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Statins and clinical outcomes with COVID-19: Meta-analyses of observational studies



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ABSTRACT

Aims. – People with cardiovascular disease or risk factors are at increased risk when exposed to SARS-CoV-2. Most are treated with statins, but the impact of these drugs on clinical outcomes of COVID-19 remains unclear. This report is therefore based on meta-analyses of retrospective observational studies aimed at investigating the impact of previous statin therapy in patients hospitalized for COVID-19. *Methods.* – In studies reporting on the clinical outcomes of COVID-19 in statin users *vs* non-users, two endpoints have been used—in-hospital death rates, and disease severity as assessed by admission to intensive care units (ICUs)—with a special focus on patients with diabetes.

Results. – Regarding mortality, 13 studies were included in the meta-analysis for a total of 10,829 statin users (2517 deaths) and 31,893 non-users (7516 deaths): univariate analysis showed no statistically significant reduction in deaths (OR: 0.97, 95% CI: 0.92–1.03), although between-study heterogeneity was high ($I^2 = 97\%$). As for disease severity, 11 studies were selected for a total of 3462 statin users (724 endpoints) and 10,560 non-users (1763 endpoints): here again, univariate analysis showed no reduction in severity (OR: 1.09, 95% CI: 0.99–1.22; $I^2 = 93\%$). Collectively, in 10 studies using multivariable analysis adjusted for the more prevalent baseline risk factors among statin users, lower OR values were reported than with univariate analyses (0.73 ± 0.31 vs 1.44 ± 0.84, respectively; P = 0.0028; adjusted OR: P = 0.0237 vs non-users). Limited but conflicting findings were observed for diabetes patients. *Conclusion.* – Although no significant reductions in either in-hospital mortality or COVID-19 severity were reported among statin users compared with non-users after univariate comparisons, such reductions were observed after adjusting for confounding factors. These highly heterogeneous observational findings now require confirmation by ongoing randomized clinical trials.

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Introduction

Several risk factors for more severe coronavirus disease 2019 (COVID-19) infection have been identified in the general population [1,2] as well as in the diabetes population [3,4]. In addition to older age, male gender and obesity as risk factors, patients with comorbidities such as diabetes mellitus (DM) [5], arterial hypertension [6] and established cardiovascular (CV) disease have been identified as being more prone to progress to more severe acute respiratory syndrome coronavirus 2 (SARS-CoV-2) infection requiring admission to intensive care units (ICUs) because of acute respiratory distress syndrome (ARDS) and leading to premature death [7]. Indeed, according to international guidelines, most patients with CV risk factors and/or CV disease are treated with

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statins to mitigate their overall CV risk [8]. Remarkably, however, in an extensive review of COVID-19 and CV disease covering the topic from basic mechanisms to clinical perspectives, statin therapy was not considered at all [9]. In fact, it is still unclear as to whether statins can positively or negatively influence prognoses of COVID-19 [10,11].

Two previous meta-analyses of observational studies compared clinical outcomes in statin users *vs* non-users and gave conflicting results: one demonstrated a positive impact [12], whereas the other failed to show any significant differences in prognosis [13]. Nevertheless, these meta-analyses were performed using a rather limited number of studies (some only available as preprints). Furthermore, a recent analysis of the Coronavirus SARS-CoV-2 and Diabetes Outcomes (CORONADO) study concluded that the routine use of statins is associated with an increased in-hospital mortality related to COVID-19 in patients with type 2 DM [14], findings that are in contrast to the reduction in death rate found in a US study of diabetes patients treated with statins

[15]. Thus, the available reported data are divergent, preventing any definite conclusions to be drawn on the effects of statin therapy on COVID-19 clinical outcomes.

For this reason, the aim of the present meta-analyses of retrospective observational studies was to investigate the impact of statin therapy on two major clinical outcomes—in-hospital mortality and disease severity requiring ICU admission, mainly for invasive mechanical ventilation (IMV)—in patients hospitalized with COVID-19. Also, in light of the recent published report by CORONADO investigators [14], a particular focus has also been placed on other studies that recruited patients with diabetes.

Materials and methods

Data sources and search strategy

Electronic searches were performed in PubMed and Scopus from December 2019 to December 2020 using the following search terms: 'statin' combined with 'COVID-19' or 'SARS-CoV-2' or 'coronavirus' (Fig. 1). Two independent researchers (A.S., M.M.) screened the literature, analyzed the selected studies and summarized the search results, using the same inclusion and exclusion criteria (see below). Any resulting discrepancies were resolved by mutual discussion. The reference lists of previous systematic reviews and meta-analyses [12,13] and of any related narrative reviews or commentaries [16], especially those involving





diabetes populations [3,4], were manually examined to identify any additional publications relevant to the present study. A search for duplicates was done manually.

Inclusion and exclusion criteria

In the absence of randomized clinical trials (RCTs), only observational studies could be included in the analyses. The inclusion criterion was observational studies in patients with COVID-19, from which the following relevant information was collected: (*i*) number of patients fulfilling criteria for severe illness (mainly assessed as admission to ICU or a need for IMV) and/or inhospital death; and (*ii*) number of patients treated with any statin before admission ('statin users') or not ('non-users'). All types of observational studies were taken into account (retrospective, single-centre, multicentre, cohort study, case series ...) in the initial literature search whatever the number of subjects involved. Also, despite the present study being mainly focused on hospitalized patients, other studies of outpatients were also considered for a separate complementary investigation.

Exclusion criteria were: (*i*) commentaries or hypothetical mechanistic papers; (*ii*) narrative or comprehensive reviews; (*iii*) preprint papers not yet peer-reviewed or officially published; and (*iv*) observational studies with no reporting of the number of statin users and non-users who developed either of the two study outcomes (in-hospital mortality or severe COVID-19 infection). If important data were not clearly reported in the original papers, their corresponding authors were contacted to obtain further valuable information if available.

Data extraction

The following data were extracted from all included studies: (*i*) total number of patients involved in the study; (*ii*) number of statin users and non-users; (*iii*) number of patients who progressed to more severe forms of COVID-19 (ICU admission and/or IMV) or who died within either 7 or 28 days of hospitalization (if this information was available); (*iv*) mean age of the overall study population (including statin users and non-users if this information was available); and (*v*) mode of comparison used in the reported findings (univariate, multivariate, propensity scorematching, logistic regression modelling). Studies including patients with diabetes were also identified along with type of DM and, finally, studies of outpatients rather than in-hospital patients were collected to be analyzed separately.

Statistical analysis

Differences between statin users *vs* non-users were examined by odds ratio (OR) and 95% confidence interval (CI), with P <0.05 considered statistically significant. Interstudy heterogeneity was assessed using Cochran's Q test statistic and the Higgins and Thompson I² index; if the latter value was > 50%, this indicated a substantial degree of heterogeneity. Meta-analyses were performed using Review Manager (RevMan) 5.3 software (The Nordic Cochrane Centre, Copenhagen, Denmark). When separate analyses were available in some studies, comparisons between univariate *vs* multivariate/adjusted results were performed using paired *t* test.

Results

Study characteristics

The study selection process is depicted in Fig. 1. A total of 325 references were identified by the initial electronic search using the MeSH (Medical Subject Headings) terms 'statin' combined with

'COVID-19' or 'SARS-CoV-2' or 'coronavirus'. As the first step, duplicates and citations based on study characteristics were excluded after title and abstract evaluation (n = 245). The second step was to exclude other citations based on the inclusion/ exclusion criteria after abstract/full-text evaluation.

Initially, 24 eligible observational studies were identified that reported valuable data on in-hospital mortality and/or COVID-19 disease severity, including three studies of patients with diabetes [14,15,17]. Of these selected studies, 14 provided data on in-hospital mortality [15,18–30] (Table 1)—although one reported only adjusted hazard ratios (HRs) with no crude numbers of deaths in statin users vs non-users and, therefore, could not be included in the meta-analysis (authors were contacted, but did not communicate their results) [26]—while 11 studies reported data on COVID-19 severity (as determined by ICU, IMV; Table 2) [18,22,23,27,28,30–35]. Also, a study by Tan et al. [36] collected very few events in a Chinese population (personal communication from the authors), thereby preventing the use of this study in the final analysis. In addition, a Japanese study by Higuchi et al. [37] included only seven patients

Table 1

treated with statins; however, as the very high reported HR could be considered an outlier compared with the other studies, this study was likewise not included in the meta-analysis.

Three studies that specifically focused on patients with DM were also identified [14,15,17] (Table 3) as well as some studies carried out in outpatients exposed to SARS-CoV-2 (see Discussion section below). However, these studies were not included in the two main meta-analyses, but were instead analyzed separately.

In-hospital mortality

The included observational studies reporting data on inhospital mortality are listed in Table 1 [15,18–30]. Results of the corresponding meta-analysis are illustrated in Fig. 2 (which includes all studies except for Bifulco et al. [26]). A total of 13 studies were included in the meta-analysis for in-hospital mortality, involving a total of 10,829 statin users (2517 deaths) and 31,893 non-users (7516 deaths). Univariate analysis found no

Reference	Country	Type of study	Patients (total n)	Mean age ^a (years)	Statin (n/N)	No statin (n/N)	HR (95% CI) ^b	HR (95% CI) ^c P	Outcome
Zhang et al. [18]	China	Retrospective, multicentre	4305	58 (66 vs 57), 65 <i>vs</i> 65 after 4:1 PSM	45/861 (5.2%)	325/3444 (9.4%)	0.53 (0.38–0.73)	0.58 (0.43-0.80), univariate after 4:1 PSM, P = 0.001	28-day mortality
Mallow et al. [19]	USA	Retrospective, multicentre	21,676	65	1039/5313 (19.6%)	3896/ 16,363 (23.8%)	0.78 (0.72–0.84)	0.54 (0.49–0.60), logistic regression, <i>P</i> < 0.001	In-hospital mortality
Krishnan et al. [20]	USA	Retrospective, single-centre	152	66	57/92 (62.0%)	35/71 (49.3%)	1.68 (0.89–3.14)	NA	ICU mortality
Rodriguez-Nava et al. [21]	USA	Retrospective, single-centre	87	68	23/47 (48.9%)	25/40 (62.5%)	0.57 (0.24–1.36	0.38 (0.18–0.77), multivariable Cox regression, <i>P</i> = 0.008	In-hospital (ICU) mortality
Saeed et al. [15]	USA	Retrospective, single-centre	4252	65	312/1355 (23.0%)	782/2897 (27.0%)	0.81 (0.70-0.94)	0.88 (0.83–0.94), PSM, P < 0.01	In-hospital mortality
Gupta et al. [22]	USA	Retrospective cohort (2 centres)	All cohort: 2626, PSM: 1296	70/62, (69/71)	NA/951, 96/648 (14.8%)	NA/1675, 172/648 (26.5%)	0.48 (0.36–0.64), univariate after 1:1 PSM	0.59 (0.38-0.63), multivariable, adjusted, all cohort, P < 0.001	30-day mortality
Song et al. [23]	USA	Retrospective, single-centre	249	62 (71 vs 54)	27/123 (22.0%)	15/126 (11.9%)	2.08 (1.05–4.14)	0.88 (0.37–2.08), fully adjusted, P = 0.781	In-hospital mortality
Grasselli et al. [24]	Italy	Retrospective, multicentre	3988	63 (66 vs 61)	479/741 (64.6%)	1411/3165 (44.6%)	2.27 (1.92–2.68)	0.98 (0.81–1.20), multivariable, P = 0.87	In-hospital mortality ^d
Rossi et al.	Italy	Retrospective, single-centre	71	72 (71 vs 73)	9/42 (21.4%)	10/29 (34.5%)	0.52 (0.18–1.50)	NA P < 0.05	In-hospital mortality
Bifulco et al. [26]	Italy	Retrospective, single-centre	541 (123 deaths)	65 (73 vs 63)	NA/117	NA/424	NA	0.75 (0.26–2.17), adjusted P = 0.593	In-hospital mortality [15]
Masana et al. [27]	Spain	Retrospective, multicentre	1162 (after genetic matching)	67 (73 vs 62)	115/581 (19.8%)	148/581 (25.4%)	0.72 (0.55–0.95), univariate (after genetic matching)	0.60 (0.39–0.92), competing-risks, P = 0.02	In-hospital mortality
Butt et al. [28]	Denmark	Cohort	4842	(73 vs 50)	292/843 (34.6%)	589/3999 (14.7%)	2.57 (2.34–2.96), unadjusted Cox regression	1.05 (0.89–1.23), fully adjusted	All-cause mortality
Alamdari et al.	Iran	Retrospective,	459	62	6/117 (5.1%)	57/342 (16.7%)	0.27 (0.11-0.64)	NA P = 0.002	In-hospital mortality
Soleimani et al. [30]	Iran	Retrospective, single-centre	254	66	17/66 (25.8%)	51/188 (27.1%)	0.93 (0.49–1.76)	NA	In-hospital mortality

n/N, number of deaths/number of patients; HR, hazard ratio; CI, confidence interval; PSM, propensity score-matching; NA, not available; ICU, intensive care unit. ^a Users vs non-users.

^b Univariate models (see Fig. 2).

^c Multivariate or adjusted models.

^d From ICU admission to hospital discharge.

Statin use and in-hospital mortality in patients with coronavirus disease 2019 (COVID-19).

Table 2

Statins and severe disease in hospitalized patients with coronavirus disease 2019 (COVID-19).

Reference	Country	Type of study	Patients (total n)	Mean age, years ^a	Statin (n/N)	No statin (n/N)	HR (95% CI) ^b	HR (95% CI) ^c P	Outcome
Zhang et al. [18]	China	Retrospective, multicentre	4305	58 (66 vs 57), 65 vs 65 after 4:1 PSM	64/861, 24/861	353/3444, 192/3444	0.80 (0.62–1.05), 0.51 (0.34–0.78)	P = 0.110, P = 0.002	icu, IMV
Yan et al. [31]	China	Retrospective, multicentre	619	49	5/16 (31.3%)	123/594 (20.7%)	1.74 (0.59–5.10)	0.98 (0.32–2.99), P = 0.97	Severe or critical disease
Gupta et al. [22]	USA	Retrospective cohort (2 centres)	All cohort: 2626, PSM: 1296	70/62, 70 (69/71)	NA/951, 179/648 (27.6%)	NA/1675, 269/648 (41.5%)	NA, 0.54 (0.43–0.68), univariate (after 1:1 PSM)	0.54 (0.44–0.67), multivariable, adjusted, all cohort, P < 0.001	IMV or mortality
Song et al. [23]	USA	Retrospective, single-centre	249	62	19/123 (15.4%)	26/126 (20.6%)	0.70 (0.37–1.35)	0.45 (0.20–0.99), fully adjusted, P = 0.048	Tracheal intubation
Argenziano et al. [32]	USA	Retrospective, single-centre, case series	1000 (850 with statin data)	63	62/218 (28.4%)	174/632 (27.5%)	1.05 (0.74–1.47)	NA	ICU
Daniels et al. [33]	USA	Retrospective, single-centre	170	59	20/46 (43.5%)	70/124 (56.5%)	0.59 (0.30–1.17)	0.29 (0.11-0.71), multivariable logistic regression, <i>P</i> = 0.009	Death or ICU
Dreher et al. [34]	Germany	Retrospective, single-centre	50	65	9/18 (50.0%)	15/32 (46.9%)	1.13 (0.36–3.60)	NA	ARDS
Masana et al. [27]	Spain	Retrospective, multicentre	1162	67 (73 vs 62)	84/581 (14.5%)	96/581 (16.6%)	0.85 (0.52–1.17), univariate (after genetic matching)	P = 0.36	IMV
Butt et al. [28]	Denmark	Cohort	4842	73 vs 50	204/843 (24.2%)	419/3999 (10.5%)	2.41 (2.04–2.85), unadjusted Cox regression	1.16 (0.95–1.41), fully adjusted	Severe disease
Meunier et al. [35]	France	Retrospective, single-centre	234	67	26/42 (61.9%)	88/192 (45.8%)	1.92 (0.97–3.81)	NA	Severe disease
Soleimani et al. [30]	Iran	Retrospective, single-centre	254	66	52/66 (78.8%)	130/188 (69.1%)	1.66 (0.85–3.23)	NA	Severe disease

n/N, number of deaths/number of patients; HR, hazard ratio; CI, confidence interval; PSM, propensity score-matching; ICU, intensive care unit; IMV, invasive mechanical ventilation; NA, not available; ARDS, acute respiratory distress syndrome.

^a Users vs non-users.

^b Univariate model (see Fig. 3).

^c Multivariate or adjusted model.

statistically significant reduction in death rate (OR: 0.97, 95% CI: 0.92–1.03) on comparing statin users with non-users, although there was substantial between-study heterogeneity ($I^2 = 97\%$).

Clinical outcomes of severity

The list of included observational studies reporting data on disease severity (ICU, IMV) is presented in Table 2 [18,22,23,27,28,30–35], while results of the corresponding meta-analysis are illustrated in Fig. 3. A total of 11 studies were included in this meta-analysis for a total of 3462 statin users (724 severity endpoints) and 10,560 non-users (1763 severity endpoints). Univariate analysis could find no reduction in the incidence of severe COVID-19 among statin users (OR: 1.09, 95% CI: 0.99–1.22) and, once again, high heterogeneity ($I^2 = 93\%$) was noted.

Multivariable analysis and adjusted comparisons

As statin users have different clinical characteristics from non-users, which might expose them to a greater risk of developing more severe COVID-19 infection, it is important to take into account the potentially confounding factors that could influence the final results. In 10 series of data—six considering in-hospital mortality (Table 1) and four considering COVID-19 severity (Table 2) [19,21,23,24,27,28,31,33]—multivariable analyses or those adjusted for covariates, if available,

resulted in more favourable results than univariate analyses in patients treated with statins compared with patients not treated with statins at admission to hospital: adjusted OR (mean \pm SD) 0.73 \pm 0.31 vs nonadjusted OR 1.44 \pm 0.84; P = 0.0028; Fig. 4). Thus, the adjusted OR revealed a statistically significant reduction by 27% of hard clinical outcomes in statin users vs non-users (P = 0.0237). Furthermore, all studies that used propensity score matching for their comparison of inhospital mortality reported hazard ratios below one in favour of statin users, respectively, 0.58 (P = 0.001) [18], 0.88 (P < 0.01) [15] and 0.48 (P < 0.001) [22] (Table 1). One well-recognized risk factor that can easily be identified is older age in statin users vs non-users, especially in studies reporting higher HRs (> 2) on univariate analyses [23,24,28] (Tables 1 and 2). However, it was not possible to identify any other risk factors for more severe COVID-19 and its associated increased death rate. although some studies reported more male patients, more diabetics, more hypertensives and/or more people with previous CV disease in statin users than in nonusers, a finding in agreement with international guidelines for optimal management of such at-risk patients [8].

Results in patients with diabetes

Only a few studies focused on the impact of statins in patients with diabetes, with divergent results (Table 3). While Saeed et al. [15] reported significant reductions in in-hospital mortality among statin-users compared with non-users in a US study (HR: 0.88, 95%)

Table 3

Statin use and clinical outcomes (tracheal intubation or mortality) in diabetes patients with coronavirus disease 2019 (COVID-19).

Reference	Country	Type of study	Patients (total n)	Mean age (years)	Statin (n/N)	No statin (n/N)	HR (95% CI)	Р	Outcome
Cariou et al. [14] (type 2 diabetes patients)	France	Retrospective, multicentre, hospitalized patients	2449	70.9	355/1192 (29.8%), 431/1192 (36.2%)	339/1257 (27.0%), 425/1257 (33.8%)	1.38 (1.04–1.83) after IPTW, 1.22 (0.98–1.58) after IPTW	<i>P</i> = 0.1338, <i>P</i> = 0.2191	7-day tracheal intubation and/or death (primary outcome), 28-day primary outcome
					220/1192 (18.5%), 229/1192 (19.2%)	235/1257 (18.7%), 248/1257 (19.7%)	1.18 (0.86–1.61) after IPTW, 1.13 (0.83–1.53) after IPTW	P = NA, P = NA	7-day tracheal intubation, 28-day tracheal intubation
					153/1192 (12.8%), 285/1192 (23.9%)	123/1257 (9.8%), 229/1257 (18.2%)	1.74 (1.13–2.65) after IPTW, 1.46 (1.08–1.95) after IPTW	<i>P</i> = 0.02, <i>P</i> < 0.001	7-day mortality, 28-day mortality
Saeed et al. [15] (diabetes patients)	USA	Retrospective, single-centre, hospitalized patients	2266	68.0	236/983 (24%)	500/1283 (39%)	0.51 (0.43-0.61), multivariable, adjusted, 0.88 (0.84-0.91) after IPTW	P < 0.001, P < 0.001	In-hospital mortality
Holman et al. [17] (type 2 diabetes patients)	UK	Population-based cohort	2,874,020	67.5	7355/ 2,099,505 (3.5%)	3086/752,245 (4.1%)	0.72 (0.69–0.75) adjusted for demographic/ clinical characteristics	<i>P</i> < 0.0001	Covid-19-related death (no focus on hospital)
Holman et al. [17] (type 1 diabetes patients)	UK	Population-based cohort	264,390	46.6	338/118,995 (2.7%)	120/142,710 (0.8%)	0.82 (0.65–1.03) adjusted for demographic/ clinical characteristics	<i>P</i> = 0.081	Covid-19-related death (no focus on hospital)

n/N, number of deaths/number of patients; HR, hazard ratio; CI: confidence interval; IPTW, inverse probability of treatment weighting (with logistic regression analysis after propensity score-matching); NA, not available.

	Statin users S		Statin non-users		Odds Ratio		Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	M-H, Fixed, 95% Cl
Zhang et al	45	861	325	3444	4.5%	0.53 [0.38, 0.73]	
Mallow et al	1039	5313	3896	16363	56.3%	0.78 [0.72, 0.84]	
Krishnan et al	57	92	35	71	0.6%	1.68 [0.89, 3.14]	
Rodriguez-Nava et al	23	47	25	40	0.5%	0.57 [0.24, 1.36]	
Saeed et al ALL	312	1355	782	2897	14.1%	0.81 [0.70, 0.94]	-
Gupta et al	96	648	172	648	5.4%	0.48 [0.36, 0.64]	-
Song et al	27	123	15	126	0.4%	2.08 [1.05, 4.14]	
Grasselli et al	479	741	1411	3165	6.9%	2.27 [1.92, 2.68]	+
Rossi et al	9	42	10	29	0.3%	0.52 [0.18, 1.50]	
Masana et al	115	581	148	581	4.4%	0.72 [0.55, 0.95]	
Butt et al	292	843	589	3999	4.9%	3.07 [2.60, 3.62]	-
Alamdari et al	6	117	57	342	1.0%	0.27 [0.11, 0.64]	
Soleimani et al	17	66	51	188	0.7%	0.93 [0.49, 1.76]	
Total (95% CI)		10829		31893	100.0%	0.97 [0.92, 1.03]	•
Total events	2517		7516				
Heterogeneity: Chi² = 382.59, df = 12 (P < 0.00001); I² = 97%							
Test for overall effect: Z	= 0.99 (P	= 0.32)					Statin users Statin non-users

Fig. 2. Meta-analysis of studies comparing in-hospital mortality in statin users vs non-users with COVID-19 infection. M-H, Mantel-Haenszel method.

CI: 0.84–0.91), the opposite results were reported by Cariou et al. [14] in the French observational CORONADO study (HR: 1.46, 95% CI: 1.08–1.95). Such differences could not be explained by different modes of comparison, as both studies used inverse probability of treatment weighting approaches for their comparisons. Thus, no clear explanation for such conflicting results have been proposed. Holman et al. [17] reported better prognoses for COVID-19 patients with type 2 DM and, to a lesser extent, with type 1 DM, who were treated with statins; however, these data were obtained in a different outpatient cohort (population-based cohort study; Table 3).

Discussion

Both meta-analyses of retrospective observational studies investigating the effects of previous statin therapy on in-hospital mortality associated with COVID-19 and severity of the infection (admission to ICU, need for IMV) revealed no reductions in statin users compared with non-users, whereas a substantial yet poorly explained heterogeneity was observed between studies. In any case, these results should be interpreted with caution because statin users are generally patients who have additional risk factors,

	Statin u	sers	Statin non-users		Odds Ratio		Odds Ratio	
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	M-H, Fixed, 95% Cl	
Zhang et al	64	861	353	3444	19.8%	0.70 [0.53, 0.93]		
Yan et al	5	16	123	594	0.7%	1.74 [0.59, 5.10]		
Gupta et al	179	648	269	648	29.4%	0.54 [0.43, 0.68]	+	
Song et al	19	123	26	126	3.3%	0.70 [0.37, 1.35]		
Argenziano et al	62	218	174	632	9.7%	1.05 [0.74, 1.47]	↓ •	
Daniels et al	20	46	70	124	3.2%	0.59 [0.30, 1.17]		
Dreher et al	9	18	15	32	0.8%	1.13 [0.36, 3.60]		
Masana et al	84	581	96	581	12.4%	0.85 [0.62, 1.17]		
Butt et al	204	843	419	3999	16.7%	2.73 [2.26, 3.29]	+	
Meunier et al	26	42	88	192	1.8%	1.92 [0.97, 3.81]		
Soleimani et al	52	66	130	188	2.2%	1.66 [0.85, 3.23]		
Total (95% CI)		3462		10560	100.0%	1.09 [0.99, 1.22]	•	
Total events	724		1763					
Heterogeneity: Chi² = 149.03, df = 10 (P < 0.00001); l² = 93%								1
Test for overall effect:	Z=1.69 (P = 0.09	3)				STATIN USERS NON STATIN USERS	
Test for overall effect: $Z = 1.69$ (P = 0.09)							0.01 0.1 1 10 100 STATIN USERS NON STATIN USERS	

Fig. 3. Meta-analysis of studies comparing outcomes according to severity of COVID-19 infection in statin users vs non-users. M-H, Mantel-Haenszel method.



Fig. 4. Comparison of hazard ratios (HRs) in statin users *vs* non-users using multivariate (adjusted) analysis compared with univariate analysis based on 10 sets of data, including six on in-hospital mortality (solid circles) and four on disease severity (open circles). For more detailed information, see Tables 1 and 2. Data are from references [19,21,23,24,27,28,31, and 33].

such as older age, male gender, DM, hypertension and CV disease [8], all of which have been reported to worsen COVID-19 prognoses [7,9,38]. This difference in patients' characteristics between statin users and non-users explains why the fully adjusted and multivariate analyses gave more favourable results than univariate analyses (Tables 1 and 2). In fact, the results using adjusted data confirm those reported in the meta-analysis by Kow et al. [12], who reported a 30% reduction in fatal or severe COVID-19 infection. On the other hand, the univariate analyses confirm the lack of significant protection among statin users reported in the meta-analysis by Hariyanto et al. [13].

Previous studies before the COVID-19 outbreak had already reported favourable statin effects on the outcome of severe pulmonary infection [39], especially in the presence of a hyperinflammatory phenotype of ARDS, that were associated with improved survival with simvastatin compared with a placebo [40]. In contrast, other studies have reported disappointing results with the use of statins as a late treatment for ARDS [41,42]. On the other hand, statins have been shown to improve prognoses and reduce mortality among patients in hospital with influenza virus infections [43–45]. An association between outpatient statin treatment and reduction of disease severity among patients hospitalized during the 2009 H1N1 pandemic has also been reported [46]. In 2015, it was even suggested, based on theoretical grounds, that statins might be able to decrease the fatality rate of Middle East Respiratory Syndrome (MERS) infection [47]. However, caution is nonetheless required in the absence of RCTs. Indeed, while observational studies have reported improved outcomes in patients with various infections (including community-acquired pneumonia) or sepsis who were taking statins, most RCTs of inpatients with sepsis or pneumonia requiring IMV failed to demonstrate any beneficial effect with statin therapy [48,49].

Some retrospective studies have suggested the possible detrimental effect of reduced serum low-density lipoprotein (LDL) cholesterol levels on COVID-19 prognoses [50]. However, reverse causality-SARS-CoV-2 infection as a cause of LDL cholesterol reduction-instead of true causality, where LDL cholesterol reduction is a factor promoting the viral infection, might explain the association between low LDL cholesterol and severe COVID-19 manifestations. As previously discussed [51], such an association may distract from refutations of the potential benefits of statin therapy in clinical settings of patients exposed to SARS-CoV-2 infection. In fact, every study included in the present meta-analyses failed to present results on blood lipid levels, and only a few [18,19,21] made clear mention of whether or not statin use identified at admission was maintained throughout the stay in hospital [16]. In one study [52] focusing on liver abnormalities associated with COVID-19, statin use was commonplace both before admission (40%) and during hospitalization (80%), with no differences in peak liver biochemistry values between users and non-users.

Cariou et al. [14], using data from the large-scale multicentre CORONADO study carried out in France, reported that routine statin treatment is significantly associated with increased mortality (based on 7-day and 28-day in-hospital death rates) in patients with type 2 DM hospitalized for COVID-19. However, the composite primary outcome, comprising tracheal intubation and/or death within either 7 or 28 days of admission, was not statistically different between statin users and non-users (Table 3). The association between statin use and outcomes was estimated by logistic regression analysis after applying inverse probability of treatment weighting using a propensity score-weighting approach. The CORONADO study included type 2 DM patients who had a higher risk of mortality than in other studies, especially the one by Holman et al. [17], whose population-based cohort study in the UK reported positive results with statins mostly in patients with type 2 DM. Thus, it was proposed that the effect of statins on COVID-19 prognosis might vary according to infection

severity [14]. However, this hypothesis is still only speculative and has yet to be confirmed.

In a US study of hospitalized diabetes patients, the mortality rate was even higher than that reported in CORONADO, although statin use was significantly associated with a reduction of inhospital mortality [15]. More important, this difference was observed despite the use of the same statistical approach in both studies (inverse probability of treatment weighting: Table 3). In addition, it should be noted that this same CORONADO study, using a similar approach to analyze collected data, found that previous metformin therapy reduced both 7-day and 28-day in-hospital mortality rates [53], a finding in agreement with those reported in other studies [54]. Nevertheless, the reason behind such mixed results in the CORONADO study of a protective effect with metformin and deleterious effects with statins remains unclear, especially as both statins [55] and metformin [54] exert pleiotropic effects that could contribute to reducing inflammation and inducing vascular protection.

In addition to the studies of patients hospitalized due to COVID-19 included in the present meta-analyses, two other studies were carried out outside of hospital and reported contrasting results. The first was performed in COVID-19-infected older adults residing in nursing homes in Belgium. This retrospective multicentre cohort study found an association between statin use and the absence of symptoms in COVID-19 that remained statistically significant even after adjusting for covariates (OR: 2.65, 95% CI: 1.13-6.68). However, the effect of statin intakes on serious clinical outcomes was not statistically significant, albeit trending in the same beneficial direction (OR: 0.75, 95% CI: 0.24-1.87) [56]. In the multinational multicentre Lean European Open Survey on SARS-CoV-2-Infected Patients (LEOSS) cohort study, a significant univariate association between statin intake and increased risk of complicated clinical stages of COVID-19 at diagnosis was found (OR: 1.40, 95% CI: 1.09–1.80; *P* = 0.009). Unfortunately, statin use was excluded from multivariable analyses (adjusted for age, gender, underlying CV diseases, DM, pulmonary diseases) due to model quality, as admitted by the authors [57]. Differences between the results of these two studies are most likely explained by the adjustment for covariates in the Belgian study [56] compared with only univariate associations in the international one [57]. Finally, in a symptom surveillance study for COVID-19 in Germany, a statistically significant inverse relationship between self-reported typical COVID-19 symptoms and self-reported statin therapy was found [58], although whether or not statin therapy had beneficial effects for combatting COVID-19 could not be deduced from the survey.

Divergent mechanisms have been discussed regarding the potential impact of statins on COVID-19 infection. On the one hand, negative effects are possible given that statins may increase cellular expression of angiotensin-converting enzyme 2 (ACE2), the primary receptor allowing entry of SARS-CoV-2 into human cells [59,60]. It has also been reported that lower LDL cholesterol levels are associated with COVID-19 severity [61]. On the other hand, evidence in silico suggests that statins might serve as efficient inhibitors of the main protease of SARS-CoV-2 [62]. In addition, statin-induced decreases of cholesterol levels in plasma membranes could alter the assembly of ACE2 receptors, resulting in failure of SARS-CoV-2 internalization [10,63]. Moreover, statins may exert anti-inflammatory [55], immunomodulatory [64] and antioxidant properties [65]. As cytokine storms and consecutive bouts of inflammation play crucial roles in the development of poorer outcomes for COVID-19 [66], positive effects may be expected from these pleiotropic effects of statins [11,51,67]. It should also be noted that statins have been described to have antiviral properties [68]. Finally, statins, through their antithrombotic properties [69], could also attenuate diffuse thrombo-

Table 4

Hypothetical negative and positive statin mechanisms capable of influencing coronavirus disease 2019 (COVID-19) outcomes.

Negative mechanisms/ concerns	Positive mechanisms
Enhancement of SARS-CoV-2 entry through increased cellular expression of ACE2 [11.60]	Reduction of viral cellular entry through decreased cell membrane cholesterol content (antiviral activity) [63,68]
Possible weakened leucocyte function [81]	ACE2-mediated conversion of angiotensin II to angiotensin [1–7,60,73]
Drug-drug interactions with antiviral agents [59]	Inhibition of SARS-CoV-2 main protease (Mpro) [62]
and rhabdomyolysis [82]	Modulation of autophagy [71]
Potential hepatotoxicity [35]	Anti-inflammatory effects [55] Immunomodulatory effects [64] Oxidative stress reduction [65] Antithrombotic and endothelial effects [69] Cardiovascular protection [8]

For more information, see Subir et al. [10]; SARS-CoV-2, severe acute respiratory syndrome coronavirus 2; ACE2, angiotensin-converting enzyme 2.

sis, a key complication characteristic of SARS-CoV-2 infection that can markedly worsen a poor prognosis [66]. Table 4 summarizes the various hypothetical mechanisms that might either negatively or positively influence COVID-19 outcomes. In general, there are more arguments favouring the continued use of statin therapy rather than its interruption in patients exposed to SARS-CoV-2 [11,67,70]. While an in-depth description of these potential mechanisms is beyond the scope of the present report, such details may be found elsewhere [10,11,51,68,71–73].

Among the strengths of the present study is that it has analyzed a much larger set of observational studies than the two previously published meta-analyses of the same topic [12,13] while including only peer-reviewed papers, with additional information obtained from some other investigators upon request. In addition, the present meta-analysis has looked at the effects of both in-hospital mortality and admission to ICUs (with or without IMV) and produced consistent results.

However, several limitations should also be acknowledged. First, none of the included studies was prospectively designed to test the study hypothesis and all were retrospective observational studies, thereby bringing the possibility of inherited biases. Second, the substantial heterogeneity observed between the selected studies may blunt the robustness of any conclusions drawn. Third, the characteristics of statin users were different from those of nonusers because of the lack of randomization and the classic recommendations for treating patients at risk of CV disease with statins [8]. If anything, the identified risk factors for more severe COVID-19 infection should be more prevalent among statin users than non-users (older age, male gender, previous CV disease, hypertension, DM) [7], yet such detailed information is missing from some reports. Fourth, as emphasized by Fedson [16], most of the papers selected for inclusion in these meta-analyses made no clear mention of whether or not statin use identified at admission was continued throughout the hospital stay, thereby allowing some speculation as to whether the maintenace of statins, or not, may have contributed to the wide heterogeneity observed between studies. In addition, certain important information was missing from most studies, such as the duration of previous statin therapy, the type and dosage of statin used, and the indication (primary or secondary CV prevention) for which it was given, all variables that might have influenced the final results. Moreover, detailed information on plasma lipid levels with statin therapy was absent from all of the selected retrospective observational studies. As a final limitation, the statistical significance of a *P* value of 0.0237, as observed for the statin-associated reduction in hard clinical outcomes related to COVID-19 on multivariate-adjusted analysis, is within the range (from 0.05 to 0.005) where any translation to strong clinical relevance remains a subject of debate [74].

Nevertheless, the present meta-analytical findings still serve to strengthen practical guidelines for the management of patients with COVID-19. especially those at high CV risk being treated with statins. Indeed, as SARS-CoV-2 infection is associated with a heavy inflammatory burden that can induce vascular damage, myocarditis and cardiac arrhythmias [9], these findings suggest that CV risk factors, including dyslipidaemias [8], should be judiciously controlled as per evidence-based guidelines [75]. On the other Hand, as there is still no evidence that lipid-lowering therapy is unsafe in patients with COVID-19, such treatments should not be interrupted just because of the pandemic or in patients at increased risk of COVID-19 infection. Patients exposed to SARS-CoV-2 who are already using statin therapy should continue the treatment if not contraindicated by European Society of Cardiology guidelines [76] and other US recommendations [38]. On the other hand, great care should be taken to avoid any adverse interactions between lipid-lowering medications and drugs that might be used to treat COVID-19 (for instance, some antiviral agents) [59], especially in patients with abnormal liver function tests or myopathies [77]. One caveat is that, while there is no reason to stop statin therapy in patients hospitalized for COVID-19, the level of evidence remains too low-in the absence of RCTs, some of which are ongoing as reported elsewhere [38,78,79], although the impact of their results may be attenuated by successful vaccine strategies in the future-to justify initiating statin therapy in statin-free patients in hopes of reducing the burden of disease and improving clinical outcomes [38,80].

Conclusion

Patients with the usual comorbidities, including hypertension, CV disease and DM (especially type 2), are at greater risk of severe COVID-19 infection and its related ARDS, requiring admission to ICUs for IMV, while also being exposed to an increased risk of mortality, and most of these patients are taking statins routinely as per DM and CV guidelines. Overall, meta-analyses of observational studies have concluded (even though caution is required because of substantial between-study heterogeneity) that statin users are at similar risk of severe illness and in-hospital mortality as are statin non-users, at least according to univariate analyses. However, having such a similar prognosis conflicts with the expected high risk in statin-treated patients, whose different baseline characteristics should confer greater risk in the presence of COVID-19 (older age, more comorbidities and/or CV antecedents). Nevertheless, multivariable adjusted findings for potential confounding factors have indicated improved prognoses among statin users compared with non-users. Thus, continued statin use is advised for such patients exposed to SARS-CoV-2, whereas the initiation of statin therapy to treat COVID-19 is as yet unsubstantiated by the evidence. In addition, the findings of retrospective observational studies still require confirmation by RCTs.

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Conflict of interest statement

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