The effect of angiopoietin-I upregulation on the outcome of acute ischaemic stroke in rodent models: A meta-analysis

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Joseph V Moxon^{1,2}, Alexandra F Trollope^{1,2,3}, Brittany Dewdney^{1,4}, Catherine de Hollander³, Domenico R Nastasi¹, Jane M Maguire⁴ and Jonathan Golledge^{1,2,5}

Abstract

Clinical studies report that low circulating angiopoietin-1 concentration at presentation predicts worse outcomes after ischaemic stroke. Upregulating angiopoietin-1 may therefore have therapeutic benefit for ischaemic stroke. This systematic review assessed whether upregulating angiopoietin-1 improved outcomes in rodent models of ischaemic stroke. Random-effects models quantified the effect of angiopoietin-1 upregulation on stroke severity in terms of the size of cerebral infarction and the extent of blood-brain barrier permeability. Eleven studies utilising rat and mouse models of ischaemic stroke fulfilled the inclusion criteria. Meta-analyses demonstrated that angiopoietin-1 upregulation significantly reduced cerebral infarction size (standardised mean difference: -3.02; 95% confidence intervals: -4.41, -1.63; p < 0.001; n = 171 animals) and improved blood-brain barrier integrity (standardized mean difference: -2.02; 95% confidence intervals: -3.27, -0.77; p = 0.002; n = 129 animals). Subgroup analyses demonstrated that angiopoietin-1 upregulation improved outcomes in models of transient, not permanent cerebral ischaemia. Six studies assessed the effect of angiopoietin-1 upregulation on neurological function; however, inter-study heterogeneity prevented meta-analysis. In conclusion, published rodent data suggest that angiopoietin-1 upregulation improves outcome following temporary cerebral ischaemia by reducing cerebral infarction size and improving blood-brain barrier integrity. Additional research is required to examine the effect of angiopoietin-1 upregulation on neurological function, size and improving blood-brain barrier integrity. Additional research is required to examine the effect of angiopoietin-1 upregulation on neurological function, and investigate the benefit and risks in patients.

Keywords

Ischaemic stroke, angiopoietin-I, rodent model, cerebral infarction, blood-brain barrier permeability

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Introduction

Stroke is the world's third leading cause of death, and primary cause of adult disability.¹ Approximately 80% of all strokes result from acute cerebral ischaemia.² An important treatment for ischaemic stroke is rapid restoration of the cerebral blood supply by mechanical or chemical thrombolysis.² Disadvantages of this therapy include a risk of substantial complications, need for specialist expertise and resources not available in many regions of the world, limited windows of effectiveness and lack of suitability for many patients.^{3,4} Identifying novel therapies for ischaemic stroke is a key research priority.⁵

¹Queensland Research Centre for Peripheral Vascular Disease, James Cook University, Townsville, Australia

²The Australian Institute of Tropical Health and Medicine, James Cook University, Townsville, Australia

³Department of Anatomy, James Cook University, Townsville, Australia ⁴Faculty of Health, University of Technology Sydney, Sydney, Australia ⁵Department of Vascular and Endovascular Surgery, The Townsville Hospital, Townsville, Australia

Corresponding author:

Jonathan Golledge, Queensland Research Centre for Peripheral Vascular Disease, College of Medicine and Dentistry, James Cook University, Townsville, QLD, Australia. Email: jonathan.golledge@jcu.edu.au

The angiopoietins (angpts) are growth factors which play an important role in maintaining vascular function.^{6–9} Angpts bind to the Tie-2 receptor which is predominantly expressed on the vascular endothelium.^{6–9} An observational study previously reported that patients who suffered a recent ischaemic stroke had significantly lower serum angpt-1 concentrations than healthy controls.¹⁰ Patients with serum angpt-1 concentration in the lowest tertile were more likely to have died or have severe disability three months after the stroke than those in the highest tertile.¹⁰ Experimental studies suggest that angpt-1 may improve the outcome of ischaemic stroke in several ways (please see Eklund et al.⁷; Moss⁸; and van Meurs et al.¹¹ for detailed reviews). First, Tie-2 activation has been reported to activate PI3/Akt and inhibit nuclear factor kappalight-chain-enhancer of activated B cells to induce pro-survivor and anti-inflammatory pathways in endothelial cells.^{7,8,11} Angpt-1 activation of the Tie-2 receptor is also important in regulating vascular integrity by maintaining the contacts between neighbouring endothelial cells and the extracellular matrix.7,8,11 Angpt-1 has also been demonstrated to induce the migration of endothelial cells.^{7,8,11} Overall, these studies suggest that angpt-1 inhibits inflammation, promotes angiogenesis and improves endothelial integrity, which may act to limit the severity of cerebral infarction following stroke. Previous studies have examined the effect of upregulating or administering angpt-1 on the outcome of ischaemic stroke in rodent models, but findings have been inconsistent.¹²⁻²² The current study systematically reviewed the efficacy of upregulating or administering angpt-1 in improving recovery from experimentally induced ischaemic stroke in rodent models.

Methods

This systematic review was performed in line with the guidelines of PRISMA and the Systematic Review Centre for Laboratory animal Experimentation statement. A protocol was developed following these guidelines²³ and was registered in the CAMARADES database (date of registration: 21 August 2017).

Search criteria

A systematic literature search was conducted to identify studies investigating the effects of upregulating angpt-1 in rodent ischaemic stroke models (Supplement 1). Eligible studies reported original experiments involving angpt-1 administration or upregulation in a rodent model of acute ischaemic stroke (regardless of dose, method, timing or frequency of administration). Studies also had to provide data on at least one of the following outcome measures: (i) cerebral infarction size; (ii) blood-brain barrier (BBB) permeability; (iii) neurobehavioural outcomes as assessed by a neurological deficit score. Review articles, editorials and publications in languages other than English were excluded. Studies which investigated the effect of angpt-1 in combination with other interventions were excluded unless data on the effects of angpt-1 upregulation alone could be extracted. In cases of ambiguity, authors were contacted to provide additional information. The final literature search was completed in January 2019.

Data extraction and quality assessment

Data extraction and quality assessment was independently conducted by three authors (JVM, AFT and DRN). Information extracted from each study included the species, strain, age and/or body weight of rodents used, duration of cerebral ischaemia, the mechanisms by which angpt-1 was up-regulated, the interventions administered to control animals and the timing of administration relative to stroke induction. Information extracted for each outcome measure included the methods used to assess the outcome, timing of the assessment relative to stroke induction and where possible detailed numerical data to quantify the results, in addition to reported p values from comparison between groups. Behavioural assessments and molecular analyses were not reported in a standard way in the identified studies, and data relating to these outcomes could therefore not be meta-analysed.

The quality and potential bias of the included studies was assessed using a modified CAMARADES score (Supplement 2 and Macleod et al.²⁴). Individual scores were calculated based on the number of 'yes' answers. Studies were allocated a maximum score of 11, a score ≤ 3 was considered to denote a poor-quality study, those with a score of 4–7 were considered to be of moderate quality and high-quality studies achieved scores ≥ 8 . No articles were excluded based on their quality assessment score. Discrepancies in the extracted data were resolved at a consensus meeting.

Data analysis

The primary outcome measure for the meta-analysis was the difference in cerebral infarction size between rodents in which angpt-1 was up-regulated compared to controls. This was chosen since it is a widely employed outcome measure used in rodent models and human studies. Secondary outcome measures were the differences in BBB permeability and neurological function (as assessed using a defined scoring system) between experimental groups. Data were extracted from publications, and, where necessary, were converted to provide common units for comparison between studies. If studies reported findings as mean and standard error, standard deviations were calculated by multiplying the standard error by the square root of the sample size. Where necessary, data were extrapolated from graphs using the analysis function of Photoshop CC 2018 (Adobe Inc, San Jose, CA).

Meta-analysis was conducted when comparable outcome data were available from at least three studies. As high inter-study variation was anticipated, random-effects models were constructed (RevMan V5, Cochrane Collaboration). Effect sizes were calculated as standardised mean difference (SMD) and 95% confidence intervals (CI). Sub-analyses assessing the impact of angpt-1 upregulation on the outcome of stroke were performed in (a) permanent versus temporary cerebral ischaemia models and (b) studies which up-regulated angpt-1 before or after stroke induction. Inter-study variation was assessed using the I^2 statistic. Leaveone-out sensitivity analyses excluding data from individual studies were conducted to assess the robustness of findings. Significance for all analyses was accepted if p values were < 0.05. Funnel plots were generated to assess publication bias.

Results

Characteristics of included studies

The initial database search identified 642 potentially eligible studies. After duplicate removal and title and abstract screening, the full-text of 18 studies was reviewed. After excluding seven studies, 11 were included in the current systematic review (Figure 1; Table 1).^{12–22} Eight studies used rats and three employed mice. Four studies induced stroke by occluding the middle cerebral artery for two hours,^{12,14,17,21} seven induced permanent cerebral ischaemia by inserting a filament into the middle cerebral artery,^{15,16,18,19} or intravenous thrombus injection.^{13,25} Venkat et al. induced embolic cerebral ischaemia but did not specify their methodologies.²⁰

Nine studies tested a single angpt-1 intervention, two^{19,25} tested two angpt-1 upregulation approaches (Table 1; total of 13 interventions). Three of the interventions were virus-based, two were stem-cell-based and eight utilised proteins or peptides. Variations in the dosage, methods of delivery and timing of the angpt-1 interventions were observed. Five interventions were administered prior to stroke induction, seven were administered following cerebral ischaemia induction and the timing of one was unclear (Table 1). Recombinant angpt-1 was co-administered with tissue plasminogen activator in one study,¹³ and animals receiving tissue plasminogen activator plus a control protein were used as controls in the current meta-analysis.

Quality assessment

Three studies were of low quality, six were of moderate quality and two were of high quality (Supplement 2). Only two studies reported randomising rodents to experimental groups,^{20,21} only one provided a sample-size calculation¹³ and only two reported blinding observers.²⁰ No studies provided data regarding the inter/intra-rater reproducibility of the reported outcome measures.

Primary outcome assessment: The effect of angpt-1 upregulation on cerebral infarction size

The effect of angpt-1 upregulation on cerebral infarction size was assessed in 10 studies^{12–21,25} (Table 2 and Supplement 3). Only 9 of the 11 studies meeting inclusion criteria for the systematic review provided data which could be included in the meta-analysis for the primary outcome. These studies (which included 171 rodents) measured cerebral infarction size by histological assessment of excised brains.^{12–15,18–21,25} Of these, three studies expressed the infarction size as a percentage of the cerebral hemisphere and six provided absolute data estimating infarct volume. In the study authored by Shin et al., cerebral infarction size was assessed by MRI, and data could not be accurately extracted from the presented graphs; and therefore this study was excluded from the meta-analysis.¹⁷

Meta-analysis suggested that upregulating angpt-1 significantly reduced cerebral infarction size (SMD for overall effect: -3.02; 95% CI: -4.41, -1.63; *p* < 0.001; $I^2 = 89\%$; Figure 2(a) and (b); number of rodents included in the meta-analysis: 171). Sub-analyses suggested that angpt-1 upregulation reduced infarct size in models of temporary, but not permanent ischaemia (SMD: -8.46; 95% CI: -12.55, -4.37; *p* < 0.001; $I^2 = 89\%$; and SMD: -0.39 (95% CI: -0.84, 0.05; p = 0.08; $I^2 = 0\%$, respectively, Figure 2a). Significant reductions in cerebral infarction sizes were found for studies that up-regulated angpt-1 before or after ischaemia induction (SMD: -1.04; 95% CI: -1.86, -0.21; p = 0.01 and SMD: -5.87, 95% CI: -8.91, -2.82; p < 0.001, respectively; Figure 2b). Funnel plots revealed an asymmetrical distribution of reported SMD, suggesting potential publication bias (Supplement 4). Exclusion of any single study did not alter the overall finding that angpt-1 upregulation significantly reduced cerebral infarction size (Supplement 5). Removal of data reported by Yu et al. did, however, lead to a marked drop in the overall effect size. When the study authored by Meng et al. was excluded, the meta-analysis suggested that upregulating angpt-1 prior to stroke induction did not significantly reduce cerebral infarction size.¹⁴ When the study authored by Yu et al. was excluded, the meta-analysis

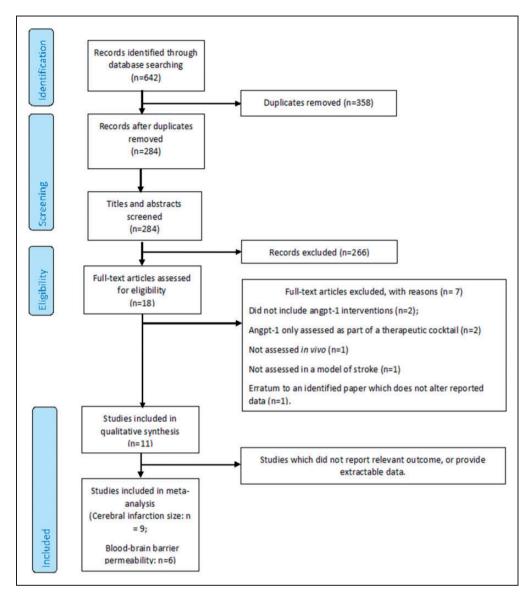


Figure 1. PRISMA diagram outlining the literature search.

suggested that upregulating angpt-1 after stroke induction did not significantly reduce cerebral infarction size (Supplement 5).²¹

Secondary outcome assessment: The effect of angpt-1 upregulation on BBB permeability

Seven of the 11 eligible studies assessed the impact of angpt-1 administration on BBB permeability (Supplement 6).^{12,14,16,19–21,25} Four studies used experimental rats, and three utilised mice. Temporary cerebral ischaemia was induced in three studies,^{12,14,21} while permanent cerebral ischaemia was induced in four studies.^{16,19,20,25} All studies assessed BBB permeability using Evan's blue staining, although the timing between stain administration and tissue collection varied.

Data reported by Valable et al. could not be extracted from the original paper; no response to correspondence was received and data from this study was excluded from the meta-analysis.¹⁹ All other studies reported data as the amount of Evan's blue dye recovered per unit weight of cerebral tissue; however, results differed between studies by several orders of magnitude (Supplement 6).

Meta-analysis of findings from a total of 129 rodents from six studies demonstrated that BBB permeability was significantly lower in animals over-expressing angpt-1 compared to controls (overall SMD: -2.02; 95% CI: -3.27, -0.77; p=0.002; I²: 85%; Figure 3(a) and (b)). Sub-analysis suggested that angpt-1 upregulation was effective in reducing BBB permeability in models of temporary, but not permanent, cerebral

Table I. Charac	Table 1. Characteristics of included studies.				
Lead		Ischaemia		Control intervention	Timing of intervention
author	Rodents used	duration	Method of Angpt-I upregulation	used for this review	relative to stroke induction
Gao ¹²	Male Wistar rats (body weight 250–300 g)	Two hours	lnjection of recombinant protein (Ιμg/kg) into carotid artery	Saline	Upon cerebral reperfusion, and every 12 h thereafter
Shin ¹⁷	Male Sprague-Dawley rats (body weight 260–300g)	Two hours	Transvenous injection of adeno- viral vector delivering a mod- ified angpt-1 (1×10^9) plaque forming units/animal)	Viruses containing LacZ gene	Three days post-stroke induction
Yu ²¹	Male Wistar rats (body weight 250–300 g)	Two hours	lnjection of recombinant protein into right intracerebral artery (Ιμg/kg body weight)	Saline	Post (immediately upon reperfusion)
Meng ¹⁴	Male Sprague-Dawley rats (body weight 180–230g)	Two hours	Intracerebral injection of a lenti- viral vector encoding human angpt1 $(2 \times 10^9$ viral particles/ animal)	Viruses containing empty vector	Two weeks prior to stroke induction
Onda ¹⁵	Male Sprague-Dawley rats (body weight 250–300g)	Permanent	Intravenous infusion of angpt-l over-expressing mesenchymal stem cells	Human mesenchymal stem cells	Six hours post-stroke induction
Venkat ²⁰	Male Wistar rats (type I diabetes induced, body weight 225–250 g, age 8–12 weeks)	Permanent	Intraperitoneal injection of mimetic peptide (3 μg/kg)	Saline	Thirty minutes pre- and 8 and 24h post-stroke induction
Shen ¹⁶	Male CD-1 mice (age and body weight not specified)	Permanent	Intracerebral adenovirus expression	Virus containing LacZ gene	One hour post-stroke induction
Toyama ¹⁸	Male Sprague-Dawley rats (body weight 250–300g)	Permanent	Intravenous infusion of angpt-1 over-expressing human mes- enchymal stem cells $(1 \times 10^6$ cells/animal)	Human mesenchymal stem cells	Six hours post-stroke induction
Valable ¹⁹	OF1 mice (age/body weight and sex not specified)	Permanent	Intracerebroventricular injection of recombinant human angpt- 1 (27 ng/animal) OR Intra-striatal infusion of recom- binant human angpt-1 via osmotic minipump (delivering 0.15 ng/h for seven days)	Saline None	Immediately prior to stroke induction Unclear

Table 1. Characteristics of included studies.

(continued)

Lead author R					
	Rodents used	lschaemia duration	Method of Angpt-1 upregulation	Control intervention used for this review	Timing of intervention relative to stroke induction
Zhang ²⁵ N	Male C57 black 6 mice (body weight 25–30g)	Permanent	Intravenous adenoviral delivery (1 × 10° plaque forming units/ animal) OR intraperitoneal injection of recombinant protein (Bow- Ang) groups (dose not specified)	Adenovirus containing green fluorescent protein Unspecified control protein	Three days prior to stroke induction; Twelve hours prior to stroke induction
Kawamura ¹³ N	Male Sprague Dawley rats (body weight 250–300g)	Permanent	Intravenous injection of COMP- angpt1 protein (30 μg/animal)	COMP (30 µg/animal)	Four hours post-stroke induction (immediately prior to tPA administration)

ischaemia (SMD: -3.74; 95% CI: -6.02, -1.45; p = 0.001; and SMD: -0.65; 95% CI: -1.52, 0.21; p = 0.14, respectively). The significant reduction in BBB permeability was not influenced by the timing of the angpt-1 upregulation (SMD: -0.95; 95% CI: -1.57, -0.34; p = 0.002 and SMD: -3.71; 95% CI: -6.37, -1.04; p = 0.006; for studies that up-regulated angpt-1 before or after stroke, respectively; Figure 3b). Funnel plots generated for this outcome revealed a symmetrical spread of reported SMD suggestive of low publication bias; however, one clear outlier was noted (Supplement 7). Removal of data reported by Yu et al. rendered the overall effect of angpt-1 on BBB permeability nonsignificant (Supplement 8). Exclusion of other studies did not influence the outcome of the overall analysis; however, sub-group analyses became non-significant following removal of the studies by Yu et al., Zhang et al., Venkat et al. and Gao et al. (Supplement 8).

Secondary outcome assessment: The effect of angpt-I upregulation on neurological function

Six studies (96 animals) assessed the effect of angpt-1 upregulation on neurological function in the stroke models (Table 2; Supplement 9) using standardised scoring systems to measure neurological deficit.^{12–14,17,20,21} All studies were performed in rats. Cerebral ischaemia was induced temporarily in four studies,^{12,14,17,21} and permanently in two studies.^{13,20} The timing of outcome assessment relative to cerebral ischaemia varied between studies. Four studies assessed neurological function repeatedly during followup.^{14,17,20,21} All studies utilised different methods to assess the severity of the neurological deficit, thereby prohibiting meta-analysis. Four studies reported that the severity of neurological deficit was significantly less in rats receiving angpt-1 upregulation during at least one of the assessed time points.^{12,14,17,20} The remaining two studies reported no statistically significant differences in neurological function between groups^{13,21} (Table 2; Supplement 9).

Discussion

COMP: cartilage oligomeric matrix protein. tPA: tissue plasminogen activator; angpt-1: angiopoietin 1.

This meta-analysis suggests that angpt-1 upregulation improves outcomes following temporary cerebral ischaemia, evidenced by significant reductions in cerebral infarction size and BBB permeability. Four studies also reported that angpt-1 up-regulation improved post-stroke neurological function; however, this outcome could not be meta-analysed. Care must be taken when interpreting these findings as inter-study heterogeneity in many aspects of study design including the choice of rodent model, methods used to induce cerebral ischaemia and mechanisms to up-regulate angpt-1

				Infarction size			BBB permeability	sability		Neurological function	tion	
	Lead authors	Species	Angpt- I timing	Method	Timing	Reported outcome	Incubation time (h)	Timing	Reported outcome	Methods	Timing	Reported outcomes
Temporary ischaemia	Gao ¹²	ĸ	Post	TTC stain	48 h	\rightarrow	2		→	Five-point scale (no reference cited)	~	~
	Shin ¹⁷	R	Post	ARI	2, 10, 17 and 31 days	, a	I	I	1	Ten-point scale based on Jeong et al. ⁴⁹	Days – 1 (baseline), 3, 6, 10, 17, 24 and 31	~
	Yu ²¹	~	Post	TTC stain	12 h, 48 h and 7 days	12 h: ↓ 48 h: ↓ 7 days: ↓	2	I2h, 48h and 7 days	12 h: ↓ 48 h: ↓ 7 days: ↓	Five-point scale based on Zea and Zuluaga ^{50,b†}	12 h, 48 h and 7 days	12 h: ↔ 48 h: ↔ 7 days: ↔
	Meng ¹⁴	ĸ	Pre	TTC stain	24 h and 14 days	24 h: ↓ I4 days: ↓	22		\rightarrow	Modified score based on Garcia et al. ⁵¹	24h and 14 days	24 h: ↔ 14 days: ↑
Permanent ischaemia	Venkat ²⁰	ĸ	Pre	H&E stain	48 h	\rightarrow	4		\rightarrow	Modified score based on Chen et al. ⁵²	2 h, 24 h, 48 h	2 h: ↑ 24 h: ↑ 48 h: ↑
	Onda ¹⁵	Я	Post	TTC stain	7 days	\$	Ι	I	I	I	I	I
	Toyama ¹⁸	К	Post	TTC stain	7 days	\$	I	I	I	I	I	I
	Zhang ²⁵	Σ	Pre	H&E stain/ TTC stain	24 h	¢. ‡¢	6		Adenovirus: \downarrow Protein: \downarrow	I	I	I
	Kawamura ¹³	۳ ۳	Post	TTC stain	24 h	\$	I	I	I	Six-point scale based on Zausinger et al. ⁵³	24 h	\$
	Valable ¹⁹	Σ	Pre	Thionin stain	24 h	₽¢	ſ		pi	I	I	I
	Shen ¹⁶	Σ	Post	I	I	I	_		\$	I	Ι	I

receiving recombinant protein.

Study or Subgroup 1.1.1 Temporary isch Gao 2015 Meng 2014 (14d) Meng 2014 (24h)	Ar	igpt-1		C	ontrol			Std. Mean Difference	Std. Mean Difference
Gao 2015 Meng 2014 (14d)	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
Meng 2014 (14d)	aemia								
	119	22	10	407	36	10	6.8%	-9.25 [-12.56, -5.93]	
Meng 2014 (24h)	6.5	4.3	5	28.2	5.8	5	8.2%	-3.84 [-6.33, -1.35]	
many wards wards	16.6	5.3	5	29.5	6.8	5	9.6%	-1.91 [-3.55, -0.27]	
Yu 2012 (12h)	160	16	8	387	17	8	4.3%	-13.00 [-18.28, -7.72]	
Yu 2012 (48 h)	123	23	8	423	25	8	4.8%	-11.81 [-16.62, -6.99]	
Yu 2012 (7d) Subtotal (95% CI)	79	4	8 44	290	20	8 44	4.0% 37.7%	-13.83 [-19.44, -8.23] -8.46 [-12.55, -4.37]	-
Heterogeneity: Tau ² = Test for overall effect :					< 0.00	0001);	* = 89%		
1.1.2 Permanent isch	aemia								
Venkat 2018	25.5	14.1	7	35.4	9.3	7	10.4%	-0.78 [-1.88, 0.33]	-+
Kawamura 2014	40.6	21.2	5	40.9	27.7	5	10.2%	-0.01 [-1.25, 1.23]	+
Onda 2008	160.7	19.7	5	173.5	41.9	5	10.2%	-0.35 [-1.61, 0.90]	-
Toyama 2009	155.7	15.8	5	158.4	18.7	5	10.2%	-0.14 [-1.38, 1.10]	+
Valable 2005	95.2	10.3	8	100	9.13	15	10.7%	-0.49 [-1.36, 0.39]	*
Zhang 2002 Subtotal (95% CI)	15.4	3.3	8 38	17.9	7.9	8 45	10.5% 62.3%	-0.39 [-1.38, 0.60] -0.39 [-0.84, 0.05]	+
Heterogeneity: Tau ² = Test for overall effect :				= 5 (P =	0.96);	l ² = 0%			
Total (95% CI)			82			89	100.0%	-3.02 [-4.41, -1.63]	•
Heterogeneity: Tau ² =	4.51; CI	hi ² = 9	7.49, dt	= 11 (F	< 0.00	0001);1	r = 89%		-20 -10 0 10 2
o) Study or Subgroup	Ar	ngpt-1	Total	Mean	ontrol	Total	Weight	Std. Mean Difference IV, Random, 95% CI	Std. Mean Difference IV, Random, 95% CI
Study or Subgroup 2.1.1 Administered pr			Total	mean	30	Total	weight	iv, Kaliuolii, 95% Ci	IV, Raildolli, 95% CI
and the second second second			6	20.2	50	-	0.00	2041022 425	
Meng 2014 (14d)	6.5	4.3	5	28.2	5.8	5	8.2%	-3.84 [-6.33, -1.35]	
Meng 2014 (24h)	16.6	5.3	5	29.5	6.8		9.6%	-1.91 [-3.55, -0.27]	
Valable 2005		10.3	8		9.13	15	10.7%	-0.49 [-1.36, 0.39]	
Venkat 2018	25.5		7	35.4	9.3	7	10.4%	-0.78 [-1.88, 0.33]	
Zhang 2002 Subtotal (95% CI)	15.4	3.3	8 33	17.9	7.9	8 40	10.5% 49.5%	-0.39 [-1.38, 0.60] -1.04 [-1.86, -0.21]	•
Heterogeneity: Tau ² =				= 4 (P =	0.07);	I ² = 549	*		
rest for overall effect.	ter stro	ke							
2.1.2 Administered at				407	36	10	6.8%	-9.25 [-12.56, -5.93]	
2.1.2 Administered at Gao 2015	119	22	10	401					
2.1.2 Administered at	1	22 21.2	10 5		27.7	5	10.2%	-0.01 [-1.25, 1.23]	T
2.1.2 Administered af Gao 2015 Kawamura 2014 Onda 2008	40.6 160.7	21.2 19.7	5 5	40.9 173.5	41.9	5	10.2%	-0.01 [-1.25, 1.23] -0.35 [-1.61, 0.90]	Ŧ
2.1.2 Administered af Gao 2015 Kawamura 2014 Onda 2008	40.6 160.7 155.7	21.2 19.7 15.8	5 5 5	40.9 173.5 158.4	41.9 18.7	5 5	10.2% 10.2%	-0.35 [-1.61, 0.90] -0.14 [-1.38, 1.10]	+
2.1.2 Administered af Gao 2015 Kawamura 2014	40.6 160.7 155.7 160	21.2 19.7 15.8 16	5 5 5 8	40.9 173.5 158.4 387	41.9 18.7 17	5 5 8	10.2% 10.2% 4.3%	-0.35 [-1.61, 0.90] -0.14 [-1.38, 1.10] -13.00 [-18.28, -7.72]	+
2.1.2 Administered at Gao 2015 Kawamura 2014 Onda 2008 Toyama 2009	40.6 160.7 155.7	21.2 19.7 15.8	5 5 5	40.9 173.5 158.4	41.9 18.7	5 5	10.2% 10.2%	-0.35 [-1.61, 0.90] -0.14 [-1.38, 1.10] -13.00 [-18.28, -7.72] -11.81 [-16.62, -6.99]	<u> </u>
2.1.2 Administered at Gao 2015 Kawamura 2014 Onda 2008 Toyama 2009 Yu 2012 (12h)	40.6 160.7 155.7 160	21.2 19.7 15.8 16	5 5 5 8	40.9 173.5 158.4 387	41.9 18.7 17	5 5 8	10.2% 10.2% 4.3%	-0.35 [-1.61, 0.90] -0.14 [-1.38, 1.10] -13.00 [-18.28, -7.72]	<u> </u>
2.1.2 Administered at Gao 2015 Kawamura 2014 Onda 2008 Toyama 2009 Yu 2012 (12h) Yu 2012 (12h) Yu 2012 (7d) Subtotal (95% Cl) Heterogeneity: Tau ² =	40.6 160.7 155.7 160 123 79 13.82; 0	21.2 19.7 15.8 16 23 4 Chi ^z =	5 5 8 8 49 88.05,	40.9 173.5 158.4 387 423 290 df = 6 (F	41.9 18.7 17 25 20	5 5 8 8 49	10.2% 10.2% 4.3% 4.8% 4.0% 50.5%	-0.35 [-1.61, 0.90] -0.14 [-1.38, 1.10] -13.00 [-18.28, -7.72] -11.81 [-16.62, -6.99] -13.83 [-19.44, -8.23]	<u> </u>
2.1.2 Administered at Gao 2015 Kawamura 2014 Onda 2008 Toyama 2009 Yu 2012 (12h) Yu 2012 (12h) Yu 2012 (48 h) Yu 2012 (7d) Subtotal (95% Cl) Heterogeneity: Tau ² = Test for overall effect 1	40.6 160.7 155.7 160 123 79 13.82; 0	21.2 19.7 15.8 16 23 4 Chi ^z =	5 5 8 8 49 88.05,	40.9 173.5 158.4 387 423 290 df = 6 (F	41.9 18.7 17 25 20	5 5 8 8 49 0001);1	10.2% 10.2% 4.3% 4.8% 4.0% 50.5%	-0.35 [-1.61, 0.90] -0.14 [-1.38, 1.10] -13.00 [-18.28, -7.72] -11.81 [-16.62, -6.99] -13.83 [-19.44, -8.23]	<u> </u>
Gao 2015 Kawamura 2014 Onda 2008 Toyama 2009 Yu 2012 (12h) Yu 2012 (48 h) Yu 2012 (7d) Subtotal (95% CI)	40.6 160.7 155.7 160 123 79 13.82; (Z= 3.78	21.2 19.7 15.8 16 23 4 Chi ² = (5 5 8 8 49 88.05, 0.0002) 82	40.9 173.5 158.4 387 423 290 df = 6 (F	41.9 18.7 17 25 20 < 0.00	5 5 8 8 49 0001);1	10.2% 10.2% 4.3% 4.8% 50.5% F = 93%	-0.35 [-1.61, 0.90] -0.14 [-1.38, 1.10] -13.00 [-18.28, -7.72] -11.81 [-16.62, -6.99] -13.83 [-19.44, -8.23] -5.87 [-8.91, -2.82]	<u> </u>

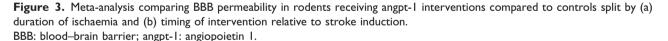
Figure 2. Meta-analysis comparing cerebral infarction size in rodents receiving angpt-1 interventions compared to controls split by (a) duration of ischaemia and (b) timing of intervention relative to stroke induction.

was observed. Only two of the 11 included studies were considered to be of high quality.^{13,21} Funnel plots suggested potential publication bias for the primary outcome (cerebral infarction size), however, this was difficult to objectively assess as relatively few studies were included. Finally, only male rodents were included

in the identified studies. It is therefore unclear whether findings of the current meta-analysis can be generalised to females.

Cerebral infarction size was the primary outcome measure for this meta-analysis as it was the most widely reported outcome measure, and has

l)		ngpt-1		C	ontrol			Std. Mean Difference	Std. Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% Cl
2.1.1 Temporary ischaemia	a								
Gao 2015	98.6	10.6	10	379.9	21.6	10	3.7%	-15.84 [-21.38, -10.29]	
Meng 2014	16.685	1.283	5	17.368	1.341	5	12.3%	-0.47 [-1.74, 0.80]	+
Yu 2012 (12h)	0.26	0.03	8	0.42	0.08	8	11.9%	-2.50 [-3.90, -1.10]	+
Yu 2012 (48h)	0.32	0.06	8	0.49	0.04	8	11.4%	-3.15 [-4.75, -1.56]	+
Yu 2012 (7d) Subtotal (95% CI)	0.13	0.03	8 39	0.26	0.04	8 39	11.1% 50.4%	-3.48 [-5.17, -1.78] -3.74 [-6.02, -1.45]	*
Heterogeneity: Tau ² = 5.47; Test for overall effect: Z = 3.3			4 (P <	0.00001;); l² = 88	%			
2.1.2 Permanent ischaemia	a								
Zhang 2002 (adenovirus)	and the second second	0.196	6	0.461	0.141	8	12.5%	-1.16 [-2.33, 0.02]	-
Zhang 2002 (protein)		0.106	7		0.442	7	12.5%	-1.21 [-2.39, -0.04]	+
Venkat 2018	17.5	5.3	4	30.6	16.1		12.1%	-0.89 [-2.20, 0.43]	-
Shen 2011	0.9	0.3	6	0.7	0.3	6	12.5%	0.62 [-0.56, 1.79]	-
Subtotal (95% CI)	0.5	0.5	23	0.7	0.5	28	49.6%	-0.65 [-1.52, 0.21]	•
Heterogeneity: Tau ^a = 0.40; Test for overall effect: Z = 1.			3 (P = 0	.10); I ² =	52%				
Total (95% CI)			62			67	100.0%	-2.02 [-3.27, -0.77]	•
Heterogeneity: Tau ² = 2.89;	Chi ² = 53	.05, df=	8 (P <	0.00001); 2 = 85	%			1 1 1 1
	17/0-0	000							-20 -10 0 10 20
Test for subaroup difference			f=1 (P	= 0.01),	* = 83.6	%			Favours angpt-1 Favours control.
Test for subaroup difference	es: Chi ² = S	6.10, dt		c	ontrol			Std. Mean Difference	Std. Mean Difference
Test for overall effect: Z = 3: Test for subgroup difference) Study or Subgroup	es: Chi ² = S Mean	6.10, dt	f=1 (P Total		ontrol		Weight	Std. Mean Difference IV, Random, 95% CI	
Test for subgroup difference) Study or Subgroup 3.1.1 Administered prior to	stroke	6.10, d troke SD	Total	C Mean	ontrol SD	Total		IV, Random, 95% CI	Std. Mean Difference
Test for subgroup difference) Study or Subgroup 3.1.1 Administered prior to Meng 2014	es: Chi ² = S Mean stroke 16.685	6.10, dt troke SD 1.283	Total 5	Mean 17.368	ontrol SD 1.341	Total 5	12.3%	IV, Random, 95% CI -0.47 [-1.74, 0.80]	Std. Mean Difference
Test for subgroup difference) Study or Subgroup 3.1.1 Administered prior to Meng 2014 Venkat 2018	stroke 16.685	6.10, dt troke SD 1.283 5.3	Total 5 4	C Mean 17.368 30.6	ontrol SD 1.341 16.1	Total 5 7	12.3% 12.1%	IV, Random, 95% CI -0.47 [-1.74, 0.80] -0.89 [-2.20, 0.43]	Std. Mean Difference
Test for subaroup difference) Study or Subgroup 3.1.1 Administered prior to Meng 2014 Venkat 2018 Zhang 2002 (adenovirus)	es: Chi ² = S Mean stroke 16.685 17.5 0.256	6.10, dt troke SD 1.283 5.3 0.196	Total 5 4 6	C Mean 17.368 30.6 0.461	0ntrol SD 1.341 16.1 0.141	Total 5 7 8	12.3% 12.1% 12.5%	IV, Random, 95% CI -0.47 [-1.74, 0.80] -0.89 [-2.20, 0.43] -1.16 [-2.33, 0.02]	Std. Mean Difference
Test for subgroup difference) Study or Subgroup 3.1.1 Administered prior to Meng 2014 Venkat 2018 Zhang 2002 (adenovirus)	es: Chi ² = S Mean stroke 16.685 17.5 0.256	6.10, dt troke SD 1.283 5.3	Total 5 4	C Mean 17.368 30.6 0.461	ontrol SD 1.341 16.1	Total 5 7	12.3% 12.1% 12.5% 12.5%	IV, Random, 95% CI -0.47 [-1.74, 0.80] -0.89 [-2.20, 0.43]	Std. Mean Difference
Test for subgroup difference) <u>Study or Subgroup</u> 3.1.1 Administered prior to Meng 2014 Venkat 2018 Zhang 2002 (adenovirus) Zhang 2002 (protein) Subtotal (95% CI) Heterogeneity: Tau ^a = 0.00;	es: Chi ² = S Mean stroke 16.685 17.5 0.256 0.131 Chi ² = 0.8	6.10, dt troke SD 1.283 5.3 0.196 0.106 37, df = 3	Total 5 4 6 7 22	C Mean 17.368 30.6 0.461 0.548	00000000000000000000000000000000000000	Total 5 7 8 7	12.3% 12.1% 12.5% 12.5%	IV, Random, 95% CI -0.47 [-1.74, 0.80] -0.89 [-2.20, 0.43] -1.16 [-2.33, 0.02] -1.21 [-2.39, -0.04]	Std. Mean Difference
Test for subgroup difference 3.1.1 Administered prior to Meng 2014 Venkat 2018 Zhang 2002 (adenovirus) Zhang 2002 (protein) Subtotal (95% CI) Heterogeneity: Tau ^s = 0.00; Test for overall effect. Z = 3.0	es: Chi ^p = <u>S</u> <u>Mean</u> 16.885 17.5 0.256 0.131 Chi ^p = 0.8 04 (P = 0.	6.10, dt troke SD 1.283 5.3 0.196 0.106 37, df = 3	Total 5 4 6 7 22	C Mean 17.368 30.6 0.461 0.548	00000000000000000000000000000000000000	Total 5 7 8 7	12.3% 12.1% 12.5% 12.5%	IV, Random, 95% CI -0.47 [-1.74, 0.80] -0.89 [-2.20, 0.43] -1.16 [-2.33, 0.02] -1.21 [-2.39, -0.04]	Std. Mean Difference
Test for subgroup difference Study or Subgroup 3.1.1 Administered prior to Meng 2014 Venkat 2018 Zhang 2002 (adenovirus) Zhang 2002 (protein) Subtotal (95% CI) Heterogeneity: Tau ^a = 0.00;	es: Chi ^p = <u>S</u> <u>Mean</u> 16.885 17.5 0.256 0.131 Chi ^p = 0.8 04 (P = 0.	6.10, dt troke SD 1.283 5.3 0.196 0.106 87, df = 1 002)	Total 5 4 6 7 22	C Mean 17.368 30.6 0.461 0.548	00000000000000000000000000000000000000	Total 5 7 8 7	12.3% 12.1% 12.5% 12.5% 49.3%	IV, Random, 95% CI -0.47 [-1.74, 0.80] -0.89 [-2.20, 0.43] -1.16 [-2.33, 0.02] -1.21 [-2.39, -0.04] -0.95 [-1.57, -0.34]	Std. Mean Difference
Test for subgroup difference) Study or Subgroup 3.1.1 Administered prior to Meng 2014 Venkat 2018 Zhang 2002 (adenovirus) Zhang 2002 (protein) Subtotal (95% CI) Heterogeneity: Tau ^a = 0.00; Test for overall effect: Z = 3.1 3.1.2 Administered after st	es: Chi ^p = <u>S</u> <u>Mean</u> stroke 16.885 17.5 0.256 0.131 Chi ^p = 0.6 04 (P = 0. troke	6.10, dt troke SD 1.283 5.3 0.196 0.106 87, df = 1 002)	Total 5 4 6 7 22 3 (P = 0	C Mean 17.368 30.6 0.461 0.548 83); P =	00000000000000000000000000000000000000	Total 5 7 8 7 27	12.3% 12.1% 12.5% 12.5% 49.3%	IV, Random, 95% CI -0.47 [-1.74, 0.80] -0.89 [-2.20, 0.43] -1.16 [-2.33, 0.02] -1.21 [-2.39, -0.04]	Std. Mean Difference
Test for subgroup difference) <u>Study or Subgroup</u> 3.1.1 Administered prior to Meng 2014 Venkat 2018 Zhang 2002 (adenovirus) Zhang 2002 (protein) Subtotal (95% CI) Heterogeneity: Tau ^a = 0.00; Test for overall effect: Z = 3.0 3.1.2 Administered after st Gao 2015 Shen 2011	es: Chi ^p = <u>Mean</u> stroke 16.685 17.5 0.256 0.131 Chi ^p = 0.8 04 (P = 0. troke 98.6	6.10, d troke SD 1.283 5.3 0.196 0.106 37, df = 3 002) 10.6	Total 5 4 6 7 22 3 (P = 0 10	C Mean 17.368 30.6 0.461 0.548 .83); I ^a = 379.9	00000000000000000000000000000000000000	Total 5 7 8 7 27	12.3% 12.1% 12.5% 12.5% 49.3%	IV, Random, 95% CI -0.47 [-1.74, 0.80] -0.89 [-2.20, 0.43] -1.16 [-2.33, 0.04] -1.21 [-2.39, -0.04] -0.95 [-1.57, -0.34]	Std. Mean Difference
Test for subgroup difference) Study or Subgroup 3.1.1 Administered prior to Meng 2014 Venkat 2018 Zhang 2002 (adenovirus) Zhang 2002 (protein) Subtotal (95% Cl) Heterogeneity: Tau ^a = 0.00; Test for overall effect: Z = 3.1 3.1.2 Administered after st Gao 2015	es: Chi ^p = <u>S</u> <u>Mean</u> stroke 16.685 17.5 0.256 0.131 Chi ^p = 0.8 04 (P = 0. roke 98.6 0.9	6.10, dt troke SD 1.283 5.3 0.196 0.106 37, df = 1 002) 10.6 0.3	Total 5 4 6 7 22 3 (P = 0 10 6	C Mean 17.368 30.6 0.461 0.548 83); I ^a = 379.9 0.7	00000000000000000000000000000000000000	Total 5 7 8 7 27 10 6	12.3% 12.1% 12.5% 12.5% 49.3%	IV, Random, 95% CI -0.47 [-1.74, 0.80] -0.89 [-2.20, 0.43] -1.16 [-2.33, 0.02] -1.21 [-2.39, -0.02] -0.95 [-1.57, -0.34] -15.84 [-21.38, -10.29] 0.62 [-0.56, 1.79]	Std. Mean Difference
Test for subgroup difference) Study or Subgroup 3.1.1 Administered prior to Meng 2014 Venkat 2018 Zhang 2002 (adenovirus) Zhang 2002 (adenovirus) Zhang 2002 (arotein) Subtotal (95% Cl) Heterogeneity: Tau ^a = 0.00; Test for overall effect: Z = 3.1 3.1.2 Administered after st Gao 2015 Shen 2011 Yu 2012 (12h) Yu 2012 (48h) Yu 2012 (7d)	es: Chi ^p = <u>S</u> <u>Mean</u> stroke 16.685 17.5 0.256 0.131 Chi ^p = 0.6 04 (P = 0. roke 98.6 0.9 0.26	6.10, d troke <u>SD</u> 1.283 5.3 0.196 0.106 87, df = : 002) 10.6 0.3 0.03	Total 5 4 6 7 22 3 (P = 0 10 6 8	C Mean 17.368 30.6 0.461 0.548 .83); I ^a = 379.9 0.7 0.42	00000000000000000000000000000000000000	Total 5 7 8 7 27 10 6 8	12.3% 12.1% 12.5% 12.5% 49.3% 3.7% 12.5% 11.9%	IV, Random, 95% CI -0.47 [-1.74, 0.80] -0.89 [-2.20, 0.43] -1.16 [-2.33, 0.02] -1.21 [-2.39, -0.04] -0.95 [-1.57, -0.34] -15.84 [-21.38, -10.29] 0.62 [-0.56, 1.79] -2.50 [-3.90, -1.10] -3.16 [-4.75, -1.56] -3.48 [-5.17, -1.78]	Std. Mean Difference
Test for subgroup difference Study or Subgroup 3.1.1 Administered prior to Meng 2014 Venkat 2018 Zhang 2002 (adenovirus) Zhang 2002 (protein) Subtotal (95% Cl) Heterogeneity: Tau ^a = 0.00; Test for overall effect: Z = 3.1 3.1.2 Administered after st Gao 2015 Shen 2011 Yu 2012 (12h) Yu 2012 (12h) Yu 2012 (7d) Subtotal (95% Cl)	es: Chi ^p = <u>S</u> <u>Mean</u> stroke 16.685 17.5 0.256 0.131 Chi ^p = 0.8 04 (P = 0. troke 98.6 0.9 0.26 0.32 0.13	6.10, d troke SD 1.283 5.3 0.196 0.106 37, df = 3 0.02) 10.6 0.3 0.03 0.03 0.03	Total 5 4 6 7 22 3 (P = 0 10 6 8 8 8 8 40	C Mean 17.368 30.6 0.461 0.548 .83); I ^a = 379.9 0.7 0.42 0.49 0.26	21.6 0.3 0.04 0.04 0.04	Total 5 7 27 10 6 8 8 8 8 40	12.3% 12.1% 12.5% 12.5% 49.3% 3.7% 12.5% 11.9% 11.4% 11.1%	IV, Random, 95% CI -0.47 [-1.74, 0.80] -0.89 [-2.20, 0.43] -1.16 [-2.33, 0.02] -1.21 [-2.39, -0.04] -0.95 [-1.57, -0.34] -15.84 [-21.38, -10.29] 0.62 [-0.56, 1.79] -2.50 [-3.90, -1.10] -3.15 [-4.75, -1.56]	Std. Mean Difference
Test for subgroup difference) <u>Study or Subgroup</u> 3.1.1 Administered prior to Meng 2014 Venkat 2018 Zhang 2002 (adenovirus) Zhang 2002 (arotein) Subtotal (95% Cl) Heterogeneity: Tau ^a = 0.00; Test for overall effect: Z = 3.1 3.1.2 Administered after st Gao 2015 Shen 2011 Yu 2012 (12h) Yu 2012 (12h) Yu 2012 (48h) Yu 2012 (7d) Subtotal (95% Cl) Heterogeneity: Tau ^a = 7.80;	es: Chi ^p = <u>S</u> <u>Mean</u> stroke 16.685 17.5 0.256 0.131 Chi ^p = 0.8 04 (P = 0. 98.6 0.9 0.26 0.32 0.32 0.13 Chi ^p = 48	6.10, d troke SD 1.283 5.3 0.196 0.106 87, df = : 002) 10.6 0.3 0.03 0.06 0.03 0.03 0.06 0.03 1.38, df =	Total 5 4 6 7 22 3 (P = 0 10 6 8 8 8 8 40	C Mean 17.368 30.6 0.461 0.548 .83); I ^a = 379.9 0.7 0.42 0.49 0.26	21.6 0.3 0.04 0.04 0.04	Total 5 7 27 10 6 8 8 8 8 40	12.3% 12.1% 12.5% 12.5% 49.3% 3.7% 12.5% 11.9% 11.4% 11.1%	IV, Random, 95% CI -0.47 [-1.74, 0.80] -0.89 [-2.20, 0.43] -1.16 [-2.33, 0.02] -1.21 [-2.39, -0.04] -0.95 [-1.57, -0.34] -15.84 [-21.38, -10.29] 0.62 [-0.56, 1.79] -2.50 [-3.90, -1.10] -3.16 [-4.75, -1.56] -3.48 [-5.17, -1.78]	Std. Mean Difference
Test for subgroup difference) Study or Subgroup 3.1.1 Administered prior to Meng 2014 Venkat 2018 Zhang 2002 (adenovirus) Zhang 2002 (adenovirus) Zhang 2002 (protein) Subtotal (95% CI) Heterogeneity: Tau ^a = 0.00; Test for overall effect: Z = 3.1 3.1.2 Administered after st Gao 2015 Shen 2011 Yu 2012 (12h) Yu 2012 (48h) Yu 2012 (48h) Yu 2012 (7d) Subtotal (95% CI) Heterogeneity: Tau ^a = 7.80; Test for overall effect: Z = 2.1	es: Chi ^p = <u>S</u> <u>Mean</u> stroke 16.685 17.5 0.256 0.131 Chi ^p = 0.8 04 (P = 0. 98.6 0.9 0.26 0.32 0.32 0.13 Chi ^p = 48	6.10, d troke SD 1.283 5.3 0.196 0.106 87, df = : 002) 10.6 0.3 0.03 0.06 0.03 0.03 0.06 0.03 1.38, df =	Total 5 4 6 7 22 3 (P = 0 10 6 8 8 8 8 40	C Mean 17.368 30.6 0.461 0.548 .83); I ^a = 379.9 0.7 0.42 0.49 0.26	21.6 0.3 0.04 0.04 0.04	Total 5 7 8 7 27 10 6 8 8 8 8 40 %	12.3% 12.1% 12.5% 12.5% 49.3% 3.7% 12.5% 11.9% 11.4% 11.1%	IV, Random, 95% CI -0.47 [-1.74, 0.80] -0.89 [-2.20, 0.43] -1.16 [-2.33, 0.02] -1.21 [-2.39, -0.04] -0.95 [-1.57, -0.34] -15.84 [-21.38, -10.29] 0.62 [-0.56, 1.79] -2.50 [-3.90, -1.10] -3.16 [-4.75, -1.56] -3.48 [-5.17, -1.78]	Std. Mean Difference
Test for subgroup difference) Study or Subgroup 3.1.1 Administered prior to Meng 2014 Venkat 2018 Zhang 2002 (adenovirus) Zhang 2002 (adenovirus) Zhang 2002 (arotein) Subtotal (95% CI) Heterogeneity: Tau ^a = 0.00; Test for overall effect: Z = 3.1 3.1.2 Administered after st Gao 2015 Shen 2011 Yu 2012 (12h) Yu 2012 (48h) Yu 2012 (48h) Yu 2012 (48h) Yu 2012 (7d) Subtotal (95% CI) Heterogeneity: Tau ^a = 7.80; Test for overall effect: Z = 2.1 Total (95% CI)	es: Chi ^p = <u>S</u> <u>Mean</u> stroke 16.685 17.5 0.256 0.256 0.4 (P = 0.8 0.4 (P = 0.8 0.9 0.26 0.9 0.26 0.32 0.13 Chi ^p = 48 73 (P = 0.8	6.10, d troke SD 1.283 5.3 0.196 0.106 0.106 0.37, df = : 002) 10.6 0.3 0.05 0.03 0.05 0.03 0.05 0.03 0.05 0.03 0.05 0.03 0.05 0.03 0.05 0.5 0.	Total 5 4 6 7 22 3 (P = 0 10 6 8 8 8 8 8 40 4 (P < 62	C Mean 17.368 30.6 0.461 0.548 .83); I ^a = 379.9 0.7 0.42 0.49 0.26 0.00001)	21.6 0.3 0.08 0.04 0.04 0.04 0.04 0.04 0.04 0.04	Total 5 7 8 7 27 10 6 8 8 8 8 40 %	12.3% 12.1% 12.5% 12.5% 49.3% 3.7% 12.5% 11.9% 11.4% 50.7%	IV, Random, 95% CI -0.47 [-1.74, 0.80] -0.89 [-2.20, 0.43] -1.16 [-2.33, 0.02] -1.21 [-2.39, -0.04] -0.95 [-1.57, -0.34] -15.84 [-21.38, -10.29] 0.62 [-0.56, 1.79] -2.50 [-3.90, -1.10] -3.15 [-4.75, -1.56] -3.48 [-5.17, -1.78] -3.71 [-6.37, -1.04]	Std. Mean Difference IV, Random, 95% CI
Test for subgroup difference) Study or Subgroup 3.1.1 Administered prior to Meng 2014 Venkat 2018 Zhang 2002 (adenovirus) Zhang 2002 (adenovirus) Zhang 2002 (protein) Subtotal (95% Cl) Heterogeneity: Tau ^a = 0.00; Test for overall effect: Z = 3.1 3.1.2 Administered after st Gao 2015 Shen 2011 Yu 2012 (12h) Yu 2012 (48h) Yu 2012 (7d)	es: Chi ^P = <u>S</u> <u>Mean</u> stroke 16.685 17.5 0.256 0.131 Chi ^P = 0.8 04 (P = 0. 104 (P	6.10, d troke SD 1.283 5.3 0.196 0.106 0.7, df = : 002) 10.6 0.3 0.03 0.03 0.03 0.03 0.03 0.03 0.03 0.05, df =	Total 5 4 6 7 22 3 (P = 0 10 6 8 8 8 8 8 40 4 (P < 62	C Mean 17.368 30.6 0.461 0.548 .83); I ^a = 379.9 0.7 0.42 0.49 0.26 0.00001)	21.6 0.3 0.08 0.04 0.04 0.04 0.04 0.04 0.04 0.04	Total 5 7 8 7 27 10 6 8 8 8 8 40 %	12.3% 12.1% 12.5% 12.5% 49.3% 3.7% 12.5% 11.9% 11.4% 50.7%	IV, Random, 95% CI -0.47 [-1.74, 0.80] -0.89 [-2.20, 0.43] -1.16 [-2.33, 0.02] -1.21 [-2.39, -0.04] -0.95 [-1.57, -0.34] -15.84 [-21.38, -10.29] 0.62 [-0.56, 1.79] -2.50 [-3.90, -1.10] -3.15 [-4.75, -1.56] -3.48 [-5.17, -1.78] -3.71 [-6.37, -1.04]	Std. Mean Difference



translational potential since cerebral infarction volume has been linked with stroke severity in clinical investigations.²⁶ Preserving hypo-perfused brain tissue following ischaemic stroke relies on rapid restoration of the cerebral blood supply before irreversible damage occurs²⁷; clinically, this can be achieved through chemical and/or mechanical thrombolysis.^{2,28} Data from the current meta-analysis suggest that angpt-1 upregulation may augment endogenous physiological responses. Neovascularisation, through angiogenesis or arteriogenesis, occurs after stroke in order to restore blood supply to the ischaemic tissue.²⁹ Angpt-1 is a key regulator of neovascularisation and stabilises blood vessels during formation and remodelling.^{7,30} Genetic data suggest a role for angpt-1 in the response to cerebral ischaemia as single nucleotide polymorphisms in the *angpt-1* gene have been associated with the severity of cerebral infarction in both mice and humans.³¹ Secondary analyses presented by Meng et al. and Shin et al. suggest that angpt-1 increases the survival and proliferation of endothelial and neuronal cells and promotes neuronal cell differentiation.^{14,17} Toyama et al. also reported that cerebral capillary density within the infarction boundary zone was significantly higher in

upregulation on cerebral infarction size may be due to enhanced revascularisation and neuronal protection. Ischaemic stroke also increases BBB permeability, and this has been suggested to promote inflammation and oedema formation.^{36,37} Clinical studies suggest that patients with the most severe BBB disruption after ischaemic stroke experience poorer functional outcomes, than those with greater BBB integrity.^{36,38} Importantly, cerebral reperfusion by mechanical or chemical thrombolysis promotes cerebral oedema.^{4,13} The ability of angpt-1 to reduce BBB leakiness after temporary cerebral ischaemia may therefore have clinical importance.

Collectively, this suggests that the benefit of angpt-1

The BBB comprises vascular endothelial cells, neurons and pericytes supported by an extracellular matrix.³⁹ Following ischaemia, BBB integrity is lost, although the mechanisms causing this are incompletely understood. Traditionally, stroke-related increases in vascular permeability are thought to result from disassembly of tight- and adherens-junctions between neighbouring endothelial cells, and angpt-1 has been demonstrated to play an important role in maintaining these junctions.^{7,8,37,39-41} Two studies included in the current review^{16,21} reported that cerebral expression of tight junction proteins such as ZO-1 and occludin was significantly higher in animals receiving angpt-1 than controls. Thus, the observed reduction in BBB permeability may be due to preservation of tight- and adherens-junctions between endothelial cells.

The severity of neurological dysfunction immediately following ischaemic stroke is strongly predictive of outcome at one year.⁴² It has been recommended that animal stroke model research should assess neurological function.^{43,44} Thus, it is surprising that the effect of angpt-1 upregulation on the severity of neurological deficits was the least widely studied outcome. Moreover, there was marked heterogeneity in the methods used to measure neurological dysfunction, complicating overall interpretation of the findings; and similar heterogeneity has been noted in clinical studies.45,46 Assessing other aspects of physical function such as behaviour or exercise capacity may also further contextualise neurological function outcomes; however, only four of the included studies presented such data.^{14,15,17,18} As with the neurological data, interstudy heterogeneity in the manner by which behavioural assessments were conducted and reported prevented detailed analysis of these outcomes. Further studies which directly assess the impact of angpt-1 upregulation on neurological and behavioural capacity are needed to provide greater insight into the translational potential of this intervention.

In conclusion, this meta-analysis suggests that angpt-1 upregulation significantly reduces cerebral infarction size and BBB leakiness following temporary cerebral ischaemia. This systematic review highlights the need for high-quality rodent model research that includes standardised assessment of neurological function and incorporates key clinical aspects, such as upregulating angpt-1 after (not before) stroke induction and including patient co-morbidities, within the experimental design.⁴⁷ Adoption of the Animal Research: Reporting of In Vivo Experiments and Stroke Therapy Academic Industry Roundtable guidelines will also improve the quality of evidence produced from future rodent studies.^{44,48}

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Supplemental material

Supplemental material for this paper can be found at the journal website: http://journals.sagepub.com/home/jcb

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