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Association of Posttraumatic Stress Disorder With rs2267735 in the *ADCYAP1R1* Gene: A Meta-Analysis

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Abstract

Recent studies point to the potential role of the (pituitary) adenylyate cyclase activating polypeptide receptor 1 (*ADCYAP1R1*) gene, which has been implicated in stress response, in posttraumatic stress disorder (PTSD). Multiple genetic association studies have examined potential PTSD risk related to this gene, with mixed results. We conducted a meta-analysis of rs2267735 in *ADCYAP1R1* in PTSD. A literature search was conducted using PubMed and PsycINFO, resulting in nine studies that met criteria for inclusion in analysis. Biostat's Comprehensive Meta-Analysis was used to conduct the main meta-analysis on the combined sex sample, as well as two subanalyses examining effects separately in female and male participants. Results indicated that the C allele of rs2267735 conferred significant risk for PTSD in the combined sex data, $OR = 1.210$, 95% CI [1.007, 1.454], $p = .042$, and in the subsample of women and girls, $OR = 1.328$, 95% CI [1.026, 1.719], $p = .031$; but not in the subsample of men and boys, $OR = 0.964$, 95% CI [0.733, 1.269], $p = .796$. These results provide evidence for an association between *ADCYAP1R1* and PTSD and indicate that there may indeed be sex differences. Implications of these findings, including the role of rs2267735 as one modulator of the stress system, are discussed.

Exposure to traumatic events is a common experience, with recent worldwide surveys documenting that around 70% of individuals have experienced at least one trauma in their lifetime (Benjet et al., 2016; Liu et al., 2017), although the range varies across countries,

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likely associated with income, development, and conflict status. Of those who experience a traumatic event, between 5 and 31% meet lifetime criteria for posttraumatic stress disorder (PTSD; Adams & Boscarino, 2006; Breslau et al., 1998), with the number of traumas experienced increasing this risk in a dose-dependent manner (e.g., Kolassa et al., 2010; Neuner et al., 2004). Those who develop PTSD are at increased risk for outcomes such as major depression (Breslau, Davis, Peterson, & Schultz, 2000), substance misuse (Breslau, Davis, & Schultz, 2003), physical health problems (Zayfert, Dums, Ferguson, & Hegel, 2002) and unemployment and marital instability (Kessler, 2000).

In an effort to inform our understanding of the biological underpinnings of PTSD, researchers have examined candidate genes influencing the stress response system, including pituitary adenylate cyclase-activating polypeptide (PACAP) receptor (PAC1), encoded by the gene *ADCYAP1R1* and specifically the single nucleotide polymorphism (SNP) rs2267735. Chronic and/or unpredictable stress leads to an increase in PACAP in the bed nucleus of stria terminalis (BNST; the “extended amygdala”), which spurs heightened physiological responses to anxiety-provoking stimuli (e.g., Walker, Toufexis, & Davis, 2003). The BNST is also implicated in communication between the limbic cognitive centers and nuclei involved in stress processing (Crestani et al., 2013). Sustained or unpredictable stress also induces PACAP in the hypothalamic paraventricular nucleus (PVN; Legradi, Hannibal, & Lechan, 1998). This process spurs the release of corticotrophin-releasing hormone (CRH; Legradi et al., 1998) and adrenocorticotrophic hormone (Sapolsky, Romero, & Munck, 2000), thus influencing an individual’s ability to regulate stress response. Recent work has found a link between rs2267735 in *ADCYAP1R1* and decreased hippocampal activity (implicated in assigning emotional valences to events; Brady & Sinha, 2005) during contextual fear conditioning tasks among female participants, but not male participants (Pohlack et al., 2015). Therefore, following contact with feared stimuli, some individuals, and in particular female participants at risk in terms of *ADCYAP1R1*, may experience more pronounced physiological symptoms of anxiety, problems extinguishing fear-related cognitions, and difficulty modulating these stress responses, placing them at increased risk for PTSD.

These basic stress response processes are thought to have multiple etiologic influences, both genetic and environmental in nature, that are mutually impactful. Preclinical and clinical evidence suggests that although heritable influences exist for the basic stress response, environmental events such as early life stressors may also contribute to subsequent stress responding, perhaps via epigenetic mechanisms that may also increase risk for distress-related phenotypes, including PTSD, following trauma exposure (Wilker & Kolassa, 2013). Given that PTSD is moderately heritable (Sartor et al., 2012), an investigation of the PACAP system in predicting PTSD is critically important. In a landmark study by Ressler et al. (2011), which examined 44 SNPs spanning PACAP and PAC1, the SNP rs2267735 (in *ADCYAP1R1*) predicted PTSD among female participants only. Since this study, numerous replication attempts have led to mixed results, with some finding significant associations and others failing to find an effect (e.g., Chang et al., 2012; Rothbaum et al., 2014).

Meta-analytic approaches have recently been applied to candidate gene studies of PTSD as a systematic method for reconciling inconsistent findings. For example, research examining

the impact of genetic variants in the dopaminergic system and PTSD has yielded mixed results (e.g., Gelernter et al., 1999; Voisey et al., 2009). However, meta-analyses suggest that the SNP rs1800498 in *DRD2* and the 3'-UTR variable number tandem repeat (VNTR) in *SLC6A3* are significantly associated with PTSD (Li et al., 2016), whereas the association between rs4680 in *COMT* and PTSD was nonsignificant. To our knowledge, a meta-analysis of the association between variants in *ADCYAP1R1* and PTSD has not been conducted. Thus, we sought to perform meta-analyses to elucidate the associations among rs2267735 (the most commonly studied *ADCYAP1R1* polymorphism in the context of risk for PTSD), traumatic event exposure, and the development of PTSD. Given the extant literature, we hypothesized that the meta-analysis would show a significant effect of rs2267735 on PTSD, but that when analyzed separately by sex, there would be a significant effect for female, but not male, participants.

Method

Procedure

Search and selection of studies for inclusion—The present study aimed to identify existing studies examining the main effects of the *ADCYAP1R1* gene and PTSD. Any variants of the gene *ADCYAP1R1* with sufficient data to calculate a minimum of 3 effect sizes were included in the final selection (note that only rs2267735 met this criteria). Potential studies were identified through the PubMed and PsycINFO databases (as of September 2015). To ensure that the meta-analysis was up-to-date, an additional date-restricted search was conducted to identify any new studies that came out between September 2015 (initial search) and November 2016. Search terms were as follows: [posttraumatic stress disorder OR PTSD OR traumatic stress] AND [gene OR genetic] AND [*ADCYAP1R1* OR PAC1 OR pituitary adenylate cyclase-activating polypeptide type 1 receptor]. Additionally, the reference sections of two recent PTSD genetics review articles (Almli, Fani, Smith, & Ressler, 2014; Voisey, Young, Lawford, & Morris, 2014) were examined to identify any articles that may have been missed through the above search strategy.

Screening of search results—Two review authors (M.J.L. and C.M.S.) independently screened search results to select studies for possible inclusion. The following inclusion criteria were applied to titles and abstracts: (a) original research; (b) use of human subjects; (c) association study including the *ADCYAP1R1* gene; (d) PTSD as an outcome; and (e) trauma-exposed control group, with one exception. Specifically, for participants within one dataset (i.e., the Grady Trauma Project), some individuals were not trauma exposed. However, the extremely high level of trauma and very rare exceptions in exposure warranted retention of studies using this sample. Only effect sizes resulting from main effects of *ADCYAP1R1* and PTSD were included (effect sizes resulting from $G \times E$ analyses were excluded). In cases where the criteria were unclear, articles were more thoroughly examined and a consensus determination was made. Supplemental material was also reviewed for identification and extraction of relevant data. Further, an existing PTSD genome-wide association study (Nievergelt et al., 2015) included additional examination of the

ADCYAP1R1 gene. This was excluded in our meta-analysis given that limited information on the specific SNP was available.

Data extraction and coding—Two review authors (M.J.L. and C.M.S.) coded each of the included articles based on a predetermined coding manual that included relevant variables (e.g., lifetime vs. current PTSD, diagnosis vs. severity, gender, race/ethnicity) and specific variables concerning the quality of studies (e.g., inclusion/exclusion criteria, statistical corrections) developed and agreed upon by all authors. Following data extraction and entry, two independent reviewers separately checked the information for agreement across coders. There was 93.0% agreement between M.J.L. and C.M.S. Discrepancies were resolved through discussion until consensus was reached.

Effect Size Calculation

An odds ratio (OR) and confidence interval (CI) were calculated for each study to examine whether the C risk allele was associated with increased likelihood of PTSD diagnoses and symptoms. Two studies (Ressler et al., 2011; Solovieff et al., 2014), modeled the number of copies (i.e., 0, 1, 2) of the G allele as the risk allele, but were transformed here to represent the inverse direction of effects, that is accounting for risk in relation to the C allele. In addition to OR data, descriptive data, including means and standard deviations or frequency of occurrence, were used to calculate effect size. When studies reported results from a multiple regression analysis, OR was calculated using the natural exponential of the unstandardized regression coefficient. In cases where no estimate of standard error was reported, the confidence interval was obtained using equations provided by Altman and Bland (2011). We contacted the authors to request further information in cases of missing data necessary for effect size computation and all studies that met inclusionary criteria were included in the analyses.

Data Analysis

Three separate meta-analyses were conducted using Biostat's Comprehensive Meta-analysis (www.meta-analysis.com; Borenstein, Hedges, Higgins, & Rothstein, 2015). First, the main meta-analysis was conducted on effect size data with combined sexes ($\kappa = 10$ samples from 9 total studies). Next, two additional subanalyses were conducted using studies with sufficient data to calculate effect sizes separately for women ($\kappa = 9$) and men ($\kappa = 4$). Models employed random effects, which account for sampling error and random effects variance (Lipsey & Wilson, 2001). For completeness, results from a fixed effects model are also presented here. For the main meta-analysis, we conducted statistical tests of two moderators (percent African American and percent female) selected a priori with weighted regression analysis (meta-regression) using a mixed effects model. Percent African American was examined as a moderator via meta-regression for subanalyses that included a sufficient number of studies.

Homogeneity of the effect size distribution was examined with visual inspection, forest plots, and the Q statistic and I^2 (95% CI) index. The I^2 statistic quantifies the amount of statistical impact of heterogeneity on the total observed variation (Borenstein, Hedges, Higgins, & Rothstein, 2009). Interpretations of low ($I^2 = 25.0\%$), moderate ($I^2 = 50.0\%$), or

high ($I^2 = 75.0\%$) have been recommended by Higgins, Thompson, Deeks, and Altman (2003). To assess publication bias, we utilized Egger's regression index (Egger, Smith, Schneider, & Minder, 1997), the funnel plot, Duval and Tweedie's (2000) trim and fill, and Rosenthal's (1979) *failsafe N*.

Many studies exploring genetic associations with PTSD reported findings from multiple outcomes and we prioritized those with sufficient data to calculate OR. For example, some studies reported effect sizes for both PTSD severity and diagnosis or reported PTSD outcomes across lifetime and isolated to specific time periods. Shown in Table S1, we implemented a protocol to handle studies with multiple effect sizes to adhere to the assumption of independence, which refers to the assumption that each measure of effect is representative of independent studies. When data were presented separately for subgroups within an individual study, we conducted a meta-analysis to compute the combined effect size across subgroups under a fixed effects model as recommended by Borenstein, Hedges, Higgins, & Rothstein, (2009).

Finally, because some studies used an additive model, examining the effect of having 0, 1, or 2 risk alleles (C) and others compared high-risk genotype status (CC) with low-risk genotype status (CG/GG), a sensitivity analysis was performed to examine mean effect sizes separately for each type of model. An additional sensitivity analysis to account for study outcome, separately examining the mean effect sizes of studies measuring PTSD diagnosis and studies measuring PTSD severity, was also conducted.

Results

Figure 1 details findings from the search steps, which resulted in 20 unique articles. The only *ADCYAP1R1* polymorphism assessed by a minimum of three studies was rs2267735, so this SNP was chosen for meta-analysis. Several manuscripts conducted analyses on multiple samples (Almli et al., 2013; Chang et al., 2012; Ressler et al., 2011; two samples each), bringing the total number of potential samples included for analysis to 12. Two samples were excluded, one from Chang (due to the use of nontrauma-exposed controls) and one from Almli (due to identical data to that used in Ressler et al., 2011). Thus, a total of 10 samples, from 9 different manuscripts (7 from the initial search results, 1 from the date-restricted literature search, and 1 from review papers), met criteria for inclusion in analysis (see Tables 1 and 2 for a summary of each study, combined and separately by sex).

Quality Assessment

Studies included in the final analysis either clearly described recruitment processes and inclusion/exclusion criteria in published manuscripts, or provided details to the current study authors in separate correspondence. All included studies identified a psychometrically sound instrument or clinical interview used to measure PTSD and seven studies (Almli et al., 2013; Chang et al., 2012; Lowe et al., 2015; Ressler et al., 2011a,b; Stevens et al., 2014; Uddin et al., 2013) assessed comorbidities. All included studies assessed deviation based on the Hardy-Weinberg equilibrium. Only one study reported a deviation for rs2267735 (Rothbaum et al., 2014), but chose to include it given that it was a trend after multiple testing correction. Finally, of the studies testing multiple comparisons (multiple outcomes or multiple SNPs),

five applied statistical corrections (Lowe et al., 2015; Ressler et al., 2011a,b; Solovieff et al., 2014; Uddin et al., 2013) whereas four did not report implementing corrections to address false-positive results (Chang et al., 2012; Rothbaum et al., 2014; Stevens et al., 2014; Wang et al., 2013).

Primary Analyses

Main meta-analysis (combined sexes)—A total of 9,630 participants from 10 samples were included in the meta-analysis examining associations between variation within the *ADCYAP1R1* gene (PACAP receptor) SNP rs2267735 and PTSD with effect sizes combined across sexes. Note that five samples only included female participants and were also included in this analysis. Participants with the high-risk allele or genotype (C or CC) were more likely than participants with the low-risk allele or genotype (G or CG/GG) to demonstrate/report PTSD symptoms or diagnosis, $OR = 1.210$, 95% CI [1.007, 1.454], $p = .042$ (see Table 3). Effect sizes ranged from $OR = 0.840$ to 2.446. The heterogeneity of variance analysis was significant, indicating significant between-study variance. A random-effects multiple meta-regression model was significant overall, $Q(2) = 11.13$, $p = .004$, R^2 analog = 1.0; however, the individual moderator variables (percent female and percent African American) did not emerge as significant (Table S2). This finding appears to be limited by low variability, with the majority of studies including predominantly female samples ($\kappa = 8$, > 70.0%) and African American samples ($\kappa = 7$, 80.0%). Trim and fill analysis indicated the imputation of three studies to reduce bias (see Figure S2a), resulting in only a slight reduction in the overall effect size, $OR = 1.139$, 95% CI [0.955, 1.358]. Egger's regression was not significant, $B = 1.167$, $SE = 0.577$, $t(8) = 2.024$, $p = .078$, and Rosenthal's N indicated a minimum of 12 null hypothesis studies to lead to a p value at or above .05. Taken together, these findings indicate minimal risk for publication bias.

Subanalysis a (female subsample)—When the meta-analysis was conducted separately in the female subsample 8,723 participants from nine samples were included. As shown in Table 3, participants with the high-risk allele or genotype (C or CC) were more likely than participants with the low-risk allele or genotype (G or CG/GG) to demonstrate/report PTSD symptoms or diagnosis, $OR = 1.328$, 95% CI [1.026, 1.719], $p = .031$. Effect sizes ranged from $OR = 0.716$ to 2.096. The heterogeneity of variance analysis was significant, indicating significant between-study variance. Meta-regression revealed a trend of increasing effect sizes, $\kappa = 7$, $B = 0.0046$, 95% CI [-0.0006, 0.0098], $p = .081$, R^2 analog = .82, for higher percent African American. Analysis of publication bias indicated minimal risk with a Rosenthal's N of 16 null hypothesis studies to lead to $p = .05$. As displayed in Figure S2b, Trim and Fill analysis did not recommend the imputation of any studies. Egger's regression was not significant, $B = 1.410$, $SE = 0.775$, $t(7) = 1.818$, $p = .112$, indicating minimal publication bias.

Subanalysis b (male subsample)—Finally, four samples, comprised of 873 participants, were included in the meta-analysis examining rs2267735 and PTSD in male participants only. The effect was not significant, $OR = 0.964$, 95% CI [0.733, 1.269], $p = .796$ (see Table 3). Effect sizes ranged from $OR = 0.836$ to 1.347. The heterogeneity of variance analysis was not significant, indicating minimal between-study variance. Due to the

low number of studies, meta-regression was not conducted. The imputation of one study to reduce bias (see Figure S2c) was recommended by Trim and Fill analysis resulting in a slightly reduced effect size of $OR = 0.959$, 95% CI [0.730, 1.260]. Egger's regression was not significant, $B = 0.235$, $SE = 0.574$, $t(2) = 0.410$, $p = .721$, indicating minimal publication bias. Because the meta-analysis was not statistically significant, Rosenthal's N was not calculated.

Sensitivity Analyses (Combined Sample)

Five samples comprised of 6,299 participants used an additive model, examining the effect of having 0, 1, or 2 risk alleles (C; see Table S3). When examined separately, the studies resulted in a mean effect size of $OR = 1.215$ that was not statistically significant, 95% CI [0.937, 1.574], $p = .142$. Results were similar for the five samples that compared high-risk genotype status (CC) to low-risk genotype status (CG/GG; $n = 3,331$), $OR = 1.186$, 95% CI [0.905, 1.554], $p = .216$. Analysis of variance (ANOVA) analog did not reveal significant differences between effect sizes from studies using an additive model (number of C alleles) and studies comparing CC to CG/GG, $Q(1) = 0.016$, $p = .900$.

Regarding PTSD outcome, five samples ($n = 6,704$) provided sufficient data to calculate effect sizes for PTSD diagnosis and five samples ($n = 2,926$) provided sufficient data to calculate effect sizes for PTSD severity (see Table S4). Results were statistically significant when separately analyzed for severity, $OR = 1.488$, 95% CI [1.056, 2.096], $p = .023$, but not for diagnosis, $OR = 1.141$, 95% CI [0.923, 1.410], $p = .221$. Note that this discrepancy is likely due to differences in power considering that dichotomization of quantitative data typically results in reduced power and effect sizes in bivariate correlations (MacCallum, Zhang, Preacher, & Rucker, 2002). Further, ANOVA analog did not reveal significant differences between the two types of studies, $Q(1) = 1.667$, $p = .197$.

Discussion

This is, to our knowledge, the first meta-analysis of the *ADCYAP1R1* gene and PTSD, filling an important gap in the literature, as findings are mixed with regard to positive associations and nonreplications. Although numerous variants of *ADCYAP1R1* have been included in individual studies across multiple phenotypes, the only polymorphism that was consistently included in PTSD investigations and thus able to be meta-analyzed was rs2267735. Findings from the present investigation provide support for an association between rs2267735 within the *ADCYAP1R1* gene (PACAP receptor) and PTSD, with minimal risk of publication bias observed. Specifically, our findings suggest that the "C" allele is associated with increased risk for PTSD. Effect sizes did not differ between studies that examined the additive effect of the C risk allele (i.e., 0, 1, or 2 copies) and for those that compared individuals who were CC homozygous relative with those who were G carriers. Note that the overall effect size for the combined sample ($OR = 1.21$) was small, but in line with what would be expected based on the genetics literature (e.g., Gibson, 2012). The effect sizes for PTSD severity and PTSD diagnostic status did not differ. Additionally, percent female and percent African American did not moderate the main effect of rs2267735 on PTSD, likely due to low variability across studies. However, consistent with existing

frameworks for understanding the biological influence of *ADCYAP1R1*, some of our analyses suggest that the effect may be especially strong in women. However, the lack of effect seen for male participants should be considered preliminary, given the smaller sample size (~ 10.0% of the overall sample) and lower power to detect the effect size from the full sample within the male subsample.

The biological implication of these findings is substantial. PACAP is a critical component of the regulation of both central and peripheral stress responses, modulating the hypothalamic–pituitary–adrenal axis through activation of CRH as well as influence on catecholamine transmission and biosynthesis (Mustafa, 2013; Stroth, Holighaus, Ait-Ali, & Eiden, 2011; Vaudry et al., 2009). PACAP is highly preserved across species, permitting the use of transgenic animal models to inform our understanding of PACAP as related to fear, stress, and anxiety. Animal models examining PACAP and PAC1 receptor function using transgenic models have implicated PACAP/PAC1R in anxiety responses, startle and fear behavior, memory, pain, and hypothalamic–pituitary–adrenal function/responsiveness (including normal and stress-evoked corticosterone responses (Hammack & May, 2015; Meloni, Venkataraman, Donahue, & Carlezon, 2016). Consistent with evidence for the effects of PACAP/PAC1R on regulation of expression of CRH activity, a critical role of PACAP expression and signaling has been observed in the bed nucleus of the stria terminalis, the central nucleus of the amygdala, and the basolateral amygdala (Hammack & May, 2015; Lezak et al., 2014; Mustafa, 2013). Importantly, the role of PACAP is shown not only during acute, but also chronic stress (Mustafa et al., 2015), with some evidence that stress effects on anxious and depressive behaviors were contingent on PACAP (Lehmann, Mustafa, Eiden, Herkenham, & Eiden, 2013). Thus, elucidating the role of genetic variation in this locus in relation to PTSD is a critical step forward in understanding the pathogenesis of this key stress-related disorder.

Our results suggest that “C” is the risk allele. However, of some importance, there is considerable variability across studies with respect to how alleles are represented in analyses. Specifically, whereas some studies have examined the effect of the high-risk genotype (CC) vs. low-risk genotypes (GC and GG groups are combined), others have assumed a dose effect (i.e., analyses assume that having 2 C alleles is “worse” than having 1 allele, which in turn is worse than having none). Interestingly, the effect sizes resulting from studies that coded the number of C alleles in an additive fashion (i.e., 0, 1, 2) compared with those who examined the effect of having two risk alleles (CC vs. CG/GG) did not differ. However, careful attention to coding of alleles in future studies, as well as testing for differential effects for type of genetic model is needed.

An important caveat to our sex-specific pattern of findings is that the number of male participants in the investigations included in the meta-analysis is modest; therefore, the male-only analyses should be considered preliminary given that we were underpowered to detect an effect size of 1.21 (from the full sample) within the male subsample (power ~50.0%). However, overall, our analyses supported a particularly strong effect of rs2267735 on PTSD in the female subsample, consistent with evidence that estrogen may regulate PACAP/PAC1R expression. Indeed, animal models have shown estradiol-evoked upregulation of PACAP and PAC1R transcripts in the BNST of ovariectomized female rats

(Ressler et al., 2011). Functional hippocampal activity during cued and contextual fear conditioning has also been examined in healthy controls, with PAC1R-dependent hippocampal activation observed during late contextual fear acquisition in female participants (Pohlack et al., 2015). Importantly, however, sex effects have not been observed in research examining dark-enhanced startle in prepubertal children (6–13 years of age), with *ADCYAP1R1* rs2267735 found to be associated with startle reactivity across both girls and boys (Jovanovic et al., 2013), suggesting perhaps a later developmental onset of the sex-dependent findings of this locus.

Moderation analyses aimed at understanding the potential for race or ancestry effects were limited by the fact that only African American self-reported race could be tested in published samples. Further, the studies with greater proportions of African American participants were also comprised of greater proportions of female participants, making it difficult to draw strong conclusions about whether there may be unique ancestry effects. HapMap data does point to slightly increased prevalence of the C allele in Yoruban (African) samples relative to European samples (International HapMap Consortium, 2005; www.hapmap.org). Research using samples in which diversity of ancestry and gender are not confounded is needed to better understand whether or how the relation between *ADCYAP1R1* and PTSD may be affected by ancestry or sex.

Although our analyses did not support the presence of a “file drawer” problem, it is impossible to completely exclude the possibility that publication bias may exist and that studies that have not detected a significant association between *ADCYAP1R1* and PTSD are unpublished. Another important limitation is that, due to characteristics of individuals within included study samples, the effects of key moderators may have been confounded with one another. For example, there is reason to believe that there may be differences as a function of sex, ancestry, and trauma type. However, the samples that are disproportionately female also report African American ancestry and report high levels of exposures to early life and more recent urban traumas whereas the samples with greater numbers of male participants have fewer African Americans and are more likely to report military trauma. There are also fewer studies that have examined men and boys, emphasizing the need for additional work to parse out sex differences. Further, one limitation of meta-analysis is reduced power for detecting variability across moderators (Hedges & Pigott, 2004), particularly for analyses including a small number of studies (e.g., $k = 4$). Given this, we were unable to test ancestry as a moderator within the male-only subsample. Thus, this line of research would benefit from replication in larger samples of varied trauma exposures and ancestries, which may allow for disentanglement of potential moderating effects. The focus on current PTSD (vs. lifetime, given the use of current by the majority of studies) represents an additional limitation, given that individuals at genetic risk may not have PTSD when assessed (Koenen et al., 2009). However, current PTSD was more commonly assessed in the extant literature, and thus, analyses of lifetime PTSD were not possible.

Another consideration and direction for future PACAP research is the potential that expression differences may be related not only to genotype, but also to epigenetic changes such as DNA methylation (Dias & Ressler, 2013). Prior research has shown differential methylation patterns in the promoter region of *ADCYAP1R1* (cg11218385) as a function of

violence exposure (Chen et al., 2013). Interestingly, in Chen et al.'s (2013) sample of children, both *ADCYAP1R1* rs2267735 genotype *and* methylation status were associated with increased odds of asthma. In addition, violence exposure was associated with methylation among children over 9 years of age. Research has also implicated PACAP in immune modulation and inhibition of inflammation (Abad & Var Tan, 2016). Indeed, there is evidence for PACAP's influence on Th2 cells, leading some to suggest PACAP as a possible therapeutic agent with respect to Th1:Th2 T-cell phenotypes (Abad & Var Tan, 2016; Delgado & Ganea, 2001; Delgado, Leceta, & Ganea, 2002). These findings are relevant to PTSD, as research has repeatedly shown evidence for immune alterations in individuals with PTSD (Gill, Saligan, Woods, & Page, 2009).

We present, to our knowledge, results of the first meta-analysis examining the role of the *ADCYAP1R1* gene, specifically rs2267735, in PTSD. Although not without limitations, our results indicate that the 'C' allele of rs2267735 may increase PTSD risk, particularly for female participants. This provides support for closer examination of the PACAP genes and their relevance to stress response in the aftermath of trauma on risk for PTSD.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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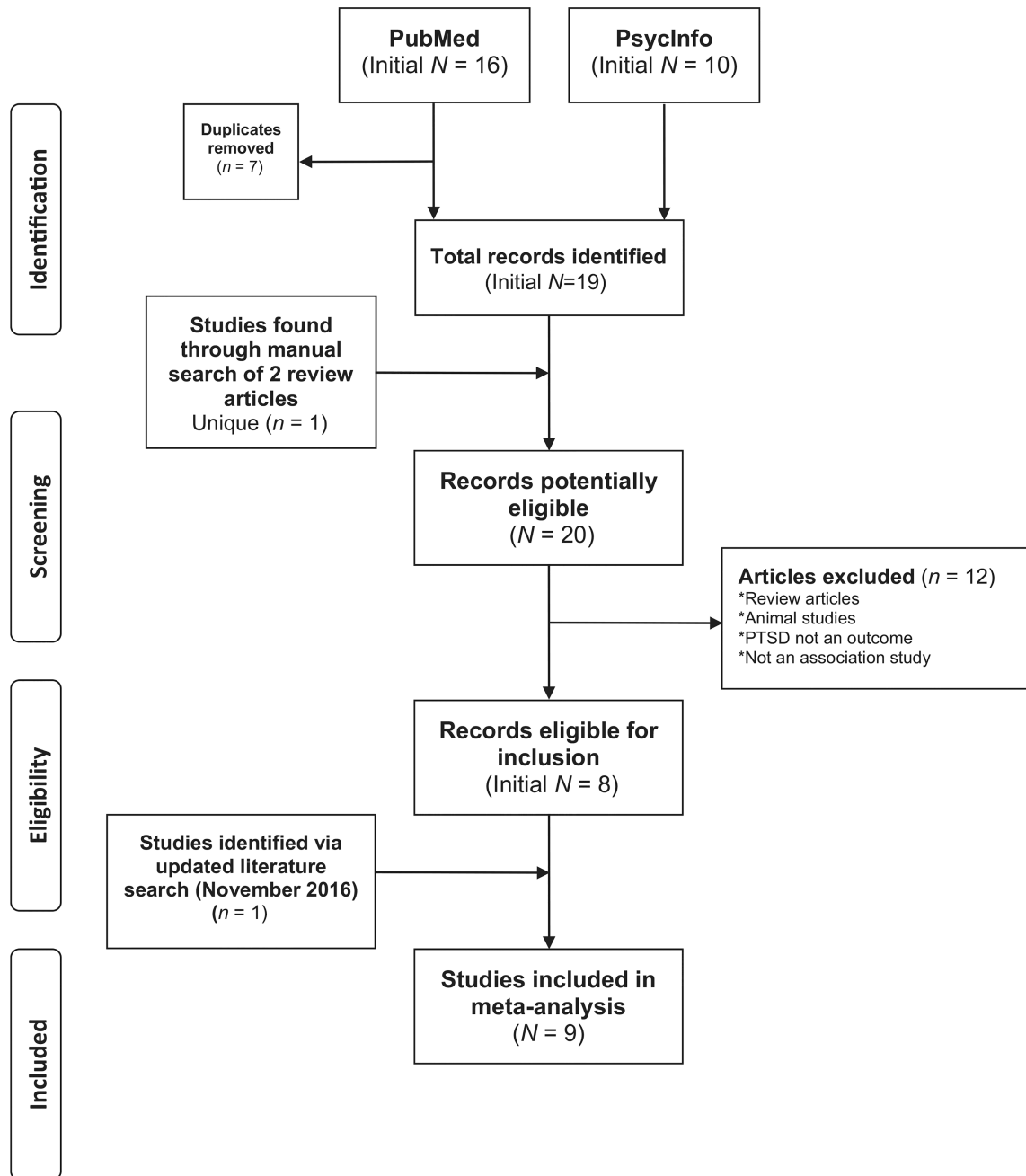


Figure 1. PRISMA-style flow chart showing selection of studies for meta-analysis of rs2267735 (*ADCYAP1R1* gene) and posttraumatic stress disorder (PTSD). Figure is consistent with guidelines provided by Moher, Liberati, Tetzlaff, Altman, and the PRISMA Group (2009).

Table 1

Included Studies and Outcome Measures (Combined Sexes)

Study	Outcome	Sample recruitment	PTSD measure	Trauma assessment		N	%AA	%F	OR	95% CI
				General	Child					
Almli et al., 2013	Current PTSD severity	Outpatient clinic-based ^a	PSS	TEI ^d	CTQ	1,163	100.0	73.8	0.84	[0.28, 2.48]
Chang et al., 2012	Current PTSD diagnosis	Population-based	Structured interview ^c	Other ^e	CTQ	2,528	0	100.0	0.94	[0.64, 1.39]
Lowe et al., 2015	Current PTSD severity	Outpatient clinic-based ^a	PSS	TEI	NA	1,361	94.1	100.0	2.10	[1.04, 4.24]
Ressler et al., 2011a	Current PTSD diagnosis	Outpatient clinic-based ^a	PSS	TEI	CTQ	798	92.3	63.0	1.43	[1.14, 1.80]
Ressler et al., 2011b	Current PTSD diagnosis	Outpatient clinic-based ^a	PSS	TEI	CTQ	439	92.4	59.2	1.20	[0.89, 1.63]
Rothbaum et al., 2014	Current PTSD severity	Outpatient clinic-based ^a	PSS	TEI	CTQ	34	82.4	61.5	2.45	[0.57, 10.53]
Solovieff et al., 2014	Lifetime PTSD diagnosis	Population-based	Structured interview ^c	Other	NA	2,538	0	100.0	0.96	[0.85, 1.08]
Stevens et al., 2014	Current PTSD severity	Hospital admission-based ^b	PSS	TEI	CTQ	49	100.0	100.0	1.29	[0.46, 3.58]
Uddin et al., 2013	Current PTSD diagnosis	Population-based	Structured interview ^c	Other	Other ^e	401	83.1	100.0	1.65	[0.71, 3.84]
Wang et al., 2013	Current PTSD severity	Epidemiologic exposure-based	PCL	NR	NR	319	0	70.5	1.39	[0.85, 2.26]

Note. All samples were primarily civilian and most did not examine a specific traumatic event (with the exception of Wang et al., 2013, which used an earthquake sample). PTSD = posttraumatic stress disorder; AA = African American; PSS = Posttraumatic Symptom Scale; PCL = Posttraumatic Stress Disorder Checklist; CTQ = Childhood Trauma Questionnaire; NA = not applicable; NR = not reported; TEI = Traumatic Events Inventory.

^aMedical or nonmedical, includes Veteran Affairs.

^bIntensive care unit/emergency department/inpatient.

^cDiagnostic measure used was unspecified structured diagnostic interview, based on criteria according to the *Diagnostic and Statistical Manual of Mental Disorders* (4th ed., *DSM-IV*; American Psychiatric Association, 1994).

^dPresented results both adjusted and unadjusted for trauma, unadjusted used in meta-analysis.

^eOnly presented trauma-adjusted results.

Table 2

Included Studies and Outcome Measures (Separated by Sex)

Study	Female participants				Male participants			
	N	% AA	OR	95% CI	N	% AA	OR	95% CI
Almli et al., 2013	858	100.0	0.72	[0.20, 2.50]	305	100.0	1.35	[0.16, 11.60]
Chang et al., 2012	2,528	0	0.94	[0.64, 1.39]	-	-	-	-
Lowe et al., 2015	1,361	94.1	2.10	[1.04, 4.24]	-	-	-	-
Ressler et al., 2011a	503	NR	1.72	[1.29, 2.30]	295	NR	1.04	[0.71, 1.52]
Ressler et al., 2011b	260	NR	1.57	[1.05, 2.34]	179	NR	0.84	[0.52, 1.33]
Rothbaum et al., 2014	-	-	-	-	-	-	-	-
Solovieff et al., 2014	2,538	0	0.96	[0.85, 1.08]	-	-	-	-
Stevens et al., 2014	49	100.0	1.29	[0.46, 3.58]	-	-	-	-
Uddin et al., 2013	401	83.1	1.65	[0.71, 3.84]	-	-	-	-
Wang et al., 2013	225	0	1.63	[0.80, 2.98]	94	0	1.01	[0.44, 2.32]

Note. AA = African American.

Table 3

Summary of All Meta-Analytic Results

Sample	n	k	Model	OR	95% CI	p	Heterogeneity	
							Q	I ²
Combined sexes	9,630	10	Fixed	1.085	[0.990, 1.190]	.082	17.279*	47.913
			Random	1.210	[1.007, 1.454]	.042		
Female participants	8,723	9	Fixed	1.098	[0.995, 1.211]	.062	24.168*	66.898
			Random	1.328	[1.026, 1.719]	.031		
Male participants	873	4	Fixed	0.964	[0.733, 1.269]	.796	0.610	0
			Random	0.964	[0.733, 1.269]	.796		

Note. NR = not reported.

* $p < .05$.