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Group Sparse Joint Non-negative Matrix Factorization on Orthogonal Subspace for Multi-modal Imaging Genetics Data Analysis

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Abstract

With the development of multi-model neuroimaging technology and gene detection technology, the efforts of integrating multi-model imaging genetics data to explore the virulence factors of schizophrenia (SZ) are still limited. To address this issue, we propose a novel algorithm called group sparse of joint non-negative matrix factorization on orthogonal subspace (GJNMFO). Our algorithm fuses single nucleotide polymorphism (SNP) data, function magnetic resonance imaging (fMRI) data and epigenetic factors (DNA methylation) by projecting three-model data into a common basis matrix and three different coefficient matrices to identify risk genes, epigenetic factors and abnormal brain regions associated with SZ. Specifically, we introduce orthogonal constraints on the basis matrix to discard unimportant features in the row of coefficient matrices. Since imaging genetics data have rich group information, we draw into group sparse on three coefficient matrices to make the extracted features more accurate. Both the simulated and real Mind Clinical Imaging Consortium (MCIC) datasets are performed to validate our approach. Simulation results show that our algorithm works better than other competing methods. Through the experiments of MCIC datasets, GJNMFO reveals a set of risk genes, epigenetic factors and abnormal brain functional regions, which have been verified to be both statistically and biologically significant.

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1. Introduction

IMAGING genetics is an emerging field in brain research [1, 2], which uses brain imaging technology evaluate the influence of genes on individuals, and discusses how genes affect the neural structure, the brain function and the resulting pathology of the nervous system[3–5]. Studying on correlation between genetic variables and imaging variables will facilitate mental disease research like schizophrenia (SZ).

Canonical Correlation Analysis (CCA)[6] is a typical method which is used to find the relationship between genetic and brain imaging data. However, high dimensionality problem featured by high dimension variables and limited sample size in both data sets remains to be a challenge for quantitative analysis. To overcome this issue, feature selection is often integrated in current computational models. For instance, Wright and Parkhomenko et al. [7, 8] built sparse canonical correlation analysis (SCCA), in which L_1 -norm aims to extract significant features through sparse canonical variables and L_2 -norm is used to relieve overfitting problems. Since SCCA doesn't consider the prior information of data, several researchers developed SCCA by adding prior information as constraint conditions. Lin However, a CCA-based model can only find a pair of canonical correlation variables (i.e. one of the modules) at the same time. In order to obtain the second pair, the third pair or even all the canonical correlation variables, we need dimensionality problem featured by high dimension variables and limited sample size in both data sets remains to be a challenge for quantitative analysis. To overcome this issue, feature selection is often integrated in current computational models. For instance, Wright and et al. [9] proposed group sparse CCA(GSCCA), in which the group information of data was considered. Du et al.[10] proposed structure CCA algorithm, in which the spatial structure information of imaging genetics datasets was employed. Moreover, to take advantage of the complementary information of multiple datasets, three or more datasets need to be integrated, Witten et al.[8] proposed sparse Multi-modal CCA (SMCCA) to find the maximum correlation among three or more datasets. Recently, Hu et al.[11] developed an adaptive SMCCA as an extension of the two-module SCCA model.

However, a CCA-based model can only find a pair of canonical correlation variables (i.e.one of the modules) at the same time. In order to obtain the second pair, the third pair or even all the canonical correlation variables, we need to gradually build new CCA model based on the results of the previous step, leading the model to be more complex and time-consuming. The extended model of the non-negative matrix factorization (NMF) algorithm can help to overcome these limitations.

Non-negative matrix factorization [12, 13] is a low-dimensionality reduction approach which has been widely used to analyze imaging genetics datasets since it is useful for learning parts-based representation. NMF decomposes the non-negative object matrix

into the non-negative basis matrix and the coefficient matrix. Considering the group and structure information as prior, Kim *et al.*[14] proposed a group sparse non-negative matrix factorization (GSNMF) algorithm. GSNMF imposed $L_{1,q}$ -norm to realize sparse representation at the level of groups[15]. On the basis of NMF, Zhang *et al.*[16] proposed joint non-negative matrix factorization (JNMF) algorithm which projected multi-modal datasets into a common space for multiple data fusion. Since JNMF hadn't taken into account data prior information, and Wang *et al.*[13] proposed group sparse joint nonnegative matrix factorization (GSJNMF) by combing JNMF and GSJNMF. In these works, each row of the coefficient matrix is equivalent to a canonical correlation variable in CCA model. That is, JNMF and GSJNMF can identify multiple sets of canonical correlation variables simultaneously. However, this NMF-based approach did not consider the correlation between basis vectors, which may lead to extract unimportant features from coefficient matrix.

To this end, we propose a novel model named group sparse of joint non-negative matrix factorization on orthogonal subspace (GJNMFO). GJNMFO enforces both orthogonal and group sparse constraints in matrix decomposition and projects multi-modal data into a low-dimensional orthogonal space. Compared to the previous NMF-based methods, our model has the following advantages. First, we use the idea of joint non-negative matrix to fuse multi-modal imaging genetics data, which can make full use of the complementary information between multi-modal data. Second, GJNMFO projects multi-modal data into a common basis matrix. We employ orthogonal constraint on basis matrix to reduce the correlation between the basis vectors, which helps discard unimportant features in the rows of coefficient matrix. Third, since the imaging genetics datasets have rich group information (i.e., brain voxel can be group by brain functional regions, SNP sites and DNA methylation CpG sites can be grouped by gene), we draw into group sparse constraint on coefficient matrices to make the extracted genetic variables and imaging variables more accurate.

We validate GJNMFO algorithm with both the simulated and real datasets. Simulation results show that our algorithm can identify significant features in multi-modal dataset, and outperform than other existing algorithms, i.e., GSJNMF and JNMF. Results by using the Mind Clinical Imaging Consortium (MCIC) data demonstrate that our algorithm find the reveals a set of risk genes (PLA2G4A, INSIG2, etc.), epigenetic factors (C10orf26, MTERF, etc.) and abnormal brain regions (Hippocampus, Occipital_Inf, etc.) closely related to schizophrenia.

The rest of this paper is organized as follows. In Section II, we propose GJNMFO algorithm and give the numerical optimization, significance estimation and parameter selection strategy. In Section III, we compare the performance of GJNMFO, GSJNMF and JNMF on the simulated data, then use GJNMFO algorithm to analyze the SNP, DNA methylation and fMRI datasets of SZ. Some concluding remarks are given in Section IV.

2 Materials and Methods

2.1 NMF and Extension

NMF has been widely used in data mining and pattern recognition [16–20]. The basic description of NMF model is: given a non-negative matrix $\mathbf{X} \in \mathbb{R}^{m \times n}_+$, we aim to find two low-rank non-negative matrices $\mathbf{W} \in \mathbb{R}^{m \times r}_+$ and $\mathbf{H} \in \mathbb{R}^{r \times n}_+$ such that

$$\min_{\mathbf{W},\mathbf{H}} \|\mathbf{X} - \mathbf{W}\mathbf{H}\|_F^2 \quad s.t. \ \mathbf{W}, \mathbf{H} \ge 0$$
⁽¹⁾

where *m* and *n* are the number of samples and features, respectively. *r* is the number of components, and $\|\cdot\|_F$ is the Frobenius norm. Here, **W** and **H** are called basis matrix and coefficient matrix, respectively. As shown in Eq. (1), **X** can be represented by a linear combination of *r* basis components with relevant coefficients.

Although NMF algorithm performed well on the low-dimensional representation via single matrix decomposition, there was an urgent need to find the connections of two or multiple matrices. Joint non-negative matrix factorization (JNMF) was then proposed to address this issue by decomposing multiple input matrices into one common shared basis matrix and multiple corresponding coefficient matrices. In this way, JNMF can find the association among multi-modal datasets via different coefficient vectors under the common basis vector. As shown in Fig.1, JNMF decomposes multiple input datasets into one common shared basis matrix and multiple different coefficient matrices.

Denote $\mathbf{X}_1 \in \mathbb{R}^{m \times n_1}_+$, $\mathbf{X}_2 \in \mathbb{R}^{m \times n_2}_+$, ..., $\mathbf{X}_K \in \mathbb{R}^{m \times n_K}_+$ as input matrices, JNMF can be formulated as:

$$\min_{\mathbf{W},\mathbf{H}_{i}}\sum_{i=1}^{K} \|\mathbf{X}_{i} - \mathbf{W}\mathbf{H}_{i}\|_{F}^{2} \quad s.t. \ \mathbf{W}, \mathbf{H}_{i} \ge 0$$
⁽²⁾

where *m* and n_i represent the number of samples and features in X_i (*i*=1, 2, 3), respectively. On the basis of JNMF, Wang *et al.*[13] proposed group sparse joint non-negative matrix factorization (GSJNMF) by considering prior group information as constrain imposed on JNMF.

However, these JNMF-based algorithms did not take the correlation between columns of basis matrix into account, so there may be a linear correlation between the basis vectors. This may cause these JNMF-based algorithms extract the insignificant features from coefficient matrix. To this end, we propose a novel algorithm, named group sparse joint non-negative matrix factorization on orthogonal subspace (GJNMFO).

2.2 The Proposed Method

GJNMFO is a JNMF-based method, aiming to discover the potential relationship among multi-modal datasets. Considering the group information of each dataset, i.e. the brain imaging is divided into different ROIs, and the nucleotides belongs to different genes,

we use group information as constraint to the coefficient matrices. Meanwhile, in order to ensure that the basis matrix is full-rank, we impose orthogonal constraint on the basis matrix. Specifically, for the *i*-th modal dataset, the features can be divided into d_i group by utilizing the prior information. Then we employ GJNMFO to decompose the multi-model input data \mathbf{X}_i to one common shared basis matrix \mathbf{W} and multiple corresponding coefficient matrices \mathbf{H}_i , as shown in Fig. 2. Since \mathbf{H}_i is composed by features which also contain the same group information as input, here we impose group sparse constraint on \mathbf{H}_i to extract the significant group of features. In addition, for the common matrix \mathbf{W} , we use $\mathbf{W}^T\mathbf{W}$ -I as orthogonal constraint guarantee full-rank and remove the linear correlation of components. Thus, the objective function of GJNMFO is formulated as follows:

$$\min_{\mathbf{W},\mathbf{H}_{K}} \sum_{i=1}^{K} \left(\frac{1}{2} \| \mathbf{X}_{i} - \mathbf{W}\mathbf{H}_{i} \|_{F}^{2} + \lambda_{i} \| \mathbf{H}_{i} \|_{\varsigma^{i}} \right) + \gamma \| \mathbf{W}^{T}\mathbf{W} - \mathbf{I} \|_{F}^{2} \ s.t. \ \mathbf{W}, \mathbf{H}_{1}, \mathbf{H}_{2}, \cdots,$$

$$\mathbf{H}_{K} \ge 0$$
(3)

where λ_i and γ are regularization parameters. $\gamma = 0$ controls orthogonality of the vectors in **W** (i.e. The larger the γ , the more orthogonal the basis vector is.[21]) and λ_i control sparsity at the group level. For the input matrix **X**_{*i*}, assuming that n_{*i*} features consist of d_i disjoint groups, we have ς^i :

$$\varsigma^{i} = \left\{\varsigma_{1}^{i}, \varsigma_{2}^{i}, \cdots, \varsigma_{d_{i}}^{i}\right\} s \cdot t \cdot \varsigma_{U}^{i} \cap \varsigma_{V}^{i} = \emptyset, U \neq V, \quad \bigcup_{J=1}^{d_{i}} \varsigma_{J}^{i} = \left\{1, 2, \cdots, n_{i}\right\}$$
(4)

where ζ_J^i is the column index set which belongs to *J*-th group in \mathbf{X}_i . The group information contained in the features of the original data \mathbf{X}_i is projected to the coefficient matrix \mathbf{H}_i by GJNMFO algorithm.

For non-negative coefficient matrix $\mathbf{H} \in \mathbb{R}^{r \times n}_+$, we assume that each row index set of **H** is divided into *d* disjoint groups according to **X**, and each group has $|\varsigma_J|$ elements respectively. The group norm $\|\mathbf{H}_{j}\|_{\varsigma_I}$ of the *J*-th group in the *j*-th row is defined as:

$$\|\mathbf{H}_{j}\|_{\zeta J} = \left(\sum_{\alpha \in \zeta J} \mathbf{H}_{j\alpha}^{2}\right)^{\frac{1}{2}}$$
(5)

where ς_J is the column index set that belongs to the *J*-th group in the *j*-th row of **H**. Then the group norm of matrix coefficient **H** can be defined in the following manner:

$$\|\mathbf{H}\|_{\varsigma} = \sum_{j,J} \|\mathbf{H}_{j}\|_{\varsigma J} = \sum_{j,J} \left(\sum_{\alpha \in \varsigma_J} \mathbf{H}_{j\alpha}^2 \right)^{\frac{1}{2}}$$
(6)

It should be noted that $\|\mathbf{H}_{j} \cdot \|_{\zeta J}$ is the $L_{1,2}$ -norm $\|\cdot\|_{1,2}$ if one row of \mathbf{H}_{j} belongs to a single group. Since the number of elements in each group in the coefficient matrix is different, calculating the group norm by Eq. (6) will have different sparsity for each element in \mathbf{H}_{j} . In

order to overcome this problem, we add the size of each group in Eq. (6) to form the final group norm, i.e. $\|H_i\|_S$ is defined in the following manner:

$$\|\mathbf{H}_{i}\|_{\varsigma} = \sum_{j,J} \left(\left| \varsigma_{J}^{i} \right| \sum_{\alpha \in \varsigma_{J}^{i}} (\mathbf{H}_{i})_{j\alpha}^{2} \right)^{\frac{1}{2}}$$
(7)

where $|\varsigma_J^i|$ is the number of elements in *J*-th group in \mathbf{H}_{i} .

It is straightforward to show that problem (3) is non-convex and it is hard to solve it directly. We propose an efficient algorithm based on ALS approach. Let $\Psi \in \mathbb{R}^{m \times r}$ and $\Phi_i \in \mathbb{R}^{r \times n_i}$ be the Lagrange multipliers to constraint $\mathbf{W} = 0$ and $\mathbf{H}_i > 0$ respectively. Then the Lagrange function L has the following expression:

$$L = \sum_{i=1}^{K} \left(\frac{1}{2} \| \mathbf{X}_{i} - \mathbf{W}\mathbf{H}_{i} \|_{F}^{2} + \lambda_{i} \| \mathbf{H}_{i} \|_{\varsigma^{i}} \right) + \gamma \| \mathbf{W}^{T}\mathbf{W} - \mathbf{I} \|_{F}^{2} + \sum_{i=1}^{K} tr \left(\mathbf{\Phi}_{i} \mathbf{H}_{i}^{T} \right)$$

+ $tr \left(\mathbf{\Psi} \mathbf{W}^{T} \right)$ (8)

where tr(·) is trace operation. Taking the partial derivative to **W** and **H**_i (i=1, 2, ..., K) of L, we then obtain:

$$\frac{\partial \mathbf{L}}{\partial \mathbf{W}} = \sum_{i=1}^{K} \left((\mathbf{W}\mathbf{H}_{i} - \mathbf{X}_{i})\mathbf{H}_{i}^{T} \right) + 4\gamma \left(\mathbf{W}\mathbf{W}^{T}\mathbf{W} - \mathbf{W} \right) + \mathbf{\Psi}$$
(9)

$$\frac{\partial \mathbf{L}}{\partial \mathbf{H}_{i}} = \mathbf{W}^{T} (\mathbf{W} \mathbf{H}_{i} - \mathbf{X}_{i}) + \lambda_{i} \frac{\partial \|\mathbf{H}_{i}\|_{\varsigma^{i}}}{\partial \mathbf{H}_{i}} + \mathbf{\Phi}_{i}$$
(10)

According to the Karush-Kuhn-Tucker conditions[22] $\Psi_{pq}W_{pq} = 0$ and $(\Phi_i)_{pq}(\mathbf{H}_i)_{pq} = 0$, we get the following equations for \mathbf{W}_{pq} and $(\mathbf{H}_i)_{pq}$.

$$\sum_{i=1}^{K} \left(\left(\mathbf{W} \mathbf{H}_{i} \mathbf{H}_{i}^{T} \right)_{pq} \mathbf{W}_{pq} + 4\gamma \left(\mathbf{W} \mathbf{W}^{T} \mathbf{W} \right)_{pq} \mathbf{W}_{pq} = \sum_{i=1}^{K} \left(\mathbf{X}_{i} \mathbf{H}_{i}^{T} \right)_{pq} \mathbf{W}_{pq} + 4\gamma \mathbf{W}_{pq} \mathbf{W}_{pq} \left(\mathbf{W}^{T} \mathbf{W} \mathbf{H}_{i} + \lambda_{i} \frac{\partial \|\mathbf{H}_{i}\|_{\zeta^{i}}}{\partial \mathbf{H}_{i}} \right)_{pq} (\mathbf{H}_{i})_{pq} = \left(\mathbf{W}^{T} \mathbf{X}_{i} \right)_{pq} (\mathbf{H}_{i})_{pq}$$

Then we can get the update equations as follows:

$$\mathbf{W}_{pq} = \mathbf{W}_{pq} \frac{\sum_{i=1}^{K} (\mathbf{X}_{i} \mathbf{H}_{i}^{T})_{pq} + 4\gamma \mathbf{W}_{pq}}{\sum_{i=1}^{K} ((\mathbf{W} \mathbf{H}_{i} \mathbf{H}_{i}^{T})_{pq} + 4\gamma (\mathbf{W} \mathbf{W}^{T} \mathbf{W})_{pq}}$$
(11)

$$(\mathbf{H}_{i})_{pq} = (\mathbf{H}_{i})_{pq} \frac{\left(\mathbf{W}^{T} \mathbf{X}_{i}\right)_{pq}}{\left(\mathbf{W}^{T} \mathbf{W} \mathbf{H}_{i} + \lambda_{i} \frac{\partial \|\mathbf{H}_{i}\|_{\varsigma^{i}}}{\partial \mathbf{H}_{i}}\right)_{pq}}$$
(12)

where

$$\left(\frac{\partial \|\mathbf{H}_i\|_{\zeta^i}}{\partial \mathbf{H}_i}\right)_{pq} = \frac{\sqrt{\left|\zeta_J^i\right| (\mathbf{H}_i)_{pq}}}{\sqrt{\sum_{\alpha \in \zeta_J^i} (\mathbf{H}_i)_{p\alpha}^2}}$$
(13)

The iteration termination condition for GJNMFO is $|L^{k+1} - L^k|/|L^{k+1}| = \tau$, where τ is a predefined tolerance error, we generally set $\tau = 10^{-6}$. L^{k+1} and L^k are the (k + 1) – th and k – th reconstruction errors, respectively. The *k*-th reconstruction error L^k is defined as follows:

$$L^{k} = \sum_{i=1}^{3} \|\mathbf{X}_{i} - \mathbf{W}^{(k)}\mathbf{H}_{i}^{k}\|_{F} / \|\mathbf{X}_{i}\|_{F}$$

Where $\mathbf{W}^{(k)}$ and $\mathbf{H}_i^{(k)}$ represent \mathbf{W} and \mathbf{H}_i obtained at the *k*-th iteration, respectively. The pseudocode of GJNMFO algorithm is shown in Algorithm 1.

As GJNMFO projects multi-model data into a common orthogonal space \mathbf{W} , we use each basis component in \mathbf{W} as the 'building block' to explore the correlation of different datasets. In original NMF, the maximum value of each column in \mathbf{H}_i is usually used to define variable's membership[23]. In this way, each variable can belong to one and only one module. However, some variable may not be active in any module or may be active in multiple modules with multiple functions [16]. Based on this problem, we calculate the z-score of each element in each row of \mathbf{H}_1 , \mathbf{H}_2 , \mathbf{H}_3 by the following formula.

$$z_{uv} = \frac{x_{uv} - \mu_u}{\delta_u} \tag{14}$$

where μ_u and σ_u represent the mean and variance of the *u*-th row in \mathbf{H}_i , respectively. The *Q*-th basis vector \mathbf{w}_Q is only related to the *Q*-th row of \mathbf{H}_i (*i*=1, 2, ..., *K*, *Q*=1, 2, ..., *r*). We select element $z_{uv} > T$ to form index set of features $S^Q = \{s_1^Q, s_2^Q, \dots, s_K^Q\}$, where s_i^Q is the sub-index set of feature selected from the *Q*-th row in \mathbf{H}_i and *T* is a predefined threshold. Each variable set S^Q can be regarded as a module (or building block).

```
Algorithm 1. The GJNMFO Algorithm
Input: \mathbf{X}_i \in \mathbb{R}^{m \times n_i}_+, \varsigma^i, \lambda_i, (i = 1, 2, \cdots K), \gamma, r, \tau
Output: W \in \mathbb{R}^{m \times r}_+, H_i \in \mathbb{R}^{r \times n_i}_+ (i = 1, 2, \dots K)
Initialize W and H<sub>i</sub> (i = 1, 2, \dots, K) with random positive matrices.
while 1 do
   Calculate the reconstruction errors L^k
  update W by Eq. (11)
   for r=1: size(W, 2)
        \mathbf{W}(:,r) = \mathbf{W}(:,r) / \| \mathbf{W}(:,r) \|_{2}
    End
for i=1: K
    for p=1: size(H_i, 1)
        for q=1: size(H, , 2)
           for J=1:size(\varsigma')
               \begin{array}{l} \text{if } \mathbf{H}_{i}(p,q) \text{ belong to } \varsigma_{j}^{i} \\ \text{ calculate the value of } \sqrt{|\varsigma_{j}^{i}|} (\mathbf{H}_{i})_{pq} / \sqrt{\sum_{\alpha \in \varsigma_{j}} (\mathbf{H}_{i})_{p\alpha}^{2}} \end{array} 
                end
                update H_i(p, q) by Eq.(12)
           end
        end
      end
    end
        Calculate the reconstruction error L^{k+1}
        if |L^{k+1} - L^k|/|L^{k+1}| \le \tau
           break
        end
end while
```

3 Experiments and Results

The purpose of GJNMFO algorithm is to explore the biomarkers of schizophrenia, i.e. to identify risk genes, epigenetic factors and abnormal brain functional regions. In this paper, we use X_1 , X_2 , and X_3 to represent SNP, fMRI, and DNA methylation datasets, respectively. We use GJNMFO to decompose X_1 , X_2 , and X_3 to obtain r (where r is the numbers of components) modules. We adopt permutation test to evaluate which modules are significant.

3.1 Significance Estimation

For the *Q*-th module $S^Q = \{s_1^Q, s_2^Q, s_3^Q\}$, assume that $\mathbf{A}^Q = [\mathbf{a}_1, \mathbf{a}_2, \dots, \mathbf{a}_{l_1}], \mathbf{B}^Q = [\mathbf{b}_1, \mathbf{b}_2, \dots, \mathbf{b}_{l_2}]$ and $\mathbf{C}^Q = [\mathbf{c}_1, \mathbf{c}_2, \dots, \mathbf{c}_{l_3}]$ where $\mathbf{a}_g \mathbf{b}_t \mathbf{c}_e$ are column vectors selected from $\mathbf{X}_1, \mathbf{X}_2$, and \mathbf{X}_3 according to s_1^Q, s_2^Q and s_3^Q , respectively. Based on the above assumption, the mean of correlations among the three types of datasets in a module can be expressed as follows:

$$\rho^{*} = \frac{1}{3} \left(\frac{1}{l_{1}l_{2}} \sum_{g=1}^{l_{1}} \sum_{t=1}^{l_{2}} \left(\rho(\mathbf{a}_{g}, \mathbf{b}_{t}) \right)^{2} + \frac{1}{l_{1}l_{3}} \sum_{g=1}^{l_{1}} \sum_{e=1}^{l_{3}} \left(\rho(\mathbf{a}_{g}, \mathbf{c}_{e}) \right)^{2} + \frac{1}{l_{2}l_{3}} \sum_{t=1}^{l_{2}} \sum_{e=1}^{l_{3}} \left(\rho(\mathbf{b}_{t}, \mathbf{c}_{e}) \right)^{2} \right)$$
(15)

For a given matrix \mathbf{C}^Q , we randomly change the order of the row vectors of the matrices \mathbf{A}^Q and \mathbf{B}^Q in *Q*-th module, and repeat this process times. For each permutation, ρ^* is used as the null hypothesis of mean correlation, and ρ_{θ}^* is the new mean correlation coefficient

calculated by Eq. (15) after the row of permutation matrix \mathbf{A}^{Q} and \mathbf{B}^{Q} . The significance of the test statistic can be estimated by

$$P-value = \left| \left\{ \theta \mid \rho_{\theta}^* \ge \rho^*, \ \theta = 1, 2, \cdots, \Delta \right\} \right| / \Delta$$
(16)

Where $|\cdot|$ denotes the number of times $\rho_{\theta}^* \ge \rho^*$. If the P-value is less than 0.05, we consider this module is significant.

3.2 Parameter Selection

In GJNMFO, we tune five parameters r, γ , λ_1 , λ_2 and λ_3 . Since determining r is still a challenging problem (if r is too small may lead to a large tolerance error and hinder the extraction of hidden skeletons in data. if r is too large, the purpose of low-rank decomposition of matrix cannot be achieved, which is also not conducive to mining the hidden skeleton in data), we generally set $r\approx 0.1^* \min(m, n_1, n_2, n_3)$. γ is used to control the orthogonality of the column vectors in the basis matrix **W**. If the value of γ is too large, the basis vectors are too sparse, so that the tolerance error of the objective function Eq. (3) is large. If the value of γ is too small, the correlation between the basis vectors cannot be reduced. We selected a variable γ based on the number of iterations. It can be defined as follows:

$$\gamma_k = \gamma_0 \frac{1 - \xi}{1 - \xi^k} \tag{17}$$

where *k* is the number of iterations and $\gamma_0 = 0.1$. A smaller ξ causes faster changes of γ_k and vice versa. The value range of ξ is $0 \quad \xi < 1$. According to experience, we usually make $\xi = 0.1$. λ_1 , λ_2 and λ_3 are used to constrain ς^i in the coefficient matrix \mathbf{H}_i (*i*=1, 2, 3) respectively. According to Ref. [13], the value range of λ_1 , λ_2 and λ_3 is $\left\{0.1 \times \frac{1}{2^n} \mid n = 1, 2, \dots, 10\right\}$. We use the grid search method to find the optimal value of the objective function Eq. (3). Finally, when $|L^{k+1} - L^k|/|L^{k+1}| = \tau$ and the reconstruction error is the smallest, the corresponding λ_1 , λ_2 and λ_3 are the optimal parameters of the model.

3.3 Simulations

In order to test the effectiveness of GJNMFO algorithm, we construct four sets of simulated datasets and then compare the performance of JNMF, GSJNMF and GJNMFO. Finally, we analyze the SNP, fMRI and DNA methylation datasets to identify risk genetic variables, epigenetic factors and abnormal brain ROIs.

3.3.1 Construction the Simulated Data—Owing to the group number and the feature number in each group may not equal, so we generate 4 cases of simulation datasets based on the different group structure information. In each case, there are 3 matrices X_i (*i*=1, 2, 3) consisting of several disjoint groups. The details of the number of groups and features in 4 cases are shown in Table I.

To imitate group structure information, features of the same group are generated by the same seed vector. Assuming that a group contains n features, it can be defined as follows:

(18)

$$\boldsymbol{\alpha}[n] = \{\boldsymbol{\beta}_i \mid \boldsymbol{\beta}_i = \boldsymbol{\alpha} + \ell \boldsymbol{\eta}_i, i = 1, 2, \dots n\}$$

where $\alpha \in \mathbb{R}^m$ is a seed vector, and it is generated from a standard normal distribution. $\eta_i \in \mathbb{R}^m$ is known as Gaussian noise. ℓ denotes a variable that represents the noise level. According to Eq. (18), we construct the simulated data of the four cases. The details are shown in table II.

The numbers (i.e. $1\sim20$, $41\sim60$, $1\sim15$ etc.) in table II indicate the column index of the feature in \mathbf{X}_i . In each case, there are two correlated modules in the three data matrices. For example, in module1 of case1, $1\sim20$ features of \mathbf{X}_1 , $41\sim60$ features of \mathbf{X}_2 , and $81\sim100$ features of \mathbf{X}_3 are generated by the same seed vector. In simulated data, the row of \mathbf{X}_i (*i*=1,2,3) in each case represents a sample, the column of \mathbf{X}_i represents a feature. The sample number *m* of \mathbf{X}_i is equal in all cases, and in this study, m = 40.

To ensure that the \mathbf{X}_i (*i*=1, 2, 3) satisfies the non-negative constraint, we pre-process the datasets in the following four steps. Firstly, we use the z-score method to standardize each column of data, so that the mean value of each column is 0 and the variance is 1. Secondly, we use $F(\mathbf{x})=\mathbf{x}-\min(\mathbf{x})$ to make each column non-negative, where \mathbf{x} is a column vector. Thirdly, we use the L_2 -norm to the unit each column vector of \mathbf{X}_i . Fourthly, we scale all the \mathbf{X}_i (*i*=1, 2, 3) matrices so that their Frobenius norms are equal. Through this step \mathbf{X}_i (*i*=1, 2, 3) have the same value level. It should be noted that the above four steps are linear transformations. Next, we use the above simulation data to verify the performance of JNMF, GSJNMF and GJNMFO.

3.3.2 Analysis of Simulation Result—In our simulation data, there are two correlated modules in each case. Since the basis matrix **W** has a total of 5 column vectors, there are a total of 5 modules. We calculate the z-score (Eq.14) in each module and then select two modules which are the closest to the ground truth as the experimental result. Furthermore, in order to compare the performance of GJNMFO, GSJNMF and JNMF algorithms at different noise levels, we let $\not = 0.5$, 1, 1.5 to construct the simulated data, respectively. Fig. 3 shows the z-score of features obtained by GJNMFO, GSJNMF and JNMF three algorithms in case1.

In Fig. 3, Ground truth represents the real z-score of the features in case1, and the color of the line represents the datasets \mathbf{H}_i (*i*=1, 2, 3), where green, red and blue color represent \mathbf{H}_1 , \mathbf{H}_2 and \mathbf{H}_3 correspondingly.

From Fig. 3(a) we can see, the z-score of GJNMFO in module 1 and module 2 are very close to ground truth and the line are very flat; The z-score of GSJNMF in module 1 and module 2 are close to the ground truth, but the line becomes fluctuant. Even though JNMF can identify most of the significant features, it unfortunately mixes some unrelated features in module 1 and module 2, making the line becomes more fluctuant.

When the noise level is $\ell = 1$ (As it shows in Fig.3(b)), although both GJNMFO and GSJNMF can identify the significant features in module 1 and module 2, GJNMFO line

is flat in two modules, and GSJNMF's line becomes fluctuant due to the incorporation of unrelated features in module 2. JNMF has not recognized the significant features in module 2.

When noise level is $\ell = 1.5$ (As it shows in Fig.3(c)), GJNMFO can accurately identify the meaningful features in module 1 and module 2, and the line is still flat. GSJNMF's module 1 is not flat because some unimportant features are selected. JNMF has not recognized the significant features whether it is module 1 or module 2.

We present the simulation results for case2, case3, and case4 in Appendix A. By comparing the simulation results in the four cases, we can conclude that GJNMFO has better anti-noise performance compared with the state-of-the-art algorithm GSJNMF and JNMF when the level of noise is high.

To analyze the reasons for the performance degradation of GJNMFO, GSJNMF, and JNMF when the noise level increases, we draw the correlation between any two columns in the basis matrix of the three algorithms at different noise levels. The results are shown in Fig. 4.

As can be seen from Fig. 4, with the increase of noise level, the correlation between the basis vectors of GJNMFO is always low, while the correlation between the basis vectors of GJNMFO and JNMF increases with the increase of noise level. When the correlation between basis vectors in GSJNMF and JNMF increases, some unrelated features are mixed and the ability to identify significant features is lost. For example, unrelated features are mixed in JNMF-module 2 in Fig. 3(a). Since too many unrelated features are mixed, some meaningful feature in GSJNMF-module 1 of Fig. 3(c) can't to be identified. However, GJNMFO algorithm overcomes above problem by imposing orthogonal constraints to the basis matrix, which makes GJNMFO algorithm has good anti-noise performance.

In order to further discuss the effect of introducing orthogonal constraints on the basis matrix, we have generated a new set of simulated data $\mathbf{X}_1 \in \mathbb{R}^{m \times n_1}$, $\mathbf{X}_2 \in \mathbb{R}^{m \times n_2}$, and $\mathbf{X}_3 \in \mathbb{R}^{m \times n_3}$. Where *m* represents sample size, n_1 , n_2 and n_3 represent features size, and we set m = 80, $n_1 = 100$, $n_2 = 120$ and $n_3 = 140$ in this paper. Specifically, the generation procedure of X_1, X_2 and X_3 are similar to that in Ref. [10] [24] with the following four steps: 1) we generate vectors \mathbf{u}_1 , \mathbf{u}_2 and \mathbf{u}_3 with lengths of 100, 120, and 140, respectively. Let \mathbf{u}_1 be a vector with 1st~10th elements as 1, 51st~60th elements as 2, 71st~90th elements as 2, and other elements as 0. We use the group information in \mathbf{u}_1 as a priori, i.e. $\varsigma =$ $\{\{1, \dots, 10\}, \{11, \dots, 50\}, \{51, \dots, 60\}, \{61, \dots, 70\}, \{71, \dots, 90\}, \{91, \dots, 100\}\}$. Let \mathbf{u}_2 be a vector with 1st~20th elements as 1, 31st~40th elements as 2, 81st~90th elements as 2, and other elements as 0. The group information of \mathbf{u}_2 is $\boldsymbol{\zeta} = \{\{1, \dots, 20\}, \{21, \dots$ 30, $\{31, \dots, 40\}$, $\{41, \dots, 80\}$, $\{81, \dots, 90\}$, $\{91, \dots, 120\}$. Let \mathbf{u}_3 be a vector with 1st~30th elements as 2, 61st~80th elements as 1, 111st~130th elements as 2, and other elements as 0. The group information of \mathbf{u}_3 is $\varsigma = \{\{1, \dots, 30\}, \{31, \dots, 60\}, \{61, \dots, 80\}, \{61, \dots, 80\},$ $\{81, \dots, 110\}, \{111, \dots, 130\}, \{131, \dots, 140\}\}$. 2) We randomly generate a latent vector **z** of length 80 from N (0, $\mathbf{I}_{80\times80}$) and normalize it to unit L_2 -norm. 3) We generate \mathbf{X}_1 with each sample $\mathbf{x}_i \sim N(z_i \mathbf{u}_1, \Sigma_{X_1})$ where $(\Sigma_{X_1})_{jk} = exp^{-|(u_1)_j - (u_1)_k|}$, \mathbf{X}_2 with each sample

 $\mathbf{x}_i \sim N(z_i \mathbf{u}_2, \Sigma_{X_2})$ where $(\Sigma_{X_2})_{jk} = exp^{-|(u_2)_j - (u_2)_k|}$, and \mathbf{X}_3 with each sample $\mathbf{x}_i \sim N(z_i u_3, \Sigma_{X_3})$ where $(\Sigma_{X_3})_{jk} = exp^{-|(u_3)_j - (u_3)_k|}$. 4) We use the linear transformation method mentioned in section 3.3.1 to make \mathbf{X}_1 , \mathbf{X}_2 and \mathbf{X}_3 non-negative.

We apply GSJNMF and GJNMFO to \mathbf{X}_1 , \mathbf{X}_2 and \mathbf{X}_3 respectively, and their experimental results are shown in Fig. 5. (a) represents the ground truth of \mathbf{u}_1 , \mathbf{u}_2 and \mathbf{u}_3 , where green represents \mathbf{u}_1 , red represents \mathbf{u}_2 , and blue represents \mathbf{u}_3 , (b) represents estimated \mathbf{u}_1 , \mathbf{u}_2 and \mathbf{u}_3 by GSJNMF(where \mathbf{u}_1 , \mathbf{u}_2 and \mathbf{u}_3 correspond to a row vector in \mathbf{H}_1 , \mathbf{H}_2 and \mathbf{H}_3 respectively), and (c) represents estimated \mathbf{u}_1 , \mathbf{u}_2 and \mathbf{u}_3 by GJNMFO. We can see from Fig. 5, \mathbf{u}_1 , \mathbf{u}_2 and \mathbf{u}_3 by GJNMFO are closer to the ground truth of \mathbf{u}_1 , \mathbf{u}_2 and \mathbf{u}_3 than those by GSJNMF, which shows that introducing orthogonal constraints on the basis matrix improves the accuracy on features selection.

We use the Pearson correlation coefficient to quantify the similarity between true \mathbf{u}_i and estimated \mathbf{u}_i by GSJNMF and GJNMFO (i = 1,2,3). As shown in Table III, the similarities between true \mathbf{u}_i and estimated \mathbf{u}_i by GSJNMF are 0.5713, 0.9073 and 0.7712, respectively. The similarities between true \mathbf{u}_i and estimated \mathbf{u}_i by GJNMFO are 0.9812, 0.9745 and 0.9370, respectively. From Table III, we can conclude that the performance of GJNMFO on similarity to the ground truth is better than that of GSJNMF. Furthermore, we use heat map drawing the correlation between any two columns in the basis matrix of GJNMFO and GSJNMF. As shown in Appendix B, the color of the heat map corresponding to GJNMFO is significantly cooler than GSJNMF, which shows that GJNMFO gets a higher incoherence between basis vectors. According to Ref. [25], high incoherence between basis vectors can guarantee that basis vectors are as discriminative as possible, which is the fundamental reason that the performance of GJNMFO is better than that of GSJNMF.

In order to make the dimension of simulated data closer to real one, we construct a new set of simulated data which dimension is 10 times the previous, i.e. $\mathbf{X}_1 \in \mathbb{R}^{80 \times 1000}$, $\mathbf{X}_2 \in \mathbb{R}^{80 \times 1200}$, and $\mathbf{X}_3 \in \mathbb{R}^{80 \times 1400}$. The experimental results of GJNMFO and GSJNMF on this simulation data are illustrated in Appendix C, shown that GJNMFO is superior to GSJNMF even when the data is in a high dimension. Next, we use GJNMFO algorithm to analyze the risk genes, epigenetic factors and the abnormal brain ROIs of SZ.

3.4 Application on Schizophrenia Dataset

3.4.1 Real Data Preparation and Preprocessing—Schizophrenia is a complex mental disease in which people often characterized by brain abnormalities, genetic variations, and environmental factors[25]. It is very important to identify how genetics and environmental factors interact and affect brain function and cognition[26].

The SNP, DNA methylation and fMRI datasets used in this paper were collected by the Mind Clinical Imaging Consortium (MCIC) from 183 subjects, including 79 SZ patients (age: 34±11, 20 females) and 104 healthy controls (age: 32±11, 38 females). The fMRI data was collected while subjects were performing sensory motion task. Then, the data was pre-processed using SPM12 software and was realigned spatially normalized and resliced

to $3 \times 3 \times 3$ mm. It was smoothed with a $10 \times 10 \times 10$ mm³ Gaussian kernel and analyzed by multiple regression considering the stimulus and their temporal derivatives plus an intercept term as repressors [27]. After these steps, the dimension of fMRI data is the 53*63*46. Then we select 41236 voxels from 116 brain regions according to the AAL brain atlas. The SNPs data was obtained from each subject's blood sample. Genotyping for all participants was performed at the Mind Research Network, covering 1140419 SNP loci. Bead Studio was used to make the final genotype calls. PLINK software package (http:// pngu.mgh.harvard.edu/~purcell/plink) was used to perform a series of standard quality control procedures, resulting in the final dataset spanning 777,635 SNP loci [9]. The DNA methylation data was obtained from the blood samples of the subjects, which assessed by the Illumina Infinium Methylation 27 k Assay. 27481 methylation CpG sites were selected after quality control[28], and 9273 methylation sites were further selected after removing sites with variance less than 1×10^{-4} . Furthermore, we use the t-test to find biomarkers that are only associated with SZ, and only retain variables with P-value < 0.05. We obtain $\mathbf{X}_1 \in \mathbb{R}^{183 \times 27041}$ SNPs loci, $\mathbf{X}_2 \in \mathbb{R}^{183 \times 2918}$ fMRI voxels and $\mathbf{X}_3 \in \mathbb{R}^{183 \times 1845}$ DNA methylation CpG sites, respectively. Moreover, as shown in Fig. 6, we group SNPs loci, DNA methylation CpG sites, and fMRI voxels based on genes and ROIs (e.g., SNPs loci belong to the same gene, fMRI voxels belong to same ROI, DNA methylation CpG sites belong to the same gene). Then SNP, fMRI and DNA methylation datasets are divided into 3412, 72 and 1845 groups, respectively.

3.4.2 Experimental Results on Real Data and Discussion—We apply GJNMFO algorithm to SNP, fMRI and DNA methylation datasets. In our experiment, the number of basis vector is set to 20 according to previous work [13]. We randomly initialize a set of non-negative **W** and \mathbf{H}_i (*i*=1, 2, 3), and use the grid search method to determine the parameters of λ_1 , λ_2 and λ_3 within the range described above. In this way, a total of 1000 sets of parameters need to be brought into the model for calculation. Fig. 7 is the reconstruction error obtained by sequentially substituting 1000 sets of regularization parameters into algorithm 1. We use the set of regularization parameters corresponding to the minimum reconstruction error for further analysis.

In Fig. 7, each point on the horizontal axis corresponds to a set of regularization parameters, and the vertical axis represents the reconstruction error obtained under the set of parameters. As can be seen from Fig. 7, when taking the 93rd group of regularization parameters, the reconstruction error is the smallest (i.e. the position of the red dot in Fig. 7). Parameters in group 93 are λ_1 =0.05, λ_2 =9.765×10⁻⁵ and λ_3 =0.0125.

We use this set of regularization parameters as the optimal one. To prevent the objective function Eq. (3) from falling into a local minimum, we repeat the whole procedure 100 times with different initialization values. When the objective function takes the minimum value, the corresponding \mathbf{W} , \mathbf{H}_1 , \mathbf{H}_2 and \mathbf{H}_3 are used for further analysis. We have drawn the values of the objective function under 100 sets of initialization value in Fig. 8.

We can see from the Fig. 8 that the 49-th initialization value corresponds to the smallest objective function value (i.e. the position of the red dot in Fig. 8), so we use \mathbf{W} , \mathbf{H}_1 , \mathbf{H}_2 and \mathbf{H}_3 obtained here to extract risk genes, epigenetic factors and abnormal ROIs. We calculate

the z-score of \mathbf{H}_1 , \mathbf{H}_2 and \mathbf{H}_3 , and set T=3. If the element with z-score is less than T, then it is 0. In this way, we get a set of variables of each module, and then we use the above-mentioned significance test method (Eq. 16) to calculate the P-value of each module.

The modules with smallest P-value are further analyzed in this study. Table IV lists the genes identified by GJNMFO algorithm from the SNPs loci.

As can be seen from Table IV, we identify a total of 54 risk genes from the SNPs data in the 19-th module. Since the protein encoded by CNTN2 has the function of adhering cells to help maintain the potassium ion channel in the correct position on the nerve fibers which plays an important role in the process of generating nerve impulses. Therefore, its mutation causes the nerve cells to behave abnormally. INSIG2 is associated with metabolic syndrome in patients with schizophrenia[29]. Lmo4 gene plays an important role in the development of the auditory nervous system[30]. OLFM3 is associated with a disorder of the metabolic pathway leading to low expression of nerve-specific genes, which causes neuronal weakness[31]. The abnormalities of PLA2G4A may be involved in a subgroup of the Schizophrenia. MaII *et al.*[32] used genome-wide association studies to demonstrate that multiple sites in the PLXNA2 gene are significantly associated with SZ.

In the 13-th module, we identify a total of 36 risk genes from the SNPs data. KCNK9 is a key factor in the excitability of cortical neurons, which may cause defects in neuronal migration. NGF is associated with the onset and clinical symptoms of schizophrenia[33]. The NTNG1 gene is located in the lpl3.3 region of chromosome 1, and it is in a linked region with the onset of schizophrenia. In the family studies and case-control studies of the Japanese population, multiple SNPs loci of this gene have been found to be associated with schizophrenia[34]. PDE4B is an important enzyme in the nervous system, and recent studies have found it to play a decisive role in DISC1 mutation-induced schizophrenia[35].

To further illustrate the biological significance of the risk genes which list in table IV, we test genes for over-representation analysis using ConsensusPathDB[11, 36]. Table V lists the gene ontology terms that obtain by gene ontology categories with P-value less than 0.01 in biological process.

In table V, 'neurogenesis' and 'neuron differentiation' are related to the neural activity. 'purinergic receptor signaling pathway' has been confirmed to be related to neuronal differentiation, neuroprotection, and brain disorders[37]. G protein-coupled purinergic receptor is widely distributed in the nervous system and peripheral tissues, performing various physiological functions such as development and regeneration of the nervous system, regulating neurons, regulating memory function, regulating transmitter release, etc. [38]. Through the above analysis, we can conclude that the risk genes selected by GJNMFO from SNPs have biological significance. Table VI lists the epigenetic factors identified by GJNMFO algorithm from the DNA methylation dataset.

As shown in Table VI, 4 genes related to environmental factors are selected from DNA methylation dataset. 'C10orf26' also selected by Ref.[13], which is reported as one of the target of miR-137 to have genome-wide significant associations with SZ[39]. When mtDNA is transcribed to produce RNA, 'MTERF' is responsible for determining the location of

mitochondrial transcription termination factors, Lauritzen *et al.*[40] demonstrated that point mutations in mitochondria may be associated with cognitive impairment of mental disease. Therefore, 'MTERF' is related to mental disease. Besides, 'LOC201164' is considered to be related to neural development in Ref.[41]. 'RAET1L' is an additional human NKG2D ligand which is related to the immune system[42], and clinical studies have shown that the immune system and brain function of patients with schizophrenia are abnormal. Table VII lists the brain ROIs selected by GJNMFO and adaptive SMCCA respectively.

In table VII, GJNMFO algorithm selects a total of 4 ROIs and adaptive SMCCA selects a total of 6 ROIs. 3 ROIs are selected at the same time by GJNMFO and adaptive SMCCA algorithm. 'Hippocampus gyrus' is important in the consolidation of information from short-term memory to long-term memory, and is one of the first brain regions to suffer damage in mental disorders, and is also selected by Ref.[26] and Ref.[43]. The 'occipital lobe' is responsible for processing visual information. Occipital injury can make it difficult for schizophrenia patients to interpret complex images, recognize other people's motives, and understand other people's emotions. The volume of the 'fusiform gyrus' in patients with first-episode schizophrenia is reduced, and the reduction in the volume of the fusiform gyrus causes memory impairment [44]. 'fusiform gyrus' is selected by Ref.[13] and Ref.[26]. The volume of amygdala in patients with schizophrenia is 6% lower than that of normal subjects, and the decrease in amygdala volume is related to the recognition of facial emotions[45]. In order to illustrate the specific location of the abnormal brain region in brain surface, we use BrainNetViewer to plot the abnormal brain ROIs that selected by GJNMFO as shown in Fig. 9.

4 Conclusion

The main contributions of this paper can be summarized as follows. Firstly, we propose GJNMFO algorithm, which projects the multi-modal datasets of SZ into a new common low-dimensional space to identify risk genes, epigenetic factors and abnormal brain ROIs. The GJNMFO model integrates the orthogonal and group sparse constraint on the basis of the JNMF model. We employ orthogonal constraint to reduce the correlation between basis vectors, which helps remove insignificant features in the rows of coefficient matrix. In addition, since the imaging genetics datasets have rich group information, we employ group sparse constraint into the model as prior knowledge to make the extracted genetic variables and imaging variables more reasonable. Secondly, we describe the numerical optimization, significance estimation and parameters selection of GJNMFO algorithm. Thirdly, our algorithm is validated on simulation data and further applied on real data analysis. Simulation results show that our algorithm performs better than state-of-the-art algorithms GSJNMF and JNMF. Experimental results on real data indicate that our algorithm can identify more significant genes (PLA2G4A, INSIG2, etc.), epigenetic factors (C10orf26, MTERF, etc.) and abnormal ROIs (Hippocampus, Occipital Inf, etc.), which are strong correlation with SZ and have been reported in many previous literatures. In conclusion, both simulation and real MCIC experiment results show that GJNMFO algorithm has outperformance in multi-modal imaging genetics data analysis.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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References

- Hariri AR, Drabant EM, and Weinberger DR, "Imaging genetics: perspectives from studies of genetically driven variation in serotonin function and corticolimbic affective processing," Biological psychiatry, vol. 59, no. 10, pp. 888–897, 2006. [PubMed: 16442081]
- [2]. Meyer-Lindenberg A, "The future of fMRI and genetics research," NeuroImage, vol. 62, no. 2, pp. 1286–1292, 2012. [PubMed: 22051224]
- [3]. Gottesman II and Gould TD, "The endophenotype concept in psychiatry: etymology and strategic intentions," American Journal of Psychiatry, vol. 160, no. 4, pp. 636–645, 2003.
- [4]. Meyer-Lindenberg A and Weinberger DR, "Intermediate phenotypes and genetic mechanisms of psychiatric disorders," Nature reviews neuroscience, vol. 7, no. 10, p. 818, 2006. [PubMed: 16988657]
- [5]. Ge T, Schumann G, and Feng J, "Imaging genetics—towards discovery neuroscience," Quantitative Biology, vol. 1, no. 4, pp. 227–245, 2013.
- [6]. Hotelling H, "Relations between two sets of variates," Biometrika, vol. 28, no. 3/4, pp. 321–377, 1936.
- [7]. Parkhomenko E, Tritchler D, and Beyene J, "Sparse canonical correlation analysis with application to genomic data integration," Statistical applications in genetics and molecular biology, vol. 8, no. 1, pp. 1–34, 2009.
- [8]. Witten DM, Tibshirani R, and Hastie T, "A penalized matrix decomposition, with applications to sparse principal components and canonical correlation analysis," Biostatistics, vol. 10, no. 3, pp. 515–534, 2009. [PubMed: 19377034]
- [9]. Lin D, Calhoun VD, and Wang Y-P, "Correspondence between fMRI and SNP data by group sparse canonical correlation analysis," Medical image analysis, vol. 18, no. 6, pp. 891–902, 2014.
 [PubMed: 24247004]
- [10]. Du L et al., "Structured sparse canonical correlation analysis for brain imaging genetics: an improved GraphNet method," Bioinformatics Medical image analysis, vol. 32, no. 10, pp. 1544– 1551, 2016.
- [11]. Hu W et al., "Adaptive sparse multiple canonical correlation analysis with application to imaging (epi) genomics study of schizophrenia," IEEE Transactions on Biomedical Engineering, vol. 65, no. 2, pp. 390–399, 2018. [PubMed: 29364120]
- [12]. Lee DD and Seung HS, "Learning the parts of objects by non-negative matrix factorization," Nature Australia, vol. 401, no. 6755, p. 788, 1999.
- [13]. Wang M, Huang T-Z, Fang J, Calhoun VD, and Wang Y-P, "Integration of imaging (epi) genomics data for the study of schizophrenia using group sparse joint nonnegative matrix factorization," IEEE/ACM transactions on computational biology and bioinformatics, 2019.
- [14]. Kim J, Monteiro RD, and Park H, "Group sparsity in nonnegative matrix factorization," in Proceedings of the SIAM International Conference on Data Mining, 2012, pp. 851–862: SIAM.
- [15]. Yuan M and Lin Y, "Model selection and estimation in regression with grouped variables," Journal of the Royal Statistical Society: Series B (Statistical Methodology), vol. 68, no. 1, pp. 49–67, 2006.

- [16]. Zhang S, Liu C-C, Li W, Shen H, Laird PW, and Zhou XJ, "Discovery of multi-dimensional modules by integrative analysis of cancer genomic data," Nucleic acids research, vol. 40, no. 19, pp. 9379–9391, 2012. [PubMed: 22879375]
- [17]. Liu W and Zheng N, "Non-negative matrix factorization based methods for object recognition," Pattern Recognition Letters, vol. 25, no. 8, pp. 893–897, 2004.
- [18]. Shahnaz F, Berry MW, Pauca VP, and Plemmons RJ, "Document clustering using nonnegative matrix factorization," Information ProcessingManagement data series. Parks and Wildlife Department (Texas), vol. 42, no. 2, pp. 373–386, 2006.
- [19]. Berry MW, Browne M, Langville AN, Pauca VP, and Plemmons RJ, "Algorithms and applications for approximate nonnegative matrix factorization," Computational Statistics & Data Analysis, vol. 52, no. 1, pp. 155–173, 2007.
- [20]. Xu Y and Qian Y, "Group sparse nonnegative matrix factorization for hyperspectral image denoising," in Geoscience & Remote Sensing Symposium, 2016.
- [21]. Li Z, Wu X, and Peng H, "Nonnegative matrix factorization on orthogonal subspace," Pattern Recognition Letters, vol. 31, no. 9, pp. 905–911, 2010.
- [22]. Kuhn HW, "Nonlinear programming: a historical view," ACM SIGMAP Bulletin, no. 31, pp. 6–18, 1982.
- [23]. Brunet J-P, Tamayo P, Golub TR, and Mesirov JP, "Metagenes and molecular pattern discovery using matrix factorization," Proceedings of the national academy of sciences, vol. 101, no. 12, pp. 4164–4169, 2004.
- [24]. Chen X, Han L, and Carbonell J, "Structured sparse canonical correlation analysis," in Artificial Intelligence and Statistics, pp. 199–207, 2012
- [25]. Fang J, Lin D, Schulz SC, Xu Z, Calhoun VD, and Wang Y-P, "Joint sparse canonical correlation analysis for detecting differential imaging genetics modules," Bioinformatics Medical image analysis, vol. 32, no. 22, pp. 3480–3488, 2016.
- [26]. Hu W, Lin D, Calhoun VD, and Wang Y.-p., "Integration of SNPs-FMRI-methylation data with sparse multi-CCA for schizophrenia study," in Medicine and Biology Society (EMBC), IEEE 38th Annual International Conference of the Engineering, 2016, pp. 3310–3313: IEEE.
- [27]. Landeau B, "Automated anatomical labeling of activations in SPM using a macroscopic anatomical parcellation of the MNI MRI single-subject brain," Neuroimage, vol. 15, no. 1, pp. 273–289, 2002. [PubMed: 11771995]
- [28]. Jingyu L et al., "Methylation patterns in whole blood correlate with symptoms in schizophrenia patients," Schizophrenia Bulletin, vol. 40, no. 4, p. 769, 2014. [PubMed: 23734059]
- [29]. Liou YJ et al., "Gene-gene interactions of the INSIG1 and INSIG2 in metabolic syndrome in schizophrenic patients treated with atypical antipsychotics," Pharmacogenomics Journal, vol. 12, no. 1, p. 54, 2012.
- [30]. Min D et al., "LMO4 functions as a negative regulator of sensory organ formation in the mammalian cochlea," Journal of Neuroscience, vol. 34, no. 30, pp. 10072–10077, 2014. [PubMed: 25057208]
- [31]. Vita A, De PL, Deste G, and Sacchetti E, "Progressive loss of cortical gray matter in schizophrenia: a meta-analysis and meta-regression of longitudinal MRI studies," Translational Psychiatry, vol. 2, no. 11, p. e190, 2012. [PubMed: 23168990]
- [32]. Mah S et al., "Identification of the semaphorin receptor PLXNA2 as a candidate for susceptibility to schizophrenia," Molecular Psychiatry, vol. 11, no. 5, pp. 471–478, 2006. [PubMed: 16402134]
- [33]. Martinotti G, Di IG, Marini S, Ricci V, De BD, and Di GM, "Nerve growth factor and brainderived neurotrophic factor concentrations in schizophrenia: a review," Journal of Biological Regulators & Homeostatic Agents, vol. 26, no. 3, p. 347, 2012. [PubMed: 23034254]
- [34]. Ohtsuki T et al., "Association of polymorphisms in the haplotype block spanning the alternatively spliced exons of the NTNG1 gene at 1p13.3 with schizophrenia in Japanese populations," Neuroscience Letters, vol. 435, no. 3, pp. 194–197, 2008. [PubMed: 18384956]
- [35]. J Kirsty M et al., "DISC1 and PDE4B are interacting genetic factors in schizophrenia that regulate cAMP signaling," Science, vol. 310, no. 5751, pp. 1187–1191, 2005. [PubMed: 16293762]

- [36]. Kamburov A, Stelzl U, Lehrach H, and Herwig R, "The ConsensusPathDB interaction database: 2013 update," Nucleic acids research, vol. 41, no. D1, pp. D793–D800, 2013. [PubMed: 23143270]
- [37]. Majumder P et al., "New insights into purinergic receptor signaling in neuronal differentiation, neuroprotection, and brain disorders," Purinergic Signalling, vol. 3, no. 4, pp. 317–331, 2007. [PubMed: 18404445]
- [38]. Illes P, Verkhratsky A, Burnstock G, and Franke H, "P2X receptors and their roles in astroglia in the central and peripheral nervous system," Neuroscientist, vol. 18, no. 5, pp. 422–438, 2012. [PubMed: 22013151]
- [39]. Kwon E, Wang W, and Tsai L-H, "Validation of schizophrenia-associated genes CSMD1, C10orf26, CACNA1C and TCF4 as miR-137 targets," Mol Psychiatry, vol. 18, no. 1, pp. 11–12, 2013. [PubMed: 22182936]
- [40]. Lauritzen KH et al., "Mitochondrial DNA toxicity in forebrain neurons causes apoptosis, neurodegeneration, and impaired behavior," Molecular and cellular biology, vol. 30, no. 6, pp. 1357–1367, 2010. [PubMed: 20065039]
- [41]. Kerkel K et al., "Altered DNA Methylation in Leukocytes with Trisomy 21," Plos Genetics, vol. 6, no. 11, p. e1001212, 2010. [PubMed: 21124956]
- [42]. Eagle RA, Traherne JA, Hair JR, Jafferji I, and Trowsdale J, "ULBP6/RAET1L is an additional human NKG2D ligand," European Journal of Immunology, vol. 39, no. 11, pp. 3207–3216, 2010.
- [43]. Ozdemir H, Ertugrul A, Basar K, and Saka E, "Differential effects of antipsychotics on hippocampal presynaptic protein expressions and recognition memory in a schizophrenia model in mice," Progress in Neuro-Psychopharmacology and Biological Psychiatry, vol. 39, no. 1, pp. 62–68, 2012. [PubMed: 22640753]
- [44]. Uk LC et al., "Fusiform gyrus volume reduction in first-episode schizophrenia: a magnetic resonance imaging study," Archives of General Psychiatry, vol. 59, no. 9, p. 775, 2002. [PubMed: 12215076]
- [45]. Namiki C et al., "Impaired facial emotion recognition and reduced amygdalar volume in schizophrenia," Psychiatry Res, vol. 156, no. 1, pp. 23–32, 2007. [PubMed: 17728113]



Fig. 1.

JNMF decomposes the multimodal imaging genetics data into a new common orthogonal space. \mathbf{w}_1 , \mathbf{w}_2 , and \mathbf{w}_3 represent basis vectors. Triangles, squares and circles represent data of different modalities. In Fig.1a, the original data \mathbf{X}_1 , \mathbf{X}_2 , \cdots , \mathbf{X}_K are projected into the respective basis vectors. In Fig.1b, the multimodal data \mathbf{X}_K is projected into the same set of basis vectors.





Fig. 2.

Schematic illustrating our method. Our method projects the multi-modal imaging genetics data into a common matrix and different coefficient matrices. The features of the original data \mathbf{X}_K can be divided into d_K groups, which are $\varsigma_1^K, \varsigma_2^K, \dots, \varsigma_{d_K}^K$. The yellow and white in $\mathbf{H}_1, \mathbf{H}_2, \dots, \mathbf{H}_k$ represent non-zero and zero respectively.



Fig. 3.

Z-score of features obtained by GJNMFO, GSJNMF and JNMF three algorithms in case1. (a), (b), and (c) are z-score of features obtained by three algorithms at a noise level of 0.5, 1, and 1.5, respectively. Ground truth represents the real z-score of the features.

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0.9

0.8

0.95

0.9

0.85

0.95

0.9

0.85





The correlations between any two columns in the basis matrix of GJNMFO, GSJNMF and JNMF at different noise levels in case1.

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Fig. 5.

(a) represents the ground truth of \mathbf{u}_1 , \mathbf{u}_2 and \mathbf{u}_3 , where green represents \mathbf{u}_1 , red represents \mathbf{u}_2 , and blue represents \mathbf{u}_3 . (b) Estimated \mathbf{u}_1 , \mathbf{u}_2 and \mathbf{u}_3 by GSJNMF. (c) Estimated \mathbf{u}_1 , \mathbf{u}_2 and \mathbf{u}_3 by GJNMFO.



Fig. 6.

SNP loci, DNA methylation CpG sites and fMRI voxels can group by gene and ROI respectively.

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The value of objective function with different initialization value

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Fig. 9. Abnormal ROIs selected by GJNMFO from fMRI data.

TABLE I

The Number of Groups and Features in Four Cases

	case 1	case 2	case3	case4
Number of groups in X_i	1	1	0	0
Number of features in each group	1	0	1	0

Statements that '1' means equal and '0' mean unequal. For example, in case 2, '1' indicates that the number of groups among X_1 , X_2 , and X_3 is equal, and '0' indicates that the features number of each group in X_i (*i*=1, 2, 3) is not equal.

TABLE II

The Column Index Set of Correlated Modules in Each Case

	column index set							
Dataset	ca	sel	case2		case3		case4	
	module1	module2	module1	module2	module1	module2	module1	module2
X1	1~20	41~60	1~15	21~29	1~20	41~60	1~100	101~230
X2	41~60	1~20	31~41	1~11	21~40	1~20	281~390	1~80
X3	81~100	61~80	50~55	7~15	81~100	121~140	410~590	171~290

TABLE III

The correlation coefficient between estimated \mathbf{u}_i and true \mathbf{u}_i (*i* = 1,2,3).

methods	The correlation coefficient			
	correlation between \mathbf{u}_1 and $\widehat{\boldsymbol{u}}_1$	correlation between \mathbf{u}_2 and $\widehat{\boldsymbol{u}}_2$	correlation between \mathbf{u}_3 and $\widehat{\boldsymbol{u}}_3$	
GSJNMF	0.5713	0.9073	0.7712	
GJNMFO	0.9821	0.9745	0.9370	

TABLE IV

The List of the Gene Select from SNPs Data

Module Index	Gene ID (SNP)
19	CNTN2 CDH4 ABCA4 AGBL4 ALDH9A1 C1orf161 C1orf163 CNTN2 DISC1 DNAH14 DNAJC11 DPYD ELAVL4 ELTD1 ESRRG FAM176A FAM46C FHL2 INSIG2FLJ35409 GPR177 HHAT KCNA3 KCND3 LMO4 KIRREL LOC100288925 LOC100289242 PLA2G4ALOC284661 LOC727944 LOC730134 LRRC38 OLFM3MYADML MYCN NPHP1 OTOF PADI4 PBX1 PLXNA2 PKP1 PRKCE QPCT RGS13 RNF144A RPS6KC1 RSBN1 SPATA1 ST6GALNAC3 TRIB2 TYW3 VAV3 VIT
13	KCNK9 ADORA3 AIM2 AJAP1 C1orf168 C2orf3NTNG1 CCDC85A CCDC88A CDC73 CYP2J2 FAM71A NGFFEZ2EZ2 FMO4 FMO9P GREM2 IL1R2 KIF26B LOC100129149 LOC126987 LOC644265 LOC728597 LOC730100 LRP8 LUZP1 MGST3 NT5C1B OLFML2B PDE4B PGBD5 RHOU SLC44A5 SRBD1 USP24 VPS13D

TABLE V

The list of the gene ontology terms

term name	p-value
neurogenesis	0.00598
neuron differentiation	0.00477
purinergic receptor signaling pathway	0.00554
G protein-coupled purinergic receptor signaling pathway	0.0032
icosanoid metabolic process	0.00628

TABLE VI

The List of the Gene Selected from DNA Methylation Data

Module Index	Gene ID (DNA methylation)		
19	LOC201164	RAET1L	
13	Clorf26	MTERF	

TABLE VII

Selected ROIs (fMRI) and Corresponding Areas

GJNMFO		Adaptive SMCCA		
ROI name	L/R volumn(cm ³)	ROI name	L/R volumn(cm ³)	
Hippocampus	1.59/*	Hippocampus	1.65/*	
Amygdala	1.3/*	Frontal_Inf_Tri	1.03/*	
Occipital_Inf	*/1.22	Occipital_Inf	*/1.22	
Fusiform_R	*/0.678	Fusiform	1.86/1.00	
		Temporal_Mid	*/1.30	
		Temporal_Inf	2.11/*	

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