

Original Contribution

Lack of Association Between Maternal or Neonatal Vitamin D Status and Risk of Childhood Type 1 Diabetes: A Scandinavian Case-Cohort Study

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Studies on vitamin D status during pregnancy and risk of type 1 diabetes mellitus (T1D) lack consistency and are limited by small sample sizes or single measures of 25-hydroxyvitamin D (25(OH)D). We investigated whether average maternal 25(OH)D plasma concentrations during pregnancy are associated with risk of childhood T1D. In a case-cohort design, we identified 459 children with T1D and a random sample (n = 1,561) from the Danish National Birth Cohort (n = 97,127) and Norwegian Mother and Child Cohort Study (n = 113,053). Participants were born between 1996 and 2009. The primary exposure was the estimated average 25(OH)D concentration, based on serial samples from the first trimester until delivery and on umbilical cord plasma. We estimated hazard ratios using weighted Cox regression adjusting for multiple confounders. The adjusted hazard ratio for T1D per 10-nmol/L increase in the estimated average 25(OH)D concentration was 1.00 (95% confidence interval: 0.90, 1.10). Results were consistent in both cohorts, in multiple sensitivity analyses, and when we analyzed mid-pregnancy or cord blood separately. In conclusion, our large study demonstrated that normal variation in maternal or neonatal 25(OH)D is unlikely to have a clinically important effect on risk of childhood T1D.

adolescent; child; diabetes mellitus, type 1; etiology; immunology; vitamin D

Abbreviations: 25(OH)D, 25-hydroxyvitamin D; BMI, body mass index; DNBC, Danish National Birth Cohort; MoBa, Norwegian Mother and Child Cohort Study; T1D, type 1 diabetes mellitus.

Type 1 diabetes mellitus (T1D) is a chronic autoimmune disease with severe long-term complications (1). There has been a marked increase in the incidence of childhood T1D worldwide during the last 4 decades (2). Genetic predisposition combined with unknown environmental factors early in life are thought to trigger a loss of self-tolerance for the insulin-producing pancreatic β -cells (3, 4).

Maternal vitamin D status during pregnancy is critical for determining fetal 25-hydroxyvitamin D (25(OH)D) concentration (5). The fetus may regulate the concentrations of both 25(OH)D and the bioactive metabolite 1,25-dihydroxyvitamin D from an early stage, suggesting an important evolutionary role for vitamin D metabolites during pregnancy. In addition, the role of vitamin D in the fetus may not be restricted to the development of healthy bones (6, 7). Experimental studies using primarily animals and in vitro human immune cell lines have demonstrated that vitamin D is involved in maintaining immunological self-tolerance (8, 9).

An inverse association was reported between a high dose of vitamin D supplements in the first year of life and the risk of childhood T1D (10). However, only 2 studies have investigated the relationship between maternal 25(OH)D concentrations during pregnancy and the risk of childhood T1D, with inconsistent results (11, 12). Partially inconsistent results were obtained from 2 additional studies that investigated the association between $25(OH)D_3$ measured in neonatal dried blood spots and the risk of childhood T1D (13, 14). These variations could be due to methodological issues including single measurements of 25(OH)D, the lack of data on potential confounders, or limited sample sizes. We tested the hypothesis that there is an association between maternal

vitamin D status and the risk of childhood T1D, using a series of 25(OH)D measurements in samples taken from early in pregnancy through until delivery, in 2 of the largest cohorts of pregnant women in the world. A secondary aim was to examine whether maternal vitamin D supplements taken during pregnancy influenced childhood T1D risk.

METHODS

Overview of study design

This binational study consists of case-cohort samples from the Danish National Birth Cohort (DNBC) and the Norwegian Mother and Child Cohort Study (MoBa), which are prospective, population-based pregnancy cohort studies conducted by the Statens Serum Institut in Denmark and the Norwegian Institute of Public Health, respectively. In DNBC, pregnant women were recruited across Denmark between 1996 and 2002. Approximately 50% of all general practitioners participated in the recruitment process, and 60% of women invited agreed to participate (15). In MoBa, pregnant women were recruited across Norway between 1999 and 2008, and 41% of eligible women participated (16).

Study sample and identification of T1D

We linked the cohorts with the Danish Childhood Diabetes Registry and the Norwegian Diabetes Childhood Registry to identify children who had developed T1D according to the World Health Organization's criteria (17, 18). These diabetes registers have nearly complete nationwide coverage and record highquality prospective data on children with T1D. We included all 459 children diagnosed with T1D (270 from DNBC and 189 from MoBa) who had available blood samples, as well as a random cohort sample of 1,561 children (985 from the DNBC and 576 from the MoBa) from 97,127 (DNBC) and 113,053 (MoBa) eligible children in our study population (Figure 1).

All participants had a minimum of 1 plasma sample assayed for 25(OH)D, and 91% of the mother/child pairs were represented by 2 or 3 blood samples.

Exposure assessment

Collection and storage of blood samples. For DNBC, maternal venous blood was drawn at approximately weeks 7–9 and 24–25 of pregnancy and from the umbilical cords of newborn infants (15). For MoBa, maternal venous blood was drawn at approximately week 17–18 of pregnancy, shortly after delivery, and from the umbilical cords of newborn infants (19, 20). DNBC plasma samples were stored at -20° C or in liquid nitrogen (15). MoBa plasma samples were stored at -80° C (20).

Assessment of vitamin D status. We used liquid chromatography–tandem mass spectrometry to measure plasma $25(OH)D_2$ and $25(OH)D_3$ separately (see Web Appendix 1 (available at https://academic.oup.com/aje) for details). Our exposure variable was defined as the sum of $25(OH)D_2$ and $-D_3$, hereafter referred to as 25(OH)D. All samples were assayed by 2 technicians in a single laboratory at the Statens Serum Institute (Copenhagen, Denmark) during July–October 2015. All samples were processed in random order (i.e., independently of their cohort or case status), and the technicians were blinded to the case status of the samples. Repeated measurements of standards gave interassay coefficients of variation for $25(OH)D_3$ of 3.4% and 7.9% for concentrations of approximately 33 nmol/L and 80 nmol/L, respectively.

Other variables. Birth weight, maternal age at delivery, and mode-of-delivery details were obtained from the nationwide Medical Birth Registry of Norway and the National Hospital Discharge Registry in Denmark (15, 16). Information regarding maternal prepregnancy body mass index (BMI, calculated as weight $(kg)/height (m)^2$ and smoking during pregnancy was obtained from telephone interviews (DNBC) and questionnaires administered mid-pregnancy (MoBa). Information on intake of vitamin D, eicosapentaenoic acid (EPA), and docosahexaenoic acid (DHA) from supplements was obtained from food frequency questionnaires administered during the second trimester in both cohorts (15, 19). See Web Appendix 1 for details on questionnaires. For DNBC, information regarding any type of maternal diabetes was obtained from the Danish National Diabetes Register. For MoBa, data on maternal T1D were obtained from questionnaires and the Norwegian Patient Registry.

Statistical analysis

Statistical analyses were performed with R, version 3.3.1 (R Foundation for Statistical Computing, Vienna, Austria), using the *lava* package (version 1.4.5) and the *survey* package (version 3.31).

All details of the analysis plan were determined a priori. The primary analysis was a 2-stage analysis, using all available 25(OH)D measurements to estimate the hazard ratio of childhood T1D per 10-nmol/L increase in estimated average 25(OH)D concentration during pregnancy and at birth. The first stage generated a measurement error (structural equation) model for the estimated average 25(OH)D concentration (i.e., "latent variable"). Concentrations of 25(OH)D were adjusted for the time of year (season) of blood sampling using cosinor modeling (21). In the second stage, we used the estimated average 25(OH)D concentration stratified by cohort (DNBC/MoBa) as the continuous exposure variable in a weighted Cox regression model, with time since birth as the baseline. The Cox model was modified to account for the case-cohort design by applying inverse probability weights (22). Further details of the statistical analysis as well as our before-study power calculations are described in Web Appendix 1.

We assessed the linearity assumption by a categorical analysis using quartiles. Based on a graphical presentation of the log cumulative-hazard functions in strata defined by quartiles of 25(OH)D, the proportional hazards assumption was valid.

The primary analysis adjusted for the following covariates: maternal diabetes, age at time of delivery (continuous), prepregnancy BMI (categorical variable with boundaries at 18.5, 25, and 30), child's sex, and birth weight (continuous).

In a series of sensitivity analyses, we examined the primary analysis after additional adjustment for mode of delivery (cesarean delivery: yes or no), maternal smoking during pregnancy (yes or no), and maternal eicosapentaenoic acid and docosahexaenoic acid intake or vitamin D supplements taken during pregnancy; the sensitivity of our results to missing covariate data, using inverse probability weighting by propensity for missing



Figure 1. Selection of the study population from the Danish National Birth Cohort (DNBC) and the Norwegian Mother and Child Cohort Study (MoBa) in a childhood type 1 diabetes mellitus (T1D) case-cohort design, 1996–2014. There was insufficient plasma for 25-hydroxyvitamin D analysis in 12 vials from the DNBC sample population. Four children in the DNBC random sample were also T1D cases, and 2 of the MoBa random sample subjects were also T1D cases.

data; the primary analysis without adjusting for season of blood sampling (to test the hypothesis that absolute 25(OH)D concentrations during pregnancy predict childhood T1D); and possible separate associations of 25(OH)D concentrations, in mid-pregnancy or in umbilical cord blood samples, with the risk of childhood T1D. We assessed the sensitivity of our results to deviation from the assumption of normality in the measurement error model by using a Gaussian mixture model with 2 or 3 components. Finally, we investigated potential heterogeneity between the DNBC and MoBa cohorts by running cohort-specific measurement error models in the first stage of the analysis.

Ethics

The DNBC study was approved by the Danish National Ethics Board and the Danish Data Protection Agency. The MoBa study was approved by the Norwegian Data Protection Authority and the Regional Ethics Committee for Medical Research of South East Norway. All women provided written informed consent.

RESULTS

The characteristics of the study participants are presented in Table 1. The median age at T1D diagnosis was 7.4 years (range, 0.7–14.9), and the median follow-up time for the random cohort sample was 12.0 years (range, 4.7–16.2). There was a positive correlation between seasonally adjusted 25(OH)D concentrations in mid-pregnancy and umbilical cord blood samples (r = 0.40;

P < 0.001). Cohort-specific correlations are presented in Web Appendix 3, Web Figure 1. The seasonal variation in 25(OH)D concentrations are shown in Web Figure 2.

Estimated average maternal 25(OH)D concentration and childhood T1D

Our primary analysis demonstrated that there was no association between estimated average seasonally adjusted 25(OH)Dconcentration and childhood T1D (adjusted hazard ratio per 10-nmol/L increase = 1.00; 95% confidence interval: 0.90, 1.10). There was also no indication of any threshold (nonlinear) association (Figure 2).

Sensitivity analyses

The lack of association between maternal/cord blood 25(OH)D and the risk of childhood T1D was demonstrated consistently in a series of sensitivity analyses. These included an analysis adjusted for maternal intake of the long chain n-3 fatty acids eicosapentaenoic acid and docosahexaenoic acid, separate analyses for mid-pregnancy and umbilical cord blood samples (Figure 3), and analyses using absolute rather than seasonally adjusted 25(OH)D concentrations (Web Figure 3). First-trimester samples were available only from DNBC and showed a suggestive but nonsignificant inverse association with childhood T1D (adjusted hazard ratio per 10-nmol/L increase = 0.95; 95% confidence interval: 0.88, 1.02). The primary association was essentially equal in boys and girls (*P* for interaction = 0.99).

Table 1.	Characteristics of Cases With Childhood Type 1 Diabetes Mellitus and Subjects Randomly Selected From the Danish National Birth Cohort and the Norwegian Mother and Child
Cohort St	rudy, 1996–2014 ^a

	DNBC								МоВа							
Characteristic	Case (n = 270)			Cohort (<i>n</i> = 985) ^b				Case (n = 189)				Cohort $(n = 576)^{\rm b}$				
	Median	IQR	No.	%	Median	IQR	No.	%	Median	IQR	No.	%	Median	IQR	No.	%
Plasma 25(OH)D, nmol/L																
Maternal first-trimester sample	50.4	36.5, 67.6			54.2	40.8, 69.4										
Maternal mid-gestation sample	61.1	41.0, 79.4			60.0	41.4, 80.9			56.2	41.6, 77.4			57.5	42.5, 74.8		
Umbilical cord blood sample	37.7	26.6, 54.9			38.9	26.6, 52.1			31.7	21.5, 45.3			31.9	21.3, 46.1		
Maternal postpartum sample									42.5	34.1, 63.2			45.7	30.6, 64.3		
Age at diagnosis of T1D, years	9.0	5.7, 11.1							5.7	3.6, 7.9						
Female children			138	51.1			497	50.4			93	49.2			285	49.5
Maternal age at delivery, years	30	26.8, 32.6			30	27, 33			30	27, 33			30	27, 33		
Birth weight, kg	3.5	3.2, 3.9			3.5	3.3, 3.9			3.7	3.3, 4.0			3.6	3.3, 4.0		
Maternal prepregnancy BMI ^c																
<18.5			17	6.9			42	4.5			7	4.0			17	3.2
18.5–24.9			155	63.0			648	70.4			90	51.7			362	68.6
25.0–29.9			48	19.5			162	17.6			49	28.2			109	20.6
≥30.0			26	10.6			68	7.4			28	16.1			40	7.6
Maternal diabetes diagnosis ^d			15	5.6			29	2.9			7	3.7			0	0.0
Maternal vitamin D supplements, $\mu g/day^e$	10.0	5.0, 10.0			9.3	5.0, 10.0			4.8	1.3, 10.0			4.6	2.2, 10.0		
Maternal vitamin D from foods, $\mu g/day^f$	2.6	1.9, 4.1			2.9	2.0, 4.1			2.7	1.8, 4.2			3.2	1.9, 4.4		
Maternal EPA supplements, mg/day ^{e,g}	63.1	336.2			25.2	157.3			214.5	317.4			199.4	290.4		
Maternal EPA from foods, mg/day ^{f,g}	89.9	77.0			94.7	80.5			155.5	150.9			170.2	171.8		
Maternal DHA supplements, mg/day ^{e,g}	52.7	318.1			17.7	106.5			232.8	328.2			223.9	304.6		
Maternal DHA from foods, mg/day ^{f,g}	223.6	177.3			238.5	195.4			250.4	205.7			270.2	226.2		
Maternal smoking in pregnancy			56	21.3			266	27.3			14	33.3			50	38.2
Cesarean delivery			39	14.4			149	15.1			36	19.0			59	10.2

Abbreviations: 25(OH)D, 25-hydroxyvitamin D; BMI, body mass index; DHA, docosahexaenoic acid; DNBC, Danish National Birth Cohort; EPA, eicosapentaenoic acid; IQR, interquartile range; MoBa, Norwegian Mother and Child Cohort Study; T1D, type 1 diabetes mellitus.

^a Participants were recruited from 1996 to 2008 and followed up to February 2014 with respect to type 1 diabetes mellitus. Data are expressed as median (IQR) for continuous variables or *n* (% of those with nonmissing data) for categorical variables, unless otherwise specified. Missing values out of 2,020 individuals: maternal age at delivery (n = 1); birth weight (n = 4); maternal prepregnancy BMI (n = 152); maternal vitamin D supplementation (n = 811); maternal vitamin D intake from foods (n = 396); maternal EPA intake from supplements (n = 811); maternal DHA intake from diet (n = 719); maternal smoking during pregnancy (n = 611).

^b A randomly selected sample from the cohort (subcohort in case-cohort design).

^c BMI was calculated as weight (kg)/height (m)².

^d Maternal T1D in the MoBa; maternal diabetes of any type in the DNBC.

^e Intake of from supplements of vitamin D, EPA, or DHA, reported during weeks 22–25 of pregnancy.

^f Maternal intake of vitamin D, EPA, or DHA from foods, estimated from food frequency questionnaires administered in the second trimester.

^g Data are mean values (standard deviations).



Figure 2. Survival curve illustrating the lack of association between quartiles of estimated average 25-hydroxyvitamin D during pregnancy and the risk of childhood type 1 diabetes mellitus (P = 0.51; degrees of freedom = 3), Danish National Birth Cohort and the Norwegian Mother and Child Cohort Study, 1996–2014. The 25-hydroxyvitamin D cutoffs for each quantile were: 0%–24.9%: 37.5–61.9 nmol/L; 25.0%–49.9%: 62.0–69.1 nmol/L; 50.0%–74.9%: 69.2–77.4 nmol/L; 75.0%–100%: 77.5–130.3 nmol/L. Note that the *y*-axis does not begin at zero.

Vitamin D supplementation and association with risk of childhood T1D

In support of our main finding, we found no association between maternal self-reported vitamin D supplementation during pregnancy and the risk of childhood T1D when used as a binary variable (hazard ratio = 0.91; 95% confidence interval: 0.57, 1.43) or when used as a continuous variable (hazard ratio per increased µg of vitamin D/day = 1.01; 95% confidence interval: 0.98, 1.03).

DISCUSSION

In this study, using 2 of the world's largest cohorts of pregnant women, we have presented novel data on an unresolved issue in T1D etiology. Our results showed that normally varying 25(OH)D concentrations in a series of maternal and umbilical cord plasma measurements were not associated with risk of childhood T1D. In addition, maternal intake of vitamin D supplements during pregnancy was not associated with risk of childhood T1D.

Comparison with other studies

Two previous studies investigated maternal 25(OH)D during pregnancy (11, 12), and 2 studies investigated 25(OH)D₃ in neonatal dried blood spots (13, 14), all relative to childhood T1D. The results were inconsistent but there were important differences and limitations to take into account. In a Norwegian nested

case-control study with 109 cases, Sørensen et al. (12) reported a 2-fold increase in T1D risk for children born to women with late-pregnancy 25(OH)D concentrations in the first compared with the fourth quartile. In an updated analysis of the same individuals, Sørensen et al. (23) found no association between first- and second-trimester 25(OH)D concentrations and child-hood T1D risk. A Finnish study of 343 case-control pairs found no association between maternal concentrations of 25(OH)D during the first trimester and the risk of childhood T1D (11). In light of this latter study (11), our suggestive but nonsignificant inverse association observed for first-trimester samples, available in DNBC only, were likely due to chance.

One small Italian case-control study (67 cases with T1D) reported that increased 25(OH)D concentrations from neonatal dried blood spots were associated with a lower risk of childhood T1D in an immigrant subgroup, but there was no significant association in the Italian subgroup or the 2 subgroups combined (14). A large Danish study, with 1,090 T1D cases, found no association between neonatal concentrations of $25(OH)D_3$ in dried blood spots and the risk of childhood T1D (13, 14). Concentrations of 25(OH)D in dried blood spots are substantially lower but correlate strongly with plasma measurements (24). In the Danish study, the median concentrations of $25(OH)D_3$ ranged from 21.1 nmol/L to 24.3 nmol/L (13), whereas in the Italian study, the mean $25(OH)D_3$ concentrations were extremely low in both groups (<5 nmol/L) (14).

The current study is, to our knowledge, the first to assess cord blood 25(OH)D in relation to childhood T1D, and our results are consistent with those of the larger Danish study on neonatal 25(OH)D concentrations (13). Importantly, we show with precision that the lack of association was not limited to a specific trimester or sample type.

As a biomarker, 25(OH)D provides an objective measurement of vitamin D that integrates both dietary intake and endogenous production in the skin in response to ultraviolet irradiation. Some previous studies have investigated dietary intake of vitamin D during pregnancy in relation to childhood T1D. In these studies, a retrospective (case-control) design would be inferior due to the high risk of recall and selection bias, and prospective designs are preferable. Vitamin D intake from food or supplements during pregnancy was not associated with islet autoimmunity (a preclinical stage of T1D) in genetically susceptible Finnish children (25). In addition, the results from a Swedish population-based study were consistent with those from our larger, prospective study in demonstrating that the use of vitamin D supplements during pregnancy was not associated with risk of childhood T1D (26).

Strengths and weaknesses

The strengths of this study included its large scale, which provided precise risk estimates and its prospective approach. In addition, multiple measurements were made during pregnancy and also from umbilical cord blood, which allowed the average 25(OH)D concentration to be estimated from pregnancy through to delivery. Another strength of the study was the concurrent assessment of maternal vitamin D supplementation.

Some study limitations should also be considered. As in any observational study, we cannot exclude the possibility that unknown confounding factors may have influenced our results. We did not have information on human leukocyte antigen



Figure 3. Sensitivity analyses for the association between maternal/cord blood vitamin D status and the risk of childhood type 1 diabetes mellitus (hazard ratio per 10-nmol/L increase in plasma 25-hydroxyvitamin D (25(OH)D) concentration), Danish National Birth Cohort and the Norwegian Mother and Child Cohort Study, 1996–2014. The primary model adjusted for maternal diabetes, age at time of delivery, prepregnancy body mass index, child's sex, birth weight, and the time of year/season that each blood sample was taken. The subsequent lines show the main association after additional adjustment: maternal eicosapentaenoic acid (EPA) and docosahexaenoic acid (DHA) intake from diet and supplements during pregnancy (both continuous variables; 40% had missing data for these covariates). Missing covariates: This shows the result of the primary model with additional adjustment for missing covariates using inverse probability weighting by propensity for missing information on at least 1 of the primary model (same covariates), using cord blood only or mid-pregnancy samples only, for 25(OH)D concentration. Details regarding missing covariates are shown in the footnote of Table 1. CI: confidence interval. HR: hazard ratio.

(HLA), the major genetic determinant of T1D, or single nucleotide polymorphisms of the vitamin D pathway. Therefore, we could not examine potential genetic-environmental interactions. However, we do not expect that our null finding could be attributable to a confounding factor, genetic variation in the vitamin D pathway (27), or HLA genotype (13, 28). Furthermore, we did not measure plasma vitamin D binding protein, which could have been helpful in estimating the free 25(OH)D fraction. On the other hand, Sørensen et al. (23) did not find any association between the estimated free maternal 25(OH)D during pregnancy and childhood T1D, and the relevance of the free 25(OH)D fraction is debated. Participants in the DNBC and MoBa studies may not be representative of the general population of pregnant women in Denmark and Norway (e.g., they may be better educated or have healthier lifestyles), but this does not necessarily confound exposure-outcome associations (29, 30). While our results should be largely generalizable to other similar European and European origin populations, we cannot exclude the possibility that results may not be generalizable to populations with much lower vitamin D status.

Implications and future perspective

While sufficient vitamin D concentrations during pregnancy could have other benefits, the results of our study do not support recommending vitamin D supplements during pregnancy to reduce the risk of T1D in the offspring. Only a large-scale, long-term randomized controlled trial could establish whether increasing 25(OH)D during pregnancy beyond the concentrations we observed would alter the risk of childhood T1D. However, our results do not favor the initiation of such a trial.

Conclusion

Our large-scale Scandinavian study shows that normal variation in maternal or neonatal 25(OH)D is unlikely to have a clinically important effect on risk of childhood T1D.

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