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Trends in end-of-life cancer care in the Medicare program

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Abstract

Objectives—To examine contemporary trends in end-of-life cancer care and geographic variation of end-of-life care aggressiveness among Medicare beneficiaries.

Materials and Methods—Using the Surveillance, Epidemiology, and End Results—Medicare data, we identified 132,051 beneficiaries who died as a result of cancer in 2006–2011. Aggressiveness of end-of-life care was measured by chemotherapy received within 14 days of death, >1 emergency department (ED) visit within 30 days of death, >1 hospitalization within 30 days of death, 1 intensive care unit (ICU) admission within 30 days of death, in-hospital death, or hospice enrollment 3 days before death. Using hierarchical generalized linear models, we

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None of the other coauthors have conflicts to report.

Study concept and design: Wang, Pollack, Bradley, and Gross.

Acquisition of data: Gross.

Author Contributions

Dr. Hall had full access to all the data in the study and takes responsibility for the integrity of the data and the accuracy of the data analysis.

Analysis and interpretation of data: Wang, Hall, Pollack, Adelson, Bradley, Long, and Gross.

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assessed potentially aggressive end-of-life care adjusting for patient demographics, tumor characteristics, and hospital referral region (HRR)-level market factors.

Results—The proportion of beneficiaries receiving at least one potentially aggressive end-of-life intervention increased from 48.6% in 2006 to 50.5% in 2011 (P < .001). From 2006 to 2011, increases were apparent in repeated hospitalization (14.1% vs. 14.8%; P = .01), repeated ED visits (34.3% vs. 36.6%; P < .001), ICU admissions (16.2% vs. 21.3%; P < .001), and late hospice enrollment (11.2% vs. 12.9%; P < .001), whereas in-hospital death declined (23.5% vs. 20.9%; P < .001). End-of-life chemotherapy use (4.4% vs. 4.5%) did not change significantly over time (P = .12). The use of potentially aggressive end-of-life care varied substantially across HRRs, ranging from 40.3% to 58.3%. Few HRRs had a decrease in aggressive end-of-life care during the study period.

Conclusions—Despite growing focus on providing appropriate end-of-life care, there has not been an improvement in aggressive end-of-life cancer care in the Medicare program.

Keywords

End-of-life care; Intensity; Geographic variation

1. Introduction

Overly aggressive care at the end of life is not consistent with patient preferences,^{1–3} incurs substantial costs, and is not associated with better outcomes.^{4–6} Since 1999, the Institute of Medicine (IOM) has released several reports calling for improvement of end-of-life cancer care.^{7,8} Efforts have identified concerning end-of-life care patterns, such as very late chemotherapy use, very short hospice enrollment, and repeated hospitalization during patients' last month of life.⁹ Such aggressive care patterns have been used by oncologists to indicate poor end-of-life care quality.^{10,11} While palliative care has been embraced by the medical community, the recent IOM report *Dying in America* highlights continued deficiencies in promoting palliative care.¹² To improve end-of-life care, several organizations, including the American Society of Clinical Oncology, have been working to improve care delivery, clinician– patient communication, and advance care planning.^{13,14}

Available literature examining end-of-life cancer care among older individuals with cancer in the United States is outdated,^{15,16} limited in scope,¹⁷ or lacks important clinical detail.^{18–20} For instance, one study of Medicare beneficiaries with ovarian cancer found that intensity of hospital-based end-of-life care increased between 1997 and 2007.¹⁷ The other analyses of Medicare beneficiaries lacked clinical detail regarding cancer characteristics, either focusing on patients with cancer who had been hospitalized during the last 6 months of life¹⁸ or comparing general end-of-life care patterns across decedents who died from cancer or other causes.²⁰ These analyses highlight the need for more comprehensive, updated information regarding end-of-life cancer care in the United States in order to assess progress after over a decade of efforts to improve care, and identify opportunities for improvement.¹⁵ To address this knowledge gap, we examined the trends in the aggressiveness of end-of-life cancer care over time in a population-based cohort of Medicare beneficiaries who had died after a cancer diagnosis. We also evaluated the trends of geographic variation of end-of-life care aggressiveness and identified the geographic regions that experienced a greater improvement in end-of-life cancer care than others. We assessed the associations of end-of-life care aggressiveness with patient characteristics and the availability of related healthcare resources. Findings from this study can not only provide a more comprehensive picture of temporal trends in the quality of end-of-life cancer care in the United States but also further our knowledge of whether certain regions or regional market factors might be more conducive to improving end-of-life care.

2. Methods

2.1. Data and Study Design

We used the Surveillance, Epidemiology, and End Results (SEER)—Medicare database, a unique data source linking Medicare enrollment and claims records to tumor registries. The SEER registries currently cover approximately 28% of the U.S. population.²¹ We used SEER data to identify baseline patient and tumor characteristics and Medicare claims to identify indicators of interaction with the healthcare system. The study was reviewed by the Institutional Review Board of Yale University who determined that this study did not directly involve human subjects.

2.2. Patients

We identified beneficiaries who had breast, prostate, lung, colorectal, pancreas, liver, kidney, melanoma, or hematologic cancer diagnosed in 2004–2011. To make the sample of each year comparable, we limited our cohort to decedents each year who died within 3 years of diagnosis as a result of cancer. This criterion of the same range of time between cancer diagnosis and death each year, consistent with prior research,⁹ allowed us to avoid the potential influence of time between cancer diagnosis and death on trend results. Consequently, only the annual results from 2006 to 2011 were compared. We limited our sample to beneficiaries who were aged 66.5–94.9 years at death and enrolled in Medicare Parts A and B during the last 18 months of life. Patients were excluded if their diagnosis occurred only according to death certificate or autopsy, if they could not be assigned to a hospital referral region (HRR), or if they lived less than 3 months after cancer diagnosis. The step-wise ascertainment of our study cohort is listed in Appendix Table A1 (online only).

2.3. Measurement

2.3.1. Outcomes—We used previously developed claims-based indicators of potentially aggressive health care within the last 30 days of life, ¹⁰ including (1) chemotherapy received within 14 days of death, (2) >1 emergency department (ED) visit within 30 days of death, (3) >1 hospitalization within 30 days of death, (4) 1 intensive care unit (ICU) admission within 30 days of death, (5) in-hospital death, and (6) hospice enrollment 3 days before death. We created a composite measure of aggressive end-of-life care, which was defined as the occurrence of at least one of the indicators above.

2.3.2. Covariates—We included candidate variables which are available in our database and have been used in research examining end-of-life care and/or healthcare market factors. Patient demographics included age, race, Hispanic ethnicity, gender, year of death, marital status, SEER registry, and metro status of residence.²² Socioeconomic status measures included median household income and percentage of adults with high school education or less, both derived from census data. We evaluated Elixhauser comorbidity conditions between 7 months and 18 months prior to death, adapting an approach that requires the diagnosis code to appear on an inpatient claim or two or more physician or outpatient claims greater than 30 days apart for the condition to be considered present.²³ We incorporated a measure of disability status, a claim-based indicator for services commonly needed by patients with poor functional performance status.²⁴ We also ascertained the number of outpatient clinic visits within 1 to 3 months before death. Tumor characteristics included tumor site, advanced stage, multiple-cancer diagnoses, and time between cancer diagnosis and death as reported by SEER.

We attributed each patient's medical care to the hospital in which each patient was admitted most frequently in the last 6 months of life. If the numbers tied between hospitals, total days of hospital stay were used. We categorized hospitals according to National Cancer Institute designation. We used data from the Area Resource File to assess the following county-level market factors assigned into quartiles each year: percentage of individuals in health maintenance organizations (HMOs), and the number of physicians, radiation oncologists, hospital beds, skilled nursing facility beds, hospices, and home health agencies per 1000 people aged 65 years and older.

2.4. Statistical Analysis

We described demographic, clinical, and social factors between patients who experienced at least one indicator of aggressive care versus those who did not experience any aggressive intervention between 2006 and 2011, using *t*-tests for continuous variables and chi-square statistics for categorical variables. The Cochran–Armitage test was used to evaluate time trends. For each HRR, we also identified the proportion receiving aggressive care each year. We then determined the trend of aggressive end-of-life care over time for each HRR and identified the HRRs with a decrease in aggressive care. We also identified HRRs which had a lower percentage of aggressive care through the study period.

In the multivariable analyses, we used hierarchical generalized linear models (HGLMs), clustering patients by HRR, to determine the independent contribution of individual and market factors to aggressiveness of end-of-life care. To reduce variability caused by low HRR volumes, we excluded from the HGLM analysis HRRs that had less than twenty decedents during the study period. We used the variance inflation factors of independent variables using multivariate linear regression models to assess potential multicollinearity.

We estimated the adjusted proportion of decedents who experienced at least one aggressive end-of-life care intervention for each HRR, controlling for patient characteristics but not market factors. We calculated the expected proportion of aggressive end-of-life care for HRR*i* each year, using HGLMs with random effects as zero. Then, for HRR*i*, we estimated its adjusted proportion as:

$$\frac{\overline{O}_{i,t}}{\overline{E}_{i,t}} \times \overline{O}_{\text{overall},t},$$

where $_{i,t}$ and $_{overall,t}$ are the observed proportion of aggressive end-of-life care in HRR*i* and in all HRRs in year_b, respectively, and $_{i,t}$ is the expected proportion of aggressive endof-life care in HRR*i* in year_b from 2006 to 2011. This measure reflects the proportion of decedents receiving aggressive end-of-life care across HRRs after adjusting for patient factors. We created HRR quintiles according to the unadjusted proportion of decedents who received at least one aggressive end-of-life care intervention in 2006, and then followed these quintiles over time in order to investigate the trends in adjusted proportion of aggressive end-of-life care in the highest and lowest quintiles over the study period. All statistical analyses were completed using SAS, version 9.4 (SAS Institute, Cary, NC) or Stata 13 (StataCorp, College Station, TX). A two-tailed P < .05 was used to define statistical significance.

3. Results

The full study sample consisted of 132,051 beneficiaries. Overall, 44.9% of the cohort was age >80 years, about 70% had comorbidities, and 35% had stage IV disease at diagnosis (Table 1). Nearly 43% had lung cancer, followed by colorectal cancer, hematologic malignancies, and pancreas cancer. Approximately one third died within 3–6 months after diagnosis, and 20% within 6 months to 1 year.

The proportion of beneficiaries receiving at least one potentially aggressive end-of-life intervention increased from 48.6% to 50.5% over time (trend test P < .001; Fig. 1). There was a slight but significant increase repeated hospitalizations (from 14.1% to 14.8%; P = . 014), repeated ED visits (from 34.3% to 36.6%; P < .001), and very short hospice enrollment (from 11.2% to 12.9%; P < .001), while increase in ICU admission was more substantial (from 16.2% to 21.3%; P < .001). In contrast, in-hospital death declined in the study period (from 23.5% to 20.9%; P < .001). End-of-life chemotherapy use (approximately 4.5% from 2006 to 2011) did not change significantly over time (trend test P = .12).

In multivariable analyses, we found decedents who were younger, male, non-white, Hispanic, married, and resided in non-metropolitan areas or areas with higher proportions with more than high school education were more likely to have aggressive care, compared with those who were older, female, white, non-Hispanic, unmarried, and resided in metropolitan areas or areas with areas with lower proportions with more than high school education (Table 2; and Appendix Table A2 for detailed results). Compared with patients without comorbidity, having three or more comorbidities was associated with increased odds of receiving aggressive care (odds ratio [OR], 1.08; 95% confidence interval [CI], 1.04 to 1.12). In contrast, having disability was associated with decreased odds of receiving aggressive care (OR, 0.92; 95% CI, 0.89 to 0.96). Cancer type, cancer stage at diagnosis, and time between cancer diagnosis and death also played a significant role in predicting aggressive care. For example, compared with patients with breast cancer, patients with

hematologic malignancies were more likely to receive aggressive end-of-life care (OR, 1.21; 95% CI, 1.14 to 1.29), whereas patients with pancreatic cancer were less likely to receive aggressive end-of-life care (OR, 0.61; 95% CI, 0.57 to 0.66).

Only two market factors that we examined were associated with aggressive end-of-life intervention in the multivariable analysis. Decedents who were in the areas with higher HMO penetration rate were less likely to receive aggressive end-of-life care (OR, 0.90; 95% CI, 0.84 to 0.97), compared with decedents who were in the areas with lower HMO penetration rate. Decedents who were in the areas with a higher number of radiation oncologists per 1000 people aged 65 years and older were less likely to receive aggressive end-of-life care (OR, 0.93; 95% CI, 0.87 to 0.99), compared with decedents in the lowest quintile area. The supply of hospice, physicians, hospitals, home health agency, or skill nursing facilities within HRRs was not associated with receipt of aggressive care (Appendix Table A2).

The use of aggressive end-of-life care varied substantially across HRRs, ranging from 40.3% to 58.3% (unadjusted). In the analysis of temporal trends, few HRRs had a decrease in aggressive end-of-life care during the study period (Fig. 2). Although 5 HRRs had an absolute decrease in end-of-life care of >2%, only one HRR (Ogden, Utah) had a statistically significantly decreased trend of 6.8% (P=.023). Seven HRRs (including Santa Cruz, CA; Cedar Rapids, Des Moines, Iowa City, and Mason City, IA; Fort Oglethorpe, GA; and Tacoma, WA) had low levels of end-of-life care intensity across the stud period. Additionally, the magnitude of regional variation in end-of-life care did not change substantially: the difference in the proportion of patients receiving potentially aggressive end-of-life care between HRRs in the lowest and the highest quintiles was 20.5% (37.4% vs. 58.0%) in 2006 and 19.8% (40.2% vs. 60.1%) in 2011. After adjusting for patient characteristics, the difference between the highest and lowest quintile areas was 15.8% (44.4% vs 60.2%) in 2011 (Fig. 3).

4. Discussion

Our analysis found that nearly 50% of Medicare beneficiaries who died as a result of cancer between 2006 and 2011 experienced some form of potentially aggressive end-of-life care. Not only has there been no improvement in the overall use of aggressive end-of-life care, there was actually a slight increase in aggressive care over time. Despite a growing focus on providing appropriate end-of-life care, our finding that aggressive end-of-life cancer care is increasing over time in the Medicare program underscores the difficulties in improving end-of-life cancer care.

Our findings build upon prior work in important ways. First, consistent with prior literature,^{18,25} there was substantial regional variation of end-of-life care. Unlike prior work, we further described the trends in variation, to identify regions that had marked improvement. Of the 92 regions in our sample, only a single region (Ogden, Utah) had a statistically significant reduction in use of aggressive end-of-life care (manifested by a decrease of 6.8% from 2006 to 2011). It has been reported that patients in Ogden, Utah, spent an average of 39.5 days per patient in hospice in the last 6 months of life in 2007, the

longest time in hospice care in the United States.²⁶ It is unclear whether the improvement over time in Ogden is associated with trends in early referral for hospice care; future research to identify factors associated with the decrease in these HRRs may help to highlight mechanisms and interventions to improve end-of-life care in other regions.

Second, we found several patient characteristics were associated with aggressive end-of-life care. Consistent with prior literature,^{9,16} older patients were less likely to receive aggressive care. As treatment decisions for older terminally ill patients with cancer are complex, it is important for patients and providers to discuss benefits and harms attributable to treatments. Recent research suggested that the effectiveness of chemotherapy in reducing cancerspecific mortality may decrease with age.²⁷ Another study, using caregiver's rating of patient quality of life near death as a proxy, indicated that palliative chemotherapy could not improve patients' quality of life near death.²⁸ These two studies highlight the importance of patient-physician communication about end-of-life treatment decisions among the older population.

Third, patients who were male, non-white, or Hispanic were more likely to receive aggressive care.¹⁵ Additionally, decedents who were unmarried were less likely to receive aggressive end-of-life care, compared with those who were married, indicating the necessity of studying the importance of cancer caregivers when it comes to patients' end-of-life care. For instance, prior research has suggested that social support can have great impact of cancer detection, treatment, and survival.²⁹ Interestingly, patients with comorbidities were more likely to receive aggressive care, whereas patients with disability were less likely to receive aggressive care. It is possible that comorbidities were associated with increased healthcare utilization, receiving medical intervention for comorbidities or for the increased side effects attributed to cancer care; thus, comorbidities led to an increase in aggressive end-of-life care. In contrast, patients with poor performance status (as the disability measure intended to capture) might be more likely to receive palliative care, which would result in a decrease in aggressive end-of-life care. Future research examining which factors lead patients with poor performance status to receive less aggressive care is needed. Nevertheless, the results suggest that comorbidities and disability measures capture different components of patient health, and should be included in related models.

Fourth, tumor characteristics were also associated with care aggressiveness. Compared with patients with breast cancer, patients with cancer of lung, colorectal, pancreas, liver, kidney, or skin were less likely to receive aggressive care. In contrast, patients with hematologic malignancy were more likely to receive aggressive care, which is consistent with prior literature.¹⁶ Also, patients with advanced cancer at diagnosis or with a shorter time between cancer diagnosis and death were associated with less aggressive care, potentially reflecting disease progression and patient perception about his/her disease. For example, patients who respond well to an initial treatment may be encouraged to receive aggressive care for a recurrence. Our study design including patients who died within 3 years after cancer diagnosis allowed us to analyze the relation between time since diagnosis and care aggressiveness. In contrast, prior literature generally examined end-of-life care patterns for decedents who died within a year after cancer diagnosis⁹; thus, proportions of receiving aggressive care among all cancer decedents may be even higher than the reported estimates.

Fifth, prior literature has indicated that higher HMO penetration was related with less aggressive EOL care in the fee-for-service markets, so-called spillover effects. It is well known that, compared with fee-for-service providers, managed care providers performed better end-of-life care, using measures of higher rates of hospice use and longer hospice stays.³⁰ We found that patients who resided in the areas with higher HMO penetration rates were less likely to receive aggressive care, suggesting that some spillover effects may exist. Literature regarding managed care's spillover effects on cancer screening is inconclusive,^{31,32} and launching a new wave of studies examining spillover effects on cancer care quality has been suggested.³³ Using a rigorous approach and validated quality measures, we demonstrated an association between HMO penetration and performance of end-of-life care. In contrast, other market factors such as numbers of hospital, hospice, or physician at county level were not associated with aggressive care. Our results indicate that healthcare supply factors may not play an important role regarding end-of-life cancer care quality among Medicare fee-for-service beneficiaries died due to cancer. Future research is needed to examine this relationship in other settings and patient populations.

Concurrent with an increase in most of the indicators of aggressive care we examined, we did find that in-hospital deaths decreased over time. The finding regarding in-hospital deaths was not surprising in that the proportion of decedents receiving hospice care continued increasing³⁴ and hospice use could decrease the likelihood of in-hospital death.⁹ However, an increase in hospice service and/or a decrease in in-hospital death did not necessarily imply a decrease in care aggressiveness. Researchers have raised concerns that hospice may be an "add-on" to a growing pattern of more utilization of intensive end-of-life care.^{17,20} Thus, efforts to improve end-of-life care may require consideration of both quantity and quality of hospice use.^{35,36}

We acknowledge several limitations. Our population is limited to elderly patients in fee-forservice Medicare with nine types of cancer. The results may not be generalizable to younger patients or those in HMO plans. Second, we lack information regarding patient preferences, although prior literature has demonstrated that patient preferences explain little of regional variation in end-of-life healthcare utilization.^{3,37} We also lack information regarding palliative consultation, which could decrease aggressive end-of-life care. Third, while our claim-based indicators of aggressive end-of-life care have been validated, the composite measure may not capture all aggressive care. For instance, late hospice enrollment (within 7 days of death) has been proposed as a measure of poor quality of end-of-life care.^{36,38} Additionally, our composite measure may miss decedents who have never enrolled in hospice. Thus, our observations may actually underestimate the rates of aggressive end-oflife care. Finally, our cohort was created backward from decedents. Such a retrospective design may lead to biased conclusions for the quality of care provided to dying patients because care received by decedents is not equivalent to care received by individuals who are dying.³⁹ Prospective designs, however, may also be limited by the difficulty of accurately assessing prognosis. Using both retrospective and prospective approaches, Setoguchi et al. found similar physician and hospital patterns of end-of-life care regarding the use of opiate analgesia.⁴⁰ Future prospective studies confirming our findings are needed.

In conclusion, approximately 50% of cancer decedents in the fee-for-service Medicare Program received aggressive end-of-life care. Despite emphasis on improving end-of-life care in the United States, end-of-life care for patients with cancer appeared more aggressive over time. Regional variation of end-of-life cancer care was substantial, with no evidence of decreased variation over time. There may be important opportunities to learn from areas with a recent decrease in aggressive care. Identifying the underlying mechanisms that led to improved end-of-life care in these areas could influence the testing and adoption of new models in other regions.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

Acknowledgments

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Appendix A. Supplementary data

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ED: Emergency department; ICU: Intensive care unit

Fig. 1.

Trends of aggressive end-of-life care: (A) proportion of patients with cancer receiving any aggressive end-of-life care; (B) proportion of patients with cancer receiving hospital-based service in last month of life; (C) proportion end-of-life medical care of patients with cancer. ED: emergency department; ICU: intensive care unit.



Fig. 2.

Trends of aggressive end-of-life care in 92 hospital referral regions among SEER-Medicare areas. Each line represents a hospital referral region (HRR), showing the regression for trend in aggressive end-of-life care. Five HRRs (highlighted in orange) had a trend of >2% decrease. SEER: Surveillance, Epidemiology, and End Results.





Fig. 3.

Proportions of decedents receiving aggressive end-of-life care between the highest and lowest quintile areas by year. The estimates have been adjusted for patient demographics and tumor characteristics, including age, race, sex, marital status, socioeconomic status, tumor type, tumor stage, and time since cancer diagnosis. Table 1

Demographic and clinical characteristics of medicare beneficiaries who died of cancer, 2006–2011.

	Patients not experiencing any a; (n = 66, 455)	iggressive care	Patients experiencing any aggres $(n = 65, 596)$	sive care P	-value
	u	%	u	%	
Age-group				V	:001
66.5–69	8,200	12.4	11,560	17.6	
70–74	13,004	19.6	18,126	27.6	
75–79	14,791	22.3	18,624	28.4	
80–84	14,900	22.4	16,076	24.5	
85–94	15,540	23.4	12,819	19.5	
Female				V	:001
No	33,317	50.1	29,640	45.2	
Yes	33,138	49.9	35,956	54.8	
Race				V	:001
White	58,819	88.5	55,848	85.1	
Black	4,855	7.3	6,318	9.6	
Other	2,781	4.2	3,430	5.2	
Hispanic				V	:001
Non-Hispanic	63,692	95.8	62,084	94.6	
Hispanic	2,763	4.2	3,512	5.4	
Marital Status				V	:001
Married	31,698	47.7	34,196	52.1	
Unmarried	31,660	47.6	28,303	43.1	
Unknown	3,097	4.7	3,097	4.7	
Metro status of residence				V	:001
Metro	54,083	81.4	54,251	81.6	
Non-metro	12,372	18.6	11,345	17.1	
Income				V	:001
< \$33,000	15,302	23.0	15,953	24.3	
\$33,000-\$39,999	10,935	16.5	10,023	15.3	
\$40,000-\$49,999	13,662	20.6	13,203	20.1	

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<i>P</i> -value			
e care		%	18.9

	Patients not experiencing any aggressive c $(n = 66,455)$	care	Patients experiencing any aggressiv $(n = 65,596)$	e care	<i>P</i> -value
	п	%	u	%	
\$50,000-\$62,999	12,688	19.1	12,413	18.9	
\$63,000	13,776	20.7	13,935	21.2	
Unknown	92	0.1	69	0.1	
High School Education					<.001
<30%	14,399	21.7	13,115	20.0	
30%–39%	10,558	15.9	9,830	15.0	
40%49%	11,541	17.4	11,225	17.1	
50%-59%	12,299	18.5	12,312	18.8	
60%	17,566	26.4	19,045	29.0	
Unknown	92	0.1	69	0.1	
Comorbidity					<.001
None	19,964	30.0	16,868	25.7	
1 to 2	26,175	39.4	26,349	40.2	
3 or more	20,316	30.6	22,379	34.1	
Disability status					<.001
No	56,790 8	85.5	56,736	86.5	
Yes	9,665	14.5	8,860	13.5	
Tumor site					<.001
Lung	28,780	43.3	27,844	42.4	
Colorectal	9,948	15.0	8,074	12.3	
Hematologic malignancies	6,560	9.6	10,180	15.5	
Pancreas	6,627	10.0	5,018	7.6	
Breast	4,227	6.4	4,217	6.4	
Prostate	3,719	5.6	4,163	6.3	
Skin	2,287	3.4	2,172	3.3	
Kidney	2,189	3.3	2,049	3.1	
Liver	2,118	3.2	1,879	2.9	
Stage IV at diagnosis					<.001
Not stage IV	42,135	63.4	43,126	65.7	
Stage IV	24,320	36.6	22,470	34.3	

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	Patients not experiencing any aggressive c $(n = 66, 455)$	care	Patients experiencing any aggressive $(n = 65, 596)$	are P-1	value
	и	%	u	%	
Multiple cancers				<.0	01
No	60,572	91.1	58,167	38.7	
Yes	5,883	8.9	7,429	11.3	
Year of death				<.0	001
2006	11,378	17.1	10,738	16.4	
2007	11,673	17.6	11,362	17.3	
2008	11,306	17.0	11,172	17.0	
2009	10,939	16.5	10,827	16.5	
2010	10,829	16.3	10,958	16.7	
2011	10,330	15.5	10,539	16.1	
Time between cancer diagnosis and death				<.0	01
3–6 months	22,344 3	33.6	23,187	35.3	
6 months-1 year	13,664	20.6	13,282	20.2	
1–2 years	19,588	29.5	18,533	28.3	
2–3 years	10,859	16.3	10,594	16.2	
Outpatient visits within 1 to 3 months before death				<.0	01
None	11,545	17.4	2,398	3.7	
1 to 4	15,355	23.1	8,680	13.2	

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Table 2

Multi-level logistic regression of factors associated with aggressive care.

	Odds ratio	95%	6 CI
		Lower	Upper
Demographic factors			
Age-group			
66.5–69	REF		
70–74	0.96	0.92	1.01
75–79	0.86	0.82	0.89
80-84	0.76	0.73	0.80
85–94	0.65	0.62	0.68
Sex			
Male	REF		
Female	0.82	0.80	0.85
Race			
White	REF		
Black	1.29	1.23	1.35
Other	1.34	1.26	1.43
Hispanic			
Non-Hispanic	REF		
Hispanic	1.20	1.13	1.28
Marital status			
Married	REF		
Unmarried	0.94	0.91	0.97
Metro status of residence			
Metro	REF		
Non-metro	1.09	1.03	1.15
Income			
<\$33,000	REF		
\$33,000-\$39,999	0.95	0.91	0.99
\$40,000-\$49,999	0.97	0.93	1.01
\$50,000-\$62,999	0.96	0.91	1.01
\$63,000	0.94	0.89	1.00
High school education			
<30%	REF		
30%-39%	1.05	1.00	1.10
40%-49%	1.06	1.01	1.12
50%-59%	1.12	1.06	1.18
60%	1.16	1.09	1.23
Clinical factors			
Comorbidity			
None	REF		

	Odds ratio	95%	6 CI
		Lower	Upper
1 to 2	1.06	1.03	1.10
3 or more	1.08	1.04	1.12
Disability			
No	REF		
Yes	0.92	0.89	0.96
Tumor site			
Lung	0.82	0.77	0.86
Colorectal	0.73	0.69	0.78
Hematologic malignancies	1.21	1.14	1.29
Pancreas	0.61	0.57	0.66
Breast	REF		
Prostate	0.92	0.86	1.00
Skin	0.81	0.75	0.89
Kidney	0.78	0.72	0.85
Liver	0.69	0.63	0.76
Stage IV at diagnosis			
Not Stage IV	REF		
Stage IV	0.85	0.83	0.87
Multiple cancers			
No	REF		
Yes	1.18	1.13	1.23
Time between cancer diagnosis and death			
3–6 months	REF		
6 months–1 year	1.12	1.07	1.16
1–2 years	1.13	1.09	1.17
2–3 years	1.14	1.09	1.20
Year of death			
2006	REF		
2007	1.04	1.00	1.09
2008	1.06	1.01	1.10
2009	1.06	1.01	1.10
2010	1.09	1.04	1.14
2011	1.10	1.05	1.15
Outpatient visits within 1 to 3 months before death			
None	REF		
1 to 4	2.92	2.76	3.08
5 to 10	5.40	5.12	5.70
11 to 19	6.46	6.12	6.82
20 or more	6.81	6.44	7.19
Market factors			

County-level HMO penetration rate

	Odds ratio	95%	- CI
		Lower	Upper
Q1 (lowest quartile)	REF		
Q2	0.95	0.91	1.00
Q3	0.88	0.83	0.94
Q4	0.90	0.84	0.97
County-level radiation oncologist number per 1000 people 65 years and older			
Q1 (lowest quartile)	REF		
Q2	0.94	0.89	1.00
Q3	0.96	0.90	1.02
Q4	0.93	0.87	0.99

CI: confidence interval; NCI: National Cancer Institute; HMO: health maintenance organization.

Model included all variables in the Table as well as hospital factor (NCI Designation) and market factors (hospice number, physician number, hospital bed number, skilled nursing bed number, and home health agency number per 1000 people 65 years and older at county level).

Detailed results can be found in the Appendix Table A2.