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# Sexual Orientation and Health Information Technology Use: A Nationally Representative Study of U.S. Adults

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#### Abstract

**Purpose:** The purpose of this study was to compare the prevalence and odds of participation in online health-related activities among lesbian, gay, and bisexual adults and straight adults aged 18–64.

**Methods:** Primary data collected in the 2013 and 2014 National Health Interview Survey, a nationally representative household health survey, were used to examine associations between sexual orientation and four measures of health information technology (HIT) use. Data were collected through face-to-face interviews (some telephone follow-up) with 54,878 adults aged 18–64.

**Results:** Compared with straight men, both gay and bisexual men had higher odds of using computers to schedule appointments with healthcare providers, and using email to communicate with healthcare providers. Gay men also had significantly higher odds of seeking health information or participating in a health-related chat group on the Internet, and using computers to fill a prescription. No significant associations were observed between sexual orientation and HIT use among women in the multivariate analysis.

**Conclusions:** Gay and bisexual men make greater use of HIT than their straight counterparts. Additional research is needed to determine the causal factors behind these group differences in the use of online healthcare, as well as the health implications for each group.

### **Keywords**

health care access; health information technology	(HIT); Internet use; National Health In	nterview
Survey; sexual orientation; sexual identity		

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Author Disclosure Statement

No competing financial interests exist.

# Introduction

As the healthcare system in the United States evolves, use of health information technology (HIT), the exchange of healthcare information through computerized systems, continues to grow. In 2012, 81% of U.S. adults used the Internet, 59% went online in search of health/medical information, and 35% used Internet-based health information to self-diagnose a medical condition for themselves or someone else. The National Health Interview Survey (NHIS) began collecting data on these outcomes in 2009. That year, 2.7% of adults aged 18–64 had scheduled an appointment with a healthcare provider through the Internet and 4.9% had communicated with a provider by email. By 2015, these percentages reached 11.3% and 11.7%, respectively. By facilitating communication between patients and providers, the Internet may be an important part of the healthcare toolbox available to consumers.

With this rise in HIT use, research has emerged to describe its correlates. Age, sex, and education have repeatedly predicted HIT use across a wide range of samples and study designs. Specifically, women, younger adults, and more educated adults are more likely to search the Internet for health information, participate in online health-related chat groups, and fill prescriptions on the Internet than men, older adults, and less educated adults. An umber of other sociodemographic, socioeconomic, health status, and healthcare access/utilization measures have been shown to predict HIT use, although findings on strength and direction of these associations are mixed: race/ethnicity; 10,19,23 income; 5,8,10,16,17,24 employment; 11,17,20 marital status; 4,16,18,19 urbanicity; 7,14 health insurance coverage; 5,8,25 general health status; 4,6,7,10,13,16,18,26 chronic conditions; 5,7,18,21 and doctor/healthcare facility visits. Surprisingly, our extensive search of the literature produced no studies that examined HIT use among lesbian, gay, and bisexual (LGB) persons.

A systematic review of research identified several possible reasons for the observed rise in HIT use and interventions.<sup>27</sup> One, stigma reduction, may be particularly relevant for the LGB population. Sexual minorities may have difficulty disclosing their sexual orientation to healthcare providers due to perceived or anticipated stigma.<sup>28–31</sup> LGB persons' efforts to hide their stigmatized identities may result in delayed/inadequate medical care and/or lead them to alternative forms of care.<sup>32–40</sup> The Internet, by contrast, allows one to search and consume health information without disclosing one's identity, offering a level of anonymity not available in face-to-face settings. Therefore, LGB persons may perceive benefits from using the Internet for health information and communication.<sup>41</sup>

We use a nationally representative sample of adults to examine whether men and women who self-identify as sexual minorities (i.e., gay/lesbian, bisexual) are more likely to use HIT than straight persons. We hypothesize that a sexual minority identity is associated with greater use of HIT, net of sociodemographic, socioeconomic, and health status covariates. While direct measures of perceived stigma were not available, we include measures of healthcare access and utilization in final multivariate models. If stigma is a causal factor, the inclusion of access and utilization measures should attenuate, if not eliminate, remaining significant relationships between sexual orientation and HIT use.

# Methods

#### Data

Data collected from 54,878 sample adults aged 18–64 who participated in the 2013 and 2014 NHIS were used in this study. The NHIS is a multipurpose, nationally representative health survey of the civilian noninstitutionalized U.S. population, conducted by the National Center for Health Statistics (NCHS). Interviewers with the U.S. Census Bureau administer the questionnaire using computer-assisted personal interviewing. Telephone interviewing is permitted to complete missing portions of the interview. 42,43

Analytic variables were drawn from the NHIS Household Composition, Family Core, and Sample Adult Core modules. Demographic and relationship information on all householders was collected with the Household Composition module. The Family Core module collected self- and proxy-reported information on each member of the selected family, including disability status, health insurance coverage, and income. The Sample Adult Core module, administered to one adult aged 18 randomly selected from each family, collected information on sexual orientation, health conditions, health status and limitations, and heath care access and use (including use of HIT). The sample adult answered for himself/herself unless mentally or physically incapacitated, in which case, a knowledgeable family member served as a proxy respondent. The final sample adult response rate for the 2 years was 60.0%.  $^{42,43}$ 

Both the 2013 and 2014 NHIS were approved by the research ethics review board of the NCHS and by the U.S. Office of Management and Budget. This specific study was exempt from review. All respondents provided oral consent before participation.

#### Measures

**HIT use.**—Four dichotomous measures of HIT use were examined. The first, "look up health information on Internet/participate in Internet health chat groups," is based on responses to questions asking if, during the past 12 months, he/she had used computers "to look up health information on the Internet" or "to use online chat groups to learn about health topics." If the adult answered "yes" to either question, he/she was defined as having sought health information on the Internet.

The second outcome is based on a yes/no question that asks if the adult, during the past 12 months, "used computers to fill a prescription." The third outcome is based on a yes/no question that ascertains if the adult, in the past year, had "used computers to schedule an appointment with a healthcare provider." The final outcome is based on a question that asks the adult if he/she had, in the past 12 months, "used computers to communicate with a healthcare provider by email" (yes/no).

**Sexual orientation.**—Sexual orientation was ascertained with the following identity question: "Which of the following best represents how you think of yourself?" For male respondents, response options were "gay"; "straight, that is, not gay"; "bisexual"; "something else"; and "I don't know the answer." For female respondents, response categories were "lesbian or gay"; "straight, that is, not lesbian or gay"; "bisexual";

"something else"; and "I don't know the answer." Given the small percentage of adults who answered "something else" (0.2%) or "I don't know the answer" (0.4%), these responses, along with "refused," were treated as missing in the analyses. Quality evaluations of the sexual identity question, including an analysis of responses to follow-up questions for "something else" and "I don't know the answer," are discussed elsewhere. 44,45

**Sociodemographic characteristics.**—Sociodemographic covariates empirically informed by the literature on HIT use included the following: sex, age, race/ethnicity, marital status, nativity, neighborhood attachment, U.S. region of residence, and urbanicity. More detail on the neighborhood attachment variable is provided elsewhere.<sup>35</sup>

**Socioeconomic characteristics.**—Education, work status, and poverty status (from NHIS imputed income files) comprised the socioeconomic characteristics. Poverty status, (total family income/federal poverty threshold [FPT])  $\times$  100, was categorized as poor (<100% FPT), near poor (100% FPT <200%), and not poor (200% FPT).

**Health status.**—Covariates included reported health status, activity limitation, multiple chronic conditions ( 2 of 10 selected conditions), and serious psychological distress (a score on the K6 scale of 13). Additional details on each of these measures can be found elsewhere. 35,46–48

Healthcare access and utilization.—Health insurance coverage, number of different providers seen/talked to in the past 12 months, whether or not the sample adult had surgery in the past 12 months, and four separate measures of barriers to healthcare encountered in the past 12 months (did not receive specific services due to cost, delayed care for noncost reasons, trouble finding a provider, and no usual source of medical care) were used to capture healthcare access and utilization. Number of different providers involved summing the number of providers seen in the past 12 months and recoding to 0, 1, 2, or 3 or more. Providers included mental health professionals; optometrist/ophthalmologist/eye doctor; foot doctor; chiropractor; physical therapist, speech therapist, respiratory therapist, audiologist, or occupational therapist; nurse practitioner, physician assistant, or midwife; medical doctor specializing in a particular medical disease/problem; and general doctor. Additional description of the barriers to healthcare measures is provided elsewhere.<sup>35</sup>

Frequency of computer use. Frequency of computer use was included in all multivariate analyses as a control.

## Statistical analyses

Since the proportion of adults aged 65 who identified as a sexual minority was <1%, our analyses are limited to adults aged 18–64. All analyses are stratified by sex. Descriptive statistics for all variables are presented first, followed by prevalence rates for HIT use by sexual orientation. Two-tailed significance tests were performed to determine whether significant differences exist by sexual orientation. Note that small sample sizes of LGB adults in the NHIS (due to the relatively small size of the larger LGB population) led to the production of certain prevalence estimates that require cautious interpretation (noted in the

Tables). This includes those with a relative standard error (i.e., [standard error/estimate]  $\times$  100) >30.0%, but 50.0%.

Two sets of multivariate logistic regression models were fit to assess the relationship between sexual orientation and each HIT-use outcome, net of covariates. The first set includes the sociodemographic, socioeconomic, health status, and frequency of computer use measures. The second set adds the healthcare access and utilization measures. The initial models allow us to determine if significant bivariate associations between sexual orientation and HIT use can be explained by factors other than the access and utilization measures. The final models allow us to assess, indirectly, the role of stigma in driving HIT use among LGB adults. We would expect stigma to impact HIT use through reduced access and utilization of conventional healthcare resources. By controlling for the aforementioned access/utilization measures, we expect an attenuation, if not elimination, of significant relationships between sexual orientation and HIT use.

To account for the stratified, complex cluster sampling design of the NHIS, analyses were conducted in SAS-callable SUDAAN 11.0 software (RTI International, Research Triangle Park, NC). To ensure the results are generalizable to the U.S. adult, civilian noninstitutionalized population aged 18–64, all analyses used final sample adult weights adjusted for nonresponse and calibrated to population control totals.

#### Results

# **Descriptive statistics**

Table 1 presents sex-stratified descriptive statistics for sexual orientation, HIT use, and covariates, for adults aged 18–64. For all four HIT outcomes, a higher percentage of women aged 18–64 used HIT compared to men. For women, percentages ranged from 8.5% for use of computers to fill a prescription to 56.5% for seeking health information on the Internet. HIT use for men ranged from 5.8% for using computers to schedule an appointment with a healthcare provider to 42.7% for seeking health information on the Internet. As for sexual orientation, 1.6% of women identified as gay, 1.2% as bisexual, and 97.2% as straight. For men, the percentages were 2.0%, 0.4%, and 97.6% respectively.

#### Bivariate analyses

Bivariate associations between sexual orientation and HIT use among men and women aged 18–64 are presented in Table 2. Among women, only two significant differences emerged. A higher percentage of bisexual women and gay/lesbian women sought health information on the Internet compared with straight women. For men, a significantly higher percentage of those who identified as gay or bisexual sought health information on the Internet compared with those who identified as straight. Similarly, a higher percentage of gay (vs. straight) men used computers to fill a prescription, schedule an appointment with a healthcare provider, and communicated with a healthcare provider by email. For both men and women, no significant differences in HIT use were observed in comparisons of gay/lesbian and bisexual adults.

# Multivariate analyses

Table 3 presents adjusted odds ratios and 95% confidence intervals for sexual orientation from two separate logistic regression models fit for each HIT use measure. The initial model for each outcome includes sociodemographic, socioeconomic, health status, and frequency of computer use measures. The full model for each outcome adds the healthcare access and utilization measures.

Focusing on the initial model results, all but one (sought health information on the Internet among gay or lesbian women) of the significant bivariate associations held. Interestingly, two new significant effects emerged for bisexual men: using computers to schedule an appointment with a healthcare provider and communicating with a healthcare provider by email. Among men, those identifying as gay or bisexual had higher odds of seeking health information on the Internet, using computers to schedule an appointment with a healthcare provider, and using computers to communicate with a healthcare provider by email compared to those identifying as straight. Gay men also had over twice the odds of using computers to fill a prescription. The lone significant association among women showed bisexuals, compared to straight adults, have significantly higher odds of seeking health information on the Internet.

In the full models, the access and utilization measures did attenuate the relationships between sexual orientation and HIT use observed in the initial models. However, while reductions in the magnitude of associations were consistent across all outcomes, six of the eight significant associations observed in the initial models held in the full models. The associations reduced to nonsignificance were sought health information on the Internet for bisexual men and bisexual women. As a consequence of the latter, no significant associations between sexual orientation and HIT use were observed for women when controlling for the full set of covariates.

Focusing on the six significant associations, gay men had higher odds of seeking health information on the Internet, over twice the odds of using computers to fill a prescription, and nearly twice the odds of using computers to schedule an appointment with a healthcare provider than straight men. In addition, bisexual men, compared to straight men, had nearly two-and-a-half times the odds of using computers to schedule an appointment with a healthcare provider. Finally, gay and bisexual men had over twice the odds of communicating with a healthcare provider by email than straight men.

## **Discussion**

In what may be the first nationally representative examination of LGB adults' use of HIT, we found, net of sociodemographic, socioeconomic, health status, and healthcare access and utilization covariates, that gay and bisexual men aged 18–64 had higher odds of using computers to schedule appointments with healthcare providers and communicating with healthcare providers by email compared to straight men. Gay men also had higher odds of seeking health information on the Internet and using computers to fill a prescription. No significant associations between sexual orientation and HIT use were observed for women in the full models.

Given the paucity of studies on HIT use by sexual orientation, we turned to research exploring reasons for HIT use and interventions, <sup>27,49–53</sup> and suggested stigma reduction as a possible reason for LGB adults to go online in search of health information and care. As expected, we identified significant associations between sexual orientation and HIT use. Not anticipated, however, were the persistent effects for sexual minority men after the addition of healthcare access and utilization measures to our models. Assuming that perceived stigma is associated with reduced utilization of traditional healthcare services, we anticipated that controlling for these measures would attenuate, if not eliminate, significant associations between sexual orientation and HIT use. While reductions in magnitude were observed, only two of eight significant associations in the initial models were reduced to nonsignificance in the final models.

The persistent associations between sexual minority status and HIT use for men are suggestive of causal factors other than, or in addition to, perceived stigma. One possible explanation involves the quality of interactions between LGB adults and traditional healthcare providers/settings. Regardless of the extent to which traditional services are used, sexual minorities may perceive the quality of their interactions as less satisfactory than straight adults, necessitating greater use of nontraditional health outlets. For example, Diamant et al.<sup>36</sup> found that 57% of straight women were very satisfied with their regular source of care compared to 45% of lesbians and 39% of bisexual women. Similarly, McNair et al.<sup>54</sup> found lesbians, compared to straight women, produced lower satisfaction ratings of their general provider, while Avery et al.<sup>55</sup> identified higher levels of dissatisfaction with mental health services among LGBT adults compared to straight adults. Eliason and Schope<sup>56</sup> found that sexual minority men provided less favorable provider ratings than their female counterparts, leading to lower levels of sexual identity disclosure. In addition, Stein and Bonuck<sup>57</sup> found 24% of male and 38% of female sexual minorities perceived their providers to be insensitive to LGB concerns. Several other studies have noted the scarcity of providers trained in and sensitive to LGB needs. 28,56-60 Not surprisingly, LGB adults often attribute their sexual orientation nondisclosure and/or delays in seeking care to fears of rejection, disrespect, and inappropriate treatment. 28,29,39,57 These issues may be compounded for bisexual men who report lower levels of self-disclosure and community connection relative to their gay peers.<sup>61</sup>

#### Limitations and future research

Even after pooling two years of data, the samples of sexual minorities remain small (especially for bisexual men), resulting in a few unreliable estimates and nonsignificant *P*-values, despite moderately sized associations. Second, the crosssectional nature of the data prohibits us from examining underlying causal mechanisms that may explain the sexual orientation—HIT use link. Finally, our models may be misspecified as we were unable to include measures of perceived stigma or quality of care.

Future research could explore the underlying reasons why LGB persons, especially sexual minority men, are more likely than straight adults to seek health information and care on the Internet. In addition to perceived stigma, this research could consider the quality of care received in traditional healthcare settings and the inability to find healthcare providers

competent in LGB needs as possible determinants of HIT use. Results of these studies would be useful for addressing shortcomings of the traditional U.S. healthcare system and may lead to improved healthcare delivery through the Internet and other nontraditional means.

# Conclusion

Gay and bisexual men were more likely to use HIT than their straight counterparts. Perceived stigma, concerns over homophobia, and greater dissatisfaction with care, likely limit the choice of providers available to these adults. Hence, Internet healthcare may be seen as a viable tool given its ability to provide anonymity and bridge large physical distances to match LGB patients with providers sensitive to their needs. However, unintended and potentially harmful consequences of delivering healthcare over the Internet should be considered, especially with regard to the consumption and exchange of health information. The provision of Internet healthcare to LGB persons who cannot find adequate support through traditional delivery mechanisms may simply shift the modality of lower quality health services these individuals receive. Furthermore, research has shown that disclosure of sexual orientation to a provider is associated with increased healthcare utilization and patient satisfaction. 28,62–66 If a goal of seeking healthcare over the Internet is to avoid disclosing one's sexual orientation, consumers may miss opportunities for healthcare providers to offer appropriate health education/counseling, perform targeted screening/treatment, and identify individual risks. 67,68

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Table 1.

CHARACTERISTICS, HEALTH STATUS MEASURES, HEALTHCARE ACCESS AND UTILIZATION MEASURES, AND FREQUENCY OF COMPUTER USE AMONG U.S. ADULTS AGED DESCRIPTIVE ESTIMATES OF HEALTH INFORMATION TECHNOLOGY MEASURES, SEXUAL ORIENTATION, SOCIODEMOGRAPHIC CHARACTERISTICS, SOCIOECONOMIC 18-64 YEARS, BY SEX

Variable	Men (n = 24,383), % (95% CI)	Women (n = $28,633$ ), % ( $95\%$ CI)
Health information technology		
Sought health information on the Internet	42.7 (41.81–43.66)	56.5 (55.62–57.35)
Used computers to fill a prescription	5.9 (5.54–6.36)	8.5 (8.07–8.95)
Used computers to schedule an appointment with a healthcare provider	5.8 (5.35–6.25)	8.7 (8.20–9.17)
Used computers to communicate with a healthcare provider by email	6.4 (5.93–6.87)	9.5 (8.97–9.95)
Sexual orientation		
Straight	97.6 (97.27–97.81)	97.2 (96.90–97.44)
Gay/lesbian	2.0 (1.77–2.30)	1.6 (1.45–1.86)
Bisexual	0.4 (0.34–0.54)	1.2 (1.01–1.37)
Sociodemographic characteristics		
Age (in years)		
18–24	16.0 (15.17–16.83)	15.3 (14.57–16.07)
25-44	42.2 (41.23–43.09)	42.0 (41.14–42.77)
45–64	41.9 (40.89–42.83)	42.7 (41.84–43.65)
Race/ethnicity		
Hispanic	17.5 (16.65–18.38)	16.3 (15.63–17.04)
Non-Hispanic white	63.5 (62.46–64.61)	62.6 (61.60–63.58)
Non-Hispanic black	11.4 (10.87–12.04)	13.0 (12.33–13.65)
Non-Hispanic other	7.5 (7.01–8.06)	8.1 (1.56–8.68)
Marital status		
Never married	28.2 (27.26–29.13)	24.4 (23.63–25.26)
Married/cohabitating	61.6 (60.61–62.53)	60.5 (59.60–61.33)
Divorced/separated/widowed	10.2 (9.78–10.72)	15.1 (14.62–15.58)
Not U.S. born	19.6 (18.81–20.48)	18.7 (18.01–19.48)
Neighborhood attachment		
Low	12.6 (12.03–13.24)	15.1 (14.48–15.74)

Variable	Men (n = 24,383), % (95% CI)	Women (n = $28,633$ ), % ( $95\%$ CI)
Medium	36.6 (35.70–37.44)	33.7 (32.87–34.53)
High	45.6 (44.60–46.55)	45.5 (44.51–46.49)
Unknown	5.2 (4.86–5.65)	5.7 (5.28–6.17)
U.S. region of residence		
Northeast	17.0 (16.09–18.02)	16.9 (15.94–17.83)
Midwest	23.4 (22.44–24.40)	22.2 (21.14–23.31)
South	36.1 (34.86–37.27)	38.0 (36.86–39.12)
West	23.5 (22.55–24.50)	23.0 (22.05–23.88)
Place of residence		
Large MSA	34.1 (32.60–35.70)	33.6 (32.07–35.23)
Small MSA	52.4 (50.58–54.20)	52.2 (50.30–54.02)
Not in MSA	13.5 (12.23–14.83)	14.2 (12.98–15.53)
Socioeconomic characteristics		
Education		
Less than high school	13.2 (12.50–13.84)	11.6 (11.06–12.17)
High school diploma/GED	27.2 (26.36–28.04)	22.9 (22.18–23.64)
Some college	30.4 (29.51–31.33)	33.9 (33.12–34.76)
Bachelor's degree or higher	29.2 (28.26–30.24)	31.6 (30.64–32.49)
Currently working	76.9 (76.00–77.71)	65.5 (64.65–66.24)
Poverty status		
Poor (<100% FPT)	12.9 (12.28–13.63)	16.1 (15.44–16.87)
Near poor (100% FPT <200%)	17.3 (16.64–18.05)	18.9 (18.21–19.63)
Not poor (200% FPT)	69.7 (68.72–70.71)	65.0 (63.96–65.93)
Health status		
Reported health status		
Poor/fair	10.4 (9.90–10.95)	11.4 (10.89–11.89)
Good	24.3 (23.52–25.08)	25.2 (24.52–25.93)
Very good/excellent	65.3 (66.19–64.40)	63.4 (62.57–64.22)
Activity limitation	11.3 (10.76–11.88)	11.9 (11.32–12.41)
Multiple chronic conditions	16.6 (15.97–17.32)	19.1 (18.44–19.70)
Serious psychological distress	3.1 (2.83–3.42)	4.2 (3.88–4.49)

Variable	Men (n = 24,383), % (95% CI)	Men $(n = 24,383)$ , % $(95\% \ CI)$ Women $(n = 28,633)$ , % $(95\% \ CI)$
Healthcare access and utilization		
Health insurance status		
Private coverage	66.6 (65.67–67.57)	65.5 (64.58–66.35)
Public/other coverage	13.5 (12.87–14.14)	18.0 (17.33–18.69)
Uninsured/no coverage	19.9 (19.15–20.63)	16.5 (15.89–17.18)
Surgery in past 12 months	8.6 (8.09–9.03)	12.1 (11.61–12.63)
No. of different providers seen in past 12 months		
0	27.4 (26.60–28.29)	16.9 (16.22–17.54)
1	29.8 (29.00–30.60)	27.4 (26.66–28.12)
2	21.9 (21.19–22.57)	25.4 (24.76–26.13)
3	20.9 (20.17–21.64)	30.3 (29.50–31.13)
Did not receive specific services due to cost	16.0 (15.37–16.71)	22.0 (21.25–22.86)
Delayed care for a noncost reason	7.7 (7.29–8.19)	11.3 (10.76–11.83)
Trouble finding a provider	4.0 (3.72–4.36)	6.2 (5.84–6.62)
No usual source of care when sick or need medical advice	23.0 (22.20–23.80)	13.6 (12.98–14.14)
Frequency of computer use		
Never/almost never	18.3 (17.48–19.07)	15.0 (14.40–15.59)
Some days/most days	18.8 (18.11–19.53)	18.6 (17.96–19.33)
Every day	62.9 (61.92–63.93)	66.4 (65.51–67.24)

Data: National Health Interview Survey, 2013-2014.

Percent distributions may not add up to 100.0% due to rounding.

CI, confidence interval; MSA, metropolitan statistical area; GED, general educational development high school equivalency diploma; FPT, federal poverty threshold

 Table 2.

 PREVALENCE OF HEALTH INFORMATION TECHNOLOGY USE AMONG U.S. Adults Aged 18–64, by Sexual Orientation and Sex

	Men % (95% CI)	Women % (95% CI)
Sought health inform	nation on the Internet	
Sexual orientation		
Straight	42.3 (41.31–43.19)	56.2 (55.35–57.10)
Gay/lesbian	62.0 <sup>a</sup> (55.53–68.01)	62.7 <sup>b</sup> (56.68–68.28)
Bisexual	62.5 <sup>a</sup> (52.05–71.96)	68.9 <sup>a</sup> (61.02–75.73)
	n = 24,362	n = 28,612
Used computers to f	ill a prescription	
Sexual orientation		
Straight	5.8 (5.35–6.18)	8.5 (8.05–8.95)
Gay/lesbian	14.5 <sup>a</sup> (11.00–18.76)	9.3 (6.62–12.82)
Bisexual	8.0 <sup>†</sup> (3.50–17.41)	8.2 (5.08–12.89)
	n = 24,374	n = 28,626
Used computers to s	chedule an appointment v	with a healthcare provider
Sexual orientation		
Straight	5.6 (5.15-6.03)	8.6 (8.16–9.13)
Gay/lesbian	14.0 <sup>a</sup> (10.39–18.65)	8.5 (6.07–11.67)
Bisexual	14.5 <sup>†</sup> (7.57–25.97)	12.4 (8.45–17.73)
	n = 24,374	n = 28,624
Used computers to c	ommunicate with a health	ncare provider by email
Sexual orientation		
Straight	6.1 (5.66–6.62)	9.4 (8.91–9.90)
Gay/lesbian	17.2 <sup>a</sup> (13.30–21.97)	9.8 (7.03–13.54)
Bisexual	14.9 <sup>†</sup> (7.83–26.44)	13.4 (8.97–19.54)
	n = 24,373	n = 28,076

Data: National Health Interview Survey, 2013-2014.

 $<sup>^{\</sup>dagger}$ Estimates marked with a dagger have a relative standard error >30.0% and 50.0% and should be interpreted with caution.

 $<sup>^</sup>aP\!<\!0.001$  for comparisons of ''gay/lesbian'' to ''straight'' and ''bisexual'' to ''straight.''

 $<sup>^</sup>bP\!<\!0.05$  for comparisons of ''gay/lesbian'' to ''straight'' and ''bisexual'' to ''straight.''

# Table 3.

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Initial model; (excludes the	ides the healthcare access and utilization measures) $AOR~(95\%~CI)$	Full model AOR (95% CI)
Sought health information on the Internet	net	
Men		
Straight	1.00 (ref.)	1.00 (ref.)
Gay	1.70 *** (1.27–2.26)	$1.50^{**}(1.12-2.01)$
Bisexual	1.95**(1.18-3.22)	1.61 (0.98–2.66)
	n = 23,829	n = 23,829
Women		
Straight	1.00 (ref.)	1.00 (ref.)
Gay/lesbian	1.11 (0.82–1.52)	1.15 (0.84–1.57)
Bisexual	1.68 *(1.11–2.54)	1.50 (1.00–2.27)
	n = 28,070	n = 28,070
Used computers to fill a prescription		
Men		
Straight	1.00 (ref.)	1.00 (ref.)
Gay	2.47 *** (1.70–3.58)	2.14 *** (1.47–3.12)
Bisexual	1.70 (0.68–4.26)	1.46 (0.59–3.62)
	n = 23,838	n = 23,838
Women		
Straight	1.00 (ref.)	1.00 (ref.)
Gay/lesbian	0.91 (0.63–1.33)	1.00 (0.68–1.47)
Bisexual	0.94 (0.54–1.62)	0.93 (0.53–1.63)
	n = 28,080	n = 28,080
Used computers to schedule an appointment with a healthcare provider	tment with a healthcare provider	
Men		
Straight	1.00 (ref.)	1.00 (ref.)
Gay	$2.30^{***}(1.55-3.40)$	1.97 *** (1.32–2.94)
Bisexual	2.83*(1.21–6.62)	2.43*(1.04–5.67)

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Initial m	Initial model; (excludes the healthcare access and utilization measures) AOR (95% CI) Full model AOR (95% CI)	Full model AOR (95% CI)
	n = 23,838	n = 23,838
Women		
Straight	1.00 (ref.)	1.00 (ref.)
Gay/lesbian	0.79 (0.54–1.15)	0.85 (0.59–1.22)
Bisexual	1.44 (0.90–2.31)	1.36 (0.83–2.23)
	n = 28,078	n = 28,078
Jsed computers to commu	Used computers to communicate with a healthcare provider by email	
Men		
Straight	1.00 (ref.)	1.00 (ref.)
Gay	2.50**** (1.68–3.72)	2.14 *** (1.41–3.25)
Bisexual	2.69*(1.23–5.88)	$2.26^*(1.02-5.02)$
	n = 23.837	n = 23,837
Women		
Straight	1.00 (ref.)	1.00 (ref.)
Gay/lesbian	0.83 (0.56–1.23)	0.90 (0.61–1.33)
Bisexual	1.57 (0.95–2.57)	1.54 (0.94–2.52)
	n = 28.079	n = 28.079

Data: National Health Interview Survey, 2013-2014.

psychological distress), healthcare access and utilization measures (health insurance status, the number of providers seen in the past 12 months, did not receive care in the past residence), socioeconomic characteristics (education, employment status, and poverty status), health status measures (reported health status, activity limitation, multiple chronic conditions, and serious The following covariates were included in the models: sociodemographic characteristics (age, race/ethnicity, marital status, nativity, neighborhood attachment, U.S. region of residence, and place of 12 months due to cost, delayed receiving care in the past 12 months due to noncost reasons, trouble finding a provider, and no usual source of care), and frequency of computer use.

 $^{***}_{P < 0.001}$ .  $^*_{\rm P}\,{<}\,0.05$  $^{**}_{\rm P}\,{<}\,0.01$ 

AOR, adjusted odds ratios; ref, reference category.