Papers

Risk of ischaemic stroke in people with migraine: systematic review and meta-analysis of observational studies

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Abstract

Objective To explore the association between migraine and risk of ischaemic stroke.

Design Systematic review and meta-analysis.

Data sources Observational studies published between 1966 and June 2004 (identified through Medline and Embase) that examined the association between migraine and risk of ischaemic stroke.

Results 14 studies (11 case-control studies and 3 cohort studies) were identified. These studies suggest that the risk of stroke is increased in people with migraine (relative risk 2.16, 95% confidence interval 1.89 to 2.48). This increase in risk was consistent in people who had migraine with aura (relative risk 2.27, 1.61 to 3.19) and migraine without aura (relative risk 1.83, 1.06 to 3.15), as well as in those taking oral contraceptives (relative risk 8.72, 5.05 to 15.05).

Conclusions Data from observational studies suggest that migraine may be a risk factor in developing stroke. More studies are needed to explore the mechanism of this potential association. In addition, the risk of migraine among users of oral contraceptives must be further investigated.

Introduction

Migraine is the most common type of headache in young adults, with an estimated prevalence of 4% before puberty and as high as 25% in women by their mid to late 30s.¹ Some observational studies have shown an increase in the risk of stroke among people with a history of migraine,² but others have failed to find this association.³ The mechanism of this potential association is believed to be in part through platelet hyperaggregability and the reduction in cerebral blood flow that usually occurs in migraine with aura.⁴ A potential association between the risk of stroke and migraine is an important public health concern, especially in young women who use oral contraceptives, which by itself may be an independent risk factor for stroke.⁵

We sought to explore the association between migraine and ischaemic stroke by conducting a meta-analysis. Specifically, we set out to quantify the risk of ischaemic stroke among people with migraine (with and without aura), as well as to quantify this risk among different age groups and users of oral contraceptives.

Methods

Search strategy

We systematically searched Medline (1966-June 2004) and Embase (1974-June 2004) for both English and non-English language articles by entering "brain ischemia," "cerebrovascular accidents," "cerebrovascular disorders," "cerebral infarction," "ischemic attack," "migraine," and "oral contraceptives" as both medical subject heading (MeSH) terms and text words. We then retrieved all relevant articles as determined by consensus among the authors and searched the reference lists of retrieved articles to find other potentially relevant articles.

Data extraction

We included studies if they used clear diagnostic criteria for migraine; had clearly stated diagnostic criteria for the outcome of ischaemic stroke; had controlled for potential confounders by using risk adjustment in the analysis or matching in the study design; and provided odds ratios or relative risks and 95% confidence intervals or provided enough data to allow us to calculate these numbers. In order to quantify the risk of stroke among people with migraine who were also using oral contraceptives, the studies had to provide data for migraine patients who were exposed to oral contraceptives compared with those who were not taking these agents. We also scored the quality of the studies by using a 10 point scale adapted from a recently published quality scale for observational studies (five criteria ranked 0, 1, or 2) and stratified the studies by score (a score of 7 or above indicates high quality; a score of 6 or below indicates low quality).⁶

Data analysis

We weighted log relative risks for cohort studies or odds ratios by the inverse of their variances to obtain a pooled measure of the relative risks. We used the assumption that an odds ratio from a case-control study approximates the relative risk in a cohort study. We combined cohort studies and case-control studies in the absence of statistical heterogeneity. When results from the fixed and random effects models were different, we presented the second as it represents a more conservative approach. We tested for heterogeneity by using the DerSimonian and Laird Q statistic. We also measured heterogeneity by using the R_i statistic, which quantifies the proportion of the total variance that is due to between study variance.⁷ We assessed publication bias graphically by using a funnel plot as well as quantitatively with Egger's regression.⁸

In order to assess publication bias further we also did a sensitivity analysis, in which we assessed the potential effect of publication bias on our pooled relative risks by using three assumptions: published studies included in our meta-analysis represent only half of the studies ever conducted; the remaining unpublished studies have found null associations (that is, relative risk = 1); the unpublished studies included as many cases and controls as the average of the published studies. This approach has been used in previously published meta-analysis.⁹ We used HEpiMA version 2.13 for all analyses.¹⁰

	Relative risk (95% CI) of migraine			Adjustment for covariates	Cases:controls or	Cases:controls with
First author, year	Any	With aura	Without aura	and confounders	cohort size	migraine
Case-control studies						
Collaborative group, 1975 ¹¹	2.0 (1.2 to 3.3)	NS	NS	Age, contraceptives, smoking	430:429	105:256
Henrich, 1989 ¹²	1.8 (0.9 to 3.6)	2.6 (1.1 to 6.6)	1.3 (0.5 to 3.6)	Not specified	89:178	17:20
Marini, 1993 ¹³	1.91 (1.05 to 3.5)	14.85 (1.8 to 124)	1.6 (0.9 to 3.0)	Diet, obesity, alcohol, smoking, contraceptives, hypertension, diabetes	308:308	46: 25
Tzourio, 1993 ¹⁴	1.3 (0.8 to 2.3)	1.3 (0.5 to 3.8)	0.8 (0.4 to 1.5)	Age, sex, hypertension, smoking, diabetes, contraceptives	212:212	28:30
Tzourio, 1995 ²	3.5 (1.8 to 6.4)	6.2 (2.1 to 18)	3.0 (1.5 to 5.8)	Age, sex, hypertension, smoking, diabetes, contraceptives	72:173	43:52
Lidegaard, 1995 ¹⁵	2.8 (2.00 to 4.25)	NS	NS	Pregnancy, diabetes, hypertension, earlier thrombotic disease, other disease	497:1370	64:66
Carolei, 1996 ³	1.9 (1.1 to 3.1)	1.0 (0.5 to 2.0)	8.6 (1.0 to 75)	Hypertension, smoking, cholesterol, diabetes, obesity, contraceptives, alcohol	308:591	46:54
Haapaniemi, 1997 ¹⁶	2.12 (1.05 to 2.95)	NS	NS	Hypertension, cardiac disease, current smoking, diabetes, alcohol	506:345	86:42
Chang, 1999 ¹⁷	3.54 (1.30 to 9.61)	2.97 (0.66 to 13.5)	2.97 (0.66 to 13.5)	Contraceptives, hypertension, smoking	291:736	71:88
Donaghy, 2002 ¹⁸	2.98 (1.24 to 7.19)	NS	NS	Age, smoking, hypertension, heart condition, education level, social class	86:214	NS
Schwaag, 2003 ¹⁹	2.16 (1.16 to 3.82)	NS	NS	Not specified	160:160	37:20
Cohort studies						
Buring, 1995 ²⁰	2.00 (1.10 to 3.64)	NS	NS	Age, smoking, hypertension, cholesterol, diabetes, heart condition, exercise frequency	1479:20 481	19:194
Merikangas, 1997 ²¹	2.1 (1.5 to 2.9)	NS	NS	Age, sex, hypertension, diabetes, heart condition, alcohol, smoking	423:11 777	NS
Nightingale, 2004*22	2.33 (1.04 to 5.21)	NS	NS	Age, hypertension, physician	190:1129	16:44
Cross sectional study						
Kruit, 2004 ²³	7.1 (0.9 to 55)	13.7 (1.7 to 112)	2.3 (0.2 to 23)	Age, sex, smoking, hypertension, cholesterol, alcohol	161:140	NS

Table 1 Relative risks of ischaemic stroke according to type of migraine

NS=not specified.

*Nested case-control study within a well defined cohort

Results

Our search resulted in 11 case-control studies,^{2-3 11-19} three cohort studies,²⁰⁻²² and one cross sectional study²³ (table 1). We excluded the cross sectional study from the analysis, as the timing of diagnosis of migraine with respect to development of ischaemic stroke was difficult to infer in this study. In total, we included 14 studies in the meta-analysis,^{2-3 11-22} Six studies provided data on the risk of ischaemic stroke and migraine with and without aura.^{2 3 12-14 17} The age of the participants in the included studies ranged from 15 to 84 years. The incidence of stroke among people with migraine in the three cohort studies ranged from 3.56 to 350 cases per 100 000 person years.²⁰⁻²²

The pooled relative risk for ischaemic stroke among patients with any type of migraine headache was 2.16 (95% confidence interval 1.89 to 2.48) (table 2, fig 1). The relative risks for people with migraine with and without aura were 2.27 (1.61 to 3.19) and 1.83 (1.06 to 3.15). This risk did not differ when we stratified our analysis by age (table 2). Users of oral contraceptives had an approximately eightfold increase in the risk of stroke compared with those not using these agents.

We did not find evidence of publication bias either graphically from the funnel plot (fig 2) or quantitatively (P = 0.685 for Egger's test of asymmetry). The pooled relative risk

for the sensitivity analysis still showed a significant increase in the risk of stroke (relative risk 1.43, 1.21 to 1.68).

Discussion

The results of our study strongly suggest that migraine may be an independent risk factor for stroke. The magnitude of this risk remained the same across all studies (case-control and cohort) as well as in those that provided data on migraine with aura, migraine without aura, and oral contraceptive users. The risk of stroke among oral contraceptive users is very high, although these data come from only three studies. Other studies in women who were users of oral contraceptives have shown that those with a history of migraine have twice the likelihood of developing an ischaemic stroke compared with those without migraine (relative risk 2.15, 95% confidence interval 0.85 to 5.45).24 Given that use of oral contraceptives is prevalent among young women, the potential risk of stroke among women with migraine who are also users of oral contraceptives must be further investigated. Possible mechanisms for this association include irregularities in blood flow,²⁵ cardiac abnormalities,²⁶ and abnormal production of prostaglandins as well as noradrenergic or cholinergic transmitters and receptors.27

	No of studies	Relative risk (95% CI)	Ri*	P value†
Migraine (any)				
All studies	14	2.16 (1.89 to 2.48)	0.00	0.77
Case-control studies	11	2.18 (1.86 to 2.56)	0.00	0.51
Cohort studies	3	2.10 (1.61 to 2.75)	0.00	0.96
Migraine with aura				
Case-control studies	7	2.27 (1.61 to 3.19)	0.49	0.08
Migraine without aura				
Case-control studies	6	1.83 (1.06 to 3.15)	0.60	0.04
Migraine among oral contraceptive users				
Case-control studies	3	8.72 (5.05 to 15.05)	0.26	0.28
Migraine among men and women <45 years				
Case-control studies	9	2.36 (1.92 to 2.90)	0.07	0.38
Migraine among women <45 years				
Case-control studies	7	2.76 (2.17 to 3.52)	0.00	0.82

*Proportion of the total variance due to between study variance. Large values (>0.75) indicate large heterogeneity between studies; small values (<0.4) indicate lack of heterogeneity.⁷ †DerSimonian and Laird Q statistic.

Limitations

Our meta-analysis is subject to several limitations. Firstly, most of the studies in the meta-analysis were case-control studies, which are subject to recall bias. Cases may have been more likely to classify their headache as a migraine headache (as opposed to a tension headache). Secondly, although almost all the studies controlled for appropriate confounders in either the design or the analysis, some important confounders may not have been controlled for. For example, use of antihypertensive drugs may be a potential confounder, as these drugs may be used to prevent migraine attacks as well as future strokes. However, use of these drugs among people with migraine would probably have produced a decrease in the risk of stroke, which was not seen in our study. Antiphospholipid antibodies have been thought to be linked to stroke and possibly to migraine.28 None of the studies included in the meta-analysis provided information on this potential confounder. People who have migraine attacks might be less likely to be diagnosed as having a stroke, as the symptoms of migraine may be confused with those of stroke. Although this remains a possibility, many of the studies included in the meta-analysis used strict criteria to define ischaemic stroke, including duration of symptoms of at least 24 hours, as well as confirmation of diagnosis by brain imaging or autopsy. Finally, we could not infer from the studies a temporal relation between the onset of migraine and the diagnosis of stroke.

Conclusion

Data from observational studies suggest that migraine may be a risk factor in developing stroke. More studies are needed to explore the mechanism of this potential association. The risk of migraine among users of oral contraceptives must be further investigated.

Contributors: ME, BT, and FCI initiated the project. ME, BT, FCI, and AS screened and extracted the data. BT analysed the data. All authors participated in discussing the results and writing the paper. ME is the guarantor.

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Competing interests: None declared. Ethical approval: Not needed.

Eulicai approvai. Not needed.

- 1 Silberstein SD. Migraine. Lancet 2004;363:381-91.
- 2 Tzourio C, Tehindrazanarivelo A, Iglesias S, Alperovitch A, Chedru F, d'Anglejan-Chatillon J, et al. Case-control study of migraine and risk of ischaemic stroke in young women. *BMJ* 1995;310:830-3.
- 3 Carolei A, Marini C, De Matteis G. History of migraine and risk of cerebral ischaemia in young adults. *Lancet* 1996;347:1503-6.

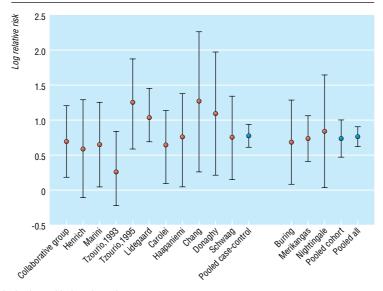


Fig 1 Forest plot of the studies of migraine and ischaemic stroke

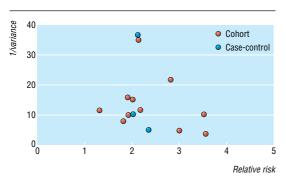


Fig 2 Funnel plot of the studies of migraine and ischaemic stroke

- Woods RP, Iacoboni M, Mazziotta JC. Brief report: bilateral spreading cerebral hypop-4 erfusion during spontaneous migraine headache. N Engl J Med 1994;331:1689-92. Lidegaard O, Kreiner S. Contraceptives and cerebral thrombosis: a five-year national
- 5
- Ease-control Study. Contracting 2002;65:197-205.
 Bhutta A, Cleves M, Casey P, Cradock M, Anand KJS. Cognitive and behavioral outcomes of school-aged children who were born preterm: a meta-analysis. JAMA 6 2002:288:728-37.
- 7 Takkouche B, Cadarso-Suarez C, Spiegelman D. Evaluation of old and new tests of heterogeneity in epidemiologic meta-analysis. *Am J Epidemiol* 1999;150:206-15. Egger M, Davey Smith G, Schneider M, Minder C. Bias in meta-analysis detected by a
- 8 simple, graphical test. *BMJ* 1997;315:629-34. Hernan MA, Takkouche B, Caamano-Isorna F, Gestal-Otero JJ. A meta-analysis of cof-9
- fee drinking, cigarette smoking, and the risk of Parkinson's disease. Ann Neurol 2002;52:276-84.
- 10 Costa-Bouzas J, Takkouche B, Cadarso-Suarez C, Spiegelman D. HEpiMA: software for the identification of heterogeneity in meta-analysis. Comput Methods Programs Biomed 2001;64:101-7.
- 11 Collaborative Group for the Study of Stroke in Young Women. Oral contraceptives and stroke in young women. JAMA 1975; 231:718-22.
- 12 Henrich JB, Horwitz RI. A controlled study of ischemic stroke risk in migraine patients. J Clin Epidemiol 1989;42:773-80.
- 13 Marini C, Carolei A, Roberts RS, Prencipe M, Gandolfo C, Inzitari D, et al. Focal cerebral ischemia in young adults: a collaborative case-control study. Neuroepiden 1993:12:70-81.

What is already known on this topic

Studies have suggested that migraine may be a risk factor for stroke, but results from these studies have been inconsistent

What this study adds

Migraine both with and without aura may be an independent risk factor for ischaemic stroke

This risk is higher among oral contraceptive users and younger adults (<45 years)

- 14 Tzourio C, Iglesias S, Hubert JB, Visy JM, Alperovitch A, Tchindrazanarivelo A, et al. Migraine and risk of ischemic stroke: a case-control study. *BMJ* 1993;307:289-92.
- Lidegaard O. Oral contraceptives, pregnancy and the risk of cerebral thromboembo-lism: the influence of diabetes, hypertension, migraine and previous thrombotic disease. *Br J Obstt Gynaecol* 1995;102:153-9. 15
- Haapaniemi H, Hillbom M, Juvela S. Life-style associated risk factors for acute brain infarction among persons of working age. Stroke 1997;28:26-30.
- mlarction among persons of working age. *Subject 1991*, 40:20-30.
 17 Chang CL, Donaghy M, Poulter N. Migraine and stroke in young women: case-control study. *BMJ* 1999;318:13-8.
 18 Donaghy M, Chang CL, Poulter N. Duration, frequency, recency, and type of migraine and the risk of ischaemic stroke in women of childbearing age. *J Neurol Neurosurg Psychiatry* 500. chiatry 2002;73:747-50
- Schwaag S, Nabavi DG, Frese A, Husstedt IW, Evers S. The association between 19 migraine and juvenile stroke: a case-control study. *Headache* 2003;43:90-5. Buring JE, Hebert P, Romero J, Kittross A, Cook N, Manson J, et al. Migraine and sub-
- sequent risk of stroke in the physicians' health study. Arch Neurol 1995;52:129-34. Merikangas KR, Fenton B, Cheng SH, Stolar MJ, Rish N. Association between migraine
- and stroke in a large-scale epidemiological study of the United States. Arch Neurol 1997;54:362-8.
- Nightingale AL, Farmer RD. Ischemic stroke in young women: a nested case-control study using the UK general practice research database. *Stroke* 2004;35:1574-8. Kruit MC, Buchem MA, Hofman P, Bakkers JT, Terwindt GM, Ferrari MD, et al. Migraine as a risk factor for subclinical brain lesions. *JAMA* 2004;291:427-34. 22
- Schwartz SM, Petitti DB, Siscovick DS, Longstreth WT Jr, Sidney S, Raghunathan TE, et al. Stroke and use of low-dose oral contraceptives in young women: a pooled analysis of two US studies. Stroke 1998;29:2277-84.
- Friberg L, Olsen TS, Roland PE, Lassen NA. Focal ischemia caused by instability of cerebro-vascular tone during attacks of hemiplegic migraine: a regional cerebral blood flow study. *Brain* 1987;110:917-34.
 Petty GW, Orencia AJ, Khandheria BK, Whisnant JP. A population based study of stroke 25
- 26 in the setting of miral valve prolapse: risk factors and infarct sub-type classification. Mayo Clin Proc 1994;69:632-4.
- 27 Ridker PM, Hennekens CH, Stampfer MJ, Manson, Vaughan DE. Prospective study of endogenous tissue plasminogen activator and risk of stroke. *Lancet* 1994;343:940-3. Tietjen GE, Day M, Norris L, Aurora S, Halvorsen A, Shultz LR, et al. Role of anticardi-
- 98 olipin antibodies in young persons with migraine and transient focal neurologic events: a prospective study. *Neurology* 1998;50:1433-40. (Accepted 9 November 2004)
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Amendment

This is Version 3 of the paper. In this version, Mahyar Etminan's first name is spelt correctly [in the previous version it was spelt Mayhar] and the University of Washington is correctly located in Seattle, WA [rather than in Washington DCl.