

Development and Evaluation of a Gentamicin Pharmacokinetic Model That Facilitates Opportunistic Gentamicin Therapeutic Drug Monitoring in Neonates and Infants

Eva Germovsek,^a Alison Kent,^b Tuuli Metsvaht,^c Irja Lutsar,^c Nigel Klein,^a Mark A. Turner,^d Mike Sharland,^b Elisabet I. Nielsen,^e Paul T. Heath,^b Joseph F. Standing,^a on behalf of the neoGent Collaboration

Inflammation, Infection and Rheumatology Section, Institute of Child Health, University College London, London, United Kingdom^a; Paediatric Infectious Diseases Research Group, Institute for Infection and Immunity, St. George's, University of London, Cranmer Terrace, London, United Kingdom^b; Department of Microbiology, University of Tartu, Tartu, Estonia^c; Department of Women's and Children's Health, Institute of Translational Medicine, University of Liverpool, Liverpool, United Kingdom^d; Department of Pharmaceutical Biosciences, Uppsala University, Uppsala, Sweden^e

Trough gentamicin therapeutic drug monitoring (TDM) is time-consuming, disruptive to neonatal clinical care, and a patient safety issue. Bayesian models could allow TDM to be performed opportunistically at the time of routine blood tests. This study aimed to develop and prospectively evaluate a new gentamicin model and a novel Bayesian computer tool (neoGent) for TDM use in neonatal intensive care. We also evaluated model performance for predicting peak concentrations and the area under the concentration-time curve from time 0 h to time t h (AUC_{0-t}). A pharmacokinetic meta-analysis was performed on pooled data from three studies (1,325 concentrations from 205 patients). A 3-compartment model was used with the following covariates: allometric weight scaling, postmenstrual and postnatal age, and serum creatinine concentration. Final parameter estimates (standard errors) were as follows: clearance, 6.2 (0.3) liters/h/70 kg of body weight; central volume (V), 26.5 (0.6) liters/70 kg; intercompartmental disposition (Q), 2.2 (0.3) liters/h/70 kg; peripheral volume V2, 21.2 (1.5) liters/70 kg; intercompartmental disposition (Q2), 0.3 (0.05) liters/h/70 kg; peripheral volume V3, 148 (52.0) liters/70 kg. The model's ability to predict trough concentrations from an opportunistic sample was evaluated in a prospective observational cohort study that included data from 163 patients and 483 concentrations collected in five hospitals. Unbiased trough predictions were obtained; the median (95% confidence interval [CI]) prediction error was 0.0004 (−1.07, 0.84) mg/liter. Results also showed that peaks and AUC_{0-t} values could be predicted (from one randomly selected sample) with little bias but relative imprecision, with median (95% CI) prediction errors being 0.16 (−4.76, 5.01) mg/liter and 10.8 (−24.9, 62.2) mg · h/liter, respectively. neoGent was implemented in R/NONMEM and in the freely available TDMx software.

The aminoglycoside antibiotic gentamicin is the most commonly used antimicrobial in neonatal units (1, 2) and is effective against Gram-negative bacteria. Gentamicin use is limited by its narrow therapeutic index and risk of toxicity, specifically, nephro- and ototoxicity (3). It is not metabolized in the liver (4) and is almost entirely eliminated by the kidneys; clearance therefore depends on renal function. During the first 2 weeks of life, renal and intrarenal blood flow increase rapidly, causing a steep rise in the glomerular filtration rate (GFR) (5, 6).

Therapeutic drug monitoring (TDM) is required to ensure maximal efficacy and, in particular, minimal toxicity, particularly in the neonatal population, where the variability in pharmacokinetic (PK) parameters is large. Dose individualization approaches focus on toxicity (7, 8) and include single-level methods and nomograms (9, 10), area under the curve (AUC) methods (11), and Bayesian methods (12). The use of nomograms is limited as they cannot readily incorporate covariates affecting PK parameters. AUC methods use a simplified 1-compartment PK model and require at least two gentamicin measurements, which is not appropriate in neonates with limited blood volumes. These drawbacks make Bayesian approaches the most attractive for newborn infants.

Deriving a Bayesian prior for TDM requires a nonlinear mixed-effect PK model, and several such studies of neonatal gentamicin were previously published (13–24). However, those studies were limited by their heterogeneity and use of sparse data (of-

ten identifying only a 1-compartment model, whereas gentamicin follows multicompartment kinetics [25, 26]) and failed to account for age-related differences in creatinine levels during the immediate newborn period. Although gentamicin is not a new drug, its dosing and monitoring are still current issues as identified in the United Kingdom National Patient Safety alert (<http://www.nrls.npsa.nhs.uk/alerts/?entryid45=66271>) and in a recent publication by Valitalo et al. (27), who used simulations to define dosing guidelines.

We aimed to investigate whether opportunistic sampling can predict trough gentamicin concentrations so that standard TDM can be performed using a blood sample taken for other purposes

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Address correspondence to Eva Germovsek, eva.germovsek.11@ucl.ac.uk.

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TABLE 1 A summary of demographics and dosing^a

Parameter	Model-building data set	Evaluation data set
No. of subjects	205	163
Wt (kg) ^b	2.12 (0.53–5.05)	2.03 (0.48–5.05)
Gestational age (wks) ^b	34.0 (23.3–42.1)	34.3 (23.9–42.3)
Postnatal age (days) ^b	5.4 (1–66)	6 (1–78)
Postmenstrual age (wks) ^b	33.0 (23.3–43.8)	34.9 (24–43.3)
No. (%) of females	89 (43%)	68 (41.7%)
No. of gentamicin samples per patient ^c	6.5	3.0
Gentamicin concn (mg/liter) ^b	3.4 (0.3–37.6)	1.0 (0.1–13.2)
Time after the dose (h) ^b	8.0 (0.02–54.1)	23.5 (0.08–79.7)
No. of occasions ^b	2 (1–22)	2 (1–7)

^a Weight and gestational age data represent values at treatment initiation; the rest of the data represent values at the time of gentamicin sampling/dosing. An occasion was defined as a dose with subsequent gentamicin samples taken; day of birth was defined as day 1.

^b Data represent medians (ranges).

^c Data represent means.

(e.g., routine blood gases). As a secondary aim, we evaluated the model's ability to predict peak gentamicin concentrations and AUC from time 0 h to time t h (AUC_{0-t}) using one randomly selected sample.

MATERIALS AND METHODS

Study population. This study used two data sets: a model-building data set and a prospectively collected evaluation data set.

To collect data for model development, the electronic bibliographic database PubMed was searched in January 2015 without time limitations. The search strategy included the following: (neonat* OR newborn*) AND (gentamicin) AND (pharmacokinetic* OR PK). Gentamicin samples had to be prospectively collected, and covariates (weight, gestational age [GA], postnatal age [PNA], serum creatinine concentration [SCr]) also had to be reported. We also searched the reference lists in identified papers. The authors of the publications that met the inclusion criteria ($n = 8$) (11, 15, 21, 22, 28–31) were then invited to contribute their data.

Data for the evaluation of the PK model were collected as a prospective observational cohort study from five United Kingdom hospitals (St George's University Hospitals NHS Foundation Trust, Liverpool Women's NHS Foundation Trust, Oxford University Hospitals, Portsmouth Hospitals NHS Trust, and Coventry & Warwickshire University Hospitals NHS Trust) from July 2012 to November 2013. Infants were eligible for inclusion if the following criteria were met: more than 36 h gentamicin therapy anticipated; postnatal age of less than 90 days; no extracorporeal membrane oxygenation, peritoneal dialysis, or hemofiltration received; and expectation of survival of the study period (as judged by the clinical team). Each patient provided a minimum of two gentamicin concentrations—a trough sample from routine TDM (i.e., a predose sample taken before a noninitial dose) and an additional study sample (taken opportunistically during a course of gentamicin when the infant required blood sampling for clinical care). These samples are referred to as routine (trough) and opportunistic study samples in this article. Exact times of gentamicin dosing and sampling were recorded, along with the patient's weight, age, and serum creatinine concentration (Table 1). Written informed consent was obtained from parents, and the study was approved by the London Central Ethics Committee (reference 12/LO/0455).

Gentamicin dosing and sampling procedure in the prospective evaluation data set. Gentamicin treatment was initiated at the discretion of the clinical team for possible infection and dosed and monitored using trough concentrations according to the standard practice at each hospital. Gentamicin was administered as a slow (<2-min) bolus via intravenous cannula, percutaneous long line, or umbilical venous catheter.

Bioanalytical techniques. An enzyme immunoassay (EMIT; Syva) (15), a fluorescence polarization immunoassay (TDx; Abbot) (15, 21), and high-performance liquid chromatography coupled with tandem mass

spectrometry (UHPLC-MS/MS) (32) were used to determine gentamicin concentrations in the model-building data set, and the Jaffe reaction (33) was used to determine serum creatinine concentrations. In the prospective evaluation data set, gentamicin serum concentrations were analyzed using immunoassay techniques (see Table S1 in the supplemental material), and creatinine concentrations were determined by either a Jaffe-based method or an enzymatic method (137 neonates and 26 neonates, respectively).

Pharmacokinetic analysis. The observed concentration-time data from the model-building studies only were pooled and simultaneously analyzed with nonlinear mixed-effects software (NONMEM version 7.3) (34). The first-order conditional estimation method with interaction was used.

Basic model. One-, 2-, and 3-compartment structural models were considered in defining the basic structural population PK model. The interindividual variability (IIV) was assumed to follow a log-normal distribution and was tested on all parameters. An additive residual error model, a proportional residual error model, and a combination of the two (equation 1) were tested:

$$y_{ij} = f(t_{ij}; \phi_i) + f(t_{ij}; \phi_i) \cdot \varepsilon_{ij(\text{proportional})} + \varepsilon_{ij(\text{additive})} \quad (1)$$

where y_{ij} is the observed gentamicin concentration at time t_{ij} , f is the function that represents the gentamicin model, ϕ_i is a vector of parameters, and ε_{ij} is a residual error term.

Interoccasion variability (IOV) was also assumed to be log-normally distributed, and it was tested for all parameters, with an occasion defined as a single dosing interval.

Covariate model. Allometric scaling was used *a priori* to standardize all PK parameters to 70 kg (35), and a maturation function, describing the maturation of the GFR with postmenstrual age (PMA) (equation 2) with fixed parameters from a previous study (5), was used to scale clearance. Allometric exponents were fixed to 0.632 for central clearance and 0.75 for intercompartmental clearances. The two different exponents were used because these values were shown to be the best for describing the maturation of renal elimination (5) and tissue blood flows (36), respectively. Allometric exponents for volumes of distribution were fixed to 1. The combination of allometric weight scaling and sigmoidal maturation function was suggested as a standard method for scaling clearance in the pediatric population in a recent comparison of different approaches (37).

$$\text{maturation function} = \frac{\text{PMA}^{\text{Hill}}}{\text{PMA}_{50}^{\text{Hill}} + \text{PMA}^{\text{Hill}}} \quad (2)$$

where Hill is the sigmoidicity coefficient and PMA_{50} is the PMA when the maturation of the GFR reaches 50% of adult values.

As it is known that PNA and serum creatinine concentration are important indicators of gentamicin clearance and also based on the *post hoc* estimates of η versus covariate plots, they were tested on clearance. These

time-varying covariates were considered to significantly improve the fit and were therefore included in the model when the difference in objective function value (ΔOFV) after their inclusion was >3.84 ($P < 0.05$). Additionally, linear extrapolations between observations were made. To account for endogenous and maternal creatinine concentrations and also for the change in renal function with age, a typical value of serum creatinine concentration for a specific PMA (TSCr), was determined using data from Cuzzolin et al. (38) for preterm (GA, <37 weeks) newborns and data from Rudd et al. (39) for term newborns. A linear decline in TSCr with increasing PMA was found according to equation 3:

$$\text{TSCr} = -2.849 \cdot \text{PMA (weeks)} + 166.48 \quad (3)$$

A possible influence of serum creatinine on clearance was tested according to equation 4, where the measured serum creatinine concentration (MSCr) was standardized using TSCr and departures from it were estimated as follows:

$$\left(\frac{\text{MSCr}}{\text{TSCr}}\right)^{\theta} \quad (4)$$

The effect of PNA was investigated with a logistic function (equation 5) to account for the rapid changes in gentamicin clearance in the first hours of life (the first day of life was defined as day 1) as follows:

$$\text{postnatal age function} = \frac{\text{PNA}}{\text{PNA}_{50} + \text{PNA}} \quad (5)$$

where PNA_{50} is the PNA when clearance has reached 50% of the clearance seen in the typical adult.

After the forward selection (ΔOFV , >3.84) of all covariates (full model), backward elimination was performed, with a P value retention cutoff value of 0.001 (ΔOFV , <10.83).

Evaluation. (i) Internal model evaluation. Basic goodness-of-fit plots for observations versus population and individual predictions and for conditional weighted residuals versus population predictions and versus time after dose were produced using R statistical software version 3.1.0 (40) and visually examined. The assumptions of normality and homogeneity of the residuals errors were investigated by inspecting a histogram and a q-q plot.

Standard errors from NONMEM covariance step and nonparametric bootstrap analysis performed with 1,000 replicates were used to determine the precision of the final PK parameter estimates.

Additionally, we simulated 1,000 data sets using parameter estimates from the final model and plotted 95% confidence intervals (CI) around the 2.5th, 50th, and 97.5th prediction percentiles of the simulated data. The observations were then overlaid on the plot, also called the visual predictive check (VPC). Perl-speaks-NONMEM (PsN) software (41) was used for the bootstrap analysis and to produce the VPC, which was visualized using R-package Xpose4 (42).

(ii) External model evaluation. The prospective evaluation data set was used to evaluate the predictive performance of the model. No additional fitting was done, and the diagnostic plots and the VPC were generated as described above.

Bayesian model-predicted trough concentrations were computed using the model as a prior and information from the opportunistic study samples only. These predictions were compared with the observed trough concentrations by calculating the prediction error (PE) (43) and also the mean PE (MPE) (i.e., a measure of bias) and the root mean square error (RMSE), a measure of precision (44) (equations 6):

$$\text{PE} = \text{observed} - \text{predicted} \quad (6)$$

$$\text{MPE} = \frac{1}{N} \cdot \sum_{i=1}^N \cdot \text{PE}_i$$

$$\text{RMSE} = \sqrt{\frac{1}{N} \cdot \sum_{i=1}^N \cdot \text{PE}_i^2}$$

Also, we counted the number of "correct" predictions that were below or above the currently recommended gentamicin trough concentra-

tion threshold of 1 mg/liter or 2 mg/liter (the National Institute for Health and Care Excellence [NICE]; <http://www.nice.org.uk/guidance/CG149/chapter/1-Guidance#therapeutic-drug-monitoring-for-gentamicin>) and British National Formulary for Children [45]).

Further analysis of paired samples (that is, both study and routine samples taken in the same dosing interval) was undertaken for the following scenarios: study samples at concentrations of ≥ 1 , ≥ 2 , and ≥ 3 mg/liter (compared with unpaired samples only).

Cross-validation. The subset with the study sample concentration above 3 mg/liter provided the most important comparison, since in this case the study sample concentration was still above the prespecified trough threshold. As there were only 18 pairs with an opportunistic study concentration of ≥ 3 mg/liter in the evaluation data set, these pairs were merged with paired samples with the same characteristics from the model-building data set. The pooled data set was then randomly split into five subsets, and cross-validation was performed (meaning that 20% of the pairs in each subset were randomly removed and the model was reestimated). The reestimated model was then used as a prior to predict the troughs, and the predicted troughs were compared to the observed trough concentrations as previously described.

Whether the model is able to predict peak concentrations from one randomly selected nonpeak sample was tested essentially as described above, using paired samples from both the model-building data set and the evaluation data set and performing cross-validations. Additionally, in recognition of the fact that a possible pharmacokinetic-pharmacodynamic target for aminoglycosides can also be $\text{AUC}_{0-24}/\text{MIC}$ (46), the model was also evaluated on how it predicts AUC_{0-t} . Only a subset of the data in which five or more samples were collected after the same dose was used for defining AUC_{0-t} , and the model-predicted versus observed (non-compartmental) AUC_{0-t} values were compared.

Comparison with other models. To compare our mechanistic model, which scales for size, age, and expected renal function, with previously published models using empirical covariate analysis, predictions of the trough concentrations were generated from the opportunistically collected samples in our prospective data set.

neoGent software. The model was implemented using R and NONMEM (see the supplemental material). It works as follows: first, an individual's data are read into R; then, Bayesian estimates generated in NONMEM are used to predict outcomes of interest (e.g., the time at which the concentration falls below 2 mg/liter).

RESULTS

Patients. We contacted eight authors identified in the literature search and obtained two large neonatal gentamicin data sets (15, 21). We received no response from four authors (11, 28–30), and, although an initial response was received from two authors (22, 31), no data were actually shared. Additionally, we obtained some previously unpublished data taken during a PK study of ampicillin and penicillin (32). The data were pooled and comprised 1,325 gentamicin concentrations from 205 neonates (Table 1). This data set was used to derive the model.

For the model evaluation, gentamicin serum concentrations were prospectively collected from a total of 194 neonates. Of the enrolled patients, 163 were included in the PK analysis (Table 1). Reasons for exclusion (31 patients) included inexact sampling times, insufficient samples, or the gentamicin opportunistic study concentration being below the limit of quantification ($n = 12$). The final evaluation data set comprised 483 gentamicin serum measurements, with 229 study and 254 routinely taken trough concentrations. Median (range) times after dosing were 13.3 (0.08 to 53.3) h and 31.1 (8.0 to 79.7) h for study and routine concentrations, respectively. Patients were on treatment for up to 20 days.

Pharmacokinetic analysis. Initially, a 2-compartment model

TABLE 2 Final parameter estimates with uncertainty from NONMEM output file and from the bootstrap analysis^a

Parameter	Value from the final model				Value from the bootstrap analysis		
	Mean	SE	% CV	η shrinkage	Median	2.5th percentile	97.5th percentile
CL (liters/h/70 kg)	6.21	0.30			6.14	5.47	6.75
θ _{SCr}	-0.13	0.055			-0.13	-0.25	-0.03
PNA ₅₀ (days)	1.70	0.30			1.68	1.15	2.30
V (liters/70 kg)	26.5	1.11			26.3	23.6	28.4
Q (liters/h/70 kg)	2.15	0.32			2.19	1.68	3.25
V2 (liters/70 kg)	21.2	1.50			20.9	17.9	24.2
Q2 (liters/h/70 kg)	0.27	0.047			0.28	0.19	0.38
V3 (liters/70 kg)	148	52.0			152	65.2	534
IIV on CL	0.175	0.038	41.8	6.9	0.170	0.104	0.254
IIV on V	0.112	0.032	33.5	15.2	0.113	0.057	0.190
CL-V covariance	0.116	0.030			0.115	0.060	0.184
IIV on V2	0.132	0.060	36.3	57.8	0.117	0.023	0.281
IIV on V3	0.177	0.216	42.1	85.0	0.114	0.00002	4.18
Interoccasion variability	0.014	0.007	11.8		0.013	0.001	0.029
Residual error (proportional)	0.036	0.006	19.0		0.036	0.025	0.049
Residual error (additive)	0.016	0.007			0.015	0.000002	0.032

^a CL, clearance; V, volume of distribution; Q (and Q2), intercompartmental CL; IIV, interindividual variability; SE, standard error obtained with NONMEM 7.3 covariance step; CV, coefficient of variation.

provided a better fit to the data ($\Delta\text{OFV} = 7.4$ with a 3-compartment model) and was therefore chosen as the basic structural model. But after the addition of the fixed allometric and renal function parameters, covariates, and IOV, a 3-compartment model described the data better (47-unit drop in OFV). The IIV was described with an exponential error structure, and the best residual error model was a combination of a proportional error and an additive error.

Postnatal age and standardized serum creatinine concentration had a significant effect on clearance ($\Delta\text{OFV} = 134.1$ and $\Delta\text{OFV} = 17.2$, respectively) and were thus included in the final model. Backward elimination ($P = 0.001$) confirmed that these covariates remained significant with the 3-compartment model. The final gentamicin population PK model is summarized with equations 7:

$$\text{CL} = \theta_{\text{CL}} \cdot \left(\frac{\text{WT}}{70}\right)^{0.632} \cdot \frac{\text{PMA}^{3.33}}{55.4^{3.33} + \text{PMA}^{3.33}} \cdot \left(\frac{\text{MSCr}}{\text{TSCr}}\right)^{\theta_{\text{SCr}}} \cdot \frac{\text{PNA}}{\theta_{\text{PNA}_{50}} + \text{PNA}} \cdot e^{(\eta_{\text{CL}} + \kappa_{\text{CL}})} \quad (7)$$

$$V = \theta_V \cdot \left(\frac{\text{WT}}{70}\right) \cdot e^{\eta_V}$$

$$Q = \theta_Q \cdot \left(\frac{\text{WT}}{70}\right)^{0.75} \cdot e^{\eta_Q}$$

where CL is gentamicin clearance, V is gentamicin volume of distribution, Q is intercompartmental gentamicin clearance, WT is body weight in kilograms, η is IIV, and κ is IOV.

There was only a small improvement in fit ($\Delta\text{OFV} = 7.6$) when the model was parameterized for time-varying covariates (linear extrapolation between observed covariate values), but as this model is more biologically plausible, it was chosen as the final model.

The OFV reduced from 2,305.0 to 1,217.5 between the basic and the final model. The inclusion of the covariates resulted in a reduction of the IIV for the PK parameters: the IIV values for CL and V with the basic model were 71.1% and 62.5%, respectively,

and with the final model were 41.8% and 33.5%, respectively. The final PK parameter estimates with uncertainty are reported in Table 2.

Evaluation. (i) Internal model evaluation. Figure 1 shows plots assessing goodness of fit by comparing observations and predictions. A VPC of the final model is shown in Fig. 2.

(ii) External model evaluation. The basic diagnostic plots are presented in Fig. 1 and the VPC performed using the evaluation data set and the final parameters from the PK model without additional fitting in Fig. 2.

Table 3 shows the number of correct predictions (for five different data sets from the evaluation data and pooled results from the cross-validation) for gentamicin trough thresholds of 1 and 2 mg/liter together with prediction errors. In the total data set, containing both paired and unpaired samples, the median (95% CI) PE was 0.0004 (-1.1, 0.8) mg/liter. The MPEs from predictions of trough and peak concentrations (using cross-validations) were 0.03 and 0.19 mg/liter and the RMSEs 1.28 and 2.55 mg/liter, respectively (Table 3). When AUC_{0-t} prediction (from one random sample) was evaluated, the MPE was 14.5 mg · h/liter and the RMSE 30.2 mg · h/liter.

Figure 3 shows the median and the range of PE for this model and previously published gentamicin population PK models.

neoGent. Figure S1 in the supplemental material shows an example of an output from neoGent.

DISCUSSION

A PK model for gentamicin in neonates was developed and evaluated with prospectively collected data. Through its use of mechanistic covariates, the model gave unbiased predictions of trough concentrations from an opportunistic sample. Using this model, concentrations from samples taken at any time can be used to generate informative TDM, potentially eliminating the need for specifically timed trough gentamicin samples and the safety concerns and inconvenience associated with them. An exploratory analysis to evaluate whether such an approach could be used for predicting individual peak concentration and AUC_{0-t} showed

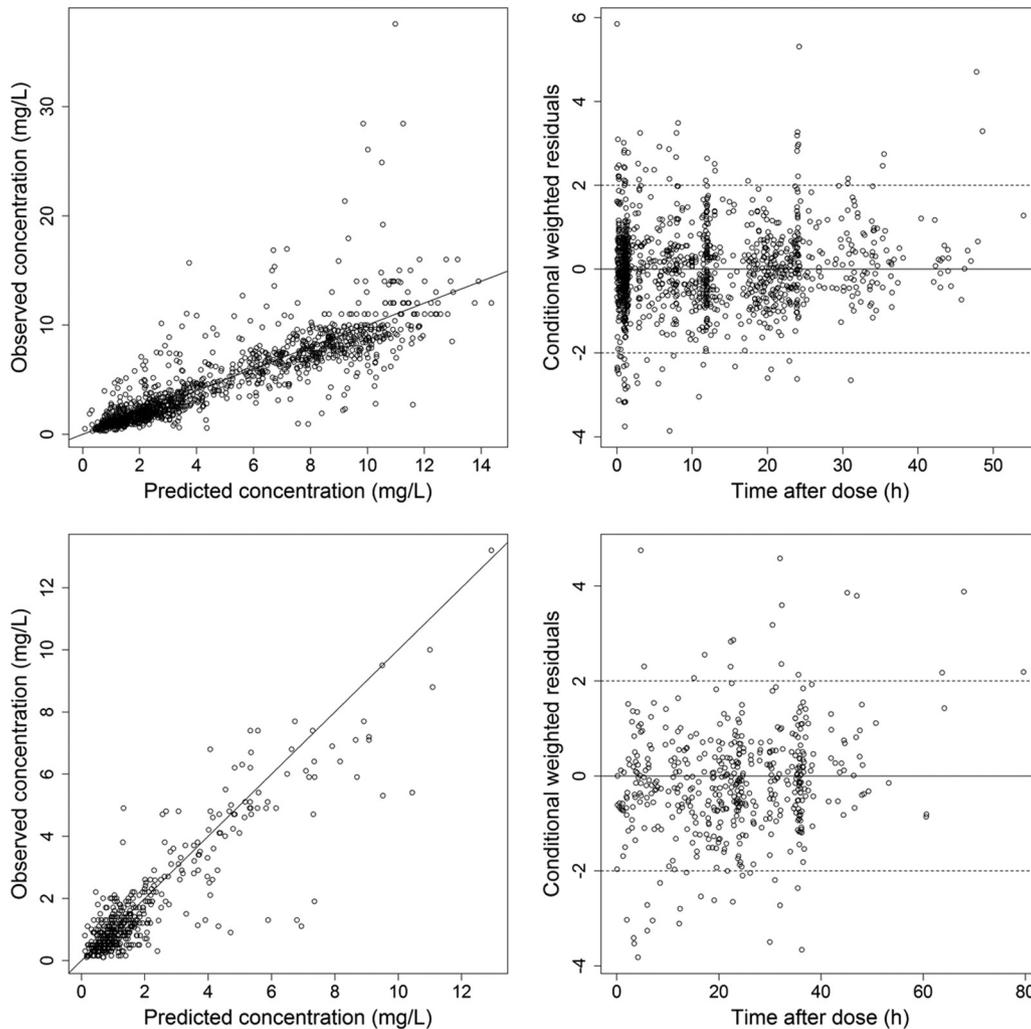


FIG 1 Observed versus population-predicted gentamicin serum concentrations (top left for the model-building data set and bottom left for the evaluation data set) and conditional weighted residuals versus time after dose (top right for the model-building data set and bottom right for the evaluation data set).

that, while the predictions were unbiased, they were relatively imprecise (Table 3).

The small median PE (0.0004 mg/liter) for trough concentrations suggests that the model implemented in neoGent performs well, although some outliers were not captured (range, -2.4 to 1.6 mg/liter). The median prediction errors were in most cases negative (Table 3), indicating that the model slightly overpredicts the trough concentrations (i.e., predicts them to be higher than they are), which might be (from a safety perspective) preferable to underpredicting. Cross-validations confirmed that samples do not need to be taken at a specific time when using this model for TDM, as the predictions of trough concentrations (as determined using an opportunistic sample) were unbiased, with a median PE of -0.04 mg/liter (Table 3). Although we did not test the effect of the sampling time on model predictions, the samples were collected from a wide range of times (0.1 to 53.3 h after the dose), as they would be in routine hospital tests.

Comparison of the developed model with the existing published models showed that the predicted trough concentrations were the least biased (i.e., the median prediction error was the smallest) when our model was used (Fig. 3). However, due to

unavailability of some covariates in our data set, three models were used without all of the covariates (Apgar score [15, 19], sepsis [19], and comedication with dopamine [23]) included, which could explain their worse predictive performance.

The rich data in our model-building data set (6.5 samples per patient) supported a 3-compartment model, where the final estimates for the third compartment were as follows: intercompartmental clearance, 0.3 liters/h/70 kg; peripheral volume of distribution, 148 liters/70 kg. Additionally, the terminal half-life for a typical subject from the prospective evaluation data set (weight, 2.0 kg; PMA, 34.9 weeks; PNA, 6 days; MSCr, 47.0 $\mu\text{mol/liter}$; TSCr, 66.4 $\mu\text{mol/liter}$) was 189.7 h. This could indicate uptake of gentamicin into the renal cortex, and slow excretion from it (47), and is in agreement with previously found evidence of deep tissue accumulation of gentamicin (26, 48).

Unfortunately, many authors were unwilling or unable to share their data, and we managed to obtain data from only two (15, 21) of eight identified studies for our model-building data set. We did obtain one further subsequent data set corresponding to results of assays performed in another pharmacokinetic study in neonates also receiving gentamicin (32). Due to differences with

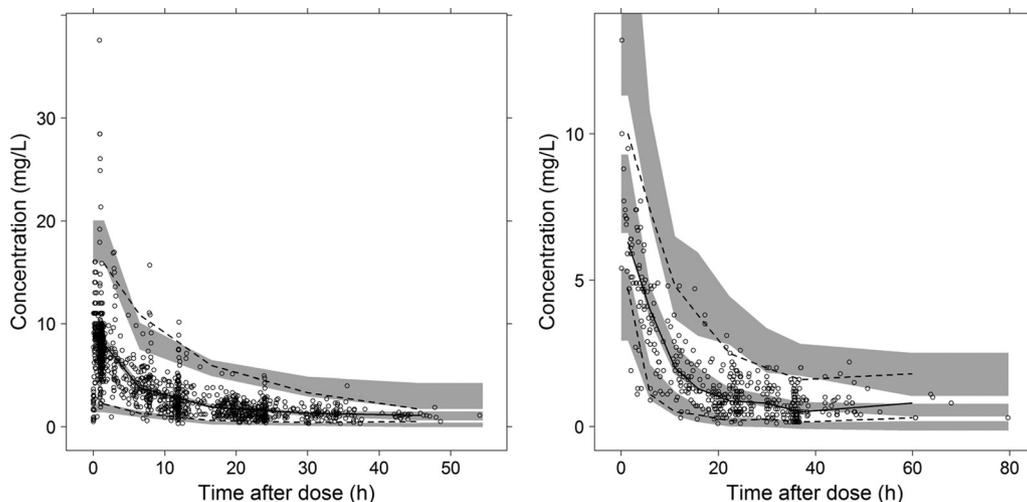


FIG 2 Visual predictive check of 1,000 simulated concentration-time data sets from the final model, using the model-building data set (left) and the evaluation data set (right). Points represent the observations, black lines represent the 2.5th, 50th, and 97.5th percentiles, and the shaded areas represent the 95% confidence intervals of the corresponding predicted gentamicin concentrations.

respect to model structure and parameterization, it was not possible to extract relevant information for model building from the published reports. However, thanks in part to the fact that the data obtained from Nielsen et al. (21) were of such high quality, with multiple samples assayed per patient, our final model described both the model-building data set and the evaluation data set well, as shown in Fig. 1 and 2. The histogram and the q-q plot of the conditional weighted residuals (data not shown) confirmed that they follow a normal distribution pattern. The mean (standard error) final estimates for clearance (CL) and volume of distribution (V) were 6.21 (0.30) liters/h/70 kg and 26.5 (1.11) liters/70 kg, respectively (Table 2). The values of the PK parameters for a typical infant from the model-building data set (weight, 2.12 kg; PMA, 33.0 weeks; PNA, 5.4 days; MSCr, 78 $\mu\text{mol/liter}$; TSCr, 71.4 $\mu\text{mol/liter}$) were 0.077 liters/h and 0.80 liters (and 0.10 liters/h and 0.78 liters for a neonate from the evaluation data set) for CL and V , respectively. These values are in agreement with clearance estimates from previous neonatal studies of gentamicin pharmacokinetics (13, 14, 18, 22–24). The reported value (0.026 liters/h) for CL from Nielsen et al. (21) may appear to be lower, but using our median demographic values in their model, the CL value becomes similar to our estimates (0.095 liters/h). The final estimate for volume of distribution is consistent with the estimates from Fuchs et al. (23) and Botha et al. (24), but it is not in accordance with what was found by García et al. (20) (0.252 liters). The probable reason for this is the use of a different studied population, because when the median weight from our data set was used in their model, the resulting V was 0.968 liters, in agreement with our estimate.

We did not attempt to estimate the allometric power exponents and constants of the maturation function, as the PMA in the studied neonates (23.3 to 43.8 weeks) was insufficient to capture the age when maturation was complete ($\text{PMA}_{50} = 55.4$ weeks [5]); instead, these constants were fixed to the values from another study in which the main focus was renal maturation (5). This type of scaling was used to improve the model usefulness by allowing it to be extrapolated to different subpopulations (for example, neonates with a different weight or a different PMA). In addition to

capturing changes in clearance due to the long-term maturation that extends throughout gestation and into the first 2 years of life, we attempted to capture the short-term changes in clearance that occur after birth regardless of gestational age. A benefit of fixing the long-term maturation based on known relationships between PMA and renal function was that this short-term maturation was apparent with our estimate of PNA_{50} of 40.8 h, indicating that clearance rapidly increases over the first few days of life. In the first day of life, the clearance was at 37% of the value for a typical adult, and it reached 95% by the end of the first month of infancy.

The typical serum creatinine concentration (used in the model) was determined using SCr concentrations from the Jaffe assay, because the same method was used to determine SCr concentrations in the model-building data set. However, for the evaluation data set, assays based on both the Jaffe and enzymatic methods were used to determine SCr concentrations. However, the goodness of fit to the evaluation data set and the predictive performance of the model were good; therefore, no correction factor was included. Also, the enzymatic assay was used in only 16% of patients. Due to the range of the data that were used to determine typical-for-PMA SCr concentrations, the model can be used for a neonate with a PMA of <44 weeks or for a term neonate of <4 weeks of age. The power exponent on the creatinine function was estimated to be -0.13 , meaning that, if observed SCr and typical SCr concentrations were 70 $\mu\text{mol/liter}$ and 60 $\mu\text{mol/liter}$, respectively, clearance would be 2% lower.

Large η shrinkage values indicate that the data did not contain enough information to make a reliable individual estimation. And while the level of shrinkage corresponding to the peripheral volumes of distribution (V_2 and V_3) was high, that corresponding to clearance was relatively low (6.9%) (Table 2), which is important for making predictions of trough gentamicin concentrations and AUC_{0-t} . The η shrinkage corresponding to the central volume of distribution was also relatively low (15%) (Table 2).

Although the main aim of the present study was to evaluate whether the model can predict trough concentrations, the ability of the model to predict the peak gentamicin concentration (from a randomly selected nonpeak sample) was also examined. Cross-

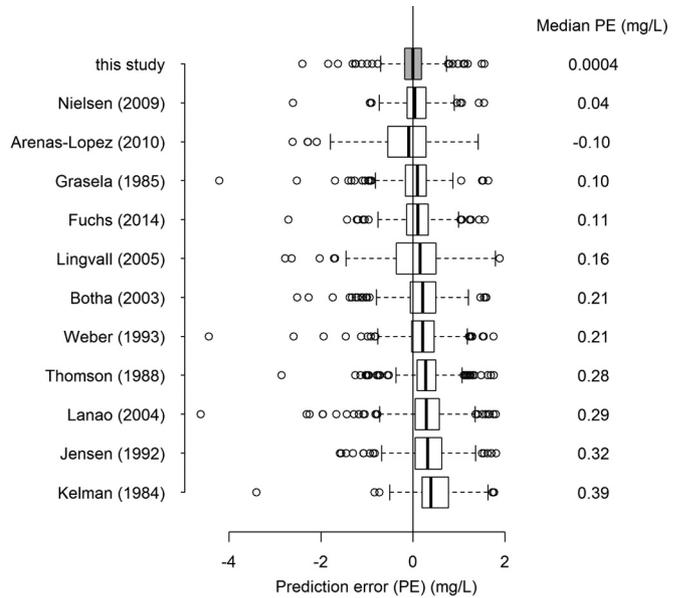


FIG 3 Comparison of predictive performances of the developed model (shaded box plot) and previously published neonatal gentamicin PK models.

validations showed that the median prediction error (95% CI) in predicting peaks was 0.16 (−4.76, 5.01) mg/liter, indicating unbiased but not very precise predictions. This is perhaps not surprising, given that the concentrations determined at a median time after dosing of 19.3 h were used to predict concentrations at a median of 1 h postdose. The prediction of AUC_{0-t} (also from one sample) was similarly unbiased (median prediction error, 10.8 mg · h/liter) but imprecise (95% CI, −24.9 to 62.2 mg · h/liter) (Table 3). However, normalized RMSEs (to the range of observed data) for peak and AUC_{0-t} prediction were 7.0% and 17.6%, respectively, indicating that, considering the range of possible values, the precision is perhaps more acceptable. Target AUC_{0-24} or peak values have not been defined in neonates, and slow clearance and a narrow therapeutic index mean that adjusting doses to target efficacy in this population may not be realistic. However, our model does now give unbiased predictions of both metrics from an opportunistically collected single sample, which should prove useful in future clinical research to define efficacy targets in this age group. At present, due to their imprecision, these predictions (for peak concentration and AUC_{0-t}) should currently be used only for research purposes and not for dose adjustment.

Conclusion. A new gentamicin model has been developed and evaluated with prospectively collected data. We used mechanistic covariate parameterization informed by principles of allometric size scaling, known scaling of glomerular filtration maturation, and standardization for age-expected serum creatinine concentration. This “biological prior” information gave a model with better predictive performance for prospectively collected external data than any previously published gentamicin model. Using this, we developed a software tool, neoGent (see the supplemental material for the provisional stand-alone version and the version implemented in the Web TDM application TDMx [<http://www.tdmx.eu/>] [49]), which can be used to predict when the trough concentration falls below 2 mg/liter and thus to guide the dosing interval. Furthermore, the peak concentration or AUC_{0-24} from any postdose sample can also be predicted with little bias.

TABLE 3 Summary of external evaluation with the evaluation data set^a

Data set	Limit = 1 mg/liter				Limit = 2 mg/liter				PE	MPE	RMSE
	No. of correct results/total no. of results (%)	No. of OP	No. of UP	No. of results (%)	No. of correct results/total no. of results (%)	No. of OP	No. of UP	No. of results (%)			
Paired + unpaired	214/254 (84.3)	20	20	242/254 (95.3)	10	2	0.0004 mg/liter (−1.07, 0.84)	0.007 mg/liter	0.45 mg/liter		
Paired study, ≥1 mg/liter	53/57 (93.0)	3	1	56/57 (98.2)	1	0	−0.04 mg/liter (−0.57, 0.70)	−0.03 mg/liter	0.32 mg/liter		
Paired study ≥2 mg/liter	31/33 (93.9)	2	0	33/33 (100)	0	0	−0.08 mg/liter (−0.50, 0.74)	−0.05 mg/liter	0.35 mg/liter		
Paired study ≥3 mg/liter	19/20 (95.0)	0	1	20/20 (100)	0	0	−0.06 mg/liter (−0.56, 0.82)	−0.02 mg/liter	0.42 mg/liter		
Unpaired	136/161 (84.5)	14	11	155/161 (96.3)	5	1	0.02 mg/liter (−1.11, 0.70)	−0.001 mg/liter	0.43 mg/liter		
XV, paired: study ≥3 mg/liter	478/502 (95.2)	12	12	460/502 (91.6)	21	21	−0.04 mg/liter (−1.77, 3.03)	0.03 mg/liter	1.28 mg/liter		
XV, peaks							0.16 mg/liter (−4.76, 5.01)	0.19 mg/liter	2.55 mg/liter		
AUC_{0-t}							10.8 mg · h/liter (−24.9, 62.2)	14.5 mg · h/liter	30.2 mg · h/liter		

^a“Correct” indicates that the predicted trough concentration value agrees with the measured concentration value (is above/below the limit). OP, overpredictions; UP, underpredictions; PE, prediction error [median (95% confidence interval)]; MPE, mean prediction error; RMSE, root mean square error; XV, cross-validation. Other than the “XV, peaks” data and the AUC_{0-t} data, all results refer to trough prediction evaluations.

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We declare that we have no conflicts of interest.

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