



Response to Guo *et al.*'s Letter to the Editor

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We thank Drs Guo, Gao, Niu, and Zhang for their comments on the article by [Fong and others \(2018\)](#). They pointed out that the more classical methods, MW-MW₂ and SR-MW₂, which only make comparisons between \mathcal{X} and \mathcal{Y} (paired observations) and between \mathcal{X}' and \mathcal{Y}' (unpaired observations) were useful alternatives to the proposed tests, MW-MW₀^l and SR-MW₀^l, which made comparisons between all x 's and all y 's. Dr Guo *et al.*'s recommendation was “to use MW-MW₀^l and SR-MW₀^l for $\{\mathcal{X}, \mathcal{Y}, \mathcal{Y}'\}$, while use MW-MW₂ and SR-MW₂ for $\{\mathcal{X}, \mathcal{Y}, \mathcal{X}', \mathcal{Y}'\}$, especially when the correlation between the samples is high.” We agree that MW-MW₂ and SR-MW₂ are important to study as alternative approaches, and aim to refine the recommendations in this response so that practitioners may find it easier to choose the appropriate methods.

Before discussing power comparison, we would like to propose a variant of the MW-MW₂ test. Since MW-MW₂ only makes comparisons within the paired subset and the unpaired subset, it is possible to perform permutation tests to obtain p -values to avoid inflated Type 1 error rates under small sample sizes (Tables A.1–A.4 of the [supplementary material](#) available at *Biostatistics* online). We will refer to this test as MW-MW₂^{perm}.

We study power comparison under four different distributional assumptions: normal (Table 1), logistic (Table B.1 of the [supplementary material](#) available at *Biostatistics* online), gamma (Table B.2 of the [supplementary material](#) available at *Biostatistics* online), and lognormal (Table B.3 of the [supplementary material](#) available at *Biostatistics* online). We also plot the results in Figure 1 and Figures B.1, B.2, and B.3 of the [supplementary material](#) available at *Biostatistics* online to help visualize these results. All estimates are based on 10⁴ Monte Carlo replicates. m, l, n refer to the number of pairs, the number of independent x 's and the number of independent y 's, respectively. Three levels of correlation between the two samples are examined: 0, 0.5, and 0.8.

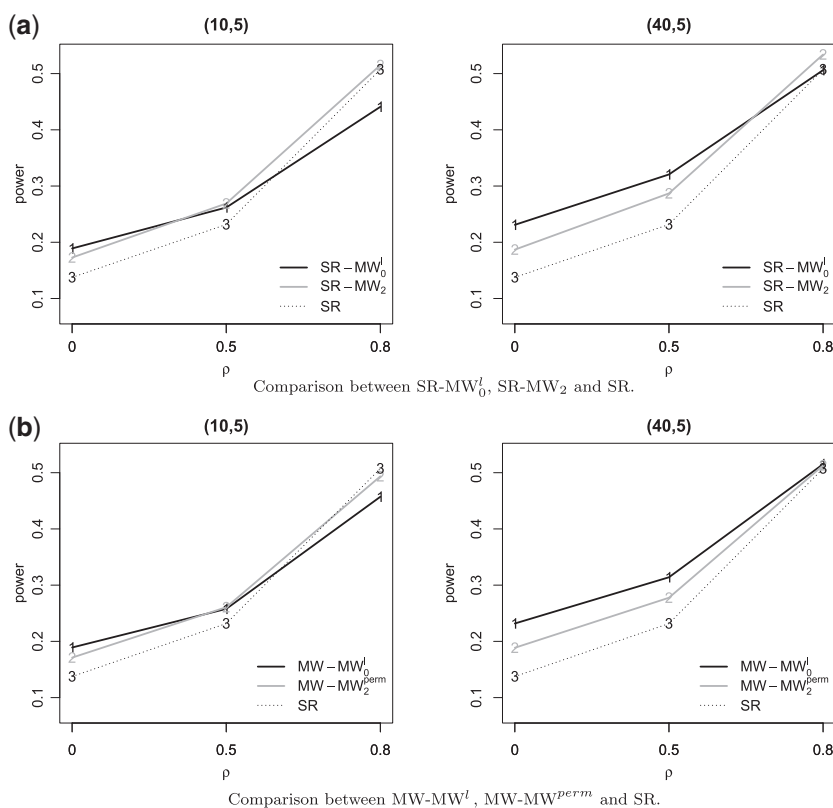
First, focusing on lines 2 and 3 in the figures, we see that SR-MW₂ and MW-MW₂^{perm} either outperform or closely match the performance of SR at all times. These empirical results are worth noting, because theoretically a test that combines two independent test statistics using weights proportional to the inverse of their variances is not always more powerful than each component test. Based on these results, we can narrow the choice down to be between SR-MW₀^l/MW-MW₀^l and SR-MW₂/MW-MW₂^{perm} when there are unpaired observations from both samples.

Now, focusing on lines 1 and 2 in the figures, we see that there is a clear trade-off between SR-MW₀^l/MW-MW₀^l and SR-MW₂/MW-MW₂^{perm} depending on ρ and sample sizes. This is true for normal,

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Table 1. *Estimated power, normal distribution, m = 20*

(l, n)	MW-MW ₀ ^l			MW-MW ₂ ^{perm}			SR			SR-MW ₂			SR-MW ₀ ^l		
	0	0.5	0.8	0	0.5	0.8	0	0.5	0.8	0	0.5	0.8	0	0.5	0.8
(10,5)	19	26	46	17	26	49	14	23	51	17	27	52	19	26	44
(10,10)	20	28	47	18	27	51	14	23	51	19	29	53	20	28	46
(40,5)	23	31	52	19	28	51	14	23	51	19	29	53	23	32	51

Fig. 1. Power comparison when the marginal distribution is normal. Sample sizes: $m = 20$ and (l, n) are given in the titles.

test. The practical implication of this observation is that we should preprocess the data by applying proper transformation if the distributions appear highly skewed.

Our recommendation for the case when there are unpaired observations from both samples has two parts. If a simple rule of thumb is desirable, our recommendation is to choose SR-MW₀^l/MW-MW₀^l when $\rho < 0.5$ and SR-MW₂/MW-MW₂^{perm} when $\rho > 0.5$. On the other hand, if an optimal choice is important, we recommend doing a simulation study to find the most powerful approach. To make this a feasible option for practitioners, we provide an easy-to-use function, *choose.test*, in the R package *chnppt*. The only information the function needs is the sample sizes and the estimated first and second moments from

the data, and it is fast, for example, it takes only 2 s to run on an Intel i7 processor clocked at 2.6GHz when $m = 20, l = 40, n = 5$.

For the case when there are only unpaired observations from one sample (thus SR-MW₂/MW-MW₂^{perm} are not applicable), we recommend choosing between SR and SR-MW₀^l/MW-MW₀^l through the *choose.test* function, since there is a trade-off in power between the two tests depending on ρ and sample sizes (Tables D.1–D.3 of the [supplementary material](#) available at *Biostatistics* online).

Lastly, given the choice between SR-MW₀^l and MW-MW₀^l, we recommend SR-MW₀^l if a monotone transformation can be performed on both samples so that the distributions from both samples are not too skewed. If that is not possible or desirable, for example, when one sample has a highly skewed distribution while the other does not, MW-MW₀^l is preferred because it is a more robust test and invariant to monotone transformations applied to both samples. When using MW-MW₀^l, one should proceed with caution as Type 1 error rates may be inflated when sample sizes are small (Tables D.4–D.6 of the [supplementary material](#) available at *Biostatistics* online). Similar arguments can be applied to the choice between SR-MW₂ and MW-MW₂^{perm}, except that there is no concern of inflated Type 1 error rates here.

The *chnopt* package is available from the Comprehensive R Archive Network, and the Monte Carlo study code can be downloaded at https://github.com/youyifong/response_to_letter_on_rank.

SUPPLEMENTARY MATERIAL

[Supplementary material](http://biostatistics.oxfordjournals.org) is available at <http://biostatistics.oxfordjournals.org>.

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