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Association of outdoor temperature with lung function in a temperate climate

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Abstract

Acute exposure to cold dry air is a trigger of bronchoconstriction, but little is known about how daily outdoor temperature influences lung function.

We investigated associations of temperature from a model using satellite remote sensing data with repeated measures of lung function among 5896 participants of the Framingham Heart Study Offspring and Third Generation cohorts residing in the Northeastern US. We further tested if temperature modified previously reported associations between pollution and lung function. We

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A higher temperature in the preceding days is associated with lower lung function, especially during the spring and winter but not in summer. These findings may be explained by more time spent outdoors on warmer days in a temperate climate. http://ow.ly/MwIE30msEWk

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constructed linear mixed-effects models, and assessed departures from linearity using penalised splines.

In fully adjusted linear models, 1-, 2- and 7-day average temperatures were all associated with lower lung function: each 5°C higher previous-week temperature was associated with a 20 mL lower (95% CI -34—6) forced expiratory volume in 1 s. There was significant effect modification by season: negative associations of temperature and lung function were present in winter and spring only. Negative associations between previous-day fine particulate matter and lung function were present during unseasonably warm but not unseasonably cool days, with a similar pattern for other pollutants.

We speculate that temperature-related differences in lung function may be explained by behavioural changes on relatively warm days, which may increase outdoor exposures.

Introduction

Climate change is increasing the magnitude of forest fire air pollution, heat-related ozone pollution, pollen production and other environmental exposures that are known to be harmful to respiratory health [1, 2]. As average temperatures rise, there is a need to quantify relationships between weather conditions and respiratory health. Cold and hot weather are both associated with more hospital visits for respiratory disease, with most studies in temperate climates finding U- or J-shaped relationships between temperature and respiratory hospitalisations [3–6]. Controlled human exposure studies provide evidence of transient increases in airway resistance after brief periods of intense exposure to cold [7], and also to heat, especially among people with asthma [7–9], but these studies do not mimic real-world daily exposure to outdoor weather conditions. Recent studies of elderly men and people with asthma have found that warmer temperatures are associated with lower lung function [10, 11], but it is unclear if the general population experiences differences in lung function in association with average outdoor temperature. Studies examining residence-specific temperature and lung function are lacking.

There is also little understanding of whether outdoor temperature modifies the known respiratory effects of exposure to air pollutants. A large number of studies, including prior work in the Framingham Heart Study, have found that higher levels of outdoor air pollutants, including fine particulate matter (particles with a diameter of <2.5 μ m (PM2.5)) and ozone (O₃), are associated with lower lung function [12–14]. One study of healthy individuals in China found that associations between PM2.5 exposure and forced expiratory volume in 1 s (FEV₁) were larger in the setting of higher temperatures [15], and a recent study of elderly men in the Northeastern US also found an interaction between warmer temperature and black carbon (a constituent of PM2.5) and lower lung function [10]. Studies of mortality have suggested that the mortality effects of air pollution, including PM2.5 and O₃, are greater in the setting of warmer temperatures [16–18]. However, despite the wealth of evidence that air pollution affects lung function, and evidence of higher respiratory hospitalisation at the extremes of temperature, interactions between outdoor temperature and pollution on lung function remain largely unexplored.

We estimated daily outdoor temperature at the home address among participants of the Offspring and Third Generation Cohorts of the Framingham Heart Study residing in the Northeastern US with a novel, validated spatiotemporal temperature model using satellitebased surface temperature measurements [19, 20]. We hypothesised that colder and warmer outdoor temperature would be associated with lower lung function in a non-linear pattern and, based on published findings of the relationship between temperature and mortality [21], that the magnitude of the effects of temperature on lung function would vary by season, owing to physiologic adaptation. We further hypothesised that the associations between outdoor pollution and lung function that we previously reported in this cohort [12] would be greater in the setting of warmer outdoor temperatures.

Methods

Study population

The study population consisted of the participants in the Framingham Heart Study Offspring and Third Generation cohorts living within the Northeastern US, including New England (Massachusetts, New Hampshire, Rhode Island, Maine) and the Tri-State area (New York, New Jersey, Connecticut). The design of these studies has previously been described [22, 23]. Offspring participants with at least one spirometry measurement after the year 2000 (when satellite-based estimates of temperature became available) at examination 7 (2000-2001) or 8 (2005–2008) and Third Generation participants with at least one measurement at examination 1 (2002–2005) or 2 (2008–2011) were included. The proportion of participants completing spirometry was 72% and 85% at Offspring examinations 7 and 8, and 94% and 95% at Third Generation examinations 1 and 2. At each examination, data were collected on demographics, medication use, smoking history and respiratory symptoms and diagnoses. There were 5896 participants with temperature data at their home address, providing 9150 spirometry measurements in our primary analysis. All participants provided written informed consent for the study examinations, and the institutional review boards of Beth Israel Deaconess Medical Center and Boston University Medical Campus approved this work.

Temperature exposure

We assigned daily temperature at the home address to each participant with a prediction model that uses satellite surface temperature data to estimate daily ambient air temperature, beginning in the year 2000. A mixed-model approach was used to calibrate daily satellite surface temperature obtained from the Moderate Resolution Imaging Spectroradiometer satellite (Terra) at 1×1 km resolution against air temperature values obtained from the National Climatic Data Center, Environmental Protection Agency (EPA) and Weather Underground Inc., as previously described [24]. Air temperature predictors (per cent urban, elevation, normalised difference vegetation index) and a random intercept and random slope were used to calibrate the model for each day. Because surface temperature values from the satellite are often missing owing to cloud cover or retrieval errors, an additional model was fit that took advantage of the association of grid cells' surface temperature values with air temperature monitoring located elsewhere, and the association with surface temperature values in neighbouring grid cells. Out-of-sample "ten-fold" cross-validation was used to

quantify the accuracy of the model's temperature predictions [24]. Model performance was excellent for days both with and without available satellite data (mean out-of-sample R^2 =0.947 and R^2 = 0.940, respectively).

Using these model predictions for each participant's geocoded home address, we calculated average temperature for the 1-, 2- and 7-day average of temperature including the day of the spirometry test, based on our *a priori* hypothesis that short-term temperature exposure would have the strongest associations with lung function. Hourly relative humidity data were obtained from the Boston Airport weather station (12 km from the central air pollution monitoring site).

Air pollution exposure

We used hourly PM2.5 concentration data from the Harvard Boston Supersite in Boston, MA, USA, five stories above ground level and 50 m from the nearest street. PM2.5 concentrations were measured using a tapered element oscillating microbalance (model 1400A; Rupprecht & Pataschnick Co., Albany, NY, USA). The gaseous pollutants O_3 and nitrogen dioxide (NO₂) were measured from state monitors in Eastern Massachusetts, and exposures were estimated by averaging data from available sites. By convention, O_3 exposure was only examined during the EPA monitoring season for O_3 in Massachusetts (April–September) [12, 25].

Lung function measurement

FEV₁ and forced vital capacity (FVC) were measured at each examination by spirometry, with the spirometer calibrated daily at room temperature, barometric pressure and humidity. For Offspring examination 7, spirometry data were obtained using a Collins Survey II spirometer, interfaced to pulmonary function data acquisition and quality control software (S and M Instruments, Doylestown, PA, USA). For Offspring examination 8 and Third Generation examinations 1 and 2, spirometry data were obtained using a Collins CPL system (nSpire Health, Inc., Longmont, CO, USA). Spirometry manoeuvres were performed according to American Thoracic Society standards [26].

Statistical analyses

We selected covariates *a priori* based on known associations with lung function or temperature, and for consistency with prior publications [12, 27]. Primary models adjusted for sex, age (as a linear and squared term), height, weight, smoking status, pack-years of smoking, any household smoking, current asthma, personal education, median household income from 2000 census tract, season, previous-day relative humidity, previous-day PM2.5, examination 8 and cohort. We defined current asthma as a report of wheeze in the past 12 months, or a current doctor diagnosis of asthma. We defined chronic obstructive pulmonary disease (COPD) based on a report of daily cough for at least three consecutive months of the year or a current diagnosis of chronic bronchitis, COPD or emphysema [12]. We used sine and cosine functions of the examination date to control for season. We assigned previousday PM2.5 using the Harvard Supersite monitor as an indicator of regional particulate pollution, and did not adjust for O₃ in our primary model because the monitors in Eastern Massachusetts may not reflect regional exposure. We did adjust for O₃ in secondary

analyses. We performed sensitivity analyses excluding Offspring examination 7 because the spirometry equipment changed after examination 7.

We fit linear mixed-effects models with a subject-specific random intercept to examine associations of 1-, 2- and 7-day average of temperature with lung function (FEV₁ and FVC), using PROC MIXED in SAS Version 9.4 (SAS Institute, Cary, NC, USA). We evaluated the linearity of each association by plotting generalised additive mixed models (GAMM) with penalised splines using R (R Foundation for Statistical Computing, Vienna, Austria). We tested if associations between temperature and lung function were modified by season or relative humidity. Seasons were defined as winter (December–February), spring (March–May), summer (June–August) and autumn (September–November). Based on evidence of differential susceptibility to temperature [9, 11, 28] and seasonal allergens such as pollen [29] among patients with asthma and people with COPD, we tested if associations with temperature differed by asthma and COPD status.

We evaluated if previous-day temperature at the home address modified associations between previous-day pollutant exposure (PM2.5, NO2 and O3) and FEV1 in a subset of noncurrent smoking participants (n=2051 participants, 3076 observations) living within 40 km of the Harvard air pollution monitoring site, to reflect average exposures within the study region. The satellite-based estimates of temperature became available in the year 2000, and therefore this subset did not include observations from the 1990s that were included in our previous publication on acute respiratory effects of pollution (n=3262, 5358 observations) [12]. For consistency with our prior work on acute respiratory effects of pollution within air quality standards [12], we included observations with previous-day pollution levels within 2013 EPA air quality standards (only 21 observations had a 24-h PM2.5 level >35 μ g·m⁻³, 37 observations had an 8-h maximum O_3 average >75 ppb, and none had an hourly NO_2 >100 ppb), and by convention we examined associations with O₃ during the O₃ monitoring season of Massachusetts (April-September). We assessed effect modification by previousday temperature as a continuous variable and, in secondary analyses, in categories of relative temperature to determine if relatively warm or cold temperature (defined within the season) modified associations between pollution and lung function. Temperature within 1 SD of the mean daily temperature for the season was defined as "seasonable", >1 SD above seasonal mean was defined as unseasonably "warm" and >1 SD below seasonal mean was defined as unseasonably "cool".

We used the Wald test for cross-product terms to determine the statistical significance of potential effect modifiers for continuous variables and the type 3 test of fixed effects to determine the significance of global interaction by categorical variables. A two-sided p-value of <0.10 was used to test for statistically significant effect modification.

Results

Participant characteristics for the 9150 observations are summarised in table 1. This is a middle- to older-aged cohort with a slight majority of women (54.2%). Approximately half of observations included in the analysis were from never smokers and current smoking was rare (13%). Nearly half of participants were college graduates and 74% completed at least

some college. Asthma and wheeze were common: nearly one quarter of observations were from people with a diagnosis of asthma or wheeze in the past 12 months. Average lung function was normal.

Season-stratified mean temperature, humidity and pollutant levels the day preceding spirometry testing are provided in table 2. Temperature at the home address on the same date varied modestly across the 5896 unique participants ($_{SD} < 1^{\circ}C$ for most sample dates tested). The geographic distribution of study participants' residences was concentrated around the state of Massachusetts, and extended to western New York and coastal Maine (supplementary figure S1).

The main associations of temperature with lung function are shown in table 3. Higher 1-, 2and 7-day average temperatures were all associated with lower FEV₁ and FVC. The penalised spline for the 2-day temperature average and FEV₁ is shown in figure 1. The other splines appeared similar and are in the supplementary material. The splines demonstrate steeper negative associations of 1-, 2- and 7-day temperatures with FEV₁ in the cooler temperature range (<10°C), and a flattening of the association at warmer temperatures. Associations were unchanged with additional adjustment for previous-day O₃ and in sensitivity analyses excluding observations from Offspring examination 7. Associations of temperature with FEV₁ did not differ by relative humidity, or by asthma or COPD status (all p_{interaction}>0.1) (supplementary table S1).

In linear mixed-effects models, there was evidence of effect modification by season for the associations of the 1-, 2- and 7-day temperature average with FEV_1 (all $p_{interaction} < 0.05$). Figure 2 plots the results of a GAMM model with an interaction term by season for the 2- day average of temperature. There were negative linear associations between temperature and FEV_1 in winter and spring, but not summer or autumn. In season-stratified linear models (table 4), associations between temperature and FEV_1 were more negative in winter and spring than for the overall sample and null in summer and autumn, except for the 7-day average of temperature in autumn, which was positively associated with FEV_1 .

Previous-day temperature did not modify associations between previous-day PM2.5, O_3 or NO_2 and FEV_1 for any of the pollutants in the subset of 2051 participants with regional air pollution and address-specific temperature data. Notably, only O_3 was significantly associated with FEV_1 in this subset: each 10 ppb higher previous-day O_3 was associated with a 19.7 mL lower FEV_1 (95% CI –37.2–2.2). When previous-day temperature was categorised as seasonal or unseasonably warm or cool (defined as 1 _{SD} higher or lower than the seasonal mean, respectively), there was a suggestion of effect modification ($p_{interaction}=0.08$) for the association of $PM_{2.5}$ with FEV_1 . Negative associations of $PM_{2.5}$ with FEV_1 were present on unseasonably warm days, but not unseasonably cool days, and a similar pattern was observed for the other pollutants (figure 3).

Discussion

In this cohort of generally healthy adults residing in the Northeastern US, warmer temperature was associated with lower lung function, after controlling for season and other

confounders. This negative relationship between temperature and lung function was steepest in the cooler temperature range (<10°C), and present during winter and spring, but not summer. People with asthma and COPD had similar associations between temperature and lung function as those without these obstructive lung diseases.

Although we hypothesised that both colder and warmer outdoor temperatures would be associated with lower lung function in a non-linear pattern, our finding of a linear negative relationship between temperature and lung function is consistent with a growing number of studies that have found linear associations between higher daily outdoor temperatures and worse respiratory outcomes, including lung function, respiratory symptoms and respiratory hospitalisation [10, 11, 15, 28, 30]. The Normative Aging Study of elderly men, also carried out in the temperate climate of the Northeastern US, found a negative association between increasing outdoor temperature (measured at a central urban monitor) averaged for the past 0-2 days and lung function [10]. This study did not detect effect modification by season, but the sample size was substantially smaller than our study (1103 versus 5896 participants). Two recent panel studies of healthy young adults in China found that higher temperatures were associated with lower lung function, measured by daily peak expiratory flow [15, 30]. In one study in Beijing conducted in the months of April–November (temperature range 1– 41°C, mean $\sim 20^{\circ}$ C), each 5°C increase in same-day outdoor temperature, regardless of season, was associated with a 6.1 L·min⁻¹ (95% CI -11.1--1.1) lower evening peak expiratory flow rate [15]. An Australian study of urban asthmatic children also detected linear relationships between increasing temperature (for up to the past 3 days) and lower lung function, controlling for season and ambient pollution [11]. This study did not assess effect modification by season, but found that associations of higher temperature with lower lung function were greatest in the cooler more southern Australian cities than the northern subtropical city of Brisbane. Warmer outdoor temperatures have also been associated with lower lung function among cystic fibrosis patients [31].

The mechanism for this association of outdoor temperature with lung function is unclear and may not be mediated by the effect of temperature itself on the airways. It has been proposed that impaired thermoregulation in the setting of higher temperatures may trigger physiologic changes that ultimately reduce lung function [10], but such a mechanism is unlikely to explain our findings because we did not detect an association with temperature during the hotter summer season. While this study did not measure daily behaviour, one possible unifying explanation is that participants spent, on average, more time outdoors (or with open windows) on warmer days during the cooler seasons. Outdoor exposures known to adversely affect lung function, such as air pollutants, environmental tobacco smoke, spring pollen and respiratory pathogens, may explain the lower lung function we observed in association with warmer temperatures. A recent National Health and Nutrition Examination Survey analysis found that lung function is on average lower in warmer parts of the USA: each 10°F $(5.56^{\circ}C)$ higher mean temperature was associated with 0.7% lower FEV₁ % pred [32]. Studies that have measured time spent outside the home have found that this time may influence lung function and respiratory symptoms. For example, a recent study of COPD patients found that higher outdoor temperature was associated with worse respiratory symptoms, and this association was only present on days when participants went outdoors [33]. We did not detect statistical evidence of differential susceptibility by asthma or COPD

Mould is another temperature-dependent exposure that may affect lung function. Outdoor concentrations of mould spores tend to be higher on warmer days (even during the winter months for some species) [34, 35], and indoor levels of moulds such as *Alternaria* have been found to correlate highly with outdoor levels [36]. Higher exposure to fungal spores, including total fungal spores and *Cladosporium* species, has been associated with lower lung function in children [37, 38]. However, it is unclear how the daily fluctuation in mould spores of various species, influenced in part by temperature, affects lung function among healthy, mostly non-sensitised adults.

Some previous studies have suggested that higher temperatures and air pollution interact to increase mortality [16, 18, 39, 40] and reduce lung function [10, 15], though findings have not been consistent across studies [30, 41]. We did not find, as we had hypothesised, that previous-day temperature augmented associations of pollutants with lung function. Our analysis of temperature–pollution interactions had less statistical power than the main analysis of temperature because it was confined to a subset of study participants with address-specific temperature and local pollution data. In secondary analyses we observed a pattern of negative associations with pollutants, in particular PM_{2.5}, following unusually warm days but not unusually cold days. This may be a consequence of greater pollution exposure on relatively warm days. More research is needed to determine if higher temperatures modify the effect of air pollutants on the lung.

This study has a number of limitations. Our study population consisted of predominantly white, middle- and older-aged men and women residing in the Northeastern US, which limits generalisability to other groups. We had a maximum of two repeated measures of lung function for each participant and cannot infer causality from these associations. In our assessment of the interaction between temperature and pollution exposure, we assigned outdoor air pollution exposure measured within 40 km of each individual's home address to estimate individual exposure levels, which may not reflect actual exposures.

Our study also has a number of strengths. We examined moving averages of outdoor temperature and repeated measures of lung function in a well-characterised cohort of men and women, controlling for individual, neighbourhood and seasonal confounders. Other studies using stationary temperature monitors ignore potentially important differences that may influence temperature at individual residences, including urban heat islands, distance from water and amount of impermeable surface [19]. This analysis leveraged a validated spatiotemporal model of temperature using satellite surface temperature data [20] to more precisely estimate temperature at the home residence and examine how daily outdoor temperature relates to lung function among generally healthy adults.

In conclusion, we found that higher daily temperature was associated with lower lung function. These negative associations were present in winter and spring, but not in summer or autumn, and were present among people with and without asthma or COPD. We speculate that the apparent effect of warmer temperature on lung function may be a consequence of

more time spent outdoors on warmer days, especially during the cooler seasons; however, we did not assess time spent outdoors in this study. It is important to understand associations between temperature and lung function because temperatures are increasing due to climate change [2]. Future research on the effect of temperature on respiratory health may consider how temperature affects activity patterns and outdoor exposure to pollutants and allergens, to distinguish health effects of temperature and other outdoor exposures.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

Acknowledgments

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Rice et al.



FIGURE 1.

Two-day average temperature and forced expiratory volume in 1 s (FEV₁). Generalised additive mixed model with penalised spline adjusted for age, sex, height, weight, smoking status and pack-years, second-hand smoke at home, asthma, education, census-level household income, cohort, examination 8, season, previous-day relative humidity and level of particles with a diameter of <2.5 μ m. Degrees of freedom=2.58. Dotted lines represent 95% CI. Black shading on the x-axis indicates the density of observations.

Rice et al.



FIGURE 2.

Two-day average temperature and forced expiratory volume in 1 s (FEV₁) by season. Generalised additive mixed model with penalised spline for the association between two-day temperature and FEV₁ in a) winter, b) spring, c) summer and d) autumn, adjusted for age, sex, height, weight, smoking status and pack-years, second-hand smoke at home, asthma, education, census-level household income, cohort, examination 8, season, previous-day relative humidity and level of particles with a diameter of <2.5 μ m. Dotted lines represent 95% CI. Black shading along the x-axis indicates the density of observations.

Rice et al.



FIGURE 3.

Previous-day pollution level and forced expiratory volume in 1 s (FEV₁) based on previousday temperature relative to seasonal mean. Particles with a diameter of <2.5 μ m (PM_{2.5}) scaled per 5 μ g·m⁻³ (among 3047 observations), nitrogen dioxide (NO₂) scaled per 10 ppb (among 3076 observations) and ozone (O₃) scaled per 10 ppb (among 1550 observations in monitoring season of April–September). Linear mixed models with an interaction term for previous-day temperature categorised as "cool" if >1 _{SD} below mean temperature for season of test, "seasonal" if within 1 _{SD} of mean temperature for season of test, and "warm" if >1 _{SD} above mean temperature of season for test. All models were adjusted for participant sex, age, age², height, weight, current asthma, smoking status, pack-years, household smoking, education, census tract median household income, cohort, date of examination, season (as sine and cosine terms), previous-day regional relative humidity and previous-day temperature at home address. Whiskers indicate the 95% CI.

TABLE 1

Characteristics of study participants

Characteristic	Mean+SD or %			
	50 4+13 8			
Age years	30.4±13.8			
	45.6			
BMI kg·m ⁻²	27.8±3.6			
Offspring cohort	34.9			
Pack-years				
Former	17.0±18.1			
Current	28.0±16.6			
Smoking status				
Never	49.4			
Former	37.1			
Current	13.4			
Smoker in household in adulthood Education	25.4			
High school or less	21.2			
Some college	31.6			
College graduate	46.8			
Median census household income \$1000	64.6±20.4			
Asthma	23.1			
COPD	4.7			
Season of study visit				
Winter	22.8			
Spring	26.4			
Summer	25.4			
Autumn	25.4			
FEV ₁ L	3.2±0.9			
FEV ₁ % pred	97.9±15.1			
FVC L	4.2±1.1			
FVC % pred	102.1±13.4			

Data calculated from 9150 observations (among 5896 participants). BMI: body mass index; COPD: chronic obstructive pulmonary disease; FEV1: forced expiratory volume in 1 s; FVC: forced vital capacity.

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Temperature °C	9150	2088	4.01 ± 3.7	2418	$9.9{\pm}4.1$	2322	17.2±2.7	2322	12.0 ± 3.7
RH %	9144	2088	60.8 ± 15.5	2418	64.9 ± 18.3	2316	71.1 ± 14.0	2322	69.7±13.7
$\rm PM_{2.5}~\mu g \cdot m^{-3}$	3047	727	8.8±4.3	824	$8.8 {\pm} 4.0$	767	11.5 ± 6.5	729	8.6 ± 5.1
NO2 ppb [¶]	3076	727	21.8 ± 7.3	824	19.5 ± 5.9	789	17.0 ± 5.0	736	17.7±5.8
$0_{3} \mathrm{ppb}^{+}$	1550			546	30.6±8.7	762	$30.0{\pm}10.0$	242	19.7±6.8

ò ;

#, 9150 observations are reported from 5896 participants residing in the Northeastern US;

🐔 standards;

 \dot{t} ozone data is reported during the EPA ozone monitoring season for Massachusetts (April–September).

TABLE 3

Associations of temperature at home address and lung function

Temperature average	Unit	Difference in FEV_1 mL	Difference in FVC mL
1-day	5°C	-16.7 (-26.66.9)	-13.4 (-24.42.4)
2-day	5°C	-15.2 (-26.24.2)	-11.1 (-23.4-1.3)
7-day	5°C	-19.8 (-33.56.1)	-16.2 (-31.60.9)

Data are presented as mean difference (95% CI). Linear mixed models adjusted for participant sex, age, age², height, weight, asthma, smoking status, pack-years, household smoking, education, census tract median household income, season (as sine and cosine terms of the date), previous-day regional relative humidity and level of particles with a diameter of <2.5 μ m, examination 8 and cohort. FEV₁: forced expiratory volume in 1 s; FVC: forced vital capacity.

TABLE 4

Within-season associations of temperature at home address and FEV1

Season	Unit	Difference in FEV ₁ mL				
		1-day average temperature	2-day average temperature	7-day average temperature		
Winter	5°C	-28.5 (-55.91.1)	-23.5 (-53.6-6.6)	-53.5 (-86.720.4)		
Spring	5°C	-23.7 (-44.82.5)	-26.3 (49.03.6)	-31.5 (-55.97.2)		
Summer	5°C	10.0 (-15.8-35.9)	17.6 (-13.7-48.9)	27.3 (-9.9-64.5)		
Autumn	5°C	8.4 (-12.5-29.3)	11.7 (-10.1-33.5)	28.1 (4.8–51.5)		

Data are presented as mean (95% CI). Linear mixed models adjusted for participant sex, age, age^2 , height, weight, asthma, smoking status, packyears, household smoking, education, census tract median household income, previous-day regional relative humidity and level of particles with a diameter of <2.5 μ m, examination 8 and cohort. FEV1: forced expiratory volume in 1 s.