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No Decline in the Risk of Heart Failure Following Incident Atrial Fibrillation: A Community Study Assessing Trends Overall and by Ejection Fraction

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Abstract

Background—Patients with atrial fibrillation (AF) experience an increased risk of heart failure (HF). However, data are lacking on current trends in the risk of HF after AF.

Objective—To describe the temporal trends in HF occurrence after AF in a community cohort of patients with incident AF from 2000–2013.

Methods—Cox regression examined the association of year of AF diagnosis with HF and the predictors of developing HF after AF.

Results—Among 3491 AF patients without prior HF, 750 (21%) developed incident HF over a mean follow-up of 3.7 years. Among those with an echocardiogram, 422 (61%) had HF with preserved ejection fraction (HFpEF) and 270 (39%) had HF with reduced ejection fraction (HFpEF). After adjusting for demographics and comorbidities, the risk of developing HF did not change over time (hazard ratio [HR] (95% CI) per year of AF diagnosis: 1.01 (0.98–1.03) overall; 1.00 (0.98–1.03) for HFpEF; 1.00 (0.96–1.03) for HFrEF). Increasing age, obesity, smoking, diabetes, chronic pulmonary disease, and renal disease were predictors of developing HF. Compared to the Olmsted County, MN population, a substantial excess risk of developing HF was observed after AF diagnosis (standardized morbidity ratio (95% CI) 9.60 (7.44–12.19), 2.13 (1.56–2.84), and 1.70 (1.34–2.14) at 90 days, 1 year, and 3 years after diagnosis).

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Keywords

atrial fibrillation; atrial flutter; heart failure; epidemiology; outcomes

INTRODUCTION

Atrial fibrillation (AF) is the most common cardiac arrhythmia and disproportionately affects the elderly. AF is associated with substantial morbidity and mortality, and one of the most common consequences of AF is the development of heart failure (HF).¹ AF and HF share common risk factors and each condition predisposes a patient to the development of the other condition; $^{2-6}$ the interaction of the 2 conditions has been described as a 'vicious' electromechanical cycle.⁷ Notably, AF patients experience an approximately 3-fold increased risk of developing HF;^{5, 6, 8–10} yet, it is unknown whether improvements over time in the risk of HF after AF have occurred. A study in Olmsted County, MN residents with incident AF between 1980 and 2000 reported no improvement in the risk of HF over time in age- and sex-adjusted models, but a small decline after adjustment for comorbidities.¹¹ Recent evidence from the Framingham Heart Study suggests that AF patients have differential risks for developing each type of HF, with greater risks for developing HF with preserved ejection fraction (HFpEF) than HF with reduced ejection fraction (HFpEF).⁵ In addition, although the incidence of HF has declined over time, a shift in case mix of HF over the recent decade has resulted in an increased proportion of HFpEF.¹² Taken collectively, these findings suggest that the trends in developing HF after AF may differ over time by type of HF. However, contemporary data on the temporal trends in the incidence of HF after AF, specifically delineating differences over time by type of HF, are not described. Our objective was to provide contemporary data on trends for all HF, HFpEF, and HFrEF in a community cohort of patients with incident AF from 2000 to 2013.

METHODS

Study Population

This study was conducted in Olmsted County, Minnesota utilizing the resources of the Rochester Epidemiology Project (REP), a records-linkage system allowing virtually complete capture of health care in county residents.¹³ This record linkage system encompasses more than 6 million person-years of follow-up in over 500,000 unique individuals since 1966, enabling virtually complete capture of outcomes in Olmsted County, MN residents.¹³ Demographic and ethnic characteristics of Olmsted County are representative of the state of Minnesota and the Midwest region of the US, and age- and sexspecific mortality rates are similar to national data, supporting the generalizability of REP data.¹⁴ This study was approved by the Mayo Clinic and Olmsted Medical Center Institutional Review Boards.

Incident Atrial Fibrillation Cohort

Adults (aged 18 years) with incident AF or atrial flutter occurring between 2000 and 2013 were identified using *International Classification of Diseases, Ninth Revision, Clinical Modification* (ICD-9-CM) codes 427.31 and 427.32 from all providers in the REP, as well as Mayo Clinic electrocardiograms (ECGs). Diagnostic codes and ECGs from both inpatient and outpatient encounters were captured. The medical records for all patients were manually reviewed to validate the events, requiring evidence of first-ever AF or atrial flutter on an ECG, monitoring device, echocardiogram, or a physician diagnosis, as previously described.¹⁵ AF that occurred within 30 days after a cardiac surgery was excluded. However, among these individuals with post-operative AF, we continued to review the medical record for another episode of AF not associated with a surgery, which we considered incident AF.

Ascertainment of Heart Failure

HF events were identified using ICD-9-CM code 428 assigned during an inpatient or outpatient encounter at any provider in the REP. Trained nurse abstractors reviewed the medical records of these patients and validated the HF events using the Framingham Criteria.¹⁶ The incident date of HF was obtained and patients with HF occurring prior to the date of incident AF were considered prevalent and excluded. Only first-ever (incident) HF events occurring after the index AF date were included as outcomes in the analysis. The closest echocardiogram within 1 year of the date of HF was obtained to characterize the type of HF. Patients with EF greater than or equal to 50% were categorized as HFpEF and those with EF less than 50% were categorized as HFrEF.¹⁷

Clinical Data Collection

Height, weight, and smoking status at the time of AF diagnosis were manually abstracted from the medical records. Body mass index (BMI) was calculated as weight (in kg) divided by height (in meters) squared. Current, former, and never smoking status were collected and those who had smoked within 6 months prior to AF diagnosis were considered current smokers. Estimated glomerular filtration rate (eGFR) was calculated using the closest serum creatinine value within 90 days of the index AF date.¹⁸ The remaining covariates were ascertained electronically, implementing rules as previously described to reduce the capture of false positives due to rule out and suspect diagnoses by requiring 2 occurrences of a code within the 5 years prior to AF.¹⁵ The ICD-9 diagnostic codes used to define the comorbidities are included in Supplemental Table 1. Finally, outpatient prescription data is not routinely available in the REP prior to 2004. In the subset of our cohort with AF diagnosed between 2004 and 2013, rhythm control treatment approaches were defined as outpatient prescriptions for antiarrhythmic medications, direct current cardioversions, catheter ablation procedures, or Maze procedures.

Statistical Analysis

Analyses were performed using SAS, version 9.4. Patients with prevalent HF at baseline or with HF occurring the same day as AF were excluded. Characteristics of the cohort are presented as means and proportions; trends across year of AF diagnosis groups (2000–2003, 2004–2007, 2008–2010, 2011–2013) were compared using linear regression with year group

modeled as a 3-level categorical variable and Mantel-Haenszel chi-square tests. Cumulative incidence curves for HF across year of AF diagnosis categories were constructed treating death as a competing risk.¹⁹ P-values comparing the cumulative incidence curves were obtained using the method by Gray.²⁰ Curves were constructed for all HF, HFpEF, and HFrEF.

Cox proportional hazards regression models were used to determine the association of calendar year of AF diagnosis with the risk of any HF, HFpEF, and HFrEF. Follow-up time was calculated from the index date of AF until diagnosis of HF, death, last clinical encounter, or December 31, 2013, whichever came first. Year was treated continuously in one set of models to obtain p-values for time trends; a second set of models categorized year of AF diagnosis into groups with 2000–2003 serving as the reference to aid in interpretation. Models were adjusted for age, sex, and comorbidities that differed across year of AF diagnosis, including BMI, hypertension, diabetes, chronic pulmonary disease, peripheral vascular disease, and renal disease. In sensitivity analyses, we adjusted for eGFR (instead of renal disease using diagnostic codes) among the patients with an eGFR measured within ± 90 days of index (n=3261). Results from our sensitivity analysis were similar. To avoid excluding patients from the analysis due to missing eGFR data, our main models included renal disease captured by diagnostic codes. We tested age*year of AF and sex*year of AF interactions and found that the trends in HF after AF did not differ by age or sex. Sub-group analyses among patients with available rhythm control treatment approaches were also conducted, adjusting for rhythm control as a time-dependent variable as well as testing a rhythm control*year of AF interaction. The proportional hazards assumption was tested using scaled Schoenfeld residuals and found to be valid. In addition, Cox regression examined associations of patient characteristics with the occurrence of any HF, HFpEF, and HFrEF after adjusting for age and sex. Finally, we determined the associations of HF with death by modeling HF as a time-dependent variable in the Cox model.

The incident HF rate after AF was compared with the incident HF rate in the population of Olmsted County, Minnesota¹² to calculate standardized morbidity ratios. Expected HF rates were calculated by applying sex-, age- and period-specific HF incidence rates in the general population of Olmsted County to the person-time follow-up of the study cohort. Standardized morbidity ratios were calculated for discrete time periods over follow-up. For each time period, the standardized morbidity ratio was calculated by dividing the number of observed incident HF cases in the AF cohort for that time period by the expected number of HF cases. Confidence intervals were calculated according to the Poisson distribution. This analysis was repeated for HFpEF and HFrEF separately.

RESULTS

Between 2000 and 2013, 4431 patients with incident AF were identified. Of these, 640 had prevalent HF and 300 had HF occurring the same day as AF diagnosis and were excluded, resulting in 3491 AF patients. The mean (standard deviation) age at time of incident AF was 71.0 (14.9) years and 54.5% were male (Table 1). The most common comorbidities were hypertension (66%), diabetes (20%), and malignancy (19%). Patients diagnosed with AF in more recent years were more likely to be male, obese, and had higher prevalences of

hypertension, diabetes, peripheral vascular disease, and renal disease compared to patients diagnosed with AF in earlier years.

Over a mean follow-up of 3.7 years, 750 (21%) developed incident HF. The type of HF was known for 692 (92%) patients: 422 had HFpEF and 270 had HFrEF. The incidence rates (95% CI) of HF per 100 person-years were 5.84 (5.43–6.27) for all HF, 3.32 (3.01–3.66) for HFpEF, and 2.13 (1.88–2.40) for HFrEF. Treating death as a competing risk, no temporal trend in the risk of developing HF after AF was observed (Figure 1).

After adjustment for demographics and comorbidities, the hazard ratio (HR) (95% CI) for HF per year of AF diagnosis was 1.00 (0.98–1.03); p-trend=0.695 (Table 2). Likewise, no difference over time was observed for HFpEF (HR 1.00, 95% CI 0.98-1.03) or HFrEF (HR 1.00, 95% CI 0.96–1.03). In a sensitivity analysis, HF occurring the same day as AF (n=300) were counted as events and results were consistent (HR (95% CI) for HF per year of AF diagnosis: 1.00 (0.98–1.02), 0.99 (0.98–1.01), and 1.00 (0.98–1.02) for all HF, HFpEF, and HFrEF, respectively). Finally, in subgroup analyses among patients with AF diagnosed between 2004 and 2013, adjustment for rhythm control treatment approaches did not change the results (HR (95% CI) for HF per year of AF diagnosis: 1.01 (0.97-1.04), 1.00 (0.95-1.05), and 1.01 (0.95–1.08) for all HF, HFpEF, and HFrEF, respectively). In addition, no significant interactions between rhythm control and year of AF diagnosis were observed, indicating that the trends in HF risk over time did not differ between patients treated with rhythm and rate control. After adjusting for sex, older age at the time of incident AF was associated with an increased risk of developing HF (Figure 2A). After adjustment for age and sex, the following conditions were significant predictors of developing subsequent HF: obesity, current smoking status, hypertension, myocardial infarction, diabetes, chronic pulmonary disease, peripheral vascular disease, stroke/transient ischemic attack (TIA), and renal disease. Predictors of developing HFpEF were similar, with the exception of myocardial infarction and stroke/TIA which were not predictors of HFpEF, and female sex which was a predictor of HFpEF (Figure 2B). Increasing age, male sex, obesity, current smoking status, myocardial infarction, diabetes, chronic pulmonary disease, stroke/TIA, and renal disease were predictors of developing HFrEF, and hypertension was borderline significant (Figure 2C).

Standardized morbidity ratios comparing the risk of incident HF for AF patients to the Olmsted County, MN population of a similar age and sex distribution showed a dramatic excess risk of HF observed early after AF diagnosis (Figure 3A). Overall, AF patients exhibited an elevated risk of incident HF through 3 years after diagnosis, but no increased risk of HF after 4 years. Standardized morbidity ratios (95% CI) were 9.60 (7.44–12.19) at 31–90 days after diagnosis, 2.13 (1.56–2.84) at 181 days-1 year and 1.70 (1.34–2.14) at 2–3 years after AF diagnosis. Similar patterns were observed for HFpEF (Figure 3B) and HFrEF (Figure 3C), except the risk of HFpEF remained elevated through 5 years of follow-up in patients with AF compared to the Olmsted County population.

We also assessed the impact of HF on death among the 3491 AF patients free of HF at the time of AF diagnosis, by treating HF as a time-dependent variable in a Cox model. After adjustment for demographics and comorbidities, patients who developed HF over follow-up

had more than a 2-fold increased risk of death than those without HF (HR 2.68, 95% CI 2.35–3.06). Associations were similar for HFpEF (HR 2.54, 95% CI 2.16–2.99) and HFrEF (HR 2.52, 95% CI 2.08–3.05).

DISCUSSION

In this geographically-defined community, the development of HF after AF was common, with more than one-fifth of AF patients developing HF over 3.7 years of follow-up. No decline in the risk of HF after AF was observed since 2000. Developing HFpEF was more common than HFrEF. A substantial excess risk of developing HF was observed for AF patients compared to the Olmsted County population of similar age and sex. Finally, AF patients who developed HF exhibited a more than 2-fold increased risk of death than AF patients who did not develop HF.

Heart Failure Risk after Atrial Fibrillation

Although HF commonly occurs in AF patients, little data exist on temporal trends in HF after AF. A prior study in Olmsted County, MN residents with incident AF between 1980 and 2000 reported no improvement in the risk of HF over time in age- and sex-adjusted models.¹¹ After adjustment for comorbidities, a decline of 2% per calendar year of AF was observed in the risk of HF (adjusted HR 0.98, 95% CI 0.97–0.99). Our updated results in Olmsted County, MN residents with AF between 2000 and 2013 indicate no difference in the risk of HF after AF since 2000. In addition, no differences were observed over time for either type of HF. However, the rates of HFpEF were higher than for HFrEF, which is consistent with findings from the Framingham Study.⁵

Interestingly, we observed that the rates of HF were more than 2-fold higher than rates of stroke/TIA observed among patients with AF in our community,²¹ with incidence rates per 100 person-years of 5.84 for HF and 2.14 for ischemic stroke/TIA. This finding is consistent with a recent study among Medicare beneficiaries that identified HF as the most common non-fatal cardiovascular event occurring after AF.²² However, in our cohort, 21% of AF patients developed HF over a mean follow-up of 3.7 years, whereas only 13.7% of Medicare beneficiaries developed HF in the first 5 years after AF.²² The larger proportion of patients developing HF in our cohort is likely due to differences in ascertainment; we identified HF occurring in both inpatient and outpatient settings, whereas only inpatient or emergency claims identified HF outcomes in the Medicare study. Furthermore, our rates are in line with recent data from the Framingham Heart Study, which reported that 28% of AF patients developed HF the same day or after AF.⁵

Factors associated with an increased risk of developing HF included increasing age, obesity, current smoking status, hypertension, prior myocardial infarction, diabetes, chronic pulmonary disease, peripheral vascular disease, prior stroke or TIA, and renal disease. These findings are consistent with the knowledge that many of the underlying risk factors and mechanisms for AF and HF are the same.²³ However, our study identified male sex as a risk factor for only HFrEF; female sex was a risk factor for HFpEF, which has been previously been reported in the Framingham Heart Study.⁵ In addition, peripheral vascular disease was associated with the development of HFpEF but not HFrEF, and myocardial infarction and

stroke/TIA were associated with the development of HFrEF only. Finally, consistent with previous reports from the Framingham Heart Study,^{5, 6} we observed that the development of HF after AF conferred more than a 2-fold increased risk of death compared to patients who did not develop HF after AF.

Clinical Implications

Our findings augment important evidence that outcomes after AF have not improved over recent years despite continued emphasis on improving AF therapies. Previous data from our cohort showed that the risk of ischemic stroke/TIA²¹ and survival¹⁵ after AF have not improved since 2000. The current study now demonstrates that the risk of HF after AF is also stagnating. The lack of improvement in major cardiovascular events after AF indicate that efforts are urgently needed to improve AF outcomes. The importance of studying coexisting conditions and addressing comorbidities in treatment guidelines for patients with cardiovascular disease has been recently emphasized.²⁴ Yet, while current AF guidelines provide treatment recommendations for patients with AF and concomitant HF,²⁵ no recommendations exist for the prevention of HF in patients with AF. To reduce adverse clinical outcomes in AF, interventions should not only treat the AF itself, but should address the many comorbidities that commonly occur in elderly AF patients. For example, interventions to improve modifiable risk factors, including weight loss²⁶ and improvements in cardiorespiratory fitness,²⁷ which are associated with reduction in AF burden and maintenance of sinus rhythm, may reduce adverse outcomes in AF. Thus, future studies should address not only short term reduction in AF burden but the influence of weight loss and improvements in cardiorespiratory fitness on long term clinical outcomes, such as the development of HF.

Limitations and Strengths

Some limitations deserve mention. First, we did not have sufficient information to stratify results by type of AF (paroxysmal, persistent, permanent). Second, we may have missed some HF events if the diagnosis occurred outside of Olmsted County. Finally, our results may not be generalizable to the entire US, although Olmsted County is representative of the state of Minnesota and the Upper Midwest region of the US.¹⁴ Importantly, our data represent the experience of a community and includes all residents with AF without restriction on age, sex, other patient characteristics, treatment setting, or insurance provider. We captured and manually validated AF and HF occurring in both inpatient and outpatient settings from multiple providers in the area, allowing for near complete ascertainment of AF and outcomes.

Conclusion

In the community, 21% of AF patients developed HF over 3.7 years of mean follow-up, with HFpEF being more common than HFrEF. Patients with AF have an excess risk of developing HF, especially early after AF diagnosis, and those who develop HF experience more than a 2-fold increased risk of death compared to AF patients who do not develop HF. Finally, the rates of developing HF after AF have not improved since 2000. Thus, continued efforts to optimize the management of AF to improve outcomes are needed.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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Figure 1. Cumulative incidence of heart failure by year of atrial fibrillation diagnosis Panel A, all heart failure; Panel B, heart failure with preserved ejection fraction; Panel C, heart failure with reduced ejection fraction. The cumulative incidence curves were adjusted for death as a competing event.

		Heart	Failure			
					Estima	ate (95% CI)
Age (per 10 years)			⊢⊷⊣		1.81	(1.69-1.93)
Male		⊢ +⊣			0.98	(0.84-1.13)
Overweight		⊢ -			1.03	(0.85-1.24)
Obese			→ →→		1.48	(1.23-1.78)
Current smoker			·		1.68	(1.30-2.16)
Hypertension		- I	- -		1.25	(1.05-1.48)
Myocardial infarction			⊢ •		1.73	(1.40-2.14)
Diabetes			⊢ → → → →		1.58	(1.33-1.87)
Chronic pulmonary disease			·		1.81	(1.49-2.18)
Peripheral vascular disease		,			1.46	(1.14-1.87)
Stroke/TIA					1.30	(1.06-1.59)
Dementia					0.78	(0.55-1.10)
Malignancy		H+++			0.92	(0.76-1.11)
Metastatic solid tumor					0.79	(0.51-1.24)
Renal disease			·		1.73	(1.30-2.29)
Rheumatologic disease			•		1.19	(0.89-1.60)
Liver disease			•		1.38	(0.76-2.50)
	0.0	0.6	1.2 1.8 Hazard Ratio	2.4	3.0	

Heart Failure with Preserved EF					
	7	Estimate (95% CI)			
Age (per 10 years)		1.93 (1.76-2.12)			
Male		0.67 (0.54-0.82)			
Overweight	⊢ - 1	1.05 (0.82-1.35)			
Obese		1.53 (1.20-1.96)			
Current smoker	·	1.54 (1.07-2.22)			
Hypertension		1.30 (1.03-1.65)			
Myocardial infarction	HH	1.31 (0.96-1.80)			
Diabetes		1.59 (1.26-2.00)			
Chronic pulmonary disease		1.99 (1.56-2.54)			
Peripheral vascular disease		1.66 (1.21-2.28)			
Stroke/TIA	⊢ 1	1.18 (0.90-1.55)			
Dementia		0.70 (0.44-1.12)			
Malignancy		0.91 (0.70-1.17)			
Metastatic solid tumor	↓ →	1.07 (0.64-1.79)			
Renal disease	· · · · · · · · · · · · · · · · · · ·	1.74 (1.20-2.54)			
Rheumatologic disease		1.19 (0.82-1.72)			
Liver disease	+ + · · · · · · · · · · · · · · · · · ·	1.53 (0.72-3.22)			
	0.0 0.8 1.6 2.4 3.2 4 Hazard Ratio	1 .0			



Figure 2. Age- and sex-adjusted predictors of heart failure

Panel A, all heart failure; Panel B, heart failure with preserved ejection fraction; Panel C, heart failure with reduced ejection fraction. COPD, chronic obstructive pulmonary disease; TIA, transient ischemic attack; EF, ejection fraction. The estimate for age is adjusted for sex and the estimate for sex is adjusted for age; all other estimates are adjusted for age and sex.







Figure 3. Standardized morbidity ratios of heart failure after incident atrial fibrillation Panel A, all heart failure; Panel B, heart failure with preserved ejection fraction; Panel C, heart failure with reduced ejection fraction. Shaded area indicates 95% confidence intervals.

Table 1

Baseline characteristics of the incident atrial fibrillation patients

	Overall (N=3491)	2000-2003 (N=808)	2004-2007 (N=944)	2008–2010 (N=769)	2011-2013 (N=970)	P-value for trend
Age, mean (SD)	71.0 (14.9)	71.1 (14.9)	71.0 (15.1)	70.5 (14.8)	71.4 (14.7)	0.671
Male	1901 (54.5)	420 (52.0)	491 (52.0)	435 (56.6)	555 (57.2)	0.007
Body mass index, kg/m ²						0.010
<25	985 (28.2)	226 (28.0)	295 (31.3)	211 (27.4)	253 (26.1)	
25 to <30	1153 (33.1)	283 (35.0)	312 (33.1)	253 (32.9)	305 (31.5)	
30	1350 (38.7)	299 (37.0)	335 (35.6)	305 (39.7)	411 (42.4)	
Smoking status						0.093
Current	372 (10.7)	90 (11.1)	92 (9.8)	87 (11.3)	103 (10.6)	
Former	1428 (40.9)	317 (39.2)	374 (39.6)	327 (42.5)	410 (42.3)	
Never	1691 (48.4)	401 (49.6)	478 (50.6)	355 (46.2)	457 (47.1)	
Hypertension	2302 (65.9)	473 (58.5)	623 (66.0)	536 (69.7)	670 (69.1)	<0.001
Myocardial infarction	331 (9.5)	74 (9.2)	79 (8.4)	70 (9.1)	108 (11.1)	0.103
Diabetes	686 (19.7)	117 (14.5)	172 (18.2)	179 (23.3)	218 (22.5)	<0.001
Chronic pulmonary disease	430 (12.3)	129 (16.0)	107 (11.3)	78 (10.1)	116 (12.0)	0.015
Peripheral vascular disease	254 (7.3)	45 (5.6)	71 (7.5)	58 (7.5)	80 (8.3)	0.045
Dementia	182 (5.2)	36 (4.5)	43 (4.6)	54 (7.0)	49 (5.1)	0.253
Malignancy	671 (19.2)	167 (20.7)	153 (16.2)	159 (20.7)	192 (19.8)	0.686
Metastatic solid tumor	178 (5.1)	40 (5.0)	44 (4.7)	51 (6.6)	43 (4.4)	0.996
Renal disease	194 (5.6)	30 (3.7)	33 (3.5)	44 (5.7)	87 (9.0)	<0.001
eGFR, mean (SD), mL/min per 1.73 ${ m m^2}^*$	65.9 (26.2)	59.0 (19.9)	63.0 (23.4)	70.6 (29.0)	70.6 (29.1)	<0.001
Rheumatologic disease	166 (4.8)	40 (5.0)	42 (4.5)	35 (4.6)	49 (5.1)	0.854
Ischemic stroke or TIA	404 (11.6)	90 (11.1)	110 (11.7)	89 (11.6)	115 (11.9)	0.672
Liver disease	65 (1.9)	9 (1.1)	20 (2.1)	13 (1.7)	23 (2.4)	0.104
CHA ₂ DS ₂ -VASc, mean (SD)	2.9 (1.8)	2.8 (1.8)	2.9 (1.8)	2.9 (1.8)	3.0 (1.8)	0.079
All results are reported as n (%) unless otherv	wise specified.					

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TIA, transient ischemic attack; eGFR, estimated glomerular filtration rate. * Among 3261 patients with eGFR measured within ± 90 days of index.

Table 2

Hazard ratios (95% confidence intervals) for incident heart failure by year of atrial fibrillation diagnosis

	All heart failure	Heart failure with preserved ejection fraction	Heart failure with reduced ejection fraction
Per year	1.00 (0.98–1.03)	1.00 (0.98–1.03)	1.00 (0.96–1.03)
P time trend *	0.695	0.805	0.919
By year group			
2000-2003	1.00 (ref)	1.00 (ref)	1.00 (ref)
2004-2007	1.07 (0.89–1.29)	1.15 (0.90–1.47)	0.96 (0.70–1.32)
2008-2010	0.94 (0.75–1.18)	0.89 (0.65–1.22)	0.96 (0.67–1.39)
2011-2013	1.07 (0.84–1.36)	1.10 (0.80–1.51)	0.93 (0.62–1.40)

Adjusted for age, sex, body mass index, hypertension, diabetes, chronic pulmonary disease, peripheral vascular disease, and renal disease.

 $\stackrel{*}{P}$ time trend is the 1df p-value for year based on the model treating year as a continuous variable.