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Inherited myeloproliferative neoplasm risk impacts hematopoietic stem cells

Erik L. Bao^{1,2,3,23}, Satish K. Nandakumar^{1,2,23}, Xiaotian Liao^{1,2,23}, Alexander G. Bick^{2,4,5}, Juha Karjalainen⁶, Marcin Tabaka², Olga I. Gan^{7,8}, Aki Havulinna⁶, Tuomo Kiiskinen⁶, Caleb A. Lareau^{1,2,9}, Aitzkoa Lopez de Lapuente Portilla¹⁰, Bo Li^{2,11}, Connor Emdin^{2,4}, Veryan Codd¹², Christopher P. Nelson¹², Christopher J. Walker¹³, Claire Churchhouse², Albert de la Chapelle¹³, Daryl E. Klein¹⁴, Björn Nilsson^{2,10}, Peter W.F. Wilson^{15,16}, Kelly Cho^{17,18}, Saiju Pyarajan¹⁷, J. Michael Gaziano^{17,18}, Nilesh J. Samani¹², FinnGen, 23andMe Research Team, Million Veteran Program, Aviv Regev^{2,19}, Aarno Palotie^{2,6}, Benjamin M. Neale², John E. Dick^{7,8}, Pradeep Natarajan^{2,4,20}, Christopher J. O'Donnell^{5,18}, Mark J. Daly^{2,6}, Michael Milyavsky²¹, Sekar Kathiresan^{2,4}, Vijay G. Sankaran^{1,2,22}

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*Correspondence and requests for materials should be addressed to sankaran@broadinstitute.org (V.G.S.).

Author contributions

E.L.B. and V.G.S. conceived the study. E.L.B., S.K.N., and V.G.S. designed the study. S.K.N., X.L., O.I.G., D.E.K., and M.M. performed experiments. E.L.B., X.L., A.B., J.K., M.T., A.H., T.K., C.A.L., A.L.P., D.E.K., B.L., and C.E. performed computational and statistical analyses. C.C. and B.M.N. contributed to genetic analysis of UK Biobank. A.B., C.E., P.N., P.W.F.W., K.C., S.P., J.M.G., C.J.O., and S.K. contributed to genetic analysis of the Million Veteran Program. J.K., A.H., T.K., A.P., M.J.D. contributed to genetic analysis of FinnGen. M.T., B.L., and A.R. contributed to analysis of the Human Cell Atlas. V.C., C.P.N., N.J.S. contributed to genetic analysis of leukocyte telomere length. C.J.W. and A.d.I.C. contributed to genetic analysis of acute myeloid leukemia. A.L.P., B.N., J.E.D., and M.M. contributed ideas and insights. V.G.S. supervised all experimental and analytic aspects of this work. E.L.B., S.K.N., X.L., and V.G.S. wrote the manuscript with input from all authors. All authors read and approved the final version of the manuscript.

Author disclosures

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Data availability

Summary statistics for variants with fine-mapped posterior probability >0.1% from the full GWAS meta-analysis (UK Biobank, FinnGen, 23andMe) are available in Supplementary Table 5. Full summary statistics from 23andMe data cannot be reported due to a clause in the 23andMe data transfer agreement, intended to protect the privacy of the 23andMe research participants. Thus, we provide full summary statistics for the MPN meta-analysis comprising UK Biobank and FinnGen cohorts on GWAS Catalog under accession code GCST90000032 (<https://www.ebi.ac.uk/gwas/downloads/summary-statistics>). To fully recreate our meta-analysis results for MPN: researchers can (1) obtain MPN summary statistics from 23andMe (<https://research.23andme.com/dataset-access/>), and (2) conduct a meta-analysis of our summary statistics with the 23andMe summary statistics. For downloads of FinnGen summary statistics, information on how to access individual level FinnGen data by application to responsible agencies (FinBB and THL), and other collaborative access inquiries, please see <https://www.finnngen.fi/en/access>. Individual genetic and phenotypic data for the following cohorts are available by application: UK Biobank (<https://www.ukbiobank.ac.uk>) and Million Veteran Program (<https://www.research.va.gov/MVP/research.cfm>).

Code availability

Code and source data required for reproducing results and figures discussed herein are available on GitHub (<https://github.com/sankaranlab/mpn-gwas>).

¹Division of Hematology/Oncology, Boston Children's Hospital and Department of Pediatric Oncology, Dana-Farber Cancer Institute, Harvard Medical School, Boston, MA, USA ²Broad Institute of MIT and Harvard, Cambridge, MA, USA ³Harvard-MIT Health Sciences and Technology, Harvard Medical School, Boston, MA, USA ⁴Center for Genomic Medicine, Massachusetts General Hospital, Boston, MA, USA ⁵VA Boston Healthcare, Section of Cardiology and Department of Medicine, Boston, MA, USA ⁶Institute for Molecular Medicine Finland, Helsinki, Finland ⁷Princess Margaret Cancer Centre, University Health Network, Toronto, ON, Canada ⁸Department of Molecular Genetics, University of Toronto, Toronto, ON, Canada ⁹Program in Biological and Biomedical Sciences, Harvard Medical School, Boston, MA, USA ¹⁰Hematology and Transfusion Medicine, Department of Laboratory Medicine, Lund University, Lund, Sweden ¹¹Center for Immunology and Inflammatory Diseases, Division of Rheumatology, Allergy, and Immunology, Massachusetts General Hospital and Harvard Medical School, Boston, MA, USA ¹²Department of Cardiovascular Sciences and NIHR Leicester Biomedical Centre, University of Leicester, United Kingdom ¹³Department of Cancer Biology and Genetics, The Ohio State University Comprehensive Cancer Center, Columbus, OH, USA ¹⁴Department of Pharmacology, Cancer Biology Institute, Yale University School of Medicine, West Haven, CT, USA ¹⁵Atlanta VA Medical Center, Atlanta, GA, USA ¹⁶Emory Clinical Cardiovascular Research Institute, Atlanta, GA, USA ¹⁷Massachusetts Veterans Epidemiology Research and Information Center (MAVERIC), VA Boston Healthcare System, Boston, MA, USA ¹⁸Department of Medicine, Brigham and Women's Hospital, Boston, MA, USA ¹⁹Howard Hughes Medical Institute, Department of Biology and Koch Institute of Integrative Cancer Research, Massachusetts Institute of Technology, Cambridge, MA, USA ²⁰Cardiovascular Research Center, Massachusetts General Hospital, Boston, MA, USA ²¹Department of Pathology, Sackler Faculty of Medicine, Tel-Aviv University, Tel-Aviv, Israel ²²Harvard Stem Cell Institute, Cambridge, MA, USA ²³These authors contributed equally: Erik L. Bao, Satish K. Nandakumar, Xiaotian Liao.

Abstract

Myeloproliferative neoplasms (MPNs) are blood cancers characterized by excessive production of mature myeloid cells, which result from the acquisition of somatic driver mutations in hematopoietic stem cells (HSCs). Epidemiologic studies indicate a substantial disease heritability that is among the highest known for cancers¹. However, only a limited set of genetic risk loci have been identified, and the underlying biological mechanisms leading to MPN acquisition remain unexplained. Here, we conducted a large-scale genome-wide association study (3,797 cases and 1,152,977 controls) to identify 17 MPN risk loci ($p < 5.0 \times 10^{-8}$), seven of which have not been previously reported. We find a shared genetic architecture between MPN risk and several hematopoietic traits spanning distinct lineages, an enrichment for risk variants mapping to accessible chromatin in HSCs, and associations of increased MPN risk with longer leukocyte telomere length and other clonal hematopoietic states, collectively implicating HSC function and self-renewal. Gene mapping identifies modulators of HSC biology and targeted variant-to-function assays suggest likely roles for *CHEK2* and *GFI1B* in altering HSC function to confer disease risk. Overall, we demonstrate the power of human genetic studies to illuminate a previously unappreciated mechanism for inherited MPN risk through modulation of HSC function.

Myeloproliferative neoplasms (MPNs) comprise a group of blood cancers characterized by excessive production of mature myeloid cells. Although MPNs arise from somatic driver mutations, recent work has uncovered a substantial heritable component to the disease. There is a 5–7-fold increased risk of MPN acquisition in first-degree relatives of individuals with the disease^{1,2}, amounting to a familial risk greater than that observed in breast (relative risk (RR) = 3.5)³, prostate (RR = 2.46)⁴, and colorectal cancer (RR=2.25)⁵. Genetic studies to date have identified a limited set of loci associated with MPN predisposition^{6,7}. However, many key aspects of inherited MPN risk remain uncharacterized, including the causal genes, cell types, and underlying biological mechanisms involved.

Genetic discovery of inherited MPN risk

To characterize the germline genetic architecture that confers risk for MPNs, we conducted a genome-wide association study (GWAS) meta-analysis using three population-based cohorts (UK Biobank (UKBB), 23andMe, and FinnGen), comprising a combined sample size of 2,949 MPN cases and 835,554 controls (Supplementary Table 1, Extended Data Fig. 1). We tested 7,343,617 well-represented variants passing central and study-specific quality control measures. Linkage disequilibrium score regression (LDSC)⁸ showed negligible inflation in test statistics due to population structure ($\lambda = 1.035$, intercept = 1.01).

Approximate conditional analysis of GWAS signals⁹ revealed 15 linkage disequilibrium (LD)-independent loci at genome-wide significance ($p < 5 \times 10^{-8}$) and an additional 10 loci with suggestive associations ($p < 1 \times 10^{-6}$) (Fig. 1a, Supplementary Table 2). To replicate these associations in an independent cohort, we assessed genotypes at 24 of 25 lead variants ($p < 1 \times 10^{-6}$) among 848 carriers of the somatic JAK2 V617F mutation, the most common driver mutation for MPNs, and up to 317,423 controls within the Million Veteran Program. While JAK2 V617F is not an exact phenotypic proxy for MPN, previous studies have shown that JAK2 V617F strongly replicates MPN risk associations⁷. Because our study focuses on inherited risk prior to MPN driver mutation acquisition, we reasoned that JAK2 V617F could model a pre-MPN state. A perfect replication for direction of effects was observed (24 of 24 variants, binomial $p = 5.96 \times 10^{-8}$), and 17 variants were significant at the 5% level (binomial $p < 2.2 \times 10^{-16}$) (Extended Data Fig. 2, Supplementary Table 3). We then combined data from MVP with the discovery GWAS to reach a total of 3,797 cases and 1,152,977 controls. This combined analysis revealed 17 independent risk associations exceeding genome-wide significance, seven of which have not been previously reported (Extended Data Table 1, Supplementary Note). All 10 previously reported loci for MPN risk^{6,7} remained genome-wide significant in our analysis. We estimate that the 17 loci explain 18.4% of the ~5-fold familial relative risk for MPN acquisition¹.

To further characterize the polygenic architecture of MPN risk, we generated a 104-variant polygenic risk score (PRS) using a pruning and thresholding method ($R^2 < 0.2$, $P < 1 \times 10^{-5}$) on summary statistics from 23andMe and FinnGen (Supplementary Table 4), and then tested this PRS in the out-of-sample UKBB cohort. Given the rarity of MPNs and the relatively common frequency of these risk variants, a PRS for MPNs is unlikely to have utility for population-based screening. Nonetheless, PRS analysis is useful for evaluating the distribution of genetic risk and comparing the predictive capacity of germline variants with

other risk factors. Individuals with MPN in the UKBB had a median PRS percentile of 67 compared to 50 for those without MPN (Fig. 1b, Extended Data Fig. 3), and the PRS explained more variance than any other covariate in our model (Fig. 1c). Individuals in the 20th to 70th PRS percentiles had similar disease risk, but significant stratification occurred at the top and bottom quintiles of the distribution (Fig. 1d). For example, the 40,824 individuals above the 90th PRS percentile had 4.20-fold (95% CI: [3.14, 5.63]) higher odds of MPN compared to those in the lowest decile, and 2.70-fold (95% CI: [2.11, 3.46]) higher odds compared to those with average genetic risk (5th PRS decile). Finally, given that the *JAK2* 46/1 haplotype is one of the largest contributors to germline MPN risk¹⁰, we sought to determine the extent to which a broader set of risk loci, as represented by our PRS, could modify the penetrance of the *JAK2* 46/1 haplotype. When we stratified individuals by both PRS and *JAK2* 46/1 carrier status, MPN risk varied substantially, even among *JAK2* 46/1 carriers (Extended Data Fig. 3). Compared to non-carriers of *JAK2* 46/1, the risk among carriers with intermediate PRS risk (middle three quintiles) was 0.87-fold (95% CI: [0.72, 1.07]), but 1.87-fold (95% CI: [1.61, 2.16]) for those with high PRS (top quintile). Validation of this PRS in external cohorts is necessary to further assess the generalizability of these risk variants in different populations. Nevertheless, these results indicate that polygenic profiles can refine stratification of MPN risk.

Next, we performed Bayesian fine-mapping to prioritize putative causal variants at each MPN risk association (Methods)¹¹. To improve detection power for downstream functional enrichments, we included all 25 suggestive loci in this analysis. Briefly, for each LD-independent signal, we extracted all variants within a 2 Mb window, calculated posterior probabilities of causality (PP) for each variant based on effect size estimates, ranked variants by decreasing PP, and added variants to credible sets until a cumulative PP > 95% was reached. Of the 25 credible sets, two mapped to a single variant, and six contained five or fewer variants. In 14 regions, the top fine-mapped variant had a PP > 0.25 (Extended Data Fig. 3). To further resolve loci containing highly correlated variants, we annotated fine-mapped variants with functional criteria including regulatory features, evolutionary conservation, target genes, and transcription factor motif disruption potential (Supplementary Table 5).

Enriched cell types underlying MPN risk

We next attempted to identify relevant cell states underlying MPN predisposition. Because MPNs arise in the hematopoietic compartment, we examined genetic correlations between MPN risk and 19 blood cell traits in 408,241 European ancestry individuals from the UKBB. MPN risk demonstrated positive genetic correlations with counts of red blood cells, platelets, total white blood cells, monocytes, neutrophils, and eosinophils – all of which derive from multipotential hematopoietic stem and progenitor cells (HSPCs) (Fig. 2a, Extended Data Fig. 4). This led us to wonder whether MPN risk variants may be pleiotropically influencing distinct blood lineages from a common progenitor population. We previously showed that variants associated with multiple blood lineages preferentially co-localize with accessible chromatin of more primitive HSPCs¹². Similarly here, a greater proportion of well fine-mapped MPN risk variants (PP > 0.10) exhibited pleiotropic blood trait associations compared to more weakly fine-mapped variants (PP < 0.10) (Fig. 2b,

21.4% vs. 1.6%, Fisher's exact test $p = 8.48 \times 10^{-7}$), supporting the concept that MPN risk variants may act in hematopoietic progenitors to influence multiple lineages.

We next leveraged chromatin accessibility (ATAC-seq) data across 18 human hematopoietic populations to elucidate causal cell types underlying MPN predisposition. 11.6% (34/294) of MPN risk variants with $PP > 0.01$ fell within accessible chromatin of one or more hematopoietic populations, compared to 4.82% (686/14,235) of variants with $PP < 0.01$ (χ^2 $p = 2.13 \times 10^{-6}$), suggesting that stronger fine-mapped variants are enriched for regulatory function in hematopoietic populations. We next used g-chromVAR¹² to compute enrichments of fine-mapped variants across these 18 hematopoietic chromatin accessibility profiles. Strikingly, in contrast to variants associated with peripheral blood cell traits, which are maximally enriched in terminally differentiated hematopoietic populations, MPN risk variants showed the strongest enrichments in multipotent progenitors (MPPs) and HSCs ($p = 5.59 \times 10^{-3}$ and 1.42×10^{-2} , respectively) (Fig. 2c). Because g-chromVAR only considers variants from significant loci, we applied LDSC to assess for enrichments in cell type-specific accessible chromatin on a genome-wide level. We again observed that HSCs and MPPs showed the highest enrichment relative to all other hematopoietic populations ($p = 7.10 \times 10^{-3}$ and 1.89×10^{-2} , respectively) (Fig. 2d).

Linking genetic MPN risk to other traits

To gain additional insights into mechanisms of MPN predisposition, we examined the relationship between MPN risk and leukocyte telomere length (LTL). LTL is correlated with telomere length of earlier hematopoietic progenitors¹³, telomere length is associated with HSC self-renewal potential¹⁴, and individuals with telomerase loss-of-function associated-diseases have fewer HSCs¹⁵. Motivated by the robust MPN risk associations at the telomerase reverse transcriptase (*TERT*) locus (Extended Data Table 1), we assessed the genetic overlap between MPN risk and LTL from a study involving 78,592 individuals¹⁶. At the *TERT* locus, the top two independent variants for increased telomere length, rs7705526 ($p = 5.34 \times 10^{-45}$) and rs2853677 ($p = 3.35 \times 10^{-31}$), were also lead variants for MPN risk ($p = 4.78 \times 10^{-54}$ and $p = 2.79 \times 10^{-44}$, respectively; $PP = 1$ for both variants) (Extended Data Fig. 4). We observed a positive genetic correlation between LTL and MPN risk both within the *TERT* locus (Pearson $r = 0.84$, $p < 0.001$) (Fig. 2e) and genome-wide (LDSC $r_g = 0.19$, s.e.m. = 0.093, $p = 0.037$). We then tested whether increased telomere length may be causally linked to MPN risk by performing two-sample Mendelian randomization (MR), using independent genetic predictors of telomere length as instruments. Applying four related MR algorithms with varying assumptions regarding horizontal pleiotropy, we found that increased telomere length was consistently associated with increased MPN risk at a conservative multiple-testing-adjusted threshold of $p < 6.25 \times 10^{-3}$ (Fig. 2f). The reverse association was significant in only one out of the four tests (see Methods). These results support a positive risk effect of increased telomere length on MPN acquisition. Importantly, however, other causal pathways may be mediating these effects. For instance, given the known association between HSC self-renewal and telomere length, a causal effect of HSC self-renewal on MPN risk could lead to similar MR estimates.

Because variants affecting telomere length have been implicated in many other cancers¹⁶, we performed a phenome-wide association study (pheWAS) to measure the effects of MPN risk variants on 1,130 well-represented case-control phenotypes from the UKBB¹⁷ (Supplementary Table 6). Consistent with previous studies, the two lead variants at *TERT* (rs7705526 and rs2853677), as well as others near *ATM*, *SH2B3*, and *MRPS3*, demonstrated associations with several other cancers (Extended Data Fig. 4). In addition, we noted a positive correlation between genetic risk for MPN and clonal hematopoiesis of indeterminate potential (CHIP) ($r_g = 0.51$, s.e.m. = 0.25, $p = 0.04$)¹⁸. Finally, we identified a greater than expected replication of MPN risk loci with acute myeloid leukemia (AML) risk – 5/17 lead MPN variants were significant for AML risk at the 5% level with concordant direction of effect (binomial $p = 1.17 \times 10^{-3}$, Supplementary Table 7). These findings suggest that MPN risk loci may promote risk not only for MPNs, but also somatic mutation acquisition in HSCs leading to other clonal disorders.

Genes underlying inherited MPN risk

Next, we sought to gain further biological insights into germline MPN predisposition by using an integrative approach to map risk variants to target genes. We first nominated three genes from risk loci containing variants at PP > 0.10 with missense consequences: rs1800057 causes the P1054R substitution in *ATM*, rs3184504 causes the R262W substitution in *SH2B3*, and rs17879961 causes the I157T substitution in *CHEK2*. For the remaining 22 non-coding loci, we nominated target genes based on five markers of biological evidence for regulatory function: 1) gene body colocalization, 2) mapping to hematopoietic promoter capture Hi-C (PCHi-C) interactions, 3) correlation between chromatin accessibility and nearby gene expression across hematopoietic cell types, 4) MAGMA gene-wise association signal, and 5) overlap with genes harboring recurrent somatic MPN driver mutations. Genes were selected if they scored in two or more non-coding criteria and were the highest scoring gene in the locus. Using this approach, combined with the three genes harboring coding variants, we identified a single high-confidence target gene in 15/25 MPN risk loci (Extended Data Fig. 5). The nominated genes displayed a much stronger set of functional interactions than expected by chance (STRING database¹⁹, 17 vs. 4 expected interactions, $p = 3.1 \times 10^{-7}$) (Extended Data Fig. 5). Remarkably, 11 of the 15 genes have known roles as modulators of HSC self-renewal and function, including *ZNF521*²⁰, *GATA2*²¹, *MECOM*²², *RUNX1*²³, *HMGA1*²⁴, *ATM*²⁵, *FOXO1*²⁶, *TET2*²⁷, *JAK2*²⁸, *SH2B3*^{29,30} and *TERT*³¹ (Fig. 3a). The most enriched biological annotations included replicative senescence, cell cycle, and hematopoietic and lymphoid development (Extended Data Fig. 5, Supplementary Table 8). Analysis of bulk RNA sequencing of 16 primary hematopoietic populations showed a maximal enrichment of target genes for expression in HSCs, MPPs, and myeloid committed progenitors (Fig. 3a). To extend these observations to single cell resolution, we examined the expression of MPN target genes in 278,978 human bone marrow cells. We observed an MPN gene signature that was higher in HSC-enriched cell clusters compared to all other hematopoietic cells. Following gene imputation, we also noted co-localization between MPN and HSC signatures across all cells (Spearman $r_s = 0.13$, $p < 2.2 \times 10^{-16}$) (Fig. 3b–c, Supplementary Figs. 1–2).

Together, these results indicate that target genes implicated in MPN risk are enriched for HSC function and expression.

Characterizing mechanisms of MPN risk

In two cases, we performed targeted variant-to-function analyses to gain mechanistic insights into MPN predisposition, starting with the I157T missense variant (rs17879961) in *CHEK2*. Importantly, this variant is known to be a hypomorphic allele with functionally impaired activation of downstream effectors^{32,33}, and has been previously linked to increased risk for several cancers. Structural analysis of the *CHEK2* protein revealed that the I157T variant would diminish hydrophobic interfaces and destabilize the *CHEK2* dimer (Extended Data Fig. 6), consistent with prior reports examining the biochemical consequences of this loss-of-function mutation³². We found that *CHEK2* inhibition³⁴ reduced cell death after irradiation of primitive human Lin⁻CD34⁺CD38⁻ cells, as compared to more differentiated progenitors (Extended Data Fig. 7). Remarkably, while we did not observe skewed lineage commitment in colony assays by *CHEK2* inhibition (Extended Data Fig. 7), suppression of *CHEK2* through RNA interference increased expansion of human cord blood Lin⁻CD34⁺ in long-term cultures in the absence of genotoxic stress (Fig. 4a). These results suggest that *CHEK2* ordinarily constrains HSPC expansion, and the I157T variant may reduce *CHEK2* function to promote self-renewal and thereby increase MPN risk.

In a second instance, we identified a compelling risk locus in a region of hematopoietic accessible chromatin located ~12 kb downstream of *GFI1B*, a master transcription factor necessary for maintaining HSC quiescence³⁵ and promoting megakaryocytic/erythroid differentiation³⁶ (Fig. 4b). The top fine-mapped variants at this locus were rs1633768 (PP = 0.28) and rs524137 (PP = 0.27). While rs1633768 lacked a target gene, rs524137 mapped to *GFI1B* by ATAC-RNA correlation (Pearson $r = 0.88$, $p = 2.0 \times 10^{-14}$) and was predicted to be deleterious (scaled CADD score: 12.73) (Supplementary Table 5). In an exogenous reporter assay in hematopoietic cells, the regulatory region containing rs524137 conferred a 40-fold increase in transcriptional activity compared to a minimal promoter construct, whereas rs1633768 did not change transcriptional activity (Fig. 4c). In otherwise matched genetic constructs, rs524137-T (risk allele) resulted in a 1.7-fold decrease in enhancer activity compared to the non-risk allele when using integrating lentiviral reporters in the presence of the endogenous *GFI1B* promoter (Fig. 4d, Extended Data Fig. 8). Deletion of a 1.6 kb region encompassing this putative enhancer in human CD34⁺ HSPCs using two largely overlapping pairs of CRISPR guide RNAs resulted in 61.5% (ENH1) and 68% (ENH2) editing efficiency and up to 36.3% reduction in *GFI1B* expression, nominating *GFI1B* as the causal gene (Fig. 4e–g). Enrichment of LT-HSCs based on surface marker phenotypes revealed a 32.3% decrease in *GFI1B* expression (Fig. 4g, Extended Data Fig. 8). Previous studies have found that knockout of *Gfi1b* in mice results in expansion of hematopoietic stem cells *in vivo*³⁵. To test whether *GFI1B* enhancer deletion also expands human LT-HSCs, we quantified total numbers of LT-HSCs *in vitro* at the end of a 7-day culture. We observed a 2.7-fold increase in LT-HSCs in the *GFI1B* enhancer-edited group compared to control AAVS1-edited cells (Fig. 4h), a higher degree of expansion compared with total cells and CD34⁺ progenitors (Fig. 4i). To assess the effects of HSPC expansion

using an independent approach, we performed methylcellulose colony re-plating assays after *GFI1B* gene disruption and enhancer deletion. *GFI1B* coding disruption greatly increased self-renewal of progenitors as demonstrated through the presence of increased secondary colonies (Fig. 4j, Extended Data Fig. 8, Supplementary Fig. 3). Loss of *GFI1B* also reduced formation of primary erythroid colonies, a phenotype consistent with its role in erythroid differentiation³⁷. Deletion of the *GFI1B* enhancer similarly increased formation of secondary colonies, but did not impact erythroid colony formation (Fig. 4k–l), supporting a mechanism by which rs524137 affects an HSPC-selective enhancer of *GFI1B*. Analogous to the results obtained for *CHEK2*, these results show that this MPN risk variant reduces *GFI1B* expression in HSPCs and thereby increases self-renewal (Extended Data Fig. 9).

Discussion

In summary, we have elucidated the germline genetic architecture of MPN risk through a large-scale GWAS and implicate HSC function as an important driver through a number of functional assays. These results support a model in which MPN risk alleles expand the baseline pool of HSCs, which increases the risk of any one HSC acquiring an MPN somatic driver mutation, consistent with previous studies showing that MPNs arise in only a small subset of HSCs^{38,39}.

MPN risk variants likely exert effects both before and after acquisition of somatic driver mutations (e.g., *JAK2* V617F). In this study, we focused on how these variants act before driver mutation acquisition in order to better understand underlying mechanisms of germline risk, which is more commonly observed than the rare occurrence of overt MPNs. Work in other cancers has demonstrated that germline susceptibility can also cooperate with pathogenic driver mutations⁴⁰, and testing for the presence of this mechanism in MPNs is an important area for future research.

Our discoveries of genetic variation underlying MPN risk complement other studies identifying the genetic determinants of clonal mosaicism in blood cells^{41–43}. Understanding the fundamental mechanisms of MPN predisposition may inform the development of novel interventions for the disease, analogous to the implementation of human papillomavirus testing and colonoscopic surveillance to reduce cervical and colorectal cancer incidence^{44,45}. In this context, improved surveillance or modulation of HSC function may enable therapies to prevent progression to clonal hematopoietic disorders like MPNs^{46–48}. More broadly, our findings serve as a model for dissecting the underlying basis of cancer predisposition alleles.

Methods

UK Biobank GWAS

We performed a GWAS on MPNs using the UK Biobank (UKBB)⁴⁹. Written consent was obtained for all participants. The study was conducted with the approval of the North-West Multi-centre Research Ethics Committee, in accordance with the principles of the Declaration of Helsinki. Individuals were genotyped using the UK Biobank Axiom Array. Variant imputation was performed as previously described using a combined 1000 Genomes Phase 3-UK10K panel and Haplotype Reference Consortium (<http://biobank.ctsu.ox.ac.uk/>

[crystal/label.cgi?id=263](https://crystal.label.cgi?id=263))⁴⁹. The following sample-level quality control filters (previously generated by UKBB) were applied to subset our samples: self-reported British ancestry with similar genetic ancestry based on a principal components analysis of the genotypes, included in kinship inference, no excess (>10) of putative third-degree relatives inferred from kinship, no outlier in heterozygosity and missing rates, and no putative sex chromosome aneuploidy. Samples were filtered from UKBB .bgen files using bgenix version 1.0.1 and qctools v2.0-rc7. 408,241 individuals satisfied all of these criteria and were included for genetic association studies.

We curated a definition for the MPN phenotype within UKBB using the following codes: polycythemia (ICD10 D45; ICD9 2384), essential thrombocythemia (ICD10 D47.3, D75.2), osteomyelofibrosis (ICD10 D47.4, D75.81), chronic myeloid leukemia (ICD10 C921, C922, C931; ICD9 2051), and chronic myeloproliferative disease (ICD10 D47.1). Individuals were also classified as cases if they had a self-reported cancer, self-reported illness code, or histology of cancer tumor code for polycythemia vera, essential thrombocythemia, myelofibrosis, chronic myeloid leukemia, or malignant mastocytosis.

We used SAIGE version 0.29.4⁵⁰ to perform the GWAS using a generalized linear mixed model that controls test statistic inflation and improves power in studies with unbalanced case:control ratios. To select for variants used to estimate the genetic relatedness matrix (GRM), genotyped variants were filtered using the following criteria: minor allele frequency (MAF) > 0.01, LD pruned with r^2 threshold of 0.2, Hardy-Weinberg p-value > 1×10^{-6} , and genotype missingness < 0.01. Variants were filtered from UKBB .bgen files using bgenix version 1.0.1 and qctools v2.0-rc7. Principal components of ancestry were calculated with these variants using PLINK2 (--pca approx). Age, sex, genotyping array, and the top 10 principal components were included as covariates when fitting the logistic mixed model. For association testing, 26,942,478 autosomal and 1,039,234 X chromosome variants (excluding the pseudo-autosomal region) with MAF > 0.0001 and Info > 0.6 were included.

We also performed GWAS on 19 continuous blood traits within the same 408,241 individuals and same Info and MAF-filtered variants from the UKBB. These associations were performed using BOLT-LMM v2.3.2⁵¹, with the same covariates of age, sex, genotyping array, and top 10 principal components.

23andMe GWAS

GWAS summary statistics on MPNs from the cohort collected by the personal genetics company 23andMe, Inc. were obtained from a previous study, whose analysis has been described in-depth elsewhere⁵². A list of 23andMe contributors is presented in the Supplementary Note. Ethical approval was obtained by Ethical & Independent Review Services, an independent external AAHRPP-accredited Institutional Review Board (IRB). Written consent was obtained for all participants. Individuals were genotyped using a custom 23andMe version of the Illumina HumanOmniExpress + BeadChip. Imputation was performed using the August 2010 release of 1000 Genomes reference haplotypes. The MPN phenotype was defined using participant self-report for the following diseases: polycythemia vera (PV), essential thrombocythemia (ET), primary myelofibrosis (PMF), post-PV/ET myelofibrosis [MF], systemic mastocytosis, chronic myelogenous leukemia, chronic

eosinophilic leukemia, and hypereosinophilic syndromes. For our meta-analysis, we used the revised phenotype in their study which also included carriers of the somatic JAK2 V617F mutation. There was a total of 1,223 cases and 252,140 controls in this population. The GWAS was performed using a logistic regression model with covariates of age, gender, and the top five principal components.

FinnGen GWAS

FinnGen is a public–private partnership project combining genotype data from Finnish biobanks and digital health record data from Finnish health registries (<https://www.finnngen.fi/en>), and has been approved by the Ethics Review Board of the Hospital District of Helsinki and Uusimaa. A list of FinnGen contributors is presented in the Supplementary Note. The GWAS on MPNs in the FinnGen cohort was performed using SAIGE version 0.29.4, modified to handle missing genotypes and complete separation in covariates. We used the FinnGen release 4 data in this project, which comprised of 640 cases of MPN and 176,259 controls. Written consent was obtained for all participants. The MPN phenotype was defined by individuals with one or more clinical codes in nationwide hospital discharge or cause-of-death registries for polycythemia vera (ICD10 D45; ICD9 238.4; ICD8 208), chronic myeloproliferative disease (ICD10 D47.1), thrombocythemia (ICD10 D47.3; ICD9 238.7B; ICD8 287.2), myelofibrosis (ICD10 C9.45; ICD9 238.7A; ICD8 209), and chronic myeloid leukemia (ICD10 C92.1; ICD9 205.1). The majority of individuals were genotyped using a custom-designed Affymetrix array for the FinnGen project, with a subset of legacy samples genotyped using various generations of Illumina GWAS arrays. Variant imputation was performed by beagle4.1 software using a reference panel of 3,775 whole-genome sequenced individuals from Finland. The GRM was calculated with 49,811 variants that were imputed in every cohort, and which satisfied the following quality control filters: INFO > 0.95, MAF > 0.05, genotype missingness < 0.05, LD pruned with r^2 threshold of 0.1. Age, sex, the top 10 principal components, and genotyping batch/cohort were applied as covariates to the SAIGE logistic regression model. For association testing, 16,289,692 autosomal and X chromosome variants with Info > 0.6 were included.

GWAS meta-analysis

We aggregated association summary statistics from the UKBB, 23andMe, and FinnGen GWAS used a fixed effects model with inverse-variance weighting of log(odds ratios), as implemented in the METAL software⁵³. We meta-analyzed 7,343,617 variants which had association statistics in at least the two largest cohorts (UKBB and 23andMe). Linkage disequilibrium score regression (LDSC) of the meta-analysis showed an LDSC intercept of 1.01 and genomic control factor of 1.035, indicating negligible inflation in test statistics due to population structure. Thus, we did not adjust test statistics using genomic control.

Approximate conditional association analysis

GCTA was used to perform approximate conditional and joint association analyses (COJO) to identify independent MPN risk loci⁵⁴. In brief, this method performs a stepwise model selection (--cojo-slc) to identify all conditionally independent risk signals at a given p-value of association, using GWAS summary statistics and estimated LD from a reference panel. For estimation of LD, we used a reference sample of 6,000 unrelated individuals of white

British origin, randomly selected from the UKBB. After excluding variants with low imputation quality ($INFO < 0.4$) or deviation from Hardy-Weinberg equilibrium ($p < 1 \times 10^{-6}$), this reference panel included ~36 million variants. The reference panel was converted from BGEN files to hard-called PLINK files using PLINK2. When running GCTA-COJO, we set the threshold p-value to $p < 1 \times 10^{-6}$ and the distance for assuming complete linkage equilibrium (--cojo-wind) at 10000 kb (i.e. 10 Mb).

Million Veteran Program replication

We performed replication of our discovery meta-analysis results in up to 318,271 individuals of European descent from the Million Veteran Program (MVP)⁵⁵. Written consent was obtained for all participants. The MVP received ethical and study protocol approval from the VA Central Institutional Review Board in accordance with the principles outlined in the Declaration of Helsinki. Individuals were genotyped using the MVP 1.0 custom Axiom array. We attempted to replicate lead variants at 24 / 25 suggestive loci identified from our discovery analysis which also had genotype information within MVP (Supplementary Table 2, Extended Data Fig. 2); the one exception was rs75405916 ($p = 7.4 \times 10^{-7}$, $RAF = 0.0004$), which was not detected in MVP. Within the MVP cohort, cases ($n = 848$) were defined as individuals carrying substantial JAK2 V617F mutation burden, determined by specifying a threshold variant mutant allele intensity for rs77375493, the genotyping probe for the V617F mutation. To set the threshold, we utilized a previously reported odds ratio of 2.04 between rs7868130 (a *JAK2* 46/1 haplotype variant) and JAK2 pV617F, determined from an out-of-sample association study of 446 V617F carriers vs. 169,021 non-carriers⁵², as an estimate of the true association between *JAK2* 46/1 and V617F cases. Setting the rs77375493 allele intensity threshold to achieve this odds ratio resulted in a V617F case prevalence of 0.27% within MVP, which was comparable to the previously reported population prevalence of ~0.2%⁵². The slightly higher point estimate may reflect the fact that the MVP cohort is predominantly male and older than other population studies, both of which have been associated with increased rates of V617F⁵⁶.

We also attempted to replicate within MVP using an MPN definition based on Phecodes (<https://phewascatalog.org/phecodes>) and ICD9 codes: polycythemia vera (Phecode 200.1), chronic myeloid leukemia (Phecode 204.22), essential thrombocythemia (ICD9 238.71), and myelofibrosis (Phecode 289.1, ICD9 238.76). The replication rate using this definition was substantially lower. One reason for this could be winner's curse bias. However, even odds ratios for known MPN risk loci showed consistently weaker effect sizes compared to all three discovery cohorts, whereas the JAK2 V617F phenotype produced effect sizes that were closer to the three discovery cohorts (Extended Data Fig. 2). To quantify this discrepancy, we calculated the M statistic, a measure of between-cohort heterogeneity that combines information across independent lead variants to reveal systematic patterns of heterogeneity⁵⁷. Across the 15 genome-wide significant lead variants from the discovery analysis, the MVP cohort exhibited significant heterogeneity using the ICD-coded MPN definition ($M = -1.26$, Bonferroni $p = 2.81 \times 10^{-9}$), but not when using the JAK2 V617F definition ($M = -0.45$, Bonferroni $p = 0.32$). This heterogeneity is unlikely to be explained by differences in cohort characteristics, as all cohorts were restricted to European ancestry and controlled for age, sex, and ancestry-informed principal components. Moreover, the

heterogeneity was resolved upon changing the phenotype to a more objective measure of a pre-MPN state (*JAK2* V617F). These reasons led us to suspect the presence of spurious ICD-based phenotype designations for MPN within the MVP cohort. Thus, we chose to use the *JAK2* V617F carrier as a proxy for a “pre-MPN” state for final replication. Notably, the use of *JAK2* V617F for replication of MPN phenotypes has also been done in a previous GWAS⁵². Logistic regression for each replication variant was performed using the PLINK2-glm function, with age, sex, and the top 5 principal components included as covariates. Inverse-variance weighted meta-analysis was used to compute joint p values combining the discovery meta-analysis and MVP replication associations.

Polygenic risk score analysis

We trained a polygenic risk score (PRS) on a meta-analysis of the 23andMe and FinnGen cohorts, which included 1,863 cases of MPN and 428,399 controls. To determine the risk variants to be used in the PRS, we performed a pruning and thresholding analysis using the LD clumping method in PLINK version 1.90⁵⁸ (--clump) with an r^2 threshold of 0.2 and a p-value threshold of 1×10^{-5} . In brief, this algorithm forms clumps around variants with association *P* values less than the specified threshold, in which each clump contains all variants within 250 kilobases of the lead variant that are correlated with the lead variant ($r^2 > 0.2$), based on an LD reference panel of 6,000 randomly selected European individuals from the UKBB. The final output was a PRS containing 104 independent ($r^2 < 0.2$), MPN-associated ($P < 1 \times 10^{-5}$) variants representing the strongest risk variants for each LD clump across the genome. We performed a sensitivity analysis by testing different p-value thresholds for variant inclusion into the PRS. Across a broad range of thresholds, the area under the receiver operating characteristic (AUROC) for the PRS was similar (Supplementary Table 4).

We applied the PRS to the UKBB, an out-of-sample test set containing 1,086 cases of MPN and 401,155 controls. The PRS was computed for each individual by multiplying the genotype dosage of each risk allele for each variant by its association estimate beta (log-odds) from the UKBB summary statistics as a weight. This was performed using the PLINK2 --score function.

We modeled the PRS using a logistic regression with MPN case-control status as the phenotype and PRS, age, sex, top 10 principal components of ancestry, and genotyping array as covariates. We calculated the area under the receiver-operator curve (AUROC) for the model using the pROC R package. The proportion of variance explained was calculated by using the Nagelkerke’s pseudo- R^2 metric. We used this metric to calculate the incremental R^2 , which quantifies the gain in R^2 when a variable is added to a logistic regression of MPN case-control status on a set of other covariates (sex, age, genotyping array, 10 principal components of ancestry). To use PRS to stratify individual risk for MPNs, we divided individuals in the UKBB cohort into deciles based on their PRS. For each non-reference decile, logistic models were fitted with the same covariates as used above, comparing MPN risk for members of the given decile compared to those in the 5th decile (i.e., those with average risk) and lowest decile (i.e., those with the lowest 10% of PRS). We also performed a logistic regression for MPN risk while stratifying on individual carrier status for *JAK2*

46/1, a major risk haplotype for MPNs, and PRS category (low = bottom quintile of PRS, intermediate = quintiles 2–4, high = top quintile).

Genetic fine-mapping

For each distinct association signal, we calculated approximate Bayes' factors (ABFs)⁵⁹ for all variants within 1-Mb of the lead variant. ABFs were calculated as:

$$ABF = \sqrt{1-r} * e^{\frac{rz^2}{2}}$$

where $r = \omega / (s.e.^2 + \omega)$ and $z = \beta / s.e.$ The parameter ω denotes the prior variance in allelic effects, estimated here as 0.035 based on formula (8) of the original publication of Wakefield's formula and the 95% interval of variant effect sizes in the GWAS. For loci with multiple distinct signals, association statistics were based on GCTA approximate conditional analysis adjusting for all other index variants in the region. We then calculated the posterior probability of being causal (PP) by dividing the ABF of each variant by the sum of ABF values over all variants in the locus. The 95% credible set for each locus was constructed by 1) ranking all variants in descending order of PP and 2) including ordered variants until the cumulative PP reached 95%.

Contribution of variants to overall familial relative risk

We estimated the proportion of the familial risk of MPNs that can be explained by variants identified in our GWAS under a log-additive model, as previously described⁶⁰. We applied the formula $\lambda_g = \sum_i p_i (1 - p_i) (\beta_i^2 - \tau_i^2) / \ln \lambda$, where λ_g is the proportion of familial risk explained, p_i is the MAF for variant i , β_i is the log(odds ratio) estimate for variant i , τ_i is the standard error of β_i , and λ is the overall familial relative risk. We assumed the overall familial relative risk for MPNs to be 4.93 based on a recent epidemiological study⁶¹.

Definition of known loci

We compiled a list of 10 previously reported genome-wide significant MPN association signals from literature^{52,62}. Loci were only included if they were identified using a similar MPN phenotype (i.e., combination of JAK2 V617F carriers and all MPN subtypes, without any sub-stratification of MPN subtypes). All loci reaching genome-wide significance ($p < 5 \times 10^{-8}$) before or after replication steps were included.

g-chromVAR cell type enrichment analysis

Bias-corrected enrichment of MPN risk variants for chromatin accessibility of 18 hematopoietic populations was performed using g-chromVAR (<https://github.com/caleblareau/gchromVAR>), whose methodology has been previously described⁶³. In brief, this method weights chromatin peaks by fine-mapped variant posterior probabilities and computes the enrichment for each cell type versus an empirical background matched for GC content and feature intensity. For the chromatin accessibility component, we used a consensus peak set for all 18 hematopoietic cell types with a uniform width of 250 bp

centered at the summit. For the variant scores, we used the fine-mapped PP for all MPN risk variants with fine-mapped PP > 0.001 across suggestive loci.

Linkage disequilibrium score regression

We used LD score regression (LDSC) to estimate the narrow-sense heritability estimate of MPN risk and compute genetic correlations between MPN risk and other phenotypes⁶⁴. Reference LD scores were computed with a subset of European individuals combined from the 1000 Genomes Phase 3 (1000GP3) and UK10K cohorts. Variants were filtered by MAF > 1%, and 5,653,963 variants were used as input to LDSC. We estimated the liability scale heritability of MPN to be ~6.5% (s.e. = 2.98%) based on common genetic variation (MAF > 1%). For this estimation, the sample prevalence of MPNs was 0.00352 in our GWAS, and the population prevalence of MPNs was estimated to be 0.000328 based on previous reports^{65,66}.

To calculate genetic correlations with blood traits, we first used BOLT-LMM to perform GWAS on 19 blood traits in 408,241 European ancestry individuals from the UKBB, the same samples used for the MPN GWAS. Imputation and variant quality control filters were the same as those applied in the MPN GWAS. To calculate cross-trait correlations, we used the same 1000GP3-UK10K reference panel used to estimate LDSC heritability. We constrained the intercept by accounting for the known sample overlap between the MPN and blood trait GWAS, as well as adjusting for phenotypic correlations between MPN case-control status and each of the 19 blood traits. To calculate genetic correlations with telomere length and CHIP, we obtained previously generated summary statistics from recent studies^{67,68} and calculated genetic correlations as described above, but without constraining the intercept.

To calculate cell type enrichments, we generated LD scores for 18 primary hematopoietic ATAC peak sets. We also generated a “pan-heme” peak set representing the union of peaks across all 18 populations. Adopting the approach previously used for LDSC tissue-specific enrichments⁶⁹, we jointly modeled the annotation for each cell type of interest, the “pan-heme” annotation for all hematopoietic peaks, as well as the 52 annotations in the baseline model.

Blood trait pleiotropy analysis

We tested whether fine-mapped MPN risk variants were more likely to demonstrate pleiotropic associations with common blood traits from distinct lineages. To do this, we performed fine-mapping on all genome-wide significant regions for each of 18 lineage-specific blood traits (all except white blood cell count) using FINEMAP v1.3.1⁷⁰. We then assigned each blood trait to a major hematopoietic lineage: basophil count to basophils; eosinophil count to eosinophils; neutrophil count to neutrophils; red blood cell count, hematocrit, hemoglobin, RDW, MCH, MCHC, MCV, reticulocyte count, and mean reticulocyte volume to red blood cells; platelet count, MPV, platelet crit, platelet distribution width to platelets; monocyte count to monocytes; and lymphocyte count to lymphocytes. We considered a variant to be pleiotropic if it had a fine-mapped PP > 0.10 for blood traits from multiple lineages. Then, we constructed a contingency table for the proportion of MPN fine-

mapped variants ($PP > 0.10$ vs. $PP < 0.10$) which were classified as pleiotropic vs. not pleiotropic, and calculated an enrichment using a two-sided Fisher's exact test.

Phenome-wide association study

To identify associations between MPN risk variants and other clinical phenotypes, we conducted a phenome-wide association study (PheWAS) using summary statistics of 1,403 ICD-based clinical phenotypes analyzed from the UK Biobank (see URLs). As input, we included all fine-mapped MPN risk variants with $PP > 0.001$. To reduce noise from phenotypes with low case numbers, we restricted our analysis to only variant-phenotype combinations with an expected case minor allele count > 50 (calculated as $2 * \text{variant MAF} * \text{number of cases}$), resulting in 1,130 unique remaining phenotypes. Variants in the HLA region were excluded. We used a Bonferroni-corrected p-value threshold to define significant pheWAS associations ($0.05 / 1130 \text{ phenotypes} = 4.42 \times 10^{-5}$).

Comparison with acute myeloid leukemia risk

We examined the effects of the 17 lead variants ($p < 5 \times 10^{-8}$) from the MPN GWAS on inherited risk for acute myeloid leukemia (AML). To do this, we obtained AML association statistics for these 17 variants from three independent cohorts: UK Biobank, FinnGen, and Alliance for Clinical Trials in Oncology (Alliance). The AML GWAS for UK Biobank and FinnGen were conducted using the same quality control metrics and covariates as were used for the MPN analysis. Association statistics for the 17 variants from Alliance were obtained from a previous study, whose analysis has been described in-depth elsewhere⁷¹. We then meta-analyzed these statistics using a fixed effects model with inverse-variance weighting of log(odds ratios), as implemented in the METAL software⁵³.

Transcription factor motif analysis

Prediction of the effects of fine-mapped variants on transcription factor binding sites (TFBS) was performed using the motifbreakR R package⁷² and a comprehensive collection of 426 human TFBS models (HOCOMOCO⁷³). For all fine-mapped variants with $PP > 0.001$, we applied the 'information content' scoring algorithm and used a P -value cutoff of 1×10^{-3} for TFBS matches; all other parameters were kept at default settings.

Mendelian randomization

Mendelian randomization (MR) analysis was performed using the R packages TwoSampleMR⁷⁴ and MRPRESSO⁷⁵. MR consists of two steps: (i) identification of proper instrumental variables or genetic predictors, i.e., variants independently associated with the exposure factor, and (ii) calculation of causal estimates⁷⁶. To achieve step 1, we first obtained GWAS summary statistics for leukocyte telomere length. We filtered for variants associated with leukocyte telomere length at a minimum $p < 1 \times 10^{-5}$, and then clumped these variants with an LD threshold of $r^2 < 0.001$ to obtain 38 independent genetic instruments.

Next, we extracted the MPN risk effect sizes of these genetic instruments from our MPN GWAS and harmonized the data to ensure the variant statistics for telomere length and MPN were oriented to the same allele. To calculate causal estimates of telomere length on MPN

risk, we implemented four non-independent methods with varying assumptions regarding horizontal pleiotropy: MR-Egger regression with bootstrap⁷⁷, the IVW method, the weighted median test, and MR-PRESSO. We adopted a conservative multiple-testing-adjusted threshold of $p < 6.25 \times 10^{-3}$ ($0.05/(4*2)$) to account for the use of four tests and testing bidirectionally. All tests found a significant causal relationship between telomere length and MPN risk: IVW, $p = 1.36 \times 10^{-4}$; outlier-corrected MR-Presso, $p = 1.05 \times 10^{-5}$; MR-Egger, $p = 4.83 \times 10^{-5}$; weighted-median, $p = 1.15 \times 10^{-5}$.

We also tested the reverse association to assess causal estimates of MPN risk on telomere length. The same parameters were used, except this time we first clumped MPN risk variants ($p < 1 \times 10^{-5}$), and then extracted the corresponding effect sizes for telomere length. The MR test statistics for the reverse association were the following: IVW, $p = 5.78 \times 10^{-3}$; outlier-corrected MR-Presso, $p = 0.057$; MR-Egger, $p = 0.058$; weighted-median, $p = 0.070$.

Target gene identification

We implemented a two-stage process to identify high-confidence target genes at MPN risk loci. In all stages, the HLA locus (chromosome 6:28866528–33775446) was excluded due to its complex linkage structure. In the first stage, we checked whether any fine-mapped risk variants ($PP > 0.10$) resulted in coding consequences or splice alterations. To check for splice variants, we annotated variants with spliceAI⁷⁸, a neural net prediction tool for splice altering variants. No variants had a delta score > 0.2 , indicating a low probability for any variant to cause splicing changes. We used Variant Effect Predictor⁷⁹ to screen for coding variants, which identified three variants in distinct loci with $PP > 0.10$ causing missense mutations: rs17879961 for *CHEK2*, rs1800057 for *ATM*, and rs3184504 for *SH2B3*. These three regions were mapped to these respective target genes and were not analyzed further using non-coding variant approaches, described below.

The second stage consisted of mapping noncoding regions to target genes. Because non-coding gene regulation is more difficult to pinpoint, we applied a lower variant inclusion threshold of $PP > 0.01$ as a high-sensitivity screen for target gene interactions. To increase specificity for high-confidence target genes, we incorporated five different functional annotations for evidence of noncoding gene regulation: 1) gene body colocalization, 2) mapping to hematopoietic promoter capture Hi-C (PCHi-C)⁸⁰, 3) correlation between chromatin accessibility and nearby gene expression across hematopoietic cell types, 4) MAGMA gene-wise association signal, and 5) overlap with a recurrent somatically mutated gene in MPNs.

To map variants to gene bodies, we used gene annotation coordinates from GENCODE release 33⁸¹ for the GRCh37 genome build. We removed all ribosomal protein genes by excluding any genes with names starting with “RP”. We then identified the nearest genes to risk variants using the nearest command in the GenomicRanges R package.

For nominating target genes by enhancer promoter interactions, we used a published PCHi-C dataset spanning 15 hematopoietic cell types⁸⁰. We filtered for looping interactions with a CHiCAGO score > 5 . If multiple gene targets were nominated for one variant, only the gene with the top CHiCAGO score was kept.

ATAC-RNA correlations were generated by computing Pearson correlations between hematopoietic ATAC peaks and RNA counts of genes within a 1-Mb window of the ATAC peak, as previously described⁶³.

We performed gene-based associations for 18,987 protein-coding genes using MAGMA⁸², implemented through FUMA⁸³. In brief, this test provides aggregate association statistics based on all variants located in a gene, adjusting for LD. Default MAGMA parameters were used, which mapped variants to genes with no window around genes (window size = 0). A gene was considered positive by MAGMA if it passed the genome-wide significance of $p = 0.05/18987 = 2.63 \times 10^{-6}$.

For overlap with genes with known somatic mutations in individuals with MPN, we used all genes identified with driver mutations in at least 5 patients from a recent large-scale study which performed targeted sequencing on a cohort of 2,035 patients with MPN⁸⁴.

Collectively, non-coding target genes were selected if they scored in two or more criteria and were also the highest scoring gene in the risk locus. Finally, target genes from the coding and non-coding annotations were combined and filtered for protein-coding genes, as annotated by Ensembl using the annotables R package.

Target gene cell type enrichments

To measure the enrichment of MPN target genes in bulk RNA-seq data in 16 hematopoietic populations, we performed a rank-sum permutation test. First, we summed the ranks of all target genes ordered by expression (log₂ counts per million) amongst all assayed protein-coding genes with non-zero expression in each cell type. Next, we randomly sampled 10,000 equally sized gene sets and obtained their rank sums within each cell type. We calculated the target gene enrichment z-score as the difference between the mean rank-sum of the permuted sets and the target gene rank sum, divided by the standard deviation of the permuted gene set rank-sums. The z-score was then converted into a two-sided p-value for each cell type.

Gene set enrichments

We used g:Profiler⁸⁵ to map putative target genes to enriched gene sets. All protein-coding genes with one or more annotations were used as the background.

Single-cell RNA sequencing analysis

Single-cell RNA-seq of 378,000 cells from human bone marrow were generated as part of the Human Cell Atlas project using 10X Genomics sequencing technology and aligned to the GRCh38 reference genome using the Cell Ranger pipeline as previously described (<https://preview.data.humancellatlas.org/>). Downstream analyses including normalization, scaling, and cell clustering were performed using the R software package Seurat⁸⁶ version 2 (<http://satijalab.org/seurat/>). We filtered out low-expressed genes expressed in fewer than 50 cells and low-quality cells with fewer than 500 detected genes, leaving 19,156 genes and 278,978 cells for downstream analysis. Raw gene expression counts of each cell were normalized over total counts and log transformed, and gene expression was then scaled to have a mean of 0 and variance of 1 across cells. We performed dimensionality reduction

using PCA, with the top 1000 most variable genes as input, and computed the top 50 principal components (PCs). To identify clusters of cells, we used the 'FindClusters' function from Seurat, which applies a shared nearest neighbor modularity optimization-based clustering algorithm to identify clusters based on their PCs (in this case, top 50). To infer the HSC population, we used a marker gene signature similar to one recently applied in a different scRNA-seq human hematopoiesis dataset⁸⁷ – *CD34*, *HLF*, and *CRHBP*. All 15 MPN target genes were detected in at least 50 cells and thus included in the aggregate MPN signature. To calculate scores based on specific gene sets (e.g., HSC marker genes, MPN target genes) for each cell, we calculated the average of the *Z*-normalized expression (across all cells) of each gene in the list. To adjust for dropout in single-cell data when estimating the correlation between MPN and HSC signatures, we applied a gene imputation approach called MAGIC⁸⁸ to infer missing transcripts in cells. To do so, gene expression values in all cells were normalized, dimensionally reduced and transformed by internal algorithms in MAGIC with the parameters: *n_pca_components* = 100, *t* = 6, *k* = 10, *alpha* = 15, *rescale_percent* = 99. Following imputation, marker gene expression was again used to calculate an MPN and HSC signature per cell, as described above, and a Spearman correlation of these scores was calculated.

Structural analysis

The X-ray crystallographic structure of the human CHEK2 protein was used to create a CHEK2 structural model. Structural superposition, analyses, and figures were rendered using PyMOL Molecular Graphics System v2.3.2⁸⁹.

Irradiation experiments

Umbilical cord blood Lineage negative (Lin⁻) and CD34⁺ cells were obtained using StemSep system according to the manufacturer's protocol (Stem Cell Technologies, Canada). Lin⁻CD34⁺ cells were sorted to obtain HSC (CD34⁺38^{-/low}CD45RA⁻CD90⁺), CMP (CD34⁺38⁺CD45RA⁻CD135⁺), GMP (CD34⁺38⁺CD45RA⁺CD135⁺), and MEP (CD34⁺38⁺CD45RA⁻CD135⁻) fractions. Then cells were resuspended in X-VIVO 10 (BioWhittaker, Walkersville, MD) medium supplemented with 1% BSA, SCF (100 ng/ml), FLT3L (100 ng/ml), TPO (15 ng/ml), G-CSF (10 ng/ml), and IL-6 (10 ng/ml) and incubated for 72–96 hours followed by irradiation with 3Gy. When indicated, cells were pre-treated with CHEK2i (CHEK2 Inhibitor II, 10uM final, Sigma 220486) or DMSO for 1hr prior to irradiation. Assessment of IR-induced cell death in the indicated populations 18hr post IR relied on double staining with Annexin and Sytox. IR-induced cell death was calculated to reduce the variability between CD34⁺ batches by subtracting the fraction of AnnexinV⁺Sytox⁺ cells scored in the untreated sample from the same fraction in the IR sample.

Viral constructs, human CD34⁺ transduction and long-term expansion assays

The following pLKO-puro lentiviral shRNA constructs from the RNAi consortium shRNA library (TRC, <https://portals.broadinstitute.org/gpp/public/>) were used: shCHEK2 (TRCN0000039946) and shControl (TRCN0000231746). This shRNA has been shown to induce knockdown of total and phosphorylated CHEK2 protein⁹⁰. Viral particles pseudotyped with VSV-G were prepared using transient transfection of 293T cells as described elsewhere⁹¹.

Lin⁻CD34⁺ cells were incubated with the indicated lentiviruses, at multiplicity of infection 50–100, in the X-VIVO 10 medium supplemented with 1% BSA, SCF (100 ng/ml), FLT3L (100 ng/ml), TPO (15 ng/ml), G-CSF (10 ng/ml), and IL-6 (10 ng/ml) for 16 hours followed by cell wash and medium replacement. Puromycin (500 ng/ml) was added to the infected cells two days post infection for the additional two days. At the end of puromycin selection CD34⁺ cells were seeded in the ex vivo expansion cultures as previously described⁹².

Briefly, CD34⁺ cells were plated in IMDM, 10% FCS (Sigma) supplemented with FLT3L (50 ng/ml), TPO (20 ng/ml), SCF (50 ng/ml), and IL-6 (10 ng/ml) at the density of 1×10^5 cells/ml. Every seven days, cells were counted, washed, and resuspended at the density of 1×10^5 cells/ml in fresh medium and cytokines.

Luciferase reporter assays

The genomic region containing risk and non-risk alleles of the variants rs524137 (~501 bp) and rs1633768 (~294 bp) were synthesized as gblocks (IDT Technologies) and cloned into the Firefly luciferase reporter constructs (pGL4.24) using NheI and EcoRV sites. Each allele-specific construct differed by only the single nucleotide of the variant of interest (i.e., all other nucleotides in the construct are the same). The Firefly constructs (500 ng) were co-transfected with pRL-SV40 Renilla luciferase constructs (50 ng) into 100,000 K562 cells using Lipofectamine LTX (Invitrogen) according to manufacturer's protocols. Cells were harvested after 48 hours and the luciferase activity measured by Dual-Glo Luciferase Assay system (Promega). For each sample, the ratio of firefly to Renilla luminescence was measured and normalized to the minimal promoter construct. Location of rs524137 and rs1633768 gblocks in hg19 are chr9:135879384–135879677 and chr9:135878924–135879424 respectively.

Lentiviral reporter assays

The lentiviral reporter constructs were designed to deliver enhancer elements containing risk and non-risk alleles of rs524137 positioned upstream of human *GFI1B* promoter (~400 bp) driving a reporter GFP. The constructs were generated from a pLKO based lentiviral construct overexpressing GFP from a PGK promoter. Fragments containing human *GFI1B* promoter and enhancer elements were synthesized as gblocks (IDT technologies) and cloned using Kfl1 and BamHI sites upstream of the GFP start site replacing the PGK promoter. Lentiviral supernatants produced by transient transfection in 293Ts were used to transduce 50,000 K562 cells by spinfection at 2000 rpm for 90 min. K562s infected with the lentiviral reporters were cultured for 6 days and reporter GFP expression was measured by flow cytometry. Location of *GFI1B* promoter gblock in hg19 is chr9:135853723–135854128.

Ribonucleoprotein (RNP) electroporation of CD34⁺ human HSPCs

Electroporation was performed using the Lonza 4D Nucleofector with 20 μ l Nucleocuvette Strips. For the *GFI1B* enhancer element deletion experiment, CD34⁺ HSPCs were thawed 24 hrs before electroporation. The RNP complex was made by mixing Cas9 (50 pmol) and modified sgRNAs from Synthego (100 pmol in total), including two pairs of guides ENH1 (sgRNA1: TAAGTCTGGGGTCTACAAAG & sgRNA2: ATGACTTGCTTAGAGCACCA) and ENH2 (sgRNA3: ATAGAAGACCACTTCTCGCA & sgRNA2: ATGACTTGCTTAGAGCACCA) targeting the five and three prime end of the enhancer

element (hg19 chr9:135878960–135880490). For negative control, a guide targeting AAVS1 site was used (GGGGCCACTAGGGACAGGAT). *GFI1B* gene expression and editing outcome of the electroporation were measured at 6 days post-electroporation. Quantitative PCR was performed using SYBR green (Bio-Rad) to access the editing efficiency of the target sequences in bulk and sorted primitive LT-HSC cell populations, with primers designed to detect the percentage of unedited wild-type cells (forward: TCTACCACTCCCAGCAGCT; reverse: CGTCTCCTCTCCTGGGTCTT). *GFI1B* RNA expression was measured using quantitative PCR with specific primers (forward: GCAGGAAGATGAACCGCTCT; reverse CCAGGCACTGGTTTGGGAA).

For the coding deletion experiment, CD34⁺ HSPCs were thawed 48 hrs before electroporation. The RNP complex was prepared by mixing Cas9 (50 pmol) and modified sgRNA from Synthego (100 pmol) and incubating for 15 min at room temperature immediately before electroporation. HSPCs (3.75×10^5) resuspended in 20 μ l P3 solution were mixed with RNP and transferred to a cuvette for electroporation with program DZ-100. The electroporated cells were resuspended with Stemspan II media with CC100 cytokine cocktail (Stem Cell Technologies). Two guides targeting *GFI1B* coding regions (sgRNA4: GGGGTCTGGGACAGCACAATG; sgRNA5: CCTTGTTGCACTTCACACAG) and control non-targeting guide (NT) was used in these experiments. Gfi-1b protein expression was measured at 5 days post-electroporation using anti-GFI1B antibody (Santa Cruz Biotechnology, sc-28356; 1:3000 dilution) and a loading control anti-Lamin B1 antibody (Santa Cruz Biotechnology, sc-374015; 1:1000 dilution).

HSC maintenance cultures

Enhancer edited HSPCs were cultured for 6 days post-editing in HSC maintenance conditions using serum free StemSpan II media (Stem Cell technologies) supplemented with CC100 cytokine cocktail (Stem Cell technologies), 50ng/ml TPO (Peprotech), and a small molecule UM171 (35 nM) shown to help maintain and potentially expand HSCs in culture^{93,94}. At the end of the culture, total cells were enumerated and stained with a panel of antibodies to quantify and enrich for LT-HSCs. The antibody panel includes anti-CD45RA-Alexa Fluor® 488 (Biolegend, #304114), anti-CD34-PerCP-CY5.5 (Biolegend, #343612), anti-CD90-PECy7 (BD, #561558), anti-CD133-superbright 436 (Ebioscience, #62-1388-42), anti-EPCR-PE (Biolegend, #351904), and anti-ITGA3-APC (Biolegend, #343808). 3 μ l of anti-CD34-PerCP-CY5.5, anti-CD90-PECy7, anti-EPCR-PE, anti-ITGA3-APC, and 6 μ l of anti-CD45RA-Alexa Fluor® 488, anti-CD133-superbright 436 were used per one hundred thousand cells in 100 μ l staining volume. LT-HSCs were enriched by FACS sorting using previously published methods⁹⁵. Both genomic DNA and RNA were extracted from sorted LT-HSCs and bulk HSPCs using AllPrep DNA/RNA Mini Kit (Qiagen) according manufacturer's recommendations. Total LT-HSC numbers were calculated as a product of frequency of LT-HSC by FACS and total cell numbers counted at the end of culture. Flow cytometry data were analyzed on FlowJo™ software (v10.6.1).

Colony-forming unit cell assays

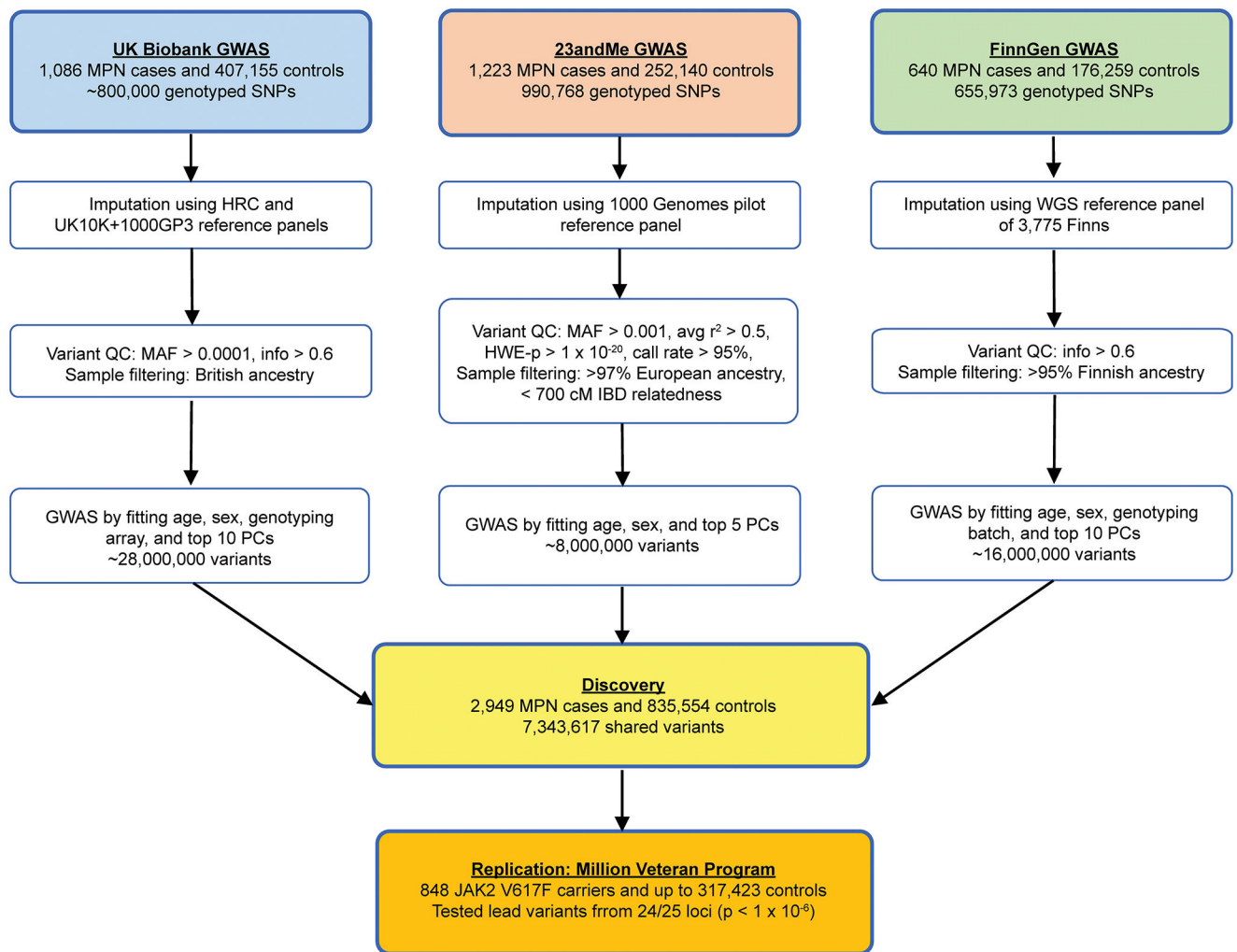
3 days RNP post-electroporation, 500 CD34⁺ HSPCs were plated in 1ml methylcellulose media (#H4034, Stem Cell Technologies). Primary CFU-C colonies were counted after 14

days. For the colony replating experiments, 2 weeks after the primary plating, the colonies from two plates were pooled, washed with PBS, and the cells were plated in new methylcellulose media at 25000 cells/ml for an additional 14 days. Images of primary and secondary colonies were taken using StemVision (Stem Cell technologies) using manufacturer’s recommendations.

URLs

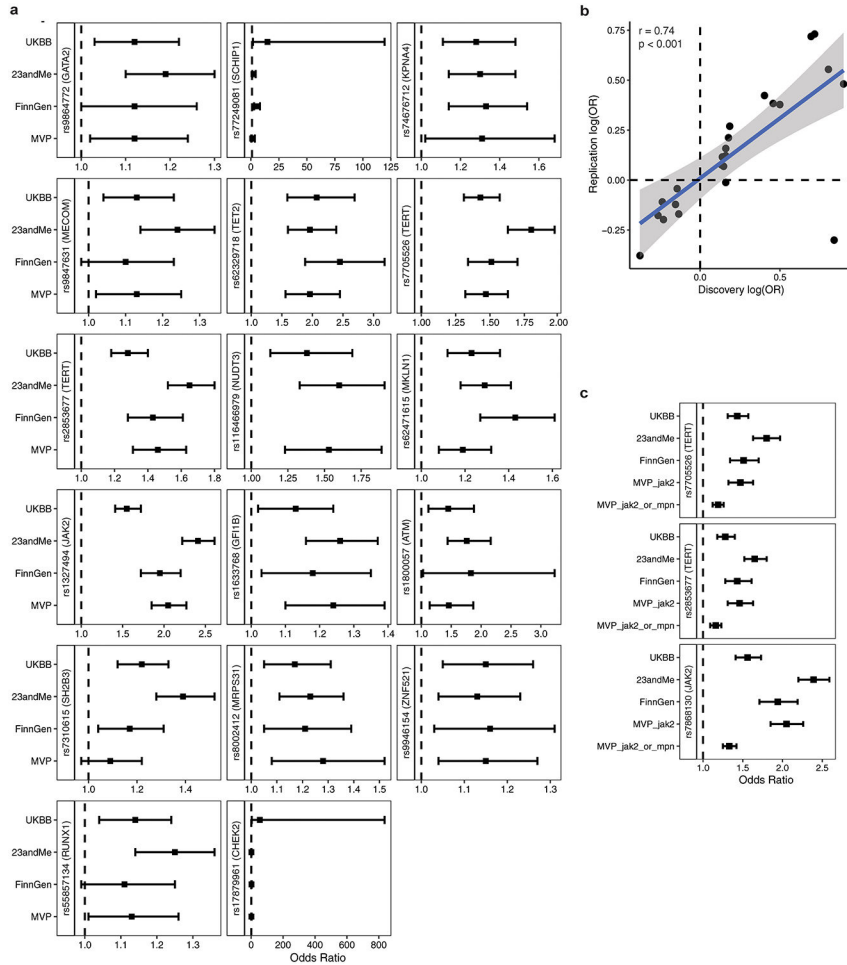
1000 Genomes Phase 3: <ftp://ftp.1000genomes.ebi.ac.uk/vol1/ftp/release/20130502/>;
 UK10K: https://www.uk10k.org/data_access.html; Haplotype Reference Consortium: <http://www.haplotype-reference-consortium.org/>; HOCOMOCO: <https://hocomoco11.autosome.ru/>; Protein Data Bank: https://www.rcsb.org/pdb/static.do?p=general_information/about_pdb/index.html; g-chromVAR: <https://github.com/caleblareau/gchromVAR>; UK Biobank phenome-wide association study: <https://www.leelabsg.org/resources>; Immune cell atlas: <https://preview.data.humancellatlas.org/>.

Extended Data



Extended Data Figure 1.

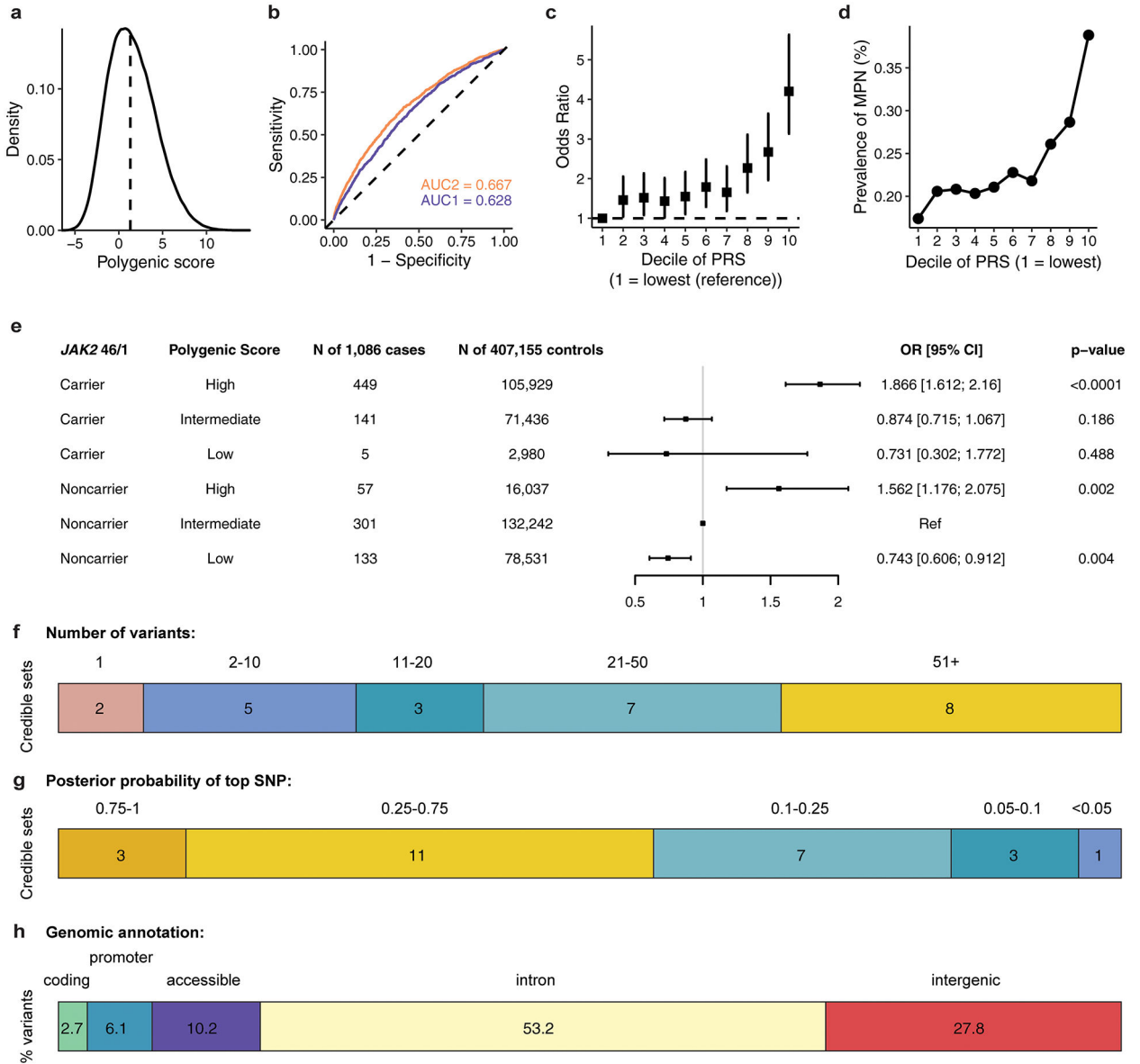
Flowchart of genetic association analyses. Flowchart of the quality control steps and analysis methods for the three discovery-phase genome-wide association studies (GWAS) in the UK Biobank, 23andMe, and FinnGen, followed by replication in the Million Veteran Program.



Extended Data Figure 2.

MPN GWAS cohort-specific effect sizes. **a**, Forest plot displaying cohort-specific odds ratios for lead variants of the 17 loci reaching genome-wide significance after replication. Sample sizes are: UKBB, n = 1,086 cases and 407,155 controls; 23andMe, n = 1,223 cases and 252,140 controls; FinnGen, n = 640 cases and 176,259 controls; MVP, n = 848 cases and 317,423 controls. Data represent odds ratios and 95% confidence intervals. **b**, Overall correlation of effect sizes between MVP cohort and combined discovery cohort (UKBB + 23andMe + FinnGen) for all 24 variants reaching suggestive significance ($p < 1 \times 10^{-6}$) which underwent replication ($p = 3.76 \times 10^{-5}$, two-tailed Pearson correlation). **c**, Forest plot displaying cohort-specific odds ratios for lead variants of the three most significant loci in the meta-analysis: the *JAK2* 46/1 haplotype and two independent signals at the *TERT* locus. MVP_jak2 = *JAK2* V617F phenotype in MVP, MVP_jak2_or_mpn = *JAK2* V617F or ICD-based MPN definition in MVP. Data are odds ratios and 95% confidence intervals. Sample sizes are: UKBB, n = 1,086 cases and 407,155 controls; 23andMe, n = 1,223 cases and

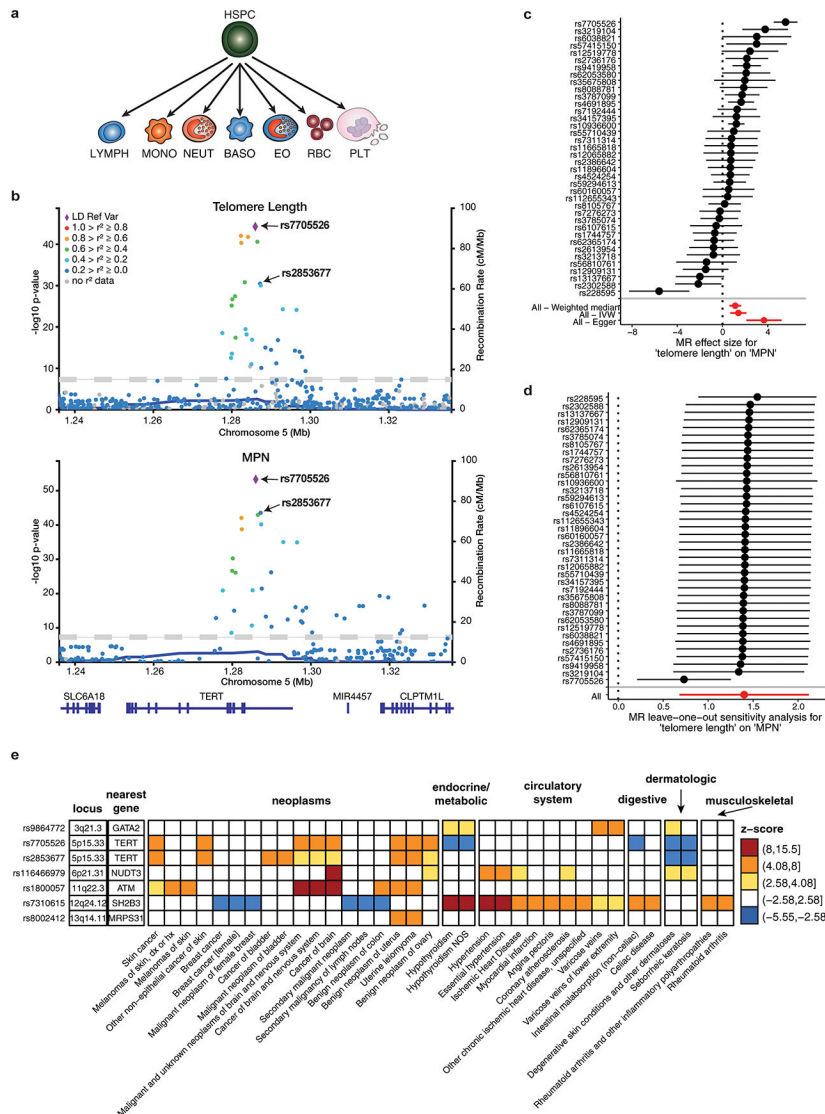
252,140 controls; FinnGen, n = 640 cases and 176,259 controls; MVP_jak2, n = 848 cases and 317,423 controls; MVP_jak2_or_mpn, n = 2,203 cases and 218,607 controls.



Extended Data Figure 3.

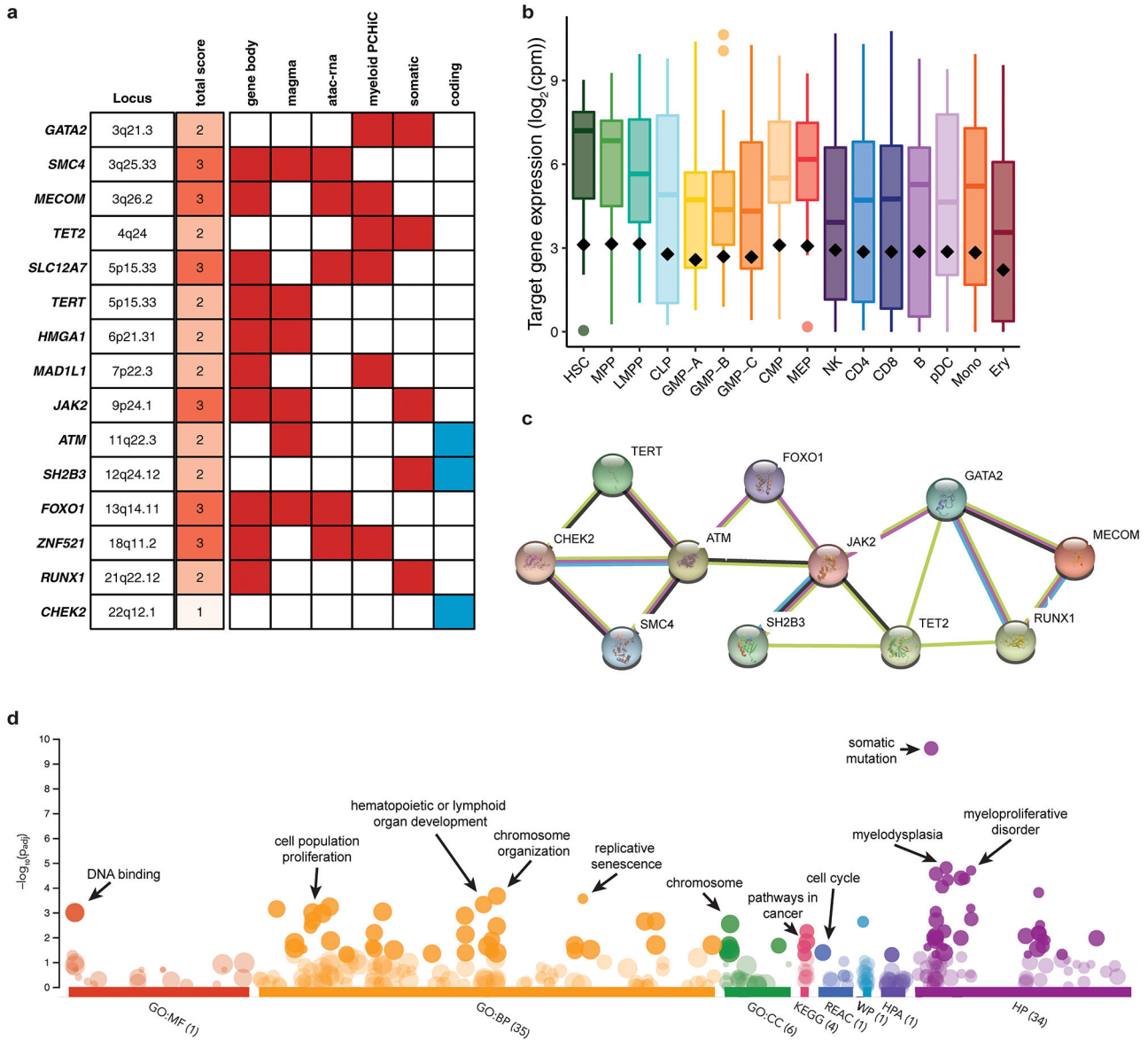
Assessing the distribution and prevalence of MPN polygenic risk score in UK Biobank. **a**, Density distribution of the MPN polygenic risk score (PRS) within the UK Biobank. **b**, Receiver operating characteristic curves for MPN predictions (n = 1,086 cases and 407,155 controls), using information from age, sex, genotyping array, and ancestry-informed principal components (AUC2, blue) alone, or with the addition of PRS (AUC1, orange). **c**, Odds ratio (mean and 95% confidence interval) for MPN acquisition according to deciles of the PRS (n = 1,086 cases and 407,155 controls), with decile 1 (10% of individuals with lowest PRS) as the reference group. **d**, Prevalence of MPN within each decile of the PRS in the UK Biobank population (n = 1,086 MPN cases, 407,155 controls). **e**, MPN cases and

controls in the UK Biobank were stratified into three groups according to their PRS – low, intermediate, or high defined as the lowest quintile, the middle three quintiles, and the highest quintile of the PRS distribution respectively. For carriers and noncarriers of the *JAK2* 46/1 haplotype, the odds ratio for MPN was calculated in a logistic regression model with PRS group, age, sex, and the top ten principal components of ancestry as covariates. Non-carriers with intermediate PRS served as the reference group. Data are odds ratios and 95% confidence intervals. **f**, Fine-mapped 95% credible sets for all 25 MPN risk loci reaching suggestive significance, stratified by the number of variants comprising each credible set. **g**, The fine-mapped posterior probability of causality for the highest fine-mapped variant in each locus credible set. **h**, Variants within the 95% credible sets and posterior probability (PP) > 0.001 across all regions, grouped by genomic annotation.

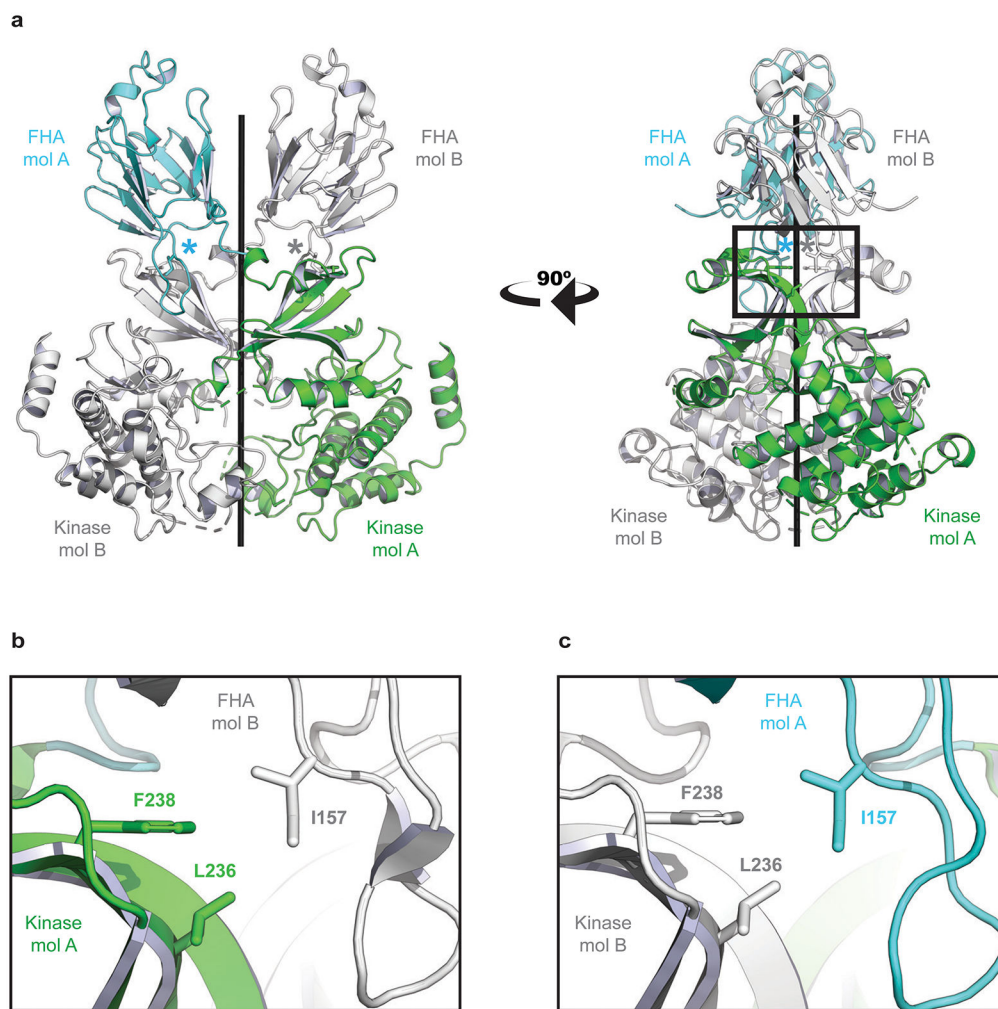


Extended Data Figure 4. Shared genetic associations between MPN risk and other phenotypes. **a**, Schematic depicting the trajectory of undifferentiated hematopoietic stem and progenitor cells (HSPCs) into

various committed cell types: lymphocytes (LYMPH), monocytes (MONO), neutrophils (NEUT), basophils (BASO), eosinophils (EO), red blood cells (RBC), and platelets (PLT). **b**, Regional association plots at the *TERT* locus (\pm 50kb from lead variant), showing the associations of variants with leukocyte telomere length and MPN. The colors of the points depict pairwise linkage disequilibrium (r^2) to sentinel variant rs7705526. The two conditionally independent lead variants for both traits, rs7705526 and rs2853677, are labeled. **c**, Individual SNPs associated with telomere length and their effect sizes on MPN risk (n = 2,949 cases and 835,554 controls), calculated using the fixed effects meta-analysis method. Aggregate mendelian randomization (MR) effects, calculated from three different methods (weighted median, inverse-variance weighted, and Egger regression), are shown at the bottom. Data are MR effect sizes and standard errors. Red color indicates significance. **d**, MR leave-one-out sensitivity analysis, showing MR effect estimates using the inverse variance weighted approach after excluding each individual SNP from the analysis (n = 2,949 cases and 835,554 controls). Data are MR effect sizes and standard errors. **e**, Phenome-wide association study (pheWAS) of MPN risk variants. We tested fine-mapped MPN risk variants (PP > 0.10 or lead variant) for associations with 1,130 well-represented case-control phenotypes from the UK Biobank, calculated by two-tailed logistic mixed model association test. Shown in this heatmap are the top MPN-associated variants at each locus with one or more associations reaching Bonferroni-corrected significance ($p = 0.05 / 1130$ phenotypes = 4.4×10^{-5} , or $\text{abs}(z\text{-score}) = 4.08$). Heatmap color indicates association z-score. All variant effects are oriented with respect to the risk-increasing MPN allele. Phenotypes are divided into major clinical categories, as listed in the annotations above the heatmap.

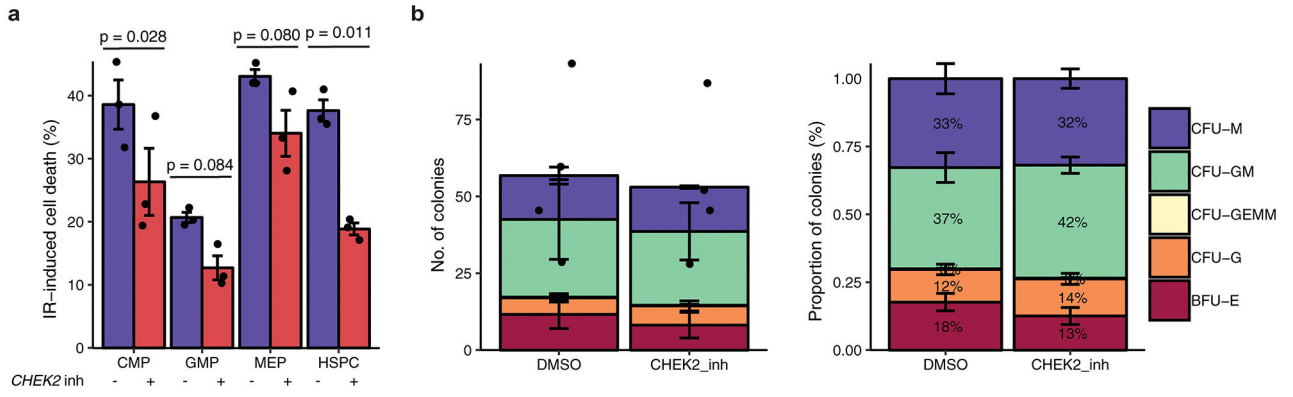


Extended Data Figure 5. Characterizing MPN target genes. **a**, Target genes prioritized based on non-coding criteria (red boxes) and coding consequences (blue boxes) and scored based on the number of criteria met. Only the highest scoring gene per locus is reported, and for non-coding loci, only genes with a score of 2 or more are reported. **b**, Average expression (\log_2 counts per million) of MPN target genes ($n = 15$) across 16 primary hematopoietic cell types. Black diamonds indicate the mean expression of all non-zero expressed protein-coding genes in each cell type. Box plots show the median at the center, with the top and bottom of the box indicating the interquartile range. Whiskers extend to either the maximum/minimum value or 1.5x the interquartile range. **c**, Protein-protein interaction network showing known and predicted associations between the protein products of MPN target genes, generated with STRING database. **d**, Top-enriched biological annotations for MPN target genes identify key pathways associated with hematopoiesis and oncogenesis.



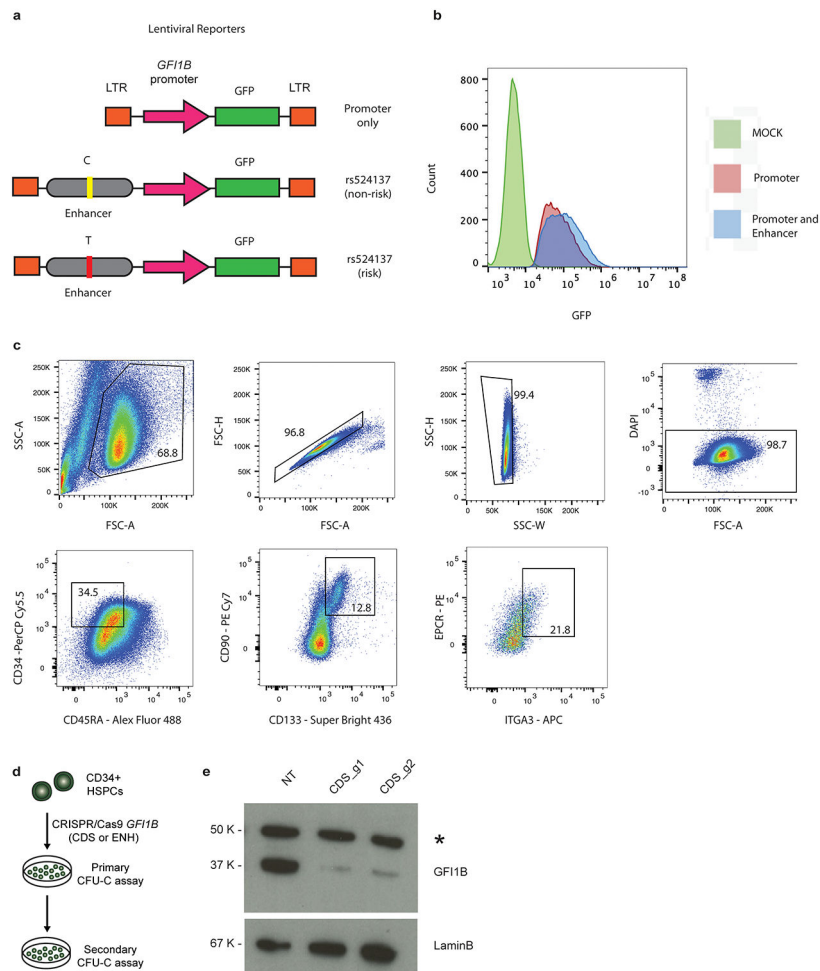
Extended Data Figure 6.

Structural basis for CHK2 homodimer disruption by Isoleucine 157 mutation. **a**, The crystal structure of the CHK2 (FHA-Kinase) homodimer (PDB: 3I6U). The FHA domain of molecule A (mol A) is shown in cyan and the kinase domain is colored green. A second CHK2 (mol B) has both domains colored white. The two CHK2 molecules are nearly symmetric – coiling around the central axis (black rod). The location of each Isoleucine 157 residue is marked with an asterisk. **b**, A zoomed window showing details of the interactions. I157 links the FHA of one CHK2 molecule (white) to the kinase domain of a second (green). The side chain of I157 mediates an FHA-Kinase hydrophobic interface, interacting with Phenylalanine 238 (F238) and Leucine 236 (L236) on the kinase domain. **c**, The second interface of the CHK2 dimer (180° rotation from panel **b**) is nearly identical. A Threonine at position 157 would diminish these hydrophobic interfaces and destabilize the CHK2 dimer, as has been previously reported.

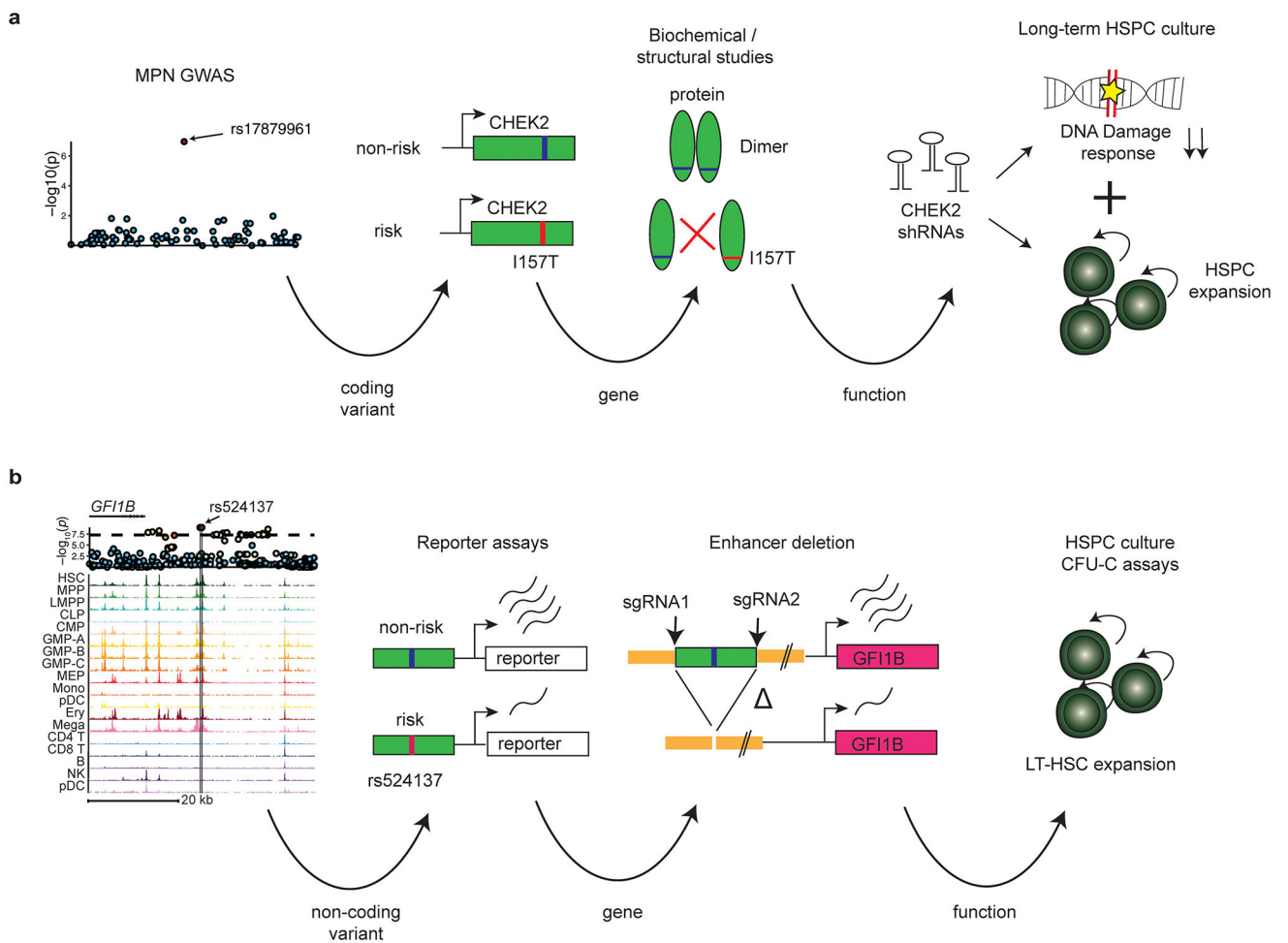


Extended Data Figure 7.

CHEK2 is required for apoptosis of cycling HSPCs, but not for lineage commitment. **a**, Assessment of IR-induced cell death of cycling HSPCs and myeloid progenitors (CMP, common myeloid progenitor; GMP, granulocyte-monocyte progenitor; MEP, megakaryocyte-erythroid progenitor) following sublethal irradiation, after treatment with *CHEK2* inhibitor (n = 3) or dimethylsulfoxide control (n = 3) (two-sided paired t-test). n is the number of biologically independent experiments. Data are mean ± s.e.m. **b**, Numbers (left) and percent (right) of HSPC colonies formed following *CHEK2* inhibition (*CHEK2* inhibitor II, Sigma 220486) (n = 4) vs. dimethylsulfoxide (DMSO) control (n = 4). n is the number of biologically independent experiments. Data are mean ± s.e.m. CFU-M, colony forming unit-macrophage; CFU-GM, granulocyte macrophage; CFU-GEMM, granulocyte erythrocyte macrophage megakaryocyte; CFU-G, granulocyte; BFU-E, burst forming unit-erythroid.



Extended Data Figure 8. Supplementary data for variant-to-function studies at *GFI1B* locus. **a**, Map of the lentiviral constructs designed to assess enhancer activity at rs524137. **b**, Histogram displays GFP mean fluorescence intensity (MFI) of hematopoietic K562 cells infected with Promoter only vs Promoter and Enhancer lentiviral constructs. Compared to mock uninfected control cells, cells infected with the construct carrying both *GFI1B* promoter and enhancer show greater GFP intensity. **c**, FACS gating for sorting and identifying the primitive CD34⁺CD45RA⁻CD90⁺CD133⁺EPCR⁺ITGA3⁺ LT-HSC population in day 7 CD34⁺ HSPCs presented in Fig. 4g, h, i. **d**, Schematic of colony-replating assays using human HSPCs edited with *GFI1B* coding (CDS) and enhancer guides (ENH). **e**, Representative western blot measuring *GFI1B* protein expression 5 days following CRISPR/Cas9 targeting with non-targeting control (NT), or coding regions of *GFI1B* (g1, g2). LaminB expression used as loading control. LaminB controls was probed on the same blot as the *GFI1B*. Similar results were obtained in 3 independent experiments. For gel source data, see Supplementary Fig. 3.



Extended Data Figure 9. Schematics illustrating the variant-to-function arcs for MPN risk loci at (a) *CHEK2* and (b) *GFI1B* demonstrated in this study.

Extended Data Table 1.

Genome-wide-significant loci from MPN GWAS. Variants shown are the most associated variant at each locus.

locus	SNP	bp	risk	non-risk	RAF	OR	OR 95CI	pvalue	meta pvalue	Nearest gene	Other genes
3q21.3	rs9864772*	128316939	G	A	0.6075	1.15	1.09–1.21	2.74E-07	2.06E-08	<i>GATA2</i>	
3q25.33	rs77249081*	159633461	G	C	0.0096	3.7	2.49–5.49	8.04E-11	5.54E-10	<i>SCHIP1</i>	
3q25.33	rs74676712*	160284736	T	C	0.1138	1.3	1.2–1.41	3.57E-10	3.64E-11	<i>KPNA4</i>	
3q26.2	rs9847631	168832107	T	G	0.3979	1.17	1.11–1.23	7.06E-09	4.89E-10	<i>MECOM</i>	
4q24	rs62329718	105758059	A	T	0.0374	2.11	1.84–2.42	7.46E-27	2.72E-34	<i>TET2</i>	

locus	SNP	bp	risk	non-risk	RAF	OR	OR 95CI	pvalue	meta pvalue	Nearest gene	Other genes
5p15.33	rs7705526	1285974	A	C	0.3358	1.58	1.49–1.67	4.78E-54	2.42E-64	<i>TERT</i>	
5p15.33	rs2853677	1287194	G	A	0.4176	1.46	1.38–1.54	2.79E-44	4.32E-54	<i>TERT</i>	
6p21.31	rs116466979*	34235378	C	T	0.0453	1.5	1.31–1.71	3.34E-09	1.86E-12	<i>NUDT3</i>	<i>HMGA1</i>
7q32.3	rs62471615	130746955	C	A	0.3037	1.3	1.23–1.38	6.55E-19	7.20E-21	<i>MKLN1</i>	
9p24.1	rs1327494	4999303	G	A	0.2735	2	1.89–2.12	3.07E-128	1.11E-170	<i>JAK2</i>	
9q34.13	rs1633768	135879138	T	C	0.2708	1.2	1.13–1.27	1.03E-09	2.15E-12	<i>GFIIB</i>	
11q22.3	rs1800057	108143456	G	C	0.0255	1.65	1.41–1.92	2.00E-10	2.94E-12	<i>ATM</i>	
12q24.12	rs7310615	111865049	C	G	0.4768	1.27	1.21–1.34	3.05E-19	2.46E-18	<i>SH2B3</i>	<i>ATXN2</i>
13q14.11	rs8002412*	41331497	C	T	0.1803	1.2	1.13–1.29	2.59E-08	5.23E-10	<i>MRPS31</i>	
18q11.2	rs9946154*	22810619	T	C	0.6436	1.15	1.09–1.21	6.95E-07	1.50E-08	<i>ZNF521</i>	
21q22.12	rs55857134*	36347627	C	T	0.3353	1.17	1.11–1.24	1.63E-08	1.93E-09	<i>RUNX1</i>	
22q12.1	rs17879961	29121087	G	A	0.0221	2.23	1.66–3	1.12E-07	3.60E-08	<i>CHEK2</i>	<i>HSCB</i>

Novel associations are denoted with an asterisk after the SNP. Coordinates based on hg19 genome build. Alleles are on the + strand. Locus, chromosome band and locus; SNP, variant RSID identifier; bp, base position; risk, risk-increasing allele; non-risk, other allele; RAF, risk allele frequency; OR, odds ratio estimate for risk allele; OR 95CI, 95% confidence interval for OR; pvalue, discovery GWAS association p-value (two-tailed logistic regression); meta pval, joint discovery + replication association p-value (inverse-variance weighted meta-analysis); nearest gene, nearest gene to variant; other genes, additional genes located within 25 kb of the variant. SNPs marked with * indicate a novel locus.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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References

1. Sud A et al. Familial risks of acute myeloid leukemia, myelodysplastic syndromes, and myeloproliferative neoplasms. *Blood* 132, 973 (2018). [PubMed: 29991558]
2. Landgren O et al. Increased risks of polycythemia vera, essential thrombocythemia, and myelofibrosis among 24,577 first-degree relatives of 11,039 patients with myeloproliferative

- neoplasms in Sweden. *Blood* 112, 2199–2204, doi:10.1182/blood-2008-03-143602 (2008). [PubMed: 18451307]
3. Brewer HR, Jones ME, Schoemaker MJ, Ashworth A & Swerdlow AJ Family history and risk of breast cancer: an analysis accounting for family structure. *Breast cancer research and treatment* 165, 193–200, doi:10.1007/s10549-017-4325-2 (2017). [PubMed: 28578505]
 4. Albright F et al. Prostate cancer risk prediction based on complete prostate cancer family history. *The Prostate* 75, 390–398, doi:10.1002/pros.22925 (2014). [PubMed: 25408531]
 5. Johns LE & Houlston RS A systematic review and meta-analysis of familial colorectal cancer risk. *American Journal Of Gastroenterology* 96, 2992, doi:10.1111/j.1572-0241.2001.04677.x (2001). [PubMed: 11693338]
 6. Tapper W et al. Genetic variation at MECOM, TERT, JAK2 and HBS1L-MYB predisposes to myeloproliferative neoplasms. *Nat Commun* 6, 6691, doi:10.1038/ncomms7691 (2015). [PubMed: 25849990]
 7. Hinds DA et al. Germ line variants predispose to both JAK2 V617F clonal hematopoiesis and myeloproliferative neoplasms. *Blood* 128, 1121 (2016). [PubMed: 27365426]
 8. Bulik-Sullivan BK et al. LD Score regression distinguishes confounding from polygenicity in genome-wide association studies. *Nature Genetics* 47, 291 (2015). [PubMed: 25642630]
 9. Yang J et al. Conditional and joint multiple-SNP analysis of GWAS summary statistics identifies additional variants influencing complex traits. *Nature Genetics* 44, 369 (2012). [PubMed: 22426310]
 10. Jones AV et al. JAK2 haplotype is a major risk factor for the development of myeloproliferative neoplasms. *Nature Genetics* 41, 446 (2009). [PubMed: 19287382]
 11. Wakefield J A Bayesian Measure of the Probability of False Discovery in Genetic Epidemiology Studies. *The American Journal of Human Genetics* 81, 208–227, doi:10.1086/519024 (2007). [PubMed: 17668372]
 12. Ulirsch JC et al. Interrogation of human hematopoiesis at single-cell and single-variant resolution. *Nature Genetics* 51, 683–693, doi:10.1038/s41588-019-0362-6 (2019). [PubMed: 30858613]
 13. Kimura M et al. Synchrony of telomere length among hematopoietic cells. *Experimental hematology* 38, 854–859, doi:10.1016/j.exphem.2010.06.010 (2010). [PubMed: 20600576]
 14. Morrison SJ, Prowse KR, Ho P & Weissman IL Telomerase Activity in Hematopoietic Cells Is Associated with Self-Renewal Potential. *Immunity* 5, 207–216, doi:10.1016/S1074-7613(00)80316-7 (1996). [PubMed: 8808676]
 15. Yamaguchi H et al. Mutations in TERT, the Gene for Telomerase Reverse Transcriptase, in Aplastic Anemia. *New England Journal of Medicine* 352, 1413–1424, doi:10.1056/NEJMoa042980 (2005). [PubMed: 15814878]
 16. Li C et al. Genome-wide Association Analysis in Humans Links Nucleotide Metabolism to Leukocyte Telomere Length. *The American Journal of Human Genetics* 106, 389–404, doi:10.1016/j.ajhg.2020.02.006 (2020). [PubMed: 32109421]
 17. Zhou W et al. Efficiently controlling for case-control imbalance and sample relatedness in large-scale genetic association studies. *Nature Genetics* 50, 1335–1341, doi:10.1038/s41588-018-0184-y (2018). [PubMed: 30104761]
 18. Bick AG et al. Inherited Causes of Clonal Hematopoiesis of Indeterminate Potential in TOPMed Whole Genomes. *bioRxiv*, 782748, doi:10.1101/782748 (2019).
 19. Roth A et al. STRING v10: protein–protein interaction networks, integrated over the tree of life. *Nucleic Acids Research* 43, D447–D452, doi:10.1093/nar/gku1003 (2014). [PubMed: 25352553]
 20. Garrison BS et al. ZFP521 regulates murine hematopoietic stem cell function and facilitates MLL-AF9 leukemogenesis in mouse and human cells. *Blood* 130, 619–624, doi:10.1182/blood-2016-09-738591 (2017). [PubMed: 28615219]
 21. Rodrigues NP et al. Haploinsufficiency of GATA2 perturbs adult hematopoietic stem-cell homeostasis. *Blood* 106, 477, doi:10.1182/blood-2004-08-2989 (2005). [PubMed: 15811962]
 22. Kataoka K et al. Evi1 is essential for hematopoietic stem cell self-renewal, and its expression marks hematopoietic cells with long-term multilineage repopulating activity. *The Journal of Experimental Medicine* 208, 2403, doi:10.1084/jem.20110447 (2011). [PubMed: 22084405]

23. Tober J, Yzaguirre AD, Piwarzyk E & Speck NA Distinct temporal requirements for Runx1 in hematopoietic progenitors and stem cells. *Development* 140, 3765, doi:10.1242/dev.094961 (2013). [PubMed: 23924635]
24. Cabezas-Wallscheid N et al. Identification of Regulatory Networks in HSCs and Their Immediate Progeny via Integrated Proteome, Transcriptome, and DNA Methylome Analysis. *Cell Stem Cell* 15, 507–522, doi:10.1016/j.stem.2014.07.005 (2014). [PubMed: 25158935]
25. Ito K et al. Regulation of oxidative stress by ATM is required for self-renewal of haematopoietic stem cells. *Nature* 431, 997 (2004). [PubMed: 15496926]
26. Tothova Z et al. FoxOs Are Critical Mediators of Hematopoietic Stem Cell Resistance to Physiologic Oxidative Stress. *Cell* 128, 325–339, doi:10.1016/j.cell.2007.01.003 (2007). [PubMed: 17254970]
27. Moran-Crusio K et al. Tet2 Loss Leads to Increased Hematopoietic Stem Cell Self-Renewal and Myeloid Transformation. *Cancer Cell* 20, 11–24, doi:10.1016/j.ccr.2011.06.001 (2011). [PubMed: 21723200]
28. Akada H et al. Critical Role of Jak2 in the Maintenance and Function of Adult Hematopoietic Stem Cells. *STEM CELLS* 32, 1878–1889, doi:10.1002/stem.1711 (2014). [PubMed: 24677703]
29. Buza-Vidas N et al. Cytokines regulate postnatal hematopoietic stem cell expansion: opposing roles of thrombopoietin and LNK. *Genes & Development* 20, 2018–2023 (2006).
30. Seita J et al. Lnk negatively regulates self-renewal of hematopoietic stem cells by modifying thrombopoietin-mediated signal transduction. *Proceedings of the National Academy of Sciences* 104, 2349, doi:10.1073/pnas.0606238104 (2007).
31. Allsopp RC, Morin GB, DePinho R, Harley CB & Weissman IL Telomerase is required to slow telomere shortening and extend replicative lifespan of HSCs during serial transplantation. *Blood* 102, 517, doi:10.1182/blood-2002-07-2334 (2003). [PubMed: 12663456]
32. Cai Z, Chehab NH & Pavletich NP Structure and Activation Mechanism of the CHK2 DNA Damage Checkpoint Kinase. *Molecular Cell* 35, 818–829, doi:10.1016/j.molcel.2009.09.007 (2009). [PubMed: 19782031]
33. Falck J, Mailand N, Syljuåsen RG, Bartek J & Lukas J The ATM–Chk2–Cdc25A checkpoint pathway guards against radioresistant DNA synthesis. *Nature* 410, 842–847, doi:10.1038/35071124 (2001). [PubMed: 11298456]
34. Zipin-Roitman A et al. SMYD2 lysine methyltransferase regulates leukemia cell growth and regeneration after genotoxic stress. *Oncotarget* 8, 16712–16727, doi:10.18632/oncotarget.15147 (2017). [PubMed: 28187429]
35. Khandanpour C et al. Evidence that Growth factor independence 1b regulates dormancy and peripheral blood mobilization of hematopoietic stem cells. *Blood* 116, 5149, doi:10.1182/blood-2010-04-280305 (2010). [PubMed: 20826720]
36. Polfus Linda M. et al. Whole-Exome Sequencing Identifies Loci Associated with Blood Cell Traits and Reveals a Role for Alternative GFI1B Splice Variants in Human Hematopoiesis. *The American Journal of Human Genetics* 99, 481–488, doi:10.1016/j.ajhg.2016.06.016 (2016). [PubMed: 27486782]
37. Vassen L et al. Growth Factor Independence 1b (Gfi1b) Is Important for the Maturation of Erythroid Cells and the Regulation of Embryonic Globin Expression. *PLOS ONE* 9, e96636, doi:10.1371/journal.pone.0096636 (2014). [PubMed: 24800817]
38. Lundberg P et al. Myeloproliferative neoplasms can be initiated from a single hematopoietic stem cell expressing JAK2-V617F. *The Journal of Experimental Medicine* 211, 2213–2230, doi:10.1084/jem.20131371 (2014). [PubMed: 25288396]
39. Mansier O et al. Description of a knock-in mouse model of JAK2V617F MPN emerging from a minority of mutated hematopoietic stem cells. *Blood* 134, 2383–2387, doi:10.1182/blood.2019001163 (2019). [PubMed: 31697834]
40. Musa J et al. Cooperation of cancer drivers with regulatory germline variants shapes clinical outcomes. *Nature Communications* 10, 4128, doi:10.1038/s41467-019-12071-2 (2019).
41. Thompson DJ et al. Genetic predisposition to mosaic Y chromosome loss in blood. *Nature* 575, 652–657, doi:10.1038/s41586-019-1765-3 (2019). [PubMed: 31748747]

42. Loh P-R, Genovese G & McCarroll SA Monogenic and polygenic inheritance become instruments for clonal selection. *Nature*. doi:10.1038/s41586-020-2430-6 (2020).
43. Terao C et al. Chromosomal alterations among age-related haematopoietic clones in Japan. *Nature*. doi:10.1038/s41586-020-2426-2 (2020).
44. Naucler P et al. Human Papillomavirus and Papanicolaou Tests to Screen for Cervical Cancer. *New England Journal of Medicine* 357, 1589–1597, doi:10.1056/NEJMoa073204 (2007). [PubMed: 17942872]
45. Løberg M et al. Long-Term Colorectal-Cancer Mortality after Adenoma Removal. *New England Journal of Medicine* 371, 799–807, doi:10.1056/NEJMoa1315870 (2014). [PubMed: 25162886]
46. Cimmino L et al. Restoration of TET2 Function Blocks Aberrant Self-Renewal and Leukemia Progression. *Cell* 170, 1079–1095.e1020, doi:10.1016/j.cell.2017.07.032 (2017). [PubMed: 28823558]
47. Chen J et al. Myelodysplastic syndrome progression to acute myeloid leukemia at the stem cell level. *Nature Medicine* 25, 103–110, doi:10.1038/s41591-018-0267-4 (2019).
48. Agathocleous M et al. Ascorbate regulates haematopoietic stem cell function and leukaemogenesis. *Nature* 549, 476 (2017). [PubMed: 28825709]
49. Bycroft C et al. The UK Biobank resource with deep phenotyping and genomic data. *Nature* 562, 203–209, doi:10.1038/s41586-018-0579-z (2018). [PubMed: 30305743]
50. Zhou W et al. Efficiently controlling for case-control imbalance and sample relatedness in large-scale genetic association studies. *Nature Genetics* 50, 1335–1341, doi:10.1038/s41588-018-0184-y (2018). [PubMed: 30104761]
51. Loh P-R et al. Efficient Bayesian mixed-model analysis increases association power in large cohorts. *Nature Genetics* 47, 284 (2015). [PubMed: 25642633]
52. Hinds DA et al. Germ line variants predispose to both JAK2 V617F clonal hematopoiesis and myeloproliferative neoplasms. *Blood* 128, 1121 (2016). [PubMed: 27365426]
53. Willer CJ, Li Y & Abecasis GR METAL: fast and efficient meta-analysis of genomewide association scans. *Bioinformatics* 26, 2190–2191, doi:10.1093/bioinformatics/btq340 (2010). [PubMed: 20616382]
54. Yang J et al. Conditional and joint multiple-SNP analysis of GWAS summary statistics identifies additional variants influencing complex traits. *Nature Genetics* 44, 369 (2012). [PubMed: 22426310]
55. Hunter-Zinck H et al. Measuring genetic variation in the multi-ethnic Million Veteran Program (MVP). *bioRxiv*, 2020.2001.2006.896613, doi:10.1101/2020.01.06.896613 (2020).
56. Nielsen C, Birgens HS, Nordestgaard BG & Bojesen SE Diagnostic value of JAK2 V617F somatic mutation for myeloproliferative cancer in 49 488 individuals from the general population. *British Journal of Haematology* 160, 70–79, doi:10.1111/bjh.12099 (2013). [PubMed: 23116358]
57. Magosi LE, Goel A, Hopewell JC, Farrall M & on behalf of the, C. D. C. Identifying systematic heterogeneity patterns in genetic association meta-analysis studies. *PLOS Genetics* 13, e1006755, doi:10.1371/journal.pgen.1006755 (2017). [PubMed: 28459806]
58. Chang CC et al. Second-generation PLINK: rising to the challenge of larger and richer datasets. *GigaScience* 4, doi:10.1186/s13742-015-0047-8 (2015).
59. Wakefield J A Bayesian Measure of the Probability of False Discovery in Genetic Epidemiology Studies. *The American Journal of Human Genetics* 81, 208–227, doi:10.1086/519024 (2007). [PubMed: 17668372]
60. Michailidou K et al. Association analysis identifies 65 new breast cancer risk loci. *Nature* 551, 92 (2017). [PubMed: 29059683]
61. Sud A et al. Familial risks of acute myeloid leukemia, myelodysplastic syndromes, and myeloproliferative neoplasms. *Blood* 132, 973 (2018). [PubMed: 29991558]
62. Tapper W et al. Genetic variation at MECOM, TERT, JAK2 and HBS1L-MYB predisposes to myeloproliferative neoplasms. *Nat Commun* 6, 6691, doi:10.1038/ncomms7691 (2015). [PubMed: 25849990]
63. Ulirsch JC et al. Interrogation of human hematopoiesis at single-cell and single-variant resolution. *Nature Genetics* 51, 683–693, doi:10.1038/s41588-019-0362-6 (2019). [PubMed: 30858613]

64. Bulik-Sullivan B et al. An atlas of genetic correlations across human diseases and traits. *Nature Genetics* 47, 1236 (2015). [PubMed: 26414676]
65. Roaldsnes C, Holst R, Frederiksen H & Ghanima W Myeloproliferative neoplasms: trends in incidence, prevalence and survival in Norway. *European Journal of Haematology* 98, 85–93, doi:10.1111/ejh.12788 (2017). [PubMed: 27500783]
66. Höglund M, Sandin F & Simonsson B Epidemiology of chronic myeloid leukaemia: an update. *Annals of Hematology* 94, 241–247, doi:10.1007/s00277-015-2314-2 (2015).
67. Li C et al. Genome-wide Association Analysis in Humans Links Nucleotide Metabolism to Leukocyte Telomere Length. *The American Journal of Human Genetics* 106, 389–404, doi:10.1016/j.ajhg.2020.02.006 (2020). [PubMed: 32109421]
68. Bick AG et al. Inherited Causes of Clonal Hematopoiesis of Indeterminate Potential in TOPMed Whole Genomes. *bioRxiv*, 782748, doi:10.1101/782748 (2019).
69. Finucane HK et al. Heritability enrichment of specifically expressed genes identifies disease-relevant tissues and cell types. *Nature Genetics* 50, 621–629, doi:10.1038/s41588-018-0081-4 (2018). [PubMed: 29632380]
70. Benner C et al. FINEMAP: efficient variable selection using summary data from genome-wide association studies. *Bioinformatics* 32, 1493–1501, doi:10.1093/bioinformatics/btw018 (2016). [PubMed: 26773131]
71. Walker CJ et al. Genome-wide association study identifies an acute myeloid leukemia susceptibility locus near BICRA. *Leukemia* 33, 771–775, doi:10.1038/s41375-018-0281-z (2019). [PubMed: 30291333]
72. Coetzee SG, Coetzee GA & Hazelett DJ motifbreakR: an R/Bioconductor package for predicting variant effects at transcription factor binding sites. *Bioinformatics* 31, 3847–3849, doi:10.1093/bioinformatics/btv470 (2015). [PubMed: 26272984]
73. Kulakovskiy IV et al. HOCOMOCO: towards a complete collection of transcription factor binding models for human and mouse via large-scale ChIP-Seq analysis. *Nucleic Acids Research* 46, D252–D259, doi:10.1093/nar/gkx1106 (2017).
74. Hemani G et al. The MR-Base platform supports systematic causal inference across the human phenome. *eLife* 7, e34408, doi:10.7554/eLife.34408 (2018). [PubMed: 29846171]
75. Verbanck M, Chen C-Y, Neale B & Do R Detection of widespread horizontal pleiotropy in causal relationships inferred from Mendelian randomization between complex traits and diseases. *Nature Genetics* 50, 693–698, doi:10.1038/s41588-018-0099-7 (2018). [PubMed: 29686387]
76. Sanna S et al. Causal relationships among the gut microbiome, short-chain fatty acids and metabolic diseases. *Nature Genetics*, doi:10.1038/s41588-019-0350-x (2019).
77. Bowden J, Davey Smith G & Burgess S Mendelian randomization with invalid instruments: effect estimation and bias detection through Egger regression. *International Journal of Epidemiology* 44, 512–525, doi:10.1093/ije/dyv080 (2015). [PubMed: 26050253]
78. Jaganathan K et al. Predicting Splicing from Primary Sequence with Deep Learning. *Cell* 176, 535–548.e524, doi:10.1016/j.cell.2018.12.015 (2019). [PubMed: 30661751]
79. McLaren W et al. The Ensembl Variant Effect Predictor. *Genome Biology* 17, 122, doi:10.1186/s13059-016-0974-4 (2016). [PubMed: 27268795]
80. Javierre BM et al. Lineage-Specific Genome Architecture Links Enhancers and Non-coding Disease Variants to Target Gene Promoters. *Cell* 167, 1369–1384.e1319, doi:10.1016/j.cell.2016.09.037 (2016). [PubMed: 27863249]
81. Frankish A et al. GENCODE reference annotation for the human and mouse genomes. *Nucleic Acids Research* 47, D766–D773, doi:10.1093/nar/gky955 (2018).
82. de Leeuw CA, Mooij JM, Heskes T & Posthuma D MAGMA: Generalized Gene-Set Analysis of GWAS Data. *PLOS Computational Biology* 11, e1004219, doi:10.1371/journal.pcbi.1004219 (2015). [PubMed: 25885710]
83. Watanabe K, Taskesen E, van Bochoven A & Posthuma D Functional mapping and annotation of genetic associations with FUMA. *Nature Communications* 8, 1826, doi:10.1038/s41467-017-01261-5 (2017).

84. Grinfeld J et al. Classification and Personalized Prognosis in Myeloproliferative Neoplasms. *New England Journal of Medicine* 379, 1416–1430, doi:10.1056/NEJMoa1716614 (2018). [PubMed: 30304655]
85. Raudvere U et al. g:Profiler: a web server for functional enrichment analysis and conversions of gene lists (2019 update). *Nucleic Acids Research* 47, W191–W198, doi:10.1093/nar/gkz369 (2019). [PubMed: 31066453]
86. Butler A, Hoffman P, Smibert P, Papalexi E & Satija R Integrating single-cell transcriptomic data across different conditions, technologies, and species. *Nature Biotechnology* 36, 411 (2018).
87. Pellin D et al. A comprehensive single cell transcriptional landscape of human hematopoietic progenitors. *Nature Communications* 10, 2395, doi:10.1038/s41467-019-10291-0 (2019).
88. van Dijk D et al. Recovering Gene Interactions from Single-Cell Data Using Data Diffusion. *Cell* 174, 716–729.e727, doi:10.1016/j.cell.2018.05.061 (2018). [PubMed: 29961576]
89. Delano WL The PyMOL Molecular Graphics System. <http://www.pymol.org> (2002).
90. Zipin-Roitman A et al. SMYD2 lysine methyltransferase regulates leukemia cell growth and regeneration after genotoxic stress. *Oncotarget* 8, 16712–16727, doi:10.18632/oncotarget.15147 (2017). [PubMed: 28187429]
91. Milyavsky M et al. A Distinctive DNA Damage Response in Human Hematopoietic Stem Cells Reveals an Apoptosis-Independent Role for p53 in Self-Renewal. *Cell Stem Cell* 7, 186–197, doi:10.1016/j.stem.2010.05.016 (2010).
92. Piacibello W et al. Lentiviral gene transfer and ex vivo expansion of human primitive stem cells capable of primary, secondary, and tertiary multilineage repopulation in NOD/SCID mice. *Blood* 100, 4391, doi:10.1182/blood.V100.13.4391 (2002). [PubMed: 12453876]
93. Cohen S et al. Hematopoietic stem cell transplantation using single UM171-expanded cord blood: a single-arm, phase 1–2 safety and feasibility study. *The Lancet Haematology* 7, e134–e145, doi:10.1016/S2352-3026(19)30202-9 (2020). [PubMed: 31704264]
94. Fares I et al. Pyrimidoindole derivatives are agonists of human hematopoietic stem cell self-renewal. *Science* 345, 1509, doi:10.1126/science.1256337 (2014). [PubMed: 25237102]
95. Tomellini E et al. Integrin- α 3 Is a Functional Marker of Ex Vivo Expanded Human Long-Term Hematopoietic Stem Cells. *Cell Reports* 28, 1063–1073.e1065, doi:10.1016/j.celrep.2019.06.084 (2019). [PubMed: 31340144]

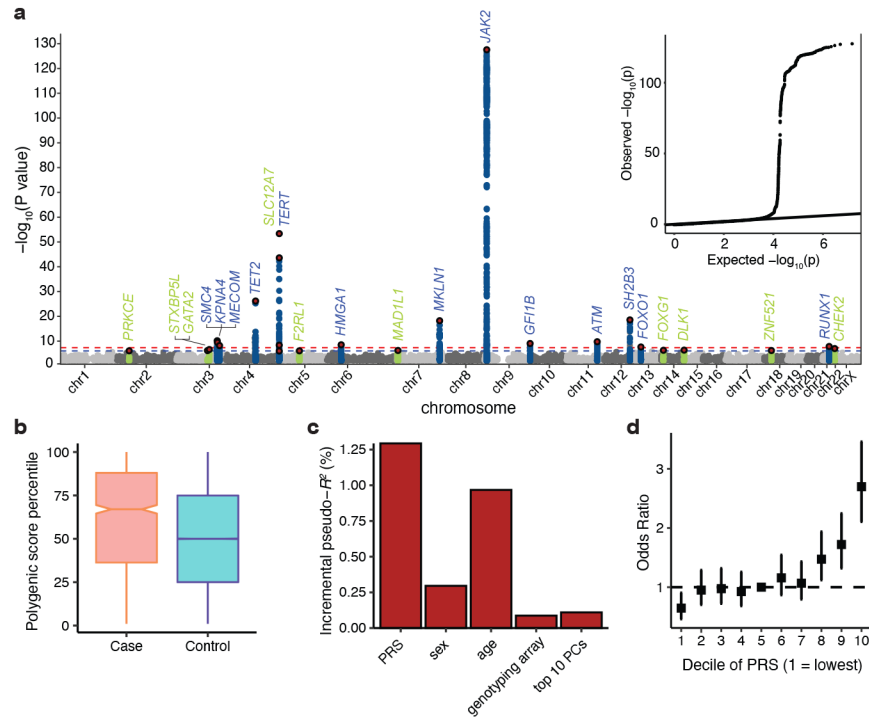


Figure 1. Genetic architecture of inherited MPN risk.
a, Manhattan plot and quantile-quantile (QQ) plot illustrating results of the genome-wide association study (GWAS) meta-analysis for MPNs (n = 2,949 cases and 835,554 controls). X axis is chromosomal position, and y axis is the $-\log_{10}(P)$ value of association (two-tailed, logistic regression). Association signals reaching genome-wide significance ($P < 5 \times 10^{-8}$) and suggestive significance ($P < 1 \times 10^{-6}$) are shown in blue and green, respectively. Red points represent conditionally independent lead variants within each locus. Labels correspond to target gene if present (Fig. 3a), or otherwise the nearest gene at each association locus (± 500 kb). The QQ plot illustrates the deviation of association test statistics (points) from the distribution expected under the null hypothesis (line). **b**, Polygenic risk score (PRS) percentile among MPN cases (n = 1,086) versus controls (n = 407,155) in the UK Biobank test set. Box plots show the median as the line in the notch, with the top and bottom of the box indicating the interquartile range. Whiskers extend to the maximum or minimum value. Notches indicate the 95% confidence interval of the medians. **c**, Additional variance in MPN risk (n = 1,086 cases and 407,155 controls) explained by PRS compared to age, sex, genotyping array, and top 10 principal components of genetic relatedness. **d**, Odds ratio for MPN acquisition (n = 1,086 cases and 407,155 controls) stratified by deciles of the PRS, with the 5th decile as the reference. Data represent odds ratios and 95% confidence intervals.

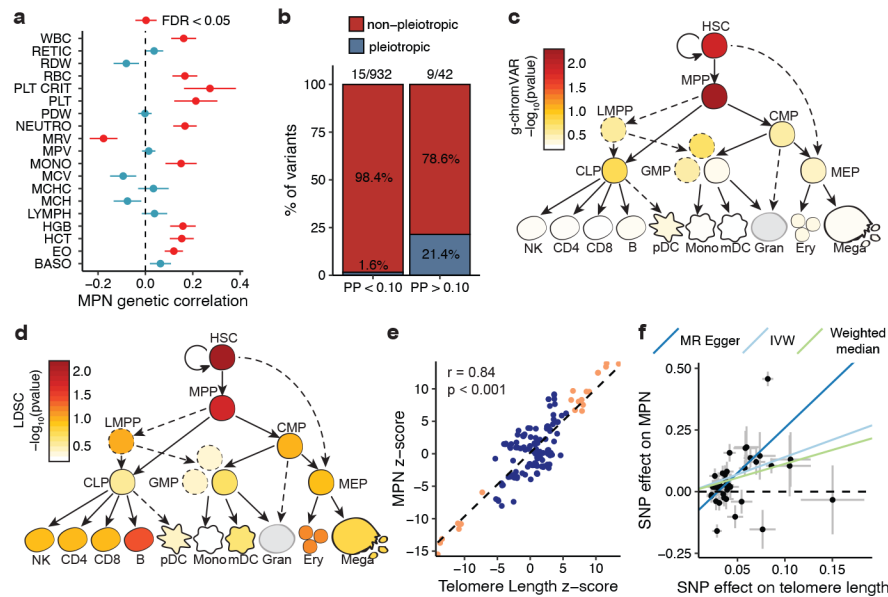


Figure 2. Functional enrichments in MPN risk.

a, Genetic correlations (\pm standard errors) between MPN risk ($n = 2,949$ cases and $835,554$ controls) and 19 blood traits ($n = 408,241$), estimated by LDSC. WBC, white blood cell count; RETIC, reticulocyte count; RDW, red cell distribution width; RBC, red blood cell count; PLT CRIT, platelet crit; PLT, platelet count; PDW, platelet distribution width; NEUTRO, neutrophil count; MRV, mean reticulocyte volume; MPV, mean platelet volume; MONO, monocyte count; MCV, mean corpuscular volume; MCHC, mean corpuscular hemoglobin concentration; MCH, mean corpuscular hemoglobin; LYMPH, lymphocyte count; HGB, hemoglobin; HCT, hematocrit; EO, eosinophil count; BASO, basophil count. Red color indicates false discovery rate-adjusted $p < 0.05$. **b**, Proportion of MPN risk variants with fine-mapped posterior probability (PP) > 0.10 vs. PP < 0.10 that exhibit pleiotropic associations with traits from two or more hematopoietic lineages (basophil, eosinophil, neutrophil, red blood cell, platelet, monocyte, lymphocyte). **c-d**, g-chromVAR and LD score regression results for the enrichment of MPN risk variants across 18 hematopoietic chromatin accessibility profiles. **e**, Correlation of GWAS effect sizes (z-scores) of variants in the *TERT* locus for MPN vs. telomere length ($p < 2.2 \times 10^{-16}$, two-tailed Pearson correlation). The dashed line is the line of best fit. Orange color indicates variants with association $P < 5 \times 10^{-8}$ in both the MPN and telomere length GWAS. **f**, Mendelian randomization (MR) plot showing LD-independent telomere length GWAS variants ($p < 1 \times 10^{-5}$, $r^2 < 0.001$) and their effects on MPN risk (outcome) versus telomere length (exposure). Lines represent slopes of three regression tests: MR-Egger ($p = 4.83 \times 10^{-5}$), inverse-variance weighted ($p = 1.36 \times 10^{-4}$), and weighted median ($p = 1.15 \times 10^{-5}$). Data represent MR effect sizes \pm standard error.

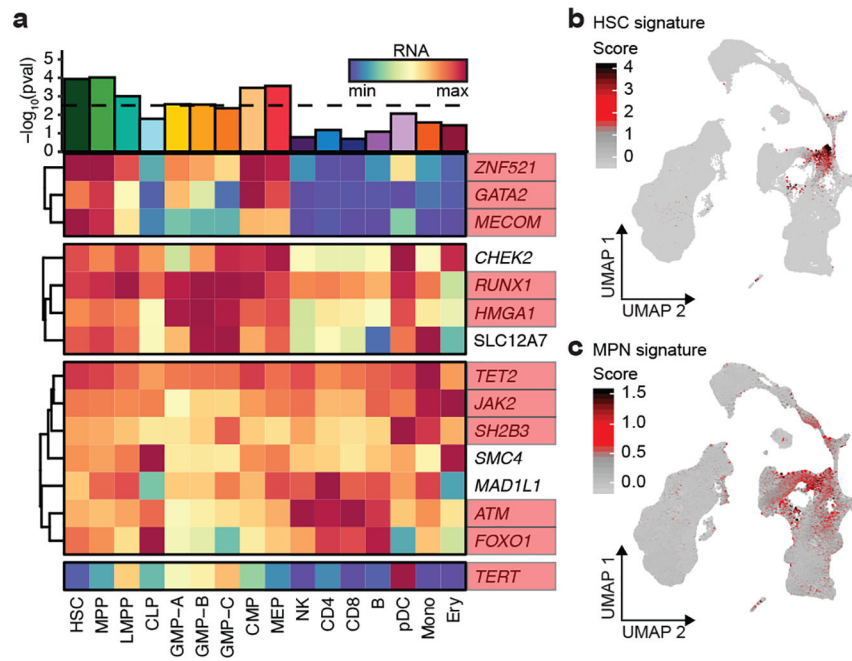


Figure 3. Target genes for MPN risk.

a, Heatmap of 15 MPN target genes, visualizing RNA expression across 16 hematopoietic populations. Bar plot depicts the enrichment of target gene expression in each cell type (two-tailed rank-sum permutation test). Genes with known involvement in hematopoietic stem cell function are boxed in red. **b-c**, UMAP projections of 278,978 single cells from human bone marrow, colored according to **(b)** HSC and **(c)** MPN target gene signatures.

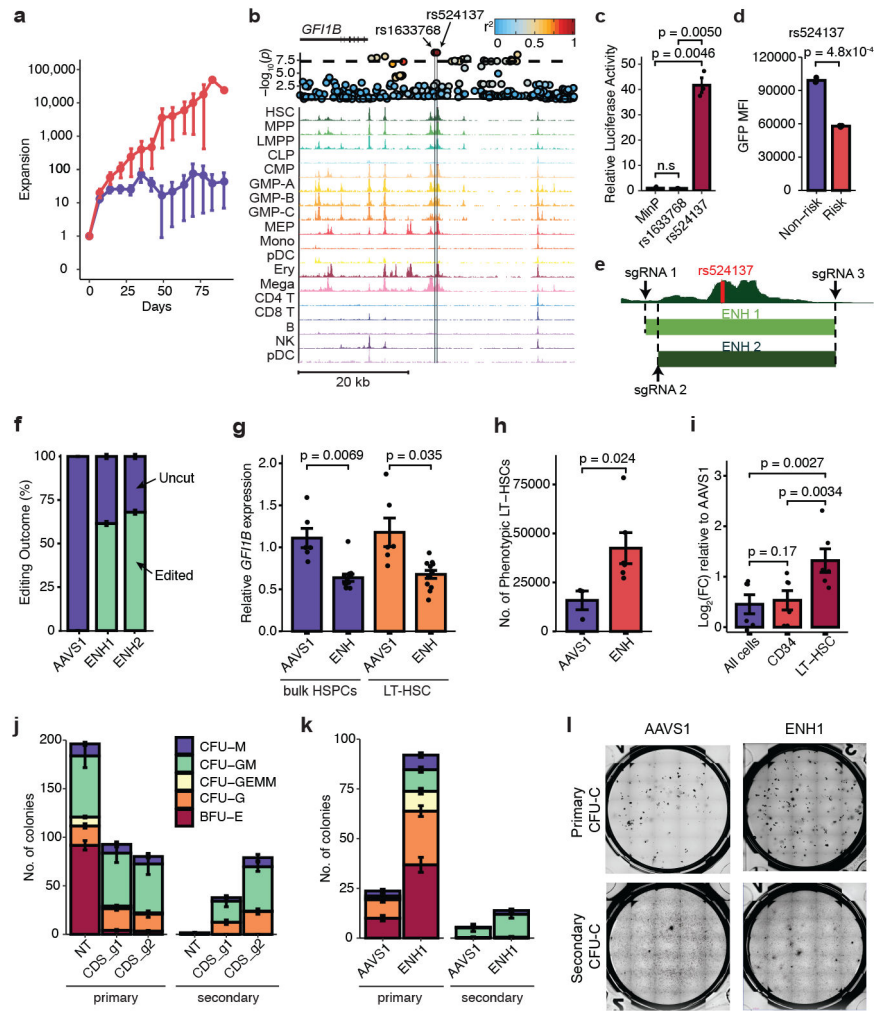


Figure 4. Characterizing the mechanisms of two MPN risk variants.
a, Expansion of Lin-CD34⁺ derived hematopoietic stem and progenitor cells (HSPCs) after short hairpin RNA knockdown of *CHEK2* vs. control (*CHEK2*, n = 4; control, n = 9). **b**, rs1633768 and rs524137 fall in a region of hematopoietic accessible chromatin downstream of *GFI1B*. A locus plot is shown above, plotting $-\log_{10}(p)$ of MPN association; color reflects linkage disequilibrium to rs524137. **c**, Luciferase reporter assay testing regulatory activity of genomic regions containing rs1622768 and rs524137 in hematopoietic cells compared to a minimal promoter (MinP) construct (n = 3). **d**, Lentiviral reporter assays testing allele-specific activity of rs524137 in hematopoietic cells (n = 3). **e**, HSC chromatin accessibility around rs524137 and the two CRISPR-Cas9 guide RNA pairs (ENH1 and ENH2) used to delete this region. **f**, Frequency of uncut and edited alleles after editing of *GFI1B* enhancer or control AAVS1 site in human CD34⁺ HSPCs (n = 6). **g**, *GFI1B* expression in bulk HSPCs and sorted phenotypic long-term HSCs (LT-HSCs) following *GFI1B* enhancer deletion (n = 12) compared to AAVS1 editing (n = 6). Due to similar editing outcomes, ENH1 and ENH2 were combined as ENH in this and subsequent experiments. **h**, Total number of phenotypic LT-HSCs in HSC maintenance culture, 6 days after editing of *GFI1B* enhancer (n = 6) or AAVS1 (n = 3). **i**, Relative expansion of cell numbers in various compartments (All cells,

CD34⁺, and LT-HSCs) upon *GFI1B* enhancer deletion (n = 6) compared to AAVS1 controls (n = 3). **j**, *GFI1B* coding disruption (CDS_g1 and CDS_g2, n = 3 each) leads to reduced erythroid primary colony formation compared to non-targeting (NT) control (n = 3), but increases secondary colony formation. CFU-M, colony forming unit-macrophage; CFU-GM, granulocyte macrophage; CFU-GEMM, granulocyte erythrocyte macrophage megakaryocyte; CFU-G, granulocyte; BFU-E, burst forming unit-erythroid. **k**, *GFI1B* enhancer deletion increases secondary colony formation without affecting erythroid colony formation (n = 3). **l**, Representative images of primary (top) and secondary (bottom) CFU-C colonies. Data from **a**, **c-d**, **f-k** are means \pm s.e.m. n denotes the number of biologically independent replicates. Statistical methods used were two-tailed unpaired *t*-test (**c**, **d**, **g**, **h**) and two-tailed paired *t*-test (**i**).