# Original Article

# Predicting long-term multicategory cause of death in patients with prostate cancer: random forest versus multinomial model

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Abstract: The majority of patients with prostate cancer die of non-cancer causes of death (COD). It is thus important to accurately predict multi-category COD in these patients. Random forest (RF), a popular machine learning model, has been shown useful for predicting binary cancer-specific deaths. However, its accuracy for predicting multi-category COD in cancer patients is unclear. We included patients in Surveillance, Epidemiology, and End Results-18 cancer registry-program with prostate cancer diagnosed in 2004 (followed-up through 2016). They were randomly divided into training and testing sets with equal sizes. We evaluated prediction accuracies of RF and conventional statistical/multinomial models for 6-category COD by data-encoding types using the 2-fold cross-validation approach. Among 49,864 prostate cancer patients, 29,611 (59.4%) were alive at the end of follow-up, and 5,448 (10.9%) died of cardiovascular disease, 4,607 (9.2%) of prostate cancer, 3,681 (7.4%) of non-prostate cancer, 717 (1.4%) of infection, and 5,800 (11.6%) of other causes. We predicted 6-category COD among these patients with a mean accuracy of 59.1% (n=240, 95% CI, 58.7%-59.4%) in RF models with one-hot encoding, and 50.4% (95% CI, 49.7%-51.0%) in multinomial models. Tumor characteristics, prostate-specific antigen level, and diagnosis confirmation-method were important in RF and multinomial models. In RF models, no statistical differences were found between the accuracies of training versus cross-validation phases, and those of categorical versus one-hot encoding. We here report that RF models can outperform multinomial logistic models (absolute accuracy-difference, 8.7%) in predicting long-term 6-category COD among prostate cancer patients, while pathology diagnosis itself and tumor pathology remain important factors.

Keywords: Prostate cancer, cause-specific mortality, machine learning, prediction, prognosis

#### Introduction

Prostate cancer is the most prevalent cancer and the second leading-cause of cancer deaths among men in the U.S.A., accounting for 174,650 new cases and 31,620 deaths in 2019 [1, 2]. More patients with prostate cancer died of non-cancer causes than of prostate cancer [3, 4]. It is thus important to understand, predict and prevent non-cancer causes of death (CODs) among these patients, particu-

larly cardiovascular disease (CVD) [5]. However, only a limited number of studies investigated multi-category COD in prostate cancer patients, and none of them were focused on the prediction of COD [3, 5, 6].

The random forest (RF) model, a widely-used machine/statistical learning model, improves the performance of decision trees through random sampling of training data when building trees and random subsetting of features when

splitting nodes [7]. The RF model often outperforms several other machine learning and conventional statistical (e.g. logistic regression) models in predicting binary cancer-specific or all-cause deaths [8-12], with exceptions in a few simulation or biomarker studies [13, 14]. It hasalsobeenusedtopredictcancer-specificdeaths in prostate cancer patients [15]. However, few studies have used RF model for predicting multi-category COD in cancer patients, or compared the prediction accuracies of RF versus conventional statistical model (e.g. multinomial logistic regression) for multi-category COD. Our research aims to fill this gap, and we designed a population-based observational study to predict 12-year multi-category COD in prostate cancer patients using RF and multinomial logistic models.

#### Methods

## Patient data

We extracted individual-patient data from the Surveillance, Epidemiology, and End Results-18 (SEER-18) Program (www.seer.cancer.gov) SEER\*Stat Database with Treatment Data using the SEER\*Stat software (Surveillance Research Program, National Cancer Institute SEER\*Stat software (seer.cancer.gov/seerstat) version <8.3.6>) [16]. SEER-18 is the largest SEER database including cases from 18 states and covers near 30% of the U.S. population [17]. The datasets have been widely used and validated for research on breast and colorectal cancers [18-20]. Any summary data involving fewer than 15 patients were statistically suppressed to protect patient identity. Since the SEER database is an existing, de-identified and publicly available dataset, this study is exempt from Institutional Review Board (IRB) review under exempt category 4.

We included all qualified invasive prostate cancer cases in SEER-18 diagnosed in 2004 (2019 data-release, followed up through December 2016). The diagnosis year of 2004 was chosen because the 6<sup>th</sup> edition of the Tumor, Node and Metastasis staging manual (TNM6) of the American Joint Commission on Cancer (AJCC) was started in 2004 and allowed 12 years of follow-up. However, the AJCC 7<sup>th</sup> edition of the Tumor, Node and Metastasis staging manual (TNM7) was started in 2010, and would allow only up to 6 years of follow-up, which was not

long enough in our view. The inclusion criteria were survival time longer than 1 month, aged 20 years and older, with known COD and first primary only.

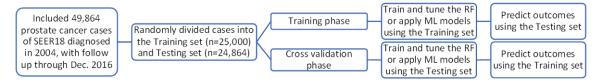
## Outcome and variables

The outcome of the statistical models was the patients' 6-category COD. The COD were originally classified using SEER's recodes of the causes of death according to the COD definition of the U.S. Mortality Data, which were extracted from underlying cause of death on the death certificates of deceased patients [21]. The underlying COD was the unique and most important etiology of the patients' death, while other causes may link to the death and be recorded as other COD on the death certificate. We simplified the SEER COD into 6 categories based on the prevalence of COD [3, 6, 15], including alive, CVD, infection, non-prostate cancer, prostate cancer and others.

The following factors were included in the analysis as variables in RF or multinomial models: age at diagnosis, race/ethnicity (non-Hispanic White, Hispanic, non-Hispanic Black, Asian and Pacific Islanders, and others) [22], T, N and M categories of TNM6, AJCC TNM6 clinical staging, prostate specific antigen level (PSA, ng/ ml), sum of the Gleason score, chemotherapy, radiotherapy, surgery, and attributes of the county where the patient resided at the time of diagnosis [23]. The PSA levels and Gleason scores were collected from medical records as site specific factors of prostate cancer since 2010 [24, 25]. Specifically, sums of the Gleason score were obtained from pathology report of resected specimen when available, or that of biopsy specimen if no surgery done. The 4 census-regions of patient's residence county were defined by the U.S. Census Bureau [26]. We converted continuous variables into 4-category variables based on their quartiles. The chemotherapy and radiotherapy data were obtained after signing a user agreement [25, 27]. It is noteworthy that no or unknown status of these treatments should be considered less reliable. while receipt of these treatments was generally confirmed and reliable [25, 27].

## Statistical analysis

We compared the accuracies of the RF and multinomial logit models after tuning the



**Figure 1.** Study flow. We randomized the patients into training and testing sets with similar sample sizes in each group. We then tuned the random forest (RF) model, chose the best-fit RF model, and fitted multinomial logistic (ML) model using the training set. Using the ML and RF models, we predicted 6-category causes of death among the patients in the testing set. During the cross-validation phase, we followed a similar protocol but with swapped datasets.

parameters of the RF model and choosing the model with the best accuracy. Using the twofold cross-validation approach, the patients were first randomly divided into two subsets of similar sizes (n=25,000 and 24,864, respectively). In each round of the validation, one subset is treated as the training data for constructing models, and the other subset is treated as the test data for evaluating prediction performance (Figure 1). We identified the RF model with the best accuracy, which is termed as tuning process in data science. Specifically, we examined prediction accuracies (i.e. 1 - classification error) of the models with various numbers of iterations (from 50 to 800 by an interval of 50) and variables (from 1 to 15), which were the number of computation rounds and the preset number of the features in RF model, respectively. After the two rounds of validation, the set of parameters that led to the RF model with the smallest average classification error were selected.

Several sensitivity analyses were performed on RF models. To exclude patients lost to follow-up, we conducted training and validation processes in the patients who died during the follow-up or was alive for >150 months (12.5 years).

A previous study has shown that one-hot encoding could sometimes outperform complex encoding systems [28]. This approach was also used in machine learning models of cancer driver genes [29]. For one-hot encoding, all multicategory variables (i.e. discrete variables with more than two categories) were transformed into a new set of binary variables. For example, the categorical variable for race/ethnicity group would be replaced by 5 binary variables representing whether the patients are non-Hispanic White, Hispanic, non-Hispanic Black, Asian and Pacific Islanders, or others, respectively. We

trained the RF or multinomial logit models using the one-hot encoded data, and compared the results with those using multicategory variables.

For the multinomial logistic regression model, we first constructed the model using the training set (the subset with sample size of 25,000) and predicted the 6-category COD using the testing set (**Figure 1**). If the predicted probability of a given COD was higher than 0.5 for a patient, the COD would be assigned to that patient. Ideally, only one COD had a predicted probability >0.5 and was allowed for each case, thus any patient with 0 or >1 predicted COD was considered mistakenly predicted using multinomial model.

We carried out the above statistical analyses using the RF package and multinomial logistic models of Stata (version 16, College Station, TX) [30-32]. The 95% confidence intervals (CI) of prediction accuracies were estimated using both binomial and Poisson models, that produced very similar results. All *P* values were two-tailed, and a *P* value < 0.05 was considered statistically significant.

#### Results

## **Patients**

We identified and analyzed 49,864 men with prostate cancer diagnosed in 2004 in the SEER-18 (**Table 1**), including 29,611 (59.4%) alive, 5,448 (10.9) died of CVD, 4,607 (9.2%) of prostate cancer, 3,681 (7.4%) of non-prostate cancer, 717 (1.4%) of infection, and 5,800 (11.6%) of other causes. The mean survival time was 117 months, while there were 31,273 patients who died during follow-up or were alive for >150 months. The majority of the cancers were of AJCC 6 stage 2 (80.9%) and not treated

**Table 1.** Baseline characteristics of included subjects

	Alive, n=29,611	CVD, n=5,448	Infection, n=717	Non-Prostate cancer, n=3,681	Other cause, n=5,800	Prostate cancer, n=4,607	Total, n=49,864
Age (yr)¶	63 (50-77)	74 (59-87)	75 (58-88)	70 (56-83)	73 (58-86)	72 (54-88)	67 (51-83)
Survival time (mo)¶	146 (131-155)	77 (7-143)	78 (6-141)	78 (12-141)	82 (11-143)	59 (4-137)	117 (16-154)
Race							
API	1,453	210	43	172	268	195	2,341
(%)	(4.9)	(3.9)	(6.0)	(4.7)	(4.6)	(4.2)	(4.7)
Hispanic	2,662	412	68	249	497	423	4,311
(%)	(9.0)	(7.6)	(9.5)	(6.8)	(8.6)	(9.2)	(8.7)
NH Black	3,830	865	143	553	807	812	7,010
(%)	(12.9)	(15.9)	(19.9)	(15.0)	(13.9)	(17.6)	(14.1)
NH White	21,093	3,920	461	2,690	4,189	3,147	35,500
(%)	(71.2)	(72.0)	(64.3)	(73.1)	(72.2)	(68.3)	(71.2)
Unknown/Other	573	41	<15*	17	39	30	702
(%)	(1.9)	(0.8)		(0.5)	(0.7)	(0.7)	(1.4)
TNM6 T category							
T1/2	26,641	4,873	635	3,245	5,201	2,917	43,512
(%)	(90.0)	(89.5)	(88.6)	(88.2)	(89.7)	(63.3)	(87.3)
T3/4	2,543	278	40	281	329	890	4,361
(%)	(8.6)	(5.1)	(5.6)	(7.6)	(5.7)	(19.3)	(8.8)
Unknown/Other	427	297	42	155	270	800	1,991
(%)	(1.4)	(5.5)	(5.9)	(4.2)	(4.7)	(17.4)	(4.0)
TNM6 N category	, ,	,	, ,	,	` ,	,	` ,
0	28,140	4,850	631	3,354	5,226	3,057	45,258
(%)	(94.7)	(88.3)	(87.2)	(90.6)	(89.4)	(65.2)	(90.2)
1	283	64	<15*	60	55	357	830
(%)	(1.0)	(1.2)		(1.6)	(0.9)	(7.6)	(1.7)
Unknown/Other	1,307	579	82	289	566	1,272	4,095
(%)	(4.4)	(10.5)	(11.3)	(7.8)	(9.7)	(27.1)	(8.2)
TNM6 M category	( /	(====)	(==)	(112)	(===)	(=:)	()
0	28,615	4,911	648	3,389	5,291	2,794	45,648
(%)	(96.3)	(89.4)	(89.5)	(91.5)	(90.5)	(59.6)	(91.0)
1	160	182	25	120	174	1,363	2,024
(%)	(0.5)	(3.3)	(3.5)	(3.2)	(3.0)	(29.1)	(4.0)
Unknown/Other	955	400	51	194	382	529	2,511
(%)	(3.2)	(7.3)	(7.0)	(5.2)	(6.5)	(11.3)	(5.0)
AJCC6 staging	(0.2)	(1.0)	(,	(0.2)	(0.0)	(22.0)	(0.0)
1	47	21	<15*	<15*	26	<15*	110
(%)	(0.2)	(0.4)	120	120	(0.5)	120	(0.2)
2	25476	4459	574	3013	4785	2054	40361
(%)	(86.0)	(81.9)	(80.1)	(81.9)	(82.5)	(44.6)	(80.9)
3	2110	173	19	200	216	328	3046
(%)	(7.1)	(3.2)	(2.7)	(5.4)	(3.7)	(7.1)	(6.1)
4	607	252	38	200	240	1595	2932
(%)	(2.1)	(4.6)	(5.3)	(5.4)	(4.1)	(34.6)	(5.9)
Unknown/Other	1371	543	81	258	533	629	3415
(%)	(4.6)	(10.0)	(11.3)	(7.0)	(9.2)	(13.7)	(6.9)
Chemotherapy	(4.0)	(±0.0)	(±±.0)	(1.0)	(3.2)	(13.7)	(0.9)
None/Unknown	20 617	5 //70	720	2 671	E 901	A 516	/Q 017
•	29,617	5,472		3,671	5,821	4,516	49,817
(%)	(99.6)	(99.6)	(99.5)	(99.1)	(99.6)	(96.4)	(99.3)
Received	113	21	<15*	32	26	170	366

(%)	(0.4)	(0.4)		(0.9)	(0.4)	(3.6)	(0.7)
Radiotherapy							
None/Unknown	18,450	3,364	446	2,094	3,537	3,257	31,148
(%)	(62.1)	(61.2)	(61.6)	(56.6)	(60.5)	(69.5)	(62.1)
Received	11,280	2,129	278	1,609	2,310	1,429	19,035
(%)	(37.9)	(38.8)	(38.4)	(43.5)	(39.5)	(30.5)	(37.9)
Surgery							
Local Excision	1,093	599	72	270	657	483	3,174
(%)	(3.7)	(11.0)	(10.0)	(7.3)	(11.3)	(10.5)	(6.4)
No surgery	15,142	4,261	578	2,649	4,413	3,666	30,709
(%)	(51.1)	(78.2)	(80.6)	(72.0)	(76.1)	(79.6)	(61.6)
Prostatectomy	13,376	588	67	762	730	458	15,981
(%)	(45.2)	(10.8)	(9.3)	(20.7)	(12.6)	(9.9)	(32.1)
Rural-urban continuum 2		, ,	, ,	, ,	, ,	, ,	, ,
Metro	26,709	4,758	635	3,178	4,958	4,039	44,277
(%)	(89.8)	(86.6)	(87.7)	(85.8)	(84.8)	(86.2)	(88.2)
Non-Metro	3,021	735	89	525	889	647	5,906
(%)	(10.2)	(13.4)	(12.3)	(14.2)	(15.2)	(13.8)	(11.8)
Census region	(===)	(====,	(==)	(= :-=)	(===)	(====)	(==:-)
Midwest	2,946	658	77	399	613	429	5,122
(%)	(10.0)	(12.1)	(10.7)	(10.8)	(10.6)	(9.3)	(10.3)
Northeast	4,797	882	123	631	874	721	8,028
(%)	(16.2)	(16.2)	(17.2)	(17.1)	(15.1)	(15.7)	(16.1)
South	5,573	1,140	176	843	1,393	1,009	10,134
(%)	(18.8)	(20.9)	(24.6)	(22.9)	(24.0)	(21.9)	(20.3)
West	16,295	2,768	341	1,808	2,920	2,448	26,580
(%)	(55.0)	(50.8)	(47.6)	(49.1)	(50.3)	(53.1)	(53.3)
Percent of education atta			(47.0)	(43.1)	(30.3)	(33.1)	(55.5)
Q1, <15.08	8,001	1,200	140	836	1,339	1,029	12,545
(%)	(26.9)	(21.9)	(19.3)	(22.6)	(22.9)	(22.0)	(25.0)
Q2, 15.09-18.15	7,538	1,287	182	898	1,448	1,193	12,546
(%)	(25.4)	(23.4)	(25.1)	(24.3)	(24.8)	(25.5)	(25.0)
Q3, 18.17-25.79	7,236	1,420	189	997	1,492	1,212	12,546
(%)	(24.3)	(25.9)	(26.1)	(26.9)	(25.5)	(25.9)	
Q4, >50.77			213	972			(25.0)
• •	6,955	1,586			1,568	1,252	12,546
(%)	(23.4)	(28.9)	(29.4)	(26.3)	(26.8)	(26.7)	(25.0)
Percent of persons in por 01, <21.18			160	865	1 205	1.044	10.610
• /	8,034	1,210	160		1,305	1,044	12,618
(%)	(27.0)	(22.0)	(22.1)	(23.4)	(22.3)	(22.3)	(25.1)
Q2, 21.33-29.81	7,655	1,258	152	929	1,364	1,129	12,487
(%)	(25.8)	(22.9)	(21.0)	(25.1)	(23.3)	(24.1)	(24.9)
Q3, 29.86-37.36	7,276	1,493	220	986	1,662	1,256	12,893
(%)	(24.5)	(27.2)	(30.4)	(26.6)	(28.4)	(26.8)	(25.7)
Q4, >67.40	6,765	1,532	192	923	1,516	1,257	12,185
(%)	(22.8)	(27.9)	(26.5)	(24.9)	(25.9)	(26.8)	(24.3)
Percent of foreign-born re		_					
Q1, <5.95	6,864	1,467	188	1,041	1,747	1,254	12,561
(%)	(23.1)	(26.7)	(26.0)	(28.1)	(29.9)	(26.8)	(25.0)
Q2, 5.98-15.22	7,739	1,399	172	922	1,498	1,157	12,887
(%)	(26.0)	(25.5)	(23.8)	(24.9)	(25.6)	(24.7)	(25.7)
Q3, 15.45-21.55	7,412	1,257	179	866	1,342	1,171	12,227
(%)	(24.9)	(22.9)	(24.7)	(23.4)	(23.0)	(25.0)	(24.4)
Q4, >38.52	7,715	1,370	185	874	1,260	1,104	12,508

(%)	(26.0)	(24.9)	(25.6)	(23.6)	(21.6)	(23.6)	(24.9)
Confirmation method of o	diagnosis						
Microscopic	29,628	5,321	697	3,652	5,688	4,223	49,209
(%)	(99.7)	(96.9)	(96.3)	(98.6)	(97.3)	(90.1)	(98.1)
Radiologic and clinic	40	122	21	43	104	285	615
(%)	(0.1)	(2.2)	(2.9)	(1.2)	(1.8)	(6.1)	(1.2)
Unknown/Other	62	50	<15*	<15*	55	178	359
(%)	(0.2)	(0.9)			(0.9)	(3.8)	(0.7)
PSA, quartiles (ng/ml)							
<4.9	8,360	765	88	665	874	367	11,119
(%)	(28.2)	(14.0)	(12.3)	(18.1)	(15.1)	(8.0)	(22.3)
5.0-6.8	7,406	829	108	735	1,023	337	10,438
(%)	(25.0)	(15.2)	(15.1)	(20.0)	(17.6)	(7.3)	(20.9)
6.9-11.3	6,331	1,199	157	804	1,239	580	10,310
(%)	(21.4)	(22.0)	(21.9)	(21.8)	(21.4)	(12.6)	(20.7)
11.3+	4,081	1,487	216	887	1,494	2,331	10,496
(%)	(13.8)	(27.3)	(30.1)	(24.1)	(25.8)	(50.6)	(21.1)
Unknown/Other	3,433	1,168	148	590	1,170	992	7,501
(%)	(11.6)	(21.4)	(20.6)	(16.0)	(20.2)	(21.5)	(15.0)
Gleason score							
5	<15*	<15*	<15*	<15*	<15*	<15*	15
6	264	37	<15*	21	26	19	374
(%)	(0.9)	(0.7)		(0.6)	(0.5)	(0.4)	(8.0)
7	219	29	<15*	22	28	32	335
(%)	(0.7)	(0.5)		(0.6)	(0.5)	(0.7)	(0.7)
8	36	15	<15*	<15*	<15*	20	94
(%)	(0.1)	(0.3)				(0.4)	(0.2)
9	<15*	<15*	<15*	<15*	<15*	<15*	65
(%)							(0.1)
10	<15*	<15*	<15*	<15*	<15*	<15*	<15*
Unknown/Other	29,069	5,358	701	3,624	5,728	4,493	48,973
(%)	(98.2)	(98.4)	(97.8)	(98.5)	(98.8)	(97.5)	(98.2)

Note: AJCC, 6th edition clinical staging of the American Joint Commission on Cancer; TNM6, 6th edition Tumor, node and metastasis staging manual of the American Joint Commission on Cancer; API, Asian Pacific Islanders; NH, Non-Hispanic; CVD, cardiovascular disease; PSA, Prostate specific antigen; \*, statistically suppressed; ¶, 95% confidence intervals in parenthesis; §, County attributes of Year 2000; Education attainment defined as percent of residents with less than high-school graduate in the county; Person in poverty defined as percent of residents with income below 200% of poverty in the county.

with prostatectomy (61.6%). We randomly divided the cases into training and testing sets ( $\underline{\text{Table S1}}$ ), and found the outcome and all covariates were similarly distributed in these sets, except radiotherapy status (P=0.047). For the sensitivity analyses on the patients who died during follow-up or were alive >150 months, CODs were similarly distributed in the training and testing sets (data not shown).

Predicting multi-category causes of death with random forests model

There were 17 variables with categorical encoding and 61 variables with one-hot encoding,

and 240 candidate models in each tuning process. Our tuning processes showed that the prediction accuracy increased with the iteration number in either conventionally or one-hot encoded data (**Figure 2**), as shown before [30]. The mean prediction-accuracy for 6-category COD were 58.6% (95% CI, 58.2%-59.1%) in the RF models with conventional encoding and 59.1% (95% CI, 58.7%-59.4%) in those with one-hot encoding. The best accuracy was reached in the model of 3 variables and 800 iterations with conventional encoding (59.2%, 95% CI [58.6%-59.8%], **Table 2** and **Figure 3**) and that of 1 variable and 700 iterations with one-hot encoding (59.6%, 95% CI [58.9%-

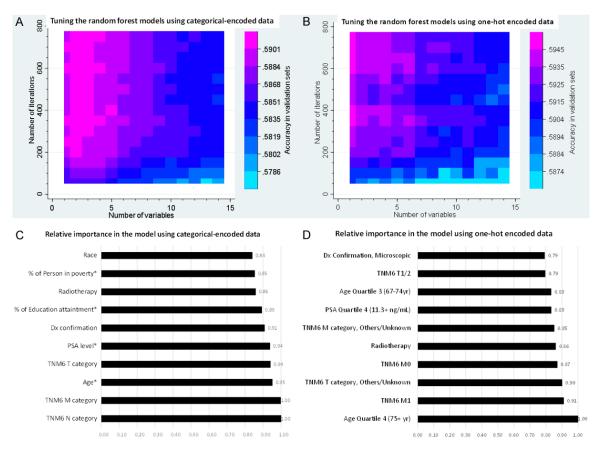


Figure 2. Characteristics of random forest models. During cross validation process, prediction accuracies of random forest models varied by the corresponding numbers of variable and iteration (Heatmap graphs: A. Categorical data encoding; B. One-hot data encoding). The random forest models provided relative importance values for all included variables (C and D. Relative importance values of the top 10 variables in the chosen random forest models using categorical encoding and one-hot encoding, respectively). Note: \*, Continuous variables were converted to 4-category variables by their respective quartiles; Dx, diagnosis; PSA, Prostate specific antigen; Education attainment defined as percent of residents with less than high-school graduate in the county; Person in poverty defined as percent of residents with income below 200% of poverty in the county.

**Table 2.** Prediction accuracy for long-term 6-category causes of death among the patients with prostate cancer diagnosis in 2004 (follow up through Dec. 2016)

Predicted classes	Alive, n=14,746	CVD, n=2,689	Infection, n=371	Non-Prostate cancer, n=1,873	Other cause, n=2,897	Prostate cancer, n=2,288	Total, n=24,864
Random forest model							
Alive, %	87.70*	52.73	52.29	67.49	55.82	39.9	73.75
CVD, %	3.79	15.88*	15.90	10.04	15.08	8.92	7.54
Infection, %	0.21	0.67	0.27*	0.32	0.69	0.31	0.33
Non-Prostate cancer, %	1.94	3.35	2.96	2.94*	3.11	3.23	2.44
Other cause, %	3.82	17.44	16.44	10.62	15.05*	10.01	7.87
Prostate cancer, %	2.54	9.93	12.13	8.60	10.25	37.63*	8.06
Multinomial model							
Alive, %	82.63*	33.51	31.27	51.84	37.04	32.87	64.34
NA, %	17.37	66.49	68.73	48.16	62.96	67.13	35.66

Note: CVD, cardiovascular disease; NA, not available; \*, correct prediction.

## **Prediction Accuracy**

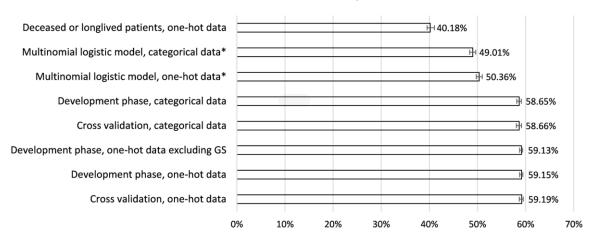


Figure 3. Summary of prediction accuracies by model and data type. In the tuning process and sensitivity analyses, we computed the validation accuracy of each random forest model (n=240), and chose the one with the best accuracy as the final model. The error bars show 95% confidence intervals of prediction accuracies in those models and data types during tuning process, except 3 models, whose 95% confidence intervals were calculated for the accuracy of a single binomial model (indicated by \*). One-hot indicates one-hot encoding of the data; balanced set refers to the sensitivity analysis with training and testing sets that had balanced distribution of all variables.

60.2%], Table S2 and Figure 3). The best RF model with one-hot encoding appeared to outperform that with conventional encoding, but no statistically significant difference was found. Alive was the COD that all RF models could predict with the best accuracy, while cancer pathological staging and age at diagnosis were topimportant factors in the RF models (Figure 3). The sensitivity analyses revealed that the prediction accuracies were statistically similar in the training phase, and cross-validation phase, but statistically lower in the models in patients who died during follow-up or was alive for >150 months (Figure 3).

Predicting multi-category causes of death with multinomial model

As the RF models, and the multinomial logistic regression models with one-hot encoding seemed to have better goodness of fit than with categorical encoding (pseudo/adjusted R-square =0.1707 versus 0.1416, respectively). Because multinomial models used a stepwise approach (i.e. identify the most possible COD versus others, then second most possible COD versus remaining possible COD...until the last possible COD) to determine the best-fit outcome, it is possible that more than one outcome (i.e. COD) had a probability ≥0.5. However, the predicted COD in multinomial model was

only unique in being alive among the 6-category COD and all other categories were of <0.5 probability (**Table 2**). The mean prediction-accuracy was 50.4% (95% CI, 49.7%-51.0%) in the multinomial models, and lower than RF models, except the RF model on the patients who died during follow-up or was alive for >150 months (**Figure 3**). Age at diagnosis, AJCC6 staging, confirmation method of diagnosis, surgery and PSA level were associated with all 6-category COD in multinomial model, while other factors were only linked to some of the 6-category COD (<u>Table S3</u>).

### Discussion

In this study, we investigated the multilevel prediction problem of prostate patients' COD using a carefully constructed RF model. In the patients with prostate cancer diagnosed in 2004, 59.4% were alive at the end of 12-year follow-up, while the top-3 CODs were CVD, prostate cancer and non-prostate cancer. We predicted 6-category COD among these patients with a mean accuracy of 59.1% (95% CI, 58.7%-59.4%) in the tuned RF model with one-hot encoding, and 50.4% (95% CI, 49.7%-51.0%) in the multinomial logit model, suggesting RF models outperformed multinomial model. Tumor characteristics, PSA level, diagnosis confirmation-method, and radiotherapy status were

the top-ranked variables in RF model, but only age, surgery, diagnosis confirmation-method, PSA level and AJCC6 stages as the factors were linked to all of the COD (versus alive) in multinomial models.

The proportions of various COD in our study are similar to those in previous reports [4]. Given the increasing proportion of deaths from COD other than prostate cancer, it is critical to accurately predict or identify the factors linked to these COD among prostate cancer patients. Several studies have attempted to predict cancer-specific or all-cause deaths in prostate cancer patients using clinical pathological and genomic/genetic factors [15, 33-36]. However, few studies to our knowledge predict the multicategory COD. Multinomial logistic regression is suitable for analyzing categorical/multi-category outcomes [31, 32]. In this study, multinomial logistic regression seems only able to predict the alive status of the 6-category COD if a unique COD was successfully identified. In the meantime, a tuned RF model outperformed multinomial logistic regression in predicting 6-category COD by 17.2% higher prediction accuracy (8.7% absolute accuracy-difference). This finding agrees with previous findings that RF's accuracy is similar to or better than support vector machines, artificial neural network and logistic regression in predicting various clinical outcomes [9-11, 37], but contrasts to a study showing that its accuracy is inferior to that of logistic regression [38]. It is plausible, but needs additional validation, that RF could also be highly useful in predicting multi-category COD or outcomes of other diseases. Despite the slightly better accuracy linked to data with one-hot encoding than standard encoding, we found no statistical differences between the two methods. This finding is inconsistent with previous reports [28, 29], and needs further validation. We also noticed that the minimal depths of trees in our best-fit RF models were usually 1 to 3. Those observations may help develop and improve machine learning models for predicting multi-category COD in cancer or other patients.

This study has several strengths that are noteworthy. First, this population-based study provides early evidence on the frequencies of various COD among the prostate cancer patients who were followed up for 12 years. Second, we

tuned RF models for predicting 6-category COD in prostate cancer patients, while existing RF models on prostate cancer only predicted binary cancer-specific death [15, 33], all-cause death [33, 39] or cancer recurrence [40]. Compared with binary death-outcomes, multiple-category COD are more informative, but much more difficult to predict. This is supported by the low success rate of multinomial models in predicting unique COD. Third, the tuned RF models in this study outperformed multinomial models in predicting 6-category COD. Indeed, the multinomial model was only able to predict alive as a unique COD, and missed other COD. Fourth, we characterized RF models and identified the model with best accuracy, while few of the prior works tuned their models [15, 33, 40]. Fifth, we are able to able to achieve a promising prediction accuracy given the large sample size of this prostate cancer dataset and the cross-validation procedure [41]. Some of previous studies on prostate cancer survivals using machine learning/RF models had either large sample sizes [15] or cross validation [42-44], but few combined both. Small sample size was indeed reported as the most common limitation of machine learning studies on cancer prognosis and prediction [41]. Finally, age, PSA level and tumor characteristics were found linked to long-term 6-category COD in prostate cancer patients in both RF and multinomial models. Three socioeconomic factors, including race, education attainment of patient's residence-county, average poverty-level of patient's residence-county, were found important in RF models, supporting a previous report on race and survival [15]. Future research is needed to further investigate these factors.

This study also has the following limitations. The prediction accuracy for 6-category COD in this study is not yet as good as prediction for binary outcomes, such as all-cause deaths [33]. Moreover, despite some shared factors, RF models did not completely agree with multinomial models on the factors linked to 6-category COD. However, RF and other machine learning models are known for their limitations in identifying associated factors [45]. In addition, an external validation dataset might be needed, but is unavailable, largely due to the lack of registry-data. SEER18 is the largest population cancer dataset in the North America [16]. Thus, it is very challenging to obtain anoth-

er population dataset of similar size for validation. However, we prospectively used the cross validation approach to validate our findings, as previously recommended [41, 45]. Finally, Gleason scores were available in a very small proportion of the patients, but might otherwise improve prediction accuracy [46].

#### Conclusions

In this population-based study, CVD, prostate cancer and non-prostate cancer were the most common long-term COD among prostate cancer patients. RF and multinomial models could predict 6-category COD among these patients with acceptable prediction accuracy, which needs further improvement. Those models enable clinicians to gain more granular prognostic information on prostate cancer patients, and target at the relevant COD to improve survival. We also show that a tuned RF model outperforms multinomial models by 8.7% (absolute difference), or 15,195 person-case for the cases diagnosed in 2019 alone in the U.S. Additional studies are needed to better predict multiple-category COD of other cancers.

## Disclosure of conflict of interest

None.

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## References

- [1] Siegel RL, Miller KD and Jemal A. Cancer statistics, 2019. CA Cancer J Clin 2019; 69: 7-34.
- [2] Miller KD, Nogueira L, Mariotto AB, Rowland JH, Yabroff KR, Alfano CM, Jemal A, Kramer JL and Siegel RL. Cancer treatment and survivorship statistics, 2019. CA Cancer J Clin 2019; 69: 363-385.
- [3] Zaorsky NG, Churilla TM, Egleston BL, Fisher SG, Ridge JA, Horwitz EM and Meyer JE. Causes of death among cancer patients. Ann Oncol 2017; 28: 400-407.
- [4] Epstein MM, Edgren G, Rider JR, Mucci LA and Adami HO. Temporal trends in cause of death among Swedish and US men with prostate cancer. J Natl Cancer Inst 2012; 104: 1335-42
- [5] Walter SD, de Koning HJ, Hugosson J, Talala K, Roobol MJ, Carlsson S, Zappa M, Nelen V, Kwi-

- atkowski M, Páez Á, Moss S and Auvinen A; ERSPC Cause of Death Committees. Impact of cause of death adjudication on the results of the European prostate cancer screening trial. Br J Cancer 2017; 116: 141-148.
- [6] Nguyen-Nielsen M, Møller H, Tjønneland A and Borre M. Causes of death in men with prostate cancer: results from the Danish Prostate Cancer Registry (DAPROCAdata). Cancer Epidemiol 2019; 59: 249-257.
- [7] Breiman L. Random Forests. Machine Learning 2001; 45: 5-32.
- [8] Sakr S, Elshawi R, Ahmed AM, Qureshi WT, Brawner CA, Keteyian SJ, Blaha MJ and Al-Mallah MH. Comparison of machine learning techniques to predict all-cause mortality using fitness data: the Henry ford exercise testing (FIT) project. BMC Med Inform Decis Mak 2017; 17: 174.
- [9] Peng SY, Chuang YC, Kang TW and Tseng KH. Random forest can predict 30-day mortality of spontaneous intracerebral hemorrhage with remarkable discrimination. Eur J Neurol 2010; 17: 945-50.
- [10] Shi M and He J. SNRFCB: sub-network based random forest classifier for predicting chemotherapy benefit on survival for cancer treatment. Mol Biosyst 2016; 12: 1214-23.
- [11] Montazeri M, Montazeri M, Montazeri M and Beigzadeh A. Machine learning models in breast cancer survival prediction. Technol Health Care 2016; 24: 31-42.
- [12] Bartholomai JA and Frieboes HB. Lung cancer survival prediction via machine learning regression, classification, and statistical techniques. Proc IEEE Int Symp Signal Proc Inf Tech 2018; 2018: 632-637.
- [13] van der Ploeg T, Austin PC and Steyerberg EW. Modern modelling techniques are data hungry: a simulation study for predicting dichotomous endpoints. BMC Med Res Methodol 2014; 14: 137.
- [14] Kim S, Park T and Kon M. Cancer survival classification using integrated data sets and intermediate information. Artif Intell Med 2014; 62: 23-31.
- [15] Hanson HA, Martin C, O'Neil B, Leiser CL, Mayer EN, Smith KR and Lowrance WT. The relative importance of race compared to health care and social factors in predicting prostate cancer mortality: a random forest approach. J Urol 2019; 202: 1209-1216.
- [16] Surveillance E; End Results (SEER) Program. SEER\*Stat Database: Incidence - SEER 18 Regs Research Data, Nov 2018 Sub (1975-2016) <Katrina/Rita Population Adjustment> - Linked To County Attributes - Total U.S., 1969-2017 Counties, National Cancer Institute, DC-CPS, based on the November 2018 submission. ed, 2019.

- [17] SEER. (Oct 27). Number of Persons by Race and Hispanic Ethnicity for SEER Participants (2010 Census Data). Available: https://web. archive.org/web/20191028021627/https:// seer.cancer.gov/registries/data.html.
- [18] Chavali LB, Llanos AAM, Yun JP, Hill SM, Tan XL and Zhang L. Radiotherapy for patients with resected tumor deposit-positive colorectal cancer: a surveillance, epidemiology, and end results-based population study. Arch Pathol Lab Med 2018; 142: 721-729.
- [19] Yang M, Bao W, Zhang X, Kang Y, Haffty B and Zhang L. Short-term and long-term clinical outcomes of uncommon types of invasive breast cancer. Histopathology 2017; 71: 874-886.
- [20] Mayo E, Llanos AA, Yi X, Duan SZ and Zhang L. Prognostic value of tumour deposit and perineural invasion status in colorectal cancer patients: a SEER-based population study. Histopathology 2016; 69: 230-8.
- [21] SEER. (Oct. 28). SEER Cause of Death Recode 1969+ (03/01/2018). Available: https://web. archive.org/web/20191028030412/https:// seer.cancer.gov/codrecode/1969\_d030120-18/index.html.
- [22] SEER. (Oct. 28). Race Recode Changes: For the 1973-2005 SEER Research Data (November 2007 Submission) and Later Releases. Available: https://web.archive.org/web/2019-1028023614/https://seer.cancer.gov/seerstat/variables/seer/race\_ethnicity/.
- [23] SEER. (2019, Oct. 27). County attributes. Available: https://web.archive.org/web/201910-28025023/https://seer.cancer.gov/seerstat/variables/countyattribs/.
- [24] SEER. (2013, May 3). Collaborative Stage Data Set: Prostate. Available: https://web.archive. org/web/20190517115038/http://web2. facs.org/cstage0205/prostate/Prostateschema.html.
- [25] Guo Y, Mao S, Zhang A, Zhang J, Wang L, Wang R, Zhang W, Zhang Z, Wu Y, Cao X, Yang B and Yao X. Survival significance of patients with low prostate-specific antigen and high-grade prostate cancer after radical prostatectomy, external beam radiotherapy, or external beam radiotherapy with brachytherapy. Front Oncol 2019; 9: 638.
- [26] Bureau UC. (Dec. 19). Geographic Terms and Concepts Census Divisions and Census Regions. Available: https://www.census.gov/geo/reference/gtc/gtc\_census\_divreg.html.
- [27] SEER. (May 3). Radiation/Chemotherapy Databases (1975-2016). Available: https://web.archive.org/save/https://seer.cancer.gov/data/treatment.html.
- [28] Waldmann P. Approximate Bayesian neural networks in genomic prediction. Genet Sel Evol 2018; 50: 70.

- [29] Agajanian S, Oluyemi O and Verkhivker GM. Integration of random forest classifiers and deep convolutional neural networks for classification and biomolecular modeling of cancer driver mutations. Front Mol Biosci 2019; 6: 44.
- [30] Zou R and Schonlau M. (2018, Oct. 15). Applications of Random Forest Algorithm. Available: https://web.archive.org/web/2019101412-5205/https://www.stata.com/meeting/canada18/slides/canada18\_Zou.pdf.
- [31] Long JS and Freese J. Regression models for categorical dependent variables using Stata. Stata Press; 2006.
- [32] (Oct. 31). Multinomial Logistic Regression Stata data analysis examples. Available: https://web.archive.org/web/20181010004634/https://stats.idre.ucla.edu/stata/dae/multinomiallogistic-regression/.
- [33] Lin YT, Lee MT, Huang YC, Liu CK, Li YT and Chen M. Prediction of recurrence-associated death from localized prostate cancer with a charlson comorbidity index-reinforced machine learning model. Open Med (Wars) 2019; 14: 593-606.
- [34] Kleppe A, Albregtsen F, Vlatkovic L, Pradhan M, Nielsen B, Hveem TS, Askautrud HA, Kristensen GB, Nesbakken A, Trovik J, Wæhre H, Tomlinson I, Shepherd NA, Novelli M, Kerr DJ and Danielsen HE. Chromatin organisation and cancer prognosis: a pan-cancer study. Lancet Oncol 2018; 19: 356-369.
- [35] Cruz JA and Wishart DS. Applications of machine learning in cancer prediction and prognosis. Cancer Inform 2007; 2: 59-77.
- [36] Carmona R, Zakeri K, Green G, Hwang L, Gulaya S, Xu B, Verma R, Williamson CW, Triplett DP, Rose BS, Shen H, Vaida F, Murphy JD and Mell LK. Improved method to stratify elderly patients with cancer at risk for competing events. J Clin Oncol 2016; 34: 1270-7.
- [37] Song Y, Gao S, Tan W, Qiu Z, Zhou H and Zhao Y. Multiple machine learnings revealed similar predictive accuracy for prognosis of PNETs from the Surveillance, Epidemiology, and End Result Database. J Cancer 2018; 9: 3971-3978.
- [38] van der Ploeg T, Nieboer D and Steyerberg EW. Modern modeling techniques had limited external validity in predicting mortality from traumatic brain injury. J Clin Epidemiol 2016; 78: 83-89.
- [39] Zhang Y, Yan L, Zeng J, Zhou H, Liu H, Yu G, Yao W, Chen K, Ye Z and Xu H. Pan-cancer analysis of clinical relevance of alternative splicing events in 31 human cancers. Oncogene 2019; 38: 6678-6695.
- [40] Shen J, Wang L and Taylor JMG. Estimation of the optimal regime in treatment of prostate cancer recurrence from observational data us-

## Random forest model for multicategory death-cause of prostate cancer

- ing flexible weighting models. Biometrics 2017; 73: 635-645.
- [41] Kourou K, Exarchos TP, Exarchos KP, Karamouzis MV and Fotiadis DI. Machine learning applications in cancer prognosis and prediction. Comput Struct Biotechnol J 2014; 13: 8-17.
- [42] Fa B, Luo C, Tang Z, Yan Y, Zhang Y and Yu Z. Pathway-based biomarker identification with crosstalk analysis for robust prognosis prediction in hepatocellular carcinoma. EBioMedicine 2019; 44: 250-260.
- [43] Hussain L, Ahmed A, Saeed S, Rathore S, Awan IA, Shah SA, Majid A, Idris A and Awan AA. Prostate cancer detection using machine learning techniques by employing combination of features extracting strategies. Cancer Biomark 2018; 21: 393-413.
- [44] Agranoff D, Fernandez-Reyes D, Papadopoulos MC, Rojas SA, Herbster M, Loosemore A, Tarelli E, Sheldon J, Schwenk A, Pollok R, Rayner CF and Krishna S. Identification of diagnostic markers for tuberculosis by proteomic fingerprinting of serum. Lancet 2006; 368: 1012-21.
- [45] Shameer K, Johnson KW, Glicksberg BS, Dudley JT and Sengupta PP. Machine learning in cardiovascular medicine: are we there yet? Heart 2018; 104: 1156-1164.
- [46] Rodrigues G, Warde P, Pickles T, Crook J, Brundage M, Souhami L and Lukka H; Genitourinary Radiation Oncologists of Canada. Pretreatment risk stratification of prostate cancer patients: a critical review. Can Urol Assoc J 2012; 6: 121-7.

 Table S1. Baseline characteristics of the training and validation sets

	Training set (n=25,000)	Testing set (n=24,864)	Total (n=49,864)	P value
Survival time (mo)	117 (16-154)	117 (16-154)	117 (16-154)	0.183
Cause of death				0.663
Alive	14865	14746	29611	
(%)	(59.5)	(59.3)	(59.4)	
CVD	2759	2689	5448	
(%)	(11.0)	(10.8)	(10.9)	
Infection	346	371	717	
(%)	(1.4)	(1.5)	(1.4)	
Non-Prostate cancer	1808	1873	3681	
(%)	(7.2)	(7.5)	(7.4)	
Other cause	2903	2897	5800	
(%)	(11.6)	(11.7)	(11.6)	
Prostate cancer	2319	2288	4607	
(%)	(9.3)	(9.2)	(9.2)	
Age at diagnosis, quartile (yr)		•	•	0.539
<61	6665	6502	13167	
(%)	(26.7)	(26.2)	(26.4)	
61-67	6303	6353	12656	
(%)	(25.2)	(25.6)	(25.4)	
67-74	6040	6043	12083	
(%)	(24.2)	(24.3)	(24.2)	
75+	5992	5966	11958	
(%)	(24.0)	(24.0)	(24.0)	
Race				0.317
API	1184	1157	2341	
(%)	(4.7)	(4.7)	(4.7)	
Hispanic	2098	2213	4311	
(%)	(8.4)	(8.9)	(8.7)	
NH Black	3500	3510	7010	
(%)	(14.0)	(14.1)	(14.1)	
NH White	17859	17641	35500	
(%)	(71.4)	(71.0)	(71.2)	
Unknown/Other	359	343	702	
(%)	(1.4)	(1.4)	(1.4)	
TNM6 T category				0.850
T1/2	21823	21689	43512	
(%)	(87.3)	(87.2)	(87.3)	
T3/4	2191	2170	4361	
(%)	(8.8)	(8.7)	(8.8)	
Unknown/Other	986	1005	1991	
(%)	(3.9)	(4.0)	(4.0)	
TNM6 N category	,	,	,	0.744
0	22638	22469	45107	
(%)	(90.6)	(90.4)	(90.5)	
1	406	420	826	
(%)	(1.6)	(1.7)	(1.7)	
Unknown/Other	1956	1975	3931	
		(7.9)	(7.9)	

TNM6 M category				0.997
0	22813	22686	45499	
(%)	(91.3)	(91.2)	(91.3)	
1	1009	1003	2012	
(%)	(4.0)	(4.0)	(4.0)	
Unknown/Other	1178	1175	2353	
(%)	(4.7)	(4.7)	(4.7)	
AJCC6 staging				0.662
1	62	48	110	
(%)	(0.3)	(0.2)	(0.2)	
2	20258	20103	40361	
(%)	(81.0)	(80.9)	(80.9)	
3	1529	1517	3046	
(%)	(6.1)	(6.1)	(6.1)	
4	1452	1480	2932	
(%)	(5.8)	(6.0)	(5.9)	
Unknown/Other	1699	1716	3415	
(%)	(6.8)	(6.9)	(6.9)	
Chemotherapy				0.793
None/Unknown	24820	24680	49500	
(%)	(99.3)	(99.3)	(99.3)	
Received	180	184	364	
(%)	(0.7)	(0.7)	(0.7)	
Radiotherapy				0.047
None/Unknown	15578	15278	30856	
(%)	(62.3)	(61.5)	(61.9)	
Received	9422	9586	19008	
(%)	(37.7)	(38.6)	(38.1)	
Surgery				0.389
Local Excision	1610	1564	3174	
(%)	(6.4)	(6.3)	(6.4)	
No surgery	15323	15386	30709	
(%)	(61.3)	(61.9)	(61.6)	
Prostatectomy	8067	7914	15981	
(%)	(32.3)	(31.8)	(32.1)	
Rural-urban continuum 2003§				0.887
Metro	22084	21974	44058	
(%)	(88.3)	(88.4)	(88.4)	
Non-Metro	2916	2890	5806	
(%)	(11.7)	(11.6)	(11.6)	
Census region				0.315
Midwest	2628	2494	5122	
(%)	(10.5)	(10.0)	(10.3)	
Northeast	4033	3995	8028	
(%)	(16.1)	(16.1)	(16.1)	
South	5081	5053	10134	
(%)	(20.3)	(20.3)	(20.3)	
West	13258	13322	26580	
(%)	(53.0)	(53.6)	(53.3)	
Percent of education attainment,	quartile§			0.456

Q1, <15.08	6450	6296	12746	
(%)	(25.8)	(25.3)	(25.6)	
Q2, 15.09-18.15	6293	6213	12506	
(%)	(25.2)	(25.0)	(25.1)	
Q3, 18.17-25.79	6233	6246	12479	
(%)	(24.9)	(25.1)	(25.0)	
Q4, >50.77	6024	6109	12133	
(%)	(24.1)	(24.6)	(24.3)	
Percent of persons in poverty,	quartile§			0.182
Q1, <21.18	6406	6173	12579	
(%)	(25.6)	(24.8)	(25.2)	
Q2, 21.33-29.81	6212	6170	12382	
(%)	(24.9)	(24.8)	(24.8)	
Q3, 29.86-37.36	6375	6448	12823	
(%)	(25.5)	(25.9)	(25.7)	
Q4, >67.40	6007	6073	12080	
(%)	(24.0)	(24.4)	(24.2)	
Percent of foreign-born resider	, ,	,	,	0.223
Q1, <5.95	6259	6080	12339	
(%)	(25.0)	(24.5)	(24.8)	
Q2, 5.98-15.22	6481	6365	12846	
(%)	(25.9)	(25.6)	(25.8)	
Q3, 15.45-21.55	6068	6124	12192	
(%)	(24.3)	(24.6)	(24.5)	
Q4, >38.52	6192	6295	12487	
(%)	(24.8)	(25.3)	(25.0)	
Confirmation method of diagno		(20.0)	(20.0)	0.686
Microscopic	24553	24407	48960	0.000
(%)	(98.2)	(98.2)	(98.2)	
Radiologic and clinic	306	302	608	
(%)	(1.2)	(1.2)	(1.2)	
Unknown/Other	141	155	296	
(%)	(0.6)	(0.6)	(0.6)	
PSA, quartiles (ng/ml)	(0.0)	(0.0)	(0.0)	0.854
<4.9	5546	5573	11119	0.004
(%)	(22.2)	(22.4)	(22.3)	
5.0-6.8	5231	5207	10438	
(%)	(20.9)	(20.9)	(20.9)	
6.9-11.3	5179	5131	10310	
(%)	(20.7)	(20.6)	(20.7)	
11.3+	5243	5253	10496	
(%)	(21.0)	(21.1)	(21.1)	
Unknown/Other	3801	3700	7501	
(%)	(15.2)	(14.9)	(15.0)	
	(15.2)	(14.9)	(15.0)	0.957
Gleason score 5	<15*	<15*	15	0.931
(%)	<b>\10</b>	/TO	(0.0)	
6	192	182	374	
(%)	(0.8)	(0.7)	(0.8)	
(70)	(0.0)	(0.1)	(0.0)	

## Random forest model for multicategory death-cause of prostate cancer

7	169	166	335	
(%)	(0.7)	(0.7)	(0.7)	
8	50	44	94	
(%)	(0.2)	(0.2)	(0.2)	
9	35	30	65	
(%)	(0.1)	(0.1)	(0.1)	
10	<15*	<15*	<15*	
(%)				
Unknown/Other	24541	24432	48973	
(%)	(98.2)	(98.3)	(98.2)	

Note: AJCC, 6<sup>th</sup> edition clinical staging of the American Joint Commission on Cancer; API, Asian Pacific Islanders; NH, Non-Hispanic; TNM6, 6<sup>th</sup> edition Tumor, node and metastasis staging manual of the American Joint Commission on Cancer; CVD, cardiovascular disease; PSA, Prostate specific antigen; \*, statistically suppressed; §, Country attributes of Year 2000; Education attainment defined as percent of residents with less than high-school graduate in the county; Person in poverty defined as percent of residents with income below 200% of poverty in the county.

**Table S2.** Prediction accuracy for long-term 6-category causes of death among the patients with prostate cancer diagnosed in 2004 (12-year follow up) using one-hot encoding

Predicted classes	Alive, n=14,746	CVD, n=2,689	Infection, n=371	Non-Prostate cancer, n=1,873			Total, n=24,864
Random forest model							
Alive, %	88.87*	56.56	57.68	69.62	59.20	41.78	75.67
CVD, %	3.54	14.95*	13.75	9.24	13.95	8.65	7.04
Infection, %	0.20	0.60	0.27*	0.27	0.62	0.39	0.31
Non-Prostate cancer, %	1.70	2.98	2.43	2.72*	2.93	2.80	2.17
Other cause, %	3.49	15.43	15.09	9.72	13.84*	9.35	7.17
Prostate cancer, %	2.20	9.48	10.78	8.44	9.46	37.02*	7.64
Multinomial model							
Alive, %	84.90*	34.85	32.35	53.66	39.28	27.93	65.79
Prostate cancer, %	0.01	0.00	0.27	0.00	0.00	0.09*	0.02
NA, %	15.09	65.15	67.39	46.34	60.72	71.98	34.19

Note: CVD, cardiovascular disease; NA, not available;  $\star$ , correct prediction.

# Random forest model for multicategory death-cause of prostate cancer

Table S3. Factors associated with long-term 6-category cause of death among men with prostate cancer in multinomial model

					Cause of deat	th				
Covariate	CVD		Infection		Non-Prostate car	ncer	Other causes		Prostate cance	er
	coefficient (95% CI)	P	coefficient (95% CI)	P	coefficient (95% CI)	Р	coefficient (95% CI)	P	coefficient (95% CI)	Р
Age at diagnosis, quartile (yr)*	0.92 (0.88 to 0.97)	<0.001	0.91 (0.78 to 1.03)	<0.001	0.62 (0.57 to 0.67)	<0.001	0.84 (0.79 to 0.88)	<0.001	0.51 (0.46 to 0.55)	<0.001
AJCC6 staging*	0.15 (0.04 to 0.27)	0.007	0.34 (0.10 to 0.58)	0.005	0.15 (0.02 to 0.29)	0.021	0.14 (0.02 to 0.25)	0.019	0.68 (0.58 to 0.77)	<0.001
Confirmation method of diagnosis*	0.54 (0.24 to 0.85)	0.001	0.70 (0.16 to 1.24)	0.011	0.29 (-0.09 to 0.68)	0.136	0.69 (0.39 to 0.98)	<0.001	0.66 (0.39 to 0.93)	<0.001
Surgery*	-0.76 (-0.86 to -0.66)	<0.001	-0.84 (-1.07 to -0.61)	<0.001	-0.45 (-0.56 to -0.34)	<0.001	-0.79 (-0.88 to -0.70)	<0.001	-1.04 (-1.14 to -0.94)	<0.001
PSA, quartiles (ng/ml)*	0.17 (0.13 to 0.20)	<0.001	0.16 (0.07 to 0.25)	<0.001	0.13 (0.09 to 0.17)	<0.001	0.14 (0.11 to 0.17)	<0.001	0.34 (0.30 to 0.38)	<0.001
Chemotherapy	0.23 (-0.45 to 0.92)	0.505	-0.11 (-2.11 to 1.88)	0.911	0.73 (0.11 to 1.36)	0.022	0.00 (-0.74 to 0.74)	0.999	2.03 (1.63 to 2.44)	<0.001
Census region	-0.12 (-0.17 to -0.07)	<0.001	-0.15 (-0.27 to -0.03)	0.016	-0.04 (-0.10 to 0.02)	0.160	-0.01 (-0.06 to 0.04)	0.829	0.06 (0.00 to 0.12)	0.048
Percent of education attainment, quartile§	0.08 (0.01 to 0.15)	0.019	0.11 (-0.06 to 0.28)	0.198	0.14 (0.06 to 0.21)	0.001	0.11 (0.04 to 0.17)	0.002	0.08 (0.01 to 0.16)	0.031
Percent of persons in poverty, quartile§	0.08 (0.01 to 0.15)	0.032	0.01 (-0.16 to 0.18)	0.929	-0.06 (-0.14 to 0.02)	0.161	0.01 (-0.06 to 0.08)	0.788	-0.03 (-0.11 to 0.05)	0.436
Percent of foreign-born residents, quartile§	-0.08 (-0.13 to -0.03)	0.003	-0.11 (-0.24 to 0.02)	0.109	-0.10 (-0.16 to -0.04)	0.001	-0.19 (-0.24 to -0.14)	<0.001	-0.13 (-0.18 to -0.07)	<0.001
Gleason score	-0.03 (-0.13 to 0.07)	0.566	-0.07 (-0.30 to 0.16)	0.565	-0.04 (-0.15 to 0.07)	0.454	0.06 (-0.05 to 0.17)	0.257	-0.05 (-0.16 to 0.06)	0.341
Race	0.03 (-0.02 to 0.08)	0.292	-0.23 (-0.34 to -0.11)	<0.001	0.04 (-0.03 to 0.10)	0.252	0.01 (-0.05 to 0.06)	0.796	-0.03 (-0.09 to 0.03)	0.267
Radiotherapy	-0.22 (-0.31 to -0.12)	<0.001	-0.40 (-0.63 to -0.16)	0.001	-0.07 (-0.18 to 0.04)	0.238	-0.22 (-0.32 to -0.13)	<0.001	-0.37 (-0.48 to -0.26)	<0.001
Rural-urban continuum 2003§	-0.03 (-0.18 to 0.13)	0.720	-0.23 (-0.64 to 0.17)	0.261	0.11 (-0.07 to 0.28)	0.239	-0.03 (-0.17 to 0.12)	0.739	-0.01 (-0.18 to 0.17)	0.948
TNM6 T category	0.08 (-0.05 to 0.21)	0.242	0.15 (-0.15 to 0.45)	0.335	0.19 (0.03 to 0.35)	0.017	0.03 (-0.11 to 0.16)	0.714	0.50 (0.39 to 0.61)	<0.001
TNM6 N category	-0.05 (-0.13 to 0.02)	0.190	-0.08 (-0.24 to 0.09)	0.365	-0.08 (-0.17 to 0.01)	0.065	-0.05 (-0.12 to 0.03)	0.218	-0.18 (-0.24 to -0.12)	<0.001
TNM6 M category	0.08 (-0.08 to 0.24)	0.322	-0.34 (-0.68 to 0.00)	0.053	0.06 (-0.13 to 0.25)	0.511	0.06 (-0.10 to 0.22)	0.447	0.06 (-0.07 to 0.19)	0.357

Note: AJCC, 6<sup>th</sup> edition clinical staging of the American Joint Commission on Cancer; TNM6, 6<sup>th</sup> edition Tumor, node and metastasis staging manual of the American Joint Commission on Cancer; CVD, cardiovascular disease; PSA, Prostate specific antigen; \*, factors linked to all causes of death; §, Country attributes of Year 2000; Education attainment defined as percent of residents with less than high-school graduate in the county; Person in poverty defined as percent of residents with income below 200% of poverty in the county.