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Gastrointestinal bleeding and intracranial hemorrhage in concomitant users of warfarin and antihyperlipidemics○

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Abstract

Background—Drug interactions, particularly those involving warfarin, are a major clinical and public health problem. Minimizing serious bleeding caused by anticoagulants is a recent major focus of the United States (US) Department of Health and Human Services. This study quantified the risk of gastrointestinal bleeding (GIB) and intracranial hemorrhage (ICH) among concomitant users of warfarin and individual antihyperlipidemics.

Methods—The authors conducted a high-dimensional propensity score-adjusted cohort study of new concomitant users of warfarin and an antihyperlipidemic, among US Medicaid beneficiaries from five states during 1999–2011. Exposure was defined by concomitant use of warfarin plus one of eight antihyperlipidemics. The primary outcome measure was a composite of GIB/ICH within the first 30 days of concomitant use. As a secondary outcome measure, GIB/ICH was examined within the first 180 days of concomitant use.

Results—Among 236,691 persons newly-exposed to warfarin and an antihyperlipidemic, the crude incidence of GIB/ICH was 13.2 (95% confidence interval 12.7 to 13.8) per 100 personyears. Users were predominantly older, female, and Caucasian. Adjusted hazard ratios (aHRs) for warfarin and individual statins were consistent with no association. Warfarin + gemfibrozil was associated with an 80% increased risk of GIB/ICH within the first month of concomitant use (aHR $= 1.8$, 1.4 to 2.4). Warfarin + fenofibrate was associated with a similar increased risk (aHR = 1.8, 1.2 to 2.7), yet with an onset during the second month of concomitant use.

Conclusions—Among warfarin-treated persons, the use of fibrates—but not statins—increases the risk of hospital presentation for GIB/ICH.

Keywords

cohort studies; drug interactions; hemorrhage; Medicaid; pharmacoepidemiology; warfarin

1. INTRODUCTION

Drug-drug interactions (DDIs) are a large and growing clinical and public health problem, especially in older adults, of whom >76% take two or more drugs[1] and >50% take five or more drugs in a given month.[2] Given the high prevalence of polypharmacy in older adults, it is not surprising that known DDIs are responsible for 13% of all adverse drug events[3] and 4.8% of hospital admissions[4] in this population. Relatively few studies have examined the health effects of specific potential DDI pairs in populations, which leave critical knowledge gaps for clinicians, patients, editors and users of DDI compendia, and those who manage and use clinical decision support systems. Recognizing these gaps, attendees of a 2009 stakeholder meeting on DDIs made the conduct of additional research on the health effects of DDIs its principal recommendation.[5]

Anticoagulants have been consistently identified as among the most common causes of serious adverse drug events.[6] Underscoring this, the United States (US) Department of Health and Human Services issued a National Action Plan for Adverse Drug Event Prevention that identified serious bleeding caused by anticoagulants as one of only three adverse drug events specifically targeted because they are common, serious, and potentially preventable.[6] Despite the rapid market uptake of the direct oral anticoagulants, warfarin remains the most widely-prescribed anticoagulant.[7] DDIs involving warfarin are of major concern, since the drug: is commonly-used;[7,8] has a narrow therapeutic index; may interact with almost every therapeutic class;[9] and is the leading cause of adverse drug event-related hospitalizations in older adults.[10] Further, clinical sequelae resulting from over-anticoagulation—particularly gastrointestinal bleeding (GIB) and intracranial hemorrhage (ICH), the vast majority of serious bleeds in warfarin-treated persons[11]—are common and can be fatal.[12] Given that most persons treated with warfarin are older adults, [8] and dyslipidemia is a prevalent comorbidity in older individuals, many such persons are concomitantly-treated with an antihyperlipidemic to manage cardiovascular disease risk. In fact, publically-available data from the US Centers for Disease Control and Prevention[13] suggest that 46% of warfarin-treated older adults take a statin or fibrate.

Drug interactions with warfarin may potentiate bleeding risk via inhibition of hepatic cytochrome P450 (CYP) enzymes responsible for its metabolism. In particular, antihyperlipidemic drugs may inhibit CYP2C9, CYP3A, and CYP1A2,[14,15] each of which is partially responsible for the inactivation of warfarin.[16] Antihyperlipidemics may also displace warfarin from plasma proteins, thereby increasing free plasma concentrations of warfarin;[17-19] these displacement effects may be transient and are of debated clinical relevance. Such mechanisms might result in over-anticoagulation in concomitant users of warfarin and certain antihyperlipidemics. Further, there is some suggestion that statins and fibrates themselves may have antiplatelet and anticoagulant properties.[20-22] Yet, statins alone have not been associated with GIB[23] or ICH.[24,25] Corresponding outcome data on fibrates alone are scant.

We therefore quantified and compared the rates of GIB/ICH among concomitant users of warfarin and individual antihyperlipidemics.

2. METHODS

2.1. Overview and study population

We conducted a high-dimensional propensity score (hdPS)-adjusted cohort study of adult users of warfarin. The cohort consisted exclusively of person-time concomitantly-exposed to warfarin plus one of the following antihyperlipidemics: atorvastatin; fenofibrate (including fenofibric acid[26]); fluvastatin; gemfibrozil; lovastatin; pravastatin; rosuvastatin; or simvastatin. Study data included demographic, enrollment, and healthcare claims from the Medicaid programs of California, Florida, New York, Ohio, and Pennsylvania from 1999– 2011.[27] These states comprise ~38% of the US Medicaid population,[28] with the 13-year dataset recording the experience of more than 69 million cumulative enrollees and nearly 222 million person-years (p-y) of observation. Because a substantive proportion of Medicaid beneficiaries are co-enrolled in Medicare,[29-31] we also obtained Medicare claims to ascertain a more complete picture of enrollees' healthcare.[32] Findings from 1999–2003 data, for a subset of the drug pairs examined herein and using different methods, were reported previously.[33]

2.2. Defining the study cohort

Persons under study were apparent new users of concomitant therapy, i.e., having 12 months of Medicaid enrollment before their first overlapping use of warfarin and an antihyperlipidemic. The day on which users were first co-exposed defined cohort entry. The 12-month period immediately preceding cohort entry served as the baseline period. Use of a fixed baseline period is standard in studies utilizing hdPS methods.[34] Persons entered the cohort without regard to the initiation order of warfarin and the antihyperlipidemic (**Figure 1**).[35]

Persons were excluded if <18 or 100 years of age. Persons with exposure to a nonwarfarin oral anticoagulant (e.g., dabigatran) during the baseline period were not excluded from the cohort, as patients can switch anticoagulants; however, baseline use of a non-warfarin oral anticoagulant was a prespecified covariate included in the PS. Baseline use of a subcutaneous or injectable anticoagulant was treated in the same manner. Neither baseline GIB nor ICH served as an exclusion criterion, as bleeding events can recur;[36] as above, prior GIB/ICH was a prespecified variable included in the PS. Further, we conducted a prespecified secondary analysis limited to incident outcomes.

Follow-up began upon cohort entry and continued until the first occurrence of the following: a) outcome of interest (defined below); b) death, as ascertained from linkage to the Social Security Administration Death Master File (National Technical Information Service: Alexandria, VA); c) the $31st$ day of follow-up (relaxed to the $181st$ day in a secondary analysis); d >15 -day gap in either warfarin or antihyperlipidemic therapy; e) prescription for a non-warfarin oral anticoagulant or antihyperlipidemic different than that upon cohort entry (i.e., indicative of switching to an alternate therapy); f) disenrollment from Medicaid; or g) the end of the dataset. Follow-up time occurring during a period of hospitalization was excluded, although hospitalization did not serve as a censoring event. This exclusion served to minimize immeasurable time bias.[37]

2.3. Exposure and covariate ascertainment

Exposure was defined by the antihyperlipidemic active on the day of cohort entry. The following antihyperlipidemics were excluded because of scant use: cerivastatin; clofibrate; and pitavastatin. Pravastatin served as the reference exposure, as it is a negligible inhibitor of CYP isozymes[38] involved in the metabolism of warfarin.[16] Therefore, pravastatin would not be expected to interact pharmacokinetically with warfarin. Further, statins alone (including pravastatin[39]) have not been shown to, themselves, impact bleeding risk. [23,25,39,40]

Potential confounders included prespecified variables and those identified via empiric methods, both of which informed the PS. Prespecified variables included demographics, baseline measures of intensity of healthcare utilization, baseline drug exposures, and baseline comorbidities. Empiric covariates included those identified during baseline via a high-dimensional approach,[41,42] which ranks and selects potential confounders (or proxies thereof) based on their empirical associations with exposure and outcome (see specifications in **Appendix Table 1**).

2.4. Outcome ascertainment

The outcome of interest was a composite of GIB or ICH within 30 days of cohort entry (relaxed to within 180 days in a secondary analysis). Operational definitions, including quantitative measures of algorithm performance,[33,43-45] are presented in **Table 1**. GIB and ICH are the most common types of serious bleeding attributed to anticoagulants (capturing the vast majority of all events and nearly all serious events) and outcomes that can be fatal $[11,12]$ —this was the rationale for studying a composite outcome of bleeding from these sites.

2.5. Statistical analysis

We calculated descriptive statistics for baseline variables and calculated incidence and unadjusted association measures, the latter via Cox proportional-hazards models. We utilized the hdPS approach to reduce the impact of measured potential confounders. However, as we wished to compare multiple antihyperlipidemic drugs to a common active comparator, matching on PS was impractical, and the hdPS algorithm has so far been developed only for pairwise comparisons.[42,46] As described below, we therefore used pairwise hdPS to identify potential confounders for each antihyperlipidemic drug versus pravastatin and included all such empirically-identified variables (plus prespecified variables) in a multinomial PS model. We first used the hdPS program[42,46] to identify the 200 most prevalent diagnosis, procedure and drug codes (excluding drug codes indicative of warfarin or antihyperlipidemic prescribing) in each of nine data dimensions, to assess their associations with the antihyperlipidemic of interest versus pravastatin, and to assess their associations with the outcome. We then used these associations to select the top 500 codes with the largest potential for causing confounding. Because of the large number of variables in the final multinomial PS model, empirically-identified covariates did not include measures of frequency (i.e., sporadic, frequent) as generated by the hdPS program. Then, the union of all confounders arising from the seven sets of 500 hdPS-identified variables (one for each antihyperlipidemic versus pravastatin) were included in the multinomial PS.

Prespecified covariates included in the multinomial PS model are presented in **Appendix Table 2**. The multinomial PSs were modeled using multinomial logistic regression,[47] generating for each subject the predicted probability of receiving each antihyperlipidemic drug. These PSs were then included in the outcome model as continuous covariates.[48] PSadjusted HRs and 95% CIs were calculated via Cox proportional-hazards regression. Association measures were examined overall within the first 30 days of follow-up and also stratified as three prespecified, mutually exclusive time periods (i.e., 1–10 days, 11–20 days, and 21–30 days).

Numerous prespecified secondary analyses (**Table 2**) were conducted to assess the robustness of our primary findings. Primary and secondary analyses were conducted using SAS v9.4 (SAS Institute Inc.: Cary, NC). The research described herein was approved by the institutional review board of the University of Pennsylvania.

3. RESULTS

3.1 Cohort characteristics and outcome frequency

We identified 236,691 concomitant users of warfarin and an antihyperlipidemic of interest; their baseline characteristics are presented in **Table 3**. Overall, users were predominantly female (63.0%) and Caucasian (54.1%), with a median age of 71.0 years. In the primary analysis examining events within the first 30 days of concomitant use, these individuals contributed 18,135p-y of concomitant exposure, during which we identified 2,401 GIB/ICH outcomes (unadjusted incidence rate = 13.2 per 100p-y [95% confidence interval 12.7 to 13.8]). Findings stratified by individual antihyperlipidemics are presented in **Table 3**. In the secondary analysis examining events within the first 180 days of concomitant use, we identified 4,621 outcomes during 51,905p-y of concomitant exposure (unadjusted incidence rate $= 8.9$ per 100p-y [8.6 to 9.2]). These incidence rates are similar to prior findings, particularly within older adult populations.[49-51]

3.2 Measures of association: primary analysis

The hdPS algorithm identified 688 covariates for inclusion in the PS model (**Appendix Table 3**). Among these, 34 variables occurred very infrequently $(N < 10$ for at least one antihyperlipidemic exposure group) and were excluded to avoid model instability. Therefore, the multinomial PS model included 654 empirically-identified covariates; each model also included 24 predefined covariates (**Appendix Table 2**). Crude hazard ratios (HRs) are presented in **Table 3**; PS-adjusted HRs are presented in **Table 3** and **Figure 2**. Notably, concomitant use of warfarin with gemfibrozil (vs. pravastatin) was associated with an elevated rate of GIB/ICH within 30 days (adjusted $HR = 1.79$, 1.36 to 2.35)—an association most pronounced within the first 10 days of concomitant use (adjusted $HR = 2.09$, 1.35 to 3.24).

3.3 Measures of association: secondary analyses

For the secondary analysis examining events within the first 180 days of concomitant use, warfarin + gemfibrozil and warfarin + fenofibrate (each vs. pravastatin) were associated with an elevated rate of GIB/ICH (adjusted HRs 1.53, 1.23 to 1.90 and 1.27, 1.03 to 1.55,

respectively). Crude and PS-adjusted HRs for all antihyperlipidemics are presented in **Appendix Table 4** and **Figure 3**, respectively. Results from the remaining secondary analyses (detailed in **Table 2**) are presented in **Table 4**.

4. DISCUSSION

We examined serious bleeding associated with potential drug-drug interactions between warfarin and individual antihyperlipidemics. The incidence of GIB/ICH among concomitant warfarin and antihyperlipidemic users was about 13 per 100p-y. While this rate is consistent with some prior findings, [49-51] it is generally many-fold greater than that reported in other cohorts.[52-57] This is likely explained by: a) the left-skewed age distribution of our study population (**Appendix Figure 1**); b) our study of economically-disadvantaged individuals with a generally poor health status and more barriers to timely care;[58] c) our inclusion of persons with dual enrollment in Medicaid and Medicare, a group enriched with the sickest and poorest individuals served by the US Centers for Medicare and Medicaid Services;[59] d) the large proportion of our population newly-initiating warfarin upon cohort entry (**Table 3**), as the period immediately following warfarin initiation carries the highest bleeding risk; [60] and e) differences in outcome definitions across studies. As only 3.1% of warfarin users had a history of GIB/ICH during the baseline period (**Table 3**), this seems unlikely to contribute to our high incidence rate.

We found an increased hazard of GIB/ICH during the *first* month of concomitant use of warfarin and gemfibrozil. The highest risk period was within the first 10 days, during which the adjusted HR was more than 2-fold as large as that for warfarin + pravastatin. Interestingly, we found an increased hazard of GIB/ICH during the second month of concomitant use of warfarin and fenofibrate, during which the adjusted HR was 1.8-fold as large as that for warfarin + pravastatin. One could hypothesize that the difference in onset of these interactions is related to the disparate half-lives of gemfibrozil (1.5 hours) and fenofibrate (19.6–26.6 hours).[61] There was no corresponding increased hazard in these months (and up through the first six months) among concomitant users of warfarin and individual statins. Importantly, these findings were broadly consistent across all secondary analyses.

The null associations for statins are consistent with case-control findings by Douketis et al in which recent statin use (considered collectively) among warfarin users was not associated with GIB/ICH (adjusted odds ratio $[OR] = 1.04$, 0.74 to 1.48). Those investigators did not examine fibrates. Our positive association for gemfibrozil is consistent with case-control findings by Schelleman et al (adjusted OR for GIB within the first 30 days of concomitant use $= 1.96, 1.19$ to 3.24). That earlier study using a subset of the data reported here was too small to generate a precise estimate for fenofibrate, reporting an adjusted OR of 2.07 (0.91 to 4.69) during days 31–60—consistent with our fenofibrate result. Interestingly, and discordant from our findings, Schelleman *et al* further reported an increased GIB risk within the first 30 days of concomitant use for simvastatin and atorvastatin (adjusted $ORs = 1.33$) [1.00 to 1.78] and 1.29 [1.04 to 1.61], respectively). Yet, this prior study adjusted for a small number of potential confounders and ORs were noticeably attenuated when comparing

minimally- to more fully-adjusted models. This may suggest that the more complete control of potential confounders would have rendered their statin findings null.

A potential explanation for the positive association with fibrates (but not statins) could be a healthy user effect among our pravastatin-exposed reference group; statin users may be healthier than fibrate users and, by virtue of their health, have an apparent lower GIB/ICH risk.[62,63] In fact, compared to fibrate users, we found that statin users were less likely to have an obesity diagnosis and to have preexisting diabetes, coagulation defects, and liver disease. Yet, we found that statin users were older, more likely to reside in a nursing home, and more likely to be dually-enrolled in Medicaid and Medicare. Regardless, we adjusted for baseline differences by including these and many other covariates in our PS. This minimizes the likelihood that a healthy user effect explains our findings.

Evidence supporting the positive association with fibrates includes clinical observations of international normalized ratio elevations and serious bleeding among concomitant users of warfarin and fibrates.[19,61,64-67] These events have occurred within the first few days of concomitant use through a month thereafter, a time frame broadly consistent with our findings. A number of mechanisms underlying this interaction, while neither fully elucidated nor agreed upon, have been proposed—including, as examples: inhibition of warfarin metabolism via CYP2C9; displacement of warfarin from protein binding sites; and inherent anticoagulant and antiplatelet effects of fibrates.[20,64,66]

Our results suggest that the apparent drug interaction between warfarin and fibrates is unlikely to be mediated primarily by CYP2C9 inhibition. Support for this conclusion is as follows. First, neither gemfibrozil nor fenofibrate are thought to be clinically-important inhibitors of CYP2C9.[68] Second, concomitant use of warfarin with rosuvastatin or fluvastatin, important competitive inhibitors of CYP2C9,[14,69] was not associated with an increased rate of GIB/ICH. Third, the increase in GIB/ICH rate with concomitant use of warfarin and fenofibrate was delayed, and interactions involving enzymatic inhibition are usually rapid-onset interactions.[70] This apparent drug interaction is also unlikely mediated by protein binding displacement. Such interactions are generally considered transient because the increased amount of unbound drug is readily metabolized or eliminated.[19] Further, concomitant use of warfarin with simvastatin or fluvastatin, examples of highly protein bound statins,[71] was not associated with an increased rate of GIB/ICH. Finally, the potential mechanism of fibrates' inherent anticoagulant and antiplatelet effects is intriguing, but these effects have also been reported with statins.[20-22] Future work should elucidate the mechanism(s) underlying this apparent drug interaction.

Our study has notable strengths. It is the largest to date to examine the association between warfarin + antihyperlipidemics and GIB. Further, it is the first pharmacoepidemiologic study of this interaction to: a) include ICH as a clinical endpoint, while elucidating antihyperlipidemic agent-specific rates; b) consider a warfarin-triggered drug interaction (**Figure 1**); c) include Medicare Part D prescription data; d) quantify agent-specific rates of GIB/ICH in time windows within the first 30 days of use; and e) fully quantify the rate of GIB/ICH among concomitant users of warfarin + fenofibrate and warfarin + fluvastatin. Our use of an active comparator, hdPS methods, and sensitivity analyses serves to mitigate

confounding. Further, our large sample sizes allow for the examination of the time-course of the interactions. Finally, our algorithms to identify GIB and ICH have very good positive predictive values.

Our study also has limitations. First, we did not have access to biosamples and were therefore unable to examine genetic CYP polymorphisms. Second, we lacked data on adherence to prescribed warfarin and antihyperlipidemic therapies. Third, although we were able to identify and consider orders for laboratory monitoring (e.g., international normalized ratio), Medicaid and Medicare claims do not include laboratory results. Fourth, administrative databases may poorly capture some lifestyle behaviors and nonprescription therapies that may modify bleeding risk. Regardless, such factors seem unlikely to differ substantially by antihyperlipidemic exposure. Finally, our results may not be generalizable beyond a US Medicaid population. Nevertheless, this population was specifically chosen because of its inherent vulnerability and inclusion of large numbers of women and minorities—groups typically understudied. Biological associations identified in Medicaid populations are often replicated in commercially insured populations and *vice versa*.[27]

Warfarin is highly prone to drug interactions via numerous complex mechanisms.[72] By far the most important consequence of such interactions is serious bleeding—an outcome of significant clinical and public health concern that is feared by clinicians, patients, and their relatives. We found that concomitant therapy with warfarin and a fibrate, but not statin, is associated with an increased rate of GIB/ICH. The mechanism underlying this apparent drug-drug interaction needs further elucidation, but is unlikely to solely involve a pharmacokinetic interaction mediated by CYP inhibition or displacement of binding from plasma proteins. Clinicians should be attuned to both immediate- and delayed-onset bleeding in their patients on this drug combination.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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Method 1 - Combination Triggered: subject newly-initiates warfarin and antihyperlipidemic on the same day

Method 2 - Warfarin Triggered: subject with ongoing antihyperlipidemic treatment newly-initiates warfarin

Method 3 - Antihyperlipidemic Triggered: subject with ongoing warfarin treat ment newly-initiates an antihyperlipidemic

time window of interest for secondary analysis

Legend

↓ cohort entry begins

antihyperlipidemic prescription dispensing

 \circ warfarin prescription dispensing

Figure 1.

Three methods by which concomitant warfarin and antihyperlipidemic users could enter the study cohort

Figure 2.

Propensity score-adjusted hazard ratios (HRs) for association between warfarin + antihyperlipidemic drug (vs. pravastatin) and gastrointestinal bleeding/intracranial hemorrhage within 30 days of concomitant use—overall and within three prespecified time windows

Propensity score-adjusted hazard ratios (HRs) for association between warfarin + antihyperlipidemic drug (vs. pravastatin) and gastrointestinal bleeding/intracranial hemorrhage within 180 days of concomitant use—overall and within four prespecified time windows

Table 1

Operational definition of composite outcome of interest Operational definition of composite outcome of interest

Prespecified secondary analyses Prespecified secondary analyses

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* events with an International Classification of Diseases 9th Revision Clinical Modification discharge diagnosis code indicative of intracranial injury (with or without skull fracture) during the same events with an International Classification of Diseases 9th Revision Clinical Modification discharge diagnosis code indicative of intracranial injury (with or without skull fracture) during the same hospitalization as the putative outcome, defined by 800.X-804.X (fractures: fracture of skull) or 850.X-854.X (intracranial injury excluding those with skull fracture) hospitalization as the putative outcome, defined by 800.X–804.X (fractures: fracture of skull) or 850.X–854.X (intracranial injury excluding those with skull fracture)

 $\tau_{\rm{venous~ttronboembolism~vs.~atrial~fbrillation/flutter~vs.~valvular heart disease~vs.~other~indication}$ venous thromboembolism vs. atrial fibrillation/flutter vs. valvular heart disease vs. other indication

t new warfarin initiator (i.e., first warfarin dispensing on cohort entry) vs. recent warfarin initiator (i.e., first warfarin dispensing within 45 days prior to cohort entry) vs. chronic warfarin user (i.e., first ‡ new warfarin initiator (i.e., first warfarin dispensing on cohort entry) vs. recent warfarin initiator (i.e., first warfarin dispensing within 45 days prior to cohort entry) vs. chronic warfarin user (i.e., first warfarin dispensing greater than 45 days prior to cohort entry) warfarin dispensing greater than 45 days prior to cohort entry)

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Table 3

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ASA = aspirin; CI = confidence interval; CPT-4 = Current Procedural Terminology-4; COX = cyclooxygenase; CYP = cytochrome P450; GIB = gastrointestinal bleeding; ICH = intracranial hemorrhage; NSAID = nonsteroidal anti-infl ASA = aspirin; CI = confidence interval; CPT-4 = Current Procedural Terminology-4; COX = cyclooxygenase; CYP = cytochrome P450; GIB = gastrointestinal bleeding; ICH = intracranial hemorrhage; NSAID = nonsteroidal anti-inflammatory drug; Q = quartile

* cell suppressed to comply with Centers for Medicare and Medicaid Services privacy policy

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procedure codes; # unique inpatient CPF-4/HCPCS procedure codes; # outpatient ICD-9 procedure codes; # other setting ICD-9 diagnosis codes; # unique other setting ICD-9 diagnosis codes; # other setting ICD-9 procedure code procedure codes; # unique impatient ICD-9 procedure codes; # unique codes; # unique outpatient ICD-9 procedure codes; # unique codes; # unique outer, setting ICD-9 diagnossis codes; # dies; # different CD-9 diagnossis code the following healthcare utilization covariates were excluded from presentation the table, as their median values were zero for the majority of antihyperlipidemics: # inpatient ICD-9 procedure codes; # unique inpatient IC the following healthcare utilization covariates were excluded from presentation the table, as their median values were zero for the majority of antihyperlipidemics: # inpatient CD-9 procedure codes; # inpatient ICD-9 proc unique other setting ICD-9 procedure codes unique other setting ICD-9 procedure codes

Table 4

Summary of findings from prespecified secondary analyses

aHR = propensity score-adjusted hazard ratio; GIB = gastrointestinal bleeding; ICH = intracranial hemorrhage; PS = propensity score

* most non-significant findings not shown; data available from the authors

† comparison vs. pravastatin

 $t_{\rm p}^{\ddag} = 0.053$