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## Identification of common genetic risk variants for autism spectrum disorder

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### Abstract

Autism spectrum disorder (ASD) is a highly heritable and heterogeneous group of neurodevelopmental phenotypes diagnosed in more than 1% of children. Common genetic variants contribute substantially to ASD susceptibility, but to date no individual variants have been robustly associated with ASD. With a marked sample size increase from a unique Danish population resource, we report a genome-wide association meta-analysis of 18,381 ASD cases and 27,969 controls that identifies five genome-wide significant loci. Leveraging GWAS results from three phenotypes with significantly overlapping genetic architectures (schizophrenia, major depression, and educational attainment), seven additional loci shared with other traits are identified at equally strict significance levels. Dissecting the polygenic architecture, we find both quantitative and qualitative polygenic heterogeneity across ASD subtypes. These results highlight biological insights, particularly relating to neuronal function and corticogenesis and establish that GWAS performed at scale will be much more productive in the near term in ASD.

### Editorial Summary

A genome-wide association meta-analysis of 18,381 autism spectrum disorder (ASD) cases and 27,969 controls identifies 5 risk loci. The authors find quantitative and qualitative polygenic heterogeneity across ASD subtypes.

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ASD is the term for a group of pervasive neurodevelopmental disorders characterized by impaired social and communication skills along with repetitive and restrictive behavior. The clinical presentation is very heterogeneous, including individuals with severe impairment and intellectual disability as well as individuals with above average IQ and high levels of

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Sample and/or data provider and processing: JG, SR, MM, RKW, EA, OAA, RA, RB, JDB, JBG, MBH, FC, KC, DD, ALD, JIG, CSH, MVH, CMH, JLM, AP, CBP, MGP, JBP, KRe, AR, ES, GDS, HS, CRS, PGC-ASD, BUPGEN, PGC-MDD, 23andMe, KS, DMH, OM, PBM, BMN, MJD, ADB

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#### Competing Interests Statement

Hreinn Stefansson, Kari Stefansson, Stacy Steinberg, and G. Bragi Walters are employees of deCODE genetics/Amgen. The 23andMe Research Team are employed by 23andMe. Daniel H Geschwind is a scientific advisor for Ovid Therapeutic, Falcon Computing and Axial Biotherapeutics. Thomas Werge has acted as scientific advisor and lecturer for H. Lundbeck A/S.

academic and occupational functioning. ASD affects 1–1.5% of individuals and is highly heritable, with both common and rare variants contributing to its etiology<sup>1–4</sup>. Common variants have been estimated to account for a major part of ASD liability<sup>2</sup> as has been observed for other common neuropsychiatric disorders. By contrast, *de novo* mutations, mostly copy number variants (CNVs) and gene disrupting point mutations, have larger individual effects, but collectively explain < 5% of the overall liability<sup>1–3</sup> and far less of the heritability. While a number of genes have been convincingly implicated via excess statistical aggregation of *de novo* mutations, the largest GWAS to date (n = 7,387 cases scanned) – while providing compelling evidence for the bulk contribution of common variants – did not conclusively identify single variants at genome-wide significance<sup>5–7</sup>. This underscored that common variants, as in other complex diseases such as schizophrenia, individually have low impact and that a substantial scale-up in sample numbers would be needed.

Here we report the first common risk variants robustly associated with ASD by more than doubling the discovery sample size compared to previous GWAS<sup>5–8</sup>. We describe strong genetic correlations between ASD and other complex disorders and traits, confirming shared etiology, and we show results indicating differences in the polygenic architecture across clinical sub-types of ASD. Leveraging these relationships and recently introduced computational techniques<sup>9</sup>, we identify additional novel ASD-associated variants that are shared with other phenotypes. Furthermore, by integrating with complementary data from Hi-C chromatin interaction analysis of fetal brains and brain transcriptome data, we explore the functional implications of our top-ranking GWAS results.

## Results

### GWAS

As part of the iPSYCH project<sup>10</sup>, we collected and genotyped a Danish nation-wide population-based case-cohort sample including nearly all individuals born in Denmark between 1981 and 2005 and diagnosed with ASD (according to ICD-10) before 2014. We randomly selected controls from the same birth cohorts (Supplementary Table 1). We have previously validated registry-based ASD diagnoses<sup>11,12</sup> and demonstrated the accuracy of genotyping DNA extracted and amplified from bloodspots collected shortly after birth<sup>13,14</sup>. Genotypes were processed using Ricopili<sup>15</sup>, performing stringent quality control of data, removal of related individuals, exclusion of ancestry outliers based on principal component analysis, and imputation using the 1000 Genomes Project phase 3 reference panel. After this processing, genotypes from 13,076 cases and 22,664 controls from the iPSYCH sample were included in the analysis. As is now standard in human complex trait genomics, our primary analysis was a meta-analysis of the iPSYCH ASD results with five family-based trio samples of European ancestry from the Psychiatric Genomics Consortium (PGC; 5,305 cases and 5,305 pseudo controls)<sup>16</sup>. All PGC samples had been processed using the same Ricopili pipeline for QC, imputation and analysis as employed here.

Supporting the consistency between the study designs, the iPSYCH population-based and PGC family-based analyses showed a high degree of genetic correlation with  $r_G = 0.779$  (SE = 0.106;  $P = 1.75 \times 10^{-13}$ ), similar to the genetic correlations observed between datasets in

other mental disorders<sup>17</sup>. Likewise, polygenicity as assessed by polygenic risk scores (PRS) showed consistency across the samples supporting homogeneity of effects across samples and study designs (see the results below regarding PRS on a five-way split of the sample). The SNP heritability ( $h_G^2$ ) was estimated to be 0.118 (SE = 0.010), for a population prevalence of 0.012<sup>18</sup>.

The main GWAS meta-analysis totaled 18,381 ASD cases and 27,969 controls, and applied an inverse variance-weighted fixed effects model. To ensure that the analysis was well-powered and robust, we examined markers with minor allele frequency (MAF)  $\geq 0.01$ , imputation INFO score  $\geq 0.7$ , and supported by an effective sample size in  $> 70\%$  of the total. This final meta-analysis included results for 9,112,387 autosomal markers and yielded 93 genome-wide significant markers in three separate loci (Figure 1; Table 1a; Supplementary Figures 1–44). Each locus was strongly supported by both the Danish case-control and the PGC family-based data. While modest inflation was observed ( $\lambda = 1.12$ ,  $\lambda_{1000} = 1.006$ ), LD score regression analysis<sup>19</sup> indicates this is arising from polygenicity ( $> 96\%$ , see Methods) rather than confounding. The strongest signal among 294,911 markers analyzed on chromosome X was  $P = 7.8 \times 10^{-5}$ .

We next obtained replication data for the top 88 loci with p-values  $< 1 \times 10^{-5}$  in five cohorts of European ancestry totaling 2,119 additional cases and 142,379 controls (Supplementary Table 2 and 3). An overall replication of direction of effects was observed (53 of 88 (60%) of  $P < 1 \times 10^{-5}$ ; 16 of 23 (70%) at  $P < 1 \times 10^{-6}$ ; sign tests  $P = 0.035$  and  $P = 0.047$ , respectively) and two additional loci achieved genome-wide significance in the combined analysis (Table 1a). More details on the identified loci can be found in Supplementary Table 4 and selected candidates are described in Box 1.

### Correlation with other traits and multi-trait GWAS

To investigate the extent of genetic overlap between ASD and other phenotypes we estimated the genetic correlations with a broad set of psychiatric and other medical diseases, disorders, and traits available at LD Hub<sup>65</sup> using bivariate LD score regression (Figure 2, Supplementary Table 5). Significant correlations were found for several traits including schizophrenia<sup>15</sup> ( $r_G = 0.211$ ,  $P = 1.03 \times 10^{-5}$ ) and measures of cognitive ability, especially educational attainment<sup>20</sup> ( $r_G = 0.199$ ,  $P = 2.56 \times 10^{-9}$ ), indicating a substantial genetic overlap with these phenotypes and corroborating previous reports<sup>5,66–68</sup>. In contrast to previous reports<sup>16</sup>, we found a strong and highly significant correlation with major depression<sup>21</sup> ( $r_G = 0.412$ ,  $P = 1.40 \times 10^{-25}$ ), and we report a prominent overlap with ADHD<sup>69</sup> ( $r_G = 0.360$ ,  $P = 1.24 \times 10^{-12}$ ). Moreover, we confirm the genetic correlation with social communication difficulties at age 8 in a non-ASD population sample reported previously based on a subset of the ASD sample<sup>70</sup> ( $r_G = 0.375$ ,  $P = 0.0028$ ).

In order to leverage these observations for the discovery of loci that may be shared between ASD and these other traits, we selected three particularly well-powered and genetically correlated phenotypes. These were schizophrenia (N = 79,641)<sup>15</sup>, major depression (N = 424,015)<sup>21</sup> and educational attainment (N = 328,917)<sup>20</sup>. We utilized the recently introduced MTAG method<sup>9</sup> which, briefly, generalizes the standard inverse-variance weighted meta-

analysis for multiple phenotypes. In this case, MTAG takes advantage of the fact that, given an overall genetic correlation between ASD and a second trait, the effect size estimate and evidence for association to ASD can be improved by appropriate use of the association information from the second trait. The results of these three ASD-anchored MTAG scans are correlated to the primary ASD scan (and to each other) but given the exploration of three scans, we utilized a more conservative threshold of  $1.67 \times 10^{-8}$  for declaring significance across these secondary scans giving an estimated maximum false discovery rate (maxFDR) of 0.021. In addition to stronger evidence for several of the ASD hits defined above, variants in seven additional regions achieved genome-wide significance, including three loci shared with educational attainment and four shared with major depression (Table 1b, Box 1, Supplementary Table 6, Supplementary Figures 49–55). We note that in these seven instances, the effect size estimate is stronger in ASD than the secondary trait, that the result is not characteristic of the strongest signals in these other scans (Supplementary Table 7–9) (and in fact 3 of these 7 were not significant in the secondary trait and constitute potentially novel findings). Moreover, we benchmarked against MTAG running two very large and heritable traits (height<sup>74</sup>,  $N = 252,288$ , and body mass index (BMI)<sup>24</sup>,  $N = 322,154$ ) with no expected links to ASD and no significant loci were added to the list of ASD-only significant associations.

### Gene and gene-set analysis

Next, we performed gene-based association analysis on our primary ASD meta-analysis using MAGMA<sup>75</sup>, testing for the joint association of all markers within a locus (across all protein-coding genes in the genome). This analysis identified 15 genes surpassing the significance threshold (Supplementary Table 10). As expected, the majority of these genes were located within the genome-wide significant loci identified in the GWAS, but seven genes are located in four additional loci including *KCNN2*, *MMP12*, *NTM* and a cluster of genes on chromosome 17 (*KANSL1*, *WNT3*, *MAPT* and *CRHRI*) (Supplementary Figures 57–71). In particular, *KCNN2* was strongly associated ( $P = 1.02 \times 10^{-9}$ ), far beyond even single-variant statistical thresholds and is included in the descriptions in Box 1.

Enrichment analyses using gene co-expression modules from human neocortex transcriptomic data (M13, M16 and M17 from Parikshak et al. 2013<sup>76</sup>) and loss-of-function intolerant genes (pLI > 0.9)<sup>77,78</sup>, which previously have shown evidence of enrichment in neurodevelopmental disorders<sup>69,76,79</sup>, yielded only nominal significance for the latter ( $P = 0.014$ ) and M16 ( $P = 0.050$ ) (Supplementary Table 11). Genes implicated in ASD by studies or rare variants in Sanders et al.<sup>56</sup> was just shy of showing nominally significant enrichment ( $P = 0.063$ ) while enrichment in the curated gene list from the SPARK consortium<sup>80</sup> was significant ( $P = 0.0034$ ). Likewise, analysis of Gene Ontology sets<sup>81,82</sup> for molecular function from MsigDB<sup>83</sup> showed no significant sets after Bonferroni correction for multiple testing (Supplementary Table 12).

### Dissection of the polygenic architecture

As ASD is a highly heterogeneous disorder, we explored how  $h_G^2$  partitioned across phenotypic sub-categories in the iPSYCH sample and estimated the genetic correlations between these groups using GCTA<sup>84</sup>. We examined cases with and without intellectual

disability (ID, N = 1,873) and the ICD-10 diagnostic sub-categories: childhood autism (F84.0, N = 3,310), atypical autism (F84.1, N = 1,607), Asperger's syndrome (F84.5, N = 4,622), and other/unspecified pervasive developmental disorders (PDD, F84.8–9, N = 5,795), reducing to non-overlapping groups when doing pairwise comparisons (see Supplementary Table 13). While the pairwise genetic correlations were consistently high between all subgroups (95% CIs including 1 in all comparisons), the  $h_G^2$  of Asperger's syndrome ( $h_G^2 = 0.097$ , SE = 0.001) was found to be twice the  $h_G^2$  of both childhood autism ( $h_G^2 = 0.049$ , SE = 0.009,  $P = 0.001$ ) and the group of other/unspecified PDD ( $h_G^2 = 0.045$ , SE = 0.008,  $P = 0.001$ ) (Supplementary Tables 14 and 15, Supplementary Figures 82 and 83). Similarly, the  $h_G^2$  of ASD without ID ( $h_G^2 = 0.086$ , SE = 0.005) was found to be three times higher than for cases with ID ( $h_G^2 = 0.029$ , SE = 0.013,  $P = 0.015$ ).

To further examine the apparent polygenic heterogeneity across subtypes, we investigated how PRS trained on different phenotypes were distributed across distinct ASD subgroups. We focused on phenotypes showing strong genetic correlation with ASD (e.g., educational attainment), but included also traits with little or no correlation to ASD (e.g., BMI) as negative controls. In this analysis, we regressed the normalized scores on ASD subgroups while including covariates for batches and principal components in a multivariate regression. Of the eight phenotypes we evaluated, only the cognitive phenotypes showed strong heterogeneity (educational attainment<sup>20</sup>  $P = 1.8 \times 10^{-8}$ , IQ<sup>23</sup>  $P = 3.7 \times 10^{-9}$ ) (Supplementary Figure 84). Interestingly, all case-control groups with or without intellectual disability showed significantly different loading for the two cognitive phenotypes: controls with intellectual disability have the lowest score followed by ASD cases with intellectual disability, and ASD cases without intellectual disability have significantly higher scores again than any other group ( $P = 2.6 \times 10^{-12}$  for educational attainment,  $P = 8.2 \times 10^{-12}$  for IQ).

With respect to the diagnostic sub-categories constructed hierarchically from ASD subtypes (Supplementary Table 13), it was again the cognitive phenotypes that showed the strongest heterogeneity across the diagnostic classes (educational attainment  $P = 2.6 \times 10^{-11}$ , IQ  $P = 3.4 \times 10^{-8}$ ), while neuroticism<sup>67</sup> ( $P = 0.0015$ ), chronotype<sup>73</sup> ( $P = 0.011$ ) and subjective well-being<sup>67</sup> ( $P = 0.029$ ) showed weaker but nominally significant degree of heterogeneity, and schizophrenia (SCZ), major depressive disorder (MD) and BMI<sup>24</sup> were non-significant across the groups ( $P > 0.19$ ) (Figure 3). This pattern weakened only slightly when excluding subjects with intellectual disability (Supplementary Figure 85). For neuroticism, there was a clear split with atypical and other/unspecified PDD cases having significantly higher PRS than childhood autism and Asperger's,  $P = 0.00013$ . Considering the genetic overlap of each subcategory with each phenotype, the hypothesis of homogeneity across sub-phenotypes was strongly rejected ( $P = 1.6 \times 10^{-11}$ ), thereby establishing that these sub-categories indeed have differences in their genetic architectures.

Focusing on educational attainment, significant enrichment of PRS was found for Asperger's syndrome ( $P = 2.0 \times 10^{-17}$ ) in particular, and for childhood autism ( $P = 1.5 \times 10^{-5}$ ), but not for the group of other/unspecified PDD ( $P = 0.36$ ) or for atypical autism ( $P =$

0.13) (Figure 3). Excluding individuals with intellectual disability only changes this marginally, with atypical autism becoming nominally significant ( $P = 0.020$ ) (Supplementary Figure 85). These results show that the genetic architecture underlying educational attainment is indeed shared with ASD but to a variable degree across the disorder spectrum. We find that the observed excess in ASD subjects of alleles positively associated with education attainment<sup>85,86</sup> is confined to Asperger's and childhood autism, and it is not seen here in atypical autism nor in other/unspecified PDD.

Finally, we evaluated the predictive ability of ASD PRS using five different sets of target and training samples within the combined iPSYCH-PGC sample. The observed mean variance explained by PRS (Nagelkerke's  $R^2$ ) was 2.45% ( $P = 5.58 \times 10^{-140}$ ) with a pooled PRS-based case-control odds ratio OR = 1.33 (CI.95% 1.30–1.36) (Supplementary Figures 89 and 91). Dividing the target samples into PRS decile groups revealed an increase in OR with increasing PRS. The OR for subjects with the highest PRS increased to OR = 2.80 (CI.95% 2.53–3.10) relative to the lowest decile (Figure 4a and Supplementary Figure 92). Leveraging correlated phenotypes in an attempt to improve prediction of ASD, we generated a multi-phenotype PRS as a weighted sum of phenotype specific PRS (see Methods). As expected, Nagelkerke's  $R^2$  increased for each PRS included attaining its maximum at the full model at 3.77% ( $P = 2.03 \times 10^{-215}$ ) for the pooled analysis with an OR = 3.57 (CI.95% 3.22–3.96) for the highest decile (Figure 4b and Supplementary Figure 93 and 94). These results demonstrate that an individual's ASD risk depends on the level of polygenic burden of thousands of common variants in a dose-dependent way, which can be reinforced by adding SNP-weights from ASD correlated traits.

### Functional annotation

In order to obtain information on possible biological underpinnings of our GWAS results we conducted several analyses. First, we examined how the ASD  $h_G^2$  partitioned on functional genomic categories as well as on cell type-specific regulatory elements using stratified LD score regression<sup>87</sup>. This analysis identified significant enrichment of heritability in conserved DNA regions and H3K4me1 histone marks<sup>88</sup>, as well as in genes expressed in central nervous system (CNS) cell types as a group (Supplementary Figures 95 and 96), which is in line with observations in schizophrenia<sup>15</sup>, major depression<sup>21</sup>, and bipolar disorder<sup>66</sup>. Analyzing the enhancer associated mark H3K4me1 in individual cell/tissue<sup>88</sup>, we found significant enrichment in brain and neuronal cell lines (Supplementary Figure 97). The highest enrichment was observed in the developing brain, germinal matrix, cortex-derived neurospheres, and embryonic stem cell (ESC)-derived neurons, consistent with ASD as a neurodevelopmental disorder with largely prenatal origins, as supported by data from analysis of rare *de novo* variants<sup>76</sup>.

Common variation in ASD is located in regions that are highly enriched with regulatory elements predicted to be active in human corticogenesis (Supplementary Figures 95–97). As most gene regulatory events occur at a distance via chromosome looping, we leveraged Hi-C data from germinal zone (GZ) and post-mitotic zones cortical plate (CP) in the developing fetal brain to identify potential target genes for these variants<sup>89</sup>. We performed fine mapping of 28 loci to identify the set of credible variants containing likely causal genetic risk<sup>90</sup> (see

Methods). Credible SNPs were significantly enriched with enhancer marks in the fetal brain (Supplementary Figure 98), again confirming the likely regulatory role of these SNPs during brain development.

Based on location or evidence of physical contact from Hi-C, the 380 credible SNPs (28 loci) could be assigned to 95 genes (40 protein-coding), including 39 SNPs within promoters that were assigned to 9 genes, and 16 SNPs within the protein coding sequence of 8 genes (Supplementary Table 16, Supplementary Figure 98). Hi-C identified 86 genes, which interacted with credible SNPs in either the CP or GZ during brain development. Among these genes, 34 are interacting with credible SNPs in both CP and GZ, which represent a high-confidence gene list. Notable examples are illustrated in Figure 5 and highlighted in Box 1. By analyzing their mean expression trajectory, we observed that the identified ASD candidate genes (Supplementary Table 16) show highest expression during fetal corticogenesis, which is in line with the enrichment of heritability in the regulatory elements in developing brain (Figure 5e–g). Interestingly, both common and rare variation in ASD preferentially affects genes expressed during corticogenesis<sup>76</sup>, highlighting a potential spatiotemporal convergence of genetic risk on this specific developmental epoch, despite the disorder's profound genetic heterogeneity.

## Discussion

The high heritability of ASD has been recognized for decades and remains among the highest for any complex disease despite many clinical diagnostic changes over the past 30–40 years resulting in a broader phenotype that characterizes more than 1% of the population. While early GWAS permitted estimates that common polygenic variation should explain a substantial fraction of the heritability of ASD, individually significant loci remained elusive. This was suspected to be due to limited sample size since studies of schizophrenia – with similar prevalence, heritability and reduced fitness – and major depression achieved striking results only when sample sizes five to ten times larger than available in ASD were employed. This study has finally borne out that expectation with definitively demonstrated significant “hits”.

Here we report the first common risk variants robustly associated with ASD by using unique Danish resources in conjunction with results of the earlier PGC data – more than tripling the previous largest discovery sample. Of these, five loci were defined in ASD alone, and seven additional suggested at a stricter threshold utilizing GWAS results from three correlated phenotypes (schizophrenia, depression and educational attainment) and a recently introduced analytic approach, MTAG. Both genome-wide LD score regression analysis and the fact that even among the loci defined in ASD alone there was additional evidence in these other trait scans indicate that the polygenic architecture of ASD is significantly shared with risk to adult psychiatric illness and higher educational attainment and intelligence. It should be noted that the MTAG analyses were carried out as three pairwise analyses. This way we avoid the complex interactions that could arise if we ran three or four correlated phenotypes at a time<sup>9</sup>. Indeed, what we get, despite the secondary summary statistics coming from large, high-powered studies, are relatively modest weights to the contributions from these statistics, because the genetic correlations are modest. The largest weight was

0.27 for schizophrenia, followed by 0.24 for major depression, and 0.11 for educational attainment. Moreover, the estimated worst case FDR was just 0.021 up just 0.001 from that of the ASD GWAS alone. Thus all loci identified by MTAG were found with an acceptable degree of certainty and have substantial contributions from ASD alone (Table 1a, b and Supplementary Table 6). Our expectation is that most or all such loci will likely be identified in future ASD-only GWAS as sample sizes are increased substantially – however, given how new these methods are, the precise phenotypic consequences of these particular variants awaits expansion of all these trait GWAS.

In most GWAS studies there has been little evidence of heterogeneity of association across phenotypic subgroups. In this study, however, we see strong heterogeneity of genetic overlap with other traits when our ASD samples are broken into distinct subsets. In particular, the excess of alleles associated with higher intelligence and educational attainment was only observed in the higher functioning categories (particularly Asperger's syndrome and individuals without comorbid intellectual disability) – and not in the other/unspecified PDD and intellectual disability categories. This is reminiscent, and logically inverted, from the much greater role of spontaneous mutations in these latter categories, particularly in genes known to have an even larger impact in cohorts ascertained for intellectual disability/developmental delay<sup>91</sup>. Interestingly, other/unspecified PDD and atypical autism also have a significantly higher PRS for neuroticism than childhood autism and Asperger's. These different enrichment profiles observed provide evidence for a heterogeneous and qualitatively different genetic architecture between sub-types of ASD, which should inform future studies aiming at identifying etiologies and disease mechanisms in ASD.

The strong differences in estimated SNP heritability between ASD cases with and without intellectual disability, and highest in Asperger's provide genetic evidence of longstanding observations. In particular, this aligns well with the observation that *de novo* variants are more frequently observed in ASD cases with intellectual disability compared to cases without comorbid intellectual disability, that IQ correlates positively with family history of psychiatric disorders<sup>92</sup> and that severe intellectual disability (encompassing many syndromes that confer high risk to ASD) show far less heritability than is observed for mild intellectual disability<sup>93</sup>, intelligence in general<sup>94</sup> and ASDs. Thus it is perhaps unsurprising that our data suggests that the contribution of common variants may be more prominent in high-functioning ASD cases such as Asperger's syndrome.

We further explored the functional implications of these results with complementary functional genomics data including Hi-C analyses of fetal brains and brain transcriptome data. Analyses at genome-wide scale (partitioned  $h_G^2$  (Supplementary Figures 95–97) and brain transcriptome enrichment (Figure 5e–g)) as well as at single loci (Figure 5a–d, Box 1) highlighted the involvement of processes relating to brain development and neuronal function. Notably, a number of genes located in the identified loci have previously been linked to ASD risk in studies of *de novo* and rare variants (Box 1, Supplementary Table 4), including *PTBP2*, *CADPS*, and *KMT2E*, which were found to interact with credible SNPs in the Hi-C analysis (*PTBP2*, *CADPS*) or contain a loss-of-function credible SNP (*KMT2E*). Interestingly, aberrant splicing of *CADPS*' sister gene *CADPS2*, which has almost identical



function, has been found in autism cases and *Cadps2* knockout mice display behavioral anomalies with translational relevance to autism<sup>95</sup>. *PTBP2* encodes a neuronal splicing factor and alterations in alternative splicing have been identified in brains from individuals diagnosed with ASD<sup>96</sup>.

In summary, we have established a first robust set of common variant associations in ASD and have begun laying the groundwork through which the biology of ASD and related phenotypes will inevitably be better articulated.

## Methods

### Subjects

**iPSYCH sample**—The iPSYCH ASD sample is a part of a population based case-cohort sample extracted from a baseline cohort<sup>10</sup> consisting of all children born in Denmark between May 1<sup>st</sup> 1981 and December 31<sup>st</sup> 2005. Eligible were singletons born to a known mother and resident in Denmark on their one-year birthday. Cases were identified from the Danish Psychiatric Central Research Register (DPCRR)<sup>12</sup>, which includes data on all individuals treated in Denmark at psychiatric hospitals (from 1969 onwards) as well as at outpatient psychiatric clinics (from 1995 onwards). Cases were diagnosed with ASD in 2013 or earlier by a psychiatrist according to ICD10, including diagnoses of childhood autism (ICD10 code F84.0), atypical autism (F84.1), Asperger's syndrome (F84.5), other pervasive developmental disorders (F84.8), and pervasive developmental disorder, unspecified (F84.9). As controls we selected a random sample from the set of eligible children excluding those with an ASD diagnosis by 2013.

The samples were linked using the unique personal identification number to the Danish Newborn Screening Biobank (DNSB) at Statens Serum Institute (SSI), where DNA was extracted from Guthrie cards and whole genome amplified in triplicates as described previously<sup>13,97</sup>. Genotyping was performed at the Broad Institute of Harvard and MIT (Cambridge, MA, USA) using the PsychChip array from Illumina (CA, San Diego, USA) according to the instructions of the manufacturer. Genotype calling of markers with minor allele frequency (MAF) > 0.01 was performed by merging callsets from GenCall<sup>98</sup> and Birdseed<sup>99</sup> while less frequent variants were called with zCall<sup>100</sup>. Genotyping and data processing was carried out in 23 waves.

All analyses of the iPSYCH sample and joint analyses with the PGC samples were performed at the secured national GenomeDK high performance-computing cluster in Denmark.

The study was approved by the Regional Scientific Ethics Committee in Denmark and the Danish Data Protection Agency.

**Psychiatric Genomic Consortium (PGC) samples**—In brief, five cohorts provided genotypes to the sample (N denoting the number of trios for which genotypes were available): The Geschwind Autism Center of Excellence (ACE; N = 391), the Autism Genome Project<sup>62</sup> (AGP; N = 2,272), the Autism Genetic Resource Exchange<sup>101,102</sup>

(AGRE; N = 974), the NIMH Repository, the Montreal<sup>103</sup>/Boston Collection (MONBOS; N = 1,396, and the Simons Simplex Collection<sup>104,105</sup>(SSC; N = 2,231). The trios were analyzed as cases and pseudo controls. A detailed description of the sample is available on the PGC web site and even more details are provided in Anney et al<sup>5</sup>. Analyses of the PGC genotypes were conducted on the computer cluster LISA at the Dutch HPC center SURFsara.

**Follow-up samples**—As follow-up for the loci with p-values less than  $10^{-6}$  we asked for look up in five samples of Nordic and Eastern European origin with altogether 2,119 cases and 142,379 controls: BUPGEN (Norway: 164 cases and 656 controls), PAGES (Sweden: 926 cases and 3,841 controls not part of the PGC sample above), the Finnish autism case-control study (Finland: 159 cases and 526 controls), deCODE (Iceland 574 cases and 136,968 controls; Eastern Europe: 296 cases and 388 controls). See Supplementary Note for details.

### Statistical analyses

All statistical tests are two-sided unless otherwise stated. Software versions etc. can be found in Life Sciences Reporting Summary.

**GWAS analysis**—RicoPili<sup>15</sup>, the pipeline developed by the Psychiatric Genomics Consortium (PGC) Statistical Analysis Group was used for quality control, imputation, principle component analysis (PCA) and primary association analysis. For details see the Supplementary Note. The data were processed separately in the 23 genotyping batches in the case of iPSYCH and separately for each study in the PGC sample. Phasing was achieved using SHAPEIT<sup>106</sup> and imputation done by IMPUTE2<sup>107,108</sup> with haplotypes from the 1000 Genomes Project, phase 3<sup>109</sup> (1kGP3) as reference.

After excluding regions of high linkage disequilibrium (LD)<sup>110</sup>, the genotypes were pruned down to a set of roughly 30k markers. See supplementary Note for details. Using PLINK's<sup>111</sup> identity by state analysis, pairs of subjects were identified with  $\hat{\pi} > 0.2$  and one subject of each such pair was excluded at random (with a preference for keeping cases). PCA was carried out using smartPCA<sup>112,113</sup>. In iPSYCH, a subsample of European ancestry was selected as an ellipsoid in the space of PC1–3 centred and scaled using the mean and eight standard deviation of the subsample whose parents and grandparents were all known to have been born in Denmark (N = 31,500). In the PGC sample the European (CEU) subset was chosen using a Euclidian distance measure weighted by the variance explained by each of the first 3 principal components. Individuals more distant than 10 standard deviations from the combined CEU and Toscani in Italy (TSI) HapMap reference populations were excluded. We conducted a secondary PCA on the remaining 13,076 cases and 22,664 controls to provide covariates for the association analyses. Numbers of subjects in the data generation flow for the iPSYCH sample are found in Supplementary Table 1.

Association analyses were done by applying PLINK 1.9 to the imputed dosage data (the sum of imputation probabilities  $P(A1A2) + 2P(A1A1)$ ). In iPSYCH we included the first four principal components (PCs) as covariates as well as any PC beyond that, which were significantly associated with ASD in the sample, while the case-pseudo-controls from the

PGC trios required no PC covariates. Combined results for iPSYCH and for iPSYCH with the PGC was achieved by meta-analysing batch-wise and study-wise results using METAL<sup>114</sup> (July 2010 version) employing an inverse variance weighted fixed effect model<sup>115</sup>. On chromosome X males and females were analyzed separately and then meta-analyzed together. Subsequently we applied a quality filter allowing only markers with an imputation info score  $\geq 0.7$ , MAF  $\geq 0.01$  and an effective sample size (see Supplementary Note) of at least 70% of the study maximum. The degree to which the deviation in the test statistics can be ascribed to cryptic relatedness and population stratification rather than to polygenicity was measured from the intercept in LD score regression<sup>19</sup> (LDSC) as the ratio of (intercept-1) and  $(\text{mean}(\chi^2)-1)$ .

MTAG<sup>9</sup> was applied with standard settings. The iPSYCH-PGC meta-analysis summary statistics were paired up with the summary statistics for each of major depression<sup>21</sup> (excluding the Danish sampled but including summary statistics from 23andMe<sup>22</sup>; 111,902 cases, 312,113 controls, and mean  $\chi^2 = 1.477$ ), schizophrenia<sup>15</sup> (also excluding the Danish samples; 34,129 cases, 45,512 controls, and mean  $\chi^2 = 1.804$ ) and educational attainment<sup>20</sup> (328,917 samples and mean  $\chi^2 = 1.648$ ). These are studies that have considerably more statistical power than the ASD scan, but the genetic correlations are modest in the context of MTAG, so the weights ascribed to the secondary phenotypes in the MTAG analyses remain relatively low (no higher than 0.27). The maximum FDR was estimated as recommended in the MTAG paper<sup>9</sup>. See Supplementary Note for details. The results were clumped and we highlighted loci of interest by selecting those that were significant at  $5 \times 10^{-8}$  in the iPSYCH-PGC meta-analysis or the meta-analysis with the follow-up sample or were significant at  $1.67 \times 10^{-8}$  in any of the three MTAG analyses. The composite GWAS consisting of the minimal p-values at each marker over these five analyses was used as a background when creating Manhattan plots for the different analyses showing both what is maximally achieved and what the individual analysis contributes to that.

**Gene-based association and gene-set analyses.**—MAGMA 1.06<sup>75</sup> was applied to the ASD GWAS summary statistics to test for gene-based association. Using NCBI 37.3 gene definitions and restricting the analysis to SNPs located within the transcribed region, mean SNP association was tested with the sum of  $-\log(\text{SNP p-value})$  as test statistic. The resulting gene-based p-values were further used in competitive gene-set enrichment analyses in MAGMA. One analysis explored the candidate sets M13, M16 and M17 from Parikshak et al. 2013<sup>76</sup>, constrained, loss-of-function intolerant genes ( $pLI > 0.9$ )<sup>77,78</sup> derived from data of the Exome Aggregation Consortium (see Supplementary Note for details), as well as gene sets found in studies of rare variants in autism by Sanders et al.<sup>56</sup> and the curated gene list from the SPARK consortium<sup>80</sup>. Another was an agnostic analysis of the Gene Ontology sets<sup>81,82</sup> for molecular function from MsigDB 6.0<sup>83</sup>. We analyzed only genes outside the broad MHC region (hg19:chr6:25–35M) and included only gene sets with 10–1,000 genes. The gene set from Sanders et al. and SPARK include only one gene in MHC and was exempt from the MHC exclusion to be as true to the set as possible. All gene sets with significant enrichment were inspected to ensure that the signal was not driven by one or a few associated loci with multiple genes in close LD.

**SNP heritability**—SNP heritability,  $h_G^2$ , was estimated using LDSC<sup>19</sup> for the full ASD GWAS sample and GCTA<sup>84,116,117</sup> for subsamples too small for LDSC. For LDSC we used precomputed LD scores based on the European ancestry samples of the 1000 Genomes Project<sup>118</sup> restricted to HapMap3<sup>119</sup> SNPs. The summary stats with standard LDSC filtering were regressed onto these scores. For liability scale estimates, we used a population prevalence for Denmark of 1.22%<sup>18</sup>. Lacking proper prevalence estimates for subtypes, we scaled the full spectrum prevalence based on the composition of the case sample.

For subsamples too small for LDSC, the GREML approach of GCTA<sup>84,116,117</sup> was used. On best guess genotypes (genotype probability > 0.8, missing rate < 0.01 and MAF > 0.05) with INDELs removed, a genetic relatedness matrix (GRM) was fitted for the association sample (i.e. the subjects of European ancestry with  $\hat{\pi} \leq 0.2$ ) providing a relatedness estimate for all pairwise combinations of individuals. Estimation of the phenotypic variance explained by the SNPs (REML) was performed including PC 1–4 as continuous covariates together with any other PC that was nominally significantly associated to the phenotype as well as batches as categorical indicator covariates. Testing equal heritability for non-overlapping groups was done by permutation test (with 1000 permutations) keeping the controls and randomly assigning the different case labels.

Following Finucane et al.<sup>87</sup>, we conducted an enrichment analysis of the heritability for SNPs for functional annotation and for SNPs located in cell-type-specific regulatory elements. Using first the same 24 overlapping functional annotations (stripped down from 53) as in Finucane et al. we regressed the  $\chi^2$  from the ASD GWAS summary statistics on to the cell-type specific LD scores download from the site mentioned above with baseline scores, regression weights and allele frequencies based on European ancestry 1000 Genome Project data. The enrichment of a category was defined as the proportion of SNP heritability in the category divided by the proportion of SNPs in that category. Still following Finucane et al. we did a similar analysis using 220 cell type-specific annotations divided into 10 overlapping groups. In addition to this, we conducted an analysis based on annotation derived from data on H3K4Me1 imputed gapped peaks data from the Roadmap Epigenomics Mapping Consortium<sup>120</sup>; more specifically information excluding the broad MHC-region (chr6:25–35MB).

**Genetic correlation**—For the main ASD samples, SNP correlations,  $r_G$  were estimated using LDSC<sup>19</sup> and for the analysis of ASD subtypes and subgroups where the sample sizes were generally small, we used GCTA<sup>84</sup>. In both cases, we followed the same procedures as explained above. For all but a few phenotypes, LDSC estimates of correlation were achieved by upload to LD hub<sup>65</sup> for comparison to all together 234 phenotypes.

**Polygenic risk scores**—For the polygenic risk scores (PRS) we clumped the summary stats applying standard Ricopili parameters (see Supplementary Note for details). To avoid potential strand conflicts, we excluded all ambiguous markers for summary statistics not generated by Ricopili using the same imputation reference. PRS were generated at the default p-value thresholds (5e-8, 1e-6, 1e-4, 0.001, 0.01, 0.05, 0.1, 0.2, 0.5 and 1) as a weighted sum of the allele dosages in the ASD GWAS sample. Summing over the markers

abiding by the p-value threshold in the training set and weighing by the additive scale effect measure of the marker ( $\log(\text{OR})$  or  $\beta$ ) as estimated in the training set. Scores were normalized prior to analysis.

We evaluated the predictive power using Nagelkerke's  $R^2$  and plots of odds ratios and confidence intervals over score deciles. Both  $R^2$  and odds ratios were estimated in regression analyses including the relevant PCs and indicator variables for genotyping waves.

Lacking a large ASD sample outside of iPSYCH and PGC, we trained a set of PRS for ASD internally in the following way. We divided the sample in five subsamples of roughly equal size respecting the division into batches. We then ran five GWAS leaving out each group in turn from the training set and meta-analyzed these with the PGC results. This produced a set of PRS for each of the five subsamples trained on their complement. Prior to analyses, each score was normalized on the group where it was defined. We evaluated the predictive power in each group and on the whole sample combined.

To exploit the genetic overlap with other phenotypes to improve prediction, we created a series of new PRS by adding to the internally trained ASD score the PRS of other highly correlated phenotypes in a weighted sum. See supplementary info for details.

To analyze ASD subtypes in relation to PRS we defined a hierarchical set of phenotypes in the following way: First hierarchical subtypes was childhood autism, hierarchical atypical autism was defined as everybody with atypical autism and no childhood autism diagnosis, hierarchical Asperger's as everybody with an Asperger's diagnosis and neither childhood autism nor atypical autism. Finally, we lumped other pervasive developmental disorders and pervasive developmental disorder, unspecified into pervasive disorders developmental mixed, and the hierarchical version of that consists of everybody with such a diagnosis and none of the preceding ones (Supplementary Table 13). We examined the distribution over the distinct ASD subtypes of PRS for a number of phenotypes showing high  $r_G$  with ASD (as well as a few with low  $r_G$  as negative controls), by doing multivariate regression of the scores on the subtypes while adjusting for relevant PCs and wave indicator variables in a linear regression. See Supplementary Note for details.

**Hi-C analysis**—The Hi-C data were generated from two major cortical laminae: the germinal zone (GZ), containing primarily mitotically active neural progenitors, and the cortical and subcortical plate (CP), consisting primarily of post-mitotic neurons<sup>89</sup>. We first derived a set of credible SNPs (putative causal SNPs) from the identified top ranking loci in the ASD GWAS using CAVIAR<sup>90</sup>. The 30 loci showing the strongest association was intersected with the Hi-C reference data resulting in 28 loci for analysis. To test whether credible SNPs are enriched in active marks in the fetal brain<sup>120</sup>, we employed GREAT as previously described<sup>89,121</sup>. Credible SNPs were sub-grouped into those without known function (unannotated) and functionally annotated SNPs (SNPs in the gene promoters and SNPs that cause nonsynonymous variants) (Supplementary Figure 98). Then we integrated unannotated credible SNPs with chromatin contact profiles during fetal corticogenesis<sup>89</sup>, defining genes physically interacting with intergenic or intronic SNPs (Supplementary Figure 98).

The spatiotemporal transcriptomic atlas of human brain was obtained from Kang et al<sup>122</sup>. We used transcriptomic profiles of multiple brain regions with developmental epochs that span prenatal (6–37 post-conception week, PCW) and postnatal (4 months–42 years) periods. Expression values were log-transformed and centered to the mean expression level for each sample using a  $scale(center=T, scale=F)+1$  function in R. ASD candidate genes identified by Hi-C analyses (Supplementary Figure 98) were selected for each sample and their average centered expression values were calculated and plotted.

## Availability of summary statistics

The summary statistics are available for download the iPSYCH and at the PGC download sites (see the URL section).

## Availability of genotype data

For access to genotypes from the PGC samples and the iPSYCH sample, researchers should contact the lead PIs Mark J. Daly and Anders D. Børglum for PGC-ASD and iPSYCH-ASD respectively.

## URLs

The GenomeDK high performance-computing cluster in Denmark, <https://genome.au.dk>; the iPSYCH project, <http://ipsych.au.dk>, the iPSYCH download page, <http://ipsych.au.dk/downloads/>; the NIMH Repository, [https://www.nimhgenetics.org/available\\_data/autism/](https://www.nimhgenetics.org/available_data/autism/); the PGC download site, <https://www.med.unc.edu/pgc/results-and-downloads>; the LISA cluster at SURFsara, <https://userinfo.surfsara.nl/systems/lisa>; plink 1.9, [www.cog-genomics.org/plink/1.9/](http://www.cog-genomics.org/plink/1.9/); LDSC and associated files, <https://github.com/bulik/ldsc>; LD hub, <http://ldsc.broadinstitute.org/ldhub/>; GTExportal, <https://gtexportal.org/home/>

## Supplementary Material

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**Box1.**

**Selected loci and candidates (ordered by chromosome).**

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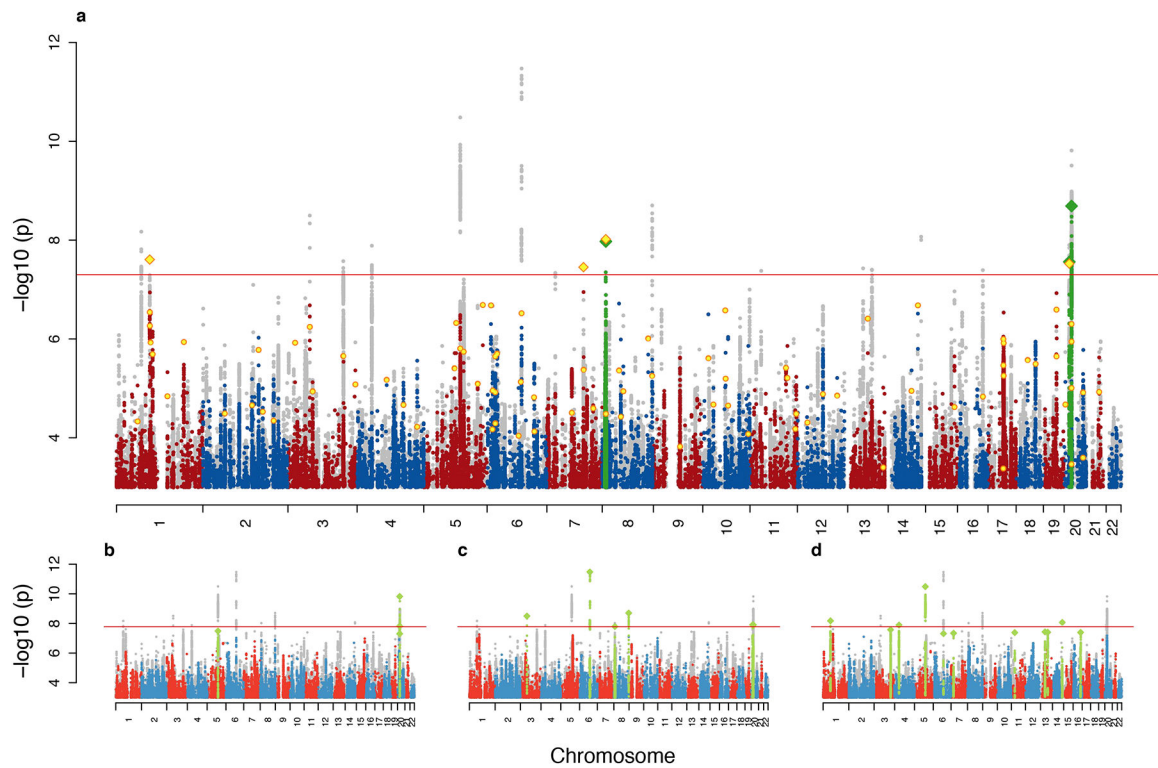
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Gene	Locus* and supporting evidence	Gene function
<i>NEGR1</i>	Chr1:72,729,142 Shared ASD-MDD locus Locus also significant in depression <sup>21,22</sup> , educational attainment <sup>60</sup> , intelligence <sup>23</sup> , obesity and BMI <sup>24-28</sup> and in an ASD-schizophrenia meta-analysis <sup>5</sup> . <i>NEGR1</i> is the only protein-coding gene in the locus <i>NEGR1</i> is supported by brain Hi-C and eQTL analyses <sup>21</sup>	<i>NEGR1</i> (neuronal growth regulator 1) is an adhesion molecule modulating synapse formation in hippocampal neurons <sup>29,30</sup> and neurite outgrowth <sup>31,32</sup> . It is a member of the IgLON protein family implicated in synaptic plasticity and axon extension <sup>33-35</sup> . Predominantly expressed (and developmentally upregulated) in hippocampus and cortex <sup>36</sup> and also hypothalamus <sup>37</sup> .
<i>PTBP2</i>	Chr1:96,561,801 ASD locus Locus also significant in BMI <sup>24,25,27</sup> weight <sup>25</sup> and educational attainment <sup>20</sup> . In schizophrenia, the locus shows a p-value of 6.5 × 10 <sup>-6</sup> –15 <i>PTBP2</i> is the nearest protein-coding gene, approx. 625 kb from index SNP. De novo and rare variants in <i>PTBP2</i> have been reported in ASD cases <sup>1,3,38</sup> . <i>PTBP2</i> is supported by Hi-C results in this study (Fig. 5d)	<i>PTBP2</i> is also known as nPTB (neuronal PTB) or bnPTB (brain PTB) and is a splicing regulator; <i>PTBP1</i> and its paralog <i>PTBP2</i> bind to intronic polypyrimidine tracts in pre-mRNAs and target large sets of exons to coordinate alternative splicing programs during development <sup>39</sup> . Several switches in the expression of <i>PTBP1</i> and <i>PTBP2</i> regulate alternative splicing during neurogenesis and neuronal differentiation <sup>40-43</sup> .
<i>CADPS</i>	Chr3:62,481,063 Shared ASD-Educational attainment locus Locus also significant in study of cognitive decline rate <sup>44</sup> <i>CADPS</i> is supported by Hi-C results in this study (Fig. 5a).	<i>CADPS</i> encodes a calcium-binding protein involved in exocytosis of neurotransmitters and neuropeptides. In line with <i>CADPS</i> mRNA being mainly expressed in brain and pituitary (GTEx portal- see URLs), immunoreactive <i>CADPS</i> -I is localized in neural and various endocrine tissues <sup>45</sup> . In hippocampal synapses, <i>CADPS</i> regulates the pool of readily releasable vesicles at pre-synaptic terminals <sup>46,47</sup> .
<i>KCNN2</i>	Chr5: 113,801,423 ASD locus (gene-wise analysis) Locus also significant in educational attainment <sup>20,48</sup> . <i>KCNN2</i> synaptic levels are regulated by the E3 ubiquitin ligase UBE3A <sup>49</sup> , of which overexpression has been linked to ASD risk <sup>49,50</sup> .	<i>KCNN2</i> is a voltage-independent Ca <sup>2+</sup> -activated K <sup>+</sup> channel that responds to changes in intracellular calcium concentration and couples calcium metabolism to potassium flux and membrane excitability. In CNS neurons, activation of <i>KCNN2</i> modulates neuronal excitability by causing membrane hyperpolarization <sup>51</sup> . Hippocampal <i>KCNN2</i> has roles in the formation of new memory <sup>52</sup> , encoding and consolidation of contextual fear <sup>53</sup> , and in drug-induced plasticity <sup>54</sup> .
<i>KMT2E</i>	Chr7:104,744,219 ASD locus Locus also significant in schizophrenia <sup>15,55</sup> and in ASD-schizophrenia meta-analysis <sup>5</sup> . <i>KMT2E</i> de novo mutations are associated with ASD at FDR < 0.1 <sup>56</sup> A <i>KMT2E</i> credible SNP is a loss-of-function variant (Supplementary Table 16)	<i>KMT2E</i> encodes Histone-lysine N-methyltransferase 2E and forms a family together with SETD5 <sup>57,58</sup> . Evidence suggest that recognition of the histone H3K4me3 mark by the <i>KMT2E</i> PHD finger can facilitate the recruitment of <i>KMT2E</i> to transcription-active chromatin regions <sup>59,60</sup> . <i>KMT2E</i> has been implicated in chromatin regulation, control of cell cycle progression, and maintaining genomic stability <sup>61</sup> .
<i>MACROD2</i>	Chr20: 14836243 ASD locus Locus found significant in previous ASD GWAS <sup>62</sup> but not supported in larger study <sup>63</sup> <i>MACROD2</i> is the only protein-coding gene in the locus	<i>MACROD2</i> is a nuclear enzyme that binds to mono-ADP-ribosylated (MARylated) proteins and functions as an eraser of mono-ADP-ribosylation <sup>64</sup> . Intracellular MARylated histones and GSK3 $\beta$ are substrates of <i>MACROD2</i> , and the removal of MAR from GSK3 $\beta$ is responsible for reactivating of its kinase activity <sup>64</sup> . This gene is expressed in lung and multiple regions of the brain. Low or no expression across most other tissue (GTEx portal- see URLs).

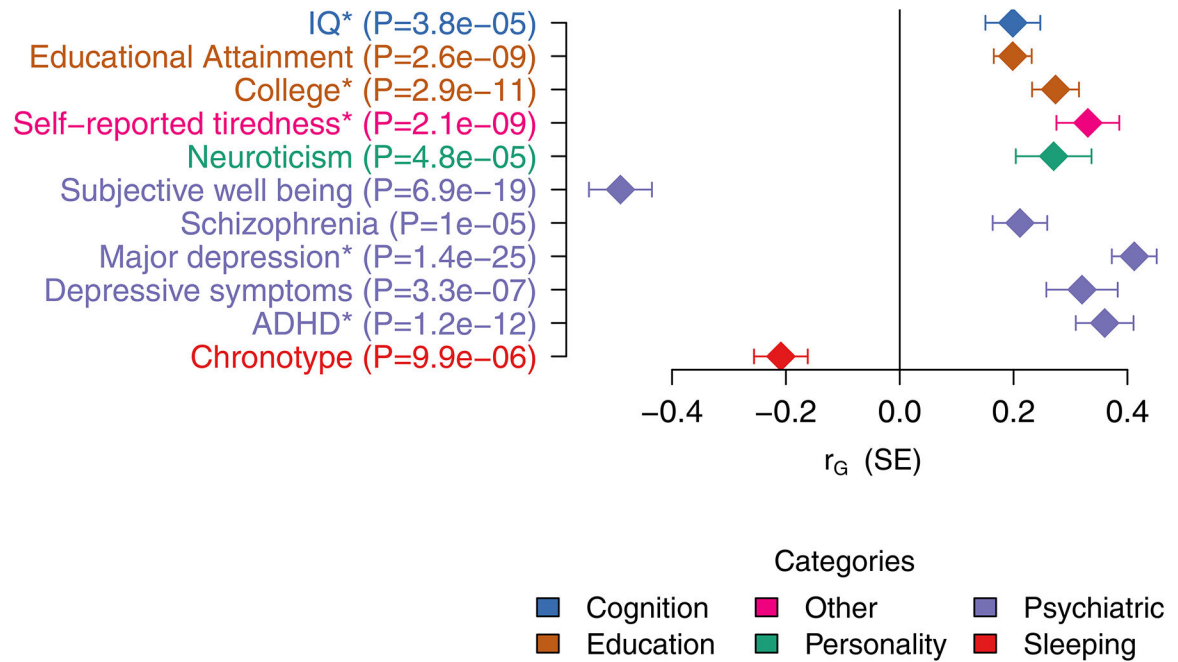
\* position of index SNP is listed.



**Figure 1. Manhattans plots:**

with the x axis showing genomic position (chromosomes 1–22) and the y axis showing statistical significance as  $-\log_{10}(P)$  of z statistics. **a:** The main ASD scan (18,381 cases and 27,969 controls) with the results of the combined analysis with the follow-up sample (2,119 cases and 142,379 controls) in yellow in the foreground. Genome-wide significant clumps are painted green with index SNPs as diamonds. **b-d:** Manhattan plots for three MTAG scans of ASD together with, respectively, schizophrenia<sup>15</sup> (34,129 cases and 45,512 controls), educational attainment<sup>20</sup> ( $N=328,917$ ) and major depression<sup>21</sup> (111,902 case and 312,113 controls). See Supplementary Figures 45–48 for full size plots. In all panels the results of the composite of the five analyses (consisting for each marker of the minimal p-value of the five) is shown in grey in the background.



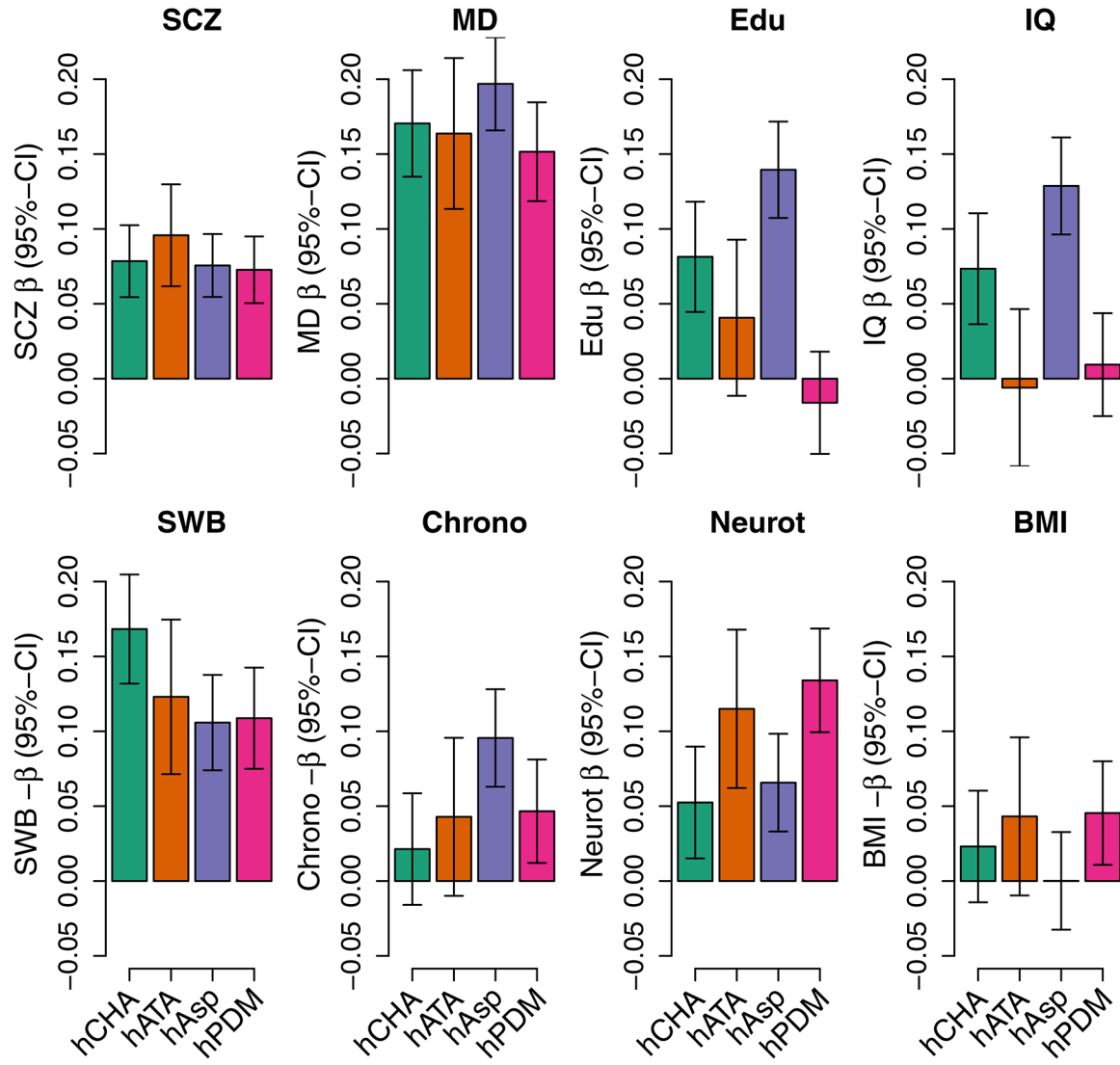


**Figure 2. Genetic correlation with other traits.**

Significant genetic correlations between ASD ( $N = 46,350$ ) and other traits after Bonferroni correction for testing a total of 234 traits available at LDhub with the addition of a handful of new phenotypes. Estimates and tests by LDSC<sup>19</sup>. The results here correspond to the following GWAS analyses: IQ<sup>23</sup> ( $N = 78,308$ ), educational attainment<sup>20</sup> ( $N = 328,917$ ), college<sup>71</sup> ( $N = 111,114$ ), self-reported tiredness<sup>72</sup> ( $N = 108,976$ ), neuroticism<sup>67</sup> ( $N = 170,911$ ), subjective well-being<sup>67</sup> ( $N = 298,420$ ), schizophrenia<sup>15</sup> ( $N = 82,315$ ), major depression<sup>21</sup> ( $N = 480,359$ ), depressive symptoms<sup>67</sup> ( $N = 161,460$ ), attention deficit/hyperactivity disorder (ADHD)<sup>69</sup> ( $N = 53,293$ ), and chronotype<sup>73</sup> ( $N = 128,266$ ). See Supplementary Table 5 for the full output of this analysis.

\* Indicates that the values are from in-house analyses of new summary statistics not yet included in LD Hub.

Multivariate regression of nRPS over Hierarchical Subtypes subtypes



**Figure 3. Profiling PRS load across distinct ASD sub-groups for 8 different phenotypes** (schizophrenia (SCZ)<sup>15</sup>, major depression (MD)<sup>21</sup>, educational attainment (Edu)<sup>20</sup>, human intelligence (IQ)<sup>23</sup>, subjective well-being (SWB)<sup>67</sup>, chronotype<sup>73</sup>, neuroticism<sup>67</sup> and body mass index (BMI)<sup>24</sup>. The bars show coefficients from multivariate multivariable regression of the 8 normalized scores on the distinct ASD sub-types of 13,076 cases and 22,664 controls, adjusting for batches and principal components. The subtypes are the hierarchically defined subtypes for childhood autism (hCHA, *N*= 3,310), atypical autism (hATA, *N*= 1,494), Asperger’s (hAsp, *N*= 4,417), and the lumped pervasive disorders developmental group (hPDM, *N*= 3,855). Please note that the orientation of the scores for subjective well-being, chronotype and BMI have been switched to improve graphical presentation. The corresponding plot where subjects with intellectual disability have been excluded can be seen in Supplementary Figure 85, and with intellectual disability as a subtype in Supplementary Figure 84. Applying the same procedure to the internally trained ASD score

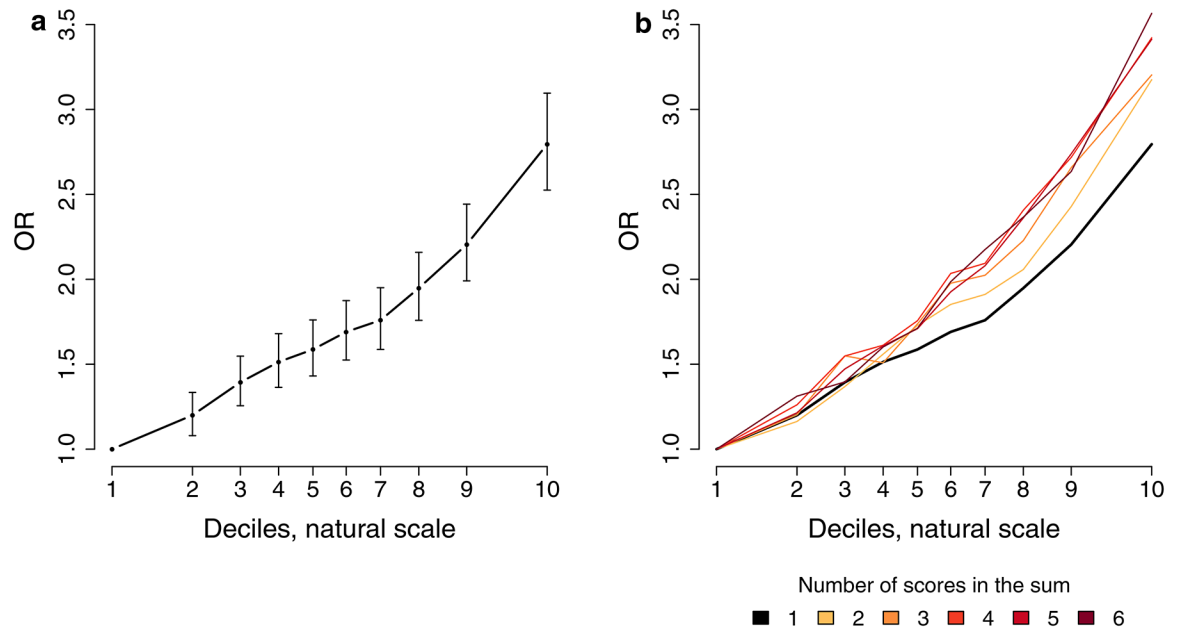
did not display systematic heterogeneity ( $P=0.068$ ) except as expected for the ID groups ( $P=0.00027$ ) (Supplementary Figure 88). Linear hypotheses tested using the Pillai test.

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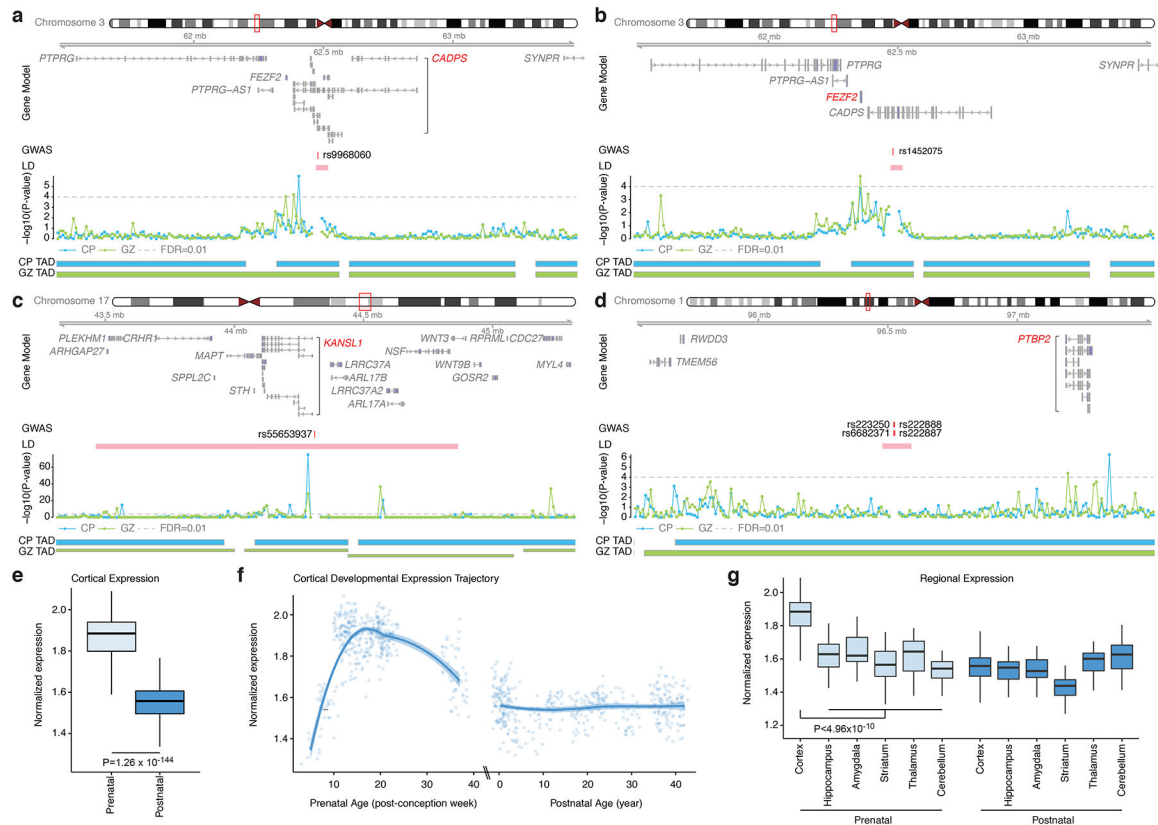


**Figure 4. Decile plots**

(Odds Ratio (OR) by PRS within each decile for 13,076 cases and 22,664 controls): **a.**

Decile plot with 95%-CI for the internally trained ASD score (P-value threshold is 0.1). **b.**

Decile plots on a weighted sums of PRSs starting with the ASD score of panel a and successively adding the scores for major depression<sup>21</sup>, subjective well-being<sup>67</sup>, schizophrenia<sup>15</sup>, educational attainment<sup>20</sup>, and chronotype<sup>73</sup>. In all instances the P-value threshold for the score used is the one with the highest Nagelkerke's  $R^2$ . Supplementary Figures 92 and 94 show the stability across leave-one out groups that was used to create these combined results.



**Figure 5. Chromatin interactions identify putative target genes of ASD loci.**

**a-d.** Chromatin interaction maps of credible SNPs to the 1 Mb flanking region, providing putative candidate genes that physically interact with credible SNPs. Gene Model is based on Gencode v19 and putative target genes are marked in red; Genomic coordinate for a credible SNP is labeled as GWAS;  $-\log_{10}(P\text{-value})$ , significance of the interaction between a SNP and each 10-kb bin, grey dotted line for FDR = 0.01 (one-sided significance test calculated as the probability of observing a higher contact frequency under the fitted Weibull distribution matched by chromosome and distance); Topologically associated domain (TAD) borders in cortical plate (CP) and germinal zone (GZ). **e-f.** Developmental expression trajectories of ASD candidate genes show highest expression in prenatal periods. Significance by t-test ( $N = 410$  and  $453$  for prenatal and postnatal samples, respectively). Box-plots showing median, interquartile range (IQR) with whiskers adding IQR to the 1<sup>st</sup> and 3<sup>rd</sup> quartile (**e** and **g**). LOESS smooth curve plotted with actual data points (**f** **g**). ASD candidate genes are highly expressed in the developing cortex as compared to other brain regions. One-way ANOVA and posthoc Tukey test, FDR-corrected. ( $N = 410/453$ , 39/36, 33/37, 48/34, 37/36, 32/39 for prenatal/postnatal cortex, hippocampus, amygdala, striatum, thalamus, and cerebellum, respectively).

**Table 1.**  
**Genome-wide significant loci from ASD scans and MTAG analyses.**

Panel **a** shows the loci reaching genome-wide significance in analysis of the ASD phenotype alone. The column “Analysis” indicates the minimum p-value arising from the original scan (ASD) and the combined analysis with the follow-up sample (Comb ASD). The columns “Support from other scans” list the other analyses (including MTAG) that further support the locus at genome-wide significance. For the ASD scan results, this shows the genome-wide significant results in the locus from the other scans, and for Comb ASD it displays the result from ASD. Panel **b** presents additional genome-wide significant loci identified in the three MTAG analyses. The three analyses are ASD with schizophrenia (SCZ)<sup>15</sup>, educational attainment (Edu)<sup>20</sup> and major depression (MD)<sup>21</sup>. Here the “Analysis” column points to which MTAG analysis gave the results (ASD-Edu or ASD-MD), and the columns “Support from other scans” provide the corresponding scan results in ASD alone. In both panels, independent loci are defined to have  $r^2 < 0.1$  and distance  $> 400$  kb and the index variant is displayed in the column “Index var”. Other columns are: chromosome (CHR), chromosomal position (BP), alleles (A1/A2), allele frequency of A1 (FRQ), estimate of effect ( $\beta$ ) with respect to A1, standard error of  $\beta$  (SE), and the association p-value of the index variant (P). The column “Nearest genes” lists nearest genes from within 50 kb of the region spanned by all SNPs with  $r^2 \geq 0.6$  to the index variant.

\*Indicate different lead SNP than the index variant

Index var	CHR	BP	Analysis	P	$\beta$	SE	A1/A2	FRQ	Support from other scans		Nearest genes	
									Scan	P		
<b>a</b>	rs910805	20	21248116	ASD	$2.04 \times 10^{-9}$	-0.096	A/G	0.76	ASD-SCZ	$1.5 \times 10^{-10}$	-0.069	KIZ, XRN2, NKX2-2, NKX2-4
	rs10099100	8	10576775	ASD	$1.07 \times 10^{-8}$	0.084	C/G	0.331	ASD-Edu*	$2.0 \times 10^{-8}$	-0.061	
	rs201910565	1	96561801	Comb ASD	$2.48 \times 10^{-8}$	-0.077	A/AT	0.689	Comb ASD	$9.6 \times 10^{-9}$	0.078	C8orf74, SOX7, PINX1
	rs71190156	20	14836243	ASD	$2.75 \times 10^{-8}$	-0.078	G/TTT	0.481	ASD-Edu	$1.6 \times 10^{-8}$	0.056	
	rs111931861	7	104744219	Comb ASD	$3.53 \times 10^{-8}$	-0.216	A/G	0.966	ASD	$3.4 \times 10^{-7}$	-0.033	LOC102723661, PTBP2
<b>b</b>	rs2388334	6	98591622	ASD-Edu	$3.34 \times 10^{-12}$	-0.065	A/G	0.517	ASD	$1.0 \times 10^{-6}$	-0.068	MMS22L, POU3F2
	rs325506	5	104012303	ASD-MD	$3.26 \times 10^{-11}$	0.057	C/G	0.423	ASD	$3.5 \times 10^{-7}$	0.071	NUD12
	rs11787216	8	142615222	ASD-Edu	$1.99 \times 10^{-9}$	-0.058	T/C	0.364	ASD	$2.6 \times 10^{-6}$	-0.030	MROH5
	rs1452075	3	62481063	ASD-Edu	$3.17 \times 10^{-9}$	0.061	T/C	0.721	ASD	$2.1 \times 10^{-7}$	0.035	CADPS
	rs1620977	1	72729142	ASD-MD	$6.66 \times 10^{-9}$	0.056	A/G	0.26	ASD	$1.2 \times 10^{-4}$	0.062	NEGR1
	rs10149470	14	104017953	ASD-MD	$8.52 \times 10^{-9}$	-0.049	A/G	0.487	ASD	$8.5 \times 10^{-5}$	-0.056	MARK3, CKB, TRMT61A, BAG5, APOPT1, KLC1, XRCC3
	rs16854048	4	42123728	ASD-MD	$1.29 \times 10^{-8}$	0.069	A/C	0.858	ASD	$5.9 \times 10^{-5}$	0.082	SLC30A9, BEND4, TMEM33, DCAF4L1