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Fine Particle Constituents and Mortality: A Time-Series Study in Beijing, China

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Abstract

There is a rising concern that fine particle (PM_{2.5}) compositions may play an important role in explaining PM_{2.5}-related mortality risks. However, PM_{2.5} constituents responsible for these risks have not yet been determined. To date, there are few PM_{2.5} constituent health studies in developing countries. We adopted a time-series approach, using generalized linear regression models to examine associations between short-term exposure to PM_{2.5} constituents and mortality. We analyzed data stratified by sex and by age groups (<65, 65–74, and >74) from 2013 to 2015 in Beijing, China. We also investigated seasonal patterns of such associations. For a 0 day lag, interquartile range increases in potassium, calcium, magnesium, and organic carbon were associated with 0.51% (95% CI: 0.17–0.85), 2.07% (95% CI: 0.71–3.44), 0.26% (95% CI: 0.08–0.44), and 2.65% (95% CI: 0.18–5.18) increases in respiratory mortality, and sulfate with a 1.57% (95% CI: 0.04–3.12) increase in cardiovascular mortality. In the season-stratified analysis, the association of some constituents (potassium, calcium, magnesium, nitrate, sulfate, and organic carbon) with respiratory mortality appeared to be stronger in cold seasons than in warm seasons. Older adults (65–74) may be susceptible to certain compositions. Our findings provide evidence that link PM_{2.5} constituents with mortality and suggest that adverse effects vary among constituents in different seasons.

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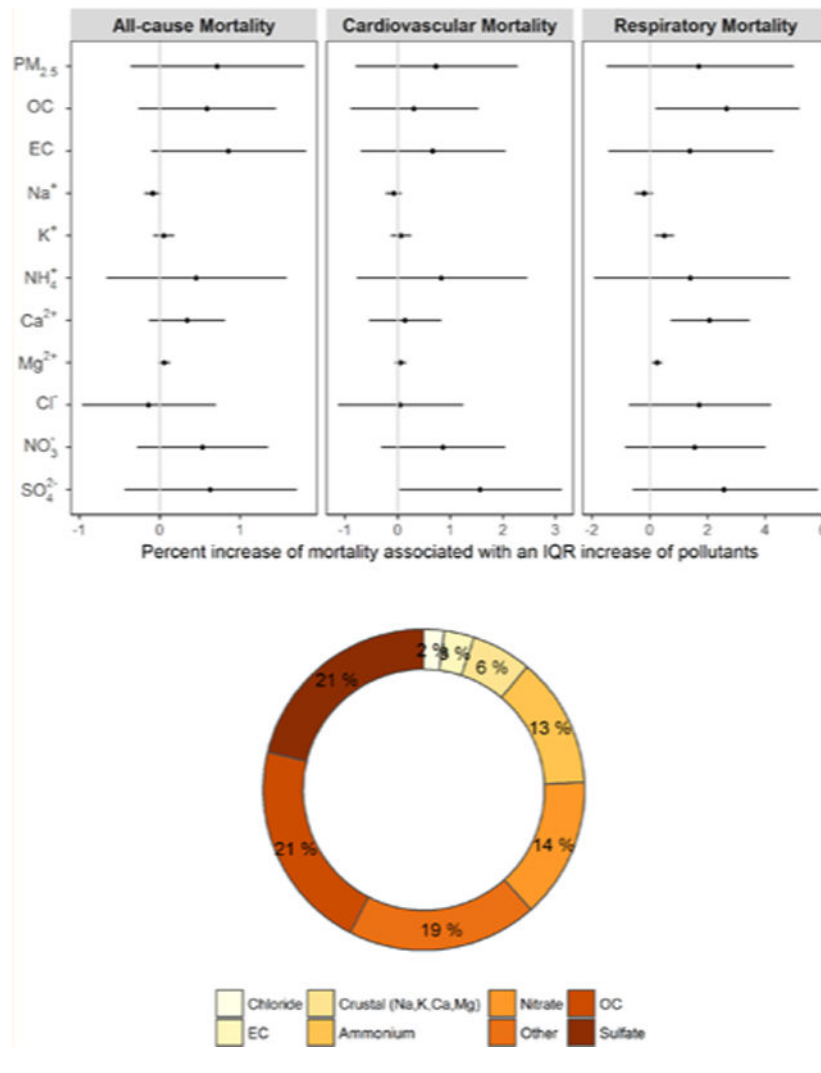
Supporting Information

The Supporting Information is available free of charge on the [ACS Publications website](https://pubs.acs.org/doi/10.1021/acs.est.8b00424) at DOI: 10.1021/acs.est.8b00424.

Map of study area and fixed-site monitors (Figure S1), table of correlation coefficients between PM_{2.5} and chemical constituents (Table S1), and results of the sensitivity analysis: using alternative values of the df for temporal trend and meteorological indicators, as well as including PM_{2.5} mass in the statistical model as a potential confounder (Table S2), replacing the BDL values with half of the detection limit (Table S3), and conducting sensitivity analysis for extreme values of each constituent (Table S4) (PDF)

The authors declare no competing financial interest.

Graphical abstract



INTRODUCTION

Recent epidemiologic studies have confirmed that short-term exposure to fine particle ($PM < 2.5 \mu m$ in aerodynamic diameter, or $PM_{2.5}$) increases the risk of human mortality.¹⁻⁴ Policy makers and regulators have sought to determine which components of the particulate mixture are most harmful to human health.⁵ The evidence indicates that the mechanism of $PM_{2.5}$ -related mortality may be related to oxidative stress and inflammatory response,^{2,6,7} but it is uncertain if the specific chemical constituents of $PM_{2.5}$ play a role. Previous research studies conducted in the U.S. and Europe have tried to determine which constituents pose the highest risks to human health,⁸⁻¹¹ mostly through single-center studies.^{10,11} The increasing availability of national-scale monitoring data of the composition of $PM_{2.5}$ has permitted some nationwide multicenter studies in recent years.¹²⁻¹⁶ These studies have reported that sulfate, nitrate, elemental carbon, and organic carbon are positively related to the risks of mortality and hospitalization¹²⁻¹⁶ and that transition metals including

zinc, nickel, chromium, and iron are also responsible for PM_{2.5}-related effects.^{12,15,16} However, because of a lack of available data for chemical components and spatial–temporal variations in their distributions,^{10,11} the evidence currently available does not allow us to directly link the epidemiological effects of specific constituents with mortality.

China now experiences nationwide air pollution, and PM_{2.5} has become the main risk factor to population health. Numerous studies have confirmed the acute health effects of PM_{2.5} in China and have found that the magnitude of the effect is much less than what has been reported in similar studies conducted in the U.S. and Europe.^{17–19} The reason for this discrepancy can be explained by different characteristics of population exposure, the shape of the exposure-response function, and regional socioeconomic factors. Another probable reason is that excess risks vary among constituents, and different proportions of them in PM_{2.5} may induce either greater or fewer PM_{2.5}-related effects. Information about how the constituents contribute to mortality in China can provide epidemiological evidence to safeguard human health and can also provide evidence for the differences in the health effects of PM_{2.5} observed in developed countries and China. In most heavy-polluting countries, it is difficult to comprehensively reduce emissions and achieve air quality standards in a short period of time. If the PM_{2.5}-related effects can be linked directly to certain chemical compositions or the sources of its constituents, PM_{2.5} pollution can be regulated more effectively by targeting these specific constituents. To date, there have been few studies of the links between constituents and mortality in China, and researchers have mainly examined individual cities such as Beijing,²⁰ Shanghai,²¹ Guangzhou,²² and Xi'an²³ using a single-center research design. Because monitoring for measuring composition in China is very limited and there is no nationwide network, no multicenter studies have yet been published. Existing studies^{20–24} have mainly considered carbonaceous particles and water-soluble ions (WSI) and have consistently shown that the levels of combustion species which have been shown to be largely responsible for PM_{2.5}-related health impacts, such as elemental carbon and organic carbon, are higher in China^{20–23} than in developed countries.

In light of the severe pollution and the knowledge gaps mentioned above, we utilized a time-series design to estimate how short-term exposure to different chemical constituents of PM_{2.5} influence the risks of mortality in Beijing. We also investigated seasonal patterns of associations between each constituent and mortality. Furthermore, we used the grouping factors sex and age to divide the study population into several subgroups and explored which sector of the population was most vulnerable to short-term exposure to the chemical constituents.

MATERIALS AND METHODS

Study Sites.

Beijing is the largest city in China and experiences the worst air pollution in the country. The study population included individuals who lived continuously in Beijing for at least six months. Our study covered seven districts of Beijing, including Dongcheng, Fengtai, Tongzhou, Mentougou, Changping, Yanqing, and Miyun, which together had a combined population of 6.61 million. We chose these districts because data for daily mortality, PM_{2.5} and its constituents, and meteorological factors were available for at least 1000 days

between 1 January 2013 and 31 December 2015. The locations of study sites are shown in Figure S1 in the Supporting Information.

Mortality Data.

Daily mortality counts were obtained from the Disease Surveillance Point System (DSPS) of the Chinese Center for Disease Control and Prevention (China CDC). We excluded external deaths and then classified the deaths by cause into three categories—all-cause (excluding external) mortality (ICD-10, A00–R99), cardiovascular diseases (ICD-10, I00–I99), and respiratory diseases (ICD-10, J00–J99)—in line with the International Statistical Classification of Disease, 10th Revision (ICD-10; World Health Organization 2007). We divided the data into subgroups and classified all-cause mortality by sex (male and female) and age (<65 years, 65–74 years, and >74 years).

Air Pollution Data.

We obtained PM_{2.5} data from the Beijing Municipal Environmental Monitoring Center (<http://zx.bjmemc.com.cn/>). We calculated PM_{2.5} mass concentrations by averaging the data from three fixed monitoring stations in Chaoyang District. We also obtained information about PM_{2.5} constituents—including organic carbon (OC), elemental carbon (EC), sodium (Na⁺), ammonium (NH₄⁺), potassium (K⁺), magnesium (Mg²⁺), calcium (Ca²⁺), chloride (Cl⁻), sulfate (SO₄²⁻), and nitrate (NO₃⁻)—from the particulate composition monitoring station in Chaoyang, which is measured as the average over a 24-h period. We measured the concentrations of ions in PM_{2.5} with an online monitoring instrument (MARGA ADI 2080, Metrohm and Applikon), and we measured organic carbon (OC) and elemental carbon (EC) with a real-time ECOC analyzer (U.S. Sunset Lab RT-4). We chose these constituents because they make up the majority of PM_{2.5} pollution in Beijing, and their concentration levels are mostly above the method detection limits. The detection limit for each constituent was 0.05 μg/m³. In our analysis, we replaced hourly reported values that were below the detection limit (BDL) with the detection limit values.

Meteorological Data.

We obtained data for the 24-h average temperature and relative humidity for Beijing Station No. 54511 from the Chinese Meteorological Data Network of the Chinese Meteorological Bureau, and we adjusted for the potential confounding impacts of meteorology on daily mortality.

Analytical Strategy.

We summarized daily mortality, PM_{2.5} mass and its chemical constituents, and meteorological data using the mean, standard deviation (SD), minimum value (min), maximum value (max), and interquartile range (IQR). In our time-series analysis, considering that the data over-dispersion (mean does not equal variance) assumption may be violated in a standard Poisson regression model, we applied a quasi-Poisson regression in a generalized linear model framework to estimate the association between daily mortality and PM_{2.5} constituents for each cause of death and each subgroup for the entire study period. Regression models were also run for the seasonal subset as well. We stratified the three- year

time series data by warm (May–October) and cold (November–April) seasons to explore potential seasonal influences.

In the basic model, we controlled for seasonal and long-term trends with a natural cubic regression spline with five degrees of freedom (df) per year, and same-day daily temperature and relative humidity with a natural cubic regression spline with three df for each. We included indicator variables to allow baseline mortality rates to vary for each day of the week. In the seasonal stratified model, we chose three df per year for the temporal trend and two df for the meteorological parameters. The following model was fitted to obtain an estimate of each PM_{2.5} constituent log-relative rate β

$$\log E(Y_t) = \text{intercept} + \text{ns}(\text{time, df}) + \text{ns}(\text{temperature, df}) + \text{ns}(\text{humidity, df}) + \text{dow} + \beta Z_t$$

where $E(Y_t)$ represents the expected number of deaths at day t , ns(time, df) is the natural-spline function of the temporal trend, ns(temperature, df) and ns(humidity, df) are the natural-spline functions of the meteorological parameters on day t , dow is a dummy variable for the day of the week, β represents the log-relative rate of mortality associated with a unit increase in a certain constituent of PM_{2.5}, and Z_t is the constituent of PM_{2.5} at day t .

We included same day (lag = 0) PM_{2.5} constituent exposure into the basic model to assess the mortality effect. Previous studies^{8,9} have reported strong associations between mortality and PM_{2.5} for lags of 1 or 2 days, so we also considered PM_{2.5} constituents with lags of up to 3 days (lag1, lag2, and lag3).

We examined the sensitivity of our key findings in the all-cause mortality relative to the df specified in the smooth functions of time (df = 6, 7) and meteorological indicators (df = 5) to assess model stability. We included linearly controlled daily PM_{2.5} mass in the statistical model as a potential confounder to adjust for potential influence of the daily variation of mass, and we replaced values that were below the detection level by half of the detection limit. In addition, to examine the role of outliers and extreme values, we subtracted the values above the 99th percentile from the three-year data and used this data set to perform regression analyses.

We used R version 3.4.2 for the data analysis. We expressed the effect estimates as a percent increase in mortality for each increase in the interquartile range (IQR) of the PM_{2.5} constituents.

RESULTS

During the period from 2013 to 2015, there were 87 843 deaths from all natural causes, of which 42 491 and 9 043 were related to cardiovascular disease (CVD) and respiratory disease, respectively, in the seven districts of Beijing. Descriptive statistics and seasonal characteristics of the deaths from specific causes, the PM_{2.5} mass and its chemical constituents, and meteorological data used in this present study are presented in Table 1. There were, on average, 81 nonaccidental deaths per day, comprising 42 men and 39 women. When divided into age groups, there were approximately 19, 14, and 48 deaths daily

for the <65, 65–74, and >74 age groups, respectively. On average, 39 people died from CVD and 8 people died from respiratory disease per day. The mean concentrations of PM_{2.5} in Beijing generally exceeded 90 µg/m³ (ranging from 5.5 to 512.2 µg/m³) over the study period (from 2013 to 2015). More daily deaths occurred in the cold seasons than in the warm seasons. The overall mean concentrations of OC and EC were 15.52 and 2.28 µg/m³, respectively. Of the ion chemical constituents, the concentrations of NH₄⁺ (11.97 µg/m³), NO₃⁻ (14.23 µg/m³), and SO₄²⁻ (18.52 µg/m³) were the highest. The daily temperature and relative humidity averaged around 13.6 °C and 58%, respectively. A seasonal pattern was exhibited in the average concentrations of PM_{2.5} and compositions, which was substantially higher in the cold seasons. Table S1 of the Supporting Information displays correlations among PM_{2.5} mass and constituents. The Pearson's coefficients of PM_{2.5} mass with OC, EC, NH₄⁺, NO₃⁻, and SO₄²⁻ were above 0.7 (*p* < 0.05). The major constituents (OC, NH₄⁺, NO₃⁻, and SO₄²⁻) were highly correlated with each other (range: 0.62–0.82).

Year-round and season-specific estimates [percentage increase and 95% confidence interval (CI)] of PM_{2.5} mass and the constituents from regression models are presented in Figure 1. In the year-round regression model, most point estimates of the constituents were positive except for Na⁺ and Cl⁻. Adjusted positive associations of the per IQR increase in OC, K⁺, Ca²⁺, and Mg²⁺ concentrations on current-day exposure with respiratory mortality and of SO₄²⁻ concentration with cardiovascular mortality were found at the *p* < 0.05 level: In terms of results in the season-stratified analysis, OC and EC were associated with all-cause mortality in the warm season at lag 0 day: 2.74% (95% CI, 0.60–4.91) and 2.02% (95% CI, 0.00–4.08), respectively. Stronger associations were observed in the cold season: an IQR increase in NO₃⁻ was associated with a 1.33% increase (95% CI, 0.38–2.29) in all-cause mortality, IQR increases in NO₃⁻ and SO₄²⁻ were associated with a 1.86% increase (95% CI, 0.51–3.23) and a 2.69% increase (95% CI, 0.90–4.51) in cardiovascular mortality, and IQR increases in OC, K⁺, Ca²⁺, Mg²⁺, and NO₃⁻ were associated with a 3.04% (95% CI, 0.36–5.79), 0.59% (95% CI, 0.24–0.94), 2.60% (95% CI, 1.05–4.17), 0.27% (95% CI, 0.08–0.46), and 2.87% (95% CI, 0.16–5.66) increase, respectively, in respiratory mortality.

The estimated risks of mortality per IQR increase in PM_{2.5} and the constituents calculated by the lag structure models from 2013 to 2015 are presented in Figure 2. Compared to other lag days, the highest estimates were observed at lag 0 or 1 in the year-round regression model.

The associations between the PM_{2.5} chemical constituents and all-cause mortality by age and sex subgroups are presented in Table 2. There were no statistical differences between the estimated risks of short-term exposure to the mass and the constituents between men and women. We also found that four constituents were responsible for greater risks in the 65–74 age group: mortality by all causes increased by 2.68% (95% CI, 0.44–4.96), 2.71% (95% CI, 0.08–5.4), 2.28% (95% CI, 0.38–4.22), and 2.51% (95% CI, 0.01–5.07) for the 65–74 age group with IQR increases in the concentrations of EC, NH₄⁺, NO₃⁻, and SO₄²⁻.

We examined the effects of the PM_{2.5} constituents using alternative values of *df* for the temporal trend and meteorological indicators in sensitivity analyses, and the results presented in Supporting Information Table S2 suggest that our results were relatively robust.

There were few changes in the estimated effects of the PM_{2.5} constituents, apart from Ca²⁺, NH₄⁺, and NO₃⁻, when the number of df per year changed between 5 and 7 in the temporal trend. Between 3 and 5 df, the estimated effects of the PM_{2.5} constituents were almost identical when the number of df for temperature and relative humidity changed. When we compared the results from single- and two-pollutant models that included PM_{2.5} mass, we found that there were large changes in the effects of OC, Cl⁻, NO₃⁻, and SO₄²⁻ on all-cause mortality at 0 lag. When the BDL values were replaced by half of the detection limit (Supporting Information Table S3), the effects of the PM_{2.5} constituents produced little change in all-cause mortality at 0 lag. By comparing the model estimates between the full data set and subset which had values below the 99th percentile, we found that the point estimates of Na⁺, K⁺, Ca²⁺, Mg²⁺, and Cl⁻ changed a lot, but their confidence intervals still overlapped each other (Table S4 of the Supporting Information).

DISCUSSION

This study is one of very few to date that has attempted to quantify the acute effects of chemical constituents of PM_{2.5} in developing countries. We found that the pollution concentrations of carbonaceous particles, secondary aerosol anions, and some WSI were much higher than those reported in developed countries, and that short-term exposure to certain constituents was associated with increases in mortality risks. Our results also suggest that those constituents might increase the risks of PM_{2.5}-related mortality in different seasonal patterns. OC and EC were significantly associated with all-cause mortality in warm seasons, and strong associations of OC, K⁺, Ca²⁺, Mg²⁺, NO₃⁻, and SO₄²⁻ were observed in cold seasons. Most commonly, effect estimates were higher for lag 0 or 1 day exposure. Our findings supplement existing evidence that shows PM_{2.5}-related effects are modified by age, and that people between 65 and 74 years old are particularly susceptible to EC, SO₄²⁻, and NH₄⁺.

As in previous studies,^{10–16} the estimates of acute effects were mostly related to the constituents that made the greatest contributions to the PM_{2.5} mass. The constituents most frequently considered in studies carried out in the U.S. and Europe were SO₄²⁻, carbonaceous particles (usually EC and OC), and some metals.^{10–16} Sulfate and carbonaceous particles were positively associated with increases in all-cause, CVD, and respiratory mortalities.^{10,11} However, the roles of other PM_{2.5} constituents are still uncertain, as those constituents were present in quantities that were insufficient for pooling estimates.¹¹ Earlier Chinese studies have generally focused on the adverse effects of sulfate and WSI.^{20–23} Our study and two other studies, Qiao et al. (2014)²¹ and Lin et al. (2016),²² included carbon-containing components to examine. It is, however, difficult to reach consistent conclusions mainly because of regional variations, such as the temporal–spatial distributions of PM_{2.5} constituents, population adaptation, incomparable methodologies, and insufficient evidence.

Carbonaceous particles are considered the most important contributors to PM_{2.5}-related health impacts.^{8–16} Most studies of mortality or hospital admissions and biomarkers have reported significant associations between CVD mortalities^{8,9,11–13} and either EC, OC, or both. Positive excess risks have also been estimated for respiratory outcomes but less

consistently. In a systematic review of epidemiological evidence for the adverse effects of particle components worldwide, Atkinson et al. (2015)¹¹ conducted a meta-analysis of 19 different studies of OC/EC and mortality and found that an increase of 1 $\mu\text{g}/\text{m}^3$ in OC and EC was significantly associated with increases of 0.56% (0.01–1.10%) and 1.66% (0.52–2.81%) in CVD mortality but not respiratory mortality. In our study, OC was positively associated with the increase in respiratory mortality risk, and this result was consistent with findings from two single-city studies in Xi'an, China²⁵ and Madrid, Spain,²⁶ which support a link between exposure to OC and increased risks of respiratory mortality. Therefore, our findings suggest that OC is one of the major constituents which is responsible for the adverse respiratory effects of $\text{PM}_{2.5}$. In addition, short-term exposure to carbonaceous particles was reported to be associated with increased concentrations of exhaled nitric oxide,^{27,28} an indicator of airway inflammation, which provides the biological plausibility that carbonaceous particles may contribute to respiratory mortality.

We found positive associations between EC and daily mortality in our main analysis, and this association became statistically significant in our age-stratified analyses. Bell et al. (2009)¹² examined people of 65 years or older in 106 U.S. counties and concluded that estimates of the short-term effects of EC were statistically and positively related to cardiovascular and respiratory hospitalizations. The magnitude of the EC effect on the older group had a greater number of mortalities in our study than in the study of the general population of the U.S., in which an IQR increase in EC was associated with a 0.22% (95% PI, 0.00–0.44) increase in mortality risk. These results suggest that EC has an important role in $\text{PM}_{2.5}$ -related acute effects on mortality and that it may have been modified by age.

Our results also suggest that WSI may impact mortality and especially respiratory-related mortality. Penttinen et al.²⁹ examined a panel of adult asthmatics from urban Helsinki and found a soil factor containing Ca^{2+} and Mg^{2+} that was negatively associated with any changes in peak expiratory flow. Our estimate of the effects of Mg^{2+} on mortality was slightly lower than that of Son et al.³⁰ in Korea, which reported that all-cause mortality increased by 1.4% (95% CI, 0.2–2.6) per IQR increase in Mg^{2+} . SO_4^{2-} had the largest association of all measured WSI species. Its adverse effect on CVD has been presented in a recent meta-analysis.³¹ One study in Guangzhou of China (2007–2010)²² reported that the mortality of the general population from CVD increased by 2.21% (95% CI, 1.05–3.38) per IQR increase in sulfate at daily average concentrations of around 18.52 $\mu\text{g}/\text{m}^3$.

There is an increased recognition of seasonal variations in constituents contributing to the seasonal pattern in association with short-term exposure to $\text{PM}_{2.5}$ and mortality. As toxic constituents appear in certain time periods, the adverse effects of $\text{PM}_{2.5}$ would differ across different seasons. Our study showed a stronger association between six constituents (OC , K^+ , Ca^{2+} , Mg^{2+} , NO_3^- , and SO_4^{2-}) and mortality in cold seasons, especially for respiratory-related mortality, when their concentrations were higher. In Beijing, Ca^{2+} was mainly derived from soil dust and construction materials in urban aerosol.³² K^+ most likely comes from open stalk burning and household heating in cold seasons.³³ Poor weather conditions in cold seasons may lead to increased concentrations of certain constituents. SO_4^{2-} and NO_3^- represent secondary aerosols formed by the transformation of their precursors, SO_2 and nitrogen oxides, in the atmosphere.³² The mass ratio of $\text{NO}_3^-/\text{SO}_4^{2-}$ has been used as

an indicator to evaluate the relative importance between mobile and stationary sources of nitrate and sulfate in the atmosphere.³³ In our study, these ratios of cold seasons and warm seasons were below 1, which may point to the importance of mobile source emissions. The ratio would be greater than 1 when stationary emissions were more prevailing.³⁴ Both increasing the number of vehicles and emission reduction of atmospheric SO₂, e.g., a strict implementation of flue gas desulfurization in coal-fired power plants, would increase the ratio.

We found greater estimates of OC and EC occurring in warm seasons when their concentrations slightly declined. Recent evidence on source apportionment and seasonal variation of carbonaceous particles in Beijing showed that OC and EC in PM_{2.5} across seasons were mainly generated from coal combustion, vehicle exhaust, and biomass burning, with vehicle exhaust contributing to 63% of emissions in the summer.³⁵ This means that OC not only includes primary OC but also secondary OC produced via photochemical reactions in the summer. EC has stable chemical properties and is mainly produced by direct emissions from incomplete combustion. EC in Beijing mainly comes from gasoline and diesel vehicle emissions.³⁵ Thus, these carbonaceous particles were not greatly different between warm and cold seasons. In addition, seasonal associations between components and mortality did not necessarily match the seasonal variations of concentrations from fixed monitors, possibly because of personal behavioral factors increasing personal exposure which would lead to a larger effect. Activities such as more open windows and ventilation in warm seasons and longer times for outdoor activities could contribute to higher air pollutant exposure.

Studies have frequently reported that people older than 65 years are more susceptible to PM_{2.5} short-term exposure than younger people^{1,8,19} because of degraded functioning of cardiopulmonary systems and increased prevalence of pre-existing chronic diseases (such as coronary heart disease, hypertension, and chronic obstructive pulmonary disease) in the elderly. However, studies of the vulnerability of subpopulations to PM_{2.5} chemical constituents are still limited. We examined our data to ascertain whether there were any subpopulation patterns between health effects and the PM_{2.5} constituents. We found that the 65–74 year old age group differed from the others and that the estimated effects for EC, SO₄²⁻, and NH⁺₄ were particularly different. The smaller magnitudes of estimated risks for the 74 years old and above age group, compared to the 65–74 years old age group, might be because of the lower intensity and shorter time spent on outdoor physical activities. Given the results of indoor/outdoor ratios of PM_{2.5} that are less than 1 in Beijing,^{36,37} the elderly population (>74 years old) is expected to be exposed to lower levels of PM_{2.5} mass and constituents. Ostro and colleagues⁸ also found that mortality was linked with certain PM_{2.5} components among those over 65 years of age. There is increasing evidence of the acute effects associated with exposure to secondary aerosol anions, especially sulfate.^{10,11} In a review of 12 single- and multicity studies, all-cause mortality was estimated to increase by 0.15% (0.06, 0.25) per 1 μg/m³ increase in sulfate.¹¹ There was insufficient epidemiologic evidence for a meaningful meta-analysis, and while these summary estimates showed positive effects in the general population, they were imprecise at the subgroup level. Our findings could be improved to provide more evidence about whether subgroups are vulnerable to specific PM_{2.5} components.

There are several limitations associated with this study, as follows. An observational study cannot identify the independent impacts of individual constituents. One common strategy to examine the association between constituents and health outcomes involves examining individual constituent concentration alone in modeling work.^{8,9,12,21} With this strategy, it is easy to interpret the question of what the effect of a specific constituent is. However, since PM_{2.5} is positively correlated with both constituent concentration and health outcomes, estimates of individual constituents on adverse effects would be confounded and biased upward.³⁸ Therefore, we included both the concentrations of individual constituent and PM_{2.5} mass in a two-pollutant modeling work to explore the estimates of specific constituents after adjusting for total PM_{2.5}. The model parameter for individual constituent concentration represents the impact of the constituent (and its correlates), holding the other constituents constant.³⁸ The model parameter for PM_{2.5} mass concentration represents the difference in mortality risk associated with the other constituents.³⁸ However, this association may be prone to overadjustment if constituents are strongly correlated with PM_{2.5}.³⁸ In our two-pollutant model, estimates of OC, Cl⁻, NO₃⁻, and SO₄²⁻ were much smaller after adjustment for PM_{2.5}, which indicate strong correlations with total PM_{2.5}. It is hard to say whether the smaller effect estimates represent an accurate portrayal of the impact of certain constituents independent of PM_{2.5} or not. This raises an issue that the current statistical strategy of an observational study cannot completely interpret the roles of specific constituents in PM_{2.5}-related risks.

Our study was based on the assumption that constituent concentrations of a fixed-monitoring site located in the center of a city are representative of population exposure and did not account for exposure misclassification. In fact, there could be two types of exposure misclassification in this study. One is a classical type, caused by variations between the ambient level and personal exposure. The true personal exposure is varied by the outdoor level, individual time–activity patterns, residential characteristics, and indoor sources of pollutants. And this exposure misclassification tends to bias associations between pollutants and health outcomes toward the null.^{39,40}

The other type occurs due to spatial heterogeneity of constituent levels. Evidence on community-level spatial heterogeneity of constituent levels showed that exposure misclassification differed by constituents.³⁹ Ambient secondary pollutants, such as OC, sulfate, nitrate, and ammonium, had less spatial variation than primary pollutants (e.g., EC) and were uniformly distributed within a spatial unit, which could have better representation and less exposure error.³⁹ Future work should focus on obtaining a more accurate picture of exposure or new methods for assessing exposure error induced by spatial variability.

In addition, when the daily concentrations of specific constituents are very low, the values of measurement may be below the detection level, resulting in an unstable evaluation of the constituents. Although we have performed a sensitivity analysis on values below the detection limits, there was still a limitation of detection limits which might result in inaccurate estimates. The degree of measurement error may differ between components. Some components with many BDL days (e.g., Ca²⁺, Na⁺, Mg²⁺, and K⁺) had a large proportion of inaccurate values in the entire time series. The uncertainty created from these measurement errors of constituents leads to large variances of the residuals from models,

which may negatively influence their estimates.³⁸ The coefficients for some models may be larger simply due to less measurement error.³⁸

In our sensitivity analysis removing extreme values, regression estimates of these components were unstable and varied greatly. Further research is needed to develop statistical methods which could adjust estimate error obtained from detection limits and extreme values. Furthermore, the study design was ecological in nature, and it was difficult to identify the role of one specific PM_{2.5} constituent as the estimates of the effects of single constituents might have been a collaborative proxy of many constituents. Finally, because of limited data availability, our study was only conducted in Beijing over a relatively short time period. To account for the differences in the temporal–spatial distribution of chemical constituents and reduction of measurement error, future epidemiological studies of large populations will be developed over longer time frames at the national scale.

In our study, we found evidence to support our hypothesis that specific chemical constituents of PM_{2.5} play a role in PM_{2.5}-related mortality risks. Our findings suggest that adverse effects vary among constituents in different seasons, and certain constituents may be more harmful to the elderly. We believe that stricter regulations on PM_{2.5} mass emissions in conjunction with targeting specific, more toxic components of PM_{2.5} will be a more effective way of protecting the public's health.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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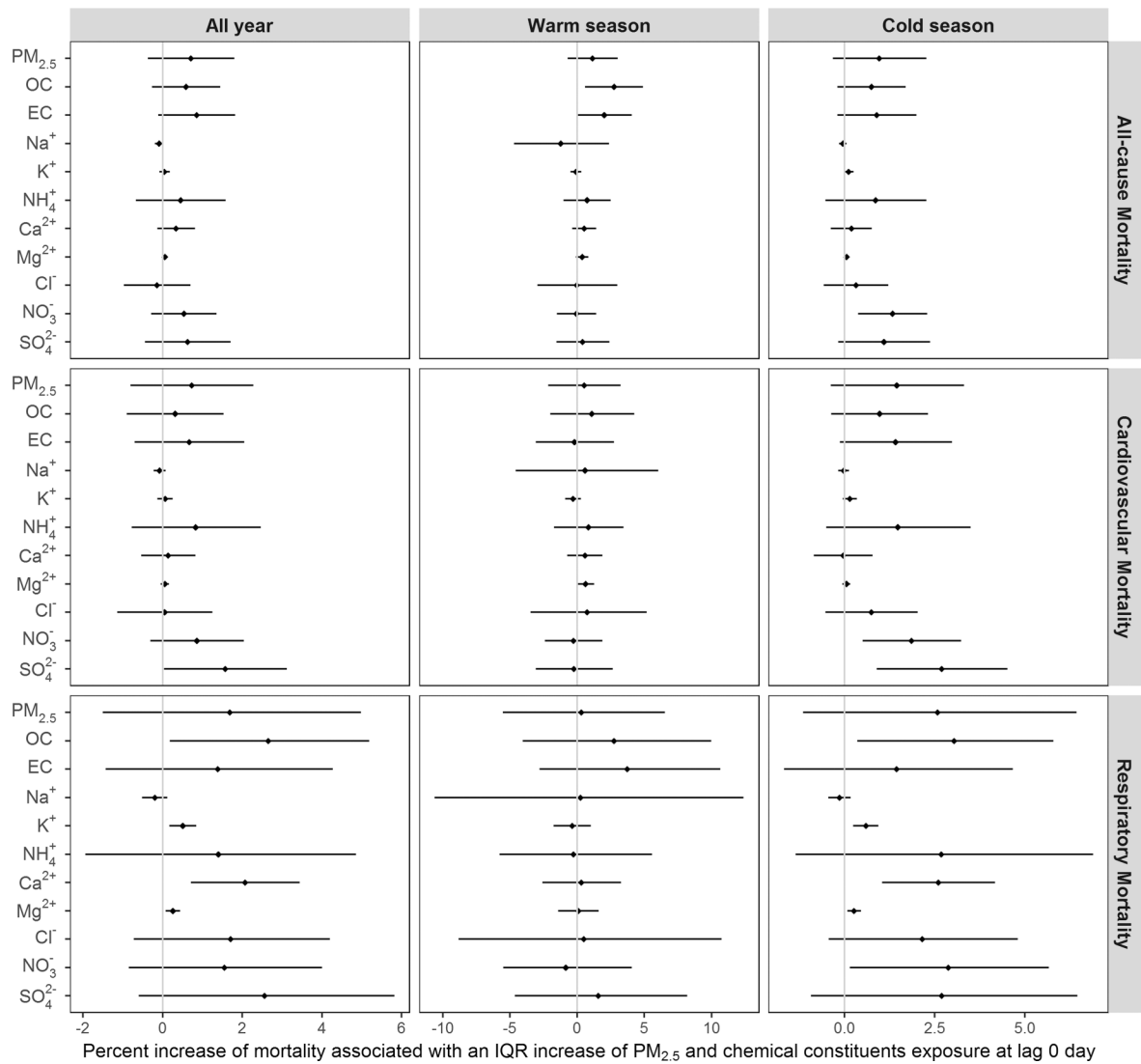


Figure 1. Estimated risk for mortality of year-round and seasonal stratification per IQR increase in PM_{2.5} mass and chemical constituents, adjusted for temporal trend, day-of-week, and same-day meteorological parameters (temperature and relative humidity).

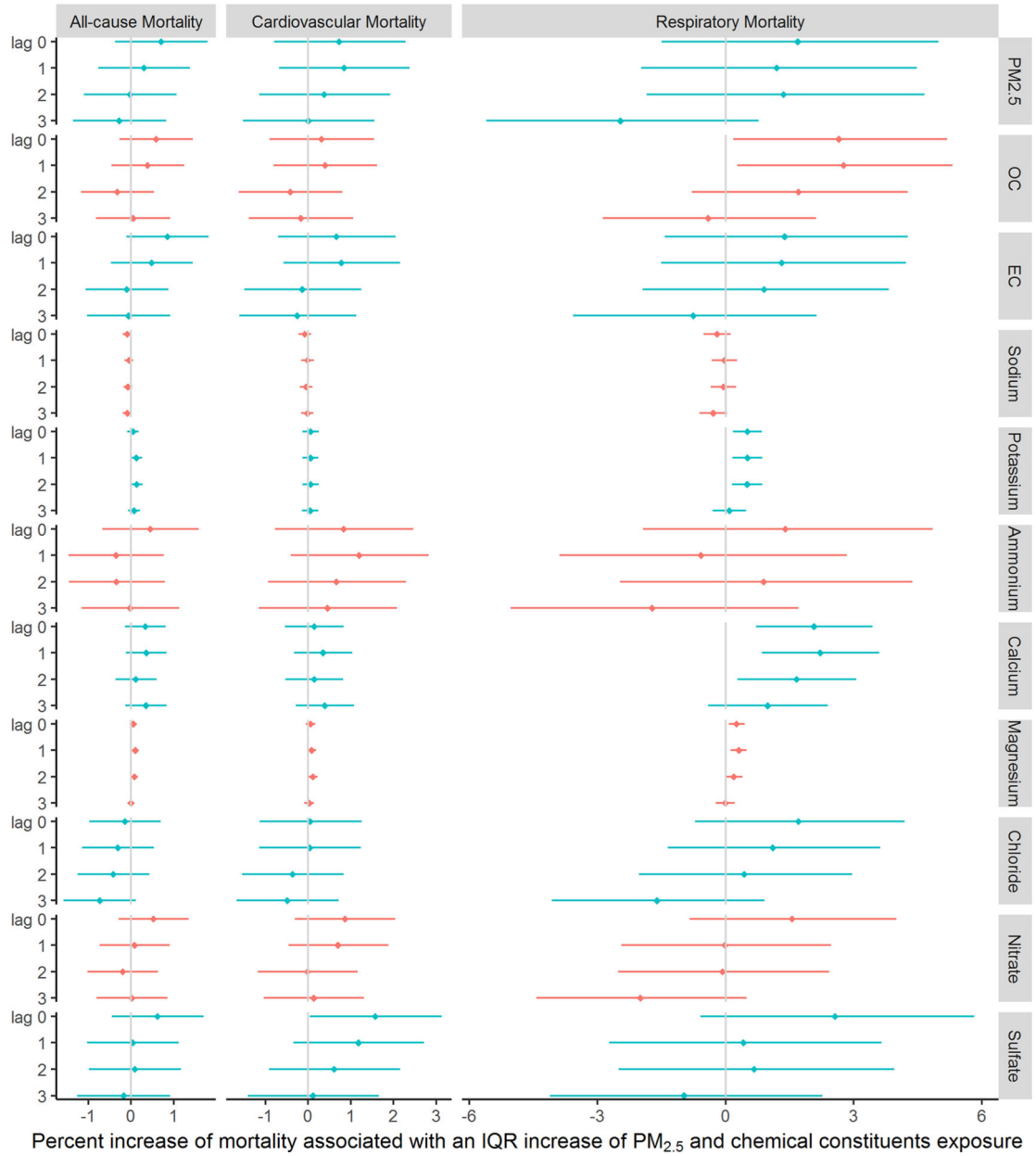


Figure 2. Estimated effects of 0–3 day lags in PM_{2.5} mass and chemical constituents.

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Table 1. Year-Round and Seasonal Stratified Summary Statistics for Daily Mortality, PM_{2.5} Mass and Chemical Constituents, and Meteorological Factors in Beijing, China (from January 2013 to December 2015)

	days of data available	year-round			warm seasons			cold seasons					
		mean	SD	IQR	min, max	mean	SD	mean	SD				
cause of mortality													
all cause (excluded external deaths)	1083	81	13	19	46, 126	74	10	88	12				
male	1083	42	9	13	20, 74	38	8	46	9				
female	1083	39	9	12	19, 70	36	8	43	9				
0-64 years old	1083	19	5	6	6, 39	18	4	21	5				
65-74 years old	1083	14	4	6	4, 27	13	3	15	4				
>74 years old	1083	48	9	12	20, 81	43	7	53	9				
cardiovascular diseases	1083	39	9	12	15, 70	35	7	44	8				
respiratory diseases	1083	8	4	4	1, 24	7	3	10	4				
air pollution (24-h average, $\mu\text{g}/\text{m}^3$)													
PM _{2.5}	1095	92.6	77.4	83.0	5.5, 512.2	78.1	60.7	107.4	89.0				
organic carbon	1078	15.52	11.49	10.11	0.05, 84.77	11.41	6.31	19.62	13.79				
elemental carbon	1078	2.28	1.95	2.03	0.05, 17.59	1.73	1.34	2.83	2.28				
Na ⁺	1079	1.46	6.72	0.89	0.05, 128.63	0.41	0.36	2.54	9.42				
K ⁺	1049	1.58	5.62	1.01	0.05, 91.65	0.63	2.61	2.61	7.50				
NH ₄ ⁺	1080	11.97	11.65	12.44	0.05, 73.88	10.38	10.02	13.58	12.90				
Ca ²⁺	1080	0.40	0.61	0.37	0.05, 6.38	0.32	0.50	0.49	0.69				
Mg ²⁺	1077	0.15	0.44	0.05	0.05, 8.55	0.09	0.12	0.21	0.60				
Cl ⁻	1082	2.34	3.28	2.86	0.05, 27.61	0.87	1.33	3.84	3.93				
NO ₃ ⁻	1082	14.23	16.72	14.50	0.05, 126.13	11.76	13.62	16.74	19.05				
SO ₄ ²⁻	1083	18.52	18.19	20.10	0.05, 120.00	14.88	13.90	21.99	21.10				
meteorological parameters													
temperature (°C)	1094	13.7	11.0	20.6	-9.7, 32.6	22.6	5.3	4.5	7.0				
humidity (%)	1095	54	20	31	8, 99	61	17	47	20				

Percent Increase (95% Confidence Interval) in All-Causes Mortality by Sex and Age Group Associated with an IQR Increase at Lag 0 in Beijing, China from 2013 to 2015

Table 2.

PM _{2.5} components	sex		age			
	male	female	<65	65-74	>74	
PM _{2.5}	0.86 (-0.63, 2.36)	0.53 (-0.97, 2.05)	-0.22 (-2.36, 1.97)	2.24 (-0.27, 4.81)	0.59 (-0.76, 1.96)	
organic carbon	0.71 (-0.48, 1.89)	0.53 (-0.67, 1.74)	0.49 (-1.22, 2.23)	1.51 (-0.48, 3.54)	0.40 (-0.67, 1.48)	
elemental carbon	1.06 (-0.23, 2.38)	0.64 (-0.71, 2.01)	0.32 (-1.59, 2.26)	2.68 (0.44, 4.96)	0.53 (-0.67, 1.74)	
Na ⁺	-0.2 (-0.34, -0.05)	0.03 (-0.11, 0.18)	-0.02 (-0.22, 0.19)	-0.16 (-0.41, 0.10)	-0.1 (-0.24, 0.03)	
K ⁺	0.01 (-0.17, 0.19)	0.10 (-0.10, 0.29)	0.05 (-0.22, 0.32)	0.20 (-0.11, 0.50)	0.01 (-0.16, 0.17)	
NH ₄ ⁺	0.69 (-0.84, 2.24)	0.20 (-1.37, 1.79)	-0.92 (-3.12, 1.32)	2.71 (0.08, 5.40)	0.34 (-1.07, 1.76)	
Ca ²⁺	0.41 (-0.23, 1.06)	0.19 (-0.48, 0.86)	0.36 (-0.58, 1.31)	-0.21 (-1.31, 0.91)	0.43 (-0.17, 1.02)	
Mg ²⁺	0.02 (-0.08, 0.12)	0.09 (-0.01, 0.20)	0.03 (-0.12, 0.19)	0.14 (-0.03, 0.31)	0.03 (-0.06, 0.13)	
Cl ⁻	-0.16 (-1.30, 1.00)	-0.10 (-1.26, 1.08)	-0.42 (-2.08, 1.27)	1.28 (-0.66, 3.26)	-0.44 (-1.48, 0.61)	
NO ₃ ⁻	0.55 (-0.57, 1.67)	0.52 (-0.62, 1.68)	-0.44 (-2.05, 1.20)	2.28 (0.38, 4.22)	0.40 (-0.62, 1.43)	
SO ₄ ²⁻	0.24 (-1.20, 1.70)	1.04 (-0.48, 2.58)	-0.83 (-2.93, 1.32)	2.51 (0.01, 5.07)	0.64 (-0.70, 2.00)	