

Optimizing Tuberculosis Diagnosis in Human Immunodeficiency Virus–Infected Inpatients Meeting the Criteria of Seriously Ill in the World Health Organization Algorithm

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Background. The World Health Organization (WHO) algorithm for the diagnosis of tuberculosis in seriously ill human immunodeficiency virus (HIV)–infected patients lacks a firm evidence base. We aimed to develop a clinical prediction rule for the diagnosis of tuberculosis and to determine the diagnostic utility of the Xpert MTB/RIF assay in seriously ill HIV-infected patients.

Methods. We conducted a prospective study among HIV-infected inpatients with any cough duration and WHO-defined danger signs. Culture-positive tuberculosis from any site was the reference standard. A priori selected variables were assessed for univariate associations with tuberculosis. The most predictive variables were assessed in a multivariate logistic regression model and used to establish a clinical prediction rule for diagnosing tuberculosis.

Results. We enrolled 484 participants. The median age was 36 years, 65.5% were female, the median CD4 count was 89 cells/ μ L, and 35.3% were on antiretroviral therapy. Tuberculosis was diagnosed in 52.7% of participants. The c-statistic of our clinical prediction rule (variables: cough \geq 14 days, unable to walk unaided, temperature $>$ 39°C, chest radiograph assessment, hemoglobin, and white cell count) was 0.811 (95% confidence interval, .802–.819). The classic tuberculosis symptoms (fever, night sweats, weight loss) added no discriminatory value in diagnosing tuberculosis. Xpert MTB/RIF assay sensitivity was 86.3% and specificity was 96.1%.

Conclusions. Our clinical prediction rule had good diagnostic utility for tuberculosis among seriously ill HIV-infected inpatients. Xpert MTB/RIF assay, incorporated into the updated 2016 WHO algorithm, had high sensitivity and specificity in this population. Our findings could facilitate improved diagnosis of tuberculosis among seriously ill HIV-infected inpatients in resource-constrained settings.

Keywords. HIV; tuberculosis diagnosis; WHO algorithm; inpatients; Xpert MTB/RIF assay.

Tuberculosis remains a major cause of death among human immunodeficiency virus (HIV)–infected adults in resource-constrained countries in the antiretroviral therapy (ART) era [1]. Delayed or missed diagnosis contributes to tuberculosis mortality [2, 3]. World Health Organization (WHO) 2007 guidelines to diagnose smear-negative tuberculosis [4] included an algorithm for seriously ill patients with cough of 2–3 weeks and \geq 1 danger sign (respiratory rate $>$ 30 breaths/minute, heart rate $>$ 120 beats/minute, temperature $>$ 39°C, and being unable to walk unaided).

A modified WHO algorithm for seriously ill patients improved clinical outcomes in 2 African cohort studies [5, 6]. However, no study has explored the ability of the 2007 WHO seriously ill algorithm to diagnose tuberculosis, which has several limitations. First, pulmonary tuberculosis commonly presents with cough duration of $<$ 14 days [7, 8]. Second, other classic tuberculosis symptoms (fever, night sweats, and weight loss), which have high negative predictive value in active case finding [9], were not included. Third, it preceded the Xpert MTB/RIF assay, which is more sensitive than smear microscopy in HIV-infected patients [10]. Finally, hemoglobin concentration and white cell count (WCC), which predict tuberculosis among HIV-infected patients [11–13], could add diagnostic value.

We conducted a prospective cohort study to attempt to improve the ability of the 2007 WHO algorithm to diagnose tuberculosis in seriously ill HIV-infected patients by evaluating any cough duration, classic tuberculosis symptoms, chest radiographic features, hemoglobin, WCC, and the Xpert MTB/RIF assay.

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METHODS

Study Population

We conducted a prospective cohort study in 2 regional hospitals in Cape Town, South Africa serving high-burden HIV and tuberculosis communities, GF Jooste Hospital (November 2011–February 2013) and Khayelitsha District Hospital (March 2013–October 2014). Inclusion criteria were as follows: HIV infected, ≥ 18 years of age, admitted within 24 hours, coughing for any duration, and ≥ 1 WHO danger sign. Exclusion criteria were antituberculosis therapy that is current or completed in the previous month or defaulted within the past 6 months, exacerbation of cardiac failure or chronic obstructive pulmonary disease, and inability to produce a spontaneous or induced sputum sample. Participants were followed up at 28 and 56 days after discharge.

Data Collection

Demographic data and clinical features (at the time of admission) were recorded on a standardized case record form. Chest radiographs performed on admission were assessed by the study medical officer for features suggestive of pulmonary tuberculosis and/or bacterial pneumonia and/or *Pneumocystis jirovecii* pneumonia (PJP). Chest radiographs were also retrospectively reviewed by a blinded specialist radiologist, who documented the presence of specific radiographic features (enlarged hilar or mediastinal lymph nodes, pleural effusions, interstitial (“ground glass”) infiltration, nodularity, cavitation, diffuse micronodular infiltration, linear/reticulonodular infiltration, consolidation, and features of previous tuberculosis), and then classified radiographs as likely, possible, or unlikely for pulmonary tuberculosis and/or bacterial pneumonia and/or PJP. CD4 cell count was obtained on admission if none was available from the prior 6 months, and hemoglobin and WCC were done on admission. Sputum induction, using an ultrasonic nebulizer and hypertonic saline, was done on participants unable to spontaneously produce sputum. Three sputum samples from each participant were sent for Gram stain, culture, and sensitivity (1 sample); and auramine smear microscopy for acid-fast bacilli, and mycobacterial culture (BACTEC MGIT 960; Becton Dickinson, Franklin Lakes, New Jersey) (2 samples). The sputum pellet on one of the samples for mycobacterial culture was split after decontamination for Xpert MTB/RIF assay (Cepheid, Sunnyvale, California). Mycobacterial blood cultures (BacT/Alert MP; bioMérieux, Durham, North Carolina) were sent from all participants. Extrapulmonary samples (eg, pleural fluid) were sent for mycobacterial culture when appropriate.

Case Definitions and Procedures

Tuberculosis was defined as a positive culture for *Mycobacterium tuberculosis* from any site. The study criteria for initiating antituberculosis therapy included positive microbiological diagnosis (auramine stain and/or Xpert MTB/RIF assay and/or culture);

and/or a radiological diagnosis (chest radiograph showing mediastinal and/or hilar lymph nodes or miliary infiltrates; abdominal ultrasound showing multiple enlarged lymph nodes and/or multiple splenic hypoechoic lesions and/or pericardial effusion); and/or pleural effusion/ascites showing a lymphocytic exudate; and/or no clinical improvement after 3–5 days of antibiotic therapy. Decisions to start antituberculosis therapy were made by study staff.

The diagnosis of bacterial pneumonia was made with consistent symptoms and evidence of consolidation on chest radiograph. Bacterial pneumonia was treated with a broad-spectrum β -lactam antibiotic (eg, ceftriaxone, amoxicillin-clavulanate), and a macrolide was added with CRB-65 score > 2 (1 point for each of: confusion, respiratory rate ≥ 30 /min, blood pressure systolic < 90 mm Hg or diastolic ≤ 60 mm Hg, age ≥ 65 years [14, 15].

PJP was diagnosed with a cough duration ≤ 12 weeks, bilateral interstitial infiltration on chest radiograph, and hypoxia (defined as oxygen saturation of $\leq 92\%$) or dyspnea. PJP was treated with high-dose trimethoprim-sulfamethoxazole, and prednisone if hypoxia was present.

Statistical Analysis

A sample size of 473 was sufficient to detect an estimated 30% prevalence of culture-positive tuberculosis with 5% precision. This sample size incorporated the need for at least 10 culture-positive tuberculosis events per predictive variable for multivariate logistic regression analysis [16].

Analyses were performed using Stata version 12.1 software (StataCorp, College Station, Texas). Missing data were imputed using chained equations and 20 iterations. Baseline characteristics were described as proportions or medians. We used univariate associations to assess the ability of the following a priori selected variables to predict tuberculosis: age, sex, cough duration, individual WHO danger signs, classic tuberculosis symptoms (fever, night sweats, and weight loss), radiologist assessment of tuberculosis on chest radiograph (categorized as likely or possible), hemoglobin, and WCC [17]. A backward stepwise approach proposed by Collet [18] was used to select the most predictive variables in establishing a multivariate logistic regression model.

The model was visually assessed by a calibration plot, and by the Hosmer–Lemeshow test. An estimate of the c-statistic was used to assess discrimination (values 0.7–0.8 are deemed acceptable and 0.8–0.9 good) [19]. Internal validation used 1000 bootstrap resamples [20]. A clinical prediction rule with score chart was constructed utilizing a standard method [21]. We used the clinical prediction rule to predict the probability of having tuberculosis, and compared it with the reference standard of culture-positive tuberculosis. We calculated the diagnostic accuracy for the range of possible scores from the clinical prediction rule.

We explored associations of individual chest radiograph features recorded by the radiologist with culture-positive tuberculosis, reported as odds ratios with 95% confidence intervals (CIs).

We calculated the diagnostic accuracy of sputum smear and Xpert MTB/RIF assay for diagnosing culture-positive tuberculosis.

Ethics Approval

Approval for the study was obtained from the University of Cape Town Human Research Ethics Committee. All eligible participants signed informed consent before enrollment into the study. Confused participants were enrolled and given the option to continue with participation once orientated; their data were removed from the study if consent was declined.

RESULTS

Participant Characteristics

We screened 2054 patients and enrolled 484 (Figure 1). Table 1 shows participants' baseline characteristics. All participants were commenced on antibiotic therapy at admission or at the referral site. Fifty-three percent (255/484) of participants had culture-positive tuberculosis, 50.6% (245/484) bacterial pneumonia, and 10.5% (51/484) PJP. Antituberculosis therapy was empirically started in 56 (11.6%) participants with negative sputum smears and Xpert MTB/RIF assays, of whom 22 (43.1%) had culture-positive tuberculosis. Other diagnoses (*Cryptococcus neoformans*, *Emmonsia parva*, nontuberculous mycobacteria, and malignancy) were made in 8.5% (41/484). Clinically diagnosed coinfections were common: tuberculosis with bacterial pneumonia in 25.8% (125/484) and tuberculosis with PJP

in 4.3% (21/484). Different sites of culture-positive specimens (Supplementary Table 1) and pulmonary/extrapulmonary/disseminated tuberculosis are given in the Supplementary Data. The diagnostic accuracy of Xpert MTB/RIF assay and smear is shown in Table 2.

The yield of sputum culture was higher with sputum induction than spontaneous sputum production (56.1% [175/312] vs 45.5% [74/163]; $P = .027$). Similarly, the yield of smear and Xpert MTB/RIF assay was higher with sputum induction than spontaneous sputum production (32.7% [102/312] vs 30.1% [49/163], $P = .553$ and 51.9% [162/312] vs 41.7% [68/163], $P = .035$, respectively).

Seventy-two participants with negative sputum smears had features of tuberculosis on chest radiograph diagnosed by medical officers, fulfilling the criteria of the 2007 WHO algorithm for seriously ill inpatients to start empiric antituberculosis therapy with a sensitivity of 49.1% and specificity of 91.5% for the diagnosis of culture-positive tuberculosis.

Clinical Prediction Rule

Univariate associations with culture-positive tuberculosis are shown in Table 3. In multivariate logistic regression (Supplementary Table 2) the most significant predictors of culture-positive tuberculosis were being unable to walk unaided, radiologist assessment of "likely tuberculosis" on chest radiograph, and anemia. Raised WCC was a significant negative predictor of tuberculosis. The only classic tuberculosis symptom that showed a significant association with the diagnosis of tuberculosis was weight loss, but it was not significant on multivariate analysis. Cough duration ≥ 14 days was predictive of tuberculosis, but 28.6% (73/255) of culture-positive tuberculosis participants had cough duration < 14 days.

The calibration curve (Figure 2A) for the final model followed the ideal calibration line, indicating good agreement between the rate of tuberculosis estimated by the model and the tuberculosis frequency observed in the study population, confirmed by the Hosmer-Lemeshow χ^2 statistic of 10.65 ($P = .385$). The model's c-statistic was 0.834 (95% CI, .798–.871) (Figure 2B). The equivalents in bootstrap validation were 0.836 (95% CI, .800–.871), with an optimism estimate of 0.011 (95% CI, .010–.012), indicating good stability of the model in internal validation.

The clinical prediction rule with score chart was developed using the 6 selected variables (Table 4). Hemoglobin and WCC were categorized as tertiles. The diagnostic accuracy assessment based on different scores obtained from the clinical prediction rule is summarized in Table 5. The clinical prediction rule model showed a c-statistic of 0.811 (95% CI, .802–.819).

Chest Radiograph Assessment

The radiologist assessed 416 participants' chest radiographs: 151 (36.3%) were categorized as likely, 223 (53.6%) as possible,

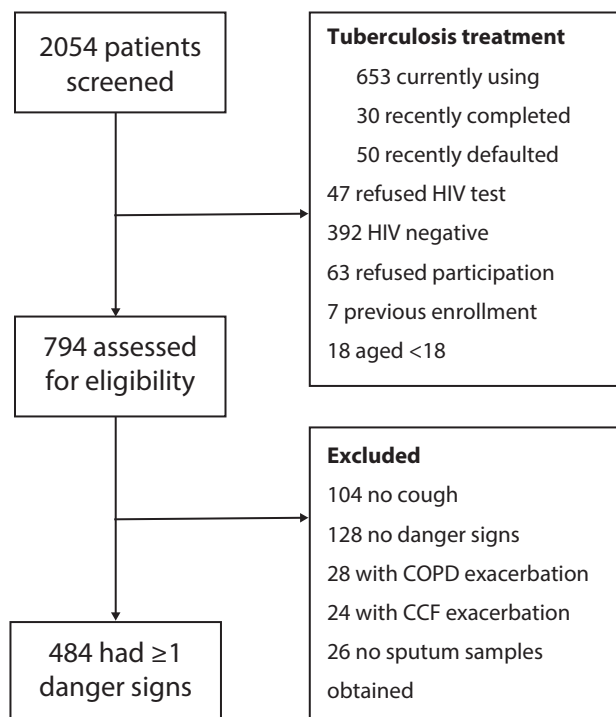


Figure 1. Flow diagram for participant inclusion into the study. Abbreviations: CCF, congestive cardiac failure; COPD, chronic obstructive pulmonary disease; HIV, human immunodeficiency virus.

Table 1. Baseline Characteristics of 484 Seriously Ill Human Immunodeficiency Virus–Infected Participants Presenting With a Cough of Any Duration and ≥1 World Health Organization Danger Sign

Characteristic	No. (%) or Median (25th–75th Percentile)
Baseline variables	
Age, y	36 (30–42)
Sex, female	317 (65.5)
Body mass index ^a , kg/m ²	20 (18–39)
CD4 count ^b , cells/μL	89 (34–210)
Cough duration ^c , d	14 (7–21)
Cough duration ≥14 d	296 (61.2)
Using ART	171 (35.3)
Duration on ART, y	2.3 (0.2–5)
Previous tuberculosis	236 (48.8)
Sputum induction ^d	312 (65.7)
WHO danger signs	
Respiratory rate >30 breaths/min	315 (65.1)
Heart rate >120 beats/min	383 (79.1)
Temperature >39°C	81 (16.7)
Unable to walk unaided	259 (53.5)
Tuberculosis symptoms	
Fever ^e	394 (81.6)
Night sweats ^f	314 (65.3)
Weight loss	442 (91.3)
Laboratory investigations	
Hemoglobin ^g g/dL	9.5 (7.8–11.2)
WCC ^h , ×10 ⁹ /L	8.6 (5.6–13.0)

Abbreviations: ART, antiretroviral therapy; WCC, white cell count; WHO, World Health Organization.

^aSeventeen missing values.

^bOne missing value.

^cTwo missing values.

^dNine missing values.

^eOne missing value.

^fThree missing values.

^gThree missing value.

^hThree missing values.

and 42 (10.0%) as unlikely tuberculosis. Univariate analysis of individual chest radiograph features revealed that diffuse micronodular infiltration, enlarged hilar or mediastinal lymph nodes, and nodularity >3 mm were predictors of culture-positive tuberculosis (Table 6). Bacterial pneumonia was a likely

Table 2. Diagnostic Accuracy Evaluation of Sputum Auramine Smear and Xpert MTB/RIF Assay by the Reference Standard of Positive Culture From Any Site (n = 255) Among 484 Seriously Ill Human Immunodeficiency Virus–Infected Participants Presenting With a Cough of Any Duration and ≥1 World Health Organization Danger Sign

Parameter	Xpert MTB/RIF	Auramine Smear
Sensitivity, %	86.3 (81.5–90.3)	57.0 (50.7–63.2)
Specificity, %	96.1 (92.6–98.2)	98.7 (96.2–99.7)
Positive predictive value, %	96.1 (92.7–98.2)	98.0 (94.2–99.6)
Negative predictive value, %	86.2 (81.4–90.2)	67.2 (61.9–72.2)
Positive likelihood ratio	21.9 (11.5–41.6)	43.3 (14.0–134.0)
Negative likelihood ratio	0.14 (.11–.19)	0.44 (.38–.50)

Values in parentheses indicate the 95% confidence interval.

diagnosis in 9.6% (40/416) and PJP in 9.4% (39/416). Medical officers' categorization of likely tuberculosis on chest radiograph had an odds ratio of 9.4 (95% CI, 5.7–15.6) for culture-positive tuberculosis.

DISCUSSION

Our study is the first to evaluate the clinical, radiographic, and laboratory diagnosis of tuberculosis in a prospective cohort of

Table 3. Univariate Associations With Culture-Positive Tuberculosis Among 484 Seriously Ill Human Immunodeficiency Virus–Infected Participants Presenting With a Cough of Any Duration and ≥1 World Health Organization Danger Sign

Variable	OR	(95% CI)	Wald P Value
Age ^a	0.98	(.96–1.00)	.044
Sex			
Female		Referent group	
Male	0.95	(.65–1.39)	.799
Cough duration ≥14 d			
No		Referent group	
Yes	2.47	(1.70–3.60)	<.001
WHO danger signs			
Respiratory rate >30 breaths/min			
No		Referent group	
Yes	0.78	(.54–1.14)	.207
Heart rate >120 beats/min			
No		Referent group	
Yes	1.53	(.98–2.37)	.060
Temperature >39°C			
No		Referent group	
Yes	1.86	(1.13–3.07)	.014
Unable to walk unaided			
No		Referent group	
Yes	2.77	(1.92–4.01)	<.001
Tuberculosis symptoms			
Fever			
No		Referent group	
Yes	1.07	(.67–1.69)	.783
Night sweats			
No		Referent group	
Yes	1.18	(.81–1.72)	.384
Weight loss			
No		Referent group	
Yes	3.08	(1.54–6.17)	.002
Increasing number of tuberculosis symptoms	1.28	(1.01–1.64)	.041
Radiographic assessment of tuberculosis			
Unlikely		Referent group	
Possible	2.13	(1.00–4.54)	.052
Likely	11.01	(4.92–24.66)	<.001
Laboratory investigations			
Hemoglobin, g/dL ^b	0.69	(.63–.76)	<.001
WCC, ×10 ⁹ /L ^c	0.90	(.87–.93)	<.001

Abbreviations: CI, confidence interval; OR, odds ratio; WCC, white cell count; WHO, World Health Organization.

^aPer 1 year increase.

^bPer 1 g/dL increase.

^cPer 1 × 10⁹/L increase.

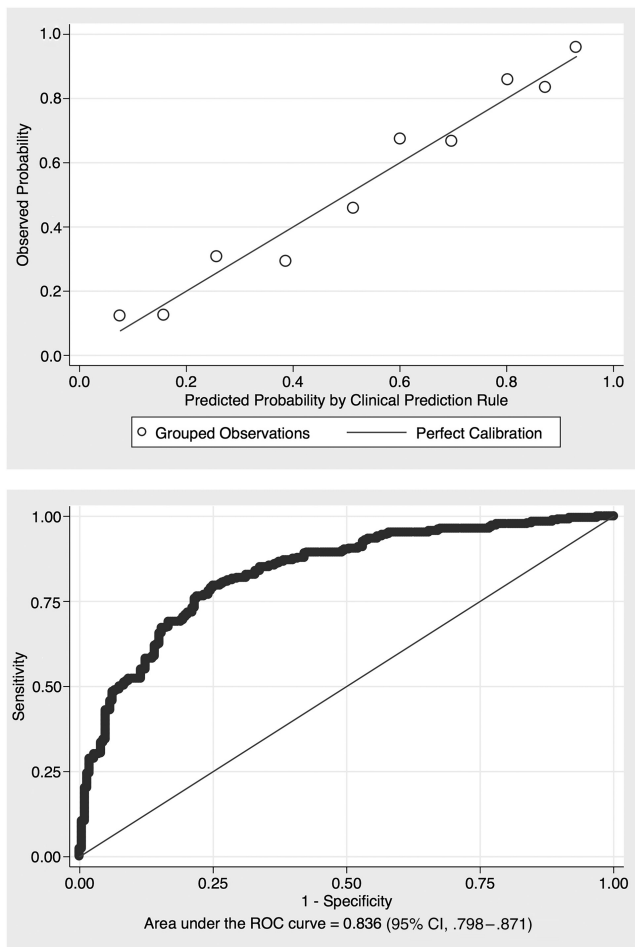


Figure 2. Upper, Calibration plot for the assessment of variables included in a multivariate logistic regression model aimed at establishing a clinical prediction rule for the diagnosis of tuberculosis among 484 seriously ill human immunodeficiency virus (HIV)-infected participants presenting with a cough of any duration and 1 or more World Health Organization (WHO) danger sign. The line shows perfect calibration between observed and predicted tuberculosis. Lower, Discrimination of the multivariate logistic regression model aimed at establishing a clinical prediction rule for the diagnosis of tuberculosis among 484 seriously ill HIV-infected participants presenting with a cough of any duration and ≥ 1 WHO danger sign. Abbreviations: CI, confidence interval; ROC, receiver operating characteristic.

inpatients fulfilling the criteria of the 2007 WHO algorithm for seriously ill HIV-infected patients. In 2016 (when enrollment into our study was completed), the WHO updated the algorithm by including classic tuberculosis symptoms and the Xpert MTB/RIF assay, and by making cough of any duration an inclusion criterion [22]. We had already incorporated all of these features into our study; therefore, we are also able to evaluate the updated WHO algorithm, but it should be noted that cough is no longer a requirement. We developed a clinical prediction rule with good diagnostic performance for tuberculosis using 6 variables, all of which are readily obtainable in most resource-constrained settings. The most significant predictors of tuberculosis were a radiologist assessment of likely tuberculosis on chest radiograph and anemia, while a raised WCC was

a strong negative predictor of tuberculosis. The classic tuberculosis symptoms added no discriminatory value in diagnosing tuberculosis. Scores of 3 or 4 in our clinical prediction rule could be used to start empiric antituberculosis therapy in seriously ill patients as the sensitivity of these scores is around 90%. However, no single feature of our clinical prediction rule should be used in decisions to treat empirically for tuberculosis. The Xpert MTB/RIF assay, which has not previously been evaluated in seriously ill inpatients, had a high sensitivity of 86.3% in our participants. Implementation of our clinical prediction rule in resource-constrained settings could augment the 2016 WHO algorithm, inform the development of future algorithms, and ensure timely initiation of empiric antituberculosis treatment in seriously ill HIV-infected patients.

The 2007 WHO algorithm for seriously ill patients requires cough duration of 2–3 weeks. Although we found that cough duration of ≥ 14 days was predictive of tuberculosis, 28.6% of participants with confirmed tuberculosis had cough duration < 14 days. Most of the data on cough duration of > 14 days as a trigger for tuberculosis investigations are from studies of ambulatory patients and may not be generalizable to seriously ill inpatients. Two African studies found a high prevalence of tuberculosis among inpatients presenting acutely with pneumonia [7, 8]. These findings together with ours suggest that any cough duration was an appropriate improvement made in the 2016 WHO algorithm for the diagnosis of tuberculosis among seriously ill patients.

The classic tuberculosis symptoms fever and weight loss were sensitive, but not specific for the diagnosis of tuberculosis among HIV-infected inpatients with negative sputum smears [23]. The differential diagnosis in HIV-infected inpatients with cough and danger signs is wide, and many of these disorders and/or advanced HIV disease could reduce the diagnostic utility of the classic tuberculosis symptoms, which could explain why these were not independently associated with tuberculosis in our study. Two of the WHO danger signs, being unable to walk unaided and temperature $> 39^\circ\text{C}$, were significant predictors of tuberculosis in our cohort.

The WHO has recently reiterated the importance of using chest radiographic assessment to facilitate the rapid initiation of antituberculosis therapy among seriously ill HIV-infected patients [24]. Our study supports this recommendation, with a radiologist assessment of likely tuberculosis being the strongest predictor of tuberculosis in our clinical prediction rule. Other studies of hospitalized HIV-infected participants with smear-negative pulmonary tuberculosis have also reported the diagnostic value of chest radiography [17, 23, 25, 26]. The commonest individual radiographic feature associated with culture-positive tuberculosis in our study was enlarged hilar or mediastinal lymph nodes. Cavitation was not a significant clinical predictor of culture-positive tuberculosis, which might be explained by the low median CD4 count and high prevalence

Table 4. Clinical Prediction Rule for the Diagnosis of Culture-Positive Tuberculosis Among 484 Seriously Ill Human Immunodeficiency Virus–Infected Participants Presenting With a Cough of Any Duration and ≥1 World Health Organization Danger Sign

Variable	Category	Points
Cough duration	≥14 d	1
Temperature >39°C	Yes	1
Unable to walk unaided	Yes	1
Tuberculosis on chest radiograph	Possible	1
	Likely	5
Hemoglobin, g/dL	3.3–8.3	3
	8.4–10.6	2
White cell count, ×10 ⁹ /L	1–6.5	1
	11.2–40.4	–2

of previous tuberculosis. Other studies of inpatients from high-burden settings have also reported that intrathoracic lymphadenopathy was a good predictor of tuberculosis, but cavitation was not [25, 26].

Our finding that anemia was associated with tuberculosis is in keeping with other studies of inpatients and outpatients [12, 26]. A study from Cameroon found anemia and leukopenia to be independent predictors of extrapulmonary involvement in patients with pulmonary tuberculosis; many patients in our cohort had disseminated disease, which could explain why hemoglobin and WCC were good predictors of tuberculosis [13].

Sputum Xpert MTB/RIF assay performed well in our cohort, with a sensitivity of 86.3%, which is somewhat higher than the sensitivity of 79% reported in HIV-infected patients in a systematic review [10]. The high sensitivity of Xpert MTB/RIF assay we found could be explained by the high proportion of participants who had sputum induction, which increased the yield of Xpert MTB/RIF assay in our study (contrary to the findings of a study in ambulatory patients) [27]. We found a higher yield of sputum culture with sputum induction than spontaneous sputum production, which is consistent with other studies [27, 28]. In view of these findings, we call for greater access to sputum induction in resource-constrained settings.

Our study has some limitations. First, our findings may not be generalizable to patients without cough or in settings with different prevalence of tuberculosis and other pulmonary opportunistic infections. Two studies evaluated the prevalence of culture-positive tuberculosis in HIV-infected inpatients with WHO danger signs and negative sputum smears, reporting 23% in a Ugandan study [6] (51% of whom had WHO danger signs) and 23% in a study from South Africa [5], which is similar to the 22.7% culture-positive prevalence in our participants with negative sputum smears, suggesting that our findings are generalizable to sub-Saharan Africa. Tuberculosis is the leading cause of hospitalization among HIV-infected adults worldwide, with similar proportions in low- to middle-income countries, and it is thus conceivable that our findings may also be generalizable outside of Africa [29]. Second, all our participants had 1 or more WHO danger signs, which limited our ability to assess their value for the diagnosis of tuberculosis. Third, a specialist radiologist’s assessments of the chest radiographs performed well in our clinical prediction rule, but in resource-constrained settings nonspecialist doctors usually read chest radiographs and their interpretation is likely to be less accurate. Fourth, the diagnoses of bacterial pneumonia and PJP were clinical and not based on microbiological reference standards. We attempted to confirm bacterial pneumonia with blood and sputum cultures, but these were almost all negative because of antibiotic use immediately before inclusion into our study. Fifth, the reference standard for diagnosing tuberculosis was culture, which is not 100% sensitive. Finally, while bootstrap resampling has several advantages over other internal validation methods, it is not enough to demonstrate the external applicability of the derived prediction rule. Because internal validation in general is optimistic, a drop in the performance of the prediction rule could be observed when it is applied to different settings. While our measure of optimism suggests that such a drop would be marginal, external validation, ideally conducted by different investigators, is needed to confirm the performance of our prediction rule before wide uptake in routine practice. Strengths of

Table 5. Diagnostic Accuracy Assessment of a Clinical Prediction Rule for the Diagnosis of Culture-Positive Tuberculosis Among 484 Seriously Ill Human Immunodeficiency Virus–Infected Participants Presenting With a Cough of Any Duration and ≥1 World Health Organization Danger Sign

Total Score	Probability of Tuberculosis	Sensitivity	Specificity	Positive Likelihood Ratio	Negative Likelihood Ratio
≤1	0.1%	100%	0.0%	1	...
2	0.2%	95.2%	28.8%	1.34	0.17
3	0.5%	92.2%	45.4%	1.69	0.17
4	1.1%	87.2%	59.3%	2.14	0.23
5	2.5%	78.6%	67.7%	2.43	0.31
6	5.6%	66.0%	80.3%	3.36	0.42
7	12.2%	50.8%	90.5%	5.35	0.54
8	24.4%	39.1%	93%	5.58	0.65
9	43.0%	31.2%	96.4%	8.67	0.71
10	63.7%	19.5%	97.9%	9.13	0.82
11	80.4%	5.9%	99.6%	13.44	0.95
12	90.5%	0.2%	100%	...	1

Table 6. Univariate Analysis of Individual Chest Radiograph Variables as Predictors of Culture-Positive Tuberculosis, Among the 416 Seriously Ill Human Immunodeficiency Virus-Infected Participants Presenting With a Cough of Any Duration and ≥ 1 World Health Organization Danger Sign and With a Specialist Radiologist Report

Chest Radiograph Variables	no./No. (%)	Crude OR	(95% CI)	Wald P Value
Consolidation	217/413 (52.5)	0.86	(.57–1.29)	.456
Diffuse micronodular infiltration	29/416 (6.9)	6.45	(2.20–18.87)	.001
Linear/reticulonodular infiltration	226/415 (54.5)	0.93	(.63–1.37)	.713
Enlarged hilar or mediastinal lymph nodes	210/414 (50.7)	2.34	(1.58–3.47)	<.001
Pleural effusion	195/414 (47.1)	1.24	(.84–1.82)	.284
Interstitial infiltration	259/415 (62.4)	0.92	(.62–1.37)	.685
Nodularity (>3 mm)	217/415 (52.3)	2.21	(1.49–3.28)	<.001
Cavitation	249/415 (60.0)	1.01	(.68–1.49)	.969

Abbreviations: CI, confidence interval; OR, odds ratio.

our study are its prospective design, the use of multiple mycobacterial cultures and sputum induction to establish a reference standard for tuberculosis, and a higher than expected number of tuberculosis events, which increased our power to develop a clinical prediction rule.

In conclusion, we developed a clinical prediction rule with good performance characteristics for diagnosing tuberculosis in seriously ill HIV-infected patients and showed that the Xpert MTB/RIF assay had high sensitivity. These findings, if validated in different settings, could contribute to the development of an improved algorithm for tuberculosis in seriously ill inpatients.

Supplementary Data

Supplementary materials are available at *Clinical Infectious Diseases* online. Consisting of data provided by the authors to benefit the reader, the posted materials are not copyedited and are the sole responsibility of the authors, so questions or comments should be addressed to the corresponding author.

Notes

Author contributions. G. M., M. M., M. N., and M. X. R. designed the study. R. G., H. v. d. P., and W. S. collected the data. G. M., M. M., R. G., A. S., and A. P. K. analyzed and interpreted the data. R. G. wrote the first draft. All authors reviewed, revised, and approved the final report.

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