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Prognostic factors of survival in patients with malignant pleural mesothelioma: an analysis of the National Cancer Database

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Abstract

Malignant pleural mesothelioma (MPM) is a rare disease with a very poor prognosis. Previous studies have indicated that women experience longer survival compared with men. We analyzed 16 267 eligible patients (21.3% females) in the National Cancer Database to evaluate which clinical factors are independently predictive of longer survival. After adjusting for all covariates, survival was significantly better in females compared with males [HR_{adj}: 0.81, 95% confidence interval (CI): 0.77–0.85]. Other factors significantly associated with better survival were younger age at diagnosis, higher income, lower comorbidity score, epithelial histology, earlier stage and receipt of surgical or medical treatment. After propensity matching, survival was significantly better for females compared with males [hazard ratio (HR): 0.86, 95% CI: 0.80–0.94]. After propensity matching within the epithelial group, survival remained significantly better for females compared with males (HR: 0.85, 95% CI: 0.74–0.97). This study adds information to the known significant gender survival difference in MPM by disentangling the effect of gender from the effect of age and histology, two known independent factors affecting survival. Circulating estrogen, present in young but not older women, and higher expression of the estrogen receptor beta in epithelial mesothelioma have been suggested to play a role in gender survival differences. These findings may lead to exploring new therapeutic options, such as targeting estrogen receptor beta, and considering hormonal therapy including estrogens for patients with otherwise limited prognosis.

Introduction

Malignant pleural mesothelioma (MPM) is a rare but aggressive form of thoracic cancer linked to asbestos exposure. Patients with MPM have historically had a very poor prognosis, with modest changes in survival observed over time, despite the introduction of modern therapeutic interventions (1). Previous analyses of the Surveillance, Epidemiology, and End Results (SEER) database (2–5) and single-center studies (6–8) have suggested that females with MPM experience longer survival compared with males. However, MPM is a rare disease and only a small proportion of affected patients are female, which limits the extent of conclusions that can be drawn from single-institution databases.

Among the possible reasons for the observed longer survival experienced by women, authors suggested that they present at earlier stage (9), have tumors with more favorable histology (4), experience a different amount or type of asbestos exposure responsible for a more indolent tumor biology (10), and may benefit from protective effects of circulating estrogen (6) interacting with estrogen receptors present in their tumors (11).

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Abbreviations	
CI	confidence interval
CoC	Commission on Cancer
ERβ	estrogen receptor beta
HR	hazard ratio
MPM	malignant pleural mesothelioma
NCDB	National Cancer Database
OS	overall survival

Because of the aggressive nature of MPM and its poor prognosis, identifying prognostic factors and characterizing their relation with therapeutic options are critical objectives. Currently, therapeutic options are still of limited efficacy and include surgery, chemotherapy, radiation and combinations of these treatments. The suggestion that female patients with MPM survive longer than male patients and the possible association with histology, stage or the presence of estrogens and their interaction with estrogen receptors, may be key elements for identifying novel therapeutic options which could ultimately improve patient outcomes.

To date, no large population-based studies have specifically evaluated MPM outcomes and their determinants in female patients. We used the National Cancer Database (NCDB) to retrospectively evaluate survival in patients with MPM according to gender and determine which clinical factors are independently predictive of longer survival in females.

Materials and methods

Data source

The NCDB is a joint project between the American College of Surgeons and the American Cancer Society which is sourced from hospital registry data collected in more than 1500 Commission on Cancer (CoC)-accredited facilities starting in 1989. Patients who receive some element of their cancer care at a CoC-accredited facility are included in the NCDB, representing ~70% of all patients newly diagnosed with cancer nationwide (12). The NCDB offers data about cancer characteristics, patient demographics, reporting facility characteristics, first course of treatment and survival, as described previously in more detail (13,14). Because the data used in the study were derived from a de-identified NCDB file, the research was considered exempt by the Institutional Review Board at the Icahn School of Medicine at Mount Sinai. The American College of Surgeons and the CoC have not verified and are not responsible for the analytic or statistical methodology used, or the conclusions drawn from these data by the investigator.

Study population

The initial NCDB dataset contained 23 414 patients diagnosed with mesothelioma between 2004 and 2013. The sample was limited to 19 134 patients with invasive pleural mesothelioma. As recommended by the NCDB, only patients (i) with at least a part of their treatment at the CoC facility that reported them and (ii) who were diagnosed after their facility's reference date for data completeness were included in this study ($n = 16\ 267$) (Figure 1).

Predictor and primary outcome

The primary predictor was gender, whereas the primary outcome was overall survival (OS) after diagnosis. The NCDB recorded the number of months of follow-up after diagnosis and the patient's vital status at that time.

Covariates

Covariates included age at diagnosis, race/ethnicity, insurance status, percent of adults without a high school diploma in the patient's zip code of residence, median income in the patient's zip code of residence, Charlson Comorbidity Index, facility type, distance from the patient's zip code to the reporting facility, histology, bilateral involvement, stage and receipt of treatment, including surgery, radiotherapy and chemotherapy. Receipt of radiotherapy was defined if the patient received 40–65 Gy of external beam radiation therapy to the chest wall, lungs or pleura as part of the first course of treatment (15).

Statistical analysis

Categorical demographic and clinical characteristics were compared between males and females using χ^2 tests. Independent associations of gender with these variables were assessed using multivariable logistic regression to estimate odds ratios and 95% CIs. Patients missing any covariate information (except stage) were excluded from analysis. As ~25% of the sample had incomplete information to define stage, a missing category was created for analysis.

The Kaplan–Meier method, along with the log-rank test was used to estimate and compare univariate OS at 2 and 5 years in males and females. A multivariable Cox proportional hazards model was used to assess the association between gender and OS, adjusted for possible confounders, including year of diagnosis and US census division of the reporting facility. The association between gender and OS was also analyzed using a 1:1 propensity score matching with the greedy algorithm, matching on all covariates. A stratified Cox proportional hazard model was then used to assess the association between gender and OS. Stratification according to age at diagnosis (<50 years, \geq 50 years) was done, and within each age group, the propensity score matching analysis was repeated. A similar propensity score matching analysis was also conducted among those with epithelial histology. All statistical analyses were performed using SAS software, v9.4 (SAS Institute, Cary, NC).

Results

We identified 16 267 patients with MPM who met the selection criteria of which 21.3% were female (Table 1). At diagnosis, females were significantly younger (<60 years: 17.2 versus 9.9%), less likely to be on Medicare (65.3 versus 70.6%) but more likely to be on Medicaid (3.3 versus 1.9%), had lower comorbidity scores (Charlson Comorbidity Index = 0: 70.5 versus 67.6%) and tended to live closer to their reporting facility (within 25 miles: 74.3 versus 71.9%) than males. Females had significantly more epithelial cancer (40.3 versus 35.0%), and were less likely to receive chemotherapy as a first course of treatment than males (45.9 versus 49.3%) (Table 1).

After adjustment, females were still diagnosed at younger age, had a lower comorbidity score and were more likely to be on Medicaid compared with males (Table 2). Furthermore, females had significantly more epithelial histology and were less likely to receive chemotherapy compared with males (Table 2).

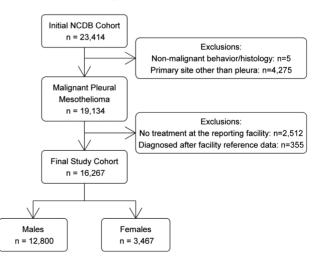


Figure 1. Selection criteria.

				Propensity-matched sample	tched sample	
	Males	Females		Males	Females	
Variable	(n = 12 800) No. (%)	(n = 3467) No. (%)	Pa	(n = 2374) No. (%)	(n = 2374) No. (%)	Ра
Age (years)			<0.0001			0.9877
<50	261 (2.0)	176 (5.1)		76 (3.2)	84 (3.5)	
50-59	1013 (7.9)	422 (12.2)		286 (12.0)	283 (11.9)	
60-69	3115 (24.3)	747 (21.5)		519 (21.9)	510 (21.5)	
70–79	4669 (36.5)	1086 (31.3)		754 (31.8)	748 (31.5)	
280	3742 (29.2)	1036 (29.9)		739 (31.1)	749 (31.6)	
Race			0.1665			0.4704
White	11 909 (93.0)	3200 (92.3)		2187 (92.1)	2198 (92.6)	
Black	561 (4.4)	158 (4.6)		116 (4.9)	113 (4.8)	
Asian/other	330 (2.6)	109 (3.1)		71 (3.0)	63 (2.7)	
Spanish/Hispanic origin			0.0748			0.6362
No	11 416 (89.2)	3064 (88.4)		2252 (94.9)	2259 (95.2)	
Yes	495 (3.9)	157 (4.5)		122 (5.1)	115 (4.8)	
Missing/unknown	889 (6.9)	246 (7.1)				
% of adults in patient zip code without a HS degree (quartiles)			0.0089			0.8237
≥21	1631 (12.7)	449 (13.0)		331 (13.9)	321 (13.5)	
13–20.9	3065 (23.9)	841 (24.3)		580 (24.4)	581 (24.5)	
7–12.9	4576 (35.8)	1145 (33.0)		805 (33.9)	818 (34.5)	
<7	3153 (24.6)	933 (26.9)		658 (27.7)	654 (27.5)	
Missing/unknown	375 (2.9)	99 (2.9)				
Patient zip code median income (\$) (quartiles)			0.0353			0.4103
<38 000	1704 (13.3)	462 (13.3)		327 (13.8)	304 (12.8)	
38 000-47 999	2925 (22.9)	735 (21.2)		513 (21.6)	501 (21.1)	
48 000–62 999	3595 (28.1)	950 (27.4)		652 (27.5)	688 (29.0)	
≥63 000	4197 (32.8)	1221 (35.2)		882 (37.2)	881 (37.1)	
Missing/unknown	379 (3.0)	99 (2.9)				
Insurance			<0.0001			0.7976
Uninsured	221 (1.7)	70 (2.0)		40 (1.7)	46 (1.9)	
Private insurance	2947 (23.0)	931 (26.9)		653 (27.5)	648 (27.3)	
Medicaid	246 (1.9)	113 (3.3)		77 (3.2)	84 (3.5)	
Medicare/other government	9033 (70.6)	2265 (65.3)		1604 (67.6)	1596 (67.2)	
Missing/unknown	353 (2.8)	88 (2.5)				
Charlson Comorbidity Index			0.0020			0.2013
0	8647 (67.6)	2443 (70.5)		1729 (72.8)	1678 (70.7)	
1	2972 (23.2)	755 (21.8)		468 (19.7)	518 (21.8)	
≥2	1181 (9.2)	269 (7.8)		177 (7.5)	178 (7.5)	
Facility type			0.2436			0.5090
Community Cancer Program	7384 (57.7)	1944 (56.1)		1390 (58.6)	1368 (57.6)	
Academic/Integrated Network	5354 (41.8)	1475 (42.5)		984 (41.4)	1006 (42.4)	
Missinghin Incite		10 /1 /1				

$ \begin{array}{cccccccccccccccccccccccccccccccccccc$					Propensity-matched sample	cched sample	
n = 2400 $n = 2400$ $n = 2$		Males	Females		Males	Females	
ce from patient to reporting facility (miles) 00117 1385 (777) 383 (776) 264 (11.1)	Variable	(n = 12 800) No. (%)	(n = 3467) No. (%)	Pa	(n = 2374) No. (%)	(n = 2374) No. (%)	Ра
$ \begin{array}{cccccccccccccccccccccccccccccccccccc$	Distance from patient to reporting facility (miles)			0.0117			0.9645
$ \begin{array}{cccccccccccccccccccccccccccccccccccc$	<25 - 25 - 2 - 2 - 2 - 2 - 2 - 2 - 2 - 2	9209 (71.9)	2577 (74.3)		1845 (77.7)	1843 (77.6)	
	25-49.9	1483 (11.6)	383 (11.0)		265 (11.2)	267 (11.2)	
ing/unknown $37 (2.9)$ $55 (2.7)$ -0.001 -0.001 $979 (41.2)$ $948 (35.9)$ inglindl $126 (13.8)$ $259 (7.5)$ $136 (40.7)$ $130 (7.6)$ $130 (7.6)$ indl $126 (13.8)$ $259 (7.5)$ $136 (4.7)$ $130 (7.6)$ $130 (7.6)$ indl $126 (13.8)$ $555 (13.4)$ $1518 (45.7)$ $130 (7.6)$ $130 (7.6)$ indl $555 (13.4)$ $1518 (45.7)$ $133 (45.2)$ $133 (45.2)$ $133 (45.2)$ indl $555 (13.6)$ $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ indl $232 (13.2)$ $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ indl $232 (13.2)$ $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ indl $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ indl $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ indl $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ indl $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ indl $143 (12.2)$ $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ indl $143 (12.2)$ $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ indl $110 (12.2)$ $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ $233 (13.2)$ indl $110 (12.2)$ $110 (12.2)$ $110 (12.2)$ $110 (12.2)$	≥50	1741 (13.6)	412 (11.9)		264 (11.1)	264 (11.1)	
	Missing/unknown	367 (2.9)	95 (2.7)				
ueilal446 (5.0)136 (4.0)97 (4.12)94 (3.9)matoid $259 (7.3)$ $156 (4.3)$ $97 (4.12)$ $94 (3.9)$ matoid $95 (7.3)$ $196 (7.3)$ $136 (7.1)$ $100 (7.6)$ therwise specified $557 (2.6)$ $136 (5.7)$ $137 (5.2)$ $1108 (6.7)$ thy 0.6770 $57 (3.2)$ $103 (5.2)$ $136 (5.7)$ $137 (5.2)$ $1108 (6.7)$ thy 0.6770 $220 (97.7)$ $220 (97.7)$ $231 (9.7)$ $232 (9.7)$ $232 (9.7)$ iny 0.6770 $220 (9.7)$ $220 (9.7)$ $231 (9.7)$ $232 (9.7)$ $232 (9.7)$ iny 0.6770 $220 (9.7)$ $223 (4.9)$ $232 (4.9)$ $232 (9.7)$ $232 (9.7)$ iny 0.6770 $223 (4.9)$ $223 (4.9)$ $232 (4.9)$ $232 (4.9)$ iny 0.1360 $1187 (9.2)$ $223 (4.9)$ $232 (4.9)$ $232 (4.9)$ iny $0.131 (4.7)$ $0.137 (4.2)$ $232 (4.9)$ $232 (4.9)$ $232 (4.9)$ iny $0.131 (4.1)$ $223 (4.9)$ $223 (4.9)$ $232 (9.7)$ $232 (9.7)$ iny $0.131 (4.1)$ $223 (4.9)$ $223 (4.9)$ $232 (9.7)$ $232 (9.7)$ iny $0.131 (4.1)$ $223 (4.9)$ $223 (4.9)$ $232 (9.7)$ $232 (9.7)$ iny $0.131 (4.1)$ $223 (4.9)$ $223 (4.9)$ $232 (4.9)$ $232 (4.9)$ iny $0.131 (4.1)$ $233 (4.9)$ $0.131 (4.1)$ $0.127 (4.9)$ iny $0.127 (4.9)$ $0.127 (4.9)$ $0.127 (4.9)$ $0.127 (4.9)$	Histology			<0.0001			0.3631
	Epithelial	4486 (35.0)	1396 (40.3)		979 (41.2)	948 (39.9)	
asic model of the form of the	Sarcomatoid	1766 (13.8)	259 (7.5)		169 (7.1)	180 (7.6)	
	Biphasic	996 (7.8)	194 (5.6)		153 (6.4)	138 (5.8)	
ity involvement/midline ity 0.6770 if $230 (97.7)$ $2316 (97.6)$ side $230 (97.7)$ $2316 (97.6)$ side $230 (97.7)$ $2316 (97.6)$ $2324 (95.6)$ $2324 (95.$	Not otherwise specified	5552 (43.4)	1618 (46.7)		1073 (45.2)	1108 (46.7)	
side the form of	Laterality			0.6770			0.6976
eral involvement/midline $320 (2.5)$ $81 (2.3)$ $54 (2.3)$ $58 (2.4)$ ing/unknown $623 (4.9)$ $223 (6.4)$ 0.1360 $1031 (43.4)$ $1022 (43.0)$ ing/unknown $5564 (44.3)$ $1475 (42.5)$ 0.1360 $1031 (43.4)$ $1022 (43.0)$ ind $2048 (15.0)$ $579 (15.7)$ $862 (2.4.9)$ $357 (15.9)$ $357 (15.9)$ int $1022 (15.0)$ $551 (15.9)$ $552 (2.4.9)$ $357 (15.0)$ int $1022 (15.0)$ $551 (15.9)$ $359 (15.1)$ $357 (15.0)$ int $1022 (15.0)$ $551 (15.9)$ $359 (15.1)$ $357 (15.0)$ int $1022 (15.0)$ $521 (15.9)$ $352 (15.9)$ $359 (15.1)$ $357 (15.0)$ int/unknown $316 (24.7)$ $862 (24.9)$ 0.2129 $1970 (83.0)$ $1977 (83.3)$ int/unknown $37 (0.3)$ 210^{0} 0.2119 $1070 (82.0)$ $1977 (83.3)$ int/unknown $324 (32.5)$ $236 (18.5)$ $238 (18.5)$ $238 (16.5)$ $238 (16.7)$ int/unknown $324 (32.5)$ $1204 (94.1)$ $324 (32.5)$ $97 (17.0)$ $97 (15.7)$ int/unknown $324 (25.9)$ $92 (2.7)$ 0.0004 $1777 (53.8)$ $1780 (53.9)$ int/unknown $393 (15.1)$ $1097 (46.2)$ $1097 (46.1)$ $1097 (46.1)$ int/unknown $393 (21.1)$ $1097 (46.2)$ $1094 (46.1)$ int/unknown $393 (21.1)$ $1097 (46.2)$ $1094 (46.1)$	One side	11 857 (92.6)	3163 (91.2)		2320 (97.7)	2316 (97.6)	
int/unknown $623 (4.9)$ $223 (6.4)$ $223 (6.4)$ 11 $5664 (4.3)$ $1475 (42.5)$ 0.1360 11 $579 (16.7)$ $902 (16.9)$ $406 (17.1)$ 11 $2048 (16.0)$ $579 (16.7)$ $359 (15.1)$ $357 (15.0)$ 11 $1923 (15.0)$ $551 (15.9)$ $402 (16.9)$ $406 (17.1)$ 11 $1923 (15.0)$ $551 (15.9)$ $552 (24.5)$ $589 (24.8)$ 1 $10427 (81.5)$ $882 (24.9)$ $522 (24.5)$ $589 (24.8)$ $97 (100)$ $3165 (24.7)$ $861 (82.5)$ $197 (83.0)$ $197 (83.3)$ $97 (100)$ $316 (12.1)$ $327 (15.2)$ $329 (15.1)$ $329 (15.7)$ $329 (15.7)$ $97 (100)$ $316 (12.1)$ $327 (12.2)$ $329 (13.1)$ $329 (15.7)$ $329 (15.7)$ $97 (100)$ $107 (12.0)$ $127 (12.0)$ $197 (12.0)$ $197 (12.0)$ $197 (12.0)$ 100 unknown $32 (12.7)$ $32 (13.2)$ $120 (12.0)$ $127 (12.0)$ $197 (12.0)$ 110 uhnnown $32 (12.7)$ $92 (13.7)$ $90 (13.8)$ $100^7 (16.2)$ $109 (16.2)$ 110 uhnnown $32 (12.5)$ $92 (12.1)$ $92 (12.1)$ $90 (13.8)$ 110 ununu $120 (100)$ $127 (100)$ $127 (100)$ $127 (100)$ 1100 unu $127 (100)$ $127 (100)$ $1007 (16.2)$ $1007 (16.2)$ 1100 unu $127 (12.0)$ $1207 (12.0)$ $1007 (127 (10))$ $1007 (16.2)$ 1100 unu $120 (12.0)$ $1207 (12.0)$ $127 (12.0)$ 127	Bilateral involvement/midline	320 (2.5)	81 (2.3)		54 (2.3)	58 (2.4)	
$ \begin{array}{cccccccccccccccccccccccccccccccccccc$	Missing/unknown	623 (4.9)	223 (6.4)				
	Stage			0.1360			0.7998
al (17.1) and (17.1) (15.2) (16.7) (16.6) (17.1) (10.2) (15.1) (Local	5664 (44.3)	1475 (42.5)		1031 (43.4)	1022 (43.0)	
tt (15) (15) (15) (15) (15) (15) (15) (15)	Regional	2048 (16.0)	579 (16.7)		402 (16.9)	406 (17.1)	
	Distant	1923 (15.0)	551 (15.9)		359 (15.1)	357 (15.0)	
$\begin{array}{cccccccccccccccccccccccccccccccccccc$	Missing/unknown	3165 (24.7)	862 (24.9)		582 (24.5)	589 (24.8)	
$\begin{array}{cccccccccccccccccccccccccccccccccccc$	Surgery			0.2129			0.7835
$\begin{array}{cccccccccccccccccccccccccccccccccccc$	No	10 427 (81.5)	2861 (82.5)		1970 (83.0)	1977 (83.3)	
Jown $37 (0.3)$ $< 10^{\circ}$ $< 10^{\circ}$ 12 044 (94.1) 3243 (93.5) 0.3115 12 044 (94.1) 3243 (93.5) $2287 (96.3)$ 2287 (95.3) $324 (3.7)$ $90 (3.8)$ nown $324 (2.5)$ $95 (2.7)$ $90 (3.8)$ 0.0004 $1.767 (51.0)$ $1.007 (46.2)$ $1.780 (53.9)$ 0.0004 $1.097 (46.2)$ $1.097 (46.1)$ $1.097 (46.1)$	Yes	2336 (18.3)	≥10 ^b		404 (17.0)	397 (16.7)	
0.3115 12 044 (94.1) 3243 (93.5) 0.3115 12 044 (94.1) 3243 (93.5) 2287 (96.3) 2284 (96.2) 432 (3.7) 90 (3.8) 324 (2.5) 95 (2.7) 0.0004 6098 (47.6) 1767 (51.0) 1277 (53.8) 1780 (53.9) 6309 (49.3) 1592 (45.9) 1097 (46.2) 1094 (46.1) 108 (3.1) 108 (3.1)	Missing/unknown	37 (0.3)	<10 ^b				
$\begin{array}{cccccccccccccccccccccccccccccccccccc$	Radiotherapy			0.3115			0.8175
432 (3.4) 129 (3.7) 87 (3.7) 90 (3.8) nown 324 (2.5) 95 (2.7) 90 (3.8) 129 (3.7) 95 (2.7) 90 (3.8) 120 (3.1) 95 (2.7) 90 (3.8) 121 (3.1) 90 (3.1) 1277 (53.8) 1780 (53.9) 122 (45.9) 1592 (45.9) 1097 (46.2) 1094 (46.1) 1000 12.1 1097 (46.2) 1094 (46.1)	No	12 044 (94.1)	3243 (93.5)		2287 (96.3)	2284 (96.2)	
nown 324 (2.5) 95 (2.7) 0.0004 6098 (47.6) 1767 (51.0) 1277 (53.8) 1780 (53.9) 6309 (49.3) 1592 (45.9) 1097 (46.2) 1094 (46.1) 1004 108 (3.1) 108 (3.1) 108 (3.1)	Yes	432 (3.4)	129 (3.7)		87 (3.7)	90 (3.8)	
0.0004 6098 (47.6) 1767 (51.0) 1277 (53.8) 1780 (53.9) 6309 (49.3) 1592 (45.9) 1097 (46.2) 1094 (46.1) 333 (3.1) 108 (3.1)	Missing/unknown	324 (2.5)	95 (2.7)				
6098 (47.6) 1767 (51.0) 1277 (53.8) 6309 (49.3) 1592 (45.9) 1097 (46.2) 393 (3.1) 108 (3.1)	Chemotherapy			0.0004			0.9269
6309 (49.3) 1592 (45.9) 1097 (46.2) 393 (3.1) 108 (3.1)	No	6098 (47.6)	1767 (51.0)		1277 (53.8)	1780 (53.9)	
393 (3.1)	Yes	6309 (49.3)	1592 (45.9)		1097 (46.2)	1094 (46.1)	
	Missing/unknown	393 (3.1)	108 (3.1)				

HS, high school. ³x² test among non-missing values. ^bCells masked to protect against identification of patients.

Table 1. Continued

Table 2. Demographic and clinical factors associated with gender (n = 12759)

Variable	Adjusted OR ^a (female versus male)	95% CI
Age (years)		
<50	1.00	Ref
50–59	0.72	0.54-0.97
60–69	0.43	0.32-0.58
70–79	0.43	0.32-0.58
≥80	0.49	0.36-0.66
Race		
White	1.00	Ref
Black	1.00	0.80-1.24
Asian/Other	1.15	0.87-1.53
Spanish/Hispanic origin	1.19	0.07 1.55
Yes versus No	0.99	0.79–1.2
% of adults in patient zip code without a HS degree (quartiles)	0.55	0.7 5-1.2
<7	1.00	Ref
7–12.9		
	0.86	0.76-0.97
13–20.9	0.96	0.83-1.11
≥21 Σ	0.96	0.80-1.17
Patient zip code median income (\$) (quartiles)		
≥63 000	1.00	Ref
48 000–62 999	0.97	0.86-1.09
38 000–47 999	0.90	0.78-1.04
<38 000	0.93	0.78–1.12
Insurance		
Medicare/other government	1.00	Ref
Medicaid	1.42	1.08–1.89
Private insurance	1.11	0.98–1.25
Uninsured	0.98	0.71-1.36
Charlson Comorbidity Index		
0	1.00	Ref
1	0.89	0.80-0.99
≥2	0.76	0.65–0.90
Facility type		
Community Cancer Program	1.00	Ref
Academic/Integrated Network Cancer Program	1.04	0.95–1.14
Distance from patient to reporting facility (miles)		
<25	1.00	Ref
25-49.9	0.96	0.84-1.11
≥50	0.89	0.77-1.03
Histology	0.05	0.77 1.05
Epithelial	1.00	Ref
Sarcomatoid	0.50	0.42-0.59
Biphasic	0.68	0.57-0.82
Not otherwise specified	0.93	0.84-1.02
	0.95	0.04-1.02
Laterality Bilateral involvement/midline versus one side	0.95	0 CF 1 12
	0.85	0.65–1.13
Stage		7
Local	1.00	Ref
Regional	1.08	0.96–1.23
Distant	1.07	0.93–1.21
Missing/unknown	1.05	0.94–1.18
Surgery		
Yes versus No	0.91	0.81-1.04
Radiotherapy		
Yes versus No	1.07	0.85–1.35
Chemotherapy		
Yes versus No	0.78	0.71-0.86

HS, high school; OR, odds ratio.

^aOdds of being female, adjusted for all other variables, year of diagnosis, and reporting facility census division.

Bold values indicate statistically significant results.

OS was significantly better for females than males at 2 years (26.5 versus 16.6%), and 5 years (9.4 versus 4.2%) which remained significant after adjusting for all covariates $[HR_{adj}; 0.81,$

95% confidence interval (CI): 0.77–0.85] (Table 3). Besides gender, other factors significantly associated with better survival were younger age at diagnosis, higher income at the zip-code level,

lower comorbidity score, diagnosis in an academic/integrated network cancer program, epithelial histology, earlier stage and receipt of treatment (Supplementary Table 1, available at *Carcinogenesis* Online). OS were better in females compared with males [hazard ratio (HR): 0.86, 95% CI: 0.80–0.94] (Table 3, Figure 2A).

After stratification by age at diagnosis, propensity matching within the <50 years age group yielded 70 matched female and male cases. The 2-year OS was 43.5 and 33.9% (P = 0.0772), and the 5-year OS was 28.8 and 16.7% (P = 0.0642) for females

Propensity matching yielded 2374 matched cases for each gender (Table 1). Both the 2- (P = 0.0019) and 5-year (P = 0.0005)

	Univariate a	analysis					Multivariate ana (n = 11 388)	lysis
	2-year survival			5-year survival				
	Rate (%)	95% CI	Р	Rate (%)	95% CI	Р	Adjusted HR ^a	95% CI
Males	16.6	15.9–17.3		4.2	3.7–4.6		1	Ref
Females	26.5	24.9-28.1	<0.0001	9.4	8.2-10.5	< 0.0001	0.81	0.77–0.85
	Propensity-matched analysis ($n = 4,748$) ^b							
Males	17.5	15.9–19.0		4.3	3.3–5.3		1	Ref
Females	26.0	24.1-27.8	0.0019	8.7	7.3-10.0	0.0005	0.86	0.80-0.94

Table 3. Univariate and multivariate overall survival according to gender

^aAdjusted for age at diagnosis, race, Hispanic origin, % of adults in the zip code with no high school, ZIP code median income, insurance, Charlson Comorbidity Index, facility type, distance from the reporting facility, histology, laterality, stage, receipt of surgery, receipt of radiotherapy, receipt of chemotherapy, year of diagnosis and US census division of the reporting facility.

^bP-values and HRs obtained from stratified Cox proportional hazards models to account for matching.

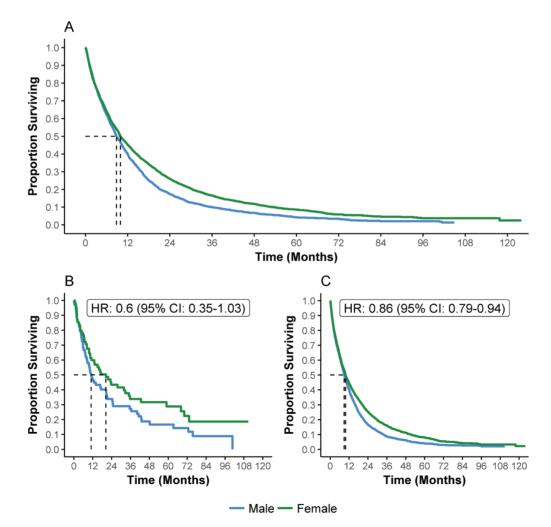


Figure 2. Overall survival according to gender in (A) propensity-matched cohort (n = 4748); (B) <50 years (n = 140); (C) ≥50 years (n = 4580).

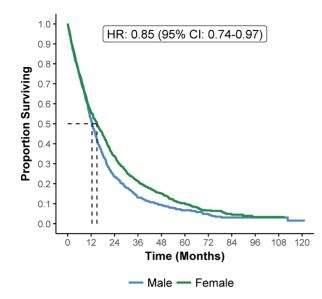


Figure 3. Overall survival in epithelia histology cases (propensity-matched cohort) according to gender (*n* = 1890).

and males, respectively, with an HR of 0.6 (95% CI: 0.35–1.03) (Figure 2B).

Propensity matching within the \geq 50 years age group yielded 2290 matched pairs. The 2-year OS was 25.3 and 16.7% (*P* = 0.0003) and the 5-year OS was 8.1 and 4.1% (*P* = 0.0007), for females and males, respectively, with an HR of 0.86 (95% CI: 0.79–0.94) (Figure 2C).

Propensity matching in the group with epithelial histology yielded two balanced groups of 945 female and male cases, and confirmed a better OS in females versus males at 2 years (P = 0.0128) and 5 years (P = 0.0152); HR: 0.85 (95% CI: 0.74–0.97) (Figure 3). Within the epithelial histology group females showed significantly better survival than males in both <50 years and \geq 50 years groups (data not shown).

Discussion

This study includes more than 16 000 patients with MPM from the National Cancer Database and is one of the largest endeavors comparing survival in males and females. Our analysis shows that females with MPM survive significantly longer than males, independently from other contributing factors.

Differences in seeking medical advice have been suggested as a potential reason for the difference in survival with gender, because women tend to consult a physician earlier in the course of a disease than men (11). The current study, however, showed that stage at diagnosis was not significantly different in the two genders, as shown by the same rate of metastatic disease at diagnosis. This makes lead-time bias or earlier stage of disease at diagnosis a less likely explanation of the observed gender differences in survival.

Another factor that may affect survival is receipt of treatment. Our study showed that males and females equally received surgery or radiotherapy, and, in contrast to the survival outcome, females received chemotherapy less frequently compared with males, making treatment a less likely explanation for the difference in survival.

Recently published literature suggested that female malignant mesotheliomas have a higher frequency of germline mutations in DNA repair genes, some of which are associated with better survival, and this could in part account for the gender differences in survival observed in epidemiologic studies (16,17). Although other studies specifically focused on BAP1 polymorphism did not indicate that the frequency of the germline mutation varies with gender (18) or is lower in females than in males (19), it is possible that BAP1 is a mediator of the association between gender and survival. The dataset we used in this analysis does not include biological information or genetic testing, thus we are unable to test this hypothesis. More research on MPM germline variant in DNA repair genes and their association with gender and survival is needed.

Previous studies already identified the association of older age and male gender with shorter survival (5,8). A SEER analysis reported a better survival rate in females compared with males, however this study did not address the role of histology because this variable was missing in almost two-thirds of the cases (5). A smaller, single-institution study producing similar results on survival was also unable to fully adjust for all different histologies because of the small sample (8). Despite the fact that histology was not specified for all patients in the NCDB, we were able to study the effect of gender within the epithelial group, and confirmed the survival advantage of females.

The stratified analysis according to age found that females had higher survival at all ages, although the difference tended to be more pronounced in younger patients. These results seem to be unaffected by the tumor histology, despite the smaller sample size of this subanalysis. Circulating estrogen, present in young but not older women, and the expression of the estrogen receptor beta (ER β) have been suggested to play a role in the survival difference between males and females. It has been shown that the ER β has antiproliferative effects in in vitro and in vivo cancer models (20). A study analyzing 78 samples of patients with MPM showed a significantly better survival in patients with high $ER\beta$ expression compared with patients with low $ER\beta$ expression (11). A follow-up study showed that reduced $ER\beta$ expression led to a phenotypic shift to a less epithelioid phenotype and vice versa. Furthermore, $ER\beta$ expression was lost in more aggressive sarcomatoid forms of MPM (21). Therefore the potential to reverse a more aggressive biphasic phenotype by inducing the $ER\beta$ expression, as shown by Manente et al. (22), could be clinically significant. Epithelial forms of MPM have a better prognosis compared with other histology types, and the interplay with $ER\beta$ expression may be the underlying reason for better survival. Our finding that younger women have a more pronounced survival advantage further supports the theory that the presence of circulating estrogens combined with higher $ER\beta$ expression in epithelial histology may be partly responsible for the survival difference between genders. These results suggest the possibility of reversing a more aggressive histotype by targeting $ER\beta$ with a selective agonist together with estrogens as a novel treatment option, as described for prostate cancer (23), thus increasing therapeutic options for patients with MPM.

Among the limitations of this analysis is the fact that differences in survival between males and females for other histology types, such as sarcomatoid and biphasic, could not be assessed because of the small number of female patients with these MPM subtypes. Although the NCDB captures ~70% of newly diagnosed cancer cases in the USA, it is not a nationally representative population-based sample. However, cases reported to the NCDB are most likely representative at regional level (24). Another limitation is that the NCDB is used by communities and participating hospitals as a self-assessment tool to assess patterns of care and outcomes (25). The dataset might therefore lack relevant information because it was not collected. For example, a proportion of cases lacked complete information on histology; the lack of difference between males and females in the percentage of histology 'not otherwise specified' makes it unlikely that bias could explain the observed survival differences. Strengths related to using a large database include the large group of eligible patients with MPM, which is an otherwise rare disease. Furthermore, the results are more generalizable to the general population compared with singleinstitution studies.

This study adds information to the known significant survival difference between male and female patients with MPM using a large cohort of patients with MPM, by disentangling the effect of gender from the effect of age and histology, two known independent factors affecting survival. These results may provide additional evidence that $ER\beta$ expression present on epithelial-type MPM and circulating estrogens, present in younger women, positively affect survival in MPM.

Supplementary material

Supplementary data are available at Carcinogenesis online.

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References

- 1. Yang, H. et al. (2008) Mesothelioma epidemiology, carcinogenesis, and pathogenesis. Curr. Treat. Options Oncol., 9, 147–157.
- Flores, R.M. et al. (2010) Frequency of use and predictors of cancer-directed surgery in the management of malignant pleural mesothelioma in a community-based (Surveillance, Epidemiology, and End Results [SEER]) population. J. Thorac. Oncol., 5, 1649–1654.
- Spirtas, R. et al. (1988) Survival patterns for malignant mesothelioma: the SEER experience. Int. J. Cancer., 41, 525–530.
- 4. Milano, M.T. et al. (2010) Malignant pleural mesothelioma: a populationbased study of survival. J. Thorac. Oncol., 5, 1841–1848.
- Taioli, E. et al. (2014) Women with malignant pleural mesothelioma have a threefold better survival rate than men. Ann. Thorac. Surg., 98, 1020–1024.
- Wolf, A.S. et al. (2010) Characteristics of malignant pleural mesothelioma in women. Ann. Thorac. Surg., 90, 949–56; discussion 956.

- Flores, R.M. et al. (2007) Prognostic factors in the treatment of malignant pleural mesothelioma at a large tertiary referral center. J. Thorac. Oncol., 2, 957–965.
- Rusch, V.W. et al. (1999) Important prognostic factors in patients with malignant pleural mesothelioma, managed surgically. Ann. Thorac. Surg., 68, 1799–1804.
- 9. Yan, T.D. et al. (2006) Sex difference in diffuse malignant peritoneal mesothelioma. Br. J. Surg., 93, 1536–1542.
- Hillerdal, G. (1999) Mesothelioma: cases associated with nonoccupational and low dose exposures. Occup. Environ. Med., 56, 505–513.
- 11. Pinton, G. et al. (2009) Estrogen receptor-beta affects the prognosis of human malignant mesothelioma. Cancer Res., 69, 4598–4604.
- 12. Bilimoria, K.Y. et al. (2009) Comparison of commission on cancerapproved and -nonapproved hospitals in the United States: implications for studies that use the National Cancer Data Base. J. Clin. Oncol., 27, 4177–4181.
- 13. Boffa, D.J. et al. (2017) Using the National Cancer Database for outcomes research: a review. JAMA Oncol., 3, 1722–1728.
- 14. Steele, G.D. Jr *et al.* (1994) The National Cancer Data Base. a mechanism for assessment of patient care. Cancer, 73, 499–504.
- Ohri N, et al. (2016) Definitive radiation therapy is associated with improved survival in non-metastatic malignant pleural mesothelioma. Int. J. Radiat. Oncol. Biol. Phys., 96, S132–S133.
- Hassan R, et al. (2018) Inherited predisposition to malignant mesothelioma (MM) due to mutations in DNA repair genes. J. Clin. Oncol., 15, 8504–8504.
- Panou, V. et al. (2018) Frequency of germline mutations in cancer susceptibility genes in malignant mesothelioma. J. Clin. Oncol., 36, 2863–2871.
- Pastorino S, et al. (2018) Subset of mesotheliomas with improved survival occurring in carriers of BAP1 and other germline mutations. J. Clin. Oncol., 36, 3485–3494.
- De Rienzo, A. et al. (2016) Gender-specific molecular and clinical features underlie malignant pleural mesothelioma. Cancer Res., 76, 319–328.
- 20. Pinton, G. et al. (2017) Expression and the rapeutic significance of estrogen receptor β in malignant pleural mesothelioma. Future Sci. OA., 3, FSO175.
- 21. Pinton, G. et al. (2010) Estrogen receptor β exerts tumor repressive functions in human malignant pleural mesothelioma via EGFR inactivation and affects response to gefitinib. PLoS One, 5, e14110.
- Manente, A.G. et al. (2015) Intracellular lactate-mediated induction of estrogen receptor beta (ERβ) in biphasic malignant pleural mesothelioma cells. Oncotarget, 6, 25121–25134.
- Ho, S.M. et al. (2011) Estrogens and prostate cancer: etiology, mediators, prevention, and management. Endocrinol. Metab. Clin. North Am., 40, 591–614, ix.
- 24. Fremgen, A.M. et al. (1999) Clinical highlights from the National Cancer Data Base, 1999. CA. Cancer J. Clin., 49, 145–158.
- 25. Menck, H.R. et al. (1997) The growth and maturation of the National Cancer Data Base. Cancer, 80, 2296–2304.