

Potential Role of Social Distancing in Mitigating Spread of Coronavirus Disease, South Korea

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In South Korea, the coronavirus disease outbreak peaked at the end of February and subsided in mid-March. We analyzed the likely roles of social distancing in reducing transmission. Our analysis indicated that although transmission might persist in some regions, epidemics can be suppressed with less extreme measures than those taken by China.

The first coronavirus disease (COVID-19) case in South Korea was confirmed on January 20, 2020 (1). In the city of Daegu, the disease spread rapidly within a church community after the city's first case was reported on February 18 (1). Chains of transmission that began from this cluster distinguish the epidemic in South Korea from that in any other country. As of March 16, a total of 8,236 cases were confirmed, of which 61% were related to the church (1).

The Daegu Metropolitan Government implemented several measures to prevent the spread of COVID-19. On February 20, the Daegu Metropolitan Government recommended wearing masks in everyday life and staying indoors (2). On February 23, South Korea raised its national alert level to the highest level (1) and delayed the start of school semesters (3). Intensive testing and contact tracing enabled rapid identification and isolation of case-patients and reduction of onward transmission (4). We describe potential roles of social distancing in mitigating COVID-19 spread in South Korea by comparing metropolitan traffic data with transmission in 2 major cities.

The Study

We analyzed epidemiologic data describing the COVID-19 outbreak in South Korea during January

20–March 16. We transcribed daily numbers of reported cases in each municipality from Korea Centers for Disease Control and Prevention (KCDC) press releases (1). We also transcribed partial line lists from press releases by KCDC and municipal governments. All data and code are stored in a publicly available GitHub repository (<https://github.com/parksw3/Korea-analysis>).

We compared epidemiologic dynamics of COVID-19 from 2 major cities: Daegu (2020 population: 2.4 million) and Seoul (2020 population: 9.7 million). During January 20–March 16, KCDC reported 6,083 cases from Daegu and 248 from Seoul. The Daegu epidemic was characterized by a single large peak followed by a decrease (Figure 1, panel A); the Seoul epidemic comprised several small outbreaks (Figure 1, panel B).

We obtained the daily number of persons who boarded the subway or monorail in Daegu and Seoul during 2017–2020. For Daegu, we used data from <https://data.go.kr> for lines 1–3; for Seoul, we used data from <https://data.seoul.go.kr> for lines 1–9 (Figure 1). Soon after the first church-related case was reported, traffic volume decreased by $\approx 80\%$ in Daegu and $\approx 50\%$ in Seoul. To our knowledge, KCDC first recommended social distancing on February 29 (1), and no official guidelines existed regarding public transportation, which suggests that distancing was, at least in part, voluntary.

We reconstructed the time series of a proxy for incidence of infection I_t , representing the number of persons who became infected at time t and reported later, and estimated the instantaneous reproduction number, R_t , defined as the average number of secondary infections caused by an infected person, given conditions at time t (5). We adjusted the daily number of reported cases to account for changes in testing criteria and censoring bias (Appendix, <https://wwwnc.cdc.gov/EID/article/26/11/20-1099-App1.pdf>) and then sampled infection dates using inferred onset-to-confirmation delay distributions from the partial line list (Appendix Figure 1) and previous

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estimated incubation period distribution (Table) to obtain our incidence proxy, I_t . Finally, we estimated R_t on the basis of the renewal equation (5):

$$R_t = \frac{I_t}{\sum_{k=1}^{14} I_{t-k} w_k}$$

where w_t is the generation-interval distribution randomly drawn from a prior distribution (Table). We weighted each sample of R_t using a gamma probability distribution with a mean of 2.6 and a SD ± 2 to reflect prior knowledge (S. Abbott, unpub. data, <https://doi.org/10.12688/wellcomeopenres.16006.1>) and took weighted quantiles to calculate medians and associated 95% credible intervals. We estimated R_t for February 2 (14 days after the first confirmed case) through March 10 (after which the effects of censoring were too strong for reliable estimates) (Appendix). All analyses were performed using R version 3.6.1 (<https://www.r-project.org>).

We reconstructed incidence proxy (Figure 2, panels A, B) and estimates of R_t (Figure 2, panels C, D) in Daegu and Seoul. In Daegu, incidence peaked shortly after the first case was confirmed (Figure 2, panel A). In Daegu, symptoms had developed in the first case-patient on February 7; this person had visited the church on February 9 and 16, indicating the disease probably was spreading within the church community earlier (1). Likewise, the estimates of R_t gradually decreased and eventually decreased to <1 approximately 1 week after the first case was reported, coinciding with the decrease in the metropolitan traffic volume (Figure 2, panel C). The initial decrease in R_t was likely to have been caused by our resampling method for infection times for each reported case, which oversmooths the incidence curve and the R_t estimates (K. Gostic, unpub. data, [858\). In Seoul, estimates of \$R_t\$ decreased slightly but remained at \$\approx 1\$ \(Figure 2, panel D\), which might be explained by less-intense social distancing. Stronger distancing or further intervention would have been necessary to reduce \$R_t\$ to \$<1\$ by March 10.](https://doi.org/10.1101/2020.06.18.20134</p>
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Although we found clear, positive correlations on a daily scale between normalized traffic and the median estimates of R_t in Daegu ($r = 0.93$; 95% credible interval 0.86–0.96; Appendix Figure 2) and Seoul ($r = 0.76$; 95% credible interval 0.60–0.87; Appendix Figure 2), these correlations are conflated by time trends and by other measures that could have affected R_t . We did not find clear signatures of lags in the correlation between R_t and traffic volume (Appendix Figure 3). Patterns in R_t were similar in directly adjacent provinces (Gyeongsangbuk-do and Gyeonggi-do), demonstrating the robustness of our analysis (Appendix Figure 4).

Conclusions

The South Korea experience with COVID-19 provides evidence that epidemics can be suppressed with less extreme measures than those taken by China (9) and demonstrates the necessity of prompt identification and isolation of case-patients in preventing spread (4). Our analysis reveals the potential role of social distancing in assisting such efforts. Even though social distancing alone might not prevent spread, it can flatten the epidemic curve (compare Figure 2, panels B, D) and reduce the burden on the healthcare system (10).

Our study is not without limitations. Because of insufficient data, we could not account for differences in delay distributions or changes in testing capacity among cities; line list data were mostly derived from outside Daegu. Nonetheless, the sensitivity analyses support the robustness of our findings (Appendix

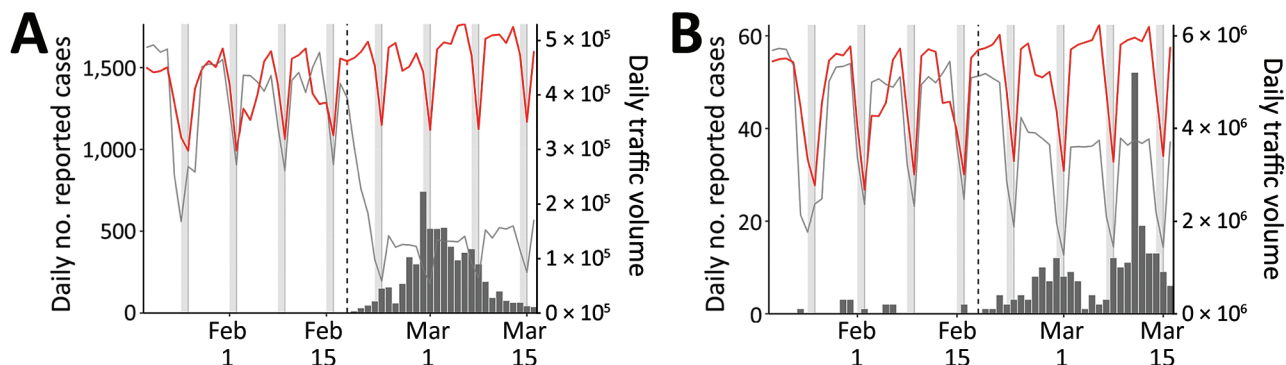


Figure 1. Comparison of daily epidemiologic and traffic data from Daegu (A) and Seoul (B) during the coronavirus disease (COVID-19) outbreak, South Korea. Black bars indicate no. COVID-19 cases; lines represent daily metropolitan traffic volume in 2020 (red) and mean daily metropolitan traffic volume during 2017–2019 (black). Daily traffic from previous years have been shifted by 1–3 days to align day of the weeks. Vertical dashed lines indicate February 18, 2020, when the first COVID-19 case was confirmed in Daegu. Gray bars indicate weekends.

Table. Assumed incubation and generation-interval distributions in an analysis of the potential role of social distancing in mitigating the spread of coronavirus disease, South Korea, 2020*

Distribution	Parameterization	Priors	Source
Incubation period distribution	Gamma ($\mu_I, \mu_I^2/\sigma^2$)	$\mu_I \approx \text{gamma}$ (6.5 d, 145); $\sigma \approx \text{gamma}$ (2.6 d, 25)	(6)
Generation-interval distribution	Negative binomial (μ_G, θ)	$\mu_G \approx \text{gamma}$ (5 d, 62); $\theta \approx \text{gamma}$ (5, 20)	(7,8)

*Gamma distributions are parameterized using its mean and shape. Negative binomial distributions are parameterized using its mean and dispersion. Priors are chosen such that the 95% quantiles of prior means and standard deviations are consistent with previous estimates.

Figures 5–8). We were unable to distinguish local and imported cases and thus might have overestimated R_t (11). Conducting a separate analysis for Seoul that accounts for imported cases did not affect our qualitative conclusions (Appendix Figure 9). Finally, although the method of resampling infection time can capture qualitative changes in R_t , estimates of R_t can be oversmoothed and should be interpreted with care (K. Gostic, unpub. data, <https://doi.org/10.1101/2020.06.18.20134858>). Nonetheless, our estimates of R_t are broadly consistent with previous estimates (12).

We used metropolitan traffic to quantify the degree of social distancing. The 80% decrease in traffic volume suggests that distancing measures in Daegu might be comparable to those in Wuhan, China (13). We were unable to directly estimate the effect of social distancing on population contacts or epidemiologic dynamics. Other measures, such as intensive testing and tracing of core transmission groups, are also likely to have affected transmission dynamics.

Our study highlights the importance of considering geographic heterogeneity in estimating epidemic potential. The sharp decrease in Daegu drove the number of reported cases in South Korea. Our analysis revealed that the epidemic remained close to the epidemic threshold in other regions, including Seoul and Gyeonggi-do. Relatively weak distancing might have assisted the recent resurgence of COVID-19 cases in Seoul (E. Shim, G. Chowell, unpub. data, <https://doi.org/10.1101/2020.07.21.20158923>).

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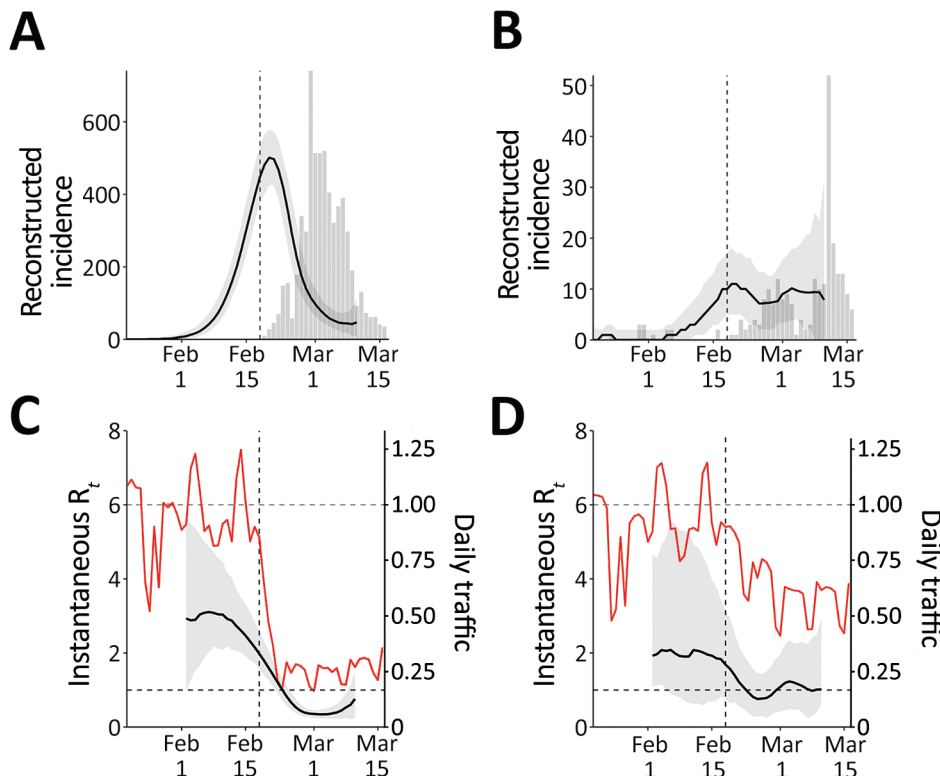


Figure 2. Comparison of reconstructed coronavirus disease incidence proxy and instantaneous reproduction number R_t in Daegu (A, C) and Seoul (B, D), South Korea. The instantaneous reproduction number R_t reflects transmission dynamics at time t . Black lines and gray shading represent the median estimates of reconstructed incidence (A, B) and R_t (C, D) and their corresponding 95% credible intervals. Gray bars show the number of reported cases. Red lines represent the normalized traffic volume (daily traffic, 2020, divided by the mean daily traffic, 2017–2019). Vertical dashed lines indicate February 18, 2020, when the first case was confirmed in Daegu.

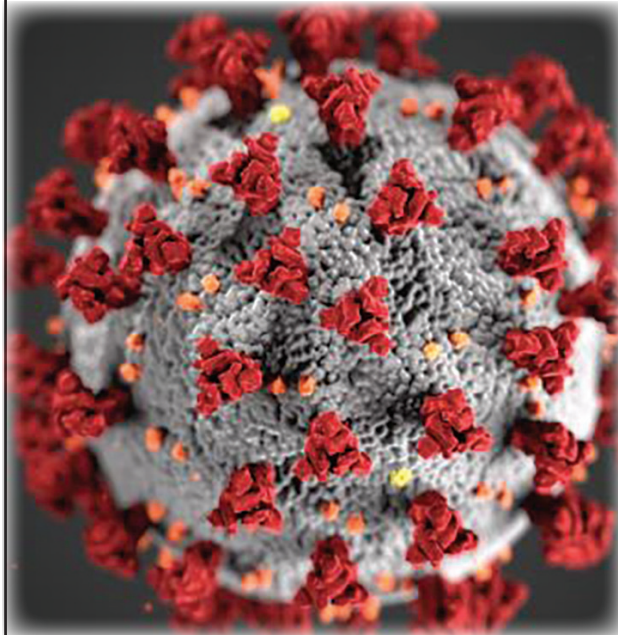
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