23-Year Trends in Life Expectancy in Good and Poor Physical and Cognitive Health at Age 65 Years in the Netherlands, 1993–2016

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Objectives. To examine 23-year trends in both physically and cognitively healthy life expectancy from age 65 years in the Netherlands.

Methods. We used 8 waves between 1993 and 2016 from the nationally representative Longitudinal Aging Study Amsterdam (12 948 observations). We calculated physically and cognitively healthy life expectancies by using the Sullivan life table method and tested prevalence trends over time by using generalized estimating equations.

Results. Total life expectancy at age 65 years rose from 14.7 to 18.7 years (men) and from 19.2 to 21.4 years (women). Life expectancy in poor physical health increased nonlinearly from 1.8 to 2.9 years for men; for women it fluctuated around 5.7 years. Meanwhile, life expectancy in good cognitive health increased linearly from 11.0 to 15.7 years (men) and from 13.4 to 18.0 years (women). The proportion of people with poor physical and poor cognitive health combined did not increase, averaging 5.9% (men) and 8.7% (women).

Conclusions. This multiwave study shows that a negative trend in physically healthy life expectancy is accompanied by a positive trend in cognitively healthy life expectancy. (*Am J Public Health.* 2018;108:1652–1658. doi:10.2105/AJPH.2018.304685)

See also Crimmins, p. 1582.

n view of population aging and the need to contain health care costs, many Western countries are reforming their long-term care system.¹ One such reform is to cut back on long-term care services. This policy assumes that in the near future, fewer older people will need these services, because they will stay healthy to older ages. Accumulated evidence on trends in healthy and unhealthy life years from age 65 years onward, however, suggests that the rise in life expectancy is not accompanied by a similar rise in healthy life years.^{2,3} In particular, years with mild physical disability have been shown to increase while years with severe physical disability remain more or less constant.⁴ If these trends continue, increasing numbers of older people will have to pay for care themselves, or go without care.

The rise in the number of very old people and the relatively high costs of dementia care call for monitoring trends in cognitive health.^{5,6} Trend studies on cognitive impairment or dementia generally show decreases in prevalence and incidence.^{7–9} Moreover, increases have been reported in years with good cognitive health that come close to or even exceed the increase in life expectancy at age 65 years.^{4,10} These studies suggest a delay of years with poor cognitive health, but evidence for this is not unequivocal.¹¹ So far, it seems that the prospects for life expectancy in good cognitive health are more favorable than for life expectancy in good physical health.

Evidence on trends in cognitively healthy life expectancy is recent and limited to a few

countries. Trend studies on physically healthy life expectancy have accumulated over several decades and have shown differences in trends across countries and across time periods within 1 country.¹² This body of research has taught us to be cautious about generalizing across periods and geographic regions.¹³ So far, no evidence exists on trends in cognitively healthy life expectancy for the Netherlands over more than a decade. The current study contributes evidence from a nationwide, representative study in the Netherlands, across a period of 23 years. Life expectancy from age 65 years is distinguished into life expectancy in good, fair, and poor physical health; life expectancy in good and fair-topoor cognitive health; and their combination. A joint study of physically and cognitively healthy life years is rare, but provides the necessary evidence to properly assess future health and long-term care costs.

METHODS

We used data from the nationally representative Longitudinal Aging Study Amsterdam (LASA). Participants were drawn from the population registries of 11 municipalities in 3 socio-culturally distinct geographic areas in the Netherlands (west, northeast, and south) and are followed at 3-year intervals with exactly the same face-to-face interviews and tests. The first cohort, recruited in 1992 to

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1993, included 3107 participants aged 55 to 85 years. After 10 and 20 years, a second and a third cohort aged 55 to 64 years were added, including 1002 and 1023 participants, respectively. The samples were weighted so that together they represented the distribution of population density of the Netherlands. Participants who refuse a face-to-face interview are offered a 15-minute telephone interview, in which either the participant or a proxy can participate. Details on the sampling and data collection have been published elsewhere.¹⁴

For the current study, we selected subsamples aged 65 years or older from participants at each wave between 1993 and 2016. (Although each wave takes place in 2 successive calendar years, in the following, for brevity, the waves are denoted by the second year, as the majority of the participants were interviewed in that year.) This study design defined a partly different cohort at each wave, because after each time interval of 3 years, new individuals aged 65 to 67 years are included. Up to 2002, these newly included participants were members of the original cohort. From 2006 to 2012, they were members of the cohort added in 2002, and in 2016, they were members of the cohort added in 2012. The number of participants at each wave varied from 1415 to 2003 for physical health and from 1240 to 2129 for cognitive health.

Measures

Physical health. The definition of physical health took into account both multimorbidity and disability. First, multimorbidity, defined as the presence of at least 2 chronic diseases, is a complex condition that may have serious consequences for functioning and quality of life.¹⁵ Second, disability is associated with health and long-term care use and quality of life.¹⁶ Thus, multimorbidity and disability combined may be considered to reflect the physical health states that are relevant to older people's use of care.

We measured multimorbidity by selfreports of major chronic diseases: chronic obstructive lung diseases (including emphysema, bronchitis, and asthma), coronary heart diseases, peripheral arterial disease, stroke, diabetes, cancer, and arthritis (including osteoarthritis and rheumatoid arthritis). In addition, participants could indicate a maximum of 2 other chronic diseases. The comparability of these self-reports with general practitioner records was satisfactory and remained stable across the period 1992 to 2009.¹⁷

We measured disability by self-reports of difficulty or need of help with 6 activities: climbing stairs, dressing, getting up from a chair, cutting one's toenails, walking 400 meters, and using own or public transportation. We defined mild disability as difficulty with at least 1 activity and severe disability as needing help with at least 1 activity.¹⁸

Chronic diseases and disability were included in both the face-to-face and the telephone interview. We derived 3 physical health states:

- Good physical health: no multimorbidity and no mild disability;
- Fair physical health: (1) multimorbidity with at most mild disability, (2) mild disability with or without multimorbidity, or (3) severe disability without multimorbidity; and
- Poor physical health: both multimorbidity and severe disability.

Cognitive health. In keeping with most previous studies, we measured cognitive functioning with the Mini-Mental State Examination (MMSE), which includes 23 items with a maximum score of 30 when all responses are correct.¹⁹ This test was included only in the face-to-face interview. We distinguished 2 cognitive health states, because the available numbers did not allow examination of a third state. The cut-off between good and fair or poor followed Jagger et al.⁴:

- 1. Good cognitive health: MMSE score of 26 or higher, and
- 2. Fair or poor cognitive health: MMSE score of 25 or lower.

Covariates. We obtained age and gender from the municipal registries. We defined level of education as number of years of schooling, ranging from 5 to 18 years.

Data Analysis

We calculated healthy and unhealthy life expectancies by using the Sullivan life table method for 8 calendar years corresponding to the LASA waves for 1993, 1996, 1999, 2002, 2006, 2009, 2012, and 2016.²⁰ We used generalized estimating equations (GEEs) to statistically test trends in the prevalence of good and poor health. Each step is explained next.

We obtained life tables for the national population for each calendar year stated from Statistics Netherlands. We derived annual survival probabilities for men and women from age 65 years onward. At age 65 years, the survival probability was set to 1. At age 66 years, the survival probability was 1 minus the proportion of deaths at age 65 years; at age 67 years, the survival probability was the survival probability at age 66 years minus the proportion of deaths at age 66 years, and so on. If a person died, this person was assumed to have lived 0.5 year. The resulting probabilities added up to the total life expectancy from age 65 years.

To obtain age-specific prevalences of good and poor health, we used LASA data from each wave. In the higher age groups the number of cases was relatively small; thus, we refrained from using age-specific prevalence rates directly derived from the data because these would be too unstable. Instead, for each calendar year and gender, we fitted a polynomial of the association of health with age to derive age-specific prevalence rates.²¹

Following the Sullivan method, for each age year from 65 years onward, we multiplied the national life table–derived survival probability by the LASA-derived prevalence of good or poor health in this age year. We summed these age-specific probabilities to the life expectancy in good or poor health. As during the study period total life expectancy increased, we calculated not only the number of years but also the proportion of remaining life with and without health problems.

A systematic statistical test for trends in healthy life expectancy across multiple waves is not available. Therefore, we tested trends in the prevalence of good and poor health in the LASA data by using GEEs after pooling the data across the 8 waves, with year as the main determinant and adjustment for age. Because there is an overlap in participants across the waves, we accounted for the interdependence of the data by including a 7-dependent correlation matrix.²² For this trend analysis, we selected same-age participants (ages 65-85 years) at each wave. To test for nonlinear trends, we added a quadratic term and reported it if significant. As the MMSE is known for its sensitivity to education, which might lead to misclassification, and as the level of education has increased in subsequent

TABLE 1—Descriptive Characteristics, Total Life Expectancy, and Proportion of Life Expectancy in Good and Poor Physical Health and in Good Cognitive Health at Age 65 Years, by Gender: Longitudinal Aging Study Amsterdam, The Netherlands, 1993–2016

	1993	1996	1999	2002	2006	2009	2012	2016	Difference 2016–199 No. (%) or Percentag Points
Men									
No. for physical health	972	881	750	665	644	613	603	650	
No. for cognitive health	1031	809	697	611	572	564	540	590	
Age, y, mean (range)	75.7 (65.0–85.6)	76.2 (65.1–88.7)	76.3 (65.0–91.8)	76.0 (65.0–94.0)	75.5 (65.0–97.9)	75.3 (65.0–100.6)	75.4 (65.0–104.2)	74.7 (65.0–102.8)	
Years of education, mean	9.2	9.5	9.8	10.0	10.3	10.8	10.9	11.4	
Total life expectancy, y	14.7	15.1	15.5	16.0	17.1	17.8	18.2	18.7	+4.0 (27.2)
Proportion life expectancy in good physical health	46.6	37.6	33.6	32.4	32.7	30.5	22.8	25.0	-21.6
Proportion life expectancy in poor physical health	12.3	13.8	15.2	14.6	19.1	17.9	17.0	15.7	+3.4
Proportion life expectancy in good cognitive health	74.8	78.0	81.5	83.8	81.1	82.8	85.5	83.7	+8.9
Women									
No. for physical health	1031	1010	950	900	863	809	812	795	
No. for cognitive health	1098	920	884	797	739	721	700	695	
Age, y, mean (range)	75.3 (65.0–85.6)	76.1 (65.0–88.8)	76.6 (65.0–91.7)	77.1 (65.0–94.3)	76.7 (65.0–98.6)	76.9 (65.1–100.1)	76.7 (65.0–102.3)	76.0 (65.0–102.0)	
Years of education, mean	7.8	8.0	8.1	8.3	8.7	9.0	9.3	10.0	
Total life expectancy, y	19.2	19.4	19.5	19.7	20.5	21.1	21.2	21.4	+2.2 (11.5)
Proportion life expectancy in good physical health	28.7	26.0	24.5	22.7	21.2	20.2	14.9	18.9	-9.8
Proportion life expectancy in poor physical health	28.3	27.8	25.2	26.0	28.8	28.7	32.9	28.9	+0.6
Proportion life expectancy in good cognitive health	69.9	73.7	76.1	78.2	78.9	82.4	82.3	84.4	+14.5

cohorts,²³ we examined all trends in additional GEE models adjusted for educational level.

We performed a series of sensitivity analyses to address (1) a potential practice effect in the MMSE in continuing participants, (2) the use of proxy information for physical health, (3) nonmortality attrition and item nonresponse for cognitive health, and (4) other cutpoints for fair physical health problems and fair or poor cognitive health (Appendix, available as a supplement to the online version of this article at http://www.ajph.org).

RESULTS

During the period 1993 to 2016, total life expectancy at age 65 years rose from 14.7 to 18.7 years (27.2%) for men and from 19.2 to 21.4 years (11.5%) for women (Table 1).

Physical Health

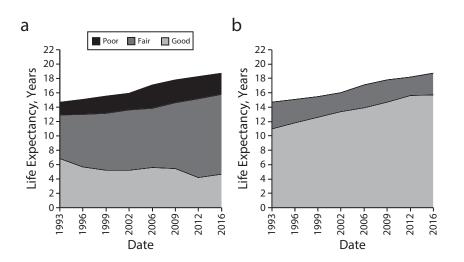
Life expectancy in good physical health decreased from 6.9 to 4.7 years for men

(Figure 1). A steady decrease occurred in the 1990s, after which the pattern was more irregular, with a low of 4.2 years in 2012. For women, physically healthy life expectancy decreased as well, from 5.5 to 4.0 years (Figure 2). This decrease continued throughout the study period, with a low of 3.2 years in 2012. Life expectancy in poor physical health increased for men from 1.8 to 2.9 years, with a steady increase occurring up to 2006, after which an improvement set in. For women, life expectancy in poor physical health showed an increase from 5.4 to 6.2 years. This increase occurred only in the late 2000s, with a high of 7.0 years in 2012.

The proportional decrease in life expectancy in good physical health continued throughout the study period, with, for both men and women, a small rise after 2012 (Table 1). By contrast, we saw a proportional increase in life expectancy in poor health for men up to 2009 with a decrease after that, whereas for women this proportion fluctuated, with a low in 1999 and a high in 2012.

Taking all 8 LASA waves into account, we tested trends in the prevalence of good and poor health (Table 2). The prevalence of good physical health showed a decrease for men that flattened off, as demonstrated by significant odds ratios (ORs) for both the linear and the quadratic terms for year (OR_{vear} = 0.912; 95% confidence interval [CI] = 0.887, 0.938, and OR_{vear²} = 1.002; 95% CI = 1.001, 1.003). For women, we found a linear decrease ($OR_{vear} = 0.971$; 95% CI = 0.962, 0.980). The prevalence of poor physical health for men showed an initial increase that also flattened off (OR_{vear} = 1.069; 95% CI = 1.035, 1.104, and ORyear² = 0.998; 95% CI = 0.996, 0.999). For women, it showed stability (OR_{vear} = 1.008; 95% CI = 0.999, 1.017).

Adding years of education to the GEE models did not change the trend for men. For women, the negative trend in good health became more negative and the nonsignificant positive trend in poor health became statistically significant (OR_{year} = 1.016; 95% CI = 1.007, 1.025).



Note. The years in good, fair, and poor health are stacked vertically so that for each calendar year they sum to the total life expectancy. Good physical health = no multimorbidity and no mild disability; fair physical health = multimorbidity, disability, or both; poor physical health = multimorbidity and severe disability; good cognitive health = Mini-Mental State Examination (MMSE) \geq 26; fair or poor cognitive health = MMSE \leq 25.

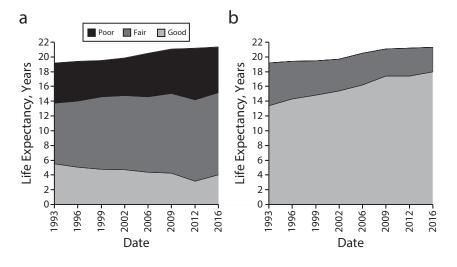
FIGURE 1—Life Expectancy for Men Aged 65 Years in (a) Good, Fair, and Poor Physical Health and in (b) Good and Fair or Poor Cognitive Health: Longitudinal Aging Study Amsterdam, The Netherlands, 1993–2016

Cognitive Health

Life expectancy from age 65 years in good cognitive health showed increases from 11.0 to 15.7 years for men and from 13.4 to 18.0 years for women; both increases continued at

more or less the same pace throughout the study period (Figures 1 and 2).

A proportional increase was concentrated for men in the 1990s, a period with relatively little change in total life expectancy from age



Note. The years in good, fair, and poor health are stacked vertically so that for each calendar year they sum to the total life expectancy. Good physical health = no multimorbidity and no mild disability; fair physical health = multimorbidity, disability, or both; poor physical health = multimorbidity and severe disability; good cognitive health = Mini-Mental State Examination (MMSE) ≥ 26 ; fair or poor cognitive health = MMSE ≤ 25 .

FIGURE 2—Life Expectancy for Women Aged 65 Years in (a) Good, Fair, and Poor Physical Health and in (b) Good and Fair or Poor Cognitive Health: Longitudinal Aging Study Amsterdam, The Netherlands, 1993–2016 65 years (Table 1). Women showed a greater and more steady increase than men, which can be attributed partly to the smaller rise in their life expectancy, and partly to their greater absolute increase in cognitively healthy life expectancy. These increases exceed those in total life expectancy.

The GEE models showed linear decreases in the prevalence of fair or poor cognitive health for both men ($OR_{year} = 0.973$; 95% CI = 0.962, 0.983) and women ($OR_{year} = 0.963$; 95% CI = 0.953, 0.973) over time. Adding years of education to these models halved the decreases both for men ($OR_{year} = 0.986$; 95% CI = 0.975, 0.997) and for women ($OR_{year} = 0.980$; 95% CI = 0.969, 0.991), but the decreases remained significant.

Physical and Cognitive Health Combined

We assessed trends in 2 extreme health categories: good in both physical and cognitive health (good overall health), and poor in physical and fair or poor in cognitive health (poor overall health). The number of participants with poor overall health per wave was too small to calculate unhealthy life years.

Averaged across waves, 5.9% of men and 8.7% of women were in poor overall health. The great majority (86.4%) of those in good physical health also had good cognitive health, whereas 34.4% of those in poor physical health also had fair or poor cognitive health.

The decline in the trend in good overall health was only slightly less negative compared with that in good physical health (OR_{vear} = 0.922; 95% CI = 0.895, 0.950, and $OR_{vear^2} = 1.002$; 95% CI = 1.001, 1.003, for men and OR_{vear} = 0.978; 95% CI = 0.968, 0.988, for women). Regarding poor overall health, the trend for men was linear and stable, but for women it showed a significant linear decrease (OR_{vear} = 0.989; 95% CI = 0.971, 1.007, and OR = 0.971; 95% CI = 0.957, 0.986, respectively). Thus, for this group in poor overall health, the decrease in fair or poor cognitive health compensated for the increase (men) or stability (women) in poor physical health.

TABLE 2—Tests of Trends Over Time in Good and Poor Physical Health, and in Fair or Poor Cognitive Health: Longitudinal Aging Study Amsterdam, The Netherlands, 1993–2016

	OR (95% CI)
Physical healt	h, men
Good physical health ^a	
Linear term	0.912 (0.887, 0.938)
Quadratic term	1.002 (1.001, 1.003)
Poor physical health ^b	
Linear term	1.069 (1.035, 1.104)
Quadratic term	0.998 (0.996, 0.999)
Physical health,	women
Good physical health ^a	0.971 (0.962, 0.980)
Poor physical health ^b	1.008 (0.999, 1.017)
Cognitive healt	:h, men
Fair or poor cognitive health ^c	0.973 (0.962, 0.983)
Cognitive health	, women
Fair or poor cognitive health ^c	0.963 (0.953, 0.973)
Physical and cogni	tive health
combined,	men
Good overall health ^d	
Linear term	0.922 (0.895, 0.950)
Quadratic term	1.002 (1.001, 1.003)
Poor overall health ^e	0.989 (0.971, 1.007)
Physical and cogni	tive health
combined. w	omen

combined, women					
Good overall health ^d	0.978 (0.968, 0.988)				
Poor overall health ^e	0.971 (0.957, 0.986)				

Note. CI = confidence interval; OR = odds ratio. Using generalized estimating equations adjusted for age and gender, with year as the independent variable. If the quadratic term of year was significant (P<.05), the OR (95% CI) for this term is reported along with the OR (95% CI) for the linear term; if the quadratic term was not significant, by default only the OR (95% CI) for the linear term is reported.

^aAt most 1 chronic disease and no difficulty with any activity.

^bMultimorbidity and needing help with at least 1 activity.

^cMini-Mental State Examination (MMSE) \leq 25. ^dAt most 1 chronic disease, no difficulty with any activity, and MMSE \geq 26.

^eMultimorbidity, needing help with at least 1 activity, and MMSE≤25.

DISCUSSION

In this study, we examined national trends in physically and cognitively healthy life expectancy over the relatively long period of 23 years. This study provides, to our knowledge, the first national estimates of long-term trends in cognitively healthy life expectancy for the Netherlands. The findings show a substantial decline in physically healthy life expectancy and a smaller increase in life expectancy in poor physical health. There were fluctuations across the study period, with, for men, the worsening trend leveling off after the mid-2000s. For women, an improvement was apparent only after 2012, but as this concerns only 1 time point, conclusions about leveling off cannot be drawn. In contrast, life expectancy in good cognitive health showed increases that continued throughout the study period.

Strengths and Limitations

Our study had several strengths. First, it was based on a nationally representative data set, the LASA. Second, exactly the same measurement instruments were used at each LASA wave. Third, we performed our estimates of healthy and unhealthy life expectancy for 8 points in time, covering more than 2 decades. With only 2 time points, which most earlier studies have used,⁴ it cannot be derived whether trends accelerate or decelerate in specific periods; also, measurement error may over- or underestimate trends. Finally, the addition to LASA of new cohorts in 2002 and 2012 allows the examination of trends over a long period in the relevant age group of 65 years and older. Had new cohorts not been added, the minimum age in 2016 would have been 78, allowing the study of trends only in participants aged 78 years and older.

This study also had limitations. First, we derived our criteria to define less-than-good physical and cognitive health from previous studies,^{4,17} but in relation to health care use, stricter criteria may be needed. People with difficulty in only 1 activity may not be at substantially higher risk of needing care than their peers with no difficulty in any activity; also, people with scores of 24 to 25 on the MMSE may have functioned at this level during their lifetime, which means that they do not have age-related cognitive impairment.²⁴ Our sensitivity analyses, however, revealed virtually the same trends over time for narrower definitions of fair physical health

and fair or poor cognitive health (Table A, available as a supplement to the online version of this article at http://www.ajph.org).

Second, during the 23 years covered by our study, the reporting behavior of successive cohorts might have changed because of higher expectations of medical care, less tolerance for mild health problems, greater awareness of chronic disease in general, and more frequent attribution of similar complaints to a disease instead of to old age.^{17,25} These developments may result in more recent cohorts reporting fair rather than good health more easily.

Third, the use of proxy information on physical health for respondents who were unable to self-report might have affected the trend over time if the association between physical health and use of proxies had changed over time. However, a sensitivity analysis showed that this association had not changed over time (see "Proxy information," in the Appendix, available as a supplement to the online version of this article at http:// www.ajph.org).

Fourth, although the MMSE is designed as a screening instrument for cognitive impairment and dementia, additional information is needed to distinguish age-related cognitive impairment from dementia. Cognitive impairment per se may not be associated with health care use to the same extent as dementia.^{26,27}

Fifth, the use of longitudinal data entails that a substantial number of participants were tested repeatedly, raising the possibility of a practice effect in the MMSE. However, participants in more recent cohorts who were tested the same number of times showed consistent improvements in cognitive health compared with earlier cohorts (see "Practice effect," in the Appendix). A further issue concerning longitudinal data is nonmortality attrition, which proved to be associated with cognitive impairment at the previous wave. Additional analyses showed that this association was of similar size across waves and that nonmortality attrition did not affect the trend in cognitive health (see "Attrition," in the Appendix). The proportion of participants with nonresponse on the MMSE did increase among participants in fair or poor cognitive health at the previous wave, and did not increase among participants in good cognitive health at the previous wave.

However, the trend in cognitive health was not affected when this proportion was included in the model. These various sensitivity analyses suggest that practice effect and nonresponse attrition had no substantial effects on the positive trends observed in cognitive health.

Relation to Other Studies

Our findings fit in with those from other western countries, in particular regarding the increase in cognitively healthy life expectancy.4,9 Regarding physically healthy life expectancy, the findings are a little more pessimistic, in that we note a small increase in life years in poor health, whereas other studies show stability or a small decrease. However, in most studies, a single measure of physical disability is used.^{3,4} Indeed, when we examined the trends in mild and severe disability alone, we found that the increase in mild disability came to a halt in the 2000s and that severe disability showed a small decline across the study period (data available upon request). Therefore, any increase in fair physical health after 2002 is likely to be driven by an increase in multimorbidity. Such increase has been found in other countries.²⁸ Meanwhile, we maintain that our health measure, which combines multimorbidity and disability, better reflects the issues that older people are confronted with than disability alone. Our reasons for this are that the prevalence of multimorbidity is expected to increase because of further advances in medical treatment and care,²⁸ that multimorbidity represents a complex condition with serious consequences for functioning and wellbeing,¹⁵ and that multimorbidity did not become less disabling during the period covered by our study.²⁹

The positive trend observed in cognitive health could be attributed more than 50% to the rise in the level of education, but remained significant after we accounted for education. Various authors have attributed the improvement in cognitive functioning not only to increased access to education but also more meaningful and active learning,³⁰ and to an increase in occupations with high complexity.³¹ More distal explanatory factors include better nutrition, smaller and more affluent families in which parents pay more attention to their children, and greater information density in society.³² Particularly relevant to later life is the recent improvement in the control of cardiovascular risk factors such as hypertension and cholesterol, which may help slow cognitive decline.^{10,33}

Thus, various developments are likely drivers of the observed improvement in cognitive health. In contrast, the negative physical health trend for women would have been more pronounced had women's level of education not increased over time. Further rises in the level of education of the older population may be expected, so that increases in life expectancy in poor physical health may continue to be counteracted by increases in the level of education.³⁴

Public Health Implications

In light of current policy to contain the costs of long-term care, this study's findings are pertinent to the question whether trends in physically and cognitively healthy life expectancy warrant reductions in the supply of elder care. Because of the increase in lessthan-good physical health, future health care needs may change toward more social and rehabilitative care as well as toward more medical care geared to multimorbidity.²⁵ The small increase in life expectancy in poor physical health was accompanied by a clear decrease in life expectancy in fair or poor cognitive health. The trend in poor physical and fair or poor cognitive health combined showed no increase for men and even a decrease for women. Although the proportion of people with poor overall health is small, it accounts for a relatively large share of health care costs.^{25,35} Thus, one might argue that an increase in demand for long-term care might be limited, were it not for the expected absolute increase in the number of oldest-old people. Interventions to improve both physical and cognitive health in the older population remain important. AJPH

CONTRIBUTORS

D. J. H. Deeg designed the study, performed the statistical analyses, and wrote drafts of the article. H. C. Comijs, E. O. Hoogendijk, M. van der Noordt, and M. Huisman contributed to the design, analysis, and interpretation, and revised successive versions of the article critically. All authors approved the final version before submission.

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HUMAN PARTICIPANT PROTECTION

The VU University Medical Centre's Medical Ethics Evaluation Committee approved the study (archive numbers 92/138 and 2002/141). Informed consent was obtained from all individual participants included in the study.

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