

# SYSTEMATIC REVIEW AND META-ANALYSIS

# Association between antidepressant medication use and epithelial ovarian cancer risk: a systematic review and meta-analysis of observational studies

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#### AIM

The aim of this paper is to clarify the inconsistent findings in the association between antidepressant use and the risk of epithelial ovarian cancer (EOC).

#### **METHODS**

This study is a meta-analysis of observational studies retrieved from the PubMed, EMBASE, and Web of Science databases prior to August 15, 2017. Two researchers independently screened studies and extracted study characteristics and risk estimates. The odds ratios (OR) and 95% confidence intervals (CI) of EOC risk were summarized using an inverse variance weighted randomeffects model. Heterogeneity between studies was assessed with the  $l^2$  statistic.

#### RESULTS

Eight case-control studies involving 7878 EOC cases and 73913 controls were identified. Compared with non-use, use of antidepressants was not significantly associated with EOC risk (summarized OR = 1.10, 95% CI: 0.91-1.32,  $I^2 = 74.4\%$ ). Similar null results were also observed in the use of selective serotonin reuptake inhibitors (OR = 1.04, 95% CI = 0.80–1.35), tricyclic antidepressants (OR = 1.01, 95% CI = 0.79–1.30), and other antidepressant drugs (OR = 0.91, 95% CI = 0.74–1.12). Subgroup analyses of study characteristics, stratified by the type of control subjects, geographic location, exposure assessment, number of cases, and adjustment for potential confounders, showed that the ORs were broadly consistent across strata. The OR per 1 yearincrement of duration was 0.99 (95% CI = 0.94–1.05,  $l^2$  = 40.0%, P = 0.154). Additionally, the OR for the greatest intensity of antidepressant use compared with never use was 0.82 (95% CI = 0.70–0.98,  $l^2 = 0\%$ , P = 0.489). Furthermore, no evidence of publication bias was detected through Funnel plots as well as Egger's and Begg's tests.

#### **CONCLUSIONS**

There is no association between antidepressant use and EOC risk. Further prospective studies are warranted to confirm these findings.



#### WHAT IS ALREADY KNOWN ABOUT THIS SUBJECT

- Antidepressants are widely prescribed to treat depression and anxiety disorders that may become chronic conditions among women.
- Epidemiological studies have yielded inconsistent results on the correlation between antidepressant use and ovarian cancer risk.

#### WHAT THIS STUDY ADDS

- Women who take antidepressants are not at an increased risk of ovarian cancer.
- Since these findings were generated on the basis of case–control studies, further large-scale prospective cohort studies are warranted.

## Introduction

Epithelial ovarian cancer (EOC) is the second most frequently occurring female reproductive malignancy and causes more deaths than any other type of gynaecological cancer, accounting for approximately 240 000 new diagnoses and 150 000 deaths worldwide in 2012 [1]. Recent studies have suggested that hormones and reproductive status are the predominant risk and protective factors for this disease [2–4].

Antidepressants are widely prescribed to treat depression and anxiety disorders that may become chronic conditions among women [5]. Of the more than 27 million people currently taking antidepressants in the United States, most are women [6], as they are twice as likely as men to be diagnosed with a major depressive disorder and up to three times more likely to be diagnosed with a dysthymic disorder [7].

Rodent experiments have shown that two distinct chemical classes of antidepressants, selective serotonin reuptake inhibitors (SSRIs) and tricyclics (TCAs), may have a growthpromoting effect on various types of tumours, including mammary tumours, fibrosarcoma, and melanoma [8-10]. Specially, these drugs may affect levels of dopamine or norepinephrine, leading to increased levels of gonadotropins and a subsequent greater risk of EOC [11]. However, since some studies have found antineoplastic properties or no effects of SSRIs and TCAs on cancer growth [7, 12, 13], evidence from experimental studies remains controversial. However, the concerns raised by a number of simulated epidemiologic investigations have questioned the possible association between antidepressant use and various types of cancer [14-19], including EOC [20]. Similarly, the evidence from epidemiological studies has been inconsistent [5, 11, 20-25].

A 2011 meta-analysis by Cosgrove et al. [7] reported that there might be a modest increase in the risk of EOC with the use of antidepressants, especially SSRIs. However, these findings were based on a combination of studies of breast and ovarian cancers, which is not ideal given that such an association could differ between these two entities. In view of this meta-analysis, several relevant studies with large sample sizes have been published [5, 21]. For example, a recent case-control study conducted in Denmark involving 4103 EOC patients and 58706 population-based controls suggested that use of SSRIs was associated with a decreased risk of EOC [21]. However, no recent meta-analysis has comprehensively summarized the evidence of a dose-response relationship between the duration of antidepressant use and EOC risk. Therefore, considering these inconsistencies, the aim of the present updated meta-analysis of observational

studies was to evaluate the association between antidepressant use and the risk of EOC.

# **Methods**

#### Data sources and search strategy

The present systematic review and meta-analysis was conducted in accordance with the Preferred Reporting Items for Systematic Reviews and Meta-Analyses statement [26]. We carried out a systematic search of the PubMed (https://www. ncbi.nlm.nih.gov/pubmed/), EMBASE (https://www.elsevier. com/solutions/embase-biomedical-research), and Web of Science (https://login.webofknowledge.com) databases (inception to August 15, 2017) for observational studies using the following combined key words and medical subject heading terms without language restriction: "antidepressants," "serotonin reuptake inhibitors," "SSRI," "fluoxetine," "paroxetine," "citalopram," "escitalopram," "sertraline," "fluvoxamine," "mianserin," "desipramine," "clomipramine," "lofepramine," "nortriptyline," "protriptyline," "imipramine," "tricyclics," "MAO inhibitors," "venlafaxine," "amitriptyline," "MAOI," "isocarboxazid," "moclobemide," "phenelzine," "tranylcypromine," "selegiline," "iproniazid," "TCAs," "ovarian," "ovary," "cancer," "tumour," "neoplasms," and "carcinoma" [27–29]. Furthermore, the references of the included studies and review articles were screened for relevant articles [30].

#### Study selection criteria

The included studies were limited to (1) observational studies that (2) defined the non-exposed group as women who did not use any kind of antidepressant, and (3) provided adjusted risk estimates (e.g., odds ratio [OR], relative risk, or risk ratios with 95% confidence intervals [CIs]) or provided data allowing the calculation of the risk estimates and 95% CIs of the association between antidepressant use and EOC risk. Review articles, systematic reviews, meta-analyses, commentaries, editorials and meeting abstracts, as well as those with other study designs (e.g., descriptive or ecological studies) were excluded from analysis. The duration of antidepressant use was defined as the day supply of the prescription [5]. The dose of antidepressant use was defined as the daily dose of the prescription. The intensity of antidepressant use was defined as the daily dose divided by the duration of use in days [21]. The selection and exclusion of studies were reviewed by two investigators (Y.-L.H. and S.G.).



Disagreements were resolved by consensus with a third author (J.-M.Q.).

#### Data extraction and quality assessment

Two researchers (Y.-L.H. and S.G.) independently retrieved the following general study information, where available, from all included studies: author, publication year, geographic location, study design, prevalence of antidepressant use, numbers of cases and controls, exposure category and measurement, outcome with risk estimates and 95% CIs, and adjusted/matched factors.

As all of the included articles were case–control studies, two researchers (Y.-L.H. and S.G.) independently used the Newcastle–Ottawa quality assessment scale for case–control studies to assess the risk of bias [31–35]. Subsequently, the studies were considered at a low risk of bias if they achieved a full rating in at least two categories of selection, comparability, or outcome assessment [36].

#### *Statistical analysis*

Since the absolute risk of EOC is low and all of the included articles were case-control studies, all results are reported as the OR for simplicity [37]. For studies that separately reported results for the use of SSRIs, TCAs, and other drugs, but not in combination, inverse variance weighted fixed effects metaanalysis was first used to generate an OR value of the overall study level of antidepressant use before random-effects meta-analysis [37]. Heterogeneity among studies was quantified with the  $I^2$  statistic [32, 38] and visually depicted using a Galbraith plot [39]. An  $I^2$  statistic of  $\leq$ 50% was considered to indicate less heterogeneity among the included studies. For all analyses, overall summary estimates were calculated using inverse variance weighted random-effects meta-analysis. Individual OR estimates and summary estimates are displayed graphically with forest plots. Pre-specified subgroup analyses were conducted according to type of control subjects (population-based vs. hospital-based), geographical location (US vs. non-US), exposure measurement (database vs. questionnaire), median number of cases (≥550 vs. <550), and adjustment for potential confounders (age, reproductive factors, and use of other non-antidepressant drugs). Furthermore, the study-specific OR was summarized for each 1-year-increment for the duration of antidepressant use. A study-specific trend reflecting correlated log OR values across different durations of drug use was computed using the generalized least-squares trend estimation method developed by Greenland et al. [40] and Orsini et al. [41]. Additionally, a potential nonlinear dose-response relationship between the duration of antidepressant use and EOC risk was modelled using restricted cubic splines with three knots at fixed percentiles (10%, 50% and 90%, respectively) of the distribution of exposure. An overall P-value was calculated by showing that these two regression coefficients simultaneously equalled zero. The P-value of nonlinearity was calculated by showing that the coefficient of the second spline equalled zero. The details of this method for calculating P-values are described elsewhere [42, 43].

The following information was required and used to conduct the dose–response meta-analysis [44–47]: (1) the distribution of cases and controls, and the risk and variance estimates for at least three quantitative exposure categories; (2) the median or mean level of these exposures in each category (if reported by ranges, mean levels were calculated by averaging the lower and upper boundaries; if the lowest category was open ended, the lowest boundary was considered to be zero; if the highest category was open ended, the open-ended interval length was assumed to be the same as the adjacent interval). Finally, six studies met the specifications when using these techniques and were included in the dose–response analysis of duration of antidepressant use and EOC risk. Given that only two studies met the criteria of intensity of antidepressant use and EOC risk, dose– response analysis was not conducted.

The sequential exclusion strategy proposed by Patsopoulos *et al.* [48] was used to examine whether the overall estimates were influenced by the substantial heterogeneity observed. Studies that accounted for the largest share of heterogeneity were sequentially and cumulatively excluded until  $I^2$  was <50%. We then examined whether the OR estimates were consistent [36]. Evidence of publication bias was examined using funnel plots. Additionally, funnel plot asymmetry was further confirmed with the Egger's [49] and Begg's [50] tests ( $P \le 0.10$ ). All statistical analyses were conducted using Stata 12.1 software (StataCorp LP, College Station, TX, USA).

#### Nomenclature of ligands

Key ligands in this article are hyperlinked to corresponding entries in http://www.guidetopharmacology.org, the common portal for data from the IUPHAR/BPS Guide to PHAR-MACOLOGY [51].

## Results

#### Characteristics of the retrieved studies

Of a total of 202 studies that were reviewed, 189 (93.6%) were excluded after screening of the titles and abstracts. After full text review of the remaining 13 (6.4%) studies, eight published studies of 7878 EOC cases and 73 913 controls detailing the associations between antidepressant use and EOC risk were included for analysis (Figure 1).

Supplementary Table S1 summarizes the characteristics of the eight included studies [5, 11, 20-25]. Notably, each was published between 1984 and 2017 and all employed a casecontrol study design. Of these, six included population-based controls [5, 11, 21-23, 25] and two had hospital-based controls [20, 24]. The number of cases in each study ranged from 150 [20] to 4103 [21], and the number of controls ranged from 250 [20] to 58706 [21]. Furthermore, five and three studies used questionnaires [11, 20, 22, 24, 25] and databases [5, 21, 23], respectively, to collect information regarding the duration of drug exposure. The majority of studies were conducted in North America (n = 5), Europe (n = 2), and Asia (n = 1), and were adjusted for potentially important confounders, such as age (n = 7), reproductive factors (n = 3), and use of other drugs (n = 3). On the basis of the Newcastle-Ottawa quality assessment scale, six studies [5, 11, 21–23, 25] were judged to be at a low risk of bias (Supplementary Table S2).





#### Figure 1

Selection of studies for inclusion in this meta-analysis

## *Ever antidepressant medication use and EOC* risk

All eight included studies, involving 7878 EOC cases and 73913 controls, examined ever antidepressant use and EOC risk. The pooled OR was 1.10 (95% CI = 0.91-1.32) (Figure 2). High heterogeneity was observed ( $I^2 = 74.4\%$ ). P < 0.001). The study by Coogan *et al.* was the farthest one from the fitted values, which contributed the most to the heterogeneity (Figure 3). Null results were also observed for the use of SSRIs (OR = 1.04, 95% CI = 0.80–1.35), TCAs (OR = 1.01, 95% CI = 0.79-1.30) and use of other non-antidepressant drugs (OR = 0.91, 95% CI = 0.74-1.12) (Table 1 and Supplementary Figure S1). Subgroup analyses of the study characteristics, stratified by the type of control subjects, geographic location, exposure assessment, number of cases, and adjustment for potential confounders, showed that the OR of outcomes were broadly consistent across strata (Table 1). The significance of the meta-regression analysis result for whether adjustment for age was observed might be attributed to a single study that only provided data for crude OR calculations [20]. Funnel plots as well as Egger's and Begg's tests showed no evidence of publication bias (Supplementary Figure S2). When we sequentially excluded the studies that contributed the largest amount to heterogeneity until  $I^2$  was less than 50%, the pooled ORs for outcomes (OR = 1.00, 95% CI = 0.86–1.16,  $I^2$  = 48.4%) were similar to the original estimates. Of note, after excluding the study by Coogan et al. [24], the result was robust but with moderate heterogeneity. Additionally, sensitivity analysis of the studies retrieved from the databases with less information bias as well as those using population-based controls was also conducted. Furthermore, the study by Tzonou et al. [20] was excluded because the authors failed to control for age in their primary analysis. Therefore, only three studies [5, 21, 23] met the inclusion criteria. The summarized OR was 0.93 (95% CI = 0.84–1.02), which was consistent with the main finding.

## Duration of antidepressant use and EOC risk

Six studies [5, 21-25] involving 7165 EOC cases and 71644 controls provided risk estimates for the longest duration of antidepressant use compared with never use. The pooled OR was 0.89 (95% CI = 0.66-1.19) with low heterogeneity  $(I^2 = 35.0\%, P = 0.174)$  (Supplementary Figure S3).

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#### Figure 2

Forest plot (random-effects model) of antidepressant use and epithelial ovarian cancer risk (ever vs. never use). The squares indicate study-specific odds ratio (ORs) (size of the square reflects the study-specific statistical weight), the horizontal lines indicate 95% confidence intervals (CIs), and the diamond indicates the summary OR estimate with its 95% CI. CI, confidence intervals; OR, odds ratio



Figure 3 Galbraith plot corresponding to the relationship between antidepressant use and epithelial ovarian cancer risk (ever vs. never use)



#### Table 1

Summary risk estimates of the association between antidepressant drug use and ovarian cancer risk (ever vs. never use)

	No. of studies	Summary OR 95% CI	<b>/<sup>2</sup> (%)</b> <sup>a</sup>	<b>P</b> h <sup>b</sup>	<b>P</b> h <sup>c</sup>
Antidepressant drugs	8	1.10 (0.91–1.32)	74.4	<0.001	
SSRIs	4	1.04 (0.80–1.35)	62.9	0.04	
TCAs	4	1.01 (0.79–1.30)	56.6	0.08	
Others	3	0.91 (0.74–1.12)	0	0.56	
Subgroup analyses of antidepressant drugs					
Type of control subjects					0.17
Population based	6	1.00 (0.86–1.16)	48.4	0.08	
Hospital based	2	2.50 (0.50–12.46)	77.3	<0.001	
Geographic location					0.50
US	5	1.16 (0.92–1.47)	61.1	0.04	
Non-US	3	1.00 (0.77–1.29)	72.7	<0.001	
Exposure assessment					0.13
Database	3	0.93 (0.84–1.02)	0	0.58	
Questionnaire	5	1.32 (0.98–1.78)	68.0	0.01	
Number of cases					0.95
<550	4	1.10 (0.90–1.36)	70.0	0.02	
≥550	4	1.27 (0.75–2.15)	82.4	<0.001	
Risk of bias					0.17
Low	6	1.00 (0.86–1.16)	48.4	0.08	
High	2	2.50 (0.50–12.46)	77.3	<0.001	
Adjustment for potential confounders					
Age					0.08
Yes	7	1.07 (0.90–1.27)	72.2	<0.001	
No	1	6.99 (1.46–33.35)	N/A	N/A	
Reproductive factors					0.78
Yes	3	1.23 (0.79–1.90)	77.8	0.01	
No	5	1.07 (0.85–1.36)	72.9	0.01	
Other non-antidepressant drugs use					0.78
Yes	3	1.23 (0.79–1.90)	77.8	0.01	
No	5	1.07 (0.85–1.36)	72.9	0.01	

CI, confidence interval; N/A, not available; OR, odds ratio; SSRIs, selective serotonin reuptake inhibitors; TCAs, tricyclic antidepressants. <sup>a</sup>/<sup>2</sup> statistic was used to quantify the magnitude between-study heterogeneity, and assigned values of 50% or less, 51–75%, and 76% or more for low, moderate and high heterogeneity, respectively.

<sup>b</sup>*P* value for heterogeneity within each subgroup. A two-tailed P < 0.1 was considered statistically significant.

 $^{c}P$  value for heterogeneity between subgroups with meta-regression analysis. A two-tailed P < 0.1 was considered statistically significant.

Additionally, of the five studies [5, 21–23, 25] included for dose–response analysis of duration, the summary OR per year was 0.99 (95% CI = 0.94–1.05) with low heterogeneity ( $I^2 = 40.0\%$ ) (Figure 4). There was no evidence of a nonlinear association between the parity number and gastric cancer risk (*P* for nonlinearity = 0.144).

## Intensity of antidepressant use and EOC risk

Two studies [21, 23] involving 4417 EOC cases and 59 496 controls examined the intensity of antidepressant use and EOC risk. The pooled OR was 0.79 (95% CI = 0.62–1.02) with low heterogeneity ( $I^2 = 17.4\%$ , P = 0.271) (Supplementary Figure S4).



#### Figure 4

Forest plot (random-effects model) of dose-response analysis of the duration (per 1-year increment) of antidepressant use and epithelial ovarian cancer risk. The squares indicate study-specific odds ratio (size of the square reflects the study-specific statistical weight), the horizontal lines indicate 95% confidence intervals (CIs), and the diamond indicates the summary OR estimate with its 95% CI. CI, confidence intervals; OR, odds ratio

## Discussion

The present updated meta-analysis with a large sample size (involving 7878 EOC cases and 73 913 controls) failed to detect a significant association between antidepressant use and the risk of EOC. These findings were consistent among different antidepressants, including SSRIs and TCAs. Null results were observed throughout the subgroup analyses stratified by study characteristics.

Although we noticed no significant result of the subgroup analysis of exposure assessment, the point estimate was relatively different (0.93 for database vs. 1.32 for questionnaire) and no heterogeneity was observed in studies using prescription record databases. Compared with the use of a questionnaire, using computerized prescription data to collect information might have eliminated recall bias, minimized selection bias, and provided a longer study period [21]. However, it must be recognized that pharmacy data also are imperfect and therefore insufficient to determine drug compliance [22]. Similar situations were also observed in the subgroup analyses stratified by exposure assessment for SSRIs and TCAs (data not shown). However, we could not rule out the possibility of limited inclusion of these two kinds of drugs in the studies (n = 4). Furthermore, misclassification of exposure status might still occur. If incorrect classification from these sources occurred similarly in both the case and control groups, thus creating bias, the risk estimates would tend to be biased towards the null [5, 52]. Consequently, the neutral association between antidepressant use and EOC risk might be partly attributed to these aforementioned issues. However, compared to misclassification of exposure status,

reporting/information bias (cases recalling exposures of use of antidepressants because of their disease) as well as controls not trying to find an explanation for their disease may possibly overestimate the aforementioned association.

The prevalence of antidepressant users varied among countries and studies, which might have resulted in the slightly stronger point estimate after summarizing the studies from the United States only. For example, in a study conducted in the US, Dublin *et al.* [23] reported that approximately 23.3% of the control group used antidepressants. By comparison, Harlow *et al.* [25] reported that only 2.2% of 454 population-based controls from the US used antidepressants, which was far less than that in a study conducted in Denmark by Morch *et al.* [21] who reported that 17.1% of 58 706 population-based controls used antidepressants. A similar situation was also observed for use of SSRIs and TCAs.

Even though the exact biological mechanism underlying the association of antidepressant with EOC risk remains unclear, several potential hypotheses have been proposed. Several studies of rodents and goldfish have indicated that certain dopaminergic medications affect the release of gonadotropins and prolactin [24, 53–56], which may play a role in ovarian carcinogenesis. Human studies have also shown that stimulation of dopaminergic receptors is involved in the release of luteinizing hormone, follicle-stimulating hormone, and prolactin [11, 57–59]. Increases in endogenous norepinephrine have been shown to induce gonadotropin secretion as well [54, 60]. Fluoxetine, a type of SSRI, was demonstrated to induce apoptosis of ovarian carcinoma OVCAR-3 cells [61]. Additionally, other studies have shown that sertraline and paroxetine can inhibit HT29 cells and induce alterations in



apoptosis-related proteins, stimulating programmed death of tumour cells [62]. Citalopram has also been suggested to control tumour growth by reducing the malignant cell cycle [63] and activating the immune system leading to apoptosis of tumour cells [12]. On the other hand, the possibility that antidepressants may exhibit a biphasic effect, characterized by 'low-dose stimulation and high-dose inhibition' of neoplastic cell proliferation, has been suggested [5, 7]. Some *in vitro* studies of rodents have indicated that antidepressant use can both promote and inhibit tumour growth [24, 64]. Considering that the evidence from biological studies are both limited and speculative, future experimental studies are certainly warranted.

An important strength of the present study was the large sample size. As compared with a previous meta-analysis conducted in 2011 [7], the present study also searched the EMBASE database. Furthermore, we included two studies [5, 21] published in the past 5 years comprising 5060 EOC patients and 68 276 controls. Hence, the numbers of patients and controls (n = 2668 and 5387, respectively) were far greater than included in the 2011 meta-analysis. Additionally, we included one study [20] that was missed in the previous metaanalysis, though only crude OR data were provided. Notably, we investigated the aforementioned association in numerous subgroup analyses stratified by study characteristics.

Nonetheless, there were also some limitations to the present study that should be addressed. First, all of the included studies were case-control studies. As compared with prospective studies, case-control studies are more susceptible to bias (e.g., recall bias, selection bias) due to their nature. However, no recent prospective study has reported the association between antidepressant use and EOC risk according to our search strategy. Secondly, we could not rule out the influence of depression besides that of antidepressant therapy. Huang et al. [65] reported that a state of depression was associated with an increased risk of EOC on the basis of two large prospective cohort studies. Only one study had adjusted for comorbid medical and psychiatric illnesses in their primary analysis [5]. Considering that severely depressed patients are more likely to be treated with antidepressants, they might just be a surrogate marker for the depression-related incidence of EOC [5]. Furthermore, since TCAs are used predominantly for manifest depression, whereas SSRIs are used for variety of other or 'off-label' indications, we cannot exclude the possibility that the specific indications for use of SSRIs and TCAs may have differentially influenced the associations of EOC risk with SSRI and TCA use. However, a limited number of the included studies elaborated on this issue and presented primary analysis data restricted to women taking some kind of antidepressant.

Thirdly, previous studies found a tendency of a decreasing risk of EOC with increasing intensity of antidepressant use, which was defined as the daily dose divided by the duration of use in days, but not with increasing duration [21]. However, as limited studies (n = 2) provided risk estimates of the intensity of antidepressant use [21, 23] and none of the included studies focused on the dose of antidepressant use, further studies are warranted to investigate this possible association as well as to confirm the findings of the present study. Furthermore, as a limited number of studies included in our analysis provided sufficient data for dose–response analysis, meta-analysis of the intensity of antidepressant use and EOC was not conducted. Notably, only three studies [11, 21, 23] carried out subgroup analyses by histological type; therefore, future studies are needed.

Lastly, although the primary analyses of the included studies were adjusted for several known risk factors for EOC, we could not adjust for all potential confounders in this meta-analysis nor rule out the possibility of residual confounding factors. For example, nulliparity and infertility may be associated with not only EOC, but also antidepressant use [11]. However, only four studies [11, 21, 23, 25] considered these factors and just one [23] conducted subgroup analysis subdivided by parity. Therefore, future studies should investigate this association after careful and thorough adjustment for potential confounders.

In conclusion, the results of the present updated metaanalysis involving the largest sample size to date and mostly included comprehensive observational studies showed no association between antidepressant use and the risk of EOC. Further prospective studies are warranted to confirm these findings.

# **Competing Interests**

There are no competing interests to declare.

# Contributors

Y.-L.H. and S.G. designed the research, analyzed the data, and wrote the draft manuscript. All authors conducted the study search, read, reviewed and approved the final manuscript. S.G. had primary responsibility for final content.

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# **Supporting Information**

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http://onlinelibrary.wiley.com/doi/10.1111/bcp.13498/suppinfo

**Table S1** Characteristics of observational studies of antidepressant drug use and epithelial ovarian cancer risk

**Table S2** Quality assessment of the included case–control studies using the Newcastle–Ottawa quality assessment scale **Figure S1** Test for publication bias of antidepressant use and epithelial ovarian cancer risk through Begg's funnel plot. OR, odds ratio; SE, standard error. The circles alone are real studies. The vertical lines represent the summary effect estimates, and the dashed lines represent pseudo-95% confidence interval limits

**Figure S2** Forest plot (random-effects model) of antidepressant use and epithelial ovarian cancer risk by types of drug (ever *vs.* never use). The squares indicate study-specific odd ratios (size of the square reflects the study-specific statistical weight), the horizontal lines indicate 95% CIs, and the diamond indicates the summary OR estimate with its 95% CI. CI, confidence intervals; OR, odds ratio; SSRIs, selective serotonin reuptake inhibitors; TCAs, tricyclic antidepressants

**Figure S3** Forest plot (random-effects model) of antidepressant use and epithelial ovarian cancer risk (longest duration *vs.* never use). The squares indicate study-specific ORs (size of the square reflects the study-specific statistical weight), the horizontal lines indicate 95% CIs, and the diamond indicates the summary OR estimate with its 95% CI. CI, confidence intervals; OR, odds ratio

**Figure S4** Forest plot (random-effects model) of antidepressant use and epithelial ovarian cancer risk (strongest intensity *vs.* never use). The squares indicate study-specific ORs (size of the square reflects the study-specific statistical weight), the horizontal lines indicate 95% CIs, and the diamond indicates the summary OR estimate with its 95% CI. CI, confidence intervals; OR, odds ratio