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# Inconsistency between univariate and multiple logistic regressions

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**Summary:** Logistic regression is a popular statistical method in studying the effects of covariates on binary outcomes. It has been widely used in both clinical trials and observational studies. However, the results from the univariate regression and from the multiple logistic regression tend to be conflicting. A covariate may show very strong effect on the outcome in the multiple regression but not in the univariate regression, and vice versa. These facts have not been well appreciated in biomedical research. Misuse of logistic regression is very prevalent in medical publications. In this paper, we study the inconsistency between the univariate and multiple logistic regressions and give advice in the model section in multiple logistic regression analysis.

**Key words**: Conditional expectation; model selection; logistic regression

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### 1. Introduction

Many medical studies have binary primary outcomes. For example, to study the treatment effect of a new intervention on patients with severe anxiety disorders, patients are randomized to the new intervention or treatment as usual (control) groups. The outcome is significant clinical improvement (yes or no) within a period such as 12 months. For this kind of outcome, we use 1 (0) to denote the occurrence or success (no occurrence or failure) of the outcome of interest such as significant (no significant) clinical improvements. The treatment effects can be measured by the difference or ratio of success rates in the two groups. Pearson's chi-square test (or Fisher's exact test) can be easily used if the treatment effect of the treatment method is better than the current method.

It is not uncommon that treatment effect is confounded by differences between treatment groups such as age, medication use and comorbid conditions.

If such confounding covariates are categorical, such as gender and smoking status, contingency table methods can be easily used to study treatment differences. For continuous covariates such as age, although still possible to apply such methods by categorizing them into categorical variables, results depend on how continuous variables are categorized such as the number of end cut-points for categories.

The multiple logistic regression<sup>[1]</sup> provides a more objective approach for studying effects of covariates on the binary outcome. It addresses both categorical and continuous covariates, without imposing any subjective element to categorize a continuous covariate. Coefficients of continuous as well as noncontinuous covariates, which are readily obtained using well-established estimation procedures such as the maximum likelihood, have clear interpretation. Also, its ability to model relationships for case-control studies has made logistic regression one of the favorite statistical models in epidemiologic studies.<sup>[2]</sup>

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Model selection offers advantages of increasing power for detecting as well as improving interpretation of effects of covariates on the binary outcome, especially when there are numerous covariates to consider. Here is how model selection was carried out in multiple logistic regression in a paper recently published in JAMA surgery<sup>[3]</sup>:

'Associations between preoperative factors and adenocarcinoma or HGD were determined with univariate binary logistic regression analysis. Variables with statistically significant association on univariate analysis were included in a multivariable binary logistic regression model.'

Such a univariate analysis screening (UAS) method to select covariates for multiple logistic regression has been widely used in research studies published in top medical journals [4-6] since it seems very intuitive, reasonable, and easy to understand. In this paper we take a closer look at this popular approach and show that the UAS is guite flawed, as it may miss important covariates in the multiple logistic regression and lead to extremely biased estimates and wrong conclusions. The paper is organized as follows. In Section 2 we give a brief overview of the logistic regression model. In Section 3 we study the relationship between the univariate regression analysis, the basis for selecting covariates for further consideration in multiple logistic regression, and the multiple logistic regression model. In Section 4 we use the theoretical findings derived, along with simulation studies, to show the flaws of the UAS. In Section 5, we give our concluding remarks.

### 2. Logistic regression model

We use Y=1 or 0 to denote 'success' or 'failure' of the outcome. Here 'success' and 'failure' only indicate two opposite statuses and should not be interpreted literally. For example, if we are interested in the relation between the exposure of high density of radiation and cancer, we can use Y=1 to denote that the subject develops cancer after the exposure. Aside from the outcome, we also observe some factors (covariates) which may have significant effects on the outcome, denoting them by  $X_1, X_2, ..., X_p$ . The relation between the outcome and the covariates is characterized by the conditional probability distribution of Y given  $X_1, X_2, ..., X_p$ . In multiple logistic regression, the conditional distribution is assumed to be of the following form

$$\Pr\{Y=1\big|X_1,\dots,X_p\}=\frac{\exp\bigl(\beta_0+\beta_1X_1+\dots+\beta_pX_p\bigr)}{1+\exp\bigl(\beta_0+\beta_1X_1+\dots+\beta_pX_p\bigr)}, \qquad (1)$$

where  $\beta_1\beta_2...\beta_p \neq 0$ . This is the model on which our following discussions will be based. The covariates may include both continuous and categorial variables. A more familiar equivalent form of (1) is

$$log \frac{Pr\{Y=1|X_{1},...,X_{p}\}}{Pr\{Y=0|X_{1},...,X_{n}\}} = \beta_{0} + \beta_{1}X_{1} + \cdots + \beta_{p}X_{p},$$

where the left hand side is called the conditional logodds.

Given a random sample, the parameters  $(\theta_0, \theta_1, \dots, \theta_p)$  in (1) can be easily estimated by maximum likelihood estimation (MLE) method, see for example.<sup>[7,8]</sup>

### 3. Univariate regression model

Suppose we are interested in the marginal relation between the outcome and a single factor  $X_1$ , i.e. we need to find  $Pr\{Y = 1|X_1\}$ .

From the property of conditional expectation<sup>[9]</sup> we know that

$$\begin{aligned} \Pr\{Y = 1 | X_1\} &= \mathbb{E}\{\Pr\{Y = 1 | X_1, \dots, X_p\} | X_1\} \\ &= \mathbb{E}\left\{\frac{\exp(\beta_0 + \beta_1 X_1 + \dots + \beta_p X_p)}{1 + \exp(\beta_0 + \beta_1 X_1 + \dots + \beta_p X_p)} \, | X_1\right\}. \end{aligned} \tag{2}$$

If the joint distribution of  $X_1$ ,  $X_2$ ,..., $X_p$  is unknown, generally it is impossible to find the analytical form of (2). In this section we consider the univariate regression model with following some specific distributions.

### 3.1 Univariate regression with categorical covariate

First assume  $X_1$  is a 0-1 valued covariate. For example, in the randomized clinical trial, we can use  $X_1$  as the group indicator (=1 for the treatment group and for the control group). It is easy to prove that there exist unique constants  $\alpha_0$  and  $\alpha_1$  such that

$$\Pr\{Y = 1 | X_1\} = \frac{\exp(\alpha_0 + \alpha_1 X_1)}{1 + \exp(\alpha_0 + \alpha_1 X_1)},\tag{3}$$

where both  $\alpha_o$  and  $\alpha_1$  are functions of  $\theta_o$ ,  $\theta_v$ , ...,  $\theta_p$ . Usually the form of these functions are complex as they depend on the joint distribution of  $X_1$ ,  $X_2$ ,..., $X_p$ . There is no obvious qualitative relation between  $\alpha_1$  in (3) and  $\theta_1$  in (1).

Equation (3) indicates the marginal relation between Y and  $X_1$  still satisfies the logistic regression model and

$$\log \frac{\Pr\{Y=1|X_1=1\}\Pr\{Y=0|X_1=0\}}{\Pr\{Y=1|X_1=0\}\Pr\{Y=0|X_1=1\}} = \alpha_{1^{j^*}}$$

which means that  $\alpha_1$  in (2) is the log odds ratio. Furthermore, if  $X_1$  is independent of  $(X_1, X_2, ..., X_p)$ , we can prove that (i)  $\alpha_1 > 0$  if and only if  $\beta_1 > 0$ , (ii)  $\alpha_1 < 0$  if and only if  $\beta_1 < 0$ , and (iii)  $\alpha_1 = 0$  if and only if  $\beta_1 = 0$ . The independent assumption is true for completely randomized clinical trials. However, it seldom holds in practice, especially in observational studies.

Now assume  $X_1$  is a covariate with k-categories, denoted by 1, ... k. Let  $Z_j=1\{X_1=j\}$ . We can also prove that there exist constant  $\alpha_0$ ,  $\alpha_1$ , ...,  $\alpha_{(k-1)}$  such that

$$\Pr\{Y = 1 | X_1\} = \frac{\exp(\alpha_0 + \sum_{j=1}^{k-1} \alpha_j Z_j)}{1 + \exp(\alpha_0 + \sum_{j=1}^{k-1} \alpha_j Z_j)}.$$

All those parameters have similar interpretation as in the binary case.

This section shows that for categorical covariate, the univariate regression still has the form of logistic regression. However, the interpretation of the parameter is different from that in multiple logistic regression.

### 3.2 Univariate regression with continuous covariate

Assume  $X_1$  is a continuous covariate, for example, the age of the patient. We want to know if  $Pr\{Y=1|X_1\}$  can be written in the (3) if (1) is the true multiple logistic regression model. Before answering this question, let's us take a look a the following example.

Example 1. Suppose there are only two covariates in the multiple logistic regression model (1), where  $X_1$  is a continuous covariate with range R,  $X_2$  is 0-1 valued random variable with  $Pr\{X_2=1\}=1/2$ , and  $X_1$  and  $X_2$  are independent. We further assume  $\beta_0=\beta_1=\beta_2=1$  in (1). Then

$$\Pr\{Y = 1 | X_1\} = \frac{1}{2} \left\{ \frac{\exp(2 + X_1)}{1 + \exp(2 + X_1)} + \frac{\exp(1 + X_1)}{1 + \exp(1 + X_1)} \right\}.$$

If (3) is true, then we should have

$$\frac{1 + \exp(\alpha_0 + \alpha_1 X_1)}{1 + \exp(2 + X_1)} + \frac{1 + \exp(\alpha_0 + \alpha_1 X_1)}{1 + \exp(1 + X_1)} = 2. \tag{4}$$

Let  $X_1 \rightarrow \infty$  in (4) we have  $\alpha_1 = 1$ . Let  $X_1 = 0$  in (4) we have

$$\alpha_0 = 1 + \log \frac{1 + e + 2e^2}{2 + e + e^2}$$

However, if  $X_1=1$  in (4), then

$$\alpha_0 = 1 + \log \frac{1 + e + 2e^3}{2 + e^2 + e^3}.$$

Since these two solutions of  $\alpha_{\scriptscriptstyle 0}$  do not match, model (3) does not hold.

This example shows that, for continuous covariate  $X_1$ , the regression of Y on  $X_1$  does not in general satisfy the univariate logistic regression model even if  $X_1$  is an essential component in the multiple logistic regression. Hence, the univariate logistic regression model should not be used to estimate the marginal relation between the outcome and a continuous covariate.

# 4. Inconsistency between univariate and multiple logistic regressions

In Section 3 we show that in multiple logistic regression, the univariate regression of the outcome on each individual covariate may not satisfy the logistic regression any more. This fact has serious implications for model selection and interpretation of results in data analysis. In this section, we demonstrate this important issue using simulation studies.

# 4.1 Significant effect in multiple but not in univariate logistic regression

In this section we show an example where a continuous covariate is a necessary part in the multiple logistic, but the univariate regression indicates that the covariate has no effects in the univariate regression. The following preliminary result will be used in our discussion.

**Lemma 1.** Suppose c is a positive constant and the random variable X has standard normal distribution. Then  $E[X/(1+cexp(\vartheta X))]=0$  if and only if  $\vartheta=0$ .

The proof of this result is available from authors upon request.

**Example 2.** Let  $X_2$  and  $X_3$  be independent random variables with standard normal distributions, and  $X_1=X_2+2X_3$ . Consider the following multiple logistic regression model

$$\Pr(Y = 1 | X_1, X_2) = \frac{\exp(\alpha_0 + \alpha_1 X_1 + \alpha_2 X_2)}{1 + \exp(\alpha_0 + \alpha_1 X_1 + \alpha_2 X_2)'} \tag{5}$$

where  $\alpha_1$ =- $\alpha_2/5$ ,  $\alpha_2 \neq 0$ . Using the result in Lemma 1 we can prove that if

$$\Pr\{Y = 1 | X_1\} = \frac{\exp(\theta_0 + \theta_1 X_1)}{1 + \exp(\theta_0 + \theta_1 X_1)},\tag{6}$$

then  $\theta_1=0$ .

What does this result mean within the current context? Although  $X_1$  and  $X_2$  both are in the multiple logistic regression, if their coefficients satisfy the condition (5), the regression of Y on  $X_1$  is no longer a univariate logistic regression. Moreover, if  $(Y_{i1}, X_{i2}, X_{i2})$ , i=1,...,n is a random sample from (5),  $X_1$  and  $X_2$  will become increasingly significant in the multiple logistic regression, but  $X_1$  will remain nonsignificant regardless of sample size, as illustrated by the following simulation results.

The data was generated according (5) with  $\alpha_0$ =1, $\alpha_1$ =-3/5, $\alpha_2$ =3. Shown in Table 1 are the estimates and standard deviations of the coefficient of  $X_1$  in both univariate and multiple logistic regression after 10,000 Monte Carlo (MC) replicates. The parameters were estimated by MLE. For a wide range of sample sizes, the maximum likelihood estimator of the coefficient of  $X_1$  in the multiple logistic regression was very close to the true value, and the standard errors decreased with the sample size, as expected. However, the estimated coefficient in the univariate analysis was consistently close to 0 in all cases.

Table 1 also reports the chance that p-value is >0.2 (or >0.1) in the univariate logistic regression. It shows that although  $X_1$  is a necessary part of the multiple logistic regression,  $X_1$  will most likely be excluded from the multiple logistic regression, if the cutoff of the p-value is set at 0.2 (or 0.1).

Reported in Table 2 are the estimates of the coefficient of  $X_2$  in the logistic regression if  $X_1$  is

Table 1. Estimate of regression coefficient of X <sub>1</sub> in Example 2									
		Univaria	Multiple regression						
n	Estimate	SD	p-value>0.2	p-value>0.1	Estimate	SD			
100	-0.0042	0.0983	0.7963	0.8991	-0.6533	0.2103			
200	-0.0015	0.0674	0.7952	0.898	-0.6173	0.1308			
500	-0.0004	0.0429	0.791	0.889	-0.6085	0.0828			
1,000	-0.0009	0.0284	0.801	0.907	-0.6056	0.0566			
1,500	-0.0002	0.0239	0.799	0.902	-0.6046	0.0465			
2,000	-0.0004	0.0205	0.809	0.905	-0.6027	0.0392			

Table 2. Estimates of coefficients of X<sub>2</sub> in logistic regression with X<sub>1</sub> being removed in Example 2

	Coefficient of X <sub>2</sub> (α <sub>2</sub> =3)				
n	Estimate	SD			
100	2.0243	0.4268			
200	1.9843	0.2903			
500	1.9579	0.1789			
1,000	1.9556	0.1231			
1,500	1.9498	0.1040			
2,000	1.9495	0.0857			

mistakenly excluded due to UAS method. The true coefficient of  $X_3$  is 3 in the multiple logistic regression, but the estimated coefficient of  $X_2$  became extremely biased if  $X_1$  was excluded.

Taken together, the results show that the UAS not only most likely misses some important covariates in the multiple logistic regression, but also leads to severely biased estimates of effects of other covariates on the response.

# 4.2 Significant effect in univariate but not in multiple regression

In this section we show a case where a continuous covariate has significant effect in the univariate regression, but is not significant if it is included in the multiple regression.

Example 3. Suppose  $X_1$ ,  $X_2$ ,  $X_4$  and  $\epsilon$  are independent standard normal random variables, and  $X_3=X_1+X_4$ . Consider the following multiple logistic regression model

$$\Pr(Y = 1 | X_1, X_2) = \frac{\exp(\alpha_0 + \alpha_1 X_1 + \alpha_2 X_2)}{1 + \exp(\alpha_0 + \alpha_1 X_1 + \alpha_2 X_2)}, \tag{7}$$

where  $\alpha_1 \alpha_2 \neq 0$ .

In the simulation study, the data was generated according model (7) with  $\alpha_0$ =0, $\alpha_1$ =2, $\alpha_2$ =1. Shown in Table 3 are the estimates of the coefficient of  $X_3$  in both univariate and multiple linear regression (with  $X_1, X_2$  and  $X_3$  as covariates) after 10000 replicates. For all sample sizes,  $X_3$  shows very significant effect on Y

Table 3. Estimate of the regression coefficient of $\mathbf{X}_{\scriptscriptstyle{3}}$								
	Univariate r	egression	Multiple regression					
n	Estimate	SD	Estimate	SD				
100	0.7120	0.2012	0.0130	0.3079				
200	0.6907	0.1320	-0.0021	0.1953				
500	0.6787	0.0800	-0.0039	0.1221				
1,000	0.6777	0.0588	0.0005	0.0865				
1,500	0.6772	0.0463	-0.0012	0.0681				
2,000	0.6771	0.0400	0.0000	0.0602				

in the univariate regression, but no significant effect in the multiple logistic regression.

### 5. Discussion

Although the logistic regression is a very powerful analytical method for binary outcome, the results from the univariate and multiple logistic regressions tend to be conflicting. A covariate may show very significant effect in the univariate analysis but has no role in the multiple logistic regression model. On the other hand, a covariate may be an essential part of the multiple logistic regression but shows no significant effect on the outcome in the univariate regression. The UAS method uses the univariate analysis as an initial step to select covariates for further consideration in the multiple regression. This method may mistakenly exclude important covariates in the multiple logistic

regression and lead to extremely biased estimates of the effects of other covariates in the multiple model. Hence the UAS is not a valid method in model selection. It should be removed from the tool kits of biomedical researchers and even some PhD statisticians. Formal model selection methods based on solid theory, such as Akaike's information criterion (AIC) and Schwarz' Bayesian information criterion (BIC) discussed in [10] should be implemented in all regression analyses.

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# 单因素与多因素逻辑回归的不一致性

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概述:逻辑回归是研究协变量对二元结果影响的一种常用的统计方法。它已被广泛应用于临床试验和观察性研究。然而,单因素回归得到的结果和多元逻辑回归得到的结果往往是相互矛盾的。在多元回归中可能对结果会显示出非常强烈的影响的一个协变量在单因素回归中可能不会,反之亦然。这些事

实在生物医学研究中并没有引起足够的重视。误用逻辑回归在医学出版物中非常普遍。在本文中,我们研究了单因素和多因素逻辑回归分析的不一致性,并在多元逻辑回归分析的模型部分中给出建议。

关键词:条件期望;模型选择;逻辑回归

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