

# **HHS Public Access**

Author manuscript Pain. Author manuscript; available in PMC 2021 September 01.

Published in final edited form as:

Pain. 2020 September 01; 161(9): 2167–2178. doi:10.1097/j.pain.00000000000001904.

## **Brain grey matter abnormalities in osteoarthritis pain: a crosssectional evaluation**

**Joana Barroso, MD**1,2,3, **Andrew D. Vigotsky, MS**4, **Paulo Branco, PhD**3, **Ana Mafalda Reis, MD PhD**5, **Thomas J. Schnitzer, MD PhD**2,6,7, **Vasco Galhardo, PhD**1, **A. Vania Apkarian, PhD**2,3,7

<sup>1</sup>Departamento de Biomedicina, Faculdade de Medicina, Universidade do Porto, Porto, Portugal. Instituto de Investigação e Inovação em Saúde - i3S, Universidade do Porto, Porto, Portugal.

<sup>2</sup>Department of Physical Medicine and Rehabilitation, Northwestern University, Feinberg School of Medicine, Chicago, IL, USA.

<sup>3</sup>Department of Physiology, Northwestern University, Feinberg School of Medicine, Chicago, IL, USA.

<sup>4</sup>Departments of Biomedical Engineering and Statistics, Northwestern University, Evanston, IL, USA.

<sup>5</sup>Unilabs Boavista, Porto, Portugal.

<sup>6</sup>Department of Rheumatology, Northwestern University, Feinberg School of Medicine, Chicago, IL, USA.

<sup>7</sup>Department of Anesthesia, Northwestern University, Feinberg School of Medicine, Chicago, IL, USA.

### **Summary**

Grey matter distortions in OA can be captured as a multidomain model, that integrates cortical and subcortical structures, associates with patient-reported neuropathic pain in knee OA and is linked to pain and motor control terms in brain imaging literature.

## **1. Introduction**

Osteoarthritis (OA) is the most common form of arthritis worldwide, one of the most prevalent sources of chronic musculoskeletal pain [43], and a leading cause of disability [14] [42]. Pain is the defining symptom and primary complaint of OA [37], and its consequences include fatigue, sleep disruption, reduced activity and psychological stress [26]. Understanding the mechanisms of pain in OA is thus critical for the management of the disease.

Competing interests

**Corresponding author:** Dr. A.V. Apkarian, PhD, Northwestern University, Feinberg School of Medicine, Tarry Bldg. 7-705, Chicago, Illinois 60611 USA, Phone Office: (312) 503-0404 Fax (312)503-5101, a-apkarian@northwestern.edu.

The authors declare no conflict of interest regarding the present paper.

Traditionally, OA was considered the hallmark of nociceptive pain, explained by mechanical and biochemical activation of afferents innervating the affected joint [28,64]. Recent studies show evidence for peripheral sensitization of nociceptors [59] and central sensitization, associated with nociceptive-mediated transmission strengthening and dysfunction of descending modulator pathways in OA [41]. Additionally, neuropathic-like symptoms [19, 27], negative affect, anxiety and catastrophizing have been related with pain severity [1,57], suggesting linkage between the brain's emotional circuitry and OA pain.

We [8, 46, 6,38] and other researchers [16, 23, 39] have provided evidence for brain anatomical and functional abnormalities in OA. Brain activity for ongoing OA pain [46] shows similar activity patterns as observed in chronic back pain (CBP) [24, 6]. Disruption of information sharing [38], and brain morphological distortions are reported both in OA and in CBP [8]. Given that transition from subacute to CBP is predicted by brain anatomical and functional parameters [7,63] and chronification of back pain is accompanied by anatomical and functional reorganization [25], a similar concept can be formulated for OA pain, namely, brain parameters in part predefine risk and consequently reorganize with OA pain. Here, we attempt to address this issue from the viewpoint of brain morphology. HOA and KOA have distinct clinical characteristics, pain properties, and joint replacement surgery outcomes. For both conditions, pain is poorly related to joint structural damage [18]. Our main hypothesis is that KOA and HOA exhibit brain grey matter (GM) distortions when contrasted with healthy controls, and GM properties should also be different between the two conditions.

GM indices indicate morphologic brain abnormalities in OA [23,39, 2,34], but their clinical import is not yet understood. Although promising, these data are primarily derived from small samples at distinct stages of OA progression, and evidence is discrepant as GM changes exhibit little overlap between studies. Moreover, previous literature fails to find robust associations between clinical properties of OA and brain structural changes [39, 2]. Our second objective is to study the association between brain morphological distortions and OA clinical properties.

We study brain morphological distortions and their relationship with clinical outcomes in a large population of advanced KOA, and a smaller group of HOA, in contrast to healthy controls. We study GM properties at whole-brain voxel-wise scale and after sub-dividing grey matter to regions of interest (ROI) and apply both univariable and multivariable machine-learning analysis to explore the multidimensional basis of these alterations. Finally, we perform a reverse inference mapping of obtained results to terms (using Neurosynth), providing a meta-analytic summary of results and linking brain anatomy to the psychological domain.

#### **2. Patients and Methods**

#### **2.1 Subjects**

Advanced HOA and KOA patients with indication for total joint replacement (TJR) were recruited in the Orthopedic Department of Centro Hospitalar e Universitário de São João, Porto, together with healthy controls from the same geographical area. Participant recruitment was finished at the time of this report. Sample size was determined by the

number of patients waiting for surgery who met the eligibility criteria during a period of 20 months. Controls were recruited in the end of the patient recruitment phase and intended to be one third of the patient number in the largest OA group.

Inclusion criteria for OA patients were: 1) age between 45 to 75 years old; 2) diagnosis of HOA and KOA according to the clinical classification criteria of the American College of Rheumatology [3]; 3) surgical indication for TJR. Exclusion criteria included: 1) secondary OA due to congenital and development diseases or inflammatory and auto-immune articular diseases; 2) bilateral OA with indication for contralateral arthroplasty in the following year or bilateral knee pain with less than or equal to 4 points difference on the numeric pain rating (NRS) scale between knees; 3) other chronic pain conditions (e.g., fibromyalgia; chronic pelvic pain), 4) chronic neurological or psychiatric disease; 5) previous history of stroke or traumatic brain injury.

Controls were subjects without chronic conditions, musculoskeletal and neurological disorders, from the same geographic area, age range and social and educational background as the patients.

Clinical, demographic and biopsychosocial measures were obtained 2–6 weeks prior to surgery and repeated at three, six and twelve months after surgery, together with brain imaging collection (magnetic resonance imaging, MRI). The study was approved by the local Ethics Committee, and all participants provided informed consent prior to partaking in the study. In this manuscript we are reporting baseline cross-sectional data from pre-surgical state. Longitudinal data results will follow in future work.

#### **2.2 Clinical and behavioral measures**

Demographic and clinical data were obtained at the initial interview and by clinical chart assessment. Bilateral radiographs of involved joints were performed as part of the standard hospital protocol and scored accordingly to the Kellgren-Lawrence classification [32] by two trained radiologists. Physical performance was assessed with two distinct tasks – timed up and go test (TUG) [47] and six-minute walking test (6MWT) [9, 50].

All patients completed the following questionnaires: Knee and Hip Injury and Osteoarthritis Score (KOOS; HOOS) [52][20][44]; Hospital Anxiety and Depression Scale (HADS) [45] [68]; Pain Catastrophizing Scale (PCS) [58, 5]; Doleur Neuropathique en 4 questions - Neuropathic pain scale (DN4) [12, 5]. All questionnaires were used in their Portuguese validated versions.

#### **2.3 Brain imaging**

High-resolution, MPRAGE type T1-anatomical brain images were acquired with a 3.0 T Siemen Magnetom Spectra scanner (Siemens Medical, Erlagen, Germany). Acquisition parameters: isometric voxel size =  $1 \times 1 \times 1$  mm, TR = 2500 ms, TE = 3.31 ms, flip angle = 9<sup>o</sup>, in-plane matrix resolution =  $256\times256$ , number of slices, = 160; field of view = 256 mm. Total time of acquisition was 6 minutes.

#### **2.4 Grey matter properties and statistical analyses**

GM properties were extracted, and analyses were performed using tools from the Oxford Center for Functional Magnetic Resonance Imaging of the Brain Software Library (FMRIB, Oxford OK; FSL version 5.0.10, [https://fsl.fmrib.ox.ac.uk/fsl\)](https://fsl.fmrib.ox.ac.uk/fsl) [30]; Matlab (MATLAB 2018a, The MathWorks, Natick, 2018), and RStudio (RStudio Team (2016). RStudio: Integrated Development for R. RStudio, Inc., Boston, MA URL [http://www.rstudio.com](http://www.rstudio.com/)).

**2.4.1 Total gray matter volume estimation—**SIENAx (2.6) from FSL [55, 56] was used to calculate estimates of volumes of interest, after automated brain extraction and tissue segmentation. We compared total gray matter (GM) volume across groups using an analysis of covariance (ANCOVA), controlling for intracranial volume, age and sex. Differential relationships between GM and age were assessed using linear regression, controlling for sex and total intracranial volume (TICV), in which a group interaction indicated a difference in slopes.

**2.4.2 Voxel based morphometry (VBM)—**Regional GM density was analyzed with VBM using tools from FSL. In brief, images were skull stripped using Brain Extraction Tool (BET) [56] and tissue-type segmented into GM, white matter (WM) and cerebrospinal fluid (CSF) using FMRIBs Automated Segmentation Tool (FAST) [67]. A study-specific GM template was created using 72 grey-matter-segmented native images (24 subjects of each group; KOA and controls were randomly selected to minimize size of population bias). Steps towards the template creation included: affine registering the GM partial volume images (FLIRT) [29], left-to-right x-axis flipping to produce a symmetrical image followed by non- linear registration (FNIRT) [4] to standard space. The native space GM volume images were then normalized into this template using a non-linear registration (FNIRT) and corrected for the contraction/enlargement due to the non-linear component of the transformation using the Jacobian. Finally, GM images were concatenated and smoothed with an isotropic gaussian kernel (FWHM=8mm). Cerebellum was excluded from the analyses as the foci of this study were neocortical and subcortical structures.

A second group for parallel VBM analysis was created by midline flipping the scans of patients with left sided OA (knee=45; hip=10), in accordance with previous OA and unilateral pain conditions studies [36, 22]. Brains were flipped along the x-axis prior to nonlinear registration to the study template. All subsequent steps were identical.

Grey matter differences in OA groups and controls were assessed through a one-way between subjects ANCOVA, using age, sex and total intracranial volume, plus laterality of OA for the flipped data analysis as covariates of no-interest. Permutation-based testing (n=5,000 permutations) was performed using FSL's randomize toolbox [65]. Threshold-free cluster enhanced (TFCE) uncorrected p-values were obtained [54], and  $\alpha$ =0.001 was used as threshold for statistical significance of group differences. A minimum cluster size of 66 voxels (number of voxels contained within the 8 mm smoothing kernel) was used as a secondary threshold. Post hoc analyses of VBM results using Tukey HSD tests for multiple comparisons of means were conducted for the identified clusters.

The association of significant clusters with clinical variables was examined using partial correlations, with age, gender and total intracranial volume (TICV) as covariates.

In addition to assessing simple group differences, we sought to understand their relationship with clinical variables. Thus, the association of significant clusters with clinical variables was examined using partial Pearson correlations, which controlled for age, sex and TICV.

**2.4.3 Regional Brodmann area parcellation based analysis of grey matter GM morphometry volume—**Next, we calculated GM volume for 90 cortical and subcortical ROIs approximating Brodmann areas, defined anatomically by the Automated Anatomical Labeling atlas (AAL, Neurofunctional Imaging Group (GIN-IMN) [http://](http://www.gin.cnrs.fr/en/tools/aal/) [www.gin.cnrs.fr/en/tools/aal/](http://www.gin.cnrs.fr/en/tools/aal/))[61]. Cerebellum and Vermis areas were excluded from the analysis. A list of regions, labeling and respective coordinates can be found in Table S1.

The volume of each ROI was calculated as the mean GM volume of all voxels within the ROI for each subject. Volumes were calculated both for native brains and x-axis-flipped brains depending on laterality of OA. Univariable analyses were conducted using one-way ANCOVA test for each region contrasting the three groups: KOA, HOA and controls, while controlling for age, sex and TICV. For the flipped brain analysis, laterality was also added as a covariate. Again, an  $\alpha$ =0.001 was used a threshold for statistical significance of group differences. Multiple comparison correction was performed using a False Discovery Rate (FDR) adjusted p-value of 0.05.

Next, a multidimensional approach using a L1-regularized logistic regression (i.e., least absolute shrinkage and selection operator, LASSO) [60] was performed. Firstly, we divided our KOA sample into two distinct sets, training (n=46) and test (n=45). This was performed using the Kennard-Stone algorithm [33], which allowed us to select samples with a uniform distribution over a multivariate predictor space (here: age, sex, pain levels and behavioral variables). Second, using the KOA training set, we performed a logistic model with a LASSO penalty to predict group membership (i.e., knee vs controls), based on the GM volumes of all 90 ROIs. In addition, we added age, sex and TICV to the full model. Ten-fold cross-validation was used to choose the hyperparameter,  $\lambda$ , that minimized binomial deviance over a default grid size of 100 [31]. Finally, models were evaluated in the KOA test group and HOA group by constructing receiver operating characteristic (ROC) curves and computing areas under the curve (AUC).

Later, Neurosynth term-based reverse inference was used to decode the identified AAL regions for the flipped brain model. This meta-analytical tool contains a database with activation coordinates for a total of 14371 functional MRI studies (September 2019), paired with their associated cognitive and anatomical terms, ([http://neurosynth.org/decode/\)](http://neurosynth.org/decode/) [21,48,66]. The decoder takes in the voxelwise representation of the ROI, cross-references it with the full database, and returns a list of terms and correlation values, which indicates the strength of terms associated with delineated brain regions, in contrast to the rest of the brain. Here, we retrieved the top 40 non-anatomical terms (of about 1700 terms generated) showing the greatest correlation with the model's ROIs. Next, to further elucidate the contribution of each region in the model, the correlation for each ROI was weighted by its

LASSO-derived coefficient and ranked. Terms were visualized using a word-cloud (MATLAB 2018a and Text Analytical Toolbox, The MathWorks, Natick, 2018).

Finally, we aimed to test if the OA classification models were associated with OA clinical variables.

For this, using the KOA testing group, we linearly modeled each clinical variable of interest with the variables retrieved from the LASSO. For statistically significant relationships  $(p<0.05)$ , the relative importance of each independent variable was calculated using the sequential sums of squares over all orderings of regressors [35], and confidence intervals computed over 1000 bootstraps for each regressor [62]. Finally, statistically significant regression models were tested in the KOA validation group and HOA group by calculating Pearson correlation coefficient between predicted and real values.

#### **3. Results**

#### **3.1 Demographic and clinical characteristics**

A total of 91 KOA patients, 24 HOA patients and 36 healthy controls were included in the study. Patient and control groups were similar concerning education, habitation profile, marital status and smoking habits. Age and gender were statistically significantly different among the three groups: age was higher in the KOA group and HOA had a lower representation of females.

Pain intensity as well as affective measures and performance-based tasks did not statistically significantly differ among OA groups. The KOA group, however, presented a longer pain duration and a higher probability of neuropathic pain than the HOA group, which in contrast exhibited higher radiographic structural damage (Table 1).

#### **3.2 Total grey matter characterization**

Total neocortical grey matter volumes (TGM) were compared between the three groups (OA and controls), while controlling for intracranial volume (TICV), age and sex. TGM was not statistically significantly different between the three groups (F  $(2,145)$  =0.07, p=0.79)). Women showed greater TGM than men (controlling for group, TIVC and age) (Figure 1. A). The effect of age on TGM was calculated for each group separately (after correcting for sex). TGM decreased with age (Figure 1.B) and the interaction between age and group was not statistically significant  $(F(2,145)=2.62, p=0.078)$ , showing approximately similar trends for the three groups.

#### **3.3 HOA patients exhibit lower GM volume in anterior cingulate/paracingulate cortex**

In order to identify regional differences in GM volume between OA groups and controls, we performed a whole-brain, voxel-wise, one-way ANCOVA for VBM outcomes (Figure 1.C). Statistically significant differences in GM volume in the right and left anterior cingulate cortex/paracingulate gyrus were identified, differentiating the three groups (Table 2). Post hoc analyses (Tukey HSD test) revealed statistically significantly reduced regional volume only for HOA when compared with KOA and controls (p<0.001). KOA and controls did not show a statistically significant difference within this cluster (p=0.34).

#### **3.4 HOA and KOA patients exhibit lower GM volume in primary motor cortex contralateral to OA pain**

Next, we sought to isolate the effects of OA pain on regional GM volumes. To this end, we first realigned the OA brains relative to the body part with OA pain. That is, brains were flipped across the mid-sagittal plane for left side OA patient (x-axis flipping), such that the left hemisphere would correspond to the body side contralateral to OA pain for all patients. We then performed whole brain, voxel-wise analyses using one-way ANCOVAs across the three groups for VBM outcomes. Using the same criteria for statistical significance as above (i.e.,  $\alpha$ =0.001), we identified five statistically significant clusters (Figure 2.A, Table 2). Post hoc analysis of these clusters revealed: (1) relative to healthy controls, both OA groups exhibited lower volumes in left primary motor cortex (precentral cortex), as well as in the left temporal pole; and (2) relative to healthy controls, KOA, but not HOA, had lower GM volume in the precuneus cortex and increased GM volume in the medial frontal gyrus. Additionally, we identified lower GM volume in the anterior cingulate cortex (ACC\_LR) for the HOA group relative to KOA and controls. This cluster considerably overlapped (Sørensen-Dice coefficient  $= 0.71$ ) with the ACC cluster we obtained in the non-flipped analysis. The unthresholded statistical maps for native space and flipped analysis can be accessed in [Neurovault.org](http://Neurovault.org) (<https://identifiers.org/neurovault.collection:6002>).

Overall, by aligning the brain in relation to body side with OA pain, we observe more statistically significant clusters and stronger effects of GM volume differences between the groups. Importantly, we are able to identify the primary motor cortex region contralateral to the OA pain with lower GM volume in both knee and HOA patients.

#### **3.5 Lack of relationship between clinical variables for clusters identified by VBM**

To elucidate the clinical relevance of the GM volume differences among HOA, KOA, and healthy controls (6 clusters, before and after brain flipping), we correlated clinical characteristics of OA with GM volumes. The relationship between cluster GM volumes and clinical variables was calculated using partial correlations, controlling for age, sex and TICV. We did not find any statistically significant relationships  $(|r| \ 0.28$ ; (Table S2.).

#### **3.6 Univariable and multivariable GM volume relationships associated with OA, after parceling the brain into Brodmann areas**

The above analyses were done at the voxel level. To reduce dimensionality, we studied GM properties across the groups and as a function of clinical characteristics after parceling the brain into regions of interest (ROI) approximating Brodmann areas (AAL atlas). GM volume differences were studied at this larger scale by subdividing the brain into 90 predefined regions (approximating 45 left and 45 right brain Brodmann areas) and calculating the mean GM volume for each ROI. Reduced dimensionality enabled studying OA related brain volume differences with a multivariable approach, which considers the inter-relationships between regional GM volumes.

Firstly, differences in GM volume across the three groups were examined for each ROI in the native and appropriately flipped brains (rendering OA pain to be left body side lateralized) using a one-way ANCOVA (Figure 3.A; 3.B). Multiple regions that attained

statistical significance were identified  $(p<0.001)$ : left putamen in the native brain analysis; bilateral pallidum, precentral left cortex and two right frontal regions (rectus and lingual gyrus) for the flipped brain analysis. After FDR correction for multiple comparisons, only left precentral cortex yielded statistical significance (F=5.9, corrected p<0.05). Post-hoc analysis for this ROI showed decreased volume for both OA groups when compared with controls (Tukey HSD tests,  $p<0.01$ ), but not between OA groups ( $p=0.73$ ), showing a pattern concordant with the results obtained in the voxel-wise analysis.

We next sought to study GM volumetric differences in a multivariable space using logistic regression to classify OA patients and controls based on 90 ROIs. To enforce sparsity and reduce overfitting, we used L1 regularization (LASSO) to drive coefficients towards zero. We took advantage of the large sample available for KOA and subdivided the KOA brains into a training and test groups to ensure generalizability and unbiasedness. Results obtained in the training group were tested on the hold-out test group. Figure 3.C and 3.D shows multivariable model results for the training KOA sample in native space (Figure 3.C left panel) and for flipped brain analyses (Figure 3.D left panel). Age and sex, although having small influences (0.24 to 0.06 log (OR)) were captured in both models. Additionally, the combination of lower GM volumes of the left thalamus and parietal superior left cortex and higher GM volumes of temporal inferior gyrus together were strong predictors for KOA classification in both models. The flipped brain model captured additional regions not observed in the native analysis, including lower GM volumes in bilateral putamen, and higher GM volumes in right precentral gyrus and in frontal inferior orbicular gyrus.

We tested the performance of the classification models in the KOA hold-out sample to test validity of obtained results, and in HOA sample to test for generalization from KOA to HOA. For the hold out KOA group an AUC of 0.82 (non-flipped analysis) and 0.89 (flipped brain analysis) was attained. When evaluated in the HOA group, logistic regression presented an AUC of 0.8, and 0.77 for nonflipped and flipped analyses respectively. ROC curves are displayed in Figure 3. C. Right panel and 3.D. Right panel. Probability plots for each model are presented in Figure S1.

Lastly, we tested the model's performance for classifying amongst OA groups. The obtained AUC's were 0.65 and 0.76 for the nonflipped and flipped model respectively (probability plots, Figure S2). Multivariable models seem both valid and generalizable from KOA to HOA, with the flipped model showing higher specificity for KOA.

#### **3.7 Neuropathic pain associated with multivariable model for KOA**

Given that the flipped brain multivariable model could better distinguish between OA and healthy controls, also showing higher specificity for knee OA, we further examined this model (figure 4.A). Firstly, we used Neurosynth database meta analytical decoding to identify psychological/task terms associated with brain activity in these regions. We decoded the regions as absolute values, reflecting only location (Figure 4.B), and also using weighted values to account relative contribution of different regions in the model (Figure 4.C). The top five terms associated with the model for location only were: dementia, personality, gestures, social interaction, and interpersonal relations. When incorporating the model

coefficient of each area, the main terms identified were pain, anticipatory, finger tapping, index finger, and integrate.

In an unbiased approach, we tested the model's association with clinical aspects of pain and mobility in OA. Using the KOA training sample (n=45), and the predictors obtained from the LASSO, multiple regression models were built for each clinical variable of interest (i.e. pain duration; pain intensity; HADS anxiety; HADS depression; PCS; 6MWT; TUG; DN4 and KOOS subscales). Only one model yielded statistical significance: neuropathic pain scale, DN4 (F  $(1,45) = 2.29$ ; Adj R2=0.29; p=0.03). Details for all generated models can be found in Table S3. Finally, because several models were generated and thus spuriousness of the DN4 model is highly feasible, we sought to better characterize and unbiasedly validate the DN4 model. To do so, we first calculated the relative importance for each predictor variable and computed its respective 95% confidence interval using 1000 bootstrap replicates (Figure 4.D). Age, parietal superior cortex, Heschl/S2 gyrus, and frontal inferior orbitalis cortex presented the highest percentage of explained variance, with smaller contributions from other brain regions. We then validated the model in the KOA holdout sample (n=45), where the correlation between DN4 and predicted DN4 was statistically significant and strong ( $r=0.47$ ,  $p<0.001$ ). The model failed to replicate in the HOA group ( $r=$ −0.018, p=0.93; Fig. 4.E). Therefore, the obtained multivariable model reflects neuropathic pain characteristics, and the model is valid for neuropathic pain in KOA but does not generalize to HOA.

#### **4. Discussion**

Our univariate analyses highlight primarily lower GM volume in HOA in the anterior cingulate, and lower GM volume in the precentral cortex in both KOA and HOA contralateral to OA pain. Additionally, the multivariable analysis reveals GM distortions captured as a multidomain system that integrates cortical and subcortical structures, shared between KOA and HOA. The reverse inference of this multidomain brain regional distortions, from functional brain space to psychological and task-related terms, identified two sets of concepts: 1) for unweighted location analysis associated terms included dementia, personality, gestures, social interaction, and interpersonal relations; 2) when the same analysis was performed taking into consideration the weights of contributing regions, the top associated concepts were: pain, anticipatory, finger tapping, index finger, and integrate. We found that the multivariable set of regions was associated with patient-reported neuropathic pain in KOA, but not in HOA patients. Importantly, the latter group has shorter pain duration and lower neuropathic pain scores. Our univariable analysis demonstrates a more widespread and stronger GM distortions when taking into account the laterality of OA; however, no statistically significant relationships were found between univariable (VBM) results and clinical variables, similarly to previous reports in the field [39,2].

Brain structural changes using GM indices in chronic pain have been reported by multiple groups, primarily using VBM. Disparate regions, both directions of GM change and even null findings are reported [13]. Our study was designed with the purpose of characterizing brain structural properties in the largest OA population reported to date. Our results are, in part, consistent with earlier reports: decreased GM in ACC seen in HOA has also been

described in the past [51]; decreased precentral GM seen here in KOA and HOA is consistent with a report of functional remodeling of the motor cortex in KOA patients [53]. Importantly, we did not observe significant GM changes (univariable analysis) in brain areas previously reported for OA VBM studies, including both subcortical (thalamus [23], amygdala and nuclei accumbens [34], caudate nuclei [39]) and cortical areas (insular cortex, dorsolateral prefrontal cortex [51]). Incongruent findings across studies may relate to the heterogeneity of OA pain at distinct stages of the disease, but also to the potential inherent limitations of VBM; e.g., the sensitivity of VBM to imaging parameters and scanner differences, in addition to analytical details, such as spatially normalizing atypical brains and limitations of univariate techniques [8, 40].

Our multivariable analysis indicates the existence of a common set of brain distortions between KOA and HOA that differentiate these patients from healthy controls: 9–13 brain areas, comprised of both cortical and subcortical regions. There is accumulating evidence that chronic pain leads to whole-brain global reorganization [13]. Tested specifically in OA patients, pain was associated with whole-brain degree rank order disruption [38], and changes in large scale brain network temporal dynamics [16]. Bridging this concept to morphologic brain reorganization in OA, we examined the GM volumetric distortions in a multidimensional space, accounting for the interrelationship between regions. Classification models were generated in a KOA sample, both in native and flipped brain space, and tested in hold-out KOA and HOA groups. Interestingly, the native space model yielded similar accuracy for knee and hip patients. Accuracy improved for the KOA group in the flipped brain space.

To better characterize this model, we identified the terms in the brain imaging literature associated with these regions using Neurosynth Image Decoder and selected the top 40 (of about 1300 terms) non-anatomical terms. The method employs meta-analysis (exploring around 14,000 brain imaging studies) to identify psychological and task-related terms associated with the collection of brain regions included in the multivariable model. The reverse association for the unweighted multivariable model identifies cognitive, interpersonal, and social interaction concepts, not obviously related to OA. Remarkably, when the weighted multivariable model was examined, the top term was pain (chance probability of this observation is near 1 in 1300= 0.0008); additional pain-related terms, including chronic pain as well as motor-related terms, were also captured. This result is noteworthy as Neurosynth explores brain activity in a heterogeneous population and yet identifies pain and motor properties in the weighted multivariable model.

When studying the relationship between clinical measures and LASSO-derived multivariable models, we show that neuropathic pain probability, measured with the DN4 scale, is associated with the set of identified variables in the KOA group. Using an unbiased approach, we demonstrated that this result validates in the KOA sample but fails to generalize for HOA patients. This finding deserves a thorough examination: 1) KOA patients show a higher probability of neuropathic pain than HOA, fact that has been previously described [11]; 2) interestingly, in our sample, KOA patients also showed a longer pain duration, but less severe radiographic damage.

These results are concordant with Dabare et al [17], who showed that HOA symptomatology clinically presents later with more advanced disease than KOA, and patients are likely to undergo arthroplasty earlier from the time of presentation. Our findings suggest that, together with longer pain duration, there is the emergence of a neuropathic pain profile that can be clinically identified and linked to brain GM morphological properties. Our present result cannot unequivocally establish whether the multivariable GM distortion is a prognostic or diagnostic of neuropathic pain. However, the temporal dissociation between knee and hip pain and radiographic damage, together with a shared GM multivariable model, suggests that the model is more likely to be prognostic risk factor for developing signs of neuropathic pain in OA.

In a recent report based on the same KOA and HOA patient cohort, we examined the relationship between OA pain and clinical parameters [10]. Although multiple models could be derived for different measures of OA pain, these were not consistent and thus were deemed unstable and uninformative. Here, we queried the univariate and multivariable GM distortions associated with OA regarding their relationship to clinical characteristics. Univariate brain distortions were more prominent in the hemisphere contralateral to OA pain, and the neuropathic pain scale was related to the multivariable model, yet the majority of clinical parameters associated with OA were unrelated to GM distortions. This raises the question whether observed GM distortions can be seen as prognostic or diagnostic. Within the conceptual construct of the four-stage model for transition to chronic pain [49], formulated based on a longitudinal study of back pain patients [7], we have proposed a set of criteria that would differentiate between prognostic and diagnostic biomarkers for chronic pain: prognostic biomarkers should be preexisting factors that influence transition to chronic pain but are not modulated by pain. Conversely, diagnostic biomarkers are emergent in time with the transition to chronic pain and should reflect pain and related clinical characteristics. If we presume that the transition to chronic pain for OA, relating to its persistence after joint replacement surgery, follows the same general model as in back pain, then a lack of relationship between clinical parameters and GM abnormalities suggests these are more likely to be prognostic in nature. This notion should be especially true for brain regions or multivariable models that show distortions shared between KOA and HOA as the impact of OA pain in HOA is distinct from that for KOA [10], which in turn should differentially distort GM in areas reflecting diagnostic consequences.

This study does not come without limitations. Results from the VBM analysis should be cautiously interpreted as identified clusters result from uncorrected whole-brain analysis; however, our report is in line with previous publications on brain morphometry in chronic pain, reporting group comparisons at uncorrected p<0.001 [2]. We limit reporting only large cluster GM abnormalities (> 66 voxels) to minimize false-positive findings. Another important limitation that has to be highlighted is the imbalance in number of subjects between OA groups and controls. While such an imbalance limits our statistical power [15], our knee group was sufficient to construct the multivariable models using half of the knee patients. However, this is in contrast to the control group, which was not large enough to split, and thus, it was used in both the training and test set. Another shortcoming of the present study, is the fact that controls are not perfectly matched regarding age and gender with our KOA group, however we account for the effect of these variables in all stages of our

analysis. Finally, socio-economic and cultural differences regarding attitudes of coping and/or suffering with OA pain may also underlie differences in VBM results when contrasting our findings with other studies.

In conclusion, our results show decreased GM in HOA in the ACC and decreased GM in precentral cortex in both KOA and HOA; however, these GM differences are not strongly correlated with clinical variables in OA. Moreover, we derived a multivariable model based on GM distortion that differentiated OA patients from healthy subjects. Importantly, this model was associated with the neuropathic scale in KOA, and in reverse inference it was linked with pain and motor control in the brain imaging literature. We reason that observed GM distortions may reflect a priori risks for chronicity of OA pain rather than being consequential to the development of the disease.

#### **Supplementary Material**

Refer to Web version on PubMed Central for supplementary material.

#### **Acknowledgements**

We thank to Dr. Paulo Oliveira, for the clinical support and supervision of patients enrolled in the study and all the members of the Orthopedic Department of Centro Hospitalar Universitário de São João, Porto, for their support with data collection. We thank to Unilabs Boavista and Dr. Nuno Martins for all the support in imaging data acquisition. We are thankful to all Apkarian lab and Galhardo lab members that contributed to this study with their time and resources.

#### Disclosures

J.B. was funded through CCDRN [Norte-08-5369-FSE-000026], OARSI Collaborative Scholarship 2018 and Luso-American Development Foundation R&D@PhD scholarship grant. Neuroimage data acquisition was provided by Unilabs Boavista and the Grünenthal Young Pain Researcher 2017 Grant, Portugal. This material is based upon work supported by the National Science Foundation Graduate Research Fellowship under Grant No. DGE-1324585.

#### **References:**

- [1]. Akin-Akinyosoye K, Frowd N, Marshall L, Stocks J, Fernandes GS, Valdes A, McWilliams DF, Zhang W, Doherty M, Ferguson E, Walsh DA. Traits associated with central pain augmentation in the Knee Pain In the Community (KPIC) cohort. Pain 2018;159:1035–1044. [PubMed: 29438225]
- [2]. Alshuft HM, Condon LA, Dineen RA, Auer DP. Cerebral Cortical Thickness in Chronic Pain Due to Knee Osteoarthritis: The Effect of Pain Duration and Pain Sensitization. PLoS One 2016;11:e0161687. [PubMed: 27658292]
- [3]. Altman R, Asch E, Bloch D, Bole G, Borenstein D, Brandt K, Christy W, Cooke TD, Greenwald R, Hochberg M, et al. Development of criteria for the classification and reporting of osteoarthritis. Classification of osteoarthritis of the knee. Diagnostic and Therapeutic Criteria Committee of the American Rheumatism Association. Arthritis Rheum 1986;29:1039–49. [PubMed: 3741515]
- [4]. Andersson JLRJM. Non-linear registration, aka spatial 524 normalisation. FMRIB Tech Rep TR07JA2 2010.
- [5]. Azevedo LCP. (Tradução, adaptação cultural e estudo multicêntrico de validação de instrumentos para rastreio e avaliação do impacto da dor crónica (Translation, cultural adaptation and multicentric validation study of chronic pain screening and impact assessment instruments). Dor 2007;15:6–65.

- [6]. Baliki MN, Geha PY, Jabakhanji R, Harden N, Schnitzer TJ, Apkarian AV. A preliminary fMRI study of analgesic treatment in chronic back pain and knee osteoarthritis. Mol Pain 2008;4:47. [PubMed: 18950528]
- [7]. Baliki MN, Petre B, Torbey S, Herrmann KM, Huang L, Schnitzer TJ, Fields HL, Apkarian AV. Corticostriatal functional connectivity predicts transition to chronic back pain. Nat Neurosci 2012;15:1117–1119. [PubMed: 22751038]
- [8]. Baliki MN, Schnitzer TJ, Bauer WR, Apkarian AV. Brain morphological signatures for chronic pain. PLoS One 2011;6:e26010. [PubMed: 22022493]
- [9]. Balke B A Simple Field Test for the Assessment of Physical Fitness. Rep 63–6. Rep Civ Aeromed Res Inst US 1963:1–8.
- [10]. Barroso J, Wakaizumi K, Reckziegel D, Pinto-Ramos J, Schnitzer T, Galhardo V, Apkarian AV. Prognostics for pain in osteoarthritis: Do clinical measures predict pain after total joint replacement? PLOS ONE 2020;15:e0222370. [PubMed: 31914126]
- [11]. Blikman T, Rienstra W, van Raay J, Dijkstra B, Bulstra SK, Stevens M, van den Akker- Scheek I. Neuropathic-like symptoms and the association with joint-specific function and quality of life in patients with hip and knee osteoarthritis. PLoS One 2018;13:e0199165. [PubMed: 29902245]
- [12]. Bouhassira D, Attal N, Alchaar H, Boureau F, Brochet B, Bruxelle J, Cunin G, Fermanian J,Ginies P, Grun-Overdyking A, Jafari-Schluep H, Lanteri-Minet M, Laurent B, Mick G, Serrie A, Valade D, Vicaut E. Comparison of pain syndromes associated with nervous or somatic lesions and development of a new neuropathic pain diagnostic questionnaire (DN4). Pain 2005;114:29–36. [PubMed: 15733628]
- [13]. Cauda F, Palermo S, Costa T, Torta R, Duca S, Vercelli U, Geminiani G, Torta DM. Gray matter alterations in chronic pain: A network-oriented meta-analytic approach. Neuroimage Clin 2014;4:676–86. [PubMed: 24936419]
- [14]. Centers for Disease Control and Prevention (CDC). Prevalence and most common causes of disability among adults--United States, 2005. MMWR Morb Mortal Wkly Rep 2009;58:421–6. [PubMed: 19407734]
- [15]. Cohen J Statistical Power Analysis for the Behavioral Sciences. Hillsdale NJ Lawrence 558 Erlbaum Assoc Publ n.d.;2nd ed.
- [16]. Cottam WJ, Iwabuchi SJ, Drabek MM, Reckziegel D, Auer DP. 559 Altered connectivity of the right anterior insula drives the pain connectome changes in chronic knee osteoarthritis. Pain 2018;159:929–938. [PubMed: 29557928]
- [17]. Dabare C, Le Marshall K, Leung A, Page CJ, Choong PF, Lim KK. Differences in presentation, progression and rates of arthroplasty between hip and knee osteoarthritis: Observations from an osteoarthritis cohort study-a clear role for conservative management. Int J Rheum Dis 2017;20:1350–1360. [PubMed: 28493422]
- [18]. Finan PH, Buenaver LF, Bounds SC, Hussain S, Park RJ, Haque UJ, Campbell CM, Haythornthwaite JA, Edwards RR, Smith MT. Discordance between pain and radiographic severity in knee osteoarthritis: findings from quantitative sensory testing of central sensitization. Arthritis Rheum 2013;65:363–72. [PubMed: 22961435]
- [19]. Fingleton C, Smart K, Moloney N, Fullen BM, Doody C. Pain sensitization in people with knee osteoarthritis: a systematic review and meta-analysis. Osteoarthritis Cartilage 2015;23:1043–56. [PubMed: 25749012]
- [20]. Goncalves RS, Cabri J, Pinheiro JP, Ferreira PL. Cross-cultural adaptation and validation of the Portuguese version of the Knee injury and Osteoarthritis Outcome Score (KOOS). Osteoarthritis Cartilage 2009;17:1156–62. [PubMed: 19303082]
- [21]. Gorgolewski KJ, Varoquaux G, Rivera G, Schwartz Y, Sochat VV, Ghosh SS, Maumet C, Nichols TE, Poline JB, Yarkoni T, Margulies DS, Poldrack RA. [NeuroVault.org](http://NeuroVault.org): A repository for sharing unthresholded statistical maps, parcellations, and atlases of the human brain. Neuroimage 2016;124:1242–1244. [PubMed: 25869863]
- [22]. Gustin SM, Peck CC, Wilcox SL, Nash PG, Murray GM, Henderson LA. Different pain, different brain: thalamic anatomy in neuropathic and non-neuropathic chronic pain syndromes. J Neurosci 2011;31:5956–64. [PubMed: 21508220]

- [23]. Gwilym SE, Filippini N, Douaud G, Carr AJ, Tracey I. Thalamic atrophy associated with painful osteoarthritis of the hip is reversible after arthroplasty: a longitudinal voxel-based morphometric study. Arthritis Rheum 2010;62:2930–40. [PubMed: 20518076]
- [24]. Hashmi JA, Baliki MN, Huang L, Baria AT, Torbey S, Hermann KM, Schnitzer TJ, Apkarian AV. Shape shifting pain: chronification of back pain shifts brain representation from nociceptive to emotional circuits. Brain 2013;136:2751–68. [PubMed: 23983029]
- [25]. Hashmi JA, Baliki MN, Huang L, Baria AT, Torbey S, Hermann KM, Schnitzer TJ, Apkarian AV. Shape shifting pain: chronification of back pain shifts brain representation from nociceptive to emotional circuits. Brain 2013;136:2751–68. [PubMed: 23983029]
- [26]. Hawker GA. Experiencing painful osteoarthritis: what have we learned from listening? Curr Opin Rheumatol 2009;21:507–12. [PubMed: 19633560]
- [27]. Hochman JR, Davis AM, Elkayam J, Gagliese L, Hawker 594 GA. Neuropathic pain symptoms on the modified painDETECT correlate with signs of central sensitization in knee osteoarthritis. Osteoarthritis Cartilage 2013;21:1236–42. [PubMed: 23973136]
- [28]. IASP. IASP TERMINOLOGY 2018 [Definitions of pain terminology]. Available from: [http://](http://www.ias-pain.org/Edication/Content.aspx?Itemnumber=1698) [www.ias-pain.org/Edication/Content.aspx?Itemnumber=1698](http://www.ias-pain.org/Edication/Content.aspx?Itemnumber=1698) 2018.
- [29]. Jenkinson M, Bannister P, Brady M, Smith S. Improved optimization for the robust and accurate linear registration and motion correction of brain images. Neuroimage 2002;17:825–41. [PubMed: 12377157]
- [30]. Jenkinson M, Beckmann CF, Behrens TE, Woolrich MW, Smith SM. Fsl. Neuroimage 2012;62:782–90. [PubMed: 21979382]
- [31]. Friedman Jerome, Tibshirani Trevor Hastie Rob. Regularization Paths for Generalized Linear Models via Coordinate Descent. J Stat Softw 2010;33:1–22. [PubMed: 20808728]
- [32]. Kellgren JH, Lawrence JS. Radiological assessment of osteo-arthrosis. Ann Rheum Dis 1957;16:494–502. [PubMed: 13498604]
- [33]. Kennard RW, Stone LA. Computer Aided Design of Experiments. Technometrics 1969;11:137– 148.
- [34]. Lewis GN, Parker RS, Sharma S, Rice DA, McNair PJ. Structural Brain Alterations Before and After Total Knee Arthroplasty: A Longitudinal Assessment. Pain Med 2018;19:2166–2176. [PubMed: 29917139]
- [35]. Lindeman RHMP. Introduction to Bivariate and Multivariate Analysis. Scott Foresman Glenview IL 1980:119.
- [36]. Maeda Y, Kettner N, Sheehan J, Kim J, Cina S, Malatesta C, Gerber J, McManus C, Mezzacappa P, Morse LR, Audette J, Napadow V. Altered brain morphometry in carpal tunnel syndrome is associated with median nerve pathology. Neuroimage Clin 2013;2:313–319. [PubMed: 23799199]
- [37]. Malfait AM, Schnitzer TJ. Towards a mechanism-based approach to pain management in osteoarthritis. Nat Rev Rheumatol 2013;9:654–64. [PubMed: 24045707]
- [38]. Mansour A, Baria AT, Tetreault P, Vachon-Presseau E, Chang PC, Huang L, Apkarian AV, Baliki MN. Global disruption of degree rank order: a hallmark of chronic pain. Sci Rep 2016;6:34853. [PubMed: 27725689]
- [39]. Mao CP, Bai ZL, Zhang XN, Zhang QJ, Zhang L. Abnormal Subcortical Brain Morphology in Patients with Knee Osteoarthritis: A Cross-sectional Study. Front Aging Neurosci 2016;8:3. [PubMed: 26834629]
- [40]. Mechelli A; Price CF. Voxel-Based Morphometry of the 627 Human Brain: Methods and Applications. Curr Med Imaging Rev 2005;1.
- [41]. Murphy SL, Phillips K, Williams DA, Clauw DJ. The role of the central nervous system in osteoarthritis pain and implications for rehabilitation. Curr Rheumatol Rep 2012;14:576–82. [PubMed: 22879060]
- [42]. Murray CJ, Vos T, Lozano R, Naghavi M, Flaxman AD, Michaud C, Ezzati M, Shibuya K, Salomon JA, Abdalla S, Aboyans V, Abraham J, Ackerman I, Aggarwal R, Ahn SY, Ali MK, Alvarado M, Anderson HR, Anderson LM, Andrews KG, Atkinson C, Baddour LM, Bahalim AN, Barker-Collo S, Barrero LH, Bartels DH, Basanez MG, Baxter A, Bell ML, Benjamin EJ, Bennett D, Bernabe E, Bhalla K, Bhandari B, Bikbov B, Bin Abdulhak A, Birbeck G, Black JA,

Blencowe H, Blore JD, Blyth F, Bolliger I, Bonaventure A, Boufous S, Bourne R, Boussinesq M, Braithwaite T, Brayne C, Bridgett L, Brooker S, Brooks P, Brugha TS, Bryan-Hancock C, Bucello C, Buchbinder R, Buckle G, Budke CM, Burch M, Burney P, Burstein R, Calabria B, Campbell B, Canter CE, Carabin H, Carapetis J, Carmona L, Cella C, Charlson F, Chen H, Cheng AT, Chou D, Chugh SS, Coffeng LE, Colan SD, Colquhoun S, Colson KE, Condon J, Connor MD, Cooper LT, Corriere M, Cortinovis M, de Vaccaro KC, Couser W, Cowie BC, Criqui MH, Cross M, Dabhadkar KC, Dahiya M, Dahodwala N, Damsere-Derry J, Danaei G, Davis A, De Leo D, Degenhardt L, Dellavalle R, Delossantos A, Denenberg J, Derrett S, Des Jarlais DC, Dharmaratne SD, Dherani M, Diaz-Torne C, Dolk H, Dorsey ER, Driscoll T, Duber H, Ebel B, Edmond K, Elbaz A, Ali SE, Erskine H, Erwin PJ, Espindola P, Ewoigbokhan SE, Farzadfar F, Feigin V, Felson DT, Ferrari A, Ferri CP, Fevre EM, Finucane MM, Flaxman S, Flood L, Foreman K, Forouzanfar MH, Fowkes FG, Fransen M, Freeman MK, Gabbe BJ, Gabriel SE, Gakidou E, Ganatra HA, Garcia B, Gaspari F, Gillum RF, Gmel G, Gonzalez- Medina D, Gosselin R, Grainger R, Grant B, Groeger J, Guillemin F, Gunnell D, Gupta R, Haagsma J, Hagan H, Halasa YA, Hall W, Haring D, Haro JM, Harrison JE, Havmoeller R, Hay RJ, Higashi H, Hill C, Hoen B, Hoffman H, Hotez PJ, Hoy D, Huang JJ, Ibeanusi SE, Jacobsen KH, James SL, Jarvis D, Jasrasaria R, Jayaraman S, Johns N, Jonas JB, Karthikeyan G, Kassebaum N, Kawakami N, Keren A, Khoo JP, King CH, Knowlton LM, Kobusingye O, Koranteng A, Krishnamurthi R, Laden F, Lalloo R, Laslett LL, Lathlean T, Leasher JL, Lee YY, Leigh J, Levinson D, Lim SS, Limb E, Lin JK, Lipnick M, Lipshultz SE, Liu W, Loane M, Ohno SL, Lyons R, Mabweijano J, MacIntyre MF, Malekzadeh R, Mallinger L, Manivannan S, Marcenes W, March L, Margolis DJ, Marks GB, Marks R, Matsumori A, Matzopoulos R, Mayosi BM, McAnulty JH, McDermott MM, McGill N, McGrath J, Medina-Mora ME, Meltzer M, Mensah GA, Merriman TR, Meyer AC, Miglioli V, Miller M, Miller TR, Mitchell PB, Mock C, Mocumbi AO, Moffitt TE, Mokdad AA, Monasta L, Montico M, Moradi-Lakeh M, Moran A, Morawska L, Mori R, Murdoch ME, Mwaniki MK, Naidoo K, Nair MN, Naldi L, Narayan KM, Nelson PK, Nelson RG, Nevitt MC, Newton CR, Nolte S, Norman P, Norman R, O'Donnell M, O'Hanlon S, Olives C, Omer SB, Ortblad K, Osborne R, Ozgediz D, Page A, Pahari B, Pandian JD, Rivero AP, Patten SB, Pearce N, Padilla RP, Perez-Ruiz F, Perico N, Pesudovs K, Phillips D, Phillips MR, Pierce K, Pion S, Polanczyk GV, Polinder S, Pope CA, Popova S, Porrini E, Pourmalek F, Prince M, Pullan RL, Ramaiah KD, Ranganathan D, Razavi H, Regan M, Rehm JT, Rein DB, Remuzzi G, Richardson K, Rivara FP, Roberts T, Robinson C, De Leon FR, Ronfani L, Room R, Rosenfeld LC, Rushton L, Sacco RL, Saha S, Sampson U, Sanchez-Riera L, Sanman E, Schwebel DC, Scott JG, Segui-Gomez M, Shahraz S, Shepard DS, Shin H, Shivakoti R, Singh D, Singh GM, Singh JA, Singleton J, Sleet DA, Sliwa K, Smith E, Smith JL, Stapelberg NJ, Steer A, Steiner T, Stolk WA, Stovner LJ, Sudfeld C, Syed S, Tamburlini G, Tavakkoli M, Taylor HR, Taylor JA, Taylor WJ, Thomas B, Thomson WM, Thurston GD, Tleyjeh IM, Tonelli M, Towbin JA, Truelsen T, Tsilimbaris MK, Ubeda C, Undurraga EA, van der Werf MJ, van Os J, Vavilala MS, Venketasubramanian N, Wang M, Wang W, Watt K, Weatherall DJ, Weinstock MA, Weintraub R, Weisskopf MG, Weissman MM, White RA, Whiteford H, Wiebe N, Wiersma ST, Wilkinson JD, Williams HC, Williams SR, Witt E, Wolfe F, Woolf AD, Wulf S, Yeh PH, Zaidi AK, Zheng ZJ, Zonies D, Lopez AD, AlMazroa MA, Memish ZA. Disability-adjusted life years (DALYs) for 291 diseases and injuries in 21 regions, 1990–2010: a systematic analysis for the Global Burden of Disease Study 2010. Lancet 2012;380:2197–223. [PubMed: 23245608]

- [43]. Neogi T The epidemiology and impact of pain in osteoarthritis. Osteoarthritis Cartilage 2013;21:1145–53. [PubMed: 23973124]
- [44]. Nilsdotter AK, Lohmander LS, Klassbo M, Roos EM. Hip disability and osteoarthritis outcome score (HOOS)--validity and responsiveness in total hip replacement. BMC Musculoskelet Disord 2003;4:10. [PubMed: 12777182]
- [45]. Pais-Ribeiro J, Silva I, Ferreira T, Martins A, Meneses R, Baltar M. Validation study of a Portuguese version of the Hospital Anxiety and Depression Scale. Psychol Health Med 2007;12:225–35; quiz 235–7. [PubMed: 17365902]
- [46]. Parks EL, Geha PY, Baliki MN, Katz J, Schnitzer TJ, Apkarian AV. Brain activity for chronic knee osteoarthritis: dissociating evoked pain from spontaneous pain. Eur J Pain 2011;15:843 e1-14. [PubMed: 21315627]

- [47]. Podsiadlo D, Richardson S. The timed "Up & Go": a test of basic functional mobility for frail elderly persons. J Am Geriatr Soc 1991;39:142–8. [PubMed: 1991946]
- [48]. Poldrack R Can cognitive processes be inferred from neuroimaging data? Trends Cogn Sci 2006;10:59–63. [PubMed: 16406760]
- [49]. Reckziegel D, Vachon-Presseau E, Petre B, Schnitzer TJ, Baliki MN, Apkarian AV. Deconstructing biomarkers for chronic pain: context- and hypothesis-dependent biomarker types in relation to chronic pain. PAIN 2019;160:S37–S48. [PubMed: 31008848]
- [50]. Rejeski WJ, Ettinger WH, Schumaker S, James P, Burns R, Elam JT. Assessing performance701 related disability in patients with knee osteoarthritis. Osteoarthritis Cartilage 1995;3:157–67. [PubMed: 8581745]
- [51]. Rodriguez-Raecke R, Niemeier A, Ihle K, Ruether W, May A. Structural Brain Changes in Chronic Pain Reflect Probably Neither Damage Nor Atrophy. PLoS ONE 2013;8:e54475. [PubMed: 23405082]
- [52]. Roos EM, Toksvig-Larsen S. Knee injury and Osteoarthritis 705 Outcome Score (KOOS) validation and comparison to the WOMAC in total knee replacement. Health Qual Life Outcomes 2003;1:17. [PubMed: 12801417]
- [53]. Shanahan CJ, Hodges PW, Wrigley TV, Bennell KL, Farrell MJ. Organisation of the motor cortex differs between people with and without knee osteoarthritis. Arthritis Res Ther 2015;17:164. [PubMed: 26080802]
- [54]. Smith S, Nichols T. Threshold-free cluster enhancement: Addressing problems of smoothing, threshold dependence and localisation in cluster inference. NeuroImage 2009;44:83–98. [PubMed: 18501637]
- [55]. Smith SM, Jenkinson M, Woolrich MW, Beckmann CF, Behrens TEJ, Johansen-Berg H, Bannister PR, De Luca M, Drobnjak I, Flitney DE, Niazy RK, Saunders J, Vickers J, Zhang Y, De Stefano N, Brady JM, Matthews PM. Advances in functional and structural MR image analysis and implementation as FSL. NeuroImage 2004;23:S208–S219. [PubMed: 15501092]
- [56]. Smith SM, Zhang Y, Jenkinson M, Chen J, Matthews PM, Federico A, De Stefano N. Accurate, Robust, and Automated Longitudinal and Cross-Sectional Brain Change Analysis. NeuroImage 2002;17:479–489. [PubMed: 12482100]
- [57]. Somers TJ, Keefe FJ, Pells JJ, Dixon KE, Waters SJ, Riordan PA, Blumenthal JA, McKee DC, LaCaille L, Tucker JM, Schmitt D, Caldwell DS, Kraus VB, Sims EL, Shelby RA, Rice JR. Pain Catastrophizing and Pain-Related Fear in Osteoarthritis Patients: Relationships to Pain and Disability. J Pain Symptom Manage 2009;37:863–872. [PubMed: 19041218]
- [58]. Sullivan MJL; B. The Pain Catastrophizing Scale: Development and validation. Psychol Assess 1995;7:524–532.
- [59]. Thakur M, Dickenson AH, Baron R. Osteoarthritis pain: nociceptive or neuropathic? Nat Rev Rheumatol 2014;10:374–380. [PubMed: 24686507]
- [60]. Tibshirani R Regression Shrinkage and Selection Via the Lasso. J R Stat Soc Ser B Methodol 1996;58:267–288.
- [61]. Tzourio-Mazoyer N, Landeau B, Papathanassiou D, Crivello F, Etard O, Delcroix N, Mazoyer B, Joliot M. Automated Anatomical Labeling of Activations in SPM Using a Macroscopic Anatomical Parcellation of the MNI MRI Single-Subject Brain. NeuroImage 2002;15:273–289. [PubMed: 11771995]
- [62]. Ulrike Grömping. Relative Importance for Linear Regression in R: The Package relaimpo. J Stat Softw 2006;17:1–27.
- [63]. Vachon-Presseau E, Tétreault P, Petre B, Huang L, Berger SE, Torbey S, Baria AT, Mansour AR, Hashmi JA, Griffith JW, Comasco E, Schnitzer TJ, Baliki MN, Apkarian AV. Corticolimbic anatomical characteristics predetermine risk for chronic 739 pain. Brain 2016;139:1958–1970. [PubMed: 27190016]
- [64]. Wieland HA, Michaelis M, Kirschbaum BJ, Rudolphi KA. Osteoarthritis an untreatable disease? Nat Rev Drug Discov 2005; 4:331–344. [PubMed: 15803196]
- [65]. Winkler AM, Ridgway GR, Webster MA, Smith SM, Nichols TE. Permutation inference for the general linear model. NeuroImage 2014; 92:381–397. [PubMed: 24530839]

- [66]. Yarkoni T, Poldrack RA, Nichols TE, Van Essen DC, Wager TD. Large-scale automated synthesis of human functional neuroimaging data. Nat Methods 2011; 8:665–670. [PubMed: 21706013]
- [67]. Zhang Y, Brady M, Smith S. Segmentation of brain MR images through a hidden Markov random field model and the expectation-maximization algorithm. IEEE Trans Med Imaging 2001; 20:45– 57. [PubMed: 11293691]
- [68]. Zigmond AS SR. The hospital anxiety and depression scale. Acta Psychiatr Scand 751 1983; 67:361–70.

Barroso et al. Page 18



#### **Figure 1. Whole brain cortical grey matter changes reflect lower volumes in anterior cingulate, paracingulate gyri in HOA.**

A. Average total neocortical grey matter (TGM) matter volume was not different between healthy controls, HOA and KOA patients (age, gender and intracranial volume were used as confound variables (F  $(2,145) = 0.07$ , p=0.79)). TGM differed by gender across groups, when controlling for total IC volume and age (F  $(1,145) = 15.7$ , p<0.001), with females showing higher volumes. C. Scatterplot represents total neocortical GM in relation to age for each subject, color-coded by group; slopes were computed for each group, showing decrease of TGM with age, for all groups (p<0.05; R2: controls=0.52; KOA=0.37; HOA=0.17). C. Grey matter morphological changes assessed by voxel-based morphometry (VBM). The three groups were contrasted together, using a 1-way between subjects'-ANCOVA, with age, gender and total ICV as covariates of no interest. Shown is the F-test statistical map, clustering determined using uncorrected p-value binarized at value of 0.001 and cluster size>66 voxels. Only cingulate/paracigulate gyri showed regional GM volume difference between the groups. D. Post hoc results for GM volume comparisons in the identified region – anterior cingulate/paracingulate gyri (ACC/PcG\_LR). Bars represent the mean grey matter volume and error bars represent the standard error. Tukey HSD test showed regional GM volume was lower in HOA from KOA and healthy subjects at p<0.001.KOA and controls did not show a statistically significant difference (p=0.34). \*p<0.001.

Barroso et al. Page 19



#### **Figure 2. Primary motor cortex, contralateral to OA pain, showed lower volume in both HOA and KOA patients.**

A. Grey matter morphological changes assessed by voxel-based morphometry (VBM) after aligning the brain in relation to OA pain (brain x axis left-to-right flipping for left sided KOA and HOA; which renders the left hemisphere to be contralateral to OA in all patients). The three groups were contrasted together, using a 1-way between subjects' ANCOVA, controlling for age, gender, total ICV and brain flipping. Shown is the F-test statistic map clustering determined using a TFCE uncorrected p-value binarized mask at value of 0.001 and cluster size >66 voxels. Multiple brain regions showed changes in regional volume B. Post-hoc results for GM volume differences for significant clusters, adjusted for age, gender and TICV (Tukey HSD tests at p<0.001); Bars represent the mean grey matter volume and error bars represent the standard error. Left primary motor cortex (precentral gyrus; PCG\_L) and right temporal pole (TP\_R) presented lower GM volume in both OA groups. Bilateral anterior cingulate cortex (ACC\_LR) showed lower GM volume only in HOA, when contrasted with the other two groups. In the KOA group, GM volume was greater for middle frontal gyrus cluster on the right (MFG\_R) and lower for right precuneus cortex (PC\_R). Further information on thresholded clusters can be found in Table 2. \*p<0.001.

Barroso et al. Page 20



**Figure 3. Univariable and multivariable analysis of regional GM differences between groups.** Cortex was subdivided into 90 ROIs approximating Brodmann areas, using the AAL atlas; A and B plots show the F-value for each ROI across group comparison in native space, without brain flipping (A) and after flipping the brains  $(B)$  based on OA body side dominance  $(x)$ axis). Age, gender, TICV and OA side in lateralized analysis were entered as covariates of no interest. Green line indicates uncorrected significance threshold at  $p<0.01$  (F<sub>(3,148)</sub>>4 and F  $(4.127)$  > 3.5 for A and B); brain regions above and close to uncorrected threshold are labeled and include subcortical regions, temporal and frontal regions. Primary motor cortex was identified solely in the flipped brain analysis (B) and was the only region surviving FDR correction for multiple comparisons (\*p<0.05). C. and D. show the coefficient estimates (log(odds)) in KOA training group after LASSO regularization with lambda selection via 10 fold cross-validation for native (C) and in flipped volumes (D). ROC for classification by logistic regression for the test KOA group and HOA group are also reported (blue=KOA; red=HOA). AUCs showed values between 0.77–0.90, suggesting that generated models were not overfit and selected parameters are relevant predictors of OA related brain morphological properties.



#### **Figure 4. Neuropathic pain profile was associated with KOA multivariable brain volume changes.**

A. Graphic representation of the AAL ROIs identified in the multivariable analysis (flipped brain). B. and C. plots show the top 40 (of about 1700) correlated psychological/task terms using Neurosynth reverse inference decoder for the brain regions (B) and weighted regions, given the model estimate coefficients (C). The size of a term in each word cloud is proportional to the correlation strength of term/region. Blue color represents top 5 correlations. D-E: A multiple regression model using the logistic regression/LASSO results for flipped brain KOA classification predicted neuropathic pain score (DN4) in a subsample of KOA patients (training sample, n=45). This result was validated on a KOA holdout sample (test sample, n=46), but not in the HOA group. C. shows the relative importance of all predictive variables where % of  $\mathbb{R}^2$  is normalized to sum 100%. Bars correspond to 1000 bootstrap confidence intervals at 95%. E. shows the validation of the obtained regression model in KOA testing sample and HOA groups. Predicted neuropathic pain scale (DN4)

strongly and significantly correlated with the actual score in the KOA test group, but not in the HOA. DN4 (Douleur Neuropathique en 4 questions); \*\*p<0.001.

#### **Table 1.**

Demographic and clinical characteristics of osteoarthritis patients and controls. To identify demographic difference between groups ANOVA and t-tests were performed for continuous data, controls and OA groups, and between OA groups, 773 respectively. Chi-square tests were performed for categorical data. 6MWT= 6 Minute Walking Test;  $BMI = Body$  mass index;  $DN4 = Douleur Neuropathique en 4 questions; HADS =$ Hospital anxiety and Depression Scale; HOOS = Hip injury and Osteoarthritis Outcome Score; KOOS = Knee injury and Osteoarthritis Outcome Score, (ADL = Activities of Daily Living; SR = Sports and Recreation; QoL=Quality of Life); PCS= Pain Catastrophizing Scale; KL = Kellgren-Lawrence scale; TUG= stand up and go test.



**PCS**, mean, SD



\* p<0.05

\*\*<br>p<0.001.

#### **Table 2.**

Clusters for GM volume differences in HOA, KOA and healthy controls for whole brain analysis with and without brain flipping based on OA body side dominance (x-axis).



All identified clusters from VBM GLM analysis, TFCE uncorrected p<0.001. Cluster peak coordinates (x,y,z) are displayed according to MNI atlas, labels accordingly to the Oxford-Harvard Structural Cortical Atlas [\(http://fsl.fmrib.ox.ac.uk/fsl/fslwiki/Atlases\)](http://fsl.fmrib.ox.ac.uk/fsl/fslwiki/Atlases). We accepted clusters greater than 66 voxels in size for post-hoc analysis, given the Gaussian Kernel (FWHM=8mm) used for smoothing.