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## Mitochondrial genetic haplogroups and depressive symptoms: A large study among people in North America

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### Abstract

**Background**—A possible relationship between mitochondrial haplogroups and psychiatric diseases (e.g. schizophrenia and bipolar disorder) has been postulated, but data regarding depression is still limited. We investigated whether any mitochondrial haplogroup carried a significant higher risk of depressive symptoms in a large prospective cohort of North American people included in the Osteoarthritis Initiative.

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#### Authors' contribution

NV, AK: made the statistical analysis; NV, BS, MS: wrote the paper; AV, JD, TT: analysis of the data; AFC, MZ, SM: critical revision of the paper. All authors have approved the final article.

**Methods**—Cross sectional data was derived from the Osteoarthritis Initiative. The haplogroup was assigned through a combination of sequencing and PCR-RFLP techniques. All the mitochondrial haplogroups were named following this nomenclature: H, U, K, J, T, V, SuperHV, I, W, X or Others. Depression was ascertained through the 20-item Center for Epidemiologic Studies-Depression (CES-D), with 16 indicating depressive symptoms.

**Results**—Overall, 3601 Caucasian participants (55.9% women), mean age of  $61.7 \pm 9.3$  years were included. No difference was observed in mitochondrial haplogroups frequency among those with depressive symptoms ( $n=285$ , =7.9% of the baseline population) compared to participants with no depressive symptoms ( $N=3316$ ) (chi-square test=0.53). Using a logistic regression analysis, adjusted for eight potential confounders, with those having the haplogroup H as the reference group (the most common haplogroup), no significant mitochondrial haplogroup was associated with prevalent depressive symptoms. The same results were evident in secondary analysis in which we matched depressed and non-depressed participants for age and sex.

**Limitations**—Cross-sectional design; only CES-D for evaluating mood; participants not totally representative of general population.

**Conclusions**—We found no evidence of any relationship between specific mitochondrial haplogroups and depressive symptoms. Future longitudinal research is required to confirm/refute these findings.

## Keywords

Mitochondrial haplogroups; Depression; Osteoarthritis initiative

## 1. Introduction

The human mitochondrial genome is a circular set of 16,569 base pairs encoding 37 genes and finally 13 proteins, involved in processes relevant to cellular function and survival. (Wang and Brinton, 2016) Similarly to nuclear DNA mitochondrial DNA (mtDNA) undergoes frequent mutations. (Copeland and Longley, 2014; Wang and Brinton, 2016) Mismatch, usually due to a recombination or repair, can lead to single nucleotide polymorphisms (SNPs) and clusters of these specific SNPs in the mitochondrial genome define the mitochondrial haplogroups. The evolution of mtDNA occurs at a much faster pace compared to the average nuclear DNA, and thus mutations have accumulated sequentially, even if only along radiating maternal lineages. (Wallace and Chalkia, 2013) Therefore, it is hardly surprising that the biology of mitochondrial DNA may partly explain the genetic predisposition to certain medical conditions, probably in a similar way to the germline mutations of nuclear DNA (Luchini et al., 2016; Mafficini et al., 2016). Recent studies have proposed that the different mitochondrial haplogroups might play a role in the development of several chronic disease states (Mishmar et al., 2003; Picard et al., 2016; Wallace and Chalkia, 2013). Each of the mitochondrial haplogroups, in fact, is probably involved not only in mitochondria physiology, but also in cell metabolism (Horan and Cooper, 2014; Wallace, 2015). Thus, organs at high consumption of energy (such as brain) are the main target of mitochondrial dysfunction (Mink et al., 1981).

From an epidemiological viewpoint, several psychiatric conditions, such as schizophrenia and mood disorders have a preferential transmission through the maternal line (Kirk et al., 1999). Moreover, more recent studies reported that mood disorders are more frequent in maternal relatives of children with mitochondrial diseases, compared with their paternal relatives, or compared with maternal relatives of the children with other metabolic diseases (Grigoriou-Serbanescu et al., 1998; Kato, 2001; Linnane et al., 1989). In this argument, some authors reported an increased mtDNA- copy number in peripheral cells and plasma in depressed (Cai et al., 2015; Nicod et al., 2016) and suicidal (Lindqvist et al., 2016) subjects. Finally, several mitochondrial encephalomyopathies are characterized by a high prevalence of bipolar disorder and depression (Kato, 2001). Taken together, these findings suggest that mitochondrial dysfunction may be associated with mood disorders, including depression.

Even if several experimental and pre-clinical pieces of research report that mitochondrial dysfunction is plausible in the development of depression (e.g. depressed people use less ATP in some brain regions) (Gardner et al., 2003; Karabatsiakos et al., 2014), only two studies have investigated an association between mitochondrial haplotypes and the presence of depression in human being not affected by mitochondrial myopathies (Rollins et al., 2009; Zhang et al., 2016). Although these studies reported a significant association between mitochondrial haplotypes and mental illness, they were limited in sample size suggesting that further research is needed.

Given this background, we aimed to investigate whether any mitochondrial haplogroup carried a significant higher risk of depressive symptoms in a large cohort of North American people included in the Osteoarthritis Initiative.

## 2. Methods

### 2.1. Data source and subjects

All participants in this cross-sectional study were recruited as part of the ongoing, publicly and privately funded, multicenter Osteoarthritis Initiative (OAI study) (<http://www.oai.ucsf.edu/>). The specific dataset used was the OAI baseline data (November 2008) (V00). The OAI recruited participants with/at high risk of knee osteoarthritis (e.g. obese, with familiarity for OA) from four clinical sites in the US (Baltimore, MD; Pittsburgh, PA; Pawtucket, RI; and Columbus, OH) between February 2004 and May 2006. For this study, specific exclusion criteria are: pregnancy, presence of active rheumatoid arthritis, bilateral total knee replacement, unable to undergo knee magnetic resonance or to give a blood sample, any co-morbidity precluding the participation to the study and unwilling to sign the informed consent.

All participants provided written informed consent. The OAI study protocol was approved by the institutional review board of the OAI Coordinating Center, University of California at San Francisco.

### 2.2. Exposure

The haplogroup assignment was performed by a previously published method (Rego-Perez et al., 2008), which consisted in a combination of sequencing and PCR-RFLP techniques.

The sequencing technique consisted in the multiplex assignment of the main 6 SNPs that contribute to the generation of the most prevalent Caucasian haplogroups (Torrioni et al., 1996) (H, V, super HV, U, K, T, J) following the single base extension (SBE) assay. All the mitochondrial haplogroups have been named following this nomenclature in agreement with those suggested by the OAI (<http://www.oai.ucsf.edu/>): H, U, K, J, T, V, SuperHV, I, W, X or Others.

### 2.3. Outcome - depressive symptoms

The presence of depressive symptoms was derived from the 20-item Center for Epidemiologic Studies-Depression (CES-D) instrument. (Radloff, 1977) The range of possible values for this scores is 0–60, where higher scores indicate more depressive symptoms (Radloff, 1977). A cut-off of 16 is commonly used for the diagnosis of depressive symptoms (Radloff, 1977).

### 2.4. Covariates

A number of variables were identified from the OAI dataset to explore the relationship between mitochondrial haplogroups and depression. These included: (1) physical activity evaluated through the Physical Activity Scale for the Elderly (PASE) (Washburn et al., 1999). The PASE is validated in older populations, covers 12 different activities, such as walking, sports, and housework, and is scored from 0 to 400 and more; (2) race was defined as “white” vs. “others”; (3) smoking habits as “previous/current” vs. never; (4) educational level was categorized as “degree” vs. others; (5) yearly income as < vs. 50,000 \$ or missing data; (6) co-morbidities assessed through the modified Charlson comorbidity score, with higher scores indicating an increased severity of conditions (Katz et al., 1996); (7) body mass index (BMI), as recorded by a trained nurse.

### 2.5. Statistical analyses

For continuous variables, normal distributions were tested using the Kolmogorov-Smirnov test. The data are shown as means and standard deviations (SD) for quantitative measures, and percentages for all discrete variables by presence of depression. For continuous variables, Student’s *T* test was used, whilst chi-square test was applied for discrete variables. Levene’s test was used to test the homoscedasticity of variances and, if its assumption was violated, then Welch’s ANOVA was used.

The strength of the association of mitochondrial haplogroups and depressive symptoms was assessed through a logistic regression analysis. Factors significantly associated with depression (taking a *p*-value < 0.05 as statistically significant) were included. Multi-collinearity among covariates was assessed through variance inflation factor, taking a cut-off of 2 as a reason of exclusion, but no variable was excluded due to this reason. The basic model was not adjusted for any confounders, while the fully adjusted model included baseline values of: age, BMI, PASE score, Charlson comorbidity index as continuous variables, and gender, education, smoking habits, yearly income as categorical variables. Data of logistic regression analysis were reported as odds ratios (ORs) with 95% confidence intervals (CIs).

In a secondary analysis, we used controls (matched for age and sex) to depressed participants in order to attenuate the presence of any selection bias (Pearce, 2016).

All analyses were performed using the SPSS 17.0 for Windows (SPSS Inc., Chicago, Illinois). All statistical tests were two-tailed and statistical significance was assumed for a p-value < 0.05.

### 3. Results

#### 3.1. Study participants

At baseline, among 4796 potentially eligible individuals, 1047 subjects did not have a mitochondrial DNA assessment, 126 did not have this assessment for technical problems and 22 did not have information regarding CES-D. Thus, 3601 participants were enrolled in the current study. Compared to those included in this analysis, the 1195 people not enrolled were not significantly different in terms of age and gender, but there was a higher presence of non-Caucasian ethnicities (chi-square test,  $p < 0.0001$ ) (details not shown).

#### 3.2. Baseline analyses

The 3601 participants included aged a mean of  $61.7 \pm 9.3$  (range: 45–79) years, with a slight higher prevalence of women (=55.9%). All participants were Caucasians.

Participants were divided into those with depressive symptoms ( $n=285$ , 7.9% of the sample) and without ( $n=3316$ ). The baseline characteristics of both groups are summarized in Table 1. Briefly, people with depressive symptoms were significantly older, more likely to be females, and obese compared to the non-depressed group. People with depressive symptoms were more frequently smokers, with a lower educational level, poorer and with more comorbidities than those with no depressive symptoms (Table 1).

#### 3.3. Association between mitochondrial haplogroups and depression

As reported in Fig. 1, no significant differences emerged regarding in the prevalence of mitochondrial haplogroups (chi-square test=0.53). For example, the most frequent haplogroup (i.e. H) had the same prevalence among the depressed and no depressed (42.5% in both groups).

Using a logistic regression analysis, adjusted for eight potential confounders, with those having the haplogroup H as the reference, no significant mitochondrial haplogroup was associated with a prevalent depressive symptoms. The OR ranged from 0.50 (95%CI: 0.18–1.43) of other haplogroup to 1.74 (95%CI: 0.75–4.07) of haplogroup W (Table 2).

In a secondary analysis, we explored whether using controls matched for age and sex, influenced our results, but the association between mitochondrial haplogroups and depressive symptoms remained not significant (Table 2).

## 4. Discussion

In this large cross-sectional paper, we found there was no significant association between mitochondrial haplotypes and depressive symptoms. Our findings remained unaltered after adjusting for potential confounders or using controls matched for sex and age. Thus, our data suggest that mitochondria haplogroups may not be related to depressive symptoms.

The interest of mitochondrial haplotypes in the development of psychiatric diseases is rapidly increasing. Several studies have reported that mitochondrial haplotypes are associated with the presence of some psychiatric conditions, particularly schizophrenia and bipolar disorder (Clay et al., 2011; Rajasekaran et al., 2015). The mechanisms linking mitochondrial haplogroups and related dysfunction and these conditions may be attributed to altered oxidative/nitrosative stress responses which could activate the immunoinflammatory pathways (Bano et al., 2017; Soysal et al.; Soysal et al., 2016) and subsequently lead to neuroprogressive changes, typical of bipolar disorder and schizophrenia (Clay et al., 2011; Rajasekaran et al., 2015). Since the same mechanisms are relevant for the development of depression, some researchers have tried to disentangle if any association exists between mitochondrial haplogroups and depressive mood. These studies, however, are very limited and included few participants. Initially Suomalainen et al. reported in a family affected by a chronic progressive ophthalmoplegia that the presence of depression was very common (Suomalainen et al., 1992). These authors interpreted this finding as a possible sign of mitochondrial encephalopathy. Similarly, Suzuki and collaborators investigated 15 people affected by a form of mitochondrial diabetes founding that four of them were affected by mood disorders (Suzuki et al., 1997). More recently, Rollins et al. showed that, in a peculiar subgroup of patients with schizophrenia, the rate of synonymous base pair substitutions in the coding regions of the mtDNA genome was significantly higher compared to controls (Rollins et al., 2009). However the same results were not confirmed in patients with depression compared with controls. These findings highlight the concept of a higher mtDNA instability in some mental disorders. Finally, in 2016, Zhang et al., reported in 744 Chinese subjects with major depression matched with 767 Chinese healthy controls that the T-C haplotype was significantly associated with depression (Zhang et al., 2016). Contrary to results reported by Zhang et al., we did not find any significant association between mitochondrial haplogroups and depressive symptoms in our study. Several hypotheses could explain our findings. First, the studies cited before (except the last two) were made among populations affected by mitochondrial diseases and without a control group. Therefore, it is difficult to say whether a particular mitochondrial haplogroup is associated or not with depression. Second, since our knowledge regarding the possible effect of mitochondrial haplotypes on depression derives from the reported similarities with other psychiatric diseases and from some indirect findings (such as people with depression have reduced blood flow and metabolism in the prefrontal cortex compared to healthy controls) (Videbech, 2000), thus it is possible that mitochondrial haplotypes do not play a role in the development of this disease. It is noteworthy, for example, that Stine et al. in postmortem samples found no significant alterations in mtDNA in nine depressed people who completed suicide compared to controls (Stine et al., 1993). However our study suggests further studies are required to disentangle the potential relationship between mitochondria haplogroups and

depression. Moreover, Zhang et al. (2016) compared patients with a full diagnosis of MDD with healthy controls, thus including people with more pronounced depressive illness than ours which included people with depressive symptoms. Also, while we focused on a Caucasian population, Zhang et al. (2016) focused on a Chinese sample, which might have different mitochondrial haplotypes' distribution. Furthermore, we did not account for stress in our analyses; since it has been suggested that stress can modify amount of mtDNA, our findings may have been biased by neglecting stress as confounding factor. Finally, it may be that there were too few participants within each mitochondria haplogroups and thus some of our analyses were underpowered.

A special consideration should also account for the close relation between mitochondrial functioning and ageing, in elderly population. Ageing is per se associated with a weakened mitochondrial efficiency (Chistiakov et al., 2014), and mitochondrial age-related impairment may also play a role in depression onset (Harper et al., 2016), beyond haplotypes' distribution.

#### 4.1. Limitations

The findings of our paper should be interpreted within its limitations. First, the diagnosis of depressive symptoms was made only using the CES-D that, however, includes many symptoms defined by the (DSM-V) for a major depressive episode and therefore is somewhat justifiable (Veronese et al., 2016). Second, we did not assess the role of medications used for depression and this could introduce an important bias. Third, the number of people in some haplogroups was limited. Fourth, the medical diseases in the OAI are self-reported, and this could introduce another potential bias. Fifth, data was only available of mitochondrial haplogroups among Caucasians and the results do not extend beyond this ethnicity. Sixth, the OAI did not contain detailed information regarding other psychiatric diseases that seem to be associated with mitochondrial haplogroups and this could affect our results. Finally, the OAI included only people with or at high risk of knee osteoarthritis and so it could be not fully representative of the general population.

#### 4.2. Conclusion

In conclusion, in this large cohort study we did not find a relationship between any mitochondrial haplogroups with depressive symptoms. Further research with different diagnostic criteria for depression and with a longitudinal design are needed to disentangle if any mitochondrial haplotype is associated with depression.

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#### Role of founding source

The funding sources did not have any role in study design; in the collection, analysis and interpretation of data; in the writing of the report; and in the decision to submit the article for publication.

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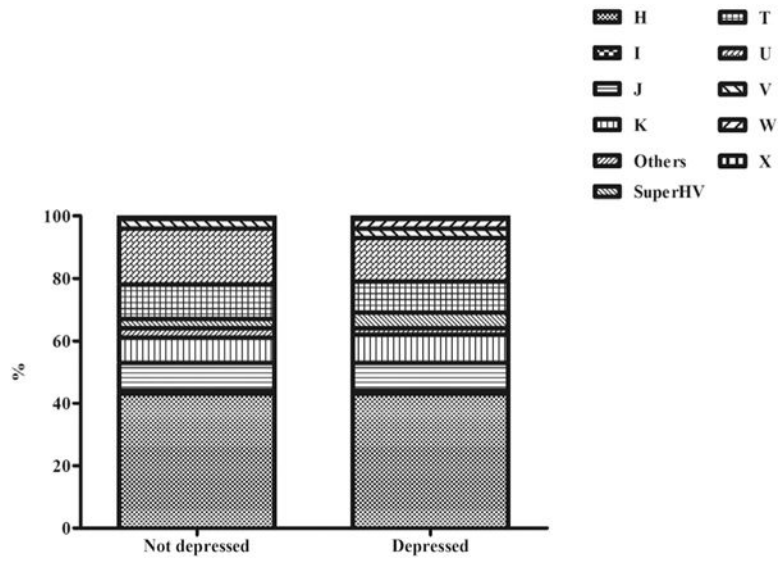
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**Fig. 1.** Prevalence of depressive symptoms by mitochondrial haplogroups. The prevalence of each haplogroup is represented as percentage in people with depressive symptoms or not.

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**Table 1**

Baseline characteristics by depression status.

Variable	Depressive symptoms (n=285)	No depressive symptoms (n=3316)	p-value
Age (years)	61.9 (9.3)	59.3 (9.3)	< 0.0001
Females (%)	63.2	55.3	0.001
BMI (kg/m <sup>2</sup> )	29.1 (5.5)	28.0 (4.5)	< 0.0001
White race (%)	All whites		
PASE (points)	156.5 (81.1)	164.1 (80.0)	0.13
Smoking (%)	55.6	45.5	0.001
Degree (%)	24.6	34.3	0.001
Yearly income (< 50,000 \$)	46.5	68.7	< 0.0001
Charlson co-morbidity score	0.5 (0.8)	0.3 (0.8)	< 0.0001

Notes:

Numbers are mean values (and standard deviations) or percentages, as appropriate.

Abbreviations: BMI: body mass index; PASE: physical activity scale for the elderly; CES-D: Center for Epidemiological Studies Depression.

**Table 2**

Association between mitochondrial haplogroups and Depressive symptoms.

	All cohort						
	Unadjusted			Matched (for age and sex)			
	Unadjusted OR (95%CI)	P value	Fully-adjusted OR <sup>a</sup> (95%CI)	P value	Unadjusted OR (95%CI)	Fully-adjusted OR <sup>a</sup> (95%CI)	P value
H	1 [reference]		1 [reference]		1 [reference]	1 [reference]	
I	0.77 (0.23–2.50)	0.66	0.57 (0.17–1.91)	0.36	0.51 (0.15–1.74)	0.37 (0.11–1.28)	0.12
J	1.05 (0.67–1.64)	0.83	0.99 (0.63–1.56)	0.97	1.00 (0.62–1.62)	0.93 (0.56–1.52)	0.76
K	1.21 (0.77–1.91)	0.41	1.32 (0.83–2.11)	0.24	1.30 (0.79–2.13)	1.36 (0.81–2.28)	0.25
Others	0.61 (0.22–1.70)	0.35	0.50 (0.18–1.43)	0.20	0.72 (0.24–2.12)	0.45 (0.14–1.43)	0.18
SuperHV	1.39 (0.72–2.67)	0.33	1.22 (0.63–2.38)	0.56	1.08 (0.54–2.18)	1.07 (0.52–2.22)	0.85
T	0.99 (0.65–1.51)	0.97	1.02 (0.66–1.57)	0.94	0.95 (0.61–1.50)	0.94 (0.58–1.50)	0.78
U	0.76 (0.52–1.11)	0.16	0.78 (0.53–1.14)	0.20	0.84 (0.56–1.27)	0.82 (0.54–1.50)	0.37
V	0.89 (0.42–1.88)	0.89	0.83 (0.39–1.79)	0.64	0.66 (0.30–1.43)	0.56 (0.25–1.27)	0.17
W	1.91 (0.84–4.35)	0.12	1.74 (0.75–4.07)	0.20	2.12 (0.83–5.44)	2.17 (0.80–5.91)	0.13
X	1.15 (0.40–3.26)	0.80	1.07 (0.36–3.15)	0.91	1.05 (0.34–3.23)	1.00 (0.32–3.21)	0.99

Notes:

All the data are presented as odds ratios (ORs) with their 95% confidence intervals.

Abbreviations: CI: confidence intervals; OR: odds ratio.

<sup>a</sup>Fully-adjusted model included as covariates: age (as continuous), sex; body mass index (as continuous); education (degree vs. others); smoking habits (current and previous vs. others); yearly income (categorized as or < 50,000\$ and missing data); Physical Activity Scale for Elderly score (as continuous); Charlson co-morbidity index (as continuous).