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Association of Blood Pressure With Outcomes in Acute Stroke Thrombectomy

Konark Malhotra,

Department of Neurology, Allegheny Health Network, Pittsburgh, PA

Nitin Goyal,

Department of Neurology, University of Tennessee, Memphis

Aristeidis H. Katsanos,

Department of Neurology, McMaster University/Population Health Research Institute, Hamilton, Canada

Angeliki Filippatou,

Second Department of Neurology, "Attikon" University Hospital, National and Kapodistrian University of Athens, Greece

Eva A. Mistry,

Department of Neurology, Vanderbilt University, Nashville, TN

Pooja Khatri,

Department of Neurology, University of Cincinnati, OH

Mohammad Anadani,

Department of Neurology, Washington University School of Medicine, St Louis, MO

Department of Neurosurgery, Medical University of South Carolina, Charleston

Alejandro M. Spiotta,

Department of Neurosurgery, Medical University of South Carolina, Charleston

Else Charlotte Sandset,

Department of Neurology, Stroke Unit, Oslo University Hospital, Norway

The Norwegian Air Ambulance Foundation, Oslo, Norway

Amrou Sarraj,

Department of Neurology, UT Houston, TX

Georgios Magoufis,

Stroke Unit, Metropolitan Hospital, Piraeus, Greece

Christos Krogias,

Department of Neurology, St. Josef-Hospital, Ruhr-University Bochum, Germany

Correspondence to Konark Malhotra, Department of Neurology, Allegheny Health Network, Pittsburgh, PA. konark.malhotra@yahoo.com.

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Lars Tönges,

Department of Neurology, St. Josef-Hospital, Ruhr-University Bochum, Germany

Apostolos Safouris,

Stroke Unit, Metropolitan Hospital, Piraeus, Greece

Lucas Elijovich,

Department of Neurosurgery, University of Tennessee/Semmes-Murphey Clinic, Memphis

Mayank Goyal,

Departments of Radiology and Clinical Neurosciences, University of Calgary, AB, Canada

Adam Arthur,

Department of Neurosurgery, University of Tennessee/Semmes-Murphey Clinic, Memphis

Andrei V. Alexandrov,

Department of Neurology, University of Tennessee, Memphis

Georgios Tsivgoulis

Department of Neurology, University of Tennessee, Memphis

Second Department of Neurology, "Attikon" University Hospital, National and Kapodistrian University of Athens, Greece

Abstract

Limited data exist evaluating the effect of blood pressure (BP) on clinical outcomes among patients with acute ischemic stroke with large vessel occlusion treated with mechanical thrombectomy (MT). We sought to evaluate the association of BP levels on clinical outcomes among patients with acute ischemic stroke with large vessel occlusion treated with MT. Studies were identified that reported the association of systolic BP (SBP) or diastolic BP levels before, during, or after MT on the outcomes of patients with acute ischemic stroke treated with MT. Unadjusted and adjusted analyses of studies reporting odds ratios (ORadi) per 10 mm Hg BP increment were performed. Our analysis included 25 studies comprising 6474 patients. Higher pre-MT mean SBP (P=0.008) and post-MT maximum SBP (P=0.009) levels were observed in patients who died within 3 months. Patients with 3-month functional independence were noted to have lower pre-MT (P<0.001) and post-MT maximum SBP levels (P<0.001). In adjusted analyses, increasing post-MT maximum SBP and diastolic BP levels were associated with 3-month mortality (OR_{adi}, 1.19 [95% CI,1.00–1.43]; I²=78%, P value for Cochran Q test: 0.001) and symptomatic intracranial hemorrhage (OR_{adj}, 1.65 [95% CI, 1.11–2.44]; I²=0%, P value for Cochran Q test: 0.80), respectively. Increasing pre- and post-MT mean SBP levels were associated with lower odds of 3-month functional independence (OR_{adj}, 0.86 [95% CI, 0.77–0.96]; I²=18%, P value for Cochran Q test: 0.30) and (OR_{adj}, 0.80 [95% CI, 0.72–0.89]; I²=0%, P value for Cochran Q test: 0.51), respectively. In conclusion, elevated BP levels before and after MT are associated with adverse outcomes among patients with acute ischemic stroke with large vessel occlusion.

Keywords

blood pressure; consensus; intracranial hemorrhages; odds ratio; thrombectomy

Large vessel occlusion (LVO) has been reported to occur in up to one-third of patients with acute ischemic stroke (AIS) and is associated with higher rates of poststroke dependence and mortality.^{1,2} Mechanical thrombectomy (MT) in patients with LVO improves functional outcomes and has profoundly changed the landscape of acute stroke therapy.³ During the hyperacute stage of AIS due to LVO, the fate of the ischemic penumbra largely depends on the maintenance of perfusion above the threshold for the infarct core. Among other factors that affect cerebral perfusion, optimizing blood pressure (BP) remains a potential target to improve neurological outcome in patients with LVO during hyperacute AIS stage. Elevated pre- and post-treatment BP levels have been adversely associated with AIS outcomes in patients treated with intravenous thrombolysis (IVT).⁴ Current American Heart Association/ American Stroke Association guidelines recommend strict, though arbitrary, thresholds of systolic BP (SBP) <180 mm Hg and diastolic blood pressure (DBP) <105 mm Hg during and after MT.⁵ However, data regarding guidance for optimal BP management among patients with LVO-AIS treated with MT largely remain scarce.

Observational data indicate that elevated BP among patients with LVO treated with MT is associated with an increased risk of symptomatic intracranial hemorrhage (sICH)^{6,7} as well as mortality^{8,9} and reduce the odds of MT-induced recanalization^{10,11} and functional independence.^{6,9,12–14} Randomized data are lacking to render a clear consensus for the optimal BP control before, during, and after MT among patients with AIS with LVO. In view of these considerations, we conducted a systematic review and a pairwise meta-analysis seeking to evaluate the association of elevated acute BP levels before, during, and after MT on different clinical outcomes among patients with LVO treated with MT.

Methods

We adopted the Preferred Reporting Items for Systematic Reviews and Meta-Analyses statement¹⁵ and reported in accordance to the Meta-analysis of Observational Studies in Epidemiology¹⁶ proposal. The protocol of the present systematic review and meta-analysis has been registered to PROSPERO (The International Prospective Register of Systematic Reviews) (CRD42019134621). The present manuscript also adheres to the American Heart Association Journals implementation of the Transparency and Openness Promotion guidelines.¹⁷ Authors declare that all supporting data are available within the article and its online-only Data Supplement.

Data Sources and Searches

Eligible studies were identified by systematically searching in Ovid MEDLINE, Ovid Embase, and Scopus databases. The combination of search strings used to query all the databases included: "mechanical thrombectomy", "stroke", "cerebral ischemia", "blood pressure," "systolic," and "diastolic". The search algorithm used for the MEDLINE database is available in the online-only Data Supplement. We restricted our search to articles in English language, and our search spanned from database inception to August 6, 2019. Additional manual search of conference abstracts and bibliographies of articles meeting study criteria for a comprehensive literature search was conducted.

Study Selection and Data Extraction

We identified all observational studies (prospective or retrospective) and post hoc studies of randomized controlled clinical trials that provided data on the association of acute BP levels with clinical outcomes in patients with AIS treated with MT. Per study protocol, we excluded studies that reported (1) outcomes not documented according to our predefined criteria such as parenchymal hematoma or asymptomatic intracranial hemorrhage, (2) treatment with intraarterial thrombolysis, (3) categorical or descriptive data for BP levels reported as median values, (4) studies reporting mean arterial pressure levels instead of SBP or DBP levels, and (5) case reports, case series, or conference abstracts. In case of overlapping data (same patient data used in >1 publication) for each outcome of interest, we used the study with the highest number of included patients. Reference lists of all articles that met the inclusion criteria and of relevant review articles were examined to identify studies that may have been missed by the initial database search. All retrieved studies were scanned independently by 2 reviewers (K. Malhotra, A. Filippatou), and in any case of disagreement, the third author (G. Tsivgoulis) was consulted to resolve any disagreements.

We primarily documented maximum, minimum, or mean SBP and DBP levels reported as mean±SD before, during, and after MT. During the post-MT interval, we recorded BP levels documented within 2 to 4 hours following MT. If these data were unavailable, BP levels within 24 to 72 hours of MT were used. Additional data on BP variability (BPV) including SD, coefficient of variation (CV), and successive variation (SV) were collected if available. In case of missing data, the authors of relevant studies were contacted, and previously unpublished data was occasionally provided according to their discretion.

Outcomes

The primary efficacy outcomes were defined as the unadjusted and adjusted for potential confounders likelihood of 3-month functional independence, defined as a modified Rankin Scale (mRS) score of 0 to 2 at 3 months or discharge. The primary safety outcomes were defined as unadjusted and adjusted for potential confounders likelihood of sICH (according to the definition provided in each study; Table S1 in the online-only Data Supplement).

Secondary outcomes included 3-month mortality, successful recanalization (defined as Thrombolysis in Cerebral Infarction scores of 2b or 3 at the end of MT), and functional improvement (assessed with ordinal [shift] analysis of the mRS scores at 3 months or at discharge).

Risk of Bias Assessment

We used the Newcastle-Ottawa Scale⁴ and the ROBINS-E (Risk of Bias in Nonrandomized Studies of Exposures)¹⁸ to assess the quality and explore the sources of bias amongst the included cohort studies. The quality control and bias identification were performed independently by 2 reviewers (K. Malhotra and A. Filippatou), and all potential disagreements were resolved by a third tie-breaking evaluator (G. Tsivgoulis).

Data Synthesis and Statistical Analysis

In the current pairwise meta-analysis, both unadjusted and adjusted for potential confounders analyses for pre- and post-MT SBP/DBP levels were handled as continuous variables, while the outcomes of interest were handled as dichotomous variables. Differences in maximum, minimum, or mean pre- and post-MT BP levels according to the outcomes of interest were reported in the form of standardized mean differences (SMDs) in all unadjusted analyses as previously described.^{3,16} We also conducted adjusted odds ratios (OR_{adj}) of these associations evaluating the association of pre- and post-MT BP levels with different clinical outcomes. The OR_{adj} of these associations are all presented per 10 mm Hg increments in SBP or DBP levels and derived as available from the original studies. For each of these associations (unadjusted and adjusted for potential confounders), the individual study effects were estimated using the random-effects model (DerSimonian and Laird).¹⁹ We used inverse variance method to calculate SMD for continuous variables. SMDs were interpreted using a general rule of thumb reported by Cohen,²⁰ in which an SMD of 0.2 represents a small effect, an SMD of 0.5 represents a medium effect, and an SMD of 0.8 or larger represents a large effect.

As per the *Cochrane Handbook for Systematic Reviews of Interventions*,²¹ we assessed for heterogeneity between the included studies using Cochran Q and I² statistics. For the qualitative interpretation of heterogeneity, I²>50% and I²>75% indicated substantial and considerable heterogeneity, respectively. We performed subgroup analyses on the outcomes of interest by including only studies that provided data on the association of BP levels and clinical outcomes in patients with LVO who achieved successful recanalization following MT.

Publication bias across individual studies was graphically evaluated using a funnel plot,²¹ while funnel plot asymmetry was assessed using the Egger linear regression test with P<0.10 significance level. For all other outcomes of interest, we performed equivalent *z* test for each pooled estimate and a 2-tailed *P*level <0.05 was considered statistically significant.

All statistical analyses were carried out with Cochrane Collaboration's Review Manager Software Package (RevMan 5.3) and the Comprehensive Meta-analysis version 2 software (Biostat, Englewood, NJ, https://www.meta-analysis.com).

Results

Study Selection and Study Characteristics

A systematic search of all the databases yielded 303 articles. After removing the duplicates, the titles and abstracts from the remaining 287 studies were screened and 43 potentially eligible studies for the meta-analysis were retained. After retrieving the full-text version of the aforementioned 43 studies, 18 studies^{22–39} were excluded (Table S2) due to nonavailability of intended data and use of endovascular reperfusion procedures other than MT (eg, Intraarterial thrombolysis). One multicenter study³⁹ was excluded due to overlapping data with other previous studies that were already included in our meta-analysis. After careful evaluation and without disagreements among the 2 reviewers, 25

studies $^{6-14,40-54}$ were included that met the study protocol's inclusion criteria. The detailed flow chart of the current meta-analysis is presented in Figure S1.

The included 25 studies comprising 6474 patients with their baseline characteristics are summarized in Table S3. Two studies were post hoc analyses^{51,54} of randomized-controlled clinical trial including 278 patients with AIS and the remaining 23 were retrospective observational^{6–14,40–50,52,53,55}). Thirteen studies were conducted in the United States, 4 in South Korea, 3 in France, 2 in Germany, and 1 each in China, Czech Republic, and Denmark.

Study Quality Assessment

We assessed the risk of bias among the included studies using Newcastle-Ottawa scale (Table S4) and Risk of Bias in Nonrandomized Studies of Exposures (Table S5). The evaluation of the risk of bias using the Newcastle-Ottawa scale disclosed that the risks of selection and comparability biases were considered low in all included studies. Outcome bias was quantified as moderate since the majority of included studies did not report data on patients lost to follow-up or outcome assessment. The overall score of Newcastle-Ottawa scale was 213/225 (94.6%), which is considered to represent an overall high quality. Notably, we also documented low and medium risk of bias in 19 and 6 studies respectively using the ROBINS-E approach.

Association Between BP Levels and Outcomes—Table 1, Table S6, and Table 2 present an overview on the overall unadjusted and adjusted analyses investigating the association of BP levels measured before, during, and after MT with various clinical outcomes, respectively.

Unadjusted Analyses

BP Levels Before MT.: Lower pre-MT mean SBP (Figure S2A) and DBP levels (Figure S2B) were observed in patients with 3-month functional independence. Similarly, lower pre-MT mean SBP levels (Figure S3A) were documented in patients who achieved successful recanalization, whereas preMT mean DBP levels did not differ between patients with and without successful recanalization following MT (Figure S3B). Higher pretreatment mean SBP (Figure S4A) and DBP levels (Figure S4B) were documented in patients who were dead within 3 months from stroke onset. However, no differences in pretreatment mean SBP (Figure S5A) and DBP (Figure S5B) levels were noted in patients with and without sICH.

<u>BP Levels During MT.</u>: During MT procedures, lower levels of maximum SBP (Figure S6A) were observed in patients who achieved 3-month functional independence (Table S6). However, no difference was noted for minimum SBP (Figure S6B) and maximum (Figure S7A) or minimum (Figure S7B) DBP levels in patients with and without 3-month functional independence. No additional data were available for other outcomes.

<u>BP Levels After MT.</u>: After MT procedures, lower levels of maximum SBP (Figure S8), mean SBP (Figure 1A), and maximum DBP (Figure S9A) were recorded among patients who achieved 3-month functional independence. However, mean DBP (Figure S9B) and

minimum SBP (Figure S10A) or DBP (Figure S10B) levels did not differ between patients with and without functional independence. No difference was observed for both maximum (Figure S11A) and mean (Figure S11B) post-MT SBP levels in patients with and without successful recanalization. Higher maximum SBP (Figure 2A) and maximum (Figure S12A) or mean (Figure S12B) DBP levels were observed among patients who died within 3 months of stroke onset, whereas no such difference was observed for mean SBP (Figure S13), minimum SBP (Figure S14A), and minimum DBP (Figure S14B) levels. No differences in maximum SBP (Figure S15A), minimum SBP (Figure S15B), mean SBP (Figure S15C), maximum DBP (Figure 3A), minimum DBP (Figure S16A), or mean DBP (Figure S16B) levels were noted in patients with and without sICH.

After MT elevated levels of maximum SBP (Figure S17A), mean SBP (Figure S17B) and maximum DBP (Figure S18A) were observed among patients with worsening 3-month mRS (1-point increase in mRS-scores in shift analyses), whereas mean DBP just failed to reach statistical significance (Figure S18B). We conducted additional analyses to assess the associations of BPV quantified by SD, CV, and SV in BP levels with various clinical outcomes. Although lower SV for both SBPBPV and DBP-BPV were observed in patients with 3-month functional independence, no differences were observed for post-MT SBP-BPV (quantified as SD or CV; Figure S19A) and DBP-BPV (quantified as SD or CV; Figure S19A) and DBP-BPV (Figure S20B) were observed in patients of SV for both SBP-BPV (Figure S20A) and DBP-BPV (Figure S20B) were observed in patients who died within 3 months, whereas no differences were observed for both SBP-BPV (Figure S21A) and DBP-BPV (Figure S21B) among patients with or without sICH.

Adjusted Analyses—We performed pairwise meta-analyses using the adjusted for potential confounders associations of BP levels before, during, and after MT with various outcomes. Adjusted variables from the individual studies are listed in Table S3. All OR_{adj} on the association of SBP and DBP increments with the respective outcomes of interest were rescaled and presented as 10 mm Hg BP increments.

<u>BP Levels Before MT.</u> In adjusted analyses, increasing mean SBP levels before MT (Table 1) were associated with lower odds of 3-month functional independence (4 studies; OR_{adj} , 0.86 [95% CI, 0.77–0.96], I²=18%, *P* value for Cochran Q test: 0.30; Figure S22). Elevated pre-MT mean SBP levels were associated with higher odds of 3-month mortality (2 studies; OR_{adj} , 1.22 [95% CI, 1.00–1.49], I²=0%, *P* value for Cochran Q test: 0.42; Figure S23A); however, there was no independent association between pre-MT mean DBP levels and 3-month mortality (2 studies; OR_{adj} , 1.19 [95% CI, 0.65–2.18], I²=10%, *P* value for Cochran Q test: 0.29; Figure S23B).

<u>BP Levels During MT.</u>: During MT, increasing levels of maximum SBP were independently associated with lower odds of 3-month functional independence (2 studies; OR_{adj} , 0.93 [95% CI, 0.90–0.96], I²=0%, *P* value for Cochran Q test: 0.78; Figure S24; Table S6). No additional data were available for other outcomes.

<u>BP Levels After MT.</u>: After MT, increasing mean SBP levels (6 studies; OR_{adj}, 0.80 [95% CI, 0.72–0.89], I²=0%, *P* value for Cochran Q test: 0.51; Figure 1B) and maximum DBP

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levels (5 studies; OR_{adj} , 0.83 [95% CI, 0.72–0.96]; $I^2=23\%$, *P* value for Cochran Q test: 0.27; Figure S25A; Table 2) were independently associated with reduced likelihood of 3-month functional independence without significant evidence of heterogeneity across included studies. Increasing maximum SBP levels was also associated (8 studies; OR_{adj} , 0.82 [95% CI, 0.73–0.93]; Figure S25B) with lower odds of 3-month functional independence, however, with considerable heterogeneity ($I^2=72\%$; *P* value for Cochran Q test: 0.0007). There was no independent association between minimum SBP (2 studies; OR_{adj} , 1.25 [95% CI, 0.79–1.99]; $I^2=64\%$, *P* value for Cochran Q test: 0.10; Figure S26A) and mean DBP (3 studies; OR_{adj} , 0.80 [95% CI, 0.61–1.05]; $I^2=44\%$, *P* value for Cochran Q test: 0.17; Figure S26B) and 3-month functional independence.

Higher maximum SBP (5 studies; OR_{adj}, 1.19 [95% CI, 1.00–1.43], I²=78%, P value for Cochran Q test: 0.001; Figure 2B) and DBP (2 studies; OR_{adi}, 1.38 [95% CI, 1.09-1.76], I²=3%, Pvalue for Cochran Q test: 0.31; Figure S27A) levels were independently associated with 3-month mortality, whereas increasing levels of mean SBP levels suggested a trend toward increased 3-month mortality (2 studies; OR_{adi}, 1.02 [95% CI, 1.00–1.04], I²=0%, P value for Cochran Q test: 0.35; Figure S27B). Higher maximum DBP levels were independently associated with sICH (2 studies; ORadi, 1.65 [95% CI, 1.11-2.44]; Figure 3B) without any evidence of heterogeneity across included studies ($I^2=0\%$; P value for Cochran Q test: 0.80). No association was noted for maximum SBP (7 studies; ORadi, 1.02 [95% CI, 0.99–1.05]; I²=58%, P value for Cochran Q test: 0.03; Figure S28A), minimum SBP (2 studies; OR_{adi}, 0.93 [95% CI, 0.78–1.12]; I²=54%, Pvalue for Cochran Q test: 0.14; Figure S28B), mean SBP (4 studies; OR_{adi}, 1.00 [95% CI, 0.99–1.02]; I²=0%, P value for Cochran Q test: 0.78; Figure S28C), minimum DBP (2 studies; OR_{adi}, 0.98 [95% CI, 0.95-1.01]; I²=0%, Pvalue for Cochran Q test: 0.89; Figure S29A), and mean DBP (2 studies; OR_{adi}, 1.27 [95% CI, 0.91–1.78]; I²=0%, P value for Cochran Q test: 0.84; Figure S29B) levels with the likelihood of sICH in adjusted analyses.

Additionally, elevated levels of maximum (5 studies; cOR_{adj} , 1.15 [95% CI, 1.09–1.21]; $I^2=76\%$, *P* value for Cochran Q test: 0.002; Figure S30A) and mean (5 studies; cOR_{adj} , 1.28 [95% CI, 1.17–1.39]; $I^2=7\%$, *P* value for Cochran Q test: 0.36; Figure S30B) SBP levels were independently associated with worsening 3-month mRS (1-point increase in mRS scores in shift analyses). No association was noted for maximum (4 studies; cOR_{adj} , 1.09 [95% CI, 0.98–1.21]; $I^2=0\%$, *P* value for Cochran Q test: 0.47; Figure S31A) or mean (4 studies; $cOR_{-1.23}$ [95% CI, 0.99–1.53]; $I^2=54\%$, *P* value for Cochran Q adj test: 0.09; Figure S31B) DBP levels in ordinal analysis.

Subgroup Analyses—We performed subgroup analyses among patients who achieved successful recanalization with MT and evaluated the association of post-MT BP with clinical outcomes. We observed that lower levels of maximum SBP (Figure S32A) and DBP (Figure S32B) were observed in patients with 3-month functional independence. Additionally, elevated levels of mean SBP (Figure S34A) were observed in patients with sICH. No difference was noted in mean SBP (Figure S33A) and DBP (Figure S33B) levels for patients with and without 3-month functional independence. Additionally, elevated levels of mean SBP (Figure S34A) were observed in patients with and without 3-month functional independence. Additionally, elevated levels of mean SBP (Figure S34A) were observed in patients with sICH. No differences were documented

In adjusted analyses, elevated levels of mean SBP (2 studies; OR_{adj} , 0.81 [95% CI, 0.71– 0.93]; I²=0%, *P* value for Cochran Q test: 0.95; Figure S36A) and maximum DBP (3 studies; OR_{adj} , 0.87 [95% CI, 0.77–0.99]; I²=0%, *P* value for Cochran Q test: 0.61; Figure S36B) were independently associated with reduced odds of 3-month functional independence, whereas no such association was noted for maximum SBP (3 studies; OR_{adj} , 0.87 [95% CI, 0.73–1.03]; I²=66%, *P* value for Cochran Q test: 0.05; Figure S37). No association was also detected for maximum SBP and 3-month mortality (3 studies; OR_{adj} , 1.04 [95% CI, 0.97–1.12]; I²=82%, *P* value for Cochran Q test: 0.004; Figure S38). Similarly, no association was documented for maximum SBP (4 studies; OR_{adj} , 1.01 [95% CI, 0.97–1.06]; I²=58%, *P* value for Cochran Q test: 0.07; Figure S39A) or mean SBP (2 studies; OR_{adj} , 1.01 [95% CI, 0.99–1.03]; I²=0%, *P* value for Cochran Q test: 0.66; Figure S39B) with sICH in adjusted analyses.

Publication Bias Assessment—We inspected funnel plot symmetry and Egger statistical test for outcomes involving 10 studies. No evidence of asymmetry or publication bias was observed in studies reporting unadjusted associations between pre-MT mean BP levels (*P*=0.572; Figure S40) or post-MT maximum BP levels (*P*=0.982; Figure S41) and 3-month functional independence. However, we detected evidence of publication bias in the adjusted associations between post-MT maximum BP parameters (*P*=0.034; Figure S42) and 3-month functional independence.

Discussion

Our systematic review and meta-analysis showed that increased BP levels before, during, and after MT are associated with adverse clinical outcomes including sICH, lack of successful recanalization, 3-month mortality, and 3-month functional dependence. These associations were consistent across adjusted and unadjusted analyses without substantial heterogeneity in the majority of independent associations. They also persisted in the subgroup of patients who achieved successful recanalization following MT.

Despite successful recanalization in \approx 70% to 90% of patients with endovascular thrombectomy, approximately half of the recanalized patients suffer poor functional outcomes or mortality at 3 months.^{3,56} Previous studies have suggested that elevated baseline BP levels are associated with higher clot burden, lower likelihood of recanalization and good functional outcomes, increased infarct volumes and early ischemic stroke recurrence.^{12,57,58} Modifiable factors including BP levels before and after endovascular procedures are potential targets to improve clinical outcomes in these patients. The findings of the present meta-analysis argue against elevated BP levels before and after MT. These observations are in line with a recent meta-analysis by our international collaborative group reporting that elevated pre and post-MT BP levels are associated with adverse functional outcomes and increased mortality in AIS patients treated with intravenous thrombolysis.⁴

Additionally, our results suggest reduced rates of recanalization after MT procedures among patients with elevated BP levels before and after the procedure. This observation can be attributed to the following plausible explanations (1) acute elevation of admission BP levels, among patients with uncontrolled hypertension, may adversely affect collateral flow,⁵⁹ and this may impart a stronger hemodynamic force resulting in a heavier mechanical clot impaction and impairment of mechanical clot retrieval³⁶ and (2) elevated admission SBP isassociated with poorer collateral flow and thereby results in lower rates of recanalization and functional improvement.¹² Consequently, the association between increased pretreatment BP levels and lower odds of recanalization after endovascular reperfusion could in part explain worse functional outcomes in patients with AIS.³⁶ This hypothesis is further supported by the present meta-analysis, as patients with AIS with successful recanalization following MT had lower pretreatment BP levels compared with patients with persistent LVO. Similar findings have been reported in patients with IVT-treated AIS.⁴ Elevated BP levels may lead to increased baseline thrombus burden and impaired endogenous capacity for fibrinolysis.⁶⁰ These findings argue against the hypothesis that higher sICH rates may represent the predominant causative link between increasing SBP levels and poor clinical outcomes.

Various hemodynamic factors including ischemic penumbra, collateral status, and clot burden play an important role in the optimization of BP in patients with AIS due to LVO.⁶¹ It is well established that the ischemic penumbra surrounding the infarct core in patients with LVO has impaired cerebral autoregulation and is sensitive to alterations in systemic blood pressure and reperfusion injury.^{7,61} Wide BP fluctuations, especially drop in BP levels during or within the first 24 hours after MT may result in penumbral tissue loss, exacerbation of reperfusion injury, and worse functional outcomes.^{62,63} Additionally, systemically elevated BP levels and increased BPV likely predispose to adverse outcomes in patients with AIS treated with MT.⁶⁴ Our study results lend support to the former consideration that wide fluctuations in BP after MT may induce infarct expansion or reperfusion injury leading to lower odds of functional independence and sICH.^{13,61} We documented an independent association of elevated maximum DBP levels following MT with sICH only in the adjusted analysis. The other associations between BP levels and sICH failed to achieve significance in the adjusted analyses due to limited number of studies with moderate sample. Nevertheless, the direction of all associations suggested the potential detrimental effect of increasing BP levels before or after MT on the risk of sICH. Additionally, we observed greater association of clinical outcomes with post-MT SBP levels and only with maximum DBP levels, likely due to small number of studies providing available data.

The findings of this meta-analysis support the recent American Heart Association/American Stroke Association guidelines for BP control in patients with AIS treated with IVT and MT.⁵ However, these recommendations also indicate that an optimal BP target which simultaneously avoids the risk of sICH and impairment of cerebral perfusion remains unknown (Class I; Level of Evidence B). The second arm of ENCHANTED trial (Enhanced Control of Hypertension and Thrombolysis Stroke Study)⁶⁵ provided randomized evidence that BP reduction in patients with AIS treated with IVT was safe but did not improve clinical outcomes despite lowering the incidence of any intracranial hemorrhage. The lack of an

association between BP lowering in patients with AIS treated with systemic reperfusion therapies and improved clinical outcome may be related to the modest BP reduction (5 mm Hg) that was achieved in the active group of ENCHANTED in comparison to the control group. Although our findings guide clinicians to avoid excessive BP elevations before, during, and after MT, these findings should be cautiously interpreted as extreme BP lowering in patients with LVO before successful revascularization with MT may exacerbate cerebral hypoperfusion and predispose to ischemic injury. Also, extreme BP reductions in patients with nonrecanalized LVO after MT with potentially viable penumbra may expand the infarct core and lead to worse outcomes. Consequently, acute decision-making warrants caution as the target SBP and DBP levels need to be individualized weighing the benefits and risks to the patients, including the extent of infarct core on postprocedural imaging, degree of recanalization, collateral blood flow, type of anesthesia, and prior history of hypertension.

Certain limitations of the present report need to be acknowledged. First, our analyses were primarily based on observational studies that were not designed in a randomized fashion to evaluate the intended associations. Thus, despite the use of adjusted outcomes for potential inherent biases related to the design of the included studies, unmeasured confounders (methodology and frequency of serial BP measurements, type and dose of antihypertensive medications used to treat excessive BP levels) cannot be eliminated. We decided a priori to exclude the studies that provided median (instead of mean) SBP or DBP levels, as well as those that reported mean arterial pressure levels or % drop in BP values whereas focused our analyses on maximum, minimum, and mean BP values. This decision was based on the fact that the majority of previous studies reported independent associations between maximum or mean SBP or DBP levels and outcomes, while mean arterial pressure goals are not included in international recommendations regarding optimal blood pressure management in patients with AIS treated with systemic or endovascular reperfusion therapies. Additionally, due to limited number of studies providing association data during MT, we failed to evaluate intraprocedural BP reduction and outcomes in patients with LVO treated with MT, and these methodological shortcomings argue against the aggressive BP reduction during the procedure. Second, although we performed an adjusted pairwise meta-analysis to account for the potential confounders available from the individual studies, the adjusted variables varied among included studies in the present meta-analysis with the exception of age and admission NIHSS-score (Table S3). Different BP metrics, time points, and techniques of BP measurements across the included studies represent another source of heterogeneity that cannot be assessed using I² and Cochran Q statistics, and this potential source of heterogeneity should be taken into account when interpreting our study findings. Third, few outcomes in adjusted analyses had substantial heterogeneity and these did not reach statistical significance, likely due to the small number of studies reporting the associations on the adjusted data for different confounding variables. Fourth, we performed no correction for multiple comparisons despite conducting a vast number of different analyses. This decision was made a priori during the preparation of our manuscript protocol. Nevertheless, it should be noted that many of the associations were highly significant (P<0.001) and reproducible in both unadjusted and adjusted analyses (Tables 1 and 2). Fifth, due to the way clinical data was presented among the included studies, it should be acknowledged that we

were unable to test the cause-effect association between increased BP levels and worse clinical outcomes of patients with LVO treated with MT. We were also unable to evaluate the hypothesis of a *U*-shaped relationship between BP parameters and clinical outcomes of patients with AIS treated with MT. Additionally, the effect of bridging therapy (IVT followed by MT) on the association of BP levels with outcomes could not be performed. These relationships could be evaluated in the settings of an individual patient data meta-analysis pooling large datasets from numerous MT registries. Sixth, the lack of independent association of SBP levels with available data. Last, due to limited available data, we were not able to test the impact of BP levels with admission BP-lowering medications, collateral status, final infarct volumes, choice of anesthesia that could potentially moderate the association of BP parameters with various clinical outcomes.^{66,67}

In conclusion, the present systematic review and meta-analysis provide preliminary evidence that elevated BP levels before, during, and after MT are likely associated with detrimental effect among patients with LVO treated with endovascular reperfusion therapies because they seem to increase the likelihood of mortality and functional dependence. These associations persisted even after adjustment for potential confounders and subgroup analyses of patients achieving successful recanalization during MT. Individual patient-data meta-analysis and randomized-controlled clinical trials are needed to assess the potential beneficial effect of moderate BP control among patients with LVO treated with MT.

Perspectives

This study evaluated the association of BP levels before, during, and after mechanical thrombectomy among patients with AIS with LVO. These findings further our understanding of the interplay between BP levels and clinical outcomes among patients with AIS with LVO undergoing mechanical thrombectomy.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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Novelty and Significance

What Is New?

• This is the first systematic review and meta-analysis involving 25 studies to evaluate the association of blood pressure (BP) on clinical outcomes among acute ischemic stroke patients with large vessel occlusion treated with mechanical thrombectomy (MT).

What Is Relevant?

- Higher SBP levels were noted among patients with functional dependenceand who died within 3 months.
- After adjustment of potential confounders, association persisted for elevated SBP levels before and after MT with lower odds of functional independence.Similarly, elevated SBP and DBP levels after MT were associated with symptomatic intracranial hemorrhage and mortality.
- Additional subgroup analyses among patients who achieved successful recanalization with MT demonstrated worse clinical outcomes with elevated BP levels.

Summary

Even after adjustment of potential confounders, elevated BP levels before and after MT are associated with worse clinical outcomes among acute ischemic stroke patients with large vessel occlusion.

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	mF	S 0-2		mF	(5 3-6		2	Std. Mean Difference	Std. Mean Difference
itudy or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
Anadani, 2019	121	11.5	131	125.5	12.5	167	14.2%	-0.37 [-0.60, -0.14]	
Cernik, 2019	133	15	332	139	16	358	17.6%	-0.39 [-0.54, -0.24]	
Chang (recanalized), 2019	125.5	12.4	49	132.7	10.1	41	7.9%	-0.63 [-1.05, -0.20]	
Chang, 2019	124.34	12.46	154	129.04	14.06	149	14.4%	-0.35 [-0.58, -0.13]	
Cho, 2019	124.2	13	149	129.9	14.6	229	15.2%	-0.41 [-0.61, -0.20]	
Goyal (post-MT, non-recanalized), 2018	140	13	22	141	17	59	6.6%	-0.06 [-0.55, 0.43]	
Goyal (post-MT, recanalized), 2017	138	15	97	135	15	120	12.7%	0.20 [-0.07, 0.47]	
Maier IL, 2018	127.2	13.8	74	131.9	14.4	94	11.3%	-0.33 [-0.64, -0.02]	
Total (95% CI)			1008			1217	100.0%	-0.30 [-0.45, -0.15]	•
Test for overall effect: $Z = 3.89 (P = 0.000)$	1)		-// -						-1 -0.5 0 0.5 Eavours [mPS 0_2] Eavours [mPS 3_6]
Test for overall effect: Z = 3.89 (P = 0.000 3	1)						Odds R	atio	-1 -0.5 0 0.5 Favours [mRS 0-2] Favours [mRS 3-6]
Test for overall effect: Z = 3.89 (P = 0.000 3 Study or Subgroup	1) log[Odds F	Ratio]	SE	Weig	ht IV,	Odds Ra Random	atio 1, 95% CI	-1 -0.3 Favours [mRS 0-2] Favours [mRS 3-6] Odds Ratio IV, Random, 95% CI
Test for overall effect: Z = 3.89 (P = 0.000 3 Study or Subgroup Anadani, 2019	1) log[Odds F	Ratio] 0.305	SE 0.104	Weig 29.3	<u>ht IV,</u> 7%	Odds Ra Random	atio 1, 95% Cl 60, 0.90]	-1 -0.3 Favours [mRS 0-2] Favours [mRS 3-6] Odds Ratio IV, Random, 95% CI
Test for overall effect: Z = 3.89 (P = 0.000 Study or Subgroup Anadani, 2019 Chang, 2019	1) log[Odds F -0	Ratio] 0.305 .0943	SE 0.104 0.1238	Weig 29.7 21.0	ht IV, 7% 0%	Odds Ra Random 0.74 [0.0 0.91 [0.7	atio 1 , 95% CI 60, 0.90] 71, 1.16]	-1 -0.3 GU -0.3 -6] Favours [mRS 0-2] Favours [mRS 3-6] Odds Ratio IV, Random, 95% CI
Test for overall effect: Z = 3.89 (P = 0.000 Study or Subgroup Anadani, 2019 Chang, 2019 Cho, 2019	1) log[Odds F -0. -0.	Ratio] 0.305 .0943 .1744	SE 0.104 0.1238 0.0991	Weig 29.7 21.0 32.7	<u>ht IV,</u> 7% 0% 7%	Odds Ra Random 0.74 [0.0 0.91 [0.1 0.84 [0.0	atio 1, 95% Cl 60, 0.90] 71, 1.16] 69, 1.02]	-1 -0.3 Favours [mRS 0-2] Favours [mRS 3-6] Odds Ratio IV, Random, 95% Cl
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Test for overall effect: Z = 3.89 (P = 0.000 Study or Subgroup Anadani, 2019 Chang, 2019 Cho, 2019 Ding, 2019 Goval (post-MT, non-recanalized), 201	1) log[8	Odds F -0. -0. -0. -0. -0.	Ratio] 0.305 .0943 .1744 .2485 .1054	SE 0.104 0.1238 0.0991 0.1754 0.4514	Weig 29.7 21.0 32.7 10.4	<u>ht</u> IV, 7% 0% 7% 4% 6%	Odds Ra Random 0.74 [0.0 0.91 [0.1 0.84 [0.0 0.78 [0.1 0.90 [0.1]	atio h, 95% Cl 60, 0.90] 71, 1.16] 69, 1.02] 55, 1.10] 37, 2.18]	-1 -0.3 Favours [mRS 0-2] Favours [mRS 3-6] Odds Ratio IV, Random, 95% CI
Test for overall effect: Z = 3.89 (P = 0.000 Study or Subgroup Anadani, 2019 Chang, 2019 Cho, 2019 Ding, 2019 Goyal (post-MT, non-recanalized), 201 Maier IL, 2018	1) log [8	Odds F -0. -0. -0. -0. -0. -0. -0.	Ratio] 0.305 .0943 .1744 .2485 .1054 0.619	SE 0.104 0.1238 0.0991 0.1754 0.4514 0.264	Weig 29.1 21.0 32.1 10.4 1.6 4.6	ht IV, 7% 0% 7% 4% 6% 6%	Odds Ra Random 0.74 [0.0 0.91 [0.1 0.84 [0.1 0.78 [0.1 0.90 [0.1 0.54 [0.1	atio h, 95% Cl 60, 0.90] 71, 1.16] 69, 1.02] 55, 1.10] 37, 2.18] 32, 0.90]	-1 -0.3 Favours [mRS 0-2] Favours [mRS 3-6] Odds Ratio IV, Random, 95% CI
Test for overall effect: Z = 3.89 (P = 0.000 Study or Subgroup Anadani, 2019 Chang, 2019 Cho, 2019 Ding, 2019 Goyal (post-MT, non-recanalized), 201 Maier IL, 2018 Total (95% CI)	1) log[8	Odds F -0. -0. -0. -0. -0. -0. -0.	Ratio] 0.305 .0943 .1744 .2485 .1054 0.619	SE 0.104 0.1238 0.0991 0.1754 0.4514 0.264	Weig 29.7 21.0 32.7 10.4 1.0 4.0	ht IV, 7% 7% 4% 6% 6% 0%	Odds Ra Random 0.74 [0.1 0.91 [0.1 0.84 [0.1 0.78 [0.1 0.90 [0.1 0.54 [0.1 0.80 [0.1	atio h, 95% Cl 60, 0.90] 71, 1.16] 69, 1.02] 55, 1.10] 37, 2.18] 32, 0.90] 72, 0.89]	-1 -0.3 Favours [mRS 0-2] Favours [mRS 3-6] Odds Ratio IV, Random, 95% CI
Test for overall effect: Z = 3.89 (P = 0.000 Study or Subgroup Anadani, 2019 Chang, 2019 Cho, 2019 Ding, 2019 Goyal (post–MT, non–recanalized), 2011 Maier IL, 2018 Total (95% CI) Heterogeneity: Tau ² = 0.00; Chi ² = 4.2:	1) log[8 8, df = 5	$\frac{\text{Odds f}}{-0} = -0.0000000000000000000000000000000000$	Ratio] 0.305 .0943 .1744 .2485 .1054 0.619	SE 0.104 0.1238 0.0991 0.1754 0.264 ² = 0%	Weig 29.1 21.0 32.1 10.4 1.0 4.0	ht IV, 7% 0% 7% 4% 6% 6% 0%	Odds Ra Random 0.74 [0. 0.91 [0.] 0.84 [0.] 0.78 [0.] 0.90 [0.] 0.90 [0.]	atio h, 95% Cl 60, 0.90] 71, 1.16] 69, 1.02] 55, 1.10] 37, 2.18] 32, 0.90] 72, 0.89]	-1 -0.3 Favours [mRS 0-2] Favours [mRS 3-6] Odds Ratio IV, Random, 95% CI

Figure 1.

Forest plot presenting the (**A**) unadjusted and (**B**) adjusted for potential confounders associations of post-mechanical thrombectomy (MT) mean systolic blood pressure levels with 3-mo functional independence. mRS indicates modified Rankin Scale.

	Mo	rtality	1	No n	nortali	ty	5	td. Mean Difference	Std. Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
Cho, 2019	160	25.2	32	157.8	25.7	346	21.2%	0.09 [-0.28, 0.45]	
Goyal (post-MT, non-recanalized), 2018	183	20	31	169	23	54	18.8%	0.63 [0.18, 1.08]	· · · · · · · · · · · · · · · · · · ·
Goyal (post-MT, recanalized), 2017	184	22	56	166	20	161	22.5%	0.87 [0.56, 1.19]	
Kim, 2019	171.9	19.9	12	155.2	24.4	199	15.5%	0.69 [0.10, 1.27]	
McCarthy, 2019	162	28	42	159	24	170	21.9%	0.12 [-0.22, 0.46]	
Total (95% CI)			173			930	100.0%	0.47 [0.12, 0.82]	
Heterogeneity: $Tau^2 = 0.12$; $Chi^2 = 15.74$, df = 4 (P = 0.4	003); 1	$^{2} = 75\%$					
Test for overall effect: $Z = 2.62$ (P = 0.00)	9)								-1 -0.5 0 0.5 1
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3							Odds Rat	0	Odds Ratio
3 Study or Subgroup	log[Oc	lds Ra	tio]	SE	Weigh	t IV,	Odds Rati Random, S	o 95% Cl	Odds Ratio IV, Random, 95% Cl
3 Study or Subgroup Cho, 2019	log[Oc	lds Ra	tio]	SE	Weigh 28.19	t IV,	Odds Rat Random, 9 1.00 [0.97	o 95% Cl 7, 1.03]	Odds Ratio IV, Random, 95% CI
Study or Subgroup Cho, 2019 Goyal (post-MT, non-recanalized), 2018	log[Oc	ids Ra 0.1	tio] 0 (0 133 (0	SE).0151).1491	Weigh	nt IV, %	Odds Rat Random, 9 1.00 [0.97 1.12 [0.84	o 95% CI , 1.03] , 1.50]	Odds Ratio IV, Random, 95% Cl
Study or Subgroup Cho, 2019 Goyal (post–MT, non–recanalized), 2018 Goyal (post–MT, recanalized), 2017	log[Oc	<mark>dds Ra</mark> 0.1 0.3	tio] 0 (0 133 (0 988 (0	SE).0151).1491).1186	Weigh 28.19 16.19 19.29	nt IV, % %	Odds Rat Random, 9 1.00 [0.97 1.12 [0.84 1.49 [1.18	50 95% CI (, 1.03] (, 1.50] (, 1.88]	Odds Ratio IV, Random, 95% CI
Study or Subgroup Cho, 2019 Goyal (post-MT, non-recanalized), 2018 Goyal (post-MT, recanalized), 2017 Kim, 2019	log[Oc	dds Ra 0.1 0.3 0.	tio] 0 (0 133 (0 988 (0 198	SE).0151).1491).1186 0.148	Weigh	nt IV, % % %	Odds Rat Random, 9 1.00 [0.97 1.12 [0.84 1.49 [1.18 1.22 [0.91	0 95% CI , 1.03] , 1.50] , 1.88] , 1.63]	Odds Ratio IV, Random, 95% CI
Study or Subgroup Cho, 2019 Goyal (post-MT, non-recanalized), 2018 Goyal (post-MT, recanalized), 2017 Kim, 2019 McCarthy, 2019	log[Oc	dds Ra 0.1 0.3 0. 0.2	tio] 0 (133 (988 (198 469 (SE).0151).1491).1186 0.148).1074	Weigh 28.19 16.19 19.29 16.29 20.39	it IV, % % %	Odds Rati Random, 9 1.00 [0.97 1.12 [0.84 1.49 [1.18 1.22 [0.91 1.28 [1.04	0 35% Cl , 1.03] , 1.50] , 1.88] , 1.63] , 1.58]	Odds Ratio IV, Random, 95% CI
Study or Subgroup Cho, 2019 Goyal (post-MT, non-recanalized), 2018 Goyal (post-MT, recanalized), 2017 Kim, 2019 McCarthy, 2019 Total (95% Cl)	_log[Oc	dds Ra 0.1 0.3 0. 0.2	tio] 0 (133 (988 (198 469 (SE).0151).1491).1186 0.148).1074	Weigh 28.19 16.19 19.29 16.29 20.39 100.09	t IV, % % % %	Odds Rati Random, 9 1.00 [0.97 1.12 [0.84 1.49 [1.18 1.22 [0.91 1.28 [1.04 1.19 [1.00	0 95% Cl , 1.03] , 1.50] , 1.88] , 1.63] , 1.58] , 1.43]	Odds Ratio IV, Random, 95% CI
S Study or Subgroup Cho, 2019 Goyal (post-MT, non-recanalized), 2018 Goyal (post-MT, recanalized), 2017 Kim, 2019 McCarthy, 2019 Total (95% CI) Heterogeneity: Tau ² = 0.03: Chi ² = 18.08	log[Oc	dds Ra 0.1 0.3 0. 0.2 (P = 0.	tio] 0 (133 (988 (198 469 (001):	SE).0151).1491).1186 0.148).1074	Weigh 28.19 16.19 19.29 16.29 20.39	nt IV, % % % %	Odds Rati Random, 9 1.00 [0.97 1.12 [0.84 1.49 [1.18 1.22 [0.91 1.28 [1.04 1.19 [1.00	0 95% Cl , 1.03] , 1.50] , 1.68] , 1.63] , 1.58] , 1.58] , 1.43]	Odds Ratio IV, Random, 95% CI
Study or Subgroup Cho, 2019 Goyal (post-MT, non-recanalized), 2018 Goyal (post-MT, recanalized), 2017 Kim, 2019 McCarthy, 2019 Total (95% Cl) Heterogeneity: Tau ² = 0.03; Chi ² = 18.08 Test for noverall effect: Z = 1.94 (P = 0.05)	$\log[Oc$	$\frac{dds \ Ra}{0.1}$ 0.1 0.3 0. 0.2 (P = 0.	tio] 0 (133 (988 (198 469 (001);	SE).0151).1491).1186 0.148).1074 ² = 78%	Weigh 28.19 16.19 19.29 16.29 20.39 100.09	nt IV, % % % %	Odds Rati Random, 9 1.00 [0.97 1.12 [0.84 1.49 [1.18 1.22 [0.91 1.28 [1.04 1.19 [1.00	io 95% CI , 1.03] , 1.50] , 1.88] , 1.63] , 1.58] , 1.43] 0.5_	Odds Ratio IV, Random, 95% CI

Figure 2.

Forest plot presenting the (**A**) unadjusted and (**B**) adjusted for potential confounders associations of post-mechanical thrombectomy (MT) maximum systolic blood pressure levels with 3-mo mortality. mRS indicates modified Rankin Scale.



Figure 3.

Forest plot presenting the (**A**) unadjusted and (**B**) adjusted for potential confounders associations of post-mechanical thrombectomy (MT) maximum diastolic blood pressure levels with symptomatic intracranial hemorrhage. mRS indicates modified Rankin Scale,

Table 1.

Overview of Primary and Secondary Analyses of Pretreatment Mean BP Association With Various Outcomes

			Una	djusted An	alyses		Adj	justed Analy	/ses*
Clinical Outcome	BP Level	Studies	SMD (95% CI)	P Value	Heterogeneity (I ² , P for Cochran Q)	Studies	OR (95% CI)	P Value	Heterogeneity (1 ² , <i>P</i> for Cochran Q)
FI	SBP	7	-0.26 (-0.34 to -0.17)	<0.001	0%, 0.88	4	0.86 (0.77 to 0.96)	600'0	18%, 0.30
	DBP	9	-0.16 (-0.25 to -0.07)	<0.001	0%, 0.74	-	I	-	I
Mortality	SBP	3	0.17 (0.04 to 0.29)	800.0	0%, 0.42	2	1.22 (1.00 to 1.49)	20.05	0%, 0.42
	DBP	3	0.16 (0.03 to 0.28)	0.01	0%, 0.64	2	1.19 (0.65 to 2.18)	0.57	10%, 0.29
sICH	SBP	3	-0.11 (-0.31 to 0.10)	0:30	19%, 0.29	-	I	-	I
	DBP	3	0.08 (-0.09 to 0.25)	0.37	0%, 0.44	Ι	I	Ι	I
Recanalization	SBP	2	-0.24 (-0.46 to -0.02)	0.03	0%, 0.91	-	I	-	I
	DBP	2	-0.15 (-0.48 to 0.18)	0.38	52%, 0.15	-	I	-	I
BD indicates blood we	DDD .	diastalia bla	lonoitonti II functional	indonondai	. O to more clock mithing to different of O	0. OD - 44	Motion CDD another M	annona boo	or of OU summaria internation

intrac SICH, symptoi blood pro -2); OR, odds ratio; SBP, systolic 5 đ ore Scale (modified BP indicates blood pressure; DBP, diastolic blood pressure; FI, functional independence hemorrhage; and SMD, standardized mean difference.

 $_{\star}^{*}$ In the adjusted for potential confounders analyses all associations of SBP/DBP with the outcomes of interest are presented per 10 mm Hg SBP/DBP increment.

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Overview of Primary and Secondary Analyses of Posttreatment BP Association With Various Outcomes

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				ijustea Ana	Iyses		AQ	Justea anal	yses	
Clinical Outcome	BP Level	Studies	SMD (95% CI)	P Value	Heterogeneity (1 ² , <i>P</i> for Cochran Q)	Studies	OR (95% CI)	P Value	Heterogeneity (1 ² , <i>P</i> for Cochran Q)	
FI	Max SBP	7	-0.47 (-0.62 to -0.31)	<0.001	61%, 0.02	8	0.82 (0.73 to 0.93)	0.001	72%, 0.0007	
	Min SBP	4	0.09 (-0.15 to 0.33)	0.47	60%, 0.06	2	1.25 (0.79 to 1.99)	0.34	64%, 0.10	
	Mean SBP	6	-0.27 (-0.42 to -0.12)	<0.001	65%, 0.004	9	0.80 (0.72 to 0.89)	<0.001	0%, 0.51	
	Max DBP	s	-0.24 (-0.39 to -0.09)	0.002	51%, 0.08	5	0.83 (0.72 to 0.96)	0.01	23%, 0.27	
	Min DBP	ε	-0.04 (-0.21 to 0.12)	0.61	0%, 0.48	I	I	I	1	
	Mean DBP	∞	-0.08 (-0.20 to 0.05)	0.23	43%, 0.09	3	0.80 (0.61 to 1.05)	0.10	44%, 0.17	i
Mortality	Max SBP	s	0.47 (0.12 to 0.82)	0.009	75%, 0.003	5	1.19 (1.00 to 1.43)	0.05	78%, 0.001	
	Min SBP	ю	-0.07 (-0.59 to 0.45)	0.79	76%, 0.01	I	I	I	I	
	Mean SBP	4	0.16 (-0.10 to 0.42)	0.22	39%, 0.18	2	1.02 (1.00 to 1.04)	0.06	0%, 0.35	
	Max DBP	4	0.45 (0.22 to 0.69)	0.0001	25%, 0.26	2	1.38 (1.09 to 1.76)	0.008	3%, 0.31	
	Min DBP	ε	-0.0 (-0.23 to 0.23)	66.0	0%, 0.49	I	I	I	1	
	Mean DBP	4	0.25 (0.01 to 0.49)	0.04	31%, 0.23	2	1.27 (0.91 to 1.78)	0.16	0%, 0.84	
sICH	Max SBP	4	0.17 (-0.16 to 0.50)	0.32	50%, 0.11	7	1.02 (0.99 to 1.05)	0.13	58%, 0.03	
	Min SBP	ю	0.17 (-0.15 to 0.48)	0.30	0%, 0.41	2	0.93 (0.78 to 1.12)	0.47	54%, 0.14	
	Mean SBP	4	0.16 (-0.15 to 0.46)	0.32	43%, 0.16	4	1.00 (0.99 to 1.02)	0.62	0%, 0.78	
	Max DBP	4	0.18 (-0.03 to 0.38)	0.10	0%, 0.73	2	1.65 (1.11 to 2.44)	0.01	0%, 0.80	
	Min DBP	ю	-0.08 (-0.39 to 0.24)	0.63	0%, 0.74	2	0.98 (0.95 to 1.01)	0.19	0%, 0.89	
	Mean DBP	4	0.07 (-0.14 to 0.27)	0.54	0%, 0.88	2	1.27 (0.91 to 1.78)	0.16	0%, 0.84	
mRS shift $^{\not au}$	Max SBP	ю	1.25 (1.13 to 1.37)	<0.001	41%, 0.18	5	1.15 (1.09 to 1.21)	<0.001	76%, 0.002	
	Mean SBP	4	1.27 (1.15 to 1.41)	<0.001	27%, 0.25	5	1.28 (1.17 to 1.39)	<0.001	7%, 0.36	
	Max DBP	ю	1.29 (1.17 to 1.43)	<0.001	0%, 0.72	4	1.09 (0.98 to 1.21)	0.13	0%, 0.47	i
	Mean DBP	4	1.23 (0.99 to 1.53)	0.07	54%, 0.09	4	1.23 (0.99 to 1.53)	0.07	54%, 0.09	
Recanalization	Max SBP	2	-0.09 (-0.62 to 0.44)	0.75	86%, 0.008	I	I	Ι	T	_
	Mean SBP	2	-0.37 (-0.96 to 0.22)	0.22	75%, 0.04	I	I	I	-	

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BP indicates blood pressure; DBP, diastolic blood pressure; FI, functional independence (modified Rankin Scale [mRS] score of 0–2); Max, maximum; Min, minimum; OR, odds ratio; SBP, systolic blood pressure; sICH, symptomatic intracranial hemorthage; and SMD, standardized mean difference.

* In the adjusted for potential confounders analyses, all associations of SBP/DBP with the outcomes of interest are presented per 10 mm Hg SBP/DBP increment.

 $\stackrel{\scriptstyle \ell}{\not}$ Defined as 1-point increase in mRS scores in ordinal logistic regression analyses.