during hospital stay. Variable direct costs are those that are dictated wholly by treatment of the specific patient, such as those for medical supplies and pharmaceuticals and imaging and laboratory expenditures. Fixed direct costs are those that do not vary with a specific patient's utilization but that nonetheless can be identified with the treatment of that patient (for example, staff salaries and equipment expense).[29] Cost data were standardized to 2011 dollars, since that year was the end point of data collection.

Generalized linear model regression (gamma distribution, log link) was performed on total direct costs against a binary intervention variable, the independent variables listed in Exhibit 1, and fixed-effects variables to control for hospital site.

#### Limitations

Our study had several limitations. First, propensity score weights ensured balance between treatment and comparison arms on observed covariates but did not control for unobserved confounders. An instrumental variable, which would have helped control for unobserved confounding, was not available within our data set.

However, a strength of our data set and what sets our study apart from previous ones[21] is its inclusion of rich patient-reported information on many important potential confounders, including demographic and socioeconomic factors, psychological and physical symptoms, functional status, and formal health care use before hospitalization or at hospital admission (Exhibit 1). Because we hypothesized the hospital site to be a weaker potential confounder than the patient characteristics included in our propensity score model, we chose to account

for site differences in costs via fixed effects in our regression model instead of including them in the propensity score models.

A second limitation of our study is that patients who received palliative care consultations may have been more inclined to elect less aggressive (and less expensive) care, even without the involvement of palliative care. However, previous reports that demonstrate that per diem hospital costs decline after palliative care consultations suggest that palliative care consultation teams have a causal impact on goals of care and treatment decisions.[30, 31]

Our models included data on advance directives—whether or not a patient had completed a living will and designated a proxy at baseline. Although we did not have data on specific patient preferences, people who wish to restrict life-prolonging treatment are more likely to complete an advance directive than those who do not wish to restrict such treatment (because the default treatment option is usually to intervene). Although patient preferences should drive care, studies have consistently demonstrated that the effect of such preferences on treatments received is small compared to the other variables included in our analyses.[32]

A third limitation is that the impact of palliative care consultation teams on patient and family outcomes has not yet been evaluated with the study data analyzed in this article; this impact will be addressed in future articles. Therefore, the cost savings reported here represent only evidence that the intervention is cost-effective, based on a "noninferiority" assumption—that is, the assumption that outcomes were at least no worse for patients in the intervention group than for those in the comparison group. This assumption is well supported by reports that hospital inpatient palliative care teams improve symptom control, quality of life, emotional burden, and caregiver and patient satisfaction.[33–36]

Our results were derived from studying data from hospitals with established palliative care teams that met both the current standards for the Joint Commission's Advanced Certification Program for Palliative Care[37] and the guidelines established by the National Consensus Project for Quality Palliative Care.[38] Thus, our results likely reflect savings that can be expected from programs of acceptable quality and provide a target for programs that are being developed. As access to high-quality palliative care teams increases, the generalizability of our results to hospitals with substandard programs will become less of a concern.

A fourth limitation is that inclusion in our study reflected patients' ability to participate throughout their hospitalization. This means that the very sickest enrolled patients may have disproportionately been omitted as a result of incomplete data. The final limitation is that our data did not include professional fees or the costs of postacute care, costs from the payer's perspective, or costs from the patient's or family's perspective, which may include an impact on family wages or savings.[21] Hospital costs reflect one portion of all costs of hospitalization.

# Study Results

#### **Patient Attributes**

There were 906 patients with advanced cancer matched for analysis, 193 (21 percent) of whom were seen by a palliative care consultation team within two days of admission during the index hospitalization. Baseline characteristics treated as covariates are reported in Exhibit 1. The prevalence of each of the thirty-one comorbidities in the Elixhauser index for each multimorbidity-defined subsample is provided in Exhibit 2.

#### Treatment Effect

Receipt of a consultation by a palliative care team within two days of admission was significantly associated with lower total direct hospital costs for advanced cancer patients with multimorbidity, and the effect size grew larger as the number of comorbidities increased (Exhibit 3). For patients with a comorbidity score of 0–1, the estimated mean treatment effect was not significant. For patients with a score of 2–3, the estimated effect was a reduction in costs of \$2,321 (22 percent). For patients with a score of 4 or higher, the reduction was \$3,515 (32 percent).

## Secondary Analyses

Elsewhere we have demonstrated a systematic relationship between time to consult and the palliative care consultation's effect on cost.[22] To examine whether earlier palliative team treatment was associated with lower hospital costs among patients with differing levels of multimorbidity, we combined the two analytical approaches. We created subsamples defined both by number of comorbidities and by definitions of treatment according to time to consult. The results demonstrate a consistent pattern: Associations between higher patient comorbidity scores and cost-saving effect of the treatment and between earlier time to consultation and cost-saving effect were both robust (Exhibit 4). For any given definition of treatment according to timing, the cost-saving effect was larger for the group with higher comorbidity scores. And for either subsample defined by comorbidity score, the cost-saving effect was larger for earlier interventions.

To examine the underlying source of the observed cost-saving effect, we examined the treatment effect on hospital length-of-stay and major utilization categories. The results showed that the intervention was significantly associated with a reduction in laboratory costs, which was inferred to result from palliative care consultation's reducing the number of patient tests, and shorter length-of-stay, which was inferred to result from patient discharge being expedited by discussions of goals of care. Both effects were larger for subsamples with higher comorbidity scores (see the Appendix).[25]

#### **Confirmatory Analyses**

Our results were robust to multiple sensitivity analyses: pooling of decedents with those discharged alive; removing high-cost utilization outliers; alternative approaches to intervention definition by timing; modeling outcomes with and without propensity score weights; and using length-of-stay to control for unobserved confounding (data not shown).

# Discussion

Our results demonstrate that palliative care consultation is significantly associated with reduced direct hospital costs for advanced cancer patients with multimorbidity, and the average effect is larger for patients with higher comorbidity scores. Previous studies have estimated that the cost-saving effect of palliative care consultation for patients discharged from the hospital alive is in the range of 5-14 percent.[21] Thus, the magnitude of cost savings for patients with multimorbidity – 22% for a comorbidity score of two or three, 32% for a score of four or more - appears much larger than previously suggested. These savings result from a combination of reduced utilization during hospital stay and reduced length-of-stay.

This is the first study we are aware of to examine if the treatment effect of palliative care consultation varied by level of patient comorbidities. Our results have a number of potentially important implications in the policy context of care for patients with advanced cancer and multimorbidity.

#### **Reforming Care For Patients With Multimorbidity**

The long-term viability not only of government-funded health programs such as Medicaid and Medicare but also of the national health system overall depends on reforming the provision of care to patients with serious illness in a way that reduces costs without compromising quality and access.[39] One piece of this jigsaw puzzle is timely access to palliative care. Patterns of improved quality and reduced costs through coordinated patientcentered palliative care are already evident in the literature.[20, 35]

Our results demonstrate for the first time that the cost-effects of palliative care consultation teams are on average larger for patients with advanced cancer and higher comorbidity scores, compared to those with advanced cancer and lower comorbidity scores. This indicates that the cost-saving scope of hospital-based palliative care programs for patients with multimorbidity, who account for a disproportionate share of health care costs, may be larger than previously realized. Early palliative consultations on the sickest patients may help reverse two trends: the increasing use of unwanted aggressive end-of-life care observed in Medicare patients with advanced cancer and the increasing percentage of patients who use hospice for less than seven days.[11] While the cost-saving effect of palliative care consultation teams appears greatest for patients with higher comorbidity scores, the intervention may also be beneficial for both patient outcomes and in-hospital utilization among people with a single serious illness. The nonsignificant result for patients with a comorbidity score of 0-1 in Exhibit 3 may arise from a sample size issue within this subsample. Only twenty-eight patients in the treatment group had such low comorbidity scores, which may have contributed to the absence of a significant association in our analysis.

#### Workforce Allocation

There are demonstrable short- and long-term gaps in the hospice and palliative care workforce.[40] The projected level of future need is such that not all patients will be seen by

specialists—who are already and will remain a scarce resource to be allocated in the most effective way. Our results strongly suggest that palliative care consultation teams are most likely to have an impact with patients who have higher numbers of co-occurring conditions. On the evidence-based assumption that patient and family outcomes are at least as good for advanced cancer patients with multimorbidity who receive palliative care as they are for those who do not, specialist palliative care would be most cost-effective with patients who have more comorbidities, and they should thus be prioritized.

#### **Case For Increased Access To Palliative Care**

Programs using palliative care consultation teams have rapidly expanded in recent years, and over 90 percent of medium-size to large hospitals in the United States now have a palliative team.[41] Yet in our primary analysis 25 percent of patients with an advanced cancer diagnosis and multimorbidity admitted to hospitals with well-established palliative care programs received a consultation with a palliative team within two days of admission (Exhibit 3). There is demonstrable scope for reducing costs and improving care through increased access to specialist palliative care for patients with advanced cancer and complex multimorbidity needs. Currently 35 percent of direct medical cancer costs in the United States are attributable to inpatient hospital stays, and these costs are expected to increase.[1]

# Importance Of Screening At Admission And Early Intervention

Palliative care is increasingly available both earlier in the care trajectory than it was in the past and concurrent with curative care, with observable benefits.[33, 34, 42] We have shown elsewhere that delivering palliative care consultations earlier to patients with advanced cancer also brings economic benefits.[22]

There was no formal system for identifying or prioritizing potential palliative care patients on the basis of comorbidity in the consultation model we studied. Screening patients with advanced cancer for palliative care needs at hospital admission may facilitate early intervention, which our results suggest would maximize the cost impact.

# **Policy Implications**

Translating a growing body of evidence on palliative care programs into improved care for seriously ill patients requires changes to policy. In addition to workforce allocation, areas requiring urgent attention include scaling up and disseminating successful models of provision of high-quality palliative care and the design of regulatory, accreditation, payment, and financing mechanisms that strengthen access.[43]

Our results relate only to the intervention's impact on hospital costs. It is not clear how the use of palliative care consultation teams affects insurance expenditures, since the extent to which reduced hospital costs are passed on to payers varies by reimbursement system.[44] Palliative care provided early in a hospitalization may change the procedures performed and thus the Medicare Severity Diagnosis-Related Group, which could result in lower expenditures by Medicare (or other case-rate payers) for that case than would otherwise have been required.

# **Future Research**

The fact that we found greater cost savings for cancer patients with more comorbidities than for those with fewer comorbidities raises the question of whether similar results would be observed in patients with other serious illnesses and multimorbidity. Further important extensions of this work will be to identify how costs and palliative care's effects on costs vary for specific combinations of comorbidities and diagnoses, and to determine when in the course of illness specialist palliative care is most cost-effective. Finally, future studies should examine whether and how the cost-saving effect of inpatient palliative care consultation teams has an impact on payer expenditures.

# Conclusion

Patients with multiple serious conditions account for a high proportion of US health care spending, and substantial health spending growth is projected over the next decade. Previous studies have established the clinical and financial benefits of palliative care, and our results supplement these studies by demonstrating that among patients with advanced cancer, the cost effect is greater for those with higher numbers of serious coexisting conditions. Increasing access to palliative care during hospitalization for patients with advanced cancer and multiple chronic conditions could improve care while complementing strategies to curb cost growth.

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# Notes

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To access the Appendix, click on the Appendix link in the box to the right of the article online

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#### Exhibit 4.

Estimated Effect Of Consultation With Palliative Care Consultation Team On Total Direct Hospital Cost For Patients With Advanced Cancer, By Elixhauser Cormorbidity Score At Admission And Timing Of Consultation

SOURCE Authors' analysis. NOTES We regressed total direct costs as explained in Exhibit 3 Notes. Of the 906 matched patients, 98 had an advanced cancer that was not lymphoma, metastatic cancer, or a solid tumor. Of these, 49 (50 percent) had no comorbidity on the Elixhauser comorbidity scale and thus had an Elixhauser comorbidity score of 0, as explained in the text. The other 49 had an Elixhauser comorbidity score of 1–8 (median: 2), with their cancer diagnosis unaccounted for in the score.

## Exhibit 1

Baseline Characteristics Of Matched Groups Of 906 Patients, Treatment (Patients Seen By A Palliative Care Consultation Team [PCCT]) And Comparison (Usual Care Only) Between 2007 And 2011

Characteristic	Compari son group $(n = 713)$	Treatme nt group (n = 193)	Standar dized differenc e (%)
Age (years)	,	,	. ,
55-75	53%	55%	5.3
More than 75	13	11	-6.7
Female	53	52	-1.7
Race			
White	60	62	4.7
Black	35	35	< 0.1
Had advance directive	49	47	-3.4
Insurance			
Medicare only	23	24	2.6
Medicaid and Medicare	26	24	-4.3
Highest level of education			
High school	55	56	1.2
College	37	35	-2.8
Used visiting nurse services 2 weeks before hospitalization	15	14	-1.9
Hours of home health aide use 2 weeks before hospitalization	1.0	0.9	-2.7
Primary diagnosis of lymphoma or myeloma	5%	5%	1.1
Patient had a complication <sup>b</sup>	1%	1%	-1.0
Mean Elixhauser comorbidity score	3.4	3.3	-6.8
Needed partial or complete help with ADLs			
Bathing (partial)	39%	39%	0.8
Transferring from chair (partial)	34	35	2.7
One or more ADL (complete)	13	13	1.3
Mean ESAS score			
Physical at admission	2	2	-0.4
Psychological at admission	1.6	1.6	3.8
Physical on the reference day	1.8	1.8	-1.8
Psychological on the reference day	1.5	1.4	-3.8
Mean CMSAS score <sup>C</sup>			
Number at admission	8.9	8.9	0.2
Number on the reference day	7.8	7.7	-4.1
Severity at admission	15.8	15.8	0.4
Severity on the reference day	12.5	12.4	-1.2

Characteristic	Compari son group (n = 713)	Treatme nt group ( <i>n</i> = 193)	Standar dized differenc e (%)
Equivalent dose of morphine (mg), $d$	21.7	22.2	1.3
In pain <sup>e</sup>			
Somewhat	9%	10%	3.0
Quite a bit	29	30	1.2
Very much	35	35	0.5
Fatigued <sup>e</sup>			
A little, somewhat, or quite a bit	38	37	-2.7
Very much	29	29	0.3

SOURCE Authors' analysis. NOTES There were 910 patients in the final sample who were eligible for the primary analysis in this article. In matching stratified subsamples, one palliative care consultation team (PCCT) patient with 2–3 comorbidities and three PCCT patients with 4 or more comorbidities were lost to matching. No patients in the comparison group were lost to matching in any subsample. There are therefore 906 patients in the primary analysis in this article. For patients in the treatment group, the reference day was the day of consultation; for patients in the comparison group, it was the day their symptom severity was most similar to that of palliative care patients. ADLs are activities of daily living. ESAS is the Edmonton Symptom Assessment Scale [45], which evaluates six physical and three psychologic symptoms on a scale of 0 to 10 (0=absence of symptom; 10=most severe presence): Pain, tiredness, nausea, drowsiness, appetite and shortness of breath; and depression, anxiety and wellbeing. CMSAS is the Condensed Memorial Symptom Assessment Scale [46], which evaluates 14 symptoms on a scale of 0 to 4 (0=absent; 4=Very much): Lack of energy, lack of appetite, pain, dry mouth, weight loss, feeling drowsy, shortness of breath, constipation, difficulty sleeping, difficulty concentrating, nausea, worrying, feeling sas and feeling nervous. Reference categories are as follows: for binary variables, no; age, younger than fifty-five; race, other; insurance, neither Medicare nor Medicaid; education, elementary school; pain and fatigue, none.

<sup>b</sup>Major or minor complication on the reference day.

 $^{C}$ Number is the number of physical symptoms on the CMSAS; severity is the number of physical symptoms multiplied by the mean severity of physical symptoms on the CMSAS.

 $^{d}$ Average daily dose of opioids in milligrams of morphine sulfate equivalents in week prior to hospitalization.

 $e^{0}$  On the reference day. Standardized differences measure the imbalance between treatment and comparison groups on baseline characteristics, taking into account both means and variances.

# Exhibit 2

Prevalence Of Specific Comorbidities Within Subsamples, By Elixhauser Comorbidity Score At Admission

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	0-1 ( 23	<b>n</b> = (5	2-3 ( 362	n = (1	4 or hi $(n = 2)$	igher 309)
Comorbidity	No.	%	N0.	%	N0.	%
Congestive heart failure	0	0	5	1	54	17
Cardiac arrhythmia	0	0	20	6	93	30
Valvular disease	0	0	3	1	26	8
Pulmonary circulation	1	0	6	2	33	11
Peripheral vascular disorders	0	0	7	2	18	9
Hypertension (uncomplicated)	9	3	166	46	169	55
Hypertension (complicated)	0	0	1	0	47	15
Paralysis	0	0	7	2	15	5
Neurologic disorders other than paralysis	0	0	18	5	25	8
Chronic pulmonary disease	1	0	36	10	106	34
Diabetes (uncomplicated)	0	0	11	3	26	8
Diabetes (complicated)	0	0	0	0	12	4
Hypothyroidism	1	0	26	7	50	16
Renal failure	1	0	7	2	55	18
Liver disease	0	0	10	3	17	9
Peptic ulcer disease	0	0	1	0	6	2
AIDS or HIV	0	0	2	1	1	0
Lymphoma	9	4	22	6	30	10
Metastatic cancer	107	46	248	69	207	67
Solid tumor without metastasis	116	49	251	69	217	70
Rheumatoid arthritis	0	0	2	1	11	4

	E	lixhau	ser cor	norbid	lity scor	e
	0–1 ( 23	(n =	2–3 ( 362	= (2	$\begin{array}{l} 4 \text{ or } h \\ (n = 2 \end{array}$	igher 309)
Comorbidity	No.	%	No.	%	N0.	%
Coagulopathy	0	0	8	2	36	12
Obesity	0	0	3	1	16	5
Weight loss	0	0	59	16	100	32
Fluid or electrolyte disorders	9	3	108	30	178	58
Anemia (blood loss)	0	0	3	1	1	0
Anemia (deficiency)	0	0	5	1	15	5
Alcohol abuse	0	0	4	1	18	9
Drug abuse	0	0	10	3	10	3
Psychoses	0	0	2	1	6	3
Depression	0	0	47	13	95	31

SOURCE Authors' analysis. NOTES Of the 906 matched patients, 98 had an advanced cancer that was not lymphoma, metastatic cancer, or a solid tumor. Of these, 49 (50 percent) had no comorbidity on the Elixhauser comorbidity scale and thus had an Elixhauser comorbidity score of 0, as explained in the text. The other 49 had an Elixhauser comorbidity score of 1–8 (median: 2), with their cancer diagnosis unaccounted for in the score. Author Manuscript

Estimated Effect Of A Consultation With A Palliative Care Team Within Two Days Of Admission On Total Direct Hospital Costs For Patients With Advanced Cancer, By Elixhauser Comorbidity Score At Admission

	<b>T</b>			
	Implied mean saving <sup>b</sup>	18%	22	32
Primary results	95% CI	-5,093, 1,544	-3,869, -773	-5,949, -1,081
	Estimated mean treatment effect <sup>a</sup>	-\$1,775	-2,321 *	-3,515 *
ary	Mean LOS per patient (days)	7.4	7.0	8.2
ation summ	Mean direct costs per patient	\$8,440	8,528	10,030
Utiliz	Proporti on in treatmen t group	12%	24	26
	All patie nts	235	362	309
Sample	Treatme nt group	28	98	79
5	Compariso n group	207	276	230
	Comorbidity score	0-1	2–3	4 or higher

SOURCE Authors' analysis. NOTES We regressed total direct costs against a binary intervention variable, the independent variables listed in Exhibit 1, and fixed-effects variables to control for hospital site, lymphoma, metastatic cancer, or a solid tumor. Of these, 49 (50 percent) had an advanced cancer diagnosis but no comorbidity on the Elixhauser comorbidity scale and thus had an Elixhauser comorbidity applying subsample-specific propensity score weights in all cases. Further details are available in the Appendix (see Note 25 in text). Of the 906 matched patients, 98 had an advanced cancer that was not score of 0, as explained in the text. The other 49 had an Elixhauser comorbidity score of 1–8 (median: 2), with their cancer diagnosis unaccounted for in the score. Costs are in 2011 dollars. Significance refers to the estimated mean treatment effect.LOS is length-of-stay. CI is confidence interval.

 $\frac{a}{2}$ The estimated effect on total direct hospital cost of moving a patient from the comparison arm to the treatment arm, holding all other values constant.

b savings in total direct hospital costs resulting from the consultation. Further details on how these savings were calculated are available in the Appendix (see Note 25 in text).

p < 0.01