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# **OPEN** White matter aberrations and agerelated trajectories in patients with schizophrenia and bipolar disorder revealed by diffusion tensor imaging

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Supported by histological and genetic evidence implicating myelin, neuroinflammation and oligodendrocyte dysfunction in schizophrenia spectrum disorders (SZ), diffusion tensor imaging (DTI) studies have consistently shown white matter (WM) abnormalities when compared to healthy controls (HC). The diagnostic specificity remains unclear, with bipolar disorders (BD) frequently conceptualized as a less severe clinical manifestation along a psychotic spectrum. Further, the age-related dynamics and possible sex differences of WM abnormalities in SZ and BD are currently understudied. Using tractbased spatial statistics (TBSS) we compared DTI-based microstructural indices between SZ (n = 128), BD (n = 61), and HC (n = 293). We tested for age-by-group and sex-by-group interactions, computed effect sizes within different age-bins and within genders. TBSS revealed global reductions in fractional anisotropy (FA) and increases in radial (RD) diffusivity in SZ compared to HC, with strongest effects in the body and splenium of the corpus callosum, and lower FA in SZ compared to BD in right inferior longitudinal fasciculus and right inferior fronto-occipital fasciculus, and no significant differences between BD and HC. The results were not strongly dependent on age or sex. Despite lack of significant group-by-age interactions, a sliding-window approach supported widespread WM involvement in SZ with most profound differences in FA from the late 20 s.

Schizophrenia (SZ) is a major cause of years lived with disability, yet the pathological mechanisms underlying the diverse manifestations of this disorder remain unclear<sup>2</sup>. In line with the notion that the symptoms partly arise from abnormal brain connectivity and functional integration of brain processes<sup>3-5</sup>, histopathological and genetic investigations have implicated lipid homeostasis, neuroinflammation, myelin and oligodendrocyte abnormalities<sup>6-11</sup>. Furthermore, brain imaging has consistently indicated anatomically widespread white matter (WM) microstructural abnormalities using diffusion tensor imaging (DTI)<sup>12–15</sup>. A recent meta-analysis across 29 cohorts of the ENIGMA consortium, including the current, documented significantly lower fractional anisotropy (FA) in patients with SZ (n = 1963) compared to healthy controls (n = 2359) in 20 of 25 regions of interest, and no significant associations with age of onset and medication status<sup>16</sup>.

WM abnormalities have been reported across a wide range of clinical traits and disorders<sup>17-22</sup>, and the diagnostic specificity remains unknown. Current diagnostic nosology treats SZ and bipolar spectrum disorder (BD) as independent categories, but genetic<sup>23</sup>, clinical and neuropsychological<sup>24,25</sup> evidence suggest partly overlapping

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pathophysiology and clinical manifestation, with higher symptom burden, poorer function and worse outcome in  $SZ^{26}$ . Including both patients with SZ and BD in the same analysis is vital for probing common and distinct etiological mechanisms across the psychosis spectrum. While DTI aberrations have consistently been documented in  $BD^{27-31}$ , the existing studies that have included both groups have not provided conclusive evidence of marked group differences between BD and  $SZ^{32-39}$ .

SZ have been conceptualized as a neurodevelopmental disorder<sup>40,41</sup> and deficient myelination during adolescence has been included among the core features of the prodromal phase<sup>42</sup> with WM aberrations present before disease onset<sup>43,44</sup>. Along with evidence of accelerated brain changes in adult SZ<sup>45-47</sup>, neurodevelopmental theories strongly support the need for a dynamic lifespan perspective. The magnitude and modulators of group differences vary across life, e.g. manifested as delayed developmental trajectories<sup>48</sup> and progressive aging-related changes in adulthood<sup>49</sup>. Moreover, whereas the largest meta-analysis to date revealed no significant group by sex interactions, effect sizes for female patients were significantly larger compared to the effect sizes for males for global FA<sup>16</sup> (but see another recent review and meta-analysis which found no significant sex-related differences in effect sizes when comparing patients and controls within males and females, respectively<sup>50</sup>). Although the evidence of strong modulating effects of sex on WM abnormalities in severe mental disorders is lacking, sexual dimorphisms in brain biology and clinical expression warrant further studies on possible sex by diagnosis interactions on the human brain.

In order to address these unresolved issues, our main aim was to compare several DTI indices across the brain between patients diagnosed with a SZ spectrum disorder (n = 128), BD (n = 61), and HC (n = 293), both within and across males and females, using tract-based spatial statistics (TBSS)<sup>51</sup>. Based on the literature we expected widespread WM microstructural alterations in SZ compared to HC, in particular lower FA. Additionally, based on clinical severity, we anticipated moderate group differences between BD and HC, with BD showing less distinct and distributed abnormalities. To comply with a dynamic age-variant perspective, we tested for group by age interactions and compared effect sizes within age cohorts using a sliding window technique<sup>52</sup>, and also tested for group by sex interactions and compared effect sizes within females and males, respectively. DTI based indices of WM microstructure are highly sensitive to differences in data quality, e.g. due to subject motion, which may bias the results<sup>53–55</sup>. Since previous studies have often failed to report quality control (QC) measures or simply omitted systematic QC altogether, it is unknown if reported group effects are biased due to differences in data quality. Hence, we employed a stringent multi-step exclusion protocol based on quantitative quality assessment, and compared groups at different levels to assess the relevance.

#### Results

**Demographics and clinical characteristics.** Table 1 summarizes demographics and clinical characteristics. There was a significant main effect of group on age (F = 4.2, p = 0.016), education (F = 34.1, p < 0.001) and IQ (F = 37.3, p < 0.001), with higher age in HC compared to SZ (Supplementary Fig. S1), longer education and higher IQ (HC > BD > SZ). Compared to BD, SZ had higher symptom severity as measured by The Positive and Negative Syndrome Scale (PANSS) total (t = 8.2, p < 0.001), positive (t = 7.6, p < 0.001), negative (t = 6.4, p < 0.001), and disorganized (t = 4.2, p < 0.001) sub-scales, and the split version of Global Assessment of Functioning Scale split version (GAF) with GAF function (t = -5.5, p < 0.001) and GAF symptom (t = -7.4, p < 0.001).

Main effects of diagnostic groups on DTI. Figure 1 and summarize results from voxelwise analyses testing for main effects of group on the DTI indices. We found significant and widespread main effects of group on FA and radial diffusivity (RD), including the corpus callosum, superior longitudinal fasciculus, fornix, cingulum, forceps major and inferior fronto-occipital fasciculus. Pairwise comparisons revealed widespread FA reductions and RD increases in SZ compared to HC, and FA reductions in SZ compared to BD. No other group comparisons yielded significant effects.

Global DTI measures and sliding window approach. Figure 2 shows mean skeleton DTI values plotted as a function of age and group. Table 2 summarizes the results from linear models accounting for age, age², sex and diagnosis. We found significant main effects of group, sex and age on FA. There was no significant group by age or group by age² interaction; therefore the model was run without the interaction terms. Pairwise comparisons revealed lower FA in SZ compared to HC, and lower FA in SZ compared to BD. Females showed significantly lower FA compared to males. Figure 2, Supplemental Fig. S2 and Table 3 summarize results from the bootstrapped age fitting procedure, yielding estimates of the mean and standard deviation of age at maximum FA and minimum RD, MD and axial diffusivity (AD) within groups. The overlap in confidence intervals (Table 3) and the comparison against empirical null distributions generated using permutation testing (see Methods and Supplemental Fig. S2) revealed no significant between-group differences in age at maximum (FA) or minimum (RD, MD, and AD). Supplemental Figs S3–S6 summarize the bootstrapped age fitting procedure for regions-of-interest (ROI). Briefly, the majority of ROIs show a consistent pattern with early peak in FA in SZ compared to BD and HC. However, the results from left and right cingulum have a more intricate pattern with BD reporting an older age peak. For RD, AD and MD the trend is less consistent and more complex.

Effect sizes within different age-bins from the sliding-window technique are presented in Supplementary Fig. S7. For FA, MD, RD, effect sizes for HC vs. SZ increased until the late 20 s. Effect sizes for SZ vs. BD showed a similar pattern for FA, and more complex non-linear associations for MD, RD and AD. Effect sizes for HC vs. BD for FA straddled around 0 throughout the sampled age range. For RD, MD, and AD the effect sizes showed more complex non-linear associations.

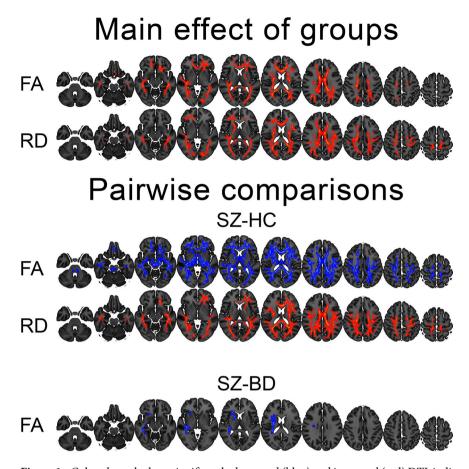
	Schizophre	onia (\$7)			Healthy Controls	ANOVA/Chi-Square Analysis/t Tests			
	(n=128)	ema (SZ)	Bipolar (B	D) $(n = 61)$	(HC) $(n=293)$	$F/\chi^2/t$	p	Post Hocb	
Demographics					'		'	•	
Age, years <sup>c</sup> 29.38 (8.59)			31.74 (11.5	8)	31.91 (7.48)	F=4.2	0.016	HC > SZ	
Sex, n, (% male)	77 (60.2%)		29 (47.5%)		172 (58.7%	$\chi^2 = 3.0$	0.221		
Handedness n, (% right)	86 (92.5%)		49 (87.5%)		221 (86.3%)	$\chi^2 = 2.4$	0.296		
Ethnicity, n, (% Caucasian)	81 (79.4%)		49 (83.1%)		252 (94.4%	$\chi^2 = 46.0$	< 0.001		
Education, years <sup>d</sup>	12.40 (2.43	)	13.78 (2.13	)	14.52 (2.11)	F=34.1	< 0.001	HC > BD > SZ	
WASI (IQ)	100.29 (14.	78)	109.32 (10.	81)	112.62 (10.73)	F=37.3	< 0.001	HC > BD > SZ	
Age of onset, yearse	20.74 (6.74	)	21.33 (7.84	:)		t = -0.49	0.628		
Duration of illness, years <sup>f</sup>	8.86 (7.74)		10.20 (9.14	:)		t = -0.96	0.342		
Symptom Ratings <sup>g</sup>							'	'	
PANSS total score	57.0 (15.4)		42.4 (8.0)			t = 8.2	< 0.001		
PANSS positive score	9.1 (4.4)		5.3 (2.2)			t = 7.6	< 0.001		
PANSS negative score	12.5 (5.3)		8.2 (3.3)			t = 6.4	< 0.001		
PANSS excited score	5.3 (2.1)		4.8 (1.4)			t = 1.6	0.102		
PANSS depressed score	7.6 (3.1)		7.4 (2.6)			t = 0.0	0.749		
PANSS disorganized score	5.2 (2.5)		4.0 (1.4)			t = 4.2	< 0.001		
GAF symptom score	46.6 (14.4)		60.5 (10.0)			t = -7.4	< 0.001		
GAF function score	47.2 (13.9)		58.6 (12.2)			t = -5.5	< 0.001		
Medication n (%) DDI		DDD	n (%)	DDD					
First-generation antipsychotic	8 (8.2)	1.9 (2.8)	1 (2.1)	0.9 (NA)					
Second-generation antipsychotic	79 (85.9)	1.2 (0.9)	30 (62.5)	0.9 (0.9)					
Lithium	8 (8.7)	0.8 (0.2)	12 (25)	1.1 (0.4)					
Antiepileptic	tiepileptic 5 (5.4) 0.5 (0.2)		14 (29.2) 1.1 (0.3)						
Antidepressants	30 (32.6)	2.2 (3.5)	15 (31.3)	1.3 (0.8)					

**Table 1.** Demographic and clinical data<sup>a</sup>. Number of missing data: Handedness: schizophrenia: 35, bipolar: 5, healthy controls: 37, Ethnicity: schizophrenia: 33, bipolar: 2, healthy controls: 26, Education: schizophrenia: 25, bipolar: 6, healthy controls: 38, IQ: schizophrenia: 37, bipolar: 4, healthy controls: 35, Age of onset: schizophrenia: 17, bipolar disorder: 3, Duration of illness: schizophrenia: 17, bipolar: 3, PANSS score: schizophrenia: 14, bipolar: 2, GAF score: schizophrenia: 15, bipolar: 2, medication: schizophrenia: 36, bipolar: 13. Abbreviations: ANOVA, univariate analysis of variance; DDD, defined daily doses, in accordance with guidelines from the World Health Organization Collaboration Center for Drug Statistics Methodology (http://www.whocc.no/atcdd); GAF, Global Assessment of Functioning, IQ, intelligence quotient; PANSS, Positive and Negative Syndrome Scale; WASI, Wechsler Abbreviated Scale of Intelligence. <sup>a</sup>Means and standars deviations are reported unless otherwise is specified. Analyses of demographics and clinical data were performed in R (www.r-project.org). <sup>b</sup>Tukey post hoc tests. <sup>c</sup>Age was defined as the age of magnetic resonance scanning. <sup>d</sup>Years of education refers to the total number of years of completed education as reported by the participant. <sup>e</sup>Age of onset was defined as age at first contact with the mental health service due to a psychiatric symptom (depression, psychosis, mania or hypomania). <sup>f</sup>Duration of illness was defined as number of years between age of onset and age at MRI. <sup>g</sup>PANSS five factor model was used <sup>103,104</sup>.

**Sex related differences.** Mean skeleton and ROI analyses revealed no significant sex-by-diagnosis interaction effects on DTI WM metrics. Supplementary Table S1 shows results from group comparisons within females and males, respectively. Briefly, the analysis revealed main effects of group on FA both in males (F = 3.36, p = 0.036) and females (F = 3.99, p = 0.02). Pairwise comparisons revealed lower FA in SZ compared to HC in both sexes and lower FA in SZ compared to BD in females.

**Associations with symptom domains.** Mean skeleton and ROI analyses revealed no significant associations with GAF and PANSS domain scores across patient groups. Global and ROI-based t- and p-statistics are summarized in Supplementary Table S2 and Supplementary Fig. S8, respectively.

**Effects of quality control.** Visual inspection of the datasets with QC summary z-score below -2.5 (n = 35, see below) indicated no clear reason for exclusion. Therefore, the main analyses were run on the entire dataset, but we also present results using varying QC levels. Figure 3 summarizes the effects of QC on the mean skeleton data. Effect sizes for HC vs. SZ and BD vs. SZ increased with QC stringency for all metrics except AD, which showed a more complex pattern. The effect size for HC vs. BD remained relative unchanged as a function of QC. Voxelwise analysis revealed highly similar patterns as those obtained using the full sample (Fig. 1 and Supplementary Fig. S9).



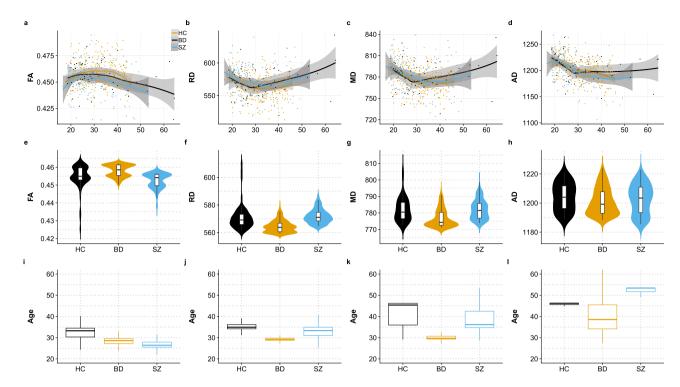
**Figure 1.** Colored voxels show significantly decreased (blue) and increased (red) DTI-indices in SZ patients relative to HC and BD. Group differences are thresholded at p < 0.05 (two-tailed) after permutation testing using threshold free cluster enhancement (TFCE). Note that the white matter skeleton has been slightly thickened to aid visualisation.

**Subgroup and age restricted analyses.** Density plots showing the distribution of the four DTI metrics for each of the subgroups within each diagnostic group are presented in Supplementary Fig. S10. The results from the subgroup analyses are presented in Supplementary Table S3, while the age restricted analyses are presented in Supplementary Table S4. Briefly, but not limited to, for the diagnostic subgroups there were significant differences between psychosis not otherwise specified (PNOS) and a strict SZ diagnosis for global MD (p = 0.011, uncorrected) and global AD (p = 0.002, uncorrected). A similar pattern was observed for the BD subgroups, with significant difference between BDI and BDII for MD (p = 0.035, uncorrected) and AD (p = 0.035, uncorrected). For the age- restricted analyses (55 years and younger) we observed main effect of group for FA only, with pairwise comparisons indicating lower global FA in SZ compared to HC.

**ROI analyses.** Figure 4 and Table 4 summarize the ROI results. Most ROIs showed main effects of group for FA ( $\eta_{ps}^2$ : 0.001–0.029), with strongest effects in the body (BCC) and splenium (SCC) of the corpus callosum and forceps major. We found substantial effects of group in 7 and 1 of the 23 ROIs in RD and AD. We found a nominal significant (p < 0.05, uncorrected) age by group interactions for MD in the BCC ( $\eta_{ps}^2$  < 0.013, p = 0.044), indicating larger group differences with increasing age. Since no age by group interactions remained after corrections for multiple comparisons, all main effects and results from pairwise comparisons were computed without the interaction term in the models.

#### Discussion

One of the major implications of the brain dysconnectivity hypothesis of psychotic disorders is that the WM microstructural layout and integrity modulates risk and give rise to a range of symptoms. In order to test this hypothesis, we compared DTI metrics between patients with SZ and HC across the brain. Our inclusion of a group of patients with BD allowed us to test for diagnostic specificity or, conversely, cross-diagnostic convergence. In line with our primary hypothesis and converging evidence<sup>16</sup>, the results revealed robust differences between patients with SZ and HC on several metrics, in particular lower FA across the brain in patients, even after careful quality assessment. In general, the results were not strongly dependent on age or sex and we found no significant associations with symptoms across groups. Adding to the accumulated evidence of brain gray matter abnormalities in patients with severe mental illness<sup>52,57</sup>, these results support converging evidence implicating WM abnormalities in SZ and suggest these abnormalities are more pronounced for SZ than for BD.



**Figure 2.** Plots (**a**–**d**) Mean skeleton DTI values plotted as a function of age and group (HC = healthy controls, BD = bipolar disorders, SZ = schizophrenia spectrum disorders). Plots (**e**–**h**) Violin plot depicting the fitted values for each group. Plots (**i**–**l**) Uncertainty estimates of the age within each group when maximum FA, minimum RD, MD and AD are reached from a bootstrap procedure with 10 000 resamples.

	SZ Mean (SD) 95%CI	BD Mean (SD) 95% CI	HC Mean (SD) 95% CI	F	<b>p</b> <sup>(a)</sup>	Pairwise comparison	Cohens d HC/ SZ,HC/BD, BD/SZ	Age <sup>2</sup>	Age <sup>2</sup>	Age t	Age p	Sex t	Sex p
FA	0.452 (0.01) 0.450-0.455	0.455 (0.01) 0.451-0.458	0.458 (0.01) 0.457-0.460	7.02	<0.001*	HC > SZ, BD > SZ	0.35, 0.03, 0.21	-3.97	< 0.001	3.49	< 0.001	4.98	< 0.001
MD	782.04 (22.01) 778.19–785.89	782.11 (20.16) 776.95–787.27	776.71 (19.22) 774.50–778.92	0.82	0.437		-0.11, -0.06, -0.03	5.06	< 0.001	-5.63	< 0.001	-1.16	0.247
RD	571.82 (22.04) 567.97-575.68	571.06 (21.65) 565.52–576.61	564.49 (20.34) 562.15–566.83	2.96	0.052		-0.23, -0.07, -0.09	5.21	< 0.001	-5.29	< 0.001	-2.24	0.025
AD	1202.49 (27.94) 1197.60-1207.38	1204.20 (22.96) 1198.32-1210.08	1201.16 (22.08) 1198.62–1203.69	0.71	0.492		0.11, -0.01, 0.09	3.69	< 0.001	-5.08	< 0.001	1.11	0.273

**Table 2.** Mean skeleton DTI metrics within groups. *Note*. Abbreviations: FA; fractional anisotropy, MD; mean diffusivity, RD; radial diffusivity and AD; axial diffusivity. MD, RD, AD multiplied by 1 000 000 to preserve precision. The p-values are not corrected for multiple testing. There was no significant group by age interaction or group by age² interaction; therefore the model was run without these interaction terms. The effect size represent the group by age interaction in the age squared model ( $x \sim \text{group} + \text{age} + \text{age}^2 + \text{sex} + \text{group:age}$ ). \*Bonferroni corrected p < 0.05.

Clinical overlaps have motivated a dimensional approach to reveal common and distinct disease mechanisms in SZ and BD. In line with cognitive and genetic studies<sup>23,24</sup> suggesting several commonalities, brain imaging has not revealed structural or functional brain characteristics unambiguously distinguishing the two disorders, and previous DTI studies comparing SZ and BD have been largely inconclusive<sup>32,34,35,58</sup>.

The neurobiological underpinnings of DTI metrics are complex and multidimensional, and our findings do not allow for interpretation regarding specific cellular processes. Previous studies have shown associations between RD and myelin related processes<sup>59,60</sup>, and higher RD may suggest reduced myelin integrity in SZ, in particular when considered in light of genetic studies reporting altered expression of genes involved in lipid homeostasis and myelination<sup>9,61</sup>. Complementary models implicate microglial inflammatory processes and oxidative stress in WM pathology<sup>62,63</sup>, and inflammation-related cytokines and growth factors have been associated with reduced FA and increased RD and MD in BD<sup>64</sup>. Further studies are needed to delineate the roles of myelination and inflammation for WM integrity and mental health across the lifespan<sup>65,66</sup>.

Considering the strong impact of age on brain WM microstructure<sup>67,68</sup>, characterizing age trajectories within groups may provide indirect information about the temporal evolution of aberrations in an ontogenetic perspective. Both neurodevelopmental and neurodegenerative models of the development and sustainment of psychosis

	Age SZ Mean (SD) 95% CI	Age BD Mean (SD) 95% CI	Age HC Mean (SD) 95% CI
FA	26.57 (2.7) 26.50–26.63	27.15 (4.4) 27.03–27.16	32.42 (2.8) 32.34–32.48
RD	33.70 (4.4) 33.63-33.77	29.44 (1.5) 29.37–29.51	36.09 (4.1) 36.02-36.16
MD	39.24 (7.0) 39.14–39.35	30.73 (3.0) 30.63-30.83	41.6 (5.1) 41.50-41.71
AD	50.69 (4.8) 50.55-50.84	42.03 (11.6) 41.89-42.17	45.87 (0.7) 45.73-46.01

**Table 3.** Age where maximum FA or minimum RD, MD or AD were reached. *Note*. Abbreviations: FA; fractional anisotropy, MD; mean diffusivity, RD; radial diffusivity and AD; axial diffusivity, SD; standard deviation, 95% CI; 95% confidence intervals are reported. The sample was bootstrapped using 10 000 iterations.

have been formulated 46,47,69, and both genetic liability and neurodevelopmental perturbations play critical roles in the modulation of risk 43,44. Indeed, the emergence of psychotic symptoms in late adolescence and early adulthood may in fact reflect late stages of the disorder 42. Patients with early onset SZ show microstructural aberrations 70, which could manifest as delayed WM development during adolescence 48. The current lack of group-by-age interactions may suggest the observed group differences are explained by events prior to the sampled age-range, which would indicate parallel age trajectories in adulthood 47.

For HC, we observed a peak of FA at approx. 32 years followed by decreases until the maximal sampled age, which is highly corresponding with previous cross-sectional studies<sup>67</sup>. The FA trajectory for SZ (approx. 27 years) and BD (approx. 27 years) showed an earlier peak, followed by a linear decrease until the maximum age. Although permutation testing revealed no significant differences in age at peak FA, the sliding window approach, which provide further insight beyond a standard age-by-diagnosis interaction test, suggested the magnitude of the group differences are not completely invariant to age, with indications of increasing group differences in FA between SZ and HC until the late 20 s. In addition to suggesting early and possibly accelerating age-related differences in SZ, moderate age-related differences in effect sizes may also reflect a combination of clinical heterogeneity, sampling issues, and power. The lack of significant age by group interactions possibly also hints at the shortcomings of simple models for delineating and comparing complex trajectories<sup>71</sup>. Future studies utilizing a longitudinal design including a wide age-range and participants at genetic or clinical risk who have still not developed psychosis are needed to characterize the trajectories of the dynamic WM aberrations during the course of brain maturation and disease development<sup>22</sup>.

Although detrimental effects of subject motion and other sources of noise on DTI metrics have been documented \$3,55,72\$, most previous DTI studies on psychotic disorders have not provided sufficient details regarding the employed QC procedures or included any quantitative QC measures as covariates. It is therefore largely unknown to which degree various sources of noise have contributed to the reported group differences. A major strength of this study is the use of an automated approach for identifying and replacing slices with signal loss due to bulk motion, considerably increasing temporal signal-to-noise ratio (tSNR<sup>21</sup>), and a comprehensive QC protocol including both manual and automated quantitative measures. Comparisons of summary statistics and group differences at different steps in the QC and exclusion procedure revealed a tendency of increasing group differences in FA between SZ and HC with the exclusion of noisy data. These results indicate that stringent QC may increase sensitivity to WM aberrations in SZ, and suggest that future studies should carefully address different sources of noise in their datasets before interpreting their findings as reflecting relevant pathophysiology.

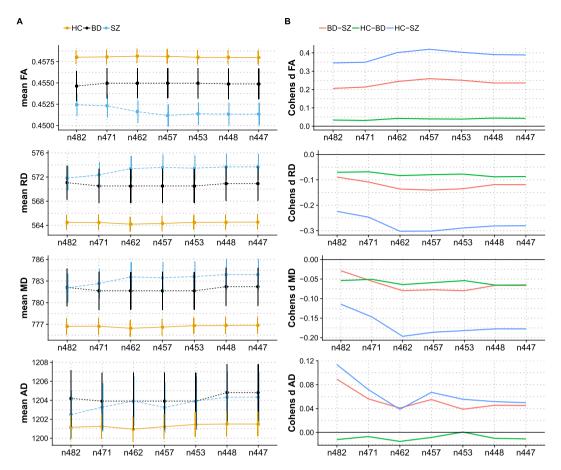
Some limitations should be considered while interpreting our findings. The influence of medication on WM is debated<sup>73–75</sup>. As most patients were medicated, with the majority of patients taking antipsychotics, confounding effects of medication cannot be ruled out. Future studies with an appropriate design for assessing medication effects are needed. Despite our current lack of significant sex by diagnosis interaction, there is a growing appreciation of sexual dimorphisms in brain and behavior both in health and disease, which warrants further investigations. Further, our cross-sectional design is not suitable for delineating dynamic individual changes in WM microstructure, and further studies utilizing a prospective design in younger children and adolescents are needed to map microstructural changes to risk and development of psychosis. Integrating a wider range of MRI modalities with clinical, cognitive and genetic features<sup>21,76</sup>, and including microstructural indices based on multi-compartment diffusion models (e.g., Neurite Orientation Dispersion and Density Imaging<sup>77,78</sup>, free water imaging<sup>79</sup>, or restriction spectrum imaging<sup>80</sup>), cortical and subcortical morphometry and functional measures, may prove helpful for increasing diagnostic sensitivity and specificity.

In conclusion, we report widespread WM microstructural aberration in patients with SZ compared to BD and HC. We found no significant differences between patients with BD and HC, suggesting the biophysical processes causing DTI based WM abnormalities in severe mental disorders are more prominent for SZ. These results are in line with converging genetic and pathological evidence implicating neuroinflammatory and lipid and myelin processes in SZ pathophysiology.

#### Methods

**Sample.** Adult patients were recruited from psychiatric units in four major hospitals Oslo. Patients had to fulfill criteria for a Structured Clinical Interview (SCID)<sup>81</sup> DSM-IV diagnosis of schizophrenia spectrum disorder, collectively referred to as SZ (n=128 including schizophrenia (n=70), schizoaffective (n=18), schizophreniform (n=7)) and psychosis not otherwise specified (n=33), or bipolar spectrum disorder, collectively referred to as BD (n=61 including BDI (n=39), BDII (n=17) and BD NOS (n=5)). The sample comprised both medicated (n=133), unmedicated (n=7) and patients missing information regarding medication status (n=49).

293 healthy controls from the same catchment area were invited through a stratified randomized selection from the national records. Exclusion criteria for both patients and HC included hospitalized head trauma,

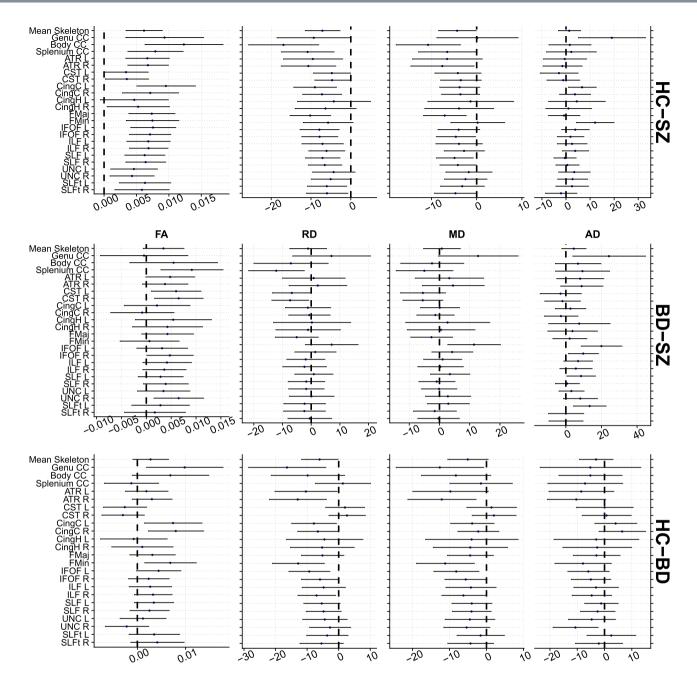


**Figure 3.** Mean of skeleton DTI metrics plotted across quality control subgroup analyses (**A**) and Cohens d for pairwise comparisons across quality control subgroup analyses. (**B**) The labels on the x-axes reflect the number of participants in each analysis. The error bars of part A represent the standard error of the mean.

neurological disorder or IQ below 70. In addition, HC were screened with a questionnaire about severe mental illness and the Primary Care Evaluation of Mental Disorders (PRIME-MD)<sup>82</sup>. Exclusion criteria included somatic disease, substance abuse or dependency the last 12 months or a first-degree relative with a lifetime history of severe psychiatric disorder (SZ, BD, or major depressive disorder). The Tematisk Område Psykoser (TOP) Study is approved by the Regional Ethics Committee (REK Sør-Øst C, 2009/2485) and the Norwegian Data Inspectorate (2003/2052). Study protocol and procedures adhered to the ethics approval and to the Declaration of Helsinki. All participants provided written informed consent, see SI for more information regarding neuropsychological and clinical assessment. Due to confidentiality and privacy of participant information data may not be shared readily online, but data can be requested by contacting the authors.

**MRI acquisition.** Imaging was performed on a General Electric (Signa HDxt) 3 T scanner using an 8-channel head coil at Oslo University Hospital. For DTI, a 2D spin-echo whole-brain echo planar imaging pulse with the following parameters was used: repetition time: 15 s; echo time: 85 ms; flip angle: 90°; slice thickness: 2.5 mm; in-plane resolution: 1.875 \* 1.875 mm; 30 volumes with different gradient directions (b =  $1000 \, \text{s/mm}^2$ ) in addition to two b = 0 volumes with reversed phase-encode (blip up/down) were acquired.

**DTI processing.** Image analyses and tensor calculations were performed using FSL<sup>83–85</sup>. Pre-processing steps included topup (http://fsl.fmrib.ox.ac.uk/fsl/fslwiki/topup)<sup>86</sup> and eddy (http://fsl.fmrib.ox.ac.uk/fsl/fslwiki/eddy)<sup>87,88</sup> to correct for geometrical distortions and eddy currents. Topup uses information from the reversed phase-encode blips, resulting in pairs of images with distortions going in opposite directions. From these image pairs the susceptibility-induced off-resonance field was estimated and the two images were combined into a single corrected one. Eddy detects and replaces slices affected by signal loss due to bulk motion during diffusion encoding, which is performed within an integrated framework along with correction for susceptibility induced distortions, eddy currents and motion<sup>88</sup>. In order to assess the effect of replacement of dropout-slices on tSNR we also processed the data using eddy without slice replacement (Supplementary Fig. S11). Briefly, mean tSNR was significantly (t = 25.76, p < 0.001) lower when running eddy without slice replacement (mean: 7.77, SD: 0.52) compared to with slice replacement (mean 8.79, SD: 0.70). There was no significant group differences in the amount of slices replaced (F = 1.046, p = 0.352, mean group slice replacement: HC: 10.92 ( $\pm$ 7.36), BD: 10.52 ( $\pm$ 7.40), SZ: 12.46 ( $\pm$ 9.42)).



**Figure 4.** Results from region of interest (ROI) analyses with mean difference and variance from pairwise comparisons plotted for each DTI metric. The error bars represent 95% confidence intervals. List of abbreviations: Genu of corpus callosum (GCC), Body of corpus callosum (BCC), Splenium of corpus callosum (SCC), Anterior thalamic radiation L (ATR L), Anterior thalamic radiation R (ATR R), Corticospinal tract L (CST L), Corticospinal tract R (CST R), Cingulum (cingulate gyrus) L (CGL), Cingulum (cingulate gyrus) R (CG R), Cingulum (hippocampus) L (CGH L), Cingulum (hippocampus) R (CGH R), Forceps major (FMI), Forceps minor (FMI), Inferior fronto-occipital fasciculus L (IFO L), Inferior fronto-occipital fasciculus R (IFO R), Inferior longitudinal fasciculus R (ILF R), Superior longitudinal fasciculus R (SLF R), Uncinate fasciculus L (UNC L), Uncinate fasciculus R (UNC R), Superior longitudinal fasciculus (temporal part) L (Temporal SLF L), Superior longitudinal fasciculus (temporal part) R (Temporal SLF R).

Diffusion tensor fitting was done using dtifit in FSL. FA is a scalar value of diffusion directionality, MD was computed as the mean of all three eigenvalues, RD as the mean of the second and third eigenvalue<sup>89</sup>, while AD represent the principal eigenvalue.

Prior to statistical analyses we employed a stepwise QC procedure, including maximum voxel intensity outlier count (MAXVOX)<sup>90</sup> and tSNR<sup>90</sup>. Since reduced data quality due to subject motion and other factors may bias the results in clinical studies, we defined various quantitative QC metrics and tested for group differences

	SZ	BD	НС	F	p	Pairwise comparison	ηp2 <sup>(a)</sup>
FA					,		Ť
Mean FA	0.45	0.45	0.46	7.02	<0.001*	HC > SZ, BD > SZ	0.029
Genu of corpus callosum	0.69	0.69	0.70	3.73	0.025		0.015
Body of corpus callosum	0.70	0.70	0.71	6.15	0.002	HC>SZ	0.025
Splenium of corpus callosum	0.79	0.79	0.79	7.18	<0.001*	HC > SZ, BD > SZ	0.029
Anterior thalamic radiation L	0.46	0.47	0.47	6.01	0.003	HC > SZ, BD > SZ	0.025
Anterior thalamic radiation R	0.47	0.47	0.47	6.99	<0.001*	HC > SZ, BD > SZ	0.029
Corticospinal tract L	0.59	0.59	0.59	3.36	0.035	BD > SZ	0.014
Corticospinal tract R	0.57	0.58	0.57	4.05	0.018	BD > SZ	0.017
Cingulum (cingulate gyrus) L	0.50	0.50	0.51	5.81	0.003	HC>SZ	0.024
Cingulum (cingulate gyrus) R	0.49	0.49	0.50	3.84	0.022	HC>SZ	0.016
Cingulum (hippocampus) L	0.41	0.42	0.42	1.72	0.181		0.007
Cingulum (hippocampus) R	0.43	0.44	0.44	1.98	0.139		0.008
Forceps major	0.54	0.54	0.54	6.98	<0.001*	HC>SZ	0.029
Forceps minor	0.50	0.50	0.51	4.46	0.012	HC>SZ	0.018
Inferior fronto-occipital fasciculus L	0.46	0.46	0.47	6.71	<0.001*	HC>SZ	0.027
Inferior fronto-occipital fasciculus R	0.47	0.47	0.48	7.51	<0.001*	HC > SZ, BD > SZ	0.031
Inferior longitudinal fasciculus L	0.43	0.44	0.44	6.12	0.002	HC>SZ, BD>SZ	0.025
Inferior longitudinal fasciculus R	0.40	0.40	0.41	7.94	<0.001*	HC > SZ, BD > SZ	0.032
Superior longitudinal fasciculus L	0.44	0.44	0.44	5.63	0.004	HC>SZ	0.023
Superior longitudinal fasciculus R	0.44	0.44	0.44	6.20	0.002	HC > SZ, BD > SZ	0.025
Uncinate fasciculus L	0.41	0.41	0.42	2.38	0.093		0.010
Uncinate fasciculus R	0.42	0.43	0.42	4.61	0.010		0.019
Superior longitudinal fasciculus (temporal part) L	0.44	0.44	0.45	4.02	0.019	HC>SZ, BD>SZ	0.017
Superior longitudinal fasciculus (temporal part) R	0.48	0.48	0.48	2.48	0.085	-	0.010
MD	<u> </u>					1	1
Mean MD	782.0	782.1	776.7	0.83	0.437	T	0.003
Genu of corpus callosum	855.3	869.3	856.5	1.01	0.364		0.004
Body of corpus callosum	822.6	820.1	811.4	2.65	0.071		0.011
Splenium of corpus callosum	722.8	717.4	715.1	1.77	0.171		0.007
Anterior thalamic radiation L	790.6	793.4	783.5	0.58	0.560		0.002
Anterior thalamic radiation R	808.2	812.5	800.5	1.98	0.139		0.008
Corticospinal tract L	709.7	703.2	703.8	1.65	0.192		0.007
Corticospinal tract R	714.0	707.3	709.0	1.94	0.146		0.008
Cingulum (cingulate gyrus) L	767.2	767.3	762.6	0.50	0.606		0.003
Cingulum (cingulate gyrus) R	769.3	767.4	764.8	0.30	0.444		0.002
Cingulum (hippocampus) L	801.6	801.9	796.8	0.10	0.904		0.003
Cingulum (hippocampus) R	818.3	815.9	811.6	0.10	0.824		0.000
	768.5	765.0	760.2	3.61	0.024		0.001
Forceps major Forceps minor	847.6		847.4		0.028		0.013
Inferior fronto-occipital fasciculus L		858.9	-	2.01	-		-
*	793.9	797.6	789.1	1.19	0.305		0.005
Inferior fronto-occipital fasciculus R	798.1	798.6	793.1	1.25	0.287		0.005
Inferior longitudinal fasciculus L	780.3	780.1	775.2	0.38	0.682		0.002
Inferior longitudinal fasciculus R	793.1	795.5	789.2	0.79	0.455		0.003
Superior longitudinal fasciculus L	753.7	752.1	747.7	1.25	0.286		0.005
Superior longitudinal fasciculus R	757.6	756.7	752.5	0.94	0.390		0.004
Uncinate fasciculus L	811.5	813.2	808.6	0.03	0.974		0.000
Uncinate fasciculus R	829.9	831.5	826.3	0.14	0.870		0.001
Superior longitudinal fasciculus (temporal part) L	744.4	742.4	739.8	0.47	0.623		0.002
Superior longitudinal fasciculus (temporal part) R	761.2	760.5	755.8	0.79	0.455		0.003
RD		1	1		1	1	
Mean RD	571.8	571.1	564.5	2.960	0.053		0.012
Genu of corpus callosum	439.3	449.8	433.0	1.740	0.177		0.007
Body of corpus callosum	417.3	411.8	401.3	5.023	0.007	HC < SZ	0.021
	290.6	279.3	279.6	5.477	0.004	HC < SZ, BD < SZ	0.023
Splenium of corpus callosum	250.0						
Splenium of corpus callosum  Anterior thalamic radiation L	577.5	578.7	568.0	1.710	0.182		0.007

	SZ	BD	нс	F	p	Pairwise comparison	ηp2 <sup>(a)</sup>
Corticospinal tract L	443.1	436.3	437.8	2.713	0.067	Turi wise comparison	0.011
Corticospinal tract E  Corticospinal tract R	453.5	446.2	448.5	3.417	0.034	BD < SZ	0.011
Cingulum (cingulate gyrus) L	527.2	527.5	518.4	3.142	0.044	HC < SZ	0.013
Cingulum (cingulate gyrus) R	539.0	538.9	531.4	2.595	0.076	110 (02	0.011
Cingulum (hippocampus) L	615.3	614.0	608.1	0.131	0.877		0.001
Cingulum (hippocampus) R	615.5	612.4	606.9	0.741	0.477		0.003
Forceps major	496.6	491.5	485.8	6.376	0.002	HC < SZ	0.026
Forceps minor	582.8	591.8	578.5	1.666	0.190	110 (02	0.007
Inferior fronto-occipital fasciculus L	576.7	578.6	568.8	3.308	0.037	HC < SZ	0.014
Inferior fronto-occipital fasciculus R	573.5	572.1	566.0	3.722	0.025	HC < SZ	0.015
Inferior longitudinal fasciculus L	581.6	579.6	574.0	2.071	0.127		0.009
Inferior longitudinal fasciculus R	611.4	612.3	604.9	2.374	0.094		0.010
Superior longitudinal fasciculus L	563.9	562.4	556.4	2.915	0.055		0.012
Superior longitudinal fasciculus R	566.1	564.8	559.5	2.635	0.073		0.011
Uncinate fasciculus L	617.7	617.9	613.1	0.439	0.645		0.002
Uncinate fasciculus R	622.9	620.7	617.6	1.092	0.336		0.005
Superior longitudinal fasciculus (temporal part) L	554.7	552.9	548.1	1.749	0.175		0.007
Superior longitudinal fasciculus (temporal part) R	543.3	543.4	537.6	1.209	0.299		0.005
AD	1		1				1
Mean AD	1202.5	1204.2	1201.2	0.71	0.492		0.003
Genu of corpus callosum	1687.3	1708.2	1703.5	4.81	0.009		0.020
Body of corpus callosum	1633.1	1636.7	1631.6	0.38	0.688		0.002
Splenium of corpus callosum	1587.2	1593.7	1586.0	0.72	0.488		0.003
Anterior thalamic radiation L	1216.8	1222.6	1214.4	0.27	0.764		0.001
Anterior thalamic radiation R	1246.4	1253.0	1243.6	0.67	0.512		0.003
Corticospinal tract L	1242.9	1237.0	1235.8	0.31	0.734		0.001
Corticospinal tract R	1235.1	1229.6	1229.9	0.21	0.808		0.001
Cingulum (cingulate gyrus) L	1247.1	1246.9	1250.9	2.65	0.072		0.011
Cingulum (cingulate gyrus) R	1229.9	1224.5	1231.4	0.86	0.425		0.004
Cingulum (hippocampus) L	1174.2	1177.8	1174.3	0.68	0.508		0.003
Cingulum (hippocampus) R	1223.7	1222.8	1221.0	0.12	0.890		0.001
Forceps major	1312.4	1312.1	1308.9	0.20	0.819		0.001
Forceps minor	1377.3	1393.0	1385.2	7.33	<0.001*		0.030
Inferior fronto-occipital fasciculus L	1228.3	1235.5	1229.6	2.33	0.099		0.010
Inferior fronto-occipital fasciculus R	1247.3	1251.7	1247.1	1.08	0.340		0.005
Inferior longitudinal fasciculus L	1177.6	1181.2	1177.6	0.78	0.458		0.003
Inferior longitudinal fasciculus R	1156.3	1162.0	1157.9	2.10	0.124		0.009
Superior longitudinal fasciculus L	1133.4	1131.4	1130.2	0.20	0.820		0.001
Superior longitudinal fasciculus R	1140.5	1140.5	1138.6	0.23	0.798		0.001
Uncinate fasciculus L	1199.1	1203.7	1199.6	1.33	0.266		0.006
Uncinate fasciculus R	1243.9	1253.1	1243.8	3.11	0.046		0.013
Superior longitudinal fasciculus (temporal part) L	1123.9	1121.5	1123.2	1.34	0.262		0.005
Superior longitudinal fasciculus (temporal part) R	1197.2	1194.9	1192.2	0.06	0.945		0.000

**Table 4.** Anatomical regions of interest (ROI) analyses. Note. Abbreviations: FA; fractional anisotropy, MD; mean diffusivity, RD; radial diffusivity and AD; axial diffusivity. MD, RD, AD multiplied by 1 000 000 to preserve precision. MD, RD, AD multiplied by 1 000 000 to preserve precision. <sup>a</sup>Partial eta squared. <sup>\*</sup>False Discovery Corrected using Benjamini and Hochberg methods.

within different QC strata. Specifically, we devised a semi-qualitative QC protocol including methods provided in DTIPrep<sup>91</sup> and tSNR<sup>90</sup>. Supplementary Fig. S12 shows a flowchart of the QC protocol. At each step the distributions of the quality metrics were visually inspected. In our step-wise exclusion protocol, datasets were excluded based on a summary score utilizing (1) maximum MAXVOX<sup>90</sup> and (2) tSNR<sup>90</sup>. The summary score was formed by first inverting the MAXVOX score, z-normalize both scores independently, add 10 to each of the z-scores (to avoid negative values), and then computing the product of the two. This product was then z-normalised, with low scores indicating worse quality. In an iterative fashion, subjects with a QC sum z-score below -2.5 were excluded, and the group statistics were recomputed. This was repeated until no datasets had a z-score below -2.5. Briefly, the slice-wise check and the MAXVOX screens the DWI data for intensity related artifacts while tSNR is a global summary measure. See SI for further details regarding the QC such as summary stats for each step of the QC procedure (Supplementary Table S5), demographic overview of excluded participants (Supplementary Table S6),

density plots of DTI metrics before and after exclusion (Supplementary Fig. S13) and voxel-wise analyses after QC (Supplementary Fig. S9). In short, a thorough inspection of the excluded and included participants after QC suggested that general quality of the data is good. Thus, we present results on the full dataset with supplemental and complementary results from a stringent QC.

Voxelwise analysis of FA, MD, AD and RD were carried out using TBSS $^{51}$ . FA volumes were skull-stripped and aligned to the FMRIB58\_FA template supplied by FSL using nonlinear registration (FNIRT) $^{92}$ . Next, mean FA were derived and thinned to create a mean FA skeleton, representing the center of all tracts common across subjects. The same warping and skeletonization was repeated for MD, AD and RD. We thresholded and binarized the mean FA skeleton at FA > 0.2 before feeding the data into voxelwise statistics.

**Statistical analyses.** Voxelwise statistical analyses were performed using permutation testing, implemented in FSL's randomise<sup>93</sup>. Main effects of diagnosis on FA, RD, MD and AD were tested using general linear models (GLM) by forming pairwise group contrasts and corresponding F-tests. Since previous studies have documented strong curvilinear relationships between DTI features and age throughout the adult lifespan<sup>67</sup>, we included age, age<sup>2</sup> and sex as covariates. The data was tested against an empirical null distribution generated by 5000 permutations and threshold free cluster enhancement (TFCE)94 was used to avoid arbitrarily defining the cluster-forming threshold. Voxelwise maps were thresholded at p < 0.05 and corrected for multiple comparisons across space. Mean FA, MD, RD and AD across the brain and within significant clusters were submitted to R<sup>95</sup> for peak estimation and to compute effect sizes and visualization. In a resampling with replacement (bootstrapping) procedure we fitted the DTI data to age using local polynomial regression function (LOESS). LOESS has previously been used in lifespan studies<sup>67</sup> and avoids some of the shortcoming of polynomial models for age fitting<sup>71</sup>. Using boot package<sup>96,97</sup> in R, we repeated the age fitting procedure for each of the 10,000 bootstrapped samples for each group to estimate the mean age at the maximum (FA) or minimum (MD, RD, AD) value across iterations, and its uncertainty with confidence intervals calculated using the adjusted bootstrap percentile method. Additionally, the group differences in age at peak FA was tested against an empirical null distribution generated by 10000 permutations, generated by randomly shuffling group labels and computing the pairwise group differences at each iteration. All pairwise differences were combined into one null distribution, and the differences in the true data were compared to this common null, enabling correction for multiple comparisons across all pairwise comparisons.

We tested for associations between GAF/PANSS domains and FA, MD, RD and AD across both patient groups (SZ and BD grouped together) in the whole brain and within specific regions, covarying for age,  $age^2$  and sex. False discovery rate (FDR)<sup>98</sup> and Bonferroni was used to correct for multiple testing.

Differences between and within subgroups (see SI for more information), group by age and group by age<sup>2</sup> interactions on the mean skeleton DTI metrics were tested. In order to account for heterogeneity in the diagnostic groups we ran subgroup analyses on the mean skeleton metrics on the largest subgroups (strict SZ, PNOS, BDI and BDII). Additionally, in a control analysis to confirm that possible differences in age distribution did not influence the main results we excluded participants over 55 years of age and ran mean skeleton analyses across groups.

We performed a sliding window technique to obtain effect sizes for each of the pairwise group comparison within different age-strata. Utilizing the zoo R package<sup>99</sup>, we slid a window of 150 participants in steps of 5 participants along the sorted age span. At each step, we computed a linear model investigating effects of diagnosis, accounting for sex. We plotted the resulting t-values and effect sizes (Cohen's d) representing pairwise group differences against the mean age of each sliding group and fit a LOESS function using ggplot2 in R<sup>100</sup>. In order to test if group differences varied between females and males we reran the analysis when including a sex-by-diagnosis interaction term for the mean skeleton and ROI analyses.

To facilitate future meta-analyses, we calculated raw mean DTI values across the skeleton and within various anatomical regions of interest (ROIs) based on the intersection between the TBSS skeleton and probabilistic atlases <sup>101,102</sup>. R was used for further analysis, including linear models with each of the ROI DTI value as dependent variable, diagnostic group and sex as fixed factors, and age and age<sup>2</sup> as covariates.

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#### **Author Contributions**

S.T. and L.T.W. designed the study. S.T., T.K. and L.T.W. performed the analyses with contributions from D.A. and N.T.D., S.T., L.T.W. and T.K. interpreted the results. S.T., L.T.W. and T.K. drafted the manuscript. All authors reviewed and approved the manuscript

## **Additional Information**

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**Competing Interests:** MSc Tønnesen, Dr Kaufmann, Dr Doan, Dr Alnæs, Dr Córdova-Palomera, Dr van der Meer, Dr Rokicki, Dr Moberget, Dr Gurholt, Dr Haukvik, Dr Ueland, Dr Lagerberg, Dr Agartz, Dr Andreassen, and Dr Westlye, report no conflicts of interest.

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