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# Sex, Amyloid, and *APOE*<sub>2</sub>4 and risk of cognitive decline in preclinical Alzheimer's disease: findings from three well-characterized cohorts

Rachel F. Buckley<sup>1,2,3,4,\*</sup>, Elizabeth C. Mormino<sup>5,\*</sup>, Rebecca E. Amariglio<sup>3,4</sup>, Michael J. Properzi<sup>3</sup>, Jennifer S. Rabin<sup>6</sup>, Yen Ying Lim<sup>1</sup>, Kathryn V. Papp<sup>3,4</sup>, Heidi I. L. Jacobs<sup>7,8</sup>, Samantha Burnham<sup>9</sup>, Bernard J. Hanseeuw<sup>3,4,10</sup>, Vincent Doré<sup>11</sup>, Annette Dobson<sup>12</sup>, Colin L. Masters<sup>1</sup>, Michael Waller<sup>12</sup>, Christopher C. Rowe<sup>13,14</sup>, Paul Maruff<sup>15</sup>, Michael C. Donohue<sup>16</sup>, Dorene M. Rentz<sup>3,4</sup>, Dylan Kirn<sup>3,4</sup>, Trey Hedden<sup>7</sup>, Jasmeer Chhatwal<sup>3</sup>, Aaron P. Schultz<sup>3,4</sup>, Keith A. Johnson<sup>7</sup>, Victor L. Villemagne<sup>13,14</sup>, Reisa A. Sperling<sup>3,4</sup>, on behalf of collaborators in the Alzheimer's Disease Neuroimaging Initiative<sup>\*</sup>, the Australian Imaging, Biomarker and Lifestyle study of ageing<sup>\*</sup>, and the Harvard Aging Brain Study

<sup>1</sup>The Florey Institute, The University of Melbourne, Victoria, Australia <sup>2</sup>Melbourne School of Psychological Science, University of Melbourne, Victoria, Australia <sup>3</sup>Department of Neurology, Massachusetts General Hospital and Harvard Medical School, Boston, MA, USA <sup>4</sup>Center for Alzheimer Research and Treatment, Department of Neurology, Brigham and Women's Hospital, Boston, Massachusetts, USA <sup>5</sup>Department of Neurology, Stanford University, CA, USA <sup>6</sup>Department of Psychiatry, Massachusetts General Hospital and Harvard Medical School, Boston, MA, USA <sup>7</sup>Division of Nuclear Medicine and Molecular Imaging, Department of Radiology, Massachusetts General Hospital/Harvard Medical School, Boston, MA 8Faculty of Health, Medicine and Life Sciences, School for Mental Health and Neuroscience, Alzheimer Centre Limburg, Maastricht University, Maastricht, The Netherlands <sup>9</sup>The Australian eHealth Research Centre, CSIRO Health & Biosecurity, Victoria, Australia <sup>10</sup>Department of Neurology, Cliniques Universitaires Saint-Luc, Institute of Neuroscience, Université Catholique de Louvain, 10 Avenue Hippocrate, 1200 Brussels, Belgium <sup>11</sup>The Australian eHealth Research Centre, CSIRO Health & Biosecurity, Queensland, Australia <sup>12</sup>The University of Queensland, School of Public Health, Faculty of Medicine, Queensland, Australia <sup>13</sup>Department of Nuclear Medicine and Centre for PET. Austin Health. Victoria, Australia <sup>14</sup>The Department of Medicine, Austin Health.

Corresponding author: Rachel F. Buckley, PhD, Address: Department of Neurology, Level 10, Athinoula A. Martinos Center for Biomedical Imaging, 149 13th St, Charlestown, MA, USA 02129, rfbuckley@mgh.harvard.edu. These authors contributed equally.

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The University of Melbourne, Victoria, Australia <sup>15</sup>Cogstate, Ltd., Australia <sup>16</sup>Department of Neurology, University of Southern California, San Diego, California, USA Words: 4122

#### Abstract

**INTRODUCTION**—Our objective was to investigate the effect of sex on cognitive decline within the context of  $\beta$ -amyloid (A $\beta$ ) burden and apolipoprotein (*APOE*) genotype.

**METHODS**—We analyzed sex-specific effects on A $\beta$ -PET, *APOE* and rates of change on the Preclinical Alzheimer Cognitive Composite-5 (PACC-5) across three cohorts, ADNI, AIBL and HABS (n=755; clinical dementia rating (CDR)=0; Age(SD)=73.6(6.5); Female=55%). Mixed-effects models of cognitive change by sex, A $\beta$ -PET and *APOE*e4 were examined with quadratic time-effects over a median of 4 years of follow-up.

**RESULTS**—*APOE*e4 prevalence and A $\beta$  burden did not differ by sex. Sex did not directly influence cognitive decline. Females with higher A $\beta$  exhibited faster decline than males. *Post-hoc* contrasts suggested that females who were A $\beta$  and *APOE*e4 positive declined faster than their male counterparts.

**DISCUSSION**—Although  $A\beta$  did not differ by sex, cognitive decline was greater in females with higher  $A\beta$ . Our findings suggest sex may play a modifying role on risk of AD-related cognitive decline.

#### Keywords

Preclinical Alzheimer's disease; amyloid; APOE; sex; gender; cognitive decline

#### **1. INTRODUCTION**

Investigating the extent to which the early pathophysiology and cognitive decline vary by sex is critical for understanding the course of Alzheimer's disease (AD) dementia [1]. Epidemiological studies suggest higher incidence rates for women relative to men for AD dementia in older ages [2], however, this is not always supported [3–5]. Studies are mixed with regard to the effect of sex on cognitive decline in clinically-normal older adults [6, 7]. Steeper decline in delayed recall [8], performance IQ and executive function [9] has been reported in female apolipoprotein ɛ4 (APOEɛ4) carriers, however, this finding is not supported by other studies [6, 10]. In addition, some studies suggest males show steeper decline than females in areas of speed, integration and visuospatial ability [11]. A recent meta-analysis found that APOEe4 carriers exhibit particular vulnerability for progression to Alzheimer's disease (AD) dementia relative to males and female APOEe4 non-carriers between the ages 65-75 years [12]. Carrying the APOEe4 allele confers higher risk for abnormal levels of  $\beta$ -amyloid (A $\beta$ ) burden [13] and accumulation [14], and results in steeper cognitive decline when accompanied by high A $\beta$  burden [15, 16]. Hence, it is possible that APOEe4 genetic risk in the presence of A $\beta$  may impart particular susceptibility for females to AD clinical symptoms.

It is unclear whether sex differences in AD dementia risk in those with *APOE*e4 is exacerbated by higher A $\beta$  burden, whether *APOE*e4 and A $\beta$  play independent roles, or whether there are other "downstream" mechanisms that account for the increased vulnerability to dementia risk. Disentangling sex-specific effects with respect to A $\beta$  and *APOE* is particularly relevant during the preclinical stage, given that abnormal levels of A $\beta$ start to accrue decades before the onset of clinical symptoms [17], and that greater focus is being placed on prevention trials [18]. If A $\beta$  burden is differentially associated with cognitive decline in males and females, this will have implications for recruitment and treatment practices in clinical trials.

Our first aim was to ascertain whether sex differences exist in relation to *APOE* and  $A\beta$  burden as estimated by positron emission tomography (PET). Our second aim was to examine the effect of sex on cognitive decline, and whether this was influenced by *APOE*e4 carriage, abnormal  $A\beta$  or both *APOE* and  $A\beta$ . To improve our ability to detect the extent to which females could increase risk for AD biomarkers and cognitive decline, we harmonized data from three well-characterized, longitudinal datasets: the Alzheimer's Disease Neuroimaging Initiative (ADNI), the Australian Imaging, Biomarker and Lifestyle (AIBL) study of ageing, and the Harvard Aging Brain Study (HABS). This allows us sufficient statistical power to determine potentially small magnitude relationships between sex and AD risk.

#### 2. METHODS

#### 2.1. Participants

Cohort-specific inclusion criteria for recruitment have been published previously [19–21]. For the current study, the baseline was considered to be an individual's first A $\beta$ -PET scan. Participants were all required to be clinically-normal at baseline (Global clinical dementia rating (CDR) score = 0, MMSE 24); ADNI's subjective cognitive decline (SCD) group was included in the current study, given that these participants attained a CDR score of 0. Participants were included if their baseline A $\beta$ -PET scan was within 1 year of a neuropsychological testing session (either before or after the scan), and they had at least 2 follow-up neuropsychological assessments after their baseline visit. We excluded participants who carried *APOE*e2/e4 and *APOE*e2/e2 (total < 2.9%), given that the effect of these genotypes on AD risk are unclear. For analysis, 755 participants (ADNI, n=330; AIBL, n=161; HABS, n=268) formed the final participant group. We conducted the procedures for this study under the ethical guidelines stipulated by the Partners Human Research Committee, which is the Institutional Review Board for the Massachusetts General Hospital and Brigham and Women's Hospital.

#### 2.2. Cognitive outcome

We examined cognitive decline using the Preclinical Alzheimer Cognitive Composite score with an additional semantic processing component (PACC-5) [22]. This composite modifies the PACC that was developed as a sensitive measure of A $\beta$ -related cognitive decline [23]. We used the PACC-5 as recent findings suggest the inclusion of a semantic component, specifically the Categories task, adds unique variance associated with A $\beta$ -related cognitive

decline beyond that provided by the original PACC components [22]. Each study used a version of the PACC-5 that has been previously published [23–25]. In each study, the PACC-5 includes some overlapping tests (Mini-Mental State Examination [MMSE] and Logical Memory Delayed Recall) and some non-overlapping tests (for ADNI: ADAS-Cog Word Recall, Trails B, and Categories (Animals); for AIBL: the California Verbal Learning Test (second edition), Digit Symbol substitution, and Categories (Animals/Names); for HABS: the Free and Cued Selective Reminding Test, and Categories (Animals/Vegetables/ Fruit). It is important to note that all three studies had two overlapping tests in the PACC-5 (MMSE and Logical Memory), however, the other three tests were unique to each study. All test scores were standardized within their own study according to the baseline mean and standard deviation of CDR=0 participants from the respective cohort. The PACC-5 was formed by averaging these z-scores. Baseline and longitudinal slopes for the PACC-5 were compared across the three studies to determine whether means and variances were similar (see Appendix A for cross-cohort distributions). ADNI and HABS participants completed these tests approximately every year, whereas AIBL participants underwent testing every 1.5 vears. All available testing sessions following the analysis-defined baseline session were used (for ADNI: 324 participants completed 3 visits, 248 completed 4 visits, 166 completed 5 visits, and 28 completed 6 visits; for AIBL: 155 completed 3 visits, 73 completed 4 visits; for HABS: 244 completed 3 visits, 234 completed 4 visits, 183 completed 5 visits, 116 completed 6 visits). In order to adjust for baseline performance in our models, our cognitive outcome variable was PACC-5 change from baseline (with baseline cognitive performance as a covariate).

#### 2.3. Aß positron emission tomography (PET)

ADNI uses the <sup>18</sup>F-AV45 (Florbetapir or FBP) Aβ-PET tracer, while AIBL and HABS use the <sup>11</sup>C-Pittsburgh Compound-B (PiB) Aβ-PET tracer. The PET acquisition parameters for each study have been published previously [21, 26, 27]. In brief, ADNI and AIBL's PET acquisition time was 50-70 minutes post-injection (http://adni.loni.usc.edu/), while for HABS, PiB-PET data were collected 40-60 minutes post-injection. All raw Aβ-PET data were processed with a standard pipeline. For this pipeline, PET data underwent reconstruction and attenuation correction, were evaluated for head motion, and were coregistered/normalized to a PET template in MNI space using the SPM12 unified segmentation, normalization routine, which applies a rigid body registration, followed by an affine registration, and a nonlinear mapping that fits the image to pre-specified 6-class tissue probability map. Summary measures for regions of interest (ROIs) were computed from a probabilistic GTM-Seg atlas in MNI space (Freesurfer v6.0 [28]) as standard uptake value ratios (SUVrs). The following ROIs that have been validated as AD regions of interest in previous publications [16] were aggregated: the frontal, lateral, and retrosplenial (FLR) regions. Values were normalized against the whole cerebellum to yield an AB FLR SUVr for each participant. To ensure cross-tracer equivalency, we applied a novel nonlinear transformation mapping approach (NLTM; further details in Appendix B). With this equating method, we extracted equivalent FBP SUVrs (FBP<sub>equiv</sub>) for all PiB data so that they conformed to the FBP distribution. Although the NTLM is bidirectional, we chose to conform PiB SUVrs to the FBP distribution due to its more limited dynamic range and scale (Appendix B shows the pre- and post- distributions after applying NLTM in comparison

with raw values and a linear-only transformation). This equivalence allowed for combined analysis of continuous A $\beta$  FLR SUVr. For a post-hoc analysis, we also dichotomized A $\beta$  using a Gaussian mixture modelling procedure [16], which gave an FBP<sub>equiv</sub> cut-off of 1.082 for A $\beta$ + (referred to as A $\beta$ <sub>status</sub> to differentiate between the continuous A $\beta$  measure).

#### 2.4. Statistical analysis

Analyses were performed using *R* version 3.3.2. To determine sex differences in A $\beta$  burden and *APOEe*4 we ran group comparisons using Wilcoxon/Mann-Whitney and chi-square ( $\chi^2$ ) tests, respectively. To investigate sex-specific effects on cognitive decline in association with A $\beta$  and *APOEe*4 status, we ran a series of hierarchical linear mixed models with subjectspecific random intercept nested within cohort. We also modeled cohort as a fixed effect. Covariates were age, years of education, and baseline cognitive performance. As PACC-5 change is best modeled with quadratic time [29, 30], interaction terms with this time effect and all other covariates were included in the models. Only quadratic time terms were modeled, as the global extremum (vertex of parabola) of PACC-5 performance passed through zero at baseline, and an assessment of goodness-of-fit parameters suggested adequate fit against models including the linear terms. We compared goodness-of-fit of increasingly complex models using a log-likelihood ratio test. Multiple independent comparisons (n=5) were accounted for according to a Sidak correction of  $\alpha = 0.01$ .

The following hierarchy of models were run:

- A. PACC-5<sub>change</sub> ~ sex<sup>\*</sup>time + covariates<sup>\*</sup>time
- **B.** PACC-5<sub>change</sub> ~  $[A\beta \text{ OR } APOEe4]^*$ sex<sup>\*</sup>time + covariates<sup>\*</sup>time
- **C.** PACC-5<sub>change</sub> ~  $A\beta^* APOEe4^*$ sex<sup>\*</sup>time + covariates<sup>\*</sup>time
- **D.** PACC-5<sub>change</sub> ~  $A\beta^*$ sex<sup>\*</sup>Age<sup>\*</sup>time + covariates<sup>\*</sup>time

where PACC-5<sub>change</sub> is the change in PACC-5 from baseline, **OR** indicates different terms used in models A and B, A $\beta$  is the continuous FBP<sub>equiv</sub> SUVr, and the covariates were age, years of education and PACC-5 performance at baseline. Note that all lower-order terms were included in each model.

To further explore interactions in Model C between the dichotomized variables  $A\beta_{status}$  (+/ -), *APOEe*4 (+/-) and sex (M/F), pairwise comparisons were performed for the groups: (A $\beta$  -/*APOEe*4-/M (n=191), A $\beta$ -/*APOEe*4-/F (n=220), A $\beta$ -/*APOEe*4+/M (n=41), A $\beta$ -/ *APOEe*4+/F (n=62), A $\beta$ +/*APOEe*4-/M (n=46), A $\beta$ +/*APOEe*4-/F (n=59), A $\beta$ +/ *APOEe*4+/M (n=46), and A $\beta$ +/*APOEe*4+/F (n=54)).

We also ran *post-hoc* analyses within each study to determine whether patterns of findings were consistent with the combined-cohort results.

#### 3. RESULTS

#### 3.1. Cohort characteristics

We examined 755 clinically-normal individuals with a median of 4 years of follow-up in their respective study (range = 3 - 7 years across the combined group), as summarized in

Table 1. ADNI had the shortest follow-up duration (6.08 years), whereas AIBL had the longest follow-up duration (6.98 years). There were no significant differences in follow-up length by sex (t = 0.25, p = 0.80), A $\beta$  status (t = 0.05, p = 0.96), APOE (t = -0.70, p = 0.48), sex<sup>\*</sup>A $\beta$  status (F = 1.61, p = 0.21), or by sex<sup>\*</sup>APOE (F = 0.92, p = 0.34). AIBL participants were significantly younger, and had fewer years of education than the other studies. There were no cohort-level differences in the frequencies of APOEe4 carriers, females/males, or those with A $\beta$ + status. Differences also did not exist in baseline cognitive performance or individual cognitive slopes between the cohorts (cognitive slopes were extracted from ordinary least squares regression models).

#### 3.2. Sex differences in A<sub>β</sub> burden and APOE carrier status

Females did not exhibit greater median A $\beta$  burden (*Diff* = -0.006, CI 95% [-0.023, 0.011], p = 0.48; see Figure 1A and cohort-level findings in Appendix D). We also did not find sex differences according to A $\beta$  status ( $\chi^2 = 0.80$ , Cramer's V = .04, p = 0.37). Females were not more likely be *APOEe*4 carriers ( $\chi^2 = 0.39$ , Cramer's V = .02, p = 0.53; see Figure 1B). Neither were female *APOEe*4 carriers more likely to exhibit greater median A $\beta$  than male *APOEe*4 carriers (*Diff* = 0.010, CI 95% [-0.045, 0.066], p = 0.76).

#### 3.3. Longitudinal change in cognition

For the first set of mixed-effect models which looked at main effects of sex over time, a main effect of sex did not associate with cognitive decline after adjusting for covariates (p = 0.05; see Table 2 for model estimates of terms of interest, with full models to be found in Appendix D).

For the next set of models looking at  $A\beta^*$ sex and *APOE*<sup>\*</sup>sex interactions over time on cognitive decline, females with elevated A $\beta$  exhibited steeper cognitive decline than males with elevated A $\beta$  over time (p = 0.003; see Figure 1C). A comparison between the A $\beta$  main effects model and A $\beta^*$ sex interaction model showed that the latter model fit significantly better than the former, Log Likelihood Ratio (15, 19) = 185.57, p < 0.001. By contrast, we did not find an interaction effect of sex and *APOE* on cognitive decline (p = 0.19).

For the final interaction model between  $A\beta^* \sec^* APOE$  over time, a significant interaction did not exist on cognitive decline (p = 0.17). *Post-hoc* contrasts, however, suggested a weak effect of female *APOEe*4 carriers with high  $A\beta_{\text{status}}$  ( $A\beta + /APOEe4 + /F$ ) having steeper cognitive decline in comparison with male *APOEe*4 carriers with high A $\beta$ status (p = 0.04; see Figure 3). This effect, however, did not survive multiple comparison adjustment.

We next examined the effect of age on the interaction between A $\beta$  and sex on cognitive decline, and found that females with higher A $\beta$  displayed steeper cognitive decline after approximately 80 years of age in comparison with males (see Figure 2). This model did not fit better than a simpler model of A $\beta$ \*sex, Log Likelihood Ratio (17, 23) = 4.55, *p* = 0.60, and as such, we interpreted this finding with caution. We did not consider an interaction between sex, amyloid and education, as our models were not statistically powered to provide reliable estimates (given that only 15 males and 28 females with high A $\beta$  had 12 or less years of education).

When considering each of the three studies in isolation, the same pattern of effects existed for the  $A\beta^*$ sex<sup>\*</sup>time interaction on cognitive decline. Due to statistical power constraints, however, many of these models did not meet the conventional threshold for statistical significance (see Appendix D). Models of  $A\beta^*$ sex<sup>\*</sup>time interactions were also conducted on each z-scored test within the PACC-5 (see Appendix A). We found that decline on listlearning delayed recall was significantly steeper in females than males for a given level of high  $A\beta$  (estimate = -0.07, SE = 0.02, t = -4.07, p < 0.001), with performance on Logical Memory delayed recall trending (estimate = -0.03, SE = 0.02, t = -1.91, p = 0.05). As listlearning tests were different across the cohorts (ADNI = ADAS-Cog Word Recall, AIBL = CVLT, HABS = FCSRT), we argue that these exploratory analyses warrant replication.

#### 4. DISCUSSION

In a large cross-cohort dataset of clinically-normal older adults, females with elevated  $A\beta$  were found to decline in cognition more rapidly than males with a comparable level of  $A\beta$ . This mirrors recent work in ADNI reporting an interactive effect of sex and CSF  $A\beta_{42}$  on episodic memory decline [31]. We did not find an interactive effect of sex and *APOE* on cognitive decline, however, there was a weak effect of female *APOEe*4 carriers with higher  $A\beta$  demonstrating faster rates of cognitive decline in comparison with their male counterparts. There was no main effect of sex on cognitive decline, although females exhibited better cognitive performance at baseline, supporting the notion that females outperform males on tests of verbal memory tasks [32], which are key components of the PACC-5 [22].

The mechanism explaining the interaction between sex and  $A\beta$  remains unclear. In line with previous studies [33–35], sex was not associated with  $A\beta$  burden at baseline (as either a continuous or dichotomous variable), even in *APOEe*4 carriers. It is important to note that sex differences in amyloid load have been reported in relation to family history; for instance, maternal history of sporadic AD dementia influences  $A\beta$ -PET retention in normal subjects [36], and closer proximity to one's parental estimated year of onset of sporadic AD is associated with elevated amyloid burden in female, but not male, subjects [37]. We did not measure family history in the current study, however, it is possible that co-varying for proximity to parental age at onset may highlight differences in amyloid burden between males and females.

An alternative possibility is that sex effects in relation to  $A\beta$  burden do exist, but perhaps at earlier ages, such as during menopause, which was not examined in the current study. Postmortem work shows extensive senile plaque build-up in women who are in the neurofibrillary stages I, II and III compared with men in similar stages, and particularly in those with *APOEe*4 [38]. Animal studies have also reported elevated risk for greater  $A\beta$ burden in females than males in a range of transgenic mouse models [39, 40]; much more so than sex differences in tauopathy [41]. Estrogen is often implicated in findings of sex differences in animal models [42], with ovariectomies increasing  $A\beta$  burden in female mice [43], and estrogen replacement reducing the risk of  $A\beta$  burden [44]. These findings mirror human studies suggesting that higher estradiol in females is associated with better memory performance [45], oophorectomies that occur prior to menopause increases risk of cognitive

impairment in females [46], and epidemiological studies that show chronic (10+ years) use of hormone replacement therapy close to menopause may be a protective factor for AD risk [47]. Clinical trials of estrogen replacement, however, have been disappointing to date [48], highlighting the complex role that sex hormones may play in AD dementia risk [1]. We did not examine menopausal onset in the current study, nor did we measure use of hormone replacement therapy. It will be important for future studies to assess cognitive changes in relation to A $\beta$ , *APOE* and sex prior to and during the full menopause phase.

Other interpretations should also be considered. If similar levels of high  $A\beta$  burden lead to steeper cognitive decline in women, this may imply a greater sensitivity in women to  $A\beta$  burden relative to men, which may be mediated by greater levels of tau and/or neurodegeneration in  $A\beta$  positive females. It is also possible that, despite similar levels of  $A\beta$  burden at baseline, females may accumulate  $A\beta$  at a faster rate, however, the literature has not comprehensively investigated this supposition. In one recent study, proximity to parental estimated year of onset for sporadic AD corresponded with higher  $A\beta$  burden cross-sectionally in females relative to males, but no evidence was found of sex differences in  $A\beta$  accumulation longitudinally [37].

As mentioned above, downstream mechanisms of Aβ burden, such as tauopathy and neurodegeneration, may occur to a greater extent in Aβ positive females than males, that would mirror our finding of greater cognitive decline among the same group. Mounting cross-sectional evidence suggests that females, particularly those with *APOEe*4, exhibit greater levels of tau pathology and neurodegeneration [33, 35, 49, 50]. Female *APOEe*4 carriers show greater levels of CSF total tau [33], hippocampal atrophy [51], cortical thinning [50], and lower intrinsic connectivity in the default network [49] in comparison with male carriers. However, other work, by Jack and colleagues [35] showed that in normal subjects, males have smaller hippocampal volumes in comparison with females, and Koran and colleagues [31] did not find an interaction between sex and CSF total tau on episodic memory decline in a model including normal, MCI and AD dementia patients. As such, further research should focus on the question of whether sex differences in tauopathy or neurodegeneration might drive risk for AD-related cognitive decline.

Unlike previous studies on clinical progression to mild cognitive impairment (MCI) or AD dementia [12, 33, 52], we did not find a sex-by-*APOE* interaction effect on cognitive decline. It is possible that previous *APOE* findings represent proxies for A $\beta$  and/or neurodegeneration effects that are not feasibly measurable in epidemiological cohorts. Previous studies report that risk for clinical progression in female *APOEe*4 carriers exists largely within the specific 65-75 age range [12, 52], suggesting that cognitive decline would need to occur prior to this age. We were unable to accurately estimate cognitive decline in those below 65 years of age due to small sample sizes in that range, and so we may have failed to capture sex-*APOE* effects on cognition. A weak three-way sex-*APOE*-A $\beta$  interaction was found, but this finding did not survive multiple comparison adjustment and thus requires replication.

Another departure from previous studies relates to the age at which sex effects appeared; we found cognitive decline was more likely at later decades, unlike epidemiological studies that

report effects in those between 65-75 years [12]. It is possible that our finding represents a 'survivor bias' effect [53], such that sex-related cognitive decline in the oldest-old is a 'second-wave' risk. Individuals who have survived past factors that reduce the age of onset to AD dementia, such as *APOE* [54], and other factors that increase mortality, such as early male death from cardiovascular disease [55] may result in differential vulnerability to sex effects. As followup time did not differ in our study by sex it is unlikely that our findings are simply driven by the fact that women were followed for longer. Our age interactive model did not significantly explain more variance than simpler models, and so any interpretation of age effects should be considered with this caveat in mind.

Our study has several limitations. We investigated three cohorts of convenience: participants are primarily educated, of higher socioeconomic status, and are not very racially diverse compared to the general population. In addition, the AIBL study is enriched for *APOEe*4 carriers, thus increasing the proportion of carriers over the population prevalence rate [20]. As such, these cohorts may not reflect sex effects in the general population. In addition, we assumed similarity between cohorts, although demographically they were slightly different at baseline. In order to account for this, we processed A $\beta$ -PET data with the same PET-pipeline, and scrutinized baseline and longitudinal cognitive performance to ensure similarity of cognitive performance. We also accounted for random cohort effects in our linear mixed effects models, and ran analyses within each cohort to confirm that the pattern of results was similar across the cohorts. As such, we argue that the advantage of combining cohorts allows idiosyncratic noise from within each cohort to be smoothed out. Nevertheless, these findings require replication in other large datasets, such as Mayo Clinic, the Framingham Heart Study, or even meta-analysis datasets such as the Amyloid Biomarker Study [56].

It is important to note that harmonizing across cohorts is a complex endeavor [57], particularly with regard to equating across neuropsychological test performance. As the PACC-5 composite included both overlapping and non-overlapping tests, we found similar baseline and slope variability across the cohorts. This does not negate the fact that biases may be introduced by including three non-overlapping cognitive tests to form the PACC-5 across the three cohorts. Our future aim will be to further harmonize across neuropsychological tests in order to probe within cognitive domains (e.g. episodic memory, executive function, processing speed, etc). Neuropsychological tests within each cognitive domain largely do not overlap between the cohorts, and as such, it will be necessary to implement item-response theory equating methods, similar to those applied in other cohortharmonization studies [58]. We found that memory, in this case, performance from list-learning and story learning tests from the PACC-5, may be the domain most affected by the sex<sup>\*</sup>A $\beta$  interaction, however, as these tests do not overlap between the studies, our analyses are highly exploratory and need to be replicated in other cohorts.

One strength of our study involves our treatment of the different  $A\beta$ -PET tracers across the studies. In order to address differences in dynamic range, sensitivity and scaling between PiB and FBP A $\beta$ -PET tracers, we employed a non-linear transformation mapping approach. We used this method to account for non-linearities that exist in the extreme ends of the PiB and FBP SUVr distributions. An assessment of cumulative distribution functions between

raw and equivalent SUVrs showed that non-linear mapping preserves the distributional properties of the original SUVr much better than linear transformations (see Appendix B). Regardless, methods to accurately translate SUVrs across A $\beta$  tracers is complex, and will require further refinement and testing in out-of-sample cohorts.

Determining sex-specific effects on rates of cognitive decline in the context of genetic risk and AD pathology will aid in more accurate detection of individuals most vulnerable to AD pathology, and further clarify recruitment and treatment approaches for AD clinical trials with reference to sex. We found that sex differences, both as a main effect and interacted with APOE $\epsilon$ 4, were not apparent on cognitive decline, however, females with high A $\beta$ burden did show steeper cognitive decline trajectories in comparison to males with similarly high A $\beta$  burden. Taken together, these findings suggest that sex effects may be salient in preclinical AD, when A\beta burden is apparent [1]. The mechanism underlying sex-specific sensitivity to  $A\beta$ , however, is yet to be elucidated. Taken together, these findings imply that the influence of sex and sex-specific risk factors on AD risk should be considered with respect to the cascade of events thought to underlie the development of AD dementia, and that the impact of sex is present during the preclinical stage of the disease. The effects observed in this study do, however, highlight the benefit of harmonizing across smaller datasets of clinically normal older adults to detect complex modifying effects in preclinical AD, which will become important at the scalable-level of large clinical trials, such as Anti-Amyloid Treatment in Asymptomatic Alzheimer's Disease (A4) [18].

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# APPENDIX A: Density plots of baseline PACC performance and longitudinal PACC slopes for each cohort

#### APPENDIX B: Visualisation of the Amyloid-PET SUVr distributions

Visualisation of the Amyloid-PET SUVr distributions of (A) raw SUVrs across the three cohorts (AIBL and HABS use PiB, while ADNI uses Florebetapir), (B) linearly transformed SUVrs, and (C) our method of non-linearly transformed SUVrs of PiB to be equated with the Florbetapir distribution

L'L

Non-linear method:

We utilize baseline A $\beta$ -PET datasets of clinically-normal participants from ADNI (FBP), AIBL (PIB), and HABS (PIB). Summary measures were computed via a PET only pipeline with direct spatial normalization of PET data to MNI space. Measurements were made as SUVr using a cortical composite (FLR) with a whole-cerebellum reference region. Crosssample mapping was performed via 10,000 bootstrapped samples of PiB and FBP matched for age, sex and apolipoprotein e4 (APOEe4) status. A transfer function for each bootstrapped sample was generated via smoothed cumulative distribution functions. An SUVr equivalency map and confidence intervals were extracted across these 10,000 fits using sliding-window PCA.

APPENDIX C. Model estimates for each mixed-effects model in the combined group, followed by the model estimates within each cohort for the sex\*amyloid model

	Estimate	Std Error	t value	p value
Model: Sex				
Time <sup>2</sup>	0.03	0.01	2.70	< 0.001
Baseline PACC-5	-0.21	0.03	-6.18	< 0.001
Age	-0.01	0.003	-3.07	0.01
Education	0.01	0.007	1.95	0.07
Sex	0.11	0.04	2.74	0.003
Cohort	-0.03	0.02	-1.28	0.42
Cohort*time <sup>2</sup>	0.02	0.001	9.26	< 0.001
Baseline PACC-5*time <sup>2</sup>	0.003	0.002	1.68	0.06
Age*time <sup>2</sup>	-0.001	0.0001	-6.25	< 0.001
Education*time <sup>2</sup>	0.0004	0.0003	1.12	0.24
Sex*time <sup>2</sup>	-0.003	0.002	-1.99	0.05
Model: Aβ <sup>*</sup> Sex				
Time <sup>2</sup>	0.07	0.01	4.59	< 0.001
Baseline PACC-5	-0.22	0.01	-6.62	< 0.001
Age	-0.009	0.003	-2.90	0.02
Education	0.02	0.007	2.32	0.02
Αβ	-0.31	0.17	-1.79	0.07
Sex	0.02	0.24	0.06	0.95
Cohort	-0.03	0.02	-1.29	0.41
Cohort*time <sup>2</sup>	0.01	0.001	8.90	< 0.001
Baseline PACC-5*time <sup>2</sup>	0.001	0.002	0.82	0.41
Age*time <sup>2</sup>	-0.0008	0.0002	-4.95	< 0.001
Education*time <sup>2</sup>	0.0002	0.0004	0.62	0.54
Aβ <sup>*</sup> Sex	0.10	0.23	0.44	0.66
$A\beta^*$ time <sup>2</sup>	-0.04	0.01	-5.09	< 0.001
Sex*time <sup>2</sup>	0.03	0.01	2.58	0.01
$A\beta^*Sex^*time^2$	-0.03	0.01	-2.96	0.003
Model: APOE e4*Sex				
Time <sup>2</sup>	0.05	0.01	3.51	< 0.001
Baseline PACC-5	-0.21	0.03	-5.98	< 0.001
Age	-0.01	0.003	-2.97	0.003
Education	0.02	0.01	2.11	0.04
APOEe4	-0.03	0.07	-0.48	0.63
Sex	0.11	0.05	2.24	0.03
Cohort	-0.03	0.02	-1.36	0.40
Cohort*time <sup>2</sup>	0.01	0.001	9.24	< 0.001
Baseline PACC-5 <sup>*</sup> time <sup>2</sup>	0.003	0.002	1.59	0.11

	Estimate	Std Error	t value	p value
Age*time <sup>2</sup>	-0.001	0.0001	-6.66	< 0.001
Education*time <sup>2</sup>	0.0003	0.0003	0.83	0.40
APOEe4*time <sup>2</sup>	-0.01	0.004	-3.95	0.001
Sex*time <sup>2</sup>	-0.006	0.003	-2.30	0.02
APOEe4*Sex*time <sup>2</sup>	0.006	0.005	1.32	0.19
Model: Aβ <sup>*</sup> Sex <sup>*</sup> APOE ε4				
Time <sup>2</sup>	0.06	0.02	3.36	< 0.001
Baseline PACC-5	-0.16	0.03	-5.47	< 0.001
Age	-0.009	0.003	-3.15	0.002
APOEe4	0.45	0.34	1.31	0.19
Education	0.01	0.006	2.04	0.04
Αβ	-0.03	0.21	-0.13	0.90
Sex	0.31	0.29	1.09	0.28
Cohort	-0.03	0.02	-1.23	0.43
Cohort*time <sup>2</sup>	0.01	0.001	8.33	< 0.001
Baseline PACC-5*time <sup>2</sup>	0.0003	0.002	0.15	0.88
Age*time <sup>2</sup>	-0.0007	0.0002	-4.42	< 0.0001
APOEe4*time2	0.02	0.02	0.99	0.32
Education*time <sup>2</sup>	0.0003	0.0003	0.91	0.36
$A\beta^*$ time <sup>2</sup>	-0.03	0.01	-2.24	0.03
Sex*time <sup>2</sup>	0.04	0.02	2.30	0.02
$A\beta^*Sex^*time^2$	-0.04	0.02	-2.44	0.01
$A\beta^* APOEe^{4}^* time^2$	-0.02	0.02	-1.17	0.23
Sex*APOEe4*time2	-0.03	0.02	-1.43	0.15
<i>APOEe</i> 4 <sup>*</sup> Aβ <sup>*</sup> Sex	0.42	0.41	1.03	0.30
$A\beta^*Sex^*APOEe4^*time^2$	0.03	0.02	1.37	0.17
Model: Aβ <sup>*</sup> Sex <sup>*</sup> Age				
Time <sup>2</sup>	0.19	0.09	2.14	0.03
Baseline PACC-5	-0.16	0.03	-5.73	< 0.001
Age	-0.001	0.02	-0.04	0.96
Education	0.01	0.006	2.10	0.04
Αβ	0.42	1.79	0.24	0.81
Sex	0.51	2.49	0.20	0.84
Cohort	-0.03	0.02	-1.30	0.42
Cohort*time <sup>2</sup>	0.01	0.001	8.89	< 0.001
Baseline PACC-5*time <sup>2</sup>	0.0009	0.002	0.52	0.60
Age*time <sup>2</sup>	-0.002	0.001	-1.91	0.06
Education*time <sup>2</sup>	0.0001	0.0003	0.45	0.65
$A\beta^* time^2$	-0.17	0.09	-2.41	0.02
Sex*time <sup>2</sup>	-0.31	0.13	-2.41	0.01
$A\beta^*Sex^*time^2$	0.32	0.12	2.60	0.009
$A\beta^*Age^*time^2$	0.002	0.001	1.54	0.12

	Estimate	Std Error	t value	p value
Sex*Age*time <sup>2</sup>	0.004	0.002	2.57	0.01
Age <sup>*</sup> Aβ <sup>*</sup> Sex	0.008	0.03	0.25	0.80
$A\beta^*Sex^*Age^*$ time <sup>2</sup>	-0.005	0.002	-2.80	0.005
Model: Comparisons across Aβ status/APOE e4/Sex groups				
$A\beta - APOEe4 - M^*Time^2 vs A\beta - APOEe4 - F^*Time^2$	0.001	0.002	0.36	0.72
$A\beta - APOEe4 + M^*Time^2 vs A\beta - APOEe4 + F^*Time^2$	-0.004	0.005	-0.78	0.43
$A\beta + APOEe4 - M^*Time^2 vs A\beta + APOEe4 - F^*Time^2$	-0.009	0.005	-1.85	0.06
$A\beta + /APOEe4 + /M^*Time^2 vs A\beta + /APOEe4 + /F^*Time^2$	-0.009	0.005	-2.04	0.04

Note: A  $\beta$  is the FBP SUVr\_equiv. A  $\beta+/-$  is formed according to the A  $\beta$  SUVr cut off = 1.082

# Comparison of A $\beta$ -amyloid\*sex model within each cohort

					Study				
Coefficients		ADNI			AIBL			HABS	
	Estimate	Conf. Int.	p-value	Estimate	Conf. Int.	p-value	Estimate	Conf. Int.	p-value
Fixed Parts									
Intercept	0.69	-0.11-1.49	.090	0.52	-0.66-1.71	.390	0.77	-0.13-1.67	.095
I(timc_yrs^2)	0.12	0.05-0.19	.001	0.04	0.00-0.08	.030	0.12	0.08-0.16	<.001
PACC_bl	-0.20	-0.290.10	<.001	-0.21	-0.370.06	.008	-0.12	- <b>0.21-</b> -0.03	.013
YrsEd	0.01	-0.01-0.03	.321	0.01	-0.04-0.05	.795	0.01	-0.01-0.03	.195
Age	-0.01	-0.02 - 0.00	.005	-0.01	-0.02 - 0.01	.417	-0.01	-0.01 - 0.00	.245
SEXF	0.08	-0.52-0.68	.796	0.18	-0.91-1.27	.749	-0.03	-0.72 - 0.67	.940
Amyloid	0.02	-0.42-0.46	.923	-0.26	-0.93-0.41	.451	-0.66	-1.200.12	.017
I(timc_yrs^2):PACC_bl	0.01	0.00-0.02	.046	-0.00	-0.01-0.00	.633	-0.00	-0.00-0.00	.994
I(timc_yrs^2):YrsEd	0.00	-0.00-0.00	.126	0.00	0.00-0.00	.015	-0.00	-0.00-0.00	.096
I(timc_yrs^2):Agc	-0.00	-0.000.00	.004	-0.00	-0.00-0.00	.269	-0.00	-0.000.00	<.001
I(timc_yrs^2):SEXF	0.01	-0.04-0.06	.665	0.01	-0.02 - 0.05	.550	0.02	-0.01 - 0.05	.140
I(timc_yrs^2):Amyloid	-0.08	-0.120.04	<.001	-0.05	-0.07 - 0.02	<.001	-0.03	-0.05 - 0.01	.005
SEXF:Amyloid	0.00	-0.57-0.57	.999	-0.08	-1.11-0.95	.883	0.14	-0.52 - 0.80	.682
I(timc_yrs^2): SEXF: Amyloid	-0.00	-0.05 - 0.04	.855	-0.01	-0.04-0.02	.550	-0.03	-0.060.00	.045
Random Parts									
$\sigma^2$		0.214			0.131			0.143	
$\tau_{00,ID}$		0.116			0.189			0.119	
N <sub>ID</sub>		330			160			265	
ICC <sub>ID</sub>		0.351			0.590			0.455	
Observations		1102			510			1164	
$R^2/\Omega_0^2$		.613/.591			.761/.749			.653/.642	

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#### **Research in context**

Systematic review: The authors reviewed the literature using Google Scholar and PubMed sources. Sex effects on cognitive decline within the context of Alzheimer's disease biomarkers are not yet as widely published, however, there have been some publications on sex differences in dementia incidence and AD biomarker risk which the authors referred to. These relevant citations are appropriately cited.

Interpretation: Although the sexes did not differ on amyloid burden or *APOE* carrier status, our findings suggest that females decline faster in cognition for a given level of amyloid in comparison with males.

Future directions: In order to examine sex-related amyloid effects on cognition, it will be important to measure neurodegenerative and tauopathy associations. Further, cognitive domains should be investigated; it is possible that certain cognitive domains may be more affected in females than males for a given level of amyloid, however, this remains to be investigated.

## Highlights

• Females are not more likely to exhibit steeper cognitive decline

- Females are not more likely to exhibit high amyloid or carry *APOE*e4 than males
- For a given level of amyloid-PET, females exhibit steeper cognitive decline than males
- For a given level of *APOE* risk, females do not exhibit steeper decline than males

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#### Figure 1.

(A) Females and males display equal proportions of *APOEe*4 carrier status and (B)  $A\beta$  burden. (C) Represents decline in global cognition (as measured by the PACC-5) by  $A\beta$  and sex. These are model estimates from a *continuous* model of  $A\beta$ ; high and low  $A\beta$  are represented by the first quartile (on the left) and third quartile (on the right) of  $A\beta$  along the continuous spectrum. Each line extends to the longest follow-up period within that group.

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#### Figure 2.

Decline in global cognition (as measured by the PACC-5) by  $A\beta$  and sex across the age span (females = pink, males = blue). These are model estimates from a *continuous* model of  $A\beta$ ; high and low  $A\beta$  is represented by the first quartile (on the top) and third quartiles (on the bottom) of  $A\beta$  along the continuous spectrum. Age is also treated continuously in the model, with this visualization showing model estimates at the following ages: 65, 70, 75, 80, 85 years. Each line extends to the longest follow-up period within that group.

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#### Figure 3.

Decline in global cognition (as measured by the PACC-5) by *APOEe*4 status, sex and A $\beta$  status (see legend for colours). These are model estimates from a factorial model with A $\beta$  status represented by cut-off 1.082. Each line extends to the longest follow-up period within that group.

Table 1

Comparison between the studies on demographics and cognitive performance

	ADNI (n = 330) Mean $(SD)$	AIBL (n = 161) Mean $(SD)$	HABS (n=268) Mean (SD)	$F, \chi^2$	Effect size $(\eta_p^2 \text{ or Cramer's } V)$	d
Follow-up (yrs)	3.29 (1.2)	4.99 (1.3)	4.56 (1.3)	129.1	0.26	<0.001
Age	74.6 (6.5)	$71.6\ (6.8)^{*}$	73.4 (6.0)	11.3	0.03	<.001
Females (n/%)	173 (52%)	86 (53%)	159 (59%)	3.1	0.06	0.21
Years of	16.4 (2.7)	13.7 (2.2)*	15.8 (3.0)	53.8	0.12	<0.001
Education						
APOEarepsilon + (n/%)	88 (27%)	49 (32%)	67 (27%)	1.13	0.04	0.56
$A\beta~SUVr~(FBP_{equiv})$	1.04 (0.2)	1.05 (0.2)	1.04 (0.2)	0.23	0.0006	0.79
$A\beta + status \ (n/\%)$	96 (29%)	46 (29%)	76 (28%)	0.01	0.004	0.99
<b>MMSE</b> <sub>baseline</sub>	29 (1.2)	29 (1.2)	29 (1.1)	0.37	0.001	0.69
<b>PACC</b> <sup>baseline</sup>	0.03 (0.6)	0.07 (0.6)	0.02 (0.7)	0.34	0.001	0.71
PACC <sub>slope</sub>	-0.05 (0.2)	-0.03 (0.2)	-0.03 (0.2)	1.01	0.003	0.36

Note: ADNI = Alzheimer's Disease Neuroimaging Initiative; Australian Imaging, Biomarker and Lifestyle (AIBL) study; HABS = Harvard Aging Brain Study (HABS); SD = standard deviation; FBPequiv = Florebetapir equivalent score; APOE = apolipoprotein

#### Table 2

Summary of linear mixed models for PACC-5 change (only terms of interest included)

	Estimate	Std Error	t value	p value
Model A: Sex				
Sex	0.11	0.04	2.73	0.006
Sex*time <sup>2</sup>	- 0.003	0.002	- 1.98	0.05
Model B1: Aβ <sup>*</sup> Sex				
$A\beta^*$ time <sup>2</sup>	- 0.04	0.01	- 5.09	< 0.001
Sex*time <sup>2</sup>	0.03	0.01	2.58	0.01
$A\beta^*Sex^*time^2$	- 0.03	0.01	- 2.96	0.003
Model B2: APOE &4*Sex				
APOEe4*time2	- 0.01	0.003	- 3.95	< 0.001
Sex*time <sup>2</sup>	- 0.005	0.002	- 2.30	0.02
APOEe4*Sex*time2	- 0.006	0.005	1.32	0.18
Model C: Aβ*Sex*APOE ε4				
APOEe4*Sex*time <sup>2</sup>	- 0.03	0.02	- 1.43	0.15
$A\beta^*Sex^*time^2$	- 0.04	0.02	- 2.44	0.01
$A\beta^*Sex^*APOEe^{4}^*time^2$	0.04	0.02	1.37	0.17
Model D: Aβ <sup>*</sup> Sex <sup>*</sup> Age				
$A\beta^*Sex^*time^2$	0.33	0.12	2.70	0.007
$A\beta^*Age^*time^2$	0.002	0.001	1.54	0.12
Sex*Age*time <sup>2</sup>	0.005	0.002	2.70	0.007
Age <sup>*</sup> Aβ <sup>*</sup> Sex	0.008	0.03	0.25	0.80
$A\beta^*Sex^*Age^*time^2$	- 0.005	0.002	- 2.80	0.005
Model: Comparisons across A <sub>βstatus</sub> /APOE e4/Sex groups				
$A\beta - / APOEe4 - / M^* Time^2   A\beta - / APOEe4 - / F^* Time^2$	0.001	0.002	0.36	0.72
$A\beta - / APOEe4 + / M^* Time^2   A\beta - / APOEe4 + / F^* Time^2$	- 0.004	0.005	- 0.78	0.43
$A\beta + APOEe4 - M^*Time^2  A\beta + APOEe4 - F^*Time^2$	- 0.009	0.005	- 1.85	0.06
$A\beta + APOEe4 + M^*Time^2  A\beta + APOEe4 + F^*Time^2$	- 0.009	0.005	- 2.04	0.04

Note:  $A\beta$  is the FBP SUVr<sub>equiv</sub>.  $A\beta$ +/- is formed according to the  $A\beta$  SUVr cut-off = 1.082. Esimates are unstandardized and the reference group for sex is female