

Single Motherhood, Alcohol Dependence, and Smoking During Pregnancy: A Propensity Score Analysis

MARY WALDRON, PH.D.,^{a,b,*} KATHLEEN K. BUCHOLZ, PH.D.,^b MIN LIAN, M.D., PH.D.,^c
CHRISTINA N. LESSOV-SCHLAGGAR, PH.D.,^b RUTH HUANG MILLER, PH.D.,^b MICHAEL T. LYNSEY, PH.D.,^d
VALERIE S. KNOPIK, PH.D.,^{e,f} PAMELA A. F. MADDEN, PH.D.,^b & ANDREW C. HEATH, D.PHIL.^b

^aDepartment of Counseling and Educational Psychology, School of Education, Indiana University, Bloomington, Indiana

^bMidwest Alcoholism Research Center, Department of Psychiatry, Washington University School of Medicine, St. Louis, Missouri

^cDepartment of Medicine, Washington University School of Medicine, St. Louis, Missouri

^dAddictions Department, Institute of Psychiatry, Psychology & Neuroscience, King's College, London, England

^eDivision of Behavioral Genetics, Department of Psychiatry, Rhode Island Hospital, Providence, Rhode Island

^fDepartment of Psychiatry & Human Behavior, Warren Alpert School of Medicine, Brown University, Providence, Rhode Island

ABSTRACT. Objective: Few studies linking single motherhood and maternal smoking during pregnancy consider correlated risk from problem substance use beyond history of smoking and concurrent use of alcohol. In the present study, we used propensity score methods to examine whether the risk of smoking during pregnancy associated with single motherhood is the result of potential confounders, including alcohol dependence. **Method:** Data were drawn from mothers participating in a birth cohort study of their female like-sex twin offspring ($n = 257$ African ancestry; $n = 1,711$ European or other ancestry). We conducted standard logistic regression models predicting smoking during pregnancy from single motherhood at twins' birth, followed by propensity score analyses comparing single-mother and two-parent families stratified by predicted probability of single motherhood. **Results:** In standard

models, single motherhood predicted increased risk of smoking during pregnancy in European ancestry but not African ancestry families. In propensity score analyses, rates of smoking during pregnancy were elevated in single-mother relative to two-parent European ancestry families across much of the spectrum a priori risk of single motherhood. Among African ancestry families, within-strata comparisons of smoking during pregnancy by single-mother status were nonsignificant. **Conclusions:** These findings highlight single motherhood as a unique risk factor for smoking during pregnancy in European ancestry mothers, over and above alcohol dependence. Additional research is needed to identify risks, beyond single motherhood, associated with smoking during pregnancy in African ancestry mothers. (*J. Stud. Alcohol Drugs*, 78, 745–753, 2017)

MATERNAL SMOKING DURING PREGNANCY (SDP) is a preventable cause of numerous adverse health outcomes for children (U.S. Department of Health and Human Services [USDHHS], 2014). Perinatal outcomes of SDP include intrauterine growth restriction, preterm delivery, and fetal and infant death (Dietz et al., 2010; Shah & Bracken, 2000; USDHHS, 2004). Congenital abnormalities consequent to smoking early in pregnancy have been documented as well (Hackshaw et al., 2011). Although debate continues regarding the direct or teratogenic effect of SDP on later outcomes (Knopik, 2009; Thapar & Rutter, 2009), elevated risk of disruptive behavior disorders, including attention-deficit/hyperactivity disorder, is also observed in children whose mothers smoked during pregnancy, compared with children

whose mothers abstained from smoking (e.g., D'Onofrio et al., 2008; Gaysina et al., 2013; Knopik et al., 2005).

Reducing prevalence of maternal SDP is among the national health objectives outlined most recently in Healthy People 2020 (USDHHS, 2011); however, current estimates of SDP are far from the 2020 target of fewer than 2% of pregnant women who smoke. Despite underreporting of SDP on birth certificates (Dietz et al., 1998), there are still a considerable number of women who admit to SDP: on 2014 birth certificates, for example, 8.4% of women reported smoking at some point during pregnancy, with 6.6% continuing to smoke in the third trimester (Curtin & Mathews, 2016). As part of ongoing surveillance by the Centers for Disease Control and Prevention, in 2009–2011, approximately 11% of mothers surveyed 2–6 months after childbirth reported SDP, defined as having smoked during the last 3 months of pregnancy (Tong et al., 2013).

Rates of maternal SDP are especially high among women who are unmarried at childbirth. More than three times as many unmarried as married mothers reported SDP on 2014 birth certificates (14.7% vs. 4.1%; Curtin & Mathews, 2016), which is broadly consistent with earlier research, including studies conducted outside the United States (e.g., Kiernan & Pickett, 2006; Raatikainen et al., 2005). Furthermore, among pregnant smokers, unmarried women are less likely to quit

Received: August 18, 2016. Revision: December 28, 2016.

Funding for this study was provided by National Institute on Alcohol Abuse and Alcoholism Grants AA09022, AA011998, AA017688, AA017915, and AA021492; National Institute on Drug Abuse Grant DA023134; National Cancer Institute Grant CA178331; and *Eunice Kennedy Shriver* National Institute of Child Health and Human Development Grant HD049024.

*Correspondence may be sent to Mary Waldron at the Department of Counseling and Educational Psychology, School of Education, Indiana University, 201 N. Rose Avenue, Bloomington, IN 47405, or via email at: mwaldron@indiana.edu.

than women who are married (or cohabiting as though married; Dejin-Karlsson et al., 1996; Kiernan & Pickett, 2006; Martin et al., 2008); unmarried women are also more likely to relapse if they do quit (Martin et al., 2008). Not surprisingly, single mothers are a frequent target of smoking prevention and cessation programs, typically as part of efforts designed for low-income pregnant women (Chamberlain et al., 2013; USDHHS, 2014).

In the present study, we examined unique risk of SDP associated with single motherhood. Unmarried women face a number of challenges (McLanahan, 2009) that might separately account for and thus confound associations with prenatal smoking, including socioeconomic disadvantage (Graham et al., 2006). Compared with married mothers, on average, unmarried mothers are younger at childbirth (Hamilton et al., 2015; Shattuck & Kreider, 2013) and have lower education and income (Shattuck & Kreider, 2013), characteristics that likewise distinguish pregnant women who smoke from those who do not (Adams et al., 2008; Ebrahim et al., 2000). In addition, unmarried women are less likely than married women to receive adequate prenatal care related to financial and other barriers (Pagnini & Reichman, 2000), which is also true of pregnant smokers (Curtin & Mathews, 2016).

Although sociodemographic characteristics are routinely modeled as control variables, few studies linking single motherhood and SDP consider correlated risk from problem substance use beyond history of smoking (Martin et al., 2008) and concurrent use of alcohol during pregnancy (Martin et al., 2008; Urquia et al., 2013). To date, potential confounding by alcohol dependence is largely unaddressed, despite elevated rates of alcohol use disorder among women who smoke during pregnancy (Agrawal et al., 2008; Fergusson et al., 1998; Knopik et al., 2005, 2006) and among unmarried and separated or divorced parents, notably mothers (Waldron et al., 2013).

To examine whether single motherhood confers unique risk for smoking during early and later pregnancy, above and beyond shared risks associated with alcohol dependence, we conducted a propensity score analysis (PSA; Imbens & Rubin, 2015; Rosenbaum & Rubin, 1983). In PSA of non-experimental data, groups are matched on important predictors or *covariates* in the language of PSA. The purpose of matching is to reduce bias from confounders (see Rubin & Rosenbaum, 1985), particularly those likely to be overrepresented in certain groups, such as alcohol dependence among single mothers who smoke during pregnancy. To the extent that single mothers and mothers in two-parent families are observed to be well matched across predictors, where covariate balance is achieved, our confidence in comparisons of single- and two-parent families is increased. Where matching on covariates cannot be achieved, the relationship between single motherhood and maternal SDP may be spurious, due to differences predating pregnancy or childbirth, including alcohol use disorder.

Method

Participants

Data were drawn from twin-families participating in the Missouri Adolescent Female Twin Study (MOAFTS; Heath et al., 1999, 2002; Waldron et al., 2013). MOAFTS is a prospective study targeting the total cohort of female like-sex twin pairs born in Missouri to Missouri-resident parents, identified from birth records, for the period July 1, 1975–June 30, 1985 [$n = 370$ African ancestry (AA), 1,999 European or other ancestry (EA) pairs]. A cohort-sequential sampling design was used. Initial cohorts of 13-, 15-, 17-, and 19-year-old twins and their families were recruited during the first 2 years of data collection, with continued recruitment of 13-year-olds in years 3 through 4. Baseline parental interviews were conducted at twins' median age 15 (Wave 1). All participants gave verbal consent (assent if minors) following procedures approved by the Institutional Review Board at Washington University.

Analyses were limited to interview data obtained from biological mothers of twins. Of 1,738 EA and 264 AA mothers completing Wave 1 interviews, 1,714 (99%) and 260 (98%) had data on SDP, parental relationship status at childbirth, maternal and paternal histories of alcohol dependence, and relevant sociodemographic covariates. Six (3 EA, 3 AA) families were excluded from analyses because the biological father of the twins died before twins' birth or a nonbiological father figure (e.g., stepfather, adoptive father) was present at twins' birth; these exclusions reduced the sample to 1,711 EA and 257 AA mothers. At Wave 1, EA and AA mothers ranged in age from 25 to 64 years, with median ages 41 and 40, respectively. Additional sample characteristics are provided in Table 1 by race/ethnicity and single motherhood.

Measures

Individual measures were drawn from telephone adaptations of the Semi-Structured Assessment of the Genetics of Alcoholism (SSAGA; Bucholz et al., 1994; Hesselbrock et al., 1999), a semi-structured diagnostic interview developed for the Collaborative Study on the Genetics of Alcoholism (Begleiter et al., 1995). The SSAGA has documented validity (Hesselbrock et al., 1999), and retest and interrater reliability is excellent (Bucholz et al., 1994, 1995). At Wave 1, parents completed the SSAGA-II, the *Diagnostic and Statistical Manual of Mental Disorders, Fourth Edition* (DSM-IV; American Psychiatric Association, 1994), update to the original DSM-III-R–based SSAGA.

Smoking during pregnancy. Parental interviews included self-report assessment of lifetime history of smoking. Mothers who reported smoking 100 or more cigarettes were asked if they smoked while pregnant and, if so, for how many

TABLE 1. Sample characteristics, by race/ethnicity and single motherhood

Variable	European ancestry families		African ancestry families	
	Two parent (<i>n</i> = 1,637)	Single mother (<i>n</i> = 74)	Two parent (<i>n</i> = 179)	Single mother (<i>n</i> = 78)
Smoking during pregnancy, <i>n</i> (%)				
First trimester	415 (25%)	34 (46%)*	45 (25%)	17 (22%)
Beyond first trimester	330 (20%)	32 (43%)*	36 (20%)	17 (22%)
Covariates				
Parental alcohol dependence symptoms, ^a <i>M</i> (<i>SD</i>)				
Maternal	0.08 (0.37)	0.24 (0.62)*	0.07 (0.34)	0.12 (0.43)
Paternal	0.82 (0.96)	0.97 (1.02)	0.71 (0.95)	0.71 (0.97)
Education, in years, <i>M</i> (<i>SD</i>)	13.38 (2.08)	12.32 (2.06)*	13.23 (1.89)	12.21 (1.64)*
Age at childbirth, <i>M</i> (<i>SD</i>)	26.91 (4.82)	22.88 (5.12)*	25.81 (4.96)	23.11 (5.35)*
Year of birth, <i>M</i> (<i>SD</i>)	1953.97 (5.47)	1959.03 (5.43)*	1955.34 (5.22)	1957.45 (5.33)*
Parity, firstborn, <i>n</i> (%)	672 (41)	47 (64)*	55 (31)	32 (41)
Neighborhood socioeconomic deprivation, ^b <i>M</i> (<i>SD</i>)	0.52 (0.27)	0.74 (0.40)*	1.09 (0.51)	1.38 (0.35)*

Notes: Within racial/ethnic group, single mother \neq two parent, * $p < .05$. ^aLog-transformed count of symptoms of DSM-IV alcohol dependence; ^blog-transformed principal component score for neighborhood socioeconomic deprivation.

weeks or months. Consistent with prior work (Lian et al., 2016), two variables were coded: (a) whether mother smoked during the first trimester, even before knowing she was pregnant and (b) whether mother continued smoking beyond the first trimester, the latter to reflect chronicity of SDP. Mothers who never smoked or smoked fewer than 100 cigarettes in their lifetime were included in the reference group, along with mothers who smoked 100 or more cigarettes but did not smoke while pregnant.

Single motherhood. Single motherhood was coded from maternal self-report of her relationship status at twins' birth. Mothers who were unmarried and living separately from the twins' father were coded as single at childbirth. Mothers who were married to or living with the twins' biological father were coded as residing in a two-parent family.

Parental alcohol dependence symptoms. Maternal and paternal symptoms of alcohol dependence were modeled as covariates in PSAs described below. Parental interviews included self-report assessment of lifetime history of DSM-IV alcohol dependence, with age at onset of individual symptoms (ranging from 0 to 7) also queried. For mothers, a count of alcohol dependence symptoms occurring before twins' birth was coded. Paternal alcohol dependence symptoms were assessed using maternal reports based on an adaptation of the Family History Assessment Module (Rice et al., 1995). We analyzed maternal reports because of reduced participation by fathers (Waldron et al., 2013) and prior findings of moderate to strong agreement between self- and coparent reports of alcohol dependence (Waldron et al., 2012). Because coparent alcohol dependence was assessed without regard to temporal clustering or age at onset, a lifetime count of alcohol dependence symptoms was coded for fathers. Given significant positive skew, symptom-count measures of both maternal and paternal alcohol dependence were log-transformed.

Sociodemographic covariates. A number of sociodemographic predictors were also modeled as covariates in PSAs: maternal educational attainment (years of schooling), maternal age at twins' birth, maternal year of birth, parity (whether twins are mother's firstborn), and socioeconomic deprivation of the neighborhood at the time of twins' birth. Using common factor principal components analysis, a neighborhood (census tract) socioeconomic deprivation variable was constructed using eight items identified from 20 census tract-level socioeconomic indicators from 1980 Census data, and matched to residential addresses listed on twins' birth records (Lian et al., 2016). Resulting composite scores for neighborhood socioeconomic deprivation ranged from -1.12 to 6.45, with higher values representing greater deprivation. Neighborhood socioeconomic deprivation was also log-transformed to correct for positive skew.

Analytic strategy

All analyses were performed in SAS Version 9.4 (SAS Institute Inc., Cary, NC) and conducted separately for EA and AA families, given racial/ethnic differences in SDP (Curtin & Mathews, 2016), family structure (Vespa et al., 2013), and their association (Perreira & Cortes, 2006). Preliminary descriptive analyses included tests of differences by single motherhood and twin-pair zygosity, the latter to identify potential limitations to the generalizability of twin data. Although identical (monozygotic) twinning occurs at random, fraternal (dizygotic) twinning is modestly related to maternal age and socioeconomic status (Bulmer, 1970); thus, monozygotic–dizygotic differences would flag domains with potential limits to generalizability.

Following preliminary analyses, we estimated a standard logistic regression model predicting SDP from single moth-

erhood. Results of the standard model provide an omnibus estimate of the association between single motherhood and SDP without regard to potential confounding. Next, PSA was conducted comparing SDP in single-mother and two-parent families across predicted probability of single motherhood. According to the Neyman–Rubin counterfactual principal (Sekhon, 2008), we are most confident about inferences from non-experimental data when sampling designs approximate the (impossible) counterfactual ideal of the same individuals evaluated simultaneously in A and not-A conditions. We approximate this ideal by matching single-mother and two-parent families on a priori risk of single motherhood, not otherwise possible in the standard regression model (even with statistical adjustment; Rosenbaum & Rubin, 1985). For covariate selection, specification of the propensity score, and construction of propensity score strata, we followed recently refined procedures outlined by Imbens and Rubin (2015), summarized below.

Predicted probabilities of single motherhood were estimated from logistic regression models predicting single motherhood from maternal and paternal alcohol dependence symptom counts, sociodemographic covariates, and second-order terms. To select second-order covariates, a baseline model was estimated predicting single motherhood from maternal and paternal alcohol dependence symptoms and sociodemographic covariates. Consistent with Imbens and Rubin (2015), quadratic and interaction terms were added to the baseline model one at a time, with a likelihood ratio statistic computed to test the null hypothesis that the coefficient on the additional covariate is equal to zero. Quadratic and interaction terms leading to the largest increase in the likelihood function greater than 2.71 were retained. For EA families, three second-order covariates were retained: (a) a quadratic of maternal age at twins' birth, (b) a two-way interaction between paternal alcohol dependence symptoms and neighborhood socioeconomic deprivation and (c) a two-way interaction between maternal year of birth and parity. For AA families, two second-order covariates were retained: (a) a quadratic of neighborhood socioeconomic deprivation and (b) a two-way interaction between paternal alcohol dependence symptoms and parity. Results of logistic regression models are provided in Supplemental Table A. As shown, for EA but not AA families, maternal alcohol dependence symptoms were associated with increased odds of single motherhood. For both EA and AA families, paternal alcohol dependence symptoms predicted reduced odds of single motherhood.

To ensure overlap in covariate distributions, we excluded or trimmed from the sample cases where there were no counterfactual matches, specifically, two-parent families with predicted probabilities less than the smallest predicted probability among single-mother families, and single-mother families with predicted probabilities greater than the largest predicted probability among two-parent families. Subsequent

comparisons were conducted using the trimmed sample, which included 1,466 EA mothers ($n = 72$ single mother, 1,394 two-parent families) and 214 AA mothers ($n = 76$ single mother, 138 two-parent families). After trimming, we tested for independence of single motherhood and predicted probability of single motherhood as the number of strata or blocks was progressively increased. For each block, we conducted t tests of the null hypothesis of equality of the average linearized propensity score by single-mother status. Following Imbens and Rubin (2015), blocks were split at the median value of the propensity score until all t statistics were below 1 or where splitting a block would lead to a new block with too few observations. For EA families, this process resulted in six strata, with prevalence of single motherhood ranging from 1% (Block 1, very low predicted probability of single motherhood) to 24% (Block 6, very high predicted probability). For AA families, three blocks were sufficient to achieve within-strata independence of single motherhood and predicted probability of single motherhood. Prevalence of single motherhood in AA families ranged from 24% (Block 1, low predicted probability of single motherhood) to 61% (Block 3, high predicted probability).

Next, we tested for within-strata differences in maternal and paternal alcohol dependence symptoms and sociodemographic covariates as a function of single motherhood. Conditional on successful matching of covariates (evident from findings of nonsignificant differences), within-strata comparisons of SDP were conducted followed by omnibus tests across strata. Two logistic regression models were estimated—one predicting smoking during the first trimester of pregnancy from single motherhood and another predicting smoking beyond the first trimester from single motherhood. For both models, we conducted follow-up analyses to produce Cochran–Mantel–Haenszel estimates of the common odds ratio (OR) (pooled across strata) and the Breslow–Day test for homogeneity of ORs across strata (Hosmer & Lemeshow, 2000).

Results

Descriptive analyses

Single-mother and two-parent families differed significantly in the distribution of covariates. As shown in Table 1, single EA mothers reported more symptoms of alcohol dependence than did married or cohabiting EA mothers. Single EA and single AA mothers reported having completed less education and they were younger at childbirth, born more recently (as assessed by maternal year of birth), and more likely to live in socioeconomically deprived neighborhoods at twins' birth compared with married or cohabiting mothers of the same race/ethnicity. In addition, single EA mothers were more likely to be first-time mothers. For both EA and AA families, twin-pair zygosity was unrelated to

TABLE 2. Prevalence of smoking during pregnancy (SDP) in two-parent and single-mother families by predicted probability of single motherhood: European ancestry families

Variable	Predicted probability of single motherhood SDP, n (%)					
	Block 1 (very low)		Block 2 (low)		Block 3 (low intermediate)	
	Two parent (n = 724)	Single mother (n = 9)	Two parent (n = 179)	Single mother (n = 4)	Two parent (n = 170)	Single mother (n = 13)
First trimester	151 (21)	4 (44)	45 (25)	3 (75) [†]	41 (24)	6 (46) [†]
Beyond first trimester	119 (16)	4 (44)*	35 (20)	3 (75)*	29 (17)	6 (46)*
Variable	Block 4 (high intermediate)		Block 5 (high)		Block 6 (very high)	
	Two parent (n = 170)	Single mother (n = 13)	Two parent (n = 81)	Single mother (n = 11)	Two parent (n = 70)	Single mother (n = 22)
	First trimester	51 (30)	6 (46)	38 (47)	6 (55)	31 (44)
Beyond first trimester	44 (26)	6 (46) [†]	27 (33)	5 (45)	26 (37)	8 (36)

Note: Within-block, single mother > two parent, * $p < .05$, [†] $p < .10$.

single motherhood and smoking beyond the first trimester of pregnancy ($p > .05$). For EA but not AA mothers, zygosity was predictive of smoking early in pregnancy, such that more EA mothers of fraternal twins reported smoking in the first trimester of pregnancy than did EA mothers of identical twins, $\chi^2(1) = 5.16, p < .05$.

Standard logistic regression

In EA families, single motherhood was associated with increased odds of smoking during the first trimester (OR = 2.50, 95% CI [1.56, 4.01]) and smoking beyond the first trimester (OR = 3.02, 95% CI [1.88, 4.85]). In AA families, the association between single motherhood and SDP was nonsignificant (smoking during first trimester: OR = 0.83, 95% CI [0.44, 1.57]; smoking beyond first trimester: OR = 1.11, 95% CI [0.58, 2.12]).

Propensity score analysis

Although associations between single motherhood and SDP in AA families were nonsignificant in the standard model, we proceeded with PSA to determine whether effects might be observed for a subset of covariate-matched AA mothers. Distribution of covariates in two-parent and single-mother families by predicted probability of single motherhood are provided in Supplemental Tables B and C. Across race/ethnicity, differences within strata between single-mother and two-parent families were nonsignificant ($p > .05$) with two exceptions: among EA families, in Blocks 1 and 2, single mothers reported fewer alcohol dependence symptoms (in fact, none) relative to mothers in two-parent families. With these exceptions, single-mother and two-parent families were well matched on maternal and paternal alcohol dependence symptoms and sociodemographic covariates in both EA and AA families.

Given successful matching, within-strata comparisons of

SDP by single-mother status were conducted. As shown in Table 2, with the exception of Block 6, we observed elevated rates of SDP in single-mother relative to two-parent EA families across the spectrum of a priori risk of single motherhood. In follow-up analyses, Breslow–Day tests for homogeneity of ORs across strata (first trimester: OR = 1.79, 95% CI [1.09, 2.93]; beyond first trimester: OR = 2.23, 95% CI [1.35, 3.37]) showed consistency of the association between single motherhood and SDP across strata, that is, across families stratified by predicted probability of single motherhood (first trimester: $\chi^2[5] = 5.80, p = .33$; smoking beyond the first trimester: $\chi^2[5] = 7.62, p = .18$). Among AA families, within-strata differences in rates of SDP were nonsignificant (Table 3), as were post hoc tests of the common OR.

Discussion

Single motherhood is a well-documented predictor of maternal SDP, although few studies of unmarried pregnant smokers consider correlated risk from problem substance use. To our knowledge, the present study is the first to examine whether the effect of single motherhood on SDP is distinct from shared risks associated with alcohol dependence. In addition to standard logistic regression, to reduce bias from potential confounders, we conducted a PSA comparing rates of SDP in single-mother and two-parent families stratified by predicted probability of single motherhood.

Our findings highlight the importance of single motherhood as a unique risk factor for SDP among EA mothers, separate from alcohol dependence and highly correlated sociodemographic risk factors. In standard logistic regression models, single motherhood was associated with more than two times increased odds of SDP during the first trimester and three times increased odds of smoking beyond the first trimester. Results of PSA largely confirm the specificity of risk associated with single motherhood: with the exception of the highest probability families, increased rates of SDP

TABLE 3. Prevalence of smoking during pregnancy in two-parent and single-mother families by predicted probability of single motherhood: African ancestry families

Variable	Predicted probability of single motherhood SDP, <i>n</i> (%)					
	Block 1 (low)		Block 2 (intermediate)		Block 3 (high)	
	Two parent (<i>n</i> = 81)	Single mother (<i>n</i> = 26)	Two parent (<i>n</i> = 36)	Single mother (<i>n</i> = 17)	Two parent (<i>n</i> = 21)	Single mother (<i>n</i> = 33)
First trimester	22 (27)	6 (23)	8 (22)	3 (18)	7 (33)	6 (18)
Beyond first trimester	18 (22)	6 (23)	5 (13)	3 (18)	6 (29)	6 (18)

were observed across the spectrum of predicted probability of single motherhood. Although it could be that EA mothers at very high risk of single motherhood will smoke regardless of relationship status, homogeneity of ORs across strata was supported, suggesting similar effects of single motherhood for EA women at varying degrees of a priori risk of single motherhood, including those at very high risk.

Among African American women, we observed little if any association between single motherhood and SDP (and symptoms of alcohol dependence), regardless of statistical approach. It is possible that single motherhood is unrelated to SDP in African American families, consistent with evidence suggesting reduced risk to African American relative to European American offspring of single mothers, because of extended familial supports but also reduced stigma (Amato, 1994). However, wide confidence intervals suggest such conclusions may be premature. Unfortunately, most studies ignore potential cultural differences, with racial/ethnic group typically modeled as a control variable if considered at all (Kiernan & Pickett, 2006; Martin et al., 2008; Page et al., 2012). Before strong inferences can be made, additional research is needed specific to African Americans. Studies of African American mothers using matching designs, based on between- or within-family contrasts (Heath et al., 2014), will necessarily require larger samples than ours for sufficiently powered analyses. This is especially true of within-family designs, where very large samples may be required to identify siblings or twins who are discordant for single motherhood.

Strengths and limitations

The present study has a number of strengths, notably use of PSA to match on parental alcohol dependence and sociodemographic risk factors, with analyses conducted separately for European and African American families. As reviewed, prior research linking single motherhood and risk of SDP has not considered potential confounding by alcohol dependence, which is predictive of both dissolution of reproductive relationships (Waldron et al., 2013) and maternal SDP (Agrawal et al., 2008; Fergusson et al., 1998; Knopik et al., 2005). Although statistical power was limited for some comparisons, results of PSA provide increased confidence in the effect of single motherhood in European American

families, relative to standard models where covariate balance is not addressed. Additional strengths of the study include analyses of data drawn from a representative sample of twins and their families ascertained from state vital records, with minimal nonresponse by mothers. Representative sampling with high rates of participation is essential to generalize findings. However, replication in a nontwin sample including male offspring remains important given observed differences by twin zygosity.

Many limitations of the study relate to assessment. We relied on maternal report for all measures, including SDP assessed approximately 15 years after childbirth. Although reporting that is more proximal to childbirth is preferable, available evidence supports the validity and reliability of retrospective self-report of SDP, as compared with other methods and sources, including prospective self-report, multiple reporters, and biological assays (Heath et al., 2003; Klebanoff et al., 2001; Knopik et al., 2016; Pickett et al., 2009). We also relied on maternal report of paternal alcohol dependence in part to reduce bias associated with nonresponse of fathers. As a consequence, paternal alcohol dependence was assessed without regard to timing of symptom onset. Although temporal precedence remains uncertain, based on national data, we anticipate that onset of alcohol dependence symptoms will predate twins' birth for most fathers. Among men in the United States, the average age at first diagnosis of alcohol dependence is in the early 20s, much earlier than the average age in our sample (of 28–29 years) of fathers at twins' birth.

Future directions

Because the primary goal of the study was to examine with greater specificity risk of SDP associated with single motherhood, underlying processes were not assessed. By ruling out important confounds upstream of single motherhood, results of PSA provide a strong basis for identification of potential mediators. Among candidates, to be examined in future studies, is partner support. Single mothers are less likely to receive financial and emotional support from non-cohabiting reproductive partners, compared with mothers who are married or cohabiting as though married (Carlson et al., 2004). Partner support also varies among noncohabit-

ing couples, who may or may not continue their romantic involvement (Kiernan & Pickett, 2006; Page et al., 2012). Romantic involvement may have particular relevance for African American families, for whom we observed no effect of single motherhood on SDP and for whom part-time cohabitation, where biological parents reside together on an irregular basis, is more common (Mott, 1990); thus, romantic involvement may be a better indicator of whether African American mothers are indeed single, as opposed to measures of marriage or cohabitation.

Conclusion and implications

Maternal SDP is a known, modifiable risk behavior associated with adverse child outcomes observed prenatally onward (USDHHS, 2014). In our study, European American mothers who were unmarried and living apart from their child's biological father were more likely than married and cohabiting mothers to smoke early and later during pregnancy, thus increasing risk to exposed children. Although European American single mothers are a potential focus of targeted smoking prevention, additional studies are needed to identify more salient predictors of maternal SDP in African American families, that is, beyond single motherhood and alcohol dependence.

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