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## Racial Discrimination, Body Mass Index, and Insulin Resistance: A Longitudinal Analysis

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### Abstract

**Objective:** To examine prospective relations of perceived racial discrimination at ages 16–18 with body mass index (BMI) at ages 19–21 and insulin resistance (IR) at ages 25 and 27 among Black youth in the rural South, and to determine whether BMI connected discrimination to IR as a mediator.

**Method:** Participants were 315 African American adolescents in rural counties in Georgia who provided data on their perceptions of discrimination during adolescence. BMI was measured during a yearly home visit, and a certified phlebotomist drew a fasting blood sample from which IR was measured.

**Results:** The data analysis, with all confounding variables controlled, revealed that, over time, (a) discrimination was associated positively with both BMI and IR; (b) BMI was associated positively with IR; and (c) BMI acted as a mediator connecting discrimination with IR.

**Conclusions:** The findings are consistent with the hypothesis that exposure to discrimination presages IR through its effects on BMI.

### Keywords

African Americans; body mass index; insulin resistance; racism; rural population

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Compared with members of other racial groups, African Americans experience more aging-related diseases earlier in life, at greater severity, and with more serious consequences (Schuster et al., 2012). According to life-course and developmental perspectives, this

disproportionate disease risk can be traced to systematic disadvantages and social inequities starting at conception and continuing throughout childhood, adolescence, and young adulthood (Krieger, 2014). The health-risk inequities that African Americans experience likely arise from more than class disadvantage alone. Psychological stressors that disproportionately affect African Americans have been suggested as a mechanism that weathers physiological systems and increases vulnerability to poor health (Geronimus, Hicken, Keene, & Bound, 2006). Racial discrimination presents daily challenges in the lives of African American youth and their families; recent reports indicate that racial discrimination remains a persistent, if not escalating, experience for African Americans (Williams & Mohammed, 2013). Discriminatory incidents are demeaning and induce stress, frustration, depression, and anxiety in adults (Pascoe & Smart Richman, 2009) and in adolescents (Priest et al., 2013).

An association between racial discrimination and outcomes relevant to physical health is also plausible. Research conducted primarily with adults has demonstrated that discrimination is associated with a range of health-relevant physiological processes, including dysregulated patterns of cortisol release and biomarkers of low-grade inflammation (Lewis, Cogburn, & Williams, 2015; Paradies et al., 2015). Mirroring these patterns, recent research with young African American populations indicated that experiences with racial discrimination were associated with flatter diurnal cortisol slopes, indicative of unhealthier outcomes, for young adults (Zeiders, Hoyt, & Adam, 2014), heightened levels of pro-inflammatory cytokines in young adults (Brody, Yu, Miller, & Chen, 2015), and of C-reactive protein in adolescents (Goosby, Malone, Richardson, Cheadle, & Williams, 2015). Together, these studies support the importance of understanding the health consequences of exposure to racial discrimination during adolescence, a developmental stage in which African American youth become keenly aware of targeted social rejection and their treatment by others (Stevenson, 2004).

Current estimates suggest that nearly 40% of Black adults living in the southeastern United States have diabetes (Centers for Disease Control and Prevention, 2014). This alarming prevalence rate is among the highest in the nation (Menke, Casagrande, Geiss, & Cowie, 2015). Many approaches can be taken in dealing with this situation, but health psychologists, pediatricians, and public health scientists agree that understanding the contributions of exposure to race-related stressors during childhood and adolescence is important (Shonkoff, Boyce, & McEwen, 2009). In response to this need, the first purpose of this study is to determine whether exposure to high levels of racial discrimination across adolescence will predict elevated rates of insulin resistance (IR) during young adulthood. In IR, muscle, fat, and liver cells do not respond properly to insulin, reducing the absorption of glucose into cells. As a result, excess glucose builds up in the bloodstream, which can lead to diabetes. The notion examined here is that recurring exposure to interpersonal discrimination leads to persistent activation of neuroendocrine and inflammatory response systems. The end products of these systems (cortisol, epinephrine, norepinephrine, cytokines, and chemokines) all contribute to the emergence of IR (Björntorp & Rosmond, 1999; Lackey & Olefsky, 2016; McEwen, 1998).

In addition to triggering sustained activation of stress responses, the experience of interpersonal discrimination may also place young people at risk for elevated body mass index (BMI), an indicator that controls for height when measuring weight and is a risk factor for IR. It has been hypothesized that frequent encounters with interpersonal discrimination are associated with overconsumption of high-calorie foods to alleviate the resulting symptoms of stress (J. S. Jackson, Knight, & Rafferty, 2010; Pascoe & Smart Richman, 2009). Such a stress-coping dynamic would be reflected in a positive association between elevated exposure to interpersonal discrimination and BMI. Surprisingly few studies have examined this association among African American adolescents and young adults. One study found a cross-sectional relationship between exposure to racial discrimination and BMI (Nelson et al., 2018). Results of studies examining this relationship among African American adults have been inconsistent. Some studies found a positive association (Stepanikova et al., 2017; Thorpe, Parker, Cobb, Dillard, & Bowie, 2017), some found no association (Hunte & Williams, 2009; Johnson, Risica, Gans, Kirtania, & Kumanyika, 2012; Shelton et al., 2009), and others found an association that is stronger among women than among men (Cunningham et al., 2013). Although the interpretation of these findings is complicated by differences in the ages of the samples and recruitment variations, in all of the studies participants are asked about their exposure to racial discrimination at only one point in time. This limits understanding about the ways in which differences in exposures to racial discrimination across time relate to BMI. To overcome this limitation, in this study we followed a cohort of African American adolescents who were within two years of each other in age and assessed their exposure to interpersonal discrimination from ages 16 to 18.

This study also addressed a methodological issue. Prior research has shown that African American youth have qualitatively different experiences of racial discrimination across adolescence. Most youth encounter low but increasing exposures to racial discrimination, whereas others experience high exposures that remain stable over time (Brody et al., 2006; Greene, Way, & Pahl, 2006; Simons, Chen, Stewart, & Brody, 2003). Among adolescents who experience exclusively high exposures to racial discrimination, prospective associations with psychosocial and health-related indicators emerge (Brody, Yu, & Beach, 2016). Thus, perceived discrimination was hypothesized to follow qualitatively different trajectories across ages 16 to 18 years. This study was designed to test the hypotheses that exposure to high levels of discrimination across ages 16 to 18 years would be associated with BMI across ages 19 to 21 years, and that BMI would serve as a mediator connecting racial discrimination during adolescence to IR during young adulthood at ages 25 and 27 years.

## Methods

### Sample

Data for this study were drawn from the Strong African American Healthy Adult Project (SHAPE; Brody et al., 2013). Starting in 2001, 667 African American children in fifth grade (mean age = 11.2 years; range 11 to 13 years) and their primary caregivers were enrolled in SHAPE. The families resided in nine rural counties in Georgia, where poverty rates are among the highest in the nation. In 2009–2010, when the youth had reached ages 19 and 20, a subgroup was randomly selected for a substudy of stress hormones and blood pressure. In

2015–2016 and 2017, when the youth's mean ages were 25 and 27 years, blood draws were obtained from 391 participants at the age 25 assessment and 388 participants at age 27. Overnight fasting blood glucose and insulin samples were assayed from the blood samples. For the current study, the analytic sample consists of the 337 participants for whom complete data were available for all study variables across ages 16 to 27. Compared with the original cohort, the analytic sample included a higher percentage of female participants (61.1% vs. 52.8%); otherwise, the samples were demographically similar. The University of Georgia's Institutional Review Board approved the protocol and, at all waves of data collection, written informed consent was obtained from caregivers of youth under age 18 and from youth 18 and older. Minor youth assented in writing.

### Data Collection Procedures

All data were collected in participants' homes using a standardized protocol. A field researcher who was also a certified phlebotomist went to each participant's home to collect self-report data and to draw a blood sample for the assessment of glucose and insulin.

### Measures

Racial discrimination was assessed at ages 16, 17, and 18. BMI was measured at ages 19, 20, and 21. Fasting blood glucose and insulin were assessed at ages 25 and 27. The covariates gender, cumulative SES risk, life stress, and depressive symptoms were assessed at ages 16 to 18, overlapping the assessment of racial discrimination. The lifestyle covariates of drug use and exercise were assessed at ages 19 to 21, concurrently with the measurement of BMI. These variables were controlled in every analysis.

**Racial discrimination.**—At ages 16 to 18 years, youth responded to nine items from the Schedule of Racist Events (Landrine & Klonoff, 1996), which was designed to be developmentally appropriate for Black adolescents (Brody et al., 2006). This scale has been used extensively in public health and biomedical research and has demonstrated predictive validity for both psychosocial and physical health indicators (Lewis et al., 2015). The scale was used to assess the frequency during the past year with which the respondent experienced specific discriminatory events. These events included racially based slurs and insults, disrespectful treatment from community members, physical threats, and false accusations from business employees or law enforcement officials. Responses to items were summed at each wave across ages 16 to 18 years. Cronbach's alphas ranged from .87 to .89 across assessments.

**Body mass index.**—Each participant's weight and height at ages 19 to 21 years were used to calculate BMI (weight in kilograms divided by the square of height in meters). Researchers measured weight using a standard home scale and height using a tape measure. BMI was averaged across ages 19 to 21 years.

**Young adult blood glucose, insulin, and insulin resistance.**—At the age 25 and 27 assessments, a phlebotomist visited each participant's home in the morning to draw an overnight fasting blood sample. Blood was drawn into serum separator tubes (Becton, Dickinson and Company, Franklin Lakes, NJ, USA). Specimens were centrifuged on site at

1500 × g for 20 minutes. The serum was harvested, divided into aliquots, and immediately frozen on dry ice. Upon arrival at the lab, it was placed in storage at –80 C until the end of the project. At that point, serum glucose was measured photometrically using a UV test on a Roche/Hitachi Cobas c502 instrument. This assay has a dynamic range of 2–750 mg/dL and intra-assay coefficient of variation of 0.7%. Serum insulin was assayed in duplicate using a multiplex, electrochemiluminescent immunoassay (Human Leptin/Insulin Kit K15164C; MesoScale Discovery) on a SECTOR Imager 2400A (MesoScale Discovery). This assay has an 8-point standard curve, with a lower limit of detection of 25 pg/ml for insulin. The intra-assay coefficient of variation for duplicate pairs averaged 3.8%. The indicator of IR was estimated according to the original homeostasis assessment model (HOMA; Matthews et al., 1985), where IR was calculated as [fasting glucose (mmol/L) × fasting insulin (mIU/L)] / 22.5. IR levels were log-transformed because of the skewed distribution and were averaged across ages 25 and 27 ( $r = .57, p < .001$ ).

**Confounder variables.**—Family SES-related disadvantage was measured as the sum of six indicators (0 if absent, 1 if present; see Brody et al., 2013): current family poverty, primary caregiver’s noncompletion of high school or an equivalent, primary caregiver current unemployment, single-parent family structure, current receipt of Temporary Assistance for Needy Families, and income rated by the caregiver as inadequate to meet all needs. The average number of indicators across ages 16 to 18 constituted the measure of SES-related disadvantage. Personal life stress was assessed with a 12-item checklist (Brody, Chen, & Kogan, 2010) on which youth indicated whether stressors, such as experiencing a serious injury or illness, acute economic stress, or parental divorce, had occurred during the previous 6 months. The number of endorsed items was tallied at each wave, and these sums were averaged across ages 17 and 18. Depressive symptoms were assessed with the Children’s Depression Inventory (Kovacs, 1979); alphas ranged from .84 to .86. Depression scores at ages 16 and 17 were averaged. Drug use was assessed at ages 19, 20, and 21 using items from the Monitoring the Future study (Johnston, O’Malley, Bachman, & Schulenberg, 2007). The items assessed past-month cigarette, alcohol, and marijuana use, as well as binge drinking. Responses to these items were summed to form a drug use composite. Exercise was measured with a single item indicating the number of days during the past week the participant was physically active for 60 minutes. The drug use and exercise items were averaged separately across ages 19 to 21.

### Statistical Approach

Black youth living in both rural and urban areas experience heterogeneity in their exposure to racial discrimination across adolescence. Thus, the first data analytic step was to identify youth who experienced distinct trajectories of exposure to racial discrimination across ages 16 to 18 years. Latent growth curve analysis and latent growth mixture modeling were used to do this. First, latent growth curve models (LGCM) were executed to characterize the overall experience of racial discrimination across ages 16, 17, and 18. Linear growth models were fit with two individual growth parameters: an intercept parameter representing the level of racial discrimination at the starting point (age 16) and a linear slope parameter representing the rate at which racial discrimination changed across the three waves of data collection. Second, latent growth mixture models (LGMM) were executed to identify

distinct groups characterized by different levels of exposure to racial discrimination. LGMM determines whether the population under study is composed of distinct classes of individuals based on variation in intercepts or slopes. Model fits were compared for one-, two-, and three-class LGMM, assessing relative fit with conventional indices, including the Bayesian information criterion (BIC) and entropy values. Lower BIC scores indicate better fit, whereas higher entropy scores reflect greater accuracy in classification. Optimal models were chosen on the basis of goodness of fit and parsimony.

After distinct classes of exposure to racial discrimination were identified, univariate analyses of covariance (ANCOVA) were conducted to determine whether members of the racial discrimination trajectory groups differed significantly on their BMI at ages 19 to 21 and IR levels at ages 25 and 27. Hierarchical linear regression analyses were then conducted to determine whether BMI would mediate the association of racial discrimination with IR. For all models, covariates were entered first, followed by racial discrimination trajectory group and BMI. Indirect (or mediation) effects were tested by calculating bias-corrected 95% CIs using bootstrapping with 1000 resamples via the Process procedure for SPSS (Hayes, 2013).

## Results

The LGCM analysis revealed that racial discrimination increased over time (mean slope = 0.784,  $p < .001$ ). The analysis also revealed significant variation in the intercept (variance = 10.390,  $p < .001$ ) and the slope (variance = 2.171,  $p < .001$ ). The results of the LGMM analyses revealed that the three-class model had the lower BIC statistics (6904.85 for three-class vs. 6921.34 for two-class); however, the entropy values indicated that the two-class model fit the data better (0.77 for three-class vs. 0.81 for two-class). The two-class model, therefore, was chosen for follow-up hypothesis testing. The majority of participants (77.7%) were assigned to the *low-and-increasing class*, which had a low intercept at age 16 (mean intercept = 2.169, 95% CI [1.804, 2.534],  $p < .001$ ) with increasing exposure to discrimination from ages 16 to 18 (mean slope = 0.961, 95% CI [0.773, 1.149],  $p < .001$ ). The second group, which comprised 22.3% of the sample, was the *high-and-stable class*. Youth in this group reported a high level of racial discrimination that remained high across time (mean intercept = 8.610, 95% CI [7.691, 9.529],  $p < .001$ ; mean slope = 0.160, 95% CI [-0.287, 0.607],  $p = .484$ ). Figure 1 illustrates the discrimination trajectories for the classes identified. Table 1 presents demographic characteristics for both racial discrimination exposure groups and for the entire sample. As Table 1 indicates, both the entire sample and the high-and-stable class included more female than male youth.

### Racial Discrimination Trajectory Group Differences on Body Mass Index and Insulin Resistance

The first ANCOVA revealed that youth in the high-and-stable racial discrimination trajectory group displayed significantly higher IR scores at ages 25 and 27 ( $M = 1.355$ , 95% CI [1.288, 1.422]) than did those in the low-and-increasing group ( $M = 1.264$ , 95% CI [1.229, 1.299]), Cohen's  $d = 0.313$ . Figure 2a depicts this finding. The second ANCOVA showed that youth in the high-and-stable group had significantly higher BMI at ages 19 to 21 ( $M = 31.211$ ,

95% CI [29.259, 33.163]) than did those in the low-and-increasing group ( $M = 28.044$ , 95% CI [27.024, 29.064]), Cohen's  $d = 0.377$ . Figure 2b illustrates this result.

### **Body Mass Index at Ages 19 to 21 as a Mediator of the Association between Racial Discrimination and Insulin Resistance at Ages 25 and 27**

Table 2 presents descriptive statistics for, and correlations among, the study variables. The previously described analyses established associations of racial discrimination with subsequent BMI across ages 19 to 21 and IR at ages 25 and 27. Indirect (i.e., mediated) effects of racial discrimination on IR through BMI levels were also hypothesized. Figure 3 provides a conceptual diagram of the mediational analyses conducted. MacKinnon, Lockwood, Hoffman, West, and Sheets (2002) suggested that mediation analyses should be conducted when a relation exists between both a predictor and a mediator (Path A in Figure 3) and a mediator and an outcome (Path B in Figure 3). To be considered a mediator, the strength of the direct relation between the predictor and outcome (Path C in Figure 3) must decrease when the mediator is entered into the analysis (Path C' in Figure 3). Covariates included in the mediational analysis were the same as those in prior analyses and only applied to the outcome variables. First, the relationship was established between the predictor (racial discrimination) and the potential mediator (BMI) and then the relationship of the mediator to the outcome (IR) was examined. Upon determining whether the mediator-to-outcome relationship was significant, BMI's possible reduction of the direct effect of racial discrimination on IR was explored.

Figure 3 also shows the results of the mediational analyses. In the regression analyses, membership in the low-and-increasing trajectory group was assigned a value of 0; membership in the high-and-stable group was assigned a value of 1. As hypothesized, the presence of the mediator, BMI at ages 19 to 21, reduced the direct effect of membership in the high-and-stable group at ages 16 to 18 on risk for elevated IR at ages 25 and 27, from  $b = 0.091$ , 95% CI [0.014, 0.167],  $p = .021$  to  $b = 0.030$ , 95% CI [-0.034, 0.095],  $p = .356$ . The indirect effect of high racial discrimination levels on risk for elevated IR via BMI was significant (estimates = .060, 95% BCA [0.014, 0.115] with 1000 bootstrapping).

To clarify the results further, life stress, depression, exercise, and drug use were considered as possible mediators. These variables were included in separate models as mediators linking racial discrimination with IR. None of these indirect/mediating pathways, however, reached statistical significance. Gender was also considered as a potential moderator of the pathways in the mediation model. None of the interactions with gender reached statistical significance.

## **Discussion**

During their high school years, African American adolescents simultaneously experience emotional and social challenges while developing independence, a sense of self, and racial identity. Adolescence is also a time when African American youth become keenly aware of their treatment relative to others and are particularly cognizant of targeted rejection. Against this developmental backdrop, and building on previous studies of exposure to racial discrimination and preclinical indicators such as elevated glucocorticoids and pro-

inflammatory cytokines (Lewis et al., 2015), exposure to high doses of racial discrimination across adolescence were found to presage IR diagnosed 7 years later. These findings are among the first to establish a link between racial discrimination and a clinical indicator of cardiometabolic health during the transition to adulthood. If their risk for IR persists, these young adults will likely experience disproportionately high rates of morbidity and mortality from type 2 diabetes and cardiovascular disease in future decades (Mottillo et al., 2010).

Evidence, shown in Figure 2, emerged to indicate that BMI mediates the linkage between exposure to relatively high levels of racial discrimination across adolescence and IR. Adolescents exposed to more race-related stress are at heightened risk for IR, and this association is largely attributable to elevated BMI. These findings extend earlier research by focusing on a group of adolescents close in age, obtaining repeated assessments of racial discrimination and BMI, and testing the hypothesis that BMI serves as a mediator linking trajectories of racial discrimination to IR. Although these are important additions to the literature on the development of IR in cardiometabolic risk, caution must be exercised in their interpretation.

BMI may have contributed to exposure to racial discrimination; even though discrimination was assessed before BMI, it was not necessarily experienced first. In future longitudinal studies, it will be important to determine whether having an elevated BMI increases exposure to racial discrimination. In addition, future research should examine potential mediators between exposure to discrimination and BMI. Stress reduction is a plausible mediator. When young people are exposed frequently to racial discrimination in their daily lives, some may increase their consumption of high-calorie foods to help alleviate negative affect and other effects of stress (J. S. Jackson et al., 2010).

Recent research has begun to establish that perceived weight stigma, principally weight discrimination, is associated with physical health consequences. Perceived weight discrimination is associated with increases in weight, waist circumference, and, for those in the overweight range, the development of obesity over time (S. E. Jackson, Beeken, & Wardle, 2014). The effects of weight discrimination are particularly pernicious among those who internalize discriminatory weight-based attitudes. This self-directed stigma, also known as weight-based internalization (WBI), is associated with increases in binge eating (Hübner et al., 2016), reductions in physical activity (Pearl, Puhl, & Dovidio, 2015), and the presence of metabolic syndrome (Pearl et al., 2017). Thus, as intersectionality paradigms have suggested (Himmelstein, Puhl, & Quinn, 2017), research on the contributions of racial discrimination to cardiometabolic outcomes should include measures of weight discrimination and WBI. This approach not only will disentangle the unique contributions of racial discrimination and weight discrimination, but also will allow the testing of a “double jeopardy hypothesis” that persons of color with higher weight are at particularly high cardiometabolic risk.

Several potential explanations for this study’s findings were eliminated. First, family SES risk, depressive symptoms, and personal life stress were assessed at the same time as racial discrimination. The results in Table 2 show that neither depressive symptoms nor personal life stress was associated with subsequent assessments of BMI or IR. Conversely, racial



discrimination was associated over time with both health indicators. Because family SES risk is associated with both BMI and IR, SES risk was controlled in all data analyses. Physical activity was not associated with IR. Drug use was negatively associated with IR, suggesting that it does not act as a mediator connecting racial discrimination with IR. These findings are consistent with results from other studies showing that, although obesity and drug use take a toll on health outcomes, they may do so in different ways. Drug use increases risks for mental health problems, including depression, anxiety, and delinquency (Elliott, Huizinga, & Menard, 2012; Grant et al., 2004). In contrast, to the extent that it leads to obesity, overeating heightens risk for IR and, later in life, cardiovascular disease (Grundy, 2004; Guh et al., 2009). Although other stress-related mediators of racial discrimination effects exist, such as sleep loss, elevated adrenocortical and sympathetic outflow, or inflammatory activity (Kelly & Ismail, 2015), the results of this study support the plausibility of the underlying model of racial discrimination → BMI → IR. This study reinforces the notion that exposure to racial discrimination during adolescence has consequences for cardiometabolic health. The results suggest that race-related stressors contribute to cardiometabolic risk in young Black adults and may exacerbate existing community health disparities.

Strengths of this study include the longitudinal, prospective design that included assessments of racial discrimination and BMI across time, as well as their associations with IR. The lack of a BMI or IR assessment at the first wave of data collection is a major limitation. Because this study was not initially designed to focus on health, it included no individual health, family health history, biological, or anthropometric measures. In addition racial discrimination before age 16 years was not assessed, even though the literature suggests that exposure to discrimination before age 12 could have an important impact on children (Lee & Ahn, 2013; Sanders-Phillips et al., 2014; Schmitt, Branscombe, Postmes, & Garcia, 2014). Thus, future research should be designed to determine whether the relations found in this study, including trajectories of discrimination, are evident at younger ages. Finally, because this study was conducted with rural African American adolescents, future research should be designed to determine whether patterns would be different among urban African American adolescents, other minority groups, or Caucasian adolescents. These cautions notwithstanding, this study is among the first to show a positive, prospective association between perceived discrimination and IR.

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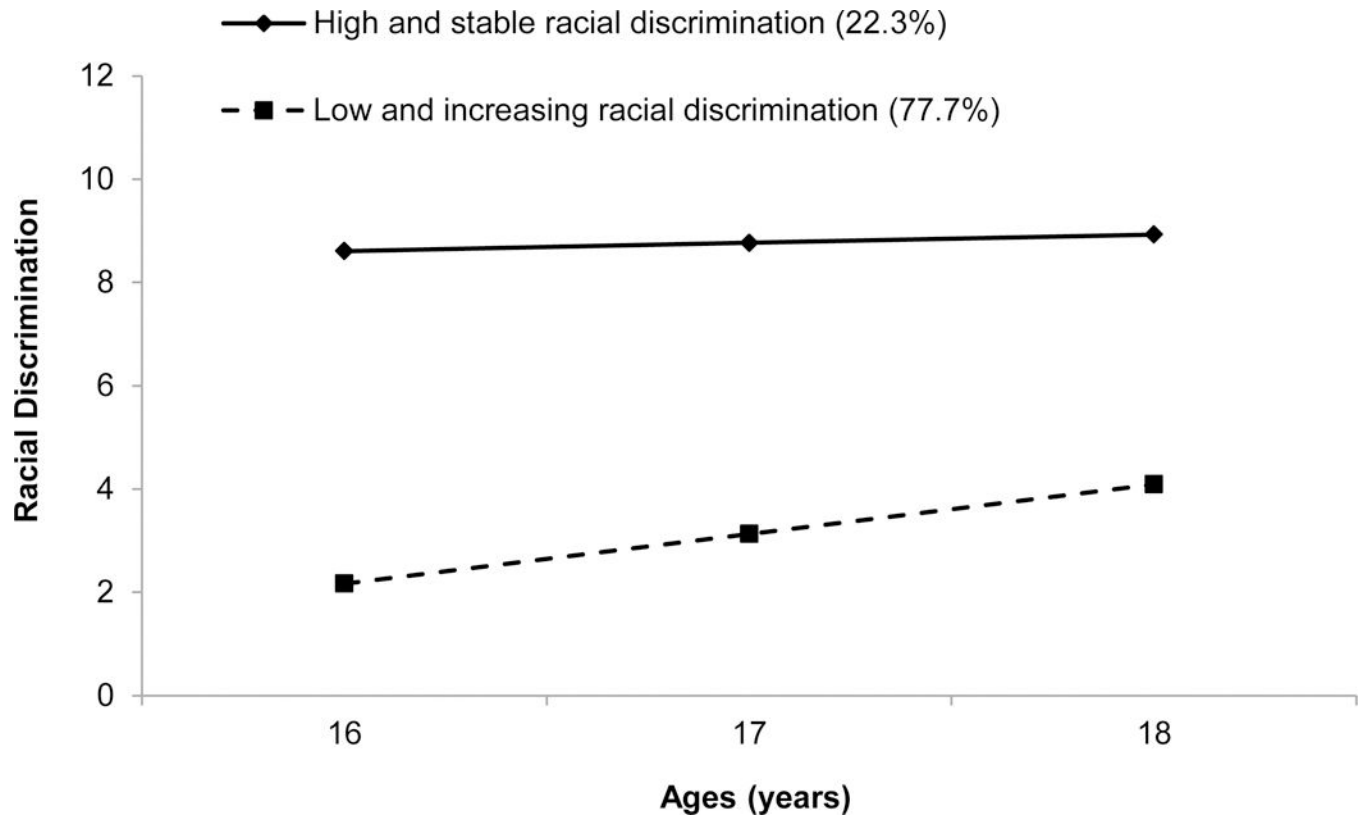
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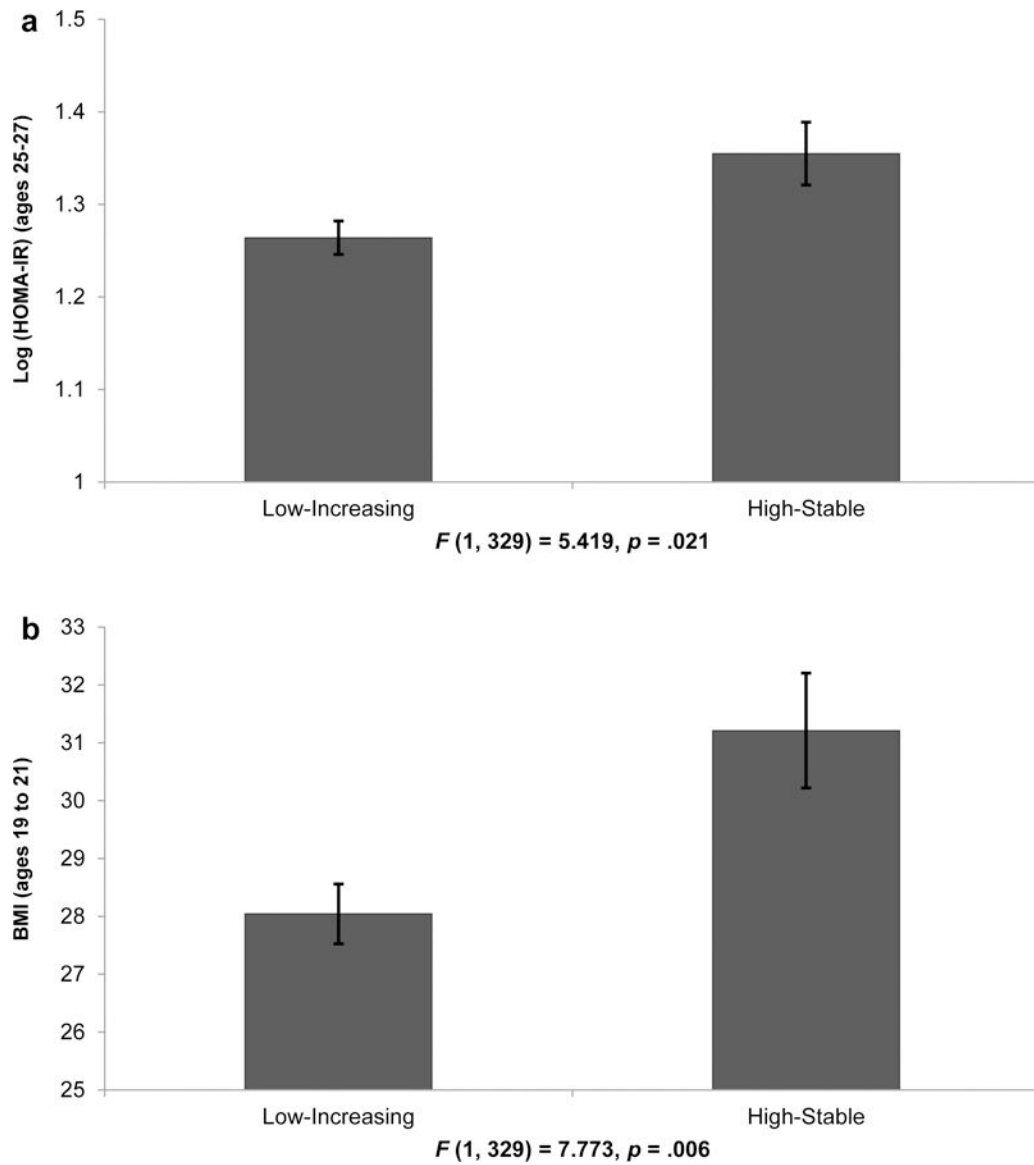
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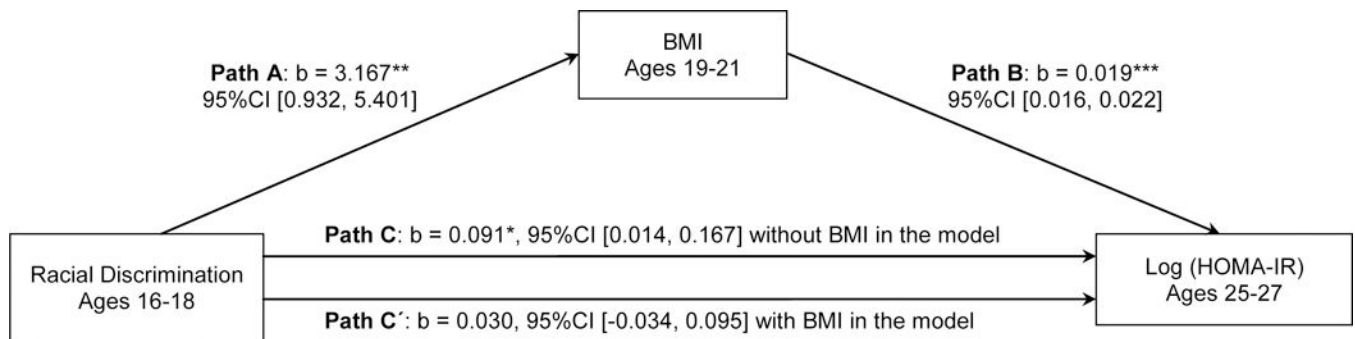
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**Figure 1.**  
The estimated trajectories of racial discrimination for each class.



**Figure 2.** Means of (a) IR levels at ages 25 to 27 and (b) BMI at ages 19 to 21 for the racial discrimination trajectory groups. Error bars =  $\pm 1$  standard error. Analyses controlled for gender, family SES disadvantage, depressive symptoms, life stress, drug use, and exercise.



**Figure 3.**

BMI at ages 19 to 21 as a mediator of the relation between racial discrimination at ages 16 to 18 and IR at ages 25 to 27. Family socioeconomic-related risk at ages 16 to 18, gender, youth depressive symptoms at ages 16 to 17, youth personal life stress at ages 17 to 18, and youth substance use and exercise at ages 19 to 21 were controlled (not shown).

Unstandardized coefficients with 95% confident intervals (CI) are presented.  $N = 337$ .  $*p < .05$ ;  $**p < .01$ ;  $***p < .001$ .

**Table 1.**

## Demographic Characteristics by Racial Discrimination Trajectory Groups

	<b>Whole Sample (<i>N</i> = 337)</b>	<b>Low-and-Increasing Class (<i>n</i> = 262)</b>	<b>High-and-Stable Class (<i>n</i> = 75)</b>
<b>Characteristics</b>	<b>% or Mean (<i>SD</i>)</b>	<b>% or Mean (<i>SD</i>)</b>	<b>% or Mean (<i>SD</i>)</b>
Age 16			
Target age (in years)	15.58 (0.49)	15.57 (0.50)	15.60 (0.49)
Parent age (in years)	42.20 (7.51)	42.26 (7.77)	41.99 (6.58)
Target gender, male	38.9%	40.5%	33.3%
Family poverty	48.2%	49.2%	44.9%
Parent unemployment	23.1%	22.8%	24.0%
Parent education			
< high school	21.9%	23.6%	16.0%
HS degree or GED	32.4%	31.4%	36.0%
Some college or technical school	38.1%	38.46%	37.3%
college graduate	7.5%	6.6%	10.7%
Single-parent status	58.7%	61.8%	48.0%



**Table 2**

Correlations and Descriptive Statistics among Study Variables

Study Variables	1	2	3	4	5	6	7	8	9
1. Log (HOMA-IR) (ages 25–27)	--								
2. BMI (ages 19–21)	.593 ***	--							
3. Discrimination class <sup>a</sup> (ages 16–18)	.091	.138 *	--						
4. Gender (male = 1)	-.303 ***	-.181 **	-.061	--					
5. Family SES disadvantage (ages 16–18)	.125 *	.107 *	-.095	-.010	--				
6. Depressive symptoms (ages 16–17)	-.069	-.034	.209 ***	-.044	.069	--			
7. Personal life stress (ages 17–18)	.024	.058	.106	.144 **	.152 **	.303 ***	--		
8. Substance use (ages 19–21)	-.215 ***	-.073	.151 **	.269 ***	-.051	.130 *	.247 ***	--	
9. Exercise (ages 19–21)	-.103	-.096	-.090	.239 ***	.027	-.095	.046	.017	--
Mean	1.28	28.75	0.22	0.39	2.35	7.17	15.70	0.73	2.82
SD	0.31	8.58	0.42	0.49	1.25	5.22	1.56	0.65	1.55

Note. N = 337.

<sup>a</sup> Membership in the high-and-stable class was coded as 1, and membership in the low-and-increasing class was coded as 0. BMI, body mass index; SES, socioeconomic status.

\*  $p < .05$

\*\*  $p < .01$

\*\*\*  $p < .001$ .