

1 **Alcohol Consumption is Associated with Poor Prognosis in Obese Patients with COVID-19: a**
2 **Mendelian Randomization Study using UK Biobank**

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22 **Financial support**

23 This work was supported in part by NIH grants: 3R01AA027456-02S1, P50AA024333,

24 U01AA021890 and U01AA026938 (LEN); RO1GM119174, R21AR71046, UO1DK061732,

25 RO1DK113196 and the Mikati foundation grant (SD); KL2TR002547 (DMR); K99AA026648

26 (KLP). XF was supported by a fellowship from the China Scholarship Council (File:201806280215)

27 and TM was supported by a JSPS Overseas Research Fellowship 201960331.

28

29 **Author contributions to manuscript**

30 X.F., L.E.N. and Z.L. formulated the project outline and analysis plan. X.F. performed all statistical
31 analyses and drafted the first manuscript draft under supervision from L.E.N. and D.M.R. L.E.N.,
32 D.M.R., S.D., K.L.P., X.W., and T.M. supervised the study. All authors contributed to interpretation
33 of results and writing of the final manuscript.

34

35 **Conflict of interest statement**

36 The authors declare no competing financial interests.

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54 **Abstract (word counts: 327)**

55 **Background:** Acute and chronic alcohol abuse have adverse impacts on both the innate and
56 adaptive immune response, which may result in reduced resistance to severe acute respiratory
57 syndrome coronavirus-2 (SARS-CoV-2) infection and promote the progression of coronavirus
58 disease 2019 (COVID-19). However, there are no large population-based data evaluating potential
59 causal associations between alcohol consumption and COVID-19.

60 **Method:** We conducted a Mendelian randomization study using data from UK Biobank to explore
61 the association between alcohol consumption and risk of SARS-CoV-2 infection and serious clinical
62 outcomes in patients with COVID-19. A total of 12,937 participants aged 50-83 who tested for
63 SARS-CoV-2 between 16 March to 27 July 2020 (12.1% tested positive) were included in the
64 analysis. The exposure factor was alcohol consumption. Main outcomes were SARS-CoV-2
65 positivity and death in COVID-19 patients. We generated weighted and unweighted allele scores
66 using three genetic variants (rs1229984, rs1260326, and rs13107325) and applied the allele scores
67 as the instrumental variables to assess the effect of alcohol consumption on outcomes. Analyses
68 were conducted separately for white participants with and without obesity.

69 **Results:** Of the 12,937 participants, 4,496 were never or infrequent drinkers and 8,441 were
70 frequent drinkers. (including 1,156 light drinkers, 3,795 moderate drinkers, and 3,490 heavy
71 drinkers). Both logistic regression and Mendelian randomization analyses found no evidence that
72 alcohol consumption was associated with risk of SARS-CoV-2 infection in participants either with
73 (OR=0.963, 95%CI 0.800-1.159; $q = 1.000$) or without obesity (OR=0.891, 95%CI 0.755-1.053; q
74 =.319). However, frequent drinking (HR=1.565, 95%CI 1.012-2.419; $q = .079$), especially heavy
75 drinking (HR=2.071, 95%CI 1.235-3.472; $q = .054$), was associated with higher risk of death in
76 patients with obesity and COVID-19, but not in patients without obesity. Notably, the risk of death
77 in frequent drinkers with obesity increased slightly with the average amount of alcohol consumed
78 weekly (HR=1.480, 95%CI 1.059-2.069; $q = .099$).

79 **Conclusions:** Our findings suggested alcohol consumption may had adverse effects on the
80 progression of COVID-19 in white participants with obesity, but was not associate with

81 susceptibility to SARS-CoV-2 infection.

82

83 **Keywords:** Alcohol consumption; COVID-19; Susceptibility; Mortality, Mendelian randomization,
84 UK Biobank.

85

86 **Introduction**

87 Coronavirus disease 2019 (COVID-19), caused by severe acute respiratory syndrome coronavirus-2
88 (SARS-CoV-2), is a highly contagious, fast-spreading, and life-threatening infectious disease[1]. So
89 far, it has spread to almost 200 countries and regions, infecting millions of people[2]. In its most
90 serious presentation, COVID-19 can progress rapidly into acute respiratory distress syndrome
91 (ARDS), multi-organ failure, and even death[3, 4]. Thus, identifying potential risk factors for
92 COVID-19 would be a substantial benefit to the public health.

93

94 In the midst of the COVID-19 pandemic, off-premise sales of alcohol have increased, according to
95 Nielsen data[5, 6]. Acute and chronic alcohol abuse have adverse impacts on both the innate and the
96 adaptive immune response[7-10] and alcohol consumption is associated with increased
97 susceptibility to pneumonia[9, 11], tuberculosis[9], respiratory syncytial virus (RSV) infection[9,
98 11], and acute respiratory distress syndrome (ARDS)[9, 12]. Chronic alcohol abuse also exacerbates
99 severity of influenza A virus infection by inhibiting influenza-specific CD8 T cell responses[10].

100 SARS-COV-2 is a positive-sense, single-stranded RNA (+ssRNA)[13]. Recent data from both
101 murine models of ethanol exposure and peripheral blood mononuclear cells (PBMCs) from patients
102 with alcohol-associated hepatitis (AH) indicate that signaling by viral ss/dsRNA is disrupted by
103 alcohol[14-16], analogous to impact of alcohol on signaling via bacterial products, such as
104 lipopolysaccharide[17]. Therefore, we hypothesized that alcohol consumption may result in reduced
105 resistance to SARS-COV-2 infection and promote the progression of COVID-19.

106

107 During the current COVID-19 pandemic, many misconceptions about the protective effects of

108 alcohol in preventing COVID-19 have appeared in social media[18, 19]. Although the World Health
109 Organization (WHO) and other public health authorities have stressed that alcohol consumption
110 does not destroy SARS-CoV-2 and may actually promote infection and accelerate disease
111 progression because of the immunosuppressive effects of alcohol[20], many people around the
112 world still believe that drinking alcohol helps prevent COVID-19[18, 19].

113

114 To date, there are no large population-based data evaluating the potential causal associations
115 between alcohol consumption and COVID-19. Therefore, in order to better understand the potential
116 impact of alcohol consumption on the risk of SARS-CoV-2 infection and the progression of
117 COVID-19, we applied the Mendelian randomization approach[21] to evaluate the causal
118 association among participants enrolled in a large national cohort, the UK Biobank, where detailed
119 information on COVID-19, alcohol consumption, other lifestyle factors, comorbidities, and
120 genotype data were rigorously collected.

121

122 **Methods**

123 This study is consistent with the Strengthening the Reporting of Observational Studies in
124 Epidemiology (STROBE) guideline. The study did not have a pre-registered or published analysis
125 plan. UK Biobank has obtained Research Tissue Bank (RTB) approval from its ethics committee
126 and this study was also approved by the Institutional Review Boards of the Cleveland Clinic (IRB
127 number:19582).

128

129 **Study population from UK Biobank**

130 The UK Biobank, a large population-based prospective cohort, recruited more than 500,000
131 participants aged 40-69 in 2006-2010 across the United Kingdom. A total of 13,502 participants
132 were tested for COVID-19 between March 16 and July 27, 2020. We excluded participants without
133 alcohol consumption data (n=67) and those without genotype data (n=515). Finally, 12937
134 participants were included in our study. Participants enrolled in UK Biobank have signed consent

135 forms.

136

137 **Exposure of interest**

138 The primary exposure of interest was alcohol consumption. Alcohol consumption data on
139 participants enrolled by the UK Biobank obtained through a self-completed touchscreen
140 questionnaire at the time of enrollment. Participants were asked about their current drinking status
141 (never, previous, current). For current drinkers, they were then asked about the frequency of intake
142 and their average weekly and monthly consumption (the unit is a standard drink) of a range of
143 beverage types (fortified wine, spirits, beer plus cider, red wine, champagne plus white wine).

144

145 According to the NIAAA's classification criteria[22], we classified participants as heavy drinkers
146 (>7 drinks per week for women; >14 drinks per week for men), moderate drinkers (4-7 drinks per
147 week for women; 4-14 drinks per week for men), light drinkers (3 drinks or fewer per week), and
148 never or infrequent drinkers (special occasions only or 1-3 times a month). We also classified heavy
149 drinkers, moderate drinkers and light drinkers as frequent drinkers and those who never drank or
150 drank infrequently as non-drinkers.

151

152 **Genetic data**

153 UK Biobank released genetic sequence data from 488,377 individuals genotyped for 847,441
154 genetic variants in July 2017. Participants were genotyped on either the UK BiLEVE genotyping
155 array (n = 49,950; 807,411 markers) or the UKB Axiom Array (n = 438,427; 825,927 markers).
156 After filtering for genetic variants available on both genotyping arrays and sample quality control
157 processes [23], 488,377 participants with 805,426 single nucleotide variants were available in the
158 release. Detailed information about genotyping and imputation in the UK Biobank has been
159 described previously[23].

160

161 Previous genome-wide association study (GWAS) of alcohol consumption in white British

162 participants of UK Biobank identified 14 single nucleotide polymorphisms (SNPs) associated with
163 alcohol consumption[24]. In addition, a study using data from UK Biobank and Genetic
164 Epidemiology Research in Adult Health and Aging (GERA) datasets identified 6 SNPs (Alcohol
165 Dehydrogenase 1B (ADH1B, rs1229984); Klotho Beta (KLB, rs13130794); Basic Transcription
166 Factor 3 Pseudogene 13 (BTF3P13, rs144198753); Glucokinase Regulator (GCKR, rs1260326);
167 Solute Carrier Family 39 Member 8 (SLC39A8, rs13107325); Dopamine Receptor D2 (DRD2,
168 rs11214609)) significantly associated with alcohol consumption[25].

169

170 The variant rs1229984 has been successfully used as an instrumental variable in the causal
171 estimation of alcohol consumption and assortative mating[26] and chronic widespread pain in the
172 participants of UK biobank[27]. However, the use of this single SNP may not adequately explain
173 genetic variance in alcohol consumption because of the low minor allele frequency (2.2%) of
174 rs1229984 in the UK Biobank[26]. In order to overcome the potential for weak instrument bias, we
175 generated an allele score utilizing three SNPs: rs1229984 (ADH1B), rs1260326 (GCKR), and
176 rs13107325 (SLC39A8), based on the results of the previous GWAS[25]. These three SNPs were
177 directly genotyped on both the UK BiLEVE and UKB Axiom Arrays and the missingness of these
178 SNPs in participants selected in this study less than 1%. Three additional SNPs (rs13130794 (KLB),
179 rs144198753 (BTF3P13), and rs11214609 (DRD2)) were identified in a previous GWAS[25] and
180 were excluded here because they were not directly genotyped (805,426 markers) on the UK
181 Biobank arrays. Detailed information about the three selected SNPs included in this study is shown
182 in **Table S1**. The allele score was calculated per individual as the weighted or unweighted sum of
183 the number of fast alcohol metabolizing alleles of each SNP, whereas the effect of each SNP on
184 alcohol consumption provided in the GERA database (**Table S1**) was used as weight for the
185 calculation of weighted allele score [25, 28].

186

187 **Other potential confounding risk factors for COVID-19**

188 A number of potential risk factors for COVID-19 were obtained through the touchscreen

189 questionnaire, inpatient hospital, death register, and genotype data: age at time of COVID-19 test,
190 sex, race (classed as white and non-white ethnic background), body mass index (BMI), blood type,
191 smoking status (no, only occasionally, most or all days), comorbidities (alcohol related diseases,
192 upper gastrointestinal diseases, chronic lower respiratory diseases, chronic heart diseases, diabetes
193 mellitus, dementia, liver cirrhosis and/or liver failure, renal failure, tumor, and acquired
194 immunodeficiency syndrome (AIDS)). In addition, BMI was categorized into four groups according
195 to the WHO classification[29]: underweight ($<18.5 \text{ kg/m}^2$), normal weight ($18.5\text{--}24.9 \text{ kg/m}^2$),
196 overweight ($25\text{--}29.9 \text{ kg/m}^2$), and obesity ($\geq 30 \text{ kg/m}^2$).

197

198 ICD-10 codes (International Classification of Diseases, Tenth Revision) were used to identify
199 comorbidities and the cause of death from medical records and death records.

200 Alcohol use disorder, alcohol liver diseases, alcohol pancreatitis, alcoholic gastritis, alcoholic
201 cardiomyopathy, alcoholic psychosis, alcoholic myopathy, alcoholic polyneuropathy, and
202 degeneration of the nervous system due to alcohol were uniformly classified as alcohol related
203 diseases. Upper gastrointestinal disease events were defined as participants with gastroesophageal
204 reflux disease (GERD), esophagitis, gastritis/duodenitis, or peptic ulcer. Chronic obstructive
205 pulmonary disease (COPD), asthma, emphysema, and bronchitis/bronchiectasis were uniformly
206 classified as chronic lower respiratory diseases events. Chronic cardiac events were defined as
207 participants with hypertensive, chronic ischemic heart disease, or heart failure. All disease-related
208 ICD 10 codes are shown in the **Table S2**.

209

210 **Ascertainment of outcomes**

211 The primary outcome was rate of positive SARS-CoV-2 tests, the secondary outcome was the
212 mortality in COVID-19 positive patients. Data on SARS-CoV-2 test results provided by the UK
213 Biobank covered England, but not Scotland and Wales. Follow-up for mortality was conducted to
214 June 27, 2020 through linkage from National Death Registries. In this study, we selected
215 participants from England in the database and defined the occurrence of outcomes as 0 =

216 non-occurrence, 1= occurrence. COVID-19 death event (n=287) was collected through latest death
217 record with ICD10 code U071 and other deceased COVID-19 positive patients without the
218 corresponding code.

219

220 **Statistical analysis**

221 All analyses were performed using Stata (Version 14.0; Stata Corp, College Station, TX). Weekly
222 alcohol consumption was natural log-transformed to meet parametric assumptions. Categorical
223 variables were tested for association using Chi-squared test or Fisher's exact test if more than 20%
224 of cells had expected frequencies < 5. For multiple comparisons correction, the false discovery rate
225 (FDR) was calculated using the Benjamini-Hochberg method[30] and an adjusted P value (*q-value*)
226 < 0.1 and an unadjusted P value < 0.05 were considered significant.

227

228 Alcohol consumption was classified in three ways: (1) A four-level exposure categorical variable
229 including never or infrequent drinkers, light drinkers, moderate drinkers, and heavy drinkers; (2) a
230 binary exposure variable comparing non-drinkers to frequent drinkers; and (3) a log-transformed
231 continuous variable of weekly self-reported alcohol consumption among frequent drinkers.

232

233 To reduce the potential confounding effects of factors other than alcohol intake on outcomes,
234 propensity score matching (PSM)[31] was applied to match non-drinkers and frequent drinkers
235 without replacement when we evaluated the associations between alcohol consumption and
236 outcomes. We included variables[32] previously reported to be associated with a higher risk of
237 COVID-19 as the matching factors for PSM. Factors including age, sex, BMI categories, current
238 smoking status, alcohol related diseases, asthma, emphysema, COPD, bronchitis/bronchiectasis,
239 esophagitis, gastritis/duodenitis, peptic ulcer, GERD, hypertensive, chronic ischemic heart disease,
240 heart failure, diabetes, dementia, renal failure, liver cirrhosis and/or liver failure, tumor and AIDS.

241

242 We applied two approaches to evaluate the relationship of alcohol consumption with outcomes:

243 logistic or Cox regression association analysis and Mendelian randomization analysis[21]. The
244 logistic regression was applied to evaluate the relationship between alcohol intake and the odds of
245 SARS-CoV-2 infection. The Cox regression analysis was applied to evaluate the relationship
246 between alcohol consumption and the risk of death in COVID-19 positive patients.

247

248 We used Mendelian randomization based on three assumptions: (1) the instrumental variable
249 (weighted or unweighted allele score) would not be associated with potentially confounding factors
250 of the risk of COVID-19; (2) the instrumental variable was significantly associated with exposure
251 factor-alcohol consumption; (3) the instrumental variable was only associated with the outcomes
252 through the exposure of interest. We evaluated these three assumptions using linear or logistic
253 regression.

254

255 For the Mendelian randomization analysis, the two-stage residual inclusion (2SRI) method[21] was
256 used to calculate the causal effect of alcohol consumption on the outcomes, using the weighted or
257 unweighted allele score as the instrumental variables. The 2SRI method requires adjustment due to
258 different classification methods and variable types (binary or continuous variable) of alcohol
259 consumption and outcomes. For the association of alcohol intake with the risk of SARS-CoV-2
260 infection, the adjusted 2SRI method was conducted. In the first stage, alcohol intake was associated
261 with the instrumental variable using a logistic regression model (binary exposure of non-drinkers
262 and frequent drinkers) or a linear regression (continuous variable of weekly alcohol consumption in
263 frequent drinkers). In the second stage, the outcome was fit using residuals from the first stage in a
264 logistic regression model. For the risk of death in COVID-19 positive patients, the residuals from
265 the first stage were included as covariates in a Cox regression model. Since the 2SRI method was
266 not suitable for evaluating the association between the four-level categorical variable of alcohol
267 consumption and outcomes, we only applied the logistic or Cox regression analysis for these
268 associations. In addition, we found obesity and race were associated with both the selected SNPs
269 and the risk of COVID-19 (all q -value<.10). These analyses were conducted separately for

270 participants with and without obesity, and in those of self-reported white ethnicity.

271

272 **Sensitivity analysis**

273 We performed two sensitivity analyses to evaluate the impact of specific subgroup of participants

274 and different definitions of outcomes, respectively, on our analysis. First, in order to clarify the

275 relationship between alcohol intake and outcomes in overweight, but not obese, patients, we

276 selected overweight participants for the sensitivity analysis. Second, because some COVID-19

277 positive patients were admitted to the intensive care unit (ICU) before they died or recovered and

278 there was no data on life support for patients other than those in ICU, we considered ICU admission

279 and death as serious clinical events and applied logistic regression and Mendelian randomization

280 analyses to evaluate the association between alcohol intake and the risk of severe clinical outcomes.

281

282 **Result**

283

284 **Characteristics of participants**

285 A total of 12,937 participants who had been tested for COVID-19 (from March 16 to July 27, 2020)

286 were selected for our analysis according to the inclusion and exclusion criteria. Of these participants,

287 1,570 (12.1%) tested positive for COVID-19 and 11,367 (87.9%) tested negative. Regarding the

288 drinking status of these participants, 4,486 (34.8%) were never or infrequently drinkers, 1156 (8.9%)

289 were light drinkers, 3,795 (29.3%) were moderate drinkers, and 3,490 (27.0%) were heavy drinkers.

290 As shown in **Table S3**, significant differences were observed in age, sex, race, blood type, current

291 smoking, and comorbidities other than dementia, tumor, and AIDS between the non-drinkers and

292 frequent drinkers (all $q < .10$). Frequent drinkers tended to have a higher proportion of participants

293 who were older than 65 years (72.0% vs. 69.3%), white (96.6% vs. 86.1%), and of normal weight

294 (28.6 % vs. 24.2%) compared to non-drinkers. In contrast to the known deleterious health effects of

295 heavy drinking [7-10], frequent drinkers had a lower rate of comorbidities than non-drinkers (all q

296 $< .10$). Since the participants in the UK Biobank are between the ages of 50 and 83, we hypothesized

297 that there may be survivor bias between exposure factor (alcohol consumption) and multiple
298 complications. Therefore, to reduce the potential for survivor bias and the effects of confounding
299 factors on outcomes, all analyses were performed in the PSM cohorts and the 2SRI method applied
300 to calculate the causal effect of alcohol consumption on the outcomes [33].

301

302 **Instrumental variable associations**

303 Characteristics of participants by rs1229984, rs1260326, and rs13107325 genotypes are shown in
304 **Table S4**. Since the number of participants with alleles rs1229984 and rs13107325, associated with
305 rapid ethanol metabolism, was relatively small, we combined participants with one or two rapid
306 metabolism alleles for comparison. Participants with one or two rapid metabolism alleles compared
307 with those with two reference alleles tended to have a higher proportion of white participants and
308 patients with obesity (all $q < .10$). No significant differences in age, sex, blood types, current
309 smoking, or comorbidities were observed (all $q < .10$).

310

311 White Participants with one or two rapid metabolism alleles were less likely to be heavy-drinkers
312 and on average drank less alcohol if they were frequent drinkers (**Table 1 and Figure S1**). These
313 associations are consistent, but the difference between those with 2 or 3 more alleles is greater than
314 those between 0 and 1 (Figure 1). Associations of rs1229984, rs1260326, rs13107325 genotypes,
315 and the weighted or unweighted allele score with alcohol consumption were similar and significant
316 (all $P < .05$). The weighted allele score (F-test=26.289) was more strongly related to alcohol
317 consumption compared with other single SNPs and the unweighted allele score. Furthermore, we
318 did not find an association between SNPs or allele scores and the odds of SARS-CoV-2 infection,
319 the risk of severe clinical outcomes or death in patients with COVID-19 (**Table S5**).

320

321 **Observational associations and instrumental variable associations**

322 The SARS-CoV-2 test positivity rate in white participants was 11.4% (1,368/11,982) and the
323 mortality rate of white patients with COVID-19 was 18.9% (258/1,368). After 1:1 PSM, 1,445

324 non-drinkers and matched frequent drinkers from the obese cohort and 2,305 non-drinkers and
325 matched frequent drinkers from the non-obese cohort were selected for evaluating the associations
326 between alcohol consumption and the odds of SARS-CoV-2 infection. In addition, 187 non-drinkers
327 and matched frequent drinkers from the obese cohort and 295 non-drinkers and matched frequent
328 drinkers from the non-obese cohort were selected for evaluating the association between alcohol
329 consumption and the risk of worse clinical outcomes of COVID-19. The association of alcohol
330 consumption with the odds of SARS-CoV-2 infection (**Figure 1**) and the risk of death in COVID-19
331 positive patients (**Figure 2**) are shown.

332

333 *Alcohol and risk of SARS-CoV-2 infection*

334 For the primary outcome, both logistic regression and Mendelian randomization analyses suggested
335 that alcohol consumption within all three classifications of drinkers --the four-level categorical
336 variable, the binary variable (non-drinkers and frequent drinkers), and the continuous variable of
337 weekly alcohol intake in frequent drinkers was not associated with the risk of SARS-CoV-2
338 infection (All $q > .10$, **Figure 1**). No association was detected in white participants either with or
339 without obesity. In addition, there was no association between average weekly alcohol consumption
340 and the risk of SARS-CoV-2 infection in either obese (OR=1.065, 95%CI 0.916-1.238; $P=.412$) and
341 non-obese (OR=0.955, 95%CI 0.864-1.057; $P=.375$) cohort before PSM from Mendelian
342 randomization analysis using the weighted allele score.

343

344

345 *Alcohol and risk of death in COVID-19 positive patients*

346 COVID-19 positive patients who were heavy drinkers with obesity had a higher risk of death
347 (HR=2.071, 95%CI 1.235-3.472; $q =.054$, **Figure 2A**). Both Cox regression (HR=1.566, 95%CI
348 1.013-2.421; $q =.066$) and Mendelian randomization analyses using unweighted allele score
349 (HR=1.564, 95%CI 1.012-2.418; $q =.057$) or weighted allele score (HR=1.565, 95%CI 1.012-2.419;
350 $q =.079$) identified that COVID-19 positive patients with obesity who reported consuming alcohol

351 weekly were more likely to die compared with those drinking none or infrequently. In addition, we
352 found higher alcohol consumption in frequent drinkers resulted in higher risk of death when
353 analyzed by either Cox regression (HR=1.460, 95%CI 1.047-2.034; $q = .058$) or Mendelian
354 randomization analyses using the unweighted allele score (HR=1.457, 95%CI 1.045-2.032; $=.078$)
355 or the weighted allele score (HR=1.480, 95%CI 1.059-2.069; $q = .099$). However, these associations
356 did not exist in non-obese patients with COVID-19 ($q > .10$, **Figure 2B**). Consistent results
357 regarding the relationship between average weekly alcohol intake and the risk of death was found in
358 both obese (HR=1.418, 95%CI 1.051-1.914; $P = .022$) and non-obese (HR=0.952, 95%CI
359 0.755-1.201; $P = .680$) cohorts before PSM from Mendelian randomization analysis using weighted
360 allele score. As shown in **Figure 3**, Kaplan–Meier survival plots illustrated that heavy drinkers with
361 obesity had a higher mortality than non-drinkers (Log rank P value= $.027$), which was not observed
362 in non-obese patients with COVID-19 (Log rank P value= $.471$).

363

364 **Sensitivity analysis**

365 *Association between alcohol consumption and outcomes in overweight but not obese patients*

366 A total of 4,869 overweight white participants were tested for SARS-CoV-2. The test positivity rate
367 was 11.7% and the mortality rate of overweight patients with COVID-19 was 18.1% (103/568).

368 Consistent with previous results for patients without obesity (**Figure 1-2**), alcohol consumption was
369 not associated with the risk of SARS-CoV-2 infection or death in overweight COVID-19 positive
370 patients ($q > .10$, **Table S6**).

371

372 *Risk of severe clinical outcomes in COVID-19 positive patients*

373 ICU admission and death in COVID-19 positive patients were grouped together as having severe
374 clinical outcomes. Because some of the patients were admitted to ICU before they were diagnosed
375 with the SARS-CoV-2 infection, we couldn't use the Cox regression model to analyze the
376 association between alcohol consumption and severe clinical outcomes in the COVID-19 positive
377 cohort without time-to-event data. Heavy drinkers with obesity had a higher likelihood of admission

378 to ICU and death compared to non-drinkers (OR=2.432, 95%CI 1.345-4.397; $q = .027$, **Figure 4A**).
379 Both logistic regression (OR=1.766, 95%CI 1.114-2.801; $q = .072$) and Mendelian randomization
380 analyses using unweighted allele score (OR=1.762, 95%CI 1.11-2.794; $q = .048$) or weighted allele
381 score (OR=1.710, 95%CI 1.077-2.715; $q = .052$) identified that COVID-19 positive patients with
382 obesity who reported consuming alcohol weekly were more likely to suffer severe clinical outcomes
383 compared with those drinking none or infrequently. In addition, we found that the likelihood of
384 serious clinical outcomes in frequent drinkers with obesity slightly increased with the average
385 amount of alcohol consumed weekly based on the result of Mendelian randomization analysis using
386 unweighted allele score (OR=1.020, 95%CI 1.003-1.038, $q = .095$) and weighted allele score
387 (OR=1.018, 95%CI 1.001-1.036; $q = .054$, **Figure 4A**). However, these associations did not exist in
388 non-obese patients with COVID-19 ($q > .10$, **Figure 4B**).

389

390 **Discussion**

391

392 Using the UK Biobank cohort, we investigated whether alcohol consumption increased
393 susceptibility to SARS-CoV-2 infection among 12,937 white participants who have been tested for
394 SARS-COV-2, as well as whether there were worse outcomes among 1,570 patients with
395 COVID-19 using regression analyses and Mendelian randomization analysis. We found that alcohol
396 consumption did not increase susceptibility to SARS-CoV-2 infection; however, frequent drinking,
397 especially heavy drinking, was associated with worse outcomes of COVID-19 in patients with
398 obesity, but not non-obese patients. Notably, the risk of worse clinical outcomes in frequent drinkers
399 with obesity increased slightly with the average amount of alcohol consumed weekly.

400

401 **Possible explanations for an interaction between alcohol consumption and worse outcomes of** 402 **COVID-19 in patients with obesity**

403 According to our findings, frequent drinking was associated with poor outcomes of COVID-19 in
404 patients with obesity, but not non-obese, patients, suggesting the potential interactions between

405 alcohol consumption and obesity to cause rapid progression of COVID-19. Consistent with a
406 previous report[34], our study found that patients with obesity were more susceptible to
407 SARS-CoV-2 infection (OR=1.260, 95%CI 1.092-1.453; $P=.002$) and had a higher risk of death
408 (HR=1.561, 95%CI=1.122-2.171; $P=.008$) from COVID-19. The high expression of
409 angiotensin-converting enzyme 2 receptor (ACE2) in adipocytes of patients with obesity may
410 promote the entry of SARS CoV2 into host cells and turn adipose tissue into a potential target and
411 reservoir of SARS CoV2[35]. Moreover, adipose tissue, a major source of pro-inflammatory
412 chemokines, cytokines, and adipokines, plays an important role in mediating inflammatory
413 responses. Patients with obesity have higher concentrations of circulating tumor necrosis
414 factor-alpha (TNF- α), interleukin-6, and C-reactive protein (CRP); this low grade chronic
415 inflammatory state may be involved in initiating cytokine storms in patients with COVID-19[34].

416

417 Alcohol consumption is associated with an increased risk of ARDS[12], which is a common
418 manifestation of severe COVID-19. This is likely due to the dysregulation of both immune and
419 non-immune host defense mechanisms in the airways, resulting in alveolar epithelial barrier
420 dysfunction and alveolar macrophage immune dysregulation in response to heavy alcohol
421 consumption[9, 12]. However, the mechanisms for the potential interaction between alcohol and
422 obesity in the progression of COVID-19 are not well understood. Previous work has found that
423 chronic ethanol exposure in murine models impacts adipose tissue, phenocopying obesity in many
424 aspects. Chronic ethanol feeding decreases glucose uptake by adipose tissue, increases immune cell
425 infiltration and expression of inflammatory cytokines and adipokines[36]. Data from pre-clinical
426 models also demonstrates that binge drinking and obesity synergistically induces steatohepatitis and
427 fibrosis in the liver of mice via the induction of hepatic chemokines, inflammatory cytokines and
428 neutrophil infiltration[37, 38]. Taken together, alcohol intake and obesity may affect the progression
429 of COVID-19 via impact on adipose tissue and inflammatory responses; the additive and/or
430 synergistic effects of obesity and alcohol may lead to a significant deterioration of COVID-19
431 patients. However, mechanisms for the potential interactions between alcohol and obesity in the

432 progression of COVID-19 are not well understood.

433

434 **Strengths and Limitations**

435 One of main advantages of this study is the application of Mendelian randomization to evaluate the
436 causal relationships between alcohol consumption and the risk of SARS CoV2 infection and the
437 severity of COVID-19 outcomes. To our knowledge, the associations between alcohol and risk of
438 COVID-19 have not been previously explored using Mendelian randomization. We generated
439 weighted and unweighted allele scores using three SNPs rs1229984 (ADH1B), rs1260326 (GCKR),
440 and rs13107325 (SLC39A8) and applied the weighted and unweighted allele scores as the
441 instrumental variables to assess the effect of alcohol consumption on outcomes. These three SNPs
442 were randomly distributed among the white participants and were not associated with other
443 confounders that may affect outcomes. However, rs1229984, rs1260326, and rs13107325 were
444 significantly correlated with obesity. This issue was well addressed by our separate analysis based
445 on whether the white participants were obese or not. In addition, the Mendelian randomization
446 approach more robustly handles measurement error and reverse causality compared to traditional
447 observational approaches. Another major advantage of our study is the detailed data in a larger
448 population-based prospective cohort including genotype data, alcohol consumption, and potential
449 confounding risk factors. Almost all the confounding factors that have been reported to be
450 associated with the risk of COVID-19 are available in the UK Biobank. Therefore, we were able to
451 balance these factors in non-drinkers and frequent drinkers using PSM to better assess the specific
452 effect of alcohol consumption on the risk of COVID-19 in the traditional observational study and
453 Mendelian randomization study.

454

455 This study also has several limitations. First, since the majority of participants in the study were
456 British and Irish descents (93%) and only 195 non-white participants were infected with SARS
457 CoV2, we could not investigate the impact of alcohol consumption on the risk of worse outcomes in
458 non-white participants with COVID-19 using Mendelian randomization analysis. However,

459 according to the results of the Chi-square test, we found heavy drinkers had a higher mortality than
460 non-heavy drinkers (57.1% vs. 18.0%, $P=.038$) in the non-white group with obesity, which was not
461 observed in non-obese group (20.0% vs. 8.0%, $P=0.220$). Second, we are unable to assess the effect
462 of changes in the alcohol consumption since subjects ($n=501,608$) were enrolled in the UK Biobank
463 in 2006-2010. Although the self-reported alcohol consumption data of some participants were
464 updated in 2012 ($n=20,336$), 2014 ($n=48,340$) and 2019 ($n=3,081$), the number of these patients
465 with COVID-19 was relatively small and the corresponding statistical analysis could not be
466 performed. Only 5 COVID-19 positive patients had their alcohol consumption data updated in 2019
467 and there was no significant change in their drinking patterns (data not shown). Finally, due to the
468 limitations of the data in the UK Biobank, we could not assess the impact of alcohol consumption
469 on the symptoms and some severe clinical outcomes (requirement of oxygen therapy, glucocorticoid
470 therapy, and administration of invasive ventilation) of patients. To minimize this limitation, we
471 classified ICU admission and death as the severe clinical outcomes of COVID-19 positive patients
472 and found heavy drinkers had a higher risk of admission to ICU and death in patients with obesity
473 compared to non-drinkers.

474

475 **Conclusions**

476 Despite the limitations, our study was the first to find that alcohol consumption, especially heavy
477 drinking, is associated with a higher risk of suffering worse COVID-19 clinical outcomes in patients
478 with obesity through both traditional regression analyses and Mendelian randomization analyses. In
479 addition, alcohol consumption was not associated with either increased or decreased risk of SARS
480 CoV2 infection. Our findings could help people understand the relationship between alcohol
481 consumption and COVID-19, especially those who may drink excessively in the mistaken belief
482 that alcohol consumption reduces the risk of SARS CoV2 infection[18, 19]. Moreover, due to the
483 possible interactions between alcohol consumption and obesity in the progression of COVID-19,
484 physicians may need to adjust and develop appropriate management and treatment strategies for
485 COVID-19 positive patients who consume alcohol and are obese.

486

487 **Acknowledgments**

488 This study has been conducted using the UK Biobank Resource (application number 59473). The
489 views expressed are those of the authors and not necessarily those of the National Institute for
490 Health Research (NIHR) or the Department of Health and Social Care.

491

492 **Data availability**

493 This study used data from the UK Biobank (application number 59473). For details please contact
494 access@ukbiobank.ac.uk. All other data are contained in the article and its supplementary
495 information or available upon reasonable request.

496

497 **Abbreviation**

498 COVID-19, coronavirus disease 2019; SARS-CoV-2, severe acute respiratory syndrome
499 coronavirus-2; HR, hazard ratio; OR, odds ratio; CI, confidence interval; N, number of participants;
500 ADH1B, alcohol dehydrogenase 1B; BMI, body mass index; GERD, gastroesophageal reflux
501 disease; COPD, chronic obstructive pulmonary disease; AIDS, acquired immunodeficiency
502 syndrome; ICU, intensive care unit; ARDS, acute respiratory distress syndrome; RSV; respiratory
503 syncytial virus; FDR, false discovery rate; SNP, single nucleotide polymorphism; PSM, propensity
504 score matching; RAF, risk allele frequency; GERA, Genetic Epidemiology Research in Adult
505 Health and Aging; ADH1B, alcohol dehydrogenase 1B; GCKR, glucokinase regulator; SLC39A8,
506 solute carrier family 39 member 8.

507

508 **Table legend**

509 **Table 1.** Association of genetic variations of ADH1B/SLC39A8/GCKR with alcohol consumption
510 in white participants.

511

512 **Figure legend**

513 **Figure 1.** Logistic regression and Mendelian randomization analyses of the association of alcohol
514 consumption with the risk of SARS-CoV-2 infection in white participants with (A) and without
515 obesity (B).

516 Analyses were performed in PSM cohort. Matching factors for PSM including age, sex, BMI
517 categories, current smoking status, alcohol related diseases, asthma, emphysema, COPD,
518 bronchitis/bronchiectasis, esophagitis, gastritis/duodenitis, peptic ulcer, GERD, hypertensive,
519 chronic ischemic heart disease, heart failure, diabetes, dementia, renal failure, liver cirrhosis and/or
520 liver failure, tumor and AIDS. q-value was calculated by false discovery rate (FDR) method.

521 Abbreviation: OR, odds ratio; HR, hazard ratio; CI, confidence interval; PSM, propensity score
522 matching; BMI, body mass index; GERD, gastroesophageal reflux disease; COPD, chronic
523 obstructive pulmonary disease; AIDS, acquired immunodeficiency syndrome.

524 **Figure 2.** Cox regression and Mendelian randomization analyses of the association of alcohol
525 consumption with the risk of death in white COVID-19 positive patients with (A) and without
526 obesity (B).

527 Analyses were performed in PSM cohort. Matching factors for PSM including age, sex, BMI
528 categories, current smoking status, alcohol related diseases, asthma, emphysema, COPD,
529 bronchitis/bronchiectasis, esophagitis, gastritis/duodenitis, peptic ulcer, GERD, hypertensive,
530 chronic ischemic heart disease, heart failure, diabetes, dementia, renal failure, liver cirrhosis and/or
531 liver failure, tumor and AIDS.

532 q-value was calculated by false discovery rate (FDR) method.

533 Abbreviation: OR, odds ratio; HR, hazard ratio; CI, confidence interval; PSM, propensity score
534 matching; BMI, body mass index; GERD, gastroesophageal reflux disease; COPD, chronic
535 obstructive pulmonary disease; AIDS, acquired immunodeficiency syndrome.

536 **Figure 3.** Survival probability of COVID-19 positive patient based on different drinking status in
537 obese (A) and non-obese (B) groups.

538 **Figure 4.** Logistic regression and Mendelian randomization analyses of the association of alcohol

539 consumption with the risk of severe clinical outcomes in white COVID-19 positive patients with (A)
540 and without obesity (B).

541 Analyses were performed in PSM cohort. Matching factors for PSM including age, sex, BMI
542 categories, current smoking status, alcohol related diseases, asthma, emphysema, COPD,
543 bronchitis/bronchiectasis, esophagitis, gastritis/duodenitis, peptic ulcer, GERD, hypertensive,
544 chronic ischemic heart disease, heart failure, diabetes, dementia, renal failure, liver cirrhosis and/or
545 liver failure, tumor and AIDS.

546 q-value was calculated by false discovery rate (FDR) method.

547 Abbreviation: OR, odds ratio; HR, hazard ratio; CI, confidence interval; PSM, propensity score
548 matching; BMI, body mass index; GERD, gastroesophageal reflux disease; COPD, chronic
549 obstructive pulmonary disease; AIDS, acquired immunodeficiency syndrome.

550

551 Supporting information

552 **Table S1.** Detailed information about the genetic variations of ADH1B/ SLC39A8/ GCKR included
553 in this study.

554 **Table S2.** Diseases diagnosis codes used by the UK Biobank.

555 **Table S3.** Association of observed confounders with alcohol consumption.

556 **Table S4.** Characteristics of participants by rs1229984, rs1260326, and rs13107325 genotypes.

557 **Table S5.** Association of genetic variations of ADH1B/SLC39A8/GCKR with outcomes of interest
558 in white participants.

559 **Table S6.** Logistic/Cox regression and Mendelian randomization analyses of the associations of
560 alcohol consumption with the risk of SARS-CoV-2 infection and the risk of death of COVID-19 in
561 white participants who were overweight but not obese.

562 **Figure S1.** Association of combined ADH1B, SLC39A8, GCKR fast-allele score with average
563 alcohol consumption levels (standard drink/weekly) in whole participants (A) and percentages of
564 heavy-drinkers (B) in whole participants by unweighted allele score. The amount of alcohol
565 consumed by non-drinkers was defined as zero.

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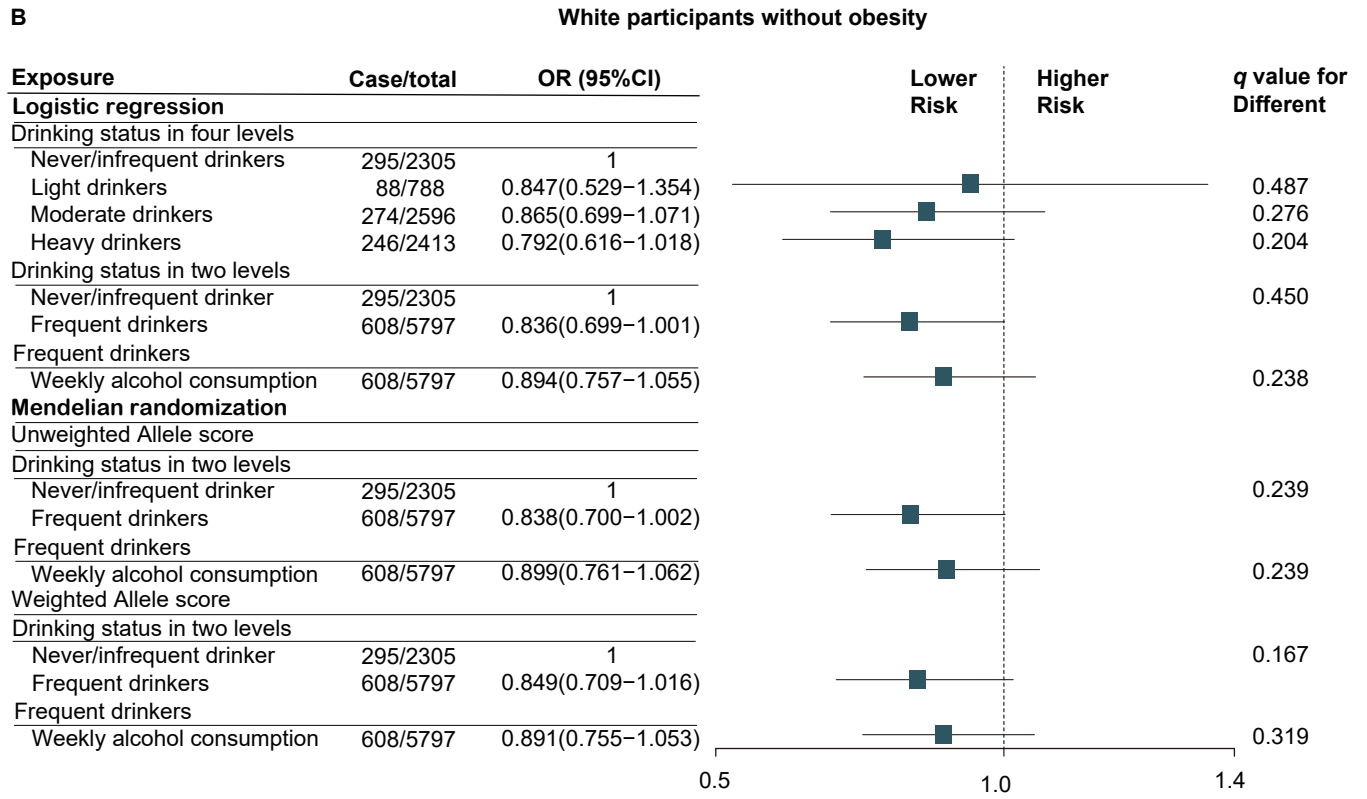
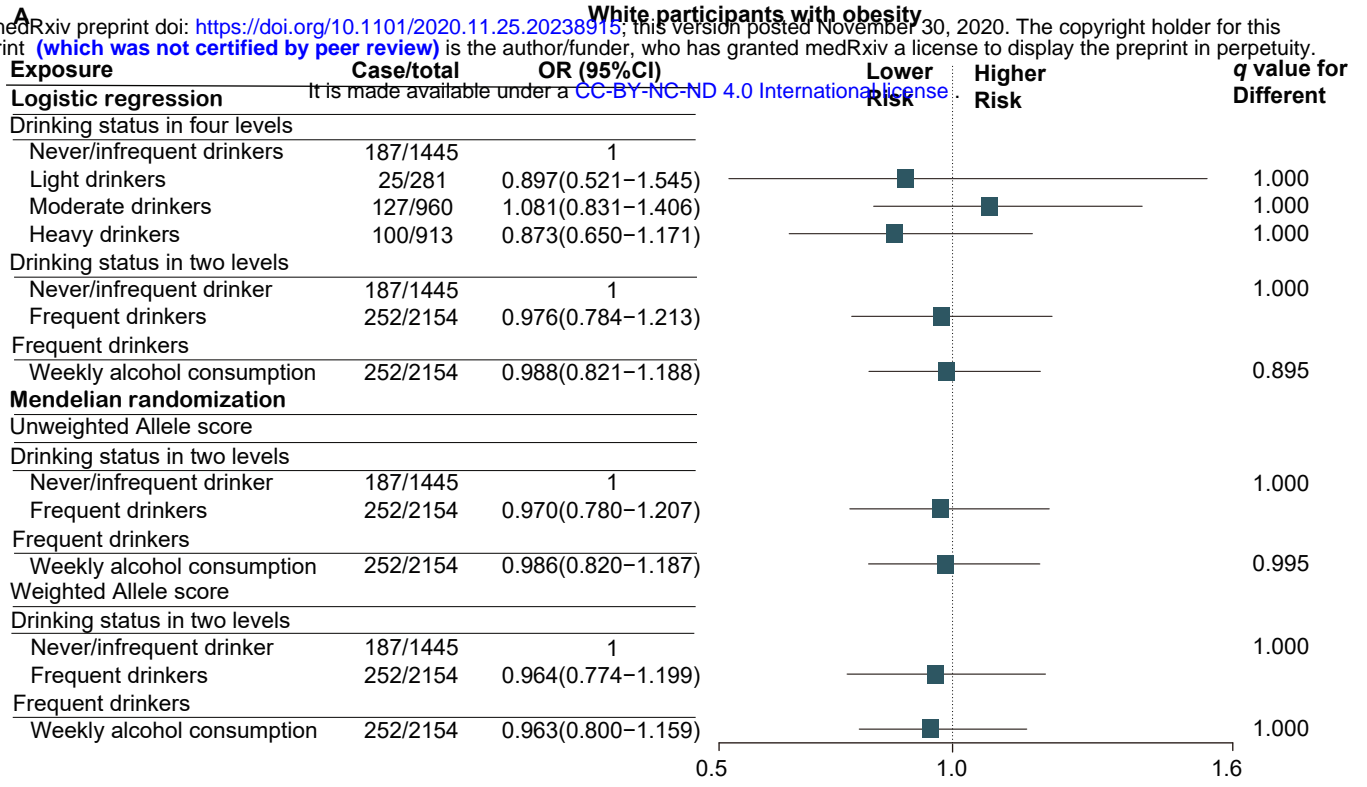
677 **Table 1 Association of genetic variations of ADH1B/SLC39A8/GCKR with alcohol consumption in white**
 678 **participants**

| Genetic variations | Alcohol consumption (standard drink/weekly) in frequent drinkers (95% CI) (n=8131) | OR of being a heavy drinker in the whole cohort (95% CI) (n=11982) |
|---|---|---|
| ADH1B one or two fast alleles vs. none | -2.777(-4.081- -1.473) | 0.525(0.425-0.647) |
| F-test | 17.424 | — |
| P-value | <0.001 | <0.001 |
| SLC39A8 one or two fast alleles vs. none | -0.919(-1.711- -0.127) | 0.874(0.778-0.981) |
| F-test | 5.168 | — |
| P-value | <0.001 | 0.022 |
| GCKR one or two fast alleles vs. none | -0.789(-1.360- -0.218) | 0.917(0.845-0.996) |
| F-test | 7.333 | — |
| P-value | <0.001 | 0.039 |
| Unweighted allele score | -0.726(-1.066- -0.385) | 0.891(0.848-0.936) |
| F-test | 17.465 | — |
| P-value | <0.001 | <0.001 |
| Weighted allele score | -15.068(-20.828- -9.307) | 0.048(0.020-0.118) |
| F-test | 26.289 | — |
| P-value | <0.001 | <0.001 |

679 **F-test** was calculated by linear regression analysis.

680 **Abbreviation:** OR, odds ratio; CI, confidence interval; BMI, body mass index.

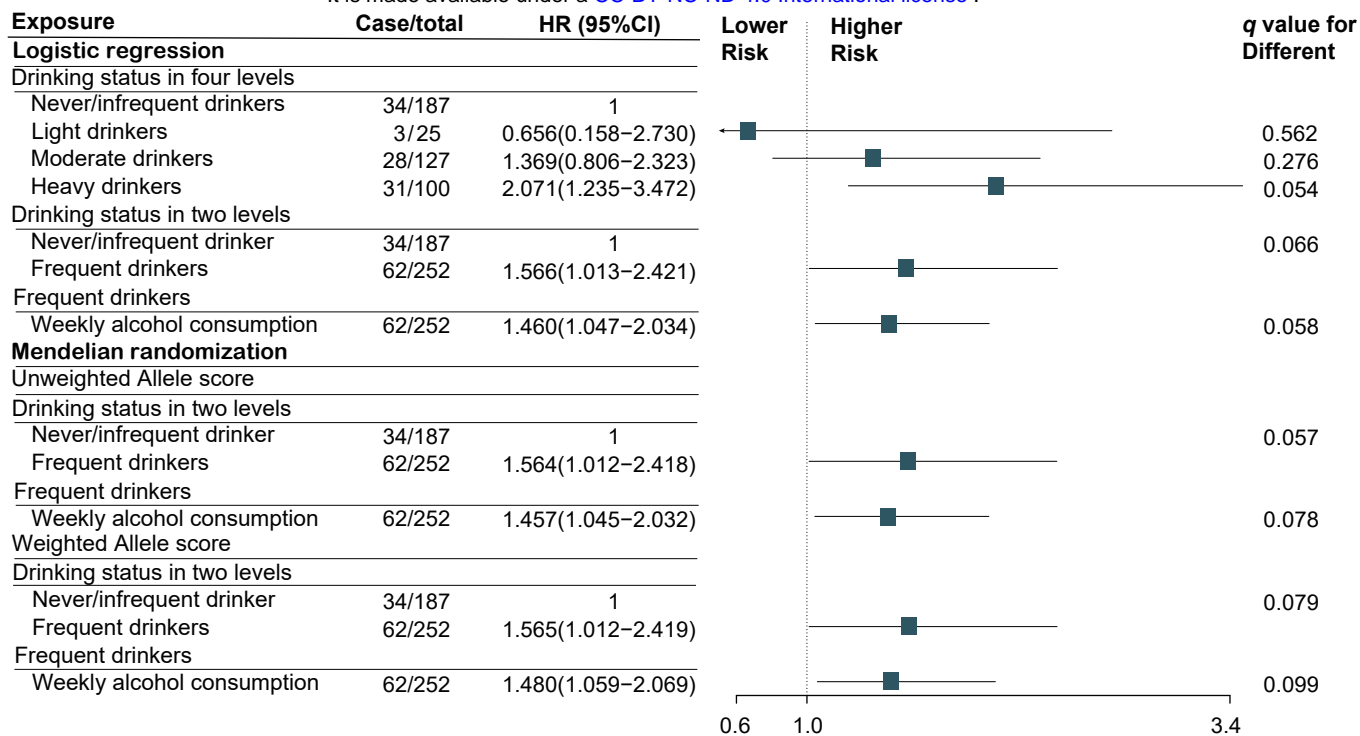
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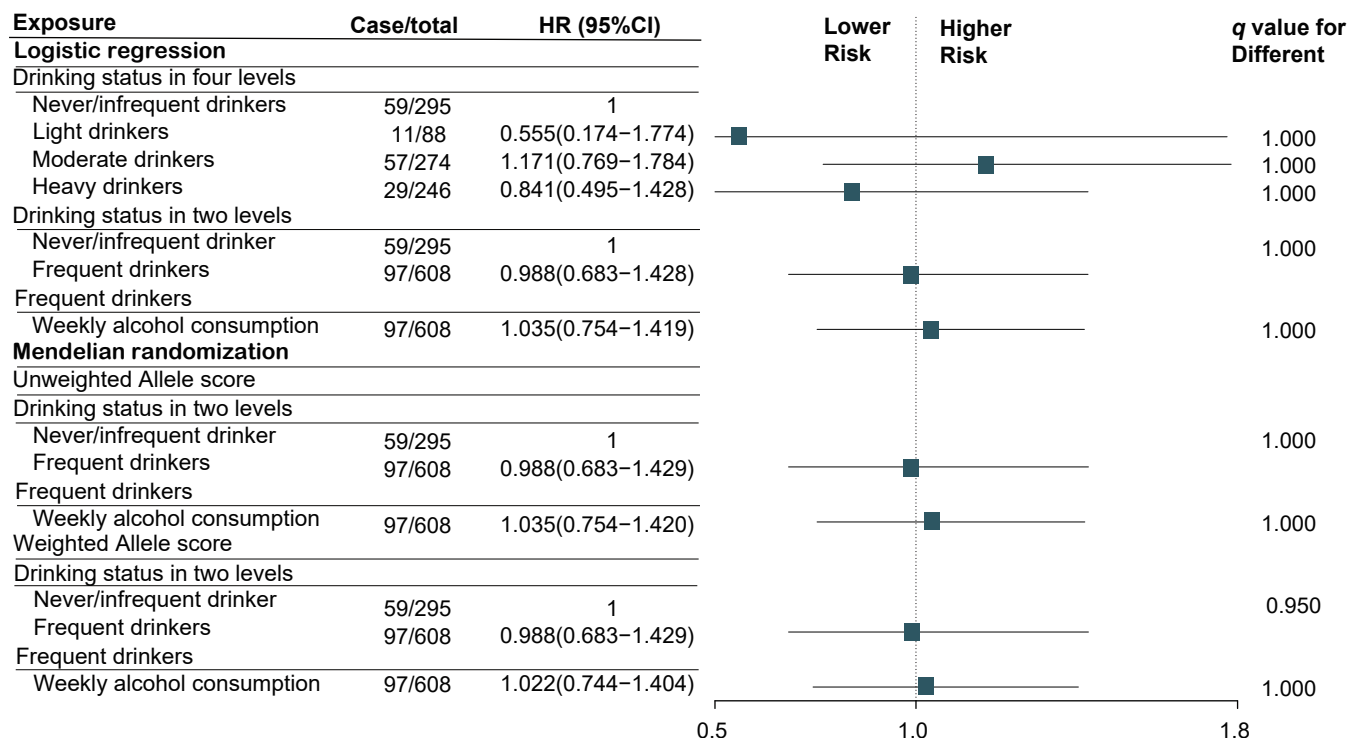
White COVID-19 positive participants with obesity

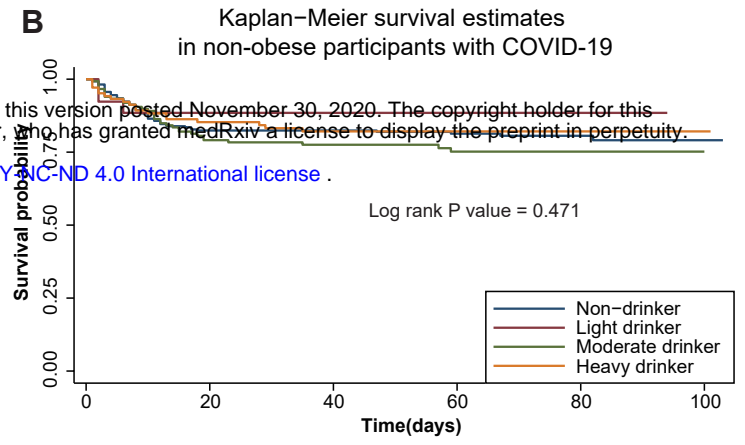
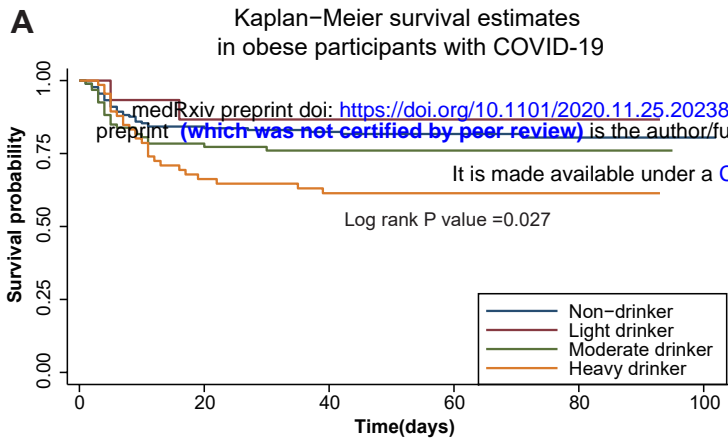
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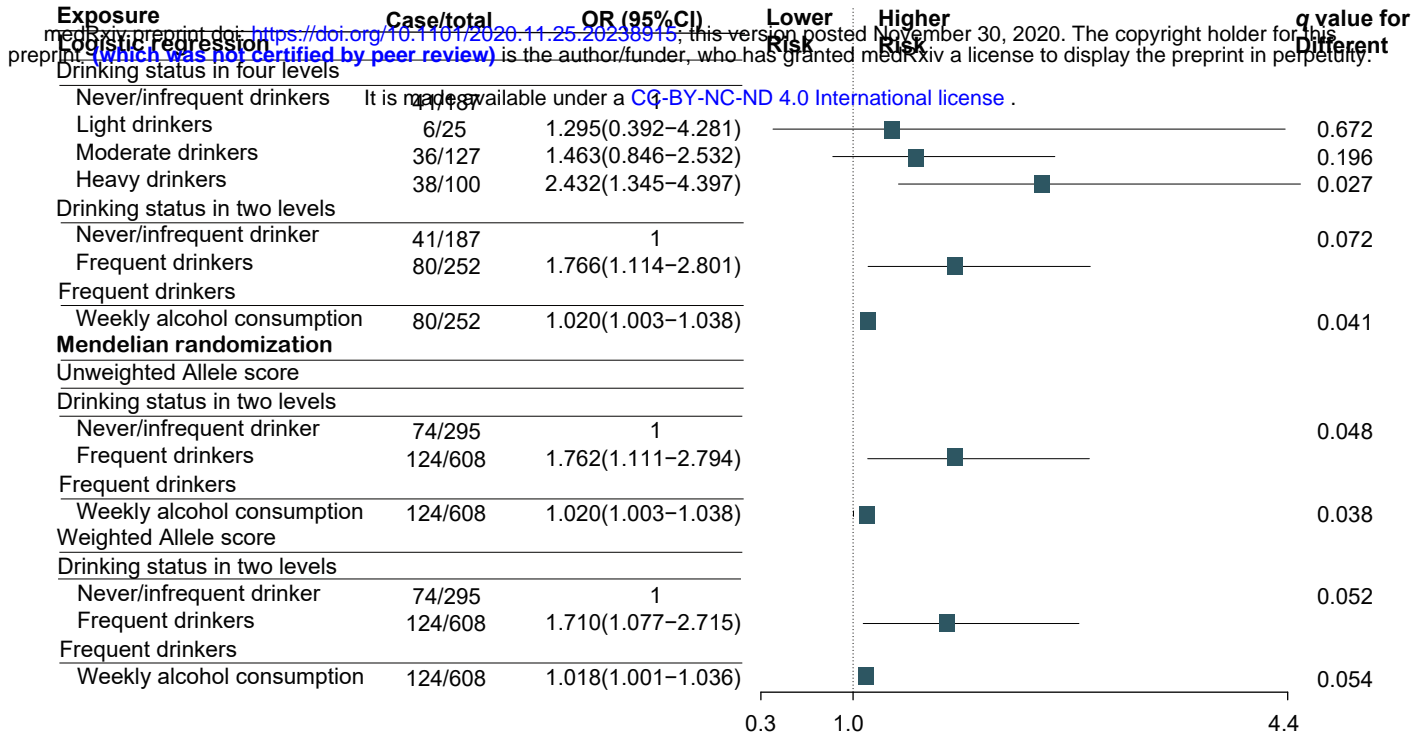
White COVID-19 positive participants without obesity





A

White COVID-19 positive participants with obesity



B

White COVID-19 positive participants without obesity

