Review

Effect of Low-Density Lipoprotein Cholesterol Lowering by Ezetimibe/Simvastatin on Outcome Incidence: Overview, Meta-Analyses, and Meta-Regression Analyses of Randomized Trials

Address for correspondence: Costas Thomopoulos, MD Helena Venizelou Square 11521 Athens, Greece thokos@otenet.gr

Costas Thomopoulos, MD; George Skalis, MD; Helena Michalopoulou, MD; Costas Tsioufis, MD; Thomas Makris, MD

Department of Cardiology (Thomopoulos, Skalis, Michalopoulou, Makris), Helena Venizelou Hospital, Athens, Greece; First Cardiology Clinic (Tsioufis), Hippokration Hospital, Athens University, Athens, Greece

ABSTRACT

This analysis investigated the extent of different outcome reductions from low-density lipoprotein cholesterol (LDL-C) lowering following ezetimibe/simvastatin treatment and the proportionality of outcome to LDL-C reductions. The authors searched PubMed between 1997 and mid-June 2015 (any language) and the Cochrane Library to identify all randomized controlled trials comparing ezetimibe/simvastatin with placebo or less intensive LDL-C lowering. Risk ratios (RR) and 95% confidence intervals (CIs), standardized to 20 mg/dL LDL-C reduction, were calculated for 5 primary outcomes (fatal and nonfatal) and 4 secondary outcomes (non-cardiovascular [CV] death, cancer, myopathy, and hepatopathy). Five ezetimibe/simvastatin RCTs (30 o51 individuals) were eligible, 2 comparing ezetimibe/simvastatin vs placebo and 3 vs less intensive treatment. Outcomes reduced almost to the same extent were stroke (RR: -13%, 95% CI: -21% to -3%), coronary heart disease (CHD; RR: -12%, 95% CI: -19% to -5%), and composite of stroke and CHD (RR: -14%, 95% CI: -20% to -8%). Absolute risk reductions: 5 strokes, 10 CHD events, and 16 stroke and CHD events prevented for every 1000 patients treated for 5 years. Residual risk was almost $7 \times$ higher than absolute risk reduction for all the above outcomes. All death outcomes were not reduced, and secondary outcomes did not differ between groups. Logarithmic risk ratios were not associated with LDL-C lowering. Our meta-analysis provides evidence that, in patients with different CV disease burden, major CV events are safely reduced by LDL-C lowering with ezetimibe/simvastatin, while raising the hypothesis that the extent of LDL-C lowering might not be accompanied by incremental clinical-event reduction.

Introduction

The lowering of low-density lipoprotein cholesterol (LDL-C) is accompanied by cardiovascular (CV) risk reduction.¹ Beyond lifestyle changes, statins represent the more established option to reduce LDL-C in patients with hypercholesterolemia, and the majority of randomized controlled trials (RCTs) so far investigated both their effects on hard endpoints and their long-term safety. Although a great amount of evidence suggests that the beneficial

effect of statins on CV outcomes is accomplished by LDL-C lowering per se, the potential pleiotropic effects of statins as modulators of risk reduction could not be definitively excluded.² In the usual clinical practice, the introduction of additional LDL-C–lowering agents on top of statins for different clinical reasons (eg, patient intolerance of high statin doses, only a small additional reduction from doubling statin dose, LDL-C target not achieved with high statin dose alone, taking advantage of a complementary mechanism of

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action to statins) may be considered, especially in high-risk patients. $^{3,4}\,$

Although more intensive LDL-C lowering (compared with less intensive LDL-C lowering) based on statin monotherapy steadily reduced all types of clinical outcomes in a wide spectrum of patients with CV risk,^{1,5} the use of different traditional lipid-modifying agents such as niacin, fibrates, and cholesteryl ester transfer protein (CETP) inhibitors as an adjunct to statins was not accompanied by incremental CV event reduction. Ezetimibe (an enterocyte cholesterol transporter Niemann-Pick C1-like 1 protein inhibitor) in combination with simvastatin has been reported to effectively reduce LDL-C levels by almost 23%,⁶ but whether this reduction, compared with less active treatment (placebo or simvastatin alone), is accompanied by a positive effect on different outcomes, and whether there is a proportional association between the LDL-C lowering and reduced event rate, still remain unclear.4

Previous overviews and meta-analyses have specifically focused on the effects of statins alone on clinical outcomes,^{1,5} whereas pooled long-term efficacy and safety results of LDL-C lowering following the use of the ezetimibe/simvastatin combination are lacking. We have done a comprehensive overview of ezetimibe/simvastatin RCTs accompanied by a series of meta-analyses and meta-regression analyses to address these issues. We quantified the relative and absolute benefits of LDL-C lowering following ezetimibe/simvastatin treatment in patients with different cardiovascular disease (CVD) burden on distinct types of CV and safety outcomes, and we investigated the effects of LDL-C reductions of different extent. Finally, we calculated the residual risk in those randomized to ezetimibe/simvastatin treatment as a measure of inadequate or suboptimal treatment.

Methods

Trial Eligibility

The present overview intended to include RCTs of LDL-C lowering with ezetimibe/simvastatin, in which (1) ezetimibe/simvastatin was compared with placebo with the intention to investigate the consequences of LDL-C differences (intentional LDL-C-lowering trials, placebo controlled); or (2) a more intensive LDL-C lowering based on the combination of ezetimibe/simvastatin was compared with a less intensive one (intentional LDL-C-lowering trials, more or less intensive).

In addition, trials had to meet the following predetermined criteria: (1) protocol including measurement of ≥ 1 type of CV event as a primary endpoint; (2) LDL-C values measured at baseline and follow-up; (3) follow-up of ≥ 6 months; (4) a minimum of 5 events during follow-up; (5) achieved LDL-C difference among arms of $\geq 5 \text{ mg/dL}$; (6) randomized allocation to treatments; (7) publication before June 15, 2015. The database search was done by 2 of the authors (C.Th and G.S.) by consulting PubMed between 1997 and mid-June 2015 (any language) and the Cochrane Library. Reference lists from retrieved articles and abstracts from international CV conferences were also sought. Recommendations of the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) statement were adhered to.⁷

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Outcomes

Data on 5 predetermined primary outcomes were extracted: (1) stroke (fatal and nonfatal); (2) coronary heart disease (CHD) events (coronary death and nonfatal myocardial infarction [MI]); (3) major CV events (composite of stroke and CHD); (4) CV death; (5) all-cause death. The definition of outcomes reported in the original article was retained, but whenever possible transient ischemic attacks, angina, and revascularization procedures were excluded. Sudden death was added to CV death, but death of unknown origin was only considered in the all-cause death outcome. Four predetermined secondary outcomes were also extracted: (1) non-CV death; (2) cancer (fatal and nonfatal); (3) myopathy (creatine kinase $>10 \times$ upper limit of normal with muscle symptoms); (4) hepatopathy (increases of alanine or aspartate transaminase $>3 \times$ upper limit of normal). Two authors (C.Th. and G.S.) independently extracted the data, with differences resolved by discussion.

Quality Assessment

Selection, detection, and attrition bias were assessed based on randomization procedure, method of blinding, and combined evaluation of lost to follow-up and therapy discontinuation ratio. Studies of higher quality were those reporting randomization generation sequence, with double blinding, and lost to follow-up ratio <10%, accompanied by therapy discontinuation of <10% per year of follow-up. We also arbitrarily assigned higher quality to studies with \geq 5000 patient-years. Moreover, we evaluated the number of primary outcomes reported in each individual trial, with those reporting >3 types of outcomes being of higher quality compared with those reporting <3 types of outcomes. The evaluation and scoring of these 5 criteria were based on a binomial integer scale ranging from 0 to 1, with 1 being better. The scores were summed up and reflected the overall study quality, with 5 being best. Trials scoring 4 or 5 were considered of higher quality than those scoring 3 or lower.

Statistical Analysis

All analyses were done using data as tabulated in the original publications. In each group, baseline patient characteristics and LDL-C differences between randomized treatments were the means of all individual trial values weighted by number of patients and follow-up duration. For every group of comparisons, the null hypothesis of no difference between randomized treatments (ezetimibe/simvastatin vs placebo or less active treatment) was tested for each of the outcomes. Relative risk estimates (with 95% confidence intervals [CI]) were combined using a random-effects model, in which the log relative risk for every trial was weighted by the reciprocal of the variance of the log relative risk. The proportion of inconsistency across studies not explained by chance was quantified with the I^2 statistic. Whenever no significant heterogeneity was detected by the χ^2 Q statistics (P > 0.05), a fixed-effect model was also implemented. The influence of individual trials on pooled effect sizes was tested by excluding one trial at a time (one-study removed analysis): If the point estimate of the combined effect size with a given trial excluded lay outside the CI of the overall estimate risk



Figure 1. Identification process for inclusion of randomized controlled trials.

with all available trials, the trial in question was considered to have an excessive influence.

Risk ratios (RR) and their 95% CIs were reported using the Mantel-Haenszel method, and the effect of LDL-C lowering on each outcome was illustrated with forest plots under the random-effects model. Risk estimates were standardized to a difference of 20 mg/dL by multiplying the relative risk estimate in each trial by the appropriate factor after having considered the effect of the inverse variance of individual trials. Five-year absolute risk reductions (weighted for follow-up period inverse variance and sample size) of standardized LDL-C-lowering treatment were also calculated, as well as numbers of patients needed to treat for 5 years to prevent 1 outcome event. Residual risk (remaining risk in the ezetimibe/simvastatin group) represented the difference between baseline outcome risk calculated in the placebo or less intensive LDL-C-lowering group and the absolute risk reduction. Random-effects meta-regression models with inverse variance weighting were constructed to explore whether the achieved LDL-C difference (independent variable) between randomized groups explained the variance of relative risk estimates for various outcomes. To correct for different levels of control LDL-C, meta-regressions were also calculated by expressing achieved LDL-C reductions as percentages of respective ongoing LDL-C values in the control groups. The presence of publication bias was investigated graphically by funnel plots of precision (random effects plotting) and the Duval and Tweedie trim-and-fill method. All statistical analyses were done using Comprehensive Meta-Analysis software, version 3 (Biostat, Englewood, NI). In each individual analysis, a P value <0.05 was considered to indicate statistical significance.

Results

Trials and Patients

Figure 1 illustrates the investigational steps to identify trials for inclusion in the meta-analysis. For the search strategy

and a list of excluded trials, see Supporting Information, tables S1 and S2, in the online version of this article.

This procedure identified 5 eligible RCTs. Characteristics of the 5 included trials,^{8–12} with a total of 30 051 participants followed up for a mean of 5.5 years (163 778 patient-years), are presented in Table 1. Eighty percent of RCTs (4/5) were of higher quality (scoring from 4 to 5). Two trials^{10,11} (11 143 participants, 961 major CV events) were of intentional LDL-C lowering vs placebo, and 3 trials^{8,9,12} (18 908 participants, 2751 major CV events) were of intentional more vs less intensive LDL-C lowering. One study reported data only for major CV events and all-cause death.¹²

Effects of Ezetimibe/Simvastatin LDL-C-Lowering Treatment on Various Outcomes

Achieved LDL-C differences between treatments were larger in placebo-controlled trials^{10,11} than in more vs less intensive LDL-C-lowering trials^{8,9,12} (-35.8 vs -17.1 mg/dL), even after standardization to ongoing LDL-C levels in the placebo or less active group (32.8% vs 22.8%). Ezetimibe-based LDL-C-lowering treatment reduced the risk of stroke, CHD, and their composite outcome but was not associated with better mortality outcomes (Figure 2A). A standardized LDL-C reduction of 20 mg/dL was found to reduce the risk of CV events by 14%. In terms of absolute risk, the same LDL-C reduction could prevent 5 strokes, 10 CHD, and 16 major CV events (composite of stroke and CHD) for every 1000 patients treated for 5 years (number needed to treat: 216, 102, and 63, respectively). The residual risk for all primary outcomes (Table 2) was higher the higher the baseline risk and the lower the absolute risk reduction. For example, residual risk of major CV events amounted to 115 events for each 1000 patients during 5 years, suggesting that suboptimal treatment was almost $7.2 \times$ greater (115/16) than absolute risk reduction (16%) from LDL-C lowering. This "suboptimal to beneficial" treatment ratio following LDL-C lowering, resulting from the same group of trials, was 6.8 (34/5) and 7.3 (73/10) for stroke and CHD, respectively.

Lowering of LDL-C by ezetimibe/simvastatin was not accompanied by different incident rates of non-CV death, cancer, hepatopathy, and myopathy as compared with placebo or less active treatment (Figure 2B). In our analysis, baseline risk for non-CV death, cancer, hepatopathy, and myopathy (9.6%, 8.3%, 1.7%, and 0.3%, respectively) remained almost unaltered after 5.5 years. By excluding the trial of lower quality,¹² no different results for both primary and secondary outcomes were obtained (data not shown).

Outcome Reductions and Extent of LDL-C Lowering by Ezetimibe/Simvastatin

Although the span of LDL-C reductions was narrower for stroke and CHD compared with that observed for their composite outcome (9.8–51.4 mg/dL vs 15.8–51.4 mg/dL), the natural logarithm of the RR of stroke, CHD, and of their composite was not significantly related to the extent of LDL-C lowering (Figure 3A). In other words, progressively lower LDL-C levels were not accompanied by incremental lowering of clinical events. When meta-regressions were calculated by using percentage changes in LDL-C (maximum reduction of 34%), regression coefficients

					0												
		Clinical	Statin	Follow-up,	Participants,	Age,	Men.	BMI,	Smoking,	DM.	HTN.	10-y Risk CV	Baseline LDL-C, EZE	Baseline LDL-C, Non-EZE	Achieved LDL-C, EZE	Achieved LDL-C, Non-EZE Group,	
Trial	Comparator	Setting	Pretreatment %	^	. =	- 	%	kg/m²	%	%	- %	Death, %	Group, mg/dL	Group, mg/dL	Group, mg/dL	mg/dL	Quality
ENHANCE, 2008 ⁸	Simvastatin	H	81	7	720	45.9	51.4	27	28.7	1.8	16	1.4	319	318	141.3	192.7	4/5
SEAS, 2008 ¹⁰	Placebo	AoS	0	4.4	1873	67.6	61.4	26.9	19	0	52	13.7	139	140	102	137.3	5/5
SHARP, 2011 ¹¹	Placebo	CKD	NR	4.9	9270	62	63	27.1	13	23	NR	19.6	107.1	107.5	69	104.9	5/5
IMPROVE-IT, 2014 ⁹	Simvastatin	Post-MI	35	9	18144	64	76	28.3	33	27	61	11.5	94	94	53.7	69.5	5/5
West et al, 2011 ¹²	Simvastatin	PAD	0	0	74	62	55	29	50	30	78	NR	118	118	67.5	87	3/5
Abbreviations: A in Hypercholest Vytorin Efficacy in Aortic Stenosi	oS, mild to m erolemia Enh International s; SHARP, Stu	ioderate as nances Ath Trial; LDI udy of Hea	symptomatic aor herosclerosis Re L-C, low-density art and Renal Pr	tic stenosis; gression; E lipoprotein otection.	BMI, body n ZE, ezetimib cholesterol; l	ass ind e; FH, M, myo	lex; CK heteroz ocardial	D, chroi ygous fa infarcti	nic kidney umilial hy on; NR, n	/ disease perchol ot repor	e; CV, c esterole ted; PA	ardiovascu emia; HTN D, periphe	ılar; DM, diał I, hypertensio eral arterial (l	oetes mellitus; E on; IMPROVE-I ower limb) dise	NHANCE, Ez T, Improved I ase; SEAS, Sir	cetimibe and Simv Reduction of Out nvastatin and Ez	astatin comes: etimibe

did not change for stroke and major CV events (composite stroke and CHD; Figure 3B). By contrast, although without statistical significance, the direction of slope for CHD was inversed.

Influence of Studies on Pooled Effect Sizes, Fixed-Effect Models, and Publication Bias

Whenever a fixed-effect model was implemented, RRs and their significance did not substantially change (data not shown). Also, by applying the one-study removed analytical procedure, no trial had an excessive influence in any analysis. Although graphic representations could not exclude publication bias for all primary outcomes, significant bias was denied by the trim-and-fill method. For this assessment, see Supporting Information, Figure S1 and Table S3, in the online version of this article.

Discussion

This is the first overview to systematically address the extent of LDL-C-lowering benefits following ezetimibe/simvastatin treatment and whether this extent is proportional to LDL-C reduction in patients with different CVD burden. Additionally, over 5.5 years of follow-up, we assessed whether the combination of ezetimibe/simvastatin is safe in terms of non-CV death, cancer, myopathy, and hepatopathy incidence.

We observed that the relative risk reduction of stroke, CHD, and of their composite, standardized for achieved LDL-C difference of 20 mg/dL (close to the mean observed in trials considered), was 13%, 12%, and 14%, respectively. However, the observed benefit might be inflated because we also included patients after acute MI,9 a condition in which ezetimibe/simvastatin treatment might have potentiated pleiotropic effects beyond LDL-C lowering. For 20-mg/dL LDL-C lowering, 5 strokes were avoided for every 1000 patients treated for 5 years (this means that 216 patients had to be treated for 5 years to prevent 1 stroke), 10 CHD events were avoided (this means that 102 patients had to be treated to prevent 1 event), and 16 major CV events (stroke and CHD) were avoided (this means that 63 patients had to be treated to prevent 1 CV event). By contrast, both CV and all-cause mortality were not associated with more intensive LDL-C lowering following ezetimibe/simvastatin treatment. Although the span of the achieved difference in LDL-C was relatively small (almost 50 mg/dL), the lack of relationship of LDL-C reductions by ezetimibe/simvastatin with the logarithm of the outcome RRs suggests that progressively greater LDL-C reductions do not result in progressively lower increments of risk reduction. Thus, based on the cross-sectional nature of our meta-regression data, we could only hypothesize that the reduction of major CV events following ezetimibe/simvastatin treatment might be independent of the LDL-C-lowering extent, in contrast to the suggestion of the Improved Reduction of Outcomes: Vytorin Efficacy International Trial (IMPROVE-IT) of "the lower the better," which, however, included only post-MI patients.9

So far, in the larger meta-analysis of LDL-C lowering,¹ in which the achieved LDL-C difference was standardized by 38.7 mg/dL (1 mmol/L), statins alone reduced the

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Outcome	Trials n	ا FU (y)*	Difference LDL (mg/dl)*	Eve (n/pa Treated	ents tients) Controls	RR (95% CI)	Standardized RR (95% Cl)	P heter	Standardized RR (95% Cl)	Absolute Risk Reduction 1000 pts/5 years (95% Cl)	NNT 5 years (95% Cl)
A. Efficacy outcom	mes (pi	rimar	y)							· · /	
Stroke	4	5.6	-22.2	511/15018	592/14989	0.86 (0.77-0.97)	0.87 (0.79-0.97)	0.79		-5 (-8, -1)	216 (134, 944)
CHD	4	5.7	-20.0	1217/15018	1387/14989	0.88 (0.81-0.95)	0.88 (0.81-0.95)	0.38		-10 (-16, -4)	102 (64, 246)
Stroke + CHD	5	5.9	-17.9	1732/15040	1980/15011	0.87 (0.82-0.93)	0.86 (0.80-0.92)	0.62	-	-16 (-23, -9)	63 (44, 111)
CV Death	4	5.7	-21.2	997/15018	1038/14989	0.96 (0.88-1.04)	0.96 (0.89-1.04)	0.61		-2 (-7, 2)	410 (149, -413)
All-cause Death	5	5.6	-22.2	2465/15040	2448/15011	1.00 (0.96-1.07)	1.00 (0.96-1.06)	0.95		0 (-6, 9)	∞(171, -115) [′]
B. Safety outcom	es (sec	onda	iry)								
Non-CV Death†	3	5.6	-22.6	1458/14661	1409/14626	1.04 (0.94-1.14)	1.04 (0.95-1.12)	0.22	+•	- 3 (-4, 12)	-291 (232, -84)
Cancer	4	5.2	-27.9	1291/15018	1242/14989	1.07 (0.93-1.24)	1.05 (0.95-1.17)	0.080	-+•	4 (-4, 14)	-251 (251, -74)
Hepatopathy	4	5.7	-20.9	269/14988	258/14985	0.97 (0.66-1.43)	0.97 (0.67-1.41)	0.12	•	→ <-1 (-5, 6)	>1000 (196, -165)
Myopathy	4	5.9	-17.2	44/14988	44/14985	1.00 (0.66-1.52)	1.01 (0.62-1.63)	0.65		→ 0 (-1, 2)	∞ (>1000, -664)
								0.4 EZE/	0.7 1.0 SIMVA better (1.25 Control better	

Figure 2. Relative and absolute risk reduction of various outcomes in EZE/SIMVA trials of LDL-C lowering. Standardized RR is to an LDL-C reduction of 20 mg/dL. Absolute risk reduction refers to the number (and 95% CI) of events prevented every 1000 patients treated for 5 years with a standardized RR. NNT is the numbers (and 95% CI) of patients needed to treat for 5 years to prevent 1 event. Abbreviations: CHD, coronary heart disease; CI, confidence interval; CV, cardiovascular; ENHANCE, Ezetimibe and Simvastatin in Hypercholesterolemia Enhances Atherosclerosis Regression; EZE/SIMVA, ezetimibe/simvastatin; FU, follow-up; LDL-C, low-density lipoprotein cholesterol; n, number; NNT, number needed to treat; P heter., *P* for heterogeneity; RR, Mantel-Haenszel risk ratios. * Weighted for the inverse variance. † ENHANCE trial reported o non-CV deaths.

Table 2. Residual Risk a in the Ezetimibe-Treated Groups by Different Primary Outcomes

Outcome	Baseline Outcome Risk, %	Residual Risk per 1000 Patients/5 Years (95% Cl)
Stroke	3.9	34 (31–38)
CHD	8.3	73 (67–79)
Stroke + CHD	13.1	115 (108–122)
CV death	6.9	67 (62–71)
All-cause death	16.3	163 (157–172)

Abbreviations: CI, confidence interval; CHD, coronary heart disease; CV, cardiovascular.

 $^a {\rm Residual}$ risk refers to the residual events every 1000 patients treated for 5 years.

incidence of CHD by 24% (compared with 22% in the present analysis), of stroke by 16% (compared with 23% in the present analysis), and of major vascular events, including revascularizations, by 22% (compared with 26% in the present analysis but excluding revascularizations). The difference in benefit of major CV events between the present analysis and the analysis of the Cholesterol Treatment Trialists' Collaboration (CTT)¹ might be driven from the relative higher number of post-MI participants in our case.⁹ Also, in another overview of the CTT including only primaryprevention trials,⁵ the reduction of major coronary events was 2% higher than our estimates. Although the higher stroke benefit by 6% to 7% between the present overview and the 2 analyses of the CTT group^{1,5} might not be explained by the baseline prevalence of hypertension (almost 60% in both the previous and the present analysis), ongoing blood-pressure levels are lacking in all LDL-C-lowering trials. A specific vascular protection on intermediate endpoints (carotid artery atherosclerotic changes) by the ezetimibe/simvastatin combination compared with simvastatin alone was rejected in previous studies, but their conclusions are limited, either because of nonrandomized allocation¹³ or because they included a very limited number of participants.¹² We have resisted performing a differential analysis between ischemic and hemorrhagic stroke, because in previous observational studies¹⁴ and trials^{15,16} not included in the CTT meta-analysis,¹ LDL-C lowering has been associated with increased incidence of hemorrhagic stroke.

At variance with the integrated evidence based on statin LDL-C-lowering trials,^{1,5} ezetimibe/simvastatin treatment was not associated with CV death reduction, possibly because the prevention of occlusive cardiac death by statins at higher doses might have an incremental functional and/or structural beneficial effect on the vasculature, such as vasodilation and/or neoangiogenesis that in turn can reduce the extent of fatal arrhythmiogenesis.¹⁷ However, our metaanalysis cannot suggest that for the same extent of LDL-C lowering, mixed statin and nonstatin treatment could be less effective compared with more aggressive statin treatment alone to modulate pivotal mechanisms of coronary death such as fatal arrhythmias. A previous meta-analysis,¹⁸ in which nonstatin LDL-C lowering intervention by diet, bileacid sequestrants, and surgery was accompanied by the same extent of CV risk benefit compared with statins, should be interpreted with caution: first, ezetimibe was not included among the nonstatin treatments; second, LDL-C levels were indirectly measured in two-thirds of nonstatin treatment studies; third, women were excluded in >60% of studies; fourth, the sample size of nonstatin studies was guite limited; and, finally, the recruitment of participants of nonstatin trials was initiated almost 30 years before the availability of statins.

Our analysis highlights once $again^{19,20}$ the importance of residual risk in patients having received the more active LDL-C-lowering treatment. Residual risk is a measure of inadequate or suboptimal treatment and should always be taken into account in relation to treatment benefit assessed by absolute risk reduction within a predetermined period of time under treatment. Our findings demonstrated that suboptimal treatment during 5 years with ezetimibe/simvastatin vs less active treatment including placebo was almost $7 \times$ higher compared with treatment benefit. Thus, to minimize "inadequate treatment" rate, more effective therapeutic strategies should be undertaken,



Figure 3. Relationships of (A) primary outcome reductions to the extent of LDL-C reductions. Meta-regressions of risk ratios on absolute LDL-C differences (active treatment group minus placebo or less active treatment group). Relationships of (B) outcome reductions to percentage LDL reductions. The meta-regressions of (A) are calculated on percentage LDL-C differences (LDL-C difference as percentage of on-treatment LDL-C in the control group). Regressions relative to stroke are in continuous lines; CHD, in dotted lines; composite of stroke and CHD, in dashed lines. Abbreviations: CHD, coronary heart disease; △-LDL, LDL-C difference; LDL-C, low-density lipoprotein cholesterol; RR, risk ratio.

such as even more aggressive LDL-C lowering^{21–23} or implementation of optimal management of traditional and emerging CV risk factors (eg, hypertension control, smoking cessation, targeting on modulation of other lipid parameters). Lowering of LDL-C by ezetimibe/simvastatin was not accompanied by changes in the rate of non-CV death and any (fatal and nonfatal) cancer, in line with previous overviews but restricted to statins.^{1,5} Additionally, incidence of clinically important side effects like myopathy and hepatopathy was quite small and not different from that reported with statin monotherapy.²⁴

Study Strengths and Limitations

One of the strengths of our meta-analysis is that we have investigated not only the effects of LDL-C lowering on relative risk reductions following ezetimibe/simvastatin compared with less active treatment, but we have also estimated the residual risk (ie, suboptimal treatment) and its relation with absolute risk reduction (ie, treatment benefits), to provide a more integrated insight for clinical decision-making. We standardized risk reductions to a 20mg/dLLDL-C decrease (similar to the mean ongoing LDL-C between arms), and not to 38.7 mg/dL (1 mmol/L), to avoid bias of extrapolation of risk estimates.

In the trials of ezetimibe/simvastatin considered in our analysis, the comparator arm was either placebo or simvastatin alone. Thus, risk reduction could not be assigned to ezetimibe alone, but only to LDL-C lowering developed by the combined treatment. We have included a large post-MI trial⁹ in which the incidence of clinical events may be independent of, and occasionally hindered by, LDL-C lowering. In our analysis, we included different CVD burden patients (eg, valvular heart disease, post-MI, chronic kidney disease) with potential high variability on outcomes' incidence. Despite the different extent of 10year CV death risk and statin pretreatment, we did not perform stratified analyses because the number of trials was quite small. Also, the limited number of trials did not allow us to perform analyses across different LDL-C thresholds. Meta-regression analyses, though instrumental at investigating quantitative relationships between risk and intervention, could not be seen as alternative to traditional meta-analyses for the estimation of the mean effect for a given intervention. Therefore, the finding that major CV events, CHD, and stroke can be reduced by LDL-C lowering should be considered stronger than the cross-sectional evidence provided by our meta-regressions that denied the association between this benefit and extent of LDL-C lowering.

Conclusions

In our meta-analysis, LDL-C lowering following ezetimibe/simvastatin treatment was associated with reduced rate of CHD, stroke, and their composite outcome. Risk reductions were not proportional with the extent of LDL-C lowering, raising the hypothesis that beneficial effects might be partially mediated by effects independent of LDL-C lowering. Cardiovascular death was not reduced, and further studies should examine whether ezetimibe-based treatment could modulate pivotal mechanisms of vascular death compared with statin monotherapy at the same extent of LDL-C lowering. We also demonstrated that ezetimibe/simvastatin treatment is safe by means of non-CV death, cancer incidence, and incidence of clinically important side effects such as myopathy and hepatopathy.

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