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## Click-evoked auditory brainstem responses and autism spectrum disorder: A meta-analytic review

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### Abstract

Behavior does not differentiate ASD risk prior to 12 months of age, but biomarkers may inform risk before symptoms emerge. Click-evoked auditory brainstem responses (ABRs) may be worth consideration due to their measurement properties (non-invasiveness; reliability) and conceptual features (well-characterized neural generators), but participant characteristics and assessment protocols vary considerably across studies. Our goal is to perform a meta-analysis of the association between ABRs and ASD. Following an electronic database search (PubMed, Medline, PsycInfo, PsycArticles), we included papers that were written in English, included ASD and typically-developing (TD) groups, and reported the information needed to calculate standardized mean differences (Hedges's  $g$ ) for at least one ABR latency component (I, III, V, I–III, III–V, I–V). We weighted and averaged effect sizes across conditions and subsets of participants to yield one estimate per component per study. We then performed random-effects regressions to generate component-specific estimates. ASD was associated with longer ABR latencies for Waves III ( $g=0.5$ , 95% CI 0.1, 0.9), V ( $g=0.7$ , 95% CI 0.3, 1.1), I–III ( $g=0.7$ , 95% CI 0.2, 1.2), and I–V ( $g=0.6$ , 95% CI 0.2, 1.0). All components showed significant heterogeneity. Associations were strongest among participants  $\geq 8$  years of age and those without middle ear abnormalities or elevated auditory thresholds. In sum, associations between ABRs and ASD are medium-to-large in size, but exhibit heterogeneity. Identifying sources of heterogeneity is challenging, however, due to power limitations and co-occurrence of sample/design characteristics across studies. Research addressing the above limitations is crucial to determining the etiologic and/or prognostic value of ABRs for ASD.

### Keywords

Evoked Potentials; Auditory; Brain Stem; Autism Spectrum Disorder

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## Introduction

Early interventions represent promising avenues for improving the functioning of children with autism spectrum disorder (ASD). While the impact of these interventions depends upon many factors (e.g., symptom severity), age at enrollment is a powerful predictor of their efficacy (Odom, Boyd et al. 2010, Wallace and Rogers 2010, Rogers, Vismara et al. 2014). Early diagnosis and/or reliable identification of ASD risk therefore represent pressing public health objectives.

To date, family history is the most clearly defined risk factor for ASD with a 10–20% recurrence rate within families and heritability estimates of 0.6–0.9 (Ronald, Happe et al. 2006, Constantino, Zhang et al. 2010, Tick, Bolton et al. 2016). This knowledge has motivated extensive work with infant siblings of children with ASD, which in turn, has informed the identification of early emerging, behavioral-level antecedents associated with diagnosis (Jones and Klin 2013, Constantino, Kennon-McGill et al. 2017). However, behavior does not reliably differentiate ASD risk prior to 12 months of age and diagnostic status is not considered reliable prior to age 2 (Lord, Risi et al. 2006, Zwaigenbaum, Thurm et al. 2007, Kleinman, Ventola et al. 2008, Rogers 2009). Thus, much effort has been devoted to identifying biomarkers (e.g., genetic, metabolic, immune) that may inform risk prior to the manifestation of behavioral-level symptoms (Newschaffer, Croen et al. 2007, Dawson 2008).

To this end, click-evoked auditory brainstem responses (hereafter, ABRs) may be a biomarker worth further consideration. ABRs are electrophysiological responses that reflect auditory pathway activation by broadband acoustic stimuli (i.e., clicks) from the cochlea through the rostral brainstem (Moore 1987a, Moore 1987b). ABRs consist of 5 waves (I–V) from which latencies and amplitudes can be derived, values that reflect the degree of dendritic branching, myelination, and synchrony of firing across populations of neurons in the central auditory pathway (Ponton, Moore et al. 1996). Well-characterized components include I, III, & V, which correspond to action potentials generated from the VIII cranial nerve, cochlear nucleus, and lateral lemniscus, respectively (Moore 1987a).

ABRs may advance our understanding of ASD for methodological and conceptual reasons. For example, ABRs are recorded using electrodes placed on the scalp and are thus a non-invasive assessment; this facilitates enrollment of participants without medical indications. ABRs also exhibit high signal-to-noise ratios. They are elicited by clicks that are presented in quick succession (e.g., 11/sec), which enables the administration of (and averaging across) thousands of trials within minutes. Because ABRs also demonstrate test-retest reliability (Yang, Stuart et al. 1993), components can be interpreted at an individual-level with clinical import (e.g., neonatal hearing screening programs) (Mason and Herrmann 1998). In addition, ABR neural generators are well-characterized, despite diverse efferent and afferent projections that converge on these generators (Winer 2005). Thus, waveform decomposition enables integration with other brain-based assessments to generate hypotheses and/or evaluate coherence of findings.

Additional features of the ABR are specifically relevant to ASD. Neuroanatomically, ASD is associated with: 1) smaller brainstem volume, driven primarily by grey matter reduction, and 2) a marked reduction in superior olivary neurons, projections from which contribute to the lateral lemniscus (i.e., Wave V) (Hashimoto, Tayama et al. 1995, Rodier 2002, Jou, Minshew et al. 2009, Jou, Frazier et al. 2013). In addition, ASD and ABRs exhibit sex differences. ASD affects 4–5 males per female (Centers for Disease Control 2014), and males produce longer ABR latencies for all major wave components across the lifespan (Jerger and Hall 1980, Li, Zhu et al. 2013). ASD and ABRs are also sensitive to perinatal health risks and exhibit family resemblance (Jiang 1998, Jerger, Chmiel et al. 1999, Maziade, Merette et al. 2000, Jiang, Brosi et al. 2005). Perhaps most importantly, given that prevailing etiologic hypotheses of ASD implicate alterations in perinatal brain development (Rodier, Ingram et al. 1996, Anderson, Jacobs-Stannard et al. 2007, Stoner, Chow et al. 2014), ABRs can be measured a time proximal to this proposed process.

Given these characteristics, it is not surprising that associations between ABRs and ASD have been explored for more than three decades. However, findings vary considerably from one study to the next – both in terms of the magnitude and the direction of associations. For example, effects range from null to large (Courchesne, Courchesne et al. 1985, Roth, Muchnik et al. 2012) and are not consistently linked to specific aspects of the ABR waveform. In addition, ASD has been associated with both slower and faster ABR wave latencies relative to non-ASD counterparts (Rumsey, Grimes et al. 1984, Kwon, Kim et al. 2007, Dabbous 2012). However, several factors currently impede a coherent synthesis of the literature. For example, studies vary in the age at ABR assessment, utilize different ABR collection methods, employ varying definitions of ASD, and do not consistently address the impact of potential confounders (e.g., sex). We therefore performed a meta-analysis of the association between ABRs and ASD to address these interpretational challenges.

## Method

### Data Sources & Search Strategy

To identify candidate papers, we searched PubMed, Medline, PsycInfo, and PsycArticles using the following terms: (“auditory brain stem” or “auditory brainstem” or “audit\$,”) and (“autism” or “autism spectrum disorder” or “PDD” or “disintegrative” or “asperger\$”). This search, most recently implemented in May 2016, yielded 80 references that were evaluated for inclusion. First, we excluded references not written in English (N=4), along with reviews/commentaries (N=5), animal studies (N=9), case studies (N=1), and duplicate papers (N=4). We also excluded papers missing either: 1) click-evoked ABR or ASD data (N=30) or 2) a typically developing comparison (TD) group (N=7). To remain sensitive to secular changes in ASD conceptualization, following diagnoses were considered indicative of and are hereafter referred to as “ASD”: Infantile Amnesia, Autistic Disorder, Asperger’s Disorder, Pervasive Developmental Disorder, Pervasive Developmental Disorder-Not Otherwise Specified, or Childhood Disintegrative Disorder. In total, 60 papers were excluded in this first stage. Next, we examined citations within the remaining 20 papers to identify references missed by our database search. We identified an additional 5 papers using this manual search strategy, resulting in 25 papers eligible for further consideration. Finally,

we obtained copies of these papers and evaluated whether standardized mean differences could be generated for at least one ABR latency component. We did not consider papers reporting only: 1) odds of ABR abnormality (N=2) (Cohen, Gardner et al. 2013, Demopoulos and Lewine 2015), because definitions for abnormality were non-comparable across studies, or 2) ABRs acquired via binaural stimulation, which are non-comparable to ABRs acquired monaurally (N=1) (Rosenblum, Arick et al. 1980). In total, we included 15 papers in this meta-analysis, all of which had a stated objective of assessing differences in ABR components between ASD and TD groups (Figure 1) (Taylor, Rosenblatt et al. 1982, Gillberg and Gillberg 1983, Rumsey, Grimes et al. 1984, Grillon, Courchesne et al. 1989, Sersen, Heaney et al. 1990, Wong and Wong 1991, Tharpe, Bess et al. 2006, Kwon, Kim et al. 2007, Tas, Yagiz et al. 2007, Russo, Nicol et al. 2009, Fujikawa-Brooks, Isenberg et al. 2010, Magliaro, Scheuer et al. 2010, Dabbous 2012, Roth, Muchnik et al. 2012, Miron, Roth et al. 2016). Because our analyses utilized published, aggregate-level data, our study is considered exempt by the Michigan State University Institutional Review Board.

### ABR components and effect size scoring

Well-characterized ABR components include waves I, III, & V, which can be reliably generated and measured across the lifespan (Jerger and Hall 1980, Skoe, Krizman et al. 2015). ABRs yield amplitudes and latencies that may reflect the processes of neuronal synchronization and myelination. However, only one study enabled effect size calculation for amplitudes (Grillon, Courchesne et al. 1989). Thus, ABR parameters of interest here include absolute (I, III, & V) and inter-peak latencies (IPLs: I–III, III–V, I–V). Absolute and inter-peak latencies differ in that the former is derived from the onset of the click (thus involving conduction and transduction) whereas the latter is derived from the onset of a particular wave.

We estimated effect sizes using Hedges's  $g$ , a standardized mean difference score corrected for inflation due to small sample sizes. Hedges'  $g$  is interpreted similarly to Cohen's  $d$ , with estimates of 0.2, 0.5, and 0.8 corresponding to small, medium, and large effects, respectively. Study- and component-specific estimates of Hedges's  $g$  were calculated to reflect latency differences between ASD and TD participants ( $g > 0$ : ASD latency > TD latency;  $g < 0$ : ASD latency < TD latency). To generate one estimate per parameter per study, effect sizes were weighted and averaged across all variable conditions (e.g., ear of stimulation) and subsets of participants (Card 2011). Exceptions included Fujikawa et al. (2010), from which we only utilized the 61/sec condition, and Miron et al. (2016), from which we only utilized infant data (see below). The first author abstracted the papers and calculated effect sizes (at the study- and component-level) on two separate occasions to identify and resolve any discrepancies. Disaggregated effect sizes by study and component are summarized in eTable 1.

### Moderator Variables

We abstracted various study characteristics to characterize heterogeneity in effects across studies and address conceptual gaps in the literature. A summary of the study characteristics and coding decisions, generated by two independent abstractors, is reported in Table 1. We did not model preterm delivery as a moderator because perinatal health information was

reported in only three of the studies included here (1 excluded preterm infants, 1 included preterm infants, and 1 excluded children with “infective prenatal conditions”) (Tas, Yagiz et al. 2007, Roth, Muchnik et al. 2012, Miron, Roth et al. 2016). A minimum of two studies was necessary to warrant interpretation of a specific moderator variable level.

**Age at ABR assessment**—Because neurodevelopmental processes impact ABR components up to 18 months of age (and perhaps again at preschool-age, adolescence, and middle to late adulthood) (Jerger and Hall 1980, Thivierge and Cote 1990, Skoe, Krizman et al. 2015, Spitzer, White-Schwoch et al. 2015), we grouped studies according to whether ABRs were assessed prior to 8 years of age. This corresponds to the age of peak prevalence for ASD (Yeargin-Allsopp, Rice et al. 2003) and occurs prior to the onset of salient pubertal events for most participants. When participant age ranges straddled this divide, the study was included in the 8 year old group. Miron et al. (2016) included separate toddler and infant samples. We only utilized the infant data from this study given that ABRs were likely assessed prior to the manifestation of behavioral-level ASD symptoms.

**ASD case definition**—We grouped studies according to whether ASD diagnoses were specified using criteria published prior to or following DSM-IV, the system that markedly broadened the conceptualization of the disorder (Volkmar, Reichow et al. 2014). For studies that did not report diagnostic criteria, this information was inferred by comparing the age range of the participants to the publication date for DSM-IV.

**Intellectual Disability**—Intellectual disability is a common comorbidity associated with ASD (Centers for Disease Control 2014). We evaluated whether studies characterized intellectual functioning, and if so, whether participants with ID were included in the ASD group, TD group, neither group, or both groups.

**Sex matching**—Males exhibit longer ABR latencies across the entire lifespan and are also more likely to have an ASD diagnosis compared to females. Thus, we classified studies according to whether the ASD and TD groups were matched on sex. Matching was inferred from the article’s text or if the calculated proportion of male participants was equivalent across the ASD and TD groups (i.e., ASD:TD ratio = 1.0).

**Middle ear characterization**—Middle ear abnormalities that impede conduction can lead to prolonged ABR latencies (particularly absolute latencies) (Gunnarson and Finitzo 1991, Hall and Grose 1993). Each study was evaluated to determine whether tympanometry and/or otoscopic examinations were performed, and if so, whether participants with abnormal findings were included or excluded from analysis.

**Elevated auditory thresholds**—Elevated auditory thresholds are associated with prolonged ABR latencies and are indicative of hearing loss (Jerger and Johnson 1988). Each study was evaluated to determine whether this information was reported and if so, whether participants with elevated thresholds were included or excluded from analysis.

**Click rate**—Click rates can be manipulated to exert varying levels of challenge to the auditory nerve, with faster click rates eliciting longer wave latencies across all ages and in

the context of some demyelinating diseases (e.g., multiple sclerosis) (Jacobson, Murray et al. 1987, Jiang, Brosi et al. 2002). Studies were grouped according to whether they utilized rates above or below 27.5 clicks/second, because rates above this threshold have been associated with longer latencies in both neonates and adults (Jiang, Brosi et al. 1998). Although Fujikawa et al. (2010) utilized 2 different click rate conditions, we utilized data from the 61/second condition here to increase the sample size of the 27.5/second group.

### Publication bias

We evaluated publication bias using Kendall's tau and Egger's intercept, and interpreted significant findings on either test as indicative of bias ( $p < 0.05$ , two-tailed). Because these tests may be underpowered (Card 2011), we also calculated the fail-safe N to estimate the minimum number of studies with an effect size of 0 needed to attenuate findings to non-significance.

### Analytic Plan

We begin by providing an overview of the studies contributing to the meta-analysis. After generating one effect size per component per study, we performed random-effects regressions (one per component) to evaluate whether ABR latencies differed between ASD and TD participants. Next, we evaluated heterogeneity in these effects using the Q statistic (Lipsey and Wilson 2001). For components exhibiting significant heterogeneity, we used mixed-effects meta-regression to evaluate the contribution of each moderator to the variability in effects. Random effects variance was based upon methods of moments estimation. To adjust for multiple comparisons, we utilized a false discovery rate of 5% (corrected  $p = 0.013$ , two-tailed) to minimize the impact of Type I error. Publication bias was evaluated only for parameters with significant effect sizes in the main (i.e., non-moderator) analysis.

### Results

Of the 15 studies included in this meta-analysis, 14 employed cross-sectional designs and 1 employed a case-control design. The number of participants per study ranged from 16 to 167, and ages ranged between 3 months and 40 years (eAppendix). Six studies (40%) involved participants  $\leq 8$  years, nine studies (60%) utilized DSM-IV or DSM-IV-TR criteria to diagnose ASD, and seven studies (47%) matched the ASD and TD groups on sex (Table 1). Seven studies (47%) did not report any information on intellectual disability (ID), whereas six studies (40%) excluded ID only from the TD group. With respect to ABR acquisition protocols, seven studies (47%) excluded children with middle ear abnormalities and nine studies (60%) excluded children with hearing loss. A majority of studies (67%) employed click rates  $< 27.5$ /second.

The number of studies contributing to component-specific effect size estimates varied from 11 (I; I-III; III-V) to 13 (V, I-V), with the number of participants ranging from 657 (I-III) to 862 (V) (Table 2). ASD was not associated with Wave I or III-V latencies. However, ASD was associated with longer latencies relative to TD counterparts for Waves III ( $g = 0.5$ , 95% CI 0.1, 0.9), V ( $g = 0.7$ , 95% CI 0.3, 1.1), I-III ( $g = 0.7$ , 95% CI 0.2, 1.2), and I-V ( $g = 0.6$ ,

95% CI 0.2,1.0). For all absolute and inter-peak latencies, we observed significant heterogeneity in these effects (all  $p < 0.001$ ; eFigure 1A–1F).

Tables 3a and 3b summarize moderator analyses for absolute and inter-peak latencies, respectively. None of the moderators were associated with Wave I latencies. For Waves III and V, age  $\geq 8$  years at ABR assessment, utilization of DSM-IV/IV-TR diagnostic criteria, exclusion of participants with middle ear abnormalities or hearing loss, and click rates  $27.5/\text{sec}$  were associated with longer latencies for ASD versus TD participants ( $0.7 < g < 1.0$ , all  $p < 0.013$ ). Sex matching was not associated with Wave III, but was associated with Wave V; specifically, ASD was associated with longer latencies in both the matched and unmatched groups ( $0.4 < g < 0.7$ , all  $p < 0.013$ ). This general pattern of findings was replicated for inter-peak latencies I–III and I–V ( $0.6 < g < 1.0$ , all  $p < 0.013$ ), except that associations: 1) were not observed with ASD diagnostic criteria for wave I–III, and 2) extended to include participants for whom the presence of middle ear abnormalities was not reported (I–III:  $g = 0.8$ , 95% CI 0.2,1.3; I–V:  $g = 0.9$ , 95% CI 0.4,1.4). In addition, associations between ASD and IPL I–V latencies were attenuated among studies that matched on sex (I–V:  $g = 0.3$ , 95% CI  $-0.1, 0.8$ ,  $p > 0.013$ ). Exclusion of participants with middle ear abnormalities was the only factor associated with IPL III–V ( $g = 0.7$ , 95% CI 0.3,1.0).

We observed no evidence of publication bias across two indices assessing this effect (Kendall's tau and Egger's test, all  $p > 0.27$ ; eTable 2). Approximately 119 (Wave III) to 290 (Wave V) studies with an effect size of zero would be required to attenuate main effects (Table 2) to non-significance.

## Discussion

We performed a meta-analysis to assess the association between ASD and click-evoked ABRs and evaluated the impact of study characteristics that currently impede synthesis of the literature. We found that ASD was associated with longer ABR latencies relative to TD participants, particularly for waves III, V, I–III, and I–V. These associations were medium-to-large in size ( $0.5 < g < 0.7$ ), but exhibited considerable heterogeneity. This variability was most consistently linked to participant age and ABR protocol characteristics.

For both absolute and inter-peak latencies, associations with ASD were limited to components involving neural transmission from the auditory nerve (wave I) to the cochlear nucleus (wave III). This raises the possibility that transmission involving wave I and wave III generators contribute to the findings observed here, given that no associations with wave I (click to auditory nerve) or III–V (cochlear nucleus to lateral lemniscus) were observed. Action potential velocity is determined primarily by degree of myelination, pathway length, and axonal diameter, but may also be influenced by the synchronization of neuronal firing or changes in synaptic efficacy (Eggermont 1988). To date, there is limited or equivocal evidence to suggest that these factors explain associations the findings observed here. For example, both hyper- and hypo-myelination of brainstem pathways have been linked to ASD (Hanaie, Mohri et al. 2016, Ouyang, Cheng et al. 2016), though we are unaware of studies that characterize these parameters for the central auditory pathway specifically. Furthermore, microscopic, imaging-based, and physiological findings that implicate brainstem-based

anomalies in ASD do not necessarily mean that this brain region drives the complex neurological and behavioral features that accompany the disorder (e.g., weaker functional connectivity in frontal cortex; stronger cortical-subcortical connectivity; rapid sensory cortical expansion) (Minshew and Williams 2007, Hazlett, Gu et al. 2017). Indeed, longer ABR latencies associated with ASD may reflect activity of more distal brain regions that converge directly (e.g., corticofugal pathways) and/or indirectly (e.g., via the pons) on neural generators of the ABR. Disentangling how this diverse network of brain-based findings relate to one another is an important direction for future research.

Despite the medium-to-large effects observed at the aggregate-level, there was great variability in the magnitude and sometimes the direction of associations across individual studies. Moderator analyses suggested that effect size was related in part to ABR study characteristics – younger age at assessment, exclusion of participants with middle ear abnormalities or hearing loss, and faster click rates. With respect to age, cross-sectional findings suggest that ABR latencies decrease markedly during the first two years of life, decrease somewhat less steeply during preschool age, and then increase during middle childhood and adolescence to approach adult values, changes hypothesized to reflect brain-based developmental processes such as myelination, synaptogenesis, and pruning (Skoe, Krizman et al. 2015, Spitzer, White-Schwoch et al. 2015). Given these age-related changes, we repeated our analyses after classifying studies according to whether participants were assessed prior to 5 years; our results for waves III, V, I–III, and I–V were unchanged ( $0.9 < g < 1.3$ , all  $p < 0.013$ ). Although it is unclear whether ABR assessment in early childhood is particularly sensitive to associations with ASD, age-related changes in ABR components underscore the importance of matching participants on this variable. One study did not match ASD and TD groups on age (Roth, Muchnik et al. 2012), and this might have contributed to the particularly large effects reported therein. However, when we excluded this study from the analysis, our main findings were altered by less than 0.2 across all components (data not shown). With respect to middle ear problems and elevated auditory thresholds, each are linked to ASD as well as longer ABR latencies (Stockard, Stockard et al. 1978, Gunnarson and Finitzo 1991, Moore, Hutchings et al. 1991, Hall and Grose 1993); however, estimates were larger following the exclusion of participants with these difficulties, suggesting that they do not account for the associations reported here. It is unclear why findings would strengthen when middle ear problems were excluded, particularly for inter-peak latencies, which do not incorporate conduction time. One possibility is that children with ASD are more likely to experience repeated occurrences of otitis media (OM) (Adams, Susi et al. 2016); repeated OM, in turn, has been linked with longer inter-peak latencies, even when middle ear problems are excluded at the time of the ABR assessment (Gunnarson and Finitzo 1991, Ferguson, Cook et al. 1998). Another possibility is that study characteristics such as exclusion of participants with middle ear problems are confounded by other factors that impact ABR latencies. For example, studies utilizing faster click rates often reported stronger associations between ABRs and ASD, but these studies were also likely to exclude participants with middle ear problems and hearing loss.

We then examined whether effect size heterogeneity was associated with ASD symptoms or comorbid conditions. When ASD was diagnosed using either DSM-IV or –IV-TR criteria, which greatly broadened the symptoms linked to the disorder, stronger associations with



several ABR latencies were reported. In addition, stronger associations were observed among studies that did not report the presence of intellectual disability in either the ASD or TD groups. The latter finding likely reflects its almost exclusive co-occurrence with the use of DSM-IV or –IV-TR criteria (Table 1), as it is unclear why lack of reporting would be related to strength in association. Indeed, two studies using DSM-IV/-IV-TR criteria but excluded ID from either the ASD group or both groups reported effect sizes comparable to aggregate-level analyses (Russo, Nicol et al. 2009, Fujikawa-Brooks, Isenberg et al. 2010). Furthermore, a comparison between mutually exclusive groups of ASD and ID participants suggested that ABR latencies were significantly longer in the ASD group (Wong and Wong 1991). Combined with the fact that earlier DSM versions identified the most severely affected children in the ASD spectrum, evidence to date suggests that associations between ASD and ABRs may not be driven by comorbid ID. Relatedly, links between ABR findings and ASD symptom dimensions (e.g., sensory hyper-/hypo-sensitivity; social communication deficits; restricted/ repetitive behaviors) are scarce. However, emerging evidence suggests that ASD with hyperacusis may be associated with *faster* ABR latencies relative to both TD counterparts and ASD children without hyperacusis (Dabbous 2012, Thabet and Zaghoul 2013). For language, studies employing more complex auditory brainstem processing protocols (e.g., speech probes; forward masking) have observed concurrent and prospective links with receptive language functioning (Russo, Nicol et al. 2009, Chonchaiya, Tardif et al. 2013), a process that is disturbed in a subset of children with ASD. Links to social communication deficits or restricted behaviors are currently unknown and under-investigated. In the end, marked improvements in the scope and depth of behavioral assessment are required to fully probe any association between ASD symptom dimensions and auditory brainstem processing findings. Indeed, most studies utilized medical record abstraction to define diagnostic status, with only one utilizing gold-standard ASD assessments (e.g., ADOS) (Fujikawa-Brooks, Isenberg et al. 2010).

There are limitations and alternative explanations that are important to consider when interpreting our results. First, as mentioned earlier, study characteristics often co-occurred. Thus, it is unclear the extent to which individual sample or study characteristics investigated here are related to ABR latencies. Separating these effects represents an important direction for future research. Second, studies to date are almost entirely cross-sectional in nature, with ABR assessment and case ascertainment for ASD taking place concurrently. Resolving the temporality of associations is critical to determining whether ABRs have etiologic and/or prognostic value in relation to ASD. Although two recent studies with infants suggest that it is possible for ABR findings to precede ASD diagnosis (Cohen, Gardner et al. 2013, Miron, Roth et al. 2016), additional evidence is needed. Third, with the exception of IPL I–V, associations between ASD and ABR latencies persisted when analyses were limited to studies that matched on sex. However, the absolute prevalence of female participants across studies was low, and this precluded a direct evaluation of effect modification by sex. Thus, it is unclear whether the findings presented here generalize to females, and this represents a very important area for future investigation. Fourth, our analyses involved many comparisons. Even though we used false discovery rates to reduce the impact of Type I error, the findings generated from our moderator analyses are in particular need of replication. Fifth, our moderator analyses do not represent the full complement of factors that may affect

the association between ASD and ABRs. For example, preterm delivery is associated with longer ABR latencies relative to full-term counterparts who are matched according to either chronological or corrected age (Jiang and Li 2015, Stipdonk, Weisglas-Kuperus et al. 2016) and preterm delivery is a well-described risk factor for ASD. However, only three studies here reported specific information regarding perinatal health; one excluded preterm children (Roth, Muchnik et al. 2012), one included preterm children (Miron, Roth et al. 2016), and one excluded children with “infective prenatal conditions” (Tas, Yagiz et al. 2007). In addition, no study reported information regarding birth size or head circumference, the latter of which is positively associated with ABR latencies and has been extensively investigated in relation to ASD. Determining whether associations between ABRs and ASD differ across risk factors for the disorder (e.g., perinatal health; family history of ASD diagnosis) may provide important insights regarding the interpretation and potential application of ABRs in risk surveillance or elucidation of symptom profiles. Relatedly, it will be important to evaluate the specificity of findings to neurodevelopmental disorders beyond ASD. Indeed, longer ABR latencies have been linked with attention deficit/hyperactivity disorder and cerebral palsy (Sano, Kaga et al. 2005, Azzam and Hassan 2010). Currently, no direct comparisons between ASD and these disorders have been evaluated within the context of the same study. For these and other reasons (e.g., the lack of prospectively gathered data), the relevance of click-based ABRs in predicting risk for ASD or elucidating symptom profiles associated with the disorder is decidedly uncertain.

In sum, ASD may be associated with longer click-evoked ABR latencies. Findings vary greatly across studies, but effect sizes reported to date are substantial. Future work that utilizes prospective designs and addresses outstanding conceptual limitations are vital to informing the etiologic or prognostic value of ABRs for ASD.

## Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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### Lay Summary

Auditory brainstem responses (ABR) may be associated with ASD, but participant characteristics and assessment protocols vary considerably across individual studies. Our goal is to combine the results across these studies to facilitate clarity on the topic. Doing so represents a first step in evaluating whether ABRs yield potential for informing the etiology of ASD risk and/or ASD symptom profiles.

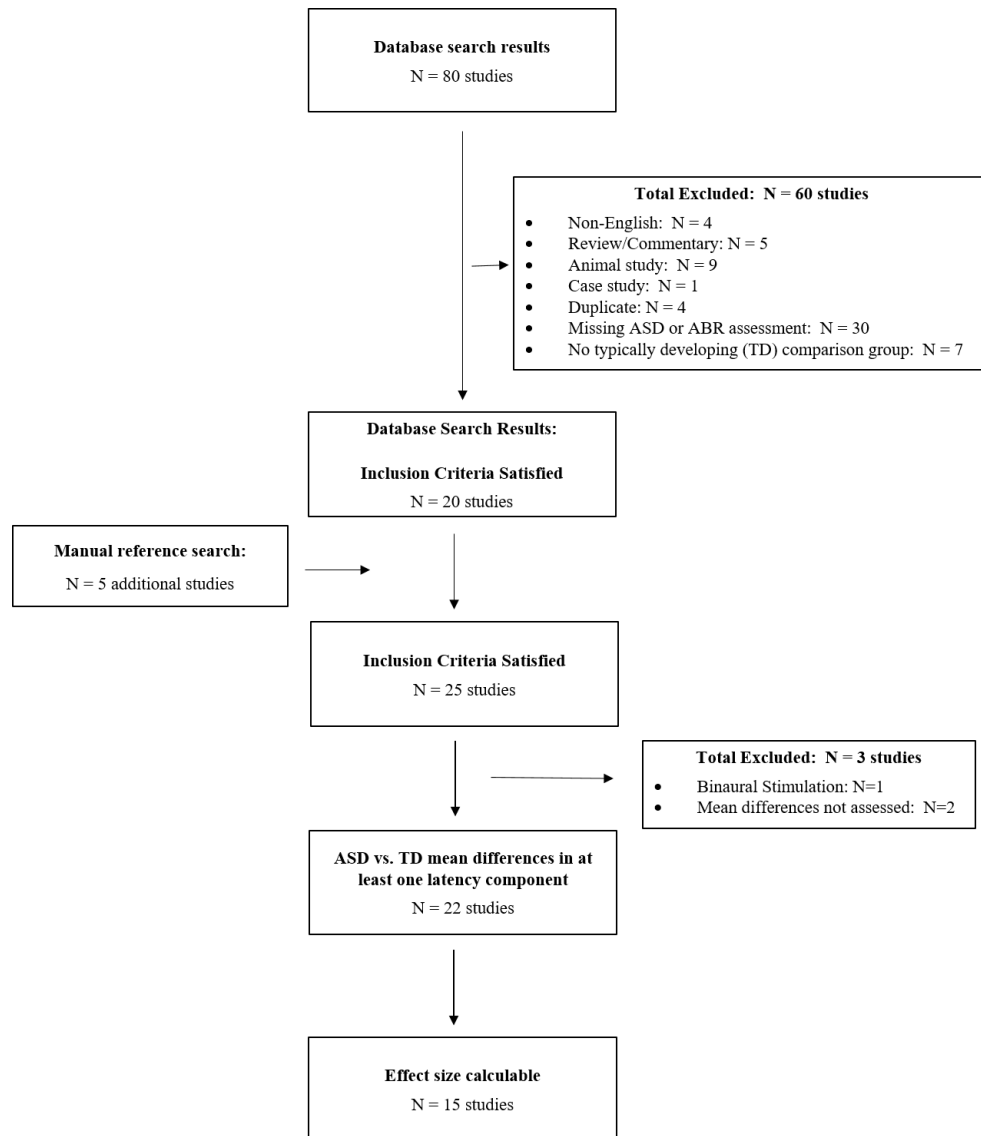


Figure 1.



**Table 1**

Coding decisions for characteristics of studies included in the meta-analysis

	Age Group	Sex Matching	ASD group	ID	Middle Ear Abnormality	Elevated Auditory Threshold	Click Rate	Waves Available
Taylor et al. (1982)	> 8 years	No	before DSM-IV <sup>1</sup>	not reported	not reported	not excluded	< 27.5/s	I-III, III-V, I-V
Gillberg et al. (1983)	> 8 years	No	before DSM-IV <sup>2</sup>	ASD only	not excluded	not excluded	not reported	V, I-V
Rumsey et al. (1984)	> 8 years	Yes	before DSM-IV	ASD only	not excluded	not excluded	< 27.5/s	III, I-III, I-V
Grillon et al. (1989)	> 8 years	Yes	before DSM-IV	none	not reported	excluded	< 27.5/s	I, III, V, I-III, III-V, I-V
Sersen et al. (1990)	> 8 years	Yes	before DSM-IV	ASD only	not reported	excluded	< 27.5/s	I, III, V
Wong and Wong (1991)	8 years	No	before DSM-IV	ASD only	not reported	excluded	< 27.5/s	I, III, V, I-III, III-V, I-V
Tharpe et al. (2006)	> 8 years	Yes	DSM IV/IV-TR	ASD only	excluded	not excluded	< 27.5/s	I, III, V, I-III, III-V, I-V
Tas et al. (2007) <sup>3</sup>	8 years	No	DSM IV/IV-TR	not reported	excluded	not excluded	< 27.5/s	I, III, V, I-III, III-V, I-V
Kwon et al. (2007) <sup>3</sup>	8 years	Yes	DSM IV/IV-TR	not reported	not reported	not reported	< 27.5/s	I, III, V, III-V, I-V
Russo et al. (2009)	> 8 years	No	DSM IV/IV-TR <sup>4</sup>	none	excluded	excluded	< 27.5/s	V
Magliaro et al. (2010)	> 8 years	No	DSM IV/IV-TR	not reported	excluded	excluded	< 27.5/s	I, III, V, I-III, III-V, I-V
Fujikawa et al. (2010)	> 8 years	No	DSM IV/IV-TR <sup>5</sup>	ASD only	excluded	excluded	27.5/s <sup>6</sup>	I, III, V, I-III, III-V, I-V
Roth et al. (2012)	8 years	Not Reported	DSM IV/IV-TR	not reported	excluded	excluded	27.5/s	I, III, V, I-III, III-V, I-V
Dabbous et al. (2012) <sup>7</sup>	8 years	Yes <sup>8</sup>	DSM IV/IV-TR <sup>9</sup>	not reported	excluded	excluded	27.5/s	I, III, V, I-III, III-V, I-V
Miron et al. (2016)	8 years <sup>10</sup>	Yes	DSM IV/IV-TR	not reported	not reported	excluded	27.5/s	I, III, V, I-III, III-V, I-V

Note. **ASD**: Autism spectrum disorder; **DSM**: Diagnostic and Statistical Manual; **ID**: Intellectual disability; **s**: second; **TD**: typically developing

<sup>1</sup>National Society for Autistic Children Criteria

<sup>2</sup>Rutter (1978) criteria

<sup>3</sup> ASD group only (n=71), because the AD group (n=20) was not described and appears to be a subset of this total

<sup>4</sup>Parent-report substantiated by medical records

<sup>5</sup>Autism Diagnostic Observation Schedule (ADOS) and Autism Diagnostic Interview-Revised (ADIR)

<sup>6</sup>Included only the 27.5/s group because click rate is a moderator variable in this analysis

<sup>7</sup>All ASD cases exhibited “intolerance to noise” or hyperacusis

<sup>8</sup>Derived from narrative text

Diagnosis by medical professional (criteria unspecified); child age within range defined by DSM-IV and article publication dates

Infancy condition only

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Autism spectrum disorder and its association with wave specific click-evoked auditory brainstem responses

**Table 2**

Wave Latency	No. of Studies	Sample Size	g	Random Effects			Q	p
				95%CI	p	p		
<b>I</b>	11	768	0.1	(-0.3, 0.5)	0.60	69.9	< 0.001	
<b>III</b>	12	818	0.5	(0.1, 0.9)	0.03	89.6	< 0.001	
<b>V</b>	13	862	0.7	(0.3, 1.1)	0.001	96.5	< 0.001	
<b>I-III</b>	11	657	0.7	(0.2, 1.2)	0.007	82.6	< 0.001	
<b>III-V</b>	11	728	0.3	(-0.1, 0.6)	0.093	39.0	< 0.001	
<b>I-V</b>	13	833	0.6	(0.2, 1.0)	0.001	75.9	< 0.001	

Note. *g* (Hedges'  $\bar{g}$ ); **Q** (test for heterogeneity); *p* (p value)

Moderator variables and effect sizes according to click evoked auditory brainstem response (ABR) absolute latencies<sup>f</sup>

**Table 3a**

	I			III			V					
	k	N	g	95% CI	k	N	g	95% CI	k	N	g	95% CI
<b>Age Group</b>												
8 years	6	541	0.0	(-0.5, 0.6)	6	541	<b>0.8</b>	<b>(0.2, 1.4)</b> *	6	297	<b>1.0</b>	<b>(0.4, 1.6)</b> *
> 8 years	5	227	0.2	(-0.4, 0.9)	6	277	0.2	(-0.4, 1.4)	7	565	0.5	(-0.1, 1.1)
<b>Sex Matching</b>												
Yes	5	341	-0.1	(-0.6, 0.3)	7	427	0.2	(-0.2, 0.5)	6	377	<b>0.4</b>	<b>(0.1, 0.7)</b> *
No	5	326	0.0	(-0.4, 0.5)	4	290	0.6	(0.1, 1.1)	6	384	<b>0.7</b>	<b>(0.4, 1.0)</b> *
Not Reported	1	101	<b>1.6</b>	<b>(1.1, 2.1)</b> *	1	101	<b>2.3</b>	<b>(1.4, 3.1)</b> *	1	101	<b>2.8</b>	<b>(2.3, 3.3)</b> *
<b>ASD definition</b>												
pre-DSM IV	3	277	0.0	(-0.8, 0.8)	4	327	0.0	(-0.7, 0.8)	3	165	0.4	(-0.5, 1.3)
DSM IV/IV-TR	8	491	0.1	(-0.4, 0.6)	8	491	<b>0.7</b>	<b>(0.2, 1.2)</b> *	10	697	<b>0.8</b>	<b>(0.3, 1.3)</b> *
<b>Intellectual disability</b>												
not reported	6	415	0.2	(-0.4, 0.7)	6	415	<b>0.8</b>	<b>(0.2, 1.4)</b> *	6	415	<b>0.9</b>	<b>(0.3, 1.5)</b> *
none	1	16	0.2	(-1.4, 1.8)	1	16	-0.8	(-2.3, 0.8)	2	55	0.1	(-1.1, 1.3)
ASD only	4	337	0.0	(-0.7, 0.7)	5	387	0.3	(-0.4, 0.9)	5	392	0.7	(0.0, 1.4)
<b>Middle Ear Abnormality</b>												
not reported	5	458	0.0	(-0.6, 0.6)	5	458	0.3	(-0.4, 0.9)	5	458	0.6	(-0.1, 1.3)
excluded	6	310	0.2	(-0.4, 0.8)	6	310	<b>0.8</b>	<b>(0.3, 1.4)</b> *	7	349	<b>0.8</b>	<b>(0.2, 1.4)</b> *
not excluded	0	0	--	----	1	50	-0.7	(-2.0, 0.6)	1	55	0.7	(-0.9, 2.3)
<b>Elevated Auditory Threshold</b>												
not reported	1	121	0.1	(-1.0, 1.2)	1	121	0.3	(-1.1, 1.6)	1	121	0.6	(-1.0, 2.2)
excluded	8	569	0.1	(-1.4, 1.6)	8	569	<b>0.7</b>	<b>(0.2, 1.2)</b> *	9	608	<b>0.8</b>	<b>(0.2, 1.3)</b> *
not excluded	2	78	0.1	(-0.4, 0.7)	3	128	-0.1	(-0.9, 0.8)	3	133	0.6	(-0.4, 1.5)
<b>Click rate</b>												
< 27.5/s	7	517	0.1	(-0.4, 0.6)	8	567	0.2	(-0.2, 0.7)	8	556	0.6	(0.0, 1.1)
27.5/s	4	251	0.1	(-0.6, 0.8)	4	251	<b>1.0</b>	<b>(0.3, 1.6)</b> *	4	251	<b>1.0</b>	<b>(0.2, 1.8)</b> *

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		A			III			I				
	95% CI	<i>g</i>	<i>N</i>	<i>k</i>	95% CI	<i>g</i>	<i>N</i>	<i>k</i>	95% CI	<i>g</i>	<i>N</i>	<i>k</i>
Not reported	(-0.9, 2.2)	0.7	55	1	-----	----	0	0	-----	----	0	0

Note: **df** (degrees of freedom); ***g*** (Hedges' *g*); ***k*** (number of studies); ***N*** (number of participants); ***Q*** (test for heterogeneity); ***p*** (p value)

*I* A minimum of two studies was necessary to warrant interpretation of a specific moderator variable level.

\* corrected  $p < 0.013$  (false discovery rate = 5%)

Moderator variables and effect sizes according to click evoked auditory brainstem response (ABR) inter-peak wave latencies<sup>f</sup>

**Table 3b**

	I-III			III-V			I-V					
	k	N	g	95% CI	k	N	g	95% CI	k	N	g	95% CI
<b>Age Group</b>												
8 years	5	420	<b>1.0</b>	<b>(0.3, 1.7)</b> *	6	541	0.3	(-0.2, 0.7)	6	541	<b>0.9</b>	<b>(0.4, 1.4)</b> *
> 8 years	6	237	0.4	(-0.2, 1.0)	5	187	0.2	(-0.3, 0.7)	7	292	0.4	(-0.1, 0.9)
<b>Sex Matching</b>												
Yes	5	212	0.5	(-0.2, 1.1)	5	283	0.1	(-0.1, 0.6)	6	333	0.3	(-0.1, 0.8)
No	5	344	0.6	(-0.1, 1.2)	5	344	0.5	(0.0, 1.0)	6	399	<b>0.7</b>	<b>(0.2, 1.1)</b> *
Not calculable	1	101	<b>2.0</b>	<b>(0.7, 3.3)</b> *	1	101	0.2	(-1.0, 1.5)	1	101	<b>1.8</b>	<b>(0.8, 2.8)</b> *
<b>ASD definition</b>												
pre-DSM IV	4	287	0.6	(-0.3, 1.4)	3	237	0.5	(-0.1, 1.1)	5	342	0.6	(-0.1, 1.2)
DSM IV/IV-TR	7	370	0.7	(0.1, 1.4)	8	491	0.2	(-0.2, 0.5)	8	491	<b>0.7</b>	<b>(0.2, 1.1)</b>
<b>Intellectual disability</b>												
not reported	6	293	<b>0.9</b>	<b>(0.3, 1.6)</b> *	7	469	0.2	(-0.2, 0.6)	7	469	<b>0.7</b>	<b>(0.2, 1.3)</b> *
none	1	16	1.0	(-0.7, 2.7)	1	16	0.0	(-1.3, 1.3)	1	16	0.6	(-1.0, 2.2)
ASD only	4	348	0.2	(-0.6, 1.0)	3	243	0.5	(-0.1, 1.1)	5	348	0.4	(-0.2, 1.1)
<b>Middle Ear Abnormality</b>												
not reported	4	297	<b>0.8</b>	<b>(0.2, 1.3)</b> *	5	444	-0.1	(-0.4, 0.3)	5	418	<b>0.9</b>	<b>(0.4, 1.4)</b> *
excluded	6	310	<b>0.9</b>	<b>(0.2, 1.6)</b> *	6	284	<b>0.7</b>	<b>(0.3, 1.0)</b> *	7	365	<b>0.6</b>	<b>(0.2, 1.1)</b> *
not excluded	1	50	-0.8	(-2.2, 0.5)	0	0	---	----	1	50	-0.7	(-1.8, 0.4)
<b>Elevated Auditory Threshold</b>												
not reported	0	0	---	-----	1	121	0.6	(-0.5, 1.6)	1	121	0.5	(-0.7, 1.6)
excluded	7	182	<b>1.0</b>	<b>(0.5, 1.5)</b> *	7	475	0.1	(-0.3, 0.5)	7	475	<b>0.9</b>	<b>(0.4, 1.4)</b> *
not excluded	4	475	0.1	(-0.6, 0.8)	3	132	0.6	(-0.1, 1.3)	5	237	0.2	(-0.3, 0.8)
<b>Click rate</b>												
< 27.5/s	7	406	0.5	(-0.1, 1.1)	7	477	0.4	(0.0, 0.8)	8	527	0.5	(0.0, 1.0)
27.5/s	4	251	<b>1.0</b>	<b>(0.2, 1.8)</b> *	4	251	0.0	(-0.5, 0.5)	4	251	<b>0.9</b>	<b>(0.2, 1.6)</b> *

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	I-V					III-V					III-I				
	<i>g</i>	<i>N</i>	<i>k</i>	95%CI		<i>g</i>	<i>N</i>	<i>k</i>	95%CI		<i>g</i>	<i>N</i>	<i>k</i>	95%CI	
Not reported	0.4	55	1	-----		---	0	0	-----		---	0	0	-----	
	(-1.0, 1.7)														

Note: **df** (degrees of freedom); **g** (Hedges'  $\bar{g}$ ); **k** (number of participants); **N** (number of studies); **Q** (test for heterogeneity); **p** (p value)

*I* A minimum of two studies was necessary to warrant interpretation of a specific moderator variable level.

\* corrected  $p < 0.013$  (false discovery rate = 5%)