



A meta-analysis on the cardiac safety profile of domperidone compared to metoclopramide

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Abstract

Background: This meta-analysis aimed to assess the cardiac safety profile of domperidone treatment for the risk of cardiovascular (CV) event and QT prolongation.

Methods: Data from nine studies involving 101,155 patients were used for the analysis of CV event risk, while data from eight studies involving 390 patients were used for the analysis of QT prolongation risk.

Results: Meta-analysis findings suggested a significant increase in CV risk under domperidone as compared to no treatment for domperidone doses of >30 mg/day (OR: 3.14, 95% CI, 1.191 to 8.304, $p = 0.021$), no significant increase in QT prolongation event rates with domperidone (3.54%, 95% CI, 1.73% to 7.10%) and a significantly lower CV risk for domperidone than for metoclopramide (OR: 0.63, 95% CI, 0.58 to 0.70, $p < 0.001$).

Conclusions: The present meta-analysis indicates that domperidone treatment may not be associated with an overall CV event risk increase at doses ≤ 30 mg/day and does not result in QT prolongation.

Keywords

Domperidone, metoclopramide, dosage, cardiovascular risk, sudden cardiac death, ventricular arrhythmia, QT prolongation, meta-analysis

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Introduction

Domperidone (DMP), a peripheral dopamine D₂-receptor antagonist with prokinetic and antiemetic properties, has been broadly prescribed for nausea and vomiting, gastroparesis and gastroesophageal reflux disease and stimulation of breast milk production.^{1–4}

DMP and metoclopramide (MCP) are dopamine antagonists with a strong affinity for dopamine receptors in the central as well as peripheral nervous systems, and specifically the gastrointestinal tract, and thus they act as antiemetics at the chemoreceptor trigger zone and as prokinetics in the upper gastrointestinal tract to accelerate gastric emptying.⁵ Although MCP readily crosses the blood-brain barrier leading to central nervous system side effects in up to 40% of patients ranging from somnolence to extrapyramidal symptoms, DMP poorly penetrates the blood-brain barrier but maintains a powerful antiemetic effect at the chemoreceptor trigger zone level as well as its peripheral prokinetic properties.⁵ Accordingly, DMP is considered a safer alternative to MCP in patients with intolerance to MCP treatment or those

requiring long-term therapy for upper gastrointestinal motility problems in whom nausea and vomiting are prominent.⁵

Although the safety profile of DMP is more acceptable than MCP and cisapride,³ hazardous cardiovascular (CV) adverse effects are considered likely because of its narrow therapeutic index.^{6,7}

The putative mechanism by which DMP delays cardiac repolarization and prolongs the QT interval is considered to involve blockage of IKr, the rapid component of the delayed rectifier potassium current,⁸ via inhibiting the potassium efflux channel, human

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ether-a-go-go-related gene, leading to a prolongation of cardiac repolarization.^{8–11} A prolonged QT interval is considered a predictive, noninvasive risk factor for sudden cardiac death (SCD) since a delay in ventricular repolarization can lead to a more chaotic cardiac depolarization/repolarization cycle and can provoke arrhythmias with high risk of SCD, such as ventricular fibrillation and torsade de pointes.^{5,12}

Published data on prolongation of QT interval and potential adverse CV events such as ventricular arrhythmias (VAs) and/or SCD under DMP treatment are largely limited to case reports and case-control studies alongside inconsistent findings.^{2,4,5,7,10,13–32} Therefore, a need for further investigation with large clinical trials is emphasized to address the safety profile of DMP more appropriately.^{2,5}

Until such a study is published, to evaluate this serious claim, we aimed with the present meta-analysis to assess the association between different doses of DMP and CV events and QT prolongation risk in comparison to MCP.

Materials and methods

Search strategy and study selection

The PubMed database was searched using the following terms: (a) “domperidone” AND at least one of the terms “cardiovascular, QT, death, arrhythmia” and (b) “domperidone” AND “metoclopramide” AND “cardiovascular,” in titles and abstracts limited to articles published in English and studies in humans. The literature was searched from inception to June 30, 2018. The authors evaluated the title and abstract of each article separately. The full text of 115 articles was reached.

Study selection

We included studies that met the following criteria: (1) a clinical study published as an original article; (2) exposure of interest was oral DMP treatment, (3) outcome of interest was a CV event or QT prolongations, and (4) available odds ratio (OR) with 95% confidence interval (CI) (or data to calculate these) was provided. Exclusion criteria were: (1) preclinical studies, (2) case reports and meta-analyses, and (3) lack of sufficient information on CV events or QT interval measurement.

Publication bias

Potential publication bias was explored using visual inspection of Begg’s funnel plot asymmetry, classic fail-safe N analysis, Begg’s rank correlation and Egger’s weighted regression tests.^{33,34}

Heterogeneity assessment

The I^2 statistics of Higgins and Thompson was used to assess heterogeneity among studies.³⁵ I^2 values of 25%, 50% and 75% represent low, moderate and high heterogeneity, respectively.³⁶ The presence of heterogeneity across studies was defined using Cochran’s Q test with a 0.10 significance level.³⁷

Quality of studies

The quality of studies was assessed using Jadad scoring³⁸ for clinical studies, using the Newcastle-Ottawa Scale for Case-Control Studies for case-control studies, and using the Newcastle-Ottawa Scale for Cohort Studies for single-arm clinical studies.^{39,40}

Jadad scoring has three items with a score range of 0 to 5. The first item is related to randomization (0, non-randomized; 1, randomized but the sequence of the randomization was not reported; 2, randomized appropriately). The second item is related to double-blinding (0, not double-blinded; 1, double-blinded but the details were not reported; 2, appropriate double-blinding techniques were performed). The third item is related to withdrawals and dropouts (0, number and reasons for withdrawals were not stated; 1, number and reasons for withdrawals were stated). An a priori cutoff value for Jadad score to include the studies was not set.³⁸

The Newcastle-Ottawa Scale for Case-Control studies is based on the selection of case and controls, comparability of cases and controls and ascertainment of exposure, while the Newcastle-Ottawa Scale for Cohort studies is based on the selection of cohorts, comparability of cohorts and assessment of outcome. “High”-quality choices are identified with a star, with a maximum of one star for each item within the “Selection” and “Exposure/Outcome” categories and a maximum of two stars for “Comparability.” For case-control studies a statement of no history of disease or incident and demonstration that outcome of interest was not present at start of study earns a star, whereas for cohort studies, adequacy of follow-up of cohorts is based on assessment of the follow-up of the exposed and nonexposed cohorts to ensure that losses are not related to either the exposure or the outcome.^{39,40}

Statistical methods

ORs with a 95% CI for CV event risk were presented based on calculated risk (using raw data provided in each study) as well as adjusted risk obtained from original data provided in each study. Fixed-effects and random-effects models were applied to all comparisons to determine corresponding overall effect sizes and related CIs when heterogeneity was absent or evident,

respectively. Only random-effects model findings are presented here since there was no significant difference between both models for all variables except for DMP vs MCP. For analysis of QT prolongation risk, event rates, difference between pretreatment and posttreatment means for QT interval values and Hedge's *g*, value with 95% CIs were used. All statistical analyses were performed with Comprehensive Meta-Analysis, v.2.2.064 (www.Meta-Analysis.com, USA). Summary statistics are expressed as mean ± standard deviation (SD), *n* (%) and 95% CI, where appropriate. A two-sided *p* < 0.05 was considered statistically significant.

Results

Included studies

Of 115 studies initially selected, 98 studies were excluded because of inclusion of preclinical data (*n* = 39), lack of data on the relationship of DMP

with CV risk or QT prolongation (*n* = 25), lack of original data or appropriate statistical analysis (*n* = 19), inclusion of data on parenteral use of DMP (*n* = 7), case reports (*n* = 4), no data on CV risk or QT prolongation (*n* = 3) and meta-analysis (*n* = 1). Of 17 studies included in the meta-analysis, data from nine studies^{17–25} were used to analyze DMP-related risk of a CV event, as well as for CV risk related to DMP vs MCP^{17–19,25} and related to different doses of DMP.^{17,18,23,25} Data from the remaining eight studies^{3,26,32} were used to analyze DMP-related risk of QT prolongation. Figure 1 illustrates the corresponding flow diagram of the study selection.

Data on quality of studies are provided in Table 1.

CV event risk

Characteristics of studies. Characteristics of studies^{17,25} selected for the analysis of DMP/MCP-related CV event risk are presented in Table 1. All studies were

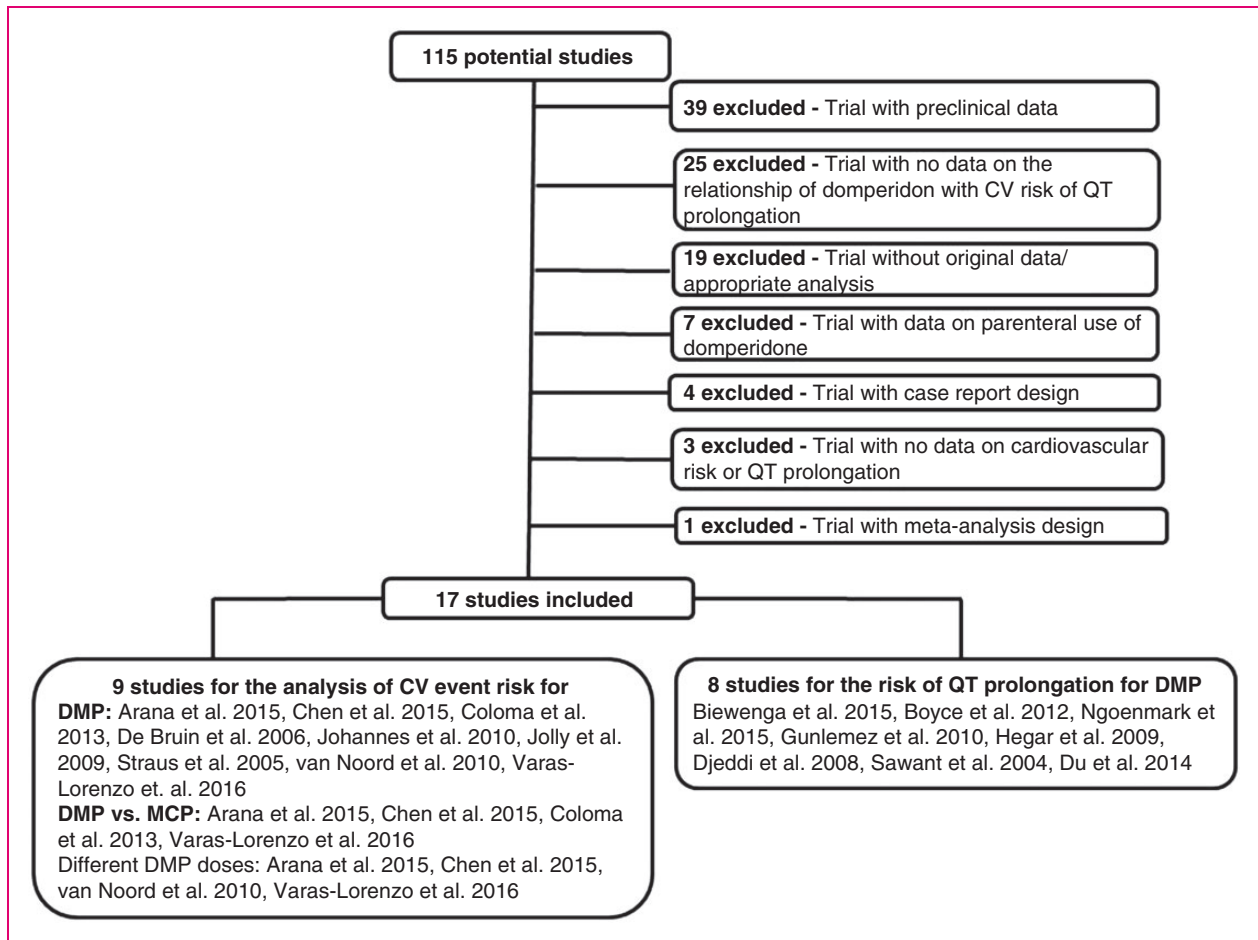


Figure 1. Flow diagram of study selection. CV: cardiovascular; DMP: domperidone; MCP: metoclopramide.

Table 1. Characteristics of studies included in the meta-analysis.

Domperidone and metoclopramide – CV event risk										
Author (year)	Country	Study design	Age (year) (mean±SD)	Treatments	included in meta-analysis for CV event risk related to			Cardiac outcome parameter(s)		
					DMP	MCP	DMP vs. MCP		DMP doses	
Arana et al. (2015) ¹³	UK	Case control ^a Crossover	44.0±23.0	DMP (<30 / 30 / >30 mg/day) & MCP	+	+	+	+	SCD	
Varas-Lorenzo et al. (2016) ²¹	China	Crossover	60.8±19.1	DMP (≤30 / >30 mg/day) & MCP	+	+	+	+	VA and SCD	
Coloma et al. (2013) ¹⁵	EU	Case-control	All ages	DMP & MCP	+	+	+ ^b	-	AMI	
De Bruin et al. (2007) ¹⁶	NL	Case-control	47.5±26.8 59.6±21.7	DMP	+	-	-	-	CA	
Johannes et al. (2010) ¹⁷	USA	Case-control	79.4	DMP	+	-	-	-	VA and SCD	
Jolly et al. (2009) ¹⁸	UK	Case-control	67.6±12.4	DMP	+	-	-	-	SCD	
Straus et al. (2005) ¹⁹	NL	Case-control	71±13 69±13	DMP	+	-	-	-	SCD	
van Noord et al. (2010) ²⁰	NL	Case-control	66.3±13.9 72.5±14.1	DMP (<30 / 30 / >30 mg/day)	+	-	-	+	VA and SCD	
Domperidone – QT prolongation risk										
Author (year)	Country	n	Study design	Mean Age	Age group	QT definition	QT value			QTc prolongation (>500 msc)
							Pre-tx	Post-tx	Difference	
Biewenga et al. (2015) ²²	Belgium	44	TQT	43.5 y	18-55 y	QTcP	408.0 ^c	400.8	-7.2	0
Boyce et al. (2012) ²³	UK	24	TQT	26.6 y	18-39 y	QTcF	401.4	407.3	5.8	0
Ngoenmak et al. (2016) ³	Thailand	22	clinical	8.5 mo	<2 y	QTcB	411.0	395.5	-15.5	2
Gunlemez et al. (2010) ²⁴	Turkey	40	clinical	32 day	Premature	QTcB	370.0	380.0	10.0	2
Hegar et al. (2009) ²⁵	Belgium	10	clinical	5.6 mo	Infant	QTcB	404.0	402.0	-2.0	0
Djeddi et al. (2008) ²⁶	France	31	clinical	18 day	Newborn	QTcB	373.2	387.2	14.0	1
Sawant et al. (2004) ²⁷	India	27	clinical	35.0 y	18-60 y	QT	320.0	310.0	-10.0	0
Du et al. (2014) ²⁸	China	192	clinical	41.3 y	18-65 y	QT	-	-	-	0

(continued)

Table 1. Continued.

Quality assessment		Quality assessment category		
	Selection	Comparability	Outcome	
Case-control studies^d				
Arana et al. (2015) ¹³ /	☆☆☆☆	☆☆	☆☆☆	
Varas-Lorenzo et al. (2016) ²¹	☆☆☆☆	☆☆	☆☆☆	
Chen et al. (2015) ¹⁴	☆☆☆	☆☆	☆☆☆	
Coloma et al. (2013) ¹⁵	☆☆☆	☆☆	☆☆☆	
De Bruin et al. (2007) ¹⁶	☆☆☆☆	☆☆	☆☆☆	
Johannes et al. (2010) ¹⁷	☆☆☆	☆☆	☆☆☆	
Jolly et al. (2009) ¹⁸	☆☆☆☆	☆☆	☆☆☆	
Straus et al. (2005) ¹⁹	☆☆☆☆	☆☆	☆☆☆	
van Noord et al. (2010) ²⁰	☆☆☆☆	☆☆	☆☆☆	
Single-arm studies^e				
	Selection	Comparability	Outcome	
Ngoenmak et al. (2016) ³	☆☆☆	-	☆☆☆	
Gunlemez et al. (2010) ²⁴	☆☆☆	-	☆☆☆	
Djeddi et al. (2008) ²⁶	☆☆☆	-	☆☆☆	
Clinical trials^f				
	Randomization	Double-blinding	Withdrawals and dropouts	
Biewenga et al. (2015) ²²	2	2	1	
Boyce et al. (2012) ²³	2	2	1	
Hegar et al. (2009) ²⁵	2	0	1	
Sawant et al. (2004) ²⁷	1	0	0	
Du et al. (2014) ²⁸	2	2	1	

AMI: Acute myocardial infarction; CA: Cardiac arrest; DMP: Domperidone; Eu: Europe; MCP: Metoclopramide; NL: Netherlands; SCD: Sudden cardiac death; UK: United Kingdom; VA: Ventricular arrhythmia; QTc: thorough QT/QTc; tx: treatment.
^aIncluded in the meta-analysis, ^bcalculated from original data provided in the study, ^cvalues in italic refer to data not provided in the publication but calculated using the other data available in the study, ^dNewcastle-Ottawa Scale for Case-Control Studies, ^eNewcastle-Ottawa Scale for Cohort Studies, ^fJadad score for clinical trials

retrospective except two studies (one with crossover and the other is a case control in a small group of patients). Cardiac outcome analysis was based both on VA and SCD in three studies,^{18,21,24} while on SCD per se in four studies,^{17,22,23,25} acute myocardial infarction risk in one study¹⁹ and cardiac arrest accompanied with or without death in one study²⁰ (Table 1).

CV event risk–DMP. A meta-analysis of 101,155 patients (33,432 in the CV event group (1873 on DMP) and 67,723 in the control group (1751 on DMP)) from nine studies^{17–25} suggested a 1.56 to 4.70 times higher rate of DMP use in patients with a CV event than those without a CV event, and suggested a significant increase in CV risk under DMP treatment (OR: 2.07, 95% CI, 1.50 to 2.84, $p < 0.001$ with a random-effects model; $Q = 68.883$, $I^2 = 89.8\%$ for adjusted risk) (Table 2, Figure 2(a)).

CV event risk–MCP. A meta-analysis of 66,523 patients (28,595 in the CV event group (1353 on MCP) and 37,928 in the control group (975 on MCP)) from four studies^{17–19,25} suggested a 1.47 to 5.70 times higher rate of MCP use in patients with a CV event than those without a CV event, and suggested a significant increase in CV risk under MCP treatment (OR: 3.27, 95% CI, 1.10 to 9.75, $p = 0.033$ with a random-effects model; $Q = 463.06$, $I^2 = 99.6\%$ for adjusted risk) (Table 2, Figure 2(b)).

CV event risk–DMP vs MCP. A meta-analysis of DMP vs MCP use in terms of CV event risk revealed the ratio of DMP/MCP use to range from 0.40 to 1.06 in patients with a CV event and suggested a lower CV risk for DMP than for MCP (OR: 0.64, 95% CI, 0.58 to 0.70, $p < 0.001$ with a fixed-effects model and OR: 0.61, 95% CI, 0.33 to 1.15, $p = 0.129$ with a random-effects model; $Q = 67.23$, $I^2 = 97.0\%$ for adjusted risk) (Table 2, Figure 2(c)).

CV event risk–DMP doses. A meta-analysis of 109,126 patients (19,801 in the CV event group (229 on DMP > 30 mg/day in 204), 79,325 in the control group (186 on DMP > 30 mg/day in 116)) from four studies^{17,18,24,25} suggested no significant risk of a CV event at doses of <30 mg/day (OR: 1.64, 95% CI, 0.509 to 5.286, $p = 0.407$) and 30 mg/day (OR: 1.37, 95% CI, 0.693 to 2.699, $p = 0.366$). These findings suggested a significant increase in CV risk at doses of >30 mg/day (OR: 2.09, 95% CI, 1.59 to 2.75, $p < 0.001$ with a fixed-effects model and OR: 3.14, 95% CI, 1.19 to 8.30, $p = 0.021$ with a random-effects model) (Table 2).

Publication bias (DMP-/MCP-CV event risk). Visual inspection of the funnel plot and Egger regression asymmetry

test (intercept (95% CI): 0.292 (–3.88 to 4.46), $p = 0.435$), classic fail-safe N analysis ($n = 426$) and Begg-Mazumdar correlation test (Kendall tau $b = 0.393$, $p = 0.087$) showed no evidence of publication bias in the analysis between DMP use and CV event risk (See Supplementary Appendix 1(a)).

Visual inspection of the funnel plot and Egger regression asymmetry test (intercept (95% CI): 12.43 (–201.2 to 226.0), $p = 0.30$), classic fail-safe N analysis ($n = 636$) and Begg-Mazumdar correlation test (Kendall tau $b = 0.333$, $p = 0.30$) showed no evidence of publication bias in the analysis between MCP use and CV event risk (See Supplementary Appendix 1(b)).

Visual inspection of the funnel plot and Egger regression asymmetry test (intercept (95% CI): 0.77 (–125.93 to 127.47), $p = 0.48$), classic fail-safe N analysis ($n = 5$) and Begg-Mazumdar correlation test (Kendall tau $b = 0.333$, $p = 0.30$) showed no evidence of publication bias in the analysis of DMP vs MCP usage in terms of CV event risk (See Supplementary Appendix 1(c)).

QT prolongation risk

Characteristics of studies. Of the eight studies^{3,26–32} included in the meta-analysis, two studies^{26,27} investigated QT/QTc (TQT) studies, and the others^{3,28–32} were clinical studies of patients with a wide age scale ranging from newborns to adults (Table 1).

DMP-QT prolongation risk: event rates. A meta-analysis of 390 patients from eight studies revealed no significant increase in QT prolongation event rates with DMP use (Table 3, See Supplementary Appendix 2(a)).

Visual inspection of the funnel plot and Egger regression asymmetry test (intercept (95% CI): –2.343 (–4.321 to –0.364); $p = 0.027$), classic fail-safe N analysis ($n = 162$) and Begg-Mazumdar correlation test (Kendall tau $b = -0.107$; $p = 0.71$) showed no evidence of publication bias in the analysis between DMP use and CV event risk based on event rates (See Supplementary Appendix 2(b)).

DMP-QT prolongation risk: QT interval length. A meta-analysis of 198 patients from eight studies on DMP-QT prolongation risk revealed no significant difference between pretreatment and posttreatment QT interval in patients receiving DMP, based on the difference between means (1.35 ms, 95% CI, 8.29 ms to 10.99 ms, $p = 0.783$), standardized difference between means (0.23 ms, 95% CI, –0.28 ms to 0.74 ms, $p = 0.376$), standardized paired difference (0.22 ms, 95% CI, –0.22 ms to 0.66 ms, $p = 0.330$) and Hedge's g value (0.22, 95% CI, –0.27 to 0.71, $p = 0.375$) (Table 3, Figure 3).

Heterogeneity was noted for the outcome of the QT prolongation for difference between means ($p = 0.000$, $I^2 = 76.8\%$), standardized difference between means

Table 2. Meta-analysis of CV event risk related to DMP, MCP and DMP vs MCP treatment; DMP-CV event risk according to DMP dose.

	Cardiac outcome	CV event (n)						Adjusted risk					
		Present			Absent			OR	95% LCL	95% UCL	Z	p	
		Total	On DMP/MCP	Total	On DMP/MCP	OR	95% LCL						95% UCL
DMP vs no treatment													
Arana et al. (2015) ¹⁷	SCD	3239	28	12,572	52	1.71	0.92	3.18	1.696	0.090			
Varas-Lorenzo et al. (2016) ²⁵	VA	25,356	1644	25,356	1148	1.56	1.41	1.72	8.771	<0.001			
Chen and Hsiao (2015) ¹⁸	AMI	-	-	-	-	2.80	2.51	3.12	18.486	<0.001			
Coloma et al. (2013) ¹⁹	CA	140	7	560	15	4.70	1.39	15.89	2.490	0.013			
De Bruin et al. (2007) ²⁰	VA/SCD	1608	169	6428	481	1.59	1.28	1.98	4.167	<0.001			
Johannes et al. (2010) ²¹	SCD	1010	6	3030	12	1.60	0.59	4.35	0.921	0.357			
Jolly et al. (2009) ²²	SCD	775	9	6297	15	3.80	1.49	9.66	2.803	0.005			
Straus et al. (2005) ²³	SCD	1304	10	13,480	28	1.99	0.80	5.0	1.478	0.139			
van Noord et al. (2010) ²⁴	RE model	33,432	1873	67,723	1751	2.07	1.50	2.84	4.459	<0.001			
Overall						Q	I ²	Tau	Tau ²	p			
<i>Heterogeneity test</i>													
						68.883	89.8	0.353	0.124	<0.001			
MCP vs no treatment													
Arana et al. (2015) ¹⁷	SCD	3239	37	12,572	44	4.31	2.33	7.98	4.652	<0.001			
Varas-Lorenzo et al. (2016) ²⁵	VA	25,356	1316	25,356	932	1.47	1.40	1.55	14.838	<0.001			
Chen and Hsiao (2015) ¹⁸	AMI	-	-	-	-	5.70	5.09	6.39	30.047	<0.001			
Coloma et al. (2013) ¹⁹	RE model	28,595	1353	37,928	975	3.274	1.10	9.75	2.131	0.033			
Overall						Q	I ²	Tau	Tau ²	p			
<i>Heterogeneity test</i>													
						463.06	99.6	0.947	0.898	<0.001			
DMP vs MCP													
Arana et al. (2015) ¹³	SCD	-	-	-	-	0.40	0.21	0.75	-2.852	0.004			
Varas-Lorenzo et al. (2016) ²⁵	VA	25,356	1644	25,356	1316	1.06	0.91	1.23	0.758	0.448			
Chen and Hsiao (2015) ¹⁸	AMI	-	-	-	-	0.49	0.44	0.55	-12.531	<0.001			
Coloma et al. (2013) ¹⁹	FE model	25,356	1644	25,356	1316	0.64	0.58	0.70	-9.925	<0.001			
Overall	RE model	25,356	1644	25,356	1316	0.61	0.33	1.15	-1.517	0.129			
<i>Heterogeneity test</i>													
						67.23	97.0	0.529	0.280	<0.001			

(continued)

Table 2. Continued.

	Cardiac outcome	CV event				Adjusted risk									
		Present Total	On DMP	Absent Total	On DMP	OR	95% LCL	95% UCL	Z	p					
											On DMP	OR	95% LCL	95% UCL	Z
DMP daily dose															
<30 mg															
Arana et al. (2015) ¹⁷	SCD	806	4	4503	10	1.96	0.44	8.75	0.882	0.378					
Varas-Lorenzo et al. (2016) ²⁵	SCD	1304	2	13,480	10	1.24	0.19	8.11	0.225	0.822					
van Noord et al. (2010) ²⁴	2110	6	17,983	20	1.64	0.51	5.29	0.830	0.407						
Overall															
30 mg															
Arana et al. (2015) ¹⁷	SCD	817	15	4528	35	1.48	0.69	3.18	1.004	0.316					
Varas-Lorenzo et al. (2016) ²⁵	SCD	1304	4	13,480	15	1.02	0.23	4.47	0.026	0.979					
van Noord et al. (2010) ²⁴	2121	19	18,008	50	1.37	0.69	2.70	0.903	0.366						
Overall															
>30 mg															
Arana et al. (2015) ¹⁷	SCD	810	8	4498	5	3.20	0.59	17.35	1.349	0.177					
Varas-Lorenzo et al. (2016) ²⁵	VA	25,356	192	25,356	108	1.98	1.50	2.62	4.769	<0.001					
Chen and Hsiao (2015) ¹⁸	SCD	1304	4	13,480	3	11.40	1.99	65.25	2.734	0.006					
van Noord et al. (2010) ²⁴	RE model	27,470	204	43,334	116	3.14	1.19	8.30	2.312	0.021					
Overall															

AMI: acute myocardial infarction; CV: cardiovascular; DMP: domperidone; FE: fixed effects; LCL: lower confidence limit; MCP: metoclopramide; OR: odds ratio; RE: random effects; SCD: sudden cardiac death; UCL: upper confidence limit; VA: ventricular arrhythmia.

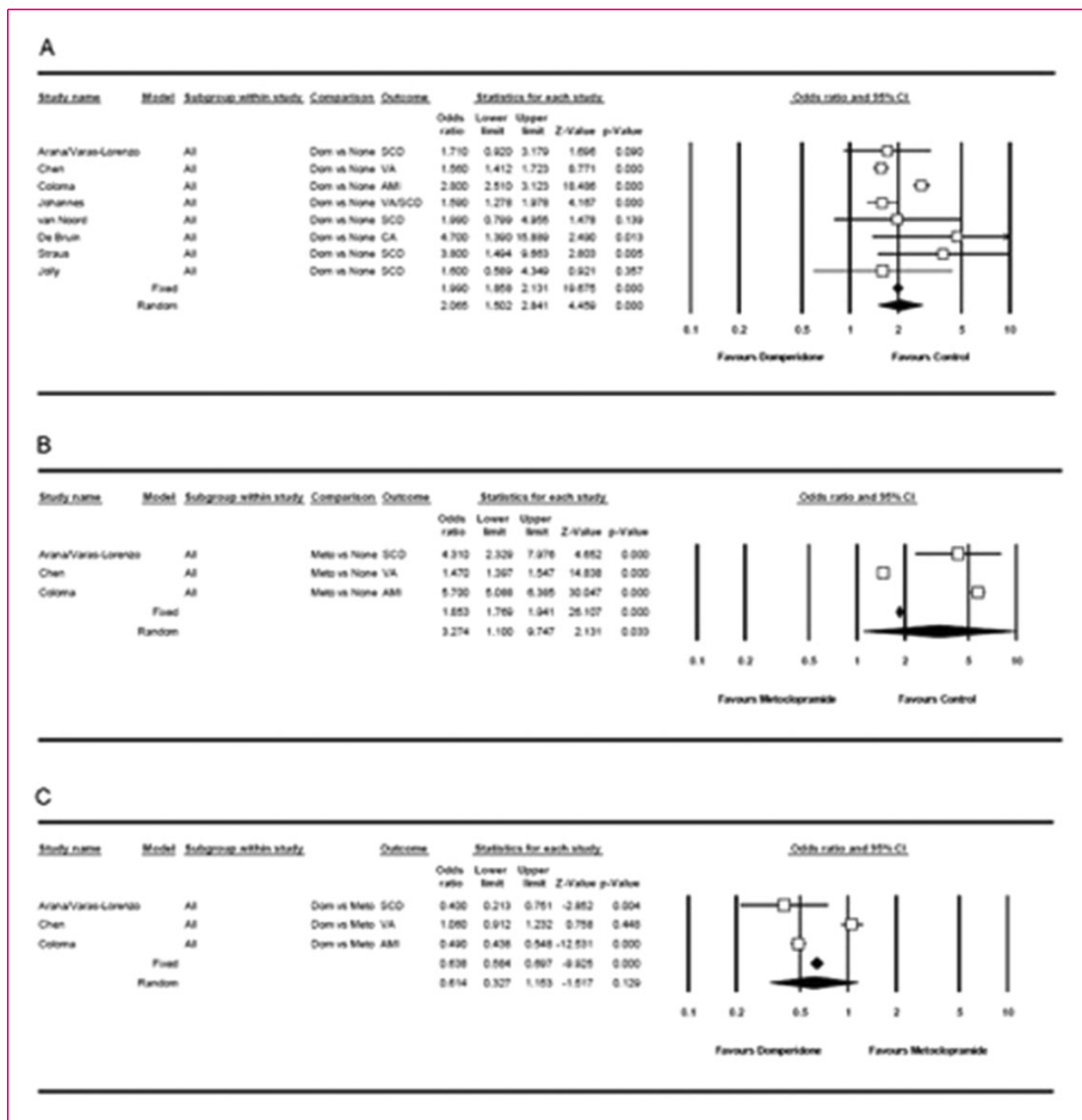


Figure 2. Forest plot displaying adjusted odds ratios (ORs) and 95% confidence intervals (CIs) for the impact of (a) domperidone treatment, (b) metoclopramide treatment, (c) and domperidone vs metoclopramide treatment on cardiovascular event risk.

($p=0.000$, $I^2=82.8\%$), standardized paired difference ($p=0.000$, $I^2=87.7\%$) and Hedge’s g value ($p=0.000$, $I^2=82.8\%$) (Table 3).

For the difference between means, standardized difference between means, standardized paired difference between means and Hedge’s g values, respectively: The visual inspection of the funnel plot and Egger regression asymmetry test (intercept (95% CI): -2.272 (-3.638 to -0.905), $p=0.0079$; 4.545 (-2.985 to

12.076), $p=0.18$; 5.358 (-4.945 to 15.660), $p=0.24$ and 4.959 (-2.678 to 12.597), $p=0.16$; respectively), classic fail-safe N analysis ($p>0.05$ for each), and Begg-Mazumdar correlation test (Kendall tau $b=-0.286$, $p=0.37$ for the difference between means; Kendall tau $b=0.190$; $p=0.55$ for others) showed no evidence of publication bias in the analysis between DMP use and the risk for OT prolongation (See Supplementary Appendix 3).

Table 3. Meta-analysis of domperidone-QT prolongation risk.

Event rate (%)									
Author (year) ^{ref}	Total n	Patients with QT prolongation	% ^a	% ^b	95% LCL	95% UCL	Z	p	
Biewenga et al. (2015) ²⁶	44	0	0.00%	1.11%	0.07%	15.43%	-3.156	0.002	
Boyce et al. (2012) ²⁷	24	0	0.00%	2.00%	0.12%	25.13%	-2.724	0.006	
Ngoenmak et al. (2016) ³	22	2	9.09%	9.09%	2.28%	29.96%	-3.105	0.002	
Du et al. (2014) ³²	192	0	0.00%	0.26%	0.02%	4.00%	-4.204	0.000	
Günlemez et al. (2010) ²⁸	40	2	5.00%	5.00%	1.25%	17.91%	-4.059	0.000	
Hegar et al. (2009) ²⁹	10	0	0.00%	4.55%	0.28%	44.83%	-2.103	0.035	
Djeddi et al. (2008) ³⁰	31	1	3.23%	3.23%	0.45%	19.64%	-3.346	0.001	
Sawant et al. (2004) ³¹	27	0	0.00%	1.79%	0.11%	22.96%	-2.808	0.005	
Overall	390	5	1.28%	3.54%	1.73%	7.10%	-8.829	0.000	
Heterogeneity test				Q	Q	Tau	Tau ²	p	
				6.717	6.717	0.000	0.000	0.459	
QT interval length									
Author (year)	n	Difference between means	SE	95% LCL	95% UCL	Z	p		
Biewenga et al. (2015) ²⁶	44	-7.20	6.31	-19.57	5.17	-1.141	0.254		
Boyce et al. (2012) ²⁷	24	5.90	8.54	-10.84	22.64	0.691	0.490		
Ngoenmak et al. (2016) ³	22	-15.50	10.29	-35.68	4.68	-1.506	0.132		
Günlemez et al. (2010) ²⁸	40	10.00	6.28	-2.31	22.31	1.592	0.111		
Hegar et al. (2009) ²⁹	10	-2.00	7.97	-17.62	13.62	-0.251	0.802		
Djeddi et al. (2008) ³⁰	31	14.00	1.18	11.69	16.31	11.883	0.000		
Sawant et al. (2004) ³¹	27	-10.00	12.94	-35.36	15.36	-0.773	0.440		
Overall	198	1.35	4.92	-8.29	10.99	0.275	0.783		
Heterogeneity test			Q	I ²	Tau	Tau ²	p		
			25.807	76.751	10.572	111.765	0.000		
Standardized difference between means									
Author (year) ^{ref}	n	Standardized difference between means	SE	95% LCL	95% UCL	Z	p		
Biewenga et al. (2015) ²⁶	44	-0.23	0.20	-0.62	0.17	-1.126	0.260		
Boyce et al. (2012) ²⁷	24	0.19	0.27	-0.35	0.72	0.685	0.494		
Ngoenmak et al. (2016) ³	22	-0.43	0.29	-1.00	0.15	-1.442	0.149		
Günlemez et al. (2010) ²⁸	40	0.33	0.22	-0.09	0.76	1.549	0.121		

(continued)

Table 3. Continued.

Event rate (%)		Patients with QT prolongation									
Author (year) ^{ref}	Total n	% ^a	% ^b	95% LCL	95% UCL	Z	p				
Hegar et al. (2009) ²⁹	10	-0.11	0.42	-0.93	0.72	-0.250	0.802				
Djeddi et al. (2008) ³⁰	31	2.83	0.53	1.78	3.87	5.317	0.000				
Sawant et al. (2004) ³¹	27	-0.20	0.26	-0.70	0.31	-0.766	0.444				
Overall	198	0.23	0.26	-0.28	0.74	0.886	0.376				
Heterogeneity test			Q	I ²	Tau	Tau ²	p				
			34.798	82.758	0.608	0.369	0.000				
Author (year) ^{ref}	n	Standardized paired difference	SE	95% LCL	95% UCL	Z	p				
Biewenga et al. (2015) ²⁶	44	-0.17	0.15	-0.47	0.13	-1.133	0.257				
Boyce et al. (2012) ²⁷	24	0.14	0.21	-0.26	0.54	0.687	0.492				
Ngoenmak et al. (2016) ³	22	-0.32	0.22	-0.75	0.11	-1.468	0.142				
Günlemez et al. (2010) ²⁸	40	0.25	0.16	-0.06	0.57	1.567	0.117				
Hegar et al. (2009) ²⁹	10	-0.08	0.32	-0.70	0.54	-0.251	0.802				
Djeddi et al. (2008) ³⁰	31	2.13	0.33	1.50	2.77	6.564	0.000				
Sawant et al. (2004) ³¹	27	-0.15	0.19	-0.53	0.23	-0.769	0.442				
Overall	198	0.22	0.22	-0.22	0.66	0.974	0.330				
Heterogeneity test			Q	I ²	Tau	Tau ²	p				
			48.913	87.733	0.549	0.301	0.000				
Author (year) ^{ref}	n	Hedge's g value	SE	95% LCL	95% UCL	Z	p				
Biewenga et al. (2015) ²⁶	44	-0.22	0.20	-0.61	0.17	-1.126	0.260				
Boyce et al. (2012) ²⁷	24	0.18	0.26	-0.34	0.70	0.685	0.494				
Ngoenmak et al. (2016) ³	22	-0.41	0.28	-0.97	0.15	-1.442	0.149				
Günlemez et al. (2010) ²⁸	40	0.33	0.21	-0.09	0.74	1.549	0.121				
Hegar et al. (2009) ²⁹	10	-0.10	0.38	-0.85	0.66	-0.250	0.802				
Djeddi et al. (2008) ³⁰	31	2.76	0.52	1.74	3.77	5.317	0.000				
Sawant et al. (2004) ³¹	27	-0.19	0.25	-0.68	0.30	-0.766	0.444				
Overall	198	0.22	0.25	-0.27	0.71	0.887	0.375				
Heterogeneity test			Q	I ²	Tau	Tau ²	p				
			34.811	82.764	0.589	0.347	0.000				

LCL: lower confidence limit; OR: odds ratio; SE: standard error; UCL: upper confidence limit. ^aCalculated based on data presented in the paper. ^bCalculated by means of a meta-analytical approach.

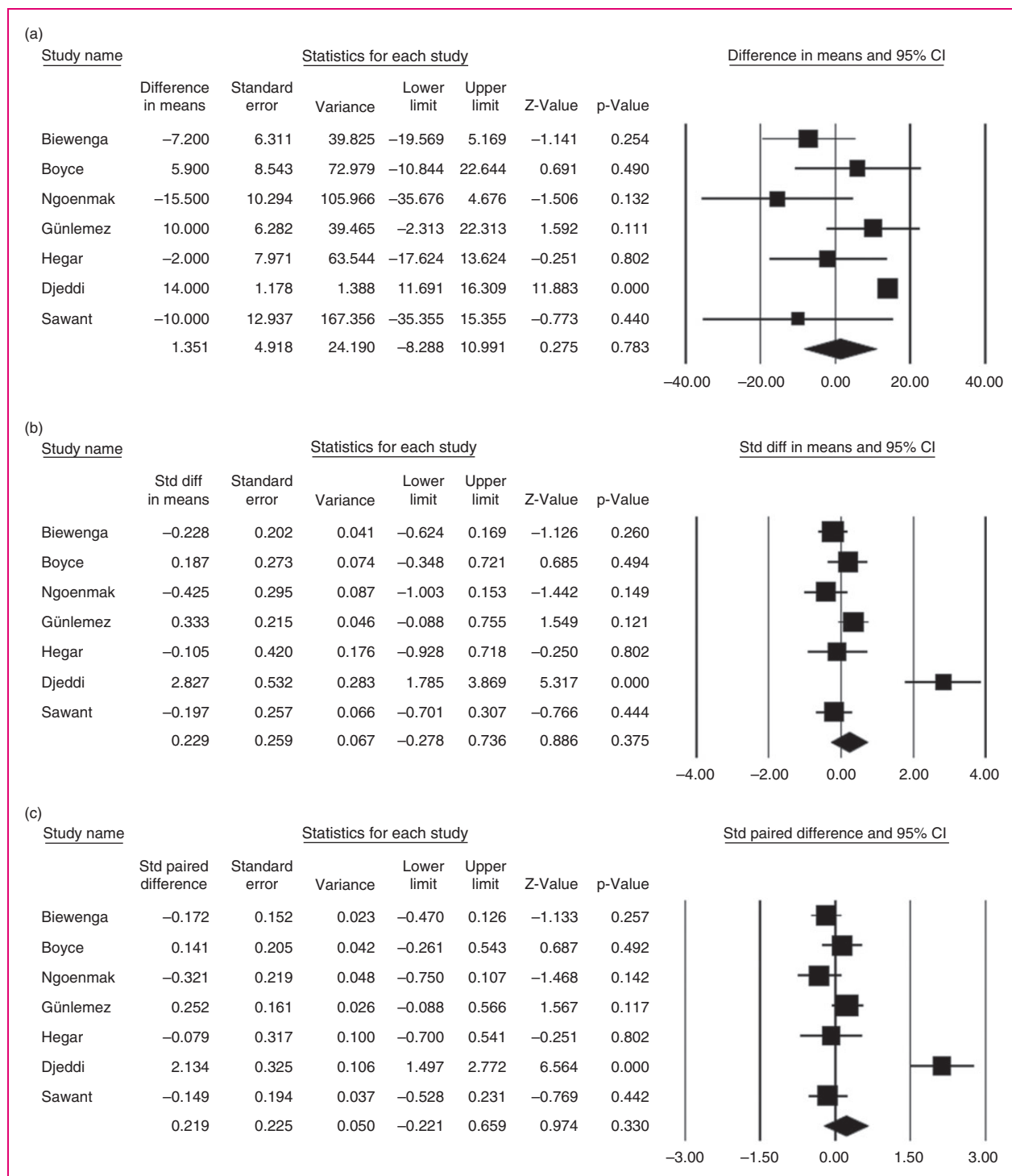


Figure 3. Forest plot displaying (a) difference, (b) standardized difference, (c) paired standardized difference between means of pre- and posttreatment QT interval, and (d) Hedges g value and 95% confidence intervals for the impact of current domperidone treatment on QT interval prolongation.

Discussion

The present meta-analysis of studies with DMP and MCP treatments^{17–19,25} suggested a significantly lower

CV event risk with DMP. When DMP treatment was compared with no treatment, a significant increase in CV event risk was suspected under DMP treatment based on adjusted risk in the overall analysis

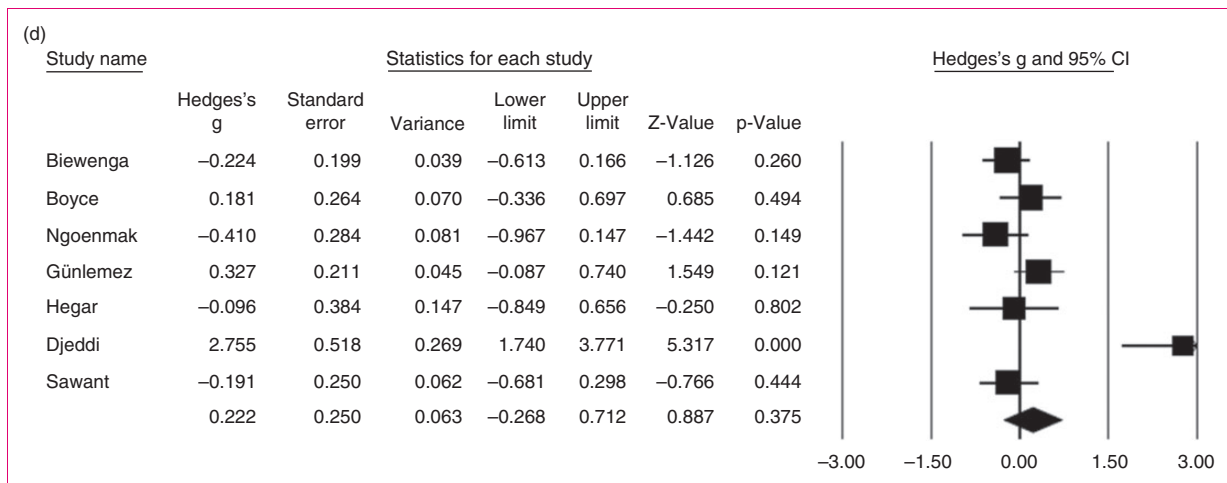


Figure 3. Continued.

population. However, subgroup analysis of studies with different DMP doses^{17,18,24,25} confirmed that this seemingly significant increase in CV event risk under DMP treatment occurred only with higher daily doses (>30 mg/day), with no CV event risk attributable to DMP use at regular (≤ 30 mg/day) daily doses.

Hence, findings from the present meta-analysis suggest that DMP may not be associated with an overall CV event risk increase given the potential CV safety of the drug when used at lower doses (≤ 30 mg/day), while DMP also seems to offer a much more favorable cardiac safety profile than MCP at any dose in accordance with the literature.³

CV risk attributable to DMP has been considered likely to differ with respect to subgroups of exposed individuals (higher in males and older individuals) as well as according to dosage (higher for doses >30 mg/day) of DMP.^{2,20,21,24,41}

These findings seem in accordance with regulatory restrictions introduced by the European Medicines Agency recommending DMP be used at a reduced dose in adults (up to 30 mg/day) and children (up to 0.75 mg/kg/day), for a restricted period (less than one week) and cautiously in high-risk groups such as the elderly (≥ 60 years).¹

Accordingly, our findings seem to indicate an association of DMP treatment with lower CV risk as compared with MCP treatment and a dose-dependent increase in CV event risk with a 2.1- to 3.1-fold increase in risk at doses >30 mg/day of MCP. This emphasizes a need for close monitoring for cardiac events and electrocardiogram (ECG) changes and adhering to appropriate cardiac risk monitoring protocols⁵ in patients receiving DMP at daily doses exceeding 30 mg and weighing the risk not only in the context of clinical decisions but also based on alternatives.²

Nonetheless, given that the present meta-analysis was based exclusively on retrospective studies and the lack of studies with a longitudinal, prospective design in the literature, our findings should be interpreted cautiously.

Notably, data from a retrospective chart review of patients with nausea and vomiting on high-dose DMP (80 to 120 mg) from 2009 to 2013 under an investigational new drug protocol revealed that despite very high dosing, DMP had a minimal risk of CV adverse events along with good clinical efficacy.⁵

Our meta-analysis suggested no significant QT prolongation risk in terms of pre- and posttreatment QT interval values in patients taking DMP.

Similarly, data from a large-scale retrospective chart review revealed prolonged QTc in seven (28.0%) out of 25 patients with a follow-up ECG in a cohort of 66 patients under DMP treatment independent of the daily dose and with no records on palpitations, chest pain, arrhythmias or changes in heart rate and electrolyte disturbances.⁵

Nonetheless, it should be noted that a significant heterogeneity was evident for the outcome of the difference between pretreatment and posttreatment means of QT interval in the studies included in our meta-analysis. In addition to data on pre- and posttreatment QT interval, measurements as well as event rates were available in only two studies,^{3,30} necessitating the calculation of these parameters using provided data in five studies,^{26–29,31} alongside the availability of only event rates with no data on QT measurements in one study.³⁰ Moreover, patients from included studies showed a wide age scale that ranged from newborns to adults, with evidence on QT prolongation >500 ms in only three studies among pediatric patients; in two out of 22 children aged <2 year,³ two out of 40

premature infants²⁸ and one out of 31 newborns,³⁰ respectively.

Hence, while the present meta-analysis revealed no significant QT prolongation with DMP overall, given the individual data from the studies in pediatric patients, our findings seem to emphasize that pediatricians should be aware of potential cardiac side effects of DMP, particularly in case of concurrent risk factors such as prescription of high doses and the presence of concomitant medications that are likely to increase the QT interval or inhibit the P450 enzyme.^{3,42}

Notably, only two studies had a “thorough QT” (TQT) design in our meta-analysis and both revealed no evidence of QTc prolongation (>500 msec) with any DMP regimen along with mean difference between pre-treatment and posttreatment QT intervals of -7.2 and 5.8 ms, respectively.^{26,27}

Available data on QT prolongation and cardiac event risk in patients under DMP treatment are largely limited to case reports and case-control studies along with inconsistent findings,^{2,4,5,7,10,13-32} and CV risk was evaluated in only one meta-analysis of six studies to date, which revealed current DMP use to increase the risk of VA and SCD by 70%.⁴ Hence, our findings emphasize a need for large clinical trials with a longitudinal, prospective design to better address the CV safety profile of DMP overall and in selected patient groups.^{2,5} Nonetheless, our findings emphasize that any risks need to be discussed in detail with patients so they are fully informed.

As a limitation, this meta-analysis was based exclusively on retrospective studies, emphasizing a need for a prospective validation to confirm data provided on DMP-related CV risk. Significant heterogeneity was noted for the outcome of the QT prolongation across included studies, which might have resulted from differences in study designs, patient characteristics and sample sizes and likely influenced the accuracy of our analysis and reduced the statistical power regarding DMP-QT prolongation risk assessment. Consistent with inclusion of studies from large clinical databases, however, classical fail-safe *N* publication bias tests of moderators indicated large numbers of studies were needed to render the effect size nonsignificant for CV event risk and QT prolongation event rate, indicating that the true effect size of these moderators seems significant enough to allow for certainty in the present meta-analysis.

Conclusion

In conclusion, while overall results of the meta-analysis of nine studies suggest an increased CV event risk in patients under DMP treatment, DMP seems to be associated with increased CV event risk only at

doses > 30 mg/day and to offer a more favorable cardiac safety profile than MCP regardless of dosage. A meta-analysis of eight studies suggested no significant QT prolongation risk of event rates and pre- and posttreatment difference in mean QT interval values in patients under DMP treatment. Accordingly, the present meta-analysis seems to indicate that DMP treatment may not be associated with an overall CV event risk increase at doses \leq 30 mg/day and does not result in QT prolongations. Our findings emphasize the need for close monitoring for cardiac events and ECG changes among those on high-dose DMP along with weighing the risk both in the context of clinical decisions and of alternatives. There is a need for well-designed, longitudinal, prospective studies to better address the CV safety profile of DMP overall and in selected patient groups.

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Declaration of conflicting interests

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Ethics approval

This type of study is exempt from ethics approval.

Informed consent

Not applicable.

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