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Association of Episodic Physical and Sexual Activity With Triggering of Acute Cardiac Events:

Systematic Review and Meta-analysis

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Abstract

Context—Evidence has suggested that physical and sexual activity might be triggers of acute cardiac events.

Objective—To assess the effect of episodic physical and sexual activity on acute cardiac events using data from case-crossover studies.

Data Sources—MEDLINE and EMBASE (through February 2, 2011) and Web of Science (through October 6, 2010).

Study Selection—Case-crossover studies investigating the association between episodic physical or sexual activity and myocardial infarction (MI) or sudden cardiac death (SCD).

Data Extraction—Two reviewers extracted descriptive and quantitative information from each study. We calculated summary relative risks (RRs) using random-effects metaanalysis and absolute event rates based on US data for the incidence of MI and SCD. We used the Fisher P value synthesis method to test whether habitual physical activity levels modify the triggering effect and

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meta-regression to quantify the interaction between habitual levels of physical activity and the triggering effect.

Results—We identified 10 studies investigating episodic physical activity, 3 studies investigating sexual activity, and 1 study investigating both exposures. The outcomes of interest were MI (10 studies), acute coronary syndrome (1 study), and SCD (3 studies). Episodic physical and sexual activity were associated with an increase in the risk of MI (RR=3.45; 95% confidence interval [CI], 2.33-5.13, and RR=2.70; 95% CI, 1.48-4.91, respectively). Episodic physical activity was associated with SCD (RR=4.98; 95% CI, 1.47-16.91). The effect of triggers on the absolute rate of events was limited because exposure to physical and sexual activity is infrequent and their effect is transient; the absolute risk increase associated with 1 hour of additional physical or sexual activity per week was estimated as 2 to 3 per 10000 person-years for MI and 1 per 10000 person-years for SCD. Habitual activity levels significantly affected the association of episodic physical activity and MI ($P₀001$), episodic physical activity and SCD ($P₀001$), and sexual activity and MI ($P₀$. 04); in all cases, individuals with lower habitual activity levels had an increased RR for the triggering effect. For every additional time per week an individual was habitually exposed to physical activity, the RR for MI decreased by approximately 45%, and the RR for SCD decreased by 30%.

Conclusion—Acute cardiac events were significantly associated with episodic physical and sexual activity; this association was attenuated among persons with high levels of habitual physical activity.

> Acute cardiac events are a major cause of morbidity and mortality, with as many as a million acute myocardial infarctions (MIs) and 300 000 cardiac arrests occurring in the United States each year.^{1,2} Regular physical activity has been identified as strongly associated with a decreased risk of cardiovascular disease and related mortality.³ Despite the wellestablished benefits of regular physical activity, anecdotal evidence has suggested that physical activity, as well as other acute exposures, such as sexual activity and psychological stress, can act as triggers of acute cardiac events. $4-7$ In fact, in the original description of MI, Obraztsov and Strazhesko⁸ observed that the acute event is often precipitated by exposure to physical or mental stressors.⁹

> Traditional epidemiological designs, such as case-control and cohort studies, are not particularly suitable for identifying acute triggers (proximal causes) of cardiac events, primarily because short-term exposures close to the time of event occurrence are likely to be confounded by patient-level factors. In the early 1990s, the case-crossover design was developed specifically to address the problem of identifying triggers of acute events.^{10,11} A case-crossover study is based on the identification of patients who have experienced the event of interest and requires the assessment of exposure during a relatively brief period preceding the event of interest (the hazard period) and during period(s) when the event did not occur (the control periods). As such, the design allows the individual's exposure status to be assessed during both the hazard and control period(s), in effect serving as his or her own control. Given this unique form of matching, case-crossover studies are not susceptible to confounding by patient characteristics that remain constant over time or to biases in control selection that can compromise case-control studies.

Several case-crossover studies have investigated the association of episodic physical and sexual activity with acute cardiac events, including MI and sudden cardiac death (SCD).¹²⁻¹⁴ These studies tend to produce wide confidence intervals (CIs), indicating considerable uncertainty around the relative risk (RR) estimates. In addition, although it is hypothesized that the triggering effect of episodic physical and sexual activity varies by habitual levels of physical activity, the consistency of this effect modification across studies as well as the magnitude of the potential modification of risk remains unclear.

We conducted a systematic review and meta-analysis of the literature to provide an overview of case-crossover studies investigating episodic physical and sexual activity as triggers of acute cardiac events and to investigate the interaction of habitual physical activity levels with the triggering effect of these exposures.

Methods

Literature Search and Eligibility Criteria

We searched the MEDLINE and EMBASE databases (through February 2, 2011) to identify studies using a case-crossover design and investigating the association of episodic physical or sexual activity with acute coronary syndromes (including MI) and SCD. We used combinations of keywords related to the exposures (exertion, exercise, physical activity, sexual activity), outcomes (myocardial infarction, acute coronary syndrome, sudden [cardiac] death), and the study design of interest (*case-cross-over*). The full search strategy is available in the eAppendix (available at <http://www.jama.com>). To increase the yield of our search, we also searched the reference lists of eligible studies and relevant narrative reviews. We also used the Institute of Scientific Information Web of Science database (through October 6, 2010) to identify all studies citing the studies we considered eligible and reviewed their titles and abstracts to determine eligibility.

One reviewer screened all abstracts (I.J.D.) to identify potentially eligible studies, and 2 reviewers (I.J.D. and J.K.P.) screened all potentially eligible studies in full text to determine eligibility for the review. Eligible studies had to have a case-crossover design, as described by Maclure, 10 and had to examine the effect of episodic physical or sexual activity (exposures) on the risk of acute coronary syndrome (including MI) and SCD (outcomes). Given that we expected studies conducted over a period of several years to be eligible for this analysis, we considered the definitions and diagnostic criteria used in the primary studies for the outcomes of interest. We excluded studies with an experimental design, ie, studies where exposure status was assigned by the investigators, including studies of induced exertion. Finally, we excluded observational studies not using a case-crossover design, including those with ecological, case-control, and cohort designs. We did not use any language restrictions, and we did not consider abstracts presented at scientific meetings, since they are frequently not peer reviewed.

Data Extraction

Two reviewers (I.J.D. and J.K.P.) independently reviewed all eligible studies and extracted data; discrepancies were resolved by consensus. From each eligible study, we extracted the

following information: first author, year and country of publication, case selection methods, criteria used for the diagnosis of the outcome, definitions and measurement methods for ascertaining exposure, duration of hazard and control periods, ascertainment of exposure during control periods, and the RR and its variance for the triggering effect of episodic physical or sexual activity for the outcomes of interest.

For studies that explored effect modification of the triggering effect by habitual physical activity levels (ie, the interaction between habitual activity levels and episodic exposures on the outcomes), we extracted information on the methods of assessing habitual activity levels, the RR with its corresponding variance for each habitual activity stratum, and the midpoint of the exposure levels. For open-ended categories, we assigned a value equal to the lower bound of the category plus half the length of the adjacent category. For example, for a study where exposure was categorized as 1 time/week or less, 2 to 3 times/week, and 4 or more times/week, we assigned values of 0.5, 2.5, and 4.5 times/week, respectively. We performed sensitivity analyses to evaluate the impact of alternative assignment methods.

Assessment of Validity

Although case-crossover studies are protected from confounding from factors that remain stable within an individual over time, there exist other threats to their validity, including potential biases in case selection, determination of hazard and control periods, and confounding due to factors that vary over time.^{10,15} To our knowledge, there are no standard criteria for the assessment of validity of case-crossover studies.16 Instead, based on general epidemiological principles and the particular features of the case-crossover design, $10,15,17,18$ we attempted to assess the validity of the eligible studies by reviewing the following information: the sampling frame used in each study, whether the study was conducted in a single center or multiple centers, the participation rate, whether definitions for outcomes were clearly reported, whether the hazard and control periods were clearly defined, and whether the selection of hazard and control period duration was justified (for example, whether they were chosen based on external evidence or were determined empirically by a sensitivity analysis). $10,15,19$

Evidence Synthesis

To assess between-study heterogeneity for each outcome of interest, we used the RR estimates reported from each study and their variance to calculate the Cochran Q statistic.²⁰ Heterogeneity was considered statistically significant at $P \le 10$. Between-study inconsistency was quantified with the \hat{P} index.²¹ \hat{P} indicates the proportion of between-study heterogeneity that is beyond chance and ranges between 0% and 100%; higher \hat{P} values indicate higher amounts of heterogeneity. We estimated summary RRs with their 95% CIs for each outcome of interest using inverse variance methods with a random-effects model.22,23

Absolute Event Rate Estimates

To demonstrate the association of the triggering effect with an individual's absolute rate of MI or SCD, we used the following methods described in Muller et al.¹⁴ We first obtained estimates of the absolute rate per 100 person-years from the Framingham Heart Study²⁴ (for MI, using data from 1980 to 2003) and the US vital statistics mortality data²⁵ (for SCD, using data from 1989 from 1998) for male and female individuals in the 55- to 64-year age bracket. We then applied the meta-analysis estimates for the RR of the triggering effect to calculate the absolute event rate, assuming an individual increased his or her exposure (to episodic physical or sexual activity) by 1 hour once per week.¹⁴

Effect Modification by Levels of Habitual Physical Activity

It has been hypothesized that the triggering effect of physical and sexual activity on cardiovascular disease is dependent on the individual's habitual activity level, ie, that habitual physical activity interacts with episodic activity. For studies that used the same scale (times/week) to quantify habitual activity, we used random-effects meta-regression to estimate the change in the RR of a cardiac event for a unit change in habitual exertion frequency. Because not all studies used the same scale to quantify habitual activity levels (for example, some used hours/week or ordinal classifications based on subjective questions), it was not possible to include all studies in the same meta-regression. To overcome this limitation, we used the following approach to test the null hypothesis that habitual activity levels do not influence the RR of adverse cardiac outcomes: for each study, we calculated the P value for the null hypothesis that habitual activity does not influence the triggering effect and then synthesized the 1-sided P values using the Fisher P value synthesis method.²⁶

Sensitivity Analyses and Assessment of Bias

Given the large time span of the eligible studies, we explored trends over time using random-effects meta-regression with the year of publication as the explanatory variable.²⁷ To explore whether studies that estimated the RR with greater precision (ie, smaller standard errors) provided different estimates compared with less precise ones, we used the Egger regression-based test for "small study effects"28 (commonly referred to as a "publication bias" test).29 Also, to ensure the robustness of the dose-response meta-analysis to different value assignment methods for the highest (open-ended) category in each study, we performed sensitivity analysis by assigning to those categories: (1) the lower bound of the category and (2) a value equal to the lower bound of the category plus the length of the adjacent category.30 This range of values is broad enough (ie, the assigned values are substantially different) that the robustness of the results of this sensitivity analysis provides convincing evidence for the dose-response relationship.

Software

Statistical analyses were conducted with Stata version SE/11.1 (StataCorp, College Station, Texas) and Meta-Analyst version 3.0 beta (Tufts Medical Center, Boston, Massachusetts).³¹ For all comparisons, except those for heterogeneity, statistical significance was defined as ^P<.05, and all tests were 2-sided.

Results

Our initial searches identified 4914 citations in MEDLINE and 717 citations in EMBASE. After screening titles and abstracts, 70 MEDLINE citations and 19 EMBASE citations were

considered potentially eligible and were retrieved in full text; after removing duplicates, 78 articles were retrieved for full text review. Of these, 66 were excluded (30 were editorials, reviews, or letters to the editor containing no primary data; 19 did not have a case-crossover design; 11 did not investigate outcomes of interest; 4 did not investigate exposures of interest; 1 was an experimental study; and 1 was a translation of an article published in English), and 12 were considered eligible for this systematic review. One additional article was identified by perusal of reference lists, thus we considered 13 articles in total.³² In addition, we screened 1017 citations from Web of Science, of which 42 were considered potentially eligible. Of these, 14 did not report primary data; 14 had been identified by the MEDLINE or EMBASE searches, 13 did not have a case-crossover design, and 1 was a case report. Thus, no additional eligible articles were identified through the Web of Science search. Figure 1 presents the overall search flow. One article reported information from 2 separate populations (geographic regions); when available, estimates from these populations were considered as independent strata in our analyses, resulting in a total of 14 studies reported in 13 articles included in the meta-analysis.¹³

Table 1 summarizes the characteristics of the eligible studies. Additional details about study design and outcome definitions appear in eTable 1. Interrater agreement for data extraction was high $(x>0.9$ both for qualitative and quantitative extraction items). Ten studies provided data on episodic physical activity, $12,13,32,33,35-39$ 3 on sexual activity, $14,40,41$ and 1 study for both exposures of interest.³⁴ Of the studies that assessed episodic physical activity, 7 studies enrolled patients with MI (1 of the articles reported on 2 independent patient populations that were analyzed separately),^{12,13,32-35} 3 enrolled patients with SCD,³⁶⁻³⁸ and 1 enrolled patients with mixed diagnoses of acute coronary syndrome,³⁹ the majority of whom had MI. All of the studies assessing sexual activity as the exposure of interest enrolled only patients with MI.14,34,40,41

Mean or median age was older than 60 years in 9 of 10 studies that reported relevant information. The majority of participants were male in 12 studies that enrolled individuals of both sexes. One study enrolled exclusively men 37 and 1 study exclusively female participants36; both these studies investigated the association of physical activity and SCD.

All 10 studies of episodic physical activity quantified the intensity of the exposure based on multiples of metabolic equivalents (METs). Moderate activity in all studies was fairly uniformly defined as exertion of at least 5 to 6 METs.

Study Validity

Most studies had well-defined sampling frames and reported the diagnostic criteria for the events of interest (eTable 2). Participation rate was variable but was at least 80% in 8 studies and at least 70% in all 13 studies that reported such information. Only 1 study³⁶ did not report the duration of the hazard period used in the primary analysis. Eight of 14 studies either performed sensitivity analyses to identify an "optimal" hazard period or explicitly stated that the choice of hazard period was based on previously published studies. In all studies of MI and in the single study of acute coronary syndrome, exposure ascertainment was performed through interviews, and the majority of studies provided details on the training of personnel collecting the exposure information and the quality control measures

that were implemented. Studies of SCD relied on medical records and third-person interviews. All studies but 1^{41} reported the duration of the control period; of 13 studies that reported relevant information, 7 studies used the habitual exposure to physical or sexual activity over long periods of time (typically 1 year preceding the event of interest) to derive the usual frequency of exposure, 1 study³⁵ only used proximal time periods (24 and 48 hours prior to the event), and 5 studies used both approaches.

Episodic Physical Activity and MI

The 7 studies (in 6 articles) investigating the effect of episodic physical activity on the risk of acute MI enrolled a total of 5503 patients and were published between 1993 and 2008.12,13,32-35 Overall the studies suggested a strong association between episodic physical activity and MI ($RR = 3.45$; 95% CI, 2.33-5.13; $P \le 0.001$), with substantial between-study heterogeneity (ℓ^2 =87%; Q statistic P<.001). Figure 2 presents a forest plot of all outcomeexposure associations assessed in this review.

One study³⁹ of episodic physical activity included patients with any diagnosis of acute coronary syndrome of whom the majority had MI (67.8% of participants had an ST-elevation MI and the remainder had a non–ST-elevation MI or unstable angina). Inclusion of this study in the metaanalysis did not substantially affect the summary estimate (RR=3.93; 95% CI, 2.63-5.89), again with substantial heterogeneity ($\hat{P} = 89\%$; Q statistic P < .001).

Episodic Physical Activity and SCD

Three studies (616 SCD events) assessed the potential of episodic physical activity to trigger SCD.³⁶⁻³⁸ One study was nested within the Physicians' Health Study and enrolled 122 male patients.37 The other study was nested within the Nurses' Health Study and enrolled 288 female patients.³⁶ The third study enrolled a mixed, predominantly male (81%) population.38 Overall there was evidence of an increase in the risk of SCD triggered by episodic physical exertion (RR=4.98; 95% CI, 1.47-16.91; $P=.01$); however, there was substantial between-study heterogeneity ($\hat{I}^2 = 95\%$; Q statistic P<.001) (Figure 2).

Sexual Activity and MI

Four studies (2960 patients) investigated the association between sexual activity and triggering of MI.14,34,41,42 Overall, sexual activity was associated with an increased risk of infarction (RR=2.70; 95% CI, 1.48-4.91; *P*=.001) with moderate heterogeneity (ℓ =64%; *Q* statistic $P = .04$) (Figure 2).

Table 2 presents a summary of the meta-analysis results for all exposure-outcome associations we investigated.

Absolute Event Rate Estimates

Although the RR estimates indicate a large and highly statistically significant increase in the risk of MI and SCD during periods of episodic physical and sexual activity, because these exposures are infrequent and their effect on the outcomes of interest is transient, their impact on an individual's absolute event rate is small. The absolute risk increase associated with an

hour of additional physical or sexual activity per week was estimated as 2 to 3 per 10 000 person-years for MI and 1 per 10 000 person-years for SCD.

Effect Modification by Habitual Activity Levels

All studies provided some measure of the effect of habitual physical activity on the triggering effect of episodic exertion (eTable 3). Overall, subgroups of patients with higher habitual activity levels tended to be less susceptible to the triggering effect of episodic physical activity. However, studies used heterogeneous scales to quantify habitual activity levels. In groups with the lowest habitual activity, the RR for the triggering effect of episodic physical activity ranged from 4.47 to 107 for MI, indicating a very substantial increase in risk during or following exertion. The corresponding range in the highest habitual activity groups was 0.86 to 3.3, indicating much smaller increases in risk. Similar patterns were observed for the associations of sexual activity with MI and episodic physical activity with SCD, although the differences were less pronounced and fewer studies contributed data.

Three studies^{12,13,33} (1897 MI patients) classified habitual exertion based on weekly frequency. We estimated that the RR of MI triggered by episodic physical activity was decreased by approximately 45% for each additional time per week a person was habitually exposed to physical activity (relative RR=0.53; 95% CI, 0.41-0.69; $P=0.001$) (Figure 3). Similarly, 2 studies $37,38$ (328 SCD cases) quantified habitual physical activity by weekly frequency. The RR of SCD triggered by episodic physical activity was decreased by approximately 30% for each additional time per week a person was habitually exposed to physical activity (relative RR=0.70; 95% CI, 0.50-0.99; P=.05) (Figure 3). Studies of sexual activity did not provide adequate data for dose-response analyses.

To further examine the effect of habitual activity by taking into account all available studies (eTable 2), we first performed interaction tests within each study and used the Fisher P value synthesis method to combine the interaction P values across studies. The combined P value tests the null hypothesis that habitual activity levels (regardless of measurement scale) do not affect the triggering RR. The combined P values were $P = 6.81 \times 10^{-20}$ for the episodic physical activity–MI association, $P=1.58 \times 10^{-4}$ for the episodic physical activity–SCD association, and $P = .04$ for the sexual activity–MI association, in all cases suggesting that higher levels of habitual physical activity reduce the acute triggering effect of episodic physical or sexual activity.

Sensitivity Analysis, Trends Over Time, and Small Study Effects

Effect sizes for the association of episodic physical or sexual activity with MI or episodic physical activity with SCD did not appear to significantly change over time (meta-regression ^P values, .71, .49, and .35, respectively). There was no evidence that smaller studies produced different results compared with larger studies. The Egger test P values for the association of episodic physical and sexual activity with MI and the association of physical activity with SCD were .76, .97, and .98, respectively.

Three of the studies investigating the association of episodic physical activity and MI, and 1 of the studies investigating the association of sexual activity and MI, reported estimates from analyses using control periods proximal to their hazard period (typically within 24 hours of

the hazard period).12,14,33,35 Using these estimates in the respective meta-analyses did not substantially modify the results: the summary RR for MI was 3.40 (95% CI, 2.25-5.15) for physical activity and 2.93 (95% CI, 1.43-6.02) for sexual activity.

In sensitivity analyses, assigning different values of weekly exercise frequency to the highest (open-ended) exposure category in the 5 studies that provided relevant data did not substantially modify our results, and the dose-response relationship remained robust both for the association of episodic physical activity with MI ($P=0.001$ both when the highest category was assigned a value equal to its lower bound and when it was assigned the sum of its lowest bound plus the length of the adjacent category) as well as for the association of episodic physical activity with SCD $(P=0.03)$ and 0.08 under the 2 assignment methods, respectively).

Comment

Several lines of evidence suggest that acute cardiac events, including MI and SCD, may be triggered by short-term exposures, such as episodic physical activity and emotional stress.12-14,43 Multiple studies have established the existence of circadian variation in the incidence of acute cardiovascular events, with the lowest rate being observed during sleep, a peak during the early morning hours, and a steady lower level during the remainder of the day.43-46 Multiple biological parameters, such as blood pressure, heart rate, blood viscosity, platelet aggregability, and adrenergic activity, follow similar circadian patterns, suggesting that exposures that can affect the levels of these factors may exert a triggering effect on cardiovascular events.47,48 The case-crossover study was designed specifically to allow the reliable identification of triggers of acute outcomes.^{10,17} Our review of 14 case-crossover studies investigating the association of episodic physical and sexual activity on acute MI, acute coronary syndrome, and SCD indicates that both exposures are associated with a statistically significant short-term increase in the risk of the outcomes of interest. In addition, we demonstrate that this increase in risk is strongly modified by habitual activity levels, with individuals with higher activity levels experiencing lower increase (and often no increase) in risk after exposure to triggering activities.

The counterfactual hypothesis of case-crossover studies is distinct from that of case-control studies, making the former suitable for answering the question "Why did the event occur now?" compared with the question "Why did the event occur in this individual?" typically the purview of case-control studies.49 Because of this unique property of the case-crossover design, our results are not incompatible with the well-established beneficial effect of regular physical activity on the risk of acute coronary events³: active individuals are overall at a lower risk of such events compared with inactive individuals; however, during the short time period of acute exposure to physical or sexual activity, an individual's risk of an event is increased compared with unexposed periods of time. This conclusion is supported by our absolute event rate calculations, which demonstrate that the short-term effect of the triggers we evaluated is unlikely to be large in absolute terms. In contrast, regular physical activity may reduce an individual's absolute risk by more than 30%.⁵⁰

This temporary increase in risk may be modified by baseline characteristics of each individual. To explore whether habitual physical activity levels modify the triggering effect

of episodic physical and sexual activity, we used data from studies reporting effect sizes stratified by habitual exertion levels. We found strong evidence to reject this null hypothesis, indicating that habitual activity levels modify the triggering effect, with individuals having the lowest habitual levels of physical activity experiencing the highest triggering risk for all exposures and outcomes. Although the Fisher P value synthesis method allowed us to combine evidence from all studies, it cannot provide an estimate of the magnitude of the interaction between habitual physical activity levels and the triggering effect of episodic exertion. Our estimates of this dose-response relationship, although based on a smaller number of studies (3 for MI and 2 for SCD), indicated that for each additional time an individual is exposed to episodic physical activity per week, the RR of the triggering effect is reduced by approximately 45% and 30%, respectively. This suggests that individuals with high levels of habitual exertion experience a much smaller increase in the risk of MI and SCD during the hazard period of episodic physical exertion compared with individuals with low levels of habitual activity. Clinically, this result suggests that physicians counseling patients regarding their exercise habits may need to tailor their advice to the patients' habitual activity levels: sedentary individuals should be counseled to increase the frequency and intensity of physical activity gradually.

Several limitations should be considered when interpreting our findings. First, there was substantial statistical heterogeneity for all outcomes and exposures of interest, and differential recall (information bias) of exposure status during hazard and control periods may also have affected the estimates of the RR, particularly in studies of SCD. However, studies followed similar protocols and had well-defined hazard and control periods (at least within each trial). In addition, methods for exposure ascertainment were generally reported in detail, and all studies classified physical activity using fairly uniform thresholds (5-6 METs). In studies of MI, exposure ascertainment was based on structured interviews and extensive efforts to standardize the process and perform quality control. In contrast, exposure ascertainment in studies of SCD was based on review of medical records and thirdperson interviews, suggesting that mis-classification or information bias may have been present. Studies were conducted over a relatively long period of time, over which the diagnostic criteria and diagnostic tests for their verification have substantially changed. Despite this, meta-regression analysis using the publication year as an explanatory variable did not reveal any trends over time. As with all syntheses of observational studies, the estimated average effect across studies and the hypothesis-generating results of our metaanalysis should not be overinterpreted.⁵¹ Further prospective studies may be necessary to confirm our findings, particularly for the association between sudden cardiac death and episodic physical activity, for which extreme between-study heterogeneity was observed.

Case-crossover studies are susceptible to confounding by factors that change over time. The co-occurrence of episodic physical or sexual activity with other exposures that can trigger MI or SCD (for example, smoking, coffee consumption, illicit drug use, or emotional stress) may account for the observed RR estimates. In addition, premonitory symptoms may affect individuals' exposure to physical or sexual activity and thus confound the association with the outcomes of interest; however, such an association would tend to bias the estimates toward the null. Control for such confounding is typically hard to implement and depends on the method of sampling control person-time and the collection of data on potential

confounders.10,18,19 The consistency of findings across studies conducted over more than 15 years in diverse populations, as well as the robustness of the estimates in sensitivity analysis using alternative control period durations, provides reassurance that our findings are indicative of a causal association. In addition, some of the studies we considered attempted to adjust for potential confounders, and in all cases the results remained unchanged. However, we caution that additional confounding by unknown (unobserved) time-related factors may also have been present.

In conclusion, based on our review of 14 case-crossover studies of acute cardiac events, we found a significant association between episodic physical and sexual activity and MI and suggestive evidence of an association between episodic physical activity and SCD. Most importantly, these associations appear to be strongly modified by habitual physical activity, with individuals with higher habitual activity levels experiencing much smaller increases in risk compared with individuals with low activity levels. In view of this, as well as the small absolute magnitude of the risk associated with acute exposure to episodic physical or sexual activity, our findings should not be misinterpreted as indicating a net harm of physical or sexual activity; instead they demonstrate that these exposures are associated with a temporary short-term increase in the risk of acute cardiac events.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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Figure 1. Search Strategy Flowchart

One publication provided results from 2 separate populations, which were considered as independent strata in our analyses. ISI indicates Institute of Scientific Information.

Figure 2. Forest Plot of Studies Assessing the Association of Episodic Physical and Sexual Activity With Myocardial Infarction and Sudden Cardiac Death

Summary results for each exposure-outcome subgroup are presented separately. Within each subgroup, studies are arranged by the point estimate of relative risk. Values greater than 1 indicate that exposure is associated with increased risk of the outcome. Squares are proportional to the weight of each study in the meta-analysis. CI indicates confidence interval.

Figure 3. Meta-regression Graph of Triggering Relative Risks for Myocardial Infarction and Sudden Cardiac Death Over Habitual Physical Activity Levels

Meta-regression graph of relative risks (RRs) for myocardial infarction is based on 3 studies that used the same scale (weekly frequency) to report habitual activity levels^{12,13,33}; graph for sudden cardiac death is based on 2 studies that used the same scale.^{37,38} Subgroup estimates are depicted as circles proportional to their precision (inverse of the variance of the log[RR]). The solid line indicates fitted values by random-effects meta-regression. rRR indicates relative RR calculated from the meta-regression.

sensitivity analysis on

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Of 699 interviewed patients, 40 were excluded because of missing, unreliable, or inaccurate information regarding time of disease onset or exposure.

 $\mathscr{E}_{\text{The study reported results for multiple hazard periods.}$ For consistency with the other studies investigating the sexual activity–MI association, we used the results of analyses based on a 2-hour hazard ^gThe study reported results for multiple hazard periods. For consistency with the other studies investigating the sexual activity-MI association, we used the results of analyses based on a 2-hour hazard
period.

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Table 2
Meta-analysis Results for the Association of Physical and Sexual Activity With the Triggering of Myocardial Infarction and Sudden Cardiac **Meta-analysis Results for the Association of Physical and Sexual Activity With the Triggering of Myocardial Infarction and Sudden Cardiac Death**

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Abbreviations: CI, confidence interval; MI, myocardial infarction; RR, relative risk; SCD, sudden cardiac death Abbreviations: CI, confidence interval; MI, myocardial infarction; RR, relative risk; SCD, sudden cardiac death

a P values from Cochran Q statistic. Heterogeneity was considered statistically significant at P< .10