

# Spatiotemporal Mapping of Sex Differences During Attentional Processing

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**Abstract:** Functional neuroimaging studies have increasingly aimed at approximating neural substrates of human cognitive sex differences elicited by visuospatial challenge. It has been suggested that females and males use different behaviorally relevant neurocognitive strategies. In females, greater right prefrontal cortex activation has been found in several studies. The spatiotemporal dynamics of neural events associated with these sex differences is still unclear. We studied 22 female and 22 male participants matched for age, education, and nicotine with 29-channel-electroencephalogram recorded under a visual selective attention paradigm, the Attention Network Test. Visual event-related potentials (ERP) were topographically analyzed and neuroelectric sources were estimated. In absence of behavioral differences, ERP analysis revealed a novel frontal-occipital second peak of visual N100 that was significantly increased in females relative to males. Further, in females exclusively, a corresponding central ERP component at around 220 ms was found; here, a strong correlation between stimulus salience and sex difference of the central ERP component amplitude was observed. Subsequent source analysis revealed increased cortical current densities in right rostral prefrontal (BA 10) and occipital cortex (BA 19) in female subjects. This is the first study to report on a tripartite association between sex differences in ERPs, visual stimulus salience, and right prefrontal cortex activation during attentional processing. *Hum Brain Mapp* 30:2997–3008, 2009. © 2009 Wiley-Liss, Inc.

**Key words:** sex; gender; attention; attention network test; event-related potentials

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## INTRODUCTION

Certain aspects of human cognition have consistently been demonstrated to differ between sexes. The majority of studies agree on sex-related performance differences within the domains of visuospatial as well as semantic-

verbal capabilities: in semantic-verbal tasks, female participants tend to outperform their male counterparts [e.g., Herlitz et al., 1997; Rossell et al., 2002; Wirth et al., 2007], whereas the opposite seems to be true for visuospatial paradigms [e.g., Astur et al., 1998, 2004; Linn and Peterson, 1985; see also Kimura, 2000 for a review].

Our knowledge of the biological mechanisms that may underlie these differences is greatly aided by rodent studies [see Jonasson, 2005 for a review]. Similar to humans, sex-related visuospatial performance differences have been described using the Morris water maze [Isgor and Sengelaub, 1998; Roof and Havens, 1992]. Lesion studies demonstrated that frontal lesions induced significantly greater disruption of spatial performance in female than in male rats [Kolb and Cioe, 1996] whereas lesions of the entorhinal cortex affected male more than female rats [Roof et al.,

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1993]. Thus, it appears promising to investigate human sex-related differences in visuospatial cognition by targeting these dimorphisms in functional neuroanatomy.

In the past few years, functional neuroimaging studies have increasingly aimed at approximating the neural substrates of human cognitive sex differences elicited by visuospatial challenge. Using mental rotation paradigms, several studies found significantly greater prefrontal cortex activation in female participants relative to their male counterparts even in the absence of significant behavioral differences [Butler et al., 2006; Hugdahl et al., 2006; Weiss et al., 2003]; this activation pattern is thought to be indicative of stronger top-down processing during visuospatial challenge in females. Using a virtual maze navigation paradigm, Riepe and coworkers have demonstrated that male participants use both hippocampi during navigation whereas women tend to use the right hippocampus only while additionally activating right prefrontal cortex [Grön et al., 2000]. While divergent cortical activation patterns have consistently been identified by these studies, the spatiotemporal dynamics of underlying cortical events has not been assessed conclusively, given the relatively coarse temporal resolution of the functional neuroimaging methods used in those studies. Yet the chronology of neural events during visuospatial processing could provide useful insights into the neurobiology of human sex differences in this cognitive domain.

A few studies using event-related potentials (ERPs) have tackled this issue so far, but the results—usually obtained applying the well-studied visual oddball paradigm—remain inconsistent. A higher visual N100 has been attributed to females, at least at temporal electrodes [Vaquero et al., 2004]; on the other hand, posterior N100 was found to be higher in pre-pubertal boys than in girls [Harter et al., 1989]. Similarly, higher amplitudes of visual P300 (P3b) have been found in females [Hoffman and Polich, 1999; Orozco and Ehlers 1998; Osterhout et al., 1997] as well as in males [Oliver-Rodriguez et al., 1999; Vaquero et al., 2004]. Thus, no “classical” ERP component has yet been identified to clearly distinguish between sexes.

Targeting the mental rotation ERP, sex differences emerged at a latency of about 100-300ms [Desrocher et al., 1995; Gootjes et al., 2008]. This is in line with other studies reporting on sex differences at this latency, although various measures were pursued, including visual evoked potentials [Emmerson-Hanover et al., 1994], potential fields and global field power [Skrandies et al., 1999], face recognition-related potentials [Proverbio et al., 2006], and event-related oscillations [Güntekin and Basar, 2007]. Despite reporting partially divergent directions of obtained results, these studies indicate where in the temporal cascade of the visual processing stream sex differences might be sought. No electrophysiological study, however, has successfully linked sex differences during visuospatial processing with differential right prefrontal cortex activation that has been consistently identified in functional neuroimaging studies.

The right prefrontal cortex is—among other functions—involved in selective spatial attention [Desimone and Duncan, 1995; Yantis and Serences, 2003]. Due to the close relationship between selective spatial attention and visuospatial cognition, we applied the Attention Network Test (ANT). This paradigm allows for assessment of attentive functions of alerting, orienting, and executive control [Fan et al., 2002] as well as examination of ERPs during selective visual attention [Neuhaus et al., 2007]. We focused our analysis on ERP components N100 and P300 which have been—albeit inconsistently—associated with sex differences during selective visual attention in previous studies; special emphasis was put on components emerging at a latency of about 100–300 ms. Source analysis was applied to allow for estimation of underlying cortical generators of differential ERP components.

## MATERIALS AND METHODS

### Subjects

Forty-four healthy subjects (22 f, 22 m) were included in this study. Participants were recruited via newspaper advertisements. Female and male participants were matched for age and education years; additionally, groups were matched for nicotine consumption since smoking status has been shown to affect cognitive ERP measures [Neuhaus et al., 2006]. None of the participants had a history of substance abuse other than tobacco smoking, of psychiatric axis I disorder according to DSM-IV [American Psychiatric Association, 1994], or of severe medical or neurological disorder. All subjects were examined by a psychiatrist and were free of pharmacological treatment.

All participants were right-handed, reported normal or corrected-to-normal vision, and were of Caucasian ethnicity. Demographic and basic neuropsychological data are provided in Table I. All subjects gave written, informed consent before participating. This study was approved by the ethics committee of the University Hospital Benjamin Franklin, Charité University Medicine Berlin, Germany, and was conducted in accordance with the Declaration of Helsinki.

### Stimuli and Task

Subjects were seated in a slightly reclined chair with a head rest and viewed the 14-inch cathode ray tube monitor from a distance of 60 cm. Behavioral responses were collected via two response keys on a keyboard resting on the subjects' lap. Visual stimuli were presented via Experimental Run Time System (ERTS; Berisoft Cooperation, Frankfurt/Main, Germany) on an IBM-compatible personal computer running Windows 98.

A fixation cross ( $0.37^\circ$  of visual angle) was visible in the center of the screen during the whole experiment. Cue stimuli ( $0.37^\circ$ ) appeared at  $1.01^\circ$  above or below the fixation cross (spatial cue), above and below the center (double cue), in the center (center cue), or were not displayed

**TABLE I. Demographic and basic neuropsychological data**

	Total N = 44	Female N = 22	Male N = 22	P —
Age [years]	30.50 ± 7.0	31.36 ± 8.2	29.64 ± 5.7	0.421 <sup>a</sup>
Education [years]	15.22 ± 2.1	15.11 ± 2.1	15.32 ± 2.2	0.755 <sup>a</sup>
Nicotine [pack years]	4.68 ± 6.3	3.66 ± 5.6	5.70 ± 7.0	0.289 <sup>a</sup>
Video playing experience <sup>b</sup>	0/28/12/4	0/15/7/0	0/13/5/4	0.107 <sup>c</sup>
LPS-IQ	114.75 ± 9.6	112.32 ± 10.1	117.18 ± 8.6	0.092 <sup>a</sup>
MWT-IQ	114.57 ± 13.8	114.55 ± 9.8	114.59 ± 17.2	0.991 <sup>a</sup>
DST	63.02 ± 9.9	63.62 ± 10.9	62.43 ± 8.9	0.701 <sup>a</sup>
TMT-A	27.00 ± 7.9	25.76 ± 5.8	28.24 ± 9.5	0.314 <sup>a</sup>
TMT-B	54.29 ± 20.1	54.33 ± 22.0	54.23 ± 18.6	0.988 <sup>a</sup>

<sup>a</sup> *t*-test.

<sup>b</sup> Factorized as: no computer experience/computer experience, but no video playing/occasional video playing (<2 h/week)/regular video playing (≥2 h/week)

<sup>c</sup>  $\chi^2$  test.

Abbreviations: LPS-IQ, Leistungsprüfsystem (non-verbal intelligence); MWT-IQ, Mehrfachwort-schatztest (verbal intelligence); DST, digit symbol test; TMT, trail making test.

(no cue). Spatial cues always validly displayed the upcoming target's location. Target stimuli consisted of five horizontally arranged arrows or lines (3.2° for horizontal target stimulus contour) presented at 1.01° above or below the fixation cross. By left or right button press, subjects had to indicate the direction of the central arrow irrespective of flanking conditions. Flankers were either lines (neutral target condition) or arrows pointing to the same (congruent) or to the opposite direction (incongruent).

Saliency of presented stimuli was parametrized as the sum of visual angle of the respective stimulus plus vertical angle from fixation in the screen's center. Thus, visual stimuli were assigned the following values: no cue 0; center cue 0.37 (0.37° visual angle only); spatial cue 1.38 (0.37° plus 1.01° vertical angle); double cue 2.76 (2 × 1.38°); and each target 4.21 (3.2° plus 1.01°).

Each trial consisted of a variable fixation period (400–1,600 ms), an invariant cue presentation (100 ms) with subsequent fixation period (400 ms), and presentation of the target (maximum duration 1,700 ms) followed by a variable fixation period immediately after response. The duration of each trial summed up to 4,000 ms (Fig. 1). After a training block of 24 trials with full feedback, subjects had to perform three experimental blocks with a total of 288 pseudo-randomized trials without feedback. Subjects were instructed to maintain focusing on the fixation cross throughout the experiment and to respond as fast and as accurately as possible.

Attention network effects were calculated as reaction time (RT) differences of the following task conditions: alerting = RT targets (no previous cue) minus RT targets (previous double cue); orienting = RT targets (previous center cue) minus RT targets (previous spatial cue); conflict = RT incongruent targets minus RT congruent targets. Additionally, mean RT and mean accuracy were assessed. Only those trials that were correctly responded to within a time window of 100–1,000 ms were taken into further analysis.

### ERP Acquisition and Analysis

EEG was recorded with 32 Ag/AgCl electrodes internally referenced to Cz using an electrode cap. The electrodes were positioned according to the International 10/20 system with the additional electrodes FC1, FC2, FC5, FC6, T1, T2, CP5, CP6, PO9, PO10, and Lo1. Electrode impedances were kept below 5 k $\Omega$ . EEG was assessed with a Neuroscan SynAmps (El Paso, TX, US) with a sampling rate of 250 Hz, gain 75,000, and an analogous 0.16 Hz high-pass filter. EEG analysis was conducted with Brain Vision Analyzer 1.05 (Brain Products, Munich, Germany). Using EEG raw data, ocular artifact correction was performed using an independent component analysis approach [Jung et al., 2000]. Data were then re-referenced to average reference and a digital low-pass filter was applied at 45 Hz. After artifact rejection (≥80  $\mu$ V at any electrode), data were segmented relative to stimulus onset (350 ms pre-stimulus to 800 ms post-stimulus) and submitted to baseline correction. At least 30 artifact-free sweeps were averaged for each analyzed experimental condition; only correctly responded trials (100–1000 ms) were analyzed.

ERP components N100 and P300 were determined semi-automatically with a visual control post hoc. N100 was identified at O1 and O2 as prominent negative deflections between 150 and 300 ms. N100 was then subdivided into two separate peaks: an earlier peak was identified between 150 and 210 ms and was computed for Fz, Pz, and Oz (interpolated: (O1+O2)/2); a later peak was identified between 200 and 300 ms and computed for Fz, Cz, Oz (interpolated). If no clear double peak was detectable, the largest deflection was accepted as first N100 peak if it occurred between 150 and 210 ms and as second N100 if it appeared between 210 and 300 ms. If only the first N100 peak was present, the amplitude of the second N100 peak was set at the corresponding latencies of the grand average. P300 was identified at Pz as a prominent positive

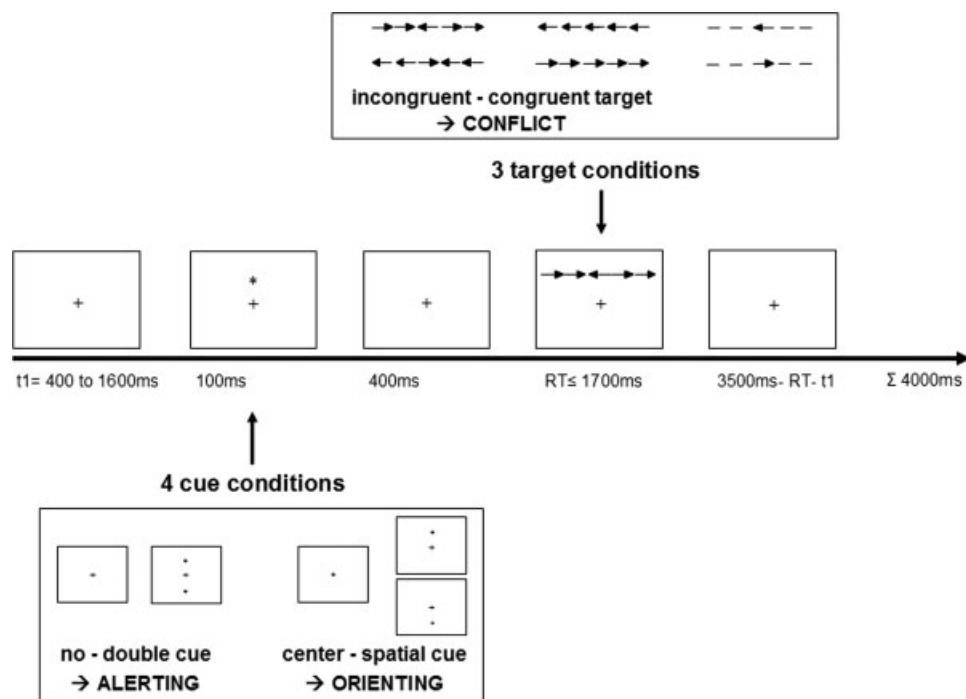


Figure 1.  
ANT design.

deflection between 300 and 600 ms and was computed for Fz, Cz, and Pz.

### Source Localization

Neuroelectric source imaging with LORETA [version 2005 March; Pascual-Marqui et al., 1994, 1999] was used to compute the cortical distribution of electrical activity as recorded from scalp electrodes. This version of LORETA employs a three-shell spherical head model registered to the Talairach atlas of the human brain [Talairach and Tournoux, 1988]. The solution space is restricted to the cortical grey matter and the hippocampus in the Talairach atlas, producing a total of 2,394 voxels. Without a priori assumptions on number and location of active sources, this solution to the inverse problem computes the current density at each voxel as the weighted sum of the scalp electric potentials. The unit of each voxel represents the electrical activity as squared magnitude (i.e., power [ $\mu\text{V}/\text{Hz}^2$ ]) of the computed current density [ $\mu\text{A}/\text{mm}^3$ ]. Current density maxima were regarded as spatially separate if the distance between the corresponding voxels was larger than 14 mm [Pascual-Marqui et al., 1994].

### Basic Neuropsychological Assessment

The German tests Mehrfachwortschatztest [MWT; Lehrl et al., 1995] and Leistungsprüfsystem [LPS; Horn, 1983]

were employed to quantify non-verbal and verbal intelligence, respectively. The Digit Symbol Test [DST; Wechsler, 1981] and Trail Making Tests A and B [TMT-A/-B; Reitan, 1959] were used to assess basic attentive and executive functions along with psychomotor function.

### Statistical Analyses

Statistical calculations were carried out with SPSS for Windows 15.0 (Chicago, IL, US). Gaussian distribution of behavioral and ERP data was assessed with Kolmogorov-Smirnov test; between-group comparisons of demographic and behavioral data were then computed with *t*-test or Pearson chi-square. Examination of ERP data was performed with repeated measures analyses of variance entering stimuli and electrodes as within-subject factors and sex and video playing experience as between-subject factors. Although not significantly different between groups, video playing experience was introduced as a co-factor since it has been shown to impact on sex differences in visuospatial cognition [Feng et al., 2007]. Video playing experience was factorized as no computer experience/computer experience, but no video playing/occasional video playing (<2 h/week)/regular video playing ( $\geq 2$  h/week). Separate repeated measures analyses were computed for first N100 peak (6 cue and target stimuli  $\times$  3 electrodes: Fz, Pz, Oz), second N100 peak (6 stimuli  $\times$  3 electrodes: Fz, Cz, Oz), and P300 (3 target stimuli  $\times$  3 electrodes: Fz, Cz, Pz). Post

TABLE II. ANT behavioral data

	Total	Female	Male	<i>P</i> <sup>a</sup>
Mean RT [ms]	537.04 ± 69.5	556.76 ± 77.5	517.31 ± 55.4	0.059
Mean accuracy [%]	98.78 ± 1.1	98.90 ± 0.9	98.66 ± 1.3	0.468
Alerting effect [ms]	44.90 ± 25.9	43.60 ± 26.4	46.20 ± 25.9	0.742
Orienting effect [ms]	52.08 ± 22.7	49.77 ± 20.0	54.38 ± 26.4	0.507
Conflict effect [ms]	99.10 ± 34.1	104.78 ± 37.7	93.41 ± 29.9	0.274

<sup>a</sup> *t*-test.

Abbreviation: RT, reaction time.

hoc tests of significant within-subjects factors were computed as post hoc ANOVAs. Correlation analysis between stimulus salience and mean sex differences of central ERP component amplitude was performed with Spearman rank order correlation, as only 7 separable stimulus conditions were available. Correlation analysis between ERP amplitude measures and RT was calculated with Pearson correlation.

Statistical imaging of current density differences was based on non-parametric voxel-by-voxel *t*-tests [Holmes et al., 1996]. This “maximum *t*-statistic” is a non-parametric analysis that offers, after a procedure of randomizations (5000 randomly created groups across conditions), a randomization distribution of the maximum statistic and will produce threshold values for single voxel *P*s. This *P* value will be < 0.01 if the maximum of the observed statistical values is in the largest 1% of the randomization values, which is the case if it is greater than the 99th percentile of the randomization values. The time frame of interest for statistical imaging (150–300 ms) was selected on the basis of prior conventional ERP analysis.

For tests of demographic and basic neuropsychological differences, a conservative approach with *P* < 0.05 was chosen; for all other tests, a *P* < 0.01 was considered significant.

## RESULTS

### Demographic and Neuropsychological Data

Demographic and basic neuropsychological data of participants are shown in Table I. There were no significant differences between groups regarding distribution of age, education years, nicotine consumption, or video playing experience. Also, no significant differences were found for performance in the basic neuropsychological tests LPS, MWT, DST, and TMT.

Mean attention network effects are summarized in Table II, mean differential RTs are illustrated in Figure 2. There was no significant difference between groups for any of these behavioral measures. However, there was a statistical trend towards longer RT in female participants (*P* = 0.059). No significant differences between groups were found for performance accuracy.

### ERP Components: N100

Grand average ERP components at midline and occipital leads stratified by sex for cue and target conditions are illustrated in Figure 3. N100 displayed a double peak that was most prominent in female participants at electrodes Oz and Fz. Mean latencies for N100 (first peak, female vs.

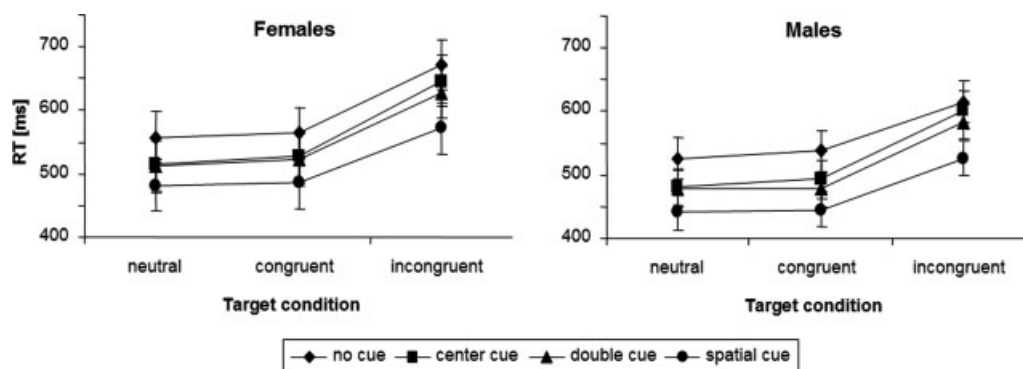
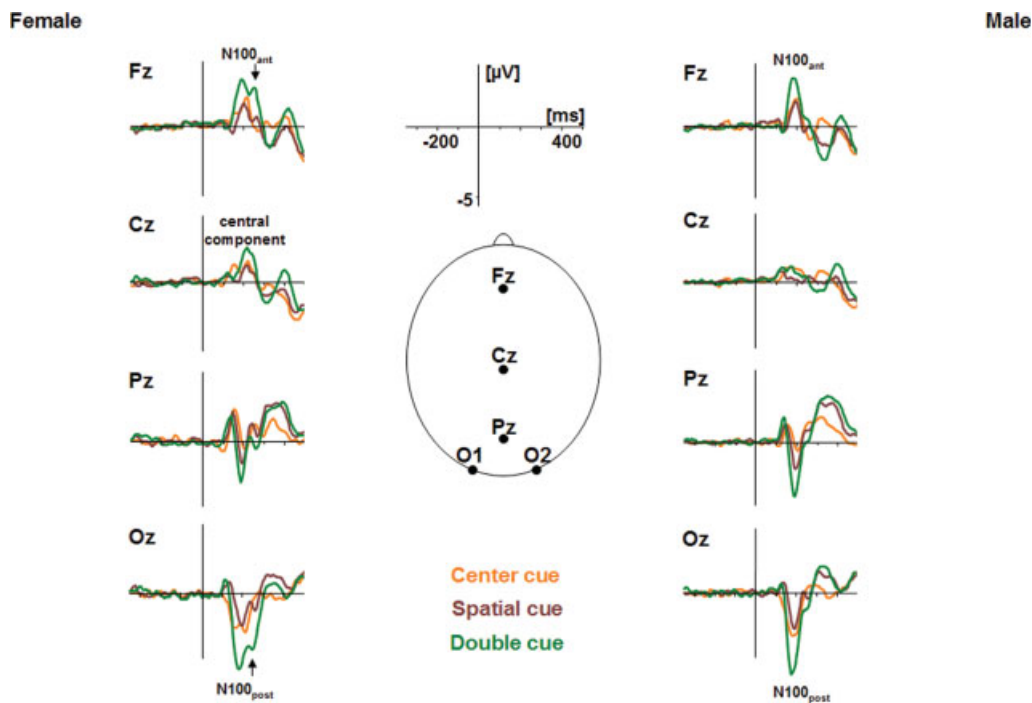


Figure 2.

Differential reaction times of female and male participants across trials for different target and preceding cue conditions.





**Figure 3.**

Grand average ERP components at midline and occipital leads following cue presentation stratified by sex. Oz is interpolated from O1 and O2. Note the polarity reversal of N100 over Cz as a consequence of reference system. A distinct N100 double peak is present at frontal (N100<sub>ant</sub>) and occipital (N100<sub>post</sub>) electrodes in female participants predominantly. At Cz, a positive shift at N100 latency is present in females exclusively. [Color figure can be viewed in the online issue, which is available at [www.interscience.wiley.com](http://www.interscience.wiley.com).]

male) were 183.6 ms vs. 184.5 ms at Oz and 188.5 ms vs. 187.2 ms at Fz; for the second peak, mean latencies were 249.6 ms vs. 257.6 ms at Oz and 249.8 ms vs. 257.1 ms at Fz, respectively. This double peak configuration was present in 18 of 22 female and in 9 of 22 male participants as assessed by a ratio of second to first peak  $\geq 0.3$  for mean N100 across conditions. Moreover, exclusively in female subjects, a central positive component was found at a latency of 223.5 ms comparable to the second peak of N100.

For the first peak of N100, significant main effects of electrode ( $F = 97.396$  (2, 38),  $P < 0.0001$ , partial  $\eta^2 = 0.837$ ) and of stimulus ( $F = 18.104$  (5, 35),  $P < 0.0001$ , partial  $\eta^2 = 0.720$ ) were observed. The interaction of electrode  $\times$  stimulus also yielded a significant effect ( $F = 15.984$  (10, 30),  $P < 0.0001$ , partial  $\eta^2 = 0.842$ ). These results indicate that the first N100 peak varies according to scalp location and stimulus applied. There was, however, no significant main effect of interaction with sex on the dependent variable.

For the second peak of N100, several significant effects were observed. Similar to the first peak, significant main effects of electrode ( $F = 63.919$  (2, 38),  $P < 0.0001$ , partial  $\eta^2 = 0.771$ ) and of interaction of electrode  $\times$  stimulus ( $F = 5.577$  (10, 30),  $P = 0.0001$ , partial  $\eta^2 = 0.650$ ) were observed. Additionally, significant main effects were

observed for interactions of electrode  $\times$  sex ( $F = 10.155$  (2, 38),  $P = 0.0003$ , partial  $\eta^2 = 0.348$ ) as well as electrode  $\times$  stimulus  $\times$  sex ( $F = 3.289$  (10, 30),  $P = 0.006$ , partial  $\eta^2 = 0.523$ ), thus indicating major modulating influences of sex. Post hoc analyses revealed significant sex differences of second N100 peak at all electrode sites analyzed (Fz:  $P = 0.001$ ; Cz:  $P = 0.004$ ; Oz:  $P = 0.003$ ). When additionally stratifying for stimulus type at these electrodes, differential effects were observed, mainly pointing at significant effects of double cue and target conditions (Table III).

The mean numbers of segments used for averaging were as follows (female vs. male): no cue  $64.09 \pm 7.2$  vs.  $66.09 \pm 4.2$ ; center cue  $64.36 \pm 6.2$  vs.  $67.3 \pm 3.4$ ; spatial cue  $64.05 \pm 7.1$  vs.  $67.45 \pm 3.1$ ; double cue  $63.77 \pm 7.7$  vs.  $67.27 \pm 3.2$ ; neutral target  $69.82 \pm 23.2$  vs.  $75.41 \pm 21.0$ ; congruent target  $67.86 \pm 25.1$  vs.  $76.27 \pm 20.8$ ; incongruent target  $62.95 \pm 21.9$  vs.  $69.50 \pm 19.0$ . None of these differences were statistically significant.

### ERP Components: P300

For P300, only a significant main effect of stimulus was observed ( $F = 22.382$  (2, 38),  $P < 0.0001$ , partial  $\eta^2 = 0.541$ ), signifying P300 amplitude modulation according to

**TABLE III. Second N100 peak and central component amplitudes [ $\mu\text{V}$ ] of visual ERP**

	Total	Female	Male	$P^a$
Fz <sub>center cue</sub>	1.16 ± 1.80	1.67 ± 1.81	0.58 ± 1.66	0.126
Fz <sub>spatial cue</sub>	1.46 ± 1.58	1.88 ± 1.51	1.04 ± 1.56	0.192
Fz <sub>double cue</sub>	2.43 ± 2.41	<b>3.56 ± 2.26</b>	<b>1.30 ± 2.04</b>	<b>0.001</b>
Fz <sub>neutral target</sub>	2.53 ± 2.02	<b>3.06 ± 2.02</b>	<b>2.00 ± 1.92</b>	<b>0.006</b>
Fz <sub>congruent target</sub>	2.24 ± 2.10	2.80 ± 2.04	1.68 ± 2.06	0.021
Fz <sub>incongruent target</sub>	2.46 ± 2.15	<b>2.98 ± 2.22</b>	<b>1.94 ± 1.98</b>	<b>0.006</b>
Cz <sub>center cue</sub>	2.52 ± 1.33	2.64 ± 1.34	2.40 ± 1.33	0.727
Cz <sub>spatial cue</sub>	2.05 ± 1.69	2.25 ± 1.81	1.85 ± 1.58	0.291
Cz <sub>double cue</sub>	3.13 ± 1.89	3.54 ± 1.96	2.73 ± 1.76	0.077
Cz <sub>neutral target</sub>	1.52 ± 2.49	2.10 ± 2.15	0.94 ± 2.71	0.013
Cz <sub>congruent target</sub>	1.58 ± 2.52	<b>2.34 ± 2.27</b>	<b>0.82 ± 2.58</b>	<b>0.009</b>
Cz <sub>incongruent target</sub>	1.39 ± 2.41	<b>2.10 ± 2.10</b>	0.70 ± 2.55	<b>0.005</b>
Oz <sub>center cue</sub>	-1.75 ± 2.45	-2.53 ± 2.67	-0.67 ± 1.96	0.080
Oz <sub>spatial cue</sub>	-1.55 ± 2.34	-2.14 ± 2.75	-0.97 ± 1.70	0.197
Oz <sub>double cue</sub>	-3.54 ± 3.09	<b>-4.99 ± 3.44</b>	<b>-2.10 ± 1.83</b>	<b>0.002</b>
Oz <sub>neutral target</sub>	-2.33 ± 2.00	-2.79 ± 2.20	-1.88 ± 1.71	0.042
Oz <sub>congruent target</sub>	-1.75 ± 2.10	-2.16 ± 2.34	-1.33 ± 1.76	0.109
Oz <sub>incongruent target</sub>	-1.87 ± 1.94	-2.49 ± 1.90	-1.25 ± 1.82	0.016

<sup>a</sup> Post hoc ANOVAS.

Significant differences at  $\alpha = 0.01$  are printed bold. Abbreviations: Fz, frontal midline electrode; Cz, central midline electrode; Oz, occipital midline electrode (interpolated).

preceding target type. There were also trends for interactions of stimulus  $\times$  sex ( $F = 4.254$  (2, 38),  $P = 0.022$ , partial  $\eta^2 = 0.183$ ) as well as electrode  $\times$  stimulus ( $F = 3.169$  (2, 38),  $P = 0.025$ , partial  $\eta^2 = 0.260$ ); however, no significant effects other than stimulus were detected in this model. Specifically, there were no significant main effects of interaction stimulus  $\times$  electrode  $\times$  sex on the depend-

**TABLE IV. P300 peak amplitudes [ $\mu\text{V}$ ] of visual ERP**

	Total	Female	Male	$P^a$
Fz <sub>neutral target</sub>	2.29 ± 2.08	2.62 ± 2.03	1.96 ± 2.12	0.195
Cz <sub>neutral target</sub>	4.58 ± 2.37	4.28 ± 2.43	4.89 ± 2.33	0.239
Pz <sub>neutral target</sub>	4.91 ± 2.03	4.57 ± 2.29	5.26 ± 1.73	0.153
Fz <sub>congruent target</sub>	2.43 ± 2.14	2.61 ± 2.02	2.24 ± 2.28	0.329
Cz <sub>congruent target</sub>	4.44 ± 2.50	4.04 ± 2.78	4.83 ± 2.18	0.138
Pz <sub>congruent target</sub>	4.56 ± 2.00	4.33 ± 2.18	4.79 ± 1.83	0.278
Fz <sub>incongruent target</sub>	2.73 ± 2.21	2.99 ± 2.35	2.48 ± 2.09	0.289
Cz <sub>incongruent target</sub>	4.61 ± 2.50	4.17 ± 2.49	5.04 ± 2.49	0.267
Pz <sub>incongruent target</sub>	3.67 ± 1.89	3.39 ± 2.06	3.95 ± 1.70	0.177

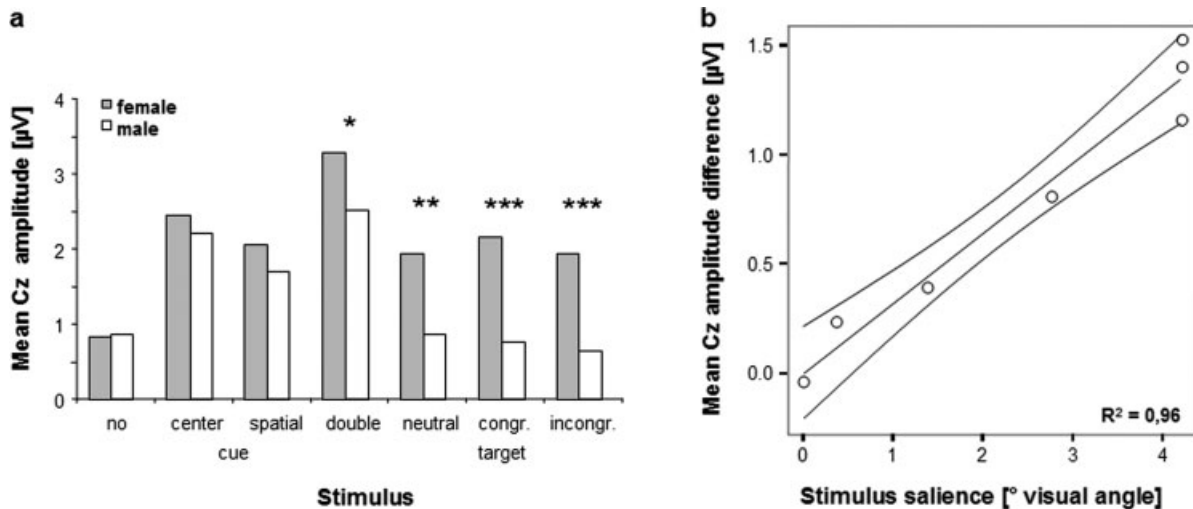
<sup>a</sup> Post hoc ANOVAS.

Abbreviations: Fz, frontal midline electrode; Cz, central midline electrode; Pz, parietal midline electrode.

ent variable. Results of our P300 analysis are summarized in Table IV.

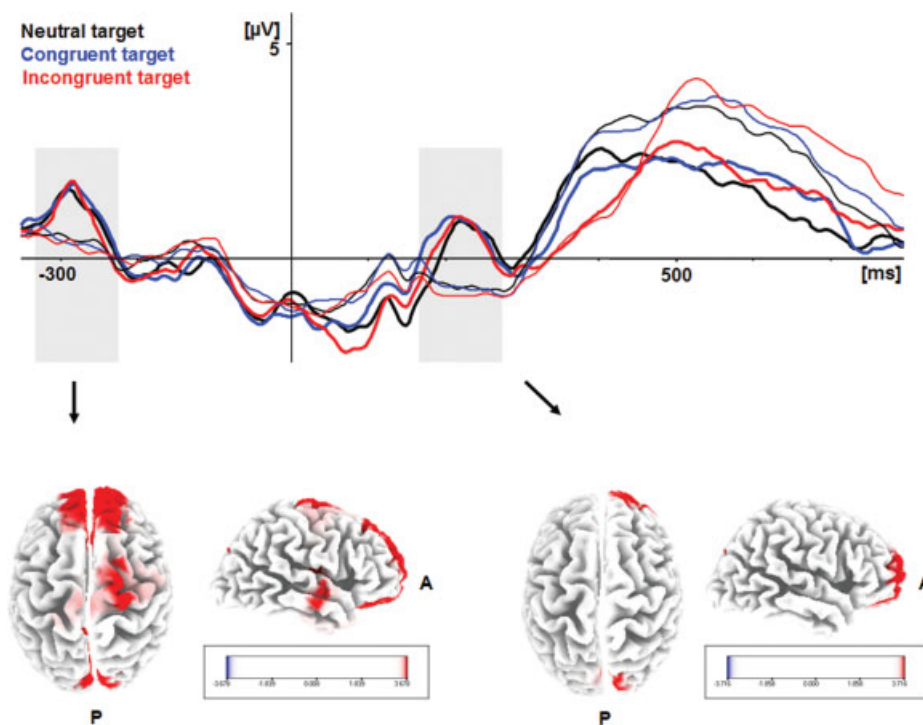
### Correlation Analyses

A non-parametric correlation analysis revealed a strong association between stimulus salience and peak differences between sexes at Cz (Spearman's  $\rho = 0.964$ ;  $P = 0.0005$ ; Fig. 4). To further validate this correlation, another repeated measures ANOVA was computed entering stimulus type with 2 levels (cues, targets). Again, significant main effects of electrode, electrode  $\times$  sex, and electrode  $\times$  stimulus ( $F = 9.358$  (2, 38),  $P = 0.0005$ , partial  $\eta^2 = 0.330$ ) were observed; however, there was no significant main effect of electrode  $\times$  stimulus  $\times$  sex ( $F = 1.448$  (2, 38),  $P =$



**Figure 4.**

(a) Mean central ERP component amplitude differences between sexes for different stimuli; \* < 0.1; \*\* < 0.05; \*\*\* < 0.01 (b) Stimulus salience (visual angle plus vertical angle from fixation) is plotted against mean central component amplitude difference (4 cues, 3 targets; female minus male); regression line with 95% mean prediction interval.



**Figure 5.**

Grand average target ERP components at Cz by sex (female, bold lines; male, thin lines) and target conditions. Time frame of interest for source localization is 170–270 ms post cue (left) and post target (right), respectively. For both contrasts between sexes, significant current density increases are observed in right ventral frontal cortex and right occipital cortex in female relative to male participants. [Color figure can be viewed in the online issue, which is available at [www.interscience.wiley.com](http://www.interscience.wiley.com).]

0.248, partial  $\eta^2 = 0.071$ ), thus ruling out a major influence of behavioral stimulus relevance, i.e., cue vs. target, on the observed sex differences.

A parametric correlation analysis was performed between RTs and second N100 peak amplitude measures at Fz, Cz, and Oz (interpolated). There was no correlation between mean RT and mean amplitudes of the second N100 peak neither were there any significant ( $\alpha = 0.01$ ) correlations for single conditions.

### Source Localization

To estimate neural sources of sex differences in visual ERP, a between-group contrast of current densities was calculated at 170–270 ms post target stimuli, corresponding to the central component observed in female participants. *t* statistics showed significant current density increases in right rostral prefrontal cortex (Brodmann area (BA) 10; Talairach *x, y, z*: 18, 59, 15;  $t = 4.32$ ) and in right occipital cortex (BA 19; *x, y, z*: 11, -81, 36;  $t = 4.22$ ) in female compared to male participants; *t*-values  $\geq 3.69$  were significant at a level of  $P = 0.01$ .

When computing corresponding cortical responses to cue stimuli, a comparable pattern emerged for female rela-

tive to male subjects with current density increases in right rostral prefrontal cortex (BA 10; *x, y, z*: 11, 59, 29;  $t = 5.93$ ) and right occipital cortex (BA 19; *x, y, z*: 4, -74, 29;  $t = 5.28$ ); additionally, increases were observed in right temporal cortex (BA 22; *x, y, z*: 46, -11, 1;  $t = 4.58$ ) and in the precentral region (BA 6; *x, y, z*: 18, -11, 71;  $t = 4.53$ ); *t*-values  $\geq 3.71$  were significant at a level of  $p = 0.01$  (Fig. 5).

An additional source analysis was calculated for the P300 responses within a time frame of 300–600 ms post target stimuli, although only a statistical trend for a differential interaction of stimulus  $\times$  sex was observed at the electrode level. *t* statistics revealed statistical trends towards higher occipital current densities in male participants for each P300 condition, i.e., neutral targets (BA 17; *x, y, z*: 11, -81, 8;  $t = 3.10$ , threshold  $t = 3.76$ ); congruent targets (BA 18; *x, y, z*: 11, -88, 8;  $t = 3.37$ , threshold  $t = 3.81$ ); and incongruent targets (BA 18; *x, y, z*: 4, -88, 22;  $t = 3.23$ , threshold  $t = 3.98$ ).

### DISCUSSION

The main results of this study are (1) a double peak of N100 and a corresponding prominent central ERP component at 170–270 ms exclusively in female participants, (2) a



strong correlation between stimulus salience and sex difference of the central ERP component amplitude, and (3) increased current density in rostral prefrontal (BA 10) and occipital (BA 19) cortices in female participants at 170–270 ms post stimulus during attentional processing. These latencies indicate that sex differences occur at a relatively early time point within the temporal cascade of the attentional processing stream. Our results are strongly supported by previous electrophysiological studies on different functional sexual dimorphisms occurring up to 300 ms [Desrocher et al., 1995; Emmerson-Hanover et al., 1994; Gootjes et al., 2008; Güntekin and Basar 2007; Proverbio et al., 2006; Skrandies et al., 1999]. Our findings were observed in the absence of significant behavioral differences and are unlikely to be confounded by group disparities other than sex as distribution of demographic variables including video playing experience and performance on basic neuropsychological tasks were not significantly different between groups in multivariate analysis. This is the first study to report on a tripartite association of sex differences in visual ERPs, stimulus salience, and right prefrontal cortex activation during attentional processing.

RT analysis showed a characteristic pattern of attention network effects that are comparable to the results published originally [Fan et al., 2002] and by our group [Neuhaus et al., 2007]. Attention network effects were comparable between sexes and no significant differences were found. However, there was a trend towards different mean RTs between sexes indicating a slightly slower overall performance in female participants. Speculatively, the effect size elicited by flanker-type paradigms might be too small to reliably evoke significant differential RTs between females and males in this study. Parametric post hoc correlation analysis indicated that RT differences are independent from differences in ERP amplitudes; thus, the present results from ERP analysis do rather not reflect behavioral differences between sexes but may be interpreted in terms of more fundamental differences.

ERP components revealed the classic componentry of posterior N100-P200 for non-target stimuli and additional P300 for target stimuli in both sexes. With the present analysis, earlier findings on parietal P300 modulation according to target complexity and on central P300 latency increase as a function of flanker conflict effect are confirmed [Neuhaus et al., 2007]; however, only statistical trends for effects of sex were obtained for the P300 component. Likewise, source analysis revealed statistical trends towards higher occipital activation in males that correspond to higher posterior P300 relative to female participants. Perhaps these results reflect the heterogeneity of findings obtained for this ERP component so far [Hoffman and Polich, 1999; Oliver-Rodriguez et al., 1999; Orozco and Ehlers 1998; Osterhout et al., 1997; Vaquero et al., 2004]. Hypothetically, there might be a sex difference in visual P300, but with the current design this question cannot be conclusively answered.

Sex differences were found for visual N100 that is considered an index of perceptual discrimination processes [Vogel and Luck, 2000]. In female participants only, a second N100 peak and a corresponding central component were present. N100 double peak was detectable in posterior as well as anterior leads for both cue and target conditions; polarity reversal at the frontal lead is attributable to the reference system with Cz used as internal reference during EEG recording. The finding of a double-peaked N100 following visual stimuli has not been described yet. Studies using auditory stimulation, however, have already reported this distinct N100 morphology at frontal electrodes and suggested that the second peak may indicate the amount of attention allocated to stimulus processing [Mulert et al., 2001]. The impact of selective attention on N100 is well known as “N1-effect” [Hillyard et al., 1973], “negative difference” [Hansen and Hillyard, 1980], or “processing negativity” [Näätänen and Picton, 1987] which refers to an augmented negativity of N100 during attentional stimulus processing. Accordingly, the second cognitive N100 component has been shown to be modulated by effort and task difficulty [Mulert et al., 2005, 2007].

In females exclusively, a central ERP component emerged at a latency that was comparable to the second N100 peak. Speculatively, this central component reflects a summation effect of higher frontal and occipital N100 activity in females as a consequence of the second N100 peak. A strong positive correlation was found between differential central ERP component amplitude as a function of sex and visual angle of presented stimuli. In other words, as visual stimulus salience increases, this central ERP component escalates in females whereas no comparable ERP modulation is observed in males. This central ERP component may thus be reflective of increasingly effortful visual stimulus discrimination in females as stimulus salience increases.

This assumption is consistent with the conceptualization of N100 as an index of visual stimulus discrimination in general [Hopf et al., 2002; Vogel and Luck, 2000]. Our hypothesis also corroborates earlier findings on the second N100 component peak as a function of effort and task difficulty in the auditory modality described by Mulert and coworkers [2005, 2007]. Further, the present findings seem to extend their observations to the visual modality, suggesting either comparable cognitive processes for these modalities or a fundamental cognitive mechanism operating at a supramodal level.

For the respective time frame (170–270 ms), source localization analyses consistently indicated greater cortical current density in right rostral prefrontal cortex (BA 10) and in extrastriate visual cortex (BA 19) in females. Due to the very similar frontal-occipital pattern, this current density configuration is particularly suitable to explain the observed voltage differences on the scalp. Moreover, a recent structural neuroimaging study found that orbitofrontal gray matter is larger in females when controlling for total intracranial volume [Gur et al., 2002]. This is also

consistent with the present results: larger cortical volumes can principally account for increased scalp potentials that are reflected in the topographic distribution of the second N100 peak in our data.

Functionally, the role of BA 19 can be clearly attributed to visual stimulus processing; however, the question of functional significance of BA 10 has been addressed only recently and remains a matter of debate. A recent meta-analysis on functional neuroimaging studies reporting activation peaks in rostral prefrontal cortex described segregation between emotional and cognitive tasks along a medial-lateral axis with lateral BA 10 being more associated with behavioral guidance during cognitive tasks [Gilbert et al., 2006]. This view is further substantiated by an increasing number of studies on top-down influences in sensory processing, especially top-down modulation of visual processing by prefrontal cortex [e.g., Gazzaley et al., 2007; Kastner and Ungerleider, 2001; Sehatpour et al., 2008; see also Gilbert and Sigman, 2007 for a review]. Top-down influences dynamically set sensory cortices in a specific working mode as behaviorally required. Although former studies did not identify BA 10 as a source of prefrontal top-down control, the results obtained here suggest a considerable role of this cortical area in top-down processes, at least in females. Specifically, the prefrontal-occipital pattern together with the correlation between scalp potentials and visual stimulus salience strongly suggests an amplification of sensory stimulus representation in BA 19 by prefrontal BA 10. It is proposed that BA 10 harbors a top-down calibration mechanism for representation and discrimination of complex visual stimuli in females. In sum, we interpret our findings as evidence for stronger top-down control in females during attentional processing, especially when stimuli are salient.

A major limitation of this study is constituted by the lack of control for hormone status in female participants. It has been shown that cognitive skills are influenced by sex hormone status across the menstrual cycle [Hampson, 1990]. Therefore, studies on cognitive sex differences are thought to require statistical control for sex hormone levels. On the other hand, our sample might be large enough to assume normal distribution of sex hormones; nevertheless, a replication study with control of sex hormone levels is needed to validate our findings. A minor limitation arises from potential lack of generalization of our results as a relatively homogenous sample in terms of age and education has been investigated. Next, our paradigm was not designed to explicitly allow for assessment of stimulus salience; thus our results warrant replication in studies employing more specific paradigms. Finally, given that processing demands for cue and target conditions are qualitatively different with respect to behavioral relevance, the integration of both stimulus types into the same correlation analysis of the central ERP component may be questionable; however, a parametric analysis indicated that stimulus type did not interact with sex differences of the central ERP component. Further, our approach might be

justified by its explorative nature and may stimulate further research with more appropriate study paradigms.

In conclusion, however, we identified a distinct ERP feature that occurs during attentional processing of increasingly complex visual stimuli and that distinguishes between healthy female and male participants. ERPs were generated in right ventral prefrontal as well as right occipital cortex suggesting a top-down influence of prefrontal on occipital cortex in females. Functionally, our findings may indicate augmentation of sensory stimulus representation and discrimination by a prefrontal mechanism. Our results indicate that we detected a complex association of sex differences in visual ERPs, stimulus salience, and right prefrontal cortex activation that point to relative differences in visual attentional processing between sexes.

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