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Dietary Inflammatory Index and Ovarian Cancer Risk in a New Jersey Case-Control Study

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Abstract

Purpose—Diet may influence the development of ovarian cancer. While inflammation has been shown to play an important etiologic role on ovarian carcinogenesis, little is known about the influence of the inflammatory potential of food consumption.

Methods—Data from a case-control study conducted in New Jersey (USA) were used to estimate the relation between a dietary inflammatory index (DII) and the risk of ovarian cancer. The study consisted of 205 cases with incident, histologically confirmed ovarian cancer, and 390 controls identified by random digit dialing, based on CMS (Centers for Medicare and Medicaid Service) lists, and area sampling. Computation of the DII was based on the intake of selected dietary factors assessed by a validated Food Frequency Questionnaire (FFQ). Logistic regression models were fit to estimate odds ratios (ORs) and 95% confidence intervals (CI) adjusted for potential covariates.

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Disclosure Dr. James R. Hébert owns controlling interest in Connecting Health Innovations LLC (CHI), a company planning to license the right to his invention of the dietary inflammatory index (DII) from the University of South Carolina in order to develop computer and smart phone applications for patient counseling and dietary intervention in clinical settings. Dr. Nitin Shivappa is an employee of CHI.

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Results—Although there was no significant association observed in pre and peri-menopausal women, a significant association was observed between the most pro-inflammatory DII scores and ovarian cancer among post-menopausal women ($OR_{Quartile4vs1}=1.89$, 95 % CI, 1.02–3.52; $P_{trend}=0.03$).

Conclusion—Our finding suggests that a pro-inflammatory diet may increase ovarian cancer risk among post-menopausal women, and warrants further study to confirm this association.

Keywords

diet; inflammatory index; ovarian cancer

INTRODUCTION

Among gynecological cancers ovarian cancer has the highest mortality rates, with dismal five-year survival rates (46% for all stages combined; 28% for advanced stage, in which 62% of the cases are diagnosed) [1]. The American Cancer Society estimates 22,440 new cases and 14,080 deaths from ovarian cancer in the United States in 2017 [1]. Risk factors for ovarian cancer include increasing age, family history of the disease (specifically mutations in BRCA1 and BRCA2 genes), obesity and nulliparity, while oral contraceptive use, higher parity, and tubal ligation have been shown to reduce risk [2,3]. Several studies have been conducted exploring the association between dietary factors and ovarian cancer with inconsistent results [4]. While there is growing evidence linking inflammation to ovarian carcinogenesis [5,6], to date there have been only two studies that have explored the role that inflammatory potential of diet plays in ovarian cancer risk [7,8]; one of them was conducted exclusively among African American women [7] and the other was among Italian women [8].

A literature-derived, population-based dietary inflammatory index (DII) was recently developed to assess the inflammatory potential of an individual's diet [9]. A proinflammatory diet is high in foods rich in saturated fat and carbohydrates, and low in foods rich in poly-unsaturated fatty acids, flavonoids, and other dietary components, include a variety of vitamins and minerals [10]. The DII has been validated in a variety of longitudinal and cross-sectional studies with various inflammatory markers, including C-reactive protein [11,12], interleukin-6 [13,14], and tumor necrosis factor- α [14]. The DII has been associated with risk of colorectal cancer in case-control studies in Spain and Italy [15,16] and in 3 cohort studies in the USA [10,17,18], and risk of pancreatic, prostate and endometrial cancers in case-control studies in Italy [19–22]. In this study we evaluate the impact of a pro-inflammatory diet, as indicated by a high DII on ovarian cancer risk in a New Jersey population.

METHODS

We evaluated the association between DII and ovarian cancer in the *NJ Ovarian Cancer Study*, described in detail elsewhere [23–26]. In brief, our study included 205 newly diagnosed, histologically confirmed cases of invasive epithelial ovarian cancer identified through rapid case ascertainment implemented by the New Jersey State Cancer Registry

(NJSCR) staff. Women older than 21 years, able to understand English or Spanish, and residing in one of six New Jersey counties (Bergen, Essex, Hudson, Middlesex, Morris, and Union) were eligible to participate. Controls (n=390) had the same eligibility criteria as the cases except that women with a history of hysterectomy and/or bilateral oophorectomy were excluded from the analysis. Controls were identified through random digit dialing for women <65 years of age and through random selection of Center for Medicare and Medicaid Services lists, complemented with area sampling for women 65 years of age.

After obtaining informed consent, a telephone interview was scheduled, during which information was collected on established and suspected risk factors for ovarian cancer as well as on demographic characteristics. Dietary data were collected using the Block 98.2 Food Frequency Questionnaire (FFQ), which included questions about usual intake during six months before diagnosis for cases or on the date of interview for controls. The Block 98.2 FFQ (NutritionQuest, Berkeley, CA) was developed from the National Health and Nutrition Examination Survey III dietary recall data includes 110 food and beverage items and queries on frequency and portion size for each item. Pictures were provided to enhance accuracy of estimation of portion size.

Dietary Inflammatory Index (DII)

FFQ-derived dietary data were used to calculate the DII for each subject. A complete description of the DII is available elsewhere [9]. Briefly, dietary data were first linked to a regionally representative global database that provided a robust estimate of the mean and the standard deviation for each food parameter included in the DII. These parameters then became the multipliers to express an individual's exposure relative to the "standard global mean" as a z-score. This was achieved by subtracting the "standard global mean" from the amount reported and dividing this value by the standard deviation. To minimize the effect of "right skewing," this value was then converted to a centered (on zero) percentile score (by taking the percentile ranking of the z-score, multiplying by 2 and subtracting 1). The centered percentile score for each food parameter for each subject was then multiplied by the corresponding food parameter effect score in order to obtain a food parameter-specific DII score. All of the food parameter-specific DII scores were then summed to create the overall DII score for each subject. The DII was calculated from foods and supplements. To control for total energy intake, the DII was calculated per 1,000 calories of food consumed, which requires using the energy-standardized version of the global database. This study had data on 29 of the 45 food parameters studied for DII development; food parameters that are available and that are unavailable in this study are shown in Appendix 1. Steps involved in calculating the DII score are described in Figure 1.

Statistical Analyses

DII scores were analyzed by quartiles of exposure in controls. Body mass index (BMI) was calculated as weight (in kg) divided by height (in meters) squared and was categorized as: underweight ($<18.5 \text{ kg/m}^2$); normal weight (BMI 18.5– 24.9 kg/m²); overweight (25.0 kg/m² BMI $<30.0 \text{ kg/m}^2$); and obese (BMI 30.0 kg/m^2). Age-adjusted means were calculated for cases and controls for pro-inflammatory food parameters (protein, saturated fat, cholesterol and carbohydrates) and anti-inflammatory food parameters (vitamin B1,

Niacin, Folate, Vitamin C and dietary fiber) and compared using analysis of covariance. Odds ratios (OR) and the corresponding 95% confidence intervals (95% CI) were estimated using logistic regression models, adjusting only for age as a continuous variable and additionally adjusting for education, race, age at menarche, menopausal status, parity, oral contraceptive use, hormone therapy use, tubal ligation, BMI categories, physical activity (in metabolic equivalents (or METs) for reported average hours per week of strenuous or moderate recreational activities), and smoking status. Effect modification by menopausal status and BMI categories was evaluated. Testing for heterogeneity was carried out by including the interaction terms in the model. Tests for trend were computed by assigning the median value to each quartile. All analyses were completed using SAS® version 9.3 (SAS Institute, Cary NC).

RESULTS

Participants in the *New Jersey Ovarian Cancer Study* were primarily White (87.3% of cases and 88.4% of controls) and approximately 25% of cases and controls had a graduate school education [23] (data not shown). Mean values of selected pro- and anti-inflammatory food parameters for cases and controls are shown in Table 1. Mean DII value among cases was $1.1 \text{ (SD}=\pm0.2)$ and among controls was $0.8 \text{ (SD}=\pm0.1)$ indicating a slightly more proinflammatory diet for cases (p=0.18). For pro-inflammatory food parameters, cases had slightly higher intakes of saturated fat and carbohydrates and for anti-inflammatory food parameters, cases had significantly lower intakes of niacin and slightly lower levels of vitamin B1, folate, vitamin C and dietary fiber compared to controls.

OR and 95% CI of ovarian cancer according to quartiles of DII are shown in Table 2. No significant associations were observed between DII and overall ovarian cancer (i.e., across all ages). In age-adjusted models results suggestive of a positive association were observed for DII with ovarian cancer ($OR_{Quartile4vs\ 1} = 1.38$, CI= 0.85-2.26, $P_{trend}=0.27$). Similarly, for multivariable analyses, suggestive positive associations were observed, with $OR_{Quartile4vs\ 1}$ of 1.39 (95% CI=0.82-2.35, $P_{trend}=0.26$). When stratified by menopausal status, a significant association was observed among post-menopausal women consuming the most pro-inflammatory diet ($OR_{Quartile4vs\ 1}=1.89$, 95 % CI, 1.02–3.52; $P_{trend}=0.03$) (Table 3). P-value for interaction was nearly significant with menopausal status (P-value=0.08). In analyses stratified by BMI, the association appeared to be stronger in overweight and obese women, but the confidence interval for both the categories included the null (p for interaction=0.46).

DISCUSSION

In this case-control study conducted in New Jersey, we found some evidence of elevated risk associated with higher DII only among postmenopausal women. No association with ovarian cancer was found in earlier reports in the same case-control study with the Healthy Eating Index or with total antioxidant capacity [24,25], while selenium from food sources reduced the risk [25] and there was suggestion of decreased risk with increased phytoestrogen consumption [23]. Results from other studies exploring dietary components that contribute to the DII score and ovarian cancer have been inconsistent. In the NIH-AARP cohort study,

sugar consumption was inversely associated with ovarian cancer [27], whereas no association was observed with sugar in this NJ case-control study [26] and in a cohort study conducted in Canada, glycemic index and carbohydrate were not associated while glycemic load increased risk of ovarian cancer [28]. In an Italian multicenter case-control study, fiber intake was associated with reduced the risk [29]. No association was observed with dietary phytoestrogens in two Australian case-control studies [30]. In relation to the DII, fiber has an anti-inflammatory effect score while simple carbohydrates have a pro-inflammatory effect score [9]. Though phytoestrogens, especially flavonoids have anti-inflammatory scores, data on flavonoids were not available in this study; hence, they could not be used for DII calculation. The DII has been shown to be associated with ovarian cancer in one study in Italy; subjects in the highest quartile of DII scores (i.e., with the most pro-inflammatory diets) had a higher risk of ovarian cancer compared to subjects in the lowest quartile (i.e., with an anti-inflammatory diet) (ORQuartile4vs1 1.47, 95% confidence interval, CI, 1.07, 2.01; p trend = 0.009) [8]. Similarly, in a study conducted in the US African-American women consuming the most pro-inflammatory diet had a statistically significant increased ovarian cancer risk in comparison to the most anti-inflammatory diet (ORQuartile4/ Quartile1 3=31.72; 95% CI3=31.18-2.51) [7,8].

We did not observe significant association between DII and ovarian cancer among premenopausal women, similar results were seen in the previous two studies [7,8]. The absence of an association between DII scores and ovarian cancer among pre-menopausal women in this study could be explained by the fact that there are strong hormonal and reproductive factors which play a more important role in the development of ovarian cancer at younger ages when the ovaries are fully functional [31,32]. By contrast, inflammation may represent relatively more important influences in in post-menopausal women. Furthermore, the pre-menopausal group may have a different type of ovarian cancer that has developed secondary to germline alterations independent of any dietary factors. For example, women with germline BRCA1/2 deleterious mutations tend to develop ovarian and other cancers at an earlier age and thus are more likely to be pre-menopausal [33]. In contrast, the postmenopausal group may develop cancer as a result of somatic mutations that happen over time and as a response to environmental factors such as exposure to an inflammatory diet.

Certain limitations of this study should be noted. Our sample size was relatively small, which may have affected our statistical power to detect associations. Additionally, the study was subjected to the limitations of case—control studies, such as recall and selection biases. However, the distribution of risk factors such as parity, tubal ligation, and oral contraceptive use of cases and controls in this study [23], is similar to that reported in other studies [3]which gives us reassurance in the validity of our data. Another limitation is the use of the FFQ, which may lead to measurement error, even in healthy individuals [34,35] and may be associated with disease- differential reporting biases [36,37]. With respect to the DII, no information was available on 16 16 food parameters. DII calculated from the 29 available food parameters has not been validated with inflammatory markers, though we have found little drop off in predictability in other studies, such as the SEASONS Study [10] and the Women's Health Initiative [17], which used essentially the same FFQ as in this study.

In conclusion, our study provided suggestive evidence that a pro-inflammatory diet, as shown by higher DII scores, increased risk of ovarian cancer in postmenopausal women. However, this finding requires replication in larger studies, including prospective cohorts, which may provide more definite evidence regarding the possible role of diet-related inflammation on ovarian cancer etiology and possible effect modification by menopausal status and body mass index.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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Highlights

- Ovarian cancer has been linked to chronic inflammation and diet. Yet, the impact of an inflammatory diet on ovarian cancer risk is unclear.
- In this study, we assessed the association between dietary inflammation and risk for ovarian cancer.
- Proinflammatory diets (as indicated by dietary scores) are associated with increased ovarian cancer risk among post-menopausal women.

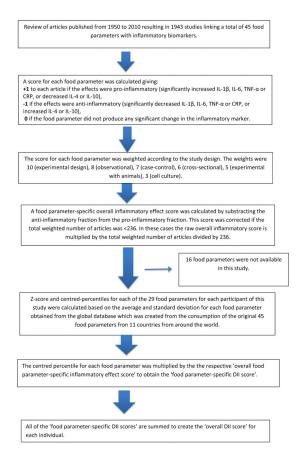


Figure 1.Sequence of steps in creating the dietary inflammatory index in the New Jersey Ovarian cancer case-control study

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Table 1

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Means for cases and controls for the Dietary	Inflammatory	Index (DII)) and some of its cor	nponents

Variable ^a	Cases (n=205)	Controls (n=390)	P-value
	Mean (SE)	Mean (SE)	
DII	1.1 (0.2)	0.8 (0.1)	0.18
Pro-inflammatory	food parameters	$_{3}b$	
Protein (g)	41.3 (2.23)	42.5 (3.5)	0.27
Saturated fat (g)	13.2 (1.82)	12.9 (1.34)	0.72
Cholesterol (mg)	114.9 (3.6)	115.0 (2.6)	0.98
Carbohydrates(g)	122.6 (1.7)	119.8 (1.2)	0.17
Anti-inflammator	y food parameter	$\mathbf{s}^{\mathcal{C}}$	
Vitamin B1 (mg)	0.78 (0.01)	0.80 (0.01)	0.26
Niacin (mg)	10.8 (0.20)	11.5 (0.14)	0.006
Folate (mcg)	221.5 (4.5)	222.5 (3.2)	0.86
Vitamin C (mg)	77.8 (3.0)	78.9 (2.1)	0.76
Dietary fiber (g)	9.9 (0.30)	10.4 (0.2)	0.16

 $^{^{}a}$ Density measure calculated as daily intake in respective units per 1,000 kcal

 $^{^{}b}$ As indicated by the positive inflammatory effect scores in the DII development manuscript (38)

^cAs indicated by the negative inflammatory effect scores in the DII development manuscript (38)

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Table 2

Odds ratios (OR) of ovarian cancer and corresponding 95% confidence intervals (CI) for quartiles of DII among 205 cases and 390 controls. New Jersey, 2004-2008.

		Id	DII quartiles		ء ا
	<-1.28	< -1.28 -1.28-0.68	0.69–2.24	>2.24	• trend
Cases/Controls	53/98	44/97	41/98	<i>L6/L9</i>	
Model 1b	1 a	0.84 (0.50, 1.41)	0.84 (0.50, 1.41) 0.71 (0.42, 1.20)	1.38 (0.85, 2.26)	0.27
Model $2^{\mathcal{C}}$	1 a	0.86 (0.49, 1.49)	0.86 (0.49, 1.49) 0.83 (0.47, 1.46) 1.39 (0.82, 2.35)	1.39 (0.82, 2.35)	0.26

 a Reference category.

bModel 1 adjusted for age

menopausal, postmenopausab), parity (0-1, 2, 3-4), oral contraceptive use (ever, never), HT use (never, unopposed estrogen only, any combined HT), tubal ligation (no, yes), BMI (continuous), smoking Model 2 adjusted for age (continuous), education (high school or less, college, graduate school), race (White, Black, Other, Hispanic), age at menarche (continuous), menopausal status (pre-/peristatus (never, past, current).

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Table 3

Odds ratios (OR) of ovarian cancer and corresponding 95% confidence intervals (CI) for quartiles of DII among 205 cases and 390 controls. New Jersey, 2004-2008.

	Cases/Controls		II	DII quartiles ^a		م	م م
		<-1.28	<-1.28 -1.28,0.68	0.69,2.24	>2.24	r trend	• interaction
Menopausal status	status						
Pre/Peri	71/49	$_1b$	0.62 (0.14, 2.67)	$0.62\ (0.14,2.67)$ $0.12\ (0.03,0.58)$ $0.74\ (0.19,2.84)$ 0.41	0.74 (0.19, 2.84)	0.41	
Post-	134/338	$_1b$	0.91 (0.47, 1.75)	$0.91 \ (0.47, 1.75) 1.30 \ (0.68, 2.52) 1.89 \ (1.02, 3.52) 0.03$	1.89 (1.02, 3.52)	0.03	0.08
$BMI\left(kg/m^{2}\right)$							
<25	91/180	1b	1.29 (0.60, 2.81)	$1.29 \ (0.60, 2.81) 1.11 \ (0.47, 2.60) 1.29 \ (0.58, 2.86) 0.60$	1.29 (0.58, 2.86)	09.0	
25	112/203	1^{b}	0.57 (0.24, 1.31)	0.57 (0.24, 1.31) 0.61 (0.27, 1.37) 1.60 (0.75, 1.37) 0.19	1.60 (0.75, 1.37)	0.19	0.46

^aModel adjusted for age (continuous), education (high school or less, college, graduate school), race (White, Black, Other, Hispanic), age at menarche (continuous), menopausal status (premenopausal), parity (0–1, 2, 3–4), oral contraceptive use (ever, never), HT use (never, unopposed estrogen only, any combined HT), tubal ligation (no, yes), BMI (continuous), smoking status (never, never) past, current).

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 $^{^{}b}$ Reference category.