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## Vitamin D as an adjunct to antibiotics for the treatment of acute childhood pneumonia (Review)

Das RR, Singh M, Naik SS

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Vitamin D as an adjunct to antibiotics for the treatment of acute childhood pneumonia (Review)

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[Intervention Review]

# Vitamin D as an adjunct to antibiotics for the treatment of acute childhood pneumonia

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## ABSTRACT

### Background

Children with acute pneumonia may be vitamin D deficient. Clinical trials have found that prophylactic vitamin D supplementation decreases the risk of developing pneumonia in children. Data on the therapeutic effects of vitamin D in acute childhood pneumonia are limited.

### Objectives

To evaluate the efficacy and safety of vitamin D supplementation as an adjunct to antibiotics for the treatment of acute childhood pneumonia.

### Search methods

We searched CENTRAL (2017, Issue 7), which includes the Cochrane Acute Respiratory Infections Group's Specialised Register; Ovid MEDLINE Epub Ahead of Print; In-Process & Other Non-Indexed Citations; Ovid MEDLINE Daily and Ovid MEDLINE (1946 to July Week 4, 2017); and Embase (2010 to 28 July 2017). We also searched ClinicalTrials.gov and the World Health Organization (WHO) International Clinical Trials Registry Platform (ICTRP) on 28 July 2017. There were no language restrictions.

### Selection criteria

Randomised controlled trials (RCTs) including children (aged over one month and up to five years) hospitalised with acute community-acquired pneumonia, as defined by the WHO acute respiratory infection guidelines, that compared vitamin D supplementation with control.

### Data collection and analysis

Two review authors independently assessed studies for inclusion and extracted data. For dichotomous data, we extracted the number of participants experiencing the outcome and the total number of participants in each treatment group. For continuous data, we used the arithmetic mean and standard deviation (SD) for each treatment group together with numbers of participants in each group. We used standard methodological procedures expected by Cochrane.

### Main results

We included seven RCTs conducted in low-income countries that involved 1529 children (780 with pneumonia and 749 with severe or very severe pneumonia). Four studies used a single 100,000 IU dose of vitamin D<sub>3</sub> at the onset of illness or within 24 hours of hospital admission;

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two used a daily dose of oral vitamin D<sub>3</sub> (1000 IU for children aged up to one year and 2000 IU for children aged over one year) for five days; and one used a daily dose of oral vitamin D<sub>3</sub> (50,000 IU) for two days. One study reported microbiological and radiological diagnosis of pneumonia.

The effects of vitamin D on outcomes were inconclusive when compared with control: time to resolution of acute illness (hours) (mean difference (MD) -0.95, 95% confidence interval (CI) -6.14 to 4.24; 3 studies; 935 children; low-quality evidence) mortality rate (risk ratio (RR) 0.97, 95% CI 0.06 to 15.28; 1 study; 193 children; very low-quality evidence); duration of hospitalisation (MD 0.49, 95% CI -8.41 to 9.4; 4 studies; 835 children; very low-quality evidence) and time to resolution of fever (MD 1.66, 95% CI -2.44 to 5.76; 4 studies; 584 children; very low-quality evidence).

No major adverse events were reported.

The GRADE assessment found very low-quality evidence (due to serious study limitations, inconsistencies, indirectness, and imprecision) for all outcomes except time to resolution of acute illness.

One study was funded by the New Zealand Aid Corporation; one study was funded by an institutional grant; and five studies were unfunded.

### Authors' conclusions

We are uncertain as to whether vitamin D has an important effect on outcomes because the results were imprecise. No major adverse events were reported. We assessed the quality of the evidence as very low to low. Several trials are ongoing and may provide additional information.

## PLAIN LANGUAGE SUMMARY

### Is vitamin D an effective and safe addition to antibiotics to treat children with acute pneumonia?

#### Review question

We wanted to find out if vitamin D helps children with acute pneumonia who are also receiving antibiotic treatment get better faster.

#### Background

Pneumonia is an acute lower respiratory tract infection that affects the lungs. Treatment for pneumonia includes antibiotics, providing supplementary oxygen to air that is breathed in through a mask, and other supportive therapies. Vitamin D boosts immune defences and reduces excessive inflammation, effects that may help children recover from an acute episode of pneumonia.

#### Search date

The evidence is current to 28 July 2017.

#### Study characteristics

We included seven studies involving a total of 1529 children (780 with pneumonia (4 studies) and 749 with severe or very severe pneumonia (3 studies)) aged under 5 years from low-income countries. In four studies, a single large dose of vitamin D was used either when the child joined the study or within 24 hours of admission to hospital; in two studies, vitamin D was used for five days; and in one study, vitamin D was used for two days. One study excluded children whose vitamin D levels were normal. One study reported the cause of children's pneumonia.

#### Study funding sources

One study was funded by the New Zealand Aid Corporation; one was funded by an institutional grant; and five studies were unfunded.

#### Key results

We are uncertain as to whether vitamin D has an important effect on outcomes due to the very-low quality of the evidence. Vitamin D may slightly decrease the time taken to get better from acute pneumonia (by 60 minutes) and the risk of death, and Vitamin D may increase the length of time in hospital (by 30 minutes) and the time taken for fever to resolve (by 90 minutes). However, there was no significant difference between groups for these outcomes. No major adverse events were reported.

#### Quality of the evidence

The quality of the evidence was very low, except for time to resolution of acute illness, which we assessed as low quality. We identified problems with the study methods and reporting, resulting in lack of precision in the included studies.

## SUMMARY OF FINDINGS

### Summary of findings for the main comparison. Vitamin D compared with placebo for acute childhood pneumonia

#### Vitamin D compared with placebo for acute childhood pneumonia

**Patient or population:** children with acute pneumonia

**Settings:** hospital setting

**Intervention:** oral vitamin D

**Comparison:** usual care

Outcomes	Illustrative comparative risks* (95% CI)		Relative effect (95% CI)	No. of participants (studies)	Quality of the evidence (GRADE)	Comments
	Assumed risk	Corresponding risk				
	Control	Vitamin D				
<b>Time to resolution of acute illness</b>	Mean time in the control group ranged from 50.8 to 119.52 hours.	The mean time in the intervention group was 0.95 hours less (6.14 hours less to 4.24 hours more).	-	935 (3 studies)	⊕⊕⊕⊕ <b>Low</b> <sup>1-4</sup>	4 studies reported the outcome (data pooled from 3 studies). The appropriate effect measure for time-to-event data is hazard ratio. However, the studies reported this information as continuous data (1 study reported as a proportion). The MD was around 1 hour less for the vitamin D group.
<b>Duration of hospitalisation</b>	Mean duration in the control group ranged from 106.82 to 168 hours.	The mean duration in the intervention group was from 0.49 hours more (8.41 hours less to 9.4 hours more).	-	835 (4 studies)	⊕⊕⊕⊕ <b>Very low</b> <sup>5-8</sup>	5 studies reported the outcome (data pooled from 4 studies; mean and standard deviation calculated for 2 studies that reported the data as median). The MD was around half an hour more for the vitamin D group.
<b>Mortality rate</b>	97 per 1000	94 per 1000 (6 to 1000)	RR 0.97 (0.06 to 15.28)	193 (1 study)	⊕⊕⊕⊕ <b>Very low</b> <sup>9</sup>	

\*The basis for the **assumed risk** (e.g. the median control group risk across studies) is provided in footnotes. The **corresponding risk** (and its 95% confidence interval) is based on the assumed risk in the comparison group and the **relative effect** of the intervention (and its 95% CI).

**CI:** confidence interval; **MD:** mean difference; **RR:** risk ratio

GRADE Working Group grades of evidence

**High quality:** Further research is very unlikely to change our confidence in the estimate of effect.

**Moderate quality:** Further research is likely to have an important impact on our confidence in the estimate of effect and may change the estimate.

**Low quality:** Further research is very likely to have an important impact on our confidence in the estimate of effect and is likely to change the estimate.

**Very low quality:** We are very uncertain about the estimate.

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<sup>1</sup>No serious study limitations: all studies had adequately concealed allocation and both participants and study staff were blinded. Considered at low risk of bias.

<sup>2</sup>No serious inconsistency: there was a low ( $I^2 = 36\%$ ) and non-significant ( $P = 0.21$ ) statistical heterogeneity.

<sup>3</sup>Serious indirectness: only studies from low-income regions with high baseline prevalence of vitamin D deficiency were assessed in this comparison (the result cannot be generalised to all locations). The characteristics of the populations in the two studies also differed slightly (one study included children aged up to three years, and another included children aged up to five years). Downgraded by one level.

<sup>4</sup>Serious imprecision: the 95% CI around the pooled effect is wide and different in the included studies. The 95% CI includes no effect, and the upper or lower confidence limit crosses the minimal important difference for benefit (here the time to resolution of illness was 5.75 hours less to 3.97 hours more). Downgraded by one level.

<sup>5</sup>Serious study limitations: one study only was assessed as at unclear risk of bias for allocation concealment; blinding of both participants and study staff was considered as at unclear risk of bias. Downgraded by one level.

<sup>6</sup>No serious inconsistency: there was no statistical heterogeneity.

<sup>7</sup>Serious indirectness: only studies from low-income regions with high baseline prevalence of vitamin D deficiency were assessed in this comparison (the result cannot be generalised to all situations). The characteristics of the populations in the two studies also differed slightly (one study included children aged from two months, and another included children aged from six months). Downgraded by one level.

<sup>8</sup>Serious imprecision: the 95% CI includes no effect, and the upper or lower confidence limit crosses the minimal important difference for benefit (here the duration of hospitalisation was 14.09 hours less to 2.64 hours more). Downgraded by one level.

<sup>9</sup>Data came from just one study: no serious study limitations or inconsistency, serious indirectness (downgraded by one level, lack of generalisability), serious imprecision (downgraded by one level, the 95% CI includes no effect).

## BACKGROUND

### Description of the condition

Pneumonia is an acute lower respiratory tract infection that primarily affects the lungs. Microscopically, the alveoli are filled with exudative fluid (pus), which compromises breathing and gas exchange in the lungs (Scott 2012). Clinical manifestations of acute childhood pneumonia include cough, fever, breathing difficulty, and lower chest wall indrawing (Scott 2012). The most common aetiological agents responsible are bacteria and viruses (e.g. *Streptococcus pneumoniae*, *Haemophilus influenzae*, and respiratory syncytial virus).

Risk factors for acute childhood pneumonia include mixed feeding or non-breastfeeding, under nutrition, indoor air pollution, prematurity, overcrowding, and no immunisation against measles (Rudan 2008). In 2010, of the 7.6 million deaths of children aged up to five years globally, 64% were due to infectious diseases (Liu 2012), and 18.3% (1396 million) were attributed to pneumonia (Liu 2012; Lozano 2011).

The management of acute childhood pneumonia includes antibiotics, oxygen, supportive therapy, and assisted ventilation (in severe cases) (Lassi 2014; Rojas-Reyes 2006). Nutritional supplementation such as zinc, vitamin A, and vitamin C have been used, although results have been unfavourable (Chang 2006; Das 2012; Haider 2011a; Haider 2011b; Hemilä 2013; Wu 2005). Similarly, the use of over-the-counter medications (such as mucolytics and antitussives) have not shown beneficial effect (Chang 2014).

### Description of the intervention

Vitamin D is a fat-soluble vitamin that plays an important role in calcium and phosphorous homeostasis and bone mineralisation (Wagner 2008). There are two types: vitamin D<sub>2</sub> (ergocalciferol), which is derived from plants, and vitamin D<sub>3</sub> (cholecalciferol), which is derived from animals. Vitamin D is primarily synthesised in the skin (as vitamin D<sub>3</sub>) after exposure to ultraviolet B radiation; less than 10% is derived from dietary sources (McAllister 2006; Misra 2008). The main form circulating in the blood is 25-hydroxy-vitamin D (25(OH)-D) (synthesised in the liver), and the active or true hormonal form is 1,25-dihydroxyvitamin D (1,25(OH)<sub>2</sub>-D or calcitriol) (synthesised in the kidneys) (Misra 2008; Walker 2009). The active form binds to the receptor present in the nucleus (vitamin D receptor, or VDR) before orchestrating a wide range of actions in the body (Walker 2009). Vitamin D receptor is present in cells of almost every organ system of the body, including immune cells. Due to a longer half-life, measurement of the blood 25(OH)-D level is the best available parameter to indicate the status of vitamin D in the body and is used to define the status of vitamin D in children as: sufficient (50 to 250 nmol/L), insufficient (37.5 to 50 nmol/L), deficient ( $\leq$  37.5 nmol/L), severely deficient ( $\leq$  12.5 nmol/L), and excessive ( $>$  250 nmol/L) (Misra 2008).

As a nutritional supplement, vitamin D is included in most multivitamins (doses vary from 50 IU to 1000 IU) as liquids, tablets, and capsules. Intramuscular injections may be used in special circumstances. A daily dose of 400 IU has been recommended for children and adolescents. The upper limit should not exceed 2000 IU per day in children aged over one year, and 1000 IU per day in infants aged up to one year. However, in deficient states, a daily dose of 1000 IU to 10,000 IU (depending on the child's age) can be

used for two to three months to enable the serum vitamin D level to normalise and replenish stores (Misra 2008).

Vitamin D deficiency or insufficiency is prevalent in low-income countries (Joshi 2014; Kelishadi 2014; Thacher 2012; Torun 2013; Zhang 2013), and also occurs in high-income countries (Grant 2009; Mansbach 2009; Ward 2007). A 10-year data series reported 126 cases of vitamin D deficient rickets in Australia (Robinson 2006), and a two-year data series reported 104 cases in Canada (Ward 2005). In the USA and the UK, there are no definitive prevalence estimates for vitamin D deficient rickets. Historical prevalence estimates from the USA indicate increases of vitamin D deficient rickets from 65 instances between 1975 and 1985 rising to 166 instances between 1986 and 2003 (Weisberg 2004); there were 62 instances between 2004 and 2006 (McAllister 2006; Mylott 2004). Vitamin D deficient rickets remains a significant public health problem in other parts of the world (Joshi 2014; Wondale 2005).

Increased urbanisation, time spent indoors, use of sunscreen, increasing latitude, seasonal change (winter or rainy season), air pollution, prematurity, and malabsorptive disorders are associated with vitamin D deficiency (Misra 2008). Vitamin D deficiency substantially increases the risk of pneumonia among children aged up to five years (Muhe 1997; Najada 2004; Rehman 1994; Wayse 2004). Studies have also found a link between single nucleotide polymorphisms (SNPs) of genes related to the VDR, and susceptibility to respiratory syncytial virus infection (Janssen 2007). A trial that provided vitamin D supplements for children with pneumonia reported a reduction in recurrent episodes (Manaseki-Holland 2010).

### How the intervention might work

Vitamin D has been reported to boost mucosal immune defences and reduce excessive inflammation (Pfeffer 2012). The immune-enhancing actions of vitamin D include induction of monocyte differentiation, inhibition of lymphocyte proliferation, stimulation of phagocytosis-dependent and antibody-dependent macrophages, and modulation of cytokine and antibody-producing lymphocytes (Liu 2006; Raloff 2006; White 2008). Hydroxylation of vitamin D to the active form occurs in the kidneys by the enzyme 1-alpha hydroxylase or CYP27B1; the same hydroxylation pathway also operates in other cells that participate in immune regulation (e.g. epithelial cells and monocyte-macrophages) (Hewison 2011). In the airways (bronchial epithelial cells), 1-alpha hydroxylase activity increases in response to any acute insult or inflammatory stimulus (Hansdottir 2008). Cathelicidin is an antimicrobial peptide whose expression is increased over immune cells in response to infection, and the whole cascade is vitamin D dependent (Gombart 2005; White 2008). As a result, the vitamin D/cathelicidin-circuit is activated in acute respiratory tract infections, including pneumonia. Experimental models have shown that vitamin D dampens adaptive immune responses and modulates key elements of innate immune response in dendritic cells following *S pneumoniae* infection (Olliver 2013). Vitamin D also confers antiviral properties that affect the role of toll-like receptors (TLR), particularly TLR7 (Alvarez-Rodriguez 2012).

Some studies have shown that daily vitamin D supplementation has better therapeutic efficacy than large bolus doses for pneumonia, tuberculosis, and fracture in the elderly (Heaney 2012; Hollis 2011; Manaseki-Holland 2012; Martineau 2012; Nielsen 2010). Clinical studies have shown that vitamin D at a higher dose could be



immunosuppressive. [Kimball 2011](#) investigated high-dose vitamin D (10,000 IU per day) and found suppressed proliferative response of peripheral blood monocytes. This finding is supported by [Manaseki-Holland 2012](#). Large bolus doses (100,000 IU) of vitamin D four times per year in infants indicated a slightly increased risk of repeat episodes of all types of pneumonia and mortality ([Manaseki-Holland 2010](#)). However, the preventive role of vitamin D in children using a single large bolus dose (100,000 IU) found reduced risk of clinically defined pneumonia by 13% up to three months after vitamin D treatment ([Manaseki-Holland 2010](#)). This may mean that the immunological effectiveness of vitamin D is blunted at extremes of dosages (either very high or low therapeutic dose).

### Why it is important to do this review

Pneumonia is a leading cause of global mortality in children aged up to five years; more than 90% of deaths occur in low-income countries ([Liu 2012](#); [Lozano 2011](#)). Although antibiotic therapy is the main treatment for pneumonia, irrational use has contributed to increasing resistance. In addition, accessibility to healthcare facilities is challenging for most children in low-income countries. Vitamin D is a relatively simple intervention that may be a helpful addition to standard treatment for acute childhood pneumonia. Due to its low cost and ease of administration, vitamin D can be administered to children with acute pneumonia. Vitamin D, as an adjunct to antibiotics, may help to reduce childhood mortality from pneumonia. In a previous review, we found no role for oral vitamin D supplementation in children aged under five with acute pneumonia ([Das 2013](#)). This Cochrane Review includes updated searches to assist healthcare providers in making informed decisions about providing vitamin D for children with acute pneumonia.

## OBJECTIVES

To evaluate the efficacy and safety of vitamin D supplementation as an adjunct to antibiotics for the treatment of acute childhood pneumonia.

## METHODS

### Criteria for considering studies for this review

#### Types of studies

Randomised controlled trials (RCTs).

#### Types of participants

Children aged from one month up to five years, hospitalised with a clinical diagnosis of community-acquired pneumonia (CAP). We defined pneumonia according to World Health Organization (WHO) acute respiratory infection guidelines ([Scott 2012](#)). We defined pneumonia in the presence of age-specific tachypnoea (> 60 breaths per minute in children aged up to 2 months; > 50 breaths per minute in children aged from 2 months to 11 months; > 40 breaths per minute in children aged from 12 months to 60 months) with or without cough or fever. We defined severe pneumonia in the presence of age-specific tachypnoea along with lower chest wall indrawing or any of the danger signs (cyanosis, inability to feed, lethargy). We excluded studies that included children with other debilitating diseases, severe wasting (weight for height z-score below -3), asthma or other respiratory diseases, hospital-acquired pneumonia and postoperative conditions.

### Types of interventions

The intervention was commenced after the diagnosis of CAP had been made. The interventions consisted of treatment with vitamin D as an adjunct to antibiotics and other supportive measures. The trials compared vitamin D with placebo only. We considered any dose schedule (low versus high dose, daily dose versus bolus dose, any duration) and route (oral or injection) of vitamin D.

### Types of outcome measures

Outcome measures frequently used to determine the clinical efficacy of any acute pneumonia treatment are time to resolution of illness, duration of hospitalisation, treatment failure, adverse events, or death.

#### Primary outcomes

1. Time to resolution of acute illness.
2. Duration of hospitalisation.

#### Secondary outcomes

1. Time to resolution of:
  - a. tachypnoea;
  - b. lower chest wall indrawing;
  - c. hypoxia (blood oxygen saturation < 95%);
  - d. fever;
  - e. inability to feed, or lethargy.
2. Treatment failure rate.
3. Mortality rate.
4. Adverse events (if any).

We defined the time to resolution of acute pneumonia as the time period to achieve the following parameters from initiation of treatment: respiratory rate less than the age-specific cut-offs, no lower chest indrawing, no danger signs or hypoxia, and ability to feed. These parameters were present for at least two consecutive days or 48 hours. We defined the duration of hospitalisation as the time period between study enrolment and discharge. We defined treatment failure as no reduction in tachypnoea over a 48-hour period compared to that detected at enrolment.

### Search methods for identification of studies

#### Electronic searches

We searched the Cochrane Central Register of Controlled Trials (CENTRAL) 2017, Issue 7, part of the Cochrane Library (accessed 28 July 2017), which contains the Cochrane Acute Respiratory Infections (ARI) Group's Specialised Register; Ovid MEDLINE(R) Epub Ahead of Print, In-Process & Other Non-Indexed Citations, Ovid MEDLINE(R) Daily and Ovid MEDLINE(R) (1946 to July Week 4, 2017); and Embase (2010 to 28 July 2017).

We used the search strategy in [Appendix 1](#) to search CENTRAL and MEDLINE. The MEDLINE search was combined with the Cochrane Highly Sensitive Search Strategy for identifying randomised trials in MEDLINE: sensitivity- and precision-maximising version (2008 revision); Ovid format ([Lefebvre 2011](#); [Schünemann 2011](#)). The search strategy was adapted to search Embase ([Appendix 2](#)).



## Searching other resources

We searched the US National Institutes of Health Ongoing Trials Register ClinicalTrials.gov ([www.clinicaltrials.gov](http://www.clinicaltrials.gov)) and the WHO International Clinical Trials Registry Platform (ICTRP) ([www.who.int/ictcp/en](http://www.who.int/ictcp/en)) on 28 July 2017. We searched the *Journal of Orthomolecular Medicine* (1986 to July 2017), which focuses on dietary supplements for illness. We contacted researchers in the relevant field to identify additional studies that may be eligible for inclusion. We also checked the reference lists of all studies identified.

## Data collection and analysis

### Selection of studies

Two review authors (RRD, SSN) independently assessed records for possible inclusion in the review. The review authors screened titles and abstracts to identify potentially relevant studies. When we could not ascertain relevance from the titles or abstracts, we retrieved the full text for assessment. Two review authors (RRD, SSN) independently assessed the eligibility of potentially relevant studies by evaluating the full-text reports and completing eligibility forms designed in accordance with the specified inclusion criteria. Any differences in opinion were resolved through discussion or by consulting the third review author (MS) when necessary. We described excluded studies in [Characteristics of excluded studies](#) tables, along with reasons for their exclusion. In the case of conference abstracts, if additional data were not forthcoming, we used the information provided in the abstract. We contacted trial authors where we required clarification or additional information.

### Data extraction and management

We extracted the following information from each study: author; year; location (country); setting (hospital or community); method of recruitment; inclusion criteria; unit of analysis; allocation ratio; risk of bias; participants (age, sex, sample size, pneumonia or severe pneumonia); intervention (dosage, duration, frequency, and co-intervention if any); outcomes (outcome definition, valid unit of measurement, time points of collection and reporting); loss to follow-up; and miscellaneous (key conclusions, references to other relevant trials, and additional data required). We designed and pilot tested the data extraction form prior to extraction of data.

Two review authors (RRD, SSN) independently extracted data from the studies. Any discrepancies were resolved through discussion with the third review author (MS). For dichotomous data (such as mortality and adverse events), we extracted the number of participants experiencing the condition and the total number of participants in each treatment group. For continuous data (such as time period and duration of illness), we used the arithmetic mean and standard deviation (SD) for each treatment group together with the number of participants in each group. If a standard error (SE) was reported, we converted it to an SD. If data were reported using geometric means, we recorded this information and extracted the SD on the log scale. If medians were used, we extracted the median and the range. If a 95% confidence interval (CI) was provided instead of a mean and SD for continuous data, we extracted the mean and SD from the 95% CI.

### Assessment of risk of bias in included studies

Two review authors (RRD, SSN) independently assessed methodological quality using Cochrane's 'Risk of bias' tool ([Higgins](#)

2011). We assessed included studies based on the following six components: method of sequence generation; allocation concealment; blinding of participants (providers and outcome assessors); incomplete outcome data; selective outcome reporting; and other biases (including detection bias, e.g. differential effort to locate death records for the intervention and control groups). We presented findings in 'Risk of bias' tables with the review authors' judgement (low, high, or unclear risk of bias) and support for the judgement. We also reported the results in a 'Risk of bias' graph. We attempted to contact the trial authors for clarification where data were missing or we needed clarification about reporting. A third review author (MS) resolved any disagreements between the review authors. Further details about the 'Risk of bias' tool are provided in the *Cochrane Handbook of Systematic Reviews of Interventions* ([Higgins 2011](#)). We looked for sources of bias originating from differences between individual RCTs and cluster-RCTs (e.g. the relationship between allocation concealment and recruitment bias may be greater in cluster-RCTs).

### Measures of treatment effect

We extracted and entered outcome data into Review Manager 5 software for statistical analysis ([Review Manager 2014](#)). We used the standard methods of the Cochrane Acute Respiratory Infections (ARI) Review Group to synthesise data. Outcomes with continuous data included time to resolution of acute illness, duration of hospitalisation, and duration of resolution of tachypnoea, chest retractions, hypoxia, fever, inability to feed or lethargy. For continuous data, we calculated mean differences (MD) with 95% CI to estimate the treatment effect. Where data were provided as medians and ranges, we presented these in a tabular form. Outcomes with dichotomous data included treatment failure rate, mortality rate, and adverse events rate. For dichotomous data, we reported the proportions of participants with these events, but were unable to calculate a pooled estimate of the treatment effect for each outcome using risk ratio (RR) with 95% CI due to insufficient numbers of included trials.

### Unit of analysis issues

In studies randomising units other than individuals (i.e. clusters), results should be presented with controls for clustering (e.g. robust SEs or hierarchical linear models). We did not include any cluster-randomised trials in our review, so we had no unit of analysis issues.

### Dealing with missing data

We described missing data, including dropouts. Differential dropout rates can lead to biased estimates of the effect size, and bias may arise if the reasons for dropping out differ across groups. We reported reasons children dropped out of studies. If data were missing, or if reasons for dropping out were not reported, we contacted the trial authors for further information. We performed quantitative analysis on an intention-to-treat basis. When analyses were reported for completers as well as controlling for dropout (e.g. imputed using regression methods), we extracted the latter.

### Assessment of heterogeneity

We assessed included studies for clinical, methodological, and statistical heterogeneity. We assessed clinical heterogeneity by comparing the distribution of important factors such as study participants, study setting, dose and duration of the intervention, and co-interventions. We evaluated methodological

heterogeneity on the basis of factors such as the method of sequence generation, allocation concealment, blinding of outcome assessment, and losses to follow-up. We used the Chi<sup>2</sup> test for statistical heterogeneity between studies and considered  $P \leq 0.10$  as indicating significant heterogeneity. We used the I<sup>2</sup> statistic to assess the magnitude of heterogeneity (Higgins 2011). We considered an I<sup>2</sup> statistic above 50% to indicate problematic heterogeneity between studies and carefully considered the value of any pooled analysis.

### Assessment of reporting biases

We undertook a comprehensive electronic search and searched trial registries to minimise the effects of publication bias. Had we identified 10 or more studies, we planned to create funnel plots of effect estimates against their standard errors (on a reversed scale) using Review Manager 5 (Review Manager 2014). However, we included only seven trials in the review.

### Data synthesis

We carried out statistical analysis using Review Manager 5 software (Review Manager 2014). We performed meta-analyses using a random-effects model due to the presence of clinical heterogeneity. We did not include any cluster-RCTs. We calculated overall effects using inverse variance methods. We presented results as the average treatment effect with 95% CI and Tau<sup>2</sup> and I<sup>2</sup> estimates.

### 'Summary of findings' table and GRADE

We created a 'Summary of findings' table using the following outcomes: time to resolution of acute illness, duration of hospitalisation, and mortality rate. We used the five GRADE considerations (study limitations, consistency of effect, imprecision, indirectness, and publication bias) to assess the quality of a body of evidence as it relates to the studies which contributed data to the meta-analyses for the prespecified outcomes (Atkins 2004). We used methods and recommendations described in Section 8.5 and Chapter 12 of the *Cochrane Handbook for Systematic Reviews of Interventions* (Schünemann 2011), employing GRADEpro GDT software (GRADEpro GDT 2015). We justified all decisions to downgrade the quality of studies using footnotes, and made comments to aid the reader's understanding of the review where necessary. Where it was not possible to carry out a meta-analysis, we summarised the data for each trial.

### Subgroup analysis and investigation of heterogeneity

If we identify more trials for inclusion in future updates of this review, we will attempt to conduct subgroup analyses and investigate sources and causes of heterogeneity. Possible subgroup analyses would involve different dose schedules of vitamin D (low or high dose, single bolus or daily dose, oral or parenteral route, duration < 5 or > 5 days); population (aged up to 1 year or 1 to 5 years, boys or girls); study setting (hospital or community, outpatient or inpatient, low-, middle-, or high-income countries); aetiology (bacterial or viral pneumonia); and severity of illness (pneumonia, severe, or very severe pneumonia).

### Sensitivity analysis

In future updates as more data become available, we will carry out sensitivity analysis to explore the effect of trial quality on results. This will involve analyses excluding studies at risk of bias (selection, performance, detection, attrition, or reporting) to assess for any substantive difference to the overall result. We will investigate the effects of the randomisation unit (individual versus cluster) on the outcomes. We will explore the effects of fixed-effect or random-effects analyses for outcomes with statistical heterogeneity, and the effects of any assumptions made such as the value of the intracluster correlation coefficient (ICC) used for cluster-randomised trials. We will use primary outcomes in sensitivity analyses.

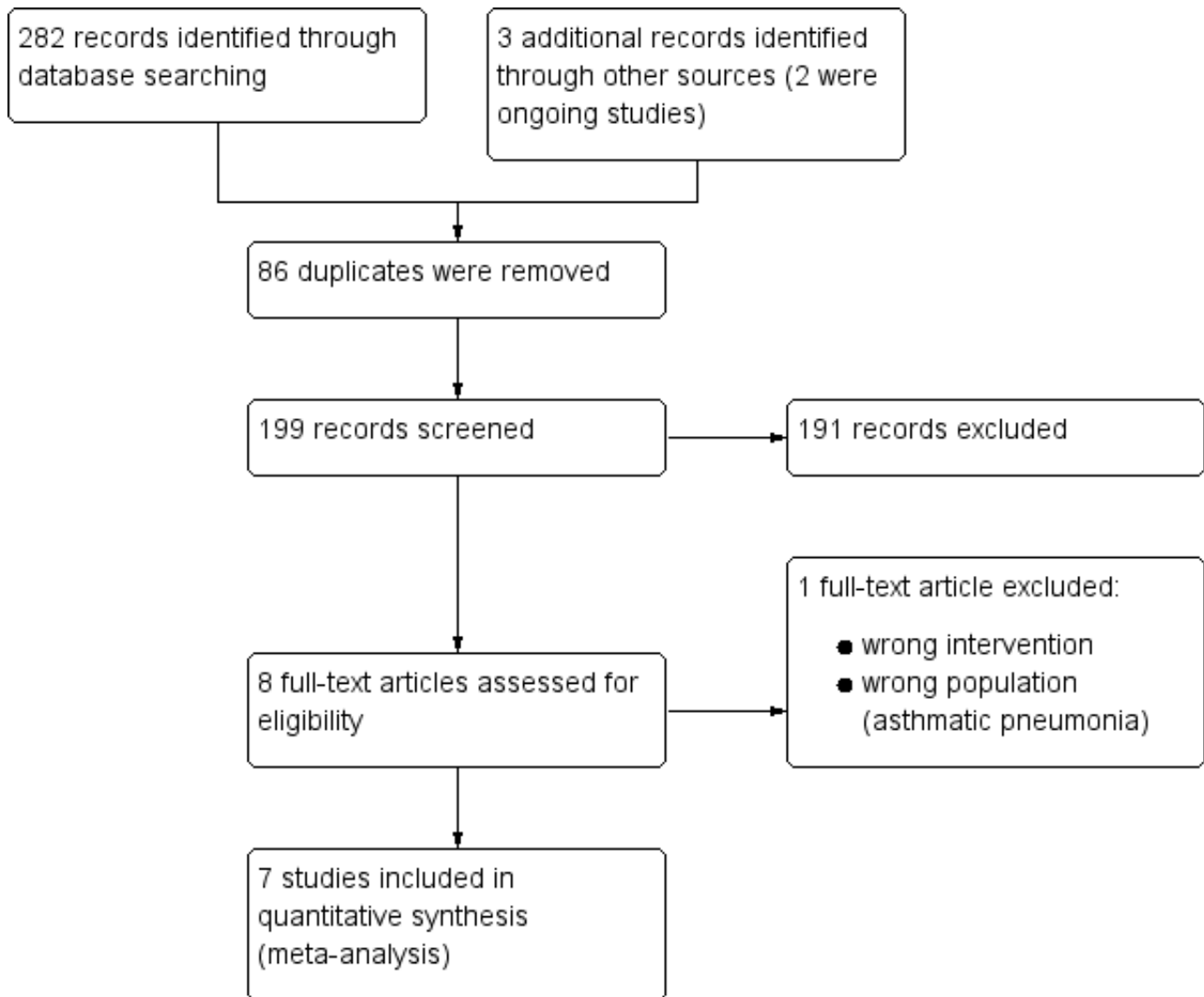
## RESULTS

### Description of studies

#### Results of the search

Our searches in July 2017 identified 285 records, of which two were ongoing trials. After removal of 86 duplicate records, we screened 199 records by title and abstract. We excluded 191 records that did not match our inclusion criteria. We obtained eight full-text studies for assessment. We included seven studies (Choudhary 2012; Dhungel 2015; Gupta 2016a; Manaseki-Holland 2010; Rahmati 2016; Rajshekhar 2016; Somnath 2017), and excluded one study that did not meet our inclusion criteria (Zhou 2016). We identified two ongoing trials (NCT02054182; NCT02185196). See Figure 1.

**Figure 1. Study flow diagram.**



**Included studies**

See [Characteristics of included studies](#).

We included seven studies involving a total of 1529 children (780 with pneumonia and 749 with severe or very severe pneumonia) (Choudhary 2012; Dhungel 2015; Gupta 2016a; Manaseki-Holland 2010; Rahmati 2016; Rajshekhar 2016; Somnath 2017). None of the included studies were conducted in high-income countries. Childrens' characteristics, aetiological agents, and prevalence of vitamin D deficiency appeared to be very similar among studies. The studies differed in terms of inclusion criteria and dose and duration of vitamin D use. One study was funded by the New Zealand Aid Corporation; one was funded by an institutional grant; and five studies were unfunded.

Choudhary 2012 was a single-centre study conducted in India. The study objective was to determine the role of oral vitamin D supplementation for resolution of severe pneumonia in children. The study enrolled 200 children (100/100 intervention/control) aged between two months and five years, who were hospitalised with WHO-defined severe pneumonia. History of recurrent pneumonia was present in 30% of children in the vitamin

D group and 33% of children in the control group. Five children had clinical evidence of rickets. A third of the children had wheezing at enrolment.

Children in the intervention arm received oral vitamin D<sub>3</sub> for five days at doses of 1000 IU for children aged up to one year and 2000 IU for children aged over one year. Children in the placebo arm received lactose alone. The children received antibiotics in accordance with the Indian Academy of Pediatrics (IAP) 2006 guidelines (IAP 2006). The guidelines recommend: infants aged up to three months: cefotaxime/ceftriaxone ± gentamicin/amikacin as first-line therapy and co-amoxyclav + gentamicin/amikacin as second-line treatment; children aged from three months to five years: ampicillin/chloramphenicol or ampicillin + chloramphenicol or co-amoxyclav as first-line therapy and co-amoxyclav or cefotaxime/ceftriaxone as second-line treatment. In children with staphylococcal infection, adjuvant therapy with cloxacillin as first-line therapy and vancomycin/teicoplanin as second-line therapy are to be added to age-specific treatment combinations.

Primary study outcomes assessed were time to resolution of severe pneumonia (absence of lower chest indrawing, hypoxia, cyanosis, lethargy and inability to feed). Secondary outcomes assessed were duration of hospitalisation, time to resolution of tachypnoea, chest retraction, and inability to feed.

**Dhungeel 2015** was a single-centre study conducted in Pakistan. The study objective was to determine if supplementation of oral vitamin D<sub>3</sub> with antibiotics reduced duration of hospitalisation in children with pneumonia. The study enrolled 200 children (100/100, intervention/control) aged between two months and 60 months, who were hospitalised with WHO-defined acute pneumonia. Among children randomised, severe pneumonia was present in 43%; the remainder had pneumonia. Children with very severe pneumonia and those with normal vitamin D levels were excluded. All children received antibiotics as per the hospital's policy (cefotaxime, ceftriaxone, ampiclox, clarithromycin, amikacin, vancomycin. The antibiotics were given as monotherapy or combination therapy).

Children in the intervention arm received a single dose (100,000 IU) of intramuscular vitamin D<sub>3</sub> within 24 hours of admission. Children in the control arm received antibiotics alone. Study outcomes assessed were duration of hospitalisation and repeat episodes of pneumonia.

**Gupta 2016a** was a single-centre study conducted in India. The study objective was to determine the role of oral vitamin D supplementation for the treatment and prevention of pneumonia in children. The study enrolled 324 children (162/162, intervention/control) aged between six months and five years, who were hospitalised with WHO-defined severe pneumonia. The prevalence of moderate malnutrition according to the WHO definition was 21.3%. The prevalence of anaemia (haemoglobin < 11 g/dL) was 82.4%; hypocalcaemia was 55.9%; hypophosphataemia was 31.4%. Vitamin D deficiency was present in 37.6% of children in the vitamin D group and 40.1% of children in the placebo group. Blood culture was positive in 8.6% (*Staphylococcus aureus* isolated in 8.3%). Around 90.1% of children had abnormal chest x-ray findings (consolidation, bilateral patchy opacities and hyperinflation or minor infiltrates). Wheezing was present in 84.6% of children in the vitamin D group and 78.4% of children in the control group.

Children in the intervention arm received a single oral dose of vitamin D<sub>3</sub> of 100,000 IU on the day of enrolment. Children in the placebo arm received a product similarly prepared and given as a single dose on day of enrolment. The children received antibiotics according to IAP 2006 guidelines (**IAP 2006**) (as specified for the study by **Choudhary 2012**).

Primary study outcomes assessed were time to resolution of severe pneumonia (duration from enrolment to chest indrawing was no longer present and maintained for 24 hours) and proportion of children with recurrent pneumonia over the next six months. Secondary outcomes assessed were change in serum level 25(OH)-D and parathyroid hormone after two weeks and three months of therapy; change in serum level of cathelicidin and immunoglobulins after two weeks of therapy; duration of hospitalisation; time to complete recovery from pneumonia (normalisation of respiratory rate), fever clearance time; and repeat episodes of pneumonia.

**Manaseki-Holland 2010** was a single-centre study conducted in Afghanistan. The study objective was to determine if supplementation of oral vitamin D<sub>3</sub> with antibiotics reduced duration of illness in children with pneumonia. The study enrolled 453 children (224/229, intervention/control) aged between one month and 36 months with WHO-defined acute pneumonia treated in outpatient and inpatient settings. Among children randomised, severe pneumonia was present in 17.4% of intervention group and 15.3% of placebo group children; the remaining children had pneumonia. The study excluded children with very severe pneumonia. All children received antibiotics as per the Integrated Management of Childhood Illness (IMCI) guidelines (children aged up to two months: intramuscular gentamicin and intramuscular benzylpenicillin; children aged from two months to five years: oral amoxicillin or intramuscular chloramphenicol (if unable to take orally)).

We included this study after reaching consensus on the following points: although **Manaseki-Holland 2010** did not comply with all predefined inclusion criteria (not all children were treated as inpatients), we decided to include the study because some study participants had pneumonia. We felt that exclusion of the study (which included children with pneumonia and severe pneumonia) may have introduced bias.

Children in the intervention arm received a single dose (100,000 IU) of oral vitamin D<sub>3</sub> at onset of pneumonia. Children in the placebo arm received olive oil alone. All children received antibiotics according to the national pneumonia treatment protocol based on IMCI guidelines.

Study outcomes assessed were time to resolution of pneumonia or recovery for 48 consecutive hours, treatment failure, and pneumonia recurrence.

**Rahmati 2016** was a single-centre study conducted in Iran. The study objective was to determine the role of oral vitamin D supplementation using the respiratory index of severity in children hospitalised with community-acquired pneumonia. The study enrolled 100 children (50/50, intervention/placebo) aged between two months and six years who were hospitalised with pneumonia. Pneumonia was defined as the presence of fever, cough, audible findings in lungs (rale), respiratory distress symptoms, and chest x-ray changes. No children had severe pneumonia.

Children in the intervention arm received oral vitamin D<sub>3</sub> 50,000 IU per day for two days. Children in the placebo arm received olive oil. Children may have received ceftriaxone as the only antibiotic. We contacted the study authors to confirm that all children received ceftriaxone, but received no response (**Das 2017**). Primary study outcomes assessed were duration of antibiotic therapy, hospitalisation, and duration of fever.

**Rajshekhhar 2016** was a single-centre study conducted in India. The study objective was to determine the role of oral vitamin D supplementation for resolution of severe pneumonia. The study enrolled 96 children (48/48, intervention/control) aged between two months and five years who were hospitalised with WHO-defined severe pneumonia. History of recurrent pneumonia was present in 35.4% of children in the vitamin D group and 39.6% of children in the control group.

Children in the intervention arm received oral vitamin D<sub>3</sub> for five days at doses of 1000 IU for children aged up to one year and 2000 IU for children aged over one year. Children in the placebo arm received lactose alone. The children received antibiotics as per the IAP 2012 guidelines (amoxicillin, amoxyclav, third-generation parenteral cephalosporins, and macrolides) (Ghosh 2012).

Study outcomes assessed were time to resolution of severe pneumonia (absence of lower chest indrawing, hypoxia, cyanosis, and inability to feed).

Somnath 2017 was a single-centre study conducted in India. The study objective was to determine the role of oral vitamin D supplementation on pneumonia in children. The study enrolled 156 children (78/78, intervention/placebo) aged between two months and five years who were hospitalised with WHO-defined pneumonia. Very severe pneumonia was present in 12.8% (N = 10) of children in the vitamin D group and 14.3% (N = 11) of children in the control group.

Children in the intervention arm received a single dose of oral vitamin D<sub>3</sub> 100,000 IU administered on Day 1 of admission as cholecalciferol granules mixed with water or milk. Children in the placebo arm received standard therapy. The children

received antibiotics according to practice guidelines at the study centre. The primary study outcome assessed was duration of hospitalisation, and secondary outcomes were time to resolution of fever, complications associated with pneumonia, need for transfer to the paediatric intensive care unit, mortality, and pneumonia recurrence in the next six months.

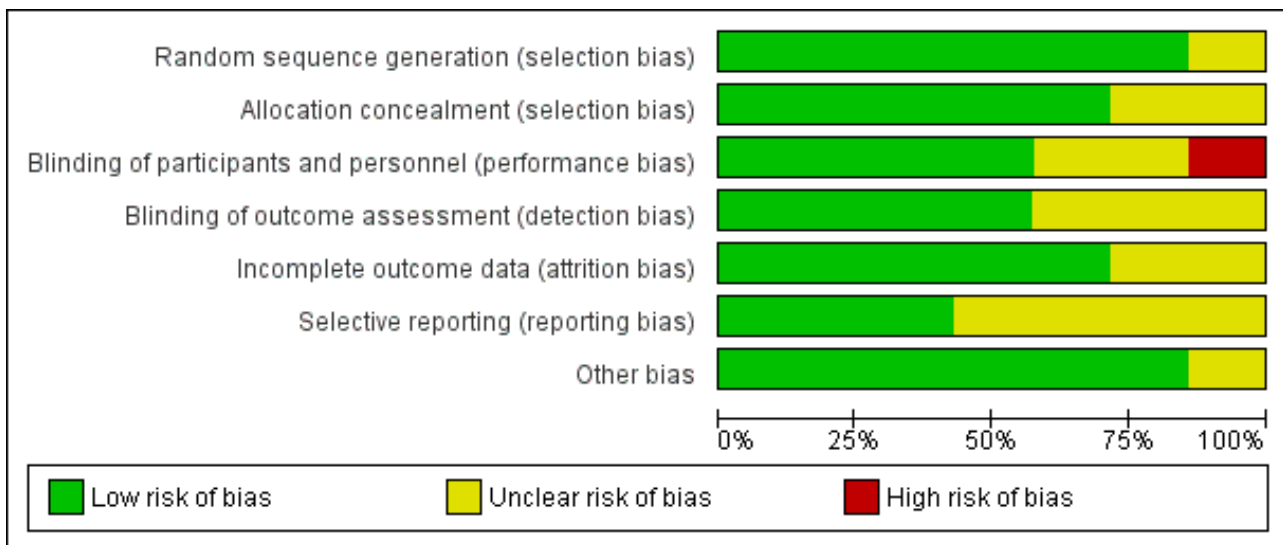
**Excluded studies**

We excluded one study (Zhou 2016). This study supplemented nutrients (vitamin A, vitamin D, zinc, and iron) for children with pneumonia who were nutritionally deficient for these vitamins and minerals. Moreover, the study included children with asthmatic pneumonia.

**Risk of bias in included studies**

We considered overall risk of bias to be low for four studies (Choudhary 2012; Gupta 2016a; Manaseki-Holland 2010; Somnath 2017), and unclear for three (Dhungel 2015; Rahmati 2016; Rajshekhar 2016). See Figure 2 and Figure 3. We contacted the authors of studies by Rajshekhar 2016 and Rahmati 2016 respectively for clarification regarding methods of blinding and allocation concealment and clarification of interventions used (Das 2017; Rajshekhar 2016 [pers comm]).

**Figure 2. Risk of bias graph: review authors' judgements about each risk of bias item presented as percentages across all included studies.**





**Figure 3. Risk of bias summary: review authors' judgements about each risk of bias item for each included study.**

	Random sequence generation (selection bias)	Allocation concealment (selection bias)	Blinding of participants and personnel (performance bias)	Blinding of outcome assessment (detection bias)	Incomplete outcome data (attrition bias)	Selective reporting (reporting bias)	Other bias
Choudhary 2012	+	+	+	+	+	?	+
Dhungel 2015	?	?	?	?	?	?	+
Gupta 2016a	+	+	+	+	+	+	+
Manaseki-Holland 2010	+	+	+	?	+	?	+
Rahmati 2016	+	?	?	?	+	+	+
Rajshekhar 2016	+	+	+	+	?	?	?
Somnath 2017	+	+	-	+	+	+	+

**Allocation**

Five studies used sequentially numbered, sealed envelopes or codes for allocation of participants to the two groups and were assessed as at low risk of bias (Choudhary 2012; Gupta 2016a; Manaseki-Holland 2010; Rajshekhar 2016; Somnath 2017). Two studies did not mention the method of allocation and were assessed as at unclear risk of bias (Dhungel 2015; Rahmati 2016).

**Blinding**

In four studies blinding of participants and carers was conducted, and it was unlikely that the blinding could have been broken as both interventions and placebos were similar in terms of appearance, taste, and colour (Choudhary 2012; Gupta 2016a; Manaseki-Holland 2010; Rajshekhar 2016). In three studies the code key was opened only after administration of the intervention,

data collection, follow-up, and tabulation were completed, thus preventing detection bias (Choudhary 2012; Gupta 2016a; Rajshekhar 2016); these studies were assessed as at low risk of bias for this domain. Dhungel 2015 and Manaseki-Holland 2010 did not describe methods to prevent detection bias and were assessed as at unclear risk of bias. Participants and personnel in Somnath 2017 were unblinded, presenting a high risk of performance bias, however outcome assessment was blinded, presenting a low risk of detection bias.

### Incomplete outcome data

We assessed five studies as at low risk of bias for this domain (Choudhary 2012; Gupta 2016a; Manaseki-Holland 2010; Rahmati 2016; Somnath 2017). Choudhary 2012 reported an attrition rate of 4.5%; Gupta 2016a reported an attrition rate of 4.63%; Manaseki-Holland 2010 reported an attrition rate of 3%; Rahmati 2016 reported no attrition; and Somnath 2017 reported an attrition rate of 1.3% before outcomes were measured. Manaseki-Holland 2010 and Gupta 2016a analysed data using the intention-to-treat principle. Dhungel 2015 and Rajshekhar 2016 did not provide this information and were assessed as at unclear risk of bias.

### Selective reporting

We assessed three studies as at low risk of reporting bias, as they were registered in the clinical trial registry (Gupta 2016a; Rahmati 2016; Somnath 2017). For the remaining four studies, protocols were neither available nor were they registered in clinical trials registries, therefore these studies were assessed as at unclear risk of bias.

### Other potential sources of bias

There were no other potential sources of bias in the included studies.

### Effects of interventions

See: [Summary of findings for the main comparison Vitamin D compared with placebo for acute childhood pneumonia](#)

### Primary outcomes

#### 1. Time to resolution of acute illness

Four studies reported time to resolution of acute illness (hours) (Choudhary 2012; Gupta 2016a; Manaseki-Holland 2010; Rajshekhar 2016). Data were provided as mean and standard deviation (SD) in Manaseki-Holland 2010 and Gupta 2016a, and as median and interquartile range (IQR) in Choudhary 2012. We contacted the corresponding author of Choudhary 2012 to request data in mean and SD, but these data were unavailable (Gupta 2016b). We therefore calculated the corresponding mean and SD as per the available statistical methods (Hozo 2005; Luo 2018; Wan 2014). The pooled result from three studies showed no significant difference for time to resolution of acute illness (mean difference (MD) -0.95, 95% confidence interval (CI) -6.14 to 4.24; N = 935; low-quality evidence; Analysis 1.1) (Choudhary 2012; Gupta 2016a; Manaseki-Holland 2010). Rajshekhar 2016 presented data as proportions for three time periods (24 hours, 24 to 48 hours, and > 48 hours); there was no difference between the vitamin D and control group in the time to resolution of severe pneumonia.

#### 2. Duration of hospitalisation

Five studies reported duration of hospitalisation (hours) (Choudhary 2012; Dhungel 2015; Gupta 2016a; Rahmati 2016; Somnath 2017). Data were provided as mean and SD in Dhungel 2015 and Gupta 2016a, and as median and IQR in Choudhary 2012 and Somnath 2017. We contacted the corresponding author of Choudhary 2012 to request data in mean and SD, but these data were unavailable (Gupta 2016b). We calculated the corresponding mean and SD as per the available statistical methods (Hozo 2005; Luo 2018; Wan 2014). In Rahmati 2016, the duration of hospitalisation was shorter in the intervention group compared to the control group (441.6 hours versus 535.2 hours; P < 0.05). Because the duration was too long for a pneumonia episode, we contacted the corresponding author to seek clarification, but did not receive a response (Das 2017). The pooled result from four studies (except Rahmati 2016) showed no significant difference in duration of hospitalisation (MD 0.49, 95% CI -8.41 to 9.40; N = 835; very low-quality evidence; Analysis 2.1).

### Secondary outcomes

Choudhary 2012 reported the following outcomes for children with severe pneumonia as median and IQR. We contacted the corresponding author of Choudhary 2012 to request data in mean and SD, but these data were unavailable (Gupta 2016b). We therefore calculated the corresponding mean and SD as per the available statistical methods (Hozo 2005; Luo 2018; Wan 2014). We assessed the quality of the evidence for all outcomes as very low.

#### 1a. Time to resolution of tachypnoea

The pooled result showed no significant difference in the time to resolution of tachypnoea (MD 4.93, 95% CI -6.15 to 16.01; N = 171; very low-quality evidence; Analysis 3.1).

#### 1b. Time to resolution of lower chest wall indrawing

The time to resolution of lower chest indrawing was the same in both groups (MD 0.00, 95% CI -10.78 to 10.78; N = 173; very low-quality evidence; Analysis 4.1).

#### 1c. Time to resolution of hypoxia

The time to resolution of hypoxia was the same in both groups (MD 0.00, 95% CI -3.59 to 3.59; N = 173; very low-quality evidence; Analysis 5.1).

#### 1d. Time to resolution of fever

Four studies reported time to resolution of fever (hours) (Choudhary 2012; Gupta 2016a; Rahmati 2016; Somnath 2017). The pooled result showed no significant difference between groups in time to resolution of fever (MD 1.66, 95% CI -2.44 to 5.76; N = 584; very low-quality evidence; Analysis 6.1).

#### 1e. Time to resolution of inability to feed or lethargy

The pooled result showed a significantly lesser time to resolution of inability to feed or lethargy in the control group (MD 8.00, 95% CI 0.60 to 15.40; N = 173; very low-quality evidence; Analysis 7.1).

### 2. Treatment failure rate

Although the failure rate after 48 hours of treatment was not specified, Choudhary 2012 reported that 23 (11.5%) children remained in the same condition, and four (2%) children had their



condition worsen. There was no description of the groups or the time period. Given that data for this outcome were from a single study, we assessed the quality of evidence as very low.

### 3. Mortality rate

Two studies reported mortality (Choudhary 2012; Somnath 2017). Choudhary 2012 reported that two children died, one each in the vitamin D and placebo groups. There was no significant difference between intervention and placebo groups (risk ratio 0.97, 95% CI 0.06 to 15.28; N = 193; very low-quality evidence; Analysis 8.1). Somnath 2017 reported no deaths in either group.

Manaseki-Holland 2010 reported three deaths (two in the vitamin D group and one in the placebo group) during the 90-day follow-up evaluation of repeated episodes of pneumonia. Because it was not clear if the children died during treatment of the index episode of pneumonia, we contacted the corresponding author (Manaseki-Holland 2010 [pers comm]). It was confirmed that the children died during the 90-day follow-up period (not during the index episode). Accordingly, we did not include mortality data from Manaseki-Holland 2010 in the analysis.

### 4. Adverse events

Choudhary 2012 reported two minor events: one child experienced vomiting and another had diarrhoea in the vitamin D group. Given that data for this outcome were from a single study, we assessed the quality of evidence as very low.

## DISCUSSION

### Summary of main results

We included seven studies involving a total of 1529 children (780 with pneumonia and 749 with severe or very severe pneumonia) (Choudhary 2012; Dhungel 2015; Gupta 2016a; Manaseki-Holland 2010; Rahmati 2016; Rajshekhar 2016; Somnath 2017). None of the included studies were conducted in high-income countries. Childrens' characteristics, aetiological agents, and prevalence of vitamin D deficiency appeared to be very similar among studies. The studies differed in terms of inclusion criteria and dose and duration of vitamin D use. One study was funded by the New Zealand Aid Corporation; one was funded by an institutional grant; and five studies were unfunded.

With the exceptions of time to resolution of acute illness, duration of hospitalisation, time to resolution of fever, and mortality rate, results could not be pooled; all secondary outcomes were reported by Choudhary 2012. Gupta 2016a, Rahmati 2016, and Somnath 2017 provided data regarding duration of fever only. The appropriate effect measure for time-to-event data (time to resolution of acute illness) is hazard ratio, but the studies reported this information as continuous data. Except for Rahmati 2016, none of the studies reported vitamin D benefits for children with acute pneumonia. Dhungel 2015 and Manaseki-Holland 2010 included children with both non-severe and severe pneumonia, but neither study reported outcomes separately for children with severe pneumonia. Though mortality rate is an important outcome for policy decisions, this outcome was reported in only two studies (Choudhary 2012; Somnath 2017). See [Summary of findings for the main comparison](#).

We questioned whether vitamin D was ineffective in treating children with acute pneumonia due to real therapeutic

ineffectiveness of vitamin D, or improper dose or mode of administration, or type or severity of the underlying illness. These are important aspects needing investigation in future studies.

The included studies differed significantly in vitamin D dose. Two studies administered oral vitamin D in a total dose of 5000 IU to 10,000 IU according to the child's age, over five days (Choudhary 2012; Rajshekhar 2016). Four studies administered a single bolus dose of 100,000 IU either at the onset of illness or within 24 hours of admission (Dhungel 2015; Gupta 2016a; Manaseki-Holland 2010; Somnath 2017). Rahmati 2016 administered 50,000 IU per day for two days. As discussed previously, a daily vitamin D supplementation has better therapeutic efficacy than large bolus doses for pneumonia, tuberculosis, and fracture in the elderly (Heaney 2012; Hollis 2011; Manaseki-Holland 2012; Martineau 2012; Nielsen 2010). Clinical studies have shown that at a higher dose, vitamin D could be immunosuppressive, and the immunological effectiveness of vitamin D is blunted at extremes of dosages (either a very high or a low therapeutic dose) (Kimball 2011; Manaseki-Holland 2010; Manaseki-Holland 2012). This was also supported by the findings of Choudhary 2012, who used a daily dose of 1000 IU to 2000 IU per day of vitamin D, but did not document any benefit. In Choudhary 2012, a low dose might possibly have been used considering the high prevalence of vitamin D deficiency along with severe illness (severe pneumonia) in the studied population. The ideal situation would have been correlation with vitamin D levels measured before and after supplementation. Two studies reported vitamin D level both pre- and post supplementation, but failed to show any benefit of vitamin D (Gupta 2016a; Somnath 2017).

The National Institutes of Health defines no observed adverse effect level (2400 IU) and lowest observed adverse effect level (3800 IU) of vitamin D (NIH Vitamin D Fact Sheet). Future trials may consider use of higher doses with monitoring of toxicity and serum vitamin D level. It is premature to hypothesise about the real therapeutic ineffectiveness of vitamin D because clinical trials are ongoing and results are not yet available.

Although the included studies were conducted in low-income countries (Afghanistan, Pakistan, Iran, and India), prevalence of malnutrition (moderate or grade II) was reported in only two studies (Dhungel 2015; Gupta 2016a). However, these two studies did not report the prevalence of malnutrition in either vitamin D or control groups, leading to difficulty in correlating the effect of vitamin D in malnourished children. This is an important aspect because underlying malnutrition may affect the state of immunity and hence blunt the vitamin D effect. Malnutrition also lowers serum vitamin D level. Furthermore, severe pneumonia may establish a systemic inflammatory response in the host that causes a decrease in the serum vitamin D level (Bang 2011). It would therefore be prudent to study the level of inflammatory markers along with vitamin D levels in children with non-severe/severe pneumonia.

None of the included studies reported the aetiology of pneumonia. This is important for any differential therapeutic effect of vitamin D (if any) in bacterial or viral pneumonia. Because there is a seasonal variation in the predominance of aetiological agents (e.g. viruses dominating over bacteria in the rainy season), any trial conducted over a year would be meaningful. This is also important because vitamin D levels may decrease during winter and the rainy season due to less sun exposure. One study was conducted over a 29-month period (Gupta 2016a); two were conducted over a 12-month

period (Rahmati 2016; Somnath 2017); two were conducted over a three-month period (Dhungel 2015; Manaseki-Holland 2010); and two did not mention the study duration (Choudhary 2012; Rajshekhar 2016).

Viral pneumonia is commonly associated with wheezing. Manaseki-Holland 2010 and Rahmati 2016 excluded children with wheezing, but almost a third of children in Choudhary 2012 and more than 80% of children in Gupta 2016a had wheezing. This finding was not reported by the other three included studies (Dhungel 2015; Rajshekhar 2016; Somnath 2017). In Choudhary 2012 and Rajshekhar 2016 almost a third of children in each study arm had recurrent pneumonia, which indicates some underlying immune dysfunction or other problem predisposing to recurrence. Vitamin D alone may not correct this, resulting in no observable clinical effect or benefit, or both. Outcome measures may also have been affected by the timing of vitamin D administration in the course of pneumonia (early versus late). Any information regarding the duration of illness prior to initiating vitamin D would be meaningful and should be included in future studies. Additionally, information regarding prior antibiotic use should be included because this can affect recovery from illness. Unfortunately, this information was not provided in any of the studies. Similarly, a large proportion of children needing second-line or broader spectrum antibiotics due to severe illness would lead to a reduction in study power, making it difficult to detect any difference attributable to vitamin D.

### Overall completeness and applicability of evidence

We included seven studies, but could pool results for three outcomes (time to resolution of illness, duration of hospitalisation, and mortality rate); most outcomes were reported in one study. None of the included studies showed any beneficial effect of vitamin D for children with acute pneumonia. Our results may not be applicable to children with acute pneumonia in high-income countries, those with HIV or severe malnutrition, and neonates (aged up to a month). Two large ongoing trials may provide additional evidence.

### Quality of the evidence

We considered the overall risk of bias to be low in four studies (Choudhary 2012; Gupta 2016a; Manaseki-Holland 2010; Somnath 2017), and unclear in three studies (Dhungel 2015; Rahmati 2016; Rajshekhar 2016). The GRADE assessment of the evidence was very-low quality for all outcomes except time to resolution of acute illness (low-quality evidence). See [Summary of findings for the main comparison](#). Our assessment of low quality was due to no serious study limitations (all studies had adequately concealed allocation and blinding), no serious inconsistency (low and non-significant statistical heterogeneity), serious indirectness (because of non-generalisation of study findings to all situations, and differing characteristics of the study population), and serious imprecision (wide 95% CI around the pooled effect, inclusion of no effect, and the upper or lower confidence limit crossed the minimal important difference for benefit). Our assessment of very low quality was due to serious study limitations (one study had unclear allocation concealment and blinding), no serious inconsistency (no statistical heterogeneity), serious indirectness (because of non-generalisation of study findings to all situations, and differing characteristics of the study population), and serious imprecision (wide 95% CI around the pooled effect, inclusion of no effect, and

the upper or lower confidence limit crossed the minimal important difference for benefit).

### Potential biases in the review process

All included studies measured the effect of vitamin D in childhood pneumonia. There was potential to miss studies that may have measured pneumonia as an acute respiratory tract infection or lower respiratory tract infection as a secondary outcome, under the broader scope of childhood infection. We aimed to avoid this scenario by conducting a wide search and carefully assessing the relevance of each paper identified. We were unable to assess aetiological agents (bacterial or viral) to inform reporting of any beneficial effects of vitamin D as none of the included studies reported these data. Furthermore, because serum vitamin D levels were not measured, we could not suggest the dose-response effect in children with acute pneumonia.

### Agreements and disagreements with other studies or reviews

We published an earlier systematic review on this topic (Das 2013), and included two studies in this review (Choudhary 2012; Manaseki-Holland 2010). The conclusions are the same.

## AUTHORS' CONCLUSIONS

### Implications for practice

We are uncertain as to whether oral vitamin D supplementation as an adjunct for the treatment of children aged under five years with acute pneumonia has an important effect on outcomes. Low- to very low-quality evidence was available from only seven studies. Future studies should focus on the limitations or weaknesses identified in this review so that good-quality evidence can be generated.

### Implications for research

Future studies may consider use of a higher dosage schedule of vitamin D along with monitoring of toxicity and serum vitamin D level. Baseline and post-treatment vitamin D levels should be measured to correlate the outcome measures with subclinical or clinical vitamin D deficiency. Prevalence of malnutrition should be reported in future trials, as this may affect the therapeutic effect of vitamin D independently of the illness.

Future studies should also report the level of inflammatory markers along with vitamin D levels in children affected with acute non-severe or severe pneumonia. The aetiology (bacterial or viral or atypical) of pneumonia should be studied, and studies should ideally be conducted over a year to include all seasons, because vitamin D levels may be affected by seasonal change, as is the aetiology of pneumonia.

Any information regarding the duration of illness prior to initiation of therapy with vitamin D would be meaningful and should be included in future studies.

Information regarding prior antibiotic use should also be included because this can modify recovery from illness. This information should be included in the baseline data to enable comparison among studies.

Future studies should also include a subgroup of children with severe malnutrition or rickets or both and those with wheezing.

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**CHARACTERISTICS OF STUDIES**
**Characteristics of included studies [ordered by study ID]**
**Choudhary 2012**

Methods

**Study design:** randomised, double-blind, placebo-controlled trial



**Choudhary 2012** (Continued)

**Study duration:** not reported

Participants	<p><b>Inclusion criteria</b></p> <ul style="list-style-type: none"> <li>• Setting: hospital inpatients department</li> <li>• Country: India</li> <li>• Health condition: children aged 2 months to 5 years hospitalised with severe pneumonia</li> <li>• Number: N = 200; treatment (100); control (100)</li> <li>• Age (mean, SD/median, range)             <ul style="list-style-type: none"> <li>◦ Treatment: 14.1 ± 12.2 months</li> <li>◦ Control: 13.8 ± 11.4 months</li> </ul> </li> <li>• Sex (m/f): not reported</li> <li>• Other information: history of recurrent pneumonia was present in 30% of intervention group and 33% of control group. 5 children had clinical evidence of rickets. One-third of children had wheezing at enrolment.</li> </ul> <p><b>Exclusion criteria</b></p> <ul style="list-style-type: none"> <li>• Children with severe wasting (weight for height &lt; 3 SD), chronic illnesses, previous history of vitamin D intake over last 4 weeks, and known asthmatics were excluded.</li> </ul>
Interventions	<p><b>Treatment group</b></p> <ul style="list-style-type: none"> <li>• Intervention: vitamin D<sub>3</sub></li> <li>• Dose, duration, frequency: oral vitamin D<sub>3</sub> was given for 5 days at doses of 1000 IU for children aged up to 1 year and 2000 IU for children aged over 1 year</li> </ul> <p><b>Control group</b></p> <ul style="list-style-type: none"> <li>• Intervention: lactose</li> <li>• Dose, duration, frequency: 200 mg once daily for 5 days</li> </ul> <p><b>Co-interventions</b></p> <ul style="list-style-type: none"> <li>• Antibiotics were administered as per the IAP 2006 guidelines (&lt; 3 months: cefotaxime/ceftriaxone ± gentamicin/amikacin as first line, and co-amoxycylav + gentamicin/amikacin as second line; 3 months to 5 years: ampicillin/chloramphenicol or ampicillin + chloramphenicol or co-amoxycylav as first line, and co-amoxycylav or cefotaxime/ceftriaxone as second line. In staphylococcal infection, cloxacillin as first line and vancomycin/teicoplanin as second line was added to any of the above combinations in the respective age groups).</li> </ul>
Outcomes	<p><b>Primary</b></p> <ul style="list-style-type: none"> <li>• Time to resolution of severe pneumonia (absence of lower chest indrawing, hypoxia, cyanosis, lethargy, and inability to feed)</li> </ul> <p><b>Secondary</b></p> <ul style="list-style-type: none"> <li>• Duration of hospitalisation, time to resolution of tachypnoea, fever, hypoxia, chest retraction, and inability to feed/lethargy</li> </ul>
Notes	<p><b>Funding source:</b> not funded by any funding agency</p> <p><b>Other:</b> microbiological and radiological diagnosis of pneumonia was not done. Serum vitamin D<sub>3</sub> level was not measured.</p> <p><b>Contact with study authors:</b> email: 24 April 2016. Data were reported in median and IQR; we requested data in mean and SD, but this information was unavailable.</p>

**Risk of bias**

**Choudhary 2012** (Continued)

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Computer randomisation
Allocation concealment (selection bias)	Low risk	Sealed envelopes were used.
Blinding of participants and personnel (performance bias) All outcomes	Low risk	Blinding of participants and caretaker was ensured and unlikely that the blinding could have been broken, as both the intervention and placebo looked alike in terms of appearance, taste, and colour.
Blinding of outcome assessment (detection bias) All outcomes	Low risk	The code key was opened only after the intervention, data collection, follow-up, and tabulation were completed.
Incomplete outcome data (attrition bias) All outcomes	Low risk	4.5% (9/200) children lost to follow-up or died
Selective reporting (reporting bias)	Unclear risk	The trial was not registered, so this was unclear, but all outcomes proposed in the methods section were reported.
Other bias	Low risk	No other known risks of bias

**Dhungel 2015**

Methods	<p><b>Study design:</b> randomised controlled trial</p> <p><b>Study duration:</b> 3 months</p>
Participants	<p><b>Inclusion criteria</b></p> <ul style="list-style-type: none"> <li>• Setting: hospital inpatients department</li> <li>• Country: Pakistan</li> <li>• Health condition: children aged from 2 to 60 months with non-severe and severe pneumonia</li> <li>• Number: treatment (100); control (100)</li> <li>• Age (mean, SD/median, range)               <ul style="list-style-type: none"> <li>◦ Treatment: 8.1 ± 11.0 months</li> <li>◦ Control: 6.9 ± 10.1 months</li> </ul> </li> <li>• Sex (m): overall 122/200 (61%)</li> <li>• Other information: severe pneumonia was present in 43% children</li> </ul> <p><b>Exclusion criteria</b></p> <ul style="list-style-type: none"> <li>• Children with very severe pneumonia, wheezing, rickets, severe malnutrition, and other comorbid severe conditions were excluded.</li> <li>• Children with very severe pneumonia, chronic respiratory disease (asthma, TB), heart disease, have received vitamin D in last 3 months, grade III malnutrition, and normal vitamin D level were excluded.</li> </ul>
Interventions	<p><b>Treatment group</b></p> <ul style="list-style-type: none"> <li>• Intervention: vitamin D<sub>3</sub></li> <li>• Dose, duration, frequency: a single 100,000 IU dose of intramuscular vitamin D<sub>3</sub> was given within 24 hours of admission</li> </ul>

**Dhungel 2015** (Continued)

**Control group**

- Intervention: none (only standard care including antibiotics)
- Dose, duration, frequency: not applicable

**Co-interventions**

- Children received antibiotics that included cefotaxime, ceftriaxone, ampiclox, clarithromycin, amikacin, vancomycin. The antibiotics were given as single or in combination.

Outcomes	<b>Primary</b> <ul style="list-style-type: none"> <li>• Duration of hospitalisation</li> </ul> <b>Secondary</b> <ul style="list-style-type: none"> <li>• The risk of children having a repeat episode of pneumonia during the 3 months post-treatment period</li> </ul>
Notes	<b>Funding source:</b> not funded by any funding agency  <b>Other:</b> microbiological diagnosis of pneumonia was not done. Only 4% children had normal chest x-ray. Serum vitamin D <sub>3</sub> level was measured and those with normal level (> 20 ng/mL) were excluded. No report of adverse events or mortality.  <b>Contact with study authors:</b> no

**Risk of bias**

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	Not mentioned
Allocation concealment (selection bias)	Unclear risk	Not mentioned
Blinding of participants and personnel (performance bias) All outcomes	Unclear risk	Not mentioned
Blinding of outcome assessment (detection bias) All outcomes	Unclear risk	Not mentioned
Incomplete outcome data (attrition bias) All outcomes	Unclear risk	Not mentioned
Selective reporting (reporting bias)	Unclear risk	The trial was not registered, so this was unclear, but all outcomes proposed in the methods section were reported.
Other bias	Low risk	No other known risks of bias

**Gupta 2016a**

Methods	<b>Study design:</b> randomised, double-blind, placebo-controlled trial
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Gupta 2016a (Continued)

**Study duration:** 29 months

Participants	<p><b>Inclusion criteria</b></p> <ul style="list-style-type: none"> <li>• Setting: hospital inpatients department</li> <li>• Country: India</li> <li>• Health condition: children aged 6 months to 5 years hospitalised with severe pneumonia</li> <li>• Number: N = 324; treatment (162); control (162)</li> <li>• Age (mean, SD)             <ul style="list-style-type: none"> <li>◦ Treatment: 16.4 ± 12.9 months</li> <li>◦ Control: 16.8 ± 11.4 months</li> </ul> </li> <li>• Sex (m): treatment: 69.7%; control: 69.7%</li> <li>• Other information: the prevalence of moderate malnutrition as per WHO definition was 21.3%. The prevalence of anaemia (Hb &lt; 11 g/dL) was 82.4%, hypocalcaemia was 55.9%, hypophosphataemia was 31.4%. Vitamin D deficiency was present in 37.6% of children in the vitamin D group and 40.1% of children in the placebo group. Wheezing was present in 84.6% of children in the vitamin D group and 78.4% of children in the control group.</li> </ul> <p><b>Exclusion criteria</b></p> <ul style="list-style-type: none"> <li>• Children with rickets, severe acute malnutrition, asthma, hypertension, complicated pneumonia or illness severe enough to require ventilation, chronic respiratory disease, heart disease, renal or hepatic insufficiency, neurological illness, immunodeficiency, received vitamin D or calcium supplements within 4 weeks prior to enrolment, and those with hypercalcaemia or allergy to vitamin D, or immunised with pneumococcal/flu vaccine were excluded.</li> </ul>
Interventions	<p><b>Treatment group</b></p> <ul style="list-style-type: none"> <li>• Intervention: vitamin D<sub>3</sub></li> <li>• Dose, duration, frequency: oral vitamin D<sub>3</sub> was given as a single dose of 100,000 IU on day of enrolment</li> </ul> <p><b>Control group</b></p> <ul style="list-style-type: none"> <li>• Intervention: placebo not described</li> <li>• Dose, duration, frequency: oral placebo was given as a single dose on day of enrolment</li> </ul> <p><b>Co-interventions</b></p> <ul style="list-style-type: none"> <li>• Antibiotics were administered as per the IAP 2006 guidelines (&lt; 3 months: cefotaxime/ceftriaxone ± gentamicin/amikacin as first line, and co-amoxycylav + gentamicin/amikacin as second line; 3 months to 5 years: ampicillin/chloramphenicol or ampicillin + chloramphenicol or co-amoxycylav as first line, and co-amoxycylav or cefotaxime/ceftriaxone as second line. In staphylococcal infection, cloxacillin as first line and vancomycin/teicoplanin as second line was added to any of the above combinations in the respective age groups).</li> </ul>
Outcomes	<p><b>Primary</b></p> <ul style="list-style-type: none"> <li>• Time to resolution of severe pneumonia (the duration from enrolment until the chest indrawing was no longer present, and continued to be absent for next 24 hours)</li> <li>• The proportion of children having a recurrence of pneumonia in next 6 months</li> </ul> <p><b>Secondary</b></p> <ul style="list-style-type: none"> <li>• Change in the serum level of 25(OH)-D and PTH after 2 weeks and 3 months of therapy</li> <li>• Change in serum level of cathelicidin and immunoglobulins (IgA, IgG, IgM) after 2 weeks of therapy</li> <li>• Duration of hospitalisation</li> <li>• Time to complete recovery from pneumonia (normalisation of respiratory rate)</li> <li>• Fever clearance time</li> <li>• Incidence rate of pneumonia during follow-up</li> </ul>

**Gupta 2016a** (Continued)

Notes

**Funding source:** not funded by any funding agency

**Other:** blood culture was positive in 8.6% (*Staphylococcus aureus* isolated in 8.3%), and 90.1% of children abnormal chest x-ray in the form of consolidation, bilateral patchy opacities, and hyperinflation or minor infiltrates.

**Contact with study authors:** no

**Risk of bias**

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Computer randomisation
Allocation concealment (selection bias)	Low risk	Sealed envelopes were used.
Blinding of participants and personnel (performance bias) All outcomes	Low risk	Blinding of participants and caretaker was ensured and unlikely that the blinding could have been broken, as both the intervention and placebo looked alike in terms of appearance, taste, and colour.
Blinding of outcome assessment (detection bias) All outcomes	Low risk	The code key was opened only after the intervention, data collection, follow-up, and tabulation were completed.
Incomplete outcome data (attrition bias) All outcomes	Low risk	4.63% (15/324) children lost to follow-up
Selective reporting (reporting bias)	Low risk	The trial was registered.
Other bias	Low risk	No other known risks of bias

**Manaseki-Holland 2010**

Methods

**Study design:** randomised, double-blind, placebo-controlled trial

**Study duration:** 2 months

Participants

**Inclusion criteria**

- Setting: hospital outpatients and inpatients department
- Country: Afghanistan
- Health condition: children aged from 1 month to 36 months with acute pneumonia
- Number: treatment (224); control (229)
- Age (mean, SD/median, range)
  - Treatment: 13.18 ± 9.1 months
  - Control: 13.19 ± 9.2 months
- Sex (m): treatment: 57.8%; control: 55.6%
- Other information: severe pneumonia was present in 17.4% of intervention group and 15.3% of placebo group. Exclusion of cases with wheeze reduced the chances of including viral respiratory diseases.

**Exclusion criteria**

**Manaseki-Holland 2010** (Continued)

- Children with very severe pneumonia, wheezing, rickets, severe malnutrition, and other comorbid severe conditions were excluded.
- Children with severe wasting (weight for height < 3 SD), chronic illnesses, previous history of vitamin D intake over last 4 weeks, and children known to have asthma were excluded.

Interventions	<p><b>Treatment group</b></p> <ul style="list-style-type: none"> <li>• Intervention: vitamin D<sub>3</sub></li> <li>• Dose, duration, frequency: a single 100,000 IU dose of oral vitamin D<sub>3</sub> was given at onset of pneumonia</li> </ul> <p><b>Control group</b></p> <ul style="list-style-type: none"> <li>• Intervention: olive oil</li> <li>• Dose, duration, frequency: a single 2 mL olive oil dose was given at onset of pneumonia</li> </ul> <p><b>Co-interventions</b></p> <ul style="list-style-type: none"> <li>• Children received antibiotics according to the national pneumonia treatment protocol based upon Integrated Management of Childhood Illnesses (IMCI) guidelines (&lt; 2 months: intramuscular gentamicin and intramuscular benzylpenicillin; 2 months to 5 years: oral amoxicillin or intramuscular chloramphenicol (if unable to take orally))</li> </ul>	
Outcomes	<p><b>Primary</b></p> <ul style="list-style-type: none"> <li>• Time to resolution of pneumonia/recovery for 48 consecutive hours, and treatment failure</li> </ul> <p><b>Secondary</b></p> <ul style="list-style-type: none"> <li>• Risk of children having a repeat episode of pneumonia during the 90-day post-treatment period</li> <li>• Risk of survival without experiencing a repeat episode of pneumonia</li> </ul>	
Notes	<p><b>Funding source:</b> funded by the New Zealand Aid Corporation. The study medications were provided by the pharmacy department of Aga Khan University Hospital, Karachi.</p> <p><b>Other:</b> microbiological and radiological diagnosis of pneumonia was not done. Serum vitamin D<sub>3</sub> level was not measured. The trial reported 3 deaths (2 in vitamin D group and 1 in placebo group) during the 90-day follow-up evaluation of repeat episodes of pneumonia.</p> <p><b>Contact with study authors:</b> email: 4 May 2017. Because it was not clear whether the children died including during the treatment of the index episode of pneumonia, we contacted the corresponding author, and her opinion was that the children died during the 90-day follow-up (not during the index episode). Mortality data were therefore not used in the analysis.</p>	
<b>Risk of bias</b>		
<b>Bias</b>	<b>Authors' judgement</b>	<b>Support for judgement</b>
Random sequence generation (selection bias)	Low risk	Random number sequence was generated in an Excel spreadsheet with no restrictions.
Allocation concealment (selection bias)	Low risk	Sealed codes were used.
Blinding of participants and personnel (performance bias) All outcomes	Low risk	Blinding of participants and caretaker was ensured and unlikely that the blinding could have been broken, as both the intervention and placebo looked alike in terms of appearance, taste, and colour.
Blinding of outcome assessment (detection bias)	Unclear risk	Not mentioned

**Manaseki-Holland 2010** (Continued)

All outcomes

Incomplete outcome data (attrition bias) All outcomes	Low risk	3% children lost to follow-up
Selective reporting (reporting bias)	Unclear risk	The trial was not registered, so this was unclear, but all outcomes proposed in the methods section were reported.
Other bias	Low risk	No known other risks

**Rahmati 2016**

Methods	<b>Study design:</b> randomised controlled trial  <b>Study duration:</b> 12 months
Participants	<b>Inclusion criteria</b> <ul style="list-style-type: none"> <li>• Setting: hospital inpatients department</li> <li>• Country: Iran</li> <li>• Health condition: children aged from 2 to 72 months with community-acquired pneumonia</li> <li>• Number: treatment (50); control (50)</li> <li>• Age (mean) <ul style="list-style-type: none"> <li>◦ Treatment: 17.6 months</li> <li>◦ Control: 16.8 months</li> </ul> </li> <li>• Sex (m): treatment: 54%; control: 61.2%</li> <li>• Other information: no children had severe pneumonia</li> </ul> <b>Exclusion criteria</b> <ul style="list-style-type: none"> <li>• Immunocompromised children, children with allergic airway diseases, those hospitalised &gt; 24 hours, and those receiving high doses of vitamin D were excluded.</li> </ul>
Interventions	<b>Treatment group</b> <ul style="list-style-type: none"> <li>• Intervention: vitamin D</li> <li>• Dose, duration, frequency: oral vitamin D 50,000 IU per day for 2 days</li> </ul> <b>Control group</b> <ul style="list-style-type: none"> <li>• Intervention: olive oil</li> <li>• Dose, duration, frequency: not mentioned</li> </ul> <b>Co-interventions</b> <ul style="list-style-type: none"> <li>• Children received a single antibiotic as ceftriaxone.</li> </ul>
Outcomes	<b>Primary</b> <ul style="list-style-type: none"> <li>• Duration of hospitalisation/antibiotic therapy</li> <li>• Duration of fever</li> </ul>
Notes	<b>Funding source:</b> not funded by any funding agency  <b>Other:</b> microbiological diagnosis of pneumonia was not done. Serum vitamin D <sub>3</sub> level was not measured. No report of adverse events or mortality.



**Rahmati 2016** (Continued)

**Contact with study authors:** emails: 8 November 2017 and 15 November 2017. We requested that the authors clarify the outcomes on duration of antibiotic therapy and duration of hospitalisation as well as the reporting measure (day or hour) for duration of patient admission. No response received before 10 January 2018.

**Risk of bias**

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Computer randomisation
Allocation concealment (selection bias)	Unclear risk	Not mentioned
Blinding of participants and personnel (performance bias) All outcomes	Unclear risk	Not mentioned
Blinding of outcome assessment (detection bias) All outcomes	Unclear risk	Not mentioned
Incomplete outcome data (attrition bias) All outcomes	Low risk	No drop outs
Selective reporting (reporting bias)	Low risk	The trial was registered.
Other bias	Low risk	No other known risks

**Rajshekhhar 2016**

Methods	<p><b>Study design:</b> randomised, double-blind, placebo-controlled trial</p> <p><b>Study duration:</b> not reported</p>
Participants	<p><b>Inclusion criteria</b></p> <ul style="list-style-type: none"> <li>• Setting: hospital inpatients department</li> <li>• Country: India</li> <li>• Health condition: children aged 2 months to 5 years hospitalised with severe pneumonia</li> <li>• Number: N = 96; treatment (48); control (48)</li> <li>• Age (mean, SD/median, range)           <ul style="list-style-type: none"> <li>◦ Treatment: 1.94 ± 1.46 years</li> <li>◦ Control: 2.08 ± 1.92 years</li> </ul> </li> <li>• Sex (m): overall 58/96 (60.4%)</li> <li>• Other information: history of recurrent pneumonia was present in 35.4% of children in intervention group and 39.6% of children in control group. More children in the vitamin D group (93.8%) were adherent to vaccination than in the control group (79.2%). Top feeding was given commonly to children in the vitamin D group (79.2%) compared to the control group (18.8%).</li> </ul> <p><b>Exclusion criteria</b></p>

**Rajshekhar 2016** (Continued)

- Children with vitamin D supplementation in the 4 weeks prior to admission, severe malnutrition, congenital heart disease, congenital deformities of chest, and immunodeficiency states were excluded.

Interventions	<p><b>Treatment group</b></p> <ul style="list-style-type: none"> <li>• Intervention: vitamin D<sub>3</sub></li> <li>• Dose, duration, frequency: oral vitamin D<sub>3</sub> tablet was given for 5 days at doses of 1000 IU for children aged up to 1 year and 2000 IU for children aged over 1 year</li> </ul> <p><b>Control group</b></p> <ul style="list-style-type: none"> <li>• Intervention: lactose</li> <li>• Dose, duration, frequency: 1 tablet once daily for 5 days</li> </ul> <p><b>Co-interventions</b></p> <ul style="list-style-type: none"> <li>• Antibiotics were administered as per the IAP 2012 guidelines and included amoxicillin, amoxycylav, third-generation parenteral cephalosporins, and macrolides.</li> </ul>
Outcomes	<p><b>Primary</b></p> <ul style="list-style-type: none"> <li>• Time to resolution of severe pneumonia (absence of lower chest indrawing, hypoxia, cyanosis, and inability to feed)</li> </ul>
Notes	<p><b>Funding source:</b> not funded by any funding agency</p> <p><b>Other:</b> microbiological and radiological diagnosis of pneumonia was not done. Serum vitamin D<sub>3</sub> level was not measured. Meyer Organics Private Ltd. was involved in the randomisation process, allotment of study and placebo medications.</p> <p><b>Contact with study authors:</b> email: 3 June 2017. Information on study methodology was sought and the authors provided the same.</p>

**Risk of bias**

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Computer-generated block randomisation by a third party
Allocation concealment (selection bias)	Low risk	Sealed, opaque envelope was used.
Blinding of participants and personnel (performance bias) All outcomes	Low risk	Blinding of participants and caretaker was ensured and unlikely that the blinding could have been broken, as both the intervention and placebo looked alike in terms of appearance, taste, and colour.
Blinding of outcome assessment (detection bias) All outcomes	Low risk	The code key was disclosed only after the intervention, data collection, follow-up, and tabulations.
Incomplete outcome data (attrition bias) All outcomes	Unclear risk	Not mentioned
Selective reporting (reporting bias)	Unclear risk	The trial was not registered, so this was unclear, but all outcomes proposed in the methods section were reported.
Other bias	Unclear risk	No other known risks

**Somnath 2017**

Methods	<p><b>Study design:</b> randomised controlled trial</p> <p><b>Study duration:</b> 12 months</p>
Participants	<p><b>Inclusion criteria</b></p> <ul style="list-style-type: none"> <li>• Setting: hospital inpatients department</li> <li>• Country: India</li> <li>• Health condition: children aged from 2 to 60 months with acute lower respiratory infection</li> <li>• Number: treatment (78); control (78)</li> <li>• Age (mean): 13 months             <ul style="list-style-type: none"> <li>◦ Treatment: ≤ 12 months (53.8%), &gt; 12 months (46.2%)</li> <li>◦ Control: ≤ 12 months (47.4%), &gt; 12 months (52.6%)</li> </ul> </li> <li>• Sex (m): treatment: 69.2%; control: 65.8%</li> <li>• Other information: very severe pneumonia in 12.8% (N = 10) children in vitamin D and 14.3% (N = 11) children in control group</li> </ul> <p><b>Exclusion criteria</b></p> <ul style="list-style-type: none"> <li>• Children with tuberculosis, bronchial asthma, congenital lung malformation, immunodeficiency states, malabsorption syndromes, chronic diarrhoea, liver disease, kidney disease, or any known contraindications for vitamin D administration were excluded.</li> </ul>
Interventions	<p><b>Treatment group</b></p> <ul style="list-style-type: none"> <li>• Intervention: vitamin D</li> <li>• Dose, duration, frequency: a single 100,000 IU dose of oral vitamin D<sub>3</sub> was given on first day of admission</li> </ul> <p><b>Control group</b></p> <ul style="list-style-type: none"> <li>• Intervention: standard therapy</li> <li>• Dose, duration, frequency: not mentioned</li> </ul> <p><b>Co-interventions</b></p> <ul style="list-style-type: none"> <li>• Children received standard pneumonia treatment as per the guideline. Details of antibiotics not mentioned.</li> </ul>
Outcomes	<p><b>Primary</b></p> <ul style="list-style-type: none"> <li>• Duration of hospitalisation</li> </ul> <p><b>Secondary</b></p> <ul style="list-style-type: none"> <li>• Decreasing the time taken for defervescence</li> <li>• Decreasing the complications associated with ALRI (syn/para-pneumonic effusion, empyema, lung abscess, pneumothorax, secondary infections, and acute respiratory distress syndrome)</li> <li>• Decreasing the need for PICU transfer</li> <li>• Decreasing the mortality associated with ALRI</li> <li>• The proportion of children having a recurrence of pneumonia in next 6 months</li> </ul>
Notes	<p><b>Funding source:</b> institute (JIPMER) Intramural Grant</p> <p><b>Other:</b> microbiological diagnosis of pneumonia was not done. Only 7% of study population had normal serum, and 71% were deficient range. No report of adverse events or mortality</p> <p><b>Contact with study authors:</b> no</p>

**Somnath 2017** (Continued)

**Risk of bias**

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Computer-generated block randomisation
Allocation concealment (selection bias)	Low risk	Sealed, opaque envelope was used.
Blinding of participants and personnel (performance bias) All outcomes	High risk	Not blinded
Blinding of outcome assessment (detection bias) All outcomes	Low risk	The outcome assessors were blinded to the intervention.
Incomplete outcome data (attrition bias) All outcomes	Low risk	1.3% (2/156) children lost to follow-up
Selective reporting (reporting bias)	Low risk	The trial was registered.
Other bias	Low risk	No other known risks

ALRI: acute lower respiratory tract infection

f: female

Hb: haemoglobin

IAP: Indian Academy of Pediatrics

Ig: immunoglobulin

IQR: interquartile range

m: male

PTH: parathyroid hormone

PICU: paediatric intensive care unit

SD: standard deviation

TB: tuberculosis

25(OH)-D: 25-hydroxy-vitamin D

WHO: World Health Organization

**Characteristics of excluded studies** [ordered by study ID]

Study	Reason for exclusion
Zhou 2016	Multiple nutrients (iron, zinc, vitamin A, and vitamin D) were supplemented in children with nutrient deficiency/malnutrition and pneumonia. The trial included children with asthmatic pneumonia.

**Characteristics of ongoing studies** [ordered by study ID]

**NCT02054182**

Trial name or title	Effect of vitamin D supplementation in young South African children hospitalised with acute lower respiratory infection
Methods	RCT
Participants	Children aged from 1 month to 5 years, admitted to hospital with acute lower respiratory tract infection, i.e. bronchiolitis or pneumonia or both
Interventions	Vitamin D: 2500 IU daily from enrolment until hospital discharge
Outcomes	Primary: change from baseline in modified Respiratory Distress Assessment Instrument score at hospital discharge between intervention and placebo groups  Secondary: duration of hospitalisation
Starting date	February 2014
Contact information	Winter-Rose S Nkosi, Univeristy of Limpopo, Medunsa Campus, czawr@webmail.co.za
Notes	ClinicalTrials.gov identifier: NCT02054182

**NCT02185196**

Trial name or title	Vitamin D supplementation: impact on severe pneumonia among under 5 children
Methods	RCT
Participants	Children of either sex aged 3 to 59 months, with a clinical diagnosis of severe pneumonia with or without diarrhoea
Interventions	Vitamin D <sub>3</sub> : 20,000 IU vitamin D <sub>3</sub> in children < 6 months; 50,000 IU in children 6 to 12 months; and 100,000 IU in children 13 to 59 months of age on first day and 10,000 IU for next 4 days  Placebo: 20,000 IU Miglyol-oil in children < 6 months; 50,000 IU in children 6 to 12 months; and 100,000 IU in children 13 to 59 months of age on first day and 10,000 IU for next 4 days
Outcomes	Primary: time taken for recovery from severe pneumonia  Secondary: duration of hospitalisation, time taken for normalisation of temperature and respiratory rate, time taken for recovery from chest indrawing, time taken for oxygen saturation, feeding, and mental status to normalise, proportion of study children who will develop new episodes of pneumonia during the follow-up period
Starting date	June 2014
Contact information	Fahmida Chowdhury, International Centre for Diarrhoeal Disease Research, Bangladesh, fahmida_chow@icddr.org
Notes	ClinicalTrials.gov identifier: NCT02185196

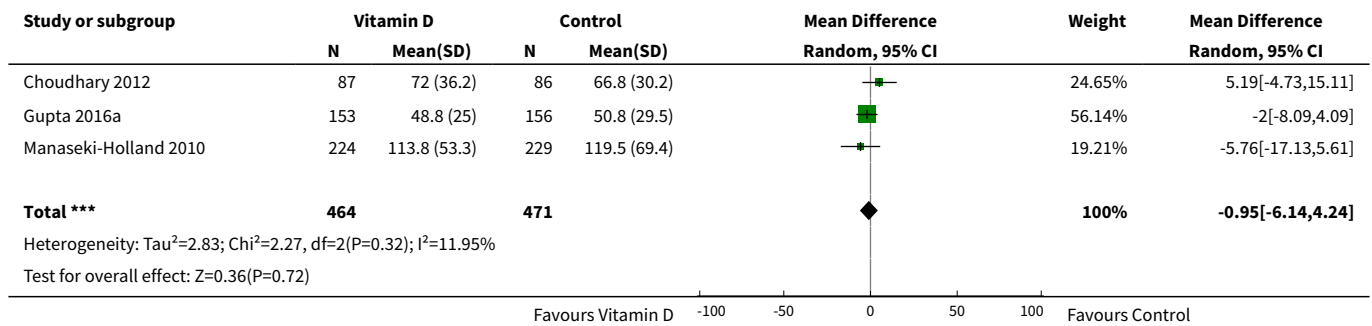
RCT: randomised controlled trial

**DATA AND ANALYSES**

**Comparison 1. Time to resolution of acute illness**

Outcome or subgroup title	No. of studies	No. of participants	Statistical method	Effect size
1 Time to resolution of acute illness	3	935	Mean Difference (IV, Random, 95% CI)	-0.95 [-6.14, 4.24]

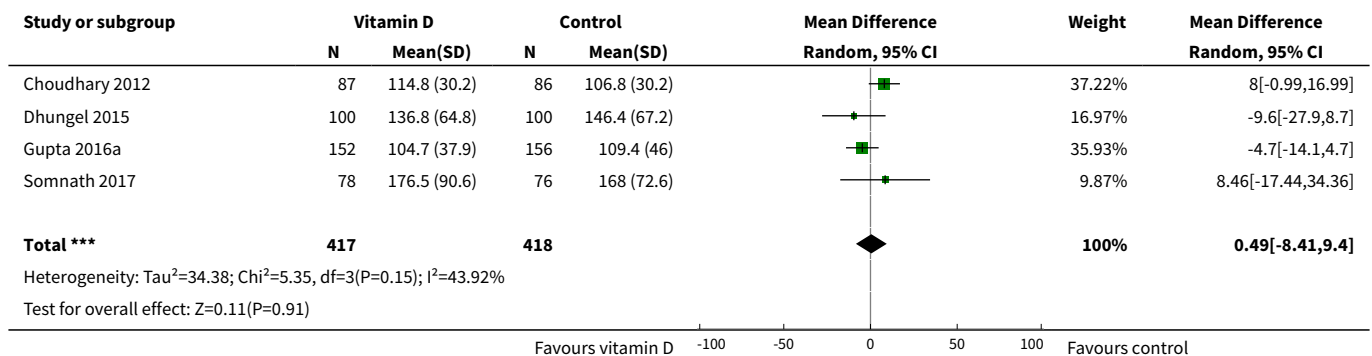
**Analysis 1.1. Comparison 1 Time to resolution of acute illness, Outcome 1 Time to resolution of acute illness.**



**Comparison 2. Duration of hospitalisation**

Outcome or subgroup title	No. of studies	No. of participants	Statistical method	Effect size
1 Duration of hospitalisation	4	835	Mean Difference (IV, Random, 95% CI)	0.49 [-8.41, 9.40]

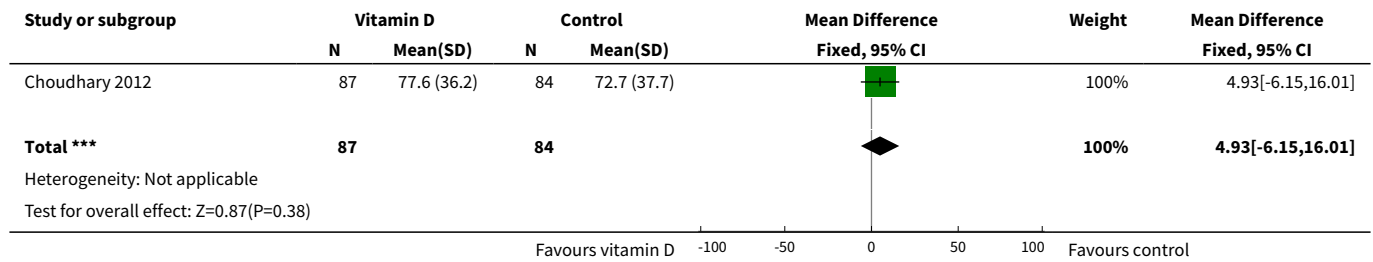
**Analysis 2.1. Comparison 2 Duration of hospitalisation, Outcome 1 Duration of hospitalisation.**



**Comparison 3. Time to resolution of tachypnoea**

Outcome or subgroup title	No. of studies	No. of participants	Statistical method	Effect size
1 Time to resolution of tachypnoea	1	171	Mean Difference (IV, Fixed, 95% CI)	4.93 [-6.15, 16.01]

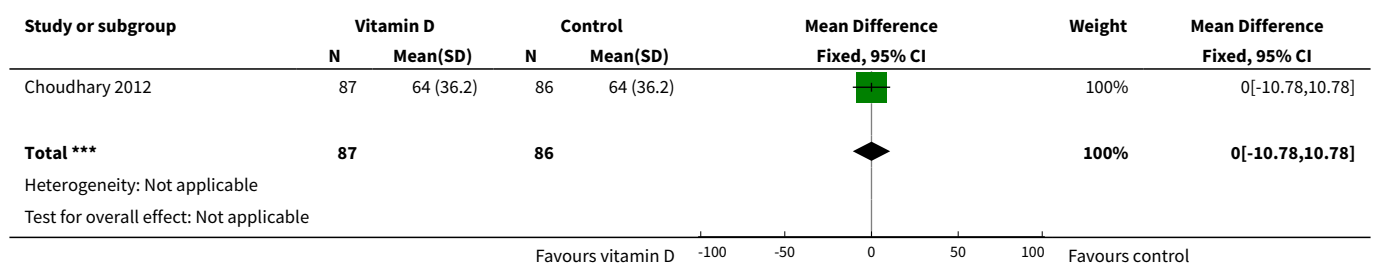
**Analysis 3.1. Comparison 3 Time to resolution of tachypnoea, Outcome 1 Time to resolution of tachypnoea.**



**Comparison 4. Time to resolution of lower chest indrawing**

Outcome or subgroup title	No. of studies	No. of participants	Statistical method	Effect size
1 Time to resolution of lower chest indrawing	1	173	Mean Difference (IV, Fixed, 95% CI)	0.0 [-10.78, 10.78]

**Analysis 4.1. Comparison 4 Time to resolution of lower chest indrawing, Outcome 1 Time to resolution of lower chest indrawing.**

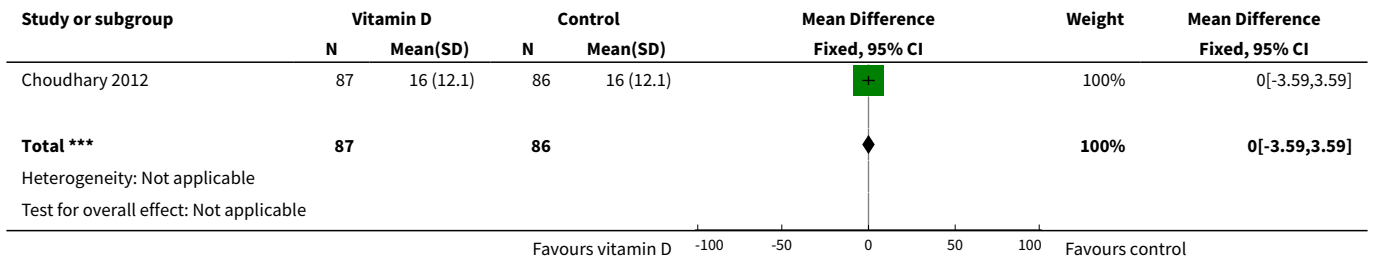


**Comparison 5. Time to resolution of hypoxia**

Outcome or subgroup title	No. of studies	No. of participants	Statistical method	Effect size
1 Time to resolution of hypoxia	1	173	Mean Difference (IV, Fixed, 95% CI)	0.0 [-3.59, 3.59]



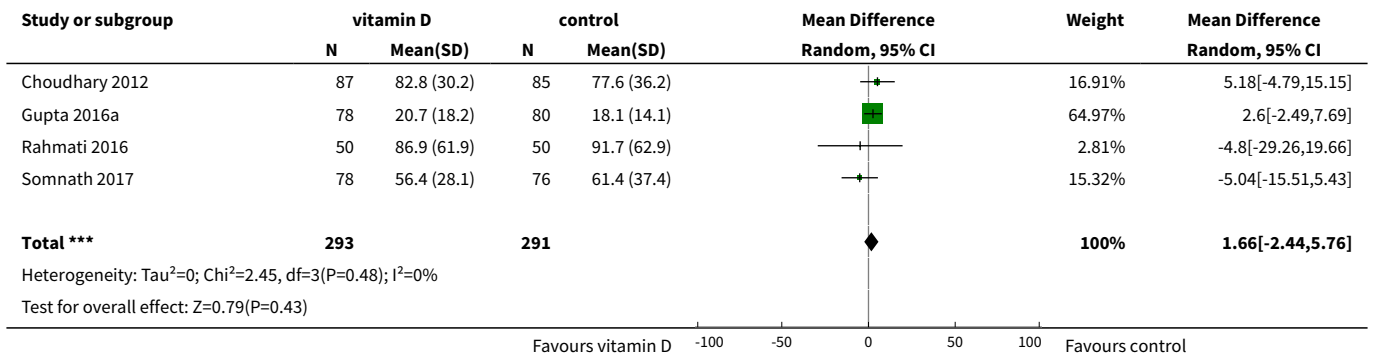
**Analysis 5.1. Comparison 5 Time to resolution of hypoxia, Outcome 1 Time to resolution of hypoxia.**



**Comparison 6. Time to resolution of fever**

Outcome or subgroup title	No. of studies	No. of participants	Statistical method	Effect size
1 Time to resolution of fever	4	584	Mean Difference (IV, Random, 95% CI)	1.66 [-2.44, 5.76]

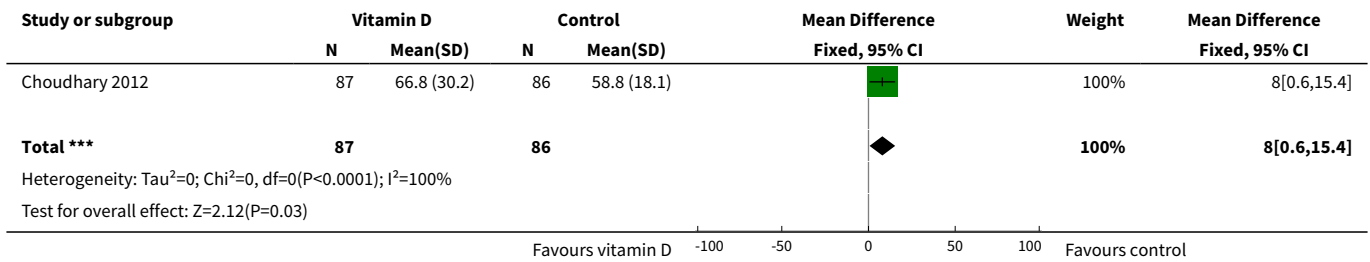
**Analysis 6.1. Comparison 6 Time to resolution of fever, Outcome 1 Time to resolution of fever.**



**Comparison 7. Time to resolution of inability to feed or lethargy**

Outcome or subgroup title	No. of studies	No. of participants	Statistical method	Effect size
1 Time to resolution of inability to feed or lethargy	1	173	Mean Difference (IV, Fixed, 95% CI)	8.00 [0.60, 15.40]

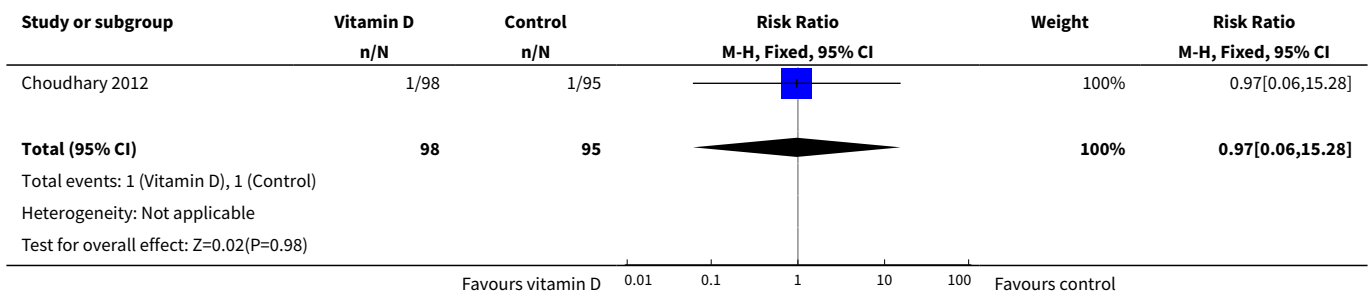
**Analysis 7.1. Comparison 7 Time to resolution of inability to feed or lethargy, Outcome 1 Time to resolution of inability to feed or lethargy.**



**Comparison 8. Mortality rate**

Outcome or subgroup title	No. of studies	No. of participants	Statistical method	Effect size
1 Mortality rate	1	193	Risk Ratio (M-H, Fixed, 95% CI)	0.97 [0.06, 15.28]

**Analysis 8.1. Comparison 8 Mortality rate, Outcome 1 Mortality rate.**



**APPENDICES**

**Appendix 1. MEDLINE (Ovid) search strategy**

- 1 exp Pneumonia/ (75695)
- 2 (pneumon\* or bronchpneumon\* or pleuropneumon\*).tw. (132976)
- 3 Respiratory Tract Infections/ (31796)
- 4 (lower respiratory tract infection\* or lower respiratory infection\* or lrti).tw. (5207)
- 5 or/1-4 (189394)
- 6 exp Vitamin D/ (45162)
- 7 ("vitamin d" or "vit d").tw,nm. (44814)
- 8 ("vitamin d2" or "vitamin d3").tw,nm. (8508)
- 9 (ergocalciferol\* or cholecalciferol\* or calcitriol\*).tw,nm. (24708)
- 10 or/6-9 (59727)
- 11 5 and 10 (214)

**Appendix 2. Embase (Elsevier) search strategy**

#13	#11 AND #12	140
#12		954414
	#12.8 #12.3 NOT #12.7 954412	
	#12.7 #12.4 NOT #12.6	
	#12.6#12.4 AND #12.5	
	#12.5'human'/de AND [embase]/lim	
	#12.4'animal'/de OR 'nonhuman'/de OR 'animal experiment'/de AND [embase]/lim	
	#12.3#12.1 OR #12.2	
	#12.2random*:ab,ti OR placebo*:ab,ti OR crossover*:ab,ti OR 'cross over':ab,ti OR allocat*:ab,ti OR trial:ti OR (doubl* NEXT/1 blind*):ab,ti AND [embase]/lim	
	#12.1'randomized controlled trial'/exp OR 'single blind procedure'/exp OR 'double blind procedure'/exp OR 'crossover procedure'/exp AND [embase]/lim	
#11	#5 AND #10	922
#10	#6 OR #7 OR #8 OR #9	96411
#9	ergocalciferol*:ab,ti OR cholecalciferol*:ab,ti OR calcitriol*:ab,ti AND [embase]/lim	6851
#8	'vitamin d2':ab,ti OR 'vitamin d3':ab,ti AND [embase]/lim	3634
#7	'vitamin d':ab,ti OR 'vit d':ab,ti AND [embase]/lim	54129
#6	'vitamin d'/exp AND [embase]/lim	87300
#5	#1 OR #2 OR #3 OR #4	244145
#4	'lower respiratory tract infection':ab,ti OR 'lower respiratory tract infections':ab,ti OR 'lower respiratory infection':ab,ti OR 'lower respiratory infections':ab,ti OR lrti:ab,ti AND [embase]/lim	6451
#3	'lower respiratory tract infection'/de AND [embase]/lim	7631
#2	pneumon*:ab,ti OR pleuropneumon*:ab,ti OR bronchopneumon*:ab,ti AND [embase]/lim	150505
#1	'pneumonia'/exp AND [embase]/lim	174841

## CONTRIBUTIONS OF AUTHORS

Dr Rashmi Ranjan Das (RRD), Dr Meenu Singh (MS), and Dr Sushree Samiksha Naik (SSN)

Conceiving the review: RRD.

Co-ordinating the review: RRD, SSN.

Undertaking manual searches: RRD, SSN.

Screening search results: RRD, SSN.

Organising retrieval of papers: RRD, SSN.

Screening retrieved papers against inclusion criteria: RRD, MS.

Appraising quality of papers: RRD, MS.

Abstracting data from papers: RRD, SSN.  
Writing to authors of papers for additional information: RRD, MS.  
Providing additional data about papers: RRD.  
Obtaining and screening data on unpublished studies: RRD, MS.  
Managing data for the review: RRD, SSN.  
Entering data into Review Manager 5: RRD, SSN.  
Calculating Review Manager 5 statistical data: RRD, MS.  
Performing other statistical analysis not using Review Manager 5: RRD, MS.  
Interpreting data: RRD, MS.  
Making statistical inferences: RRD, MS.  
Writing the review: RRD, SSN, MS.  
Performing previous work that served as the foundation of the present study: RRD, MS.  
Serving as guarantor for the review (one author): RRD.  
Taking responsibility for reading and checking the review before submission: RRD, MS, SSN.

## DECLARATIONS OF INTEREST

Rashmi R Das: none known.  
Meenu Singh: none known.  
Sushree S Naik: none known.

## DIFFERENCES BETWEEN PROTOCOL AND REVIEW

In the protocol, we indicated that if we found cluster-randomised trials for inclusion, we would adjust for clustering by intracluster correlation coefficient. However, we did not identify any cluster-randomised trials for inclusion.

We also indicated that we would perform additional analyses (subgroup analysis, and sensitivity analysis). However, there were insufficient studies and data to inform conducting these analyses for this review.

## INDEX TERMS

### Medical Subject Headings (MeSH)

Anti-Bacterial Agents [\*therapeutic use]; Chemotherapy, Adjuvant; Community-Acquired Infections [drug therapy]; Fever [drug therapy]; Pneumonia [\*drug therapy]; Randomized Controlled Trials as Topic; Vitamin D [\*administration & dosage]; Vitamins [\*administration & dosage]

### MeSH check words

Child, Preschool; Humans; Infant