SYSTEMATIC REVIEW AND META-ANALYSIS



Effect of low-dose aspirin on health outcomes: An umbrella review of systematic reviews and meta-analyses

Nicola Veronese^{1,2} | Jacopo Demurtas^{3,4} | Trevor Thompson⁵ | Marco Solmi⁶ | Gabriella Pesolillo⁷ | Stefano Celotto⁸ | Tommaso Barnini⁹ | Brendon Stubbs^{10,11} | Stefania Maggi¹ | Alberto Pilotto² | Graziano Onder¹² | Evropi Theodoratou^{13,14} | Alberto Vaona¹⁵ | Joseph Firth^{16,17} | Lee Smith¹⁸ | Ai Koyanagi^{19,20} | John P.A. Ioannidis^{21,22,23,24} | Ioanna Tzoulaki^{13,25}

¹¹Department of Psychological Medicine, Institute of Psychiatry, Psychology and Neuroscience, King's College London, De Crespigny Park, London, UK

- ¹²Fondazione Policlinico Universitario A. Gemelli, IRCCS and Università Cattolica del Sacro Cuore, Rome, Italy
- ¹³Centre for Global Health Research, Usher Institute of Population Health Sciences and Informatics, University of Edinburgh, Edinburgh, UK
- ¹⁴Edinburgh Cancer Research Centre, Institute of Genetics and Molecular Medicine, University of Edinburgh, Edinburgh, UK
- ¹⁵Primary Care Department, Azienda ULSS20 Verona, Verona, Italy
- ¹⁶NICM Health Research Institute, University of Western Sydney, Penrith, Australia
- ¹⁷Division of Psychology and Mental Health, University of Manchester, Manchester, UK

¹⁸The Cambridge Centre for Sport and Exercise Sciences, Anglia Ruskin University, Cambridge, UK

- ¹⁹Instituto de Salud Carlos III, Centro de Investigación Biomédica en Red de Salud Mental, CIBERSAM, Madrid, Spain
- ²⁰Research and Development Unit, Parc Sanitari Sant Joan de Déu, Universitat de Barcelona, Barcelona, Spain
- ²¹Stanford Prevention Research Center, Department of Medicine, Stanford University Medical School, Stanford, California, USA
- ²²Department of Health Research and Policy, Stanford University School of Medicine, Stanford, California, USA
- ²³Department of Statistics, Stanford University School of Humanities and Sciences, Stanford, California, USA
- ²⁴Department of Biomedical Data Science, Stanford University School of Medicine, Stanford, California, USA

²⁵MRC-PHE Centre for Environment, School of Public Health, Imperial College London, UK

Correspondence

Lee Smith, The Cambridge Centre for Sport and Exercise Sciences, Anglia Ruskin University, Cambridge, UK. Email: lee.smith@anglia.ac.uk **Aims:** This study aimed to use an umbrella review methodology to capture the range of outcomes that were associated with low-dose aspirin and to systematically assess the credibility of this evidence.

¹National Research Council, Neuroscience Institute, Aging Branch, Padua, Italy

²Geriatrics Unit, Department of Geriatric Care, Ortho Geriatrics and Rehabilitation, Frailty Area, E.O. Galliera Hospital, Genoa, Italy

³Primary Care Department, Azienda USL Toscana Sud Est, Grosseto, Italy

⁴Clinical and Experimental Medicine PhD Program, University of Modena and Reggio Emilia, Modena, Italy

⁵Faculty of Health, Social Care and Education, Anglia Ruskin University, Chelmsford, UK

⁶Department of Neuroscience, University of Padova, Padua, Italy

⁷Primary Care Service, ASL 02, Chieti, Italy

⁸Primary Care Department, Alto Friuli–Collinare–Medio Friuli, Udine, Italy

⁹Primary Care Service, AUSL Toscana Centro, Florence, Italy

¹⁰Physiotherapy Department, South London and Maudsley NHS Foundation Trust, Denmark Hill, London, UK

Nicola Veronese, Jacopo Demurtas and Trevor Thompson equally contributed to the work.

Methods: Aspirin is associated with several health outcomes, but the overall benefit/risk balance related to aspirin use is unclear. We searched three major databases up to 15 August 2019 for meta-analyses of observational studies and randomized controlled trials (RCTs) including low-dose aspirin compared to placebo or other treatments. Based on random-effects summary effect sizes, 95% prediction intervals, heterogeneity, small-study effects and excess significance, significant meta-analyses of observational studies were classified from convincing (class I) to weak (class IV). For meta-analyses of RCTs, outcomes with random effects *P*-value < .005 and a moderate/high GRADE assessment, were classified as strong evidence. From 6802 hits, 67 meta-analyses (156 outcomes) were eligible.

Results: Observational data showed highly suggestive evidence for aspirin use and increased risk of upper gastrointestinal bleeding (RR = 2.28, 95% CI: 1.97–2.64). In RCTs of low-dose aspirin, we observed strong evidence for lower risk of CVD in people without CVD (RR = 0.83; 95% CI: 0.79–0.87) and in general population (RR = 0.83; 95% CI: 0.79–0.89), higher risk of major gastrointestinal (RR = 1.47; 95% CI: 1.26–1.72) and intracranial bleeding (RR = 1.34; 95% CI: 1.18–1.53), and of major bleedings in people without CVD (RR = 1.62; 95% CI: 1.26–2.08).

Conclusion: Compared to other active medications, low-dose aspirin had strong evidence for lower risk of bleeding, but also lower comparative efficacy. Low-dose aspirin significantly lowers CVD risk and increases risk of bleeding. Evidence for multiple other health outcomes is limited.

KEYWORDS

aspirin, cancer, cardiovascular disease, meta-analysis, umbrella review

1 | INTRODUCTION

Low-dose aspirin, defined as less than 325 mg daily, is widely used worldwide, particularly for the prevention of cardiovascular disease (CVD).¹ The United States Preventive Services Task Force (USPSTF) recommends aspirin for primary CVD prevention in adults with a 10-year risk of heart attack or stroke exceeding 10% in individuals who are not at increased risk of bleeding and after individualized informed decisions,^{2,3} whilst other societies recommend low-dose aspirin use for secondary CVD prevention only.⁴ In the US, over 30% of adults take aspirin for CVD prevention, but use in recent years is probably decreasing.⁴

Aspirin irreversibly inhibits cyclo-oxygenase 1 (COX-1) which leads to inhibition of platelet thromboxane A2 and thrombus formation in arteries.⁵ Beyond CVD, low-dose aspirin use has been linked to lower risk of cancers, overall mortality and other chronic conditions.⁴ The veracity of these claimed non-cardiovascular effects is unclear. European and American guidelines currently do not support aspirin for cancer prevention,^{4,6} but the issue is unsettled.^{7,8} However, despite possible benefits, low-dose aspirin is also associated with an increased risk of bleeding⁹ and any clinical benefits need to be balanced against adverse effects. At the same time, prescription

of low-dose aspirin in many primary care settings is suboptimal^{10,11} and many patients who would probably benefit remain untreated.¹² Improving appropriate use of aspirin is therefore essential.¹¹ The body of research on low dose aspirin is constantly increasing with new studies and meta-analyses thereof being published during the last few years. At the same time, the breadth of outcomes examined has expanded to cover a wide range of outcomes not limited to CVD. We used the umbrella review methodology in order to capture the breadth of outcomes reported and assess the totality of evidence of low dose aspirin on a number of outcomes.¹³ In this sense, umbrella reviews (i.e. reviews of previously published systematic reviews/meta-analyses consisting in the replication of the meta-analyses following a uniform statistical approach for all factors to allow their comparison) have been created for overcoming the inherent limitations of meta-analyses.¹³

Here we aimed to capture the breadth of outcomes that have been associated with low-dose aspirin intake and systematically assess the quality, strength and credibility of the associations. We used the umbrella review methodology to combine evidence from a wide range of outcomes and populations and we present results separately for observational studies and randomized controlled trials (RCTs).

2 | METHODS

2.1 | Data sources and searches

We conducted an umbrella review,¹⁴ searching the MEDLINE, Scopus, Embase databases from inception until 15 August 2019 with: "(Meta-Analysis[ptyp] OR metaanaly*[tiab] OR meta-analy*[tiab] OR Systematic review [ptyp] OR "systematic review" [tiab]) AND (aspirin [tiab])". In addition, we hand-searched the reference lists of eligible articles.

2.2 | Study selection

Eligible articles were systematic reviews with meta-analyses of observational/intervention studies, which investigated low-dose aspirin in relation to any clinical outcome. Four authors (J.D., G.P., T.B., S.C.) independently performed title and abstract screening in couples. Disagreements were resolved through consensus with another independent author (N.V.). Full texts of all potentially eligible articles were then retrieved by the same four authors and any disagreement was resolved with another independent author (M.S.).

We included meta-analyses that investigated effects of low-dose aspirin, defined as at least 75 mg and less than 325 mg daily^{15,16} or use of the regular 325 mg aspirin dose three or more times a week (but not daily) for at least six months.¹⁷ Both meta-analyses of observational studies that investigated the association of low-dose aspirin with any clinical outcome and meta-analyses of RCTs were considered. Meta-analyses were included only if they reported study-specific information (i.e. effect size, 95% confidence intervals [Cls], sample size) or if those metrics could be inferred from the data presented. The RCT meta-analyses were divided into meta-analysis of placebo/no active control and active control groups (e.g. heparins, vitamin K antagonists). Studies were excluded if aspirin was accompanied by additional co-administered medications (e.g. clopidogrel, heparins).

2.3 | Data extraction

Four independent investigators (J.D., G.P., T.B., S.C.) extracted the following information for each meta-analysis, independently, in pairs: first author name; publication year; number of studies; study population; type of effect size; study design; number of participants with (cases) and without (controls) events for each study. We also extracted the study-specific estimated relative risk for health outcome (risk ratio, RR; odds ratio OR; hazard ratio, HR; mean difference, MD; standardized mean difference, SMD) and 95% CIs. We finally extracted the data for the Assessment of Multiple Systematic Reviews (AMSTAR)-2 tool.¹⁸

When more than one meta-analysis on the same research question using the same study design (observational or RCTs) was identified, the one with the largest number of participants was selected.

2.4 | Data synthesis and analysis

For each meta-analysis, we estimated the summary effect size and its 95% CI by using the random-effects Hartung-Knapp-Sidik-Jonkman (HK) estimator.¹⁹ This estimator consistently results in more adequate error rates than the DerSirmonian-Leird method, especially when the number of studies is small.¹⁹ We also estimated the prediction interval (PIs) and its 95% CI, which further accounts for between-study effects and estimates the certainty of the association if a new study addresses that same association.^{20–22} In order to estimate whether any large studies were available, for the largest study of each meta-analysis, we calculated the standard error (SE) of the effect size. If the SE was less than 0.10, then the 95% CI would be lower than 0.20. Between-study inconsistency was estimated with the l^2 metric, with values \geq 50% indicative of high heterogeneity.²³

We calculated the evidence of small-study effects (i.e. whether small studies inflated effect sizes) using the regression asymmetry test²⁴ with a *P*-value < $.10.^{25}$

Finally, we applied the excess of significance test.²⁶ Because of the limited statistical power of this test, a lenient significance threshold (P < .10) was adopted.²⁷ We considered the effect size of the largest dataset and based on this we estimated the power of each constituent study with an algorithm using a non-central *t* distribution. Excess significance for each meta-analysis was considered whenever P < .10.

All statistical analyses were conducted in Stata, version 14.0 (StataCorp), and R, version 3.3.0 (R Foundation for Statistical Computing).

2.5 | Grading the evidence

For observational studies, using the criteria mentioned above, significant associations (i.e. P < .05) were categorized into strong, highly suggestive, suggestive, or weak evidence, following a grading scheme that has already been applied in various fields,²⁸⁻³⁵ as reported in Table 1. We assessed the methodological quality of the included meta-analyses of observational studies using AMSTAR-2,^{18,36} which ranks the quality of a meta-analysis from critically low to high according to 16 predefined items. For each association in the convincing or highly suggestive categories, we reassessed the evidence keeping only prospective observational studies in an attempt to address reverse causality and applying the credibility ceiling at 10%. However, application of the 10% credibility ceiling did not affect any class I associations. Finally, for each association in the convincing category, we reassessed the evidence taking in account the AMSTAR-2 evaluation.

Evidence from meta-analyses of RCTs was assessed in terms of the significance of the summary effect, using a *P*-value of <.005 as the threshold for statistical significance, as recently proposed.^{37,38} We used stringent *P*-values when evaluating the findings of RCTs in order to decrease the possibility of "false-positives" (i.e. to claim that an effect is present when there is none in reality).³⁹ When the *P*-value

TABLE 1 Credibility assessment criteria for meta-analyses of observational studies

Evidence classification	Criteria
Convincing (class I)	Associations with <i>P</i> < .000001; >1,000 cases (or >20 000 participants for continuous outcomes) having the event of interest; the largest component study reporting a nominal statistically significant result (<i>P</i> < .05); a 95% PI that excluded the null; no large heterogeneity (<i>I</i> ² < 50%); no evidence of small-study effect (<i>P</i> > .10); no excess significance bias (<i>P</i> > .10); Studies survived the 10% credibility ceiling test.
Highly suggestive (class II)	Associations with <i>P</i> < .000001; >1000 cases (or >20 000 participants for continuous outcomes) having the event of interest; the largest component study reporting a statistically significant result (<i>P</i> < .05).
Suggestive (class III)	Associations with P < .001; >1000 cases (or >20 000 participants for continuous outcomes) having the event of interest
Weak (class IV)	Remaining statistically significant associations with <i>P</i> < .05.

Abbreviations: PI = prediction interval; RCT = randomized controlled trial.

for the random effect was <.005, we evaluated the evidence using the GRADE (Grading of Recommendations, Assessment, Development and Evaluation) assessment.⁴⁰ Outcomes having a *P*-value of <.005 and a moderate/high GRADE assessment were classified as strong evidence. We also considered 95% PIs (excluding the null or not), the presence of large heterogeneity ($l^2 > 50\%$), small study effects (P > .10), and excess significance (P > .10) as possible indicators of heterogeneity and bias in the available evidence.

3 | RESULTS

3.1 | Literature review

Overall, we identified 6802 papers (Figure 1); 578 publications were selected as potentially eligible and 67 meta-analyses (corresponding to 156 different outcomes) were finally included in this study (references in **Supplementary Material**).

3.2 | Meta-analyses of observational studies

The median number of studies of meta-analyses including observational studies for each outcome was three (range 2–32), the median number of participants was 11 894 (range 520 to 1059 682), and the median number of cases was 1114 (range 10 to 144 373) (Table S1). The majority of the meta-analyses included studies on general populations, followed by patients with cancer or diabetes. Overall, 11 out of the 41 outcomes reported nominally significant summary results (P < .05), but only two associations survived the application of the more stringent *P*-value ($P < 10^{-6}$), i.e. higher risk of upper gastro-intestinal bleeding in general population and in people undergoing coronary artery bypass graft.

The study with the largest number of participants had an SE of less than 0.10 in 12 outcomes and a more conservative effect compared to the random-effects model in 11 of these 12 outcomes. Heterogeneity among studies was modest and 24 outcomes presented low heterogeneity ($l^2 < 50\%$). Three associations presented 95% PIs excluding the null value. Evidence for excess statistical significance was present in five out of 41 outcomes and small-study effects were also seen in five out of 41 of the outcomes. Publication bias was present in six of the 41 outcomes.

Based on the above criteria, no outcome presented convincing evidence, only one outcome presented highly suggestive evidence (class II: higher incidence of upper gastrointestinal bleeding in the general population; RR = 2.28, 95% CI: 1.97–2.64), two outcomes presented suggestive evidence (class III: lower incidence of prostate cancer and cancer specific death in the general population) and eight outcomes weak evidence. Using the AMSTAR-2, all the meta-analyses included were evaluated as having a critically low rating mainly because the risk of bias was not accurately assessed and the sources of funding for the included studies were not reported (Table S2).

In a sensitivity analysis, we included only prospective cohort studies in each meta-analysis (Table S3). Two outcomes presented suggestive evidence (lower cancer-specific death in people affected by colorectal cancer and higher risk of upper gastrointestinal bleeding in the general population) and four were classified as weak evidence. Both outcomes with suggestive evidence had a low AMSTAR-2 score.

3.3 | Meta-analyses of RCTs (vs placebo/no treatment)

The median number of RCTs meta-analyses using placebo/no treatment for each outcome was five (range 2–23), the number of participants was, in median, 12 184 (71 to 1126 384), and the median number of cases was 377 (4 to 7087) (Table S4).

Overall, 76 outcomes were included. Of these, 25 reported significant results (P < .05), but only five survived the application of a more stringent *P*-value (P < .005): lower risk of serious CVD in people without CVD (RR = 0.83; 95% CI: 0.79-0.87) and of CVD in the general population (RR = 0.83; 95% CI: 0.78-0.89), higher risk of major gastrointestinal (RR = 1.47; 95% CI: 1.26-1.72) and intracranial bleeding (RR = 1.34; 95% CI: 1.18-1.53) in the general population, and major bleeding in people without CVD at baseline (RR = 1.62; 95% CI: 1.26-2.08). Using the GRADE assessment, as reported in Table 2, we observed strong evidence for all the



outcomes in primary prevention, having a P-value of <.005, except for major bleeding in primary prevention (presence of publication bias).

The largest study, in terms of sample size, had an SE of less than 0.10 in only 15 outcomes. Heterogeneity among studies was low in 60 out of 76 studies, with 40 reporting an l^2 of 0%. Nine outcomes presented 95% PIs excluding the null value. Finally, evidence for excess statistical significance was present in five out of 74 outcomes and small-study effects were present in six out of 76 outcomes.

As reported in Table S5, only two out of 76 rated "high", six rated "low" according to the AMSTAR-2 criteria, whilst the other metaanalyses were rated as "critically low".

Meta-analyses of RCTs (vs active controls) 3.4

As reported in Table S6, the median number of studies of metaanalyses including intervention studies using active controls for each outcome was three (range 2-15), the median number of participants was 3607 (193 to 33 435), and the median number of cases was 121 (4 to 1364).

Only animals (n=1)

1469

In these meta-analyses, 16 (41%) out of the 39 outcomes reported nominally significant summary results (P < .05), and, of them, two treatment effects had a summary effect with a P-value of <.005 (Table 3). Using the GRADE assessment, we observed strong evidence for associations between aspirin use and higher risk of subarachnoid bleeding in cerebrovascular conditions (compared to cilostazol) and higher incidence of pulmonary embolism in cancer under chemotherapy (compared to heparins). Three meta-analyses were rated as low quality according to the criteria suggested by the AMSTAR-2, and the others were rated as critically low (Table S7).

Heterogeneity among studies was low, with the majority of outcomes (32 out of 39) (82%) having an $I^2 < 50\%$. However, nine out of 39 outcomes (23%) presented summary effects with 95% PIs excluding the null value. Only one study showed evidence of excess significance, whilst no outcomes showed evidence for statistically significant small-study effects. No meta-analysis showed evidence of publication bias.

uo
/enti
ter
i oc
J)oc
lacel
us p
pirin
e as
, dos
_ ≷
ring
mpa
s co
trial
lled
ntro
с р
mize
opu
or ra
ce fi
iden
е
QP
G
7
BLE
TAI

Certainty asse	ssment						Summary of fin	dings				
							Study event rat	tes (%)		Anticipated abs	solute effects	
No. of participants (studies)	Risk of bias	Inconsistency	Indirectness	Imprecision	Publication bias	Overall certainty of evidence	With placebo/no intervention	With aspirin	Relative effect (95% CI)	Risk with placebo/no intervention	Risk difference with aspirin	Median follow-up in years
CVD events ^a ii	n general pop	oulation										
113 204 (23 RCTs)	Not serious	Not serious	Not serious	Not serious	None	⊕⊕⊕⊕ high	3836/56646 (6.8%)	3243/56558 (5.7%)	RR 0.831 (0.779 to 0.885)	68 per 1000	11 fewer per 1,000 (15 fewer to 8 fewer)	5z
Major bleedin	gs in people v	vithout CVD										
98 311 (10 RCTs)	Not serious	Not serious	Not serious	Not serious	None	⊕⊕⊕⊕ high	285/49157 (0.6%)	456/49154 (0.9%)	RR 1.618 (1.258 to 2.080)	6 per 1000	4 more per 1,000 (2 more to 6 more)	4.5
Intracranial ble	seding (primaı	ry prevention)										
160 404 (11 RCTs)	Not serious	Not serious	Not serious	Not serious	None	⊕⊕⊕⊕ high	256/79419 (0.3%)	349/80985 (0.4%)	RR 1.341 (1.178 to 1.528)	3 per 1000	1 more per 1,000 (1 more to 2 more)	Ŷ
Major gastroin	itestinal bleec	ding (primary preve	ention)									
140 792 (10 RCTs)	Not serious	Not serious	Not serious	Not serious	None	⊕⊕⊕⊕ high	308/70456 (0.4%)	593/70336 (0.8%)	RR 1.557 (1.355 to 1.790)	4 per 1000	2 more per 1,000 (2 more to 3 more)	Ŷ
Major bleedin	g (primary pre	vention)										
157 248 (10 RCTs)	Not serious	Not serious	Not serious	Not serious	Publication bias strongly suspected a	⊕⊕⊕⊖ moderate	901/77835 (1.2%)	1301/79413 (1.6%)	RR 1.474 (1.263 to 1.721)	12 per 1000	5 more per 1,000 (3 more to 8 more)	Ŷ
Serious CVD ii	n people with	out CVD at the bi	aseline									
107 074 (11 RCTs)	Not serious	Not serious	Not serious	Not serious	None	⊕⊕⊕⊕ high	2505/52587 (4.8%)	2392/54487 (4.4%)	RR 0.831 (0.791 to 0.872)	48 per 1000	8 fewer per 1,000 (10 fewer to 6 fewer)	Ŋ
Abbreviations: C ^a CVD events inc	ABG: coronal Iuded coronal	ry artery disease t ry artery disease e	oypass graft; Cl: c events, coronary ¿	confidence inter artery disease m	val; CVD: cardiov nortality, stroke, s	ascular disease; ¹ erious vascular e	MD: mean differe vents, serious va	ence; OR: odds ra scular events mor	tio; RCT: rand rtality.	omized controlle	d trial; RR: risk rat	.o

BRITISH PHARMACOLOGICAL SOCIETY VERONESE ET AL.

BJCP

controls
/s. active
aspirin v
ow dose
paring lo
rials com
trolled ti
iized con
r random
dence fo
GRADE evic
TABLE 3

		Median follow-up in years		1.7		1.2
	ed absolute	Risk difference with aspirin		15 more per 1000 (4 more to 38 more)		19 more per 1000 (3 more to 81 more)
	Anticipate effects	Risk with active control		7 per 1000		3 per 1000
		Relative effect (95% CI)		RR 3.121 (2.885-3.376)		RR 8.488 (1.653 to 43.59)
of findings	it rates (%)	With aspirin		32/1369 (2.3%)		7/396 (1.8%)
Summary c	Study even	With active control		10/1371 (0.7%)		1/385 (0.3%)
		Overall certainty of evidence		⊕⊕⊕⊕ high		⊕⊕⊕⊕ high
		Publication bias		Publication bias strongly suspected strong association ^a		Publication bias strongly suspected very strong association ^a
		Imprecision		Not serious	vs LMWH	Not serious
		Indirectness	in vs cilostazol	Not serious	therapy: aspirin v	Not serious
sment	Risk of Inconsistency		oke or TIA: aspiri	Not serious	cer under chemo	Not serious
			bleeding in st	Not serious	bolism in can	Not serious
Certainty asse		No. of participants (studies) follow-up	Subarachnoid	2740 (2 RCTs)	Pulmonary em	781 (2 RCTs)

Abbreviations: CI: confidence interval; DAPT: dual antiplatelet therapy; LMWH: low molecular weight heparins; OR: odds ratio; RCT: randomized controlled trial; RR: risk ratio; TIA: transient ischaemic attack; VKA: vitamin K antagonists. ^aFew studies

BJCP

3.5 | Comparison of findings from observational studies and clinical trials

As reported in Table S8, ten outcomes were examined by both meta-analyses of observational studies and meta-analyses of RCTs using placebo/no intervention as controls. The direction of the association/effect was concordant for seven of the 10 outcomes. In one case (stroke in patients with type 2 diabetes), the 95% Cls of RCTs excluded the null from the estimated effect size, but they were in the opposite direction, while in the other nine topics, the 95% Cls overlapped.

4 | DISCUSSION

With this work, we provide a comprehensive overview of the associations between low-dose aspirin and a wide range of health outcomes.

In a large epidemiological study, it is reported that about one third of American people take low-dose aspirin for primary and secondary prevention, even if the prevalence of people taking aspirin has been declining in the last years.⁴ In a more recent study, a consistent proportion of American people were taking aspirin without a physician's recommendation, corresponding to about 6.6 million adults. Nearly half of people at least 70 years of age in the survey, 44.6%, were on aspirin for primary CVD prevention.⁴¹ Therefore, systematic knowledge of the efficacy and risk of low-dose aspirin use is of great clinical importance.

The topic of the use of low-dose aspirin in primary prevention is of great interest. Our umbrella review found that for primary prevention, use of low-dose aspirin was associated with 17% lower CVD incidence (including serious events, i.e. non-fatal myocardial infarction, non-fatal stroke or vascular death). Moreover, low-dose aspirin was associated with 34% higher risk of bleeding in primary prevention (major and intracranial). These risks and benefits need to be weighted in formal decision analysis to guide aspirin use in primary prevention. Taken together, these findings suggest that in the balance between prevention and risk one should consider the risk of bleedings. Three recent RCTs were published last year.42-44 The ARRIVE (Aspirin to Reduce the Risk of Initial Vascular Events) trial enrolled participants at high risk for CVD events without diabetes.⁴² The results of this trial suggested no significant effect of low-dose aspirin on the reduction of CVD incidence, but a significantly increased risk of gastrointestinal bleedings.⁴² Another recent randomized controlled trial, the ASCEND study (A Study of Cardiovascular Events in Diabetes), documented a significant benefit of aspirin among people with diabetes in preventing CVD events (of about 12%), but at the cost of a significant increase in the rate of major bleeding events,⁴³ in a manner similar to our findings. In a third randomized controlled trial, data from ASPREE (Aspirin in Reducing Events in the Elderly) showed no benefit of aspirin on CVD events in older people and also demonstrated a significantly increased risk of major bleeding in those taking aspirin in this older age group.⁴⁴ Taken together these findings and those from the present umbrella review suggest that the benefits and risks of low-dose aspirin for the primary prevention of CVD events in the modern era of preventive management in middle-aged people (i.e., involving statins, anti-hypertension medications, smoking cessation, obesity management and other similar interventions) are closely balanced, calling into question the use of aspirin in those without a prior cardiovascular disease event.

A topic of great clinical relevance in clinical settings is low-dose aspirin in primary prevention specifically for those at high CVD risk, such as people with diabetes. Observational and intervention studies in this review show little evidence that low-dose aspirin prevents overall and specific CVD events in diabetes (Table S1), indicating that the widespread use of this medication may not be justified in this population.⁴⁵ As shown in Table S9, these findings can be applied in other conditions at higher risk of CVD such as women with antiphospholipid antibodies. Unfortunately, we were unable to take the same conditions into consideration for secondary prevention due to the limited data available.

The prevention of cancer is another topic of interest.⁴⁶ In our umbrella review, we identified several meta-analyses, including observational studies, which investigated the effect of aspirin on risk of cancer/cancer progression/cancer-specific death. Low-dose aspirin was associated with a reduced risk of prostate cancer with suggestive evidence in observational studies, whilst the evidence regarding mortality in colorectal cancer patients was weak. The USPSTF guidelines suggested that low-dose aspirin is efficacious in reducing the incidence of colorectal cancer, even if this benefit is not apparent until 10 years after aspirin therapy is started.⁴⁶ However, this work was not eligible for our umbrella review, since, among four eligible studies, two used doses of aspirin ≥325 mg and another one used a vitamin supplementation together with aspirin.⁴⁶ Other guidelines specifically from cancer-related societies suggest that low-dose aspirin should not be used in the general population, but only in some specific conditions, such as Lynch's syndrome.⁴⁷ In meta-analyses of the RCTs (vs placebo/no intervention), low-dose aspirin was associated with a nominally statistically significant reduction in cancer-specific death in people affected by cancer at baseline or in the general population, but this evidence remains poor.

As reported in Table S9, we found, in observational studies, highly suggestive evidence or at least suggestive evidence that low-dose aspirin is associated with a higher risk of upper and overall gastrointestinal bleeding in the general population, i.e. in the primary prevention setting. The risk of these events is more than doubled compared to non-users,⁴⁸ an observation confirmed in the RCTs *vs* placebo/no intervention. However, the risk of bleeding (overall, major, gastrointestinal) was lower in people taking low-dose aspirin compared to several other medications, including clopidogrel.

Our study has some shortcomings that we should acknowledge. First, we used evidence assessment criteria, which can be biased, as they are based on already established tools for observation and interventional studies.^{33,49} Moreover, since the meta-analyses included studies with significant differences in design, population and other basic characteristics, large heterogeneity may be worrisome. We consequently used an l^2 of 50% as one of the criteria for having

convincing outcomes. However, l^2 estimates can also carry substantial uncertainty⁵⁰ and often clinical heterogeneity might be of importance, even in the absence of statistical heterogeneity. It is known that meta-analyses have important limitations⁵¹ and their results may also depend on choices made about what estimates to select from each study and how to report them in the meta-analysis (e.g. in our umbrella review several meta-analyses did not report information regarding aspirin dosage).⁵² Applying the criteria suggested by the AMSTAR-2 for evaluating the quality of meta-analyses, we observed the presence of low/critically low rating. This evidence is mainly due to missing information in items 2 (protocol published before the metaanalysis), 7 (list of excluded studies), or 11 (no appropriate metaanalytic approach, particularly the absence of investigation in case of high heterogeneity). Furthermore, low-dose aspirin covers a substantial range of dosing regimens and these may not have exactly the same efficacy and harms, but this was beyond the discerning ability of our study design.⁵³ The umbrella review was limited to outcomes studied in the respective meta-analyses and does not provide in-depth data on disease severity, dose-response effects, or specific subgroups such as by sex or age. At the same time, it is also possible that some studies were included in two or more outcomes (e.g. in intracranial and major bleedings) in major bleeding that is a cumulative outcome: however, we believe that this includes a limited portion of the studies included, over 156 outcomes. We decided to include the data from observational studies that are, per se, biased in their nature. As a number of outcomes (34/41) were only examined in observational settings, we included in this review data from observational studies acknowledging their limitations. However, a large majority of the outcomes included in the observational studies (34/41) were not included in those of RCTs, highlighting the importance of their inclusion. Finally, this umbrella review could not explore fully the possibility of risk stratification for clinical use, especially taking into account potential risk factors for adverse outcomes.

In conclusion, in this umbrella review including 67 independent meta-analyses and 156 outcomes, we found that low-dose aspirin decreased the risk of CVD events in the general population (when compared to placebo/no intervention) with strong evidence according to GRADE criteria, whilst the data for individual CVD outcomes are limited. Moreover, when limiting to only observational studies, moderate evidence for associations between aspirin intake and lower risk of specific cancers in the general population was observed. However, this finding should be interpreted with caution given the inherent bias of observational study designs. The risk of bleeding (particularly gastrointestinal and intracranial) is, however, also strong and substantial, suggesting that physicians should accurately consider the risks and benefits of prescribing aspirin. Despite many dozens of other clinical outcomes having been assessed, evidence for them remains weak and therefore should not be a major determinant in decision-making regarding use of low-dose aspirin.

COMPETING INTERESTS

Dr Demurtas received an honorary consultancy from Bayer; Dr Stubbs is supported by Health Education England and the National



Institute for Health Research HEE/NIHR ICA Programme Clinical Lectureship (ICA-CL-2017-03-001), and is part supported by the National Institute for Health Research (NIHR) Collaboration for Leadership in Applied Health Research and Care South London at King's College Hospital NHS Foundation Trust. The views expressed in this publication are those of the author(s) and not necessarily those of the NHS, the National Institute for Health Research or the Department of Health and Social Care; Dr Firth is supported by a Blackmores Institute Fellowship; Dr Koyanagi's work was supported by the Miguel Servet contract financed by the CP13/00150 and PI15/00862 projects, integrated into the National R+D+I and funded by the ISCIII - General Branch Evaluation and Promotion of Health Research - and the European Regional Development Fund (ERDF-FEDER); Dr Theodoratou is supported by a Cancer Research UK (CRUK) Career Development Fellowship (C31250/A22804). The other authors don't have any financial arrangements, organizational affiliations or other relationships that might give rise to any conflict of interest regarding the subject matter of the manuscript submitted.

CONTRIBUTORS

N.V. and I.T. conceived the study. N.V., M.S. and A.K. did the statistical analyses. N.V., M.S., A.K., A.P. and B.S. wrote the first draft of the paper; B.S., N.V., J.D., G.P., S.C. and T.B. conducted the searches, carried out screening and data extraction, and contributed in writing the paper; B.S. and S.M. contributed to data interpretation; T.T., B.S., E.T., G.O., A.V., J.F., L.S. and J.P.A.I. contributed towards the intellectual conception of the review, and revised the manuscript; I.T. supervised the study and contributed in writing the paper and interpreting the findings.

ORCID

Lee Smith D https://orcid.org/0000-0003-0890-7566

REFERENCES

- Kim C, Beckles GL. Cardiovascular disease risk reduction in the behavioral risk factor surveillance system. Am J Prev Med. 2004;27(1): 1-7.
- Ittaman SV, VanWormer JJ, Rezkalla SH. The role of aspirin in the prevention of cardiovascular disease. *Clin Med Res.* 2014;12(3-4): 147-154.
- Arnett DK, Blumenthal RS, Albert MA, et al. ACC/AHA Guideline on the primary prevention of cardiovascular disease. J Am Coll Cardiol. 2019;2019:26029.
- Stuntz M, Bernstein B. Recent trends in the prevalence of low-dose aspirin use for primary and secondary prevention of cardiovascular disease in the United States, 2012–2015. *Prev Med Rep.* 2017;5: 183-186.
- Bartolucci AA, Tendera M, Howard G. Meta-analysis of multiple primary prevention trials of cardiovascular events using aspirin. *Am J Cardiol*. 2011;107(12):1796-1801.
- 6. Piepoli MF, Hoes AW, Agewall S, et al. 2016 European guidelines on cardiovascular disease prevention in clinical practice: the Sixth Joint Task Force of the European Society of Cardiology and other societies on cardiovascular disease prevention in clinical practice (constituted by representatives of 10 societies and by invited experts) developed with the special contribution of the European Association for

1474 BJCP BRITISH PHARMACOLOGICA SOCIETY

Cardiovascular Prevention & Rehabilitation (EACPR). *Eur Heart J.* 2016;37(29):2315-2381.

- Patrignani P, Patrono C. Aspirin and cancer. J Am Coll Cardiol. 2016; 68(9):967-976.
- Sutcliffe P, Connock M, Gurung T, et al. Aspirin for prophylactic use in the primary prevention of cardiovascular disease and cancer: a systematic review and overview of reviews. *Health Technol Assess (Winch Eng)*. 2013;17(43):1-253.
- 9. Cryer B. Gastrointestinal safety of low-dose aspirin. *Am J Manag Care*. 2002;8:S701-S708.
- McCallum AK, Whincup PH, Morris RW, Thomson A, Walker M, Ebrahim S. Aspirin use in middle-aged men with cardiovascular disease: are opportunities being missed? *Br J Gen Pract.* 1997;47(420): 417-421.
- Short D, Frischer M, Bashford J, Ashcroft D. Why are eligible patients not prescribed aspirin in primary care? A qualitative study indicating measures for improvement. *BMC Fam Pract*. 2003;4:9.
- Sweeney K. How can evidence-based medicine help patients in general practice? Fam Pract. 1996;13(6):489-490.
- 13. Fusar-Poli P, Radua J. Ten simple rules for conducting umbrella reviews. *Evid Based Ment Health*. 2018;21:95-100.
- Ioannidis JP. Integration of evidence from multiple meta-analyses: a primer on umbrella reviews, treatment networks and multiple treatments meta-analyses. CMAJ. 2009;181(8):488-493.
- Lloyd J, Bochner F. Aspirin: how low is low dose? Australian Prescriber. 1996;19:79-81.
- 16. Goodman LS. Goodman and Gilman's the pharmacological basis of therapeutics. New York: McGraw-Hill; 1996.
- 17. Sostres C, Lanas A. Epidemiology of low dose aspirin damage in the lower gastrointestinal tract. *Curr Pharm Des.* 2015;21:5094-5100.
- Shea BJ, Reeves BC, Wells G, et al. AMSTAR 2: a critical appraisal tool for systematic reviews that include randomised or non-randomised studies of healthcare interventions, or both. *BMJ*. 2017;358:j4008.
- IntHout J, Ioannidis JP, Borm GF. The Hartung-Knapp-Sidik-Jonkman method for random effects meta-analysis is straightforward and considerably outperforms the standard DerSimonian-Laird method. BMC Med Res Methodol. 2014;14:25.
- IntHout J, Ioannidis JPA, Rovers MM, Goeman JJ. Plea for routinely presenting prediction intervals in meta-analysis. BMJ Open. 2016;6(7): e010247.
- Higgins JP, Thompson SG, Spiegelhalter DJ. A re-evaluation of random-effects meta-analysis. J R Stat Soc Ser A Stat Soc. 2009;172: 137-159.
- 22. Serghiou S, Goodman SN. Random-effects meta-analysis: summarizing evidence with caveats. JAMA. 2018;321(3):301-302.
- Higgins JP, Thompson SG. Quantifying heterogeneity in a meta-analysis. Stat Med. 2002;21(11):1539-1558.
- Egger M, Davey Smith G, Schneider M, Minder C. Bias in metaanalysis detected by a simple, graphical test. BMJ. 1997;315(7109): 629-634.
- Carvalho AF, Kohler CA, Brunoni AR, et al. Bias in peripheral depression biomarkers. Psychother Psychosom. 2016;85(2):81-90.
- Ioannidis JP, Trikalinos TA. An exploratory test for an excess of significant findings. *Clin Trials*. 2007;4(3):245-253.
- Ioannidis JP. Clarifications on the application and interpretation of the test for excess significance and its extensions. J Math Psychol. 2013; 57:184-187.
- Aromataris E, Fernandez R, Godfrey CM, Holly C, Khalil H, Tungpunkom P. Summarizing systematic reviews: methodological development, conduct and reporting of an umbrella review approach. *Int J Evid Based Healthc*. 2015;13(3):132-140.
- Belbasis L, Savvidou MD, Kanu C, Evangelou E, Tzoulaki I. Birth weight in relation to health and disease in later life: an umbrella review of systematic reviews and meta-analyses. *BMC Med.* 2016; 14:147.

- Bellou V, Belbasis L, Tzoulaki I, Evangelou E, Ioannidis JP. Environmental risk factors and Parkinson's disease: an umbrella review of meta-analyses. *Parkinsonism Relat Disord*. 2016;23:1-9.
- Dinu M, Pagliai G, Casini A, Sofi F. Mediterranean diet and multiple health outcomes: an umbrella review of meta-analyses of observational studies and randomized trials. *Nutr Metab Cardiovasc Dis.* 2017; 27(1):e21.
- Kyrgiou M, Kalliala I, Markozannes G, et al. Adiposity and cancer at major anatomical sites: umbrella review of the literature. *BMJ*. 2017; 356:j477.
- Li X, Meng X, Timofeeva M, et al. Serum uric acid levels and multiple health outcomes: umbrella review of evidence from observational studies, randomised controlled trials, and Mendelian randomisation studies. *BMJ*. 2017;357:j2376.
- Theodoratou E, Tzoulaki I, Zgaga L, Ioannidis JPA. Vitamin D and multiple health outcomes: umbrella review of systematic reviews and meta-analyses of observational studies and randomised trials. *BMJ*. 2014;348:g2035.
- Veronese N, Solmi M, Caruso MG, et al. Dietary fiber and health outcomes: an umbrella review of systematic reviews and meta-analyses. *Am J Clin Nutr.* 2018;107(3):436-444.
- Towheed TE, Hochberg MC, Shea BJ, Wells G. WITHDRAWN: analgesia and non-aspirin, non-steroidal anti-inflammatory drugs for osteoarthritis of the hip. *Cochrane Database Syst Rev.* 2007;(1): CD000517.
- Ioannidis JA. The proposal to lower p value thresholds to .005. JAMA. 2018;319(14):1429-1430.
- Benjamin DJ, Berger JO, Johannesson M, et al. Redefine statistical significance. Nat Hum Behav. 2018;2(1):6-10.
- Ioannidis JP. Publishing research with P-values: prescribe more stringent statistical significance or proscribe statistical significance? *Eur Heart J.* 2019;40:2553.
- Guyatt GH, Oxman AD, Vist GE, et al. GRADE: an emerging consensus on rating quality of evidence and strength of recommendations. *BMJ*. 2008;336:924.
- O'Brien CW, Juraschek SP, Wee CC. Prevalence of aspirin use for primary prevention of cardiovascular disease in the United States: results from the 2017 National Health Interview Survey. *Ann Intern Med.* 2019;171(8):596.
- 42. Gaziano JM, Brotons C, Coppolecchia R, et al. Use of aspirin to reduce risk of initial vascular events in patients at moderate risk of cardiovascular disease (ARRIVE): a randomised, double-blind, placebo-controlled trial. *The Lancet*. 2018;392:1036-1046.
- 43. Group ASC. Effects of aspirin for primary prevention in persons with diabetes mellitus. *N Engl J Med.* 2018;379:1529-1539.
- McNeil JJ, Nelson MR, Woods RL, et al. Effect of aspirin on all-cause mortality in the healthy elderly. N Engl J Med. 2018;379: 1519-1528.
- Simpson SH, Gamble JM, Mereu L, Chambers T. Effect of aspirin dose on mortality and cardiovascular events in people with diabetes: a meta-analysis. J Gen Intern Med. 2011;26:1336-1344.
- 46. Karmali KN, Huffman MD. I do not have heart disease—should I be taking aspirin? JAMA Cardiol. 2017;2:824.
- Lok P, Dijk S. Offer daily aspirin to cut risk of colorectal cancer in people with Lynch syndrome, says NICE. BMJ (Online). 2019;366: 15010.
- Garcia Rodriguez LA, Martin-Perez M, Hennekens CH, Rothwell PM, Lanas A. Bleeding risk with long-term low-dose aspirin: a systematic review of observational studies. *PLoS ONE*. 2016;11:e0160046.
- Ioannidis JP, Trikalinos TA. The appropriateness of asymmetry tests for publication bias in meta-analyses: a large survey. CMAJ. 2007;176 (8):1091-1096.
- Ioannidis JP, Patsopoulos NA, Rothstein HR. Reasons or excuses for avoiding meta-analysis in forest plots. *BMJ*. 2008;336(7658):1413-1415.

- Ioannidis JP. The mass production of redundant, misleading, and conflicted systematic reviews and meta-analyses. *Milbank Q*. 2016;94(3): 485-514.
- 52. Kavvoura FK, Liberopoulos G, Ioannidis JP. Selection in reported epidemiological risks: an empirical assessment. *PLoS Med*. 2007;4:e79.
- Fisher M, Knappertz V. The dose of aspirin for the prevention of cardiovascular and cerebrovascular events. *Curr Med Res Opin*. 2006;22 (7):1239-1248.

SUPPORTING INFORMATION

Additional supporting information may be found online in the Supporting Information section at the end of this article.

How to cite this article: Veronese N, Demurtas J, Thompson T, et al. Effect of low-dose aspirin on health outcomes: An umbrella review of systematic reviews and meta-analyses. *Br J Clin Pharmacol*. 2020;86:1465–1475. https://doi.org/10.1111/bcp.14310