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Psychometric Properties of an Abbreviated Childhood Family Mealtime Questionnaire among Overweight and Obese Hispanic Adolescents

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Abstract

The current family mealtime literature shows that assessments of the mealtime environment are typically self-report, yet few studies discuss validation techniques or report using validated scales. As such, the current analysis was conducted to validate one of the only published measures to assess the mealtime environment from the adolescent perspective. Specifically, the Childhood Family Mealtime Questionnaire (CFMQ) was evaluated in a sample of 280 overweight and obese Hispanic adolescents to address the need for a validated measure of the family mealtime environment in a demographic that is disproportionately affected by the current obesity epidemic. Results of an exploratory and confirmatory factor analysis to evaluate the optimal factor structure, reliability, and validity for a revised, abbreviated CFMQ are presented here. The concurrent validity of the CFMQ was evaluated using correlations between the factor structures and the previously used, culturally appropriate comparable measure of family functioning. Correlations were also computed between factor scores and obesogenic outcomes (fruit and vegetable intake, added sugar intake, and physical activity). Analyses produced a revised, abbreviated version that includes 22 items (reduced from a total of 69 items) and consists of the following 4 factors: family mealtime communication (5 items), family mealtime stress (7 items), appearance weight control (5 items), and mealtime structure (6 items). Cronbach's alphas are reported for reliability. When examining CFMQ concurrent validity with the family functioning latent variable, results showed the family mealtime communication subscale ranked highest. Additionally, the family mealtime communication subscale was associated with all three obesogenic outcomes. This abbreviated CFMQ may be a useful tool for those studying family mealtime environments and their influence on obesity and its associated lifestyle behaviors.

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INTRODUCTION

Recent national prevalence estimates show that dietary intake and physical activity behaviors among the United States (US) adolescent population are not conducive to healthy weight development and are likely to promote adiposity (Kann et al., 2016). As such, pediatric obesity rates remain unacceptably high, particularly among Hispanic children and adolescents whose prevalence rates are higher than non-Hispanic whites and Asians, (38.2% compared to 19.1%, 12.1%, respectively) and about the same as non-Hispanic black adolescents (35.0%) (Skinner, Ravanbakht, Skelton, Perrin, & Armstrong, 2018). This disproportion among ethnic minority youth indicates the need for culturally competent strategies to target healthful eating and physical activity practices.

Studies have highlighted the important influence that parents and the family home environment have on youth dietary choices and other health-related behaviors (Burns, Parker, & Birch, 2011; Larsen et al., 2015; Lloyd, Lubans, Plotnikoff, Collins, & Morgan, 2014; Niemeier, Hektner, & Enger, 2012; Vaughn et al., 2015; Yee, Lwin, & Ho, 2017). The family meal has been shown to be influential in the consumption of a healthy diet and may be an important avenue for obesity treatment and prevention. Meta-analysis of 17 studies (15 reported cross-sectional findings and 5 reported longitudinal findings) with children and adolescents ages 2.8 to 17.3 years (N=182,836) found that youth from families who shared meals together three or more times per week were 12% less likely to be overweight, 20% less likely to consume unhealthy foods, and 24% more likely to consume healthy foods (Hammons & Fiese, 2011). No study to our knowledge has been conducted to examine how family mealtime is associated with physical activity. Yet studies have demonstrated the impact family functioning has on adolescent physical activity (Berge, Wall, Larson, Loth, & Neumark-Sztainer, 2013; Lebron et al., 2018). For example, in one study discrepancies between parent and adolescent report of family functioning resulted in less reported adolescent physical activity (Lebron et al., 2018). If family functioning constructs like communication and positive parenting influence are given the opportunity to play out during mealtime, it follows that mealtime behaviors would also have an impact on physical activity. Few studies have examined the relationship between mealtime frequency and obesity in diverse racial/ethnic groups. While some have reported that the association was not significant for Hispanic children (Rollins, Belue, & Francis, 2010) and adolescents (Fulkerson, Neumark/ Sztainer, Hannan, & Story, 2008; Sen, 2006), others have found that family meal frequency and family support of healthful eating was associated with outcomes like consuming more fiber and fruits and vegetables in both Hispanic children and adolescents (Andaya, Arredondo, Alcaraz, Lindsay, & Elder, 2011; Ayala et al., 2007). Inconsistent results have been attributed to differences in gender, level of education, and types of foods and portions consumed at family meals.

Although the literature supports family meal frequency as a protective factor for several adverse adolescent health outcomes, much less is known about how the family mealtime environment directly or indirectly contributes to weight status, and especially among Hispanic families. Parents who encourage shared meals may have better relationships with their adolescents (Goldfarb, Tarver, Locher, Preskitt, & Sen, 2015). Similarly, older adolescents may be given the choice by their parents to eat dinner together, and those that

choose to may place different meaning or value on shared mealtime than those who chose to opt out (Goldfarb et al., 2015). One study conducted with children ages 5–12 years showed that families of healthy weight children considered mealtime important and meaningful for gathering and conversing (Fiese, Hammons, & Grigsby-Toussaint, 2012). The benefits of sharing family meals may be diminished, however, if the family consumes fast food, sits around the television, or if it creates an argumentative, stressful environment (Martin-Biggers et al., 2014). Therefore, questions have been raised about the protective nature of mealtimes and specifically if findings are due to the frequency of family meals or if there are other systematic differences between families who eat together and those who do not (D. Neumark-Sztainer, Larson, Fulkerson, Eisenberg, & Story, 2010).

To arrive at the meaning of the questions, validated measures are needed to evaluate multiple domains of adolescent mealtime experiences. Previous empirical studies of the familial mealtime atmosphere have typically collected data via parent-report of their children's behaviors (Anderson, Must, Curtin, & Bandini, 2012; Burnier, Dubois, & Girard, 2011; McCurdy & Gorman, 2010), parent report of their own behavior (Hendy, Williams, Camise, Eckman, & Hedemann, 2009), or observation (Fiese et al., 2012). Other instruments have focused on family or child eating habits or behaviors (Boquin, Smith-Simpson, Donovan, & Lee, 2014; Golan & Weizman, 1998). For example, one widely validated scale is the Behavioral Pediatric Feeding Assessment Scale, which includes items like "Gets up from table during meal", "Comes readily to mealtime", and "Tantrums at mealtimes" (Crist & Napier-Phillips, 2001; Davis, Canter, Stough, Gillette, & Patton, 2013; Patton, Dolan, & Powers, 2006). The body of literature around family mealtimes and its impact on adolescents has been greatly influenced by Project EAT which has shown how family mealtime frequencies, food choices, and activities affect adolescent health outcomes (Arcan et al., 2007; D. Neumark-Sztainer, Eisenberg, Fulkerson, Story, & Larson, 2008; D. Neumark-Sztainer et al., 2010; D. Neumark-Sztainer, Wall, Perry, & Story, 2003; D. Neumark-Sztainer, Wall, Story, & Standish, 2012; Neumark/Sztainer, Wall, Story, & Sherwood, 2009; Watts, Berge, Loth, Larson, & Neumark-Sztainer, 2018). However, one of the foremost limitations of all of the studies discussed is the limited evidence of discussion of development and psychometrics (Martin-Biggers et al., 2014).

The current analysis was conducted to validate one of the only published measures to assess the mealtime environment from the adolescent perspective, the Childhood Family Mealtime Questionnaire (CFMQ) (Miller, McCluskey-Fawcett, & Irving, 1993). The CFMQ is a 69-item scale developed to assess childhood mealtime experiences and was inspired by the literature correlating various eating problems and disorders with early food experiences (Bruch, 1969; Frank, 1991; Marcus & Wiener, 1989; Miller et al., 1993). In a validation study of the CFMQ, Miller et al. (1993) reported that a 35-item subset of the questions could successfully predict respondent as being bulimic, repeat dieters, or non-bulimic. In one study, the CMFQ subscales were used to investigate the disturbed eating behaviors and related-psychographic characteristics of young adults (ages 18 to 26 years). Findings demonstrated that participants with diet-related chronic health conditions recalled more emphasis on mother's weight and less mealtime structure (V. M. Quick, McWilliams, & Byrd-Bredbenner, 2012). This study seeks to extend the use of the CFMQ to include mealtime experiences in the context of overweight/obesity. Childhood overweight has been

associated with an increased risk of disordered eating symptoms including body dissatisfaction, dieting, and other unhealthy weight control, and binge eating (Haines & Neumark-Sztainer, 2006; D. R. Neumark-Sztainer et al., 2007). The shared symptomology of overweight and disordered eating suggests that each condition may perpetuate the other (Goldschmidt, Aspen, Sinton, Tanofsky-Kraff, & Wilfley, 2008), and therefore, deem the CFMQ appropriate for use in overweight/obesity.

This study has three aims conducted sequentially. Aim 1 was to identify the optimal factor structure of the CFMQ scales as applied to a sample of Hispanic overweight/obese youth (mean age= 13.01) using an exploratory factor analysis (EFA) with items from the CFMQ. Aim 2 was to use confirmatory factor analysis (CFA) to evaluate the construct validity and reliability of the CFMQ factor structure results in Aim 1. Aim 3 was to evaluate the subscales produced and confirmed in Aims 1 and 2 by testing concurrent validity with a family functioning subscale and by analyzing the associations between the subscales and three obesogenic outcomes (fruit and vegetable intake, added sugar intake, and physical activity).

METHODS

Data for this analysis were derived from the baseline assessment of Familias Unidas for Health and Wellness, an ongoing randomized clinical trial evaluating the relative efficacy of a family-based intervention on obesity-related lifestyle behaviors among overweight and obese Hispanic adolescents (St. George, 2018). Study staff recruited participants from 18 middle schools in Miami-Dade County distributing letters with the study description to 7th and 8th graders. If parents were interested, they were instructed to provide their contact information. Students who returned a signed letter were visually screened by study staff who identified those who might have a BMI 85th percentile for their age and sex based on body silhouette images (Stunkard, Sorensen, & Schlusinger, 2006). This procedure was used to reduce the stigma associated with obesity (i.e., letters did not mention the BMI criterion and described the intervention broadly as one designed to “promote healthy choices and prevent risky behaviors in Hispanic youth”). To be eligible for this study families had to have a Hispanic adolescent who (1) was in the 7th or 8th grade, (2) had a BMI 85th percentile for their age and sex, (3) lived with an adult primary caregiver willing to participate in the two-year study, (4) have plans to remain a resident of South Florida during the two-year study period. If parent responses on a physical activity readiness questionnaire indicated a serious health issue (e.g., chest pain) for parents and/or adolescents, they were required to acquire physician approval to participate. Details of the intervention are described elsewhere (St George et al., 2018). This study was approved by both the University of Miami and Miami Dade County Public School System (MDCPS-S) Institutional Review Board.

Measures

Family functioning was used to evaluate concurrent validity because interaction patterns during a routine mealtime are thought to be consistent with healthy family functioning (Dickstein, 2002; Fiese, Foley, & Spagnola, 2006). Furthermore, positive family functioning has been associated with lower body mass index z score, more frequent family meals, and

less sedentary behavior in adolescents (Berge et al., 2013). Adolescent self-report of measures were used as adolescent and parent reports on variables such as family functioning have been found to be incongruent (Ohannessian, Lerner, Lerner, & von Eye, 1995; Schwartz, Mason, Pantin, & Szapocznik, 2008; Stuart & Jose, 2012) and impactful in the relationship of obesogenic behaviors (Lebron et al., 2018). All measures were translated and back-translated and were offered to participants in English or Spanish.

Childhood Family Mealtime Questionnaire (CFMQ)—Although originally the CFMQ had 69-items, Miller et al (1993) conducted a factor analysis conducted on all items resulted in a subset of 35-items forming seven different factors. Those seven factors, in turn, were categorized as subscales of the CFMQ and labeled: 1) Mealtime Communication Based Stress, 2) Mealtime Structure, and 3) Appearance Weight Control, 4) Parental Mealtime Control, 5) Emphasis on Mother’s Weight, 6) Present Parental Meal Influence, 7) and Traditional Family (Miller et al., 1993). Other studies have used the original subscales of the CFMQ to analyze how family mealtimes influence disturbed eating patterns and eating attitudes in general (Meno, Hannum, Espelage, & Low, 2008; V Quick & Byrd/ Bredbenner, 2013; V. M. Quick et al., 2012; Worobey, 2002a, 2002b). The original CFMQ has inter-item reliability of 0.91, with items scored on a scale ranging from 1 = “never” to 5 = “always” (Miller et al., 1993).

To examine factor structures of CFMQ in Hispanic youth population, the current study used the reduced scales (35 items)(Miller et al., 1993). Examples of each scale include: Mealtime Communication Based Stress (11 items) *“In my family, everyone could speak their views at dinner time;”* Mealtime Structure (7 items) *“During meals, I was told not to waste food;”* Appearance Weight Control (6 items) *“I remember worrying about my weight when I was young;”* Parental Mealtime Control (3 items) *“We ate foods my father liked;”* Emphasis on Mother’s Weight (3 items) *“My mother dieted when I was young;”* Present Parental Meal Influences (2 items) *“Presently, when I am at home, my parents influence what I eat;”* and Traditional Family (3 items) *“I saw one of my mother’s main roles as that of a cook.”*

Family Functioning

Family functioning was assessed using adolescent and parent reports of five indicators, including positive parenting, parental involvement, family communication, parental monitoring of peers, and parent-adolescent communication. Subscales from the Parenting Practices Scale (Gorman-Smith, Tolan, Zelli, & Huesmann, 1996) were used to assess positive parenting (9 items; adolescent report $\alpha = .79$, parent report $\alpha = .68$) and parental involvement (16 items; adolescent report $\alpha = .84$, parent report $\alpha = .73$). Alpha coefficients were previously reported as ranging from .68 to .81 (Gorman-Smith et al., 1996). The positive parenting subscale measures parent behaviors characterized by rewarding and acknowledging adolescent positive behaviors. Sample items for the adolescent and parent, respectively include *“When you have done something that your parents like or approve of, how often does your mother say something nice about it?”*, and *“When your child has done something that you like or approve of, do you say something nice about it; praise or give approval.”* The response options ranged from “0 = never” to “4 = always.” The parental involvement subscale included adolescent questions such as *“How often do you and your*

mom do things together at home?”, and for the parent, “*Do you and your child do things together at home?*”, with a response range of “0 = never to “4 = always.” The Parent Relationship with Peer Group Scale (Pantin, 1996) was used to assess parental monitoring of peers (5 items; adolescent report $\alpha = .84$, parent report $\alpha = .80$); this measure asks parents to indicate the extent to which they supervise adolescents’ friends, activities, and whereabouts. A five-point Likert-type scale, ranging from “1 = not at all” to “5 = extremely well (often),” was used to record responses. Sample adolescent items included “*How well do you personally know your child’s best friends?*”, and for parents, “*How well do your parents know your best friends?*” Parent-adolescent communication (20 items; adolescent report $\alpha = .86$, parent report $\alpha = .84$) was assessed using the two subscales, open communication, and problems in family communication, of Parent-Adolescent Communication Scale (Barnes & Olson, 1985). Alpha reliabilities were previously reported for each subscale as .87 and .78 (Barnes & Olson, 1985). A sample adolescent question included “When I ask questions, I get honest answers from my mother/father” and for parents, “When I ask questions I get honest answers from my child.” Response choices ranged from “1 = strongly disagree” to “5 = strongly agree.”

Dietary Intake and Physical Activity

Dietary intake (i.e., fruit and vegetable intake, added sugar intake) was assessed using the Dietary Screener Questionnaire of the National Health and Nutrition Examination Survey (NHANES, 2008). The questionnaire asks participants how much of 22 specific foods or beverages they have had in the past month. There are eight responses choices: Never, 1 time last month, 2–3 times last month, 1 time per week, 2 times per week, 3–4 times per week, 5–6 times per week, 1 time per day, 2 or more times per day. The fruit and vegetable dietary factor consists of the following food items: fruit, fruit juice, salad, fried potatoes, other potatoes, dried beans, other vegetables, tomato sauce, salsa, and pizza. The added sugars dietary factor consists of the following food items: soda, fruit drinks, cookies, cake and pie, doughnuts, ice cream, sugar/honey in coffee/tea, candy, and cereal and cereal type. For the current analyses, we used algorithms developed by the National Cancer Institute (NCI) for use with the DSQ (available at <https://epi.grants.cancer.gov/nhanes/dietscreen/scoring/current/#scoring>) to calculate daily fruit and vegetable consumption (unit: a cup; Mean [SD] = 3.38 [2.53]; skewness = 1.67) and daily added sugar consumption (unit: tsp; Mean [SD] = 17.80 [17.00]; skewness = 2.78). Following George and Mallery (2010), daily sugar intake was positively skewed and was log-transformed for analyses. Development and evaluation of the DSQ have been described elsewhere (Thompson, Midthune, Kahle, & Dodd, 2017).

Physical activity was assessed by asking adolescents “During the past 7 days, on how many days were you physically active for a total of at least 60 minutes per day? Add up all the time you spend in any kind of physical activity that increases your heart rate and make you breathe hard some of the time” as is asked in the NHANES Physical Activity and Physical Fitness Questionnaire (NHANES, 2012).

Statistical Analysis

The analyses were conducted in three sequential steps. First, to identify the optimal factor structure of CFMQ scales, an exploratory factor analysis (EFA) with 35 CFMQ items was

conducted using a principle component analysis (PCA) with geomin rotation (Sass & Schmitt, 2010). Principal Components' Analysis (PCA) has been widely used as a multivariate technique for data reduction and identifying latent factor structures (Westfall, Arias, & Fulton, 2017).

To determine the number of factors, a scree plot was used to show eigenvalues across all extracted factors. Though the criterion suggests retaining all factors that are above the eigenvalue of 1 (Yong & Pearce, 2013), this could potentially overestimate the number of true factors (Franklin, Gibson, Robertson, Pohlmann, & Fralish, 1995). To eliminate any overestimated factors, the current study used the scree plot criterion from both EFA and parallel analysis (PA). Parallel analysis is a method for creating eigenvalues from random data sets that have the same sample size and number of variables as the actual data set (Fabrigar, Wegener, MacCallum, & Strahan, 1999). Factors are retained when the eigenvalues estimated from a principal component analysis set exceeds the 95th percentile of the random data eigenvalue estimated from parallel analysis (O'Connor, 2000) because it suggests that the amount of variance explained by these factors is greater than that by chance (Lahey et al., 2004). In other words, if factor eigenvalues were lower in EFA compared to those in PA, those factors were considered unreliable and thus excluded from subsequent analyses. We derived eigenvalues from random data sets by running 100 simulations.

We investigated the communality of each item, reflecting the variance of an item in common with all other items together (Yong & Pearce, 2013). Previous studies suggest that items that obtain low commonality (< .30) have less contribution to the explanation of the variances of extracted factors (Yong & Pearce, 2013). When communality values are close to zero, the associated items are more likely to be outliers and thus distract from the model.

Communality values between .30 and 1.0 indicate that these variables should be retained, as much of their common variance can be explained by the extracted factors (Pett et al., 2003). Consistent with other reports, the current study excluded items that obtained low commonality (< .30) from subsequent analyses (Morales, Yubero, & Larranaga, 2016). For the EFA model evaluation, two fit indices were used: the Kaiser-Meyer-Olkin (KMO) measure of sampling adequacy (> .80 is an acceptable fit; (Kaiser, 1970)) and Bartlett's test of sphericity (significant *p*-value indicates the EFA fits to the data).

Next, we evaluated construct validity and reliability of the optimal factor structure that was estimated in step 1. Construct validity was evaluated by using confirmatory factor analysis (CFA) with a structural equation modeling (SEM) approach (Brown, 2006). Construct validity was tested by investigating correlations between items and corresponding factors and among extracted factors (Brown, 2006). Standardized coefficients were used as effect sizes. In addition, the scale reliabilities of factors in CFMQ were evaluated by using internal reliability (i.e., Cronbach's alphas were tested here). Concurrent validity was analyzed by measuring the correlations of family functioning with the final factors.

Finally, we evaluated the correlations between the factor structures and health behavior outcomes (self-reported fruit and vegetable intake, added sugar intake, and levels of physical activity). To estimate unique effects, the current study specified four demographic variables as controls: family annual income (continuous), marital status (married vs. non-married),

parental education (continuous), and years in the US (more than 10 years vs. less than 10 years). Standardized path coefficients that assess the direct effect can be used as correlation coefficients (r) in structural equation modeling. Indirect effects can also be calculated with standardized estimates corresponding to effect-size estimates (Kline, 2011). As such, effect sizes were interpreted based on small ($r = .10$), medium ($r = .30$), and large effects ($r = .50$) (Cohen, 1988).

For the model evaluation, we used several fit indices including the comparative fit index (CFI; acceptable fit $>.90$; Little, 2013) and the root mean square error of approximation (RMSEA; acceptable fit $<.08$; Little, 2013). We used the same sample ($n = 280$) to use EFA and CFA. The average missing percent for all 35 items in CFMQ was 3.2% out of total sample size ($n=280$). Using Little's (1988) missing completely at random (MCAR) test, missing data patterns were analyzed to see whether any significant missing case patterns across 35 items of CFMQ. Chi-square tests showed no statistically reliable missing patterns for 35 items ($\chi^2[df] = 1432.563 [1550], p = .984$). To handle missing cases, we used a full-information maximum likelihood (FIML) estimator with robust standard errors (Enders & Bandalos, 2001), implemented as MLR in Mplus (version 8.00; (Muthen & Muthen, 1987–2017). The final analytic sample was 280.

RESULTS

Participants

The study sample consisted of 280 Hispanic overweight ($< 85^{\text{th}}$ percentile for age and sex) and obese ($> 95^{\text{th}}$ percentile for age and sex; (Barlow & Dietz, 1998)) 7th and 8th-grade youth and their primary caregivers recruited from middle schools in the MDCPS-S (Table 1). Forty-eight percent of the adolescents were male and 52% were female, and the mean age was 13.01 (range: 11–15; $SD=0.83$). Adolescents were mostly born in the U.S. (64%), Cuba (19.3%), Honduras (4.3%), and Venezuela (3.6%). The majority of adolescents not being born in the U.S. (66.1%) reported living in the USA for more than 10 years ($n=185$). Table 1 includes demographic data on parents as well as adolescents.

CFMQ Factor Structure—To examine the factor structure of CFMQ in our study sample, an exploratory factor analysis (EFA) was conducted with the reduced scale (35 items). The value of the Kaiser-Meyer-Olkin (KMO) measure of sampling adequacy was .82. In addition, Bartlett's test of sphericity yielded statistically significant results, $\chi^2 (595) = 3224.85, p < .001$. Taken together, these results indicate that the fit of the data to the factor model was acceptable. Next, to select the optimal factor structure, we conducted an exploratory factor analysis (EFA) with the 35 items of CFMQ. In Figure 1, the scree plot indicates that the eigenvalues for the first four factors of the exploratory factor analysis were larger than the corresponding parallel analysis (eigenvalues of EFA = 7.58, 4.27., 2.19., and 1.69.; eigenvalues of EFA = 1.86, 1.72., 1.64., and 1.58.), suggesting that the four extracted factors were reliable. However, within the four-factor structure, we found that 13 items contained low communality ($< .30$) (see corresponding items in Table 3). Therefore, after removing the 13 items that contained low communality, an additional EFA was conducted. The scree plot also suggests the same four factor model by showing the eigenvalues for the

first four factors of the exploratory factor analysis were larger than the corresponding parallel analysis (eigenvalues of EFA = 5.34, 3.72, 1.73, and 1.54, respectively; eigenvalues of EFA = 1.55, 1.45, 1.38, and 1.32, respectively). Therefore, a four-factor structure reduced to 22 CFMQ items was selected as the optimal factor structure and was used for subsequent analyses.

Table 3 shows the standardized loadings of items (only those items greater than .40 are reported). The total explained variance of these four factors was 56.61% (not shown in Table 2). The first factor included five items (> standardized loading .40) referring to family mealtime communication and explained 25.03% of the variance. The second factor comprised seven items (>.40 standardized loading) and for the most part, referred to family mealtime stress and explained 16.75% of the variance. The third factor consisted of five items (>.40 standardized loading) related to appearance weight control (i.e. the importance of weight management) and explained 8.03% of the variance. Finally, the fourth factor consisted of six items (>.40 standardized loading) and mostly referred to mealtime structure and explained 6.81% of the variance.

Construct Validity and Reliability of Factor Structure in CFMQ—Using a four-factor structure reduced to 22 CFMQ items, we again conducted a CFA to confirm the construct validity of the extracted four factors. The model had an acceptable fit (χ^2 (df) = 387.91 (186), $p < .001$; CFI / TLI = .91 / .90, RMSEA [90% CI] = .06 [.05, .07]). Similar to the findings from the EFA, standardized factor loadings between the reduced 22 items and their respective factors ranged from .45 to .82 ($p < .001$), indicating moderate-to-large correlation (r) effects. This suggested evidence of good construct validity for all four constructs in CFMQ. Among the four latent factors, most correlations have relatively small-to-medium effect sizes (ranged from $-.28$ between childhood mealtime communication and stress, $p < .001$ to $.39$, and between childhood mealtime communication and appearance weight control, $p < .001$). The correlation between appearance weight control and mealtime structures was moderate-to-high ($r = .48$, $p < .001$). In addition, we also evaluated scale reliability of each construct using Cronbach's alphas, which have acceptable ranges (alphas ranged from .73 to .82; see alphas in Table 3).

Concurrent Validity of CFMQ—Before investigating concurrent validity of CFMQ's four factors, we first conducted CFA to assess the construct validity of the family functioning variables (one latent variable capturing family communication, parent-adolescent communication, positive parenting, parental monitoring of peer, and parental involvement) (Lebron et al., 2018; Malcolm et al., 2013; Schwartz, Pantin, Prado, Sullivan, & Szapocznik, 2005). The model fit indices were acceptable (χ^2 (df) = 15.24 (5), $p = 1.13$; CFI / TLI = .97 / .94; RMSEA [90% CI] = .06 [.04, .08]). All standardized factor loadings ranged from .55 to .88 indicating good construct validity of the indicators.

The results of concurrent validity are shown in Figure 2. The model fit indices were acceptable (CFI = .90 / RMSEA [90% CI] = .05 [.04, .06]). After adjusting for the demographic characteristics (i.e. family income, marital status, parental education, years living in the US), two significant associations were found. That is, results showed that family functioning was positively associated with childhood mealtime communication ($\beta = .$

58, 95% CI = .47, .69, $p < .001$), but negatively associated with childhood mealtime stress ($\beta = -.32$, 95% CI = $-.46, -.19$, $p < .001$).

Correlations with Variables of Interest—The results of correlations between the newly identified factors and select obesogenic variables are shown in Figure 3. The model fit indices were acceptable (CFI = .90 / RMSEA [90% CI] = .05 [.04, .06]). Results show that childhood mealtime communication is positively associated with two healthy weight behaviors (physically active days: $\beta = .20$, 95% CI = .07, .32, $p < .01$; fruits and vegetable intake: $\beta = .29$, 95% CI = .15, .45, $p < .001$) but also an unhealthy weight behavior (added sugar intake: $\beta = .23$ (Lebron et al., 2018), 95% CI = .09, .37, $p < .001$). Interestingly, childhood mealtime stress was also positively associated with fruits and vegetable intake ($\beta = .26$, 95% CI = .08, .45, $p < .01$) and added sugar intake ($\beta = .38$, 95% CI = .21, .57, $p < .001$). There were no significant control variable effects in the estimated model. The revised version of the validated subscales are presented in the Appendix 1 & 2.

DISCUSSION

We identified the optimal factor structure of the CFMQ scales as applied to our sample of Hispanic overweight/obese youth using EFA with a subset of 35 CFMQ items and evaluated the construct validity and reliability of the CFMQ factor structure that are estimated in step 1 using CFA. Factor analyses produced a CFMQ four factor model, generating a measure with 22 items that assessed four domains within the family mealtime environment: family mealtime communication (5 items), family mealtime stress (7 items), appearance weight control (5 items), and mealtime structure (6 items). We also investigated the concurrent validity between the factor structures and the family functioning measure. Results showed that family functioning was positively associated with childhood mealtime communication, but negatively associated with childhood mealtime stress. Finally, we evaluated the correlations between the factor structures and obesogenic outcomes (self-reported fruit and vegetable intake, added sugar intake, and level of physical activity) and found that the family mealtime communication subscale was associated with all three obesogenic outcomes.

Factor Structure in CFMQ

Contrary to the 35 items (7 factors) result of the original CFMQ, our EFA and CFA produced a revised, abbreviated version that concluded with 23 items and consisted of the aforementioned four factors (family mealtime communication, family mealtime stress, appearance weight control, and mealtime structure). This version of the CFMQ does not support the inclusion of the subscales ‘Emphasis on Mother’s Weight’, ‘Present Parental Meal Influence’, and ‘Traditional Family.’ This may be attributed to the fact that the CFMQ was originally developed to assess how early mealtime experiences correlated with disturbed eating practices and was tested among bulimic women (Miller et al., 1993). However, the inclusion of 4 comparable factors highlights the similarities there are among the risk and protective factors of obesity and disordered eating, such as dieting and binge-eating, among adolescents and how family mealtime can influence weight-related conditions (Fagundo et al., 2012; Goldschmidt et al., 2008; Neumark-Sztainer et al., 2009; Thompson & Smolak, 2001). Previous studies have consistently documented disordered eating predicting further

weight gain (Marcus & Wiener, 1989; Neumark-Sztainer et al., 2012; Tanofsky-Kraff et al., 2009) and a substantial portion of adolescents seeking treatment for eating disorders have a history of overweight/obesity (Lebow, Sim, & Kransdorf, 2015). In fact, researchers in the fields of obesity and eating disorders have proposed an integrated approach to the co-occurrence of these health conditions and the shared risk factors for different eating- and weight-related problems (Irving & Neumark-Sztainer, 2002; Sanchez-Carracedo, Neumark-Sztainer, & Lopez-Guimera, 2012). Furthermore, several organizations including the American Academy of Pediatricians, the American Dietetic Association, and the Society for Adolescent Health and Medicine have pointed out the potential overlap of overweight/obesity and eating disorders and have encouraged physicians to conduct early screening of eating disorders regardless of weight status (Daniels & Hassink, 2015; Golden, Katzman, Sawyer, & Ornstein, 2015; Golden, Schneider, & Wood, 2016; Ozier & Henry, 2011).

The CFMQ was not originally developed for Hispanics or to identify obesogenic behaviors, but we chose it because the items specified around the general atmosphere and attitudes were applicable to our population. This analysis is an initial step to quantify and learn about these intricate relationships especially among our target population who demonstrate very high rates of unhealthy weight compared to other racial/ethnic groups. A better understanding of how family mealtime behaviors affect Hispanic adolescents can contribute to the established literature about culture and eating (Counihan & Van Esterik, 2012). For example, Kumanyika (2008) describes various ethnic and cultural influences of childhood obesity like how cultural acceptability of overeating has been conditioned to avoid hunger. Cultural wisdom dictates that one should “feast” when food is available. Note that a recent report administered by the USDA states that 1 in 4 Latino children live in a food insecure household (Coleman-Jensen, Rabbit, Gregory, & Singh, 2017).

Concurrent Validity of CFMQ

Conducting concurrent validity analysis of the CFMQ and family functioning highlights similarities in terms of (1) family mealtime communication and (2) mealtime stress. This might be explained by the fact that family functioning is treated here as a latent variable that captures, among multi-domain family factors, family communication, and parent-adolescent communication. Furthermore, others have reported an association between positive family mealtime communication and lower stress (Offer, 2013). Interestingly, no significant associations were found between family functioning and the subscales of appearance weight control or mealtime structure.

Correlations with Variables of Interest

Increased mealtime communication was correlated with three weight-related behavioral outcomes: physical activity, fruit and vegetable intake, and added sugar intake. Similarly, Fiese et al. (2012) found that families with children of healthy weight engaged in more positive interpersonal communication at mealtime than those with overweight children. Although the association to added sugar may seem to be an anomaly, this may be explained by the cultural relationship there is between family and food in Hispanics (Kaufman & Karpati, 2007; Kumanyika, 2008; Perez-Escamilla, 2009). For example, family celebrations are often and common in the Hispanic culture and these gatherings revolve around an

abundance of foods of all types. This places social pressure to participate in eating; nonattendance or food refusal is seen as rude and disrespectful (Castro, Shaibi, & Boehm-Smith, 2009; Weiler & Crist, 2009).

Interestingly, mealtime stress was also associated with fruit and vegetable intake and added sugar intake, suggesting that poor family functioning associated with higher stress during the mealtime results in higher dietary consumption of all foods regardless of their nutritional quality. Similarly, in a study among racially/ethnically diverse adolescents, 25% of the sample was using food as a coping mechanism for stress related to parents and self and romantic interests (MartynD Nemeth, Penckofer, Gulanick, Velsor-Friedrich, & Bryant, 2009). Furthermore, using food to cope was associated with higher body weight. In another study, the CFMQ was administered to college-aged women (Worobey, 2002a) and found that mealtime communication-based stress was actually highest among underweight woman compared to their overweight and normal weight counterparts. Conversely, another consideration is that better mealtime communication and lower mealtime stress may make mealtime more enjoyable for families and may be linked to higher consumption of all food types, simply given that families are eating together on a regular basis. Similarly, it has been demonstrated that eating fast food at family mealtime diminishes or negates the nutritional benefits associated with family mealtime (Martin-Biggers et al., 2014). Although these findings show a relationship between stress, family, and disordered eating, further research could determine how the specific subscales of the CFMQ can be used to predict how the family mealtime environment influences dietary outcomes over time.

Strengths and Limitations

This revised, abbreviated CFMQ was developed for researchers interested in understanding the mealtime environment in families with adolescents. However, it is important to recognize the limitations of the questionnaire. It does not include the question of mealtime frequency that has been cited in the literature (Hammons & Fiese, 2011). Additionally, it was not originally intended for obesogenic populations, but, as has been demonstrated in the literature, obesity and eating disorders share many of the same risk and protective factors (Haines & Neumark-Sztainer, 2006; D. R. Neumark-Sztainer et al., 2007). A further limitation of the questionnaire is that it is subject to social desirability bias and it is not known whether participants differed in their interpretation or understanding of the questions. The entire study sample consisted of overweight/obese adolescents which is likely to skew the measurement of obesogenic behaviors. Lastly, although concurrent validity was assessed with all four subscales, family functioning may not be the best measure to use as a comparison measure with appearance weight control or mealtime structure. However, this study has several methodological strengths. First, participants were recruited from a large range of Hispanic origins increasing the diversity and representativeness of the sample. Second, we employed rigorous exploratory and confirmatory factor analytical approaches to ascertain the inclusion of every individual item. Lastly, the validation of the CFMQ has been shown to be good, resulting in a valid measure that is suitable for use in future research.

CONCLUSION

In sum, the findings of this analysis provide a validated, abbreviated CFMQ as an adolescent measure for the report of family mealtime environment in overweight/obese Hispanics. Further use and evaluation of the questionnaire is needed, but these promising findings of reliability and validity suggest that researchers interested in understanding the adolescent perspective of the family mealtime environment, especially among Hispanics, may find this version of the CFMQ useful. Indeed, the literature focused on mealtime frequency may be overlooking these types of relationships, especially among adolescents of unhealthy weight who have adopted poor relationships with food. Our study supports a role for further research to understand individuals' and families' environments, beliefs, and sociocultural influences mealtime in order to foster the development of relevant interventions across diverse populations.

Appendix 1: Mealtime Communication from CFMQ (English & Spanish)

Instructions: Think back to family mealtimes during your childhood (up to the age of 13). Please select the answer to each question according to how you felt at that time.

Mealtime Communication	Never	Rarely	Sometimes	Usually	Always
I felt mealtimes were a warm and sharing time in my family.					
My family talked during dinner.					
I liked to eat dinner with my family.					
In my family, everyone could speak their views at dinner time.					
I felt able to speak my mind during mealtimes.					

Instrucciones: Trate de recordar sus cenas familiares (de niñez hasta los 13 años) y por favor escoge la respuesta de acuerdo a como se sintió en esos momentos.

Comunicación a la hora de comer	Nunca	Raramente	A veces	En general	Siempre
Sentía que las cenas familiares eran placenteras y eran un momento agradable para compartir con la familia.					
Durante la hora de la cena, mi familia conversaba.					
Me gustaba cenar con mi familia.					
A la hora de la cena, todos podíamos compartir nuestros puntos de vista.					
Durante la hora de la cena me sentía suficientemente cómodo(a) de compartir mis puntos de vista.					

Appendix 2: Mealtime Stress from CFMQ (English & Spanish)

Instructions: Think back to family mealtimes during your childhood (up to the age of 13). Please select the answer to each question according to how you felt at that time.

Mealtime Stress

Never Rarely Sometimes Usually Always

I remember feeling nervous during dinner.
 Dinner times were silent in my family.
 Because of stress during meals, I liked or wished to eat alone.
 There was yelling during dinner.
 It was a relief when my father was not at dinner.
 My father commented about my mother's weight when I was young.
 In my family, mealtimes were set by my father's schedule.

Instrucciones: Trate de recordar sus cenas familiares (de niñez hasta los 13 años) y por favor escoge la respuesta de acuerdo a como se sintió en esos momentos.

Estrés a la hora de comer

Nunca Raramente A veces En general Siempre

Recuerdo sentirme nervioso(a) durante la hora de la cena
 En mi familia, la hora de la cena se pasaba en silencio.
 Prefería comer solo(a) para evitar el estrés de cenar en familia.
 Durante la hora de la cena, se gritaba.
 Era un alivio cuando mi padre no estaba presente en la hora de la cena.
 De joven, mi padre hacía comentarios sobre el peso de mi madre.
 El horario de mi padre establecía la hora de la cena para mi familia.

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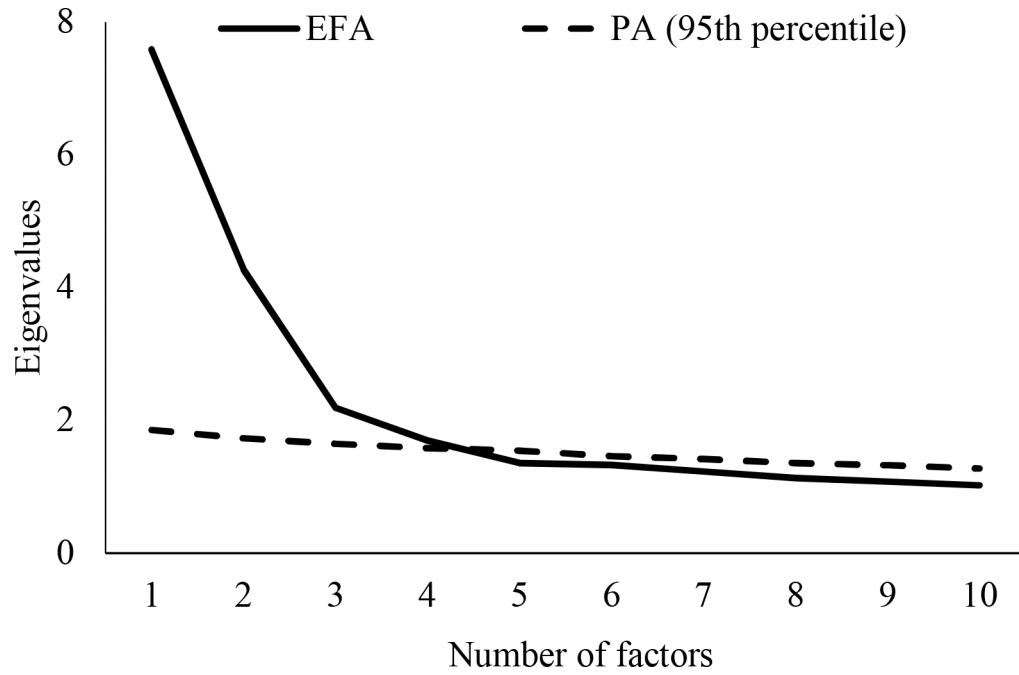
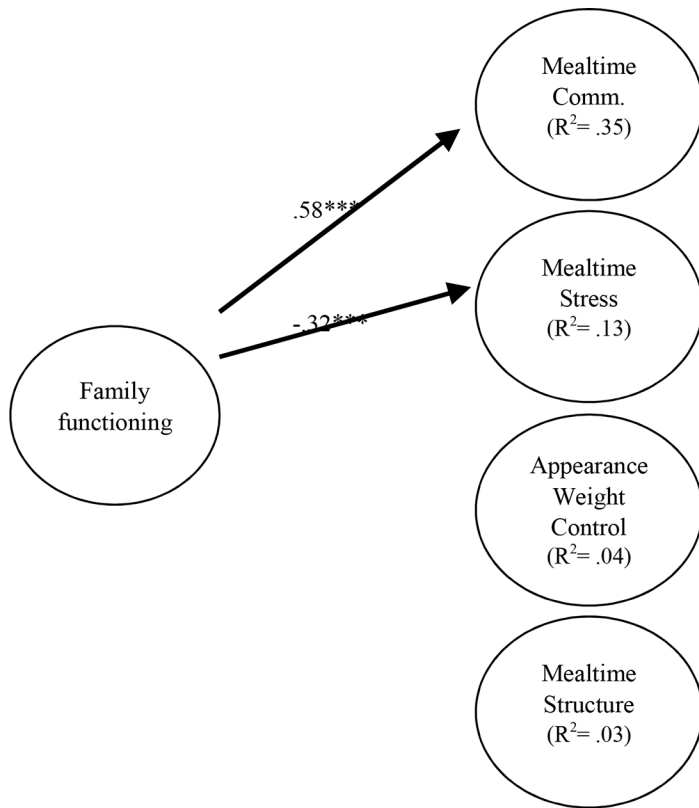


Figure 1. Scree plot of CFMQ exploratory factor analysis.

Note. CFMQ = Childhood Family Mealtime Questionnaire. EFA = Exploratory Factor Analysis. PA = Parallel Analysis.



Controls:

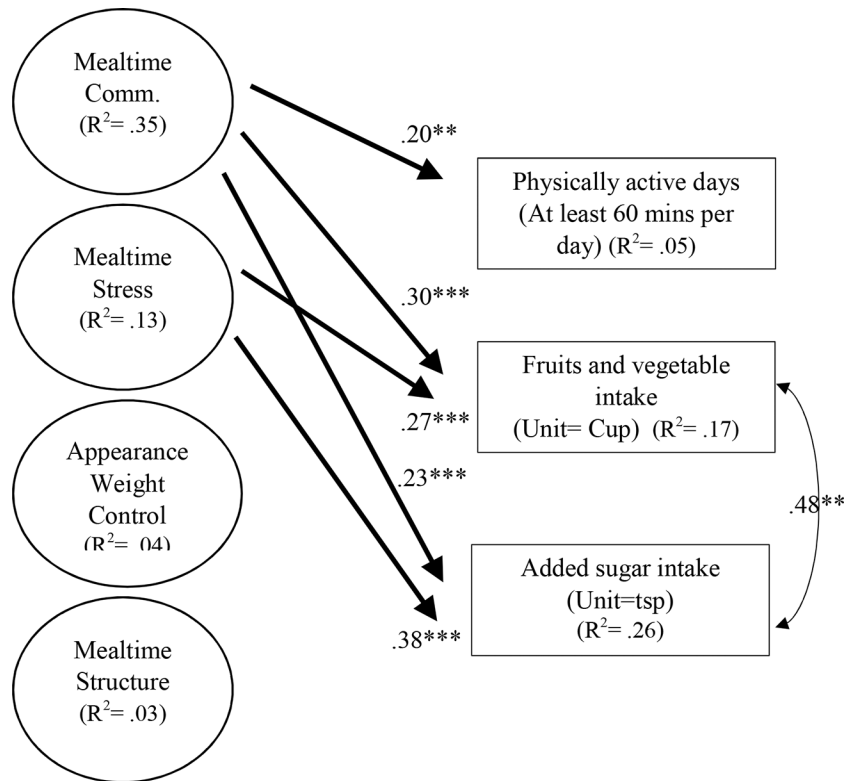
Family income (continuous),
 Martial status (Married vs. Non-married),
 Parental education (continuous),
 Years in US (More than 10 years vs. less than 10 years)

Figure 2. Concurrent validity of four factors in CFMQ.

Notes. Standardized coefficients were shown. Only significant standardized coefficients were shown in the figure.

Comm = Communication. Non-significant coefficients were not shown in figure. CFI = .90. RMSEA=.05.

** $p < .01$. *** $p < .001$.



Controls:

Family income (continuous),
 Martial status (Married vs. Non-married),
 Parental education (continuous),
 Years in US (More than 10 years vs. less than 10)

Figure 3. Predictive validity of four factors in CFMQ.

Notes. Standardized coefficients were shown. Only significant standardized coefficients were shown in the figure.

Comm = Communication. Non-significant coefficients were not shown in figure. CFI = .90. RMSEA=.05.

** $p < .01$. *** $p < .001$.

Table 1.

Sociodemographic Characteristics of Parent & Adolescent Participants

Variable	Mean (SD) or %
<i>Adolescent</i>	
Female	52%
Age	13.01 (11–15; 0.83)
<i>Country of Origin</i>	
United States	64.29%
Cuba	19.29%
Honduras	4.29%
Venezuela	3.47%
BMI	28.06 (6.07)
Obesity (> 30 BMI)	23.92%
Percentile *	94.63 (4.08)
<i>Parent</i>	
Female	88.21%
Age	41.88 (25–59; 6.0)
<i>Country of Origin</i>	
United States	8.93%
Cuba	34.29%
Nicaragua	15.00%
Honduras	11.43%
<i>Marital Status</i>	
Married	57.86%
Divorced	12.9%
Living with someone	10.00%
Separated	10.00%
Never married, not living with someone	8.47%
Widowed	0.71%
Full- time employment	50.71%
<i>Annual Income</i>	
Less than \$30,000	65.41%
Greater than \$30,000	20.68%
Greater than \$50,000	13.91%
<i>BMI</i>	
Obesity (> 30 BMI)	47.14%

* adjusted for age and sex

Table 2.Community (h^2) of full 35 items from initial EFA model

Items	h^2
I felt mealtimes were a warm and sharing time in my family.	.40
My family talked during dinner.	.31
<i>I looked forward to mealtimes.</i>	<i>.26</i>
I liked to eat dinner with my family.	.51
In my family, everyone could speak their views at dinner time.	.72
I remember feeling nervous during dinner.	.54
Dinner times were silent in my family.	.36
I felt able to speak my mind during mealtimes.	.50
Because of stress during meals, I liked or wished to eat alone.	.36
There was yelling during dinner.	.63
It was a relief when my father was not at dinner.	.34
<i>My family was conscious of wasting food.</i>	<i>.15</i>
During meals, I was told not to waste food.	.33
If I did not like what we were having for dinner, I had to eat it anyway.	.42
<i>At home, I had to clean my plate (i.e. eat all the food on it).</i>	<i>.21</i>
<i>My parents made me eat foods I did not like.</i>	<i>.27</i>
Table manners were important to my parents.	.48
Table manners were brought up at dinner.	.29
I remember thinking about my weight when I was young.	.71
I remember worrying about my weight when I was young.	.77
When I was young, I was encouraged to diet.	.45
<i>In my family, we thought and/or talked about physical appearance.</i>	<i>.29</i>
In my family, we talked about our own or each other's weight.	.32
In my family, we thought of beauty as depending a lot on weight.	.37
<i>We ate foods my father liked.</i>	<i>.14</i>
<i>During meals, you could tell who was in control in my family</i>	<i>.20</i>
<i>When my family wanted to celebrate, food was a part of the celebration.</i>	<i>.14</i>
<i>My mother worried about her weight when I was young.</i>	<i>.28</i>
<i>My mother dieted when I was young.</i>	<i>.23</i>
My father commented about my mother's weight when I was young.	.51
Presently, when I am at home, my parents influence what I eat.	.47
Presently, when I am at home, my parents influence the way I eat.	.36
<i>I saw my mother's main role as that of a homemaker.</i>	<i>.23</i>
<i>I saw one of my mother's main roles as that of a cook.</i>	<i>.19</i>
In my family, mealtimes were set by my father's schedule.	.33

Note. h^2 = Communality, reflecting the amount of variance in each variable explained by the extracted factors. Grayed and italic items (total 6 items) were low communality items (< .30), which we excluded from subsequent analyses.

Structure coefficients, means, and standard deviations of reduced CFMQ 23 items in the extracted four factor model in a sample of 280 overweight/obese Hispanic adolescents,

Table 3.

	Structural coefficients					M (SD)
	F1	F2	F3	F4	R ²	
Reduced 23 items	F1	F2	F3	F4	R²	M (SD)
I felt mealtimes were a warm and sharing time in my family.	.59				.37	3.86 (1.18)
My family talked during dinner.	.54				.31	3.91 (1.17)
I liked to eat dinner with my family.	.71				.52	4.21 (1.06)
In my family, everyone could speak their views at dinner time.	.86				.75	3.89 (1.26)
I remember feeling nervous during dinner.		.74			.56	1.71 (1.10)
Dinner times were silent in my family.		.59			.41	2.06 (1.20)
I felt able to speak my mind during mealtimes.	.70				.50	3.50 (1.25)
Because of stress during meals, I liked or wished to eat alone.		.57			.37	2.32 (1.37)
There was yelling during dinner.		.77			.63	1.72 (1.07)
It was a relief when my father was not at dinner.		.58			.35	1.70 (1.17)
During meals, I was told not to waste food.			.53		.31	3.61 (1.26)
If I did not like what we were having for dinner, I had to eat it anyway.			.45		.35	2.67 (1.29)
Table manners were important to my parents.			.71		.56	3.89 (1.24)
Table manners were brought up at dinner.			.55		.33	2.94 (1.43)
I remember thinking about my weight when I was young.			.86		.75	3.01 (1.36)
I remember worrying about my weight when I was young.			.87		.77	3.04 (1.41)
When I was young, I was encouraged to diet.			.62		.41	2.71 (1.39)
In my family, we talked about our own or each other's weight.			.41		.35	2.41 (1.29)
In my family, we thought of beauty as depending a lot on weight.			.48		.33	2.40 (1.29)
My father commented about my mother's weight when I was young.		.75			.57	1.76 (1.15)
Presently, when I am at home, my parents influence what I eat.				.73	.54	3.43 (1.26)
Presently, when I am at home, my parents influence the way I eat.				.56	.36	3.01 (1.30)
In my family, mealtimes were set by my father's schedule.		.45			.30	2.04 (1.21)
Eigenvalues	5.54	3.73	1.78	1.56		
% of explained variance	25.15	16.04	7.82	6.63		
Cronbach's alpha	.81	.82	.81	.76		

Note. F = Factor. R^2 = Communality. M = Mean. SD = Standard Deviation. Structural coefficients are standardized. Only standardized loadings over .40 are listed.

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