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## Duration of opioid prescriptions predicts incident nonmedical use of prescription opioids among U.S. veterans receiving medical care

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### Abstract

**Background/Aims:** Although nonmedical use of prescription opioids (NMUPO) is a public health problem, few studies have examined the new-onset NMUPO in clinical populations. We estimated NMUPO incidence among veterans in medical care who had received prescription opioid medication and examined correlates of new-onset NMUPO.

**Design:** Prospective cohort study.

**Setting:** Veterans Health Administration primary care and infectious disease clinics in Atlanta, Baltimore, Bronx, Houston, Los Angeles, Manhattan, Pittsburgh, and Washington, DC.

**Participants:** Patients enrolled in the Veterans Aging Cohort Study wave 3 (2005–2007) who received prescription opioids in the previous year and without lifetime NMUPO were followed at waves 4 and 5 (2008–2011).

**Measurements:** Cox proportional hazards regression was used to examine the relationship between duration of prescription opioid receipt and incident NMUPO, adjusting for demographics,

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Contributors

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Conflict of Interest

No conflict declared.

alcohol and tobacco use, substance use disorders, psychiatric and medical diagnoses, and medication-related characteristics.

**Findings:** Among eligible participants (n = 815), the median age was 52 (IQR = 47–58) and 498 (59.8%) were Black; 122 (15.0%) reported new-onset NMUPO, for an incidence rate of 5.0 per 100 person-years. In a multivariable Cox model, compared to < 30 days, receipt of prescription opioids for 30–180 days (adjusted hazard ratio [AHR] = 1.65 95% CI: 1.06, 2.58) or > 180 days (AHR = 1.99, 95% CI: 1.21, 3.29) was associated with incident NMUPO.

**Conclusions:** Duration of prescription opioid receipt is a risk factor for incident NMUPO among veterans receiving medical care. Providers who prescribe opioids should monitor for NMUPO, especially among those with a longer duration of opioid therapy.

## Keywords

Opioids; Analgesics; Pain; HIV; Veterans

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## 1. Introduction

Considerable debate surrounds the prescribing of opioid analgesics internationally for non-cancer chronic pain (Häuser et al., 2017; Katz, 2016; Novak et al., 2016; Okie, 2010). There is limited evidence supporting the efficacy of prescription opioids in managing chronic pain (Chou et al., 2015). Concerns have also emerged about the safety of long-term opioid therapy (i.e., greater than 90 days) because it may increase the risk of overdose death (Dunn et al., 2010), all-cause mortality (Ray et al., 2016), and nonmedical use of prescription opioids (NMUPO) (Becker et al., 2008). NMUPO is often defined as taking someone else's opioid medication or taking the medication only for the experience it causes (SAMSHA, 2011), and is associated with the initiation and use of heroin (Banerjee et al., 2016; Compton et al., 2016) and psychiatric, medical, and non-opioid substance use problems (Becker et al., 2008; Campbell et al., 2018; Katz et al., 2013). NMUPO is also linked to pain complaints among untreated individuals with opioid use disorder (Barry et al., 2009, 2013). NMUPO comprises a challenge for clinicians in different settings, including office-based physicians and HIV providers (Barry et al., 2010; Keller et al., 2012; Lum et al., 2011; Starrels et al., 2016).

Veterans comprise a high-risk group for pain (Institute of Medicine, 2011). Among veterans in primary care, pain is associated with both receipt of opioid medication and NMUPO (Becker et al., 2009). Similar to other healthcare systems, rates of opioid prescribing escalated at the Veterans Health Administration (VHA) in the 1990s (Kuehn, 2007). The Department of Veterans Affairs/Department of Defense, Centers for Disease Control and Prevention (CDC), and some professional organizations have recently issued guidelines to promote appropriate prescribing practices for pain management (Chou et al., 2017; Department of Veterans Affairs Department of Defense, 2010; Dowell et al., 2016; Manchikanti et al., 2012).

To date, most research on NMUPO has focused on prevalence (using cross-sectional designs). Some longitudinal studies have examined NMUPO trajectories among middle and

high school students (McCabe et al., 2016, 2014). One study of claims data examined incidence or new-onset of opioid use disorder among individuals prescribed opioid analgesics (Edlund et al., 2014). However, no studies to our knowledge have examined NMUPO incidence among individuals receiving prescription opioids. Estimating incident NMUPO in patients prescribed opioids require a relatively large sample that is systematically assessed for NMUPO over time, as well as access to detailed pharmacy information on prescription opioids. The Veterans Aging Cohort Study (VACS) meets these requirements (Justice et al., 2006). Previous studies involving the VACS have found that one-third of participants had been prescribed opioids and, of these individuals, more than one-third received opioids long-term (Edelman et al., 2013). Thirteen percent of all VACS wave-3 participants reported lifetime NMUPO (Barry et al., 2011). Substance use, medical status, and pain interference (but not HIV status) in this cross-sectional investigation were independent correlates of prevalent NMUPO (Barry et al., 2011).

The current study aimed to estimate the incidence of NMUPO among veterans with and without HIV who were prescribed opioids and to examine demographic, substance use and substance use disorder, psychiatric, medical, and medication predictors of new-onset NMUPO. Given that incident opioid use disorder risk among patients with chronic pain has been associated with longer durations of opioid use and with higher average daily opioid doses prescribed (Chou et al., 2015; Edlund et al., 2014), we hypothesized that incident NMUPO would be associated with these two opioid-medication characteristics. An enhanced understanding of the incidence of NMUPO and its associated risk factors among veterans with and without HIV may inform ongoing initiatives at the VHA and elsewhere to optimize opioid therapy benefits and minimize risks (Lin et al., 2017; Oliva et al., 2017).

## 2. Methods

### 2.1. Data sources

The Veterans Aging Cohort Study (VACS) (Justice et al., 2006) is a National Institute on Alcohol Abuse and Alcoholism (NIAAA)-funded prospective, longitudinal, multisite observational study of patients with and without HIV-infection receiving care in VHA primary care and infectious disease clinics. Uninfected participants were matched to HIV-infected ones by age, race, and site of care (Justice et al., 2006). Data for the current study were drawn from waves 3–5 of VACS follow-up surveys (hereafter referred to as waves 3–5). Since June 2002, VACS has enrolled over 7000 patients in Atlanta, Baltimore, Bronx, Houston, Los Angeles, Manhattan, Pittsburgh, and Washington, D.C. Wave 3–5 data were collected September 2005–January 2007, February 2008–August 2009, and September 2009–January 2011, and included 4133, 4182, and 3762 participants respectively. Survey data were linked to data on prescribed medications, medical and substance use diagnoses, and laboratory findings from VHA electronic medical records (EMR). Opioid prescription data were retrieved from the VHA Pharmacy Benefits Management records system. Other than NMUPO, all other data from self-report measures were collected at wave 3 (2005–2007). The VACS was approved by the institutional review boards at participating VHA facilities and affiliated academic institutions. More detailed information concerning the VACS,

including the design and data collection procedures, are provided elsewhere (Justice et al., 2006).

## 2.2. Participant eligibility

The flowchart summarizing the selection of eligible participants for this study is illustrated in Fig. 1. Among the 4133 VACS participants at wave 3, a total of 1308 received at least one prescription opioid in the previous year (based on pharmacy data). Among those who received a prescription opioid, 486 (37.1%) reported lifetime NMUPO at wave 3 (see definition below) and were thus excluded. Among those who received a prescription opioid who reported no lifetime NMUPO ( $n = 822$ ), seven (0.9%) participants provided invalid responses and were excluded, resulting in a final sample of 815 participants for analyses involving NMUPO incidence (at waves 4–5).

## 2.3. Measures

**2.3.1. Nonmedical use of prescription opioids—**NMUPO was assessed by two self-report items at waves 3–5. The first question was derived verbatim from the National Survey on Drug Use and Health (NSDUH) (Substance Abuse and Mental Health Services Administration, 2007): “Have you ever, even once, used one of the medications listed below that was NOT prescribed for you or that you took only for the experience or feeling it caused?” For each medication listed, respondents checked whether or not they had “ever used” and “used in the past 12-months.” The list included the following prescription analgesics: buprenorphine, codeine, Darvocet, Darvon, Demerol, Dilaudid, Fioricet, Fiorinal, hydrocodone, methadone, morphine, Oxycontin, Percocet, Percodan, propoxyphene, Talwin, Tylenol with codeine, Tylox, Ultram, and Vicodin. As done previously, we excluded respondents whose only nonmedical use involved Fioricet and/or Fiorinal because these medications are not opioids (Becker et al., 2008). The list of analgesics presented to participants at waves 4 and 5 also included Fentanyl. The second question that assessed NMUPO at waves 3–5 was: “Now think about the past 12-months. On average, how many days each week in the past 12-months did you use any prescription pain reliever that was not prescribed for you or that you took only for the experience or feeling it caused?” Participants who responded in the affirmative to either the first or second question were considered to exhibit NMUPO.

**2.3.2. Demographics, substance use and substance use disorder, and psychiatric status—**Participants provided information about their age, sex, and race/ethnicity. Alcohol use was assessed using the 3-item Alcohol Use Disorder Identification Test (AUDIT-C) (Bush et al., 1998). The following questions assessed smoking status (never, current, former): “Have you smoked at least 100 cigarettes in your entire life?” (yes/no) and “Do you smoke cigarettes (as of 1-month ago)?” (yes/no). Diagnostic information related to alcohol use disorder, drug use disorder, and specific psychiatric disorders (major depression, bipolar disorder, schizophrenia, and post-traumatic stress disorder [PTSD]) were collected from the VHA EMR using the International Classification of Diseases, 9th Revision (ICD-9) codes (World Health Organization, 1975). Level of depression severity was assessed using the Patient Health Questionnaire (PHQ-9) (Spitzer et al., 1999).

**2.3.3. Medical status and pain status**—HIV data were collected from the VHA EMR system using ICD-9 codes (World Health Organization, 1975). Hepatitis C status was determined by the presence of an ICD-9 code or a positive HCV lab test. Pain intensity was assessed by the following 0–10 item: “On a scale of 0 to 10, where 0 means no pain, and 10 equals the worst possible pain, what is your current pain level?” Pain interference was measured by the following 0 (“no interference”)-6 (“extreme interference”) scored item: “In general, how much does your pain problem interfere with your day to day activities?”

**2.3.4. Medication characteristics**—Prescription opioid receipt was based on pharmacy data for all outpatient oral and transdermal opioids (Edelman et al., 2013). Medications used as part of opioid agonist treatment for opioid use disorder were not included. Days of opioid receipt were calculated from prescription information with the assumption that the prescription was taken as directed. Based on pharmacy data in the year preceding the wave 3 survey date, days of opioids supplied were analyzed both continuously and also categorized as fewer than 30, 30–180, and greater than 180. (This categorization was constructed in part based on the distribution of the data.) Morphine equivalent dose (MED) was calculated using a standard conversion chart, as described elsewhere (Von Korff et al., 2008). Total MED was calculated by multiplying the strength of the prescription (milligram of opioid per unit dispensed) by the prescription quantity. Average daily MED was calculated by dividing the total MED in the year by days supplied. Mean daily MED was categorized as follows: 0–49, 50–79, 80–119, 120 mg. Fifty milligrams were used as a threshold for high daily MED since the CDC opioid prescribing guideline recommends a reassessment of individual risks and benefits when this dosage is reached (Dowell et al., 2016). We also created a composite days supply/high-low dose variable, which combined data on both average daily MED (  $\leq 50$  is low,  $> 50$  is high) and days of opioids supplied (using the same classification as described above) to create mutually exclusive groups. Schedule II opioids were categorized as short-acting or long-acting per DEA classification (US Drug Enforcement Administration Controlled Substance Schedules, 2010).

Data about receipt and days supplied of benzodiazepines and non-benzodiazepine hypnotics were abstracted from the pharmacy data system. Benzodiazepines included alprazolam, chlordiazepoxide, clonazepam, estazolam, flurazepam, lorazepam, oxazepam, temazepam, and triazolam, whereas non-benzodiazepine hypnotics included eszopiclone, zaleplon, and zolpidem. Any benzodiazepine/non-benzodiazepine hypnotic receipt was defined as receipt of at least one prescription for an outpatient benzodiazepine/non-benzodiazepine hypnotic in the year preceding wave 3. Long-term receipt of benzodiazepines/non-benzodiazepine hypnotics was defined as at least 90 consecutive days of therapy, allowing for a 30-day refill window (Von Korff et al., 2008) in the year prior to the wave 3 survey date.

## 2.4. Data analysis

As a first step, we estimated the bivariable associations between each independent variable of interest (collected at wave 3) and incident NMUPO at waves 4 or 5 using chi-square tests, *t*-tests, and Wilcoxon rank sum tests for non-normally distributed data. We calculated the incidence of NMUPO over the study period using the standard Kaplan Meier method. For incident cases, we used the survey date during which the first instance of NMUPO was

reported as the event time. All other individuals were right-censored at the date of their last survey completed or the end of the study period. Next, the independent variables of interest—days supplied (as a categorical variable) and high average daily dose—were entered into a Cox proportion hazards regression model, along with other potentially confounding covariates from the bivariate analyses whose associations were significant at  $p < 0.20$ . We adjusted for demographics, alcohol and tobacco use, substance use disorders, psychiatric and medical diagnoses, and medication-related characteristics. We assessed the assumption of proportional hazards by visual inspection of the Schoenfeld residual plots and by examining time-by-covariate interactions.

In a post hoc analysis, we used spline regression to examine further the relationship between days of opioid medication prescribed reported at wave 3 and risk of incident NMUPO at waves 4 or 5. Specifically, we used the %LGTPHCURV9 macro to fit a restricted cubic spline for proportional hazards regression models to examine the potentially non-linear relationship between days supplied and risk of NMUPO initiation (Li et al., 2011). All analyses were performed using SAS version 9.4 (SAS Institute, Cary, NC).

### 3. Results

#### 3.1. Participant characteristics

Demographic, substance use and substance use disorder, psychiatric, medical, and medication characteristics of wave 3 participants without a history of NMUPO who had received a prescription opioid in the prior year ( $n = 815$ ) are summarized in Table 1. Wave-3 data are also presented separately for participants who did or did not exhibit incident NMUPO at waves 4 or 5. Participants were on average 52 years old and were predominantly male (93.3%) and African American (59.8%). The rates of the past-12 month and lifetime NMUPO reported among wave 3 respondents who were prescribed opioid medication in the previous year ( $n = 1308$ ) were 20.0% and 37.2%, respectively (see Fig. 1).

#### 3.2. NMUPO incidence at waves 4–5

Of the 815 participants at wave 3 who received opioid medication in the previous 12 months and reported no past-year or lifetime NMUPO (see Fig. 1), 15.0% ( $n = 122$ ) reported new-onset NMUPO at waves 4 or 5 (9.2% [ $n = 75$ ] reported past 12-month NMUPO at wave 4, and 11.2% [ $n = 91$ ] reported past 12-month NMUPO at wave 5). The incidence rate of NMUPO among the 815 eligible participants was 5.02 per 100 person-years (95% confidence interval [CI] 4.18–5.97 per 100-person-years).

#### 3.3. Characteristics associated with NMUPO incidence

As summarized in Table 1, of the demographic, substance use and substance use disorder, psychiatric, medical, and medication characteristics examined in the bivariable model, only race/ethnicity and days of opioid medication supplied were significantly associated with NMUPO incidence (at  $p < 0.05$ ). Respondents with incident NMUPO compared to those without were more likely to be Hispanic (13.9% vs. 7.4%), to have higher mean days of opioid medication prescribed (56 vs. 40 days), and to have been supplied opioid medication for 30 days or more (74.6% vs. 64.7%). Opioid dosing characteristics, including median or

mean MED, high dose opioid receipt, composite days/dose, and presence of schedule II short-acting or long-acting medications, were not significantly associated with NMUPO. In the multivariable Cox regression analysis, days of opioids supplied remained positively associated with incident NMUPO (adjusted hazard ratio [AHR] = 1.65, 95% CI: 1.06–2.58 for 30–180 days; AHR = 1.99 95% CI: 1.21–3.29 for > 180, compared to < 30 days) (see Table 2). Fig. 2 summarizes the results of the spline regression model, which demonstrates a non-linear relationship between days of opioid medication supplied and risk of incident NMUPO. Specifically, the risk of new-onset NMUPO increases sharply until 75 days and then plateaus, followed by a slower increase in the risk after approximately 200 days of opioid medication supplied.

#### 4. Discussion

The current study is among the first to assess the incidence of NMUPO among veterans with and without HIV who received prescription opioids, and to examine demographic, substance use and substance use disorder, psychiatric, medical, and medication predictors of new-onset NMUPO. Three main findings emerged. First, of wave-3 participants prescribed opioids in the previous year, 37% reported lifetime (and 20% past-year) NMUPO. Second, among wave-3 participants without prior NMUPO who were prescribed opioids in the previous year, 15% reported incident NMUPO at waves 4 or 5. Third, in multivariable analyses, NMUPO incidence was associated with duration of opioids prescribed.

In a prior study of wave-3 VACS participants (irrespective of opioid medication receipt), the estimate of lifetime NMUPO was 13% (Barry et al., 2011). By contrast, in the current study when the sample was restricted to those prescribed opioids in the previous year, the lifetime rate of NMUPO was markedly higher (37.2%). These findings support those from previous studies regarding the NMUPO risk associated with prescribed opioids. For example, in a study of veterans in primary care, pain was found to be associated with both receipt of opioids and NMUPO (Becker et al., 2009), while another study of patients with chronic pain found higher rates of opioid use disorder among patients prescribed (compared to those not prescribed) opioids (Edlund et al., 2014).

Prior studies of NMUPO in individuals prescribed opioids have generally used cross-sectional designs to estimate lifetime prevalence and associated risk factors (e.g., history of substance use disorder) (Barry et al., 2011; Morasco and Dobscha, 2008). Findings from the current longitudinal study highlight the importance of considering NMUPO risk at multiple time-points. Among wave-3 participants prescribed opioids in the previous year with no prior history of NMUPO, 15% reported incident NMUPO at waves 4 or 5. Consistent with the recent CDC opioid prescribing guidelines, our findings highlight the importance of clinicians' continually assessing for NMUPO (and not just at opioid treatment initiation).

As predicted, in multivariable analyses, NMUPO was independently associated with duration of prescribed opioid receipt. Specifically, we found that the risk of new-onset NMUPO among veterans prescribed opioids increases until about 75 days and then plateaus, and is followed by a slower increase after approximately 200 days. While prior studies have demonstrated that the duration of initial opioid prescriptions is associated with greater odds

of long-term opioid therapy among previously opioid-naïve patients (Deyo et al., 2017; Shah et al., 2017) and incident opioid use disorder among patients with chronic pain (Chou et al., 2015; Edlund et al., 2014), findings from the current study indicate that the duration of prescribed opioids among veterans in care also predicts new-onset NMUPO.

Notably, while substance use, medical status, and pain interference were previously found to be independent predictors of NMUPO prevalence among all wave-3 participants (Barry et al., 2011), none of these variables were significantly associated in the current study with incident NMUPO at the subsequent two waves. Contrary to the study hypothesis, in multivariable analyses, no significant association emerged between NMUPO incidence and average prescribed daily opioid dose. In prior studies of patients with chronic pain, higher average daily prescribed opioid doses have been found to be associated with longer duration of opioid use (Chou et al., 2015; Edlund et al., 2014). Similarly, in a recent study of an integrated healthcare system that used prescription opioid registry data, higher daily opioid doses were associated with increased risk of opioid misuse (Campbell et al., 2018). Differences in study design (e.g., cross-sectional vs. longitudinal, middle-aged vs. younger participants) and definition of opioid exposure (e.g., self-report vs. pharmacy records) may explain some of these discrepant findings. Our findings also raise the possibility that predictors for NMUPO prevalence and incidence are not identical.

Prior studies on the risk of NMUPO or opioid use disorder in clinical populations have generally focused on patients with chronic pain who were prescribed opioids for pain relief. The median pain intensity and pain interference ratings among veterans in care were relatively low and somewhat surprising. The reasons for these generally low scores (e.g., self-medication) and the extent to which the use of opioid medication (especially among the approximately 31% receiving long-term opioid therapy) was clinically indicated are unclear and merit further investigation.

Our study had several limitations. The main study variable, NMUPO, was measured by self-report, which may be subject to recall bias or under-reporting because of social desirability. The NSDUH has a purposeful community sampling strategy that targets adolescents and adults; in contrast, the current study sample comprised primarily middle-aged veterans in medical care, which restricts age-related comparisons. While the definition of NMUPO used in this study has been widely used by researchers (Barry et al., 2011), it does not address additional factors that are likely to be important to clinicians (e.g., frequency, motivation). The duration and patterns (e.g., continuous vs. non-continuous) of lifetime exposure to the medical use of prescription opioids were not assessed and consequently were not included in our analyses; thus, the potential impact of these factors on incident NMUPO is unclear. Data on opioid prescriptions filled were restricted to those drawn from the VHA Pharmacy Benefits Management records system; consequently, any opioid medications filled outside of the VHA were not captured (Gellad et al., 2018). Whereas we adjusted in the Cox regression analysis for clinician diagnoses of substance use disorders and self-reported alcohol and tobacco use, we did not adjust for illicit drug use. Neither were objective measures of substance use collected or used in the data analyses (e.g., urine toxicology). Although the VHA is one of the largest healthcare systems internationally, its patients are predominantly men who have been exposed to US military training and military conflict. Additionally,



VHA policies on opioid prescribing for non-cancer chronic pain have shifted considerably in recent years. It is also important to note that the extent to which physicians incorporated opioid medications into the routine management of non-cancer chronic pain has varied considerably internationally. Consequently, both the characteristics of VHA patients and fluctuations in VHA opioid prescribing policies suggest caution in generalizing the findings of the current study to other countries, settings, and patient populations.

#### 4.1. Conclusions

Despite these limitations, this study has several strengths. Unlike prior published studies on NMUPO, this study used a relatively large clinical sample of veterans in care with and without HIV infection, systematically assessed for NMUPO over time, and had access to data from participants' pharmacy, medical, and psychiatric records. Consequently, unlike previous studies, we were able to assess NMUPO incidence and correlates among a clinical population with and without HIV who were prescribed opioid medications.

In summary, we found relatively high rates of both lifetime NMUPO among veterans in care who were prescribed opioids in the previous year (37%) and NMUPO incidence over the study period among those reporting no prior history of NMUPO (15%). In multivariable analyses, new-onset NMUPO was positively and independently associated with receipt of prescribed opioids for 30–180 days. Our findings suggest that prescribers should be aware of this risk when prescribing long-term opioids and should assess for NMUPO even among those without evidence of aberrant behavior.

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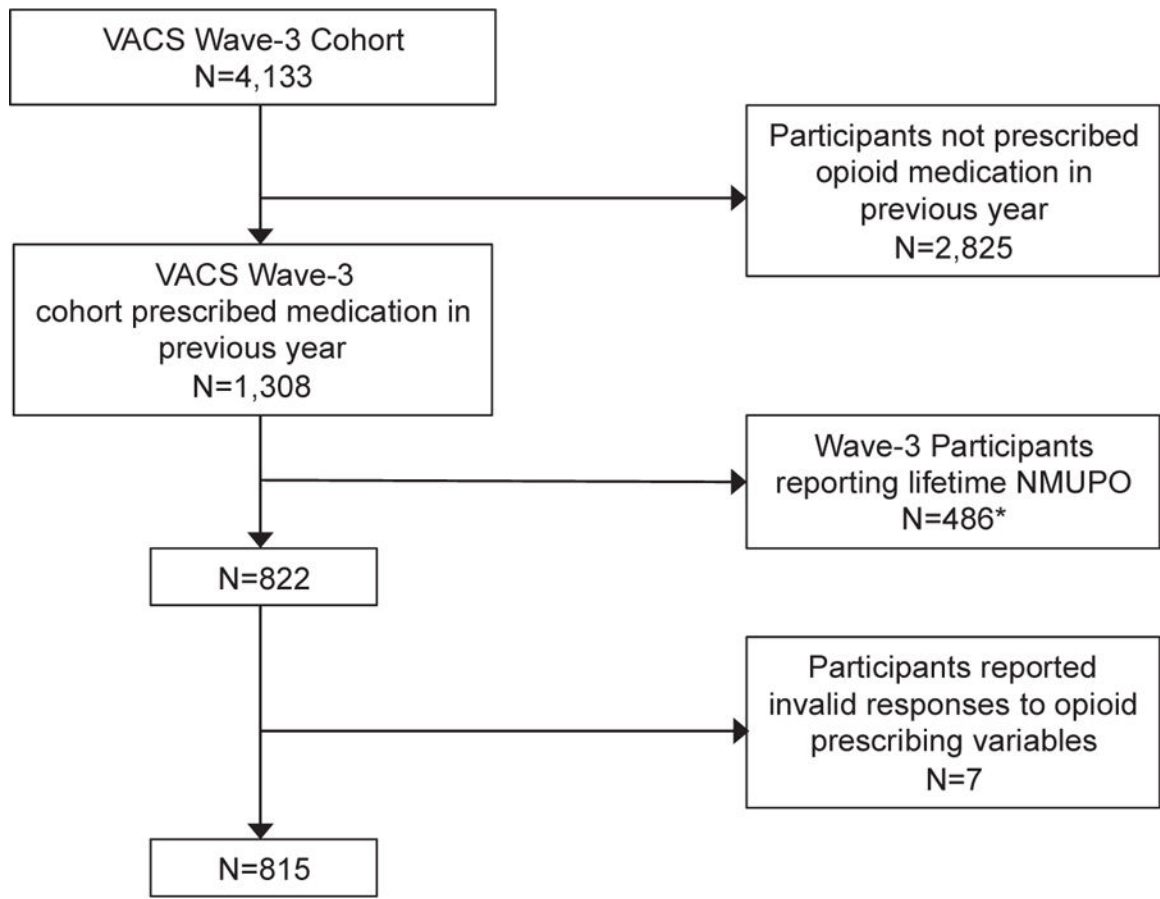
#### References

- Banerjee G, Edelman EJ, Barry DT, Becker WC, Cerdá M, Crystal S, Gaither JR, Gordon AJ, Gordon KS, Kerns RD, 2016 Non-medical use of prescription opioids is associated with heroin initiation among US veterans: a prospective cohort study. *Addiction* 111, 2021–2031. [PubMed: 27552496]
- Barry DT, Beitel M, Joshi D, Schottenfeld RS, 2009 Pain and substance-related pain-reduction behaviors among opioid dependent individuals seeking methadone maintenance treatment. *Am. J. Addict* 18, 117–121. [PubMed: 19283562]

- Barry DT, Irwin KS, Jones ES, Becker WC, Tetrault JM, Sullivan LE, Hansen H, O'Connor PG, Schottenfeld RS, Fiellin DA, 2010 Opioids, chronic pain, and addiction in primary care. *J. Pain* 11, 1442–1450. [PubMed: 20627817]
- Barry DT, Goulet JL, Kerns RK, Becker WC, Gordon AJ, Justice AC, Fiellin DA, 2011 Nonmedical use of prescription opioids and pain in veterans with and without HIV. *Pain* 152, 1133–1138. [PubMed: 21354703]
- Barry DT, Savant JD, Beitel M, Cutter CJ, Moore BA, Schottenfeld RS, Fiellin DA, 2013 Pain and associated substance use among opioid dependent individuals seeking office-based treatment with buprenorphine-naloxone: a needs assessment study. *Am. J. Addict* 22, 212–217. [PubMed: 23617861]
- Becker WC, Sullivan LE, Tetrault JM, Desai RA, Fiellin DA, 2008 Non-medical use, abuse and dependence on prescription opioids among US adults: psychiatric, medical and substance use correlates. *Drug Alcohol Depend* 94, 38–47. [PubMed: 18063321]
- Becker WC, Fiellin DA, Gallagher RM, Barth KS, Ross JT, Oslin DW, 2009 The association between chronic pain and prescription drug abuse in veterans. *Pain Med* 10, 531–536. [PubMed: 19425211]
- Bush K, Kivlahan DR, McDonell MB, Fihn SD, Bradley KA, 1998 The AUDIT alcohol consumption questions (AUDIT-C): an effective brief screening test for problem drinking. *Arch. Intern. Med* 158, 1789–1795. [PubMed: 9738608]
- Campbell CI, Bahorik AL, VanVeldhuisen P, Weisner C, Rubinstein AL, Ray GT, 2018 Use of a prescription opioid registry to examine opioid misuse and overdose in an integrated health system. *Prev. Med* 110, 31–37. [PubMed: 29410132]
- Chou R, Turner JA, Devine EB, Hansen RN, Sullivan SD, Blazina I, Dana T, Bougatsos C, Deyo RA, 2015 The effectiveness and risks of long-term opioid therapy for chronic pain: a systematic review for a National Institutes of Health Pathways to Prevention Workshop. *Ann. Intern. Med* 162, 276–286. [PubMed: 25581257]
- Chou R, Deyo R, Friedly J, Skelly A, Hashimoto R, Weimer M, Fu R, Dana T, Kraegel P, Griffin J, 2017 Nonpharmacologic therapies for low back pain: a systematic review for an American College of Physicians clinical practice guideline nonpharmacologic therapies for low back pain. *Ann. Intern. Med* 166 (7), 493–505. [PubMed: 28192793]
- Compton WM, Jones CM, Baldwin GT, 2016 Relationship between nonmedical prescription-opioid use and heroin use. *N. Engl. J. Med* 374, 154–163. [PubMed: 26760086]
- Department of Veterans Affairs Department of Defense, 2010 VA/DoD Clinical Practice Guideline for Management of Opioid Therapy for Chronic Pain, v2.0 Washington, DC.
- Deyo RA, Hallvik SE, Hildebran C, Marino M, Dexter E, Irvine JM, O'Kane N, Van Otterloo J, Wright DA, Leichtling G, 2017 Association between initial opioid prescribing patterns and subsequent long-term use among opioid-naïve patients: a statewide retrospective cohort study. *J. Gen. Intern. Med* 32, 21–27. [PubMed: 27484682]
- Dowell D, Haegerich TM, Chou R, 2016 CDC guideline for prescribing opioids for chronic pain—United States, 2016. *JAMA* 315, 1624–1645. [PubMed: 26977696]
- Dunn KM, Saunders KW, Rutter CM, Banta-Green CJ, Merrill JO, Sullivan MD, Weisner CM, Silverberg MJ, Campbell CI, Psaty BM, 2010 Opioid prescriptions for chronic pain and overdose: a cohort study. *Ann. Intern. Med* 152, 85–92. [PubMed: 20083827]
- Edelman EJ, Gordon K, Becker WC, Goulet JL, Skanderson M, Gaither JR, Brennan Braden J, Gordon AJ, Kerns RD, Justice AC, 2013 Receipt of opioid analgesics by HIV-infected and uninfected patients. *J. Gen. Intern. Med* 28, 82–90. [PubMed: 22895747]
- Edlund MJ, Martin BC, Russo JE, Devries A, Braden JB, Sullivan MD, 2014 The role of opioid prescription in incident opioid abuse and dependence among individuals with chronic non-cancer pain: the role of opioid prescription. *Clin. J. Pain* 30, 557. [PubMed: 24281273]
- Gellad WF, Sileanu FE, Hale JA, Radomski TR, Gordon AJ, Good CB, Fine MJ, 2018 Impact of dual use of Department of Veterans Affairs and Medicare Part D drug benefits on potentially unsafe opioid use. *Am. J. Public Health* 108, 248–255. [PubMed: 29267065]
- Häuser W, Schug S, Furlan AD, 2017 The opioid epidemic and national guidelines for opioid therapy for chronic noncancer pain: a perspective from different continents. *Pain Rep* 2, e599. [PubMed: 29392214]

- Institute of Medicine, 2011 *Relieving Pain in America: a Blueprint for Transforming Prevention, Care, Education and Research* The National Academies Press, Washington, DC.
- Justice AC, Dombrowski E, Conigliaro J, Fultz SL, Gibson D, Madenwald T, Goulet J, Simberkoff M, Butt AA, Rimland D, 2006 Veterans Aging Cohort Study (VACS): overview and description. *Med. Care* 44, S13–S24. [PubMed: 16849964]
- Katz MH, 2016 Opioid prescribing for chronic pain: not for the faint of heart. *JAMA Intern. Med* 176, 599–601. [PubMed: 26977893]
- Katz C, El-Gabalawy R, Keyes KM, Martins SS, Sareen J, 2013 Risk factors for incident nonmedical prescription opioid use and abuse and dependence: results from a longitudinal nationally representative sample. *Drug Alcohol Depend* 132, 107–113. [PubMed: 23399466]
- Keller CE, Ashrafioun L, Neumann AM, Van Klein J, Fox CH, Blondell RD, 2012 Practices, perceptions, and concerns of primary care physicians about opioid dependence associated with the treatment of chronic pain. *Subst. Abuse* 33, 103–113. [PubMed: 22489582]
- Kuehn BM, 2007 Opioid prescriptions soar: increase in legitimate use as well as abuse. *JAMA* 297, 249–251. [PubMed: 17227967]
- Li R, Hertzmark E, Louei M, Chen L, Spiegelman D, 2011 The SAS LGTPHCURV9 Macro Accessed 20 April 2017. <https://www.hsph.harvard.edu/donna-spiegelman/software/igtphcurv9/>.
- Lin LA, Bohnert AS, Kerns RD, Clay MA, Ganoczy D, Ilgen MA, 2017 Impact of the Opioid Safety Initiative on opioid-related prescribing in veterans. *Pain* 158, 833–839. [PubMed: 28240996]
- Lum PJ, Little S, Botsko M, Hersh D, Thawley RE, Egan JE, Mitty J, Boverman J, Fiellin DA, 2011 Opioid-prescribing practices and provider confidence recognizing opioid analgesic abuse in HIV primary care settings. *J. Acquir. Immune Defic. Syndr* 56, S91–S97. [PubMed: 21317601]
- Manchikanti L, Abdi S, Atluri S, Balog CC, Benyamin RM, Boswell MV, Brown KR, Bruel BM, Bryce DA, Burks PA, 2012 American Society of Interventional Pain Physicians (ASIPP) guidelines for responsible opioid prescribing in chronic noncancer pain: part I—evidence assessment. *Pain Phys* 15, S1–65.
- McCabe SE, Schulenberg JE, O'Malley PM, Patrick ME, Kloska DD, 2014 Non-medical use of prescription opioids during the transition to adulthood: a multi-cohort national longitudinal study. *Addiction* 109, 102–110. [PubMed: 24025114]
- McCabe SE, Kloska DD, Veliz P, Jager J, Schulenberg JE, 2016 Developmental course of non-medical use of prescription drugs from adolescence to adulthood in the United States: national longitudinal data. *Addiction* 111, 2166–2176. [PubMed: 27338559]
- Morasco BJ, Dobscha SK, 2008 Prescription medication misuse and substance use disorder in VA primary care patients with chronic pain. *Gen. Hosp. Psychiatry* 30, 93–99. [PubMed: 18291290]
- Novak SP, Håkansson A, Martinez-Raga J, Reimer J, Krotki K, Varughese S, 2016 Nonmedical use of prescription drugs in the European Union. *BMC Psychiatry* 16 (274). 10.1186/s12888-016-0909-3.
- Okie S, 2010 A flood of opioids, a rising tide of deaths. *N. Engl. J. Med* 363, 1981–1985. [PubMed: 21083382]
- Oliva EM, Bowe T, Tavakoli S, Martins S, Lewis ET, Paik M, Wiechers I, Henderson P, Harvey M, Avoundjian T, 2017 Development and applications of the Veterans Health Administration's Stratification Tool for Opioid Risk Mitigation (STORM) to improve opioid safety and prevent overdose and suicide. *Psychol. Serv* 14, 34–49. [PubMed: 28134555]
- Ray WA, Chung CP, Murray KT, Hall K, Stein CM, 2016 Prescription of long-acting opioids and mortality in patients with chronic noncancer pain. *JAMA* 315, 2415–2423. [PubMed: 27299617]
- SAMSHA, 2011 Results From the 2010 National Survey on Drug Use and Health: Summary of National Findings, NSDUH Series H-41, HHS Publication No. (SMA) 11–4658 Substance Abuse and Mental Health Services Administration, Rockville, MD <https://www.samhsa.gov/data/sites/default/files/SDUHNationalFindingsResults2010-web/2k10ResultsRev/NSDUHresultsRev2010.pdf>
- Shah A, Hayes CJ, Martin BC, 2017 Characteristics of initial prescription episodes and likelihood of long-term opioid use—United States, 2006–2015. *MMWR Morb. Mortal. Wkly. Rep* 66, 265–269. [PubMed: 28301454]

- Spitzer RL, Kroenke K, Williams JB, Group PHQPCS, 1999 Validation and utility of a self-report version of PRIME-MD: the PHQ primary care study. *JAMA* 282, 1737–1744. [PubMed: 10568646]
- Starrels JL, Peyser D, Haughton L, Fox A, Merlin JS, Arnsten JH, Cunningham CO, 2016 When human immunodeficiency virus (HIV) treatment goals conflict with guideline-based opioid prescribing: a qualitative study of HIV treatment providers. *Subst. Abus* 37, 148–153. [PubMed: 26860130]
- Substance Abuse and Mental Health Services Administration, 2007 Results From the 2006 National Survey on Drug Use and Health: National Findings Office of Applied Studies, NSDUH Series H-32, DHHS Publication No. SMA 07–4293. Accessed on 30 June 2010. <http://www.oas.samhsa.gov/nsduh/2k6nsduh/2k6results.pdf>.
- US Drug Enforcement Administration Controlled Substance Schedules Accessed 20 April 2017. <https://www.deadiversion.usdoj.gov/schedules>.
- Von Korff M, Saunders K, Ray GT, Boudreau D, Campbell C, Merrill J, Sullivan MD, Rutter C, Silverberg M, Banta-Green C, 2008 Defacto long-term opioid therapy for non-cancer pain. *Clin. J. Pain* 24, 521–527. [PubMed: 18574361]
- World Health Organization, 1975 Manual of the International Statistical Classification of Diseases, Injuries, and Causes of Death, vol 1 ninth revision Geneva <http://www.who.int/iris/handle/10665/40492>.

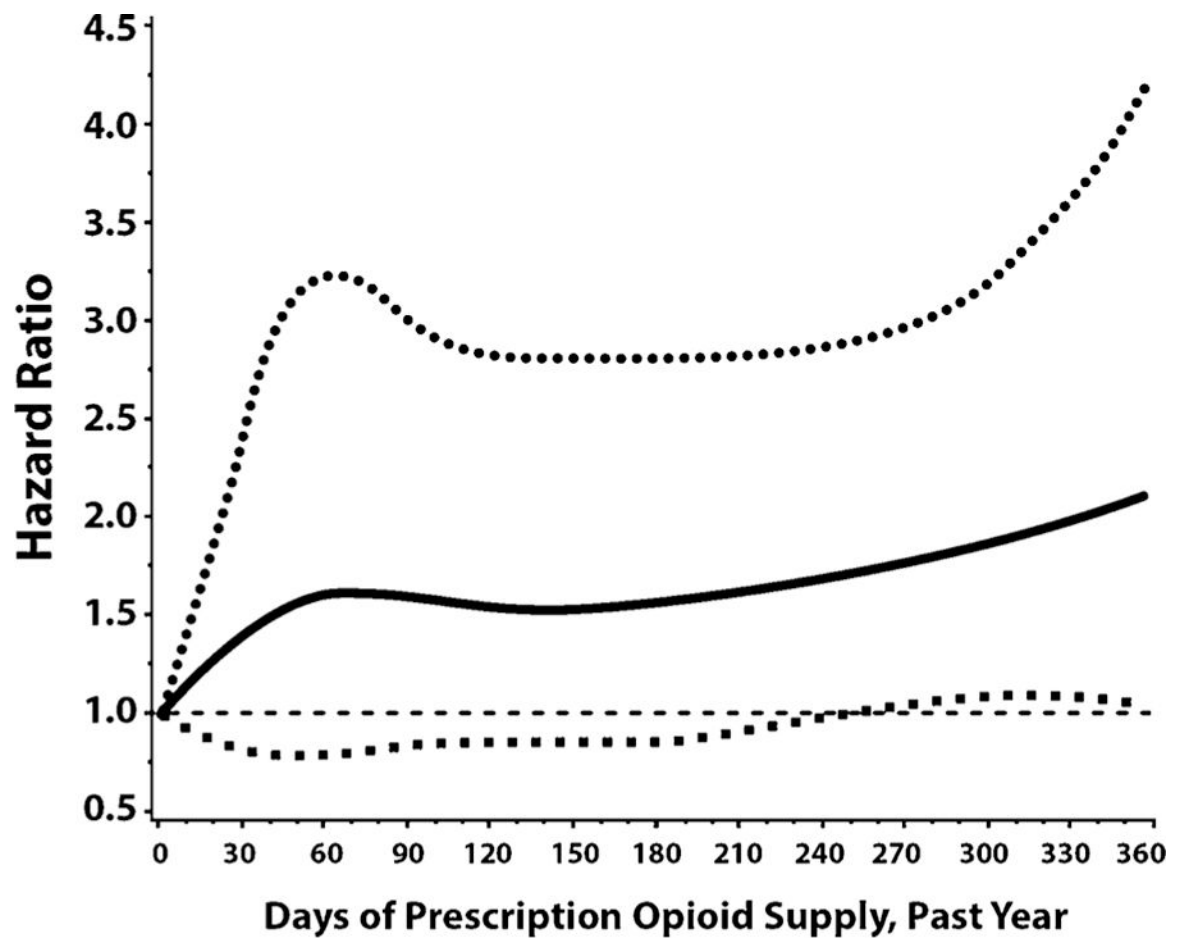


**Fig. 1.**

Flow Chart of eligible participants for incidence-related analyses.

Note: VACS: Veteran Aging Cohort Study, NMUPO: nonmedical use of prescribed opioids.

\*262 participants reported past-year NMUPO.



**Fig. 2.** Spline regression showing non-linear relationship between days of prescription opioid supply and risk of NMUPO initiation at follow-up 4 or 5. Note: solid line indicates the hazard ratio estimate for risk of NMUPO initiation at a given number of days prescription opioid supply in the year prior to wave 3 survey date, compared to persons with no opioids prescribed during the same time period. The dashed lines indicate 95% confidence intervals around this estimate.

**Table 1**  
 Characteristics of participants prescribed opioids with and without incident NMUPO.

	Overall (N = 815)	Incident NMUPO (N = 122, 15.0%)	No incident NMUPO (N = 693, 85.0%)	p-value <sup>a</sup>
<b>Demographics</b>				
Age (median, IQR)	52 (47–58)	52 (47–57)	52 (47–58)	0.94
<b>Sex</b>				
Female	54 (6.6)	4 (3.3)	50 (7.2)	0.12
Male	766 (93.3)	118 (96.7)	648 (92.7)	
<b>Race/Ethnicity</b>				
White	222 (27.2)	28 (23.0)	194 (28.0)	<b>0.049</b>
Black	498 (59.8)	74 (60.7)	413 (59.6)	
Hispanics	68 (8.3)	17 (13.9)	51 (7.4)	
Other	38 (4.7)	3 (2.5)	35 (5.0)	
<b>Substance Use and SUD</b>				
AUDIT-C (median, IQR)	1 (0–4)	2 (0–4)	1 (0–4)	0.23
Risky alcohol use <sup>b</sup>	212 (26.0)	32 (26.2)	180 (26.0)	0.95
<b>Smoking Status</b>				
Never	202 (25.0)	27 (22.3)	177 (25.5)	0.51
Current	381 (47.2)	63 (52.1)	318 (46.4)	
Former	224 (27.8)	31 (25.6)	193 (28.1)	
<b>Alcohol use disorder</b>				
Drug use disorder	104 (12.8)	15 (12.3)	89 (12.8)	0.87
<b>Psychiatric Status</b>				
Major depression	121 (14.9)	19 (15.6)	102 (14.7)	0.81
Bipolar disorder	86 (10.6)	9 (7.4)	77 (11.1)	0.22
Schizophrenia	43 (5.3)	9 (7.4)	34 (4.9)	0.26
Posttraumatic stress disorder	23 (2.7)	3 (2.5)	19 (2.7)	1.00
PHQ-9 (median, IQR)	97 (11.9)	12 (9.8)	85 (12.3)	0.46
Medical Status	4 (0–9)	3 (0–7)	4 (0–9)	0.26
HIV	396 (48.6)	62 (50.8)	334 (48.2)	0.59
Hepatitis C	307 (37.7)	49 (40.2)	258 (37.2)	0.54

	Overall (N = 815)	Incident NMUPO (N = 122, 15.0%)	No incident NMUPO (N = 693, 85.0%)	p-value <sup>a</sup>
Pain intensity (median, IQR)	2 (0–6)	3 (0–6)	2 (0–5)	0.51
Pain interference (median, IQR)	1 (0–4)	1 (0–4)	0.5 (0–3)	0.26
Medication-related Characteristics				
Opioid Medication				
Days supplied, continuous	45 (15–176)	56 (27–221)	40 (15–162)	<b>0.03</b>
Days supplied, categorical				<b>0.045</b>
< 30 days	276 (33.9)	31 (25.4)	245 (36.4)	
30 to 180 days	344 (42.2)	53 (43.4)	291 (42.0)	
> 180 days	195 (23.9)	38 (31.2)	157 (22.7)	
Median daily MED (IQR)	20 (13.5–33.0)	19.3 (12.6–30.0)	20 (13.5–33.4)	0.51
Mean daily MED (mg)				0.54
0–49	692 (84.9)	105 (86.1)	587 (84.7)	
50–79	60 (7.4)	11 (9.0)	49 (7.1)	
80–119	26 (3.2)	3 (2.5)	23 (3.3)	
120	37 (4.5)	3 (2.5)	34 (4.9)	
High-dose receipt (> 50MED)	123 (15.1)	17 (13.9)	106 (15.3)	0.69
Composite Days/Dose <sup>c</sup>				0.09
Low-dose, < 30 days	260 (31.9)	30 (24.6)	230 (33.2)	
Low-dose, 30–180 days	308 (37.8)	51 (41.8)	257 (37.1)	
Low-dose, > 180 days	125 (15.2)	24 (19.7)	100 (14.4)	
High-dose, < 30 days	16 (2.0)	1 (0.8)	15 (2.2)	
High-dose, 30–180 days	36 (4.4)	2 (1.6)	34 (4.9)	
High-dose, > 180 days	71 (8.1)	14 (11.5)	57 (8.2)	
Schedule II short-acting	291 (35.7)	47 (38.5)	244 (35.2)	0.48
Schedule II long-acting	116 (14.2)	19 (15.6)	97 (14.0)	0.65
Benzodiazepine Medication <sup>d</sup>				
Any benzodiazepine receipt	172 (21.0)	24 (19.7)	148 (21.2)	0.71
Long-term benzodiazepine receipt	108 (13.3)	14 (11.5)	94 (13.6)	0.53

Note: NMUPO = non-medical use of prescription opioids; IQR = interquartile range; SUD = substance use disorder; AUDIT-C = Alcohol Use Disorder Identification Test; PHQ-9 = Patient Health Questionnaire; MED = morphine equivalent dose.



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<sup>a</sup> **Bolded values were significant at  $p < 0.05$ .**

<sup>b</sup> Established AUDIT-C cutoffs ( 4 for men, 3 for women) were used to determine risky alcohol use.

<sup>c</sup> Low-dose refers to 50MED; high-dose refers to > 50MED.

<sup>d</sup> This medication category included both benzodiazepines and non-benzodiazepine hypnotics.

**Table 2**

Cox proportional hazard regression model of factors associated with incident NMUPO among veterans participating in VACS, 2005–2011.

Characteristic	Adjusted HR* (95% CI)	P-value
Sex (ref: male)	0.57 (0.21–1.55)	0.27
Race (ref: White)		
Black	1.13 (0.73–1.75)	0.57
Hispanic	1.57 (0.86–2.89)	0.14
Other	0.74 (0.22–2.44)	0.62
Days of opioid medication supplied (ref: < 30)		
30–180	1.65 (1.06–2.58)	0.03
> 180	1.99 (1.21–3.29)	0.01
High-dose opioid receipt (ref: none) <sup>†</sup>	1.35 (0.78–2.33)	0.29

*Note:* NMUPO = nonmedical use of prescription opioids; VACS = Veterans Aging Cohort Study.

\* HR = Hazard Ratio

<sup>†</sup> refers to > 50 mg daily morphine equivalent dose.