GENETIC DIVERSITY AND EVOLUTION



No Substantial Evidence for Sexual Transmission of Minority HIV Drug Resistance Mutations in Men Who Have Sex with Men

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ABSTRACT During primary HIV infection, the presence of minority drug resistance mutations (DRM) may be a consequence of sexual transmission, de novo mutations, or technical errors in identification. Baseline blood samples were collected from 24 HIV-infected antiretroviral-naive, genetically and epidemiologically linked source and recipient partners shortly after the recipient's estimated date of infection. An additional 32 longitudinal samples were available from 11 recipients. Deep sequencing of HIV reverse transcriptase (RT) was performed (Roche/454), and the sequences were screened for nucleoside and nonnucleoside RT inhibitor DRM. The likelihood of sexual transmission and persistence of DRM was assessed using Bayesian-based statistical modeling. While the majority of DRM (>20%) were consistently transmitted from source to recipient, the probability of detecting a minority DRM in the recipient was not increased when the same minority DRM was detected in the source (Bayes factor [BF] = 6.37). Longitudinal analyses revealed an exponential decay of DRM (BF = 0.05) while genetic diversity increased. Our analysis revealed no substantial evidence for sexual transmission of minority DRM (BF = 0.02). The presence of minority DRM during early infection, followed by a rapid decay, is consistent with the "mutation-selection balance" hypothesis, in which deleterious mutations are more efficiently purged later during HIV infection when the larger effective population size allows more efficient selection. Future studies using more recent sequencing technologies that are less prone to single-base errors should confirm these results by applying a similar Bayesian framework in other clinical settings.

IMPORTANCE The advent of sensitive sequencing platforms has led to an increased identification of minority drug resistance mutations (DRM), including among antiretroviral therapy-naive HIV-infected individuals. While transmission of DRM may impact future therapy options for newly infected individuals, the clinical significance of the detection of minority DRM remains controversial. In the present study, we applied deep-sequencing techniques within a Bayesian hierarchical framework to a cohort of 24 transmission pairs to investigate whether minority DRM detected shortly after transmission were the consequence of (i) sexual transmission from the source, (ii) *de novo* emergence shortly after infection followed by viral selection and evolution, or (iii) technical errors/limitations of deep-sequencing methods. We found no clear evidence to support the sexual transmission of minority resistant variants, and our results suggested that minor resistant variants may emerge *de novo* shortly after transmission, when the small effective population size limits efficient purge by natural selection.

KEYWORDS minority drug resistance mutation, deep sequencing, Bayesian hierarchical framework, transmission, human immunodeficiency virus

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The transmission of drug resistance mutations (DRM) may limit therapy options for the newly infected individuals (1) and thwart efforts to control the HIV epidemic with antiretroviral therapy (ART) (2–5). Using ultrasensitive techniques, such as deep sequencing or allele-specific PCR, several studies have reported a high prevalence of "transmitted" minority DRM among the ART-naive population (6–8). The exact origin and clinical relevance of these minority variants carrying DRM are still debated (1, 5, 8–10). We previously performed deep sequencing to detect minority DRM in blood plasma collected from 32 HIV-infected individuals during primary infection, and we found that most low-frequency DRM in our cohort were likely identified because of methodological errors or were a consequence of error-prone HIV replication in the recipient (11). Our previous study and other similar studies have been limited by lack of characterization of the source partner viral population (6–8, 10, 11), lack of characterization of the primary source compartment of transmitted viral populations (e.g., genital secretion), and lack of longitudinal characterization of the viral population in recipients after transmission.

Here, we applied deep-sequencing techniques within a Bayesian hierarchical framework to a well-characterized cohort of 24 epidemiologically and genetically linked sexual transmission partners to investigate whether minority DRM detected shortly after transmission were the consequence of (i) sexual transmission from the source, (ii) *de novo* emergence shortly after infection followed by viral selection and evolution, or (iii) technical errors or analytical bias of deep-sequencing methods.

(Part of this study has been presented at the annual Conference on Retroviruses and Opportunistic Infections [CROI], Boston, MA, USA, 2016.)

RESULTS

Population characteristics and biological samples. All study participants were men who have sex with men (MSM) infected with HIV-1 subtype B and without evidence of dual infection (12). Three identified source partners transmitted HIV to more than one recipient, leading to a total of 21 unique source partners (relationships between individuals are illustrated in Fig. 1). Among the 21 unique source partners, the median age was 32 years (interquartile range [IQR], 22 to 51), and the median CD4 T-cell count was 400 cells/mm³ (IQR, 270 to 821). The median HIV RNA levels in blood and seminal plasma were 4.7 \log_{10}/ml (IQR, 4.4 to 5.1) and 3.6 \log_{10}/ml (IQR, 2.7 to 5.5), respectively. All potential source partners were ART naive at the estimated time of transmission. Among the 24 recipients, baseline blood plasma samples were available within a median of 78 days (IQR, 18 to 130) from their estimated date of infection (EDI). The median age at baseline was 32 years (IQR, 26 to 41). The median HIV RNA level and CD4 T-cell count were 4.9 log₁₀/ml (IQR, 4.2 to 5.8) and 507 cells/mm³ (IQR, 248 to 1,382), respectively. The median elapsed time between collection of paired source and recipient blood was 9 days (IQR, 2 to 29). Only 4 of the 21 source partners (19.0%) had paired seminal samples available, and longitudinal blood samples were available for 11 of the 24 recipients (45.8%), with a median of 3 time points (IQR, 2 to 3) collected over a median of 106 days (IQR, 56 to 219). Characteristics of the study populations are summarized in Table 1.

Identification of DRM. The median coverage for the pol/RT region in the blood and seminal plasma samples from the source and in blood plasma from recipients of 2,429 reads (IQR, 1,790 to 6,034), 2,478 reads (IQR, 1,711 to 4,969), and 5,248 reads (IQR, 3,840 to 6,648), respectively. Using these data, we determined the presence of DRM at each site and calculated the relative frequency in the source and recipient partners' samples. Minority variants with a prevalence below the background error rates were excluded. We found medians of 3.0 (IQR, 2.0 to 4.0) and 3.0 (IQR, 1.0 to 4.0) distinct DRM in the source and recipient partners, respectively. The median relative frequencies (DRM_{Freq}) of the DRM in source and recipient partners were 2.3% (IQR, 0.5% to 7.1%) and 2.3% (IQR, 0.2% to 12.0%), respectively. Details of the identified DRM and their relative frequencies are summarized in Table 2.

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FIG 1 Approximate maximum-likelihood phylogenetic tree of the partial *pol* regions for the 24 transmission pairs. A tree of the entire data set was reconstructed specifying a GTR+Gamma model. Each transmission pair clusters in a distinct monophyletic subtree (bootstrap support of >0.70) and is depicted in a different color. Shapes denote the status (i.e., source/recipient for each individual). The external group (black) is the subtype B HXB2 reference sequence. Pairs 3, 25, and 27, pairs 10 and 11, and pairs 16 and 17 shared a unique source partner. The recipient partner from pair 10 was also the source partner for pair 9 and is depicted with circles. S, source; R, recipient.

TABLE 1 Population characteristics

Characteristic at baseline (24 pairs)	Source	Recipient
No. of participants (no. with longitudinal data)	21 (0)	24 (11)
Median age, yr (IQR)	32 (22–51)	32 (20–59)
No. of paired seminal plasma sample	4	NA ^a
Median time from EDI (IQR) at baseline, days	NA	78 (18–130)
Median CD4 cells/ml (IQR)	400 (270–821)	507 (248–1382)
Median HIV RNA level, log ₁₀ copies/ml (IQR)		
Blood plasma	4.7 (4.4–5.1)	4.9 (4.2–5.8)
Seminal plasma	3.6 (2.7–5.5)	NA
Median duration of follow-up, days (IQR)	NA	106 (52–219)

^aNA, not available.

Evidence against preferential transmission of viral populations harboring DRM. We sought to investigate the potential sexual transmission of minority DRM by comparing the identified DRM in confirmed source and recipient pairs. The probability of a DRM being present in the recipient at baseline when absent in the source (P_{R+1S-}) was low at 0.05, with a 95% credible interval of 0.01 to 0.13 (Fig. 2A, left panel). When

TABLE 2 Identification screening of NRTI and NNRTI minority DRM in source and recipient samples at baseline

	Source			Recipient			
					Time from		
Pair	Origin ^a	VL ^ø	DRM (%) ^c	VL	EDI (days)	DRM (%)	
1	BP	4.4	F227L (0.1), K103N (100)	5.6	11	K65N (1.3), D67N (2.4), K103N (100)	
2	BP	4.6	K65R (5.6), L74V (0.1), Y115F (0.1), K219W (1.3), F227W (0.1)	4.9	133	K219W (3.7)	
3 ^{<i>d</i>}	BP	3.8	Y188H (0.8), T215S (3.2), K219W (7.8), F227L (0.1)	3.8	133	K101Q (2.5)	
4	BP	4.4	K219W (7.5), F227L (0.1)	4.7	11	K70R (5.7), T215V (62.6), K219W (21.8), F227L (0.2)	
5	BP	4.7	T215V (31), K219W (6.9), F227L(0.1)	6.4	11	F227L (0.6)	
6	BP	5.0	D67N (10.8), T215S (1.2), K219W (3.5), F227L (0.3), P236L (0.7)	4.2	121	G190S (0.3), T215V (13.9), K219W (62.5), F227L (0.1), P236L (0.1)	
7	SP BP	5.9	D67N (2.0), E138A (95.6) NA ^e	5.9	27	K70R (6.8), K219W(15), F227L(0.1), L324I (52)	
8	BP	4.2	D67N (4.2), T215E (3.1), K219W (12), F227L (0.5)	3.5	133	K65N (3.2), D67N (2.1), K70R (8.6), L74I (0.1), V75L (0.2), K103R(11.4), Y188H (0.5), T215V (6.1), K219W (40.1), F227L (0.2)	
9	BP	3.5	K103N (100), K219W (36), T215V (6.9)	7.2	17	K103N (100), T215V (8), F227L (1.4)	
10 ^d	BP	5.8	K103N (100), T215V (1.3), F227L (1.0)	6.1	88	K103N (100), T215V (6.9), K219W (36.1)	
11 ^d	BP	5.8	K103N (100), T215V (1.3), F227L (1.0)	5.2	70	K65N (8.1), K103N (100), F227L (0.2)	
12	BP	6.2	F227L (0.8)	5.9	19	K65N (15), D67N (4.3), Y181S (12), M184I (1.6), T215I	
	SP	3.1	D67N (2.3)			(1.0), K219W (5.1)	
13	BP	3.7	None	3.9	14	T215V (5.1), K219W (16.7), F227L (0.1)	
14	BP	5.2	G190S (0.5), T215I (3.8), F227L (0.7)	5.2	85	None	
15	BP	5.5	None	3.7	87	M184I (1.8), T215V (1.3), K219W (24.6), F227L (0.1)	
16 ^{<i>d</i>}	BP	4.7	K65R (5.8), D67N (96.5)	4.9	137	K65R (8.8), D67N (91), T215S (100), K219E (88.4) , F227L (0.1)	
17 ^d	BP	4.7	K65R (5.8), D67N (96.5)	5.2	32	K65R (4.8), D67N (99), K219E (68.6)	
19	BP	4.7	D67N (1.0)	4.6	140	K103N (1.4)	
21	SP	4.0	D67N (2.8), K219W (77.4)	4.6	85	F227L (0.1), K219W (24.6)	
	BP	4.9	K101E (8.3), T215I (1.5), K219W (4.9), F227L (0.1)				
22	BP	4.4	None	4.4	70	T215V (5.7), F227L (0.1), K219W (32.4)	
23	BP	4.5	D67N (5.3), K219W (4.4), F227L (0.3)	5.1	114	None	
24	SP	2.6	F227L (1.1)	4.9	96	F227L (0.3)	
	BP	2.6	K65R (10.1), K70R (2.3), T215V (0.2), K219W (5.4)				
25 ^d	BP	4.8	Y188H (0.8), T215S (3.2), K219W (7.8), F227L (0.1)	6.1	11	F227L (1.5)	
27 ^d	BP	4.8	Y188H (0.8), T215S (3.2), K219W (7.8), F227L (0.1)	3.8	70	None	

^aBP, blood plasma; SP, seminal plasma.

^bVL, viral load in log HIV RNA copies/ml.

 ${}^c\!\mathsf{DRM}$ with frequencies above 20% are in bold.

^dShared source partner for transmission pairs 3, 25, 27, pairs 16 and 17, and pairs 10 and 11. The recipient partner from pair 10 was also the source partner for pair 9 (also see Fig. 1).

^eNA, not available because of PCR failure.



FIG 2 (Continued)

a minority DRM was present in the source at a given site, the probability of detecting a minority DRM at the same site in the recipient ($P_{R+ISmin}$) was similarly low at 0.08 (0.02 to 0.24) (Fig. 2A, middle panel). As a comparison, the probability of DRM detection in the recipient when the source had a majority DRM ($P_{R+ISmaj}$) was much higher at 0.73



FIG 2 Probability of observing DRM in the recipient when the DRM in the source partner is absent (left panels), present at a frequency of \leq 20% (middle panels), or present at a frequency of >20% (right panels) for the 24 transmission pairs (A), for pair 9 as an example (B), and after exclusion of mutation K219W (C). The colored diamonds indicate actual observations of DRM in the recipient at baseline (0, absence of mutation; 1, presence of mutation [right *y* axis]). Red diamonds, nonnucleoside reverse transcriptase inhibitors (NNRTIs); blue diamonds, nucleoside reverse transcriptase inhibitors (NRTIs). The horizontal line and the large dark gray dot indicate the median of the posterior distribution of the probability of DRM being present in the recipient for each scenario in the source (left *y* axis). The shaded area and the vertical line represent a 95% credible interval for each condition and each pair, respectively. (A) The 24 transmission pairs. (B) As an example, in pair 9, F227V was detected at baseline in the recipient but was absent in the source (left *y* and), T215V was present in the source (<20%) and in the recipient at baseline (middle panel), and K103N was present (>20%) in the source and in the recipient while K219 was detected only in the source (>20%) and was absent in the recipient at baseline (right panel). Mutations present in the source and/or in the recipient action of K219W. Excluding K219W did not qualitatively change other effects

in the model, and the BFs remained above substantial (BF = 6.37 and 4.41 before and after exclusion of K219W, respectively).

(0.29 to 0.96) (Fig. 2A, right panel). One example transmission pair (pair 9), where the individual dots are labeled with their respective mutations, is proposed in Fig. 2B. Altogether, Bayesian models confirmed that there was no increased probability of detecting a minority DRM in the recipient at a given site when present in the source (Bayes factor [BF] = 6.37 for $P_{R+IS-} = P_{R+ISmin}$) but were strongly indicative of an association between the presence of majority DRM in the source and in the recipients' partner P_{R+IS-} (BF \approx 0). Due to the high frequency of the rare K219W mutation, which is likely not clinically relevant in contrast to other mutations at the same position (e.g., K219QE), we performed a new analysis to ensure that our conclusions were not affected by this mutation. Excluding K219W in fact did not qualitatively change other effects in the model; the percent changes in the regression coefficients were less than 5%, and all (non)significant effects remained so. The BFs remained above substantial as well (BF = 6.37 and 4.41 before and after exclusion of K219W, respectively) (Fig. 2C and Table 3).

To adjust for variations in HIV RNA levels, we investigated whether the absolute copy numbers of HIV RNA carrying a DRM (DRM_{copies}) in the source were predictive of the presence of DRM in the recipient at each specific site when DRM was not a majority population. Similarly to the case for $DRM_{Freq'}$ higher DRM_{copies} in the source also was not associated with increased probability of observing a DRM in the recipient (odds

TABLE 3 Adjusted and unadjusted Bayes factor values for transmission of minority and majority DRM and covariates

		BF					
Model	Adjustment ^a	Minority ^b	Majority ^c	Covariate			
Base	Unadjusted	6.37	< 0.01	NA ^d			
Base without K219		4.41	<0.01				
Base	Class of DRM	6.86	<0.01	4.84			
	Sequencing coverage	3.23	< 0.01	3.65			
	EDI	5.36	< 0.01	5.61			
	Semen compartment	4.94	<0.01	0.88			

^aUnadjusted BFs are derived from models without covariates, while adjusted BFs are from models with covariates.

^bThe BFs indicate substantial evidence against transmission of the minority DRM with or without adjustment (i.e., BF > 3).

The BFs indicate decisive evidence for transmission of the majority DRM with or without adjustment. d NA, not applicable.

ratio [OR] [95% credible interval] = 1.03 [0.83 to 1.26] and BF = 48.7 for DRM_{Freq}; OR = 1.04 [0.89 to 1.2] and BF = 60.2 for DRM_{copies}). Finally, to ensure that our initial results were not due to spurious detection of very low-level minority DRM, we performed sensitivity analyses with various DRM minimum thresholds by excluding minority DRM with frequencies below 10%, 5%, and 1% from our initial model. Bayesian models confirmed our observations regardless of the DRM minimum threshold (BFs for $P_{R+IS-} = P_{R+ISmin}$ were 4.3, 5.1, and 3.6 after discarding minority DRM with frequencies below 10%, 5%, and 1%, respectively).

Sensitivity analysis. Using the baseline model described above, we examined the potential confounding effects of relevant covariates, i.e., time from EDI to sampling, compartment of sampling (i.e., blood or seminal plasma), class of drug resistance mutation (i.e., nonnucleoside reverse transcriptase inhibitors [NNRTIs] versus nucleoside reverse transcriptase inhibitors [NNRTIs] versus nucleoside reverse transcriptase inhibitors [NNRTIs] versus nucleoside reverse transcriptase inhibitors [NRTIs], or depth of sequencing coverage (Table 3). None of these covariates significantly impacted the probability of the recipient having a minority DRM. That is, Bayes factors for minority DRM all remained >3 after adjustment, indicative of substantial evidence against transmission of the minority DRM with or without adjustment: EDI (OR = 1.71 [0.65 to 4.65], BF = 5.61, change in leave-one-out cross validation information criterion [Δ LOOIC] with standard error [SE] = -0.2 ± 2.3 , compartment of sampling (OR = 51.6 [10.3 to 371.1], BF = 0.88, Δ LOOIC = 2.5 ± 3.4), class of DRM (OR = 1.19 [0.15 to 9.37], BF = 4.84), and sequencing coverage (OR = 1.78 [0.93 to 3.81], BF = 3.65, Δ LOOIC = 1.66 ± 3.0). Similarly, the BFs for majority DRM remained at <0.01, indicative of decisive evidence for transmission of the majority DRM with or without adjustment for the covariates mentioned above (Table 3).

No persistence of minority DRM during the course of infection. Finally, we evaluated the dynamics of DRM over a median duration of follow-up of 106 days (IQR, 56 to 219) among the 11 recipients with longitudinally collected samples. Using deep-sequence data from these samples, we fitted a mixed-effects exponential decay model. The decay rate was significantly greater than 0 (λ = 3.21 [1.19 to 11.64], BF = 0.05), indicating an exponential decay of the DRM frequency (Fig. 3). We found similar results after adjusting DRM frequencies to HIV RNA levels (i.e., by using DRM_{copies}). To further explore the relationship between decay and the EDI, we included EDI in the exponential decay model. EDI was not associated with either the baseline DRM frequency or the rate of decay, with the BFs not indicating strong enough evidence (BF =6.91 and 0.9 for the baseline DRM frequency and the rate of decay, respectively). These results confirmed the significant decay of DRM over time regardless of the different sampling times from EDI across participants. Finally, the HIV pol/RT molecular diversity was modeled with a Bayesian hierarchical linear model with a log-normal link function. This model revealed a significant increase of the mean viral diversity over the study period (regression coefficient $\beta = 0.67$ [0.07 to 1.31], standard deviation = 0.2, BF = 0.02) (Fig. 4).



FIG 3 Exponential decay of DRM_{Freq} over time. The thin lines indicate means of DRM_{Freq} across sites for each recipient. The thick lines and the gray shade represent the median and its 95% credible interval for the posterior distributions of the exponential-decay parameters. The Bayes factor indicates decisive evidence for exponential decay of DRM_{Freq} over time (~0.05).

DISCUSSION

FDA-approved genotyping methods can reliably identify mutations present in >20% of the circulating HIV population (13, 14). The advent of more-sensitive sequencing platforms has led to an increased identification of minority DRM among ART-naive HIV-infected individuals (6–8). Similarly to previous reports (6, 15), in this analysis of epidemiologically and genetically linked partner pairs, the presence of majority DRM in the recipient was associated with the presence of the same DRM in the source, consistent with sexual transmission. Conversely, we found no increased probability of detecting a minority DRM in the recipient at any given site when the same minority DRM was detectable in the source (compared to no DMR in the source). This observation suggests that the detection of minority DRM in recently infected individual is not likely to be the consequence of sexual transmission of variants harboring DRM, consistent with our previous report (11).

In the 11 recipient partners with longitudinal sampling, we found a significant exponential decay of the DRM frequency while viral *pol* diversity increased over time. This rapid reversion of viruses harboring minority DRM to the wild type supports the hypothesis that the detected DRM at low levels are unlikely the consequence of



FIG 4 Increased HIV *pol* diversity over time. The thin lines indicate means of diversity across sites for each recipient. The thick lines and the gray shade represent the median and its 95% credible interval for the posterior distributions of the model parameters (Bayesian hierarchical linear model with a log-normal link function). The Bayes factor indicates very strong evidence for an increase in diversity over time (0.02).

false-positive detection due to technical errors. Indeed, this exponential decay in DRM frequency in the face of increasing diversity suggests the presence of negative selection. The presence of minority DRM mainly during early infection, when the effective population size (N_e) is low, is consistent with the mutation-selection balance hypothesis, in which deleterious mutations (i.e., DRM) are more efficiently purged later during HIV infection, when the larger effective population size allows selection to more efficiently remove slightly deleterious mutations. At early stages of infection, the inefficiency of selection under low- N_e conditions allows DRM to appear and persist at low frequencies. The rapid disappearance of minority DRM over the course of infection is consistent with the importance of N_e as a factor shaping the patterns of molecular evolution.

Although our study offers new and interesting data about the appearance of minority DRM, it has some important limitations. First, given the time elapsed between the EDI and sampling of the recipient partner's blood and between sampling of the recipient and the identified source, we cannot rule out that unobserved selective pressures and other sampling-related biases may have driven the emergence and/or disappearance of minority DRM. Second, while the genital tract is a known reservoir for distinct HIV variants (16–18), our analysis found no significant compartment effect, most likely because of the small samples size of 4 sources with genital secretion (19), which limited our ability to rule out transmission of minority DRM from the genital compartment.

When assessing viral variants related to HIV drug resistance at low levels, one must also consider the absolute level of HIV to assess template resampling and its effects on frequency determination and error associated with PCR amplification. In our study, HIV-1 RNA levels exceeded 1,000 copies/ml for all but one sample, limiting the risk of oversampling by deep sequencing. In addition, we confirmed our observations after adjusting for the level of HIV RNA and after excluding DRM at very low levels in our sensitivity analysis.

Another important limitation includes the use of the 454/Roche platform for sequencing: this high-throughput sequencing platform generates massively parallel short reads that are prone to homopolymer-associated, per-base errors (20). Variability between replicated deep-sequencing runs on the 454 platform has previously been reported (21, 22), with minority variants detected at levels between 1 and 5% in one replicate while being undetectable in another. In our previous work, we also observed discordance in the detection of amino acid residues in more than half of the replicate runs (11). To overcome these limitations, we apply rigorous quality control procedures for deep sequencing (23, 24). Our validated bioinformatics pipeline (25-27) includes strict quality filtering steps, and only quality-controlled reads were included in the iterative alignment procedure. Ideally, the relatedness of existing DRM variants within transmission pairs could be assessed in a phylogenetic framework. However, the 454 platform used here produced short individual reads that were not suitable for proper phylogenetic analysis. Therefore, we relied on a Bayesian-based statistical framework to model the sexual transmission and persistence of DRM. One major advantage of using this Bayesian framework is that it allows us to directly evaluate a null hypothesis (i.e., the probability of no transmission given the data). Evaluating the null hypothesis based on the more commonly used frequentist framework is almost impossible to do, since a P value (the probability of obtaining the data given the null hypothesis) of greater than the typical 5% alpha level does not support the null hypothesis.

Finally, our study assessed the putative transmission of DRM from treatment-naive source partners and did not include treatment-experienced sources who could have been more relevant/informative for study of transmitted DRM. However, most HIV transmission events occur when the source is still ART naive and often is unaware of their HIV status (28, 29), and investigating transmission of minority variants in this context remains enlightening for assessment of the transmission of minority variants and more broadly for understanding the biology concerning male-to-male sexual transmission of HIV.

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Parameter	Value
Minimum PHRED score for inclusion of sites	15
Minimum no. of sites with PHRED score higher than the minimum for inclusion	75
of reads/fragments of reads	
Minimum read length to be included in subsequent analyses	100
Minimum coverage	250
Window size for sliding-window estimation of nucleotide diversity	150
Length of sliding-window stride	20
Minimum no. of copies for a read to be considered a variant	10

In summary, in this study we applied deep-sequencing methods within a Bayesian framework in a uniquely characterized population of source and recipient HIV-infected partners. Despite the important limitations, we found no clear evidence to support the sexual transmission of minority resistant variants, although it should be noted that the absence of evidence does not rule out potential transmission of minority DRM via sexual exposure. Overall, our results suggested that minor resistant variants may emerge *de novo* shortly after transmission, when the small effective population size limits efficient purge by natural selection (30, 31). Future studies using recent sequencing technologies that provide longer reads and are less prone to single-base errors (MiSeq or PacBio with or without Primer ID) (32, 33) or single-genome amplification/ Sanger sequencing in different settings (i.e., ART-experienced HIV-infected source partners) within a similar Bayesian framework would be informative for detecting clinically relevant minority variants and investigating their propensity for transmission.

MATERIALS AND METHODS

Ethics statement. The UCSD Human Research Protections Program approved the study protocol, consent, and procedures for consent. All study participants provided written informed consent before any study procedures were undertaken.

Subject characteristics and data collection. A total of 24 genetically and epidemiologically linked transmission pairs were included. Partnerships were identified through contact tracing of individuals newly diagnosed with primary or early HIV infection as part of the San Diego Primary Infection Resource Consortium (SD-PIRC). Each newly diagnosed participant was asked about his most recent sexual partners. HIV transmission was confirmed by sequence analysis, and phylogenetic linkage was inferred when the two viral strains were >98.5% similar in the HIV-1 *pol/reverse* transcriptase (*pol/RT*)-coding region (34). The direction of transmission (i.e., source versus recipient) was defined based on the estimated date of infection (EDI) for each individual (35), based on participant reporting and standardized serologic algorithms (36). Demographics, laboratory data (CD4 T-cell count and HIV RNA levels), and data on HIV risk factors (e.g., intravenous drug usage or sexually transmitted infections) and sexual behavior (receptive and insertive anal intercourse and number of sex partners) were collected. Blood plasma from both partners was collected at the time of recruitment; a subset of 11 recipients had additional longitudinal samples collected prior to ART initiation. Seminal plasma was collected at baseline for 4 out of the 21 (19.0%) unique source partners (37).

HIV RNA extraction and deep sequencing from blood and seminal plasma samples. HIV-1 coding regions of *pol*/RT (HXB2 coordinates 2708 to 3242) were amplified from blood and seminal plasma by PCR with region-specific primers (11). Deep sequencing was performed on a 454 GS FLX Titanium instrument (454 Life Sciences/Roche, Branford, CT) (12). For all but one seminal plasma sample, HIV-1 RNA levels exceeded 1,000 copies/ml.

Deep-sequencing processing and bioinformatics analysis. Read (FASTA) and quality score files produced by the 454 instruments were analyzed using a purpose-built bioinformatics pipeline (25–27). The pipeline is available at https://github.com/veg/HIV-NGS, and the key steps are as follows. (i) The first filtering step included exclusion of low-quality reads (*q* score of <15) and correction for random sequencing errors or homopolymers (25). (ii) High-quality reads were aligned to the reference sequence (HXB2) using a codon-based algorithm (12). (iii) We applied a Bayesian inferences with Dirichlet multinomial mixtures model (38) to distinguish true low-frequency variants from sequencing errors (posterior probabilities of \geq 99.99%). (iv) Representative reads were screened for evidence of recombination using GARD (39), APOBEC signatures, and frameshifts (40). (v) All sequences were screened for in-house cross-contamination using BLAST (41). Quality filtering parameters for deep-sequencing data are presented in Table 4. We computed the mean of all pairwise Tamura-Nei 93 distances between reads with at least 100 overlapping base pairs to quantify nucleotide diversity (42).

Identification of DRM. We identified mutations associated with known resistance to nucleoside reverse transcriptase inhibitors (NRTIs) and nonnucleoside reverse transcriptase inhibitor (NNRTIs) according to the Stanford Drug Resistance Database (score of >35) (http://hivdb.stanford.edu). Filtered

deep-sequencing reads were screened for the presence of amino acid mutations at these sites, and the relative frequency of each identified DRM was obtained as part of the bioinformatics pipeline (25, 26).

Estimation of lower limit of detection for DRM. Since the lower limit of detection for each DRM is dependent on the sampling error and background error rate for each site, we included only DRM with residue frequencies greater than the background error rate, estimated using our published binomial mixture model (11) implemented in the online molecular sequence analysis server DataMonkey (43).

Statistical analyses. Our objective was to evaluate the rate at which a DRM at a particular site was transmitted from source partners to their respective recipients and whether the transmission was affected by the relative frequency (DRM_{\rm Freq}) or the absolute copy numbers (DRM_{\rm copies} [DRM frequency imesHIV RNA copies/ml]) of each mutation. Minority DRM was defined as DRM with frequency below 20% (DRM_{Freq} \leq 20%). We tested whether the probability of observing a DRM in the recipient's blood plasma HIV RNA population sampled at baseline at a particular site differed across three conditions: (i) the source had no DRM (R+IS-), (ii) a minority DRM was present in the source at a particular site (R+IS_{min}), and (iii) a majority DRM (DRM_{Freq} > 20%) was present in the source at a particular site (R+IS_{mai}). Transmission of HIV variants harboring DRM can be affected by (i) between-pair variability (DRM might have transmitted within a particular pair more or less easily) and (ii) between-site variability (some DRM at a particular site might have transmitted more or less easily). To incorporate these sources of variability into a model, we fit Bayesian hierarchical Bernoulli logistic regression models with the three conditions as a within-pair fixed effect and two random intercepts for pairs and sites as crossed random effects. For modeling longitudinal data, we were mainly concerned with modeling two outcomes: DRM and diversity. For modeling DRM (DRM_{Freq} and DRM_{copies}), we fit a Bayesian hierarchical exponential-decay model that is represented by the function $\mu_i = N_0 \exp(-\lambda t_i)$, where t, μ_i , $N_{o'}$ and λ represent the time point, the mean at time point *i*, the baseline value at t = 0, and the rate of change, respectively. At each time point, the DRM was modeled with a normal distribution. Recipients were allowed to have their own random intercepts for N_{0} , and λ . The mean of diversity at each time point was modeled with a linear model with a log-normal link function. Diversity was assumed to be log-normally distributed at each time point because of its positive skew, while its change in the mean at each time point was assumed to be linear. Again, each recipient had his own random intercept and time slope.

The model convergence was evaluated with R hat, which is the ratio of the average variance of 1,000 Markov-chain Monte Carlo (MCMC) samples within each of the 4 chains to the variance of the pooled MCMC samples across chains. An R hat of <1.1 is considered acceptable convergence (44). We evaluated whether relevant covariates improved the base model by comparing the differences in the leave-one-out cross validation information criterion and the standard error (Δ LOOIC \pm 1 standard error [SE]) between two models. A positive ΔLOOIC value greater than 2 SE indicates an improvement in model fit. A weakly informative normal distribution with a mean of 0 and a standard deviation of 5 was used as a prior distribution for regression coefficients, and noninformative uniform distributions were used as prior distributions of the variances and random intercepts. While Bayesian modeling is not concerned with the traditional null-hypothesis significance testing, we provided the BF, a ratio of the probability of obtaining data given null and alternative hypotheses as a reference. BFs of 1 to 3, 3 to 10, 10 to 30, 30 to 100, and >100 are considered anecdotal, substantial, strong, very strong, and decisive evidence for a null hypothesis, respectively, while their reciprocals (1 to 1/3, etc.) indicate evidence against a null hypothesis (45). In other words, BFs of >3 are indicative of substantial evidence against transmission of the DRM, while BFs of \approx 0 indicate decisive evidence for transmission of a variant harboring a DRM. To examine potential confounding effects of relevant covariates on DRM transmission, we computed BFs in a model that included those covariates separately (Table 3). We used the R statistical language (46) and the rstan package (47) to fit Bayesian models.

Accession number(s). The sequencing data are available at the NCBI Sequence Read Archive under accession numbers SAMN04914210 to SAMN04914278.

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