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Methadone versus buprenorphine for opioid use dependence and risk of neonatal abstinence syndrome

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Abstract

Background—Our objective was to estimate the association between methadone and neonatal abstinence syndrome compared with buprenorphine using a probabilistic bias analysis to account for unmeasured confounding by severity of addiction.

Methods—We used a cohort of live-born infants exposed in utero to methadone or buprenorphine for maternal opioid maintenance therapy at Magee-Womens Hospital in Pittsburgh, PA from 2013–2015 (n=716). We determined exposure and outcome status using pharmacy billing claims. We used log-binomial regression models to assess association of treatment with neonatal abstinence syndrome after adjusting for parity, maternal race, age, delivery year, employment, hepatitis c, smoking, marital, and insurance status. We implemented probabilistic bias analysis, informed by an internal validation study, to assess the impact of unmeasured confounding by severity of addiction.

Results—Infants exposed to methadone in utero were more likely to experience neonatal abstinence syndrome compared with those exposed to buprenorphine [RR: 1.3, 95% CI: 1.2, 1.5]. After adjustment, infants exposed to methadone were more likely (adjusted RR 1.3, 95% CI: 1.1, 1.5) than infants exposed to buprenorphine to have the syndrome. In the validation cohort (n=200), severe addiction was more common in methadone- versus buprenorphine-exposed deliveries (77% vs. 32%). However, adjustment for severe addiction in the bias analysis only slightly attenuated the association (RR 1.2, 95% CI: 1.0, 1.4), supporting conventional analysis.

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The statistical code is available from Dr. Lemon upon request; authors must obtain permission from our Institutional Review Board to analyze the data directly

Conclusions—Methadone is associated with increased risk of neonatal abstinence syndrome compared with buprenorphine in infants exposed in utero. This association is subject to minimal bias due to unmeasured confounding by severity of addiction.

Keywords

opioid maintenance therapy; methadone; buprenorphine; pregnancy; probabilistic bias analysis

INTRODUCTION

Pregnant women are not immune from the opioid epidemic in the U.S.^{1–4} The trend of rising opioid use in pregnancy parallels simultaneous increases in the number of cases of neonatal abstinence syndrome. Neonatal abstinence syndrome is a clinical condition in which the infants exposed to opioids in utero manifest symptoms of withdrawal from the drug postnatally.^{5–7} Neonatal abstinence is costly to treat⁸ and it has long term sequelae for the child, including neurocognitive and behavioral issues⁹ along with decreases in visual acuity.¹⁰ To reduce the risk of neonatal abstinence syndrome and a host of other poor maternal and child health outcomes, pregnant women with opioid use dependence are treated with either methadone or buprenorphine as opioid maintenance therapy.¹¹ Literature has consistently shown that buprenorphine use is associated with lower risk of neonatal abstinence syndrome and shorter duration of neonatal treatment compared with the use of methadone.^{12–16} However, these findings may be biased because large databases often used for this research typically do not contain data on the severity of the mother’s addiction, a potential confounder.^{17, 18}

In the U.S., women who suffer from more severe opioid addiction are often assigned to methadone treatment, while women with lower risk of relapse and drug diversion tend to be treated with buprenorphine. This prescribing preference exists in part because methadone and buprenorphine are delivered with different systems of care in the U.S. Women prescribed methadone must attend a clinic daily to obtain medication under direct observation, eliminating the chance of illegal distribution. Alternatively, women treated with buprenorphine are legally permitted a supply of medication for administration at home through outpatient providers.¹⁹ Therefore, it is critical to account for factors that determine this prescribing preference in comparative treatment studies.

Our objective was to estimate the association between methadone versus buprenorphine exposure as opioid maintenance therapy and neonatal abstinence syndrome after accounting for unmeasured confounding by severity of addiction.

METHODS

We used data on all singleton pregnancies delivered at 20 to 42 weeks of gestation with live-born infants exposed to in utero methadone or buprenorphine opioid maintenance therapy at Magee-Womens Hospital (MWH) in Pittsburgh, PA from 2013–2015. MWH delivers over 10,000 infants annually and cares for opioid-addicted mothers with treatment protocols similar to those at other U.S. institutions.^{14, 15, 20, 21} Buprenorphine is administered through

prescription by a certified buprenorphine provider while methadone treatment requires daily visits to an opioid treatment clinic.²² The protocol is described in detail in the eAppendix 1.

International Classification of Diseases

Ninth (ICD-9) and Tenth Revision (ICD-10) codes in pharmacy billing claims were used to identify drug-dependent (ICD-9 64831) or drug-complicated deliveries (ICD-10 O99324). Billing claims that specifically documented exposure to methadone or buprenorphine as opioid maintenance therapy were then confirmed with dosing information extracted from the medical chart. Buprenorphine-exposed infants were those whose mothers were treated with Subutex® (buprenorphine, n=299) (Reckitt Benckiser Pharmaceuticals Inc., VA) or Suboxone® (buprenorphine + naloxone, n=10) (Reckitt Benckiser Pharmaceuticals Inc., VA). The exposure window of interest was the day of delivery because medication effect on neonatal abstinence syndrome is most influential closest to delivery²³ and we lacked access to the entire treatment trajectories including treatment initiation dates.

We identified cases of neonatal abstinence syndrome from pharmacy billing codes indicating treatment with morphine after delivery. At MWH, all infants with known or suspected opioid exposure in utero are kept for neonatal abstinence syndrome observation for 5 to 7 days. Infants are scored using the Finnegan Scale every 3 to 4 hours.²⁴ When the average of 3 consecutive scores is ≥ 8 on the Finnegan Scale, infants are given morphine treatment. In our cohort, morphine treatment was highly correlated with ICD code indicative of “Drug Withdrawal Syndrome in Newborn” ($\kappa > 0.99$).

Maternal characteristics and birth outcomes were obtained first from the MWH electronic pharmacy records comprised primarily of billing and ICD codes, and were informed with data provided by the birth record when data were missing. These data are a combination of self-report, clinical billing codes, and chart documentation by a health professional. Information on maternal race (Black, White, other), education level (less than high school, high school or equivalent, some college, college graduate), employment (yes, no), marital status (married, unmarried), insurance type (private, public), pre-pregnancy weight and height, parity, smoking during pregnancy (yes, no), and hepatitis c status (positive, negative) were available. Body mass index (BMI) was calculated as weight (kilograms) divided by height (meters) squared and was categorized as underweight (< 18.5), normal weight (BMI 18.5 to < 25), overweight (BMI 25 to < 30), or obese (BMI ≥ 30).²⁵ Birth outcome data included gestational age at delivery, infant length of stay (days), birthweight, congenital anomalies (yes, no), admission into the neonatal intensive care unit (yes, no), and number of prenatal visits. Gestational age was determined using the best obstetric estimate in the chart from ultrasound or last menstrual period when ultrasound was not available. This study was approved by our Institutional Review Board.

Validation Cohort

Severity of addiction is a potential confounder that was unmeasured in our dataset. We therefore performed a validation study to collect indicators of addiction severity from medical chart abstraction on a random sample of 100 buprenorphine- and 100 methadone-treated women in our cohort. The study team identified four indicators of severity of

addiction that were based on literature^{17, 18, 26, 27} and clinical expertise (details in eAppendix 2). One reviewer (LSL), who was blinded to outcome but not exposure status, performed the medical chart abstractions and entered data into an electronic database. The majority of information was abstracted from physicians' notes and the social workers' discharge plans.

We defined severe addiction as having any one of the four following indicators documented in the chart: 1) conversion to opioid maintenance therapy during pregnancy, 2) documented relapse during pregnancy, 3) use of illicit substances at delivery, and 4) use of benzodiazepines in pregnancy. When there was no documentation of conversion to opioid maintenance therapy in the chart, women were assumed to have conceived on the same treatment noted at delivery. All other lack of documentation was recorded as missing unless explicitly noted that the patient did not have the indicator (e.g. "patient did not relapse in this pregnancy"). Reconversion to therapy within one pregnancy was recorded as a relapse. Illicit substance use at the time of delivery included any of the following: marijuana, benzodiazepines, illicit buprenorphine, cocaine, nondescript intravenous drugs, heroin, or illicit opiate pills.

Statistical Analysis

Multivariable log-binomial regression models were used to estimate the independent association between neonatal abstinence syndrome and methadone compared with buprenorphine while accounting for clustering within each woman (25 who contributed multiple pregnancies).²⁸ We calculated risk ratios (RR), risk differences (RD), and their 95% confidence intervals (CI). Risk differences were calculated using marginal standardization.²⁹ We identified potential confounders *a priori* using theory-based conceptual models: maternal indication for opioid maintenance therapy, gestational age at opioid maintenance therapy initiation, duration of opioid dependence, maternal age, race, employment status, smoking status, marital status, insurance type, hepatitis c status, parity and year of delivery. The final model was limited to maternal age, race, employment status, smoking status, marital status, insurance type, hepatitis c status, parity and year of delivery, based on availability of data. We did not adjust for adequacy of prenatal care, total visits to the emergency room during the pregnancy, and gestational age because they are likely on the causal pathway.¹⁷

Probabilistic Bias Analysis—To quantify the extent to which unmeasured confounding by severity of addiction biased the association between opioid maintenance therapy and neonatal abstinence syndrome, we performed a probabilistic bias analysis. This approach is based on a set of methods developed and described in detail previously by Lash et al.^{31, 32} The parameters for this analysis were informed using data indicative of addiction severity from our internal validation study. We defined the limits of the relative risk due to confounding using the Flanders and Khoury method³³. This method involved fitting two logistic regressions in the subcohort: the first modeling the odds of treatment type by severity of addiction, the second modeling the odds of neonatal abstinence syndrome by severity of addiction. The Flanders and Khoury method also incorporates the prevalence of severity in each treatment group (eAppendix 3). This information was used to determine the limits of the trapezoidal distribution used to parameterize the risk. We sampled the risk due

to confounding from 100,000 simulated data sets using a Monte Carlo approach. Results were presented as bootstrapped point estimates with an interval defined as the 2.5th and 97.5th percentiles. This interval corresponds to the 95% confidence interval obtained in a conventional analysis but incorporates both systematic and random error. The results from the probabilistic bias analysis were then compared with the risk ratios and 95% confidence intervals from the conventional model.

RESULTS

There were a total of 872 drug-dependent pregnancies in the study period. Of these, 745 (85%) received either methadone or buprenorphine as opioid maintenance therapy on the day of delivery and were eligible for this study (Figure 1). We excluded nine women with multi-fetal gestations (18 infants), six with a fetal death, and five who had stopped all medication due to relapse or weaning prior to delivery. Our final sample consisted of 716 pregnancies.

Slightly more than half of pregnancies on opioid maintenance therapy were treated with methadone (57%) and the remaining with buprenorphine (43%). Women treated with methadone were more likely than their buprenorphine-treated counterparts to be unmarried, unemployed, hepatitis c positive, multiparous, and to have less than a high school education (Table 1). Methadone-treated pregnancies on average had shorter gestations and infants with lower birthweights. Race, age, pre-pregnancy BMI, and smoking status were not meaningfully different between treatment groups. These relationships were consistent in the validation cohort (Table 2).

Neonatal abstinence syndrome occurred in 58% of the infants (n=415). Infants with treatment for the syndrome were more likely to be born to unmarried, unemployed, hepatitis c positive mothers with less than a high school education and a normal pre-pregnancy BMI (eTable 1). Infants diagnosed with neonatal abstinence syndrome were also more likely to be born at a later gestational age without a congenital anomaly compared with their counterparts not requiring treatment.

The incidence of neonatal abstinence syndrome was 65% in infants exposed in utero to methadone compared with 49% in infants exposed to buprenorphine. Infants exposed to methadone in utero were 30% more likely than infants exposed to buprenorphine to be treated for neonatal abstinence syndrome (unadjusted RR 1.3, 95% CI: 1.2, 1.5). After adjustment for parity, maternal race, employment status, hepatitis c status, age, year of delivery, smoking status, marital status, and insurance, the association did not change (adjusted RR 1.3, 95% CI: 1.1, 1.5). On the absolute scale, the adjusted RD was 0.14 (95% CI: 0.059, 0.22), indicating that methadone was associated with 14 excess cases of NAS for every 100 live-born infants born to mothers treated with methadone compared with buprenorphine (Table 3).

Though there were expected differences comparing the study cohort of opioid dependent mothers to all births at MWH from 2013–2014 (eTable 2), the validation subsample was similar to the full study cohort (eTable 3). In the validation subsample, methadone-treated

women were more likely than buprenorphine-treated women to have converted to opioid maintenance treatment during pregnancy (58% vs 12%; median gestational age at conversion: 12 weeks vs before conception), relapsed in pregnancy (23% vs 4%), used any illicit substance at delivery (24% vs 15%), or used benzodiazepines during pregnancy (28% vs 8%) (Table 4). Prevalence of having any one of the indicators of severe addiction was higher in the methadone group compared with buprenorphine (77% vs. 32%). This composite of addiction severity was associated with a slightly higher risk of neonatal abstinence syndrome (odds ratio 1.2, 95% CI: 0.7, 2.1). These results were robust to removal of each individual factor included in the severity index (data not shown).

There was a large amount of missing data in the validation cohort that varied by treatment (eTable 4). Women treated with buprenorphine were more likely than methadone-treated women to have missing data for more than one indicator of severity. Despite the difference in rate of missing data, women treated with buprenorphine also had documentation indicating less severity (e.g. “patient did not relapse in pregnancy”) more often than methadone-treated women. This is true for each severity indicator excluding benzodiazepine use (Table 4).

After accounting for unmeasured confounding by severity of addiction in the probabilistic bias analysis, the association between methadone and risk of neonatal abstinence syndrome was slightly attenuated from the conventional results [point estimate 1.2 (95% simulation interval: 1.0, 1.4; Table 5)]. The bootstrapped 5th and 95th percentiles in the bias analysis were slightly wider than the conventional confidence intervals as they accounted for both systematic and random error.

DISCUSSION

There is agreement in the literature that buprenorphine confers benefits over methadone for opioid maintenance therapy in pregnancy, including decreased risk of neonatal abstinence syndrome in the exposed infants.^{14–16, 18} Nonetheless, there is a potential for these findings to be biased due to unmeasured confounding.^{17, 18} Our conventional analysis results suggested that the risk of neonatal abstinence syndrome in infants exposed to in-utero methadone was 30 percent higher compared with buprenorphine-exposed infants. The results from the probabilistic bias analysis suggest that unmeasured confounding by severity of addiction only slightly biased the conventional results away from the null. Although we found that women receiving methadone had more indicators of severe addiction than women receiving buprenorphine, the relatively weak relationship between addiction severity and neonatal abstinence syndrome reduced the potential for prescribing differences to confound the primary association.

The ideal approach to eliminate unmeasured confounding is to conduct a randomized controlled trial. However, the largest double-blinded, flexible-dosing, randomized controlled trial comparing methadone and buprenorphine use in pregnancy (Maternal Opioid Treatment: Human Experimental Research trial) was plagued with the same biases faced in observational research.¹² Analyzing only women who remained on randomized treatment, Jones et al.¹² found no difference in percent of infants requiring treatment for NAS between

treatment groups, though more morphine (mean dose 10.4 vs. 1.1 mg) and longer hospital stays (17.5 vs. 10.0 days) were needed for infants exposed to methadone in utero. Importantly, investigators found that 33% of women randomized to buprenorphine discontinued treatment, with 71% of them reporting “dissatisfaction” with treatment. This is in stark contrast to only 18% of methadone patients discontinuing treatment, of whom only 13% reported “dissatisfaction” with treatment. Only those women who continued allocated treatment were included in the final analyses. Furthermore, despite randomization, women who remained on methadone treatment had longer cumulative lifetime drug use. Together, these findings demonstrate a similar bias to unmeasured confounding as addiction severity may have influenced treatment choice and continuation regardless of randomization.

Our results are consistent with a large meta-analysis of 11 studies including 855 methadone-treated women and 515 buprenorphine-treated women for opioid dependence and risk of neonatal abstinence syndrome.¹⁸ These authors described a summary estimate of 1.11 (95% CI: 1.02, 1.23) reported as an increased risk of neonatal abstinence syndrome by 10% in infants exposed to methadone compared with buprenorphine in utero. The authors conducted a sensitivity analysis for unmeasured confounding by indication applying the VanderWeele and Arah³⁴ approach for unmeasured confounding. Unlike our analysis, which was informed by an internal validation study, these authors used bias parameters informed by the extant literature. They found that after accounting for unmeasured confounding by indication, the risk of neonatal abstinence syndrome associated with methadone treatment in the conventional analysis was biased away from the null (50th percentile adjusted RR 1.01, 95% CI: 0.92, 1.11). Consistent with our conceptual model, bias parameters reflected values for unmeasured confounding that conferred increased risk for poor neonatal outcomes in the methadone-treated women [RR of confounder-NAS association (RR_{CD}) 1.05–1.25] that was reversed in the buprenorphine patients (RR_{CD} 0.80–0.95). Prevalence of unmeasured confounding by indication was assumed to be 40% in both treatment groups. Inputs for this bias analysis have been previously questioned as the assumptions informing these are subjective and results vary by slight changes in their inputs.³⁵ Our findings extend this work by using an internal validation study to inform the bias parameters and draw conclusions from one study center limiting heterogeneity in treatment practices. Using more conservative bias parameters informed from the validation cohort slightly weakened the impact of unmeasured confounding on our results by comparison.

In our probabilistic bias analysis, informed from the validation cohort, the RR for neonatal abstinence syndrome associated with methadone compared with buprenorphine marginally decreased from 1.3 (95% CI: 1.1, 1.5) to 1.2 (95% CI: 1.0, 1.4) when limits were defined by the Flanders and Khoury method.³³ We therefore maintain that the risk of neonatal abstinence syndrome associated with methadone treatment even after accounting for severity, may not be fully explained by unmeasured confounding.

It was surprising that accounting for severity of addiction did not further attenuate the association between methadone treatment and neonatal abstinence syndrome compared with buprenorphine. However, the impact of addiction severity on the association between opioid maintenance therapy and neonatal abstinence syndrome is likely limited by the weak relationship between addiction severity and the syndrome. Of note, infants born to women

actively abusing heroin during pregnancy have a lower risk of neonatal abstinence syndrome compared with those women receiving methadone as replacement therapy.^{16, 36} Therefore, behaviors associated with more severe addiction such as relapse and later conversion to opioid maintenance therapy may not increase the risk of neonatal abstinence syndrome. It is important to note that the lower risk of neonatal abstinence with active abuse does not negate other potential risks such as reduced prenatal care. Opioid maintenance therapy should remain the standard care in accordance with recommendations from the American Congress of Obstetricians and Gynecologists.^{11, 37}

Of note, the rate of congenital birth anomalies was relatively high in this cohort (11%). Though this variable has not been validated, nor can it differentiate minor from major defects, it did demonstrate a clustering which may be of interest for future study. An estimated 26% of all anomalies affected the heart or circulation, 22% were classified as genitourinary, and 14% were classified as orofacial anomalies. Furthermore, the relationship between prematurity and neonatal abstinence syndrome is deserving of future study. Infants born preterm (<37 weeks), had lower rates of the syndrome compared with those born at full term. Currently, we cannot be certain if the association between preterm delivery and neonatal abstinence syndrome is causal as the pathophysiology of is unknown. Though a commonly accepted plausibility is immature opioid receptor development in the neonate (and therefore decreased risk of dependence and subsequent neonatal abstinence syndrome), it could also be an artifact as the Finnegan Scale was developed to assess neonatal abstinence syndrome in term infants.

Our findings must be interpreted within the bounds of their limitations. We used a large administrative database that lacked detailed information on treatment and addiction histories. Without information on the initiation, timing, and duration of exposure to medication, we were unable to appropriately assess how these factors influence the development of neonatal abstinence syndrome. We relied on the dose and medication treatment on the day of delivery as a relatively crude measure of exposure, as it is thought that treatment closest to the time of delivery has the strongest impact on neonatal abstinence syndrome.²³ Though using this approach allows for misclassification of exposure, this is unlikely to affect our findings as only six of 200 women in our validation cohort had documentation of ever changing treatment (including prior to pregnancy). The lack of information on addiction history contributes considerably to the unmeasured confounding remaining in the analysis. Furthermore, by using treatment for neonatal abstinence syndrome as our outcome measurement, we restricted our analysis to only the more severe cases of the syndrome. Though having a gradient of Finnegan scores or morphine dose may be informative, those receiving treatment incur the largest costs and this approach is subject to less misclassification due to the subjectivity of the Finnegan Scale.

The lack of adjustment for prescribing preferences by severity of addiction, which is typically unmeasured, is one of the greatest shortcomings in the current literature. Our probabilistic bias analysis aimed to minimize this limitation using information from our internal validation cohort. As was expected due to the nature and sensitivity of this topic, upon chart review there was a substantial amount of missing data in the validation cohort with a missingness that differed by opioid maintenance type. Differential missingness was

likely driven by more buprenorphine-treated patients entering into pregnancy on treatment and potentially having less interaction with the healthcare system due to an overall superior health profile. Both may contribute to less documentation in their charts. Though to our knowledge we are the first to use an internal validation cohort to derive information on severity to adjust for unmeasured confounding, our findings are subject to the limitations of the data available to us and to the parameterization the severity index. Future research with the aim of developing a robust severity index is warranted. Nevertheless, this approach is preferable to deriving effect estimates exclusively from the literature.

Prescribing preferences for opioid maintenance therapy are often warranted as many women benefit from the different methods of delivery of care in the U.S. However, in many places in the U.S. patients do not have access to both treatment options due to both a lack of clinics and licensed providers along with limitations imposed by insurance. Lack of treatment options can result in structural confounding in other studies. In our study population, it is unlikely that this impacted our results as women had access to both treatment options and both were covered under Pennsylvania Medicaid, the primary insurer of this population.

As both observational studies and randomized trials are subject to the biases inherent in opioid maintenance treatment choices, it is imperative to account for this unmeasured confounding when comparing methadone with buprenorphine exposures in pregnancy to advocate for availability of both options if one is superior. Our results suggest that the previous findings that buprenorphine is associated with lower risk of neonatal abstinence syndrome compared with methadone in infants exposed in utero are subject to minimal bias from unmeasured confounding. Applying similar bias analyses to the association of these treatments with other neonatal outcomes is necessary to fully inform treatment decisions.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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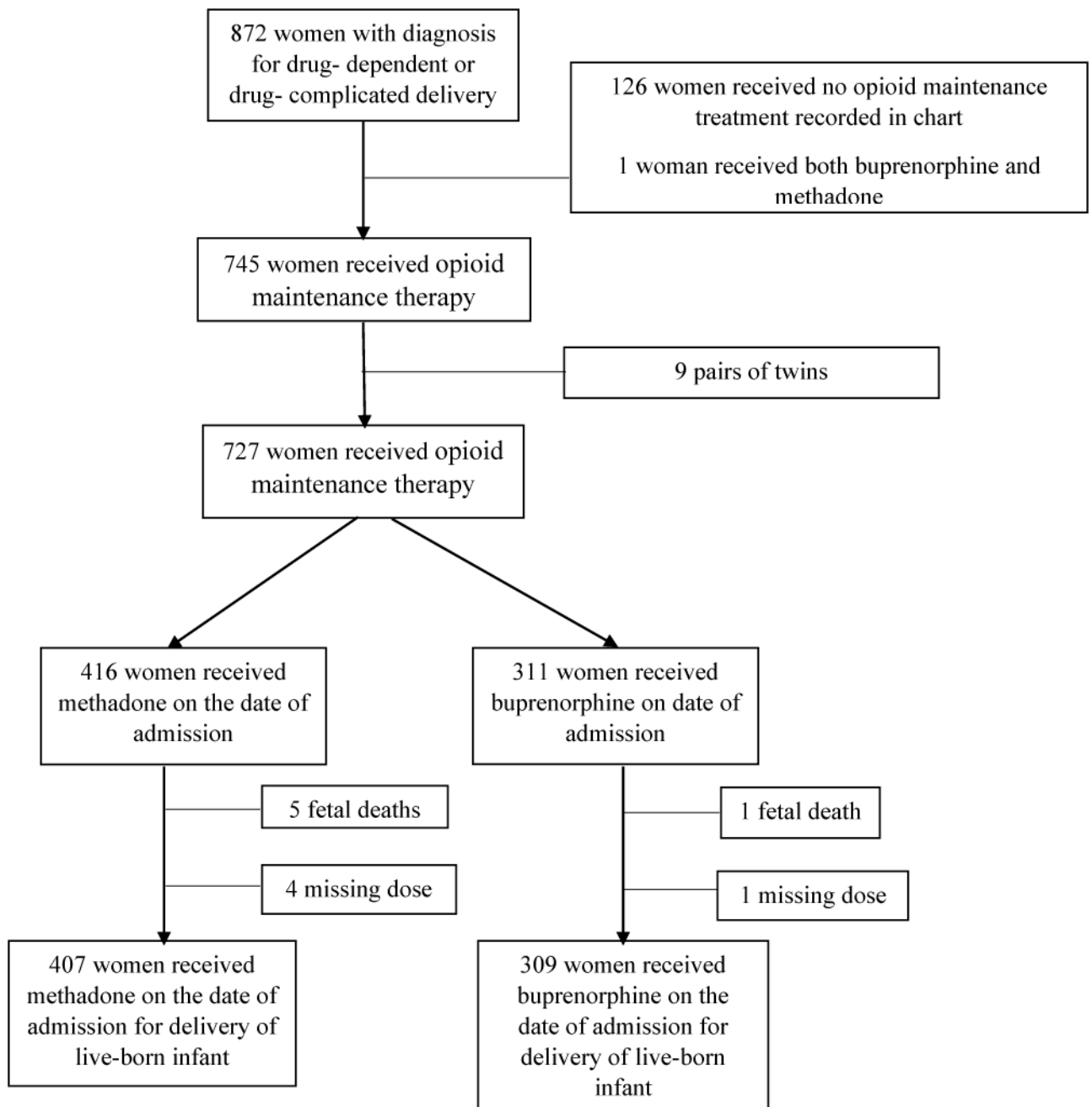


Figure 1. Flow diagram describing sample population (n=716, 2013–2015*Note: 25 women with 2 pregnancies).

Table 1

Maternal Characteristics by Opioid Maintenance Treatment Type, Magee-Womens Hospital, 2013–2015 (n=716).

Characteristic	Methadone N (%) n=407	Buprenorphine N (%) n=309
Race		
White	381 (93.6)	294 (95.1)
Black	19 (4.7)	8 (2.6)
Missing	7 (1.5)	7 (2.3)
Mother's Education		
Less than high school	83 (20.4)	45 (14.6)
High school graduate or GED completed	165 (40.5)	139 (45.0)
Some college credit	78 (19.2)	68 (22.0)
College graduate	66 (16.2)	54 (17.4)
Missing	15 (3.7)	3 (1.0)
BMI category ^a		
Underweight (<18.5kg/m ²)	11 (2.7)	17 (5.5)
Normal weight	116 (28.5)	94 (30.4)
Overweight	35 (8.6)	30 (9.7)
Obese	29 (7.1)	24 (7.7)
Missing	216 (53.1)	144 (46.6)
Married	35 (8.6)	58 (18.8)
Employed	139 (34.2)	135 (43.7)
Smoked during pregnancy	336 (82.6)	250 (80.9)
Parity		
Nulliparous	118 (29.0)	106 (34.3)
1–2 previous pregnancies	208 (51.1)	151 (48.9)
Greater than 2 pregnancies	81 (19.9)	52 (16.8)
Infant with congenital anomaly	50 (12.3)	27 (8.7)
Hepatitis c positive	61 (15.0)	31 (10.0)
Mother's age [Mean (SD)]	29.1 (4.7)	28.5 (4.9)
Prepregnancy BMI [Mean (SD)] ^b	24.6 (5.3)	24.2 (6.1)
Gestational age at delivery [Mean (SD)]	37.4 (2.9)	38.5 (2.5)
Birthweight [Mean (SD)]	2734 (619.3)	2999 (591.2)

^aPrepregnancy BMI defined as underweight (<18.5 kg/m²), normal weight (18.5 to <25 kg/m²), overweight (25 to <30 kg/m²), obese (>30 kg/m²).

^bPrepregnancy BMI based on n=356.

GED=general education development, SD=standard deviation, BMI=body mass index

Table 2

Maternal Characteristics by Opioid Maintenance Treatment Type in a Validation Subcohort, Magee-Womens Hospital, 2013–2015 (n=200).

Characteristic	Methadone N (%) n=100	Buprenorphine N (%) n=100
Race		
White	97	97
Black	3	2
Missing	0	1
Mother's Education		
Less than high school	19	9
High school graduate or GED completed	40	52
Some college credit	23	18
College graduate	14	19
Missing	4	2
Married	8	21
Employed	31	43
Smoked during pregnancy	84	80
Parity		
Nulliparous	31	39
1–2 previous pregnancies	47	46
Greater than 2 pregnancies	22	15
Hepatitis c positive	12	10
Infant with congenital anomaly	15	10
Severe maternal addiction	77	32
Mother's age [Mean (SD)]	28.6 (5.1)	28.2 (5.2)
Prepregnancy BMI [Mean (SD)] ^a	24.6 (5.9)	23.7 (5.3)
Gestational age at delivery [Mean (SD)]	37.3 (3.2)	39.1 (1.8)
Birthweight [Mean (SD)]	2695 (631.6)	3147 (472.4)

^aPrepregnancy BMI based on n=43 in methadone treated women and n=54 in buprenorphine treated women.

GED=general education development, SD=standard deviation, BMI=body mass index

Table 3

Results from conventional analyses of the risk of neonatal abstinence syndrome associated with methadone compared with buprenorphine as opioid maintenance therapy, at Magee-Womens Hospital, Pittsburgh, Pennsylvania (2013–2015).

Opioid maintenance therapy	Events (n)	Population at risk	Unadjusted risk per 100 livebirths	Unadjusted risk difference per 100 live-born infants (95% confidence interval)	Adjusted ^a risk difference per 100 live-born infants (95% confidence interval)
Buprenorphine	152	309	49	Reference	Reference
Methadone	263	407	65	15 (8.1, 23)	14 (5.9, 22)
				Unadjusted relative risk (95% confidence interval)	Adjusted^a relative risk (95% confidence interval)
Buprenorphine				Reference	Reference
Methadone				1.3 (1.2, 1.5)	1.3 (1.1, 1.5)

^a Adjusted for parity, maternal race, employment status, hepatitis c status, age, year of delivery, smoking status, marital status, and insurance.

Table 4

Characteristics of a subsample of opioid use dependent singleton pregnancies with severity of addiction indicators abstracted from medical charts at Magee-Womens Hospital in Pittsburgh, 2013–2015 (n=200).

Characteristic	Methadone n=100	Buprenorphine n=100
Converted to opioid maintenance therapy in pregnancy		
Yes	58	12
No	18	30
Missing	24	58
Gestational age at conversion (Median, IQR), weeks	12 (5, 22)	Prior to conception (prior, 4)
Relapse in pregnancy		
Yes	23	4
No	9	24
Missing	68	72
Using illicit substance at time of delivery		
Yes	24	15
No	34	71
Missing	42	14
Used benzodiazepines in pregnancy		
Yes	28	8
No	33	22
Missing	39	70
Neonatal abstinence syndrome		
Yes	61	54
No	39	46

IQR=interquartile range

Table 5

Comparison of results from adjusted conventional and probabilistic bias analyses accounting for unmeasured confounding by severity of addiction on the risk of neonatal abstinence syndrome associated with methadone compared with buprenorphine as opioid maintenance therapy, at Magee-Womens Hospital, Pittsburgh, Pennsylvania (2013–2015).

Opioid maintenance therapy	Conventional analysis: Adjusted ^a relative risk (95% confidence interval)	Bias Analysis 1: Adjusted ^a point estimate (95% bootstrapped simulation interval) ^b
Buprenorphine	Reference	Reference
Methadone	1.3 (1.1, 1.5)	1.2 (1.0, 1.4)

^aAdjusted for parity, maternal race, employment status, hepatitis c status, age, year of delivery, smoking status, marital status, and private vs. public insurance.

^bminimum RRC=1.0, mode 1=1.02, mode 2=1.11, maximum RRC=1.13

RRC=relative risk due to confounding