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# Breastfeeding and Endometrial Cancer Risk: An Analysis From the Epidemiology of Endometrial Cancer Consortium

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# Abstract

**Objective**—To investigate the association between breastfeeding and endometrial cancer risk using pooled data from 17 studies participating in the Epidemiology of Endometrial Cancer Consortium.

**Methods**—We conducted a meta-analysis with individual-level data from three cohort and 14 case-control studies. Study-specific odds ratios (ORs) and 95% confidence intervals (CI) were estimated for the association between breastfeeding and risk of endometrial cancer using multivariable logistic regression, and pooled using random-effects meta-analysis. We investigated between-study heterogeneity with I<sup>2</sup> and Q statistics and meta-regression.

**Results**—After excluding nulliparous women, the analyses included 8981 women with endometrial cancer and 17241 control women. Ever breastfeeding was associated with an 11% reduction in risk of endometrial cancer (pooled OR=0.89, 95% CI 0.81–0.98). Longer average duration of breastfeeding per child was associated with lower risk of endometrial cancer, although there appeared to be some levelling of this effect beyond 6–9 months. The association with ever breastfeeding was not explained by greater parity and did not vary notably by body mass index or histological subtype (grouped as endometrioid and mucinous versus serous and clear cell).

**Conclusions**—Our findings suggest that reducing endometrial cancer risk can be added to the list of maternal benefits associated with breastfeeding. Ongoing promotion, support and facilitation of this safe and beneficial behavior might therefore contribute to the prevention of this increasingly common cancer.

# Introduction

Endometrial cancer is the fourth most common cancer in women in high income countries such as the USA, Canada and Australia where collectively more than 57000 women are diagnosed each year<sup>1</sup> and incidence has been increasing over the last 20–30 years. A better understanding of its causes and ways it could be prevented is therefore required.

Exclusive breastfeeding generally suppresses ovulation and therefore maternal estrogen levels.<sup>2</sup> Reduced estrogen levels reduce endometrial mitoses<sup>3</sup> and might therefore reduce risk of endometrial cancer; however, clear evidence for this association is lacking. The most recent World Cancer Research Fund/American Institute for Cancer Research endometrial cancer report classified the evidence for an association with lactation as 'limited-no conclusion'.<sup>4</sup>

Some previous studies indicate risk reduction with longer breastfeeding durations;<sup>5</sup> some show risk reduction only for recent breastfeeding episodes;<sup>6,7</sup> and others found no association.<sup>8,9</sup> So, while two meta-analyses of published data suggest risk reduction with

longer total durations of breastfeeding, they were unable to account for significant heterogeneity across studies.<sup>10,11</sup> They also could not investigate factors that might modify breastfeeding associations such as body mass index, time since last pregnancy or menopausal status or how the duration of individual breastfeeding episodes might influence endometrial cancer associations. We hypothesise that, since estrogen levels during lactation are lowest while ovulation is suppressed, ongoing breastfeeding beyond this point might not confer additional benefit.

To more comprehensively assess the association between breastfeeding and endometrial cancer risk, we pooled data from 17 studies participating in the Epidemiology of Endometrial Cancer Consortium.<sup>12</sup>

### Materials and Methods

Ethical approval was obtained from each study's institutional review board. All participants provided informed consent to take part in the respective studies.

We conducted a meta-analysis with individual-participant data from 17 independent studies (14 case-control, three cohort) from the Epidemiology of Endometrial Cancer Consortium (Table 1) that provided data on breastfeeding and confounding factors. Cohort studies were analysed as nested case-control studies with four controls matched on birth year randomly selected for each case from cohort members who had not had a hysterectomy or an endometrial cancer diagnosis (or, in The Nurses' Health Study, any other invasive cancer) by the case's diagnosis date.

The data harmonisation process for the Consortium was described elsewhere<sup>13</sup> but, briefly, studies provided information on participants' demographics (cases' diagnosis age, reference date for controls, race, education) and reproductive, health and lifestyle factors (e.g., parity, oral contraceptive use (OCP), height, weight) according to specified definitions. Studies also provided information on whether women breastfed, their total duration of breastfeeding (sum of duration of all of their breastfeeding episodes) and, for some studies, breastfeeding duration for each child. For cases, studies provided tumor histology information, where available. Women with sarcomas were excluded.

For the cohort studies the data were collected in slightly different ways. For one,<sup>14</sup> data were mostly obtained from baseline questionnaires (1991–1992), with body mass index (BMI) data updated in 2003 for cases diagnosed after this and their matched controls. For the other two<sup>15,16</sup>, most data were from questionnaires returned in the period prior to a participant becoming a case or being selected as a control, although breastfeeding data were collected only once (1986) in the Nurses' Health Study.<sup>16</sup>

Analyses were restricted to parous women with breastfeeding information. Less than 2% of participants were missing breastfeeding/covariate data so a complete case analysis was undertaken. Study-specific odds ratios (ORs) and 95% confidence intervals (CIs) were estimated for the associations between breastfeeding variables (see below) and risk of endometrial cancer using multivariable logistic regression (conditional logistic regression for matched studies). Models were adjusted for parity (continuous), OCP use, BMI (around

reference age; continuous) and education level. Study-specific ORs were pooled (pORs) using random effects models giving overall pORs and 95%CIs. Between-study heterogeneity was assessed using  $I^2$  and Q statistics.

The women were classified as having breastfed or not. Total breastfeeding duration was modelled continuously and in categories: never breastfed, 3 months, >3 to 6 months, >6 to 9 months, >9 to 12 months, and then in six-month categories up to >36 months.

Because we hypothesised that risk reduction with breastfeeding might plateau as duration of a breastfeeding episode increased (beyond return of ovulation), we examined associations with duration of individual breastfeeding episodes. First, to allow inclusion of all the studies, we divided total breastfeeding duration by the number of births for each woman to give average duration of breastfeeding/child. Then, for studies with information on individual breastfeeding episodes (<sup>8,14, 17–25</sup> plus Turin Study), we divided total breastfeeding duration by the number of children actually breastfed. Both variables were analysed in categories (as above, but with the highest category 12–18 months) weighting the first variable by the number of births and the second by the number of children breastfed to give the risk per child for each average duration.

We assessed whether associations differed between case-control and cohort studies by stratifying the meta-analysis by study type. We also stratified analyses by number of births to assess potential for residual confounding by parity; and by BMI at reference age (<30 kg/m<sup>2</sup> vs. 30 kg/m<sup>2</sup>) and early adulthood (<25 kg/m<sup>2</sup> vs. 25 kg/m<sup>2</sup>, where available, as this measure was generally closer to the time of breastfeeding) to assess whether associations varied by BMI. We conducted analyses in strata of histological subtype (Type I -endometrioid and mucinous adenocarcinomas vs. Type II -serous and clear cell cancers) as etiology may vary by subtype.<sup>13</sup> Differences between strata were assessed using random-effects meta-regression.<sup>26</sup>

We explored heterogeneity between studies by assessing whether associations varied by participant characteristics, including those already mentioned plus race (white, black, Asian); menopausal status; participant's birth year (<1950 vs. 1950); and years since last pregnancy (<30 vs. 30 years). We used random-effects meta-regression<sup>26</sup> to evaluate whether these factors explained between-study heterogeneity.

We estimated the proportion of endometrial cancers (population attributable fraction) that could be attributed to breastfeeding for 6 months/child using the standard formula:

Population Attributable Fraction =  $\frac{\sum (p_x \times \text{ERR}_x)}{1 + \sum (p_x \times \text{ERR}_x)}$ , where the ERR was the excess relative risk for each category of average breastfeeding/child breastfed below the >6 to 9 months category. The prevalence  $(p_x)$  was the proportion of women in each category. Sensitivity analyses were conducted using prevalence estimates from the studies with the lowest<sup>23</sup> and highest prevalence<sup>22</sup> of breastfeeding >6 months/child.

We used SAS version 9.4 (SAS Institute, Cary, NC) and STATA version 13 (StataCorp LP, College Station, TX) for statistical analyses. All statistical significance tests are two-sided.

# Results

Table 1 shows details of the included studies. The analyses included 8981 parous women with endometrial cancer and 17241 parous control women. Overall, 68% of the control women had breastfed at least one child but the percentage varied considerably across studies ranging from 43%<sup>15</sup> to 92%.<sup>14</sup> Median total duration of breastfeeding among controls who had ever breastfed ranged across studies from five<sup>15,16,24</sup> to 13 months.<sup>22</sup> As expected, compared to cases, the controls had higher parity, were more likely to have used OCPs and, on average, had a lower BMI (Table 2). Control women were also somewhat more likely to have post-high school education.

Table 3 shows the crude and adjusted associations between the various breastfeeding measures and endometrial cancer risk – for brevity we refer only to the adjusted pooled ORs (pORs) in this section. Having ever breastfed was associated with a statistically significant 11% reduced risk of endometrial cancer compared with never breastfeeding (pOR=0.89, 95%CI 0.81–0.98) (Figure 1.), although there was moderate between-study heterogeneity  $(I^2=45\%, p=0.02)$ . For total duration of breastfeeding, estimates for durations beyond three months suggested an inverse association that became more pronounced for very long durations (pOR=0.67, 95% CI 0.53–0.83 for >36 months total duration versus never breastfeeding). However, the risk reduction associated with increasing total duration of breastfeeding was not clearly linear. Our analyses of individual episodes of breastfeeding (average breastfeeding/child, average breastfeeding/child breastfed) also showed that individual breastfeeding durations beyond three months were associated with statistically significant reductions in risk. However, beyond the >6 to 9 months category the odds ratios did not appear to decrease much further although the numbers of women who breastfed individual children for longer than this were small and therefore these estimates are less precise.

Figure 2 shows the results of analyses of ever-breastfeeding stratified by participant characteristics. The association did not vary substantially by tumor type (I vs. II), parity, recent BMI, or menopausal status. The magnitude of association differed somewhat by race/ ethnicity, BMI in early adulthood (no association in heavier women) and time since last pregnancy but the differences were not statistically significant. For the race analyses, the number of studies included in each stratum was small; only two studies contributed to the estimate for Asian women,<sup>22,29</sup> including one in which all women were Asian. Estimates for ever-breastfeeding differed significantly by women's birth year (p=0.03). For women born since 1950 (likely to have breastfed after ~1970), ever breastfeeding was associated with a 28% reduction in endometrial cancer risk (OR=0.72, 95% CI 0.59-0.87) compared with never breastfeeding, whereas for women born before 1950 the risk reduction was smaller and no longer statistically significant (OR=0.93, 95% CI 0.85–1.02). It is notable that the study with the highest OR associated with ever-breastfeeding (OR=1.12, 95%CI 0.97-1.29<sup>16</sup> included no women born post-1950 and the participants had the highest mean age at diagnosis (65 years). Removing this study from the pooled analysis reduced the I<sup>2</sup> from 45% to 24%.

Results from univariable meta-regression to evaluate potential sources of between-study heterogeneity in estimates for ever-breastfeeding are shown in Table 4. Of the factors considered, only the proportion of women whose last birth was <30 years prior to study participation (p=0.02, adjusted R<sup>2</sup>=82.3%) and the median age at diagnosis of case women in each study (p=0.03, adjusted R<sup>2</sup>=68.5%) were significantly associated with the strength of association. Although these factors were no longer statistically significant in a metaregression excluding the NHS, the adjusted R<sup>2</sup> for each (last birth 30 years prior; median age at diagnosis) remained high at 41% and 68% respectively. No factors remained statistically significant when included in multivariable meta-regression, likely because of the relatively small number of studies involved.

We estimated the proportion of endometrial cancers among parous women (population attributable fraction) that could be attributed to breastfeeding for 6 months per child to be 11% (range from sensitivity analyses 5–15%). The prevalence of nulliparity across the studies was 16% giving a population attributable fraction for all women of ~9% (possible range 4–13%).

# Discussion

We observed a modest reduction in risk of endometrial cancer associated with breastfeeding that was not explained by greater parity and did not vary by BMI or endometrial cancer type (Type I versus Type II). The association appeared stronger with increasing duration of breastfeeding episodes up to between 6–9 months per child breastfed but thereafter the decline in risk was smaller. These results suggest that reduction in endometrial cancer risk could be added to the list of maternal benefits associated with breastfeeding for more than six months.

Strengths of our study include the large sample size, our ability to define exposure levels consistently across studies, to adjust consistently for potential confounders and the inclusion of studies from different US populations and different countries. However, some limitations should be considered. We had no information on factors that predispose women to breastfeed or data on when menstruation recommenced so could not consider these in our analyses. Most of the studies were of case-control design with potential self-selection of more health-conscious controls, perhaps more likely to have breastfed. Nevertheless, the pooled estimates for ever-breastfeeding did not vary by study design (cohort versus case-control) making selection bias less likely. Also, all studies relied on retrospective self-report of breastfeeding, which for many women occurred years prior to study participation and thus was possibly subject to recall error. This would be non-differential for the cohort studies and, again, the lack of difference in pooled estimates by study design makes substantial recall bias unlikely. Finally, despite the inclusion of >26000 women, relatively few women breastfed for long durations making estimates for these categories less precise.

An inverse association between breastfeeding and endometrial cancer risk is biologically plausible. Breastfeeding can suppress gonadotrophin-releasing hormone inhibiting ovarian follicular growth and reducing estradiol levels to within the postmenopausal range.<sup>2</sup> At these levels, endometrial cell mitoses are virtually absent.<sup>3</sup> Estrogen levels in breastfeeding

women appear to depend on suckling stimuli, with the lowest levels found in women breastfeeding exclusively.<sup>2</sup> Most guidelines recommend introduction of non-milk foods when infants are around six months old, so, notwithstanding considerable variation in breastfeeding practices between women, it is likely that the suckling stimulus decreases around this time and estrogen levels increase. This is consistent with our finding of levelling in risk reduction with longer durations of individual episodes of breastfeeding.

We had expected that BMI might modify the association with breastfeeding because obesity lowers sex hormone binding globulin levels, increasing bioavailable estrogen and testosterone. This, with estrone production in adipose tissue in more obese women, might negate the relative hypo-estrogenic state induced by breastfeeding. We did not observe significant effect-modification but our BMI measures may not closely reflect adiposity during breastfeeding. Most BMI data were from around diagnosis, generally years after breastfeeding. We had BMI estimates from early adulthood from 12 studies but this may have been poorly recalled since it was further in the past and may still not reflect a woman's BMI during breastfeeding. Furthermore, few women were overweight on early adulthood measures so power may have been insufficient to detect effect-modification.

We observed moderate between-study heterogeneity in some associations, most notably for ever-breastfeeding ( $I^2=45\%$ ). This is consistent with published meta-analyses investigating this relation,<sup>10,11</sup> although only two studies<sup>8,20</sup> included in those analyses were also in ours. We found the different proportions across studies of women who gave birth further in the past significantly contributed to heterogeneity. We also found the inverse association with breastfeeding was weaker in women born pre-1950. These factors may reflect attenuation of breastfeeding effects over time or differences in breastfeeding practices across birth cohorts causing different physiological effects in the endometrium. Breastfeeding rates, at least in the USA, were much lower in the 1950s and 1960s than more recently<sup>27</sup> but whether breastfeeding practices (e.g., less demand or exclusive feeding with potentially less ovulation suppression<sup>28</sup>) also differed is not clear.

Assuming the associations we observed are causal and that prevalence of breastfeeding and nulliparity among women diagnosed with endometrial cancer in 2015 is similar to our study, it may be that, of the estimated 345,000 women diagnosed with endometrial cancer worldwide in 2015,<sup>1</sup> around 31,000 (9%, possible range 14,500 (4%) to 43,500(13%)) might have been prevented if all parous women had been able to breastfeed their babies for more than six months each.

For health practitioners, our study suggests that promoting breastfeeding and providing support to women to breastfeed for six months and beyond might have the added benefit of contributing to the prevention of this increasingly common cancer.

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STUDY		% Weigh
Case-control		
PEDS	0.62 (0.45, 0.85)	5.63
WNYDS	0.66 (0.44, 0.98)	
SECS	0.70 (0.55, 0.90)	
CONN	0.79 (0.60, 1.04)	
ALTA	0.80 (0.61, 1.05)	
FHCRC	0.87 (0.69, 1.10)	
USC	0.91 (0.69, 1.21) (	6.34
USE	0.93 (0.64, 1.35)	4.45
WISE	0.96 (0.77, 1.19) 8	8.08
EDGE	0.97 (0.70, 1.34)	5.39
PECS	1.00 (0.75, 1.35)	6.08
ANECS	1.04 (0.77, 1.40)	5.93
TURIN — + +		3.41
IMS	- 1.19 (0.86, 1.63) \$	5.55
Subtotal (I-squared = 29.3%, p = 0.143)	0.87 (0.80, 0.96) 8	82.96
Cohort		
SWLHS *	0.65 (0.32, 1.33)	
BWHS	0.75 (0.53, 1.07)	
NHS	1.12 (0.96, 1.29)	
Subtotal (I-squared = 65.0%, p = 0.057)	0.89 (0.63, 1.26)	17.04
Overall (I-squared = 44.8%, p = 0.024)	0.89 (0.81, 0.98)	100.00
NOTE: Weights are from random effects analysis		
.3 1	2	
Odds Ratio		

#### Figure 1.

Forest plot showing study-specific estimates and 95% confidence intervals (CI) for the association between ever-breastfeeding and endometrial cancer risk amongst parous women. Estimates are stratified by study design (case-control/cohort) and ordered smallest to largest. The *box size* indicates the study weight, the *line* represents the 95% confidence interval, and the *diamonds* represent the pooled estimates. Higher I<sup>2</sup> values and *P* values <.05 suggest statistically significant between-study heterogeneity. Study acronyms are defined in Table 1.

Characteristic	Studies (N)	Breastfed Cas Ever	es / Controls Never		Pooled OR (95% Cl)	l²%
Tumour Type <sup>a</sup>						
Type 1	15	4718 / 11675	2546 / 5566	- <b>-</b> -	0.89 (0.82, 0.98)	31.1
Type 2	13	486 / 11675	261 / 5566		0.86 (0.67, 1.09)	34.4
Parity						
One	16	1108 / 1750	839 / 1179		0.81 (0.66, 1.00)	28.0
Two	17	2177 / 4202	1064 / 1967	-+-	0.94 (0.82, 1.09)	20.6
Three	16	1416 / 2923	681 / 1312		0.88 (0.74, 1.02)	0.0
Race						
White <sup>b</sup>	13	4497 / 9221	2457 / 4093		0.97 (0.90, 1.05)	
Black <sup>c</sup>	6	155 / 611	209 / 776		0.81 (0.60, 1.09)	0.
Asian <sup>d</sup>	2	993 / 1037	200 / 172	<b>—</b>	0.71 (0.55, 0.90)	0.
Young Adult E						
BMI ≥ 25	10	520 / 612	328 / 319		<ul> <li>1.08 (0.77, 1.51)</li> </ul>	
BMI < 25	12	4103 / 8032	1825 / 3481	-•-	0.93 (0.84, 1.03)	22.
Recent BMI						
BMI ≥30	15	2117 / 2026	1224 / 1169		0.87 (0.73, 1.05)	35.
BMI < 30	17	3762 / 9551	1794 / 4373		0.89 (0.79, 0.99)	42.
Menopausal S	tatus					
Pre/peri	14	1100 / 2979	548 / 1406		0.89 (0.75, 1.05)	4.
Post	15	4115 / 7077	2174 / 3729		0.92 (0.82, 1.02)	39.
Year of birth				10		
Pre 1950 <sup>e</sup>	17	4959 / 9556	2647 / 4889	-+-	0.93 (0.85, 1.02)	25.4
1950 +	11	973 / 2112	402 / 675	<b>→</b>	0.72 (0.59, 0.87)	0.
Last birth						
<30 yrs ago	15	2326 / 5017	1029 / 1994		0.87 (0.74, 1.02)	38.
≥30 yrs ago	15	3123 / 5328	1743 / 2971	+	1.01 (0.91, 1.12)	5.4
			I 0.	I	1 1	

Odds ratio

#### Figure 2.

Pooled odds ratios (OR) and 95% confidence intervals (CI) for the association between everbreastfeeding and endometrial cancer risk stratified by participant characteristics. The number of studies included in each stratum varies because some studies did not have information on the variable of interest or because there were too few women in specific strata to calculate study-specific estimates. <sup>a</sup>Only case group was stratified. <sup>b</sup>Includes five studies in which all women were white. <sup>c</sup>Includes one study in which all women were black. <sup>d</sup>Includes one study in which all women were Asian. <sup>e</sup>The estimate for ever breastfed in women born before 1950, including only the 11 studies that also included women who were born after 1950 is 0.93 (95% CI 0.84–1.03). Author Manuscript

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# Table 1

Characteristics of 17 E2C2 studies pooled to investigate the association between breastfeeding and endometrial cancer risk among parous women

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Study Study name	Acronym	Country	Years of recruitment/diagnosis	Cases (8981) N	Controls (17241) N	Ever brea	Ever breastfed N (%)	Median tota (months) <sup>*</sup>	Median total breastfeeding (months)*
Case-control studies						cases	controls	cases	controls
Alberta Case-Control Study on Endometrial Cancer <sup>17</sup>	ALTA	Canada	2002–2006	441	922	298 (68)	692 (75)	9	6
Australian National Endometrial Cancer Study <sup>29</sup>	ANECS	Australia	2005–2007	1119	659	913 (82)	561 (85)	10	12
Connecticut Endometrial Cancer Study <sup>18</sup>	CONN	USA	2004–2009	505	558	218 (43)	299 (54)	8	12
Estrogen, Diet, Genetics, and Endometrial Cancer <sup>19</sup>	EDGE	USA	2001–2005	370	407	159 (43)	188 (46)	9	6
Fred Hutchinson Cancer Research Center <sup>25</sup>	FHCRC	NSA	1994–2005	703	737	452 (64)	514 (70)	9	7
Italian Multi Centre Study <sup>20</sup>	IMS	Italy	1992–2006	386	780	303 (79)	592 (76)	6	10
Patient Epidemiologic Data System <sup>30</sup>	PEDS	USA	1982–1998	401	423	195 (49)	261 (62)	9	L
Polish Endometrial Cancer Study <sup>21</sup>	PECS	Poland	2000–2003	447	1701	366 (82)	1388 (82)	8	L
Shanghai Endometrial Cancer Study <sup>22</sup>	SECS	China	1997–2004	1089	1161	896 (83)	1003 (86)	14	13
Turin Case Control Study	TURIN	Italy	1998–2008	219	260	170 (78)	196 (75)	8	6.5
USC LA case control <sup>24</sup>	USC	USA	1987–1993	660	709	411 (63)	468 (66)	3	5
US Endometrial Cancer Study <sup>8</sup>	USE	USA	1987–1990	316	290	191 (60)	181 (62)	4	9
Western New York Diet Study <sup>23</sup>	WNYDS	USA	1986–1991	185	532	91 (49)	277 (52)	3	9
Women's Insight and Shared Experience <sup>31</sup>	WISE	NSA	1999–2002	512	1418	217 (42)	677 (48)	6	8
Cohort studies									
Black Women's Health Study <sup>15</sup>	BWHS	USA	1995–2013	202	899	79 (39)	388 (43)	5	5
Nurses' Health Study <sup>16</sup>	SHN	USA	1986–2012	1197	4788	771 (64)	3069 (64)	5	5
Swedish Women's Lifestyle and Health Study <sup>14</sup>	SMLHS	Sweden	1993-2012	229	266	202 (88)	921 (92)	∞	6

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\* Amongst women who breastfed Page 12

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Table 2

Prevalence of covariates among parous women in the pooled analysis.

	Cases N=8981		Controls N=17241	s 1	
Covariate	Mean ±SD	%	Mean ±SD	*%	$\mathbf{P}^{\dagger}$
Age, years	61 ±8.9		$60 \pm 8.8$		<0.001
Body mass index $(kg/m^2)$	$29.3 \pm 7.0$		$25.9 \pm 5.0$		<0.001
Race					< 0.001
White		77.4		73.1	
Black		4.1		4.7	
Asian		12.6		12.4	
Other		3.8		3.0	
Missing		2.1		6.8	
Education					<0.001
High school or less		49.4		46.0	
Post high school		49.5		52.5	
Missing		1.1		1.5	
Parity					<0.001
1		21.7		18.2	
2		36.1		35.5	
3		23.3		24.7	
4		11.3		12.0	
5		7.6		10.5	
Oral contraceptive use					<0.001
No		57.5		50.4	
Yes		41.9		49.0	
Missing		0.6		0.6	

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\* weighted by numbers of case women in each study.

 $\dot{f}$  differences between cases and controls were tested with t-tests for continuous variables and  $\chi 2$  tests for categorical variables.

# Table 3

Associations between breastfeeding and endometrial cancer risk among parous women in the pooled analysis

	$N(\%)^*$	N(%)*	OR (95%CI)	<b>OR</b> <sup>†</sup> (95%CI)	¥%-1
Ever breast fed					
No	3049 (33.9)	5566 (32.3)	1.00	1.00	4
Yes	5932 (66.1)	11675 (67.7)	0.82 (0.75–0.90)	$0.89\ (0.81-0.98)$	5
Total duration of breastfeeding (months)	feeding (month:	s)			
Never breastfed	3049 (34.1)	5566 (32.3)	1.00	1.00	
>0 to 3	1647 (18.4)	3059 (17.8)	$0.98\ (0.88{-}1.09)$	1.01 (0.91–1.12)	18
>3 to 6	915 (10.2)	2017 (11.7)	0.83 (0.74–0.92)	0.84 (0.73–0.98)	39
>6 to 9	710 (7.9)	1555 (9.0)	0.81 (0.70–0.95)	0.90 (0.77–1.05)	32
>9 to 12	667 (7.5)	1000(5.8)	0.89 (0.73–1.07)	0.96 (0.79–1.16)	47
>12 to 18	723 (8.1)	1413 (8.2)	0.82 (0.73–0.93)	0.93 (0.82–1.05)	0
>18 to 24	506 (5.7)	998 (5.8)	0.74 (0.64–0.85)	0.80 (0.69–0.93)	3
>24 to 30	248 (2.8)	598 (3.5)	$0.66\ (0.51{-}0.84)$	$0.67\ (0.51{-}0.87)$	40
>30 to 36	179 (2.0)	319 (1.9)	0.65 (0.53–0.81)	$0.69\ (0.54-0.88)$	0
>36	309 (3.5)	686 (4.0)	$0.58\ (0.48-0.70)$	0.67 (0.53–0.83)	18
Per 3 months of total breastfeeding	eastfeeding		0.97 (0.96–0.98)	0.97 (0.96–0.98)	26
Average duration of breastfeeding/child (months) $^{S}$	astfeeding/child	l (months) $§$			
Never breastfed	3049 (34.3)	5566 (32.6)	1.00	1.00	
>0 to 3	2777 (31.3)	5603 (32.9)	0.99 (0.96–1.02)	1.01 (0.98–1.03)	0
>3 to 6	1230 (13.9)	2618 (15.4)	0.92 (0.89–0.97)	0.95 (0.91–0.99)	25
>6 to 9	844 (9.5)	1649 (9.7)	0.92 (0.88–0.97)	0.93 (0.88–0.98)	31
>9 to 12	752 (8.5)	1184 (6.9)	0.90 (0.85–0.95)	0.92 (0.86–0.97)	L
>12 to 18	225 (2.5)	430 (2.5)	0.86 (0.78–0.95)	0.88 (0.79–0.97)	22
Average duration of breastfeeding/child breastfed $^{S\!/\!/}(\mathrm{months})$	astfeeding/child	l breastfed <sup>§//</sup> (n	nonths)		
Never breastfed	1793 (32.9)	2335 (26.4)	1.00	1.00	
>0 to 3	1394 (25.6)	2324 (26.3)	0.98 (0.92–1.04)	0.99 (0.93–1.05)	29
>3 to 6	798 (14.6)	1674 (18.9)	0.92 (0.86–0.98)	0.93 (0.88–0.98)	15
>6 to 9	616(113)	1200 (13.6)	0.90 (0.85-0.96)	0 89 (0 83–0 95)	17

Breastfeeding variables	Cases N(%)*	Controls N(%) <sup>*</sup>	Crude Pooled OR (95%CI)	$\begin{array}{ll} Adjusted \ Pooled & I^{2\%_{0} \ddagger} \\ OR^{ \dagger} & (95\% CI) \end{array}$	$I^{20/0}$
>9 to 12	673 (12.3)	993 (11.2)	0.88 (0.81–0.96)	0.88 (0.81–0.96) 0.87 (0.80–0.95)	21
>12 to 18	176 (3.2)	310 (3.5)	0.86 (0.77–0.96)	0.86 (0.77–0.96) 0.86 (0.78–0.95)	-

OR, odds ratio; CI, confidence interval;

\* May not add to total due to missing data

<sup>7</sup>/Adjusted for age (years), parity (continuous), oral contraceptive duration (continuous; or ever use (yes/no)<sup>1723</sup>, BMI (around reference age, continuous), education (level achieved or number of years).

 $t_{\rm From}$  random effects meta-analysis models. The 12% measures variation across studies that is due to heterogeneity rather than chance.

 ${}^{g}\!\!\!$  OR is for the risk change per child or per child breast fed.

 $l_{\text{Included data only from 12 studies (Turin study, 17-228, 14, 23-25).}$ 

### Table 4

Results of univariable meta-regression assessing factors associated with between-study heterogeneity in estimates for the association between ever-breastfeeding and endometrial cancer risk

Factor	Adjusted R <sup>2</sup> % <sup>*</sup>	P-value from univariable meta-regression
Study design (case-control/cohort)	9.4	0.48
Proportion of white women	31.0	0.1
Median breastfeeding duration amongst women who breastfed	4.4	0.5
Median age at diagnosis	68.5	0.03
Proportion of women whose last birth was $<30$ years before reference date	82.3	0.02
Proportion of women born post-1950	24.9	0.2

Proportion of between-study variance explained