

Induction chemoradiotherapy versus induction chemotherapy for potentially resectable stage IIIA (N2) non-small cell lung cancer: a systematic review and meta-analysis

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Background: Non-small cell lung cancer (NSCLC) accounts for 85% to 90% of lung cancer cases. At diagnosis, around 30% of NSCLC patients are already at stage IIIA (N2). One standard treatment for this stage is induction chemotherapy followed by surgery, whether induction chemoradiotherapy is superior to induction chemotherapy remains uncertain. We therefore performed a systematic review and meta-analysis of published randomized control trials to evaluate the therapeutic efficacy and toxicity of induction chemoradiotherapy versus induction chemotherapy for potentially resectable stage IIIA (N2) NSCLC.

Methods: We systematically searched for relevant studies in PubMed, Embase, Web of Science and Cochrane Library from the inception of each database to September 10, 2017. The primary endpoints were objective response rate (ORR), pathological complete response (pCR) rate of mediastinal lymph nodes, toxicity (grade 3–4 adverse events, i.e., nausea and vomiting, infections, leukopenia and anemia), overall survival (OS) and progression-free survival (PFS). Statistical analyses were performed using Review Manager v5.3.

Results: Four studies, containing 461 patients in total, were included for meta-analysis. Our analyses suggest that compared with induction chemotherapy, induction chemoradiotherapy improved ORR [odds ratio (OR) =1.97, 95% confidence interval (CI): 1.25–3.10, P<0.05] and pCR rate of mediastinal lymph nodes (OR =1.97, 95% CI: 1.00–3.86, P=0.05); but it did not significantly improve OS [hazard ratio (HR) =0.91, 95% CI: 0.73–1.14, P=0.42] or PFS (HR =1.01, 95% CI: 0.81–1.26, P=0.91); also it did not exacerbate the toxicity.

Conclusions: Induction chemoradiotherapy may have limited value concerning tumor response and pCR of mediastinal lymph nodes. However, current evidence does not support that addition of radiotherapy to induction chemotherapy followed by surgery can bring significant benefits to operable stage IIIA (N2) NSCLC patients. More studies are required to draw a better conclusion.

Keywords: Non-small cell lung cancer (NSCLC); induction chemotherapy; induction chemoradiotherapy; meta-analysis; IIIA (N2)

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Introduction

Nowadays, lung cancer is the most common cancer diagnosed worldwide and the leading cause of cancer-related deaths (1). Non-small cell lung cancer (NSCLC) accounts for 85% to 90% of lung cancer cases (2). At diagnosis, around 30% of NSCLC patients are already at stage IIIA (N2). Currently, definitive concurrent chemoradiation remains as the main treatment strategy for stage IIIA NSCLC. However, for potentially resectable stage IIIA (N2) NSCLC, induction therapy (also known as neoadjuvant therapy) followed by surgery is recommended as an alternative. Induction therapy includes chemotherapy alone, sequential chemoradiation, concurrent chemoradiation and concurrent chemoradiation after chemotherapy. For stage IIIA (N2) NSCLC, several studies have shown that induction chemotherapy followed by surgery improved survival, compared with surgery alone (3-5). Meanwhile, induction by concurrent chemoradiotherapy followed by surgery resulted in 5-year survival rates of 30% to 40%, appearing superior to surgery alone (6-8). However, a consensus has not been reached on which induction therapy should be administered to stage IIIA (N2) patients—50% of the National Comprehensive Cancer Network (NCCN) member institutions choose induction chemoradiotherapy, while another 50% choose induction chemotherapy (9).

A recently published phase III randomized clinical trial evaluated the survival benefit of induction chemoradiotherapy versus induction chemotherapy followed by surgery for stage IIIA (N2) patients, showing that radiotherapy did not add any benefit. The authors suggest that induction chemotherapy followed by surgery is an adequate treatment option for stage IIIA (N2) NSCLC (10). To incorporate currently available data to better evaluate the efficacy and toxicity of addition of radiotherapy to induction chemotherapy followed by surgery for stage IIIA (N2) NSCLC patients, we carried out a systematic review and meta-analysis.

Methods

Eligibility criteria

The following criteria for eligibility were set before searching for studies: (I) patients had pathologically proven, locally advanced T1-3N2M0, stage IIIA (N2) NSCLC based on the TNM classification; (II) the trials were randomized clinical trials of induction chemoradiotherapy versus induction chemotherapy followed by surgery; (III) studies reported any of the following information: tumor

responses, pathological complete response (pCR) of mediastinal lymph nodes, adverse reactions, overall survival (OS) and progression-free survival (PFS); (IV) the full-text articles were published in English. The exclusion criteria were: (I) letters, meeting abstracts, reviews and case reports; (II) duplicate reports of the same patient cohorts; (III) studies without useful data or endpoints.

Search strategy

A comprehensive literature search was conducted to identify all relevant studies. PubMed, Embase, Web of Science and Cochrane Library were systematically searched for studies published from the inception of each database to September 10, 2017. The following key words were used in the search strategy: “non-small cell lung cancer” AND “IIIA (N2)” AND “induction therapy”. In addition, the search was extended to identify any other potential studies by review of references of articles included in the final selection.

Data extraction

We used the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) statement to report the materials and methods (11). Two investigators independently searched for eligible studies and extracted the data. Disagreements between investigators on data extraction were resolved through discussion with a third investigator. For the meta-analysis, the following data were extracted: objective response rate (ORR), pCR rate of mediastinal lymph nodes, incidence of grade 3–4 adverse events (nausea and vomiting, infections, leukopenia and anemia), the hazard ratio (HR) with a 95% confidence interval (CI) for OS and PFS. Meanwhile, the following information of the eligible studies was recorded: author, year, study period, groups, number of patients and treatment strategies.

Risk of bias assessment

Assessment of risk of bias was performed according to the Cochrane Collaboration's tool (12). Two investigators evaluated risk of bias of all eligible studies by random sequence generation (selection bias), allocation concealment (selection bias), blinding of participants and personnel (performance bias), blinding of outcome assessment (detection bias), incomplete outcome data (attrition bias), selective reporting (reporting bias) and other bias. Each

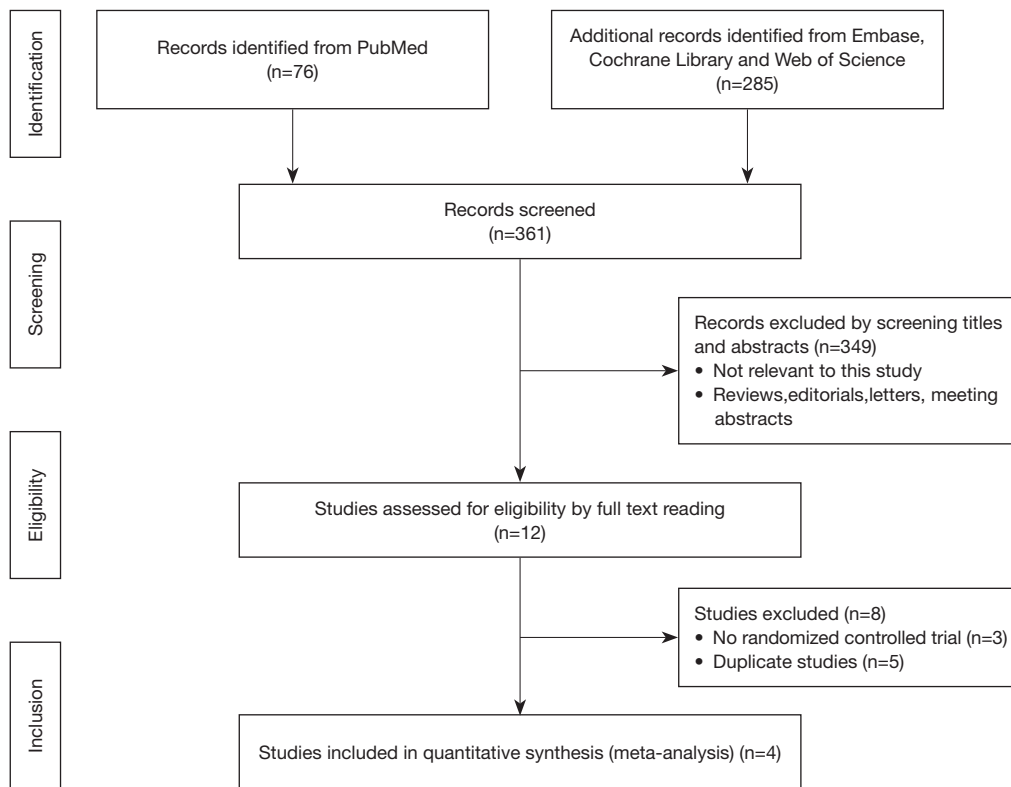


Figure 1 Flow diagram of study search and selection.

domain of bias was classified as “high risk”, “unclear risk” or “low risk”. The risk of bias of a trial was rated “high” when any of the key domains was at high risk of bias; conversely, it was rated “low” when all the key domains were at low risk of bias; otherwise, it was rated “unclear”.

Statistical analyses

Statistical analyses were performed using Review Manager v5.3 (RevMan, The Cochrane Collaboration). Primary outcome measures were odds ratios (ORs) for ORR, risk ratios (RRs) for grade 3–4 adverse events, HRs for OS and PFS. 95% CI was used for each endpoint. P value of <0.05 was considered significant. Statistical heterogeneity was evaluated using the Cochrane χ^2 test and the I^2 statistic. $I^2 < 50\%$ was considered as low heterogeneity, whereas $I^2 \geq 50\%$ was considered as potentially significant heterogeneity. The meta-analysis was performed either via a random effects model or a fixed effects model according to the degree of heterogeneity. A random effects model was used in case of potentially significant heterogeneity ($P < 0.10$, $I^2 \geq 50\%$); otherwise, a fixed effects model was used.

Results

Study selection

As summarized in *Figure 1*, 361 records of relevant studies were obtained from PubMed ($n=76$) and other databases ($n=285$). Twelve studies passed the screening of title and abstract. After full text screening, eight studies were excluded due to nonrandomized controlled trials or duplicate data. Finally, four randomized controlled trials, containing 461 patients, were included in this meta-analysis (10,13–15).

Characteristics of the included studies

These studies were conducted in Japan, Germany, Switzerland, Serbia and France and published between 2008 and 2015. All the IIIA (N2) NSCLC patients were divided into intervention groups (induction chemoradiotherapy followed by surgery, ChRT+S) and control groups (induction chemotherapy followed by surgery, ChT+S). The characteristics and the treatment schedules were listed in *Table 1*.

Table 1 Characteristics of included studies

Author, years	Study period	Group	Number of patients	Treatment strategy
Pless <i>et al.</i> 2015 (10)	2001–2012	ChRT+S	117	Three cycles of neoadjuvant chemotherapy (100 mg/m ² cisplatin and 85 mg/m ² docetaxel) followed by radiotherapy with 44 Gy in 22 fractions over 3 weeks. Surgery was scheduled 21–28 days after completion of radiotherapy
		ChT+S	115	Three cycles of 100 mg/m ² intravenous cisplatin and 85 mg/m ² docetaxel were given every 3 weeks. Surgery was scheduled 21 days after the last chemotherapy cycle
Katakami <i>et al.</i> 2012 (14)	2000–2006	ChRT+S	29	Two cycles of Docetaxel (60 mg/m ²) and carboplatin (AUC =5) plus concurrent radiation therapy (40 Gy in 20 fractions of 2 Gy over 4 weeks) followed by surgery
		ChT+S	29	Two cycles of Docetaxel (60 mg/m ²) and carboplatin (AUC =5) followed by surgery
Girard <i>et al.</i> 2010 (13)	2003–2007	ChRT+S	32*	One cycle of cisplatin (80 mg/m ²) and vinorelbine (25 mg/m ²) or carboplatin (AUC =6) and paclitaxel (200 mg/m ²) followed by radiation (46 Gy) with concurrent cisplatin (80 mg/m ²) and vinorelbine (15 mg/m ²) or carboplatin (AUC =2) and paclitaxel (40 mg/m ²), and then surgical resection
		ChT+S	14	Three cycles of cisplatin (80 mg/m ²) and gemcitabine (1,250 mg/m ²) followed by surgery
Thomas <i>et al.</i> 2008 (15)	1995–2003	ChRT+S	55	Three cycles of cisplatin (55 mg/m ²) and etoposide (100 mg/m ²) followed by radiation (45 Gy) with concurrent carboplatin (100 mg/m ²) and vindesine (3 mg absolute), and then surgical resection
		ChT+S	70	Three cycles of cisplatin (55 mg/m ²) and etoposide (100 mg/m ²) followed by surgical resection

*, arm C (15 patients) was not chosen into meta-analysis of overall survival due to lack of available survival data. ChRT+S, chemoradiotherapy + surgery; ChT+S, chemotherapy + surgery; Gy, Gray; AUC, area under curve.

Risk of bias in included studies

The quality of the studies was evaluated by risk of bias assessment. Two studies described the methods of generating random assignment sequence in detail (10,15), while the other two did not (13,14). Allocation concealment was performed adequately in all four studies. Blinding of participants and personnel was difficult to achieve due to the nature of these studies, of which the risk of bias was high. Blinding of outcome assessment was described in two studies (10,13), but not in other two studies (14,15). The outcome data of all four studies were complete and objective, and the risk of attrition bias, as well as selective reporting bias, was low. Other potential biases such as specific study design and fraudulent action were not clear in all the studies. Overall, the methodological quality of these studies was good. The details of risk of bias assessment were summarized in *Figure 2*.

Tumor responses

WHO criteria were used to assess tumor responses (16). ORR after induction treatment was reported in three studies (10,13,14). A fixed effects model was adopted, as no significant heterogeneity was observed ($\chi^2=2.75$, $P=0.25$, $I^2=27\%$). The pooled ORR in the intervention group was significantly higher than that in the control group (OR =1.97, 95% CI: 1.25–3.10, $P<0.05$) (*Figure 3A*).

pCR of mediastinal lymph nodes

Three studies provided data for pCR of mediastinal lymph nodes (10,13,14). A fixed effect model was used, as no significant heterogeneity was observed ($\chi^2=1.47$, $P=0.48$, $I^2=0\%$). The meta-analysis revealed an improved pCR rate of mediastinal lymph nodes in the intervention group (OR =1.97, 95% CI: 1.00–3.86, $P=0.05$) (*Figure 3B*).

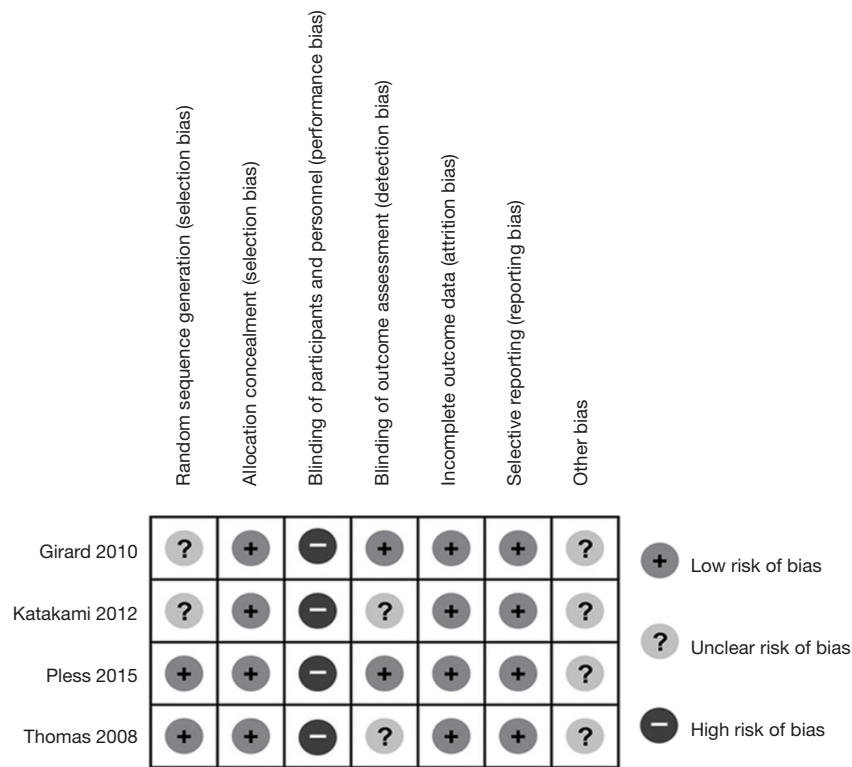


Figure 2 Summary of risk of bias assessment.

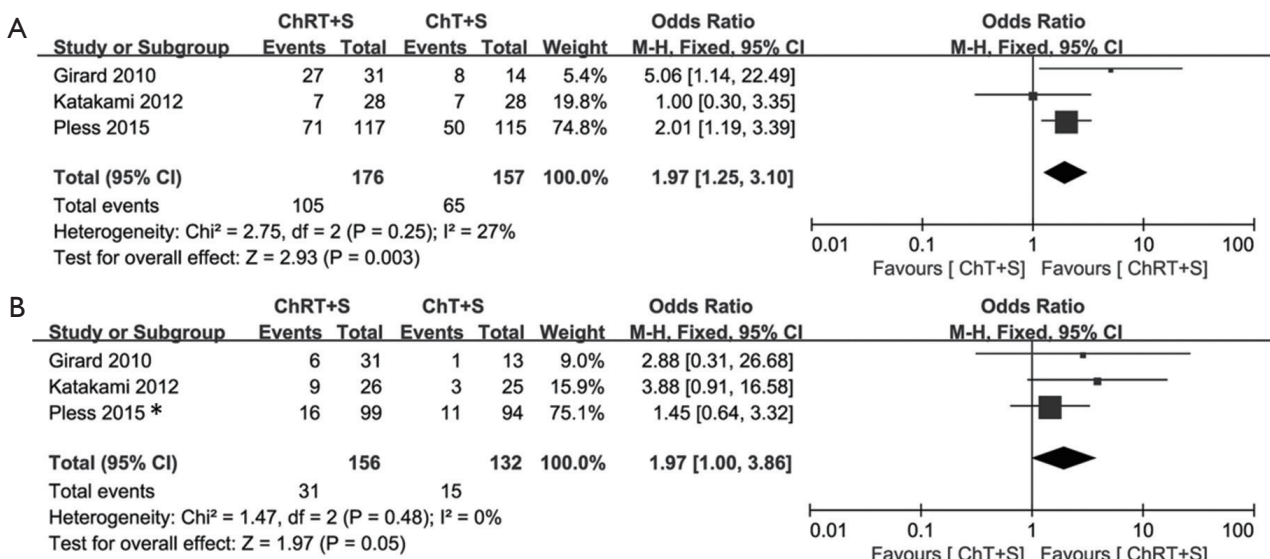


Figure 3 Meta-analysis of tumor response and pCR of mediastinal lymph nodes in induction chemoradiotherapy vs. induction chemotherapy. (A) ORR; (B) pCR rate of mediastinal lymph nodes. *, pCR here referred to as the pCR of tumors and lymph nodes, because the data for pCR of lymph nodes were not provided. ORR, objective response rate; pCR, pathological complete response; ChRT+S, chemoradiotherapy + surgery; ChT+S, chemotherapy + surgery.

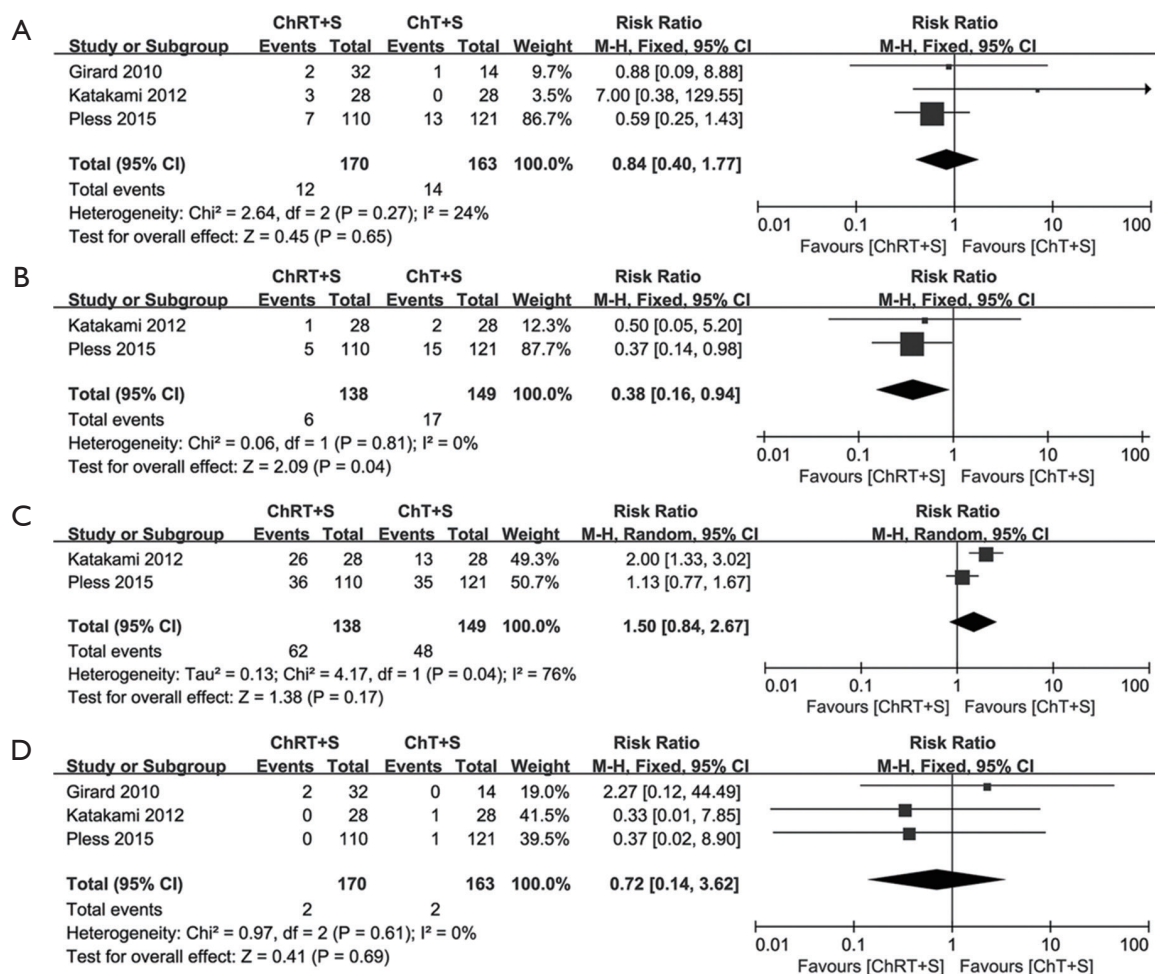


Figure 4 Meta-analysis of adverse reactions in induction chemoradiotherapy *vs.* induction chemotherapy. (A) Nausea and vomiting; (B) infections; (C) leukopenia; (D) anemia. ChRT+S, chemoradiotherapy + surgery; ChT+S, chemotherapy + surgery.

Adverse reactions

Nausea and vomiting, infections, leukopenia and anemia were included as adverse reactions in this meta-analysis. We focused on grade 3–4 adverse reactions, as they often require intervention in clinical practice. Grade 3–4 nausea and vomiting data from three studies (10,13,14) were incorporated into this meta-analysis. No significant heterogeneity was observed ($\chi^2 = 2.64$, $P = 0.27$, $I^2 = 24\%$). There was no significant difference between the two groups (RR = 0.84, 95% CI: 0.40–1.77, $P = 0.65$) (Figure 4A). Grade 3–4 infections data from two studies (10,14) were incorporated into this meta-analysis. No significant heterogeneity was observed ($\chi^2 = 0.06$, $P = 0.81$, $I^2 = 0\%$). It appears that the incidence of infections in the intervention group was lower than that in the control group (RR = 0.38,

95% CI: 0.16–0.94, $P = 0.04$) (Figure 4B); however, this was very likely an artifact (see in discussion). Grade 3–4 leukopenia data from two studies (10,14) were incorporated into this meta-analysis. A random effects model was adopted due to statistically significant heterogeneity ($\chi^2 = 4.17$, $P = 0.04$, $I^2 = 76\%$). There was no significant difference between the two groups (RR = 1.50, 95% CI: 0.84–2.67, $P = 0.17$) (Figure 4C). Grade 3–4 anemia data from three studies (10,13,14) were incorporated into this meta-analysis. No significant heterogeneity was observed ($\chi^2 = 0.97$, $P = 0.61$, $I^2 = 0\%$). There was no significant difference between the two groups (RR = 0.72, 95% CI: 0.14–3.62, $P = 0.69$) (Figure 4D).

Survival

Survival was analyzed based on OS and PFS. OS was

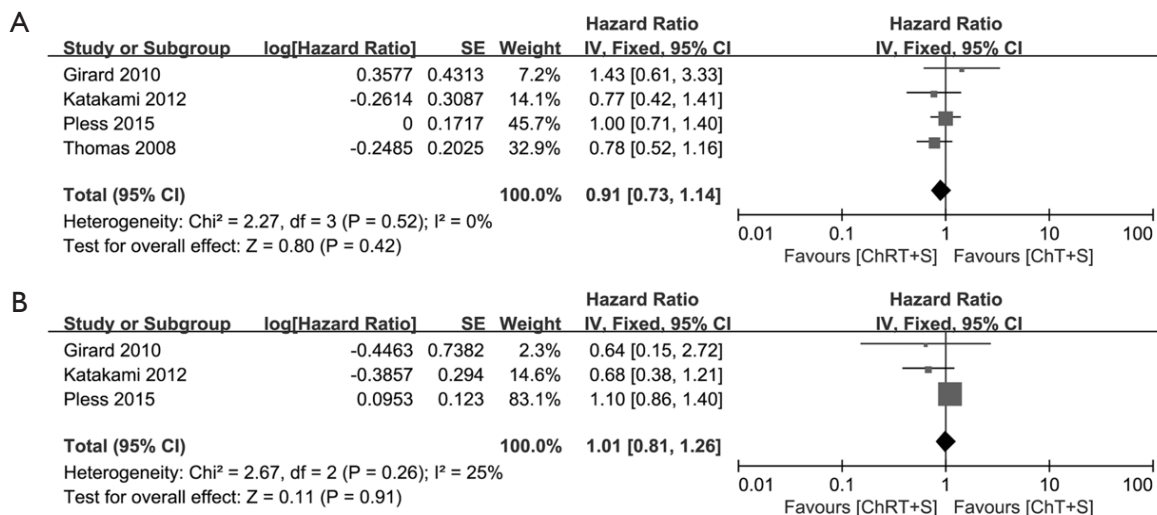


Figure 5 Meta-analysis of survival data in induction chemoradiotherapy vs. induction chemotherapy. (A) OS; (B) PFS. OS, overall survival; PFS, progression-free survival; ChRT+S, chemoradiotherapy + surgery; ChT+S, chemotherapy + surgery.

reported in all four studies with a total of 446 patients (10,13-15), no significant heterogeneity was observed ($\chi^2=2.27$, $P=0.52$, $I^2=0\%$), and there was no significant difference concerning OS between the two groups (HR = 0.91, 95% CI: 0.73–1.14, $P=0.42$) (Figure 5A). PFS from three studies (10,13,14) were incorporated into the pooled analysis, no significant heterogeneity was observed ($\chi^2=2.67$, $P=0.26$, $I^2=25\%$), and also there was no significant difference concerning PFS between the two groups (HR = 1.01, 95% CI: 0.81–1.26, $P=0.91$) (Figure 5B). Of note, a fixed effects model was adopted in all the meta-analysis for survival data.

Discussion

Over the past decades, several studies have demonstrated that preoperative induction treatment plus surgery, compared with surgery alone, significantly improved the survival of stage IIIA (N2) NSCLC patients (3-5,17). However, the optimal induction treatment for potentially resectable stage IIIA (N2) NSCLC remains unclear. Particularly, whether induction chemoradiotherapy is superior to induction chemotherapy has not been clearly demonstrated.

In this work, we attempted to resolve this issue by a systematic review and meta-analysis, which may be helpful for clinicians to determine the optimal induction treatment. We compared induction chemoradiotherapy with induction chemotherapy from four aspects: tumor responses,

pCR of mediastinal lymph nodes, adverse reactions and survival data. Our meta-analysis suggests that induction chemoradiotherapy indeed improved ORR and pCR rate of mediastinal lymph nodes. Unexpectedly, our meta-analysis indicates that induction chemoradiotherapy decreased the incidence of grade 3–4 infection. However, after having carefully inspected the two studies, we conclude that this result was not credible. Because in the WJTOG9903 study (14), wherein concurrent chemoradiotherapy was performed in the intervention group, there was no significant difference in the incidence of grade 3–4 infection between the two groups; while in the SAKK trial (10), radiotherapy alone actually did not increase the risk of grade 3–4 infection, thereby mitigating the incidence of infection in the intervention group. Importantly, our meta-analysis indicates that induction chemoradiotherapy did not result in an improved OS or PFS, compared with induction chemotherapy. A similar conclusion has been reached by two previous meta-analyses (18,19), which included different trials thus different patient cohorts, and two retrospective studies (9,20).

Of note, in the most weighted SAKK trial (10), before surgery, sequential chemoradiotherapy was conducted, whereas concurrent chemoradiotherapy was conducted in the other three trials (10,13-15). We also performed the meta-analyses without the SAKK trial to evaluate concurrent chemoradiotherapy versus chemotherapy, while the conclusions on pCR rate of mediastinal lymph nodes, OS and PFS remain unchanged, concurrent

chemoradiotherapy showed a trend to improve ORR, but the difference did not reach statistical significance, due to only two trials with small cohorts and inconsistent data on this issue were included (Figure S1-S3). Meta-analysis for the incidence of grade 3–4 infection or leukopenia was not conducted, due to only one trial could be included.

Compared with two previous meta-analyses (18,19), our meta-analysis was completely based on high-quality randomized clinical trials and included the latest study not available for previous meta-analysis. Moreover, unlike previous meta-analyses that only included survival analysis, our meta-analysis also included tumor responses, pCR of mediastinal lymph nodes and adverse reactions.

Meanwhile, our meta-analysis also has several drawbacks. First, blinding method was not adopted in all four randomized clinical trials due to practical difficulties, therefore performance bias was inevitable. Second, the differences in the regimens/protocols of radiotherapy and/or chemotherapy in all included studies might lead to certain degree of clinical heterogeneity, which might affect the outcome of the analysis. Third, only a small number of studies were included in the meta-analysis, and some studies had relatively small sample sizes, therefore weakened the strength of the meta-analysis.

More importantly, the conclusions of our meta-analysis are dictated by the quality of the included data. In the SAKK trial by Pless *et al.* (10), a high percentage of the patients (~16% of each arm) did not receive the allocated treatment. Also, the high weight of this trial has a great impact on the results of our meta-analysis. In the WJTOG9903 study by Katakami *et al.* (14), the insufficient radiation dose and the short interval between induction therapy and surgery resulted in a low response rate to induction chemoradiotherapy (~25% *vs.* ~65% in average in other concurrent chemoradiotherapy studies) (6–8). In the French IFTC-0101 trial by Girard *et al.* (13), median survival was not reached in arm C (one of the two induction chemoradiotherapy groups) due to insufficient follow-up time, which could have affected the OS analysis. In the GLCCG trial by Thomas *et al.* (15), both stage IIIA and IIIB NSCLC patients were included, the endpoints were not reported separately, precluding this trial from most of our analyses, despite it contained the largest cohorts.

The reason why addition of radiotherapy did not result in an improved OS or PFS may lie in the setting that only operated patients were taken into account for the analysis in both the original trials and the meta-analysis, the initial benefit from radiotherapy was likely mitigated by the

following surgery. However, in reality, not all the potentially resectable stage IIIA (N2) NSCLCs turn out to be operable after the induction therapy; also, not all the surgery can achieve a complete resection of the tumors. Addition of radiotherapy may still be beneficial for such patients. Thus, before treatment, it is essential to carefully evaluate the disease stage and to discuss whether surgery is appropriate by a multidisciplinary team (21,22). If the confidence for a complete resection is high, we do not recommend including radiotherapy in the induction therapy; if the situation is opposite, radiotherapy should also be considered. More studies are required to draw a better conclusion.

Conclusions

In conclusion, induction chemoradiotherapy may have limited value concerning tumor response and pCR of mediastinal lymph nodes. Compared with induction chemotherapy, induction chemoradiotherapy does not exacerbate the toxicity. Current evidence does not support that addition of radiotherapy to induction chemotherapy followed by surgery can bring a significantly prolonged OS or PFS to operable stage IIIA (N2) NSCLC patients.

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Footnote

Conflicts of Interest: The authors have no conflicts of interest to declare.

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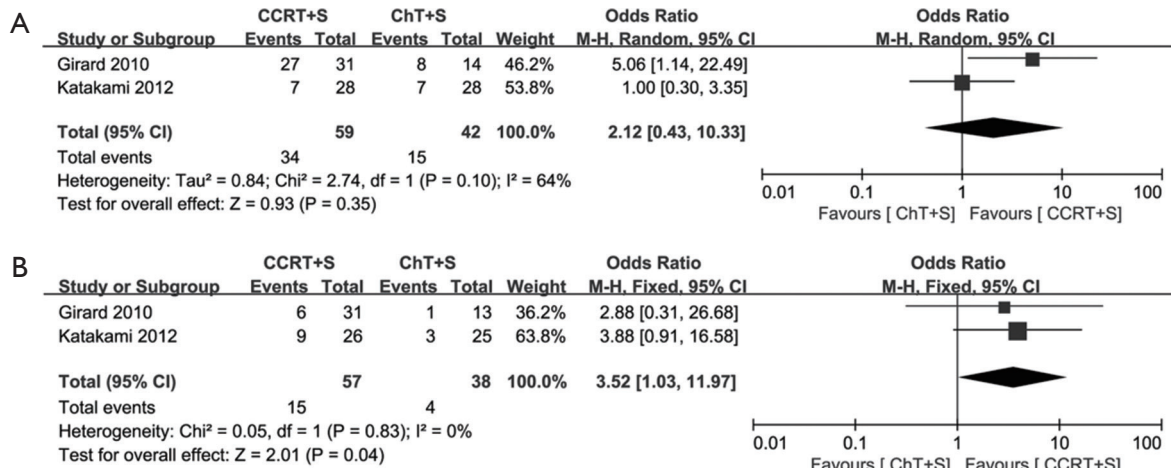


Figure S1 Meta-analysis of tumor response and pCR of mediastinal lymph nodes in induction concurrent chemoradiotherapy *vs.* induction chemotherapy. (A) ORR; (B) pCR of mediastinal lymph nodes. ORR, objective response rate; pCR, pathological complete response; CCRT+S, Concurrent chemoradiotherapy + surgery; ChT+S, chemotherapy + surgery.

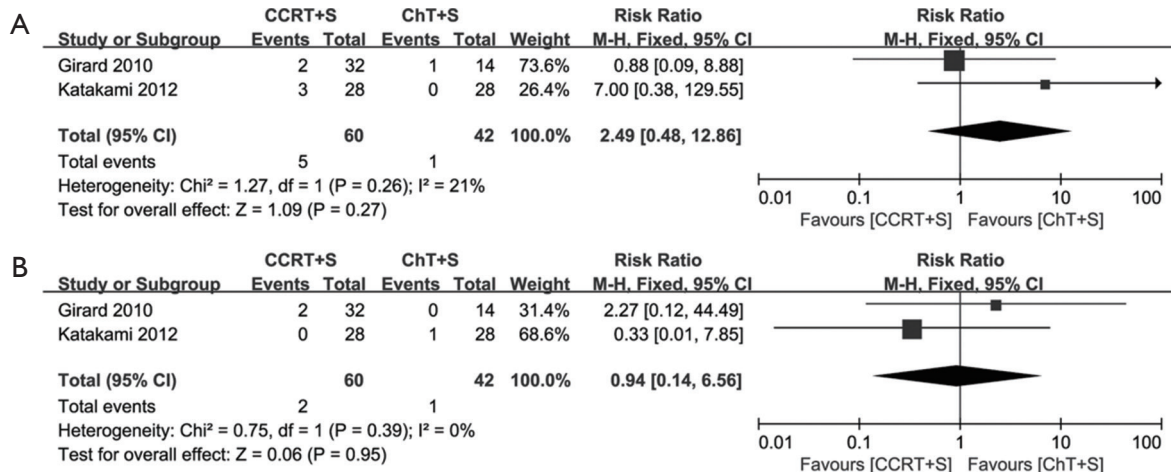


Figure S2 Meta-analysis of adverse reactions in induction concurrent chemoradiotherapy *vs.* induction chemotherapy. (A) Nausea and vomiting; (B) anemia. CCRT+S, concurrent chemoradiotherapy + surgery; ChT+S, chemotherapy + surgery.

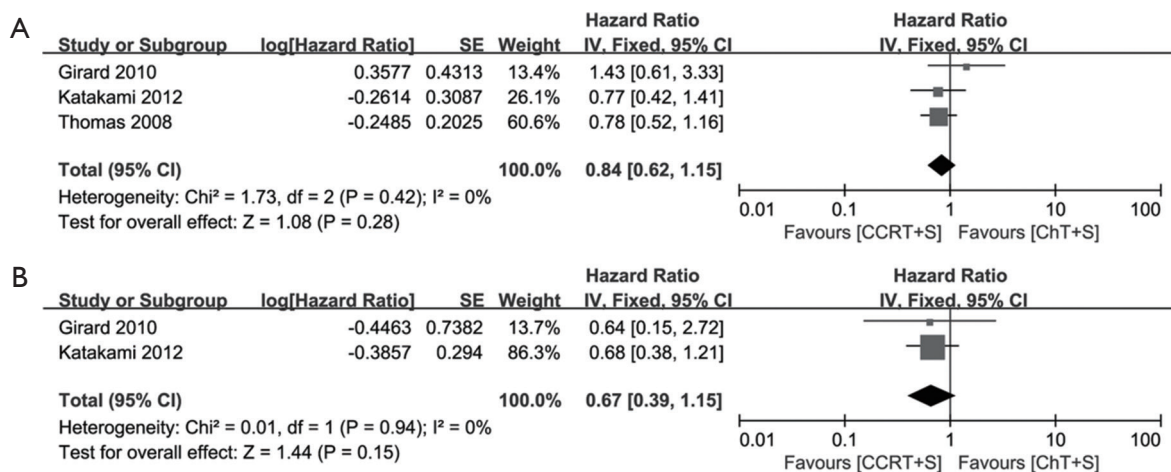


Figure S3 Meta-analysis of survival data in induction concurrent chemoradiotherapy *vs.* induction chemotherapy. (A) OS; (B) PFS. OS, overall survival; PFS, progression-free survival; CCRT+S, concurrent chemoradiotherapy + surgery; ChT+S, chemotherapy + surgery.