


BMJ Open Sex-specific temporal trends in ambulatory heart failure incidence, mortality and hospitalisation in Ontario, Canada from 1994 to 2013: a population-based cohort study

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ABSTRACT

Objectives To examine the temporal trends in mortality and heart failure (HF) hospitalisation in ambulatory patients following a new diagnosis of HF.

Design Retrospective cohort study

Setting Outpatient

Participants Ontario residents who were diagnosed with HF in an outpatient setting between 1994 and 2013.

Primary and secondary outcome measures The primary outcome was all-cause mortality within 1 year of diagnosis and the secondary outcome was HF hospitalisation within 1 year. Risks of mortality and hospitalisation were calculated using the Kaplan-Meier method and the relative hazard of death was assessed using multivariable Cox proportional hazard models.

Results A total of 352 329 patients were studied (50% female). During the study period, there was a greater decline in age standardised 1-year mortality rates (AMR) in men (33%) than in women (19%). Specifically, female AMR at 1 year was 10.4% (95% CI 9.1% to 12.0%) in 1994 and 8.5% (95% CI 7.5% to 9.5%) in 2013, and male AMR at 1 year was 12.3% (95% CI 11.1% to 13.7%) in 1994 and 8.3% (95% CI 7.5% to 9.1%) in 2013. Conversely, age standardised HF hospitalisation rates declined in men (11.4% (95% CI 10.1% to 12.9%) in 1994 and 9.1% (95% CI 8.2% to 10.1%) in 2013) but remained unchanged in women (9.7% (95% CI 8.3% to 11.3%) in 1994 and 9.8% (95% CI 8.6% to 11.0%) in 2013).

Conclusion Among patients with HF over a 20-year period, there was a greater improvement in the prognosis of men compared with women. Further research should focus on the determinants of this disparity and ways to reduce this gap in outcomes.

INTRODUCTION

Heart failure (HF) is a significant cause of morbidity and mortality for both women and men.^{1 2} Despite the current era of guideline directed medical therapy, HF continues to be a leading cause of admission to hospital. It is associated with a poor

Strengths and limitations of this study

- First and largest population-based study to examine temporal, sex-specific trends in heart failure (HF) outcomes in an ambulatory setting.
- The nature of our publicly funded healthcare system allowed for analysis of all patients diagnosed with HF in Ontario without selection bias.
- Information on ejection fraction was not available in the databases used.

prognosis and contributes to 35% of cardiovascular mortality in women.³ Despite this, HF remains poorly understood in women, and women continue to be underrepresented in HF clinical trials.⁴ The underlying mechanism of HF is often different in women and men, with women suffering more often from HF of a hypertensive rather than ischaemic aetiology.^{5 6} Important trends in the incidence and outcomes of hospitalised HF patients have been recently published^{7 8}; these studies suggest that the incidence of HF has declined in many inpatient cohorts, however, the prognosis of this disease remains poor. An in-depth understanding of the temporal trends in HF incidence and outcomes is also needed in the ambulatory setting, where the majority of HF cases are diagnosed and managed. Also, given the sex differences in comorbidities and outcomes in HF, it is not known if these temporal changes are modified by sex. We, therefore, examined the sex differences in HF comorbidities, incidence, mortality and hospitalisation in a population-based ambulatory cohort from fiscal years 1994 to 2013.

Table 1 Temporal trends in characteristics of men and women with incident heart failure over time

Variable	Historic cohort (1994–1998)		1999–2003		2004–2009		Modern cohort (2009–2013)		P value
	Women n=47 676	Men n=43 907	Women n=43 190	Men n=42 108	Women n=41 043	Men n=43 698	Women n=42 746	Men n=47 961	
Sociodemographics									
Age (mean±SD)	76.3±11.2	72.0±11.3	76.2±11.8	71.9±11.8	76.0±12.3	71.6±12.4	75.8±12.9	71.5±12.8	<0.001
Rurality	8256 (17.3)	8518 (19.4)	6781 (15.7)	6996 (16.6)	5602 (13.6)	6667 (15.3)	5358 (12.5)	6788 (14.2)	<0.001
Income quintile									
1 (lowest)	11 651 (24.4)	9297 (21.2)	10 121 (23.4)	8529 (20.3)	9661 (23.5)	9096 (20.8)	9613 (22.5)	9541 (19.9)	<0.001
2	10 377 (21.8)	9418 (21.4)	9814 (22.7)	9319 (22.1)	9070 (22.1)	9317 (21.3)	9243 (21.6)	9833 (20.5)	<0.001
3	9506 (19.9)	8983 (20.5)	8571 (19.8)	8683 (20.6)	7891 (19.2)	8576 (19.6)	8325 (19.5)	9444 (19.7)	<0.001
4	8208 (17.2)	8024 (18.3)	7403 (17.1)	7770 (18.5)	7274 (17.7)	8308 (19.0)	8017 (18.8)	9775 (20.4)	<0.001
5 (highest)	7698 (16.1)	7956 (18.1)	7168 (16.6)	7690 (18.3)	6976 (17.0)	8233 (18.8)	7388 (17.3)	9188 (19.2)	<0.001
Comorbidities									
Uncomplicated HTN	31 751 (66.6)	26 068 (59.4)	31 381 (72.7)	27 909 (66.3)	31 336 (76.3)	30 730 (70.3)	32 809 (76.8)	34 103 (71.1)	<0.001
Complicated HTN	2507 (5.3)	2006 (4.6)	2178 (5.0)	1921 (4.6)	3254 (7.9)	3551 (8.1)	5669 (13.3)	6516 (13.6)	0.15
Atrial fibrillation	4660 (9.8)	4769 (10.9)	5034 (11.7)	5313 (12.6)	5180 (12.6)	5956 (13.6)	5957 (13.9)	6919 (14.4)	0.035
Myocardial infarction	2954 (6.2)	4697 (10.7)	3629 (8.4)	5937 (14.1)	4323 (10.5)	8030 (18.4)	5051 (11.8)	9617 (20.1)	<0.001
Valvular disease	1835 (3.8)	1864 (4.2)	1981 (4.6)	2087 (5.0)	1759 (4.3)	2163 (4.9)	2099 (4.9)	2841 (5.9)	<0.001
PAD	3136 (6.6)	3994 (9.1)	2928 (6.8)	3684 (8.7)	1584 (3.9)	2405 (5.5)	1272 (3.0)	2257 (4.7)	<0.001
CVD	4429 (9.3)	4401 (10.0)	3841 (8.9)	3834 (9.1)	2808 (6.8)	3003 (6.9)	2506 (5.9)	2713 (5.7)	0.18
PCD	296 (0.6)	286 (0.7)	491 (1.1)	362 (0.9)	807 (2.0)	588 (1.3)	1113 (2.6)	882 (1.8)	<0.001
COPD/asthma	13 091 (27.5)	14 352 (32.7)	14 557 (33.7)	14 988 (35.6)	15 209 (37.1)	15 996 (36.6)	16 261 (38.0)	17 164 (35.8)	<0.001
Alcohol abuse	128 (0.3)	398 (0.9)	187 (0.4)	562 (1.3)	313 (0.8)	965 (2.2)	337 (0.8)	1123 (2.3)	<0.001
Renal disease	1127 (2.4)	1527 (3.5)	1410 (3.3)	1935 (4.6)	1849 (4.5)	2588 (5.9)	1954 (4.6)	2865 (6.0)	<0.001
Diabetes	9600 (20.1)	10 462 (23.8)	10 359 (24.0)	12 182 (28.9)	12 204 (29.7)	15 243 (34.9)	14 751 (34.5)	18 721 (39.0)	<0.001
Hypothyroidism	2155 (4.5)	541 (1.2)	2212 (5.1)	597 (1.4)	1421 (3.5)	428 (1.0)	1300 (3.0)	479 (1.0)	<0.001
Liver disease	559 (1.2)	772 (1.8)	541 (1.3)	800 (1.9)	540 (1.3)	775 (1.8)	579 (1.4)	895 (1.9)	<0.001
Anaemia	2569 (5.4)	1403 (3.2)	4734 (11.0)	3204 (7.6)	4629 (11.3)	3382 (7.7)	5182 (12.1)	3929 (8.2)	<0.001
Dementia	1935 (4.1)	1188 (2.7)	2007 (4.6)	1258 (3.0)	1816 (4.4)	1121 (2.6)	1791 (4.2)	1194 (2.5)	<0.001
Depression	1788 (3.8)	911 (2.1)	1617 (3.7)	872 (2.1)	1364 (3.3)	859 (2.0)	1213 (2.8)	756 (1.6)	<0.001
Psychosis	1932 (4.1)	1134 (2.6)	1695 (3.9)	1106 (2.6)	486 (1.2)	350 (0.8)	223 (0.5)	171 (0.4)	<0.001
Primary tumour	3055 (6.4)	4363 (9.9)	3026 (7.0)	4045 (9.6)	2947 (7.2)	4060 (9.3)	3376 (7.9)	4419 (9.2)	<0.001
Metastatic Malignancy	869 (1.8)	804 (1.8)	925 (2.1)	775 (1.8)	850 (2.1)	723 (1.7)	1000 (2.3)	840 (1.8)	<0.001
Paraplegia	840 (1.8)	899 (2.0)	854 (2.0)	903 (2.1)	521 (1.3)	572 (1.3)	434 (1.0)	475 (1.0)	0.71
VTE	1161 (2.4)	1017 (2.3)	945 (2.2)	714 (1.7)	557 (1.4)	505 (1.2)	450 (1.1)	512 (1.1)	0.83
Frailty	4302 (9.0)	2700 (6.1)	5598 (13.0)	4384 (10.4)	10 163 (24.8)	7175 (16.4)	11 223 (26.3)	8409 (17.5)	<0.001

Values are expressed number (%) unless otherwise indicated.

COPD, chronic obstructive pulmonary disease; CVD, cerebrovascular disease; HTN, hypertension; PAD, peripheral arterial disease; PCD, pulmonary circulatory disease; VTE, venous thromboembolism.

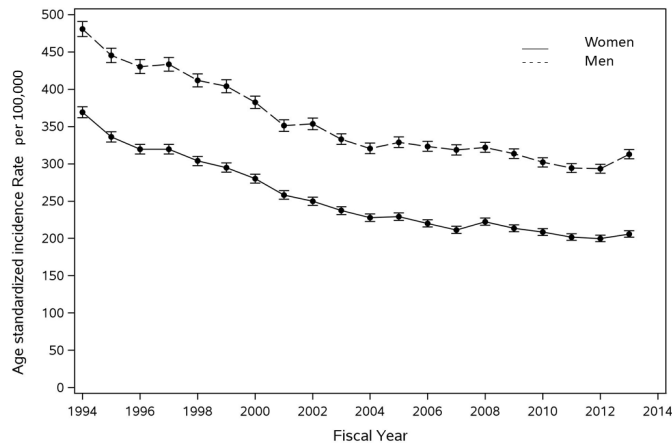


Figure 1 Sex-specific temporal trends in the incidence of heart failure in Ontario from 1 April 1994 to 31 March 2014. Incidence rates were directly standardised by age and expressed per 100 000. The solid line represents incidence trends in women. The dotted line represents incidence trends in men.

METHODS

Design and study population

We conducted a population-based, retrospective cohort study of Ontario residents who were diagnosed with HF in an outpatient setting over a 20-year period, using linked administrative databases.

Included were adult patients ≥ 40 years of age, who were newly diagnosed with HF in an ambulatory setting between 1 April 1994 and 31 March 2014. We excluded non-Ontario residents, those who were ≥ 105 years of age on the date of HF diagnosis, those who were diagnosed with HF in an inpatient setting and in whom HF had developed as a postadmission complication. Ontario is Canada's most populous and ethnically diverse province with a public funded healthcare system that reimburses all medically necessary physician and hospital services.

Patient and public involvement

Patients and the public were not involved in the design and conception of this study. However, the results will be publicly disseminated.

Data sources

Databases were linked deterministically using unique encoded identifiers. Ambulatory incident HF cases were identified using the Institute for Clinical Evaluative Sciences Congestive HF database, based on two outpatient billing claims for HF within 1 year. This algorithm was validated in primary care patient records and shown to have 85% sensitivity and 97% specificity in identifying HF events.⁹ The congestive HF database allowed us to study a validated cohort of HF patients with consistent entry criteria over time. Our analyses were conducted by linking the Congestive HF database with the Registered Persons Database, which contains demographic and vital statistics information, the Canadian Institute for Health Information Discharge Abstract Database, which contains

data on all hospitalisations and comorbidities, and Same Day Surgery database for comorbidities. Physician fee-for-service claims data were obtained from the Ontario Health Insurance Plan database. While lacking physiologic and laboratory measures, these databases have been validated for many outcomes, exposures and comorbidities.¹⁰⁻¹³

Outcome

The primary outcome was all-cause mortality within 1 year of HF diagnosis. Mortality was ascertained by using the Registered Persons Database. Secondary outcome was HF hospitalisation within 1 year of HF diagnosis, which was ascertained using the Discharge Abstract Database.

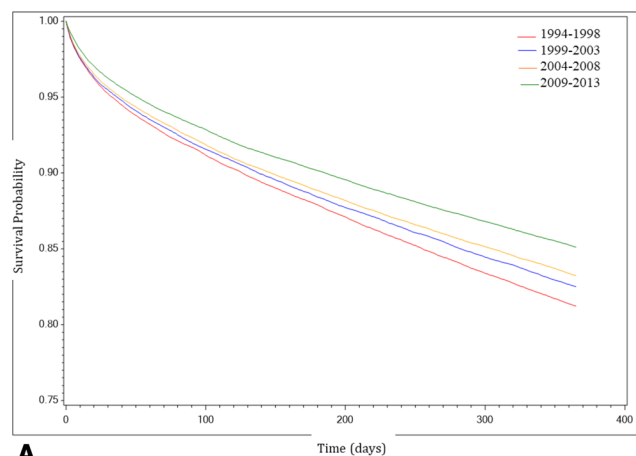
Covariates

Demographic variables were obtained from the Registered Persons Database. We estimated socioeconomic status based on patients' neighbourhood median income in the Canadian census, and determined rural versus urban residence using Statistics Canada definitions.¹⁴ We identified hypertension,¹⁰ asthma,¹⁵ chronic obstructive pulmonary disease (COPD)¹⁶ and diabetes mellitus¹² using validated algorithms applied on patient encounters within 5 years of HF diagnosis. Other comorbidities were identified using Discharge Abstract Database, Same Day Surgery and Ontario Health Insurance Plan databases based on International Classification of Diseases 10th Revision codes within 5 years of HF diagnosis, using previously described methods.¹⁷⁻²⁹ Frailty was identified using the Johns Hopkins Adjusted Clinical Groups frailty-defining diagnoses indicator, which is an instrument designed and validated for research of frailty-related outcomes and resource utilisation using administrative data.^{25 28 30-34}

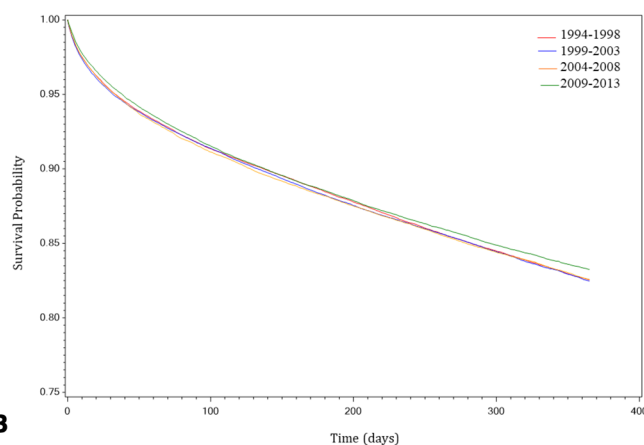
Statistical analysis

Analyses were stratified by sex. Continuous variables were expressed as mean (SD) and categorical variables as number (proportion). Mortality was assessed at 1-year post-HF diagnosis. Survival time was defined as the date of HF diagnosis until date of death or last follow-up. Patients were censored when they lost possession of a valid Ontario health insurance number for two consecutive eligibility quarters (ie, have left the province of Ontario). Probability of death within given durations of follow-up were calculated using the Kaplan-Meier method, with the significance of the difference between sexes assessed using the log-rank test. We estimated the cumulative incidence of HF hospitalisations using cumulative incidence functions, which treated death as a competing risk. We constructed age standardised plots of HF incidence, 1-year mortality and HF hospitalisation in men and women over the 20-year period. These rates were directly standardised by age using the 1991 Canadian population aged ≥ 40 years as the reference population.

We used linear regression with fiscal year as the independent variable to assess for temporal changes in HF incidence and outcomes in women and men across the



A



B

Figure 2 (A) Stacked Kaplan-Meier curves representing temporal trends in 1-year survival after heart failure diagnosis in men. The red line represents survival in the historical cohort (1994–1998). The blue line represents survival in the 1999–2003 cohort. The orange line represents survival in the 2004–2008 cohort. The green line represents survival in the modern cohort (2009–2013). (B) Stacked Kaplan-Meier curves representing temporal trends in 1-year survival after heart failure diagnosis in women. The red line represents survival in the historical cohort (1994–1998). The blue line represents survival in the 1999–2003 cohort. The orange line represents survival in the 2004–2008 cohort. The green line represents survival in the modern cohort (2009–2013).

20-year period. To examine the temporal changes in comorbidities, we divided the 20-year period into four temporal cohorts: those diagnosed with HF between 1 April 1994 and 31 March 1999 (the historical cohort), between 1 April 1999 and 31 March 2004, between 1 April 2004 and 31 March 2009 and between 1 April 2009 and 31 March 2014 (the modern cohort). The hazard of death in the historical cohort and the modern cohort were assessed using Cox proportional hazard models with and without multivariable adjustment. To justify sex-specific analyses, we also tested for the presence of any interaction between sex and each of the mortality risk factors in

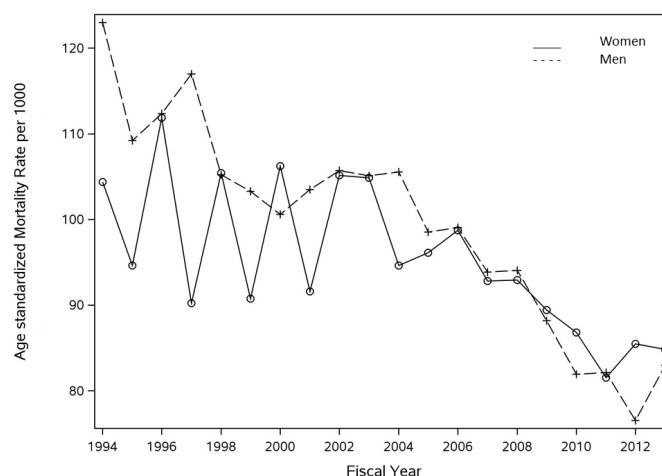


Figure 3 Sex-specific temporal trends in mortality within 1-year of ambulatory heart failure diagnosis. Mortality rates were directly standardised by age and expressed per 1000. The solid line represents mortality trends in women. The dotted line represents mortality trends in men.

these two cohorts using multiplicative interaction terms. The measure of association was HRs with 95% CI. Analyses were performed using SAS V.9.4 (SAS Institute), with statistical significance defined by a two-sided $p < 0.05$.

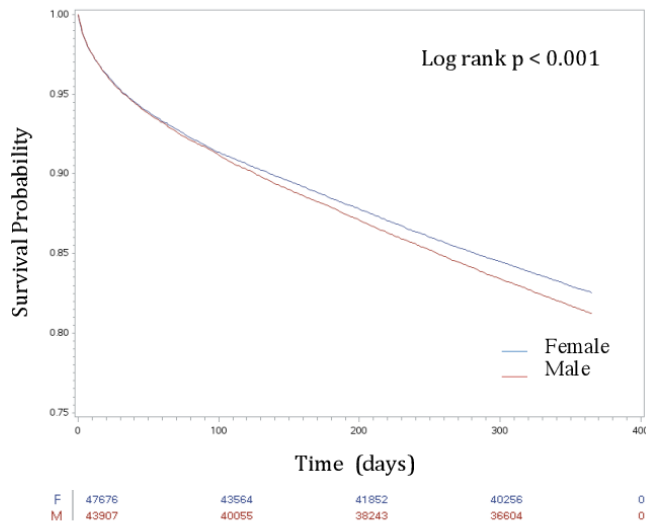
RESULTS

Over a 20-year period in Ontario, a total of 352 329 patients were diagnosed with HF in an ambulatory setting (50% women). There were 91 583 incident HF cases in the historical cohort (52% women) and 90 707 in the modern cohort (47% women). Throughout the study period, women with HF were more likely to be older, more frail, of lower income status, have comorbid conditions such as hypertension, hypothyroidism, anaemia, dementia and depression, but were less likely to have myocardial infarction (MI), peripheral arterial disease, diabetes and alcohol abuse compared with men (table 1). Compared with the historical cohort, modern HF patients were less likely to have peripheral and cerebral vascular disease, psychosis, paraplegia and venous thromboembolic disease; but more likely to be urban dwellers, have hypertension, atrial fibrillation, MI, valvular heart disease, pulmonary circulatory disorder, COPD/asthma, alcohol abuse, renal disease and are increasingly frail.

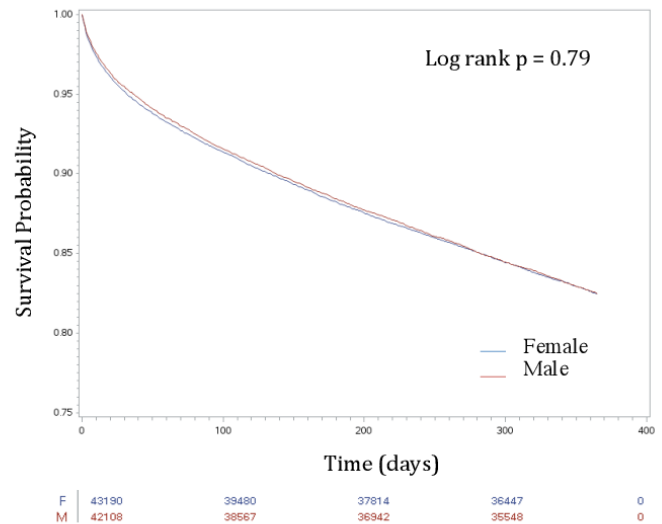
Trends in HF incidence

During the historical period (1994–1998), a total of 47 676 (0.36%) incident HF cases were identified in women and 43 907 (0.36%) in men. During the modern era (2009–2013), 42 746 (0.24%) incident cases were identified in women and 47 961 (0.29%) in men. Although the incidence of HF declined in both sexes over the 20-year period (linear regression slope, -0.031 ; $p < 0.0001$ in women and -0.025 ; $p < 0.0001$ in men), it remained higher in men than in women (figure 1). Age-standardised incidence for women decreased from 369.0 (95% CI 361.7 to

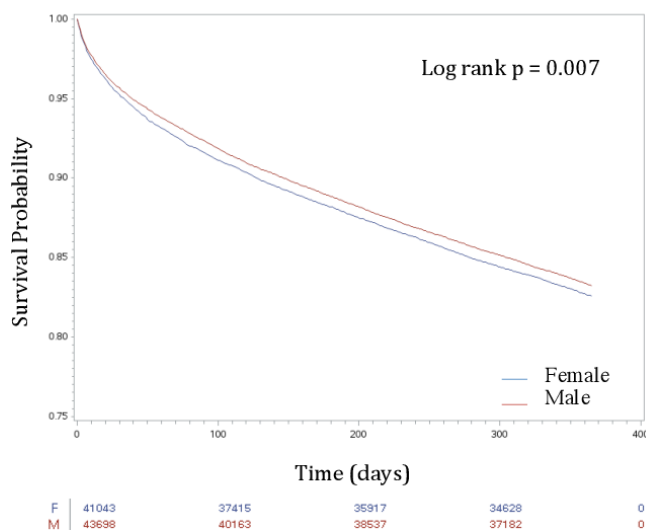
April 1, 1994 – March 31, 1999:



April 1, 1999 – March 31, 2004:



April 1, 2004 – March 31, 2009:



April 1, 2009 – March 31, 2014:

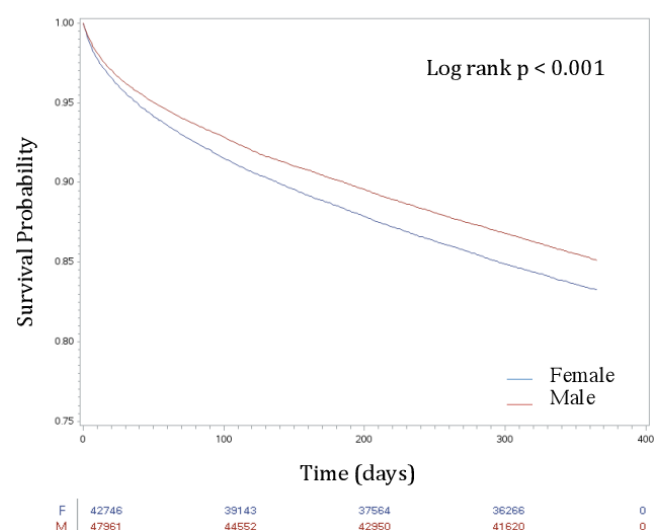


Figure 4 Kaplan-Meier curves representing temporal trends in 1-year survival after heart failure diagnosis, in each of the 5-year cohorts. The red line represents survival in men. The blue line represents survival in women.

376.3) per 100 000 population in 1994 to 205.8 (201.4–210.3) in 2013. For men, HF incidence decreased from 480.6 (470.6–490.9) per 100 000 in 1994 to 312.8 (306.8–318.9) in 2013 (online supplemental table 1).

Trends in mortality

One-year mortality occurred in 8319 (17.5%) women and 8238 (18.8%) men during the historical period; it occurred in 7156 (16.8%) women and 7138 (14.9%) men during the modern period (online supplemental table 2). These survival patterns are reinforced by the Kaplan-Meier survival curves in figure 2, as well as the stacked Kaplan-Meier curves in figure 3 that demonstrate an improvement in male survival (linear regression slope, -0.020 ; $p < 0.0001$) but relatively little change in female survival over time (linear regression slope, -0.010 ;

$p = 0.001$). Age-standardised 1-year mortality rates (AMR) also declined in both sexes but the magnitude of reduction was greater in men than in women. Men had higher AMR than women at most time points prior to 2009 (figure 4). Specifically, the female AMR was 10.4% (95% CI 9.1% to 12.0%) in 1994 and 8.5% (95% CI 7.5% to 9.5%) in 2013, representing a 19% reduction. Conversely, male AMR was 12.3% (95% CI 11.1% to 13.7%) in 1994 and 8.3% (95% CI 7.5% to 9.1%) in 2013, representing a 33% reduction (online supplemental table 3).

In the unadjusted analysis, female sex was protective against 1-year mortality in the historical cohort (unadjusted HR 0.93, 95% CI 0.90 to 0.95) but was associated with a higher risk of mortality (unadjusted HR 1.14, 95% CI 1.10 to 1.18) in the modern cohort. Adjusted analysis

Table 2 Evolution of multivariable risk factors of 1-year mortality over time

Variable	Adjusted HR (95% CI)			
	1994–1998	1999–2003	2004–2008	2009–2013
Age category				
40–64 years	Reference	Reference	Reference	Reference
65–74 years	1.56 (1.47 to 1.65)	1.64 (1.53 to 1.74)	1.66 (1.56 to 1.77)	1.54 (1.44 to 1.64)
75–84 years	2.26 (2.14 to 2.39)	2.36 (2.23 to 2.5)	2.48 (2.34 to 2.63)	2.58 (2.43 to 2.73)
>85 years	3.68 (3.48 to 3.90)	4.25 (4.00 to 4.51)	4.57 (4.31 to 4.85)	4.89 (4.62 to 5.18)
Female	0.85 (0.82 to 0.88)	0.89 (0.86 to 0.92)	0.91 (0.88 to 0.94)	0.97 (0.93 to 1.00)
Rural	0.94 (0.90 to 0.98)	1.00 (0.96 to 1.05)	0.99 (0.95 to 1.04)	1.01 (0.96 to 1.06)
Income quintile				
1 (low)	1.06 (1.01 to 1.11)	1.08 (1.03 to 1.14)	1.12 (1.06 to 1.18)	1.08 (1.02 to 1.14)
2	1.03 (0.98 to 1.08)	1.04 (0.99 to 1.09)	1.04 (0.99 to 1.10)	1.04 (0.98 to 1.09)
3	1.06 (1.01 to 1.11)	1.03 (0.98 to 1.09)	1.04 (0.98 to 1.10)	0.99 (0.94 to 1.04)
4	1.02 (0.97 to 1.08)	0.99 (0.94 to 1.05)	1.08 (1.03 to 1.14)	1.01 (0.95 to 1.06)
5 (high)	Reference	Reference	Reference	Reference
Benign hypertension	0.76 (0.74 to 0.79)	0.74 (0.72 to 0.77)	0.73 (0.70 to 0.76)	0.78 (0.75 to 0.81)
Complicated hypertension	0.93 (0.86 to 1.00)	0.95 (0.88 to 1.02)	0.87 (0.82 to 0.93)	0.77 (0.73 to 0.81)
Atrial fibrillation	0.89 (0.85 to 0.94)	0.92 (0.88 to 0.96)	0.96 (0.92 to 1.01)	0.99 (0.95 to 1.04)
Myocardial infarction	1.12 (1.06 to 1.18)	0.99 (0.94 to 1.05)	1.02 (0.97 to 1.06)	0.98 (0.94 to 1.03)
Valvular disease	1.00 (0.92 to 1.08)	0.98 (0.91 to 1.06)	0.97 (0.90 to 1.05)	0.93 (0.86 to 1.00)
Peripheral arterial disease	1.22 (1.16 to 1.29)	1.2 (1.14 to 1.27)	1.28 (1.20 to 1.37)	1.34 (1.24 to 1.44)
Cerebrovascular disease	1.29 (1.23 to 1.35)	1.25 (1.19 to 1.32)	1.23 (1.16 to 1.30)	1.20 (1.13 to 1.28)
Pulmonary circulatory disease	1.67 (1.43 to 1.96)	1.73 (1.52 to 1.97)	1.54 (1.38 to 1.71)	1.24 (1.12 to 1.37)
COPD/asthma	1.15 (1.12 to 1.19)	1.18 (1.14 to 1.22)	1.21 (1.17 to 1.25)	1.17 (1.13 to 1.21)
Alcohol abuse	1.16 (0.97 to 1.38)	1.34 (1.16 to 1.55)	1.51 (1.34 to 1.70)	1.62 (1.45 to 1.82)
Renal disease	2.04 (1.90 to 2.18)	1.86 (1.74 to 1.98)	1.73 (1.63 to 1.84)	1.67 (1.58 to 1.77)
Diabetes	1.25 (1.21 to 1.30)	1.19 (1.15 to 1.24)	1.13 (1.09 to 1.17)	1.13 (1.10 to 1.17)
Hypothyroidism	0.99 (0.91 to 1.07)	1.01 (0.93 to 1.09)	1.03 (0.93 to 1.13)	1.06 (0.96 to 1.17)
Liver disease	1.90 (1.72 to 2.10)	1.86 (1.68 to 2.06)	1.99 (1.79 to 2.21)	2.34 (2.12 to 2.59)
Dementia	1.74 (1.63 to 1.85)	1.97 (1.86 to 2.09)	2.03 (1.91 to 2.16)	1.96 (1.84 to 2.09)
Depression	1.15 (1.07 to 1.25)	1.18 (1.09 to 1.28)	1.19 (1.09 to 1.30)	1.16 (1.06 to 1.28)
Psychosis	1.29 (1.20 to 1.38)	1.36 (1.26 to 1.46)	1.14 (0.99 to 1.31)	1.43 (1.17 to 1.76)
Primary tumour	1.56 (1.49 to 1.64)	1.69 (1.60 to 1.77)	1.61 (1.53 to 1.70)	1.64 (1.56 to 1.73)
Metastatic malignancy	3.47 (3.22 to 3.74)	3.59 (3.33 to 3.87)	3.14 (2.90 to 3.41)	3.42 (3.17 to 3.70)
Paraplegia	1.36 (1.24 to 1.50)	1.33 (1.21 to 1.46)	1.52 (1.35 to 1.70)	1.53 (1.35 to 1.75)
Venous thromboembolism	1.07 (0.98 to 1.17)	1.18 (1.07 to 1.30)	1.16 (1.03 to 1.31)	1.40 (1.24 to 1.58)

Risk factors that have changed in magnitude between the historical (1994–1998) and modern (2009–2013) times are indicated in bold. COPD, chronic obstructive pulmonary disease.

demonstrated that the protective effect conferred by female sex had diminished over time (adjusted HR 0.85, 95% CI 0.82 to 0.88 in the historical and 0.97, 95% CI 0.93 to 1.00 in the modern cohort).

Table 2 lists the multivariable predictors of 1-year mortality in the modern cohort. Compared with the historical cohort, the mortality risk associated with age >75 years, liver disease, dementia and venous thromboembolism had increased while the risk associated with male sex,

complicated hypertension, diabetes, renal disease and pulmonary circulatory disease had diminished. Of note, alcohol abuse and venous thromboembolism emerged as new mortality risk factors while urban residence and MI were no longer risk factors.

Sex-specific mortality risk factors have evolved over time. Tables 3 and 4 illustrate the sex-specific HRs in the historical and modern cohorts, respectively. In the modern cohort, low income was associated with a higher

Table 3 Sex-specific risk factors of 1-year mortality in the historical cohort (1994–1998)

Variable	Women	Men	Multiplicative interaction
	HR (95% CI)	HR (95% CI)	P value*
Age category			
40–64 years	Reference	Reference	0.004
65–74 years	1.45 (1.32 to 1.59)	1.63 (1.51 to 1.75)	
75–84 years	2.06 (1.89 to 2.24)	2.42 (2.25 to 2.60)	
>85 years	3.53 (3.24 to 3.85)	3.70 (3.41 to 4.00)	
Benign hypertension	0.70 (0.67 to 0.73)	0.82 (0.79 to 0.86)	<0.001
Myocardial infarction	1.26 (1.16 to 1.36)	1.02 (0.95 to 1.10)	<0.001
COPD/asthma	1.10 (1.05 to 1.16)	1.20 (1.15 to 1.25)	0.01
Renal disease	2.26 (2.04 to 2.50)	1.90 (1.74 to 2.07)	0.01
Diabetes	1.31 (1.25 to 1.38)	1.20 (1.14 to 1.26)	0.01

*Multiplicative interaction terms were formed by multiplying sex by each of the covariates in the multivariable Cox proportional hazard model for 1-year mortality. Only significant interaction terms (ie, ones demonstrating sex-specific risk factors) were reported in this table. COPD, chronic obstructive pulmonary disease.

risk of mortality in men but not in women. Conversely, MI had a mild protective effect on men but not in women. In addition, women with peripheral arterial disease had a higher risk of death while men with COPD/asthma, dementia, primary and metastatic malignancies had a higher risk of mortality than women with similar comorbidities. Compared with the historical cohort, most sex-specific risk factors have evolved over time, with the exception of COPD/asthma.

Trends in HF hospitalisation

HF hospitalisations occurred in 5271 (13.4%) women and 5169 (14.4%) men within 1 year of HF diagnosis in the historical cohort. During the modern period, there

were 5420 (15.6%) HF hospitalisations in women and 5503 (13.8%) hospitalisations in men. Age-standardised HF hospitalisation rates declined in men (linear regression slope, -0.010 ; $p=0.0002$) but remained unchanged in women (linear regression slope, -0.005 ; $p=0.11$) during the 20-year period (online supplemental figure 1 and table 4). Specifically, male age-standardised HF hospitalisation rates were 11.4% (95% CI 10.1% to 12.9%) in 1994 and 9.1% (95% CI 8.2% to 10.1%) in 2013. Female rates were 9.7% (95% CI 8.3% to 11.3%) in 1994 and 9.8% (95% CI 8.6% to 11.0%) in 2013. The temporal trends in the cumulative incidence of HF hospitalisations are illustrated in (online supplemental figure 2).

Table 4 Sex-specific risk factors of 1-year mortality in the modern cohort (2009–2013)

Variable	Women	Men	Multiplicative interaction
	HR (95% CI)	HR (95% CI)	P value*
Income quintile			
1 (low)	1.02 (0.94 to 1.09)	1.15 (1.06 to 1.23)	0.001
2	0.97 (0.90 to 1.05)	1.10 (1.02 to 1.18)	
3	0.99 (0.92 to 1.07)	0.98 (0.91 to 1.06)	
4	1.01 (0.93 to 1.09)	1.01 (0.93 to 1.08)	
5 (high)	Reference	Reference	
Myocardial infarction	1.05 (0.98 to 1.12)	0.94 (0.89 to 1.00)	0.02
Peripheral arterial disease	1.48 (1.32 to 1.66)	1.25 (1.14 to 1.37)	0.02
COPD/asthma	1.12 (1.07 to 1.17)	1.23 (1.17 to 1.29)	0.005
Dementia	1.87 (1.72 to 2.02)	2.10 (1.92 to 2.30)	0.05
Primary tumour	1.44 (1.34 to 1.56)	1.79 (1.68 to 1.91)	<0.001
Metastatic malignancy	3.05 (2.75 to 3.38)	3.85 (3.49 to 4.26)	<0.001

ed in this table.

*Multiplicative interaction terms were formed by multiplying sex by each of the covariates in the multivariable Cox proportional hazard model for 1-year mortality. Only significant interaction terms (ie, ones demonstrating sex-specific risk factors) were reported in this table. COPD, chronic obstructive pulmonary disease.

DISCUSSION

This population-based study evaluated 352 329 individuals with a first-time diagnosis of HF from 1994 to 2013 in the ambulatory care setting. There are four main findings reported in this study: (1) HF mortality declined over time, (2) The reduction in mortality is greater in men than in women, (3) Rates of hospitalisation decreased for men but remained unchanged for women and (4) The incidence and significance of comorbidities associated with HF have changed over time, and suggest that women continue to experience a greater burden of comorbid disease when compared with men.

Trends in HF incidence and mortality

Population-based temporal trends in HF incidence and mortality have been previously reported across many cohorts, however, many of these studies have been limited to patients hospitalised with a diagnosis of HF or have not provided detailed, sex-stratified analyses. Temporal trends in the incidence and survival of HF patients were first reported by the Framingham group over a 50 year period from 1950 to 1999.³⁵ These authors reported that the incidence of HF had declined in women but not men, with improving survival in both sexes.³⁵ This pivotal study was followed by findings from a community-based cohort of 4537 patients from 1979 to 2000, which reported that although HF incidence remained unchanged for both sexes, mortality declined—with greater survival gains in men than women.³⁶ A recent study by our group demonstrated that among ambulatory Ontario residents from 2009 to 2013, the incidence of HF decreased more rapidly in men than women. At the same time, HF associated deaths and hospitalisations remain higher in women than men within a year of HF diagnosis.¹⁹ The present study extends these findings by demonstrating a continued disproportionate decrease in HF mortality for men compared with women from 1994 to 2013. Our findings corroborate with our previous study of HF incidence and 1-year mortality in rural and urban Eastern Ontario from 1994 to 2013.²⁴ They also corroborate the work of Tu *et al*⁷ who used similar administrative databases to report on the HF incidence and mortality of Ontario patients ≥ 20 years of age from 1997 to 2007. Tu *et al* evaluated both admitted and ambulatory HF patients and reported declines in HF incidence over this time period, a finding that was most evident in the older cohorts.⁷ Sex stratified mortality rates were not reported in this study. A recent study from Denmark demonstrated a decrease in HF incidence over time only in cohorts >50 years of age, but an increase in HF incidence in younger patients. Although detailed sex-specific outcomes were not provided, sex-stratified models showed similar trends in incidence and mortality over time with men having a higher incidence overall.⁸

The present study extends these observations by providing detailed sex-specific data on mortality trends over time. Our findings suggest that in Ontario, 1-year mortality rates have decreased over the past 20 years. However, this mortality reduction was greatest for men, and observed to a lesser extent for women. This translates

to the observation that women had better AMR than men in the first three temporal cohorts of this study (1994–2008); however, in the most recent cohort (2009–2013) we observed mortality to be higher in women than men for the first time. The basis for this sex-based difference is unclear but may be explained in part by the observation that women are more likely than men to have a diagnosis of HF with preserved ejection fraction (HFpEF), a disease for which there remains no evidence-based therapies which can improve survival, in contrast to the significant advances in medical therapy for HF with reduced ejection fraction (HFrEF).³⁷ In addition, female patients with HF have a higher comorbidity burden than their male counterparts. Complex comorbid conditions, coupled with atypical presentation of cardiac disease in women, may also have lead to delays in diagnosis and differences in management or response to medical therapy.³⁸ Further work is needed to determine whether the other sex-based differences in management, response to treatment or underlying pathophysiology remain to explain these sex-based trends in HF mortality over time.

Trends in HF hospitalisation

Rates of hospitalisation for HF decreased only for men in this time period. This is consistent with recent reports of sex and race differences in hospitalisation trends over a similar time period.^{19 39} It is possible that this sex-based difference may be due to death being a competing risk for hospitalisation in men, such that men with HF may suffer earlier deaths whereas women with HF survive to an older age and are more likely to become hospitalised. Alternatively, it is possible that the rates of hospitalisation in men and women reflect the underlying HF type, since men are more likely to have HFrEF (for which there are several treatments known to improve outcomes and decrease hospitalisation) while women are more likely to have HFpEF (for which there are no substantial evidence-based therapies). Nonetheless, the observed sex differences may also be attributed to the greater comorbidity burden in women, differences in social determinants of health or genetic or physiological differences that cannot be explained within the observational context of this study; all of these point to the need for further exploration to determine the adverse trends for mortality and hospitalisation in women with a diagnosis of HF.

Trends in HF comorbidities

Sex-based differences in comorbidities have been previously reported in hospitalised patients; women with HF are older and more likely than men to have comorbid hypertension, renal failure, obesity and depression. Men with HF are more often smokers, and tend to have more ischaemic heart disease, COPD and HFrEF.⁴⁰ Our study is the first to report on the relationship between sex and comorbidity in ambulatory HF patients over time. Compared with our historical cohort, our most recent cohort of patients demonstrates an overall increase in important comorbidities such as frailty, diabetes, renal

disease, MI, atrial fibrillation, COPD and hypertension. This has been observed in other population-based studies⁸ and speaks to the increased complexity of the patient with HF in the current era. The increased prevalence of these comorbidities over time was seen in both women and men. Certain comorbidities remained more common in women than in men in both the historical and recent cohorts; including depression, hypertension, advanced age, frailty, dementia and thyroid disease. Interestingly, frailty, chronic pulmonary disease, and metastatic cancer became more common in women than men in the recent cohort. Collectively these findings suggest that the comorbidity of the HF patient is increasing over time, and that women continue to experience a greater comorbidity burden than men. This observation may also explain in part the sex difference in mortality trends.

Important trends in the risk associated with these comorbidities were also observed. Hypertension conferred a greater protective effect in the modern era. This may actually reflect the known adverse prognosis associated with low blood pressure in HF.^{41 42} In addition, the risk associated with diabetes, renal disease and pulmonary circulatory disease has decreased over time. In addition, MI was no longer a mortality risk factor in the recent cohort when compared with the historical cohort. These changes over time may be due to significant advances in the medical management of these comorbidities, which have influenced overall survival.

Sex-based differences in the risks associated with certain comorbidities were also observed. In the most recent cohort, MI had a mild protective effect in men but not in women. This may be due to a lower detection rate of ischaemic heart disease in women due to atypical presentation,⁴³ which leads to missed management and poorer outcomes. Ischaemic heart disease is a leading cause of HF in both sexes. When diagnosed with ischaemic heart disease, women are less likely than men to undergo cardiac catheterisation and revascularisation; whether this is wholly attributed to the increased microvascular disease in women is not well understood.⁴⁴ In the most recent cohort, peripheral arterial disease was associated with a higher risk of mortality in women, while COPD, dementia and malignancy posed a greater risk of mortality in men. Whether these differences are clinically relevant, or help to explain the variability in mortality risk associated with HF, remains to be determined. There remains a significant knowledge gap on sex specific differences in epidemiology, pathophysiology, management and prognosis of comorbidities related to HF.⁴⁰ Such knowledge could determine if HF management should be targeted to specific sex-based comorbidities to improve outcomes and narrow the gap in mortality improvement between women and men.

Limitations and strengths

Our study has several limitations. First, cases of HF were identified in the ambulatory care setting based on the requirement of two claims for HF within 1 year. Although

this method may have led to an underestimate of HF, it has been validated previously and shown to improve the specificity of our case selection.^{7 9} Second, our algorithm for ascertainment of HF is validated in patients who are 40 years of age and older, thus limiting the generalisability of our findings. Third, information on ejection fraction was not available in the databases used, which precluded analyses in subtypes of HF based on ventricular function. Fourth, the diagnostic criteria for HFpEF have become more specific over time; whether this may influence incidence and prognosis cannot be determined from this study. Finally, cohort studies are by nature subjected to residual confounding. Despite these limitations, our study is the first to address the epidemiology of HF in a validated cohort of ambulatory patients, and one of the first to report on detailed sex-based outcome and comorbidity differences within a large universal health-care system, using the same entry criteria over a 20-year time period. The nature of our publicly funded health-care system allowed for complete analysis of all Ontario HF patients, which minimised selection bias and greatly improved the generalisability of our findings.

CONCLUSIONS

Over a 20-year window, we found an overall reduction in all-cause mortality in the year following HF diagnosis. However, there was a much larger reduction of mortality in men than in women, and HF hospitalisation rates have decreased for men but remained unchanged in women. Specifically, mortality and hospitalisation rates were higher in men than women at the start of the study period and were similar between sexes towards the end of this period. Female patients with HF continue to experience a greater burden of comorbidities than male patients with HF in the modern era. Further research should focus on the determinants of this disparity such as sex differences in medical and device management, to better characterise incidence and outcomes by HF type, and ways to reduce this gap in outcomes.

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