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Clopidogrel Drug Interactions and Serious Bleeding: Generating Real World Evidence via Automated High-Throughput Pharmacoepidemiologic Screening

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Abstract

Few population-based studies have examined bleeding associated with clopidogrel drug-drug interactions (DDIs). We sought to identify precipitant drugs taken concomitantly with clopidogrel (an object drug) that increased serious bleeding rates. We screened 2000–2015 Optum commercial health insurance claims to identify DDI signals. We performed self-controlled case series studies for clopidogrel + precipitant pairs, examining associations with gastrointestinal bleeding or intracranial hemorrhage. To distinguish native bleeding effects of a precipitant, we reexamined associations using pravastatin as a negative control object drug. Among 431 analyses, 28 clopidogrel + precipitant pairs were statistically significantly positively associated with serious

Statement of Integrity

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Dr. Leonard had full access to study data and takes responsibility for its integrity and that of the data analysis.

bleeding. Ratios of rate ratios ranged from 1.13–3.94. Among these pairs, 13 were expected given precipitant drugs alone increased and/or were harbingers of serious bleeding. The remaining 15 pairs constituted new DDI signals, none of which are currently listed in two major DDI knowledge bases.

Keywords

Clopidogrel; drug interactions; pharmacoepidemiology; pharmacovigilance

INTRODUCTION

Clopidogrel is a widely-used antiplatelet drug, prescribed or continued during ~19 million ambulatory care office visits in the United States in 2015.¹ Serious bleeding is the most prominent adverse effect of clopidogrel, occurring at an incidence of 5.6–15.7% per patient-year.² While several studies have examined drug-drug interactions (DDIs) that may interfere with clopidogrel's effectiveness in preventing thrombosis,³ few studies have examined drugs that may potentiate clopidogrel's bleeding risk, which may be of great clinical and public health importance. Because potentiation of clopidogrel's bleeding risk might occur through many known or unknown actions of concomitant medications and thus be difficult to predict, and because spontaneous reporting databases may be ineffective in identifying an increased risk of an expected adverse event (e.g., bleeding in persons receiving an antiplatelet drug),⁴ we aimed to conduct high-throughput pharmacoepidemiologic screening to identify signals of potentially clinically-significant DDIs involving clopidogrel that increase the rate of serious bleeding so that these potential DDIs can be confirmed and elucidated in future studies.

RESULTS

Characteristics of persons constituting the study cohorts are summarized in Table 1. The clopidogrel cohort (all of whom experienced serious bleeding while on clopidogrel) consisted of 21,577 persons. The plurality (26.7%) were female Caucasian older adults. The pravastatin cohort consisted of 5,785 persons experiencing serious bleeding while on pravastatin. The plurality (28.5%) were also female Caucasian older adults. In both cohorts, substantial proportions of individuals had diagnosed cerebrovascular and/or ischemic heart disease, consistent with indications for clopidogrel and pravastatin. More than half of individuals had a prior serious bleed. Very few observation days included concomitant exposure to anticoagulants, other antiplatelet agents, or nonsteroidal anti-inflammatory drugs (NSAIDs). In contrast, approximately one-quarter of observation days included concomitant exposure to a gastro-protective agent such as a histamine-2 receptor antagonist or proton pump inhibitor.

We identified 536 and 436 precipitant drugs frequently co-prescribed with clopidogrel and with pravastatin, respectively. We conducted a self-controlled case series study for each of these object-precipitant pairs. The object drug cohorts had 431 precipitant drugs in common. Summary data on unadjusted, adjusted, and semi-Bayes adjusted rate ratios for gastrointestinal bleeding/intracranial hemorrhage are presented in Table 2. See Figure S1 for

a heat map depicting overall and risk period-specific semi-Bayes adjusted rate ratios and ratios of rate ratios. The secondary analysis permitting a wider range of true rate ratios during semi-Bayes shrinkage had similar findings (Figure S2).

When using pravastatin as the quantitative negative control object drug, 37 of the 431 candidate DDIs had statistically significantly elevated ratios of rate ratios for serious bleeding when examined over the entire observation period. We excluded nine candidates solely driven by protective findings among pravastatin users, i.e., the rate ratio for the precipitant drug was not statistically significantly elevated in users of clopidogrel but was statistically significantly protective in users of pravastatin, as these unlikely indicated a clopidogrel DDI. Twenty-eight remained (6.5% of all candidate DDIs) and were deemed potential signals of DDIs involving clopidogrel that lead to serious bleeding. Interacting precipitants included central nervous system agents (N = 12, including four NSAIDs and two opioids), cardiovascular agents (N = 3), endocrine/metabolic agents (N = 3), renal/ genitourinary agents (N = 3, including two phosphate binders), nutritional agents (N = 3, including two vitamins), hematologic agents (N = 2, both anticoagulants), an anti-infective (N = 1), and a gastrointestinal agent (N = 1). Rate ratios ranged from 1.13 (95% confidence interval [CI]: 1.01–1.26) for potassium chloride to 3.94 (1.69–9.20) for cholecalciferol. Six (21.4%) of these 28 potential DDI signals are each listed in clopidogrel's label, Micromedex, and Lexicomp as potentially increasing bleeding risk.

DISCUSSION

This pharmacoepidemiologic screening study of potential DDIs involving clopidogrel leading to serious bleeding yielded expected results for oral anticoagulants and NSAIDs, although not all of the semi-Bayes adjusted association measures were statistically significant. For anticoagulants, concomitant use of clopidogrel (vs. pravastatin) with warfarin or dabigatran was associated with statistically significant 1.3-fold and 1.2-fold rates of serious bleeding respectively, while the 1.3- and 1.5-fold rates for apixaban and rivaroxaban were not statistically significant. DDIs between clopidogrel and oral anticoagulants would be expected mechanistically given the independent cumulative effects of antiplatelet agents and anticoagulants on hemostasis,^{5,6} and the magnitude of these associations were generally consistent with prior epidemiologic studies.^{2,7,8} Concomitant use of clopidogrel (vs. pravastatin) with the NSAIDs meloxicam, piroxicam, nabumetone, and etodolac was associated with statistically significant 1.6-, 2.5-, 2.8-, and 3.2-fold rates (respectively) of serious bleeding, while the 1.1-, 1.8-, 1.9-, 2.8-, 2.8-, and 4.4-fold rates associated with naproxen, diclofenac, ibuprofen, oxaprozin, indomethacin, and sulindac (respectively) were not statistically significant. These findings are mechanistically plausible given independent effects of clopidogrel and NSAIDs on bleeding risk,⁹ and also generally consistent with prior epidemiologic studies. $^{9-11}$ We believe that the consistency of the results for oral anticoagulants and NSAIDs with both mechanistic expectations and prior epidemiologic findings support the validity of our screening approach, and that the lack of statistical significance for some expected pairs may suggest that the assumptions we employed in our semi-Bayes adjustment were appropriately conservative for use in a hypothesis-generating screening context.

We also identified increased rates of serious bleeding upon concomitant use of clopidogrel with several drugs that are commonly used in persons with chronic kidney disease (CKD): bumetanide; cholecalciferol; cinacalcet; lanthanum; paricalcitol; sevelamer; and torsemide. Use of these precipitants may portend worsening renal function and/or dialysis. Advanced CKD and end stage kidney disease increase one's risk of bleeding due to platelet dysfunction¹² and heightened risk of angioectasias (i.e., thin-walled, dilated, ectatic blood vessels).¹³ which may be exacerbated in the setting of clopidogrel treatment. 5,14,15 Cinacalcet, lanthanum, paricalcitol, and sevelamer are used almost exclusively in the treatment of abhorrent bone and mineral metabolism in persons with advanced CKD and end stage kidney disease. These disorders of bone and mineral metabolism can lead to gastrointestinal mucosal calcinosis and calciphylaxis,¹⁶ which may increase the risk of bleeding. Several case reports also describe deposition of sevelamer crystals in the gastrointestinal mucosa, directly resulting in ulceration and necrosis.^{17,18} It is unclear whether findings for these precipitants likely reflect confounding by indication or important DDIs that may place these patients at greater than expected risk of bleeding. Thus, the risk of clopidogrel-associated DDIs in advanced renal disease is an important area for future investigation.

The remaining 15 pairs associated with serious bleeding were deemed potentially clinicallyrelevant DDI signals (Table 3). These have neither been described in published case reports nor examined in population-based studies. It is therefore understandable that none are listed as an interaction (leading to a serious bleed) in clopidogrel's label, Micromedex, or Lexicomp. Our automated screening approach did not consider preexisting mechanistic knowledge, which may be a poor predictor of clinically-important DDIs because of incomplete knowledge of off-target drug effects, failure to identify complex multi-pathway interactions,¹⁹ and traditional over-reliance on commonly considered mechanisms, most notably cytochrome P450 (CYP) inhibition. This is exemplified by the fact that some of our DDI signals may have identifiable putative mechanisms (e.g., primidone induces CYP2C19 and other hepatic isozymes²⁰), while others do not. Future work should seek to elucidate mechanisms underlying these signals.

Our study has notable strengths. First, it utilized a self-controlled case series design, ideal for DDI screening,²¹ to minimize confounding. Second, we used a bi-directional implementation of the design to minimize exposure trend bias.²² Third, we utilized pravastatin as a negative control object drug to which clopidogrel findings were quantitatively compared. Fourth, we studied a clinically meaningful outcome identified by algorithms with excellent performance metrics. Finally, we minimized false positive findings by using semi-Bayes shrinkage to account for multiple estimation.

Our study also has limitations. First, it did not examine higher order (i.e., beyond pairwise) DDIs. Such findings may be of future interest given bleeding rates with dual antiplatelet and triple antithrombotic therapies reported in ISAR-TRIPLE, PIONEER, REDUAL, and WOEST and under investigation in AUGUSTUS and ENTRUST, as examples. Second, because clopidogrel + precipitant pairs *and* pravastatin + precipitant pairs were required for the parameter of interest, candidate DDI signals were identified among the intersection of concomitantly used drugs identified for both objects. This prohibited us from examining

ratios of rate ratios for ~19% of precipitant drugs concomitantly prescribed with clopidogrel, but not pravastatin. Third, we did not examine time-invariant covariates as potential effect modifiers. Fourth, the bi-directional self-controlled case series design may be susceptible to reverse causality, especially for suspected DDIs. If a clinician posited that a precipitant induced a serious bleed in an object drug user (even if it had no effect on the bleeding rate), the precipitant may be subsequently discontinued. This may result in a spuriously elevated rate ratio for that precipitant. However, it seems unlikely to us that reverse causality is responsible for associations with newly-identified DDI signals because: a) DDIs are often overlooked in clinical practice and therefore clinicians would unlikely attribute a serious bleed to an interaction and discontinue the precipitant to reduce future risk; b) such precipitant discontinuation would only have the potential to cause bias if differential among users of clopidogrel and pravastatin; and c) a post hoc analysis employing a right-censored uni-directional self-controlled case series design (resistant to reverse causality, but vulnerable to exposure trend bias) replicated the signals described herein (Table S1). Fifth, our reliance on a prescription dispensing as a surrogate for drug consumption and inability to assess adherence raise concerns of exposure misclassification. Sixth, residual confounding may be present; we did not adjust for precipitant drug dose, severity of chronic diseases, frailty, or socioeconomic status-factors not always static throughout an individual's observation. Finally, our findings may not be generalizable beyond a commercially-insured, ambulatory care population.

We used longitudinal health insurance data to identify 15 previously undescribed and/or unappreciated clopidogrel DDIs associated with serious bleeding. Vigilance during clopidogrel prescribing is warranted, since these potentially clinically-relevant interactions are not documented in two major DDI knowledge bases.

METHODS

Overview

We conducted automated, high-throughput pharmacoepidemiologic screening of commercial health insurance claims to identify signals of DDIs with clopidogrel. First, we identified drugs that were frequently co-prescribed with clopidogrel as candidate interacting precipitants. Second, we identified DDI signals by performing confounder-adjusted self-controlled case series studies for clopidogrel + precipitant (i.e., interacting drug) pairs, with hospital presentation for serious bleeding as the study outcome. To help distinguish native bleeding effects of a precipitant drug from a DDI involving clopidogrel, we repeated these steps for pravastatin, which served as a quantitative comparator (i.e., negative control object drug).²³ Pravastatin was selected because it is a widely-used cardiovascular drug that does not affect the risk of serious bleeding,²⁴ minimally inhibits human carboxylesterase 1,²⁵ and lacks substantive CYP-based effects²⁶ that could affect other drugs' bleeding risk.

Data source

We used 2000–2015 data from the Optum Clinformatics Data Mart (OptumInsight: Eden Prairie, MN, United States).²⁷ Optum includes enrollment and healthcare billing data from >71 million commercially-insured and Medicare Advantage beneficiaries of a large United

States-based insurer. Data elements include: demographics (e.g., age, sex, race); enrollment periods; medical encounters (e.g., ambulatory care visits, emergency department visits, inpatient hospitalizations) and their accompanying diagnoses and procedures; pharmacy dispensings; and laboratory orders and results. We selected Optum as our data source because of its generalizability to the United States population, as ~65% of Americans receive healthcare coverage via commercial health plans or Medicare.²⁸ The University of Pennsylvania's Office of Regulatory Affairs determined that research using Optum was exempted from institutional review board review.

Identifying candidate interacting precipitant drugs

We used pharmacy claim dates and days' supply values to identify all pharmacy dispensings of clopidogrel (the object drug of interest) and all oral drugs concomitantly used with clopidogrel (precipitant drugs of interest). This step was repeated for pravastatin, as it served as the negative control object drug.²³ The intersection of drugs identified for clopidogrel and pravastatin served as candidate precipitant drugs. We selected the intersection rather than union of drugs because measures of association for clopidogrel + precipitant pairs *and* pravastatin + precipitant pairs were required for the parameter of interest (described later).

Identifying DDI signals via automated pharmacoepidemiologic screening

For each object + candidate precipitant pair, we conducted a bi-directional self-controlled case series study to examine the rate of serious bleeding for an individual receiving the object drug of interest after initiating vs. not receiving a precipitant. Although the "case series" phrase within self-controlled case series may seem to imply the absence of a comparator, the approach is a rigorous, controlled epidemiologic study design that is the cohort analogue of the case-crossover design.²² The self-controlled case series design has the following advantages that make it ideal for DDI screening: 1) it is highly computationally-efficient,²⁹ since it includes only persons who experienced the outcome of interest; 2) the causal contrast is made *within* individual and thus inherently controls for confounding by both measured and unmeasured factors that remain constant within an individual over the observation period (e.g., sex, genetics, chronic diseases, frailty, socioeconomic status); 3) the underlying statistical model can accommodate time-varying factors;³⁰ and 4) a high-throughput approach has been developed and used previously.²¹

Creating study cohorts of new users of object drugs

Separate cohorts were constructed for clopidogrel and pravastatin. For persons 18–90 years of age, we utilized pharmacy claim dates and days' supply values to build object drug exposure episodes consisting of 1 object drug dispensing(s). We permitted a 7-day grace period between contiguous dispensings and at the end of the terminal dispensing to account for imperfect adherence; this approximated 80% adherence for each 30-day dispensing. We then selected new users of the object drug by requiring a baseline period (defined below) that was devoid of a dispensing for that object or a therapeutic alternative (i.e., anagrelide, cangrelor, cilostazol, dipyridamole, prasugrel, ticagrelor, ticlopidine, or vorapaxar for clopidogrel, and other 3-hydroxy-3-methylglutaryl-CoA reductase inhibitors for pravastatin). Since the self-controlled case series design is a "case-only" approach, we also required new users to experience a serious bleeding event (defined below) during their observation period.

Defining observation and baseline periods

For each new user meeting inclusion criteria, the observation period began upon object drug initiation and was censored upon the earliest of: a) lapsed exposure to the object (accounting for the 7-day grace period); b) a switch from the object to a therapeutic alternative; c) health plan disenrollment; or d) the end of the study dataset. We did not censor observation time upon outcome occurrence since this would violate an assumption underlying the self-controlled case series design.^{29,31}

The baseline period was defined as the six months immediately before the start of the observation period. It was required to be devoid of: a) a dispensing for the object drug of interest or a therapeutic alternative; and b) an interruption in health plan coverage.

Categorizing observation period time based on precipitant drug exposure

Each person-day of the observation period was dichotomized as a precipitant-exposed or precipitant-unexposed day. Precipitant-exposed days were defined by concomitant exposure to the candidate interacting precipitant drug using days' supply values without a grace period. Precipitant-unexposed days consisted of all other person-days of observation time. Consistent with a bi-directional self-controlled case series design, precipitant-unexposed days were permitted before and/or after precipitant-exposed days; this approach helped to minimize exposure trend bias.²² Although observation periods were not required to have both precipitant-exposed and precipitant-unexposed person-days, only observation periods with such discordance contributed to the estimation of the parameter of interest. See Figure 1 for a graphical representation of the design.

Several studies have shown that the risk of an adverse event due to a pharmacokinetic DDI often peaks shortly after initiating concomitant therapy and declines thereafter.^{32–34} Therefore, we examined a duration-response relationship for each object–precipitant pair. Precipitant-exposed observation time was divided into the following risk periods following the initiation of concomitance: 0–15, 16–30, 31–60, 61–120, and >120 days.

Defining the exposure of interest and covariates

The exposure of interest was use of the candidate interacting precipitant drug. Timeinvariant covariates (e.g., genetic polymorphisms, race, ethnicity) are inherently controlled for by the self-controlled case series design.³¹ In each regression model, we included the following time-varying covariates putatively associated with serious bleeding risk among antiplatelet drug users, as adapted from S₂TOP-BLEED³⁵ and its predecessors³⁶: a) prior history of or current ischemic heart disease and/or cerebrovascular disease; b) prior gastrointestinal bleeding and/or intracranial hemorrhage; c) ongoing concomitant therapy with an anticoagulant (e.g., warfarin), a non-clopidogrel antiplatelet drug (e.g., aspirin), a gastroprotective agent (e.g., lansoprazole), and/or an NSAID (e.g., ibuprofen); and d) average daily dispensed object drug dose. Because some of these drugs are available without a prescription, reliance on claims data may lead to under-ascertainment. Table S2 includes additional detail on covariates.

Identifying outcomes

The composite outcome of interest was serious bleeding, defined as hospital presentation for gastrointestinal bleeding or intracranial hemorrhage identified by International Classification of Diseases 9th Revision Clinical Modification discharge diagnosis codes. Operational definitions, including quantitative measures of algorithm performance,^{37–40} are presented in Table 4. Gastrointestinal bleeding and intracranial hemorrhage are the most common types of serious bleeding in persons taking antiplatelet drugs, and can be fatal.⁴¹ This was the rationale for studying a composite outcome of serious bleeding from these sites.

Statistical analysis

We constructed an analytic file for clopidogrel and one for pravastatin in which the unit of observation was the person-day of time covered by an active prescription to that object drug. The binary dependent variable was whether serious bleeding occurred on that day. Independent variables included a unique subject identifier, the subject's observation period, the observation day (categorized as precipitant-exposed vs. precipitant-unexposed), and time-varying covariates discussed above. The parameter of interest was the outcome occurrence rate during precipitant-exposed vs. precipitant-unexposed days, i.e., rate_{object+precipitant} / rate_{object}. In a secondary analysis, we examined outcome occurrence separately for the five risk periods discussed above. We used conditional Poisson regression models (*xtpoisson*, Stata v.15: College Station, TX, United States) to estimate rate ratios and 95% CIs.^{29,31,42} To avoid statistically unstable estimates, we did not estimate rate ratios when there were fewer than five precipitant-exposed patients or no events during precipitant-exposed time. Further, we did not report rate ratios from nonconverged conditional Poisson regression models or if the variance of the beta estimate for the parameter of interest was >10.

To account for the multiple estimation inherent in calculating hundreds of rate ratios and CIs, we used a semi-Bayes shrinkage method. This increases the validity of effect estimates and preserves the nominal type-1 error rate.^{43,44} Operationally, we prespecified a variance ($\sigma^2 = 0.25$) to assume that 95% of true rate ratios would be within an unspecified 7-fold range, then shrunk outlying effect estimates toward their geometric mean. In a secondary analysis, we increased the variance ($\sigma^2 = 0.67$) to assume that 95% of true rate ratios would be within an unspecified 25-fold range.

As discussed above, we used pravastatin as the negative control object drug to which clopidogrel findings were quantitatively compared. Therefore, we divided the semi-Bayes adjusted rate ratio for each clopidogrel-precipitant pair by the semi-Bayes adjusted rate ratio for the corresponding pravastatin-precipitant pair, and calculated 95% CIs using the delta method.⁴⁵ To contextualize findings, we compared DDI signals generated by our automated approach to putative interactions described in: clopidogrel's label; and Micromedex (IBM Watson Health: Cambridge, MA, United States) and Lexicomp (Wolters Kluwer: Alphen aan den Rijn, South Holland, Netherlands) DDI knowledge bases.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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Conflict of Interest Statement

Dr. Leonard serves on the Executive Committee of and Dr. Hennessy directs the University of Pennsylvania's Center for Pharmacoepidemiology Research and Training. The Center receives unrestricted funding from Pfizer and Sanofi. Dr. Leonard's spouse is employed by a health technology company that receives funding from AbbVie, Adamas, Celgene, Lilly, Lundbeck, Novartis, and Sunovion. All other authors declared no competing interests for this work.

ABBREVIATIONS LIST

CI	confidence interval
CKD	chronic kidney disease
СҮР	cytochrome P450
DDI	drug-drug interaction
NSAID	nonsteroidal anti-inflammatory drug

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STUDY HIGHLIGHTS

What is the current knowledge on the topic?

While several studies have examined drug-drug interactions that may interfere with clopidogrel's effectiveness in preventing thrombosis, few studies have examined drugs that may potentiate clopidogrel's bleeding risk. Both are of great clinical and public health importance.

What question did this study address?

Which precipitant drugs taken concomitantly with clopidogrel can increase the rate of hospital presentation for gastrointestinal bleeding or intracranial hemorrhage?

What does this study add to our knowledge?

Our high-throughput pharmacoepidemiologic screening of longitudinal healthcare data identified fifteen previously undescribed and/or unappreciated clopidogrel drug interactions that may be associated with serious bleeding. These potentially clinically-important interactions deserve further investigation.

How might this change clinical pharmacology or translational science?

The goals of drug-drug interaction research include screening for previously unanticipated interactions, elucidating their potential pharmacokinetic and/or pharmacodynamic mechanisms, predicting and examining their effects on pharmacokinetic and clinical outcomes, and developing and evaluating approaches to manage their risks in clinical settings. Future research on clopidogrel drug-drug interaction signals identified herein should contribute to broader pharmacologic knowledge of the drugs involved and the biological pathways involved in their kinetics and dynamics, thus yielding generalizable biologic knowledge.

Panel A: Individual with precipitant-exposed (P-E) and precipitant-unexposed (P-U) person-days of observation



Panel B: Individual solely with precipitant-exposed (P-E) person-days of observation



Figure 1. Examples of object drug exposure episodes potentially eligible for inclusion. Panel A depicts an individual with P-E and precipitant-unexposed P-U person-days of observation. Panel B depicts an individual solely with P-E person-days of observation. P-E = precipitant-exposed; P-U = precipitant-unexposed.

Table 1.

Descriptors of the object drug cohorts.

		Object drug cohort	
		Clopidogrel	Pravastatin
Persons	Persons		5,785
Days of observation, sum		16,285,312	5,112,014
Days of observation, median	n (quartile 1-quartile 3) per person	490 (183–1,089)	758 (331–1,427)
Gastrointestinal bleed/intrac	cranial hemorrhage outcomes, sum	32,839	8,575
Demographics		•	•
Age in years, median (quart	ile 1-quartile 3)	73.7 (65.5–78.7)	72.9 (65.1–78.8)
Sex, sum (%) female		10,125 (46.9)	2,999 (51.8)
Race, sum (%)	African American	2,709 (12.6)	788 (13.6)
	Asian	409 (1.9)	117 (2.0)
	Caucasian	15,075 (69.9)	4,025 (69.6)
	Hispanic	1,771 (8.2)	440 (7.6)
	Unknown	1,613 (7.5)	415 (7.2)
Dose covariate, time-varyin	g	•	•
Object drug average daily dose (quartile 1-quartile 3), median, in milligrams		75.0 (75.0–75.0)	40.0 (20.0-40.0)
Drug covariates, time-varyi	ng*	•	•
Anticoagulant, person-days (%)		677,620 (4.2)	514,442 (10.1)
Aspirin, person-days (%)		45,043 (0.3)	32,489 (0.6)
Cilostazol, person-days (%)		138,014 (0.8)	28,817 (0.6)
Dipyridamole, person-days (%)		21,138 (0.1)	28,845 (0.6)
Gastroprotective agent, person-days (%)		4,129,210 (25.4)	1,203,886 (23.6)
Nonsteroidal anti-inflammatory drug, person-days (%)		596,580 (3.7)	209,489 (4.1)
Disease covariates, time varying			
Cerebrovascular disease $\dot{\vec{r}}$, person-days (%)		9,856,287 (60.5)	2,331,869 (45.6)
Gastrointestinal bleed/intracranial hemorrhage $\stackrel{\ddagger}{\neq}$, person-days (%)		8,502,055 (52.2)	2,766,370 (54.1)
Ischemic heart disease $\stackrel{\uparrow}{}$, person-days (%)		13,541,836 (83.2)	2,994,267 (58.6)

* dispensed (pursuant to a prescription) on the day of observation or within the prior 30 days

 \dot{f} diagnosis (any position, any claim type) on the day of observation or ever prior

 $\vec{\tau}_{\rm diagnosis}$ (any position, any claim type) ever prior to the day of observation

Table 2.

Summary data on rate ratios for serious bleeding, by object drug cohort.

			Object drug cohort			
			Clopidogrel	Pravastatin	Ratio of Clopidogrel to Pravastatin	
	Unadjusted analyses					
	Candidate interacting precipitant drugs examined, number		536	436	433	
		DDI signals, number (%)	189 (35.3)	141 (32.3)	98 (22.6)	
	RR	Increased rate *	124 (23.1)	94 (21.6)	48 (11.1)	
		Decreased rate †	65 (12.1)	47 (10.8)	50 (11.5)	
		Geometric mean ± standard deviation	1.14 ± 1.97	1.14 ± 2.30	0.99 ± 2.23	
Poforo comi Povoc odiuctmont		Range, minimum to maximum	0.08 - 24.67	0.07 - 28.19	0.05 - 14.21	
Delole selli-Dayes aujustinem	Confounder	r-adjusted analyses				
	Candidate interacting precipitant drugs examined, number		536	434	431	
		DDI signals, number (%)	174 (32.5)	125 (28.8)	97 (22.5)	
	RR	Increased rate *	107 (20.0)	81 (18.7)	51 (11.8)	
		Decreased rate †	67 (12.5)	44 (10.1)	46 (10.7)	
		Geometric mean ± standard deviation	1.14 ± 1.98	1.12 ± 2.26	0.99 ± 2.26	
		Range, minimum to maximum	0.06 - 24.43	0.05 - 42.22	0.04 - 16.92	
	Unadjusted analyses					
	Candidate interacting precipitant drugs examined, number		536	436	433	
	RR	DDI signals, number (%)	150 (28.0)	109 (25.0)	72 (16.6)	
		Increased rate *	101 (18.8)	78 (17.9)	33 (7.6)	
		Decreased rate †	49 (9.1)	31 (7.1)	39 (9.0)	
		Geometric mean ± standard deviation	1.12 ± 1.43	1.14 ± 1.47	0.98 ± 1.52	
Aftar sami Payas adjustment		Range, minimum to maximum	0.36 - 4.53	0.45 - 6.95	0.20 - 3.81	
Aner senn-bayes aufusinent	Confounder-adjusted analyses					
	Candidate interacting precipitant drugs examined, number		536	434	431	
	RR	DDI signals, number (%)	139 (25.9)	90 (20.7)	73 (16.9)	
		Increased rate *	90 (16.8)	63 (14.5)	37 (8.6)	
		Decreased rate †	49 (9.1)	27 (6.2)	36 (8.4)	
		Geometric mean ± standard deviation	1.11 ± 1.42	1.12 ± 1.48	0.99 ± 1.54	
		Range, minimum to maximum	0.35 - 4.15	0.40 - 8.11	0.17 - 3.95	

DDI = drug-drug interaction; RR = rate ratio

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 * lower bound of the 95% confidence interval for the RR of interest was greater than the null value

 $\dot{\tau}$ upper bound of the 95% confidence interval for the RR of interest was less than the null value

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Table 3.

Clopidogrel DDI signals^{*} of potential clinical concern given statistically significantly increased rates of serious bleeding, by therapeutic class of precipitant drug.

Precipitant used concomitantly with clopidogrel	Ratio of rate ratios for clopidogrel to pravastatin	95% confidence interval	Putative mechanism(s) of interaction with clopidogrel ^{***}		
Anti-infective					
demeclocycline	3.06	1.10-8.49	unknown		
Cardiovascular					
bisoprolol	1.75	1.09–2.81	unknown		
gemfibrozil	1.64	1.12-2.42	induction of CYP3A4		
Central nervous system	Central nervous system				
diazepam	1.85	1.26–2.73	unknown		
eszopiclone	1.99	1.12–3.55	unknown		
hydromorphone	2.00	1.06–3.77	unknown		
pramipexole	2.57	1.54-4.29	unknown		
primidone	2.08	1.14–3.78	induction of CYP2C19, 2C9, 3A4, 1A2, 2B6		
prochlorperazine	2.27	1.23-4.18	unknown		
quetiapine	1.47	1.06–2.05	unknown		
tramadol	1.26	1.05–1.51	unknown		
Endocrine and metabolic					
megestrol	1.46	1.02-2.09	unknown		
pioglitazone	1.47	1.05-2.05	induction of CYP3A4		
Gastrointestinal					
ursodiol	2.82	1.24-6.43	unknown		
Nutritional					
potassium chloride	1.13	1.01–1.26	unknown		

CYP = cytochrome P450; DDI = drug-drug interaction

operational definition of a signal: statistically significantly elevated ratio of rate ratios for serious bleeding, excluding instances in which the rate ratio for the precipitant drug was not statistically significantly elevated in users of clopidogrel but statistically significantly protective in users of pravastatin

per DrugBank version 5.1.1 (The Metabolomics Innovation Centre: Edmonton, Alberta, Canada) enzyme, carrier, and transporter pathways

Table 4.

Operational definition of serious bleeding.

Outcome component	Discharge diagnosis descriptor	Discharge diagnosis ICD-9- CM code(s)	Discharge diagnosis position and claim type	Positive predictive value
Gastrointestinal bleeding	esophageal ulcer, with hemorrhage	530.21		~81%
	gastric ulcer, with hemorrhage	531.0X, 531.2X, 531.4X, 531.6X		
	duodenal ulcer, with hemorrhage	532.0X, 532.2X, 532.4X, 532.6X		
	peptic ulcer, with hemorrhage	533.0X, 533.2X, 533.4X, 533.6X		
	gastrojejunal ulcer, with hemorrhage	534.0X, 534.2X, 534.4X, 534.6X		
	gastritis and duodenitis, with hemorrhage	535.01, 535.11, 535.21, 535.31, 535.41, 535.51, 535.61, 535.71	Any-position discharge diagnosis on an inpatient hospitalization claim	
	other specified disorder of stomach and duodenum, with hemorrhage	537.83, 537.84		
	diverticula of intestine, with hemorrhage	562.02, 562.03, 562.12, 562.13		
	other disorders of intestine, with hemorrhage	569.85, 569.86		
	gastrointestinal hemorrhage	578.X		
Intracranial	subarachnoid hemorrhage	430	Any-position discharge diagnosis on an	~77–94%
hemorrhage	intracerebral hemorrhage	431	hospitalization claim	

ICD-9-CM = international classification of diseases 9^{th} revision clinical modification